

Multivariate Statistical Analysis

Lecture 12

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Outline

- 1 The Density of Wishart Distribution
- 2 T^2 -Statistic and F -Distribution
- 3 The Inverted Wishart Distribution

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1 The Density of Wishart Distribution

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The Density of Wishart Distribution

Theorem

The density of $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$ is

$$w_p(\mathbf{A} | \boldsymbol{\Sigma}, n) = \frac{(\det(\mathbf{A}))^{\frac{n-p-1}{2}} \exp\left(-\frac{1}{2}\text{tr}(\boldsymbol{\Sigma}^{-1}\mathbf{A})\right)}{2^{\frac{np}{2}} \pi^{\frac{p(p-1)}{4}} (\det(\boldsymbol{\Sigma}))^{\frac{n}{2}} \prod_{i=1}^p \Gamma\left(\frac{1}{2}(n+1-i)\right)}.$$

for positive definite \mathbf{A} and 0 elsewhere.

Sketch of the proof:

- ① Observe that $\mathbf{B} = \boldsymbol{\Sigma}^{-1/2}\mathbf{A}\boldsymbol{\Sigma}^{-1/2} \sim \mathcal{W}_p(\mathbf{I}_p, n)$.
- ② Find the density of $\mathbf{B} \sim \mathcal{W}_p(\mathbf{I}_p, n)$ by induction.
- ③ Recall that the Jacobian of transform from \mathbf{A} to \mathbf{B} has determinant

$$(\det(\boldsymbol{\Sigma}^{-1/2}))^{p+1} = (\det(\boldsymbol{\Sigma}))^{-\frac{p+1}{2}}.$$

- ④ Achieve the density of $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$.

The Wishart Distribution

The multivariate gamma function is defined as

$$\Gamma_p(t) = \pi^{\frac{p(p-1)}{4}} \prod_{i=1}^p \Gamma\left(t - \frac{1}{2}(i-1)\right).$$

We also write the density function of Wishart distribution as

$$w_p(\mathbf{A} | \boldsymbol{\Sigma}, n) = \frac{(\det(\mathbf{A}))^{\frac{n-p-1}{2}} \exp\left(-\frac{1}{2}\text{tr}(\boldsymbol{\Sigma}^{-1}\mathbf{A})\right)}{2^{\frac{np}{2}} \Gamma_p\left(\frac{n}{2}\right) (\det(\boldsymbol{\Sigma}))^{\frac{n}{2}}}.$$

Properties of Wishart Distribution

Let $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$ and partition \mathbf{A} and $\boldsymbol{\Sigma}$ into q and $p - q$ rows and columns as

$$\mathbf{A} = \begin{bmatrix} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{A}_{22} \end{bmatrix} \quad \text{and} \quad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{bmatrix},$$

then we have

- (a) $\mathbf{A}_{11} \sim \mathcal{W}_q(\boldsymbol{\Sigma}_{11}, n)$ and $\mathbf{A}_{22} \sim \mathcal{W}_{p-q}(\boldsymbol{\Sigma}_{22}, n)$;
- (b) if $q = 1$, then

$$\mathbf{a}_{21} \mid \mathbf{A}_{22} \sim \mathcal{N}_{p-q}(\mathbf{A}_{22}\boldsymbol{\Sigma}_{22}^{-1}\boldsymbol{\sigma}_{21}, \sigma_{11.2}^2 \mathbf{A}_{22})$$

$$\text{where } \sigma_{11.2}^2 = \sigma_{11} - \boldsymbol{\sigma}_{21}^\top \boldsymbol{\Sigma}_{22}^{-1} \boldsymbol{\sigma}_{21};$$

- (c) if $n > p - q$, then

$$\mathbf{A}_{11.2} = \mathbf{A}_{11} - \mathbf{A}_{12}\mathbf{A}_{22}^{-1}\mathbf{A}_{21} \sim \mathcal{W}_q(\boldsymbol{\Sigma}_{11.2}, n - p + q)$$

is independent on \mathbf{A}_{22} and \mathbf{A}_{12} , where $\boldsymbol{\Sigma}_{11.2} = \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\boldsymbol{\Sigma}_{21}$.

