

Antibody Response Induced by HIV Vaccines and T-cell Suppression Treatments in Rhesus Macaques

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1 Introduction

1.1 About the Study

A dominant vaccine development strategy is to induce neutralizing antibodies by immunizing humans with the virus' glycoproteins. However, HIV vaccines that adopted this strategy mostly failed due to the fact that HIV is an RNA virus, which mutates rapidly to escape the inhibition of neutralizing antibodies. By the time the body generates neutralizing antibodies against the glycoproteins of some HIV strains, the RNA virus has already mutated. Thus, the existing neutralizing antibody fails to recognize, bind with, and neutralize the HIV virus. One possible solution is to increase the number of potential neutralizing antibodies that will cycle in the body by releasing a variety of antibodies after glycoprotein immunization. Regulatory T (Treg) cells prevent autoimmune diseases and suppress allergic reactions by inhibiting adaptive antibody immune response in the germinal center. Theoretically, this adaptive response lowers the effectiveness of vaccines. Thus the experiment used T-cell suppression treatments to investigate the effect on immunization. These drugs are widely used in post transplant immuno-suppression treatment to prevent rejection.

During the study, 20 rhesus macaques were given glycoprotein immunization and supplemental antibody doses, as well as one of three treatments (two experimental regulatory T-cell suppression treatments and one control). Our dataset includes measurements of antibodies after the 20 macaques were given the same HIV vaccine at three different time-points and one of three randomly selected anti-Treg treatments(drugs). Blood samples were collected two weeks after vaccine dosing, and antibodies were isolated from those samples. A different number of antibodies were collected from each blood sample, limited by assay yield. Each observation contains information about the antibody isolated post the glycoprotein immunization.

A human antibody is formed by a heavy chain and light chain. For heavy chain, a human has about 51 V-gene segments, 25 D-gene segments and 6 J-gene segment. For light chain there are 71 V-gene segments and 9 J-gene segments[ref.5]. Any heavy chain V-D-J combination and light chain V-J combinations can randomly happen in germline center. Theoretically, there can be $51 * 25 * 6 * 71 * 9 = 4.88835 \times 10^6$ combinations of gene segments. Considering the frequently happened mutation and other factors, each individual can have over **10 billion** different antibodies. Thus, we decided to follow the convention of vaccine studies and treat each antibody as independent.

Below is the list of variables with a brief description from our dataset. Please note that in each antibody, there are two sets of heavy chain and light chain, all of which forming a Y-shape immunoglobulin. Thus many of the variables start with H or L, indicating which chain the information comes from.

1.2 List of variables

- **Monkey_id:** Lists the identity of monkey
- **Treatment(Drug):** Treatment A is the mock control, and treatment B and C are two different kinds of Treg inhibitor treatments.
- **Time_Points:** 0 represents before immunization; 1 represents 2 weeks post 1st immunization; 2 represents 2 weeks post 2nd immunization; and 3 represents 2 weeks post 3rd immunization, respectively.
- **Isotype:** The category of antibody type; there are 5 kinds of immunoglobulin isotypes: IgG, IgA, IgM, IgE, IgD. The two most important kinds are IgG and IgM. IgM occurs in the acute stage of infection and perform an role of primary response. The secondary response IgG appears later in serum with higher binding affinity, and neutralizing potentials against toxins and virus. IgA mostly found in mucosal tissues such as Nasal mucosa. Non-dominant IgD and IgE are typically lower than 1% in blood.
- **H_ID and L_ID:** heavy chain and light chain IDs for the particular observation
- **H_VBase:** the number of nucleotide of the heavy chain variable region
- **H_Substitutions:** the number of relative nucleotide mutations in heavy chain.
- **HMufreq:** calculated by $H_Substitutions / H_VBase$
- **H_CDR3:** the number of amino acid of the heavy chain's third complementarity determining region
- **L_VBase:** the number of nucleotide of the light chain variable region
- **L_Substitutions:** the number of relative nucleotide mutations in light chain.
- **LMufreq:** calculated by $L_Substitutions / L_VBase$
- **L_CDR3:** the number of amino acid of the light chain's third complementarity determining region. H_CDR3 and L_CDR3 indicates the length of the third complementarity-determining

region on the variable heavy chain and light chain. The longer they are, the more potential there is to produce diverse antibodies. [Kan, could you check to see if this is correct?] In other words, we want the values to be higher.

- Binding: affinity of antibodies against a selected HIV glycoprotein. The larger value indicates stronger binding. Binding indicates the rate of neutralizing, meaning how much the antibodies bind with the virus and thus make the virus ineffective. This is the most important measure of the study.

The dataset, which can be found here, was provided by Kan Luo, as he was one of authors for the following four publications that used the dataset. After stating our research questions, the methods section includes exploratory data analysis, multivariate data analysis, and longitudinal data analysis. The statistical results are included in the methods section and summarized in the results section. We provide implications and limitations in the discussions and some final thoughts in the conclusions.

1.3 Research questions

The main focus of the current project is to understand whether isotypes (Isotype), the number of vaccine injections (Time_Point), and the different Treg treatments (Drug) cause changes in the antibody characteristics and if the changes are related to the immune responses against HIV virus. Specifically, we evaluate:

Q1: Do drugs and isotypes have effects on the mutation frequency and/or the amino acid count in the third complementarity determining region (CDR3)?

Q2: How does the binding strength of the antibodies develop in response to the number of vaccine dosages by treatment and drug types?

2 Methods

This section first presents some exploratory data analyses and summaries, and then it uses multivariate and longitudinal data analyses to address the research questions.

2.1 Data Summaries

2.1.1 An Overview of Antibodies by Time Points, Drug Types, and Isotypes

A total of 2465 antibodies, from 20 rhesus monkeys, were collected at four different time points (0, 1, 2, 3) and each monkey was given one of three drugs (1 and 2 are immuno-suppressing drugs and 3 is the control). Figure 1 shows the histograms of antibody counts, and Table 1 and Table 2 show the antibody counts in different combinations of drugs, time points, and isotypes.

As shown in Figure 2, the histograms of Isotype, we observed that IgG and IgM occupied the biggest proportion of antibodies in all time points. Before immunization (time point 0), there were similar weight of IgG and IgM found in blood. After the first immunization (time point 1), primary immune response resulted an increase of IgM, followed by an IgG increase at later time points 2 and 3.

2.1.2 Outlier Detection

Our response variables for the multivariate analysis include five variables: Binding, H_CDR3, HMuFreq, L_CDR3, and LMuFreq. As shown in Figure 3, L_CDR3 in one point seems an outlier. As the summary statistics of standardized L_CDR3 shown in Table 3, the maximum value is greater than 30, which is quite unusual. Figure 4 shows the Mahalanobis distances and Z scores of L_CDR3, and the data point again appears to be an apparent outlier. The value for L_CDR3 is quite unlikely. Since we are not able to go back to the original data, we remove the data point and will use the new dataset Data3.

2.1.3 Response Variables

Here we examine our responses: H_CDR3, HMuFreq, L_CDR3, LMuFreq and Binding. We observe that for H_CDR3 the distributions are approximately normal with the center around 13 at different time-points (Figure 5) without taking into account different treatments. Figure 6 represents the distribution of H_CDR3 with respect to treatments at different time-points, and slightly centered around 9 for L_CDR3 at different time points. With L_CDR3, Figure 7 and Figure 8 show approximately normal distribution with a longer right tail. The Q-Q plots in Figure 9 show that H_CDR3 and L_CDR3 are both approximately normal.

HMuFreq and LMuFreq are calculated by dividing H_Substitution by H_VBase for heavy chain and similarly for light chain. These two variables show how much the antibodies mutate. A higher mutation rate is usually indicative of better virus neutralization. Below we present comparison of mutation rate between heavy chain and light chain. Figure 10, Figure 11, Figure 12, and Figure 13 show that HMuFreq and LMuFreq are both approximately normally distributed, each with a longer right tail. The Q-Q plots in Figure 14 confirm the approximate normality of HMuFreq and LMuFreq.

Next, a histogram of Binding with respect to treatments at different time points and Q-Q plot are shown in Figure 15 and Figure 16. We observe that Binding is not normally distributed. However, since our sample size is larger than 2000, we can use the Central Limit Theorem and assume normality. Lastly, we check whether response variables could be correlated, as shown in Figure 17. In these plots, we observe that none of the response variables are highly correlated.

2.2 Multivariate Data Analysis

To answer **Q1** (Do drugs and isotypes have effects on the mutation frequency and/or the amino acid count in the third complementarity determining region (CDR3)?), we test whether predictors Time_Point, Drug, and Isotype had effects on the five responses: H_CDR3, HMuFreq, L_CDR3, LMuFreq, and Binding.

2.2.1 MANOVA

Since there are more than two populations we are comparing, we use MANOVA to test for effects. We check that the normality assumption is met due to large sample size ($n = 2464$). Each antibody is assumed to be independent. However, when we check for equal variance-covariance structures—by Fligner-Killeen Test of Homogeneity of Variances and by checking the variance-covariance matrices of different populations—many of the response variables do not meet the equal variance-covariance matrix assumption. The only populations that meet the assumption is the response variable vector $(L_CDR3, LMuFreq)^T$. Thus we perform a MANOVA test on $(L_CDR3, LMuFreq)^T \sim \text{Time_Point} + \text{Drug} + \text{Isotype}$. The output below shows that the main effects of Drug and Isotype have p-values greater than 0.05, suggesting that these main effects are not significant to the two traits, L_CDR3 and LMuFreq, of antibodies. The p-value for Time_Point is quite small. Thus we conclude that

Time_Point has significant effect to the two traits, L_CDR3 and LMuFreq, of antibodies. The next section, longitudinal data analysis, will examine the effects of time points and reach a similar conclusion.

```
##           Df    Wilks approx F num Df den Df    Pr(>F)
## drug       2 0.99808    1.1768      4  4906    0.3188
## it         4 0.99462    1.6559      8  4906    0.1039
## tp         3 0.98773    5.0610      6  4906 3.47e-05 ***
## Residuals 2454
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

2.2.2 Pairwise Comparison

The pairwise comparison results show that the changes in LMuFreq are significant between time points 0 and 2, 0 and 3, and 1 and 3. [need revisions]

```
## [1] "L_CDR3 pairwise CI's"
## contrast estimate      SE   df lower.CL upper.CL
## 0 - 1      -0.0169 0.0652 2457   -0.222    0.188
## 0 - 2      -0.0287 0.0667 2457   -0.239    0.181
## 0 - 3      -0.0034 0.0764 2457   -0.244    0.237
## 1 - 2      -0.0117 0.0447 2457   -0.153    0.129
## 1 - 3       0.0135 0.0581 2457   -0.169    0.197
## 2 - 3       0.0253 0.0598 2457   -0.163    0.214
##
## Results are averaged over the levels of: rep.meas
## Note: contrasts are still on the [.: scale
## Confidence level used: 0.998333333333333
## [1] "LMuFreq pairwise CI's"
## contrast estimate      SE   df lower.CL upper.CL
## 0 - 1       0.00853 0.00369 2457 -3.09e-03    0.0202
```

```
## 0 - 2      0.01465 0.00378 2457  2.75e-03   0.0265
## 0 - 3      0.01889 0.00432 2457  5.28e-03   0.0325
## 1 - 2      0.00611 0.00253 2457 -1.86e-03   0.0141
## 1 - 3      0.01036 0.00329 2457 -4.12e-06   0.0207
## 2 - 3      0.00424 0.00339 2457 -6.42e-03   0.0149
##
```

Results are averaged over the levels of: rep.meas

Note: contrasts are still on the [.: scale

Confidence level used: 0.9983333333333333

[1] "H_CDR3 pairwise CI's"

