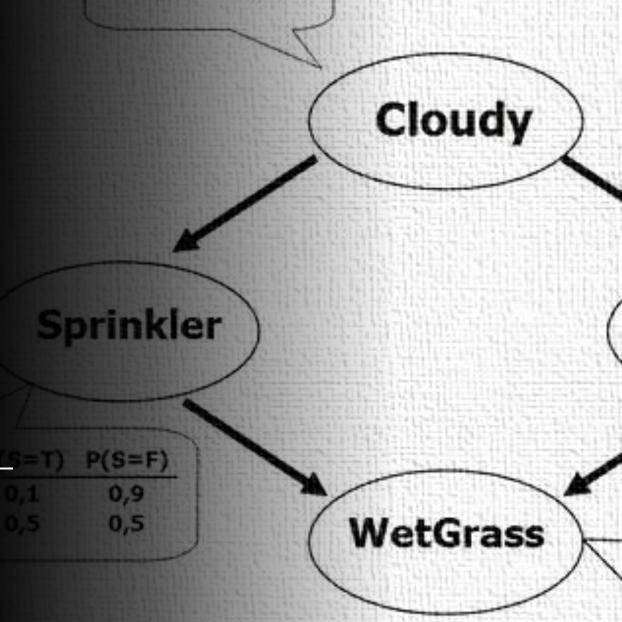
CS 5/7320 Artificial Intelligence

Probabilistic Reasoning (Bayesian networks)

**AIMA Chapter 13** 

Slides by Michael Hahsler based on slides by Svetlana Lazepnik





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## **Probability Recap**

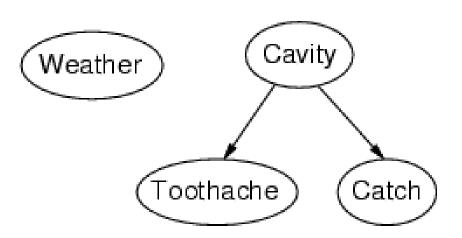
■ Conditional probability 
$$P(x|y) = \frac{P(x,y)}{P(y)} = \alpha P(x,y)$$

- Product rule P(x,y) = P(x|y)P(y)
- Chain rule  $P(X_1, X_2, ... X_n) = P(X_1)P(X_2|X_1)P(X_3|X_1, X_2)...$ =  $\prod_{i=1}^n P(X_i|X_1, ..., X_{i-1})$
- X, Y independent if and only if:  $\forall x, y : P(x,y) = P(x)P(y)$
- $\blacksquare X$  and Y are conditionally independent given Z if and only if:

$$\forall x, y, z : P(x, y|z) = P(x|z)P(y|z)$$
 Written as  $X \perp \!\!\! \perp Y|Z$ 

Notation: P(X = x) = P(x)

Bayesian networks (aka Belief Networks)





A type of graphical model.



A way to specify dependence between random variables.

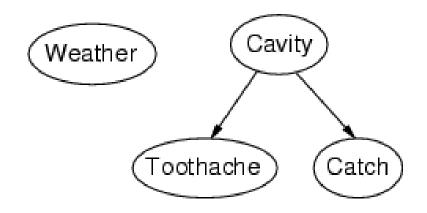


A compact specification of a full joint distributions.

## Structure of Bayesian Networks

#### **Nodes:** Random variables

 Can be assigned (observed) or unassigned (unobserved)

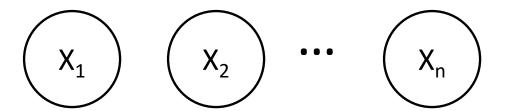


### **Arcs:** Dependencies

- An arrow from one variable to another indicates direct influence.
- Show independence
  - Weather is independent of the other variables (no connection).
  - Toothache and Catch are conditionally independent given Cavity (directed arc).
- Must form a directed acyclic graph (DAG)

## Example: N independent coin flips

Complete independence: no interactions between coin flips

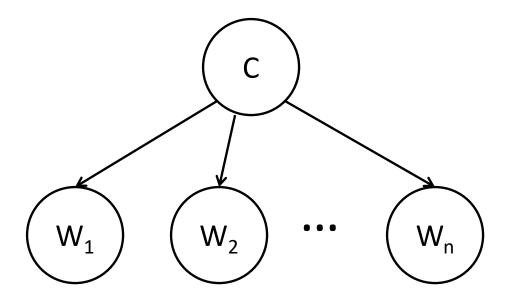


## Example: Naïve Bayes spam filter

#### Random variables:

- C: message class (spam or not spam)
- W<sub>1</sub>, ..., W<sub>n</sub>: presence or absence of words comprising the message

