

Robust high-dimensional Gaussian and bootstrap approximations for trimmed sample means

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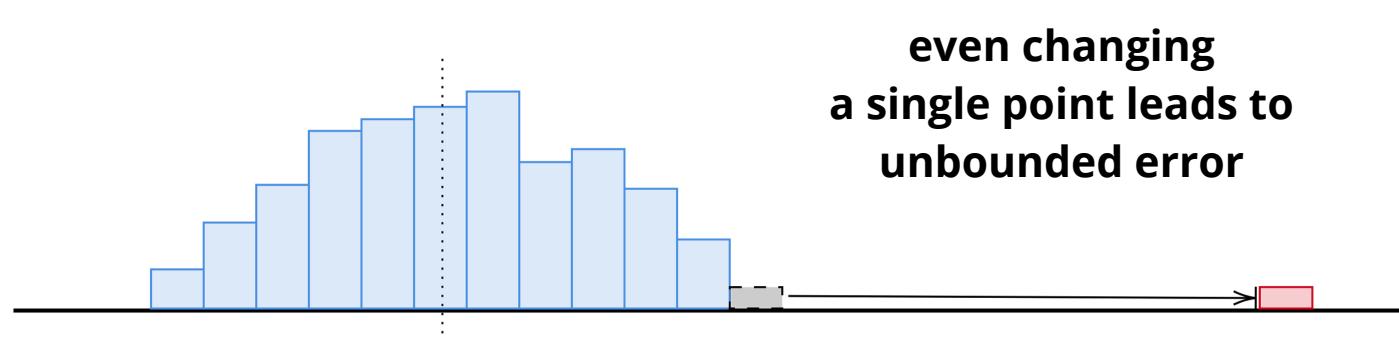
Sample Means are everywhere...

$$\mu \approx \frac{1}{n} \sum_{i=1}^n X_i \quad \Sigma \approx \frac{1}{n} \sum_{i=1}^n (X_i - \mu)(X_i - \mu)^T$$

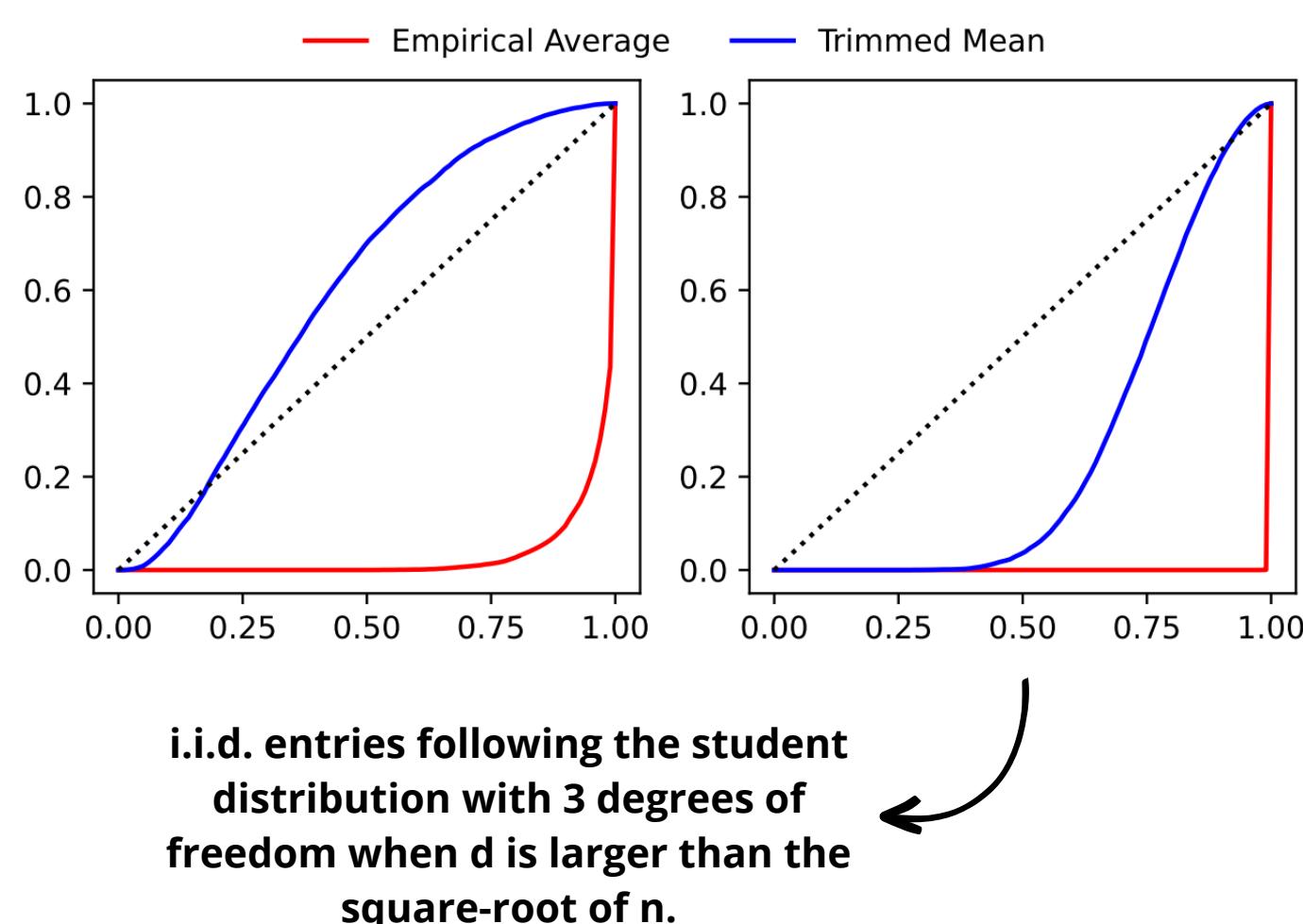
$$\hat{\theta}_n \in \operatorname{argmin}_{\theta \in \Theta} \frac{1}{n} \sum_{i=1}^n l(Z_i, \theta)$$

M-estimators

... but they are **not robust** against **contamination** or **outliers**...