```
## contrast estimate      SE   df lower.CL upper.CL
## 0 - 1          0.489 0.231 2457   -0.238    1.217
## 0 - 2          0.194 0.237 2457   -0.550    0.939
## 0 - 3         -0.608 0.271 2457   -1.460    0.244
## 1 - 2         -0.295 0.159 2457   -0.795    0.204
## 1 - 3         -1.097 0.206 2457   -1.746   -0.448
## 2 - 3         -0.802 0.212 2457   -1.469   -0.134
##
```

Results are averaged over the levels of: rep.meas

Note: contrasts are still on the [.: scale

Confidence level used: 0.9983333333333333

[1] "HMuFreq pairwise CI's"

```
## contrast estimate      SE   df lower.CL upper.CL
## 0 - 1          0.00870 0.00288 2457 -0.000379   0.0178
## 0 - 2          0.01872 0.00295 2457  0.009434   0.0280
## 0 - 3          0.02356 0.00338 2457  0.012930   0.0342
## 1 - 2          0.01002 0.00198 2457  0.003796   0.0163
## 1 - 3          0.01486 0.00257 2457  0.006768   0.0230
## 2 - 3          0.00484 0.00265 2457 -0.003491   0.0132
##
```



```
## Results are averaged over the levels of: rep.meas
## Note: contrasts are still on the [.: scale
## Confidence level used: 0.9983333333333333
## [1] "Binding pairwise CI's"
## contrast estimate      SE    df lower.CL upper.CL
## 0 - 1      0.00161 0.0689 2457  -0.2151    0.218
## 0 - 2     -1.06381 0.0704 2457  -1.2855   -0.842
## 0 - 3     -0.76734 0.0806 2457  -1.0211   -0.514
## 1 - 2     -1.06542 0.0472 2457  -1.2141   -0.917
## 1 - 3     -0.76895 0.0614 2457  -0.9622   -0.576
## 2 - 3      0.29647 0.0632 2457   0.0976    0.495
##
## Results are averaged over the levels of: rep.meas
## Note: contrasts are still on the [.: scale
## Confidence level used: 0.9983333333333333
```

2.3 Longitudinal Analysis

To answer our **Q2** (How does the binding strength of the antibodies develop in response to the number of vaccine dosages by treatment and drug types?), we use longitudinal data analysis, including general linear models and linear mixed models. For the longitudinal analysis of binding strength vs number of vaccine doses, we use the `glms` and `lme` functions from the `nlme` package^[7].

2.3.1 One Covariate: Time Point

We first take a look at the data over time. As seen in Figure 18 and Figure 19, the mean trend is not linear, and the different time points have different variances. This information suggests that we should use piecewise linear models and set variances as unequal over time.

We first consider a model with time point as the only covariate:

$$Y_{ij} = \beta_0 + \beta_1 Time_{ij} + e_{ij}$$

