Change point estimation for a stochastic heat equation

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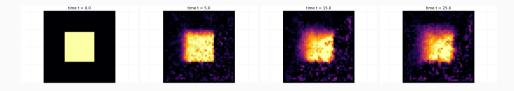
Lukas Trottner based on joint works with Markus Reiß, Claudia Strauch and Anton Tiepner 04 June 2025

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Change point model for stochastic heat equations



· Stochastic heat equation

$$dX(t) = \Delta_{\partial}X(t) dt + dW(t), \quad \Delta_{\partial} = \nabla \cdot \partial \nabla,$$

with Dirichlet boundary conditions, and broken diffusivity

$$\vartheta(x) = \vartheta_{-} \mathbf{1}_{\Lambda_{-}}(x) + \vartheta_{+} \mathbf{1}_{\Lambda_{+}}(x), \quad x \in [0, 1]^{d} = \Lambda_{-} \uplus \Lambda_{+}, \vartheta_{-} \wedge \vartheta_{+} > 0.$$

• special case for d=1: $\Lambda_+=(\tau,1]$ with change point τ



The univariate case

 $-\Delta_{\vartheta}$ is induced by Dirichlet form

$$\mathcal{E}(u,v) := \langle \partial \partial_x u, \partial_x v \rangle = \int_0^1 \partial(x) \, \partial_x u(x) \, \partial_x v(x) \, \mathrm{d}x, \quad u,v \in H_0^1((0,1)),$$

and generates C_0 -semigroup $S_{\vartheta}(t) = \exp(t\Delta_{\vartheta}), t \in [0, T]$, having transition densities that satisfy the heat kernel bound

$$p_t^9(x,y) \le c_1 t^{-1/2} \exp\left(-\frac{|x-y|^2}{c_2 t}\right), \quad (x,y) \in (0,1)^2, t \in (0,T].$$

Mild solution

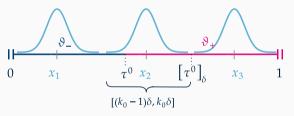
$$X(t) = \int_0^t S_{\theta}(t-s) \, \mathrm{d}W(s), \quad t \in [0,T], \quad \text{(assume } X(0) \equiv 0)$$

is $L^2((0,1))$ -valued and we have

$$\langle X(t), z \rangle = \int_0^t \langle X(s), \Delta_{\vartheta} z \rangle \, \mathrm{d}s + \langle W(t), z \rangle, \quad \forall \, z \in D(\Delta_{\vartheta}) = \big\{ u \in H_0^1((0, 1)) \, : \, \vartheta \partial_X u \in H^1((0, 1)) \big\}.$$

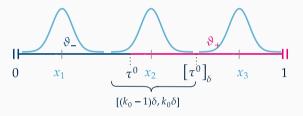
Observation model

- let $K : \mathbb{R} \to \mathbb{R}$ be a smooth kernel with supp $K \subset [-1/2, 1/2], \|K\|_{L^2} = 1$ and for $\delta = n^{-1}$, $x_i = (i 1/2)\delta(i \in \{1, ..., \delta^{-1}\})$, define $K_{\delta, i}(x) = \delta^{-1/2}K(\delta^{-1}(x x_i))$
- local observations $(X_{\delta,i}(t))_{t\in[0,T]} = (\langle X(t),K_{\delta,i}\rangle)_{t\in[0,T]}$ and $(X_{\delta,i}^{\Delta}(t))_{t\in[0,T]} = (\langle X(t),\Delta K_{\delta,i}\rangle)_{t\in[0,T]}$



