

Forecast Adjustment Under Shocks: Similarity-based Solutions to Unprecedented Events

Doctoral Defense

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A High-level Summary Uniting All My Projects

Forecasting under non-ideal conditions like

- 1 a rupture in the DGP
- 2 lack of past information for units we wish to forecast, i.e. a so-called "cold start" (Fatemi et al. 2023)

What should we do? Is there a meaningful distinction between not knowing the DGP and knowing it but lacking the data?

Central technique explored: *correction of model parameters using an aggregation strategy*, where aggregation means that we make use of external data.

We began with particular models; then built upwards, establishing an abstract, general framework for correcting forecasts that is widely applicable for forecasting and beyond.

Why is this valuable?

We are weighing-in on a puzzle:

conflict between the intuitive notion that more relevant information should help in forecasting, and the hard reality that attempts to make it do so have not been uniformly successful (Clements and Hendry [2005](#))

Why is this valuable?

We provide a precise proposal to a well-specified challenge:

incomplete information by itself is unlikely to play a key role in forecast failure (except if that information would forecast breaks). Consequently, using large amounts of data may not correct one of the main problems confronting forecasters, namely [location shifts](#), unless that additional information is directly pertinent to forecasting breaks (Castle, Clements, and Hendry [2013](#))

Why is this valuable?

- Canonical problems: predicting based on a limited feature space and/or small n
- The framework we offer is something others can 'plug-in' to and add to.
- Software developed

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- 1 What does it resemble from the past?
- 2 What past events are most relevant for our objectives?
- 3 Can we incorporate past events in a systematic, principled manner?

Example (Weekend of March 6th - 8th, 2020)



Oil nose-dives as Saudi Arabia and Russia set off 'scorched earth' price war

PUBLISHED SUN, MAR 8 2020-9:01 AM EDT | UPDATED MON, MAR 9 2020-5:33 PM EDT

Oil crashes by most since 1991 as Saudi Arabia launches price war



By [Matt Egan](#), CNN Business

🕒 3 minute read · Updated 3:21 PM EDT, Mon March 9, 2020

Punchline

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Table: A 2x2 Schema of Forecast Information, With Examples

	Conventional Econometric Models	Outside Conventional Econometric Models
internal	Lags of the series itself; past shocks	Polynomial expansion of the feature space and other transformations without solid theoretical motivation
external	Macro variables like interest rates, commodity prices; weather-related variables	Google Trends, high-frequency data like prediction markets; past shocks under similar conditions

Outline

- 1 Introduction
- 2 The Idea and Methodology
- 3 Formal Results
- 4 Applications
- 5 Software and LLMs
- 6 How can we trust this?
- 7 Future directions for Forecasting Amid Shocks
- 8 Directions and Limitations

Premise: There is a breaking news at some fractional lag $t-\epsilon$

- After-hours trading provides a poor forum in which to digest news
- News constitutes public, material information for one or more traded assets
- The **qualitative aspects** of the news provide a basis upon which to
 - match to past news shocks
 - match in a p -dimensional covariate space

Model Setup

For $t = 1, \dots, T_i$ and $i = 1, \dots, n + 1$, the model \mathcal{M}_1 is defined as

$$\begin{aligned} y_{i,t} &= F(\mathcal{F}_{i,t-1}) + \alpha_{i,t} + \epsilon_{i,t} \\ \alpha_{i,t} &= \mathbf{x}_{i,t}^T \lambda_{i,t} \\ \mathcal{M}_1: \quad \mathbf{x}_{i,t}^T &= (1, x_{i,t}^1, \dots, x_{i,t}^p) \\ \lambda_{i,t}^T &= (u_{i,t}, \lambda_t^1, \dots, \lambda_t^p), \end{aligned} \tag{1}$$

with time-varying and observable covariate vector $\mathbf{x}_{i,t}^T$, time-varying, unobservable, and potentially random vector

$$\lambda_{i,t} \sim \mathcal{F}_\lambda \text{ with } \mathbb{E}_{\mathcal{F}_\lambda}(\lambda) = \mu_{\lambda_t}, \text{Var}_{\mathcal{F}_\lambda}(\lambda) = \Sigma_{\lambda_t},$$

and time-invariant error structure

$$\epsilon_{i,t} \stackrel{iid}{\sim} \mathcal{F}_\epsilon \text{ with } \mathbb{E}_{\mathcal{F}_\epsilon}(\epsilon) = 0, \text{Var}_{\mathcal{F}_\epsilon}(\epsilon) = \sigma_\epsilon^2,$$

