

# Change point and change interface estimation for a stochastic heat equation

Stochastic Analysis and Statistics – University of Tokyo

---

Lukas Trottner

based on joint work with Markus Reiß, Claudia Strauch and Anton Tiepner

07 February 2024

Aarhus University

Humboldt-Universität zu Berlin



## Some generalities on statistics for SPDEs

- Let  $-A_\vartheta$  be a non-negative self-adjoint operator on  $L^2(\Lambda)$  for some domain  $\Lambda \subset \mathbb{R}^d$  and consider the SPDE

$$\begin{cases} dX(t) = A_\vartheta X(t) dt + dW(t), & t \in (0, T], \\ X(0) = X_0 \in L^2(\Lambda), \\ X(t)|_{\partial\Lambda} = 0, & t \in (0, T], \end{cases}$$

where  $W$  is a cylindrical Brownian motion (that is, a random linear transformation

$z \mapsto (W_z(t))_{t \in [0, T]} = (\langle W(t), z \rangle)_{t \in [0, T]}$  on  $L^2$  s.t. for any  $z$  the rhs is a BM with variance  $\|z\|_{L^2}^2$  and  $\mathbb{E}[\langle W(t), z \rangle \langle W(s), z' \rangle] = t \wedge s \langle z, z' \rangle$ )

## Some generalities on statistics for SPDEs

- Let  $-A_\vartheta$  be a non-negative self-adjoint operator on  $L^2(\Lambda)$  for some domain  $\Lambda \subset \mathbb{R}^d$  and consider the SPDE

$$\begin{cases} dX(t) = A_\vartheta X(t) dt + dW(t), & t \in (0, T], \\ X(0) = X_0 \in L^2(\Lambda), \\ X(t)|_{\partial\Lambda} = 0, & t \in (0, T], \end{cases}$$

where  $W$  is a cylindrical Brownian motion (that is, a random linear transformation

$z \mapsto (W_z(t))_{t \in [0, T]} = (\langle W(t), z \rangle)_{t \in [0, T]}$  on  $L^2$  s.t. for any  $z$  the rhs is a BM with variance  $\|z\|_{L^2}^2$  and  $\mathbb{E}[\langle W(t), z \rangle \langle W(s), z' \rangle] = t \wedge s \langle z, z' \rangle$ )

- for  $S_\vartheta(t) = \exp(tA_\vartheta)$  the **mild solution** process is defined by

$$X(t) = S_\vartheta(t)X_0 + \int_0^t S_\vartheta(t-s) dW(s), \quad t \in [0, T],$$

which for  $d > 1$  is typically only realized in a space of distributions

## Some generalities on statistics for SPDEs

- Let  $-A_\vartheta$  be a non-negative self-adjoint operator on  $L^2(\Lambda)$  for some domain  $\Lambda \subset \mathbb{R}^d$  and consider the SPDE

$$\begin{cases} dX(t) = A_\vartheta X(t) dt + dW(t), & t \in (0, T], \\ X(0) = X_0 \in L^2(\Lambda), \\ X(t)|_{\partial\Lambda} = 0, & t \in (0, T], \end{cases}$$

where  $W$  is a cylindrical Brownian motion (that is, a random linear transformation

$z \mapsto (W_z(t))_{t \in [0, T]} = (\langle W(t), z \rangle)_{t \in [0, T]}$  on  $L^2$  s.t. for any  $z$  the rhs is a BM with variance  $\|z\|_{L^2}^2$  and  $\mathbb{E}[\langle W(t), z \rangle \langle W(s), z' \rangle] = t \wedge s \langle z, z' \rangle$ )

- for  $S_\vartheta(t) = \exp(tA_\vartheta)$  the **mild solution** process is defined by

$$X(t) = S_\vartheta(t)X_0 + \int_0^t S_\vartheta(t-s) dW(s), \quad t \in [0, T],$$

which for  $d > 1$  is typically only realized in a space of distributions

- the mild solution is a **weak solution** in the sense

$$\langle X(t), z \rangle = \langle X_0, z \rangle + \int_0^t \langle X(s), A_\vartheta z \rangle ds + \langle W(t), z \rangle, \quad z \in D(A_\vartheta).$$

