HW4 Luke Beebe 2024-04-06 5.1 T/F An interaction effect in the model from a factorial experiment involving quantitative factors is a way of incorporating curvature into the response surface model representation of the results. False, curvature can be represented by adding higher-order terms. 5.2 A factorial experiment may be conducted as a RCBD by running each replicate of the experiment in a unique block **True**, this helps to control potential sources of variation. 5.3 If an interaction effect in a factorial experiment is significant, the main effects of the factors involved in that interaction are difficult to interpret individually. True, interpreting the main effects of the individual factors is challenging because the effect of one factor depends on the level of the other factor(s) involved. 5.7 The following output was obtained from a computer program that performed a two-factor ANOVA on a factorial experiment. Two-way ANOVA: y versus A, B DF SS MS Source 0.0002 180.378 8.479 0.932 Interaction ? ? 158.797 ? Error 347.653 Total 15 a. Fill in the blanks in the ANOVA table. You can use bounds on the P-values. Two-way ANOVA: y versus A, B DΕ SS MS Ρ Source ? 0.0002 0.0002 20.00001 20.998 Α 180.378 20.094 В ?3 ? 60.126 **?**3.029 8.479 **?0.1420.932** ? 2.826 Interaction 158.797 **?**19.8496 Error 15 347.653 Total b. How many levels were used for factor B? 4 c. How many replicates of the experiment were performed? d. What conclusions would you draw about this experiment? We do not reject the null. None of the factors are statistically significant from one another. 5.9 An engineer suspects that the surface finish of a metal part is influenced by the feed rate and the depth of cut. He selects three feed rates and four depths of cut. He then conducts a factorial experiment and obtains the following data: surface <-c(74,79,82,99,64,68,88,104, 60,73,92,96, 92,98,99,104, 86,104,108,110, 88,88,95,99, 99,104,108,114, 98,99,110,111, 102,95,99,107) feed_rate <- c(0.2,0.2,0.2,0.2, 0.2,0.2,0.2,0.2, 0.2,0.2,0.2,0.2, 0.25, 0.25, 0.25, 0.25, 0.25,0.25,0.25,0.25, 0.25, 0.25, 0.25, 0.25, 0.3,0.3,0.3,0.3, 0.3,0.3,0.3,0.3, 0.3,0.3,0.3,0.3) depth <-c(0.15, 0.18, 0.2, 0.25,0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15, 0.18, 0.2, 0.25data <- data.frame(surface, feed_rate, depth)</pre> a. Analyze the data and draw conclusions. Use $\alpha = 0.05$.

model <- aov(surface ~ factor(feed rate)*factor(depth), data)</pre> summary(model) Df Sum Sq Mean Sq F value Pr(>F) ## factor(feed rate) 2 3160.5 1580.2 55.018 1.09e-09 *** ## factor(depth) 3 2125.1 708.4 24.663 1.65e-07 *** ## factor(feed rate):factor(depth) 6 557.1 3.232 0.018 * 92.8 ## Residuals 24 689.3 28.7 ## ---## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 The feed_rate, depth, and interaction term feed_rate:depth are all significant.

01 180 0 0 2 0 0

Residuals vs Fitted

b. Prepare appropriate residual plots and comment on the model's adequacy.

80

plot(model)

10

0

-10

70

0

0

80

70

9.0

0.4

0.2

0.0

df = NULL

for(rate in unique(feed_rate)){

d. Find the P-values for the tests in part (a).

200, 215, 230, 200, 215, 230,

200, 215, 230, 200, 215, 230) data <- data.frame(yield, temp, pressure)</pre>

TukeyHSD(model)\$`factor(pressure)`

strength < c(196.6,197.7,199.8,

5.23

diff

model <- aov(yield ~ factor(temp)*factor(pressure), data)</pre>

215-200 0.3166667 0.1017380 0.53159536 0.0066518030 ## 230-200 -0.1833333 -0.3982620 0.03159536 0.0944905140 ## 230-215 -0.5000000 -0.7149287 -0.28507131 0.0002951232

The pressure of **215-200** and **230-215** are significantly different.

factor(hardwood_conc):factor(cooking_time)

factor(hardwood_conc):factor(pressure) ## factor(pressure):factor(cooking_time)

factor(hardwood_conc):factor(cooking_time) .

