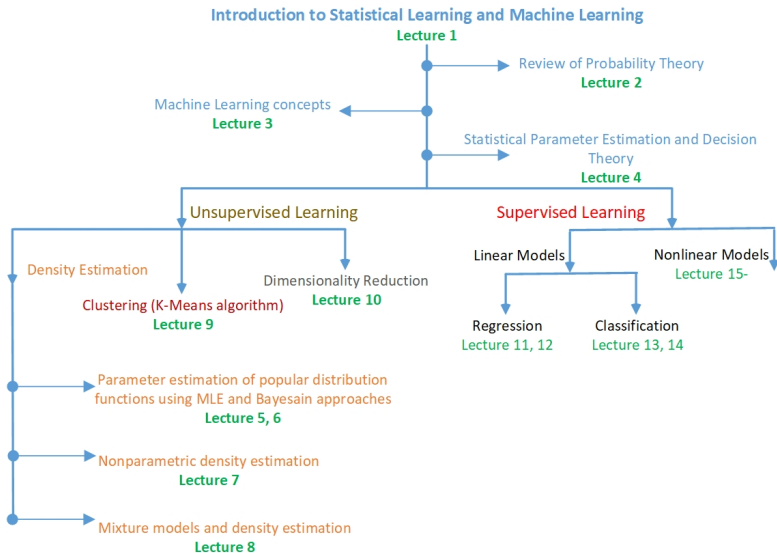


Statistical Learning and Machine Learning

Lecture 7 - Probability Distributions 3

September 18, 2021

Course overview and where do we stand



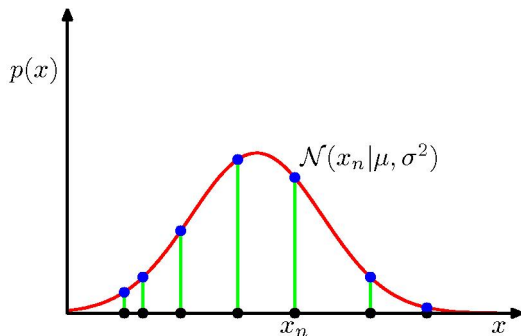
Objectives of the lecture

- Introduction to Gaussian mixture models
- Nonparametric density estimation techniques
 - ① Histogram method
 - ② Kernel density estimators
 - ③ Nearest-neighbour method



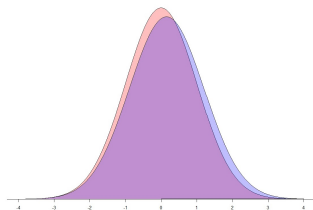
Density Estimation I

- Suppose that we have a set of data points $\{x_1, \dots, x_N\}$ which correspond to (random) measurements of a process.
- Next, assume that the process is governed by the underlying probability density function $p(x)$. The value of $p(x)$ indicates how probable is to obtain the measurement x .



Density Estimation II

- The problem of estimating the probability distribution $p(x)$ given a set of data points x_1, \dots, x_N is called **density estimation**.
- The density estimation problem is **ill-posed** as there are infinitely many probability distributions that could give rise to the specific set of data points $\{x_1, \dots, x_N\}$.



- To solve that problem, we take an *informed decision* regarding the probability distribution that generated the data $\{x_1, \dots, x_N\}$.

- **Example:** A widely used probability distribution in machine Learning is the **Gaussian distribution**: $p(x) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$ The Gaussian distribution, denoted by $\mathcal{N}(x|\mu, \sigma^2)$, is uniquely described by two parameters, the μ and the σ .
- We call distributions which are uniquely defined by a small set of parameters, as **parametric distributions**.

Frequentist (MLE) vs Bayesian Approach

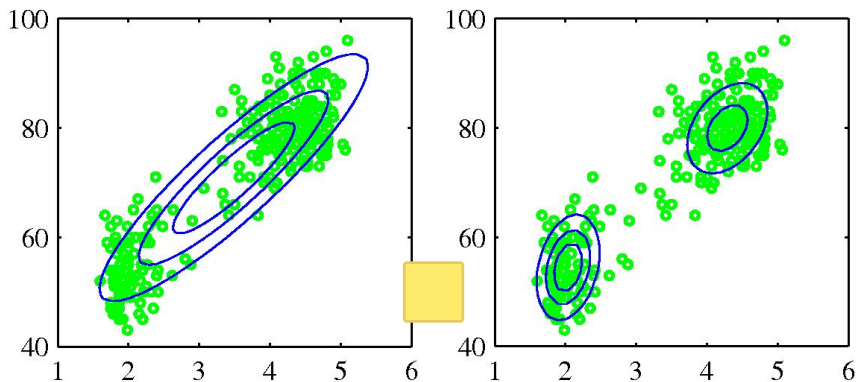
Two different ways to estimate the value of the parameter of the density function $p(x)$, e.g., μ