- **spectral observations**: provided  $A_\vartheta$  has an orthonormal eigenbasis  $(e_j)$  **that is independent of  $\vartheta$**  (think  $A_\vartheta = \vartheta \Delta$ ), observe  $(j = 1, \dots, n, t \in [0, T])$

$$\langle X(t), e_j \rangle = \langle X_0, e_j \rangle + \int_0^t \langle X(s), A_\vartheta e_j \rangle ds + \langle W(t), e_j \rangle = \langle X_0, e_j \rangle - \lambda_j(\vartheta) \int_0^t \langle X(s), e_j \rangle ds + \langle W(t), e_j \rangle,$$

---

<sup>1</sup>Altmeyer, Reiß (2021). Nonparametric estimation for linear SPDEs from local measurements. *Ann. Appl. Prob.*

- **spectral observations**: provided  $A_\vartheta$  has an orthonormal eigenbasis  $(e_j)$  **that is independent of  $\vartheta$**  (think  $A_\vartheta = \vartheta \Delta$ ), observe  $(j = 1, \dots, n, t \in [0, T])$

$$\langle X(t), e_j \rangle = \langle X_0, e_j \rangle + \int_0^t \langle X(s), A_\vartheta e_j \rangle ds + \langle W(t), e_j \rangle = \langle X_0, e_j \rangle - \lambda_j(\vartheta) \int_0^t \langle X(s), e_j \rangle ds + \langle W(t), e_j \rangle,$$

~ observations consist of  $n$  independent Ornstein–Uhlenbeck processes, asymptotics:  $n \rightarrow \infty$

---

<sup>1</sup>Altmeyer, Reiß (2021). Nonparametric estimation for linear SPDEs from local measurements. *Ann. Appl. Prob.*

- **spectral observations**: provided  $A_\vartheta$  has an orthonormal eigenbasis  $(e_j)$  **that is independent of  $\vartheta$**  (think  $A_\vartheta = \vartheta \Delta$ ), observe  $(j = 1, \dots, n, t \in [0, T])$

$$\langle X(t), e_j \rangle = \langle X_0, e_j \rangle + \int_0^t \langle X(s), A_\vartheta e_j \rangle ds + \langle W(t), e_j \rangle = \langle X_0, e_j \rangle - \lambda_j(\vartheta) \int_0^t \langle X(s), e_j \rangle ds + \langle W(t), e_j \rangle,$$

~ observations consist of  $n$  independent Ornstein–Uhlenbeck processes, asymptotics:  $n \rightarrow \infty$

- a more flexible approach is the **local observation model**<sup>1</sup>: for smooth  $K$  with  $\|K\|_{L^2} = 1$ , **resolution**  $\delta > 0$  and measurement locations  $x_i \in \Lambda$  define the **point-spread function**  $K_{\delta,i} := \delta^{-d/2} K(\frac{\cdot - x_i}{\delta})$

---

<sup>1</sup>Altmeyer, Reiß (2021). Nonparametric estimation for linear SPDEs from local measurements. *Ann. Appl. Prob.*

- **spectral observations**: provided  $A_\vartheta$  has an orthonormal eigenbasis  $(e_j)$  **that is independent of  $\vartheta$**  (think  $A_\vartheta = \vartheta \Delta$ ), observe  $(j = 1, \dots, n, t \in [0, T])$

$$\langle X(t), e_j \rangle = \langle X_0, e_j \rangle + \int_0^t \langle X(s), A_\vartheta e_j \rangle ds + \langle W(t), e_j \rangle = \langle X_0, e_j \rangle - \lambda_j(\vartheta) \int_0^t \langle X(s), e_j \rangle ds + \langle W(t), e_j \rangle,$$

~ observations consist of  $n$  independent Ornstein–Uhlenbeck processes, asymptotics:  $n \rightarrow \infty$

- a more flexible approach is the **local observation model**<sup>1</sup>: for smooth  $K$  with  $\|K\|_{L^2} = 1$ , **resolution**  $\delta > 0$  and measurement locations  $x_i \in \Lambda$  define the **point-spread function**  $K_{\delta,i} := \delta^{-d/2} K(\frac{\cdot - x_i}{\delta})$
- observations consist of  $(\langle X(t), K_{\delta,i} \rangle)_{t \in [0, T], i=1, \dots, n}$ , where for  $K_{\delta,i} \in D(A_\vartheta)$ ,

$$d\langle X(t), K_{\delta,i} \rangle = \langle X(t), A_\vartheta K_{\delta,i} \rangle dt + \langle K_{\delta,i}, dW_t \rangle$$