We then turn the model above into a piecewise linear model, in which each segment has different intercepts and slopes. We use three indicator variables: $S1, S2, S3$ as the indicator variables, where

$$S1 = \begin{cases} 1 & \text{if } 0 \leq \text{Timepoint} < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$S2 = \begin{cases} 1 & \text{if } 1 \leq \text{Timepoint} < 2 \\ 0 & \text{otherwise} \end{cases}$$

$$S3 = \begin{cases} 1 & \text{if } \text{Timepoint} \geq 2 \\ 0 & \text{otherwise} \end{cases}$$

The new model is thus

$$Y_{ij} = S1(\beta_0 + \beta_1 \text{Time}_{ij}) + S2(\beta_2 + \beta_3 \text{Time}_{ij}) + S3(\beta_4 + \beta_5 \text{Time}_{ij}) + e_{ij}$$

We also want to make sure that the trend is continuous at time points 1 and 2. Our first complete model (fit.gls) is now $Y_{ij} = \beta_0(S1+2S2-S2\text{Time}_{ij}) + \beta_1(S1\text{Time}_{ij}+2S2-S2\text{Time}_{ij}) + \beta_4(-S2 + S2\text{Time}_{ij} + S3) + \beta_5(-2S2 + 2S2\text{Time}_{ij} + S3\text{Time}_{ij}) + e_{ij}$ where

$$\mathbf{e}_i \sim N(0, \sigma^2 I)$$

The model can also be written as

$$Y_{ij} = S1(\beta_0) + S1\text{Time}_{ij}(\beta_1) + S2(2\beta_0 + 2\beta_1 - \beta_4 - 2\beta_5) + S2\text{Time}_{ij}(-\beta_0 - \beta_1 + \beta_4 + 2\beta_5) \\ + S3(\beta_4) + S3\text{Time}_{ij}(\beta_5) + e_{ij}$$

We run the model and plug in coefficients to find the intercepts and slopes for all three segments of the mean trend:

- S1: $-0.2221651 + 0.2432183 * \text{time}$
- S2: $(2 * -0.2221651 + 2 * 0.2432183 - 0.7699600 + 2 * 0.2432756) + (0.2221651 -$

- S3: $0.7699600 - 0.2432756 * time$

We make a plot of the line segments in Figure 20, which shows the two segments S1 and S2 have very similar slopes. So we can refit the model with only two line segments between time points 0 and 2 and between time points 2 and 3. We'll call them S4 and S5. The new model is therefore

$$Y_{ij} = S4(\beta_0 + \beta_1 Time_{ij}) + S5(\beta_2 + \beta_3 Time_{ij}) + e_{ij}$$

$$S4 = \begin{cases} 1 & \text{if Timepoint} < 2 \\ 0 & \text{otherwise} \end{cases}$$

$$S5 = \begin{cases} 1 & \text{if Timepoint} \geq 2 \\ 0 & \text{otherwise} \end{cases}$$

Again, we want to make sure that the trend is continuous at Time_Point = 2. Our second complete model (`fit.gls2`) is then $Y_{ij} = \beta_1(-2S4 + S4Time_{ij}) + \beta_2(S4 + S5) + \beta_3(2S4 + S5Time_{ij}) + e_{ij}$ where

$$\mathbf{e}_i \sim N(0, \sigma^2 I)$$

The model can also be written as $Y_{ij} = S4(-2\beta_1 + \beta_2 + 2\beta_3) + S4Time_{ij}(\beta_1) + S5(\beta_2) + S5Time_{ij}(\beta_3) + e_{ij}$

We first find the mean trend for S4 and S5:

- S4: $(-2*0.5310975+0.5720853+2*-0.0000723)+0.5310975*time = -0.4902543+0.5310975 * time$
