

Forecast Adjustment Under Shocks: A Unification

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Abstract

When should one’s default forecasting model be adjusted, augmented, or abandoned completely? Should it be done algorithmically or via human discretion, and if the latter, where and how should discretion play a role? Should an adjustment be carried out using external information, and what exactly should be adjusted? This work further systematizes and unifies the rich landscape of model adjustment and model correction methods, with a special focus on forecast adjustment under the presence of news shocks, when unanticipated events may give an observer reason to doubt the credibility of the default forecasting function. We demonstrate the usefulness of similarity-based methods in forecasting and present a general framework dubbed Similarity-based Parameter Correction (SPC). We highlight several specific time series models that can benefit from SPC, along with formal results for some of those special cases.

1 Introduction

For various modeling and prediction tasks in time series and panel data, the salient challenge is not predicting when an event will occur or begin but what its key properties will be. In the familiar case of scalar time series, those properties can include the time series’ post-event direction, magnitude of change, moments, duration of the event, and correlation structure, all over an arbitrary horizon or perhaps multiple horizons. This is not to say that predicting the arrival of an event is easy. In some cases, it may be difficult or impossible, and the event might not even be conceivable to many before it happens, as was the case with the COVID-19 pandemic. There, reacting sensibly to a shock or structural change may best one can hope for.

On the other end of the spectrum of events for which a temporal prediction is far less valuable than other kinds of predictions is scheduled news, like the releases of economic data or election results. Consider the analysis of Mike Wilson, Morgan Stanley’s CIO and Chief US Equity Strategist, in October 2024 regarding the upcoming 2024 US Presidential US [Wilson, 2024]:

While some argue a Trump win would be a headwind for growth and equity markets, due to tariff risks and slowing immigration, we think there’s an additional element from the 2016 experience that is worth considering—rising animal spirits. More specifically, in 2016 Trump’s pro-business approach led to the largest 3-month positive impact on small business confidence in the past 40 years. It also translated into a spike in individual investor sentiment. It appears to me that markets may be trying to front-run a repeat of this outcome as Trump’s win in 2016 came as a surprise to pundits and markets alike.

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That the event is scheduled and its result is binary simplifies the task of conditional forecasting, to a degree. At the same time, conditional forecasting concretizes one of the fundamental questions of reacting to news [Lundquist and Eck, 2024]: what does this event resemble from the past? Wilson clearly has a view on what is happening as of October 2024 and how market participants are incorporating current information as well as information about past forecast failures, and possibly overweighting more recent failures.

Herein we attempt to unify a range of conceptual approaches to forecasting amid shocks that have developed across the broad ecosystem of the econometric and forecasting literatures. We attempt to locate our main contribution, Similarity-based Parameter Correction, at the intersection of several important strands of thought. We in fact delineate a specific type of SPC called post-shock forecasting, which is induced by nuanced and interesting choices that fall under the general SPC framework. In particular, this work focuses on model adjustment amid news shocks that undermine the reliability of the model at hand. Forecasting under shocks raises unavoidable questions: should the forecast model be abandoned in favor of a discretionary or ad-hoc or one-off adjustment? Does the discretion of a forecaster rule out a quantitative method for making the adjustment? What is the ultimate purpose of the adjustment, and how it is to be used? For how long is the adjustment necessary or reliable?

Forecast model adjustment received its earliest attention under the auspices of the term “intercept correction” [Hendry and Clements, 1994, Clements and Hendry, 1996, 1998, 1999]. Of fundamental importance is the distinction between discretionary and automated intercepts corrections. Hendry and Clements [1994] define scalar intercept corrections to be automated when they follow the simple rule of adding an estimation or prediction residual e_t to subsequent (possibly but not necessarily all) forecasts $\hat{f}_{t+1}, \hat{f}_{t+2}, \dots$. “[S]etting the model back on track” is the informal label they provide for that procedure. It must be noted here that adding a past residual to future forecasts assumes implicitly that whatever caused the residual has some persistence. Hence, intercept correction might be thought more appropriate for time series of levels, or at least times series of rates that are not necessarily mean-reverting, like inflation. Of course, if one has a strong grasp of how long the intercept correction will be necessary, then all of this is moot, but in practice, knowledge of that persistence may be harder to come by than foreknowledge of the need for intercept correction.

The term intercept correction itself can be both a bit deceptive in limiting, in that it need not pertain to intercepts only. Later, Clements and Hendry [1999] add seven additional interpretations for intercept correction, and Hendry and Clements [1994], Clements and Hendry [1999] themselves note the possibility of correcting a coefficient in a forecast model specification, and Guerrón-Quintana and Zhong [2017] developed a similarity-based procedure for correcting β using past information of the time series itself. For a review of more recent work in similarity-based forecasting, we refer the reader to Lundquist and Eck [2024].

What if the break in the DGP is brought about by a very particular kind of event, a news shock? What if we could predict well the intercept shift that occurs between T^* and $T^* + 1$? In Castle et al. [2016], the authors list six necessary conditions for reacting to what we in this current paper call a news shock. Of greatest interest to us are two of the six:

- “the forecasting model already embodies that source of information” and
- “there is an operational method for selecting an appropriate model”.

For such cases, forecasting procedures have been explored in Lin and Eck [2021], Lundquist and Eck [2024], where the AR(1) and GARCH(m, s) cases, respectively, are treated. Both works target additive parameters in scalar time series, predicting those random effects by aggregating information from other time series. The authors leave several stones unturned, including a more general, dare say comprehensive treatment of how to forecast under any sort of shock.

As the literature has developed, a current practitioner of forecast adjustment can now choose between (i) procedures that are discretionary or automated, (ii) internal data (i.e. from the time series itself) or

external, (iii) the term to be corrected (e.g. intercept, coefficients), if any, (iv) as well as the specification of the correction function (i.e. the mapping from the donor unit data to the correction term in the time series under study), including the aggregation mechanism applied to the assembled data (e.g. Nearest-Neighbor, arithmetic mean, kernel methods). We discuss this in depth in Section 5.