---

<sup>1</sup>Altmeyer, Reiß (2021). Nonparametric estimation for linear SPDEs from local measurements. *Ann. Appl. Prob.*

- **spectral observations**: provided  $A_\vartheta$  has an orthonormal eigenbasis  $(e_j)$  **that is independent of  $\vartheta$**  (think  $A_\vartheta = \vartheta \Delta$ ), observe  $(j = 1, \dots, n, t \in [0, T])$

$$\langle X(t), e_j \rangle = \langle X_0, e_j \rangle + \int_0^t \langle X(s), A_\vartheta e_j \rangle ds + \langle W(t), e_j \rangle = \langle X_0, e_j \rangle - \lambda_j(\vartheta) \int_0^t \langle X(s), e_j \rangle ds + \langle W(t), e_j \rangle,$$

~ observations consist of  $n$  independent Ornstein–Uhlenbeck processes, asymptotics:  $n \rightarrow \infty$

- a more flexible approach is the **local observation model**<sup>1</sup>: for smooth  $K$  with  $\|K\|_{L^2} = 1$ , **resolution**  $\delta > 0$  and measurement locations  $x_i \in \Lambda$  define the **point-spread function**  $K_{\delta,i} := \delta^{-d/2} K(\frac{\cdot - x_i}{\delta})$
- observations consist of  $(\langle X(t), K_{\delta,i} \rangle)_{t \in [0, T], i=1, \dots, n}$ , where for  $K_{\delta,i} \in D(A_\vartheta)$ ,

$$d\langle X(t), K_{\delta,i} \rangle = \langle X(t), A_\vartheta K_{\delta,i} \rangle dt + \langle K_{\delta,i}, dW_t \rangle$$

~ observations are generalized Itô processes (but not independent for  $i \neq j$ ), asymptotics:  $\delta \rightarrow 0$ ,  $n$  may be fixed or increase with  $\delta^{-1}$

---

<sup>1</sup>Altmeyer, Reiβ (2021). Nonparametric estimation for linear SPDEs from local measurements. *Ann. Appl. Prob.*

## A change point model for a stochastic heat equation

We consider the SPDE on  $\Lambda = (0, 1)$  given by

$$\begin{cases} dX(t) = \Delta_\vartheta X(t) dt + dW(t), & t \in (0, T], \\ X(0) \equiv 0, \\ X(t)|_{\{0,1\}} = 0, & t \in (0, T], \end{cases}$$

for a weighted Laplacian  $\Delta_\vartheta := \nabla \vartheta \nabla$ , with **discontinuous diffusivity**

$$\vartheta(x) = \vartheta_- \mathbf{1}_{(0,\tau)}(x) + \vartheta_+ \mathbf{1}_{[\tau,1)}(x), \quad x, \tau \in (0, 1), 0 < \vartheta_- \wedge \vartheta_+.$$



## A change point model for a stochastic heat equation

We consider the SPDE on  $\Lambda = (0, 1)$  given by

$$\begin{cases} dX(t) = \Delta_\vartheta X(t) dt + dW(t), & t \in (0, T], \\ X(0) \equiv 0, \\ X(t)|_{\{0,1\}} = 0, & t \in (0, T], \end{cases}$$

for a weighted Laplacian  $\Delta_\vartheta := \nabla \vartheta \nabla$ , with discontinuous diffusivity

$$\vartheta(x) = \vartheta_- \mathbf{1}_{(0,\tau)}(x) + \vartheta_+ \mathbf{1}_{[\tau,1)}(x), \quad x, \tau \in (0, 1), 0 < \vartheta_- \wedge \vartheta_+.$$



Classical heat kernel bounds for analytic semigroup  $(S_\vartheta(t) = \exp(t\Delta_\vartheta))_{t \in [0, T]}$  implies that

