FISH 604 Module 4: Statistical estimation I

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Review of Module 3: Exploratory data analysis

- Visualizing data
- Assessing distributions
- Outlier detection
- Standardization
- Transformations
- Correlations



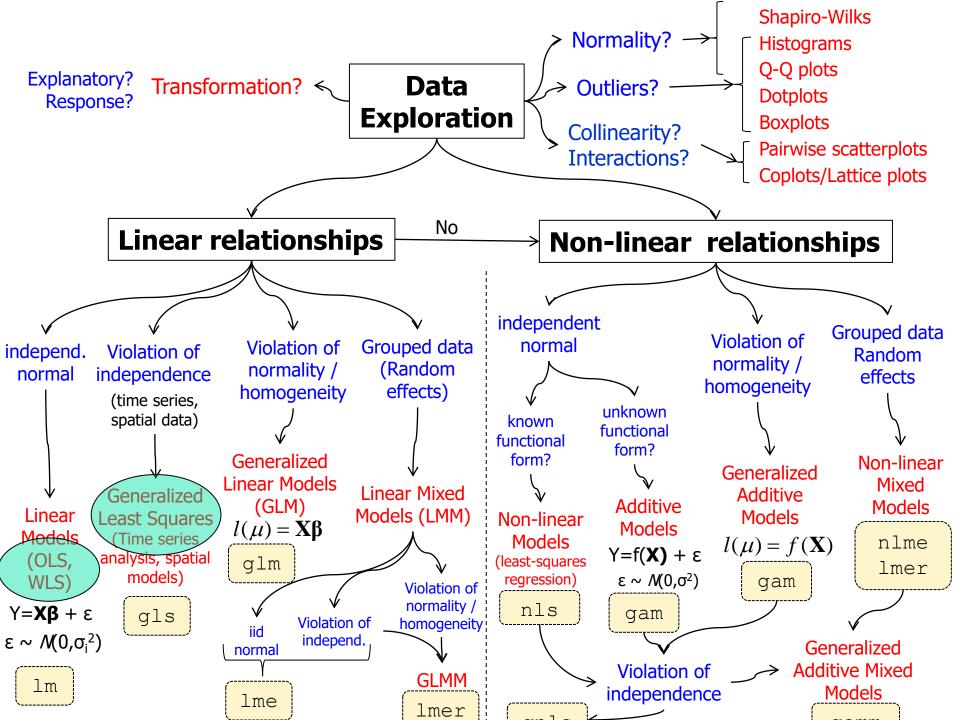
Preview - Student Project

- How to format your data?
 - Observations / samples (rows) x variables (columns)
 - Include any potentially relevant information
 - Aspects of sampling design (incl. spatial / temporal info)
 - Experimental design aspects (Treatment, Replicate)
 - Environmental covariates (measured, derived)

Case.ID	Chl. a	Station	Lat	Long	Location	Stratum	Year	Month	Season	DOY	
/ Sample	2										
1	12.6	FB1	58.25	-134.9	FunterB	surface	2018	3	Spring	86	
2	9.2	FB1	58.19	-134.9	FunterB	surface	2018	3	Spring	86	
3	23.2	FB2			FunterB						

•		
	•	••

Case.ID /	Depth	Temperature	Salinity	PAR	Weather	Fishery	Precip
Sample							(lagged)
1	7	2.3	26.3	12.6	rainy	open	23
2							
3							





Estimation methods

- Ordinary least-squares (OLS)
 Weighted least-squares (WLS)
 Generalized least-squares (GLS)

 - Maximum likelihood estimation (MLE)
 - (Bayesian estimation)
 - (Bootstrap estimation)



Objectives & Outcomes

- You should understand ...
 - ... the principle of least-squares regression
 - ... the difference between ordinary leastsquare regression, weighted LS regression, and generalized LS regression
 - ... potential pitfalls of non-linear LS regression
- You should be able to...
 - ...choose appropriate weights and do a weighted LS regression
 - ... fit a generalized LS regression model



Statistical estimation

- Methods for estimating model parameters
 - Ordinary least-squares (OLS)
 - Weighted least-squares (WLS)
 - Generalized least-squares (GLS)
 - Maximum likelihood estimation (MLE)
 - (Bayesian estimation)



Least-squares estimation

Method of fitting a curve (or surface)
to data points so as to minimize the sum
of the squares of the distances of the
points from the curve (or surface)