Model Details

Note

- F maps objects from the past into future objects of $y_{i,t}$
- $\epsilon_{i,t}$ uncorrelated across time and donors
- The dot product $\alpha_{i,t} = \mathbf{x}_{i,t}^T \lambda_{i,t}$ includes the term $u_{i,t}$ that is not shared among donors. It encodes the known possibility that the shocks will vary across donors both in ways that can predicted based on pre-shock covariates as well as in ways that cannot be predicted.
- That the parameter vector is time-varying allows us to capture that most time points are without news shocks (in which case $\alpha_{i,t}$ is of negligible effect), but conditional upon information arriving between T_1^* and $T_1^* + 1$, we may know with near certainty that λ_{1,T_1^*+h} will be nonzero in norm for some $h > 0$.

Significance of the Covariates $\mathbf{x}_{i,t}$

Covariates chosen for inclusion may be any \mathcal{F}_t -measurable function, for example

- levels
- differences in levels
- log returns
- percentage returns
- measurable transformations of the above

Key criterion for inclusion: how plausible is the covariate as a **proxy for risk conditions** for the volatility series to be forecasted?

Forecasting

We now present two one-step-ahead forecasts for the time series under study.

The first is the unadjusted forecast. The second is the adjusted forecast, which differs by the predicted correction term. The forecasts are:

$$\text{Forecast 1: } \hat{y}_{unadjusted, T_1^*+1} = \hat{\mathbb{E}}[\mathcal{F}_{T_1^*}]$$

$$\text{Forecast 2: } \hat{y}_{adjusted, T_1^*+1} = \hat{\mathbb{E}}[\mathcal{F}_{T_1^*}] + \hat{\alpha}_{T_1^*+1} .$$

Distance-based Weighting in Action

- Observe the pair $(\{\hat{\omega}_i^*\}_{i=2}^{n+1}, \{\mathbf{v}_i\}_{i=2}^{n+1})$.

Distance-based Weighting in Action

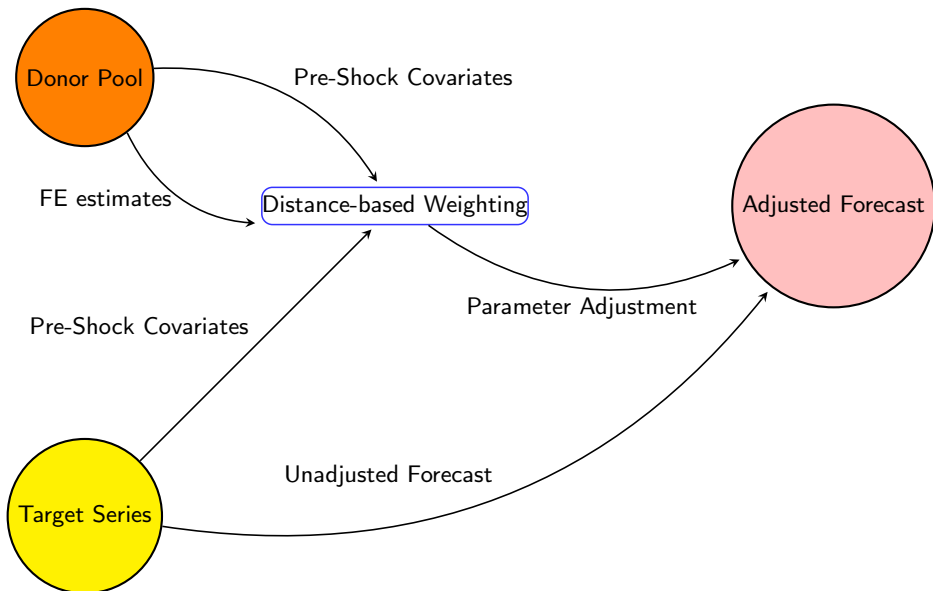
- Observe the pair $(\{\hat{\omega}_i^*\}_{i=2}^{n+1}, \{\mathbf{v}_i\}_{i=2}^{n+1})$.
- Goal: recover weights $\{\pi_i\}_{i=2}^{n+1} \in \Delta^n$ and compute $\hat{\omega}^* := \sum_{i=2}^{n+1} \pi_i \hat{\omega}_i^*$, our forecast adjustment term.

