

Competing under Information Heterogeneity: Evidence from Auto Insurance

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- Firms increasingly differ in *information precision* (data access/analytics) and in *cost structures*.
- This creates information asymmetries *between* firms (beyond classic buyer–seller asymmetry).
- Policy interest: regulations that equalize or share consumer risk information (e.g., centralized “risk bureau”, Open Banking policies).

- Research Questions
 - How does heterogeneous information across insurers shape market equilibrium?
 - What are the equilibrium impacts of establishing a centralized bureau to equalize information access?
- Contributions
 - A tractable model of imperfect competition with firm-specific information precision and costs.
 - New identification/estimation strategy using offered-price distributions and demand to recover signals.
 - Evidence from Italian auto liability insurance with rich panel linking consumers across insurers.
 - Counterfactuals: centralized risk bureau, full information, and privacy/high-variance restrictions.

1. Selection markets

- Competitive models: Einav et al. (2010, 2011), Azevedo & Gottlieb (2017); Asymmetric information: Cabral et al. (2018), Crawford et al. (2018), Cuesta et al. (2021) and Tebaldi (2024)
→ Incorporates multidimensional cost heterogeneity

2. Demand estimation

- Berry (1994), BLP(1995) and D'Haultfœuille et al. (2019)
→ Extends demand estimation when prices are not observed.

3. Antitrust/Consumer protection and Big Data

- Einav et al. (2013), Chatterjee et al. (2023), Blattner & Nelson (2021), Lam & Liu (2020), Jin & Wagman (2021)
→ Policies that equalize information access can improve competition

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Descriptive evidence

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Institutional Background: Italian Auto Liability (RCA)

- Mandatory, annual, exclusive contracts; insurers cannot reject consumers.
- Large market: ≈ 31 M contracts in 2018; ≈ 50 national competitors.
- Key contract features widely standardized; little use of deductibles.

- Nationally representative matched insurer–insuree panel with claims frequency/severity, premiums, coverage.
- Tracks policyholders across insurers and time \Rightarrow measure risk using ex-post claims panel.
- Focus sample: new customers in Rome (2013–2021); top 10 firms + fringe group.

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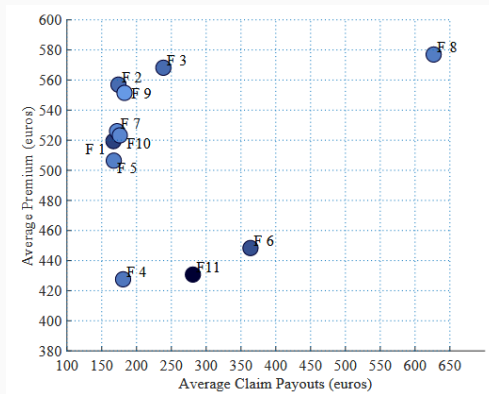
Sample & Summary Statistics

- $N \approx 124,428$ contracts; avg premium ≈ 478 ; within-year claim rate ≈ 0.08 .
- Demographics/vehicle: 56% male; avg age 48; BM class ≈ 2 ; car age ≈ 8.3 years.

Table 1: Summary statistics

Variables	Mean	Std. Dev.	Min	Max	N
Premium (€)	477.68	208.79	133.68	1335.05	124,428
Claim size (€)	260.89	10217.58	0	2521014	124,428
No. of claims (within contract year)	0.08	0.29	0	4	124,428
No. of accidents in last 5 years	0.81	1.22	0	3	124,428
BM class	2.06	2.51	1	15	124,428
Age	48.24	14.11	18	99	124,428
Man	0.56	0.50	0	1	124,428
Median city	0.10	0.30	0	1	124,428
Big city	0.62	0.49	0	1	124,428
Car age	8.30	5.27	0	19	124,428
Horsepower	66.88	26.84	0	493	124,428
Petrol vehicle	0.52	0.50	0	1	124,428
One installment	0.67	0.47	0	1	124,428

Stylized Facts: Price Variation & Sorting



- Large cross-firm variation in average premiums even at similar average risks/market shares.
- Firms with higher average claim costs attract riskier consumers \Rightarrow sorting across firms.