Observation model

- let $K: \mathbb{R} \to \mathbb{R}$ be a smooth kernel with supp $K \subset [-1/2, 1/2]$, $||K||_{L^2} = 1$ and for $\delta = n^{-1}$, $x_i = (i 1/2)\delta(i \in \{1, ..., \delta^{-1}\})$, define $K_{\delta,i}(x) = \delta^{-1/2}K(\delta^{-1}(x x_i))$
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we have

$$X_{\delta,i}(t) = \begin{cases} \int_0^t \vartheta_{\pm}^0 X_{\delta,i}^{\Delta}(s) \, \mathrm{d}s + B_{\delta,i}(t), & i \gtrless k_0, \\ \int_0^t \int_0^s \langle \Delta_{\vartheta^0} S_{\vartheta^0}(s-u) K_{\delta,i}, \mathrm{d}W(u) \rangle \, \mathrm{d}s + B_{\delta,k_0}(t), & i = k_0, \end{cases}$$

for independent Brownian motions $(B_{\delta,i})_{i \in [\delta^{-1}]}$

Estimation approach

· modified local log-likelihood:

$$\ell_{\delta,i}(\partial_{-},\partial_{+},\partial_{\circ},k) := \partial_{\delta,i}(k) \int_{0}^{T} X_{\delta,i}^{\Delta}(t) \, \mathrm{d}X_{\delta,i}(t) - \frac{\partial_{\delta,i}(k)^{2}}{2} \int_{0}^{T} X_{\delta,i}^{\Delta}(t)^{2} \, \mathrm{d}t, \quad \partial_{\delta,i}(k) := \begin{cases} \partial_{-}, & i < k, \\ \partial_{\circ}, & i = k, \\ \partial_{+}, & i > k \end{cases}$$

• set $(\hat{\partial}_-,\hat{\partial}_+,\hat{\partial}_\circ,\hat{\tau}):=(\hat{\partial}_-,\hat{\partial}_+,\hat{\partial}_\circ,\hat{k}\delta)$, where

$$\begin{split} (\hat{\partial}_{-}, \hat{\partial}_{+}, \hat{\partial}_{\circ}, \hat{k}) &:= \underset{(\partial_{-}, \partial_{+}, \partial_{\circ}, k)}{\operatorname{arg\,max}} \sum_{i \in [\delta^{-1}]} \ell_{\delta, i}(\partial_{-}, \partial_{+}, \partial_{\circ}, k) \\ &= \underset{(\partial_{-}, \partial_{+}, \partial_{\circ}, k)}{\operatorname{arg\,min}} \Big\{ \frac{1}{2} \sum_{i=1}^{\delta^{-1}} (\partial_{\delta, i}(k) - \partial_{\delta, i}^{0})^{2} I_{\delta, i} - \sum_{i=1}^{\delta^{-1}} (\partial_{\delta, i}(k) - \partial_{\delta, i}^{0}) \mathcal{M}_{\delta, i} - \partial_{\delta, k_{0}}(k) R_{\delta, k_{0}}(\partial_{\circ}^{0}) \Big\}, \end{split}$$

for

$$M_{\delta,i} := \int_0^T X_{\delta,i}^{\Delta}(t) \, \mathrm{d}B_{\delta,i}(t), \quad I_{\delta,i} := \int_0^T X_{\delta,i}^{\Delta}(t)^2 \, \mathrm{d}t,$$

and $R_{\delta,k_0}(\vartheta^0_\circ)$ is an error term resulting from $K_{\delta,k_0} \notin D(\Delta_{\vartheta})$ in general

Basic estimates

Lemma (Reiß, Strauch and T., 2023+)

• For any $i \in [\delta^{-1}] \setminus \{k_0\}$,

$$\mathbb{E}[I_{\delta,i}] = \frac{T}{2\vartheta(x_i)} \|K'\|_{L^2}^2 \delta^{-2} + \mathcal{O}(1),$$

and, moreover, $\mathbb{E}[I_{\delta,k_0}] \sim \delta^{-2}$;

• for any vector $\alpha \in \mathbb{R}^n$ s.t. $\alpha_{k_0} = 0$,

$$\operatorname{Var}\left(\sum_{i=1}^{\delta^{-1}} \alpha_i I_{\delta,i}\right) \leq \frac{T}{2\underline{\vartheta}^3} \delta^{-2} \|\alpha\|_{\ell^2}^2 \|K'\|_{L^2}^2;$$

$$\mathbb{E}[|R_{\delta,k_0}(\theta_{\circ})|] \lesssim \delta^{-2}, \quad Var(R_{\delta,k_0}(\theta_{\circ})) \lesssim \delta^{-2},$$

and, moreover,

$$\exists \, \vartheta_{\circ}^{0} \, : \quad |\mathbb{E}[R_{\delta,k_{0}}(\vartheta_{\circ}^{0})]| \leq \delta^{-1}.$$