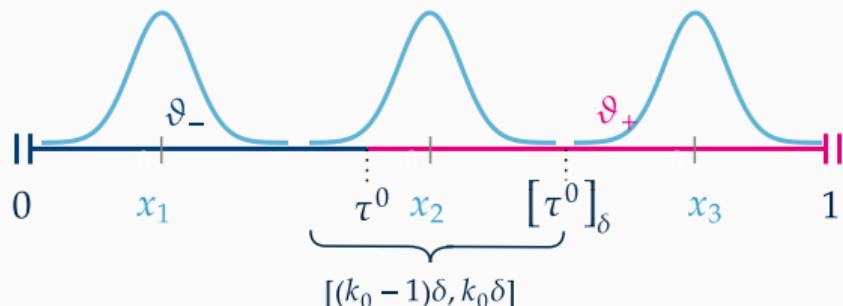
$$X(t) = \int_0^t S_\vartheta(t-s) dW_s, \quad t \in [0, T]$$

lives in  $L^2((0, 1))$  and we have

$$\langle X(t), z \rangle = \int_0^t \langle X(s), \Delta_\vartheta z \rangle ds + \langle W(t), z \rangle, \quad z \in D(\Delta_\vartheta) = \{u \in H_0^1((0, 1)) : \vartheta \nabla u \in H^1((0, 1))\}.$$

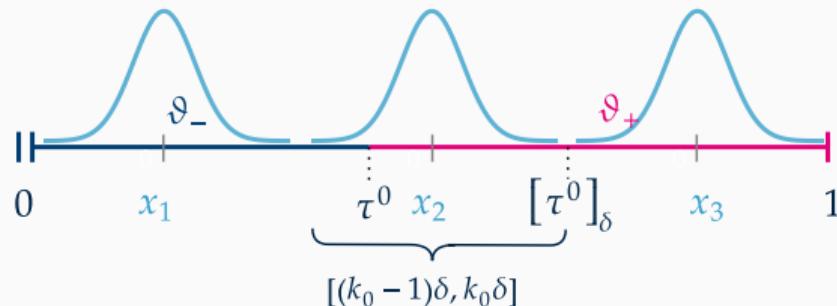
## Estimation approach

- let  $K: \mathbb{R} \rightarrow \mathbb{R}$  be a smooth kernel with  $\text{supp } K \subset [-1/2, 1/2]$ ,  $\|K\|_{L^2} = 1$  and for  $\delta = n^{-1}$ ,  $x_i = (i - 1/2)\delta$  ( $i \in \{1, \dots, \delta^{-1}\}$ ), define  $K_{\delta,i} = \delta^{-1/2}K(\delta^{-1}(x - x_i))$
- local observations  $(X_{\delta,i}(t))_{t \in [0, T]} = (\langle X(t), K_{\delta,i} \rangle)_{t \in [0, T]}$  and  $(X_{\delta,i}^\Delta(t))_{t \in [0, T]} = (\langle X(t), \Delta K_{\delta,i} \rangle)_{t \in [0, T]}$



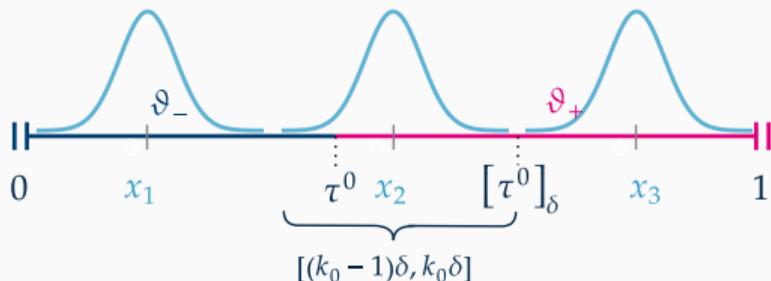
## Estimation approach

- let  $K: \mathbb{R} \rightarrow \mathbb{R}$  be a smooth kernel with  $\text{supp } K \subset [-1/2, 1/2]$ ,  $\|K\|_{L^2} = 1$  and for  $\delta = n^{-1}$ ,  $x_i = (i - 1/2)\delta$  ( $i \in \{1, \dots, \delta^{-1}\}$ ), define  $K_{\delta,i} = \delta^{-1/2}K(\delta^{-1}(x - x_i))$
- local observations**  $(X_{\delta,i}(t))_{t \in [0, T]} = (\langle X(t), K_{\delta,i} \rangle)_{t \in [0, T]}$  and  $(X_{\delta,i}^{\Delta}(t))_{t \in [0, T]} = (\langle X(t), \Delta K_{\delta,i} \rangle)_{t \in [0, T]}$



