

Teaching material is all online!

- On Minerva <http://minerva.leeds.ac.uk>
- On GitHub
<https://github.com/luisacutillo78/Statistical-Methods-Lecture-Notes>

Resources

- Mathematical Statistics and Data Analysis - 3rd ed. (by J. A. Rice);
- <http://www1.maths.leeds.ac.uk/statistics/R/Rintro.pdf>;
- <https://www.datacamp.com/courses/free-introduction-to-r>.

Where We've Been, Where We're Going

In the previous Lecture

- Random Samples from Normal Distributions
- Socrative Quiz

Today

- Confidence intervals
- examples and exercises at the whiteboard

Why do we need confidence intervals?

EXAMPLE

- Assume we are given a set of data from a normal distribution
- we wish to find a point estimate of the mean μ .
- We have seen that \bar{X} is an obvious candidate

Questions

We also need to know what is the likely error range. What If we had a different set of data? How reliable is our estimate, can we trust it? To within what error bounds? We need some theory, making use of the previous lectures!

Definition

A $100(1 - \alpha)\%$ confidence interval for an unknown parameter θ is defined as the random interval

$$(\hat{\theta}_1, \hat{\theta}_2),$$

where $\hat{\theta}_1 = g_1(\underline{X})$ and $\hat{\theta}_2 = g_2(\underline{X})$ are statistics (random variables) such that

$$p(\hat{\theta}_1 < \theta < \hat{\theta}_2) = 1 - \alpha.$$

Note 1: CI are not unique, since there are infinitely many choices for these random variables.

Note 2: θ is the true parameter value, and is not random. $\hat{\theta}_1 = g_1(\underline{X})$ and $\hat{\theta}_2 = g_2(\underline{X})$ are random variables.

Note 3: Usual value $\alpha = 0.05$; that is, 95% confidence intervals.

Interpretation of a confidence interval

If we have a 95% (i.e., $\alpha = 0.05$) confidence interval for a parameter θ , the interpretation is:

If we do many samplings, and for each observed random sample \underline{x} we construct $(g_1(\underline{x}), g_2(\underline{x}))$, we should expect to have the true value θ within this interval 95% of the times.

Usually statistics $\hat{\theta}_1$ and $\hat{\theta}_2$ are both obtained as a function of a point estimator $\hat{\theta}$ of θ .

CI for μ , σ known, using Z

Recall

A *Z-statistic* is a statistic with a standard normal distribution. The main use of *Z*-statistics stems from the facts that, for a general distribution, the Central Limit Theorem implies asymptotically that

$$\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1),$$

and that the standard normal distribution involves **no unknown parameters**.

Confidence interval for μ with σ^2 known

We can use the *Z*-statistic to calculate a range of plausible values for μ , under the assumption that σ^2 is known!

CI for μ , σ known, using Z

Remembering that $Z \sim N(0, 1)$, choose $z_{\alpha/2}$ such that

$$P(Z \leq z_{\alpha/2}) = 1 - \frac{\alpha}{2} \implies P(-z_{\alpha/2} \leq Z \leq z_{\alpha/2}) = 1 - \alpha.$$

If $Z = \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma}$ as above, then

$$\begin{aligned} P\left(-z_{\alpha/2} \leq \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \leq z_{\alpha/2}\right) &= 1 - \alpha \\ \implies P\left(\bar{X} - \frac{\sigma}{\sqrt{n}} z_{\alpha/2} \leq \mu \leq \bar{X} + \frac{\sigma}{\sqrt{n}} z_{\alpha/2}\right) &= 1 - \alpha. \end{aligned}$$

Hence the $100(1 - \alpha)\%$ confidence interval is

$$\left(\bar{x} - \frac{\sigma}{\sqrt{n}} z_{\alpha/2}, \bar{x} + \frac{\sigma}{\sqrt{n}} z_{\alpha/2}\right).$$

The most common value of α in use is 0.05, in which case

$$z_{\alpha/2} = z_{0.025} = 1.960.$$

CI for μ , σ known, using Z

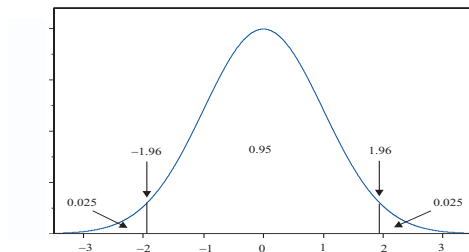


Figure 1 95% interval for $Z \sim N(0, 1)$

Note: We could also go backwards, and try to compute the minimum n in order to ensure that the width of the CI is lower than a maximum threshold.

Whiteboard: Examples 1 and 2

CI for μ , σ unknown, using t

We know that if X_1, X_2, \dots, X_n is iid $N(\mu, \sigma^2)$, then

$$T = \frac{\sqrt{n}(\bar{X} - \mu)}{S} \sim t(n-1)$$

We can now look for a confidence interval by replacing the Z -statistic with the t -statistic. Writing $t_{\alpha/2}(n-1)$ for the $1 - \frac{\alpha}{2}$ quantile from the distribution $t(n-1)$,

$$P\left(-t_{\alpha/2}(n-1) < \frac{\sqrt{n}(\bar{X} - \mu)}{S} < t_{\alpha/2}(n-1)\right) = 1 - \alpha.$$

Re-arranging gives the random interval

$$\left(\bar{X} - \frac{t_{\alpha/2}}{\sqrt{n}} S, \bar{X} + \frac{t_{\alpha/2}}{\sqrt{n}} S\right),$$

and the $100(1 - \alpha)\%$ confidence interval is the realisation of this interval.

σ^2 unknown

If $X_i \sim N(\mu, \sigma^2)$ with both μ and σ^2 unknown:

- 95% CI for μ : $\bar{X} \pm t_{0.975, n-1} \frac{S}{\sqrt{n}}$

$$T \sim t_{n-1} : \quad p(T < t_{0.975, n-1}) = 0.975$$

- 95% CI for σ^2 : $\left(\frac{(n-1)S^2}{\chi_{0.975, n-1}^2}, \frac{(n-1)S^2}{\chi_{0.025, n-1}^2} \right)$

$$Y \sim \chi_{n-1}^2 : \quad p(Y < \chi_{0.975, n-1}^2) = 0.975$$

Whiteboard: Explain why, and Example 3.

Two Sample Problems

We consider two populations $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2)$, with two independent random samples. Thus,

$$\text{Var}[\bar{X}_1 - \bar{X}_2] = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}.$$

We are interested in inferring how μ_1 and μ_2 compare.

- **Two Means:** Consider the two sample means \bar{X}_1 and \bar{X}_2 :

- **If σ_1^2 and σ_2^2 known:** Then, a $100(1 - \alpha)\%$ CI for $\mu_1 - \mu_2$ is

$$(\bar{X}_1 - \bar{X}_2) \pm z_{1-\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}, \quad \text{where } p(Z < z_{1-\frac{\alpha}{2}}) = \left(1 - \frac{\alpha}{2}\right).$$

- **If $\sigma_1^2 = \sigma_2^2$ unknown:** Then, a $100(1 - \alpha)\%$ CI for $\mu_1 - \mu_2$ is

$$(\bar{X}_1 - \bar{X}_2) \pm t_{1-\frac{\alpha}{2}, n_1+n_2-2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}},$$

where $S_p^2 = \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1+n_2-2}$ is the *Pooled Variance*.

Whiteboard: Example 4.