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The Distribution of Sample Covariance

Recall that we define

$$\mathbf{A} = \sum_{\alpha=1}^N (\mathbf{x}_\alpha - \bar{\mathbf{x}})(\mathbf{x}_\alpha - \bar{\mathbf{x}})^\top \quad \text{and} \quad \mathbf{S} = \frac{1}{n} \sum_{\alpha=1}^N (\mathbf{x}_\alpha - \bar{\mathbf{x}})(\mathbf{x}_\alpha - \bar{\mathbf{x}})^\top$$

where $\mathbf{x}_1, \dots, \mathbf{x}_N$ are independent, each with the distribution $\mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, and $n = N - 1 \geq p$.

We have $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$, then

$$\mathbf{S} \sim \mathcal{W}_p \left(\frac{1}{n} \boldsymbol{\Sigma}, n \right)$$

T^2 -Statistic

In hypothesis testing for the mean with unknown variance, we consider the t -student variable

$$t = \frac{\bar{x} - \mu}{s/\sqrt{N}},$$

where $\bar{x} = \frac{1}{N} \sum_{\alpha=1}^N x_\alpha$ and $s^2 = \frac{1}{N-1} \sum_{\alpha=1}^N (x_\alpha - \bar{x})^2$.

We have $t^2 = \frac{N(\bar{x} - \mu)^2}{s^2}$ and its multivariate analog is

$$T^2 = N(\bar{\mathbf{x}} - \boldsymbol{\mu})^\top \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}),$$

where $\bar{\mathbf{x}} = \frac{1}{N} \sum_{\alpha=1}^N \mathbf{x}_\alpha$ and $\mathbf{S} = \frac{1}{N-1} \sum_{\alpha=1}^N (\mathbf{x}_\alpha - \bar{\mathbf{x}})(\mathbf{x}_\alpha - \bar{\mathbf{x}})^\top$.

F -Distribution

The F -distribution with d_1 and d_2 degrees of freedom is the distribution of

$$z = \frac{y_1/d_1}{y_2/d_2} = \frac{d_2 y_1}{d_1 y_2},$$

where $y_1 \sim \chi_{d_1}^2$ and $y_2 \sim \chi_{d_2}^2$ are independent, written as

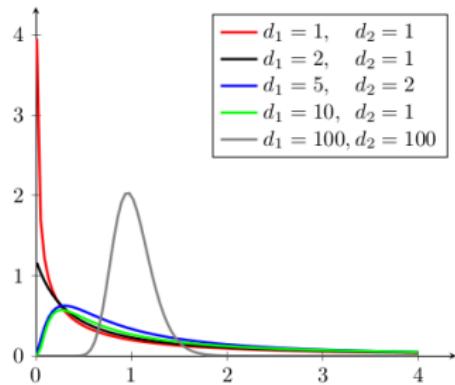
$$z \sim F_{d_1, d_2}.$$

F -Distribution

The density function of F -distribution is

$$f(z; d_1, d_2) = \frac{1}{B\left(\frac{d_1}{2}, \frac{d_2}{2}\right)} \left(\frac{d_1}{d_2}\right)^{\frac{d_1}{2}} z^{\frac{d_1}{2}-1} \left(1 + \frac{d_1 z}{d_2}\right)^{-\frac{d_1+d_2}{2}}$$

for $z > 0$, where $B(\alpha, \beta) = \int_0^1 t^{\alpha-1} (1-t)^{\beta-1} dt$.



T^2 -Statistic and F -Distribution

Theorem

Let $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$ and $\mathbf{y} \sim \mathcal{N}_p(\mathbf{0}, \boldsymbol{\Sigma})$ be independent with $n \geq p$, then

$$\frac{n-p+1}{p} \cdot \mathbf{y}^\top \mathbf{A}^{-1} \mathbf{y} \sim F_{p, n-p+1}.$$

For T^2 -statistic

$$T^2 = N(\bar{\mathbf{x}} - \boldsymbol{\mu})^\top \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}),$$

we have

$$\frac{N-p}{(N-1)p} \cdot T^2 \sim F_{p, n-p+1}.$$

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The Inverted Wishart Distribution

