# Colgate vs Crest

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## Cargamos las librerías necesarias

```
library (openxlsx)
library(skimr)
library(fpp2)
library (ggplot2)
library (zoo)
library (ggfortify)
library (tseries)
require (forecast)
require(xts)
library (readr)
library (tidyverse)
library (dplyr)
library (TSA)
library (Hmisc)
library (astsa)
library (tsoutliers)
library (normtest)
```

## Marcamos la base de datos

```
library(readxl)
data <- read_excel('data.xlsx')
View(data)</pre>
```

### Indicamos las semanas

```
data$Date <- as.Date(paste(data$Year, data$Week, 1, sep = "-"), "%Y-%U-%u")
data <- dplyr::select(data, -Year, -Week)</pre>
```

## Summary

skim(data)

#### Data summary

 Name
 data

 Number of rows
 276

 Number of columns
 3

 Column type frequency:
 1

 Date
 1

 numeric
 2

Group variables None

Variable type: Date

skim_variable	n_missing	complete_rate min	max	median	n_unique
Date	0	1 1958-01-06	1963-04-22	1960-08-25	276

#### Variable type: numeric

skim_variable	n_missing	complete_rate	mean	sd	p0	p25	p50	p75	p100 hist
Crest	0	1	0.26	0.13	0.05	0.13	0.25	0.37	0.53
Colgate	0	1	0.31	0.07	0.16	0.26	0.31	0.36	0.50

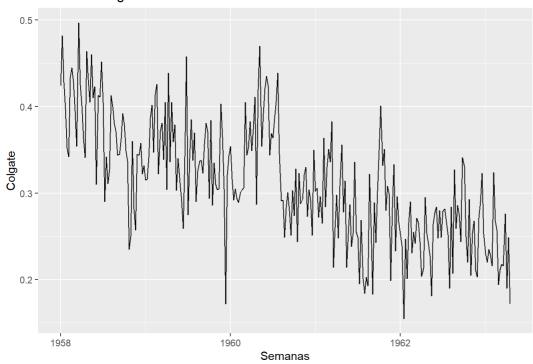
## Dividimos la serie en 2 (Colgate y Crest).

Convertimos los datos.

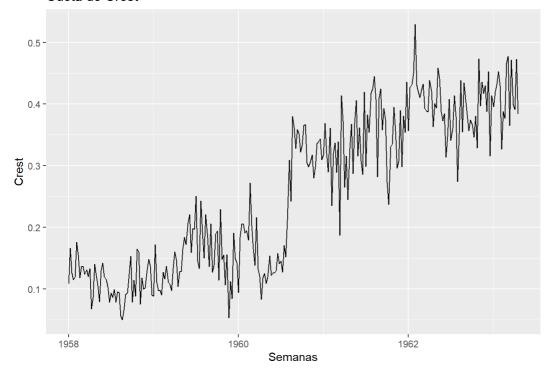
```
acolgate <- as.zoo(colgate)
acrest <- as.zoo(crest)</pre>
```

#### Visualizamos ambas series

### Cuota de Colgate



#### Cuota de Crest



# Test Dickey-Fuller

Test Dickey-Fuller (La Prueba de Dickey-Fuller busca determinar la existencia o no de raíces unitarias en una serie de tiempo. La hipótesis nula de esta prueba es que existe una raíz unitaria en la serie.)

```
##
## Augmented Dickey-Fuller Test
##
## data: acolgate
## Dickey-Fuller = -4.1783, Lag order = 6, p-value = 0.01
## alternative hypothesis: stationary
```

```
##
## Augmented Dickey-Fuller Test
##
## data: acrest
## Dickey-Fuller = -3.4715, Lag order = 6, p-value = 0.04591
## alternative hypothesis: stationary
```

Ambos valores están cerca de superar el margen de significación marcado, pero sin sobrepasarlo, por lo que concluimos que existe estacionalidad.

Eliminamos las 16 semanas del año 1963, para las dos cuotas de mercado.

```
cOmit=16

nObsColgate=length(acolgate)
nObsCrest= length(acrest)
```

### Seleccionamos el training set

```
colgatetrain <- window(acolgate, start=index(acolgate[1]),end = index(acolgate[nObsColgate- cOmit]))
cresttrain <- window(acrest, star= index(acrest[1]), end = index(acrest[nObsCrest-cOmit]))</pre>
```

## **MODELO ARIMA**

Usamos la funcion autoarima, que realiza las transformaciones necesarias para el modelo estacionario.