Some notation

Data: y_i, \mathbf{x}_i where: $\mathbf{x_i} = \{x_1, x_2, ..., x_p\}$

Parameters: α, β where: $\beta = \{\beta_1, \beta_2, ..., \beta_p\}$

Predicted values: $\hat{\alpha}, \hat{\beta}, \hat{\beta}_i, \hat{y}_i$



Ordinary least-squares (OLS)

Assumptions

- Independent, identically distributed errors (iid)!!
- Normality NOT required for fitting!
- Normality assumed for drawing inferences (testing coefficients & constructing confidence intervals)



Ordinary least-squares (OLS)

Minimize squared differences (residuals) between observed and expected values:

$$RSS = \sum_{i=1}^{n} (y_i - f(x_i))^2.$$
Data Model

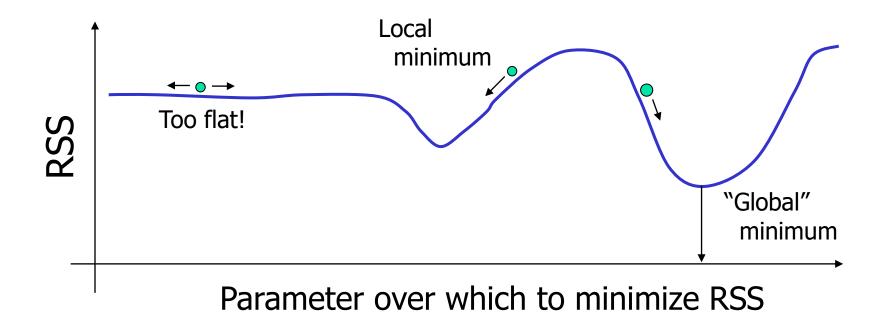
If the model f(.) is linear in the parameters, e.g.: $f(x) = a + b x + c x^{2}$

- the problem simplifies considerably (to a system of linear equations) and can be solved analytically
- If f(.) is not linear in the parameters, an algorithm for general optimization is used, such as Newton's method or gradient descent (iterative methods)



Ordinary least-squares (OLS)

- Non-linear minimization
 - Beware of local minima!
 - Use several different starting values!





Least-squares fitting in R

R functions for OLS (& WLS) estimation:

Function	Purpose				
lsfit	Find least-squares fit for linear models				
lm	Fits linear models via least-squares				
nls	Fits <u>non-linear</u> models via least-squares				
nlm	Function minimization via Newton-type algorithm (If f is RSS = least-squares)				
optim	General-purpose optimization (find min				
optimize	or max of <u>any</u> function using one of several algorithms).				



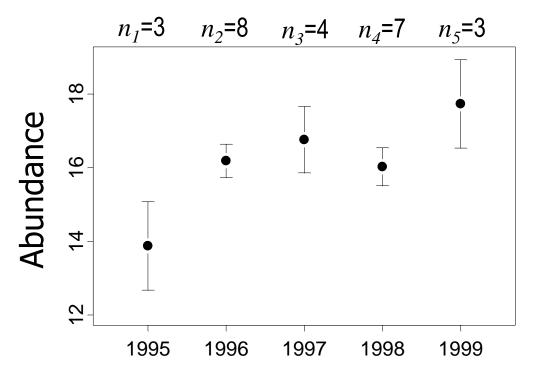
Method of regression similar to least squares in that it uses the same minimization of the sum of the residuals, but instead of weighting all points equally, they are weighted such that points with a greater weight contribute more to the fit:

$$RSS = \sum_{i=1}^{n} w_i (y_i - f(x_i))^2.$$

How to choose appropriate weights? \rightarrow Optimal weights are those that weigh each observation by the inverse of its variance, giving points with a lower variance (higher precision) a greater statistical weight: $w_i = 1/\sigma_i^2$.



- Main purpose: Deal with responses that have unequal variances!
- Example: What is the trend in abundances of scallops over time?