1. Step 1: Construct individual risk measures

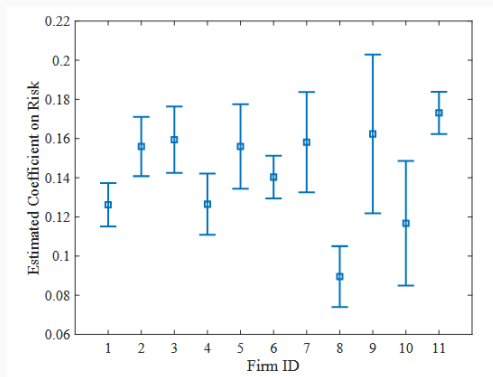
- Panel regression of claim counts and size with individual fixed effects
- Risk measure = $\underbrace{\exp(X_{it}\delta_c + \zeta_i)}_{\text{Claim Count}} \cdot \underbrace{\exp(\delta_0 + X_{it} + \eta_{it})}_{\text{Claim Size}}$

2. Step 2: Firm-specific premium-risk regressions

- For each firm j : $\text{Premium}_{ij} = \alpha_j + \beta_j \cdot \text{Risk}_i + \varepsilon_{ij}$
- Higher β_j suggests firm's prices are more responsive to actual risk

Heterogeneity in Risk Sensitivity

- Strong cross-firm differences in premium–risk slopes \Rightarrow heterogeneous precision?
- Prices are equilibrium outcomes \Rightarrow need structural model to recover information precision.



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- J insurers; standardized product; no outside option.
- Consumer true risk θ (expected cost/year) unobserved ex ante.
- Firm j observes a private signal $\hat{\theta}_j$ with precision that differs across firms.
- D denotes the contract chosen by the consumer.

$$\hat{\theta}_j \sim \mathcal{N}(\theta, \sigma_j^2), \quad \text{iid} \tag{1}$$

- Denote by $\phi(\hat{\theta}_j | \theta, \sigma_j)$ the density
- Lower $\sigma_j^2 \Rightarrow$ higher information precision for firm j .
- Signals are used to form posterior beliefs about θ *conditional on selection*.

$$E(\theta | \hat{\theta}_j, D = j) = \int_{\theta} \theta f(\theta | \hat{\theta}_j, D = j) d\theta = \frac{\int_{\theta} \theta \overbrace{\Pr(D = j | \hat{\theta}_j, \theta)}^{\text{Selection Prob.}} \phi(\hat{\theta}_j; \theta, \sigma_j) f_0(\theta) d\theta}{\int_{\theta} \Pr(D = j | \hat{\theta}_j, \theta) \phi(\hat{\theta}_j; \theta, \sigma_j) f_0(\theta) d\theta} \tag{2}$$

- Assumes:

$$p_j(\hat{\theta}_j) = \alpha_j + \beta_j \mathbb{E}[\theta \mid \hat{\theta}_j, D = j], \quad (3)$$

- α_j : baseline markup; β_j : sensitivity to risk rating.
- $\mathbb{E}[\theta \mid \hat{\theta}_j, D = j]$ embeds selection \Rightarrow nonlinearity in $\hat{\theta}_j$.

- Consumers choose one insurer (no outside option); utility depends on price and **observable** characteristics.
- Preference parameters allowed to vary with **observables** and risk type. Consumer i 's utility from firm j :

$$U_{ij} = -\gamma(\theta)p_j(\tilde{\theta}_j) + \xi_j(\theta) + \varepsilon_{ij} \quad (4)$$

$$\Pr(D = j | \hat{\theta}, \theta) = \frac{\exp(-\gamma(\theta)p_j(\hat{\theta}_j) + \xi_j(\theta))}{\sum_{j'=1}^J \exp(-\gamma(\theta)p_{j'}(\hat{\theta}_{j'}) + \xi_{j'}(\theta))} \quad (5)$$

Firm profits

- Firms simultaneously choose pricing coefficients (α_j, β_j) to maximize expected profits, given common knowledge of all firms' primitives (e.g., signal distributions, costs) and $f_0(\theta)$.
- Firm j 's profit

$$\pi_j(\alpha, \beta) = \int_{\hat{\theta}} \int_{\theta} \underbrace{(p_j(\hat{\theta}_j) + c_j - k_j\theta)}_{\text{net profit}} \underbrace{\Pr(D = j|\hat{\theta})}_{\text{choice prob.}} \underbrace{\left(\prod_{j'=1}^J \phi(\hat{\theta}_{j'}; \theta, \sigma_{j'}) \right)}_{\text{signal dist.}} \underbrace{f_0(\theta)}_{\text{type dist.}} d\theta d\hat{\theta}.$$

- c_j : “net benefits” of contracting with a customer irrespective of the risk.
- k_j : efficiency at processing claims.

