

# STAT 235 FINAL EXAM FORMULA SHEET

## Summaries:

$$\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n}, s^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{1}{n-1} \left[ \sum x_i^2 - \frac{(\sum x_i)^2}{n} \right]$$

IQR =  $Q_3 - Q_1$ , outliers are observations  $1.5 \times \text{IQR}$  below  $Q_1$  or  $1.5 \times \text{IQR}$  above  $Q_3$ .

## Probability:

Conditional Probability:  $P(A | B) = \frac{P(A \cap B)}{P(B)}$

Complement Law:  $P(A') = 1 - P(A)$

Multiplication Law:  $P(A \cap B) = P(A) \times P(B | A) = P(B) \times P(A | B)$

If  $A$  and  $B$  are independent,  $P(A \cap B) = P(A) \times P(B)$  since  $P(A | B) = P(A)$ .

Addition Law:  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

If  $A$  and  $B$  are mutually exclusive,  $P(A \cup B) = P(A) + P(B)$

For any two events  $A$  and  $B$ :  $P(A) = P(A \cap B) + P(A \cap B')$

Permutations:  $P_k^n = n \times (n-1) \times (n-2) \times \dots \times (n-k+1) = \frac{n!}{(n-k)!}$

Combinations:  $C_k^n = \binom{n}{k} = \frac{n!}{k!(n-k)!}$ , where  $n! = n \times (n-1) \times (n-2) \times \dots \times 2 \times 1$

## Linear Combinations of Random Variables:

If  $Y = a_1X_1 + a_2X_2 + \dots + a_nX_n + b$ ,  $E(Y) = a_1E(X_1) + a_2E(X_2) + \dots + a_nE(X_n) + b$

If  $X_1, X_2, \dots, X_n$  are independent,  $\text{Var}(Y) = a_1^2\text{Var}(X_1) + a_2^2\text{Var}(X_2) + \dots + a_n^2\text{Var}(X_n)$

## Discrete Variables:

$\mu = E(X) = \sum_{i=1} x_i p_i$ ,  $F(x) = P(X \leq x) = \text{Sum of probabilities } p_i \text{ for } x_i \leq x$ ,

$\sigma^2 = \text{Var}(X) = E(X^2) - [E(X)]^2 = \sum_{i=1} (x_i - \mu)^2 p_i = \sum_{i=1} x_i^2 p_i - \mu^2$

Distribution	Probability mass function	Mean	Variance
Binomial	$f(x) = P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$	$E(X) = np$	$\text{Var}(X) = np(1-p)$
Geometric	$f(x) = P(X = x) = (1-p)^{x-1} p$	$E(X) = \frac{1}{p}$	$\text{Var}(X) = \frac{(1-p)}{p^2}$
Negative Binomial	$f(x) = P(X = x) = \binom{x-1}{r-1} (1-p)^{x-r} p^r$	$E(X) = \frac{r}{p}$	$\text{Var}(X) = \frac{r(1-p)}{p^2}$
Poisson	$f(x) = P(X = x) = \frac{e^{-\lambda} \lambda^x}{x!}$	$E(X) = \lambda$	$\text{Var}(X) = \lambda$

### Continuous Variables:

$$\int_{-\infty}^{\infty} f(x)dx = 1, P(a \leq X \leq b) = \int_a^b f(x)dx, F(x) = P(X \leq x) = \int_{-\infty}^x f(t)dt$$

$$\mu = E(X) = \int_{-\infty}^{\infty} xf(x)dx, \sigma^2 = Var(X) = E(X^2) - [E(X)]^2 = \int_{-\infty}^{\infty} x^2 f(x)dx - \mu^2$$

Distribution	Density function	Mean	Variance	F(x)
Uniform	$f(x) = \frac{1}{b-a}$	$E(X) = \frac{b+a}{2}$	$Var(X) = \frac{(b-a)^2}{12}$	$F(x) = \frac{x-a}{b-a}$
Exponential	$f(x) = \lambda e^{-\lambda x}$	$E(X) = \frac{1}{\lambda}$	$Var(X) = \frac{1}{\lambda^2}$	$F(x) = 1 - e^{-\lambda x}$

Normal: If  $X$  has a normal distribution with mean  $\mu$  and variance  $\sigma^2$ , then  $X \sim N(\mu, \sigma^2)$ .

Standard normal: If  $X \sim N(\mu, \sigma^2)$ , then  $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$  and  $\Phi(z) = P(Z \leq z)$ .

Normal approximation to binomial: If  $X$  follows a binomial distribution with parameters  $n$  and  $p$ , then

$$P(a \leq X \leq b) \approx P(a - 0.5 \leq Y \leq b + 0.5),$$

where  $Y$  is normal with the mean  $np$  and the variance  $np(1-p)$ , under the assumption that  $np \geq 15$  and  $n(1-p) \geq 15$ .