Each annual estimate based on surveys at n_i randomly selected sites

If estimates of variance $(\hat{\sigma}_i^2)$ are available, use:

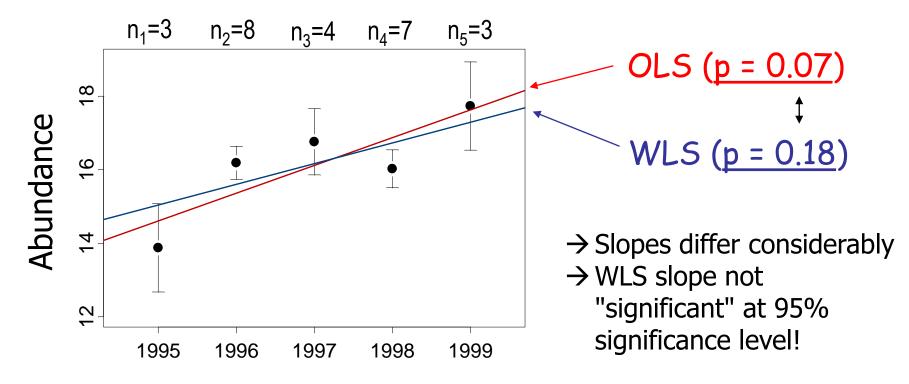
$$w_i = 1/\hat{\sigma}_i^2$$

Otherwise: $W_i = n_i$

(because $\hat{\sigma}_i^2$ is proportional to $1/n_i$)



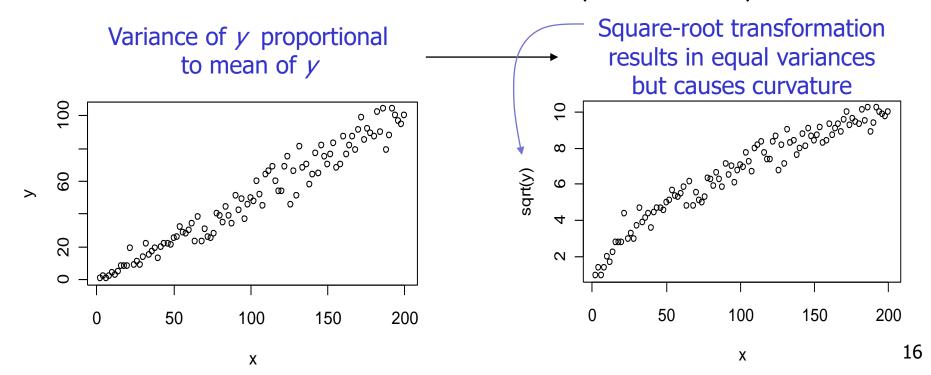
 Unequal variances can have strong influence on regression results!!





Weighting vs. transformation

- Both weighting and transformation of response variable deal with heteroscedasticity (unequal variances)
- Transformations may <u>also</u> change nature of relationship
- It may be difficult to find transformations (for x and/or y) that result in both homoscedasticity and linearity

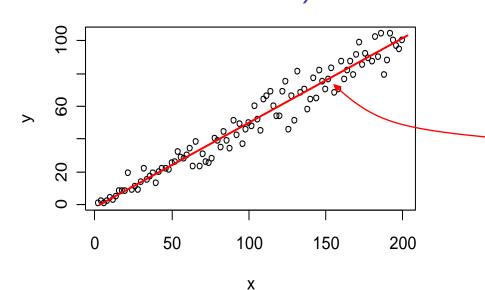




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Variance of y proportional to mean of y



→Use weighting approach

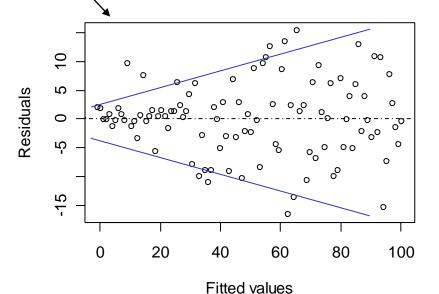
1. Start by fitting OLS regression and taking residuals



Weighting approach

- Fit OLS regression (see previous slide)
- Plot residual vs. fitted values (\hat{y})
- Determine relationship between $\hat{\sigma}^2$ and \hat{y} : $\hat{\sigma} = f(\hat{y})$
- 4. Choose appropriate weights
- 5. Fit WLS regression

Some common weighting schemes:



If <u>standard deviation</u> is proportional to fitted values:

$$\hat{\sigma} \sim \hat{y} \implies w_i = 1/\hat{y}_i^2$$

If <u>variance</u> is proportional to fitted values:

$$\hat{\sigma}^2 \sim \hat{y} \implies w_i = 1/\hat{y}_i$$



• Weights w_i essentially transform the y values such that they have equal variances (without distorting nature of relationship!):

$$\widetilde{y}_i = \sqrt{w_i} \cdot y_i$$

 $var(\tilde{y}_i) = constant$

Or, equivalently:

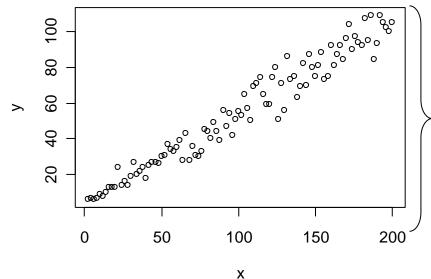
$$\widetilde{\mathcal{E}}_i = \sqrt{W_i} \cdot \mathcal{E}_i$$

$$var(\tilde{\varepsilon}_i) = constant$$



Pearson's residuals

Weighted least-squares (WLS)



Fitted values

R code for the case where σ^2 is proportional to \hat{y} :

(\hat{y} must be positive!!)

```
r <- resid(fit.wls, type = "pearson")

plot(fitted(fit.wls), r)

\sqrt{w_i \cdot \hat{\mathcal{E}}_i}
```



WLS fitting in R

- The main function for linear leastsquares regressions (1m) has a "weights = values" argument (as do most other model fitting functions)
- The main non-linear regression function (nls) also allows weights!

Or: use $nls(\sim sqrt(W)*(y-f(x)))$



- OLS assumes independent, identically distributed errors
- WLS assumes independent errors, but allows unequal variances
- GLS is a further generalization and allows both dependent errors and unequal variances
 - Dependent (correlated) errors produce standard errors that are too small and have fewer d.f. than expected (but have no effect on bias!)
 - Solution: specify correlation structure of errors and account for dependence AND unequal variances

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 Variance-covariance structure of errors is critical in least-squares fitting and is summarized by var-cov matrix:

$$\operatorname{var}(e_{i}) = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} & \cdots & \sigma_{1n} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} & \sigma_{24} & \sigma_{2n} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} & \sigma_{34} & \sigma_{3n} \\ \sigma_{41} & \sigma_{42} & \sigma_{43} & \sigma_{44} & \sigma_{4n} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{n1} & \sigma_{n2} & \sigma_{n3} & \sigma_{n4} & \cdots & \sigma_{nn} \end{bmatrix}$$



Independent, identically distributed errors (equal variances):

$$\operatorname{var}(e_i) = \begin{bmatrix} \sigma^2 & 0 & 0 & 0 \\ 0 & \sigma^2 & 0 & 0 \\ 0 & 0 & \sigma^2 & 0 \\ 0 & 0 & 0 & \sigma^2 \end{bmatrix} = \sigma^2 \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

Independent, unequal variances:

$$var(e_i) = \begin{bmatrix} \sigma_1^2 & 0 & 0 & 0 \\ 0 & \sigma_2^2 & 0 & 0 \\ 0 & 0 & \sigma_3^2 & 0 \\ 0 & 0 & 0 & \sigma_4^2 \end{bmatrix}$$

$$\operatorname{var}(e_i) = \begin{bmatrix} \sigma_1^2 & 0 & 0 & 0 \\ 0 & \sigma_2^2 & 0 & 0 \\ 0 & 0 & \sigma_3^2 & 0 \\ 0 & 0 & 0 & \sigma_4^2 \end{bmatrix} \qquad \begin{array}{c} \operatorname{Find \ weights,} \\ \operatorname{such \ that:} \\ \operatorname{var}(\sqrt{w_i} \ e_i) = \begin{bmatrix} \sigma^2 & 0 & 0 & 0 \\ 0 & \sigma^2 & 0 & 0 \\ 0 & 0 & \sigma^2 & 0 \\ 0 & 0 & 0 & \sigma^2 \end{bmatrix} \end{array}$$

$$\rightarrow$$
 Weighted regression: minimize $RSS_W = \sum_{i=1}^n w_i (y_i - \mu_i)^2$

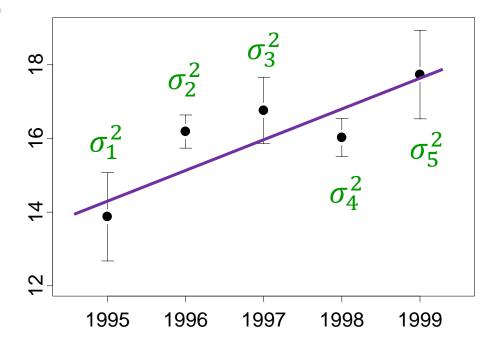


Example: Unequal residual variances

 Recall time series from above: Each residual has a different (but in this case) know variance.