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Estimation Overview

- Joint distribution of premium and risk types [▶ Details](#)
 - Uses panel of claim records
 - Claim size estimated via regression
 - Accident rate identified from panel of accident counts (repeated noisy measurements)
 - Output $\hat{g}(p|\theta, D = j)$
- Demand $(\gamma(\theta), \xi_j(\theta))$
 - $\xi_j(\theta)$: matched to observed market shares
 - $\gamma(\theta)$: identified from sorting patterns
- Supply $(\alpha_j, \beta_j, \sigma_j^2, c_j, k_j)$
 - Average price conditional on risk identifies (α_j, β_j) [▶ Details](#)
 - σ_j is identified from price dispersion given risk [▶ Details](#)
 - Cost parameters (c_j, k_j) recovered from first-order conditions [▶ Details](#)

Demand Estimation

Challenge:

- Observe **accepted prices**, not offered prices
- Lower prices over-represented in data

Solution:

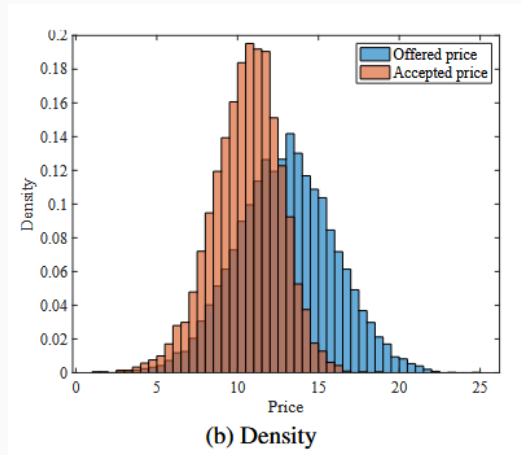
- Use selection rule (demand model) to recover distribution of offered prices

Relationship:

$$g(p|\theta, D=j) \propto g_j(p|\theta) \Pr(D=j|p_j=p, \theta)$$

Recover offered prices:

$$g_j(p|\theta) = \frac{g(p|\theta, D=j) / \Pr(D=j|p_j=p, \theta)}{\int_{p'} g(p|\theta, D=j) / \Pr(D=j|p_j=p', \theta) dp'}$$



Iterative Algorithm for Demand Estimation

Nested Fixed-Point Algorithm: Jointly estimate sorting probabilities, price distributions, and demand parameters

Inner Loop: For fixed γ (price sensitivity)

1. Initialize: $\xi_j^1 = 0$ and $\Pr^1(D = j | p_j = p, \theta) = \exp(\alpha p)$ for all j
2. Update offered-price density, using Bayes' rule:

$$g_j^r(p | \theta) = \frac{\hat{g}(p | \theta, D = j) / \Pr^r(D = j | p_j = p, \theta)}{\int_{p'} \hat{g}(p' | \theta, D = j) / \Pr^r(D = j | p_j = p', \theta) dp'}$$

3. Update ξ^{r+1} to match observed market shares:

$$\hat{s}_j = \int_{\theta} \int_p \frac{\exp(\alpha p_j + \xi_j^{r+1})}{\sum_{j'} \exp(\alpha p_{j'} + \xi_{j'}^{r+1})} \left(\prod_{j'} g_{j'}^r(p_{j'} | \theta) \right) f_0(\theta) dp d\theta$$

4. Update selection probabilities $\Pr^{r+1}(D = j | p_j = p, \theta)$ using demand model
5. Iterate until g^r converges. Output $\{g_j(p | \theta; \gamma)\}_j$, $\Pr(D = j | p, \theta; \gamma)$, and $\xi(\gamma)$.