Concentration results

 $\sum_{i=1}^{\delta^{-1}} \alpha_i(I_{\delta,i} - \mathbb{E}[I_{\delta,i}])$ belongs to second Wiener chaos for an isonormal Gaussian process associated to $(X_i^{\Delta}(t))_{t \in [0,T], i \in [\delta^{-1}]} \Rightarrow$ relate to Bernstein-type concentration result from Nourdin and Viens (2009)¹

Proposition (Reiß, Strauch and T., 2023+)

Let $\alpha \in \mathbb{R}^n_+ \setminus \{0\}$ s.t. $\alpha_{k_0} = 0$. Then, for any z > 0, we have

$$\mathbb{P}\Big(\Big|\sum_{i=1}^{n} \alpha_{i}(I_{\delta,i} - \mathbb{E}[I_{\delta,i}])\Big| \geq z\Big) \leq 2 \exp\left(-\frac{\underline{\vartheta}^{2}}{4\|\alpha\|_{\infty}} \frac{z^{2}}{z + \sum_{i=1}^{n} \alpha_{i} \mathbb{E}[I_{\delta,i}]}\right).$$

¹Nourdin, I., and F.G. Viens (2009). Density formula and concentration inequalities with Malliavin calculus. *Electron. J. Prob.*

Concentration results

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Proposition (Reiß, Strauch and T., 2023+)

Let

$$\overline{M}_{\delta,i} := \int_0^{\sigma_i} X_{\delta,i}^{\Delta}(t) \, \mathrm{d}B_{\delta,i}(t), \quad \text{where } \sigma_i := \inf \Big\{ t \geq 0 \, : \, \int_0^t X_{\delta,i}(s)^2 \, \mathrm{d}s > \mathbb{E}[I_{\delta,i}] \Big\}.$$

Then, $(\overline{M}_{\delta,i})_{i \in [\delta^{-1}]} \sim N(0, \operatorname{diag}((\mathbb{E}[I_{\delta,i}])_{i \in [\delta^{-1}]}))$ and

$$\mathbb{P}\Big(\Big|\sum_{i=1}^{n} \alpha_{i} (M_{\delta,i} - \overline{M}_{\delta,i})\Big| \ge z, \sum_{i=1}^{n} \alpha_{i}^{2} |I_{\delta,i} - \mathbb{E}[I_{\delta,i}]| \le L\Big) \le \exp(-z^{2}/2L), \quad \alpha \in \mathbb{R}^{n}, z, L > 0$$

Rate of convergence

Define the jump height $\eta := \partial_+^0 - \partial_-^0$.

Theorem (Reiß, Strauch and T., 2023+)

Suppose that $\partial_{\pm}^0 \underset{\delta \to 0}{\longrightarrow} \partial_{\pm}^*$ and that $|\eta| \ge \underline{\eta} > 0$ for all $\delta \in 1/\mathbb{N}$. Then,

$$|\hat{\tau} - \tau^0| = \mathcal{O}_{\mathbb{P}}(\delta)$$
 and $|\hat{\partial}_{\pm} - \partial_{\pm}^0| = \mathcal{O}_{\mathbb{P}}(\delta^{3/2})$.

- the estimation rate for τ^0 cannot be improved due to discretisation effects
- the estimation rate for ϑ^0_\pm is the minimax optimal rate for parametric estimation from multiple local measurements in the model $A_\vartheta=\vartheta\Delta$ without change point¹

¹Altmeyer, R., Tiepner, A. and M. Wahl (2024). Optimal parameter estimation for linear SPDEs from multiple measurements. *Ann. Stat.*

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Proof outline:

