

Applied Stochastic Analysis

Homework assignment 7

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Exercise 1

Since every Y_t^i , $i = 1, \dots, d$, is a linear combination of continuous independent (centered) Gaussian processes then every Y_t^i is a continuous (centered) Gaussian process too. Hence, we only need to examine its covariance function. We have

$$\begin{aligned}\mathbb{E}[Y_s^i Y_t^j] &= \sum_{k=1}^d \sum_{n=1}^d Q_{ik} Q_{jn} \mathbb{E}[W_s^k W_t^n] = (s \wedge t) \sum_{k=1}^d \sum_{n=1}^d Q_{ik} Q_{jn} \delta_{kn} \\ &= (s \wedge t) \sum_{k=1}^d Q_{ik} Q_{jk} = (s \wedge t) (Q \cdot Q^t)_{ij} = (s \wedge t) \delta_{ij}.\end{aligned}$$

The above concludes the proof that $\mathbf{Y}_t = (Y_t^1, \dots, Y_t^d)$ is a d -dimensional Brownian motion.

Exercise 2

(a) First of all, we notice that, by definition of that T_a , \tilde{B}_t is a.s. continuous. Hence, to show that \tilde{B}_t is a B.m. (Brownian motion), it is sufficient to show that

$$P(\tilde{B}_s \leq x, \tilde{B}_t \leq y) = P(B_s \leq x, B_t \leq y)$$

for every $0 \leq s \leq t$ and $x, y \in \mathbb{R}$ (i.e. that $(\tilde{B}_t, \tilde{B}_s)$ and (B_t, B_s) have the same distribution). Let $\mathcal{F}_t = \sigma(B_s : 0 \leq s \leq t)$. By the strong Markov property, we know that $W_t \doteq B_{t+T_a} - B_{T_a}$ is a B.m. independent by \mathcal{F}_{T_a} . Notice that

$$\tilde{W}_t \doteq \tilde{B}_{t+T_a} - \tilde{B}_{T_a} = 2a - B_{t+T_a} - B_{T_a} = B_{T_a} - B_{t+T_a} = -W_t,$$

so \tilde{W}_t is a B.m. independent by \mathcal{F}_{T_a} too. We can write

$$B_t = B_t \mathbb{1}_{\{t \leq T_a\}} + (W_{t-T_a} + B_{T_a}) \mathbb{1}_{\{t > T_a\}} \quad \text{and} \quad \tilde{B}_t = B_t \mathbb{1}_{\{t \leq T_a\}} + (\tilde{W}_{t-T_a} + B_{T_a}) \mathbb{1}_{\{t > T_a\}}.$$

Now

$$\begin{aligned}P(\tilde{B}_s \leq x, \tilde{B}_t \leq y) &= P(B_s \leq x, B_t \leq y, t \leq T_a) + P(B_s \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y, s \leq T_a < t) \\ &\quad + P(\tilde{W}_{s-T_a} + B_{T_a} \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y, T_a < s).\end{aligned} \tag{1}$$

We can evaluate the second term in the rhs as follows:

$$P(B_s \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y, s \leq T_a < t) = \mathbb{E}[P(B_s \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y \mid \mathcal{F}_{T_a}) \mathbb{1}_{\{s \leq T_a < t\}}],$$

where

$$P(B_s \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y \mid \mathcal{F}_{T_a}) = \tilde{f}(B_s, T_a)$$

with

$$\tilde{f}(z, r) \doteq \mathbb{1}_{\{z \leq x\}} P(\tilde{W}_{t-r} + a \leq y) = \mathbb{1}_{\{z \leq x\}} P(W_{t-r} + a \leq y) \doteq f(z, r).$$

Hence

$$\begin{aligned} P(B_s \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y, s \leq T_a < t) &= \mathbb{E}[f(B_s, T_a) \mathbb{1}_{\{s \leq T_a < t\}}] \\ &= P(B_s \leq x, W_{t-T_a} + B_{T_a} \leq y, s \leq T_a < t). \end{aligned} \quad (2)$$

Similarly one can prove that

$$P(\tilde{W}_{s-T_a} + B_{T_a} \leq x, \tilde{W}_{t-T_a} + B_{T_a} \leq y, T_a < s) = P(W_{s-T_a} + B_{T_a} \leq x, W_{t-T_a} + B_{T_a} \leq y, T_a < s). \quad (3)$$

From (1), (2) and (3) one gets that

$$\begin{aligned} P(\tilde{B}_s \leq x, \tilde{B}_t \leq y) &= P(B_s \leq x, B_t \leq y, t \leq T_a) + P(B_s \leq x, W_{t-T_a} + B_{T_a} \leq y, s \leq T_a < t) \\ &\quad + P(W_{s-T_a} + B_{T_a} \leq x, W_{t-T_a} + B_{T_a} \leq y, T_a < s) \\ &= P(B_s \leq x, B_t \leq y). \end{aligned}$$

This completes the proof that \tilde{B}_t is a B.m.

(b) It holds that

$$P(M_t > a) = P(M_t > a, B_t > a) + P(M_t > a, B_t \leq a) = P(B_t > a) + P(M_t > a, B_t \leq a),$$

since $B_t > a$ implies $M_t = \sup_{s \leq t} B_s \geq B_t > a$.

