Multivariate Statistical Analysis

Lecture 16

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Factor Analysis

Let the observable vector $\mathbf{y} \in \mathbb{R}^p$ be written as

$$\mathbf{y} = \mathbf{W}\mathbf{x} + \boldsymbol{\mu} + \boldsymbol{\epsilon},$$

where

- **1** $\mathbf{W} \in \mathbb{R}^{p \times q}$ is the loading matrix (parameter),
- $\mathbf{2} \mathbf{x} \in \mathbb{R}^q$ is the common factor (parameter/random vector),
- $oldsymbol{0} oldsymbol{\mu} \in \mathbb{R}^p$ is the mean vector (parameter),
- $\bullet \epsilon \in \mathbb{R}^p$ is the specific factor (random vector).

The model is similar to regression, but \mathbf{x} is unobserved.

Factor Analysis

Example of sports games:

$$y = Wx + \mu + \epsilon$$
.

- **9** y: performance in real-world
- **W**: system of the game
- 3 x: attributes in the game
- \bullet μ : average attributes
- \bullet : noise/exception









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Probabilistic Principle Component Analysis

Let $\mathbf{y}_1, \dots, \mathbf{y}_N \in \mathbb{R}^p$ be N independent observations and we have

$$\mathbf{y}_{\alpha} = \mathbf{W}\mathbf{x}_{\alpha} + \boldsymbol{\mu} + \epsilon_{\alpha},$$

where

$$\mathbf{x}_{lpha} \sim \mathcal{N}_{q}(\mathbf{0}, \mathbf{I})$$
 and $\epsilon_{lpha} \sim \mathcal{N}_{p}(\mathbf{0}, \sigma^{2}\mathbf{I})$

are independent for some $\sigma^2 > 0$ and $q < \min\{N, p\}$.

We target to estimate parameters

$$\mathbf{W} \in \mathbb{R}^{p \times q}, \quad \boldsymbol{\mu} \in \mathbb{R}^p \quad \text{and} \quad \sigma \in (0, +\infty)$$

by maximum likelihood estimation for given $\mathbf{y}_1, \dots, \mathbf{y}_N$.

Probabilistic Principle Component Analysis

Consider that

$$\mathbf{y}_{lpha} \sim \mathcal{N}_{p}(\boldsymbol{\mu}, \mathbf{W} \mathbf{W}^{\top} + \sigma^{2} \mathbf{I}).$$

We construct the likelihood function

$$\begin{split} & L(\boldsymbol{\mu}, \mathbf{W}, \sigma^2) \\ &= \prod_{\alpha=1}^N \frac{1}{\sqrt{(2\pi)^p \det(\boldsymbol{\Sigma})}} \exp\left(-\frac{1}{2}(\mathbf{y}_\alpha - \boldsymbol{\mu})^\top (\mathbf{W} \mathbf{W}^\top + \sigma^2 \mathbf{I})^{-1} (\mathbf{y}_\alpha - \boldsymbol{\mu})\right), \end{split}$$

then we have

$$\ln L(\mu, \mathbf{W}, \sigma^2)$$

$$\propto -\frac{N}{2} \ln \det(\mathbf{W} \mathbf{W}^\top + \sigma^2 \mathbf{I}) - \frac{1}{2} \sum_{\alpha=1}^N (\mathbf{y}_{\alpha} - \mu)^\top (\mathbf{W} \mathbf{W}^\top + \sigma^2 \mathbf{I})^{-1} (\mathbf{y}_{\alpha} - \mu).$$

The Maximum Likelihood Estimators

The maximum likelihood estimators of μ , **W** and σ^2 are

$$\hat{\boldsymbol{\mu}} = \overline{\mathbf{y}} = \frac{1}{N} \sum_{\alpha=1}^{N} \mathbf{y}_{\alpha}, \quad \hat{\mathbf{W}} = \mathbf{U}_{q} (\mathbf{\Lambda}_{q} - \hat{\sigma}^{2} \mathbf{I}) \mathbf{R} \quad \text{and} \quad \hat{\sigma}^{2} = \frac{1}{d-q} \sum_{j=q+1}^{d} \lambda_{j},$$

where

 $oldsymbol{0} oldsymbol{U}_q \in \mathbb{R}^{p imes q}$ is orthogonal column consisting of principal eigenvectors of

$$\hat{oldsymbol{\Sigma}} = rac{1}{N} \sum_{lpha=1}^N (\mathbf{y}_lpha - ar{\mathbf{y}}) (\mathbf{y}_lpha - ar{\mathbf{y}})^ op,$$