Outer Loop: Estimate α by matching risk sorting patterns via maximum likelihood

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Results: Supply-Side

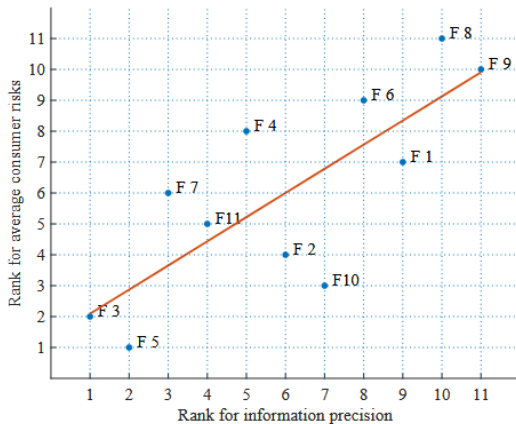
- Large differences in information precision (σ_j^2) and price sensitivity (β_j) across firms
- Baseline markups (α_j) differ, consistent with market power from information advantages
- Large contracting net benefits (e.g., for F8 is four times average premium)
- Heterogeneity in claim efficiency (e.g. F8 roughly 2.5 times less efficient than F6)

Table 3: Estimates of supply-side parameters

Firm ID	Pricing Coefficients		Signal Std. Dev.	Net Benefits	Claim Efficiency
	α_j	β_j	σ_j	c_j	k_j
1	-342.22 (48.56)	1.72 (0.10)	1339.05 (56.12)	1165.31 (346.43)	1.90 (0.34)
2	-333.44 (80.47)	1.81 (0.17)	1217.08 (83.71)	1087.50 (349.50)	1.98 (0.37)
3	-163.18 (43.31)	1.65 (0.10)	1053.16 (72.66)	1110.60 (280.35)	2.29 (0.26)
4	-315.64 (58.59)	1.45 (0.12)	1178.77 (72.82)	1034.33 (237.37)	1.60 (0.20)
5	-194.72 (61.41)	1.65 (0.16)	1117.57 (85.78)	922.39 (290.10)	1.86 (0.31)
6	-310.08 (43.47)	1.47 (0.09)	1301.49 (60.30)	943.08 (273.87)	1.37 (0.24)
7	-220.93 (90.81)	1.50 (0.19)	1118.80 (115.56)	858.29 (293.74)	1.54 (0.33)
8	-1404.66 (252.91)	3.00 (0.39)	1580.41 (119.79)	2132.07 (371.91)	3.16 (0.49)
9	-688.85 (312.62)	2.15 (0.57)	1637.52 (172.54)	1440.84 (424.11)	2.45 (0.74)
10	-246.68 (138.25)	1.59 (0.28)	1245.79 (107.35)	972.73 (317.01)	1.79 (0.39)

Results: supply-side

- More(less) precise firms, attract less(more) risky consumers



- **Centralized Risk Bureau:** aggregate firms' signals (weighted by precision), share equally with all.

$$E(\theta|\hat{\theta}) = \int_{\theta} \theta f(\theta|\hat{\theta}) d\theta = \frac{\int_{\theta} \theta \left(\prod_j \phi(\hat{\theta}_j; \theta, \sigma_j) \right) f_0(\theta) d\theta}{\int_{\theta} \left(\prod_j \phi(\hat{\theta}_j; \theta, \sigma_j) \right) f_0(\theta) d\theta} \quad (6)$$

- **Full Information Benchmark:** firms observe true θ (eliminate information asymmetry).
- **Privacy/Restriction:** firms can only use basic information; set σ_j^2 to the worst observed.

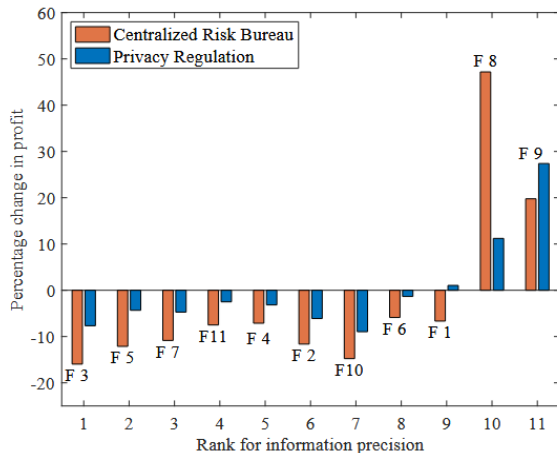
Counterfactuals: results

- Eliminating information asymmetries helps consumers.
- More(less) info. benefits less(more) risky consumers, due to less(more) cross-subsidies.