- 1. verify basic consistency of $(\hat{\partial}_{\pm},\hat{\tau})$
- 2. determine appropriate empirical process $(\mathcal{L}_{\delta})_{\delta \in 1/\mathbb{N}}$ with $[\underline{\vartheta}, \overline{\vartheta}]^3 \times (0, 1] \ni \chi \mapsto \mathcal{L}_{\delta}(\chi)$ such that

$$(\hat{\partial}_{-}, \hat{\partial}_{+}, \hat{\partial}_{\circ}, \hat{\tau}) \in \underset{\chi \in [\underline{\partial}, \overline{\partial}]^{3} \times (0,1]}{\arg \min} \mathcal{L}_{\delta}(\chi)$$

- 3. control local fluctuations of centered empirical process $\mathcal{L}_{\delta} \widetilde{\mathcal{L}}_{\delta}(\chi)$ around χ^{0} , where $\widetilde{\mathcal{L}}_{\delta}(\chi) = \mathbb{E}[\mathcal{L}_{\delta}(\chi)] + \mathcal{O}(\delta^{2})$
- 4. exploit (non-standard) peeling device to prove convergence rate

Vanishing signal

- for the previous consistency result it was crucial that the jump height η does not vanish
- assume now that $\eta \longrightarrow 0$ and that $\vartheta_{\pm}^0 = \vartheta_{\pm}^0(\delta)$ are known
- set $\hat{\tau} = \hat{k}\delta$, where

$$\begin{split} \hat{k} &\coloneqq \underset{k=1,\dots,\delta^{-1}}{\text{arg max}} \sum_{i=1}^k \left(\vartheta_-^0 \int_0^T X_{\delta,i}^\Delta(t) \, \mathrm{d} X_{\delta,i}(t) - \frac{(\vartheta_-^0)^2}{2} \int_0^T X_{\delta,i}^\Delta(t)^2 \, \mathrm{d} t \right) \\ &\quad + \sum_{i=k+1}^{\delta^{-1}} \left(\vartheta_+^0 \int_0^T X_{\delta,i}^\Delta(t) \, \mathrm{d} X_{\delta,i}(t) - \frac{(\vartheta_+^0)^2}{2} \int_0^T X_{\delta,i}^\Delta(t)^2 \, \mathrm{d} t \right) \\ &= \underset{k=1,\dots,\delta^{-1}}{\text{arg min}} \, Z_k, \end{split}$$

for

$$Z_k = \begin{cases} 0, & k = k_0, \\ -\eta \sum_{i=k+1}^{k_0} \int_0^T X_{\delta,i}^{\Delta}(t) \, \mathrm{d}B_{\delta,i}(t) + \frac{\eta^2}{2} \sum_{i=k+1}^{k_0} \int_0^T X_{\delta,i}^{\Delta}(t)^2 \, \mathrm{d}t + \eta R_{\delta,k_0}(\theta_-^0), & k < k_0, \\ \eta \sum_{i=k_0+1}^k \int_0^T X_{\delta,i}^{\Delta}(t) \, \mathrm{d}B_{\delta,i}(t) + \frac{\eta^2}{2} \sum_{i=k_0+1}^k \int_0^T X_{\delta,i}^{\Delta}(t)^2 \, \mathrm{d}t, & k > k_0, \end{cases}$$

Limit theorem for vanishing signal

Reformulate the estimator again in terms of an M-estimator: Let $v_{\delta} \rightarrow 0$, and define

$$M_{T,\delta}^{\tau^0}(h) = M_{T,\delta}(\tau^0 + hv_{\delta}) - M_{T,\delta}(\tau^0), \quad \text{for } M_{T,\delta}(z) := \sum_{i=1}^{\lfloor z/\delta \rfloor} M_{\delta,i}, \ z \in [0,1]$$

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s.t.

$$\mathcal{Z}_{\delta}(v_{\delta}^{-1}(\hat{\tau}-\tau^0)) = \min_{h \in [-\tau_0/v_{\delta},(1-\tau^0)/v_{\delta}]} \mathcal{Z}_{\delta}(h) + \mathcal{O}_{\mathbb{P}}(\eta^2\delta^{-2}), \quad \text{for } \mathcal{Z}_{\delta}(h) := \eta \mathcal{M}_{T,\delta}^{\tau^0}(h) + \frac{\eta^2}{2} I_{T,\delta}^{\tau^0}(h)$$