- S5: $0.5720853 - 0.0000723 * time$

We can make the plot again to see if the model is reasonable, as shown in Figure 21. Indeed, there is a linear line between Time_Point 0 and 2 and one between Time_Point 2 and 3. The two lines are continuous at Time_Point 2. A comparison of AIC And BIC of these two models, shown in Table 4, indicates that the second model (`fit.gls2`) is indeed a better model.

Next we check whether adding random effects improve our second complete model (`fit.gls2`). We assume that random effects exist in the intercept and slope. Our linear mixed model is then: $Y_{ij} = \beta_1(-2S4 + S4Time_{ij}) + \beta_2(S4 + S5) + \beta_3(2S4 + S5Time_{ij}) + b_{0i} + b1iTime_{ij} + e_{ij}$ where

$$\mathbf{b}_i \sim N\left(0, \mathbf{D} = \begin{bmatrix} D_{11} & D_{12} \\ & D_{22} \end{bmatrix}\right)$$

and

$$\mathbf{e}_i \sim N(0, \sigma^2 I)$$

We fit two models with random effects: `fit.a1` assumes random intercept and slope for time point, compound symmetric correlation structure, and unequal variances over time; and `fit.a2` assumes random intercept and slope for time point, AR1 correlation structure, and unequal variances over time. As shown in Table 5, the model `fit.a2` has the lowest AIC And BIC, so it seems the best model. We now check residuals for three models: `fit.gls2`, `fit.a1`, `fit.a2`, as shown in Figure 22. All three Q-Q plots show approximate normality. To further investigate the effects of drugs, We now use `fit.a2` for further analysis.

Now we would like to know if the slopes between Time_Point 0 and 2 and between Time_Point 2 and 3 equal zero.

$$H_0 : \text{slope of } S4 = 0 \text{ and slope of } S5 = 0$$

which means

$$H_0 : \beta_1 = 0 \text{ and } \beta_3 = 0$$

Thus, we can check for two tests:

$$\mathbf{L}_1 \beta = 0$$

where $\mathbf{L}_1 = (1, 0, 0)$ and $\beta = (\beta_1, \beta_2, \beta_3)^T$ and

$$\mathbf{L}_2 \beta = 0$$

where $\mathbf{L}_2 = (0, 0, 1)$ and $\beta = (\beta_1, \beta_2, \beta_3)^T$

As shown in Table 6, the slop of S4 has a very small p-value, while the slope of S5 is quite large,

indicating that the change in Binding rate between time point 0 and time point 2 is significant while the change between time point 2 and time point 3 is not significant. We conclude that time point 2, when the monkeys had received two vaccines, had the highest Binding rate, while the last vaccine shot at time point 3 did not make a difference to the Binding rate.

2.3.2 Two Covariates: Time Point and Drug

Next we add Drug as a covariate to the model `gls.a2` to see if it has effects on Binding. We use two indicator variables: `D2` and `D3`, where