If intercept correction is the correction of a forecast by adding the residual at time $T^* + 1$ to the forecasts at some subset of $\{T^* + 1, T^* + 2, \dots\}$, we are doing something rather close to that by adding the predicted correction function output at time t to the forecasts at some subset of $\{T_1^* + 1, T_1^* + 2, \dots\}$, where T_1^* is first integer index following the time series under study. The procedure presented, post-shock forecasting, is a discretionary procedure for intercept correction that integrates data internal or external to the time series under study in a systematic manner. The correction function uses the notion of similarity between breaks in different time series, it and borrows from the causal inference literature to aggregate information from those similar circumstances. Beyond of [Lin and Eck \[2021\]](#), [Lundquist and Eck \[2024\]](#), we are not aware of prior work uses a parametric specification for nonzero errors as well as a method for aggregating similar shocks occurring outside the time series under study.

The structure of this manuscript is as follows. We first provide a review of several disparate time series literatures, including intercept correction, model-evaluation, and similarity-based forecasting. We then provide a canonical setting in which we attempt to focus our work. This setting will be presented at a level of generality that showcases the broad applicability of our method. We then introduce the method from a global perspective, abstracting away from familiar applications. We then show several specific applications, including novel applications that cannot be found elsewhere in the literature. We close with a discussion, including possible extensions.

1.1 History of and Motivation for Intercept Correction

[\[Hendry and Clements, 1994, p. 177\]](#) note that a break in the DGP (and hence the misspecification of the default forecasting method) is the “most obvious and best understood rationale” for intercept correction. To be sure, this manuscript will make its contributions largely in this area. However, before we get there, let’s follow Clements, Hendry and other authors in asking whether we would ever want to intercept-correct an unconditional forecast as opposed to the conditional. This is what Clements and Hendry call “Exploiting the information in unconditional forecasts”. [\[Hendry and Clements, 1994, p. 170-1\]](#) recall how unconditional forecasts can have lower forecast variance than conditional forecasts. There seems to be two different questions that we should distinguish: we forecast would minimize the conditional expectation of MSE, and what forecast would have the lowest variance.

Mismeasured data is discussed in [\[Hendry and Clements, 1994, p. 166\]](#) as a motivator for intercept correction. Could similarity-based correction help? Here is an idea: if we believe that our most recent measurement of the series is noisy, we can disregard the point itself and instead take a convex combination of that point and the [Lin and Eck \[2021\]](#)-style prediction based on aggregation.

Yet another potential motivation springs from the way we evaluate model performance. In [Clements and Hendry \[2005a\]](#), the authors discuss six dichotomies regarding forecast evaluation as a way of advising caution about simplistic notions of performance evaluation. These ideas matter for the current manuscript for at least one reason. We must entertain the possibility that breaks in the DGP will invalidate the standard forecast evaluation framework under which we had been working. If, for example, we have a family of models estimated used a Gaussian loss function, and the post-shock residuals are not conditionally Gaussian (e.g. they could be conditionally Cauchy), then why even start with the default forecasting function? Why not throw it out entirely? If only the distribution of the innovations changes, then would that even matter for a forecast that is purely concerned with the conditional mean?

2 Forecasting Amid Shocks

In this section, we wish to locate the task of forecasting amid shocks within the broader literature on forecasting under breaks in the GDP. [Castle et al. \[2011\]](#) discusses necessary conditions for forecasting under breaks as well as a taxonomy of information available for doing so. They also make a fundamental distinction between “regular forces” and alternative sources of information for making forecasts. Most interesting, their discussion of “forecasting facing a shift in the intercept of a model of a variable y induced by another observable variable z ” represents the most similar idea to [Lin and Eck \[2021\]](#), [Lundquist and Eck \[2024\]](#) — with one important disclaimer. In those later two works, the variable z is observable and yet not quantitative.

A terminological note is in order before we proceed. We use the term “shock” as a special case of a break in the DGP, even though, strictly speaking, some shocks will neither result from a break in the DGP nor herald the emergence of a new regime. The news shocks examined in this manuscript, however, should be considered breaks in the GDP, no matter whether they are transient or permanent.

Additionally, in this work, we will be focused on shocks that occur strictly between two discrete time points, or informational available at a “fractional lag” [[Castle et al., 2011](#)]. This raises an important terminological point. Errors are a staple of stochastic modeling, and certainly no proper inventory of the wide range of uses of “noise”, “shocks”, and so can be accomplished here. Whereas in the psychometric and social-scientific fields, errors may represent unmeasured or latent capacities of a unit, in other fields like the natural sciences, errors may be used to account for variability in the instruments used to collect data. “Innovations” or “shocks”, in econometric literature, are generally understood to represent the arrival of previously unknown (but not necessarily unanticipated) information. Shocks are called structural when they can be linked back to some key feature of a scientific (usually economic) theory; otherwise, they are idiosyncratic. However, there exist other subcategories of shocks that crosscut the distinction just made. For example, a news shock is a shock in which the information material to the market is delivered via newsmedia and moves markets no earlier than the news. Shocks can also be permanent or transitory, supply-related or demand-related, and so forth. The taxonomy is vast. What matters for this work is the news shocks we posit occur between two discrete time points, so that a practitioner is faced with the question of what action is most advisable for the following time point, given the news shock.

2.1 The Role of Outside Information

When is more information better? That is a central question with regard to forecasting. In [Clements and Hendry \[2005b\]](#) (cited in [Castle et al. \[2013\]](#)), the authors note that when forecasts fail, it often has to do with location (i.e. intercept shifts), and therefore using additional information is not likely to help unless it can address potential location shifts.

We should clarify what we mean by “additional information”. By “additional information”, do we mean all information not included in the default forecast function? Could that include additional lags of the time series being forecast? Perhaps a more precise distinction to make is that between information internal to the time series being forecast and information external to the time series being forecast. Are there cases where it is ambiguous whether information should count as internal or external to the time series under study?

