Forecast Adjustment Under Shocks: Similarity-based Solutions to Unprecedented Events Doctoral Defense

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A High-level Summary Uniting All My Projects

Forecasting under non-ideal conditions like

- a rupture in the DGP
- lack of past information for units we wish to forecast, i.e. a so-called "cold start" (Fatemi et al. 2023)

Central technique explored: *correction* of model parameters using an aggregation strategy, where aggregation means that we make use of external data.

We began with particular models; then built upwards, establishing an abstract, general framework for correcting forecasts that is widely applicable for forecasting and beyond.

Why is this valuable?

We are weighing-in on a puzzle:

conflict between the intuitive notion that more relevant information should help in forecasting, and the hard reality that attempts to make it do so have not been uniformly successful (Clements and Hendry 2005)

Why is this valuable?

We provide a precise proposal to a well-specified challenge:

incomplete information by itself is unlikely to play a key role in forecast failure (except if that information would forecast breaks). Consequently, using large amounts of data may not correct one of the main problems confronting forecasters, namely location shifts, unless that additional information is directly pertinent to forecasting breaks (Castle, Clements, and Hendry 2013)

Why is this valuable?

- ullet Canonical problems: predicting based on a limited feature space and/or small n
- The framework we offer is something others can 'plug-in' to and add to.
- Software developed



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What does it resemble from the past?



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- Oan we incorporate past events in a systematic, principled manner?



Example (Weekend of March 6th - 8th, 2020)



Oil nose-dives as Saudi Arabia and Russia set off 'scorched earth' price war

PUBLISHED SUN, MAR 8 2020+9:01 AM EDT | UPDATED MON, MAR 9 2020+5:33 PM EDT

Oil crashes by most since 1991 as Saudi Arabia launches price war



Punchline

 Credible forecasting is possible under news shocks, so long as we incorporate external information to account for the nonzero errors.



Outline

- Introduction
- 2 The Idea and Methodology
- Formal Results
- Applications
- Software and LLMs
- 6 How can we trust this?
- Future directions for Forecasting Amid Shocks
- Directions and Limitations

Premise: There is breaking news at some fractional lag t- ϵ

- Market (i.e. crowd-sourced) mechanisms for evaluating the news are either offline, too thin, or otherwise unavailable. They could also be simply unreliable.
- The qualitative aspects of the news provide a basis upon which to
 - match to past news shocks
 - match in a p-dimensional covariate space

Model Setup

For $t = 1, ..., T_i$ and i = 1, ..., n + 1, the model \mathcal{M}_1 is defined as

$$\begin{aligned} y_{i,t} &= F(\mathcal{F}_{i,t-1}) + \alpha_{i,t} + \epsilon_{i,t} \\ \alpha_{i,t} &= \mathbf{x}_{i,t}^T \lambda_{i,t} \\ \mathbf{x}_{i,t}^T &= (1, x_{i,t}^1, ..., x_{i,t}^p) \text{ (observable and deterministic w.r.t to } \mathcal{F}_{i,t-1}) \\ \lambda_{i,t}^T &= (u_{i,t}, \lambda_t^1, ..., \lambda_t^p) \text{ (unobservable and potentially random),} \end{aligned} \tag{1}$$

$$\lambda_{i,t} \sim \mathcal{F}_{\lambda} \text{ with } \mathbb{E}_{\mathcal{F}_{\lambda}}(\lambda) = \mu_{\lambda_t}, \operatorname{Var}_{\mathcal{F}_{\lambda}}(\lambda) = \Sigma_{\lambda_t},$$

and time-invariant error structure

$$\epsilon_{i,t} \stackrel{iid}{\sim} \mathcal{F}_{\epsilon} \text{ with } \mathbb{E}_{\mathcal{F}_{\epsilon}}(\epsilon) = 0, \operatorname{Var}_{\mathcal{F}_{\epsilon}}(\epsilon) = \sigma_{\epsilon}^{2},$$



Model Details

Note that
$$\lambda_{i,t}^T = (u_{i,t}, \lambda_t^1, ..., \lambda_t^p)$$

- includes the term $u_{i,t}$ that is not shared among donors.
- is time-varying allows us to capture that most time points are without news shocks (in which case $\alpha_{i,t}$ is of negligible effect), but conditional upon information arriving between T_1^* and T_1^*+1 , we may know with near certainty that λ_{1,T_1^*+h} will be nonzero in norm for some h>0.