Often, we make assumptions about the variance of residual i
 (e.g. its relationship with the mean) and then estimate it! If
 variance is constant, we simply estimate a constant residual

variance σ^2 (OLS)





Dependent errors, unequal variances:

$$var(e_i) = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} & \sigma_{24} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} & \sigma_{34} \\ \sigma_{41} & \sigma_{42} & \sigma_{43} & \sigma_{44} \end{bmatrix}$$
 (most general case)

<u>Typically</u>, <u>covariances are not all different</u> (to simplify further, we will look at <u>correlation matrix</u> instead):

Example 1: "adjacent" values have fixed correlation ϕ , which induces correlations at larger d'istances'

This is known as first-order autocorrelation (first-order autoregressive process) <u>if</u> observations are equidistant in time or space:

$$\mathcal{E}_t = \phi \mathcal{E}_{t-1} + \mathcal{V}_t$$
 where: $\mathcal{V}_t \sim N(0, \sigma^2)$

$$\rho_{e} = \begin{bmatrix} 1 & \phi & \phi^{2} & \phi^{3} \\ \phi & 1 & \phi & \phi^{2} \\ \phi^{2} & \phi & 1 & \phi \\ \phi^{3} & \phi^{2} & \phi & 1 \end{bmatrix}$$



Example 2: Groups of observations are correlated (fixed within-group correlations r_i), no correlation (=0) between groups (2 groups with 3 and 2 observations in this example)

$$ho_e = egin{bmatrix} 1 & r_1 & r_1 & 0 & 0 \ r_1 & 1 & r_1 & 0 & 0 \ r_1 & r_1 & 1 & 0 & 0 \ 0 & 0 & 0 & 1 & r_2 \ 0 & 0 & 0 & r_2 & 1 \end{bmatrix}$$

Example 3: Correlations decrease with "distance" (e.g. spatial data)
This matrix is for a case where observations are equally spaced along a line. In general, correlation is a function of the distance between each pair of observations.

$$\rho_e = \begin{bmatrix} 1 & 0.8 & 0.5 & 0.1 & 0 \\ 0.8 & 1 & 0.8 & 0.5 & 0.1 \\ 0.5 & 0.8 & 1 & 0.8 & 0.5 \\ 0.1 & 0.5 & 0.8 & 1 & 0.8 \\ 0 & 0.1 & 0.5 & 0.8 & 1 \end{bmatrix}$$



- As long as the variance-covariance matrix has a certain structure (positive-definite), GLS will result in "best linear unbiased estimates"
 - Generalized LS estimate minimizes the generalized residual sum of squares:

$$RSS_G = (\mathbf{y} - \mu)' \mathbf{W} (\mathbf{y} - \mu)$$

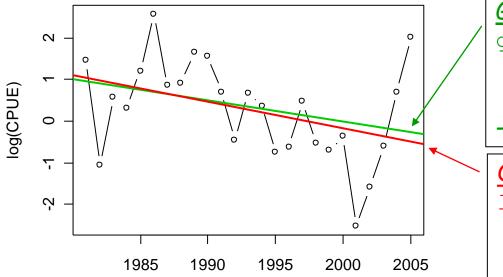
- Variance estimates can be obtained using generalized var-cov matrix W
- F-tests / t-tests are valid (using appropriate generalized RSS_G)



GLS fitting in R

Package: nlme

- Functions: gls (linear), gnls (non-linear)
- Example: GLS regression of rock sole catchper-unit-effort on year (test for linear trend)
 - minor impact on value of coefficients (both unbiased)
 - strong impact on variances (higher uncertainty!)



```
GLS fit:

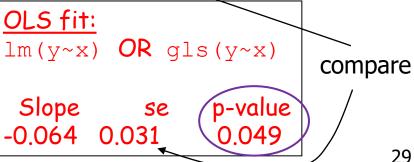
gls(y~x, correl = corAR1(),

method="ML")

Slope se

-0.051 0.045 p-value

0.261
```





Assigned reading

- Jennrich (1995). An introduction to computational statistics (p. 67-71 and p. 188-190)
- Faraway (2005). Linear models with R (p. 89-94)

→ pdf files in 'Readings'



Further reading

Least-squares regression (& more):

- Jennrich, R.I., 1995. An introduction to computational statistics: Regression analysis. Prentice Hall, Englewood Cliffs, NJ.
- Hilborn, R., Mangel, M., 1997. The ecological detective: Confronting models with data. Princeton University Press, Princeton, NJ. (Chpt. 5, 7)