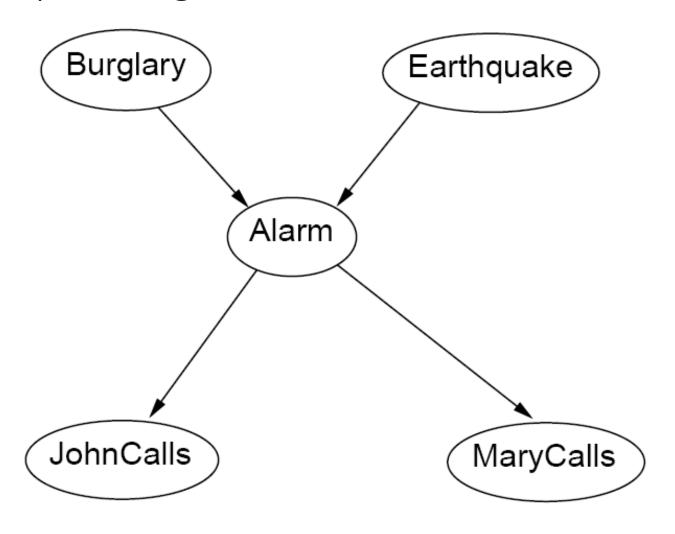
Words depend on the class, but they are modeled conditional independent of each other given the class (= no direct connection between words).



### Example: Burglar Alarm

- **Description**: I have a burglar alarm that is sometimes set off by minor earthquakes. My two neighbors, John and Mary, promised to call me at work if they hear the alarm
- Example inference task: suppose Mary calls and John doesn't call.
   What is the probability of a burglary?
- What are the random variables?
  - Burglary, Earthquake, Alarm, JohnCalls, MaryCalls
- What are the direct influence relationships?
  - A burglar can set off the alarm
  - An earthquake can set off the alarm
  - The alarm can cause Mary to call
  - The alarm can cause John to call

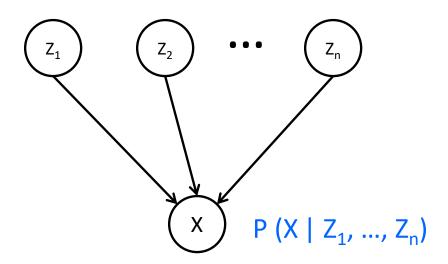
## Example: Burglar Alarm



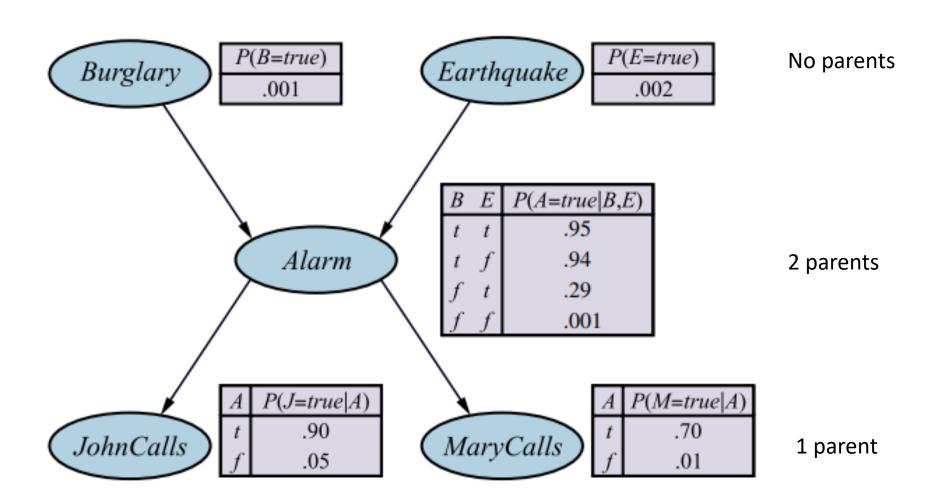
What are the model parameters?