Significance of the Covariates $x_{i,t}$

Covariates chosen for inclusion may be any \mathcal{F}_t -measurable function, for example

- levels
- differences in levels
- log returns
- percentage returns
- measurable transformations of the above

Key criterion for inclusion: how plausible is the covariate as a proxy for risk conditions for the time series under study?

Forecasting

We now present two one-step-ahead forecasts. First is the unadjusted forecast. The second is the adjusted forecast, which differs by the predicted correction term:

Forecast 1:
$$\hat{y}_{unadjusted, T_1^*+1} = \hat{\mathbb{E}}[y_{1, T_1^*+1} | \mathcal{F}_{T_1^*}]$$

Forecast 2:
$$\hat{y}_{adjusted, T_1^*+1} = \hat{\mathbb{E}}[y_{1, T_1^*+1} | \mathcal{F}_{T_1^*}] + \hat{\alpha}_{T_1^*+1}$$
.

Distance-based Weighting in Action

(For ease of exposition, we omit time indices)

 \bullet Observe the pair $(\{\hat{\alpha}_i\}_{i=2}^{n+1},\{\mathbf{x}_i\}_{i=2}^{n+1}).$

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- Goal: recover weights $\{\pi_i\}_{i=2}^{n+1} \in \Delta^n$ and compute $\hat{\alpha} := \sum_{i=2}^{n+1} \pi_i \hat{\alpha}_i$, our predicted correction term (more about this later).

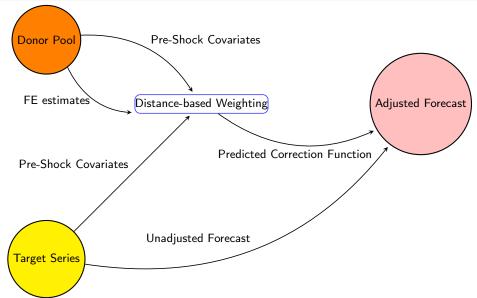
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- Following Abadie and Gardeazabal 2003, Abadie, Diamond, and Hainmueller 2010, let $\|\cdot\|_{\mathbf{S}}$ denote any semi-norm on \mathbb{R}^p , and define

$$\{\pi\}_{i=2}^{n+1} = \operatorname*{arg\,min}_{\pi} \|\mathbf{v}_{\mathbf{1},T^*} - \mathbf{X}_{T^*}\pi\|_{\mathbf{S}} \ .$$

Visuals That Tell The Story



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The forecaster's decision tree

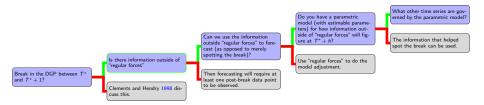


Figure: Forecast Model Adjustment: A Decision Tree

Visuals That Tell The Story

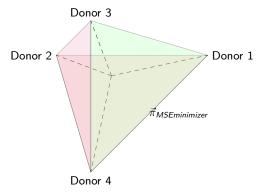


Figure: The 3-Simplex, Δ^3 , where hypothetical minimizer is a convex combination of Donors 1 and 4.

Table: A 2x2 Schema of Forecast Information, With Examples

	Conventional Econometric Models	Outside Conventional Econometric
		Models
internal	Lags of the series itself; past shocks	Polynomial expansion of the fea-
		ture space and other transforma-
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To give a name to what we're doing, we can call it distance-based weighting, which is a special case of an approximation-based *n*-ary similarity metric.