We now proceed to the topic to similarity-based methods, since the manuscript under development is at the intersection of post-break forecasting and similarity-based forecasting. In particular, it uses the former to motivate the latter. The notion of similarity appears in various statistical contexts, including matching, synthetic control, nearest-neighbor methods, not to mention the massive area of approximation theory. The uses of similarity are too sprawling to recount in detail. Below we talk about how it could matter in this work.

Our objective is to make forecast corrections by exploiting information from previous time series that

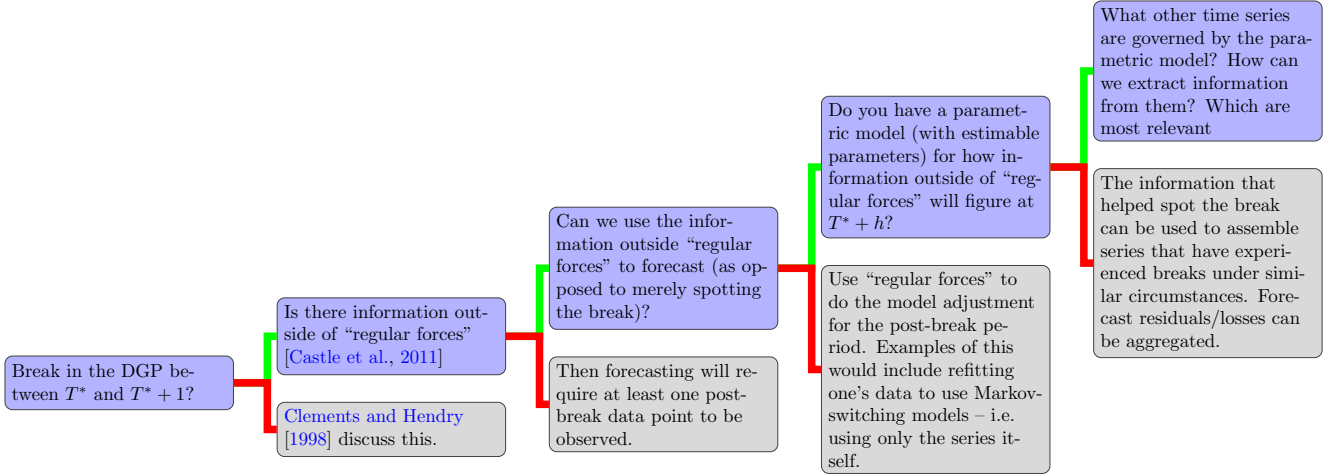


Figure 1: Forecast Model Adjustment: A Decision Tree

have undergone shocks under similar circumstances. The first use of that term is an appropriate point to attempt a definition. As a provisional definition, we say that a break in the DGP of a time series is similar to a break in another if the causes of the breaks share key, material qualities. Note here that we have not said anything about the time series themselves (their DGPs or moments) being similar, nor the entities that they represent (e.g. the firms, the financial instruments, the governments, etcetera) being similar. Why that is so will become clear shortly.

It matters where we look for similarity and how we estimate similarity. Quantitative ways of determining similarity include matching a target object to other objects along one or more variables, where exact matches may be likely or may be statistically impossible. In metric spaces, the notion of a metric that quantifies nearness is crucial to fundamental concepts of proximity. This suggests an optimization problem such that, for any target object Q , the search for similarity means a search for an approximator A s.t. $\delta(Q, A)$ is smallest, where δ is an appropriate metric.

Asymmetric distance functions (which are of course not metrics) are useful when we want to explore differences between donors and weight different contributors to that function differently. Consider an asymmetric Jaccard Index $\delta_{\mathcal{J}}(\cdot, \cdot)$ that maps from any two finite-cardinality sets to the interval $[0, 1]$ [Garg et al., 2015]. So far, we have discussed only binary measures of similarity. The distance-based weighting scheme found in Lin and Eck [2021] and repurposed in Lundquist and Eck [2024] can be thought of as an n -ary similarity measure, where the n weights means little in isolation but can tell a greater story when viewed alongside the approximation they produce.

What about qualitative ways of determining similarity? In Lundquist and Eck [2024], donors are identified by matching the qualitative aspects of the shock in the time series under study to shocks previously witnessed in other series. Only after the donor pool has been identified based on qualitative facts about the shocks is there a role for quantitative tools.

The role of the covariate vector in our model appears to be two-fold: the covariates appear in the shock distribution, and the covariates also help us determine which donors are most important.

What if a given entry of the covariate vector was helpful at only one of those. For example, what if an entry was particularly useful in yielding good weights but was not important for the shock distribution — perhaps because the corresponding random effect was deterministically zero. Conversely, suppose a given entry was known to be part of the shock distribution, but the inclusion of that covariate in distance-based weighting was often not worth it, perhaps because it was hard to approximate with using donors, and the corresponding random effect was small in magnitude with high probability.

3 Setting

In order to introduce the general framework for adjusting forecasts under shocks, we provide an illustration. In Figure 2, we borrow from Lundquist and Eck [2024] to show how the aggregation of estimated excess volatilities from donors in the donor pool works when the correction function is a specially-chosen convex combination of fixed effects from the donor pool.