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

b. Prepare appropriate residual plots and comment on the model's adequacy.

The hardwood_conc, pressure, cooking_time, and hardwood_conc:pressure are all significant.

Residuals

Residuals

plot(model)

-7

4.

1.2

1.0

0.8

9.0

0.4

0.2

0.0

this experiment?

Residuals

the blocks.

31 965.9 31.2

Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

surface <-c(74,79,82,99,

64,68,88,104, 60,73,92,96,

92,98,99,104,

88,88,95,99,

86,104,108,110,

99,104,108,114,

0

0

/Standardized residuals

05

-2

-1

05

0

0

0

0

07

0

factor(hardwood_conc)

factor(cooking_time)

factor(pressure)

lwr

experiment with two replicates is conducted, and the following data are obtained:

feed_rate, p = 1.09e-09

depth, p = 1.65e-07

5.10

in∕min.

interaction, p = 0.018

Residuals 00 0 0 00 0 -5 0 0 0 0 0

220

100

00

0

100

0

0

110

0

0

00

90

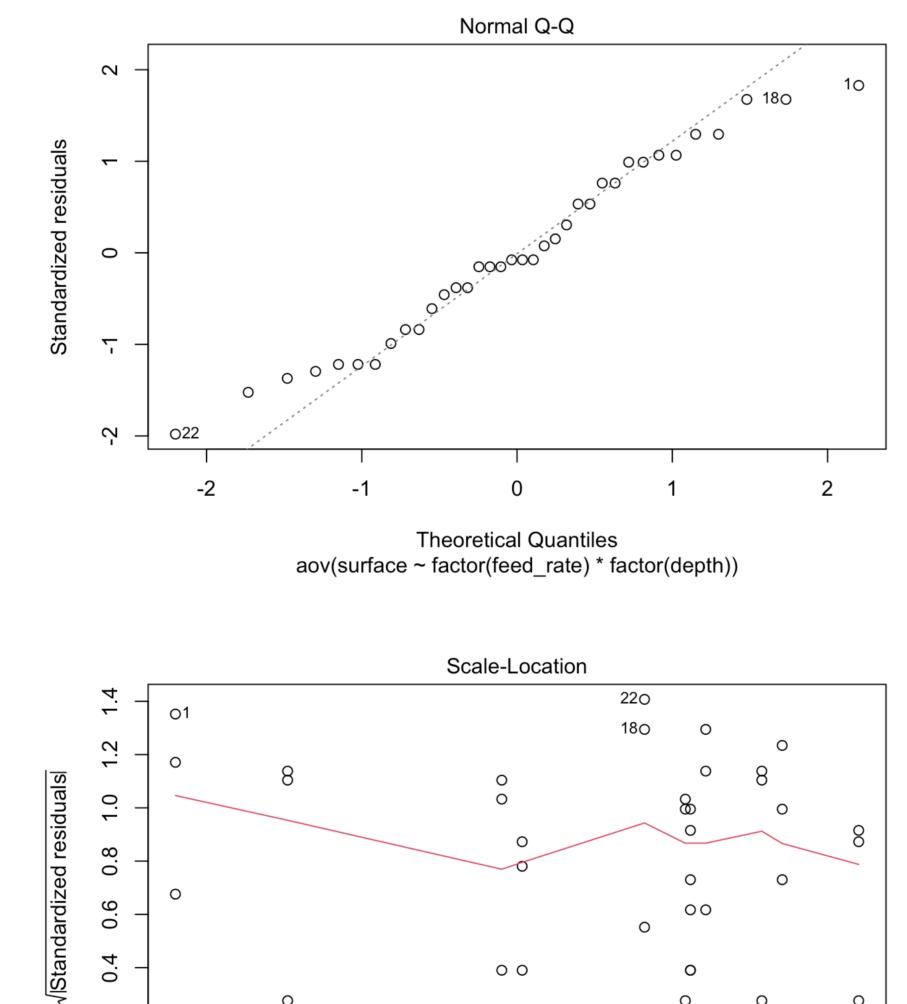
90

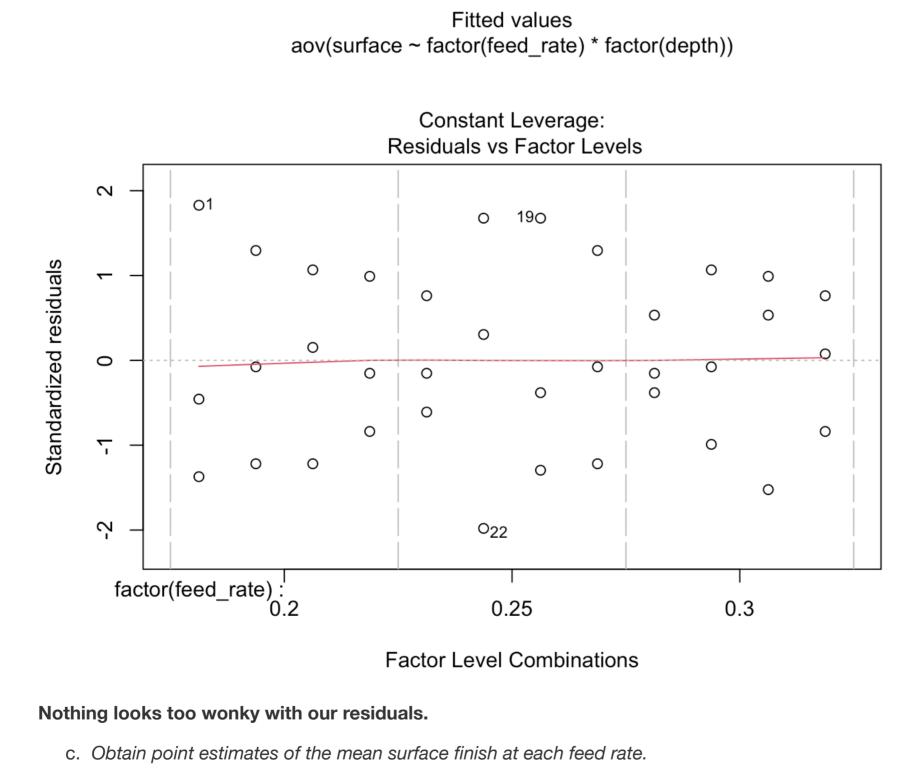
Fitted values aov(surface ~ factor(feed_rate) * factor(depth))