The term $\mathbf{x}_{i,t}^T \lambda_{i,t}$ is deemed the *correction term* for donor i at time t, and the function

$$\xi \colon \mathcal{F}_{2,T_{\mathbf{2}}^*+h} \times \ldots \times \mathcal{F}_{n+1,T_{n+1}^*+h} \to \mathcal{F}_{1,T_{\mathbf{1}}^*+h}$$

that estimates or predicts α_{1,T_1^*+h} will be deemed the correction function.

The correction function maps donor information — be it observables or inferences — to the prediction function for the time series under study.

Abstracting away from particular models, what does our method require?

 Object-to-predict Random object (indexed over time and possibly space, as well) obeying specification with additive errors

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- **Q** Reliable and Shared Model-Fitting Procedure There must exist a reliable model-fitting procedure for the n+1 units.
- Reliable Correction Term Estimation
- Reliable Correction Function Estimation There must exist a correction function (presumably based on the correction term) that maps data from the donor pool to the predicted correction term in the time series under study based on similarity.

Do we have theoretical guarantees?

We do!

Our formal results fall into the following buckets:

- convergence in distribution results
- consistency results
- asymptotic loss

Proposition

Assume

- For each i, $1 \le i \le n+1$, let $\{y_t\}_{i=1}^{T_i}$ follow an AR(p)-X as laid out in Section ??.
- **3** Assume for each i, $1 \le i \le n+1$, the shocks $\alpha_{i,t}$ are uncorrelated across donors.
- **3** Assume for each i, $1 \le i \le n+1$, the shocks $\alpha_{i,t}$ are uncorrelated with α_{i,T_i^*+1}

Then the tuple of estimators $(\hat{\rho}_{i,1},...,\hat{\rho}_{i,p},\hat{\alpha}_{i,T^*+1})$ is consistent as $t \to \infty$.



Formal Results

Proposition

Assume

- All conditions listed in Proposition 1.
- **②** There exist weights $\{\pi_i\}_{i=2}^{n+1} \in \Delta^{n-1}$ such that $\mathbf{v}_{1,T_1^*} = \sum_{i=2}^{n+1} \pi_i \mathbf{v}_{i,T_i^*}$.
- **3** For all i, the $\{u_{i,t}\}$ are equal in distribution.

Then the aggregated estimator $\alpha_{T_1^*+1}$ converges in distribution to $\alpha_{T_1^*+1}$ as $t \to \infty$. Furthermore, if the $\{u_{i,t}\}$ are constant with probability 1, the convergence is in probability.

Formal Results

Proposition

Let $\{\hat{y}_{1,T_{1}^{*}+r}\}_{r=1}^{h}$ denote the vector of adjusted predictions (adjusted through h steps ahead) in the time series under study. Assume all conditions listed in Propositions 1 and 2. Then $\{\hat{y}_{1,T_{1}^{*}+r}\}_{r=1}^{h} \xrightarrow{d} \{y_{1,T_{1}^{*}+r}\}_{r=1}^{h}$. Furthermore, if the $\{u_{i,t}\}$ are constant with probability 1, the convergence is in probability.



An inventory of models

• GARCH (paper is under review at the International Journal of Forecasting)



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- Exponential Shocks (to be submitted in the coming weeks as part of second paper)

GARCH Volatility Forecasts

The least you need to know here: GARCH is the premier volatility modeling family for daily data.

We demonstrate our method's usefulness with an easy-to-approach example.



Why apply our method to the 2016 US Election?

- You can win the US Presidency without a majority.
- No incumbent candidate
- Donald J. Trump espoused unorthodox, populist positions on healthcare, trade, foreign policy
- Donald J. Trump had no record to assess or criticize
- It was not predicted hence it delivered news.