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- Following Abadie and Gardeazabal [2003](#), Abadie, Diamond, and Hainmueller [2010](#), let $\|\cdot\|_S$ denote any semi-norm on \mathbb{R}^P , and define

$$\{\pi\}_{i=2}^{n+1} = \arg \min_{\pi} \|\mathbf{v}_{1,T^*} - \mathbf{V}_{T^*} \pi\|_S .$$

Visuals That Tell The Story



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The forecaster's decision tree

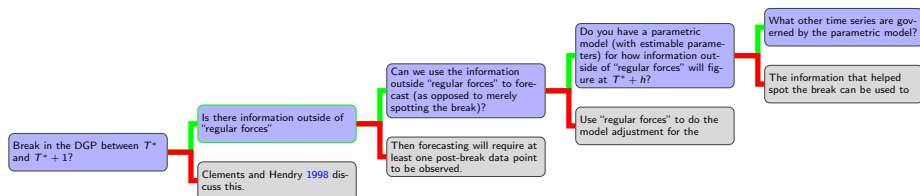


Figure: Forecast Model Adjustment: A Decision Tree

Visuals That Tell The Story

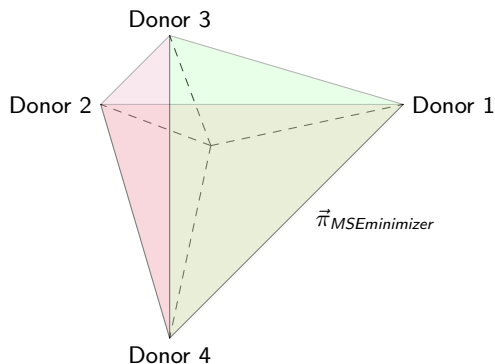


Figure: The 3-Simplex, Δ^3 , where hypothetical minimizer is a convex combination of Donors 1 and 4.

Global Overview

- 1 **Object-to-predict** Most fundamentally, the method requires a random object (indexed over time and possibly space, as well) that obeys a specification with additive errors, or, at the very least, a specification that can be transformed to have additive errors.
- 2 **Common Model Family on the Shocks** The method requires that residuals be governed by a model that is shared across all units. This ensures that in the estimation of news shocks in the donor pool, the estimators will enjoy similar properties that will produce a good aggregated shock estimator. This condition is satisfied by the parametric shock distributions found in Lin and Eck 2021; Lundquist and Eck 2024.
- 3 **Reliable and Shared Model-Fitting Procedure** There must exist a reliable model-fitting procedure for the $n + 1$ units.
- 4 **Reliable Correction Term Estimation**
- 5 **Reliable Correction Function Estimation** There must exist a correction function (presumably based on the correction term) that maps data from the donor pool to the *predicted* correction term in the time series under study based on some notion of similarity. In some cases, there may exist a posited DGP that the correction term estimates. In other cases, there may be no posited DGP.

Formal Results

Proposition

Assume

- 1 For each i , $1 \leq i \leq n+1$, let $\{y_t\}_{t=1}^{T_i}$ follow an $AR(p)$ -X as laid out in Section ??.
- 2 Assume for each i , $1 \leq i \leq n+1$, the shocks $\alpha_{i,t}$ are uncorrelated across donors.
- 3 Assume for each i , $1 \leq i \leq n+1$, the shocks $\alpha_{i,t}$ are uncorrelated with α_{i,T_i^*+1}

Then the tuple of estimators $(\hat{p}_{i,1}, \dots, \hat{p}_{i,p}, \hat{\alpha}_{i,T_i^*+1})$ is consistent as $t \rightarrow \infty$.