- **Maximum Likelihood Estimation:** Maximizing the (log-)likelihood function of a set of data points $\mathcal{D} = \{x_1, \dots, x_N\}$ as a function of parameter(s) (μ); it is assumed that \mathcal{D} correspond to the IID data.
- **Bayesian Approach:** Defining a prior distribution $p(\mu)$ for parameter (μ), and obtaining posterior distribution by using: posterior \propto likelihood \times prior

Note that the likelihood function plays a key role in both the approaches!


Mixture of Gaussian

Multi-modal data can be better described with mixtures of distributions rather than a single distribution.



Mixture of Gaussians

When using a superposition of K Gaussians, the resulting distribution has the form:


$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

and is called a **mixture of Gaussians**.

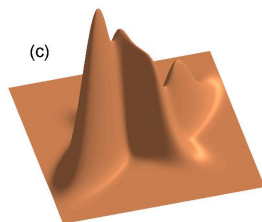
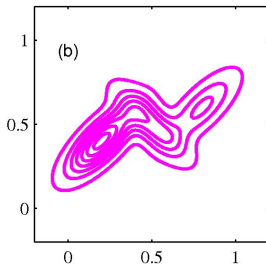
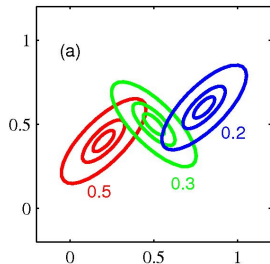
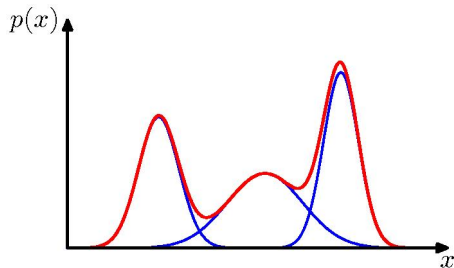
The parameters π_k are called **mixing coefficients** and satisfy:

$$0 \leq \pi_k \leq 1, \quad \sum_{k=1}^K \pi_k = 1.$$

Thus, the mixing coefficients have the form of probabilities ($\pi_k = p(k)$).

Mixture of Gaussians

The resulting distribution can take complex forms:



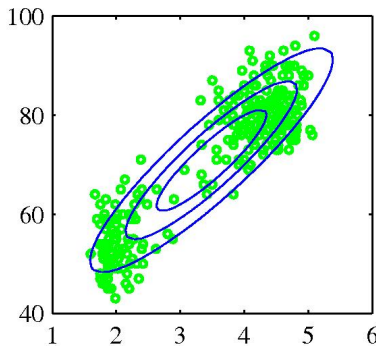
Parameter Estimation in Mixture of Gaussian

The form of the Gaussian mixture is governed by the parameters: π , μ_k and Σ .

- ① How can maximum likelihood estimation procedure be applied to obtain the above parameters, given N IID data points $\{\mathbf{x}_1, \dots, \mathbf{x}_n\}$?
- ② What problems do you see in this approach?

Nonparametric Density Estimation

- **Nonparametric density estimation** approaches make very few assumptions about the form of the distribution
- **Parametric methods**: efficient (dependence on few parameters) but not flexible (e.g., poor performance on multimodal data)
- **Nonparametric methods**: no assumptions on data hence are flexible; may be computationally expensive.