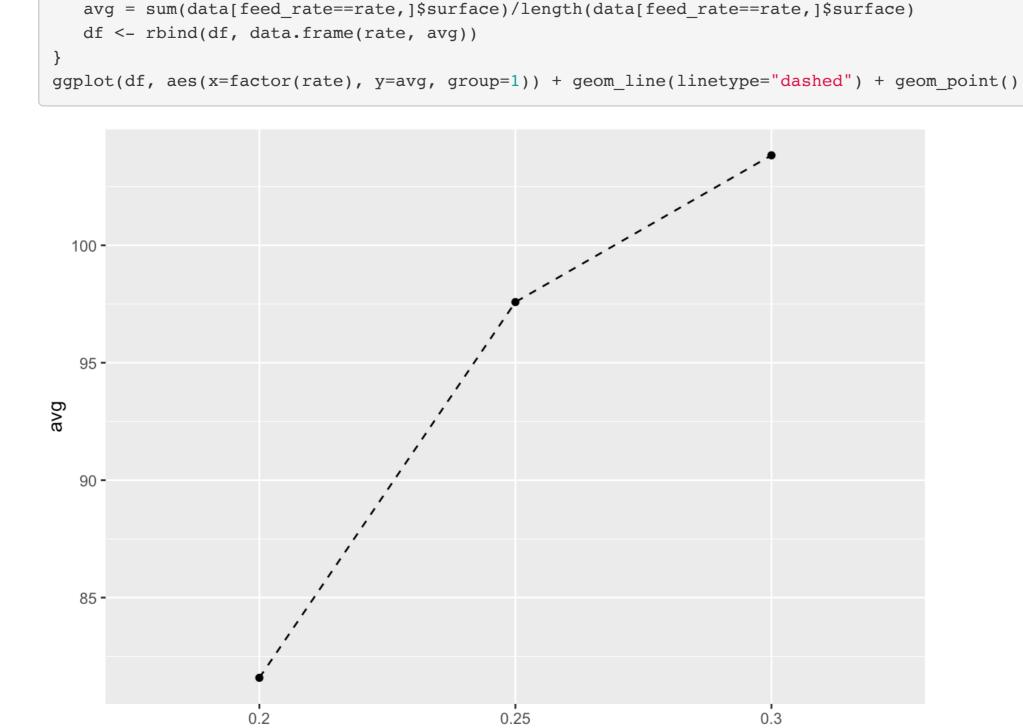
0

0

110

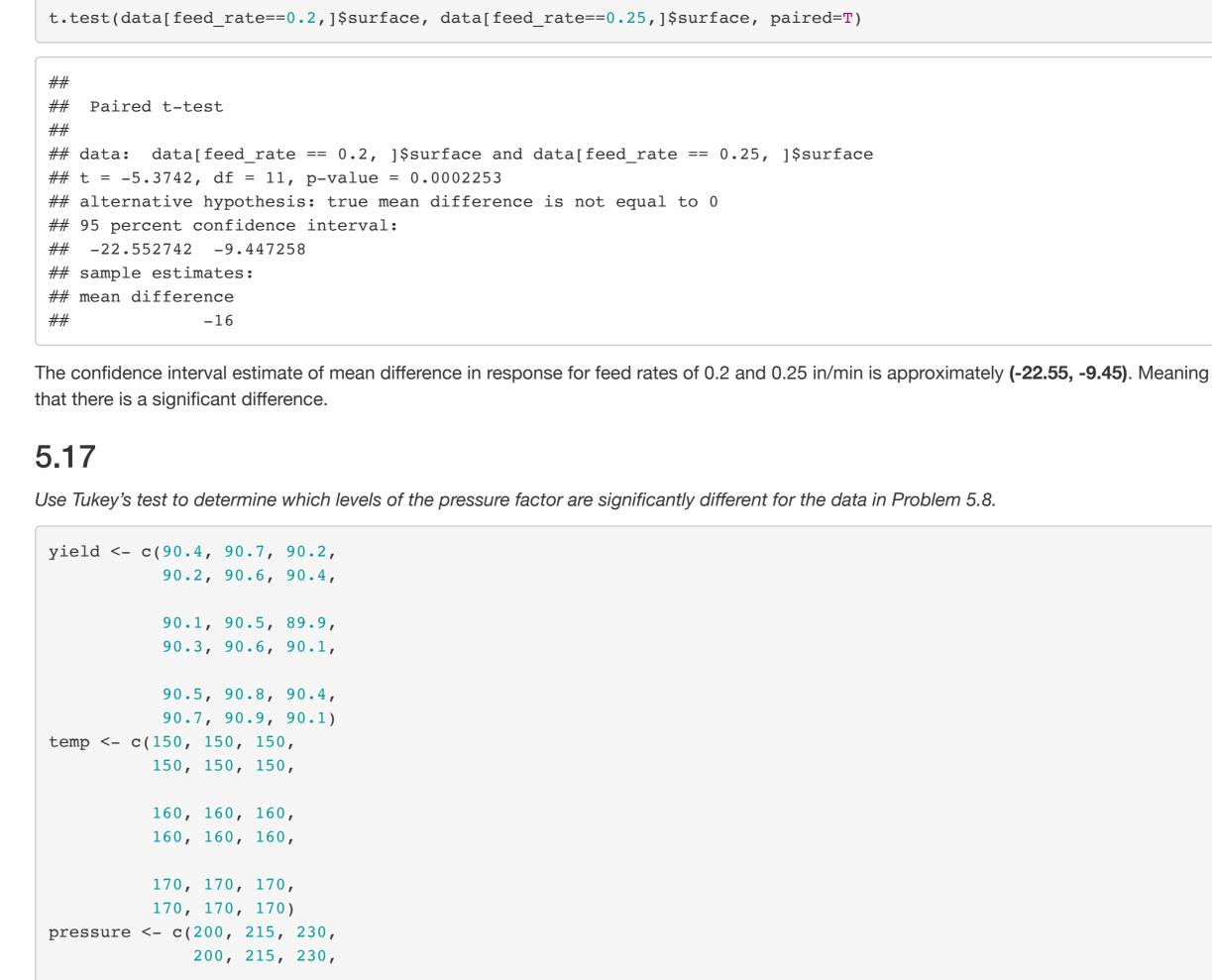


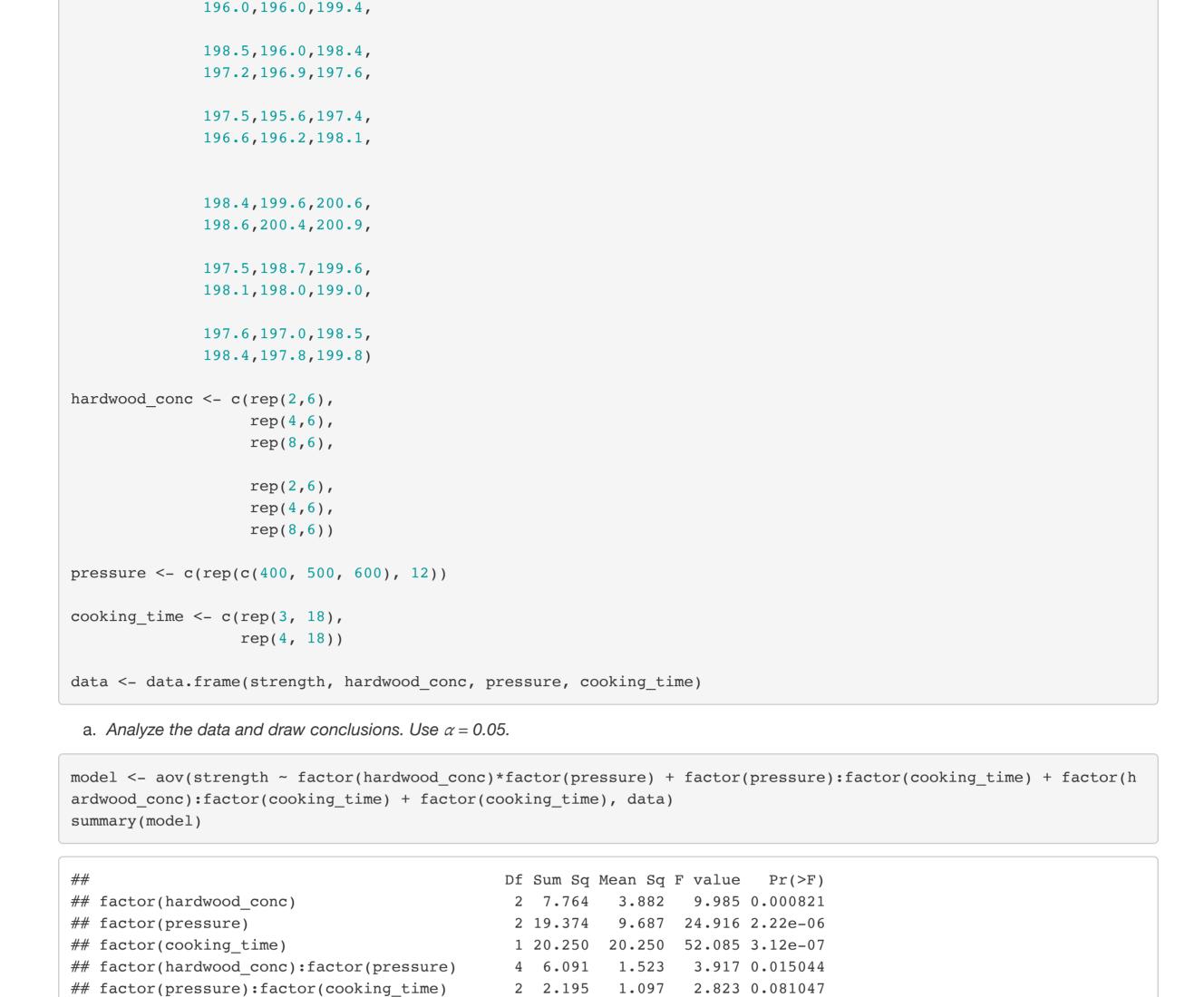




factor(rate)

For the data in Problem 5.9, compute a 95 percent confidence interval estimate of the mean difference in response for feed rates of 0.20 and 0.25





2 2.082

22 8.553

1.041

0.389

2.677 0.091070

p adj

The percentage of hardwood concentration in raw pulp, the vat pressure, and the cooking time of the pulp are being investigated for their effects

on the strength of paper. Three levels of hardwood concentration, three levels of pressure, and two cooking times are selected. A factorial

upr

Residuals vs Fitted 1.0 07 360 0 0.5 0 0 Residuals 0 0 0.0 0 0 0 0 0 -0.5 0 0 00 -1.0 05 196 198 200 201 197 Fitted values aov(strength ~ factor(hardwood_conc) * factor(pressure) + factor(pressure): ... Normal Q-Q 70 7 Standardized residuals 0 7

2

0

0

0

Theoretical Quantiles aov(strength ~ factor(hardwood_conc) * factor(pressure) + factor(pressure): ...