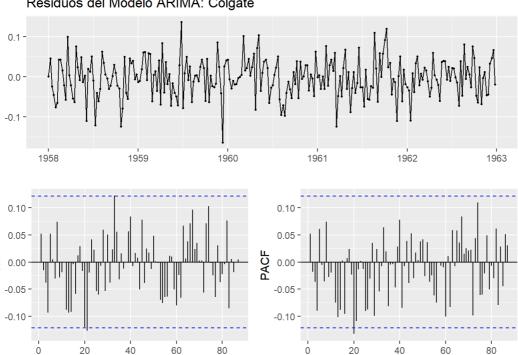
```
## Series: colgatetrain
## ARIMA(0,1,1)(1,0,0)[52]
##
## Coefficients:
       mal sar1
##
       -0.7571 0.0232
##
## s.e. 0.0470 0.0682
## sigma^2 estimated as 0.00232: log likelihood=418.63
## AIC=-831.25 AICc=-831.16 BIC=-820.58
##
## Training set error measures:
##
                      ME
                               RMSE
                                          MAE MPE MAPE
                                                                     MASE
## Training set -0.002534821 0.04788855 0.03773876 -3.019377 12.77398 0.6038201
##
## Training set 0.05216089
```

```
## Series: cresttrain
## ARIMA(0,1,1)
##
## Coefficients:
##
           ma1
        -0.6494
##
        0.0448
  sigma^2 estimated as 0.002054: log likelihood=434.08
  AIC=-864.15 AICc=-864.1 BIC=-857.04
##
## Training set error measures:
##
                      ME
                              RMSE MAE MPE
                                                          MAPE
                                                                    MASE
##
  Training set 0.003018071 0.04514444 0.03445351 -2.924537 17.27068 0.3800355
##
                     ACF1
## Training set -0.04605961
```

# Análisis de los residuos

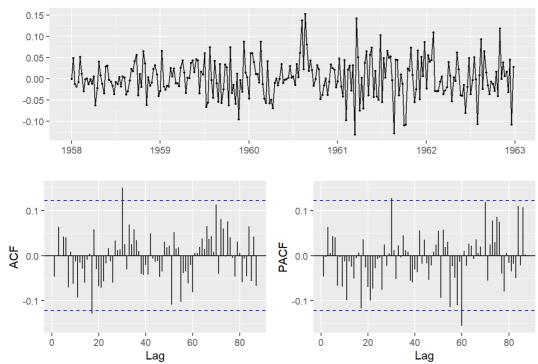
### Residuos del Modelo ARIMA: Colgate

Lag



Lag

#### Residuos del Modelo ARIMA: Crest



Se puede observar en las funciones de autocorrelación, que existen retardos fuera de las bandas de confianza, debido a Outliers.

las series son estacionarias por lo que los residuos son "ruido blanco". Ambos modelos son ARIMA(0,1,1) y los valores AIC son muy representativos, aunque como hemos observado hay información desajustada, eso es debido a los valores atípicos.

# Box-Ljung Test (comprueba correlaciones)

```
Box.test(fit_colgate$residuals,lag=3, fitdf=1, type="Lj")

##
## Box-Ljung test
##
## data: fit_colgate$residuals
## X-squared = 1.1567, df = 2, p-value = 0.5608

Box.test(fit_crest$residuals,lag=3, fitdf=1, type="Lj")

##
## Box-Ljung test
##
## data: fit_crest$residuals
## X-squared = 1.6314, df = 2, p-value = 0.4423

#los valores son superiores al p-value, por lo que no rechazamos la correlación en los datos, lo que result a en aleatoriedad del proceso de muestreo o independencia)
```