- if  $i \neq k_0$ , then  $X_{\delta,i}$  solves  $dX_{\delta,i}(t) = \vartheta(x_i)X_{\delta,i}^{\Delta}(t)dt + dB_{\delta,i}(t)$  for independent BMs  $(B_{\delta,i}, i \in [\delta^{-1}])$
- we define a **modified local log-likelihood** by

$$\ell_{\delta,i}(\vartheta_-, \vartheta_+, \vartheta_0, k) \coloneqq \vartheta_{\delta,i}(k) \int_0^T X_{\delta,i}^{\Delta}(t) dX_{\delta,i}(t) - \frac{\vartheta_{\delta,i}(k)^2}{2} \int_0^T X_{\delta,i}^{\Delta}(t)^2 dt, \quad \vartheta_{\delta,i}(k) \coloneqq \begin{cases} \vartheta_-, & i < k, \\ \vartheta_0, & i = k \\ \vartheta_+, & i > k \end{cases}$$



- modified local log-likelihood:

$$\ell_{\delta,i}(\vartheta_-, \vartheta_+, \vartheta_o, k) := \vartheta_{\delta,i}(k) \int_0^T X_{\delta,i}^\Delta(t) dX_{\delta,i}(t) - \frac{\vartheta_{\delta,i}(k)^2}{2} \int_0^T X_{\delta,i}^\Delta(t)^2 dt, \quad \vartheta_{\delta,i}(k) := \begin{cases} \vartheta_-, & i < k, \\ \vartheta_o, & i = k, \\ \vartheta_+, & i > k \end{cases}$$

- CUSUM-approach:  $(\hat{\vartheta}_-, \hat{\vartheta}_+, \hat{\vartheta}_o, \hat{k}) := (\hat{\vartheta}_-, \hat{\vartheta}_+, \hat{\vartheta}_o, \hat{k}\delta)$ , where

$$\begin{aligned} (\hat{\vartheta}_-, \hat{\vartheta}_+, \hat{\vartheta}_o, \hat{k}) &\doteq \arg \max_{(\vartheta_-, \vartheta_+, \vartheta_o, k)} \sum_{i \in [\delta^{-1}]} \ell_{\delta,i}(\vartheta_-, \vartheta_+, \vartheta_o, k) \\ &= \arg \min_{(\vartheta_-, \vartheta_+, \vartheta_o, k)} \left\{ \frac{1}{2} \sum_{i=1}^{\delta^{-1}} (\vartheta_{\delta,i}(k) - \vartheta_{\delta,i}^0)^2 I_{\delta,i} - \sum_{i=1}^{\delta^{-1}} (\vartheta_{\delta,i}(k) - \vartheta_{\delta,i}^0) M_{\delta,i} - \vartheta_{\delta,k_0}(k) R_{\delta,k_0}(\vartheta_o^0) \right\}, \end{aligned}$$

for

$$M_{\delta,i} \doteq \int_0^T X_{\delta,i}^\Delta(t) dB_{\delta,i}(t), \quad I_{\delta,i} \doteq \int_0^T X_{\delta,i}^\Delta(t)^2 dt,$$

and  $R_{\delta,k_0}(\vartheta_o^0)$  is an error term resulting from  $K_{\delta,k_0} \notin D(\Delta_\vartheta)$  in general

## Basic estimates

**Lemma** (Reiß, Strauch and T., 2023+)

- For any  $i \in [\delta^{-1}] \setminus \{k_0\}$ ,

$$\mathbb{E}[I_{\delta,i}] = \frac{T}{2\vartheta(x_i)} \|K'\|_{L^2}^2 \delta^{-2} + \mathcal{O}(1),$$

and, moreover,  $\mathbb{E}[I_{\delta,k_0}] \sim \delta^{-2}$ ;

- for any vector  $\alpha \in \mathbb{R}^n$  s.t.  $\alpha_{k_0} = 0$ ,

$$\text{Var}\left(\sum_{i=1}^{\delta^{-1}} \alpha_i I_{\delta,i}\right) \leq \frac{T}{2\underline{\vartheta}^3} \delta^{-2} \|\alpha\|_{\ell^2}^2 \|K'\|_{L^2}^2;$$

- For  $\eta := \vartheta_+^0 - \vartheta_-^0$ ,

$$\mathbb{E}[|R_{\delta,k_0}(\vartheta_\circ)|] \lesssim \delta^{-2}, \quad \text{Var}(R_{\delta,k_0}(\vartheta_\circ)) \lesssim \delta^{-2},$$

and  $\exists \vartheta_\circ^0$  s.t.  $|\mathbb{E}[R_{\delta,k_0}(\vartheta_\circ^0)]| \leq \delta^{-1}$ .