Histogram Method

For 1-dimensional data, we can use **histogram density models**. Standard histograms are obtained by:

- partitioning \mathbb{R} into distinct bins of width Δ_i
- counting the number of data points n_i falling in bin i
- normalizing using N and Δ_i to obtain a probability density:

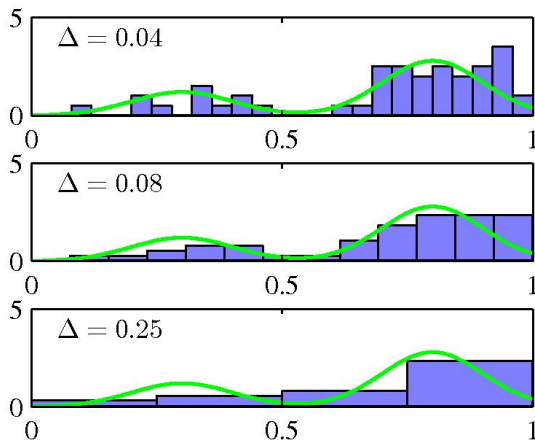
$$p_i = \frac{n_i}{N\Delta_i}.$$

This gives a model for the density $p(x)$ that is constant over the width of each bin (piece wise constant model).

Histogram Method

Bin-width as smoothing parameter

Histogram-based density estimation of the distribution following the green line using $N = 50$ data points for a varying number of (equal-length) bins.



Histogram Method: Advantages and disadvantages

- Advantages

- ① Provides quick visualization of data (as approximate density estimate)
- ② Suitable for large data: data can be discarded once histogram is computed
- ③ amenable for sequential processing

- Disadvantages

- ① Discontinuity at bin edges
- ② Curse of dimensionality

Foundations for nonparametric density estimation methods

Let us assume that observations are drawn from an unknown probability density $p(\mathbf{x})$ in some D -dimensional space $\mathbf{x} \in \mathbb{R}^D$. We wish to estimate $p(\mathbf{x})$ from data.

Consider a small region \mathcal{R} containing \mathbf{x} . The probability mass within \mathcal{R} is $P = \int_{\mathcal{R}} p(\mathbf{x}) d\mathbf{x}$. Suppose we collect N observations drawn from $p(\mathbf{x})$. Since each data point has probability P of being within \mathcal{R} , the total number K of points that lie within \mathcal{R} follow binomial distribution.

- For $N \rightarrow \infty$ data points drawn from $p(\mathbf{x})$, the approximate number of data points falling inside \mathcal{R} is NP .
- For small \mathcal{R} , we can assume that $p(\mathbf{x})$ is constant over \mathcal{R} leading to $P \sim p(\mathbf{x})V$ for $\mathbf{x} \in \mathcal{R}$.
- Combining the above two cases we have:

$$p(\mathbf{x}) = \frac{K}{NV}. \quad (3)$$

Nonparametric density estimation



Two approaches for nonparametric density estimation based on (3)

- **Kernel density estimation**: Fix V and determine K from the data
- **K nearest-neighbour**: Fix K and determine V from the data

Kernel Density Estimation

Let \mathcal{R} be a small hypercube centered at the point \mathbf{x} . To count the number K of points inside \mathcal{R} we define the function:

$$k(\mathbf{u}) = \begin{cases} 1, & |u_i| \leq 1/2, \quad i = 1, \dots, D \\ 0, & \text{otherwise} \end{cases}$$

which expresses a hypercube centered at the origin. $k(\mathbf{u})$ is called **kernel function** (or *Parzen window*).

The quantity $k\left(\frac{\mathbf{x} - \mathbf{x}_n}{h}\right)$ will be 1 if the data point \mathbf{x}_n lies inside cube of side h centered at \mathbf{x} .

The number of points inside the hypercube of side h centered at \mathbf{x} :

$$K = \sum_{n=1}^N k\left(\frac{\mathbf{x} - \mathbf{x}_n}{h}\right)$$

Kernel Density Estimation

Remember that $p(\mathbf{x}) = \frac{K}{NV}$. Thus by substituting $V = h^D$ and K from the above:

$$p(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N \frac{1}{h^D} k\left(\frac{\mathbf{x} - \mathbf{x}_n}{h}\right).$$

Reinterpret the above not as single hypercube centered at \mathbf{x} but N hypercubes centered at \mathbf{x}_n .