- ② $\mathbf{\Lambda}_q \in \mathbb{R}^{q imes q}$ is diagonal with corresponding eigenvalues $\lambda_1, \dots, \lambda_q$,
- **3** $\mathbf{R} \in \mathbb{R}^{q \times q}$ is any orthogonal matrix.

The Maximum Likelihood Estimators

The maximum likelihood estimators also minimize the error with respect to Frobenius norm

$$\left(\hat{\mathbf{W}},\ \hat{\sigma}^2\right) = \underset{\mathbf{W} \in \mathbb{R}^{p \times q}, \sigma^2 \in \mathbb{R}^+}{\arg\min} \left\|\hat{\mathbf{\Sigma}} - \left(\mathbf{W}\mathbf{W}^\top + \sigma^2\mathbf{I}\right)\right\|_F.$$

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The Expectation-Maximization Algorithm

For the model

$$y = Wx + \mu + \epsilon$$

where $\mathbf{x} \sim \mathcal{N}_q(\mathbf{0}, \mathbf{I})$ and $\epsilon \sim \mathcal{N}_p(\mathbf{0}, \sigma^2 \mathbf{I})$ are independent.

We regard $\{\mathbf{x}_{\alpha}\}_{\alpha=1}^{N}$ as missing data and $\{\mathbf{x}_{\alpha},\mathbf{y}_{\alpha}\}_{\alpha=1}^{N}$ as the complete data, then we can achieve

$$\mathbf{y}_{lpha} \, | \, \mathbf{x}_{lpha} \sim \mathcal{N}_{p}(\mathbf{W}\mathbf{x}_{lpha} + oldsymbol{\mu}, \sigma^{2}\mathbf{I})$$

and

$$\mathbf{x}_{lpha} \, | \, \mathbf{y}_{lpha} \sim \mathcal{N}_{q}(\mathbf{M}^{-1}\mathbf{W}^{ op}(\mathbf{y}_{lpha} - oldsymbol{\mu}), \sigma^{2}\mathbf{M}^{-1}),$$

where $\mathbf{M} = \mathbf{W}^{\mathsf{T}} \mathbf{W} + \sigma^2 \mathbf{I}$.

The Expectation-Maximization Algorithm

The update of the EM algorithm

1 In E-step, we take the expectation

$$I_C = \mathbb{E}\left[\ln\left(\prod_{lpha=1}^N f(\mathbf{x}_lpha \,|\, \mathbf{y}_lpha)
ight)
ight].$$

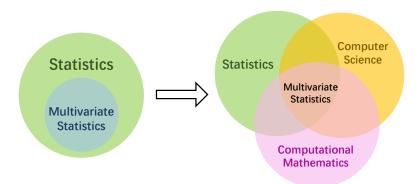
② In the M-step, we maximized I_C with respect to **W** and σ^2 :

$$\begin{aligned} \mathbf{W}_{+} = & \hat{\mathbf{\Sigma}} \mathbf{W} (\sigma^{2} \mathbf{I} + \mathbf{M}^{-1} \mathbf{W}^{\top} \hat{\mathbf{\Sigma}} \mathbf{W})^{-1}, \\ \sigma_{+}^{2} = & \frac{1}{\rho} \mathrm{tr} \left(\hat{\mathbf{\Sigma}} - \hat{\mathbf{\Sigma}} \mathbf{W} \mathbf{M}^{-1} \mathbf{W}_{+}^{\top} \right). \end{aligned}$$

Note that the computational complexity of EM is $\mathcal{O}(Npq)$, while the spectral decomposition in MLE requires $\mathcal{O}(Np^2 + p^3)$.

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Multivariate Statistics



Good Luck on Finals!