## Parameters: Conditional probability tables

To specify the full joint distribution, we need to specify a conditional distribution for each node given its parents as a conditional probability table (CPT): P (X | Parents(X))



## Example: Burglar Alarm with CPTs



## The joint probability distribution

- For each node X<sub>i</sub>, we know P(X<sub>i</sub> | Parents(X<sub>i</sub>))
- How do we get the full joint distribution  $P(X_1, ..., X_n)$ ?
- Using chain rule:

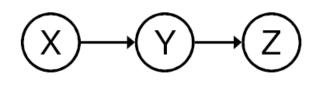
$$P(X_1,...,X_n) = \prod_{i=1}^n P(X_i \mid X_1,...,X_{i-1}) = \prod_{i=1}^n P(X_i \mid Parents(X_i))$$

• Example:

$$P(j, m, a, b, e) = P(b) P(e) P(a | b, e) P(j | a) P(m | a)$$

## Conditional independence

- Key assumption: X is conditionally independent of every non-descendant node given its parents
- Example: causal chain



X: Low pressure

Y: Rain

Z: Traffic

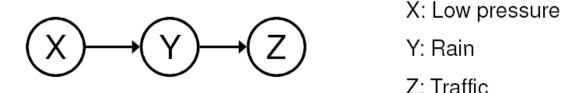
• Are X and Z independent? P(X,Y,Z) = P(X)P(Y|X)P(Z|Y)

Conditioning

$$P(X,Z) = \sum_{y} P(X)P(y|X)P(Z|y)$$
 Marginalize over Y
$$= P(X)\sum_{y} P(Z|y)P(y|X) \neq P(X)P(Z) \Rightarrow X \text{ and Z are not independent!}$$

## Conditional independence

- Key assumption: X is conditionally independent of every non-descendant node given its parents
- Example: causal chain



Is Z independent of X given Y?

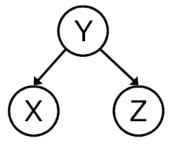
$$P(X,Z|Y) = \frac{P(X,Y,Z)}{P(Y)} = \frac{P(X)P(Y|X)P(Z|Y)}{P(Y)}$$

$$= \frac{P(X)\frac{P(X|Y)P(Y)}{P(X)}P(Z|Y)}{P(Y)}$$
Bayes' rule
$$= \frac{P(X)\frac{P(X|Y)P(Y)}{P(Y)}}{P(Y)}$$
X and Z

= P(X|Y)P(Z|Y) = Definition of conditional independenceX and Z are conditional independence independent given Y

## Conditional independence

Common cause



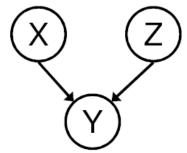
Y: Project due

X: Newsgroup busy

Z: Lab full

- Are X and Z independent?
  - No
- Are they conditionally independent given Y?
  - Yes

Common effect



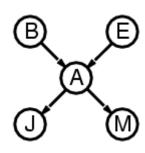
X: Raining

Z: Ballgame

Y: Traffic

- Are X and Z independent?
  - Yes
- Are they conditionally independent given Y?
  - No

## Compactness



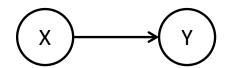
- Suppose we have a Boolean variable X<sub>i</sub> with k Boolean parents. How many rows does its conditional probability table have?
  - $2^k$  rows for all the combinations of parent values, each row requires one number p for  $X_i$  = true
- If each variable has no more than k parents, how many numbers does the complete network require?
  - $O(n \cdot 2^k)$  numbers vs.  $O(2^n)$  for the full joint distribution
- Example: How many nodes for the burglary network?

$$1 + 1 + 4 + 2 + 2 = 10$$
 numbers (vs. specification of the complete joint probability  $2^5-1 = 31$ )

## Constructing Bayesian networks

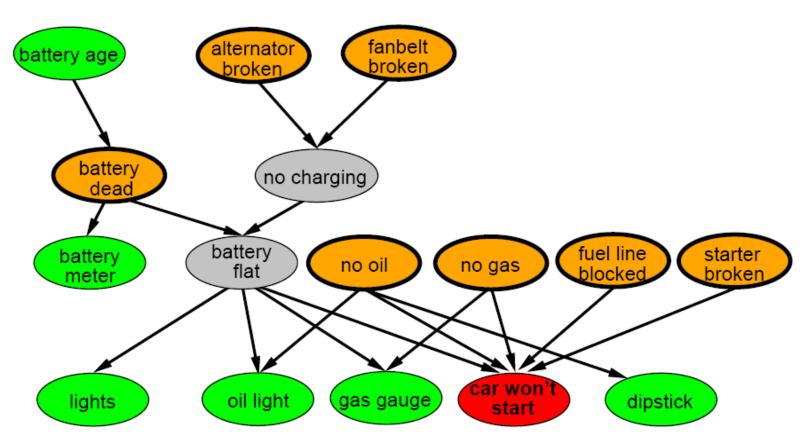
- 1. Choose an ordering of variables X<sub>1</sub>, ..., X<sub>n</sub>
- 2. For i = 1 to n
  - add X<sub>i</sub> to the network
  - select parents from  $X_1, ..., X_{i-1}$  such that  $P(X_i \mid Parents(X_i)) = P(X_i \mid X_1, ..., X_{i-1})$

