Cognition Paper: 1a

8/12/2020

1 Analysis 1a RERUN

All the code is in src/1a_RERUN_analysis.R.

1.1 Hypotheses

- H1_ATT: For the ATTRIBUTE-items there is a significant difference between TNC (thick negative concepts) and TPC (thick positive concepts), such that Contradiction ratings are higher for TNC than for TPC.
- H1_BEH: For the BEHAVIOUR-items there is a significant difference between TNC (thick negative concepts) and TPC (thick positive concepts), such that Contradiction ratings are higher for TNC than for TPC.
- H_aux1: Semantic entailment (SE), for which average cancellability ratings should be significantly above the midpoint.
- H_aux2: Conversational implicatures (CI), for which average cancellability ratings should be significantly below the midpoint

Based on .05-alpha level: $H1_ATT$ has to be rejected. $H1_BEH$ cannot be rejected. H_aux1 cannot be rejected. H_aux2 cannot be rejected.

1.2 Data Distribution

Fig. 1 (next page) shows the distributions of cancellability ratings per item.

Fig. 2 (next page) shows the same ratings aggregated per item group and polarity.

We already see that the group means between pooled TPC and TNC are most probably significant.

1.3 Testing for ANOVA assumptions

1.3.1 Homogeneity of variance

We performed a classical Levene's test based on the absolute deviations from the mean with correction factor and an alpha trimming of .25:

Test Statistic = 20.061, p-value = 1.665e-12

The H0 has to be rejected and we can assume significant differences between the group variances in the population. This ANOVA assumption is violated.

1.a) Distribution of Contradiction Ratings per Item

The distributions are futher divided by sentiment polarity of the TC (<NA> for CI and SE) and the item group (ATT, BEH, CI, SE).

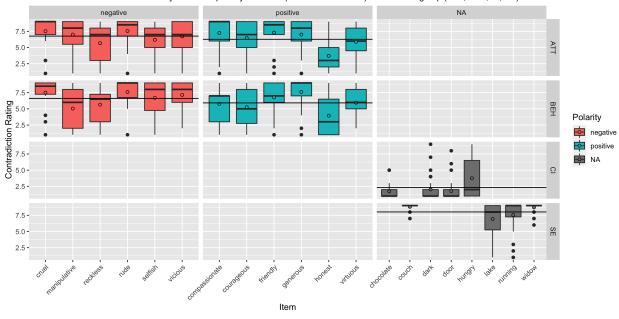


Figure 1: per item

1.a) Distribution of Contradiction Ratings per Group and Polarity

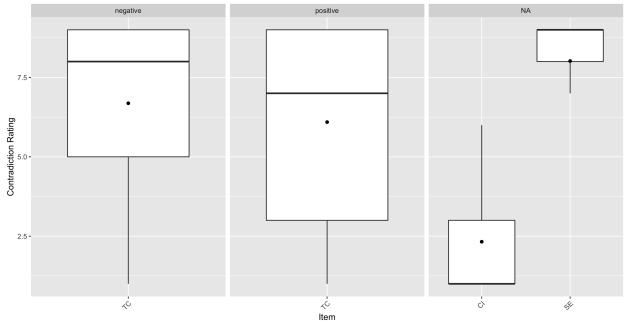


Figure 2: per item

1.3.2 Normality

To test whether the cancellability ratings are normally distributed, we performed a Shapiro-Wilk normality test, both globally and per group. For the global distribution we get a test statistic of

```
W = 0.8174, p-value < 2.2e-16
```

This means that the cancellability ratings are not normally distributed. The same holds true on group level (results omitted for brevity, see 199-100 in src/1a_RERUN_analysis.R). This ANOVA assumption is also violated.

Since both ANOVA assumptions are violated, we will use non-parametric alternatives.

1.4 Hypothesis testing

Disclaimer: There were no outliers in the overall distribution. Consequently outlier-subsetting was not necessary.

1.4.1 H1_ATT & H1_BEH: two- and one-tailed

For H1_ATT and H1_BEH we performed planned contrasts between TNC and TPC based on a mixed model which includes the subject ID as random effect (= within-subject design) as well as controlling for item order (additional random effect). Since we are interested in the contrasts between different TC polarities (pos. -neg.) by item description (ATT / BEH), the mixed model is based on the following formula:

```
value ~ group*polarity + (1|id) + (1|item_step)
```

where group is the item description (ATT / BEH) and polarity codes for TNC and TPC, and (1|id) is the random effect for the within-subjects design and (1|item_step) controls for the order effect.