	Baseline	Observing True Risk	Centralized Risk Bureau	Privacy Regulation
Average CS (€)	-542.83	-451.05 (+16.91%)	-457.59 (+15.70%)	-523.42 (+3.57%)
Average CS: Low risk (€)	-477.24	153.74 (+132.21)	-103.63 (+78.28%)	-480.25 (-0.63%)
Average CS: High risk (€)	-608.42	-1055.83 (-73.54%)	-811.55 (-33.39%)	-566.60 (+6.87%)
Average premium (€)	461.25	342.56 (-25.73%)	361.61 (-21.60%)	441.79 (-4.22%)
Average profit (€)	849.58	802.11 (-5.59%)	799.29 (-5.92%)	834.38 (-1.79%)
HHI	2297.26	2241.40 (-2.42%)	2264.29 (-1.44%)	2303.38 (+0.27%)

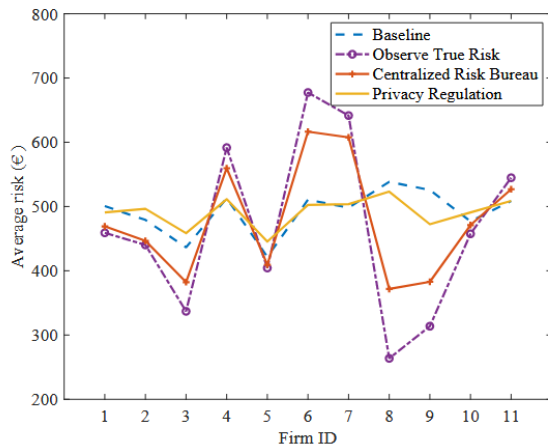
Change in firm profits

- Equalizing information weakens market power of precise firms.
- Common risk evaluation \Rightarrow more effective undercutting \Rightarrow stronger price competition.
- Perfect correlation in signals (uniform risk evaluation) leads to higher price competition.



Who goes where? Sorting patterns

- With equal access to risk, firms more efficient at processing claims re-target higher-risk consumers.
- Sorting shifts from info advantages (baseline) to cost specialization (bureau).
→ increases efficiency, avg cost ↓ by ~3.7%
- Sorting is lowest under privacy regulation.



- This paper develops empirical framework for studying competition under information asymmetries
- This paper studies the impacts of a credit bureau, it finds:
 - Centralized information can lower prices, raise consumer surplus and reorient sorting toward cost efficiency.
 - Distributional trade-offs: low-risk consumers gain more under information sharing; high-risk under privacy.
 - Industry composition effects: advanced-screening firms lose profits; potential dynamic innovation effects.

- What is the impact of a credit bureau on a dynamic world with learning-switching costs?
 - Cost of risk missclassification is low \rightarrow can update prices.
 - c_j should be risk specific to capture inertia?
- What is a credit bureau doing in the model
 - Pooling Data: if there are economies of scale, what is the market failure?
 - Sharing algorithms: impact of investment in algorithms.

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Identification of Risk Types

- Risk type $\theta_i = \underbrace{\mu}_{\text{claim size}} \underbrace{\lambda_i}_{\text{accident rate}}$,
- Claim size μ estimated via regression
- Accident rate λ_i (latent variable) identified from panel of accident counts (repeated noisy measurements)
- Observe accident counts y_{it} over T periods: $y_{it} \sim \text{Poisson}(\lambda_i)$
- Joint distribution of counts identifies $f(\lambda_i | p_i, D_i = j)$:

$$f(y_{i1}, y_{i2}, \dots, y_{iT} | p_i, D_i = j) = \int \prod_{t=1}^T \frac{\lambda_i^{y_{it}} e^{-\lambda_i}}{y_{it}!} f(\lambda_i | p_i, D_i = j) d\lambda_i$$

- Output $\hat{g}(p | \theta, D = j)$ for demand estimation [► Go back](#)

- Demand is produced by the model:

$$\Pr(D = j | p_j = p, \theta) = \int_{\mathbf{p}_{-j}} \frac{\exp(-\gamma p + \xi_j)}{\exp(-\gamma p + \xi_j) + \sum_{j' \neq j} \exp(-\gamma p_{j'} + \xi_{j'})} \left(\prod_{j' \neq j} g_{j'}(p_{j'} | \theta) \right) d\mathbf{p}_{-j}.$$

► Go back

Identification of Pricing Coefficients (α_j, β_j)

- Pricing strategy: $p_j(\hat{\theta}_j) = \alpha_j + \beta_j E(\theta | \hat{\theta}_j, D = j)$
- **Key Insight:** Linear structure allows identification from first and second moments
- **Objects already recovered:** within-firm offered-price density $g_j(p | \theta)$

