## More on MLE; Bayes Estimators

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#### Restricted MLE

Special attention is needed to make sure  $\widehat{\theta} \in \Theta$ 

- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1)$ , where  $\theta \geq 0$ .
- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} Bin(1, p)$ , where  $0 \le p \le 1$ .
- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ , where  $\mu \geq 0$ .



### Other cases: monotone likelihood

- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(0, \theta)$ , where  $\theta > 0$ .
- Example:  $X_1, \dots, X_n \stackrel{\textit{iid}}{\sim}$  exponential location family with pdf

$$f(x) = \exp^{-(x-\theta)}, \quad \text{if} \quad x \ge \theta$$

- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(\theta 1/2, \theta + 1/2)$
- Example: Let X by a single observation taking values from  $\{0,1,2\}$  according to  $P_{\theta}$ , where  $\theta=\theta_0$  or  $\theta_1$ . The probability of X is summarized



#### Practical issues with MLE

#### • Multimodality:

- Likelihood function can be multimodal, often have to use numerical techniques to try to maximize (no closed-form max).
- Can get stuck in local modes.
- Solutions:
  - **1** Choose models such that L is convex in  $\theta$ .
  - Heuristic search, multiple starting points.
  - Satisfied with a local maximum.
- Flatness and sensitivity:
  - $L(\theta; \mathbf{x})$  can be pretty flat near the max.
  - So a slightly different sample x may give a very different MLE.



#### Remarks on the MLE:

- The MLE  $\widehat{\theta}(\mathbf{x})$  is the value for which the observed sample  $\mathbf{x}$  is most likely; possess some optimal properties (Chapter 10)
- The MLE can be numerically sensitive to the variation in the data. Example: Bin(k, p).
- If T is (minimal) sufficient for  $\theta$ , then the MLE  $\widehat{\theta}$  must be a function of T. By factorization theorem, we have

$$L(\theta|\mathbf{x}) = f(\mathbf{x}|\theta) = g(T(\mathbf{x})|\theta)h(\mathbf{x}),$$

and the MLE  $\widehat{\theta}$  should maximize  $g(T(\mathbf{X})|\theta)$ . Therefore, the MLE is a function of the (minimal) sufficient statistic.



## **Invariance Property of MLE**

#### Theorem

If  $\widehat{\theta}$  is the MLE of  $\theta$ , then for any function  $\tau(\theta)$ , the MLE of  $\tau(\theta)$  is  $\tau(\widehat{\theta})$ .

#### Examples:

- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} Bin(1, p)$ . Find the MLE of  $\sqrt{p(1-p)}$
- Example:  $X_1, \dots, X_n \stackrel{iid}{\sim} Poi(\lambda)$ . Find the MLE of  $P(X \le 1)$
- Example:  $X_1, \dots, X_n \stackrel{\textit{iid}}{\sim} N(\mu, \sigma^2)$ . Find the MLE of  $\mu/\sigma$
- Example: Find the MLE of the population median
- Find the MLE for c such that  $P(\bar{X} > c) = 0.025$ . (the 97.5% percentile of the distribution of  $\bar{X}$ .



## Bayes Estimators

- Different from classical approaches, in the Bayesian approach  $\theta$  is considered as a random quantity whose variation can be described by a probability distribution (called the prior distribution).
- A sample is then taken from a population indexed by  $\theta$  and the prior distribution is updated with this sample information. The updated prior is called the posterior distribution.

## Important Concepts in Bayes Estimation

- prior distribution of  $\theta$ :  $\theta \sim \pi(\theta)$
- sampling distribution of **y** given  $\theta$ :  $\mathbf{y}|\theta \sim f(\mathbf{y}|\theta)$
- posterior distribution of  $\theta$ :  $\pi(\theta|\mathbf{y}) = f(\mathbf{y}|\theta)\pi(\theta)/m(\mathbf{x})$
- marginal distribution of **y**:  $m(\mathbf{y}) = \int f(\mathbf{y}|\theta)\pi(\theta)d\theta$
- posterior mean of  $\theta$ :  $E(\theta|\mathbf{y}) = \int \theta \pi(\theta|\mathbf{y}) d\theta$  (Bayes estimator of  $\theta$ )

## Bayes Estimation: Examples

- Example: Assume  $X_1, \dots, X_n$  iid Bin(1, p). Assume the prior distribution on p is  $Beta(\alpha, \beta)$  with known parameters  $(\alpha, \beta)$ . Find the posterior distribution of p and the Bayes estimator of p Special case:  $\pi(p) \sim Unif(0, 1)$
- Remark: If  $T(\mathbf{x})$  is a sufficient statistic, then the posterior density of  $\theta$  is  $\pi(\theta|\mathbf{x}) = \pi(\theta|T(\mathbf{x}))$

## Bayes Estimation: Conjugate family

Let  $\mathcal F$  denote the class of pdfs or pmfs  $f(x|\theta)$ . A class  $\Pi$  of prior distributions is a *conjugate family* for  $\mathcal F$  if the posterior distribution is in the class  $\Pi$  for all  $f\in \mathcal F$ , all priors in  $\Pi$ , and all  $x\in \mathcal X$ 

#### Examples:

- Example:(Beta-Binomial Conjugate)
- Example:(Gamma-Poisson Conjugate)
- Example: (Normal-Normal Conjugate) Let  $X_1, \dots, X_n$  be iid  $\sim N(\theta, \sigma^2)$ , with  $\theta$  unknown and  $\sigma^2$ known. Suppose that the prior distribution of  $\theta$  is  $N(\mu, \tau^2)$ . Here we assume both  $\mu$  and  $\tau^2$  are given. Find the posterior distribution of  $\theta$ .

# **Bayesians vs.Frequentists**

You are no good when sample is small



You give a different answer for different priors