If $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, m)$, then $\mathbf{B} = \mathbf{A}^{-1}$ has the inverted Wishart distribution with m degrees of freedom and scale parameter $\boldsymbol{\Psi} = \boldsymbol{\Sigma}^{-1}$, written as

$$\mathbf{B} \sim \mathcal{W}_p^{-1}(\boldsymbol{\Psi}, m).$$

The density function of \mathbf{B} is

$$w^{-1}(\mathbf{B} | \boldsymbol{\Psi}, m) = \frac{(\det(\boldsymbol{\Psi}))^{\frac{m}{2}} (\det(\mathbf{B}))^{-\frac{m+p+1}{2}} \exp(-\frac{1}{2}\text{tr}(\boldsymbol{\Psi}\mathbf{B}^{-1}))}{2^{\frac{mp}{2}} \Gamma_p(\frac{m}{2})},$$

where

$$\Gamma_p(t) = \pi^{\frac{p(p-1)}{4}} \prod_{i=1}^p \Gamma\left(t - \frac{1}{2}(i-1)\right).$$

Quiz

Define $\bar{\mathbb{S}}^p \rightarrow \mathbb{R}^{p \times p}$ as

$$\mathbf{F}(\mathbf{X}) = \mathbf{X}^{-1},$$

where $\bar{\mathbb{S}}^p = \{\mathbf{X} \in \mathbb{R}^{p \times p} : \mathbf{X} = \mathbf{X}^\top \text{ and } \mathbf{X} \text{ is non-singular}\}.$

What is the determinant of Jacobian of $\mathbf{F}(\mathbf{X})$?

Quiz

Let $\mathbf{x} \sim \mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. We define $\boldsymbol{\Psi} = \boldsymbol{\Sigma}^{-1}$ as the precision matrix.

- ① It is well-known that

$$\sigma_{ij} = 0 \text{ if and only if } x_i \text{ and } x_j \text{ are independent.}$$

- ② What is the meaning of $\psi_{ij} = 0$?

The Conjugate Prior for the Covariance

Theorem

If $\mathbf{A} \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n)$ and $\boldsymbol{\Sigma}$ has a prior distribution $\mathcal{W}^{-1}(\boldsymbol{\Psi}, m)$, then the conditional distribution of $\boldsymbol{\Sigma}$ given \mathbf{A} is the inverted Wishart distribution

$$\mathcal{W}^{-1}(\mathbf{A} + \boldsymbol{\Psi}, n + m).$$

Let each of $\mathbf{x}_1, \dots, \mathbf{x}_N$ has distribution $\mathcal{N}_p(\mathbf{0}, \boldsymbol{\Sigma})$ independently and $n = N - 1$, then the sample covariance

$$\mathbf{S} = \frac{1}{n} \sum_{i=1}^n (\mathbf{x}_\alpha - \bar{\mathbf{x}})(\mathbf{x}_\alpha - \bar{\mathbf{x}})^\top \sim \mathcal{W}_p(\boldsymbol{\Sigma}, n).$$

If $\boldsymbol{\Sigma} \sim \mathcal{W}_p^{-1}(\boldsymbol{\Psi}, m)$, then we have

$$\boldsymbol{\Sigma} | \mathbf{S} \sim \mathcal{W}^{-1}(n\mathbf{S} + \boldsymbol{\Psi}, n + m).$$

The Inverted Wishart Distribution

Theorem

Let $\mathbf{x}_1, \dots, \mathbf{x}_N$ be observations from $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. Suppose $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ have prior densities

$$n \left(\boldsymbol{\mu} \mid \nu, \frac{\boldsymbol{\Sigma}}{K} \right) \quad \text{and} \quad w^{-1}(\boldsymbol{\Sigma} \mid \boldsymbol{\Psi}, m)$$

respectively, where $n = N - 1$. Then the posterior density of $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ given

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{\alpha=1}^N \mathbf{x}_{\alpha} \quad \text{and} \quad \mathbf{S} = \frac{1}{N-1} \sum_{\alpha=1}^N (\mathbf{x}_{\alpha} - \bar{\mathbf{x}})(\mathbf{x}_{\alpha} - \bar{\mathbf{x}})^{\top}$$

is

$$n \left(\boldsymbol{\mu} \mid \frac{N\bar{\mathbf{x}} + K\nu}{N+K}, \frac{\boldsymbol{\Sigma}}{N+K} \right) \cdot w^{-1} \left(\boldsymbol{\Sigma} \mid \boldsymbol{\Psi} + n\mathbf{S} + \frac{NK(\bar{\mathbf{x}} - \nu)(\bar{\mathbf{x}} - \nu)^{\top}}{N+K}, N+m \right).$$