## Predicción

#### Aplicamos forecast

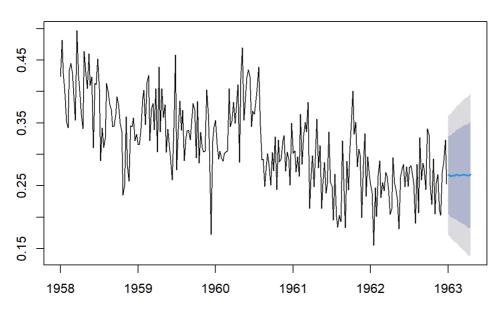
```
summary(cuota.arima.colgate)
```

```
##
## Forecast method: ARIMA(0,1,1)(1,0,0)[52]
##
## Model Information:
## Series: colgatetrain
## ARIMA(0,1,1)(1,0,0)[52]
## Coefficients:
##
   mal sar1
##
       -0.7571 0.0232
## s.e. 0.0470 0.0682
##
## sigma^2 estimated as 0.00232: log likelihood=418.63
## AIC=-831.25 AICc=-831.16 BIC=-820.58
##
## Error measures:
                     ME RMSE
                                        MAE
                                                 MPE
##
                                                         MAPE
## Training set -0.002534821 0.04788855 0.03773876 -3.019377 12.77398 0.6038201
##
                  ACF1
## Training set 0.05216089
##
## Forecasts:
## Point Forecast Lo 80 Hi 80 Lo 95 Hi 95
## 1963.019
             0.2666376 0.2031138 0.3301615 0.1694863 0.3637890
## 1963.038
             0.2646847 0.1994152 0.3299542 0.1648636 0.3645058
              0.2668236 0.1998539 0.3337933 0.1644023 0.3692449
## 1963.058
## 1963.077
              0.2657542 0.1971264 0.3343819 0.1607971 0.3707113
## 1963.096
              0.2672654 0.1970186 0.3375121 0.1598323 0.3746985
             0.2678234 0.1959942 0.3396525 0.1579701 0.3776766
## 1963.115
## 1963.135
             0.2664516 0.1930741 0.3398292 0.1542304 0.3786729
             0.2670096 0.1921157 0.3419035 0.1524693 0.3815500
## 1963.154
## 1963.173
             0.2667074 0.1903272 0.3430875 0.1498940 0.3835207
             0.2673816 0.1895436 0.3452196 0.1483386 0.3864246
## 1963.192
## 1963.212
             0.2672654 0.1879963 0.3465345 0.1460338 0.3884970
## 1963.231
             0.2667539 0.1860791 0.3474287 0.1433725 0.3901353
## 1963.250
             0.2658239 0.1837675 0.3478803 0.1403295 0.3913183
## 1963.269
             0.2660332 0.1826180 0.3494483 0.1384607 0.3936056
## 1963.288
              0.2679396 0.1831875 0.3526917 0.1383225 0.3975567
```

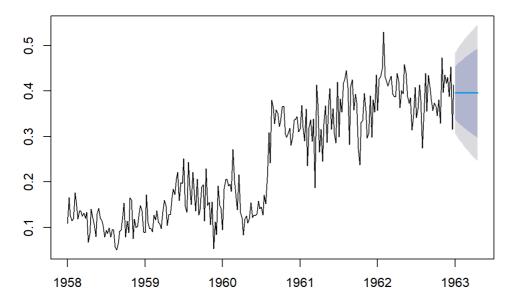
summary(cuota.arima.crest)

```
##
## Forecast method: ARIMA(0,1,1)
\#\,\#
## Model Information:
## Series: cresttrain
## ARIMA(0,1,1)
##
## Coefficients:
##
\# \#
       -0.6494
## s.e. 0.0448
##
## sigma^2 estimated as 0.002054: log likelihood=434.08
## AIC=-864.15 AICc=-864.1 BIC=-857.04
##
## Error measures:
                     ME
                             RMSE
\#\,\#
                                        MAE
                                                 MPE
                                                        MAPE
                                                                  MASE
## Training set 0.003018071 0.04514444 0.03445351 -2.924537 17.27068 0.3800355
##
   ACF1
## Training set -0.04605961
##
## Forecasts:
## Point Forecast Lo 80 Hi 80 Lo 95
                                                    Hi 95
0.3958305 0.3342850 0.4573760 0.3017048 0.4899563
## 1963.019
             0.3958305 0.3310033 0.4606577 0.2966859 0.4949751
## 1963.038
              0.3958305 0.3278800 0.4637810 0.2919092 0.4997519
## 1963.058
## 1963.077
              0.3958305 0.3248940 0.4667670 0.2873425 0.5043185
             0.3958305 0.3220288 0.4696322 0.2829605 0.5087005
## 1963.096
             0.3958305 0.3192707 0.4723903 0.2787424 0.5129186
## 1963.115
## 1963.135
             0.3958305 0.3166086 0.4750525 0.2746710 0.5169900
             0.3958305 0.3140330 0.4776280 0.2707321 0.5209290
## 1963.154
             0.3958305 0.3115362 0.4801249 0.2669134 0.5247476
## 1963.173
             0.3958305 0.3091111 0.4825499 0.2632047 0.5284563
## 1963.192
## 1963.212
             0.3958305 0.3067521 0.4849089 0.2595969 0.5320642
## 1963.231
             0.3958305 0.3044540 0.4872070 0.2560822 0.5355789
             0.3958305 0.3022122 0.4894488 0.2526537 0.5390073
## 1963.250
             0.3958305 0.3000229 0.4916381 0.2493054 0.5423556
## 1963.269
## 1963.288
              0.3958305 0.2978825 0.4937785 0.2460320 0.5456290
```