$$D2 = \begin{cases} 1 & \text{if Drug} = 2 \\ 0 & \text{otherwise} \end{cases}$$

$$D3 = \begin{cases} 1 & \text{if Drug} = 3 \\ 0 & \text{otherwise} \end{cases}$$

Adding on to the model `gls.a2` and assuming that the random effects are the same for each drug, our model (`fit.a3`) with the extra covariate Drug is:

$$Y_{ij} = \beta_1(-2S4 + S4Time_{ij}) + \beta_2(S4 + S5) + \beta_3(2S4 + S5Time_{ij}) + \\ \beta_4D2(-2S4 + S4Time_{ij}) + \beta_5D2(S4 + S5) + \beta_6D2(2S4 + S5Time_{ij}) + \\ \beta_7D3(-2S4 + S4Time_{ij}) + \beta_8D3(S4 + S5) + \beta_9D3(2S4 + S5Time_{ij}) + b_{0i} + b_{1i}Time_{ij} + e_{ij}$$

where

$$\mathbf{b}_i \sim N\left(0, \mathbf{D} = \begin{bmatrix} D_{11} & D_{12} \\ & D_{22} \end{bmatrix}\right)$$

and

$$\mathbf{e}_i \sim N(0, \sigma^2 I)$$

With the model constructed, we want to make inference on β to find whether the drugs have any effects. To see whether Drug 1 and Drug 2 have any difference, we want to perform a hypothesis

test on $H_0 : \beta_4 = \beta_5 = \beta_6 = 0$, thus we can do the test

$$\mathbf{L}_3\beta = 0$$

where

$$\mathbf{L}_3 = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \end{bmatrix}$$

and $\beta = (\beta_1, \beta_2, \beta_3, \beta_4, \beta_5, \beta_6, \beta_7, \beta_8, \beta_9)^T$

To see whether Drug 1 and Drug 3 have any difference, we want to perform a hypothesis test on $H_0 : \beta_7 = \beta_8 = \beta_9 = 0$, thus we can do the test

$$\mathbf{L}_4\beta = 0$$

where

$$\mathbf{L}_4 = \begin{bmatrix} 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix}$$

and $\beta = (\beta_1, \beta_2, \beta_3, \beta_4, \beta_5, \beta_6, \beta_7, \beta_8, \beta_9)^T$

To see whether Drug 2 and Drug 3 have any difference, we want to perform a hypothesis test on $H_0 : \beta_4 = \beta_7, \beta_5 = \beta_8, \beta_6 = \beta_9$, thus we can do the test

$$\mathbf{L}_5\beta = 0$$

where

$$\mathbf{L}_5 = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 & -1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & -1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & -1 \end{bmatrix}$$

and $\beta = (\beta_1, \beta_2, \beta_3, \beta_4, \beta_5, \beta_6, \beta_7, \beta_8, \beta_9)^T$

We found that, as shown in Table 7, Drug 1 and Drug 2 do not have significantly different effects on Binding rates. As shown in Table 8, Drug 1 and Drug 3 do not have significantly different

effects on Binding rates. Also, as shown in Table 9, Drug 2 and Drug 3 do not have significantly different effects on Binding rates. In other words, drug groups do not have significant effects on our longitudinal model. Thus we will retain `fit.a2` as our best model.