Formal Results

Proposition

Assume

- ① All conditions listed in Proposition 1.
- ② There exist weights $\{\pi_i\}_{i=2}^{n+1} \in \Delta^{n-1}$ such that $\mathbf{v}_{1,T_1^*} = \sum_{i=2}^{n+1} \pi_i \mathbf{v}_{i,T_i^*}$.
- ③ For all i , the $\{u_{i,t}\}$ are equal in distribution.

Then the aggregated estimator $\alpha_{T_1^*+1}^{\hat{}}$ converges in distribution to $\alpha_{T_1^*+1}$ as $t \rightarrow \infty$. Furthermore, if the $\{u_{i,t}\}$ are constant with probability 1, the convergence is in probability.

Formal Results

Proposition

Let $\{\hat{y}_{1, T_1^*+r}\}_{r=1}^h$ denote the vector of adjusted predictions (adjusted through h steps ahead) in the time series under study. Assume all conditions listed in Propositions 1 and 2. Then $\{\hat{y}_{1, T_1^*+r}\}_{r=1}^h \xrightarrow{d} \{y_{1, T_1^*+r}\}_{r=1}^h$. Furthermore, if the $\{u_{i,t}\}$ are constant with probability 1, the convergence is in probability.

My Prelim: An Incredibly Brief Recap

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- Event-driven investing strategies (unscheduled news shock)
- Scheduled macroeconomic news possibly pre-empted by a news leak

ECONOMY

Fed Likely to Consider 0.75-Percentage-Point Rate Rise This Week

Officials had signaled plans to raise interest rates in half-point increments before recent deterioration in data

By Nick Timiraos [Follow](#)

Updated June 13, 2022 7:47 pm ET

Why apply our method to the 2016 US Election?

- You can win the US Presidency without a majority.
- No incumbent candidate
- Donald J. Trump espoused unorthodox, populist positions on healthcare, trade, foreign policy
- Donald J. Trump had no record to assess or criticize
- It was not predicted – hence it delivered news.

IYG

iShares U.S. Financial Services ETF

Figure: IYG includes JPM, BAC, WF, CITI, among other financial majors

- ➊ **Model choice** GARCH(1,1) on the daily log return series of IYG in each donor
- ➋ **Covariate Choice**
 - previous 30 log returns of IYG (large pre-treatment period, in the language of SC)
 - log return Crude Oil (CL.F)
 - VIX
 - log return of the VIX
 - log returns of the 3-month, 5-year, 10-year, and 30-year US Treasuries
 - return of the most recently available monthly spread between AAA and BAA corporate debt
 - log return in the trading volume of the ETF IYG itself
- ➌ **Donor pool construction** US Elections from 2004, 2008, 2012
- ➍ **Choice of estimator for volatility** Sum of 77 squared five-minute returns generated between 9:35am and 4pm on November 9th, 2016.

2016 Election

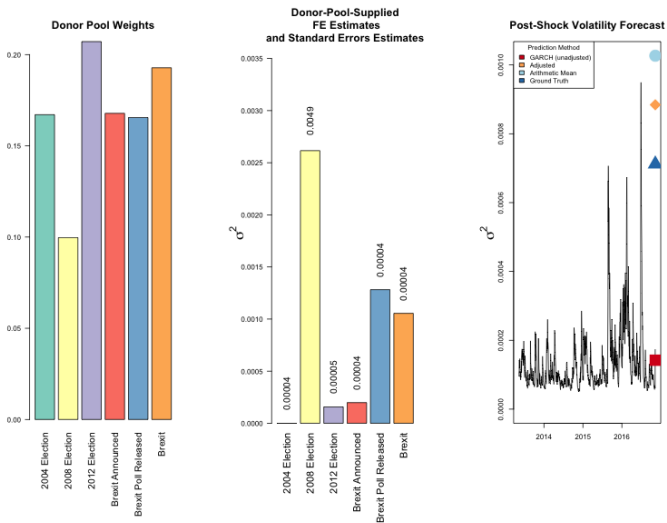


Figure: The volatility induced by the 2016 US election

The Post-prelim Direction of my research

- Extending similarity-based parameter correction to the state-of-the-art HAR model
- Extending similarity-based parameter correction to non-linear shock models
- Building out a general framework for parameter correction