Scale-Location

00

0

00

0

0

0

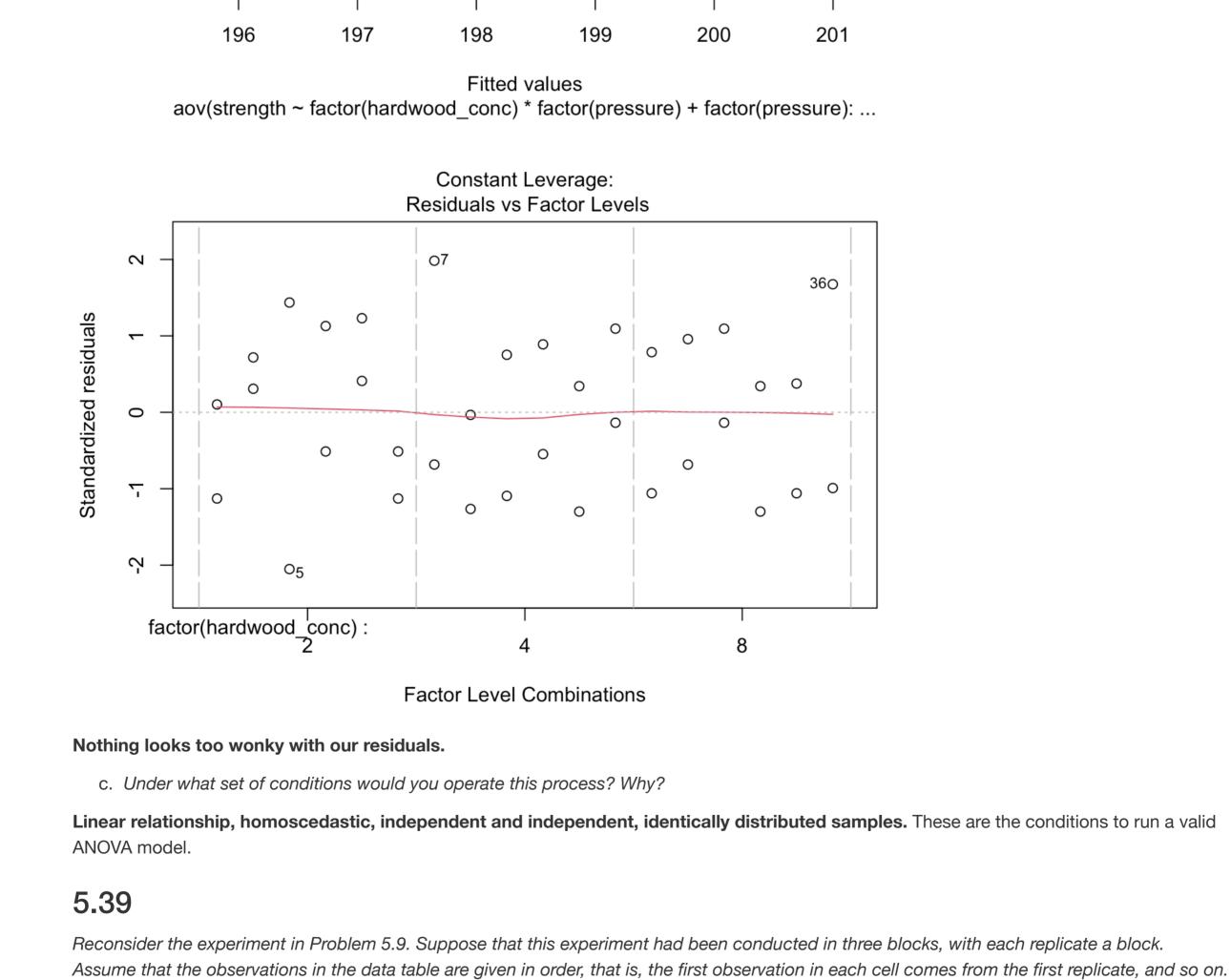
360

0

0

0

0



98,99,110,111, 102,95,99,107) feed_rate <- c(0.2, 0.2, 0.2, 0.2,0.2,0.2,0.2,0.2, 0.2,0.2,0.2,0.2, 0.25,0.25,0.25,0.25,

Reanalyze the data as a factorial experiment in blocks and estimate the variance component for blocks. Does it appear that blocking was useful in

0.25,0.25,0.25,0.25, 0.25,0.25,0.25,0.25, 0.3,0.3,0.3,0.3, 0.3,0.3,0.3,0.3, 0.3,0.3,0.3,0.3) depth <-c(0.15,0.18,0.2,0.25,0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15,0.18,0.2,0.25, 0.15, 0.18, 0.2, 0.25block <- c(rep(1,4),rep(2,4), rep(3,4), rep(1,4), rep(2,4), rep(3,4), rep(1,4), rep(2,4), rep(3,4)) data <- data.frame(surface, feed_rate, depth, block)</pre> model <- aov(surface ~ feed_rate*depth + block)</pre> summary(model) Df Sum Sq Mean Sq F value Pr(>F) ## feed_rate 1 2970.4 2970.4 95.328 5.66e-11 *** ## depth 1 2042.3 2042.3 65.543 3.84e-09 *** ## block 1 140.2 140.2 4.498 0.042029 * ## feed_rate:depth 1 413.2 413.2 13.262 0.000978 ***

Yes, it does appear that blocking was useful in this experiment. There is a statistically significant difference in variance between random error and