Limit theorem for vanishing signal

Reformulate the estimator again in terms of an M-estimator: Let $v_{\delta} \rightarrow 0$, and define

$$M_{T,\delta}^{\tau^0}(h) = M_{T,\delta}(\tau^0 + h\nu_\delta) - M_{T,\delta}(\tau^0), \quad \text{for } M_{T,\delta}(z) := \sum_{i=1}^{\lceil z/\delta \rceil} M_{\delta,i}, \ z \in [0,1]$$

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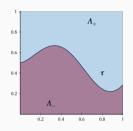
$$\mathcal{Z}_{\delta}(v_{\delta}^{-1}(\hat{\tau}-\tau^0)) = \min_{h \in [-\tau_0/v_{\delta},(1-\tau^0)/v_{\delta}]} \mathcal{Z}_{\delta}(h) + \mathcal{O}_{\mathbb{P}}(\eta^2\delta^{-2}), \quad \text{for } \mathcal{Z}_{\delta}(h) \coloneqq \eta \mathcal{M}_{T,\delta}^{\tau^0}(h) + \frac{\eta^2}{2} I_{T,\delta}^{\tau^0}(h)$$

Theorem (Reiß, Strauch and T., 2023+)

Assume $\eta = o(\delta)$ and $\delta^{3/2} = o(\eta)$. Then, for a two-sided Brownian motion ($B^{\leftrightarrow}(h)$, $h \in \mathbb{R}$), we have

$$\underbrace{\frac{\eta^2}{\delta^3}}_{=V_\delta^{-1}} \frac{T \|K'\|_{L^2}^2}{2\vartheta^*} (\hat{\tau} - \tau^0) \stackrel{\mathsf{d}}{\longrightarrow} \underset{h \in \mathbb{R}}{\operatorname{arg\,min}} \Big\{ B^{\leftrightarrow}(h) + \frac{|h|}{2} \Big\}, \quad \text{as } \delta \to 0.$$

The multivariate case



Recall:

$$dX(t) = \Delta_{\vartheta}X(t) dt + dW(t), \quad \Delta_{\vartheta} = \nabla \cdot \vartheta \nabla,$$

with

$$\vartheta(x) = \vartheta_{-} \mathbf{1}_{\Lambda_{-}}(x) + \vartheta_{+} \mathbf{1}_{\Lambda_{+}}(x), \quad x \in [0, 1]^{d} = \Lambda_{-} \uplus \Lambda_{+}, \vartheta_{-} \wedge \vartheta_{+} > 0.$$

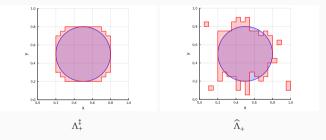
- we call Λ_+ a change domain
- · structural similarities to image reconstruction problem

$$Y_i = \partial_{-} \mathbf{1}_{\Lambda_{-}}(X_i) + \partial_{+} \mathbf{1}_{\Lambda_{+}}(X_i) + \varepsilon_i,$$

for (possibly random) measurement locations X_i and noise ε_i

Local observations

• put regular δ -grid on $[0,1]^d$ with grid centers x_α , $\alpha \in [n]^d = [\delta^{-1}]^d$ and aim for estimation of minimal tiling Λ^{\updownarrow}_+ of Λ^0_+



- set $K_{\delta,\alpha} = \delta^{-d/2} K((\cdot x_{\alpha})/\delta)$
- local observations $X_{\delta,\alpha}(t) = \langle X(t), K_{\delta,\alpha} \rangle$ and $X_{\delta,\alpha}^{\Delta}(t) = \langle X(t), \Delta K_{\delta,\alpha} \rangle$ given for $\alpha \in [n]^d, t \in [0, T]$