**Note**: Networks are typically constructed by domain experts with causality in mind. E.g., X causes Y:

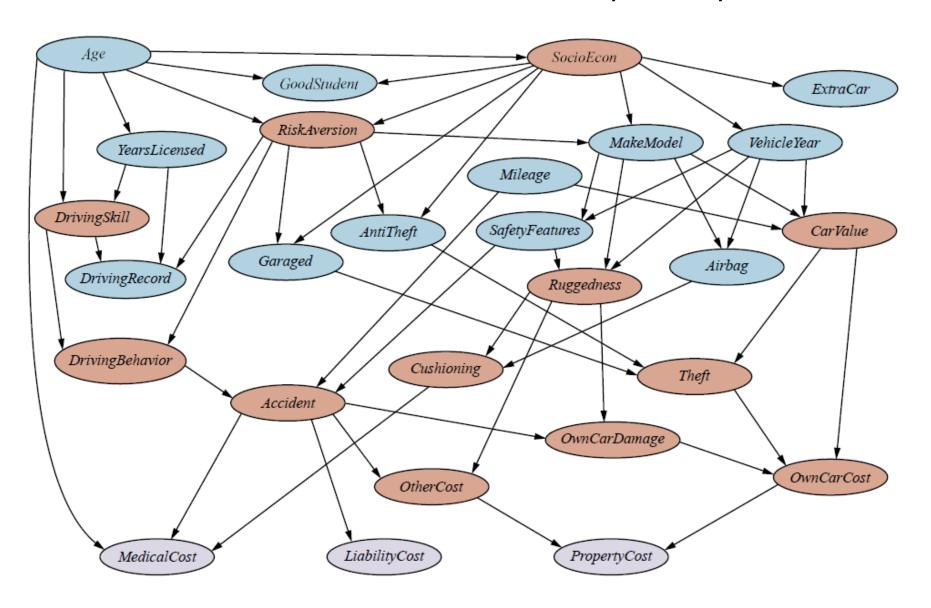


### A more realistic Bayes Network: Car diagnosis

- Initial observation: car won't start
- Green: testable evidence
- Orange: "broken, so fix it" nodes
- Gray: "hidden variables" to ensure sparse structure, reduce parameters

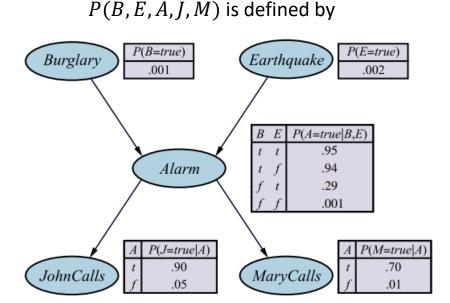


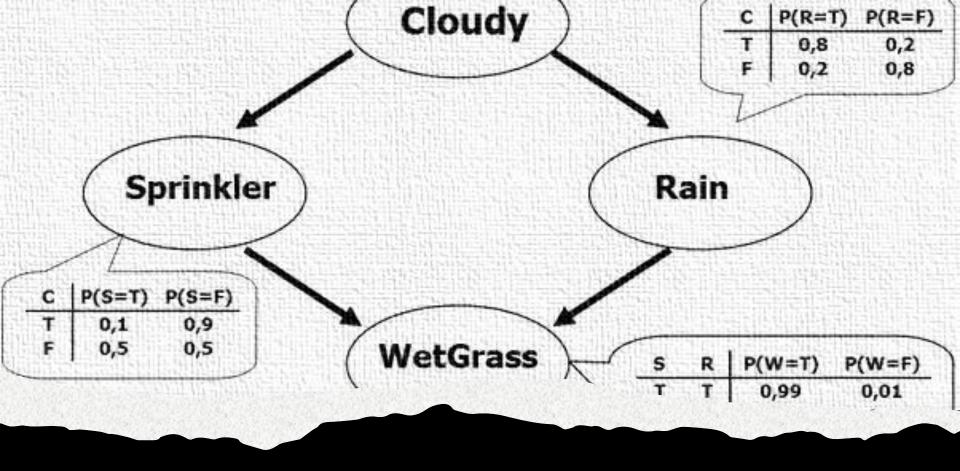
## Car insurance: Cost is affected by many factors



### Summary

- Bayesian networks provide a natural representation for joint probabilities.
- Conditional independence (induced by causality) reduces the number of needed parameters.
- Representation
  - Topology
  - Conditional probability tables
  - Generally easy for domain experts to construct





Inference

Calculate the posterior probability given evidence

### Inference

#### Goal

- Query variables: X
- Evidence (observed) variables: E = e
- Set of unobserved variables: Y
- Calculate the probability of X given e.