These are the results for the two-tailed hypothesis:

```
group = ATT:
polarity emmean
                    SE
                         df lower.CL upper.CL
negative
            6.77 0.264 51.3
                                 6.24
                                          7.30
            6.27 0.264 51.4
                                          6.80
positive
                                 5.74
group = BEH:
                         df lower.CL upper.CL
polarity emmean
                    SE
            6.62 0.265 51.0
                                 6.08
                                          7.15
negative
                                          6.46
            5.93 0.264 51.3
                                 5.39
positive
Degrees-of-freedom method: kenward-roger
Confidence level used: 0.95
$contrasts
group = ATT:
                                  SE df t.ratio p.value
contrast
                     estimate
negative - positive
                        0.501 0.297 362 1.687
                                                 0.0924
group = BEH:
 contrast
                     estimate
                                 SE df t.ratio p.value
negative - positive
                        0.690 0.298 365 2.313
                                                 0.0213
Degrees-of-freedom method: kenward-roger
```

Both our main hypotheses are formulated as one-tailed hypotheses, so that TPC < TNC, on average. To test this we simply use the left-tailed p-values:

```
group = ATT:
```

```
contrast estimate SE df t.ratio p.value plus minus midpt positive - negative -0.501 0.297 362 -1.687 0.0462 positive negative 1.5 positive - negative -0.501 0.297 362 -1.687 0.0462 negative positive 1.5
```

group = BEH:

```
contrast estimate SE df t.ratio p.value plus minus midpt positive - negative -0.690 0.298 365 -2.313 0.0106 positive negative 1.5 positive - negative -0.690 0.298 365 -2.313 0.0106 negative positive 1.5
```

Degrees-of-freedom method: kenward-roger

P values are left-tailed

H1_ATT: For H1_ATT, the difference in group means between TNC and TPC show that TNC (6.77) have a higher mean than TPC (6.27). The difference (0.501) is not significant on the two-tailed approach. Our directed hypothesis, however, i.e. TPC < TNC, cannot be rejected, since the left-tailed p-value of 0.0462 is significant. Consequently, H1_ATT cannot be rejected.

H1_BEH: For H1_BEH, the difference in group means between TNC and TPC also show that TNC (6.62) have a higher mean than TPC (5.93). In contrast to the ATT-group, the difference within the BEH-group (0.690) is indeed significant for the two-tailed test. The one-tailed test is also significant, so that we cannot reject TPC < TNC. Thus, H1_BEH cannot be rejected.

Pooled data:

For the pooled data we also see significantly higher contradiction ratings for TNC (6.69) than for TPC (6.10) in the two-sided test:

```
polarity emmean SE df lower.CL upper.CL negative 6.69 0.219 24.2 6.24 7.14 positive 6.10 0.219 24.1 5.64 6.55
```

Degrees-of-freedom method: kenward-roger

Confidence level used: 0.95

\$contrasts

```
contrast estimate SE df t.ratio p.value negative - positive 0.596 0.21 365 2.838 0.0048
```

Degrees-of-freedom method: kenward-roger

For the one-sided test we can halve the p-values accordingly:

```
contrast estimate SE df t.ratio p.value plus minus midpt positive - negative -0.596 0.21 365 -2.838 0.0024 positive negative 1.5 positive - negative -0.596 0.21 365 -2.838 0.0024 negative positive 1.5
```

Degrees-of-freedom method: kenward-roger

P values are left-tailed

1.4.2 H aux1 & H aux2

In order to test H_aux1 and H_aux2 we computed the estimated marginal means (Least-Squares Means) and compared their confidence intervals with the midpoint of the contradiction-scale. The basis for the estimated

means was the following mixed model:

```
value ~ group + (1|id) + (1|item_step)
```

Since CI and SE both have no sentiment polarity values, we drop this factor (in the formula above we used it in the interaction). The estimated means are thus only computed on group level and only for CI and SE. These are the results:

```
group emmean SE df lower.CL upper.CL CI 2.30 0.250 64.7 1.80 2.80 SE 8.03 0.249 65.4 7.53 8.53
```

Degrees-of-freedom method: kenward-roger

Confidence level used: 0.95

H_aux1: For SE we expect confidence intervals above 5, which is the case (7.53|8.53). Thus SE are on average significantly above the midpoint, which means we cannot reject H_aux1.

H_aux2: For CI we expect confidence intervals below 5, which is also the case (1.80)|2.80). Consequently, CI is on average significantly below the midpoint. Thus, we cannot reject H aux2.