3.1 Formal Modeling Setup

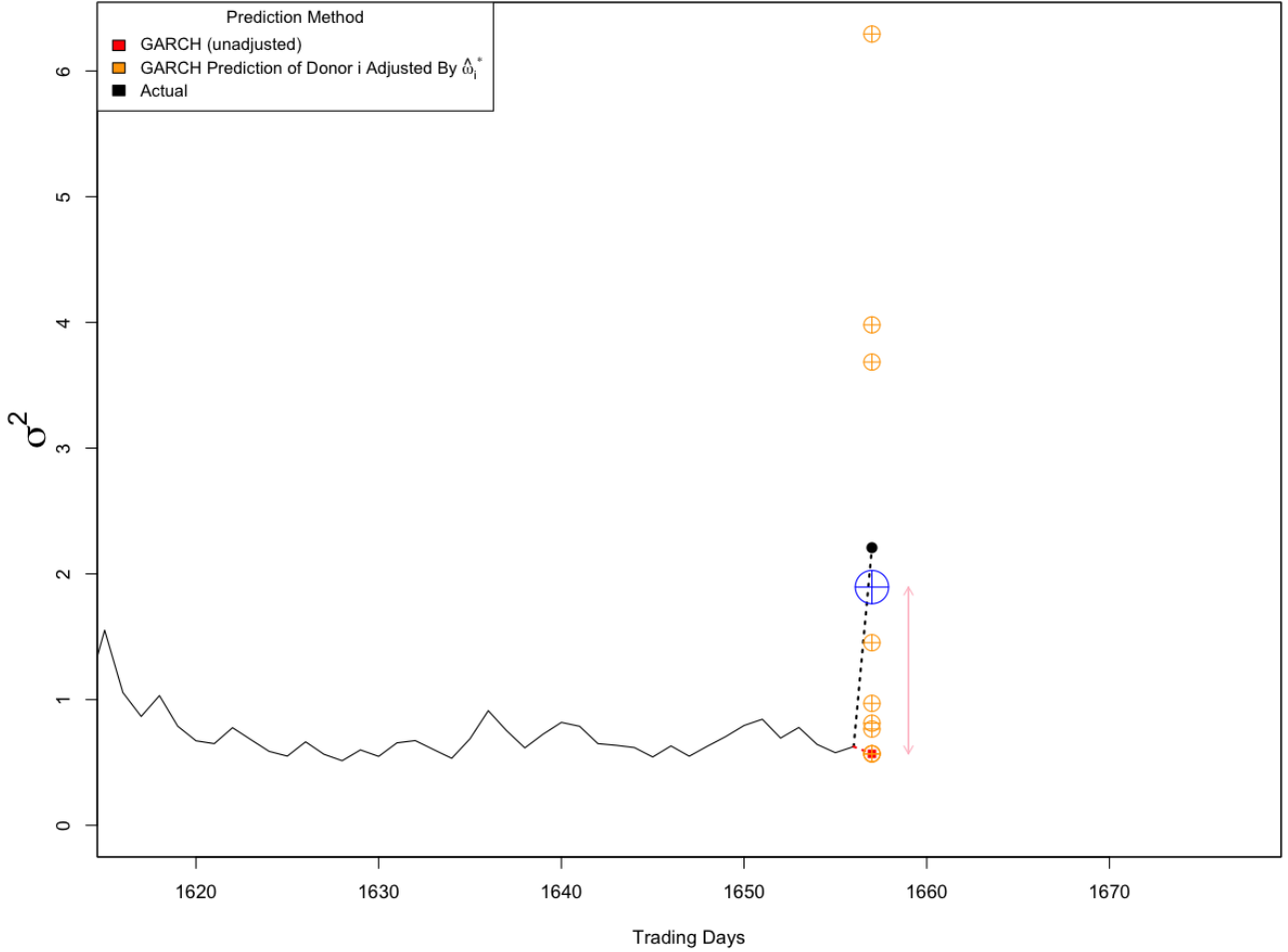


Figure 2: The simulated time series experiences a volatility shock between trading days 1,656 and 1,657. The GARCH prediction, in red, fails even to approach the volatility spike at $T^* + 1$, as do several adjusted predictions, which are orange. In contrast, the GARCH forecast adjusted by $\hat{\omega}^* = \sum_{i=2}^{n+1} \pi_i \hat{\omega}_i^*$, a convex combination of the estimated shocks in the donor pool, achieves directional correctness as well as a smaller absolute loss in its prediction. The pink vertical line serves to indicate the adjustment of size $\hat{\omega}^*$ that allows the blue bullseye to approach more closely the ground truth.

We will suppose that a researcher has a multivariate time series data $(y_{i,t}, \mathbf{x}_{i,t})$, $t = 1, \dots, T_i$, $i = 1, \dots, n + 1$, where $r_{i,t}$ is scalar and $\mathbf{x}_{i,t}$ is a vector of covariates such that $\mathbf{x}_{i,t} | \mathcal{F}_{i,t-1}$ is deterministic and observed. Suppose that the analyst is interested in forecasting $y_{1,t}$, the first time series in the collection,

which we will denote *the time series under study*. We assume the following general data-generating process

$$y_{i,t} = F(\mathcal{F}_{i,t-1}) + \mathbf{x}_{i,t}^T \lambda_t + \epsilon_{i,t} \quad (1)$$

$$\alpha_{i,t} = \mathbf{x}_{i,t}^T \lambda_t, \quad (2)$$

adjust definition

$$\mathcal{M}_1: y_{i,t} = F(\mathcal{F}_{i,t-1}) + \mathbf{x}_{i,t}^T \lambda_t + \epsilon_{i,t}$$

with time-invariant error structure

$$\lambda_t \sim \mathcal{F}_\lambda \text{ with } \mathbb{E}_{\mathcal{F}_\lambda}(\lambda) = \mu_{\lambda_t}, \text{Var}_{\mathcal{F}_\lambda}(\lambda) = \Sigma_{\lambda_t}$$

$$\epsilon_{i,t} \stackrel{iid}{\sim} \mathcal{F}_\epsilon \text{ with } \mathbb{E}_{\mathcal{F}_\epsilon}(\epsilon) = 0, \text{Var}_{\mathcal{F}_\epsilon}(\epsilon) = \sigma_\epsilon^2,$$

$$u_{i,t} \stackrel{iid}{\sim} \mathcal{F}_u \text{ with } \text{Var}_{\mathcal{F}_u}(u) = \sigma_u^2,$$

$$\epsilon_{i,t} \perp\!\!\!\perp u_{i,t}.$$

where F maps objects from the past into future objects of $y_{i,t}$, and $\epsilon_{i,t}$ is mean-zero and uncorrelated across time and donors. Note that this specification above is capacious enough to include data-generating processes of the form

$$y_{i,t} = \log Y_{i,t} = \log a_{i,t} y_{i,t-k} e^{\mathbf{x}_{i,t}^T \lambda_t + \epsilon_{i,t}} = \log a_{i,t} + \log y_{i,t-k} + \mathbf{x}_{i,t}^T \lambda_t + \epsilon_{i,t}$$

with $k \geq 1$ and $a_{i,t}, y_{i,t}$ supported on \mathbb{R}^+ . The term $\mathbf{x}_{i,t}^T \lambda_t$ will be deemed the *correction term*, and the function

$$\xi: \mathcal{F}_{2,T_2^*+h} \times \dots \times \mathcal{F}_{n+1,T_{n+1}^*+h} \rightarrow \mathcal{F}_{1,T_1^*+h}$$

that estimates or predicts $\mathbf{x}_{1,T_1^*+h}^T \lambda_{T^*+h}$ will be deemed the *correction function*. The correction function maps observable, deterministic events and unobservable events into a space so that something might be learned about λ_t . If $\mathbf{x}_{i,t}$ were not observable or not deterministic with respect to \mathcal{F}_{t-1} , then the correction term would represent an additional error term (not necessarily idiosyncratic).