If we know the full joint distribution P(X, E, Y), we can infer X by:

$$P(X|E=e) = \frac{P(X,e)}{P(e)} \propto \sum_{y} P(X,e,y)$$

Sum over values of unobservable variables = marginalizing them out.

## Inference: Bayesian network

$$P(X|E=e) = \frac{P(X,e)}{P(e)} \propto \sum_{v} P(X,e,y)$$

#### **Problems**

1. Full joint distributions are too large to store.

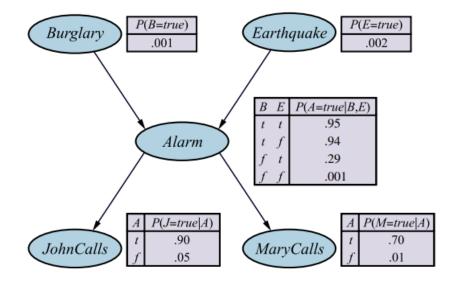
Bayes nets provide significant savings for representing the conditional probability structure.

2. Marginalizing out many unobservable variables Y may involve too many summation terms.

This summation is called **exact inference by enumeration**. Unfortunately, it does not scale well (#p-hard).

In praxis approximate inference by sampling is used.

## Exact inference: Example



Assume we can observe being called. And want to know the probability of a burglary.

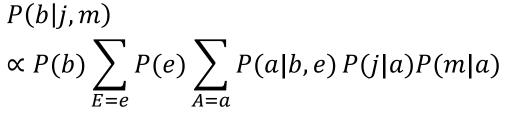
Query: P(B | j, m) with unobservable variables: Earthquake, Alarm

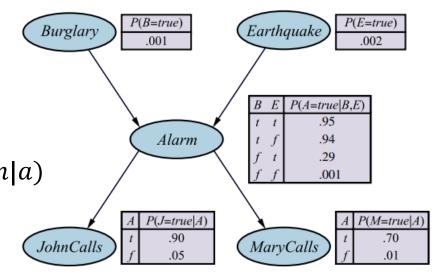
$$P(b|j,m) = \frac{P(b,j,m)}{P(j,m)} \propto \sum_{E=e}^{e} \sum_{A=a}^{e} P(b,e,a,j,m)$$

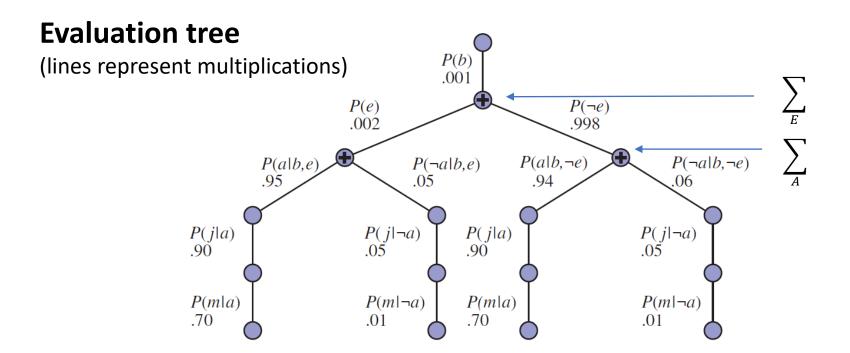
$$= \sum_{E=e}^{e} \sum_{A=a}^{e} P(b)P(e)P(a|b,e)P(j|a)P(m|a)$$