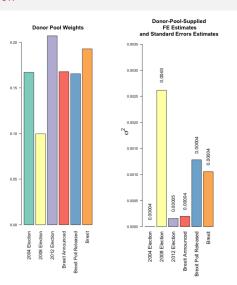
iShares U.S. Financial Services ETF

Figure: IYG includes JPM, BAC, WF, CITI, among other financial majors

- Model choice GARCH(1,1) on the daily log return series of IYG in each donor
- Covariate Choice
 - previous 30 log returns of IYG (large pre-treatment period, in the language of SC)
 - log return Crude Oil (CL.F)
 - VIX
 - log return of the VIX
 - log returns of the 3-month, 5-year, 10-year, and 30-year US Treasuries
 - return of the most recently available monthly spread between AAA and BAA corporate debt
 - log return in the trading volume of the ETF IYG itself
- 3 Donor pool construction US Elections from 2004, 2008, 2012
- Choice of estimator for volatility Sum of 77 squared five-minute returns generated between 9:35am and 4pm on November 9th, 2016.



2016 Election



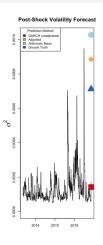


Figure: The volatility induced by the 2016 US election

Heterogeneous Autoregression (HAR)

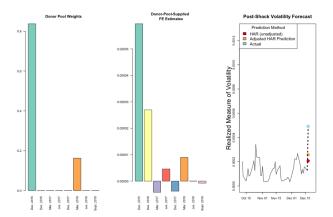


Figure: Click the image above to see a video of the R package's functionality.



Simulations: Parameter Correction Using An Aggregated Decay Parameter from Donors

Recall an \mathcal{M}_1 model on the volatility, which is characterized by an exogenous shock to the volatility equation generated by an affine function of the covariates.

Now consider, for i = 1, 2, ..., n + 1, the family of models

$$y_{i,t} = \alpha_i + \eta [1 - e^{-\psi_i [t - T_i^*]}] \mathbf{1}_{t \ge T_i^* + 1} + \epsilon_{i,t}, \epsilon_{i,t}$$

$$\psi_i = \mathbf{x}_{i,t}^T \lambda_{i,t}$$

$$\mathcal{M}_{exp}: \quad \mathbf{x}_{i,t}^T = (1, \mathbf{x}_{i,t}^1, ..., \mathbf{x}_{i,t}^p)$$

$$\lambda_{i,t}^T = (u_{i,t}, \lambda_1^1, ..., \lambda_t^p),$$
(2)

with time-varying and observable covariate vector $\mathbf{x}_{i,t}^T$, time-varying, unobservable, and potentially random vector

$$\lambda_{i,t} \sim \mathcal{F}_{\lambda} \text{ with } \mathbb{E}_{\mathcal{F}_{\lambda}}(\lambda) = \mu_{\lambda_t}, \operatorname{Var}_{\mathcal{F}_{\lambda}}(\lambda) = \Sigma_{\lambda_t},$$

and time-invariant error structure just as in \mathcal{M}_1 .



Simulations: Parameter Correction Using An Aggregated Decay Parameter from Donors

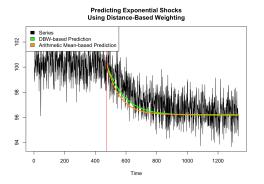


Figure: Simulated random walk centered at $\mu=100$, subject to a shock at approximately time index 440. Shock of size -4 is not realized over a single index. Instead, shock is governed by a decay exponential decay parameter ψ_1 , as are exponential shocks in n=10 donor series. We estimate ψ_1 using a convex combination of the estimated decay parameters $\{\psi_i\}_{i=2}^{n+1}$, resulting in a distance-based weighting prediction of the shock. We also illustrate the arithmetic-mean-based prediction of the shock using the color orange. This prediction is based on an overestimate of ψ_1 and hence results in residuals that undershoot the time series under study.