... and **their Gaussian and bootstrap approximations scale poorly in high-dimensions** [1].



We want to **replace the empirical average** to deal with heavy-tailed distributions in a high-dimensional setup. We also want to deal with contamination.

We can use **trimmed sample means** on the corrupted sample. And for a given family of d functions

$$f \in \mathcal{F}, f : \mathcal{X} \rightarrow \mathbb{R}$$

we can show that

$$T_{n,k}^\varepsilon = \frac{1}{n-2k} \sum_{i=k}^{n-k} X_{(i)}^\varepsilon$$

discard the least k values and the highest k

order statistics

Why is it relevant?

- Gaussian and bootstrap approximations are a cornerstone on the **construction of confidence intervals**;
- Our method works even when the dimension is exponential on the sample size. This is the scenario where the number of feature far exceeds the sample size (e.g. **genetic and financial data** [3]). A scenario where the empirical mean is infeasible.
- Finite-sample Gaussian approximations are also useful on causal inference. For instance, it is useful to identify subgroups where treatments may have effect [4].

Theorem 1 (informal). The cutoff k can be chosen, as a function of n and d , to satisfy

$$\sup_{\lambda \in \mathbb{R}} \left| \mathbb{P} \left[\sup_{f \in \mathcal{F}} \sqrt{n} \left(\hat{T}_{n,k}(f, X_{1:n}) - Pf \right) \leq \lambda \right] - \mathbb{P} \left[\sup_{f \in \mathcal{F}} G_P f \leq \lambda \right] \right| \leq \varrho$$

with

the error given by the trimmed mean

the Gaussian limit

$$\varrho \leq C \left(\nu_p \vee \nu_p^{\frac{1}{2}} \right) \left(\frac{\ln^{6-\frac{4}{p}}(nd)}{n^{\frac{2p-4}{2p-1}}} \right)^{\frac{1}{4}} + 15 \frac{\nu_p}{\sigma_{\mathcal{F}, P}} \varepsilon n^{\frac{1}{2} + \frac{3}{4p-2}} \ln^{\frac{1}{2} - \frac{1}{p}}(nd)$$

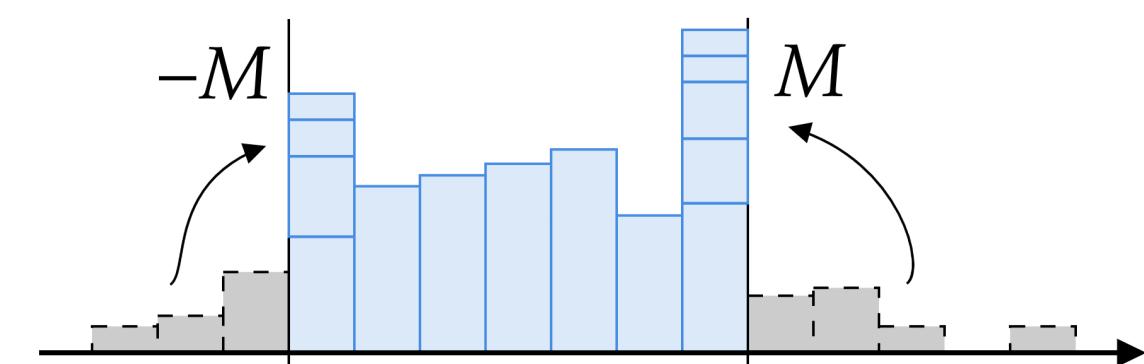
allows for d exponential on n (as is the case of the empirical mean on light-tails)

the number of contaminated sample points can grow almost as the square of n .

Similar results **also apply for infinite dimension** under some regularity conditions. Using these infinite dimension results it was also possible to obtain optimal bounds for the problem of vector mean estimation under arbitrary norms.

REFERENCES

- [1] Anders Bredahl Kock and David Preinerstorfer (2024). A remark on moment-dependent phase transitions in high-dimensional gaussian approximations. *Statistics & Probability Letters*.
- [2] Roberto I. Oliveira and Lucas Resende (2023). Trimmed sample means for robust uniform mean estimation and regression. *arXiv preprint*.
- [3] Victor Chernozhukov, Denis Chetverikov, Kengo Kato, and Yuta Koike (2023). High-dimensional data bootstrap. *Annual Review of Statistics and Its Application*.
- [4] Xinzhou Guo and Xuming He (2021). Inference on selected subgroups in clinical trials, *JASA*.



And a bound of the same order also holds for the bootstrap approximation, where one bounds

$$\sup_{\lambda \in \mathbb{R}} \left| \mathbb{P} \left[\sup_{f \in \mathcal{F}} \sqrt{n} \left(\hat{T}_{n,k}(f, \tilde{X}_{1:n}) - \hat{T}_{n,k}(f, X_{1:n}) \right) \leq \lambda \right] - \mathbb{P} \left[\sup_{f \in \mathcal{F}} G_P f \leq \lambda \right] \right| \leq \varrho$$