Within-firm moments:

$$\text{Mean premium: } E(p | D = j) = \alpha_j + \beta_j E(E(\theta | \hat{\theta}_j, D = j) | D = j) = \alpha_j + \beta_j E(\theta | D = j)$$

$$\text{Premium and Risk covariance: } \text{cov}(p, \theta | D = j) = \beta_j \text{var}(E(\theta | \hat{\theta}_j, D = j) | D = j) = \frac{\text{var}(p | D = j)}{\beta_j}$$

Solution: Solve system of two linear equations

$$\beta_j = \frac{\text{var}(p | D = j)}{\text{cov}(p, \theta | D = j)}, \quad \alpha_j = E(p | D = j) - \beta_j E(\theta | D = j),$$

Advantage: estimated separately from other parameters \rightarrow reduces computational burden [▶ Go back](#)

Identification of Signal Variance (σ_j)

- **Key Insight:** monotonicity of prices on the signal (similar to GPV(2000))

1. Using monotonicity and then inverting:

$$G_j(p_j(\hat{\theta}_j)|\theta) = \Phi\left(\frac{\hat{\theta}_j - \theta}{\sigma_j}\right) \implies p_j^o(\hat{\theta}_j; \theta, \sigma_j) = G_j^{-1}\left(\Phi\left(\frac{\hat{\theta}_j - \theta}{\sigma_j}\right)\right)$$

where $G_j(p|\theta)$ = CDF of offered prices (recovered from demand estimation)

2. For a given σ_j , evaluate risk rating at equilibrium prices:

$$E(\theta|\hat{\theta}_j, D = j; \sigma_j) = \frac{\int_{\theta} \theta \Pr(D = j | p_j^o(\hat{\theta}_j; \theta, \sigma_j), \theta) \phi(\hat{\theta}_j; \theta, \sigma_j) f_0(\theta) d\theta}{\int_{\theta} \Pr(D = j | p_j^o(\hat{\theta}_j; \theta, \sigma_j), \theta) \phi(\hat{\theta}_j; \theta, \sigma_j) f_0(\theta) d\theta}$$

3. Using $p_j(\hat{\theta}_j; \sigma_j) = \hat{\alpha}_j + \hat{\beta}_j E(\theta|\hat{\theta}_j, D = j; \sigma_j)$ obtain the model-implied joint distribution of (p, θ) within firm j
4. Match model implied dist. of (p, θ) to the empirical dist. recovered from the data.

Supply parameters: c_j, k_j

Net profit from servicing consumer: $p_j(\hat{\theta}_j) + c_j - k_j\theta$

- c_j = net benefit from contracting (inertia, dynamic pricing, cross-selling)
- $k_j\theta$ = claim processing cost (efficiency parameter k_j times expected claims)

First-Order Conditions: Firm j optimally chooses (α_j, β_j)

$$\frac{\partial \pi_j}{\partial \alpha_j} = \int_{\theta} \int_{\hat{\theta}} \Pr(D = j \mid \hat{\theta}, \theta) f(\hat{\theta} \mid \theta) f_0(\theta) d\hat{\theta} d\theta + \int_{\theta} \int_{\hat{\theta}} (\alpha_j + \beta_j \theta + c_j - k_j \theta) \frac{\partial \Pr(D = j \mid \hat{\theta}, \theta)}{\partial \alpha_j} f(\hat{\theta} \mid \theta) f_0(\theta) d\hat{\theta} d\theta,$$

$$\frac{\partial \pi_j}{\partial \beta_j} = \int_{\theta} \int_{\hat{\theta}} \theta \Pr(D = j \mid \hat{\theta}, \theta) f(\hat{\theta} \mid \theta) f_0(\theta) d\hat{\theta} d\theta + \int_{\theta} \int_{\hat{\theta}} (\alpha_j + \beta_j \theta + c_j - k_j \theta) \frac{\partial \Pr(D = j \mid \hat{\theta}, \theta)}{\partial \beta_j} f(\hat{\theta} \mid \theta) f_0(\theta) d\hat{\theta} d\theta.$$

Solution: Given previously recovered $\{g_j, E(\theta \mid \hat{\theta}_j, D = j), \alpha_j, \beta_j, \sigma_j\}$ and $\Pr(D = j \mid \cdot)$, it is a system of 2 linear equations in (c_j, k_j) where all other terms are known [▶ Go back](#)