(c) Since $\{M_t > a\} = \{T_a \leq t\}$ and if $T_a \leq t$ then $\tilde{B}_t = 2a - B_t$, then

$$P(M_t > a, B_t \leq a) = P(T_a \leq t, 2a - B_t \geq a) = P(T_a \leq t, \tilde{B}_t \geq a) = P(\tilde{B}_t \geq a),$$

since $\{T_a > t, \tilde{B}_t \geq a\} = \{T_a > t, B_t \geq a\} = \emptyset$.

(d) Thanks to part (a), (b) and (c) we have that

$$P(M_t > a) = P(B_t > a) + P(\tilde{B}_t \geq a) = 2 \int_a^\infty \frac{1}{\sqrt{2\pi t}} e^{-x^2/2t} dx = \int_a^\infty \sqrt{\frac{2}{\pi t}} e^{-x^2/2t} dx.$$

Exercise 3

Let

$$Y_t^n = \frac{1}{\sqrt{n}} S_{[nt]} + \frac{1}{\sqrt{n}} (nt - [nt]) X_{[nt]+1}.$$

By Donkser's theorem it holds that $Y^n \Rightarrow B$ as $n \rightarrow \infty$, where $Y^n = (Y_t^n)_{t \geq 0}$ and $B = (B_t)_{t \geq 0}$ is a B.m. In particular this implies that

$$\sup_{t \in [0,1]} Y_t^n \Rightarrow M_1 = \sup_{t \in [0,1]} B_t \quad (4)$$

as $n \rightarrow \infty$. We notice that

$$\begin{aligned} \sup_{t \in [0,1]} Y_t^n &= \frac{1}{\sqrt{n}} \sup_{t \in [0,1]} (S_{[nt]} + (nt - [nt])X_{[nt]+1}) \\ &= \frac{1}{\sqrt{n}} \sup\{S_i + tX_{i+1}, i = 0, \dots, n-1, t \in [0,1], S_n\} = \frac{1}{\sqrt{n}} \sup_{i=0, \dots, n} S_n = \frac{1}{\sqrt{n}} G_n. \end{aligned}$$

Therefore, (4) tells us that (heuristically), for n large, G_n is distributed as $\sqrt{n}M_1$. Noticing that exercise 2 tells us that $\sqrt{n}M_1 \sim \sqrt{n}|B_1|$, we can (heuristically) say that

$$\mathbb{E}[G_n] \simeq \mathbb{E}[\sqrt{n}|B_1|] = \sqrt{\frac{2}{\pi n}} \int_0^\infty x e^{-x^2/2n} dx = \sqrt{\frac{2n}{\pi}}$$

and

$$\text{Var}(G_n) = \mathbb{E}[G_n^2] - \mathbb{E}[G_n]^2 \simeq n \mathbb{E}[B_1^2] - \frac{2n}{\pi} = n \left(1 - \frac{2}{\pi}\right)$$

for n large.

Exercise 4

(a) We showed in class that

$$Q_t^\sigma(W_t) \xrightarrow{L^2} t \quad \text{as } |\sigma| \rightarrow 0. \quad (5)$$

Hence it is sufficient to notice that

$$\begin{aligned} Q_t^\sigma(aW_t + bt) &= \sum_{i=0}^{n-1} |aW_{t_{i+1}} + bt_{i+1} - aW_{t_i} - bt_i|^2 = \sum_{i=0}^{n-1} |a(W_{t_{i+1}} - W_{t_i}) + b(t_{i+1} - t_i)|^2 \\ &= a^2 Q_t^\sigma(W_t) + ab \sum_{i=0}^{n-1} (W_{t_{i+1}} - W_{t_i})(t_{i+1} - t_i) + b^2 \sum_{i=0}^{n-1} (t_{i+1} - t_i)^2. \end{aligned}$$

Now

$$\mathbb{E} \left[\left(\sum_{i=0}^{n-1} (W_{t_{i+1}} - W_{t_i})(t_{i+1} - t_i) \right)^2 \right] = \sum_{i=0}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^2](t_{i+1} - t_i)^2 = \sum_{i=0}^{n-1} (t_{i+1} - t_i)^4 \leq |\sigma|^3 t,$$

where the first equality holds since $\mathbb{E}[(W_{t_{i+1}} - W_{t_i})(W_{t_{j+1}} - W_{t_j})] = 0$ for every $i \neq j$, and so

$$ab \sum_{i=0}^{n-1} (W_{t_{i+1}} - W_{t_i})(t_{i+1} - t_i) \xrightarrow{L^2} 0 \quad \text{as } |\sigma| \rightarrow 0. \quad (6)$$

Combining (5) and (6) and the fact that $\sum_{i=0}^{n-1} (t_{i+1} - t_i)^2 \leq |\sigma| t \xrightarrow{L^2} 0$ as $|\sigma| \rightarrow 0$, we get that

$$Q_t^\sigma(aW_t + bt) \xrightarrow{L^2} t \quad \text{as } |\sigma| \rightarrow 0.$$