3 Results

For multivariate analyses, we performed a MANOVA test on the main effects of Isotype and Drug on the response variable vector $(H_CDR3, HMuFreq, L_CDR3, LMuFreq, Binding)^T$. We found that both Isotype and Drug have very small p-values. We also performed pairwise comparison to see where the effects were. [need revisions]

For longitudinal analyses, we found that the best linear mixed model is $Y_{ij} = \beta_1(-2S4 + S4Time_{ij}) + \beta_2(S4 + S5) + \beta_3(2S4 + S5Time_{ij}) + b_{0i} + b1iTime_{ij} + e_{ij}$

where

$$\mathbf{b}_i \sim N\left(0, \mathbf{D} = \begin{bmatrix} D_{11} & D_{12} \\ & D_{22} \end{bmatrix}\right)$$

and

$$\mathbf{e}_i \sim N(0, \sigma^2 I)$$

We performed F-tests to make inferences on β for the two line segments, We rejected the hypothesis that the slope of the first line segment (between time points 0 and 2) is zero, but we failed to reject the hypothesis that the slope of the second line segment (between time points 2 and 3) is zero. We then added Drug as another covariate to the above model and made inference on β . The comparison between three drug groups was done with three F-tests (between Drug groups 1 and 2, 2 and 3, and 1 and 3), which led to three p-values greater than 0.05. Thus We failed to reject the null hypotheses that the mean trend of either of the two drug groups was equal.

4 Discussion

4.1 Implications

Our findings show that the number of vaccine injections do contribute to higher binding rates, although we do not determine whether the increase can be translated into immunity against HIV. The lack of evidence for effects of Treg medications, on the other hand, suggest that the theory of using immuno-suppressing drugs to enhance the efficacy of HIV vaccines has not been proven.

4.2 Limitations

While our analyses reached some findings, some further investigations could improve our analyses. A common method to analyze antibody traits is to treat antibodies (rows in our data) as independent from each other. We considered perform statistical analyses to prove or disprove such a convention, but it requires more biological knowledge and the investigation would most likely be beyond the scope of a final report. However, this remains an interesting topic that could be explored.

[about MANOVA]

For longitudinal data analyses, we did not try out more combinations for models. For example, we only tried two correlation structures (compound symmetry and AR1); other structures might achieve better results. When we added drug as another covariate, we did not go back to test which correlation structure might perform better or whether the piecewise model should include two or three line segments.