We require that each time series $\mathbf{y}_{i,t}$ is subject to an observed news event following $T_i^* \leq T_i + 1$ and before witnessing $T_i^* + 1$. The reasons for this represent the heart and soul of the framework proposed. We are not proposing a method to correct for unanticipated shocks that have been reflected in the observed data. Instead, we are proposing a method to correct for shocks for which the qualitative kernel of information is known and the observed quantitative instantiation of that which is not yet known.

So far we have reviewed a long and rich literature about model adjustment, and we have also introduced a data-generating process upon which the rest of this manuscript will be based. We now proceed to introduce a particular framework of solutions to the circumstances assumed in Section 3. What we are doing here is therefore not out of the ordinary but distinct in at least two ways: (1) we are incorporating outside information faster, and (2) we are proposing a principled, systematic way to do it that defies that curse of dimensionality. If one wanted to forecast using linear models and high-dimensional information set, then some kind of method for reducing model complexity would be required. We avoid that.

4 Methodology for Similarity-based Parameter Correction

4.1 Forecasting and Correction Functions

We now present two one-step-ahead forecasts for the time series under study. The first is the unadjusted forecast. The second is the adjusted forecast, which differs by the predicted correction term. These forecasts are:

$$\begin{aligned} \text{Forecast 1: } \hat{y}_{unadjusted, T_1^*+1} &= \hat{\mathbb{E}}[\mathcal{F}_{T_1^*}] \\ \text{Forecast 2: } \hat{y}_{adjusted, T_1^*+1} &= \hat{\mathbb{E}}[\mathcal{F}_{T_1^*}] + \hat{\omega}^* . \end{aligned}$$

The problem of aggregating shocks or residuals begins with the data constraints. Let us first introduce useful notation. Let $\hat{\omega}_{i,*}^*$ denote estimated correction term for donor i . Taking the estimated correction terms as a given, we observe the tuple of information $(\{\hat{\omega}_{i,*}^*\}_{i=2}^{n+1}, \mathcal{F}_{2, T_2^*} \times \dots \times \mathcal{F}_{n+1, T_{n+1}^*})$ that is the extracted residuals as well as information important to aggregation.

One such idea comes from [Lin and Eck \[2021\]](#), where weights belonging to the simplex, $\{\pi_i\}_{i=2}^{n+1} \in \Delta^{n-1}$, are chosen via a distance-based weighting procedure, resulting in the correction function

$$\hat{\omega}^* = \sum_{i=2}^{n+1} \pi_i \hat{\omega}_{i,*}^*,$$

which is just one way among many to build a correction function. Another interesting correction function can be found in [Foroni et al. \[2022\]](#). For yet another example, consider a correction function that aggregates survey panel of forecasts by the Survey of Professional Forecasts [[Croushore, 1993](#)]. An obvious advantage of human forecasters is that they are able to integrate information revealed arbitrarily close to the scheduled release of the forecast, even if human forecaster’s lack of systematicity or closed-form forecasts, for example, leads to non-optimal forecast performance in general. By calculating forecast differences between the low-frequency information forecast, $\tau_{i,t} = \hat{y}_{unadjusted,t} - \hat{y}_{i,t}$, we prepare ourselves for late-breaking events – e.g. shocks – for which we might want a correction function

$$\xi: \tau \rightarrow \mathcal{F}_{1, T_1^*}$$

that weights the i -th forecaster according to that forecaster’s past performance or according to the similarity of the forecaster’s stated information set to the information available for the default forecasting model.

5 Model Adjustment Using Similarity-Based Parameter Correction: A Global Overview

In this section, we introduce and discuss a general framework for model adjustment that generalizes and is motivated by the circumstances laid out in [Section 3](#) by boiling down SPC to its essential five elements. The essential elements of similarity-based parameter correction are

1. **Object-to-predict** Most fundamentally, the method requires a random object (indexed over time, of course) that obeys a specification with additive errors, or, at the very least, a specification that can be transformed to have additive errors. This requirement is suitably weak, so as to include

models that are not linear or not linear in each of their parameters. It also includes multidimensional objects as well as objects from non-Euclidean probability spaces, like some function spaces.

Notice that we did not say *parametric* specification. The reason for this is that, given any forecast function f from an information set \mathcal{F}_t , and given any appropriate loss function L , we can define a residual $e_t = y_t \odot \hat{y}_t$, where \odot is subtraction in the simple case of mean squared-error. The residuals (or transformations of those residuals) of those n models can be weighted as part of a correction function. This fact is especially useful when the forecast function f is nonparametric, black-box, or stochastic with respect to \mathcal{F}_t .

