

# Multivariate Statistical Analysis

## Lecture 14

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- 1 Likelihood Ratio Criterion and  $T^2$ -Statistic
- 2 Multivariate Analysis of Variance
- 3 Multivariate Linear Regression

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# Likelihood Ratio Criterion and $T^2$ -Statistic

Let  $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N$  constitute a sample from  $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  with  $N > p$ .

We shall derive  $T^2$ -Statistic

$$T^2 = N(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top \mathbf{S}^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)$$

from likelihood ratio criterion

$$\lambda = \frac{\max_{\boldsymbol{\Sigma} \in \mathbb{S}_p^{++}} L(\boldsymbol{\mu}_0, \boldsymbol{\Sigma})}{\max_{\boldsymbol{\mu} \in \mathbb{R}^p, \boldsymbol{\Sigma} \in \mathbb{S}_p^{++}} L(\boldsymbol{\mu}, \boldsymbol{\Sigma})}.$$

# Likelihood Ratio Criterion and $T^2$ -Statistic

We have

$$\lambda^{\frac{2}{N}} = \frac{1}{1 + T^2/(N-1)},$$

where

$$T^2 = N(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top \mathbf{S}^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}_0), \quad \bar{\mathbf{x}} = \frac{1}{N} \sum_{\alpha=1}^N \mathbf{x}_\alpha$$

and

$$\mathbf{S} = \frac{1}{N-1} \sum_{\alpha=1}^N (\mathbf{x}_\alpha - \bar{\mathbf{x}})(\mathbf{x}_\alpha - \bar{\mathbf{x}})^\top.$$

# Likelihood Ratio Criterion and $T^2$ -Statistic

The condition  $\lambda^{2/N} > c$  for some  $c \in (0, 1)$  is equivalent to

$$T^2 < \frac{(N-1)(1-c)}{c}.$$

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# Multivariate Analysis of Variance

We consider testing the equality of means with common covariance.

Let  $\mathbf{x}_\alpha^{(g)}$  be an observation from the  $g$ -th population  $\mathcal{N}_p(\boldsymbol{\mu}^{(g)}, \boldsymbol{\Sigma})$  for  $\alpha = 1, \dots, N_g$  and  $g = 1, \dots, q$ . We wish to test the hypothesis

$$H_0 : \boldsymbol{\mu}_1 = \dots = \boldsymbol{\mu}_g.$$



# Multivariate Analysis of Variance

The likelihood function is

$$L(\boldsymbol{\mu}^{(1)}, \dots, \boldsymbol{\mu}^{(g)}, \boldsymbol{\Sigma}) \\ = \prod_{g=1}^q \frac{1}{(2\pi)^{\frac{pN_g}{2}} (\det(\boldsymbol{\Sigma}))^{\frac{N_g}{2}}} \exp \left( -\frac{1}{2} \sum_{\alpha=1}^{N_g} (\mathbf{x}_{\alpha}^{(g)} - \boldsymbol{\mu}^{(g)})^{\top} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_{\alpha}^{(g)} - \boldsymbol{\mu}^{(g)}) \right).$$

- 1 We let  $\boldsymbol{\theta} = \{\boldsymbol{\mu}^{(1)}, \dots, \boldsymbol{\mu}^{(g)}, \boldsymbol{\Sigma}\}$  be the parameters.
- 2 The set  $\Omega$  is the space in which  $\boldsymbol{\Sigma}$  is positive definite and each  $\boldsymbol{\mu}^{(g)}$  is any  $p$ -dimensional vector.
- 3 The set  $\omega$  is the space in which  $\boldsymbol{\mu}^{(1)} = \dots = \boldsymbol{\mu}^{(g)}$  ( $p$ -dimensional vectors) and  $\boldsymbol{\Sigma}$  is positive definite matrix.

# Multivariate Analysis of Variance

The likelihood ratio criterion is

$$\lambda = \frac{\sup_{\boldsymbol{\theta} \in \omega} L(\boldsymbol{\theta})}{\sup_{\boldsymbol{\theta} \in \Omega} L(\boldsymbol{\theta})} = \frac{(\det(\hat{\boldsymbol{\Sigma}}_{\Omega}))^{\frac{N}{2}}}{(\det(\hat{\boldsymbol{\Sigma}}_{\omega}))^{\frac{N}{2}}},$$

where

$$\hat{\boldsymbol{\Sigma}}_{\Omega} = \frac{1}{N} \sum_{g=1}^q \sum_{\alpha=1}^{N_g} (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}}^{(g)}) (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}}^{(g)})^{\top}$$

and

$$\hat{\boldsymbol{\Sigma}}_{\omega} = \frac{1}{N} \sum_{g=1}^q \sum_{\alpha=1}^{N_g} (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}}) (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}})^{\top}.$$

# Multivariate Analysis of Variance

We can write

$$N\hat{\Sigma}_{\omega} = \mathbf{A} + \mathbf{B},$$

where

$$\mathbf{A} = N\hat{\Sigma}_{\Omega} = \sum_{g=1}^q \sum_{\alpha=1}^{N_g} (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}}^{(g)}) (\mathbf{x}_{\alpha}^{(g)} - \bar{\mathbf{x}}^{(g)})^{\top} \sim \mathcal{W}_p(\mathbf{\Sigma}, N - q)$$

and

$$\mathbf{B} = \sum_{g=1}^q N_g (\bar{\mathbf{x}}^{(g)} - \bar{\mathbf{x}}) (\bar{\mathbf{x}}^{(g)} - \bar{\mathbf{x}})^{\top} \sim \mathcal{W}_p(\mathbf{\Sigma}, q - 1)$$

are independent.

# Wilks' Lambda distribution

For two independent random matrices  $\mathbf{A} \sim \mathcal{W}_p(\mathbf{\Sigma}, n)$  and  $\mathbf{B} \sim \mathcal{W}_p(\mathbf{\Sigma}, m)$  with  $n \geq p$ , the ratio

$$\frac{\det(\mathbf{A})}{\det(\mathbf{A} + \mathbf{B})}$$

has Wilks' Lambda distribution with degrees of freedom  $n$  and  $m$ , which is typically written as

$$\frac{\det(\mathbf{A})}{\det(\mathbf{A} + \mathbf{B})} \sim \Lambda_{p,n,m}.$$

## Theorem

Let  $\mathbf{A} \sim \mathcal{W}_p(\mathbf{\Sigma}, n)$  and  $\mathbf{B} \sim \mathcal{W}_p(\mathbf{\Sigma}, m)$  be two independent Wishart distributed variables, then we can write

$$\frac{\det(\mathbf{A})}{\det(\mathbf{A} + \mathbf{B})} = \prod_{i=1}^p u_i \sim \Lambda_{p,n,m},$$

where  $u_1, \dots, u_p$  are independent distributed as

$$u_i \sim \text{Beta} \left( \frac{n+1-i}{2}, \frac{m}{2} \right).$$

Let  $\mathbf{A} \sim \mathcal{W}_p(\mathbf{\Sigma}, n)$ , we can follow above theorem to show

$$\det(\mathbf{A}) = \det(\mathbf{\Sigma}) \prod_{i=1}^p v_i$$

with some independent random variables  $v_1, \dots, v_p$ ?

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