$$= P(b) \sum_{E=e}^{e} P(e) \sum_{A=a}^{e} P(a|b,e)P(j|a)P(m|a)$$









## Approximate inference: Sampling



Bayesian networks can be used as *generative models* 

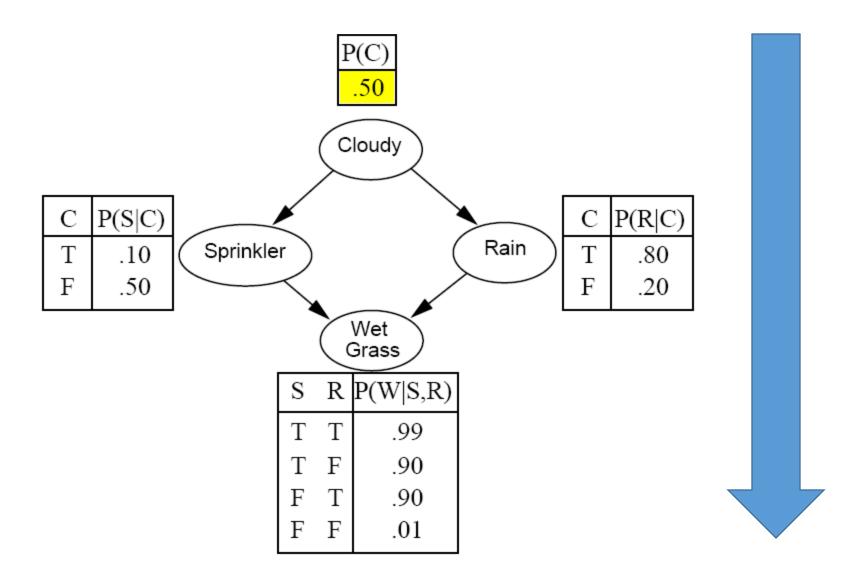
Allows us to efficiently generate samples from the joint distribution

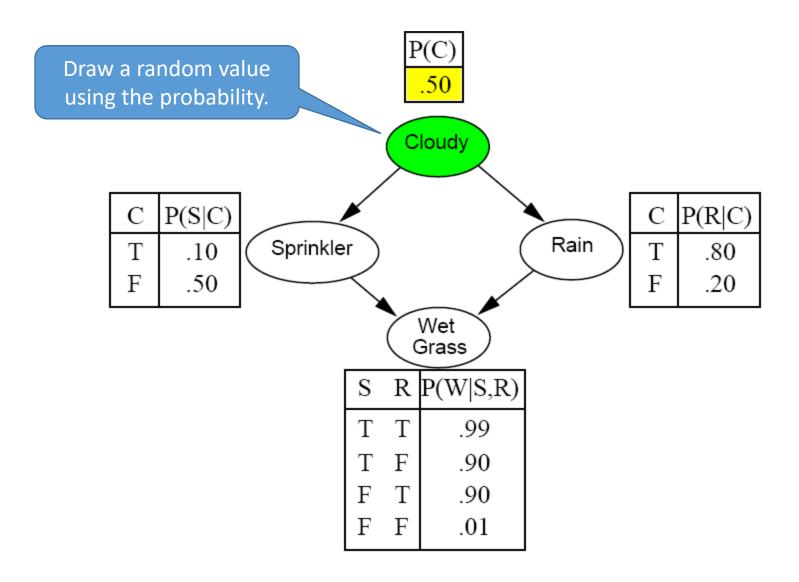


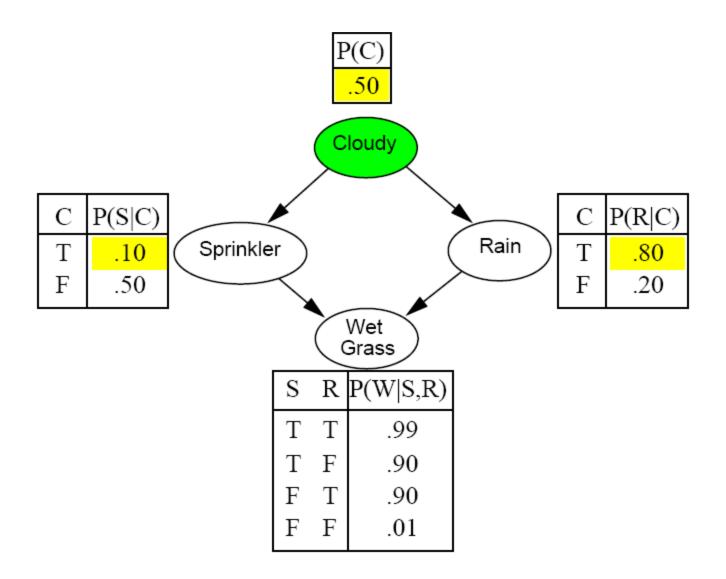
**Idea**: Sample from the network to estimate joint and conditional probability distributions.