Estimation approach

- \mathcal{A}_+ is a family of tiling sets such that $\Lambda_+^{\updownarrow} \in \mathcal{A}_+$
- Θ_{\pm} are η -separated sets such that $\vartheta_{\pm}^0 \in \Theta_{\pm}$
- set

$$(\hat{\partial}_{-},\hat{\partial}_{+},\widehat{\Lambda}_{+}) \in \mathop{\arg\max}_{(\partial_{-},\partial_{+},\Lambda_{+}) \in \Theta_{-} \times \Theta_{+} \times \mathcal{A}_{+}} \sum_{\alpha \in [n]^{d}} \ell_{\delta,\alpha}(\partial_{-},\partial_{+},\Lambda_{+}),$$

where

$$\ell_{\delta,\alpha}(\theta_-,\theta_+,\Lambda_+) = \theta_{\delta,\alpha}(\Lambda_+) \int_0^T X_{\delta,\alpha}^{\Delta}(t) \, dX_{\delta,\alpha}(t) - \frac{\theta_{\delta,\alpha}(\Lambda_+)^2}{2} \int_0^T X_{\delta,\alpha}^{\Delta}(t)^2 \, dt,$$

for

$$\vartheta_{\delta,\alpha}(\Lambda_{+}) = \begin{cases} \vartheta_{+}, & x_{\alpha} \in \Lambda_{+}, \\ \vartheta_{-}, & \text{else.} \end{cases}$$

Convergence rate

Theorem (Tiepner and T., 2025+)

Suppose that the number of hypercubes intersecting $\partial \Lambda^0_+$ is of order $\delta^{-d+\beta}$ for some $\beta \in (0,1]$. Then,

$$\mathbb{E}\big[\operatorname{vol}_d(\widehat{\Lambda}_+ \triangle \Lambda^0_+)] \lesssim \delta^{\beta},$$

and $\hat{\vartheta}_{\pm}$ are consistent.

In particular, if

- $\bullet \ \ \Lambda^0_+ \ \text{epigraph of a} \ \beta\text{-H\"older function} \ \Longrightarrow \ \mathbb{E}\big[\operatorname{vol}_d(\widehat{\Lambda}_+ \bigtriangleup \Lambda^0_+)] = \mathbb{E}\big[\|\widehat{\tau} \tau^0\|_{L^1}\big] \lesssim \delta^\beta;$
- $\Lambda^0_+ \operatorname{convex} \implies \mathbb{E} [\operatorname{vol}_d(\widehat{\Lambda}_+ \triangle \Lambda^0_+)] \lesssim \delta$

and we can choose \mathcal{A}_+ s.t. in the first case $|\mathcal{A}_+| = \delta^{-d}$ and in the second case $|\mathcal{A}_+| = \delta^{-(d+1)}$.

Improved convergence results when there is no geometric bias

- assume now that Λ^0_+ can be perfectly built from hypercubes, i.e., $\Lambda^0_+ \in \mathcal{A}_+$
- for fixed $\Lambda_+ \in \mathcal{A}_+$, the function $\mathbb{R}^2 \ni (\partial_-, \partial_+) \mapsto \sum_{\alpha \in [n]^d} \ell_{\delta,\alpha}(\partial_-, \partial_+, \Lambda_+)$ is maximised in

$$\vartheta_{\pm}^{\Lambda_{+}} = \frac{\sum_{\mathsf{Sq}(\alpha)^{\circ} \subset \Lambda_{\pm}} \int_{0}^{T} X_{\delta,\alpha}^{\Delta}(t) \, \mathsf{d} X_{\delta,\alpha}(t)}{\sum_{\mathsf{Sq}(\alpha)^{\circ} \subset \Lambda_{\pm}} I_{\delta,\alpha}}$$

→ set

$$\widetilde{\Lambda}_{+} \in \operatorname*{arg\,max}_{\Lambda_{+} \in \mathcal{A}_{+}} \sum_{\alpha \in [n]^{d}} \ell_{\delta,\alpha} \big(\vartheta_{\pm}^{\Lambda_{+}}, \Lambda_{+} \big), \quad \widetilde{\Lambda}_{-} = \widetilde{\Lambda}_{+}^{c}, \quad \widetilde{\vartheta}_{\pm} = \vartheta_{\pm}^{\widetilde{\Lambda}_{+}}$$

and (for identification purposes)

$$\widehat{\Lambda}_{+}^{*} = \begin{cases} \widetilde{\Lambda}_{+}, & \text{if } \operatorname{vol}(\widetilde{\Lambda}_{+} \cap \widehat{\Lambda}_{+}) \geq \operatorname{vol}(\widetilde{\Lambda}_{+} \cap \widehat{\Lambda}_{-}) \\ \operatorname{cl} \widetilde{\Lambda}_{-}, & \text{else} \end{cases}, \quad \widehat{\vartheta}_{\pm}^{*} = \begin{cases} \widetilde{\vartheta}_{\pm}, & \text{if } \widehat{\Lambda}_{+}^{*} = \widetilde{\Lambda}_{\pm}, \\ \widetilde{\vartheta}_{\mp}, & \text{else} \end{cases}$$