2. **Common Model Family on the Shocks** The method requires that residuals be governed by a model that is shared across all units. This ensures that in the estimation of news shocks in the donor pool, the estimators will enjoy similar properties that will produce a good aggregated shock estimator. This condition is satisfied by the parametric shock distributions found in [Lin and Eck \[2021\]](#), [Lundquist and Eck \[2024\]](#).
3. **Reliable and Shared Model-Fitting Procedure** There must exist a reliable model-fitting procedure for the $n + 1$ units, one that will allow the assumption of a shared family for the residuals to be reflected in the residuals. Here, reliable means several things. It means that the estimation procedure must produce a credible description of each data-generating process or a good prediction function, so as to aid in estimating residual in each unit. ‘Reliable’ may mean the estimators have low variance, and it may also mean that the estimation procedure is robust, to some degree, to misspecification bias. However, in theory, the lack of these properties is not harmful to the method unless the lack of these properties harms the correction term’s estimation. When we use fixed effect estimation (under ordinary assumptions), we can construct confidence intervals for the fixed effect estimates, and then assuming independence, we can get confidence intervals for convex combinations of fixed effect estimates.
4. **Reliable Correction Term Estimation** Fourth, there must exist a reliable procedure for modeling and estimating the correction term in each unit. Again, here we care about low variance as well as robustness to misspecification. The very simplest correction term is the residual itself. Alternatively, the correction term could be an inner product that depends upon external covariates, as is found in [Lin and Eck \[2021\]](#), [Lundquist and Eck \[2024\]](#).

This might not always be straightforward. Some models like GARCH, for example, might deliver very noisy estimates for indicator variables that occur just once.

5. **Reliable Correction Function Estimation** There must exist a correction function (presumably based on the correction term) that maps data from the donor pool to the *predicted* correction term in the time series under study based on some notion of similarity. In some cases, there may exist a posited DGP that the correction term estimates. In other cases, there may be no posited GDP. Similarly, if there exists a posited DGP for the correction term, it may depend upon data internal to the time series under study, or it may depend on external data, e.g. outside data could be aggregated (e.g. Nearest-Neighbor, arithmetic mean, kernel methods) to estimate the correction function.

6 Formal Properties and Model-Specific Considerations

In this section discuss particular models in which our approach has been implemented as well as other in which it has yet to be implemented. For those for which an implementation exists, we portray the model in a more general light, while also commenting on the model-specific considerations.

Recall the specification 1, which models a rare shock as an affine function of covariates:

$$y_{i,t} = F(\mathcal{F}_{i,t-1}) + \mathbf{x}_{i,t}^T \lambda_t + \epsilon_{i,t}$$

6.1 ARIMA

We begin our recounting of model-specific cases by recalling [Lin and Eck \[2021\]](#), in which it was established that for a family of AR(1)-distributed scalar time series with a common shock distribution, forecast risk can be reduced via a similarity-based adjustment procedure. AR(p) is an easy model to work with, in part because it can be consistently estimated via OLS, and an AR(p) can approximate an ARMA(p,q) with arbitrarily-small error through via a truncation of an AR(∞) representation. Setting aside the estimation method, consider a one-step-ahead for an AR(p) at time t with an adjustment estimator, $\hat{\alpha}_t$:

$$\hat{y}_t = f(\mathcal{F}_{t-1}) = \hat{\mu} + \sum_{i=1}^p \hat{\rho}_i y_{t-i} + \hat{\alpha}_t \quad (3)$$

which can be understood as an AR(p) with time-varying coefficients. The mean squared-error of the forecast above, i.e.

$$\mathbb{E}[(y_t - \hat{y}_t)^2] = \mathbb{E}[(y_t - \hat{\mu} + \sum_{k=1}^p \hat{\rho}_k y_{t-k} + \hat{\alpha}_t)^2] \quad (4)$$

$$= \mathbb{E}[(\mu + \sum_{k=1}^p \rho_k y_{t-k} + \epsilon_t + \alpha_t - \hat{\mu} + \sum_{k=1}^p \hat{\rho}_k y_{t-k} + \hat{\alpha}_t)^2] \quad (5)$$

admits of the typical bias-variance decomposition, with the special proviso that with more terms in the prediction function, more covariance terms arise. Similarly, the covariance structure of the random effects in $\alpha_t = \mathbf{x}_t^T \lambda_t$ will contribute to the variance. In the real-data examples [Lin and Eck \[2021\]](#), the entries in λ_t are proxies of general market risk, which one would commonly take to be correlated not only with each other but also with a financial time series y_t . All of this is to say that control of the MSE may require tight control of the variance in the prediction function, which in turn will gesture in the direction of a sparse $\hat{\alpha}_t$ and sparse estimation of the autoregressive structure.

Proposition 1. *Let $2 \leq i \leq n+1$. The tuple of estimators $(\hat{\rho}_1, \dots, \hat{\rho}_p, \hat{\alpha}_{i,t})$ is consistent at $t \rightarrow \infty$.*

Proposition 2. *Let $\{\hat{y}_{1,T_1^*+r}\}_{r=1}^{T_1^*+h}$ denote the vector of predictions in the time series under study. $\{\hat{y}_{1,T_1^*+r}\}_{r=1}^{T_1^*+h} \xrightarrow{d} \{y_{1,T_1^*+r}\}_{r=1}^{T_1^*+h}$.*

When forecasting financial time series using the method above, the question of prices versus returns (or in the language of economics or engineering, levels versus flows) is relevant. Consider, for example, the persistence of a shock. A shock to an asset's return series at time t implies a persistent shock to the asset's price series at time t and beyond, as one can verify by inverting the discrete-differencing operation on a time series. However, if one forecasts the price series of the asset directly, that persistence beyond the first instantiation of the shock must be modeled explicitly. This suggests interesting opportunities in the analysis and forecasting of shock persistence for situations in which level-forecasting is appropriate, e.g. for stationary time series.

We now turn to the limitations of linear time series model class like ARIMA. The first is that linearity is a nontrivial restriction, and for the purpose of modeling financial time series, applications of SPC to TAR and SETAR would be welcome. Additionally, the efficient market hypotheses implies that market prices reflect all publicly-available information, which in turn implies that ARIMA models should be of

no predictive value for market returns. This critique, in a technical sense, alleges that market returns are random walks centered at zero, and for that reason, the coefficients of an ARMA approximation to a financial return series are uniformly zero, and that any incipient deviation from that equilibrium would not go unnoticed by market participants, who would exploit those deviations and push the coefficients back to zero, leaving only idiosyncratic noise. Such concerns about forecasting in markets with heavy feedback loops motivates our next section.