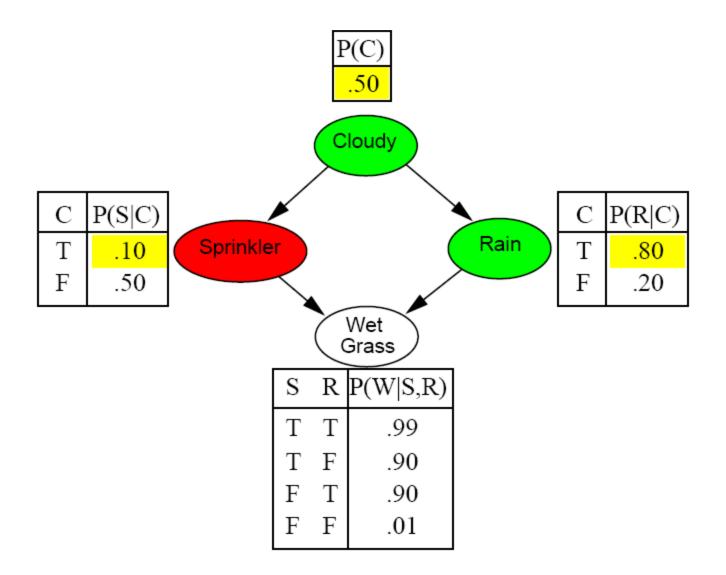
## Prior-Sample Algorithm to Create a Sample (Event)

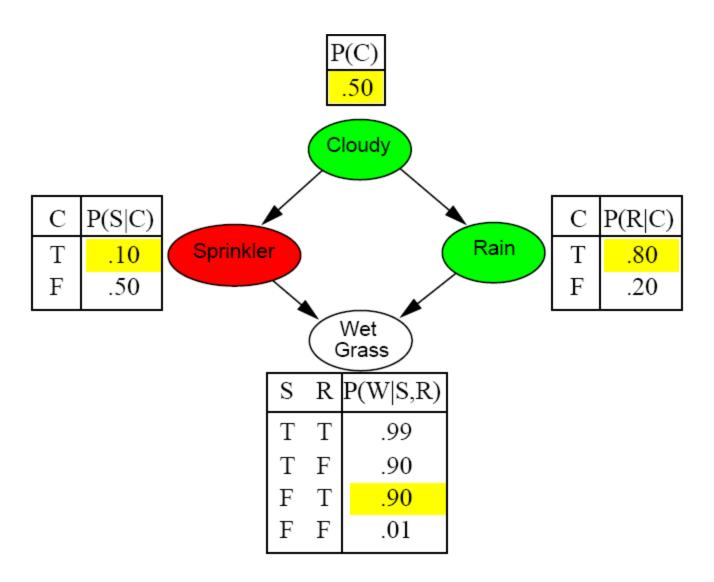
**function** PRIOR-SAMPLE(bn) **returns** an event sampled from the prior specified by bn**inputs**: bn, a Bayesian network specifying joint distribution  $P(X_1, \ldots, X_n)$  $\mathbf{x} \leftarrow$  an event with n elements for each variable  $X_i$  in  $X_1, \ldots, X_n$  do  $\mathbf{x}[i] \leftarrow \text{a random sample from } \mathbf{P}(X_i \mid parents(X_i))$ return x P(C=.5)We need to start with Cloudy the random variables Sprinkler that have no parents. Rain P(S|c).10 WetGrass

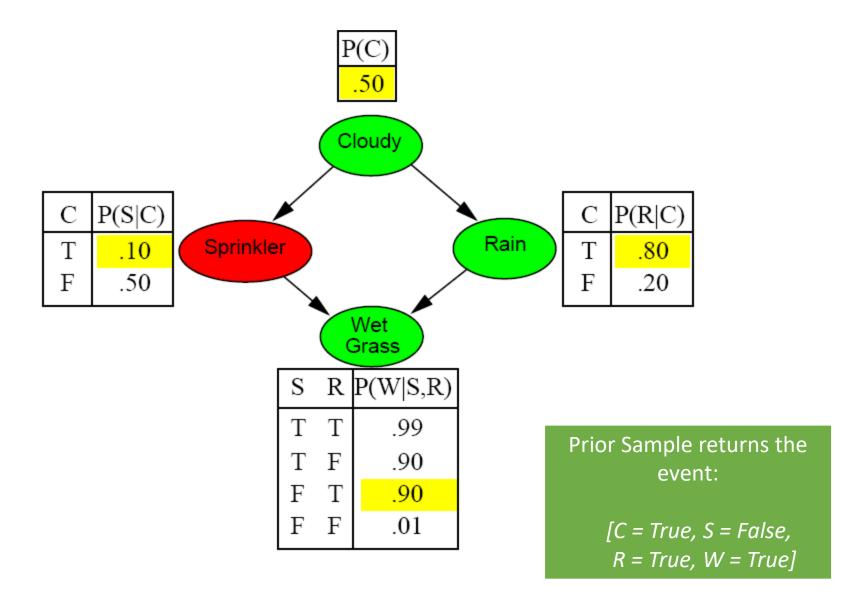












## Estimating the joint probability distribution

Sample *N* times and determine  $N_{PS}(x_1, x_2, ..., x_n)$ , the count of how many times Prior-Sample produces event  $(x_1, x_2, ..., x_n)$ .