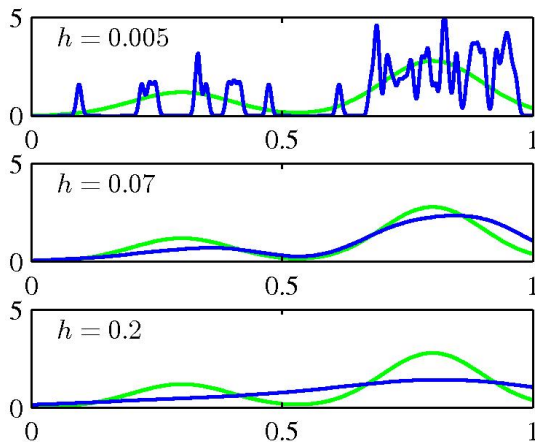
To avoid artificial discontinuities (at the borders of the hypercube), we can use a **Gaussian kernel** instead:

$$p(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N \frac{1}{(2\pi h^2)^{1/2}} \exp\left\{-\frac{\|\mathbf{x} - \mathbf{x}_n\|^2}{2h^2}\right\}$$

where h is the standard deviation of the Gaussian components.

Kernel Density Estimation

Illustration of Kernel density model applied on real data; note the role of h as smoothing parameter



Nearest-neighbour methods

In kernel density estimators, the use of a fixed parameter h causes problems

- for regions with high data density, we need h to be small; large h will lead to over-smoothing
- but for regions with low density of points, small h will lead to noisy/spikier estimates
- h should depend on data (adaptive)

Nearest-neighbour method fixes the number of neighbours K (instead of fixing V). Then:

$$p(\mathbf{x}) = \frac{K}{NV} \quad (4)$$

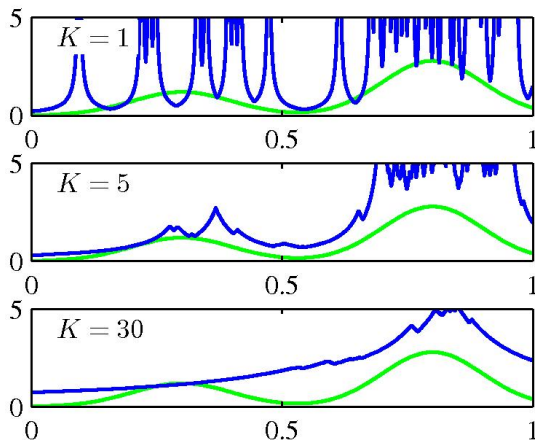
where V corresponds to the volume of the space around \mathbf{x} containing K neighbours.

Operation:

- Consider a small sphere around \mathbf{x} where we need to calculate $p(\mathbf{x})$
- Allow the radius of the sphere to grow until it contains precisely K data points
- The estimate of $p(\mathbf{x})$ is then given by (4) with V set to the volume of the sphere

Nearest-neighbour methods

Example: Illustration of Nearest-neighbour density model applied on real data.



K-NN method for classification

Let us assume that we have a dataset formed by N_k points in class \mathcal{C}_k , and $N = \sum_k N_k$ points in total. To classify a new point \mathbf{x} :

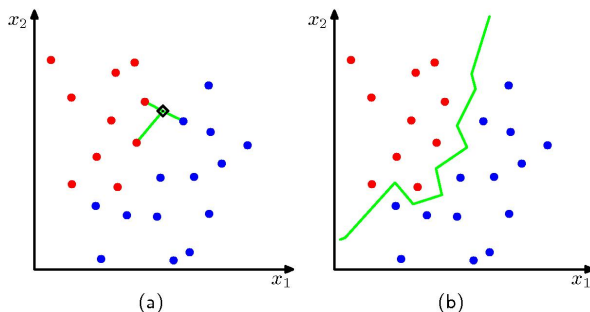
- Draw a sphere centered at \mathbf{x} containing exactly K (labeled) points irrespective of their class
- The sphere has a volume V and contains K_k points from class \mathcal{C}_k
- Then we define the following probabilities:
 - $p(\mathbf{x}|\mathcal{C}_k) = K_k/(N_k V)$
 - $p(\mathbf{x}) = K/(NV)$
 - $p(\mathcal{C}_k) = N_k/N$
- Using the [Bayes formula](#):

$$p(\mathcal{C}_k|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{p(\mathbf{x})} = \frac{K_k}{K}$$

To minimize the [misclassification probability](#), a test point \mathbf{x} is assigned to the class having largest posterior probability.

K-NN method for classification

A test point is assigned according to the majority class membership of the K nearest **training data points**. (left) $K=3$ and (right) $K=1$.



K-NN method for classification

Illustration of K as controlling the degree of smoothness of the estimated density