6.2 GARCH

The methods of [Lin and Eck \[2021\]](#) are adapted in [Lundquist and Eck \[2024\]](#) to forecast volatility. This is in response to concerns we raised above about the usefulness of ARIMA models for forecasting financial returns as well as concerns about heteroskedasticity. Financial time series are known to have both time-varying volatility but also volatility that *clusters*. A GARCH model, under weak conditions, is an ARMA model on the squared residuals of a time series, and empirical evidence strongly suggests that those ARMA (i.e. GARCH) coefficients are not zero, in general. The predictability of either those squared returns or σ_t^2 (in this context, the two tasks reduce to one) is bounded, however, by a function of the kurtosis of a GARCH process [\[Francq and Zakoian, 2019\]](#). That the GARCH model accommodates excess kurtosis in both the unconditional and conditional returns of a financial asset is a feature of the model, and perhaps for that reason, it has been noted that GARCH models do a much better job at modeling the evolution of volatility (a descriptive or inferential objective) as opposed to forecasting volatility.

In this way, [Lundquist and Eck \[2024\]](#) bridged a gap between the descriptive and predictive roles of GARCH. Fixed effects (correction terms) are estimated in the donor pool, and those fixed effects are aggregated via corrected function. A forecast adjustment is then made in the time series under study.

6.3 HAR

The Heterogeneous Autoregressive (HAR) model for realized volatility [\[Corsi, 2009\]](#) uses ordinary least-squares (OLS), WLS, or even a logarithmic specification to estimate an autoregressive model. It explores influence of previous volatility on current volatility, of course, but uses a specification motivated by theory: that a heterogeneous collection of market participants interacts at different frequencies. Therefore, the volatility can be well-modeled as a linear combination of the average volatility over the previous day, previous week, and previous month (usually taken to be 22 days). Formally, HAR assumes that the log price of an asset is governed by a diffusion process (a model beyond our scope herein), and HAR estimates

$$RV_t = \beta_0 + \beta_{1\text{-day}}RV_{t-1} + \beta_{5\text{-day}}\overline{\{RV_{t-1}, \dots, RV_{t-5}\}} + \beta_{22\text{-day}}\overline{\{RV_{t-1}, \dots, RV_{t-22}\}} + \epsilon_t$$

6.3.1 A Real Data Application of Shock-Adjusted HAR Forecasting

6.4 LSTM/GRU

[Hendry and Clements \[2004\]](#) discuss that intercept correction can improve forecasts under structural breaks as well as under “deterministic misspecification”, by which they mean misspecification not caused by stochastic events like location shifts. The same two justifications exist for SPC: we can use SPC when we know there to be a structural break, and we can also use it when we know our model is wrong but lack knowledge of precisely how it is wrong, and yet nevertheless wish to correct a parameter for some critical forecasting task.

One reason we may believe that our forecasting model is misspecified is that our forecast function is non-parametric, and hence the deterministic or stochastic mechanism through which we produce our forecasts does not involve parametric assumptions. Here we demonstrate the tremendous capaciousness

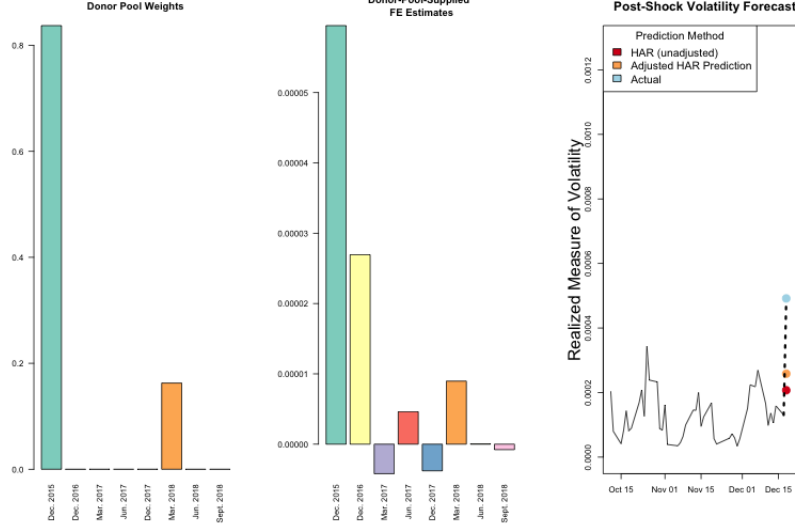


Figure 3: Over the course of the 2010s, the US FOMC began a rate-hiking cycle. Left pane: weights of the donor pool of FOMC meetings; right-pane: fixed effect estimates; right pane: the unadjusted forecast, adjusted forecast, and actual forecast

of SPC by applying it to predictions generated by a pair of non-parametric forecasting functions, Long Short Term Memory and GRU. We will re-examine an application found in [Lin and Eck \[2021\]](#).

We borrow code from [Brownlee \[2022\]](#)

- [LSTM in Python](#)

6.5 “Exponential Breaks”

In [Castle et al. \[2011\]](#), the authors discuss intercept shifts that decay deterministically. Those shifts are parameterized by a scalar λ and a decay rate ψ . We adapt their model for our own purposes:

$$y_{i,t} = \alpha_i + \lambda[1 - e^{-\psi[t-T_i^*+1]}]\mathbf{1}_{t \geq T_i^*+1} + \epsilon_{i,t}, \epsilon_{i,t} \stackrel{iid}{\sim} [0, \sigma^2]. \quad (6)$$

This poses a novel case for our method, in that the model is parametric and yet without estimability of all shock parameters, aggregation may not be straightforward or easy. Consider, for any i , $2 \leq i \leq n+1$, the estimates $\hat{\alpha}_{i,OLS}$ obtained via regression on the first $T_i^* + 1$ points as well as $\hat{\lambda}_{OLS}$ obtained by estimating a fixed effect at $t = T_i^* + 1$. Consider also a subsequent estimator, an estimator for ψ , that is derivable from inverting the model specification and plugging-in the estimates $\hat{\alpha}_{i,OLS}, \hat{\lambda}_{OLS}$:

$$\hat{\psi}_{i,plugin} = \frac{1}{T_i - (T_i^* + 2)} \sum_{t \geq T_i^*+2}^{T_i} -\log \left[- \left(\frac{y_{i,t} - \hat{\alpha}_i}{\hat{\lambda}} - 1 \right) \right] / [t - T_i^* + 1]. \quad (7)$$