5 Conclusions

[need revisions] In this report, we performed multivariate data analyses and longitudinal data analyses to understand whether time points, drugs, and isotypes have effects on characteristics of antibodies. Our statistical analyses provide answers to our two research questions. We performed a MANOVA test to our first research question, “Do drugs and isotypes have effects on the mutation frequency and/or the amino acid count in the third complementarity determining region (CDR3)?” and found significant main effects for both drugs and isotypes. However, pairwise comparisons

reveal that drugs increased mutation rates for heavy chain but did not increase binding rates, which are the key to better efficacy of vaccines. [Kan: how to interpret IgG’s higher mutation rates and binding rates?]

To answer our second research question, How does the binding strength of the antibodies develop in response to the number of vaccine dosages by treatment and drug types?”, we first used time point as the only covariate and constructed with two general linear models with two and three line segments. We then added random effects for intercept and slope for time point as well as different correlation structure—compound symmetry and AR1—and found the model with two line segments (time points 0 to 2 and time points 2 to 3), random effects of intercept and slope of time point, AR1 correlation structure, and unequal variances over time performs best. F-tests for inferences for β reveals that time point 2 have the highest binding rates, suggesting that two vaccine injections improved the binding rates while the third injection did not make a difference. We also found that adding drug as a covariate did not improve the model.

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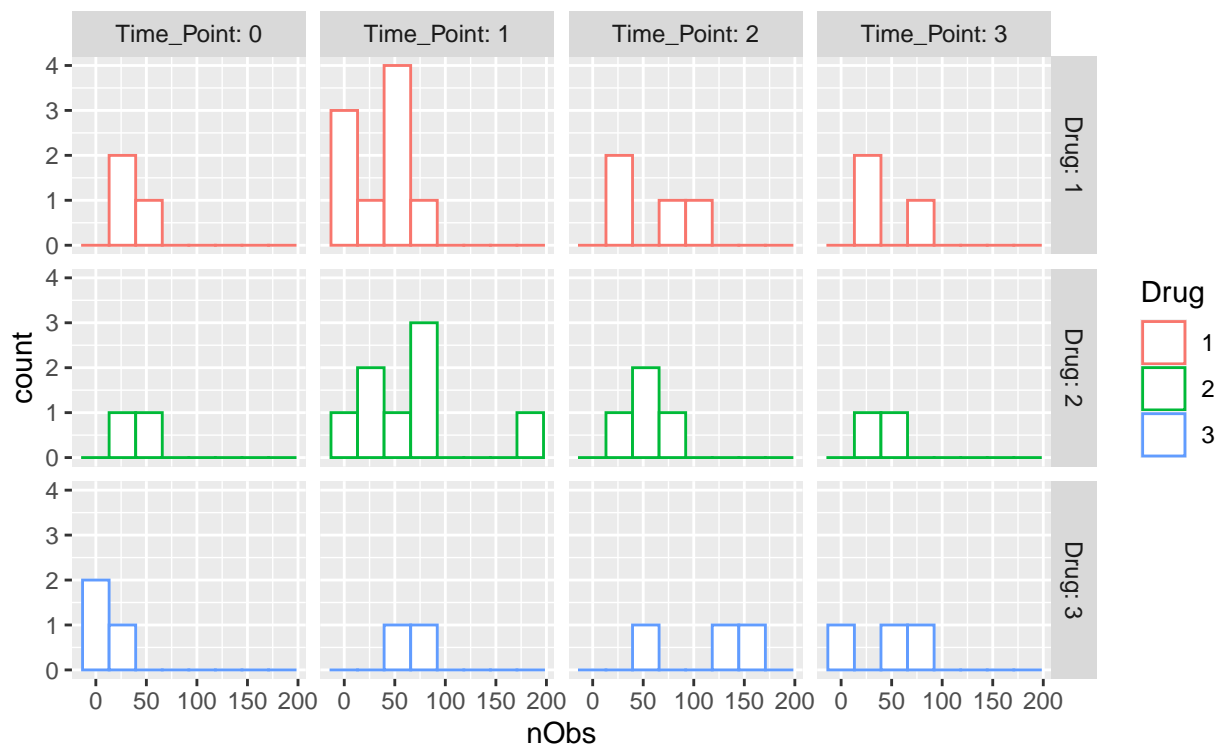


Figure 1: Histograms of Antibodies

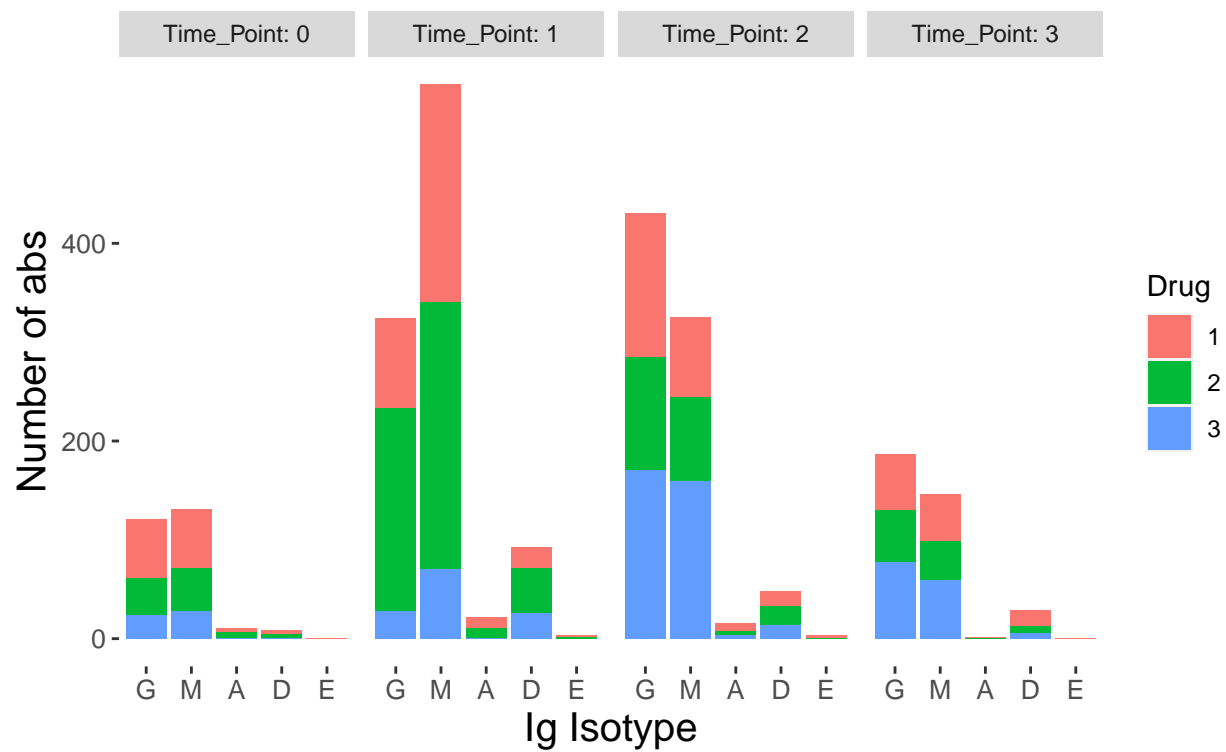


Figure 2: Histograms of Isotypes

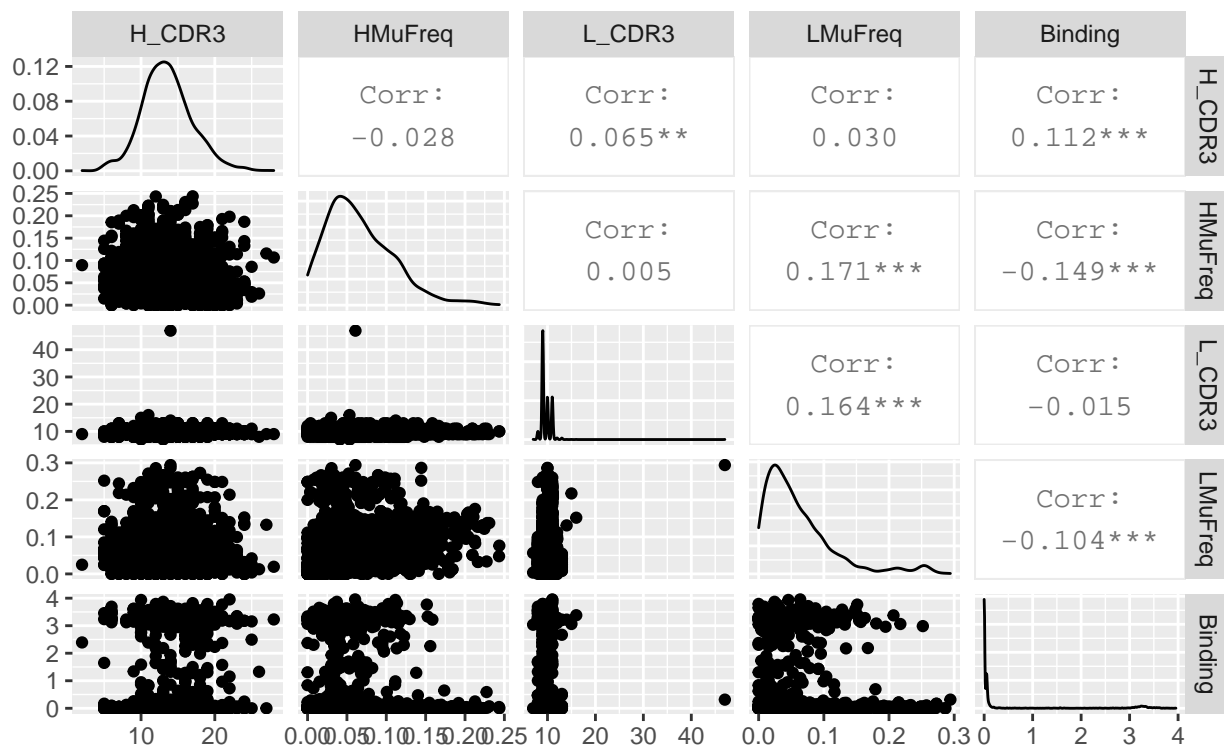


Figure 3: Histogram of Response Variables

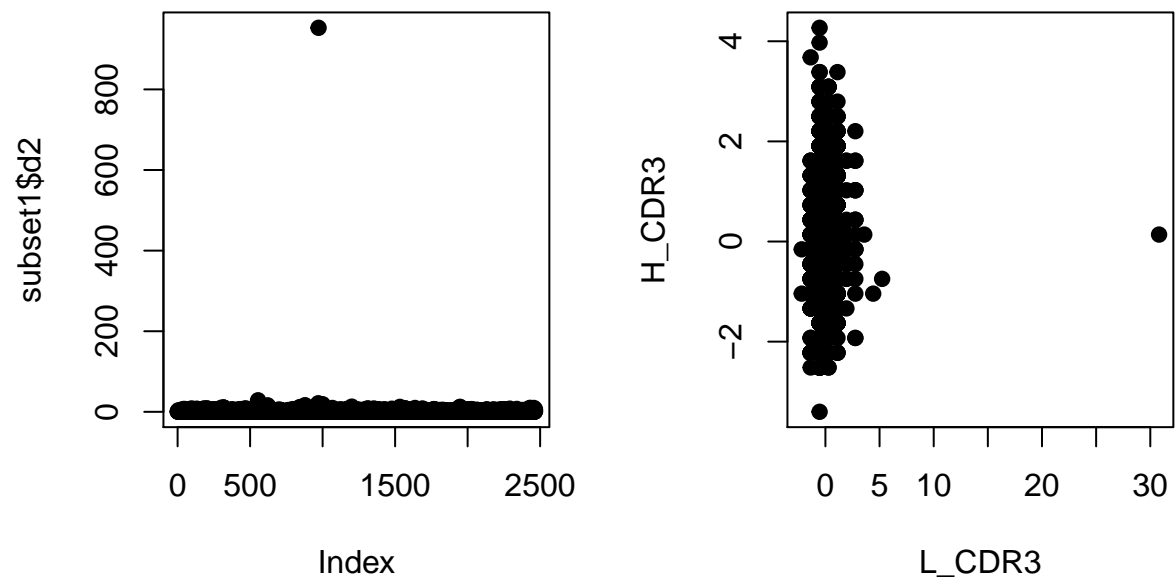


Figure 4: Mahalanobis distances and Z scores

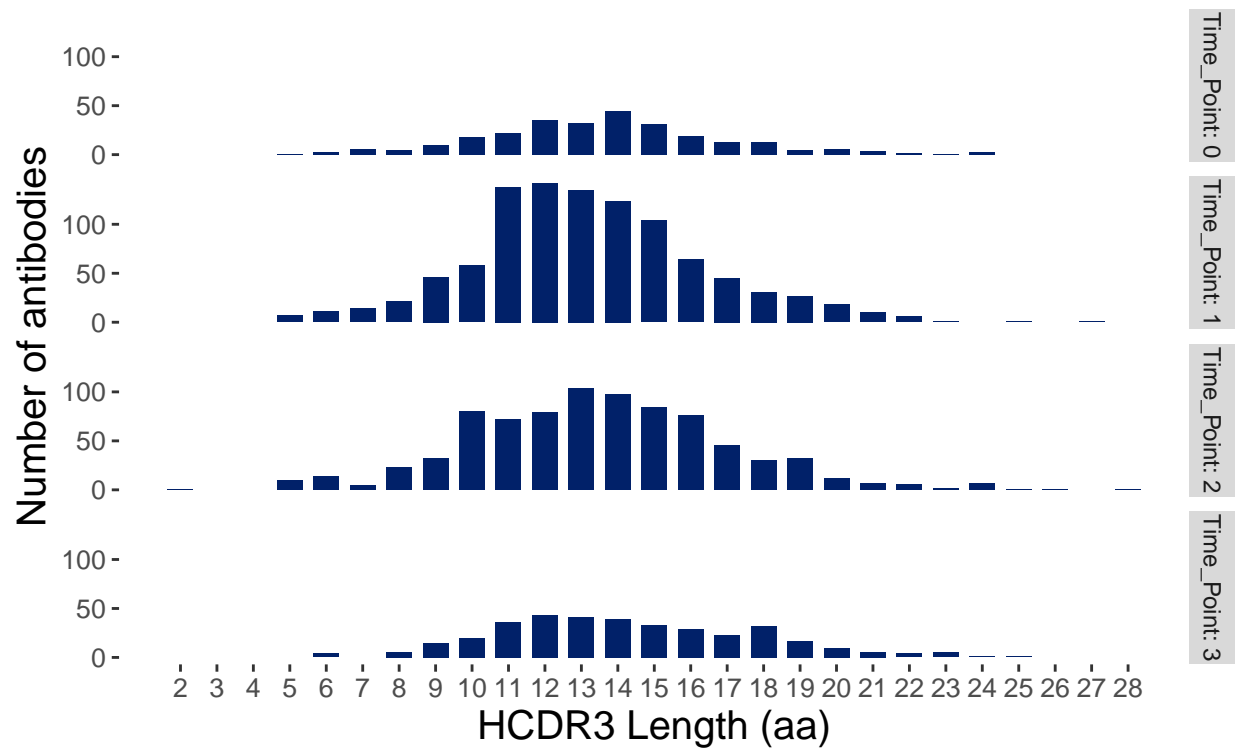


Figure 5: Histogram H_CDR3

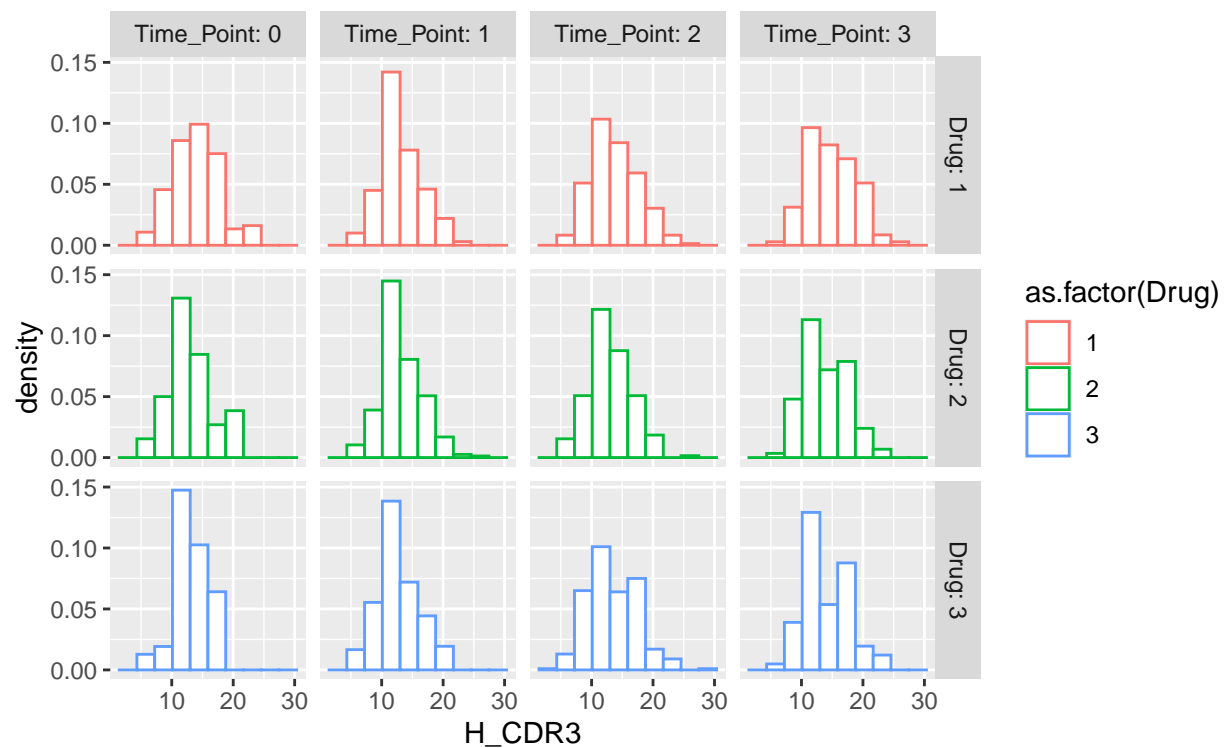


Figure 6: Histograms of H_CDR3 vs Treatment and Timepoint

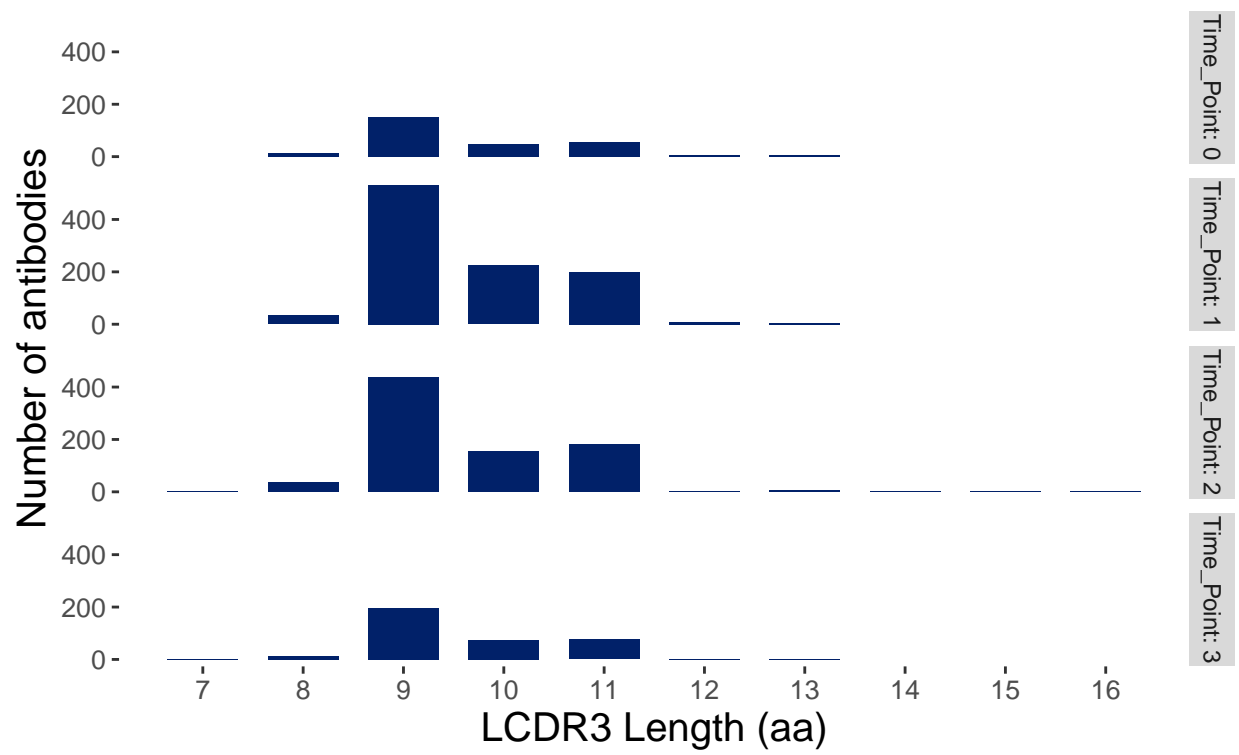


Figure 7: Histogram L_CDR3

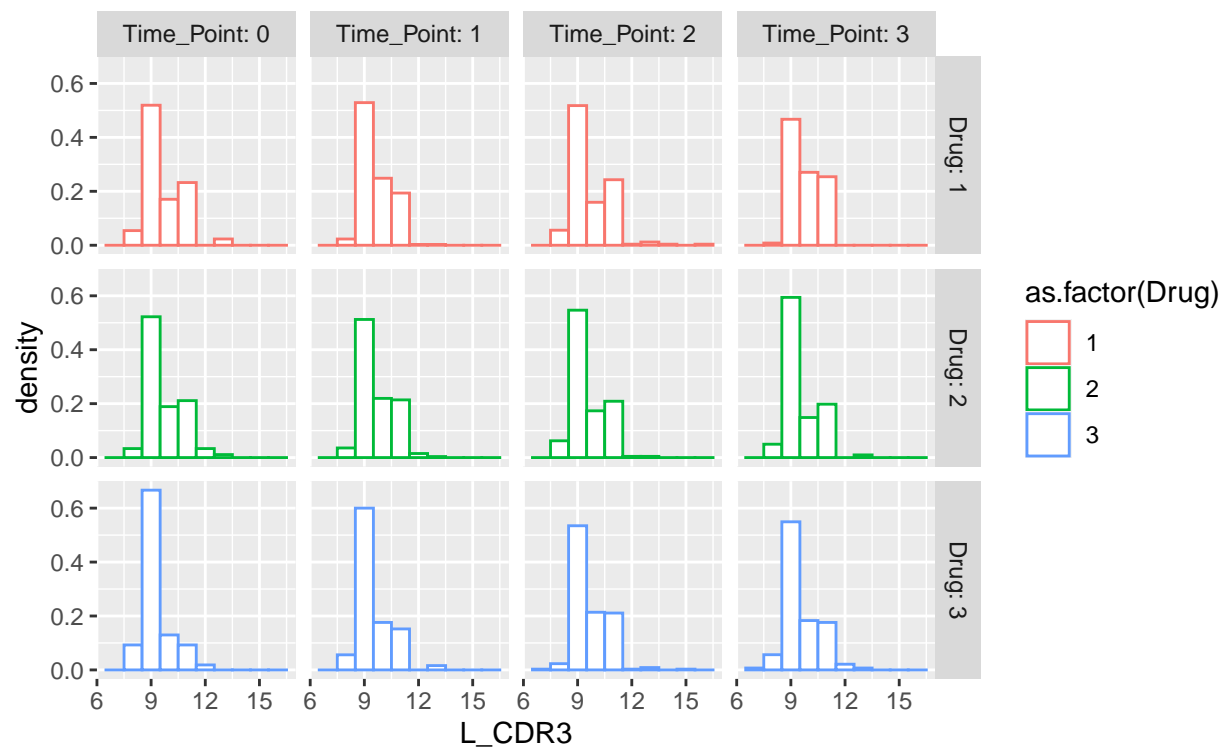