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$$\widetilde{\Lambda}_{+} \in \operatorname*{arg\,max}_{\Lambda_{+} \in \mathcal{A}_{+}} \sum_{\alpha \in [n]^{d}} \ell_{\delta,\alpha} \big(\vartheta_{\pm}^{\Lambda_{+}}, \Lambda_{+} \big), \quad \widetilde{\Lambda}_{-} = \widetilde{\Lambda}_{+}^{c}, \quad \widetilde{\vartheta}_{\pm} = \vartheta_{\pm}^{\widetilde{\Lambda}_{+}}, \quad \text{for } \vartheta_{\pm}^{\Lambda_{+}} = \frac{\sum_{\operatorname{Sq}(\alpha)^{\circ} \subset \Lambda_{\pm}} \int_{0}^{T} X_{\delta,\alpha}^{\underline{\Lambda}}(t) \, \mathrm{d}X_{\delta,\alpha}(t)}{\sum_{\operatorname{Sq}(\alpha)^{\circ} \subset \Lambda_{\pm}} I_{\delta,\alpha}}$$

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Theorem (Tiepner and T., 2025+)

Suppose that $\overline{\Lambda_+^0} \in \mathcal{A}_+$, $|\mathcal{A}_+| \leq \delta^{-c}$ for some c > 0 and $\lim_{\delta \to 0} \operatorname{vol}(\Lambda_\pm^0) = v_\pm^0 > 0$. Then,

$$\lim_{\delta \to 0} \mathbb{P} \Big(\widehat{\Lambda}_+^* = \Lambda_+^0 \Big) = 1$$

and

$$\delta^{-(d/2+1)} (\hat{\vartheta}_{\pm}^* - \vartheta_{\pm}^0) \stackrel{d}{\longrightarrow} \mathcal{N} \left(0, \frac{2\vartheta_{\pm}^0}{T \|\nabla K\|_2^2, \gamma_{\pm}^0}\right).$$

If cl $\Lambda_{-}^{0} \notin \mathcal{A}_{+}$ the above continues to hold with $\widehat{\Lambda}_{+}^{*}$ and $\widehat{\vartheta}_{\pm}$ replaced by $\widetilde{\Lambda}_{+}$ and $\widetilde{\vartheta}_{\pm}$, resp.

Improved convergence results when there is no geometric bias

Theorem (Tiepner and T., 2025+)

Suppose that $\overline{\Lambda_+^0} \in \mathcal{A}_+$, $|\mathcal{A}_+| \lesssim \delta^{-c}$ for some c > 0 and $\lim_{\delta \to 0} \text{vol}(\Lambda_\pm^0) = \nu_\pm^0 > 0$. Then,

$$\lim_{\delta \to 0} \mathbb{P} \Big(\widehat{\Lambda}_+^* = \Lambda_+^0 \Big) = 1$$

and

$$\delta^{-(d/2+1)}(\hat{\vartheta}_{\pm}^* - \vartheta_{\pm}^0) \stackrel{d}{\longrightarrow} \mathcal{N}\left(0, \frac{2\vartheta_{\pm}^0}{T\|\nabla K\|_{L^2}^2 \nu_{\pm}^0}\right).$$

If cl $\Lambda_{-}^{0} \notin \mathcal{A}_{+}$ the above continues to hold with $\widehat{\Lambda}_{+}^{*}$ and $\widehat{\vartheta}_{\pm}$ replaced by $\widetilde{\Lambda}_{+}$ and $\widetilde{\vartheta}_{\pm}$, resp.

Thank you for your attention!