Software

R package

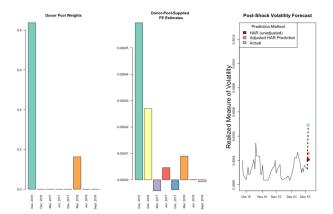


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LLMs

Use NLP to identify donors.



What are some events that are similar to Chinese government's ban on iPhones for government officials in September 2023?

ChatGPT

The Chinese government's ban on iPhones for government officials in September 2023 is a notable example of government-led restrictions or bans on certain technology products or services due to security concerns or political reasons. Here are some events or actions that are somewhat similar:

- Huawel Ban in the United States: The United States government banned Huawel, a Chinese telecommunications company, from selling its equipment and infrastructure in the U.S. due to concerns about national security and espionage. This ban was initiated in 2019 and has had significant repercussions for Huawel's business olobal.
- Russian Ban on Telegram: In 2018, the Russian government banned Telegram, a
 popular messaging app, due to the company's refusal to provide encryption keys to
 Russian security agencies. The ban resulted in the disruption of Telegram services
 within Russia and sparked debates about privacy and government surveillance.
- 3. Indisk Ban on Chinese Apps: In 2020, the Indisin poverment banned several Chinese mobile apps, including TikTolk and WeChat, citing concerns about data privacy and national security after a border clash between Indian and Chinese toops. The ban affected millions of users and led to diplomatic tensions between India and Chin.
- 4. Turkey's Wilepedia Ban: No/IT, the Turkish government banned access to Wilepedia, alleging that the website was involved in a smear campaign against Turkey. The ban stemmed from Wilepedia's refusal to remove content that the Turkish government deemed offensive. Despite international criticism, the ban remained in place for over two years before being little in 2020.
- 5. Iran's Ban on Messaging Apps: The Iranian government has intermittently banned messaging apps such as Telegram and WhatsApp over concerns about the spread or insinformation, political dissent, and national security threats. These bans have been part of broader efforts to control online communication and social media polatforms within Iran.

These examples demonstrate how governments may impose bans or restrictions on technology products, services, or platforms for various reasons, including national security, political control, and regulatory compliance.

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Robustness of the Approach

What assurances do we have that the method will be good?

Some ideas here:

- multiverse
- permute donors

We shall group the extensions into five buckets:

- How much can we automate?
- Alternatives for fixed effect estimation
- Alternative estimators and estimands
- What can you do with a volatility forecast?
- Where else is distanced-based weighting useful?
- Can we extend the results of (Bodilsen and Lunde 2023)

How much can we automate?

What if the covariates are difficult to specify?

Proposed solution:

Use shrinkage estimation to detect fleeting signals in the cross section of a_t^2 (Chinco, Clark-Joseph, and Ye 2019).

Limitations of what we're currently doing

- Our real data examples cannot be scaled up due to the need to for human involvement in donor and covariate curation
- ullet We assume that the exogenous variables (distinct from the covariates) are known at time T_i^*
- Heterogeneity of DGPs
- Forecast evaluation
- We could only truly explore a subset of the vast parameter space.
- The failure of distance-based weighting to extrapolate

New Frontiers in Distance-based Weighting

- Forecast conditional upon the shock prior to the shock's arrival
- Integrate lessons from literature on under/over reactions to information shocks (Jiang and Zhu 2017)
- Distance-based Weighting of Impulse Response Functions

Distance-based Weighting of Impulse Response Functions

Suppose

- We have a collection of *p*-variate time series of lengths T_i , i = 1, 2, ...n + 1.
- We are interested in the response of variable r to shocks in variable j, $1 \le r \le j \le p$.

There many ways to estimate $IRF_1(r, j)$.

Can we somehow aggregate the estimates $\widehat{IRF}_i(r,j)$, i=2,3,...,n+1? Additional research questions:

- What DGP would best motivate/justify such a method?
- Which method of IRF estimation would perform best?

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