$$\hat{P}(x_1, x_2, ..., x_n) = \frac{N_{PS}(x_1, x_2, ..., x_n)}{N}$$

The marginal probability of partially specified event (some x values are known) can also be calculates. E.g.,

$$\widehat{P}(x_1) = \frac{N_{PS}(x_1)}{N}$$

# Estimating conditional probabilities: **Rejection sampling**

Sample N times and ignore the samples that are not consistent with the evidence e.

$$\widehat{P}(X|e) = \alpha N_{PS}(X,e) = \frac{N_{PS}(X,e)}{N_{PS}(e)}$$

**Issue**: What if e is a rare event?

- Example: burglary ∧ earthquake
- Rejection sampling ends up throwing away most of the samples. This is very inefficient!

## Estimating conditional probabilities: **Rejection sampling**

```
function REJECTION-SAMPLING(X, \mathbf{e}, bn, N) returns an estimate of \mathbf{P}(X \mid \mathbf{e})
  inputs: X, the query variable
            e, observed values for variables E
            bn, a Bayesian network
            N, the total number of samples to be generated
  local variables: C, a vector of counts for each value of X, initially zero
  for j = 1 to N do
                                                 We throw away many samples
       \mathbf{x} \leftarrow \mathsf{PRIOR}\text{-}\mathsf{SAMPLE}(bn)
                                                            if e is rare!
       if x is consistent with e then
          C[j] \leftarrow C[j] + 1 where x_j is the value of X in x
  return NORMALIZE(C)
```

## Estimating conditional probabilities: **Importance sampling** (likelihood weighting)

Fix the evidence E = e for sampling and estimate the probably for the non-evidence variables

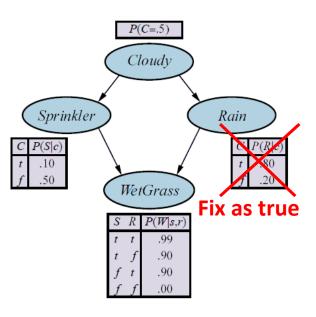
$$Q_{WS}(x)$$

Correct the probabilities using weights  $P(x|e) = w(x)Q_{WS}(x)$ 

Turns out the weights in this case can be easily calculated

$$w(x) = \alpha \prod_{i=1}^{m} P(e_i | parents(E_i))$$

Example: Evidence = it rains



## Estimating conditional probabilities: Markov Chain Monte Carlo Sampling (MCMC)

- Generates a sequence of samples instead of creating each sample individually from scratch.
- Creates new states by making random changes to the current state which forms a Markov Chain and its stationary distribution turns out to be the posteriori distribution of the non-evidence variables.
- Count how often each state is reached and normalize to obtain probability estimates.
- Algorithms:
  - 1. Gibbs sampling
  - 2. Metropolis-Hastings sampling

Note: Simulated annealing is also a MCMC algorithm.

## Gibbs sampling: One variable at a time

```
function GIBBS-ASK(X, \mathbf{e}, bn, N) returns an estimate of \mathbf{P}(X \mid \mathbf{e}) local variables: \mathbf{C}, a vector of counts for each value of X, initially zero \mathbf{Z}, the nonevidence variables in bn \mathbf{x}, the current state of the network, initialized from \mathbf{e} initialize \mathbf{x} with random values for the variables in \mathbf{Z} for k = 1 to N do choose any variable Z_i from \mathbf{Z} according to any distribution \rho(i) set the value of Z_i in \mathbf{x} by sampling from \mathbf{P}(Z_i \mid mb(Z_i)) \mathbf{C}[j] \leftarrow \mathbf{C}[j] + 1 where x_j is the value of X in \mathbf{x} return NORMALIZE(\mathbf{C})
```

- $mb(Z_i)$  is the Markov blanket of random variable  $Z_i$  using the CPTs of  $Z_i$  and its children.
- The Markov chain converges to a stationary distribution which is the asked for conditional probability.



### Conclusion

- Bayesian networks provide an efficient way to store a probabilistic model by exploiting conditional independence.
- Inference (estimating conditional probabilities) is still difficult, for all but tiny models.
- State of the art is sampling from the model.