Figure 8: Histograms of L_CDR3 vs Treatment and Timepoint

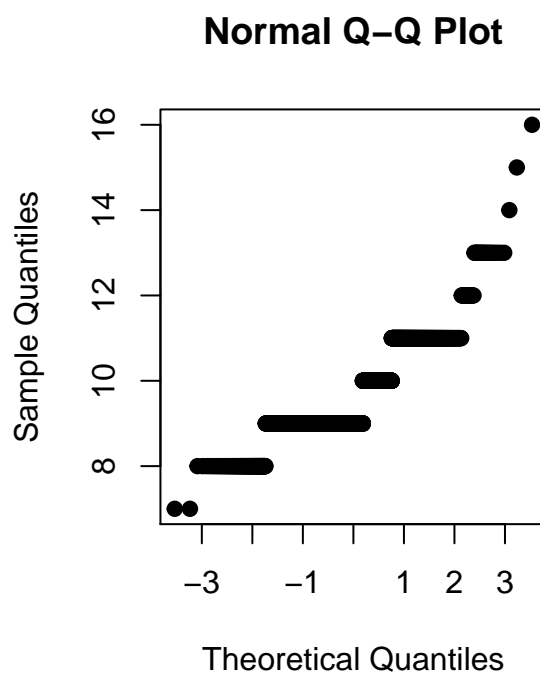
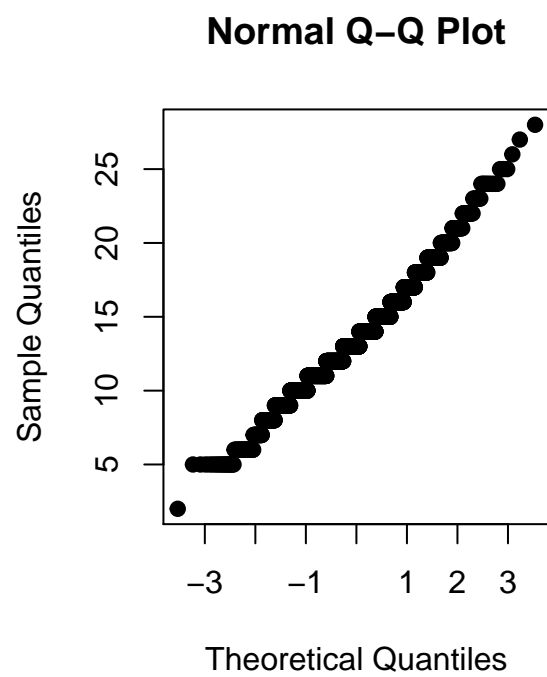


Figure 9: Q-Q Plots of H_CDR3 and L_CDR3

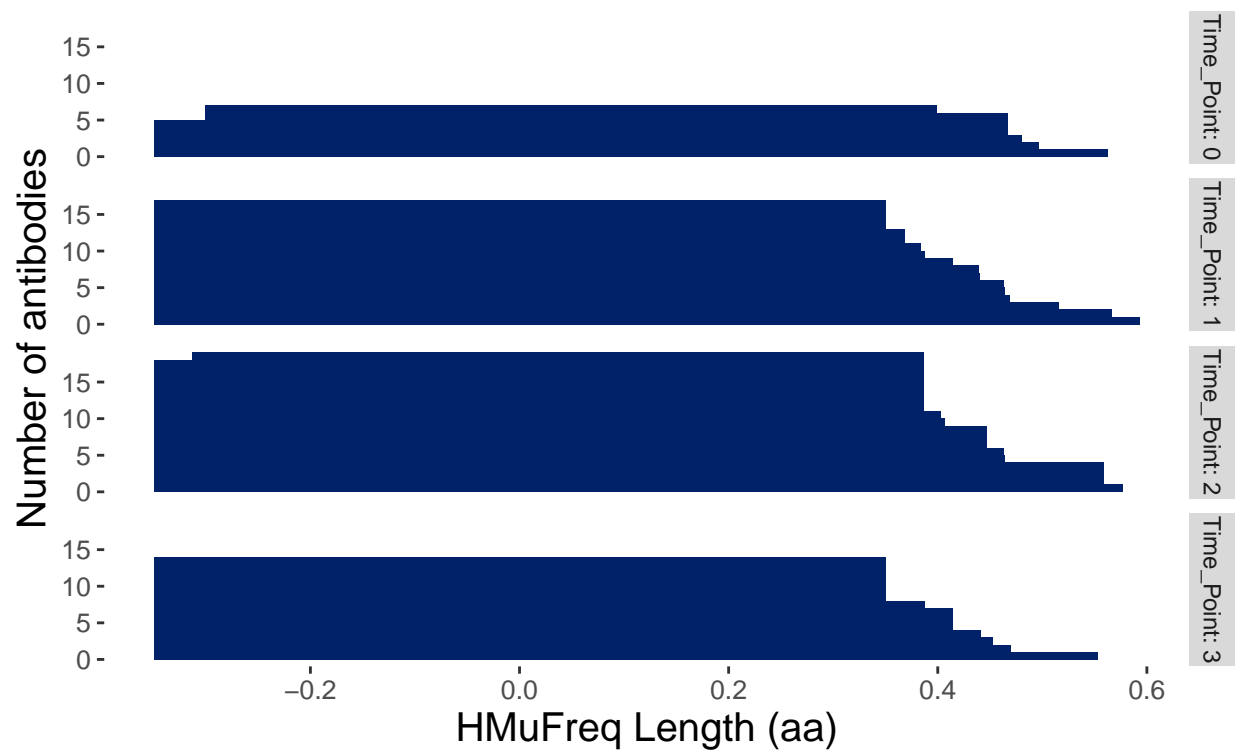


Figure 10: Histogram HMuFreq

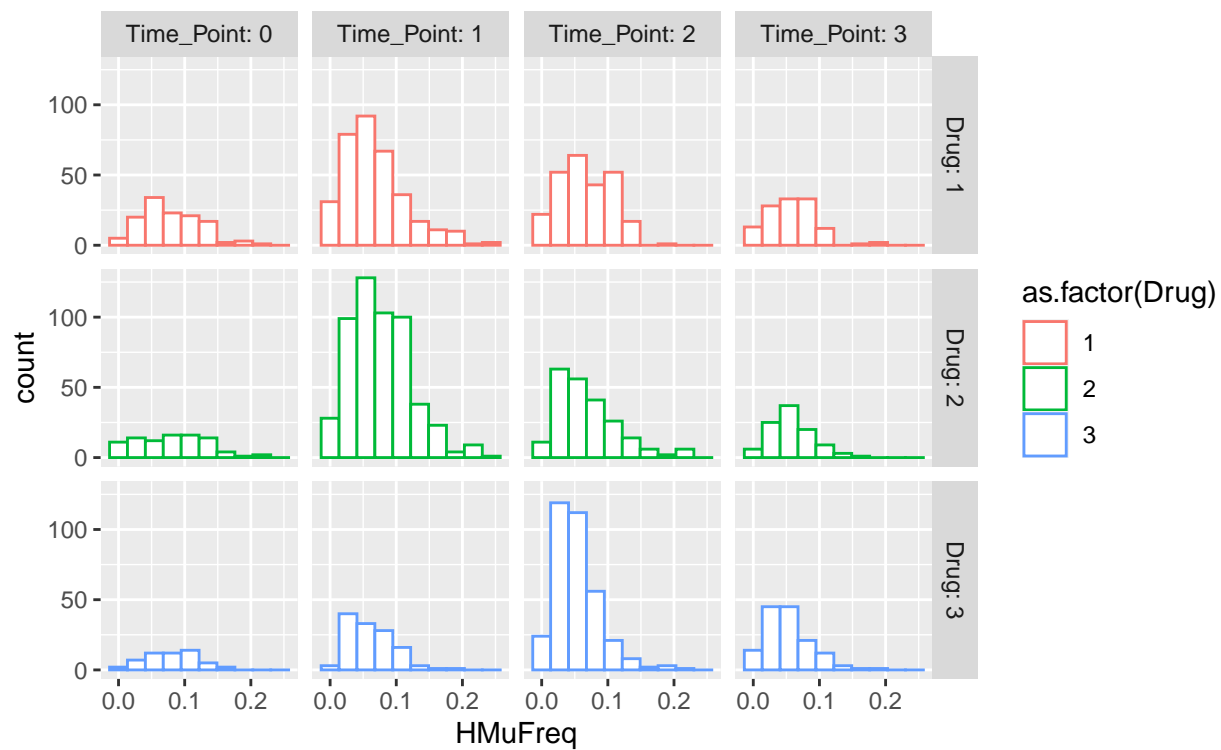


Figure 11: Histograms of HMuFreq vs Treatment and Timepoint

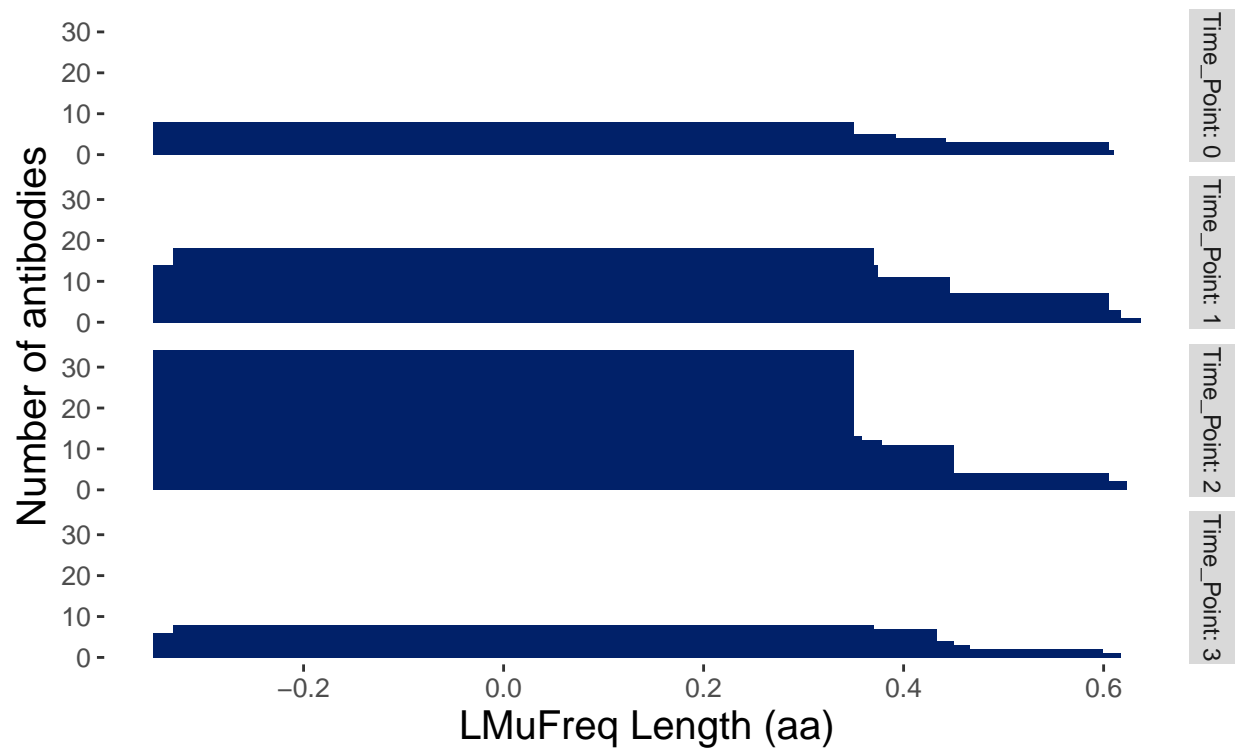


Figure 12: Histogram LMuFreq

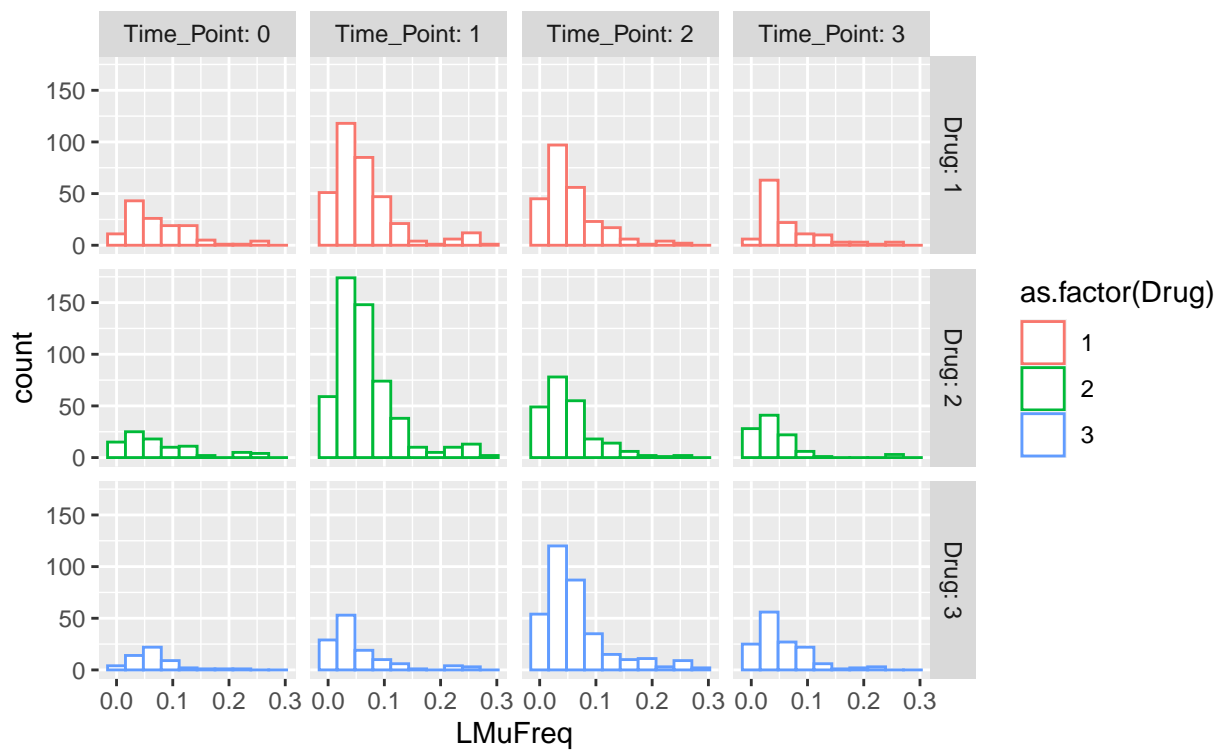


Figure 13: Histograms of LMUFreq vs Treatment and Timepoint

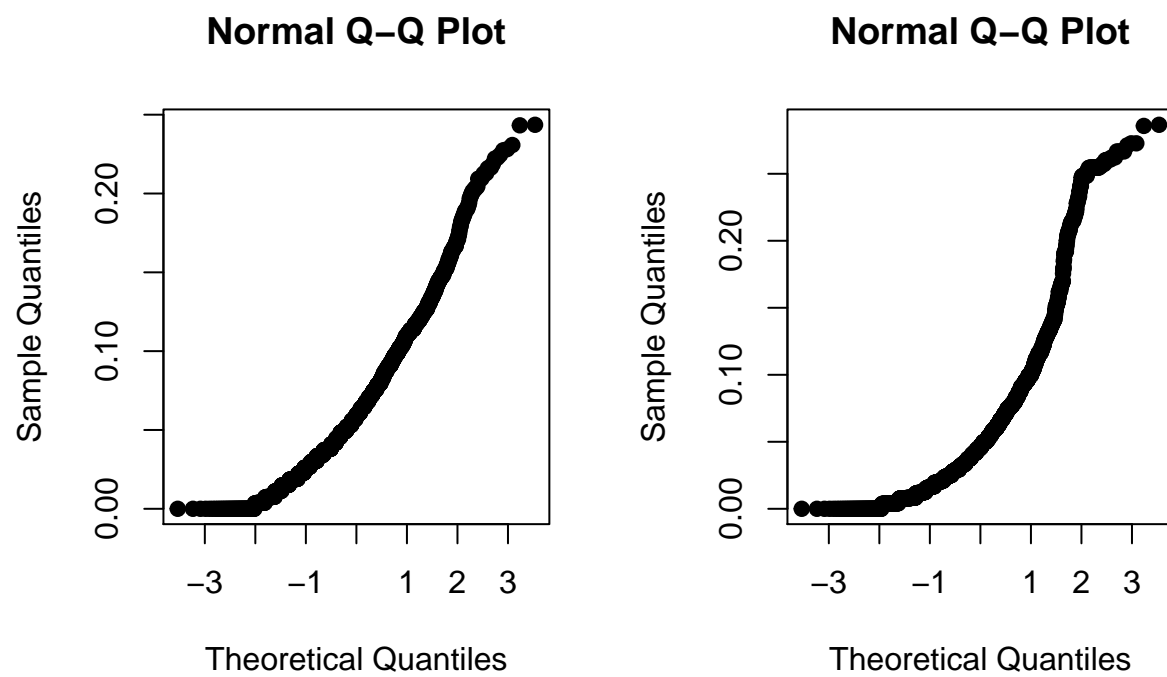


Figure 14: Q-Q Plot of HMuFreq and LMuFreq

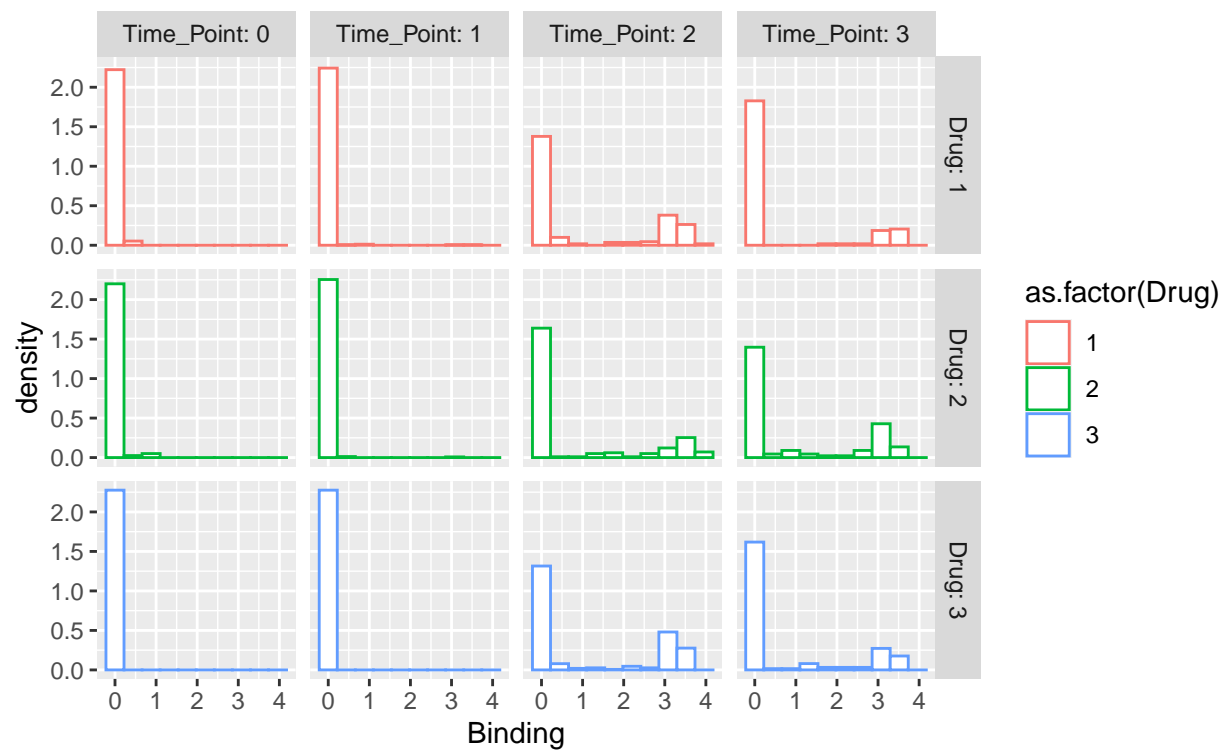


Figure 15: Histograms of Binding Strength vs Treatment and Timepoint

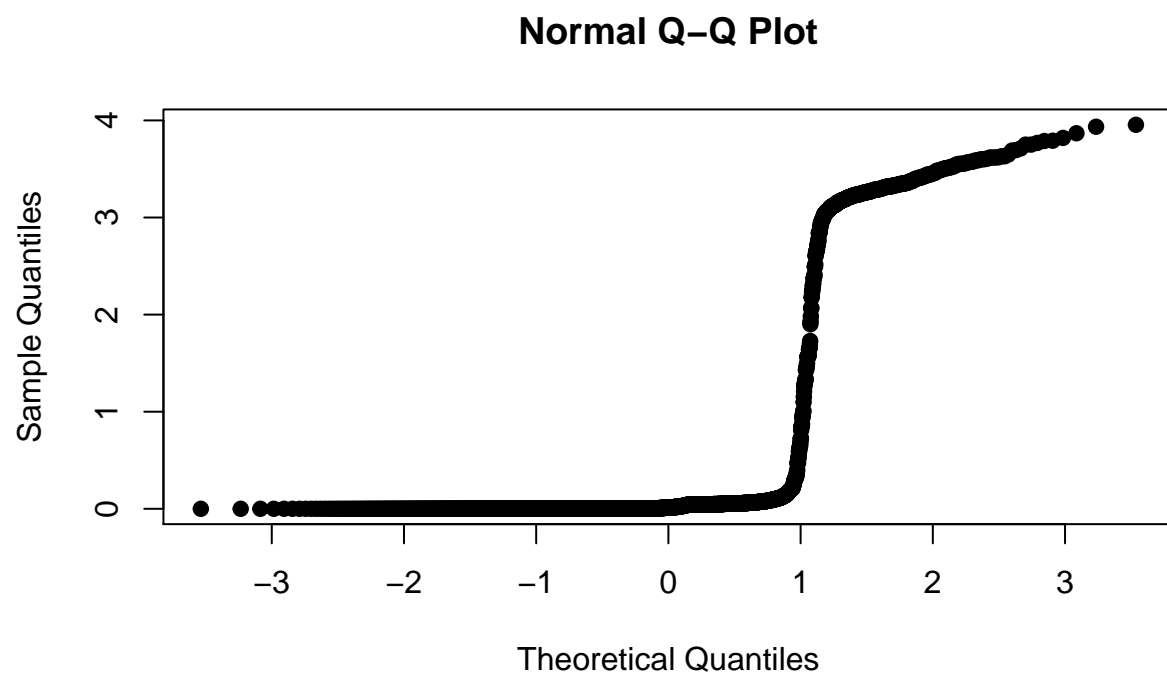


Figure 16: Q-Q Plot of Binding

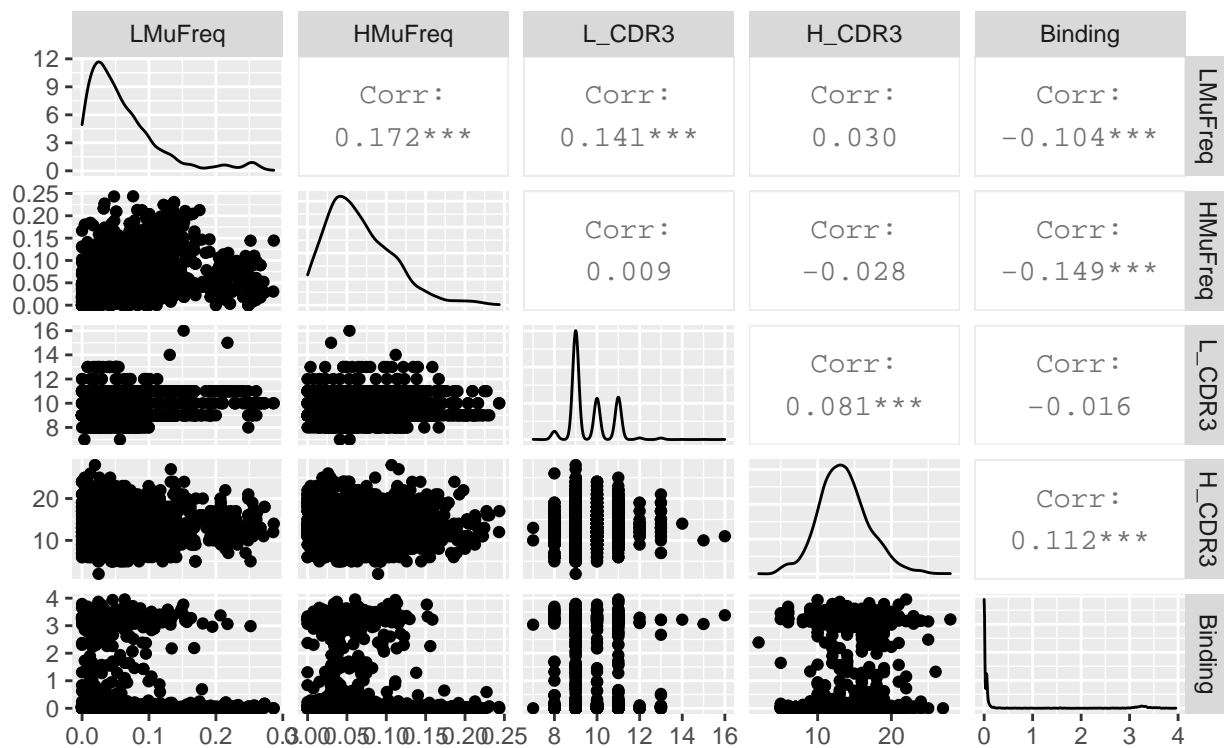


Figure 17: Plots of response variables

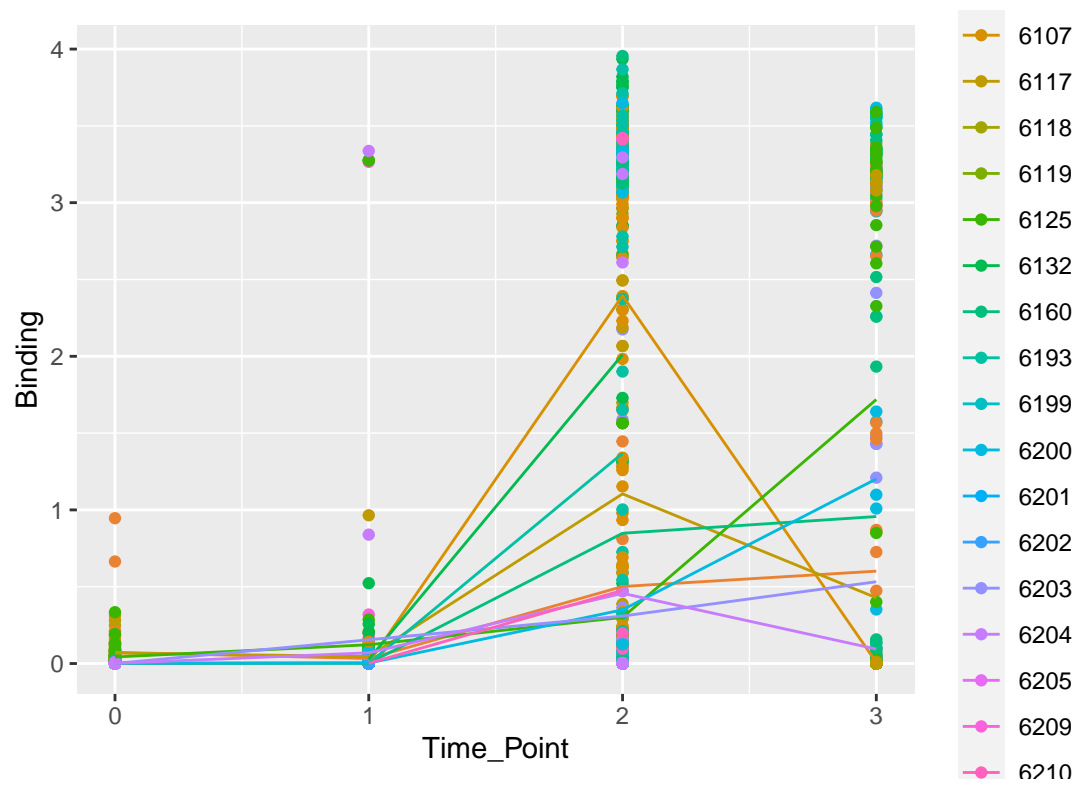


Figure 18: Mean trend by monkey

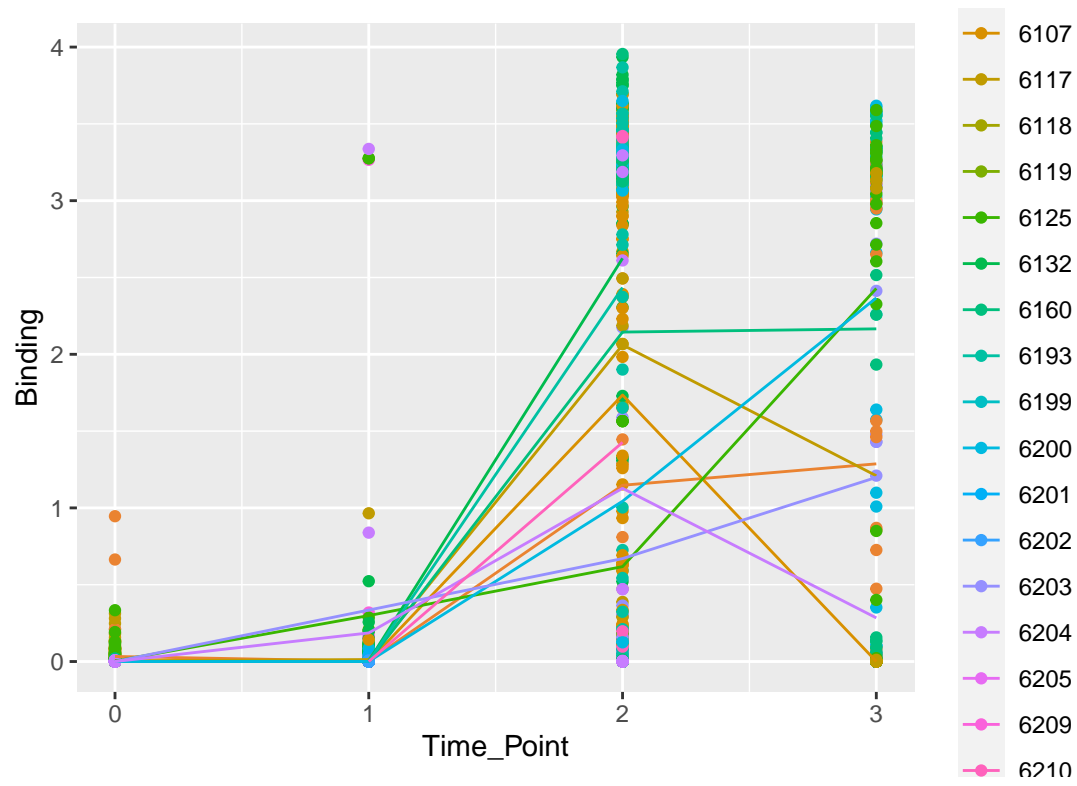


Figure 19: Variances over time by monkey

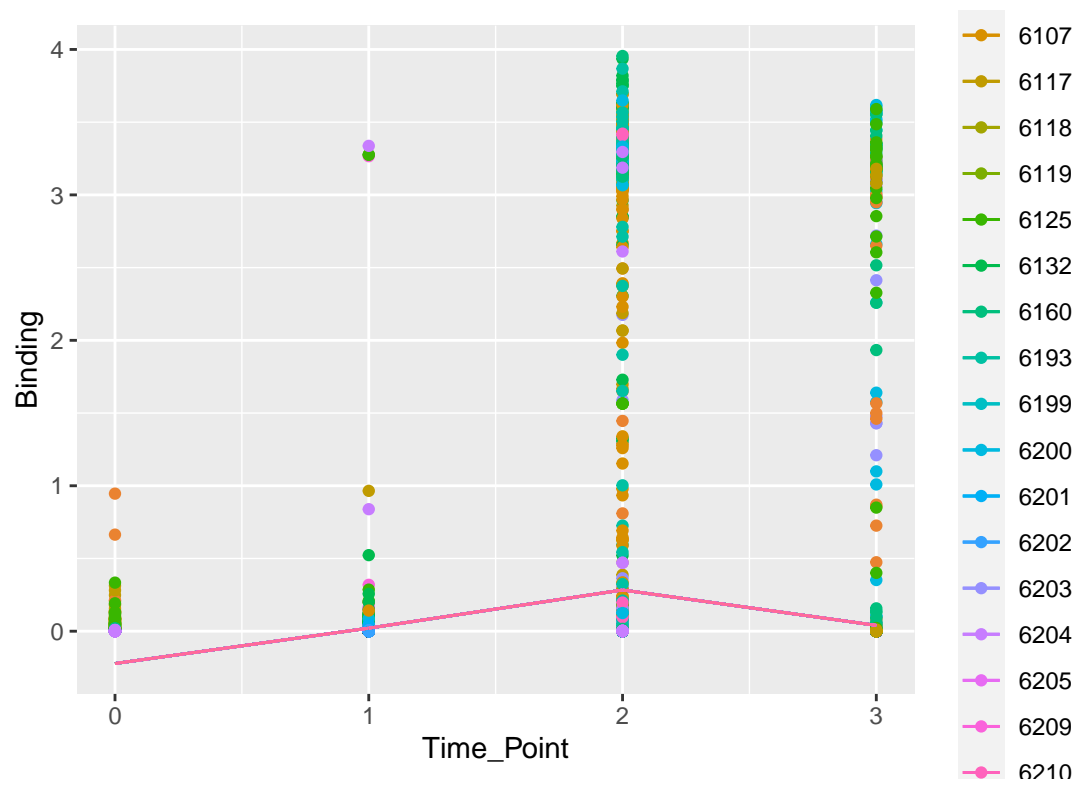


Figure 20: Piecewise Linear Function—three segments

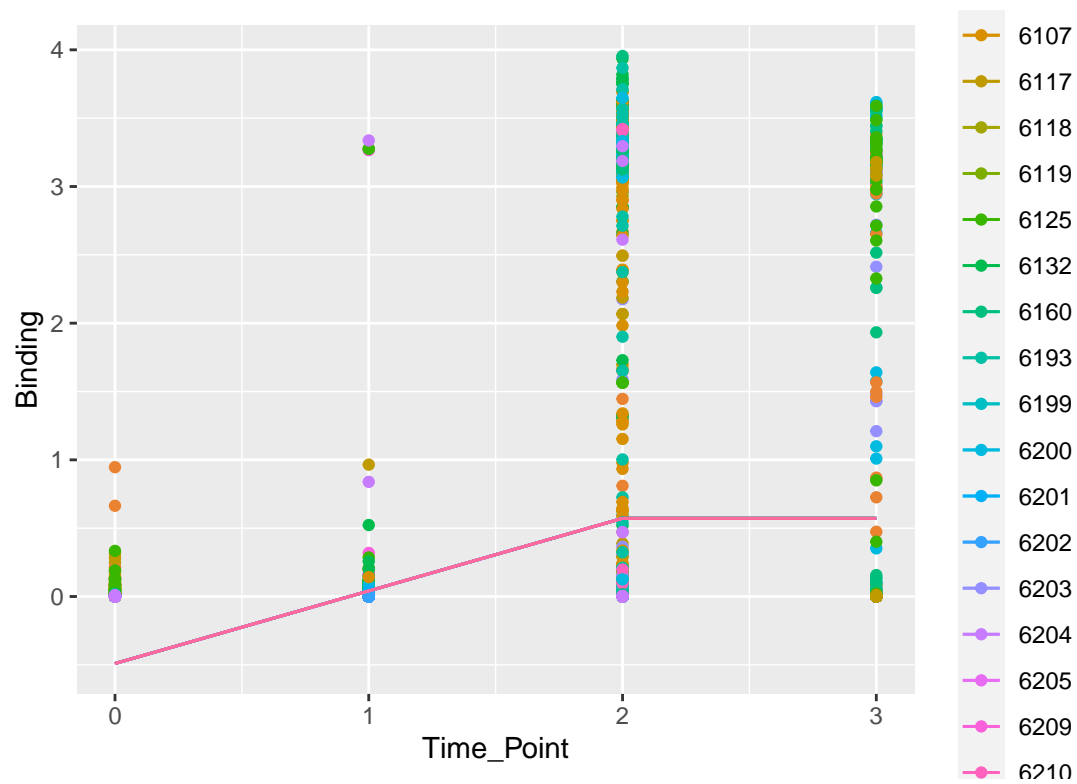


Figure 21: Piecewise Linear Function—two segments

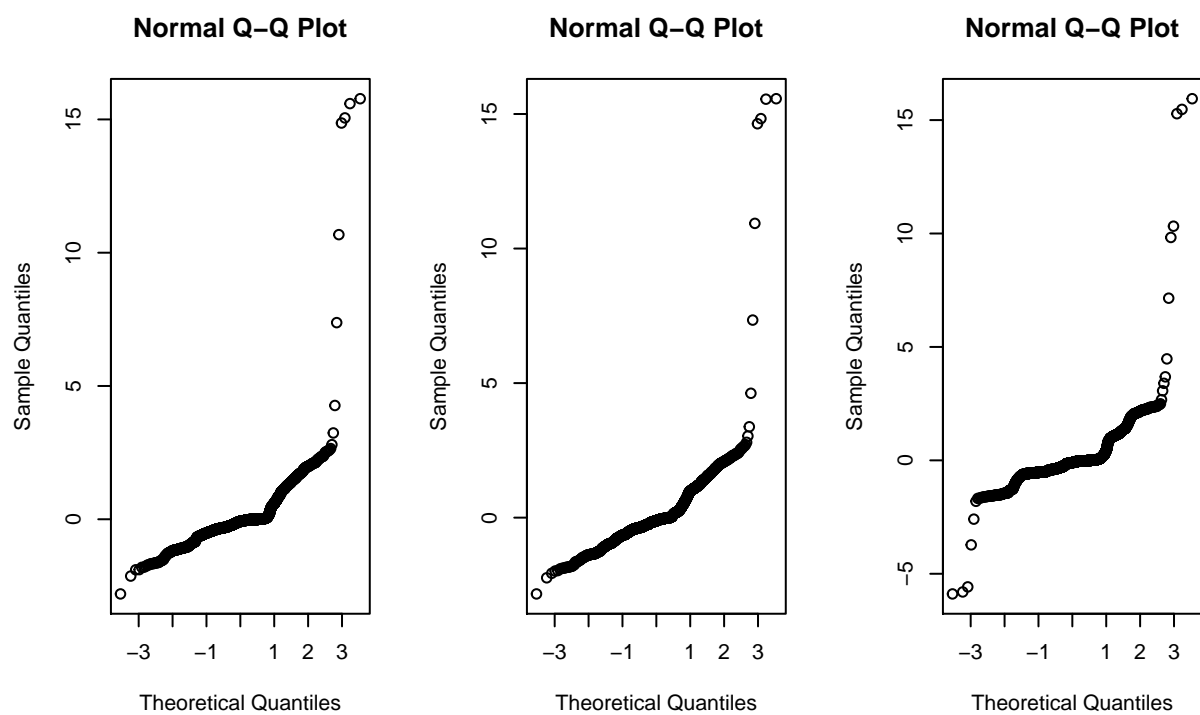


Figure 22: Q-Q plots of models: GLS, compound symmetry, AR1

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Table 1: Frequency tables of drug vs. timepoints

	0	1	2	3
1	129	346	251	122
2	90	533	225	101
3	54	125	347	142

Table 2: Frequency tables of timepoints vs. isotypes for drug = 1 (left), 2 (middle), 3 (right)

	A	D	E	G	M		A	D	E	G	M		A	D	E	G	M
0	4	4	1	60	60	0	6	4	0	37	43	0	1	1	0	24	28
1	11	22	2	91	220	1	10	45	2	205	271	1	1	26	0	28	70
2	8	15	3	145	80	2	4	19	1	115	86	2	4	14	0	170	159
3	1	16	1	57	47	3	1	7	0	53	40	3	0	6	0	77	59

Table 3: Summaries of standardized LCDR3

V1
Min. :-2.1860
1st Qu.: -0.5361
Median :-0.5361
Mean : 0.0000
3rd Qu.: 0.2888
Max. :30.8110

Table 4: AIC and BIC between two gls models

	df	AIC	df.1	BIC
fit.gls	9	3323.050	9	3375.322
fit.gls2	8	3315.264	8	3361.730

Table 5: AIC and BIC for three models

	df	AIC	df.1	BIC
fit.gls2	8	3315.264	8	3361.730
fit.a1	11	3234.628	11	3298.520
fit.a2	11	3063.290	11	3127.182

Table 6: Inference of S4 ad S5 slopes

numDF	denDF	F.value	p.value
1	2442	244.2324506	0.0000000
1	2442	0.0317192	0.8586602

Table 7: Test whether drug 1 = drug 2

Fstat	p_value
1.065151	0.3626666

Table 8: Test whether drug 1 = drug 3

Fstat	p_value
1.231968	0.2964737

Table 9: Test whether drug 2 = drug 3

Fstat	p_value
1.255448	0.288075