## Concentration result

Main observation:  $\sum_{i=1}^{\delta^{-1}} \alpha_i (I_{\delta,i} - \mathbb{E}[I_{\delta,i}])$  belongs to second Wiener chaos for an appropriate isonormal Gaussian process associated to  $(X_i^\Delta(t))_{t \in [0, T], i \in [\delta^{-1}]} \rightsquigarrow$  verify conditions for **Bernstein-type concentration inequality** from Nourdin and Viens (2009)<sup>2</sup>

### Proposition (Reiß, Strauch and T., 2023+)

Let  $\alpha \in \mathbb{R}_+^n \setminus \{0\}$  s.t.  $\alpha_{k_0} = 0$ . Then, for any  $z > 0$ , we have

$$\mathbb{P}\left(\left| \sum_{i=1}^n \alpha_i (I_{\delta,i} - \mathbb{E}[I_{\delta,i}]) \right| \geq z\right) \leq 2 \exp\left(-\frac{\underline{\vartheta}^2}{4\|\alpha\|_\infty} \frac{z^2}{z + \sum_{i=1}^n \alpha_i \mathbb{E}[I_{\delta,i}]}\right).$$

---

<sup>2</sup>I. Nourdin and F.G. Viens (2009). Density formula and concentration inequalities with Malliavin calculus. *Electron. J. Prob.*, 14: no. 78, 2287–2309.

Define the jump height  $\eta := \vartheta_+^0 - \vartheta_-^0$ .

**Theorem** (Reiß, Strauch and T., 2023+)

Suppose that  $\chi^0(\delta) \xrightarrow[\delta \rightarrow 0]{} \chi^*$  and that  $|\eta| \geq \underline{\eta}$  for all  $\delta \in 1/\mathbb{N}$ . Then,

$$\hat{\tau} - \tau^0 = \mathcal{O}_{\mathbb{P}}(\delta) \quad \text{and} \quad \hat{\vartheta}_{\pm} - \vartheta_{\pm}^0 = \mathcal{O}_{\mathbb{P}}(\delta^{3/2}).$$

- the estimation rate for  $\tau^0$  is the same as in classical discrete change point models
- the estimation rate for  $\vartheta_{\pm}^0$  is the same as the minimax optimal rate<sup>3</sup> for parametric estimation from multiple local measurements in the model  $A_{\vartheta} = \vartheta \Delta$

---

<sup>3</sup>Altmeyer, Tiepner, Wahl (2023+) Optimal parameter estimation for linear SPDEs from multiple measurements

## Limit theorem for vanishing jump height

- for the previous consistency result it was crucial that the jump height  $\eta$  **does not vanish**
- assume now that  $\eta \xrightarrow{\delta \rightarrow 0} 0$  and that the nuisance parameters  $\vartheta_{\pm}^0 = \vartheta_{\pm}^0(\delta)$  are known
- **CUSUM estimator:**  $\hat{\tau} = \hat{k}\delta$ , where

$$\begin{aligned}\hat{k} := \arg \max_{k=1, \dots, \delta^{-1}} & \sum_{i=1}^k \left( \vartheta_-^0 \int_0^T X_{\delta,i}^{\Delta}(t) dX_{\delta,i}(t) - \frac{(\vartheta_-^0)^2}{2} \int_0^T X_{\delta,i}^{\Delta}(t)^2 dt \right) \\ & + \sum_{i=k+1}^{\delta^{-1}} \left( \vartheta_+^0 \int_0^T X_{\delta,i}^{\Delta}(t) dX_{\delta,i}(t) - \frac{(\vartheta_+^0)^2}{2} \int_0^T X_{\delta,i}^{\Delta}(t)^2 dt \right)\end{aligned}$$