(b) Let's consider

$$Q_{t,4}^\sigma(W_t) = \sum_{i=0}^{n-1} (W_{t_{i+1}} - W_{t_i})^4.$$

We have

$$\mathbb{E}[(Q_{t,4}^\sigma(W_t))^2] = \sum_{i=0}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^8] + \sum_{\substack{i,j=0 \\ i \neq j}}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^4 (W_{t_{j+1}} - W_{t_j})^4]$$

Now

$$\sum_{i=0}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^8] = 105 \sum_{i=0}^{n-1} (t_{i+1} - t_i)^4 \leq 105 |\sigma|^3 t \xrightarrow{L^2} 0 \quad \text{as } |\sigma| \rightarrow 0 \quad (7)$$

and

$$\begin{aligned} \sum_{\substack{i,j=0 \\ i \neq j}}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^4 (W_{t_{j+1}} - W_{t_j})^4] &= \sum_{\substack{i,j=0 \\ i \neq j}}^{n-1} \mathbb{E}[(W_{t_{i+1}} - W_{t_i})^4] \mathbb{E}[(W_{t_{j+1}} - W_{t_j})^4] \\ &= 225 \sum_{i=0}^{n-1} (t_{i+1} - t_i)^2 \sum_{\substack{j=0 \\ j \neq i}}^{n-1} (t_{j+1} - t_j)^2 \\ &\leq 225 \sum_{i=0}^{n-1} (t_{i+1} - t_i)^2 |\sigma| t \\ &\leq 225 |\sigma|^2 t^2 \xrightarrow{L^2} 0 \quad \text{as } |\sigma| \rightarrow 0. \end{aligned} \quad (8)$$

Hence (7) and (8) imply that

$$Q_{t,4}^\sigma(W_t) \xrightarrow{L^2} 0 \quad \text{as } |\sigma| \rightarrow 0.$$

Exercise 5

Let us denote by X_t^x the solution to $\dot{X}_t = a(X_t)$, $X_0 = x$. If $f \in C_b^2(\mathbb{R})$, then

$$(\mathcal{L}f)(x) = \lim_{t \rightarrow 0^+} \frac{f(X_t^x) - f(x)}{t} = \frac{d}{dt} f(X_t^x) \Big|_{t=0} = f'(X_0^x) \frac{d}{dt} X_t^x \Big|_{t=0} = f'(x) a(X_0^x) = a(x) f'(x),$$

i.e. $\mathcal{L} = a \partial_x$.

Exercise 6

Given $f \in C_b^2$, we have, for $x \in K$, where $K \subset \mathbb{R}$ is a compact set such that $\text{supp}\{p(\cdot, t|\cdot, 0)\} \subseteq K^2$,

$$(\mathcal{L}f)(x) = \lim_{t \rightarrow 0^+} \frac{\mathbb{E}_x[f(X_t)] - f(x)}{t} = \lim_{t \rightarrow 0^+} \frac{1}{t} \int_K (f(y) - f(x)) p(y, t|x, 0) dy. \quad (9)$$

If $\varepsilon > 0$ we can split the integral in (9) as

$$\begin{aligned} t^{-1} \int_K (f(y) - f(x)) p(y, t|x, 0) dy &= t^{-1} \int_{y \in K, |x-y| \geq \varepsilon} (f(y) - f(x)) p(y, t|x, 0) dy \\ &\quad + t^{-1} f'(x) \int_{y \in K, |x-y| < \varepsilon} (y - x) p(y, t|x, 0) dy \\ &\quad + t^{-1} / 2 \int_{y \in K, |x-y| < \varepsilon} f''(\xi_y) (y - x)^2 p(y, t|x, 0) dy, \end{aligned}$$

where ξ_y is a point between x and y . Now,

$$t^{-1} \int_{y \in K, |x-y| \geq \varepsilon} (f(y) - f(x)) p(y, t|x, 0) dy \xrightarrow[t \rightarrow 0^+]{} \int_{y \in K, |x-y| \geq \varepsilon} (f(y) - f(x)) W(y|x) dy + O(\varepsilon),$$

$$t^{-1} f'(x) \int_{y \in K, |x-y| < \varepsilon} (y - x) p(y, t|x, 0) dy \xrightarrow[t \rightarrow 0^+]{} f'(x) O(\varepsilon),$$

$$\begin{aligned} t^{-1}/2 \int_{y \in K, |x-y| < \varepsilon} f''(\xi_y) (y - x)^2 p(y, t|x, 0) dy &\leq \\ &\leq \|f''\|_{\infty} t^{-1}/2 \int_{y \in K, |x-y| < \varepsilon} (y - x)^2 p(y, t|x, 0) dy \xrightarrow[t \rightarrow 0^+]{} \|f''\|_{\infty} O(\varepsilon)/2. \end{aligned}$$

hence, if we send $\varepsilon \rightarrow 0^+$, we get

$$(\mathcal{L}f)(x) = \int_K (f(y) - f(x)) W(y|x) dy = \int_{\mathbb{R}} f(y) W(y|x) dy - f(x) \lambda(x).$$