### Forecasts from ARIMA(0,1,1)(1,0,0)[52]



### Forecasts from ARIMA(0,1,1)



### **Detectamos Outliers**

```
## [1] "No AO detected"

## [,1] [,2] [,3]
## ind 135.000000 136.000000 138.000000
## lambda2 3.918954 4.372891 4.005427

## [1] "No IO detected"

## [1] "No IO detected"
```

El AO detectado en 135 es el impulso explicado por el comunicado hecho el 1 de agosto de 1960, cuando el Consejo de Terapéutica Dental de la American Dental Association (ADA) aprobó a Crest como una "ayuda importante en cualquier programa de higiene dental". Los restantes, ambos 136 y 138 pueden interpretarse como efecto rebote, sin justicacion clara.

## **ARIMAX**

Realizamos el modelo arimax con los datos obtenidos con el modelo ARIMA

```
##
## Call:
## arimax(x = as.double(acolgate), order = c(0, 1, 1), method = "ML", xtransf = data.frame(seqx = 1 *
## (seq(acolgate)), seqy = 1 * (seq(acolgate))), transfer = list(c(0, 0), c(1,
## 0)))
##
## Coefficients:
```

```
## Warning in sqrt(diag(x$var.coef)): Se han producido NaNs

## mal seqx-MAO seqy-AR1 seqy-MAO
## -0.7730 -4e-04 0 -0.0004
## s.e. 0.0455 NaN NaN 0.0072
##
```

Introducimos los AOs dentro del modelo Arimax.

 $sigma^2 estimated as 0.002256$ : log likelihood = 447.31, aic = -886.61

```
##
\#\# arimax(x = as.double(acrest), order = c(0, 1, 1), xreg = data.frame(I136 = 1 *
     (seq(acrest) == 136), I138 = 1 * (seq(acrest) == 138)), method = "ML", xtransf = data.frame(seqx = 1
##
##
     (seq(acrest) >= 135), seqy = 1 * (seq(acrest))), transfer = list(c(0, 0),
##
    c(0, 0)))
##
## Coefficients:
##
        ma1 I136 I138 seqx-MA0 seqy-MA0
        -0.7674 0.0205 0.0766 0.1346 5e-04
##
## s.e. 0.0476 0.0436 0.0420 0.0318
                                          6e-04
##
## sigma^2 estimated as 0.001899: log likelihood = 470.97, aic = -931.94
```

## Comprobamos si hemos conseguido limpiar los outliers

```
## [1] "No AO detected"

## [1] "No IO detected"

## [1] "No AO detected"

## [1] "No IO detected"
```

Podemos observar como han desaparecido

## Funcion de transferencia

```
##
## Call:
## arimax(x = as.double(acolgate), order = c(0, 1, 1), include.mean = TRUE, method = "ML",
## xtransf = acrest, transfer = list(c(0, 0)))
##
## Coefficients:
## mal xtransf-MAO
## -0.8381 -0.4884
## s.e. 0.0346 0.0530
##
## sigma^2 estimated as 0.001751: log likelihood = 481.99, aic = -959.97
```

```
##
## Call:
\#\# arimax(x = as.double(acolgate), order = c(0, 1, 1), include.mean = TRUE, method = "ML",
     xtransf = acrest, transfer = list(c(0, 0)))
##
## Coefficients:
##
    ma1 xtransf-MA0
                 -0.4884
##
       -0.8381
## s.e. 0.0346
                    0.0530
##
\#\# sigma^2 estimated as 0.001751: log likelihood = 481.99, aic = -959.97
##
## Training set error measures:
                                     MAE
                                                MPE
##
                   ME
                             RMSE
                                                        MAPE
## Training set -0.00136496 0.0417676 0.03229819 -2.196369 10.97291 0.6993702
##
                   ACF1
## Training set 0.06132245
```