

Markov chain Monte Carlo

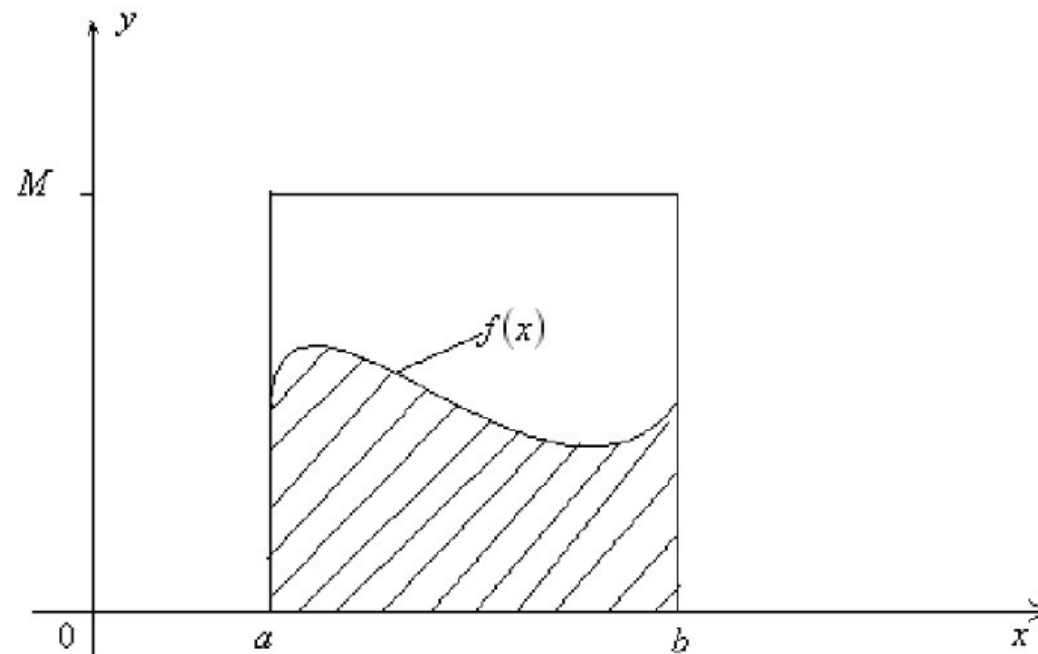
shenchuyun

CONTENTS

- Monte Carlo method
 - Importance Sampling
 - accept-reject sample method
- Markov Chain
- Markov Chain Monte Carlo
 - MH sampling
 - Gibbs sampling

Monte Carlo method

- 假设概率分布的定义已知，通过抽样获得概率分布的随机样本，再通过得到的样本对概率分布的特征进行分析。
- $\theta = \int_a^b f(x)dx$ 原函数难求
- 随机采样 n 个点
- $\theta = \int_a^b f(x)dx = \frac{b-a}{n} \sum_{i=0}^{n-1} f(x_i)$



Monte Carlo method

- 求数学期望

- 在 $p(x)$ 上随机采样 n 个样本点, 然后计算均值 $\hat{f}_n = \frac{1}{n} \sum_{i=1}^n f(x_i)$

- 作为数学期望 $E_{p(x)}[f(x)]$

- $E_{p(x)}[f(x)] = \frac{1}{n} \sum_{i=1}^n f(x_i), n \rightarrow \infty$

- 求积分

- $\int_D h(x) dx = \int_D f(x)p(x)dx = E_{p(x)}[f(x)]$ 其中 $f(x) = \frac{h(x)}{p(x)}$

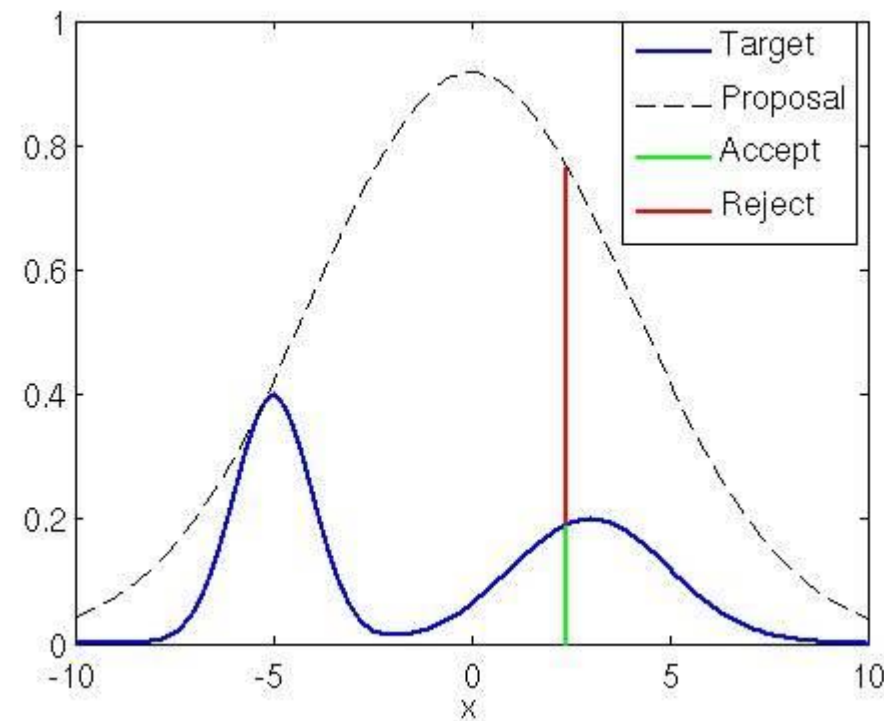
- $\int_D h(x) dx \approx \frac{1}{n} \sum_{i=1}^n f(x_i)$