**Theorem** (Reiß, Strauch and T., 2023+)

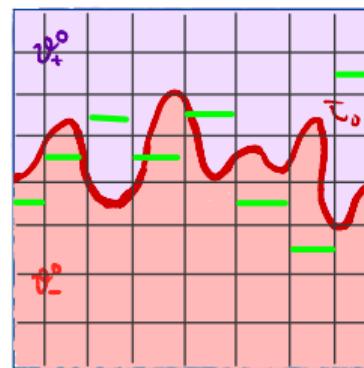
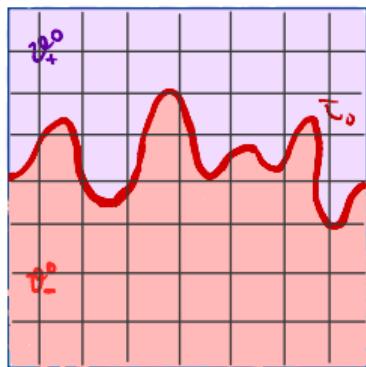
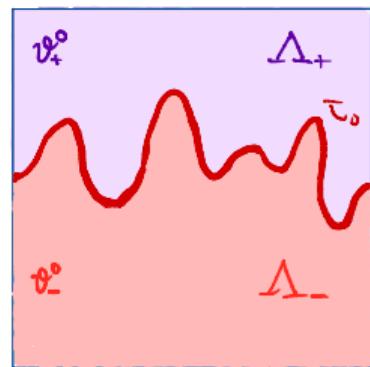
Assume  $\eta = o(\delta)$  and  $\delta^{3/2} = o(\eta)$ . Then, for a two-sided Brownian motion  $(B^{\leftrightarrow}(h), h \in \mathbb{R})$ , we have

$$\frac{\eta^2}{\delta^3} \frac{T \|K'\|_{L^2}^2}{2\vartheta^*} (\hat{\tau} - \tau) \xrightarrow{d} \arg \min_{h \in \mathbb{R}} \left\{ B^{\leftrightarrow}(h) + \frac{|h|}{2} \right\}, \quad \text{as } \delta \rightarrow 0.$$

$$dX_t = \Delta \vartheta^0 X_t dt + dW_t, \quad \vartheta^0(x) = \vartheta_-^0 1_{\Lambda_-}(x) + \vartheta_+^0 1_{\Lambda_+}(x), x \in \Lambda \subset [0,1]^d$$

$$\Delta \vartheta^0 u = \nabla \cdot \vartheta^0 \nabla u = \sum_{i=1}^d \partial_i (\vartheta^0 \partial_i u)$$

$$\partial \Lambda_- \cap \partial \Lambda_+ \cap [0,1]^d = \{(x, \tau^0(x)) : x \in [0,1]^{d-1}\}$$



$\delta^{-d}$  measurement locations,

loc. meas:  $\langle X(t), V_{S,i} \rangle, \langle X(t), \Delta V_{S,i} \rangle$   
over time interval  $[0, \bar{T}]$

estimate best piecewise  
constant approximation  $\hat{\tau}$   
based on CUSUM approach

## Convergence result in the multivariate case

**Theorem** (Tiepner and T., 2024+)

Let  $(\widehat{\vartheta}_-, \widehat{\vartheta}_+, \widehat{\tau})$  be the CUSUM estimator of  $(\vartheta_-^0, \vartheta_+^0, \tau^0)$ , designed under the assumption that we know two sets  $\Theta_{\pm} \subset [\underline{\vartheta}, \bar{\vartheta}]$  that contain  $\vartheta_{\pm}^0$  and are  $\underline{\eta}$ -separated,  $\underline{\eta} > 0$ . If  $\tau^0$  is  $\beta$ -Hölder continuous on  $[0, 1]^{d-1}$ , then

$$\mathbb{E}\left[\|\widehat{\tau} - \tau^0\|_{L^1([0,1]^{d-1})}\right] \lesssim \delta^\beta.$$

Given the necessary identifiability assumption  $\|\tau^0\|_{L^1([0,1]^{d-1})} \in (0, 1)$ ,  $(\widehat{\vartheta}_-, \widehat{\vartheta}_+)$  is a consistent estimator of  $(\vartheta_-^0, \vartheta_+^0)$ .