### Sampling Distributions

Sample mean	Sample proportion
$E(\bar{X}) = \mu$	$E(\hat{p}) = p$
$Z = \frac{\bar{X} - \mu_{\bar{X}}}{\sigma_{\bar{X}}} = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} \sim N(0,1)$	$Z = \frac{\hat{p} - \mu_{\hat{p}}}{\sigma_{\hat{p}}} = \frac{\hat{p} \pm \frac{0.5}{n} - p}{\sqrt{\frac{p(1-p)}{n}}} \sim N(0,1)$
Normal if $n \geq 30$ OR population is normal	Normal if $np \geq 15$ and $n(1-p) \geq 15$

### Inferences about $\mu$ ( $\sigma$ known)

$$z_0 = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}, \bar{x} \pm z_{\alpha/2} \times \left( \frac{\sigma}{\sqrt{n}} \right), \mu \leq \bar{x} + z_{\alpha} \times \frac{\sigma}{\sqrt{n}}, \bar{x} - z_{\alpha} \times \frac{\sigma}{\sqrt{n}} \leq \mu$$

### Inferences about $\mu$ ( $\sigma$ unknown)

$$t_0 = \frac{\bar{x} - \mu_0}{s / \sqrt{n}}, \bar{x} \pm t_{\alpha/2, n-1} \times \left( \frac{s}{\sqrt{n}} \right), \mu \leq \bar{x} + t_{\alpha, n-1} \times \frac{s}{\sqrt{n}}, \bar{x} - t_{\alpha, n-1} \times \frac{s}{\sqrt{n}} \leq \mu$$

### Inferences about $\mu_d$ for paired data

$$t_0 = \frac{\bar{d} - \delta_0}{s_d / \sqrt{n}}, \bar{d} \pm t_{\alpha/2, n-1} \times \left( \frac{s_d}{\sqrt{n}} \right)$$

### Inferences about $\mu_1 - \mu_2$ (independent samples, variances known)

$$z_0 = \frac{(\bar{x}_1 - \bar{x}_2) - \delta_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}, (\bar{x}_1 - \bar{x}_2) \pm z_{\alpha/2} \times \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

### Inferences about $\mu_1 - \mu_2$ (independent samples, variances unknown, equal variances)

$$s_p = \sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}}, s.e.(\bar{x}_1 - \bar{x}_2) = s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

$$t_0 = \frac{(\bar{x}_1 - \bar{x}_2) - \delta_0}{s.e.(\bar{x}_1 - \bar{x}_2)}, (\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2, v} \times s.e.(\bar{x}_1 - \bar{x}_2), t\text{-distribution with } v = df = n_1 + n_2 - 2$$

### Inferences about $\mu_1 - \mu_2$ (independent samples, variances unknown, variances not assumed equal)

$$t_0 = \frac{(\bar{x}_1 - \bar{x}_2) - \delta_0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}, (\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2, v} \times \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}, t\text{-distribution with } v = df = \min\{n_1, n_2\} - 1$$

### Simple Linear Regression (SLR)

Simple linear regression model:  $Y = \beta_0 + \beta_1 x + \text{ERROR}$ ,  $\text{ERROR} \sim N(0, \sigma^2)$ .

Estimated Regression Line:  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}}, S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}), \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}, \hat{\sigma}^2 = \frac{SSE}{n - 2},$$

$$SSE = \sum e_i^2 = \sum (y_i - \hat{y}_i)^2$$

$$SST = SSR + SSE, F_0 = \frac{SSR/1}{SSE/(n - 2)} = \frac{MSR}{MSE}, R^2 = \frac{SSR}{SST}$$

### Inferences in Regression

$$t_0 = \frac{\hat{\beta}_1 - b_1}{s.e.(\hat{\beta}_1)}, \hat{\beta}_1 \pm t_{\alpha/2, v} \times s.e.(\hat{\beta}_1), s.e.(\hat{\beta}_1) = \sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}, t \text{ distribution with } v = n - 2$$

$$(\hat{\beta}_0 + \hat{\beta}_1 x^*) \pm t_{\alpha/2, v} \times \hat{\sigma} \sqrt{\frac{1}{n} + \frac{(x^* - \bar{x})^2}{S_{xx}}}; (\hat{\beta}_0 + \hat{\beta}_1 x^*) \pm t_{\alpha/2, v} \times \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{S_{xx}}}$$

### Inferences about $p$

$$z_0 = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}, \text{ under the assumption that } np_0 \geq 15 \text{ and } n(1-p_0) \geq 15$$

$$\hat{p} \pm z_{\alpha/2} \times \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, \quad p \geq \hat{p} - z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, \quad p \leq \hat{p} + z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}},$$

under the assumption that  $n\hat{p} \geq 15$  and  $n(1-\hat{p}) \geq 15$ .

### Inferences about $p_1 - p_2$

$$\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}, \quad z_0 = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}, \quad (\hat{p}_1 - \hat{p}_2) \pm z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

Assumptions:  $n_1\hat{p}_1$ ,  $n_1(1-\hat{p}_1)$ ,  $n_2\hat{p}_2$ ,  $n_2(1-\hat{p}_2)$  are all larger than 15.

### One-Way ANOVA

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k, \quad SST = SSTr + SSE$$

$$F_0 = \frac{SSTr / (k-1)}{SSE / (N-k)} = \frac{MSTr}{MSE} \sim F_{k-1, N-k}$$