Heterogeneous Autoregression (HAR)

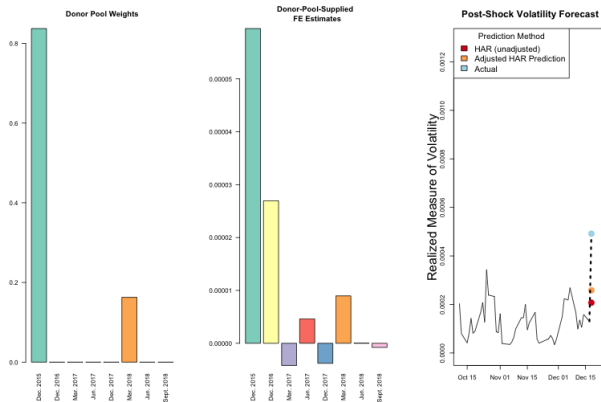


Figure: Click the image above to see a video of the R package's functionality.

Simulations: Parameter Correction Using An Aggregated Decay Parameter from Donors

Recall an \mathcal{M}_1 model on the volatility, which is characterized by an exogenous shock to the volatility equation generated by an affine function of the covariates.

Now consider, for $i = 1, 2, \dots, n + 1$, the family of models

$$\begin{aligned} y_{i,t} &= \alpha_i + \eta[1 - e^{-\psi_i[t - T_i^*]}]\mathbf{1}_{t \geq T_i^*+1} + \epsilon_{i,t}, \epsilon_{i,t} \\ \psi_i &= \mathbf{x}_{i,t}^T \lambda_{i,t} \\ \mathcal{M}_{exp}: \quad \mathbf{x}_{i,t}^T &= (1, x_{i,t}^1, \dots, x_{i,t}^p) \\ \lambda_{i,t}^T &= (u_{i,t}, \lambda_t^1, \dots, \lambda_t^p), \end{aligned} \tag{2}$$

with time-varying and observable covariate vector $\mathbf{x}_{i,t}^T$, time-varying, unobservable, and potentially random vector

$$\lambda_{i,t} \sim \mathcal{F}_\lambda \text{ with } \mathbb{E}_{\mathcal{F}_\lambda}(\lambda) = \mu_{\lambda_t}, \text{Var}_{\mathcal{F}_\lambda}(\lambda) = \Sigma_{\lambda_t},$$

and time-invariant error structure just as in \mathcal{M}_1 .

Simulations: Parameter Correction Using An Aggregated Decay Parameter from Donors

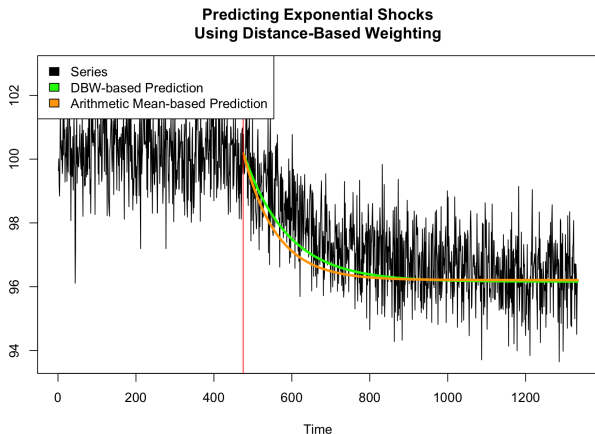


Figure: We simulate a random walk centered at $\mu = 100$, subject to a shock at approximately time index 440. Shock of size -4 is not realized over a single index. Instead, shock is governed by a decay exponential decay parameter ψ_1 , as are exponential shocks in $n = 10$ donor series. We estimate ψ_1 using a convex combination of the estimated decay parameters $\{\psi_1, \dots, \psi_{n+1}\}$ resulting in

Software

R package

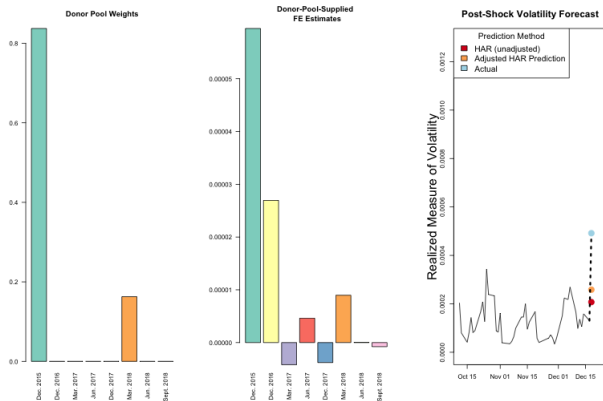


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LLMs

Use NLP to identify donors.