## Convergence result in the multivariate case

**Theorem** (Tiepner and T., 2024+)

Let  $(\widehat{\vartheta}_-, \widehat{\vartheta}_+, \widehat{\tau})$  be the CUSUM estimator of  $(\vartheta_-^0, \vartheta_+^0, \tau^0)$ , designed under the assumption that we know two sets  $\Theta_{\pm} \subset [\underline{\vartheta}, \bar{\vartheta}]$  that contain  $\vartheta_{\pm}^0$  and are  $\underline{\eta}$ -separated,  $\underline{\eta} > 0$ . If  $\tau^0$  is  $\beta$ -Hölder continuous on  $[0, 1]^{d-1}$ , then

$$\mathbb{E}\left[\|\widehat{\tau} - \tau^0\|_{L^1([0,1]^{d-1})}\right] \lesssim \delta^\beta.$$

Given the necessary identifiability assumption  $\|\tau^0\|_{L^1([0,1]^{d-1})} \in (0, 1)$ ,  $(\widehat{\vartheta}_-, \widehat{\vartheta}_+)$  is a consistent estimator of  $(\vartheta_-^0, \vartheta_+^0)$ .

In terms of the number of measurement location  $N = \delta^{-d}$ , the change profile estimation rate is given by

$$N^{-\beta/d} \leq \underbrace{N^{\frac{-\beta}{2\beta+d-1}}}_{\text{nonparametric rate}}, \quad \beta \gtrless 1/2.$$

This is the same rate that is observed in [image reconstruction problems with regular design](#).

## Summary

- for a stochastic heat equation with piecewise constant diffusivity, we construct a simultaneous M-estimator for the conductivities  $\vartheta_{\pm}^0$  and the change point  $\tau^0$  from multiple local measurements
- in case of non-vanishing jump height, we show that

$$\hat{\tau} - \tau^0 = \mathcal{O}_{\mathbb{P}}(\delta) \quad \text{and} \quad \hat{\vartheta}_{\pm} - \vartheta_{\pm}^0 = \mathcal{O}_{\mathbb{P}}(\delta^{3/2})$$

- in case of vanishing jump height and known parameters  $\vartheta_{\pm}^0$  we construct a change point estimator  $\hat{\tau}$  obeying the limit theorem

$$\frac{\eta^2}{\delta^3} \frac{T \|K'\|_{L^2}^2}{2\vartheta^*} (\hat{\tau} - \tau^0) \xrightarrow{d} \arg \min_{h \in \mathbb{R}} \left\{ B^{\leftrightarrow}(h) + \frac{|h|}{2} \right\}, \quad \text{as } \delta \rightarrow 0,$$

provided  $\eta = o(\delta)$  and  $\delta^{3/2} = o(\eta)$

- in the multivariate change interface estimation problem we construct an estimator that converges at rate  $\delta^{\beta}$  under  $\beta$ -Hölder smoothness assumptions

## Summary

- for a stochastic heat equation with piecewise constant diffusivity, we construct a simultaneous M-estimator for the conductivities  $\vartheta_{\pm}^0$  and the change point  $\tau^0$  from multiple local measurements
- in case of non-vanishing jump height, we show that

$$\hat{\tau} - \tau^0 = \mathcal{O}_{\mathbb{P}}(\delta) \quad \text{and} \quad \hat{\vartheta}_{\pm} - \vartheta_{\pm}^0 = \mathcal{O}_{\mathbb{P}}(\delta^{3/2})$$

- in case of vanishing jump height and known parameters  $\vartheta_{\pm}^0$  we construct a change point estimator  $\hat{\tau}$  obeying the limit theorem

$$\frac{\eta^2}{\delta^3} \frac{T \|K'\|_{L^2}^2}{2\vartheta^*} (\hat{\tau} - \tau^0) \xrightarrow{d} \arg \min_{h \in \mathbb{R}} \left\{ B^{\leftrightarrow}(h) + \frac{|h|}{2} \right\}, \quad \text{as } \delta \rightarrow 0,$$

provided  $\eta = o(\delta)$  and  $\delta^{3/2} = o(\eta)$

- in the multivariate change interface estimation problem we construct an estimator that converges at rate  $\delta^\beta$  under  $\beta$ -Hölder smoothness assumptions

Thank you for your attention!