Monte Carlo method (Importance Sampling)

- $\mu = \int_D f(x)p(x)dx = \int_D \frac{f(x)p(x)}{q(x)} q(x)dx = E_q \left(\frac{f(X)p(X)}{q(X)} \right)$
- $p(x)$ 不好直接采样, 在 q 上采样 n 个点
- $\mu \approx \frac{1}{n} \sum_{i=1}^n \frac{f(x_i)p(x_i)}{q(x_i)}$

Monte Carlo method (accept-reject sample method)

- $P(x)$ 不可以直接抽样， 找一个可以直接抽样的分布 (proposal distribution)
- 重复以下步骤， 直到采样到 m 个样本：
 - 从 $q(x)$ 采样一个 x^* ， 满足 $c * q(x) \geq p(x)$
 - 从 $U(0,1)$ 均匀分布中得到 u
 - if $u \leq \frac{p(x^*)}{cq(x^*)}$, 则将 x^* 作为抽样结果
 - else 拒绝该样本



Markov Chain

- $P(X_t|X_0, X_1, \dots, X_{t-1}) = P(X_t|X_{t-1})$
- $P(X_t|X_{t-1})$ 称为马尔科夫链的转移概率分布

- 离散状态马尔可夫转移矩阵 $P = \begin{bmatrix} P_{11} & P_{12} & P_{13} & \dots \\ P_{21} & P_{22} & P_{23} & \dots \\ P_{31} & P_{32} & P_{33} & \dots \\ \dots & \dots & \dots & \dots \end{bmatrix}$

- X 在时刻 t 的概率分布 $\pi(t) = \begin{bmatrix} \pi_1(t) \\ \pi_2(t) \\ \dots \end{bmatrix}$

- $\pi_i(t) = P(X_t = i), i = 1, 2, \dots$

Markov Chain

- X 在 t 时刻的状态分布可以由时刻 $(t-1)$ 的状态分布和转移矩阵决定
- $\pi(t) = P\pi(t-1)$
- 平稳分布 $\pi = P\pi$
- 遍历定理：直观解释：满足一定条件的马尔科夫链，当 t 趋于无穷，马尔科夫链的状态分布趋于平稳分布

Markov Chain Monte Carlo

- 基本思想：我们的目标分布是 $p(x)$ ，定义一个马尔科夫链 $X = \{X_0, X_1, X_2, \dots, X_t, \dots\}$ ，使其平稳分布就是抽样的目标 $p(x)$ ，然后在马尔科夫链上随机游走，每个时刻就可以得到一个样本
- **Detailed Balance Condition:** $\pi(i)P_{ij} = \pi(j)P_{ji}$ for all i, j
- 充分性证明：
 - $\sum_i \pi(i)P_{ij} = \sum_i \pi(j)P_{ji} = \pi(j) \sum_i P_{ji} = \pi(j)$
 - $\pi P = \pi$

Metropolis-Hastings

- $p(x^*|x) = q(x^*|x)\alpha(x^*|x)$ (*proposal distribution*) * (*acceptance distribution*) $q(x^*|x)$ 是容易抽样的分布
- $\alpha(x^*|x) = \min \left\{ 1, \frac{\pi(x^*)q(x|x^*)}{\pi(x)q(x^*|x)} \right\}$
- MH
 - $u \sim U(0,1)$
 - $x^* \sim q(x^*|x^i)$
 - if $u < \alpha(x^*|x)$ $x^{i+1} = x^*$
 - else $x^{i+1} = x^i$

Gibbs sampling

- Given a starting sample $(x_1, y_1, z_1)^T$
- You want to sample
 - $(x_2, y_2, z_2)^T, (x_3, y_3, z_3)^T, \dots, (x_N, y_N, z_N)^T \sim P(x, y, z)$
- Sampling:
 - $x_2 \sim P(x|y_1, z_1)$
 - $y_2 \sim P(y|x_2, z_1)$
 - $z_2 \sim P(z|x_2, y_2)$
 - ...

Gibbs sampling

- A special case of M-H
 - Let $\mathbf{x} = x_1, \dots, x_D$
 - When sampling k^{th} component, $q_k(\mathbf{x}^* | \mathbf{x}) = \pi(\mathbf{x}_k^* | \mathbf{x}_{-k})$
 - When sampling k^{th} component, $\mathbf{x}_{-k}^* = \mathbf{x}_{-k}$

$$\frac{\pi(\mathbf{x}^*)q(\mathbf{x} | \mathbf{x}^*)}{\pi(\mathbf{x})q(\mathbf{x}^* | \mathbf{x})} = \frac{\pi(\mathbf{x}^*)\pi(x_k | \mathbf{x}_{-k}^*)}{\pi(\mathbf{x})\pi(x_k^* | \mathbf{x}_{-k})} = \frac{\pi(x_k^* | \mathbf{x}_{-k}^*)\pi(x_k | \mathbf{x}_{-k}^*)}{\pi(x_k | \mathbf{x}_{-k})\pi(x_k^* | \mathbf{x}_{-k})} = 1$$