



# T Confidence Intervals

Statistical Inference

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# Confidence intervals

- In the previous, we discussed creating a confidence interval using the CLT
- In this lecture, we discuss some methods for small samples, notably Gosset's  $t$  distribution
- To discuss the  $t$  distribution we must discuss the Chi-squared distribution
- Throughout we use the following general procedure for creating CIs
  - a. Create a Pivot or statistic that does not depend on the parameter of interest
  - b. Solve the probability that the pivot lies between bounds for the parameter

# The Chi-squared distribution

- Suppose that  $S^2$  is the sample variance from a collection of iid  $N(\mu, \sigma^2)$  data; then

$S^2$  is the sum of squares

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

which reads: follows a Chi-squared distribution with  $n - 1$  degrees of freedom

- The Chi-squared distribution is **skewed** and has support on **0** to  $\infty$
- The **mean** of the Chi-squared is its degrees of freedom
- The **variance** of the Chi-squared distribution is twice the degrees of freedom

# Confidence interval for the variance

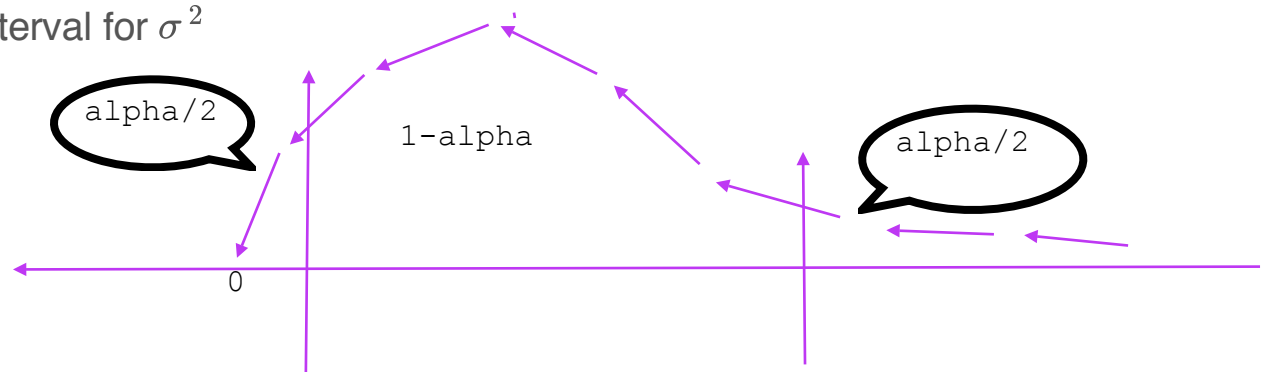
Note that if  $\chi_{n-1,\alpha}^2$  is the  $\alpha$  quantile of the Chi-squared distribution then

$$\begin{aligned} 1 - \alpha &= P\left(\chi_{n-1,\alpha/2}^2 \leq \frac{(n-1)S^2}{\sigma^2} \leq \chi_{n-1,1-\alpha/2}^2\right) \\ &= P\left(\frac{(n-1)S^2}{\chi_{n-1,1-\alpha/2}^2} \leq \sigma^2 \leq \frac{(n-1)S^2}{\chi_{n-1,\alpha/2}^2}\right) \end{aligned}$$

So that

$$\left[ \frac{(n-1)S^2}{\chi_{n-1,1-\alpha/2}^2}, \frac{(n-1)S^2}{\chi_{n-1,\alpha/2}^2} \right]$$

is a  $100(1 - \alpha)\%$  confidence interval for  $\sigma^2$



# Notes about this interval

- This interval relies heavily on the assumed normality
- Square-rooting the endpoints yields a CI for  $\sigma$

# Example

Confidence interval for the standard deviation of sons' heights from Galton's data

```
library(UsingR)
data(father.son)
x <- father.son$height
s <- sd(x)
n <- length(x)
round(sqrt((n - 1) * s^2/qchisq(c(0.975, 0.025), n - 1)), 3)
```

```
## [1] 2.701 2.939
```

# Gosset's $t$ distribution

- Invented by William Gosset (under the pseudonym "Student") in 1908
- Has thicker tails than the normal
- Is indexed by a degrees of freedom; gets more like a standard normal as df gets larger
- Is obtained as

$$\frac{Z}{\sqrt{\frac{\chi^2}{df}}}$$

where  $Z$  and  $\chi^2$  are independent standard normals and Chi-squared distributions respectively

# Result

- Suppose that  $(X_1, \dots, X_n)$  are iid  $N(\mu, \sigma^2)$ , then: a.  $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$  is standard normal b.  $\sqrt{\frac{(n-1)S^2}{\sigma^2(n-1)}} = S/\sigma$   
is the square root of a Chi-squared divided by its df

$$(n-1)S^2/\sigma^2 \text{ is Chi-Sqr}$$

- Therefore

$$\frac{\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}}{S/\sigma} = \frac{\bar{X} - \mu}{S/\sqrt{n}}$$

follows Gosset's  $t$  distribution with  $n - 1$  degrees of freedom



# Confidence intervals for the mean

- Notice that the  $t$  statistic is a pivot, therefore we use it to create a confidence interval for  $\mu$
- Let  $t_{df,\alpha}$  be the  $\alpha^{th}$  quantile of the  $t$  distribution with  $df$  degrees of freedom

$$\begin{aligned} & 1 - \alpha \\ &= P\left(-t_{n-1,1-\alpha/2} \leq \frac{\bar{X} - \mu}{S/\sqrt{n}} \leq t_{n-1,1-\alpha/2}\right) \\ &= P\left(\bar{X} - t_{n-1,1-\alpha/2} S/\sqrt{n} \leq \mu \leq \bar{X} + t_{n-1,1-\alpha/2} S/\sqrt{n}\right) \end{aligned}$$

- Interval is  $\bar{X} \pm t_{n-1,1-\alpha/2} S/\sqrt{n}$

# Note's about the $t$ interval

- The  $t$  interval technically assumes that the data are iid normal, though it is robust to this assumption
- It works well whenever the distribution of the data is roughly symmetric and mound shaped
- Paired observations are often analyzed using the  $t$  interval by taking differences
- For large degrees of freedom,  $t$  quantiles become the same as standard normal quantiles; therefore this interval converges to the same interval as the CLT yielded
- For skewed distributions, the spirit of the  $t$  interval assumptions are violated
- Also, for skewed distributions, it doesn't make a lot of sense to center the interval at the mean
- In this case, consider taking logs or using a different summary like the median
- For highly discrete data, like binary, other intervals are available

# Sleep data

In R typing `data(sleep)` brings up the sleep data originally analyzed in Gosset's Biometrika paper, which shows the increase in hours for 10 patients on two soporific drugs. R treats the data as two groups rather than paired.

# The data

```
data(sleep)
head(sleep)
```

```
##      extra group ID
## 1    0.7        1  1
## 2   -1.6        1  2
## 3   -0.2        1  3
## 4   -1.2        1  4
## 5   -0.1        1  5
## 6    3.4        1  6
```

```
g1 <- sleep$extra[1:10]
g2 <- sleep$extra[11:20]
difference <- g2 - g1
mn <- mean(difference)
s <- sd(difference)
n <- 10
mn + c(-1, 1) * qt(0.975, n - 1) * s/sqrt(n)
```

```
## [1] 0.7001 2.4599
```

```
t.test(difference)$conf.int
```

```
## [1] 0.7001 2.4599
## attr(,"conf.level")
## [1] 0.95
```