

HE3021 Tutorial 8 Attempt

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```
warningsoff <- getOption("warn")
options(warn = -1)
library(haven) # read dta files
library(tidyverse) # manipulate dataframes
library(ivreg) # 2 stage linear regression and its associated statistical tests
library(aod) # for chi2 and F tests

filpaths <- list.files(path = paste0(getwd(),"//rawdata//"),
                        full.names = TRUE)
wage2 <- read_dta(file = filpaths[1])
options(warn = warningsoff)
```

1

a

$H_0 : \pi_{sibs} = \pi_{brthord} = \pi_{meduc} = 0, H_1 : \text{otherwise}, \alpha = 0.05, LM - test$

```
instr_exo_test_model <- lm(
  formula = educ ~ sibs + brthord + meduc + married + exper + tenure, data = wage2)
instr_exo_test <- wald.test(
  Sigma = vcov(instr_exo_test_model), b = coefficients(instr_exo_test_model), Terms = c(2,3,4))
# p-value below printed value
# instrumental relevance true
```

p-value == 0 < 0.05, thus we have sufficient evidence to reject H_0 .

b

```
twostage_model <- ivreg(
  formula = lwage ~ married + exper + tenure + educ | married + exper + tenure
  + sibs + brthord + meduc, data = wage2)
twostage_model_summary <- (summary(twostage_model, diagnostics = TRUE))
summary(twostage_model, diagnostics = FALSE)
```

```
##
## Call:
## ivreg(formula = lwage ~ married + exper + tenure + educ | married +
##       exper + tenure + sibs + brthord + meduc, data = wage2)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.72191 -0.23654  0.01206  0.25728  1.33428
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  4.079660   0.351142  11.618 < 2e-16 ***
## married      0.205496   0.048986   4.195 3.04e-05 ***
## exper        0.035371   0.006378   5.546 4.01e-08 ***
## tenure       0.006522   0.003090   2.111  0.0351 *
## educ         0.153649   0.021205   7.246 1.04e-12 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4 on 779 degrees of freedom
## Multiple R-Squared: 0.05733, Adjusted R-squared: 0.05249
## Wald test: 22.26 on 4 and 779 DF, p-value: < 2.2e-16
```

c

Regressing the residuals on each instrumental variable, we have:

$$\hat{u}_{IV} = \gamma_0 + \gamma_1 \text{married} + \gamma_2 \text{exper} + \gamma_3 \text{tenure} + \gamma_4 \text{sibs} + \gamma_5 \text{brthord} + \gamma_6 \text{meduc}$$

We test: $H_0 : \gamma_4 = \gamma_5 = \gamma_6 = 0, H_1 : \text{otherwise}, \alpha = 0.05, F - \text{test}$. The J-statistic is: 0.600554. From the χ^2 -table, we have the critical value as $5.991 > 0.600554$. The p-value is $0.7406131 > 0.05$, so we have insufficient evidence to reject the null hypothesis. We assume that at least 1 of the instrument variables is exogenous, then not rejecting H_0 means that all instrumental variables are exogenous.

d

From the Sargan test, it is likely that the 2SLS model's instrumental variables fulfill instrumental exogeneity. Then, we test $H_0 : \beta_{educ,2SLS} \text{ and } \beta_{educ,OLS} \text{ are consistent, } Var(\beta_{educ,2SLS}) > Var(\beta_{educ,OLS}), H_1 : \text{only } \beta_{educ,2SLS} \text{ is consistent}, \alpha = 0.05, \chi^2 - \text{test}$. If the coefficients estimated in OLS are significantly different from those in 2-stage regression, then it is likely that the OLS coefficient is inconsistent and that $Cov(x, u) \neq 0$. The $\chi^2 - \text{statistic}$ is $16.6912726 > 5.991$ and the p-value is $4.8543614 \times 10^{-5} < 0.05$. Thus we have sufficient evidence to reject H_0 . It is likely that $Cov(x, u) \neq 0$.

2

a

$$\hat{\beta}_1 = \frac{\hat{Cov}(x, y)}{\hat{Var}(x)} \quad (1)$$

$$= \frac{\hat{Cov}(x, \beta_0 + \beta_1 x + u)}{\hat{Var}(x)} \quad (2)$$

$$= \beta_1 \frac{\hat{Var}(x)}{\hat{Var}(x)} + \frac{\hat{Cov}(x, u)}{\hat{Var}(x)} \quad (3)$$

$$= \beta_1 + \frac{\hat{Cov}(x, u)}{\sqrt{\hat{Var}(x) \cdot \hat{Var}(u)}} \cdot \frac{\sqrt{\hat{Var}(u)}}{\sqrt{\hat{Var}(x)}} \quad (4)$$

$$\lim_{n \rightarrow \infty} \left[\beta_1 + \frac{\hat{Cov}(x, u)}{\sqrt{\hat{Var}(x) \cdot \hat{Var}(u)}} \cdot \frac{\sqrt{\hat{Var}(u)}}{\sqrt{\hat{Var}(x)}} \right] = \beta_1 + Corr(x, u) \cdot \frac{\sigma_u}{\sigma_x} \quad (5)$$

b

$$\hat{\beta}_{1,IV} = \frac{\hat{Cov}(z, y)}{\hat{Cov}(z, x)} \quad (6)$$

$$= \frac{\hat{Cov}(z, \beta_0 + \beta_1 x + u)}{\hat{Cov}(z, x)} \quad (7)$$

$$= \beta_1 \frac{\hat{Cov}(z, x)}{\hat{Cov}(z, x)} + \frac{\hat{Cov}(z, u)}{\hat{Cov}(z, x)} \quad (8)$$

$$= \beta_1 + \frac{\hat{Cov}(z, u)}{\hat{Cov}(z, x)} \cdot \frac{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(x)}}{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(u)}} \cdot \frac{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(u)}}{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(x)}} \quad (9)$$

$$\lim_{n \rightarrow \infty} \left[\beta_1 + \frac{\hat{Cov}(z, u)}{\hat{Cov}(z, x)} \cdot \frac{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(x)}}{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(u)}} \cdot \frac{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(u)}}{\sqrt{\hat{Var}(z)}\sqrt{\hat{Var}(x)}} \right] = \beta_1 + \frac{Corr(z, u)}{Corr(z, x)} \cdot \frac{\sigma_u}{\sigma_x} \quad (10)$$

c

$Cov(z, u) \neq 0$, $Cov(z, x)$ small, so $\frac{Corr(z, u)}{Corr(z, x)} \cdot \frac{\sigma_u}{\sigma_x} \neq 0$, so

$$\lim_{n \rightarrow \infty} \left[\beta_1 + \frac{\hat{Cov}(x, u)}{\sqrt{\hat{Var}(x) \cdot \hat{Var}(u)}} \cdot \frac{\sqrt{\hat{Var}(u)}}{\sqrt{\hat{Var}(u)}} \right] = \beta_1 + Corr(x, u) \cdot \frac{\sigma_u}{\sigma_x} \neq \beta_1$$

and the estimator is inconsistent.

Neither estimator would work, since $Cov(x, u) \neq 0$, $Cov(z, u) \neq 0$; they are both endogenous.