(7) is, to put it lightly, a complicated estimator with potentially dubious properties. To begin, (7) is defined only when the term inside the logarithm is positive for each $t \geq T_i^* + 2$. We could get around this by dropping any summands for which that term is nonpositive and instead calculate an average over only

the remaining subset of summands. Additionally, variance estimates for $\hat{\psi}_{i,plugin}$ may require resampling methods or the delta method. Estimation error from $\hat{\alpha}$ and $\hat{\lambda}$ would naturally propagate to $\hat{\psi}_{i,plugin}$, and the expected estimation error would depend upon the size of the pre-shock sample.

Now consider a practitioner in need of a conditional forecast for $T_1^* + h, h \geq 1$. Could aggregating the prediction residuals $r_{i,T_i^*+h} = y_{T_i^*+h} - \hat{y}_{i,T_i^*+h}, h \geq 1$ lead us to a superior predictions in the time series under study?

Consider also these complicating scenarios

1. We can estimate λ using a cross-sectional regression, and that may be better estimator than doing it one-by-one in each donor.
2. Here we have taken λ to be fixed and shared across donors. If λ is a function of a random effect, however, then estimating it becomes harder.

6.6 The Method of Multiple Shocks, Dual Shocks, and Shock-Partitioning

Sometimes, we might have reason to believe that a forecast requires adjustment due to not one but multiple shocks. [Lin and Eck \[2021\]](#) explored a “dual” shock in their analysis of oil company equity price shocks due to the March 2020 COVID lockdowns.

Method 1: for each source of the k shocks, assemble n_j donors and p_j covariates, with $j = 2, \dots, n_j + 1$, compute weights $\{\pi_{i,j}\}_{i=2}^{n_j+1}$. With these k weight sets, a set of k aggregated shock estimators can be constructed and averaged (or aggregated in some other way).

Method 2: the k shocks can be represented using one application of distance-based weighting, with the vector δ and the vector \mathbf{x}_{i,T_i^*+1} can encode a partition of shocks by using zeros in appropriate entries.

7 Discussion

The forecast horizon — does it matter? If so, how so? [Lin and Eck \[2021\]](#) has a one-period horizon. [Clements and Hendry \[1998\]](#)p. 203 discuss how long to keep the forecast adjustment in place. For a corrected “slope parameter”, the effect of h is not so clear.

7.1 Shock propagation and estimation accuracy

Consider [Lundquist and Eck \[2024\]](#), in which a real world example shows how volatility can be forecasted following surprising election outcomes. There, the estimation of fixed effects is done around elections in a donor pool. That means that for each past election or poll, etcetera, the method must accurately estimate the surprise-induced volatility boost. That means picking a window around such an effect.

proposals

1. For a given data-generating process, simulate the accuracy and variability of shock estimates using post-shock measurement periods of difference lengths.
2. Develop formal results to quantify (estimate, bound, etc) the risk of using different post-estimation measurement lengths. This matters because it might be that a donor has experienced a shock in the last m days, and if m is too small, using that donor may be risky.

7.2 Aggregating residuals as opposed to aggregating fixed effect estimates?

Cases:

1. Fixed effect is estimated with accuracy
2. Fixed effect is estimated with considerable noise. Hence, the idiosyncratic error is estimated with considerable noise.
3. idiosyncratic error large
4. idiosyncratic error small

8 Extensions

- Forecast for series besides the mean or conditional mean **If you have a family of models estimated used a Gaussian loss function, and the post-shock residuals are not conditionally Gaussian (e.g. they could be conditionally Cauchy), then why even start with the default forecasting function? Why not throw it out entirely? If only the distribution of the innovations changes, then would that even matter for a forecast that is purely concerned with the conditional mean?**
- Bias-variance decomposition for ARIMA, as introduced in [6.1](#)
- Shrinkage estimation of autoregressive structure on all donors, then use our method
- TAR
- SETAR
- Function forecasts
- Binary Outcome Forecasts
- Density Forecasts
- Quantile Forecasts

8.1 Forecast Combination

what we are talking about here is not forecast combination, but there may be, nevertheless, a role for forecast combination: combining the forecasts generated by small differences in covariate and/or donor choice, as is done in [Lundquist and Eck \[2024\]](#).

Consider the leave-one-out method used on pairs of donors and covariates in [Lundquist and Eck \[2024\]](#). By averaging over $(n + 1) \cdot (p + 1)$ predictions, we average over $n + 1 \cdot p + 1$ convex combinations (where, when a donor is omitted, its weight can be imputed as zero).

8.2 Limitations

8.2.1 How important is a shared DGP?

In particular, can we relax this assumption and proceed with weaker assumptions on the donors and time series under study?

8.2.2 The assumption of a shared family can be relaxed in certain circumstances. For example, if a certain donor is not like the time series under study with respect to the covariates, then the lack of a shared family will not matter much.

In Scope

- Introduce an existing diffuse set of approaches to adjusting forecasts (1)
- Explain when/how similarity can help us forecast. 2.1
- distinguish the method from various tools that inspired it (1)
- show a few special cases, both examples and formal results (6)
- discuss limitations of the method (8.2)
- propose extensions (8)

Beyond Scope (but mentioned briefly)

- Wade too deeply into distance-based weighting details. State that it's simply one way to do it. Allude to [Lin and Eck \[2021\]](#), [Lundquist and Eck \[2024\]](#).
- wade too deeply into any of the special cases
- Tackle non-scalar random quantities (density forecasts, etc)
- Forecast combination

9 Supplement

9.1 Formal Results

Proof of Proposition 1 We claim

$$\mathbb{E}[\sigma^2] \tag{8}$$

□

Proof of Proposition 2 We claim

$$\mathbb{E}[\sigma^2] \tag{9}$$

□

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