You

What are some events that are similar to Chinese government's ban on iPhones for government officials in September 2023?

ChatGPT

The Chinese government's ban on iPhones for government officials in September 2023 is a notable example of government-led restrictions or bans on certain technology products or services due to security concerns or political reasons. Here are some events or actions that are somewhat similar:

- Huawei Ban in the United States:** The United States government banned Huawei, a Chinese telecommunications company, from selling its equipment and infrastructure in the U.S. due to concerns about national security and espionage. This ban was initiated in 2019 and has had significant repercussions for Huawei's business globally.
- Russian Ban on Telegram:** In 2018, the Russian government banned Telegram, a popular messaging app, due to the company's refusal to provide encryption keys to Russian security agencies. The ban resulted in the disruption of Telegram services within Russia and sparked debates about privacy and government surveillance.
- India's Ban on Chinese Apps:** In 2020, the Indian government banned several Chinese mobile apps, including TikTok and WeChat, citing concerns about data privacy and national security after a border clash between Indian and Chinese troops. The ban affected millions of users and led to diplomatic tensions between India and China.
- Turkey's Wikipedia Ban:** In 2017, the Turkish government banned access to Wikipedia, alleging that the website was involved in a smear campaign against Turkey. The ban stemmed from Wikipedia's refusal to remove content that the Turkish government deemed offensive. Despite international criticism, the ban remained in place for over two years before being lifted in 2020.
- Iran's Ban on Messaging Apps:** The Iranian government has intermittently banned messaging apps such as Telegram and WhatsApp over concerns about the spread of misinformation, political dissent, and national security threats. These bans have been part of broader efforts to control online communication and social media platforms within Iran.

These examples demonstrate how governments may impose bans or restrictions on technology products, services, or platforms for various reasons, including national security, political control, and regulatory compliance.



Robustness of the Approach

What assurances do we have that the method will be good?

Some ideas here:

- 1 multiverse
- 2 permute donors

We shall group the extensions into five buckets:

- How much can we automate?
- Alternatives for fixed effect estimation
- Alternative estimators and estimands
- What can you do with a volatility forecast?
- Where else is distanced-based weighting useful?
- Can we extend the results of (Bodilsen and Lunde [2023](#))

How much can we automate?

What if the covariates are difficult to specify?

Proposed solution:

Use shrinkage estimation to detect fleeting signals in the cross section of a_t^2 (Chinco, Clark-Joseph, and Ye [2019](#)).

Limitations of what we're currently doing

- Our real data examples cannot be scaled up due to the need to for human involvement in donor and covariate curation
- We assume that the exogenous variables (distinct from the covariates) are known at time T_i^*
- Heterogeneity of DGPs
- Forecast evaluation
- We could only truly explore a subset of the vast parameter space.
- The failure of distance-based weighting to extrapolate

New Frontiers in Distance-based Weighting

- Forecast conditional upon the shock *prior* to the shock's arrival
- Integrate lessons from literature on under/over reactions to information shocks (Jiang and Zhu [2017](#))
- Distance-based Weighting of Impulse Response Functions

Distance-based Weighting of Impulse Response Functions

Suppose

- We have a collection of p -variate time series of lengths $T_i, i = 1, 2, \dots, n + 1$.
- We are interested in the response of variable r to shocks in variable j , $1 \leq r \leq j \leq p$.

There many ways to estimate $IRF_1(r, j)$.

Can we somehow aggregate the estimates $\widehat{IRF}_i(r, j), i = 2, 3, \dots, n + 1$?

Additional research questions:

- What DGP would best motivate/justify such a method?
- Which method of IRF estimation would perform best?

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