

Problem 2:

Q1. (1). Complete statistics: $\max_{1 \leq i \leq n} |X_i|$.

Let $Y_i = |X_i|$. then $Y_{(n)}$ is complete for θ .

$$E_{\theta}(Y_{(n)}) = \int_0^{\theta} \left(1 - \left(\frac{y}{\theta}\right)^n\right) dy = \frac{n\theta}{n+1}.$$

$$(Or \ P_{\theta}(Y_{(n)} \leq y) = \left(\frac{y}{\theta}\right)^n \Rightarrow E_{\theta}(Y_{(n)}) = \int_0^{\theta} \theta^{-n} \cdot n y^{n-1} \cdot y dy).$$

$\Rightarrow \frac{n+1}{n} Y_{(n)}$ is unbiased for θ .

$$Also: E_{\theta}(Y_{(n)}^{-1}) = \int_0^{\theta} \theta^{-n} \cdot n y^{n-1} y^{-1} dy$$

$$= \frac{n}{n-1} \theta^{-1} \Rightarrow \frac{n-1}{n} Y_{(n)}^{-1} \text{ is unbiased for } \theta^{-1}.$$

$\Rightarrow \frac{1}{n}(n+1) Y_{(n)} + \frac{1}{n}(n-1) Y_{(n)}^{-1}$ is the UMVUE.

(2). (a) Note that the MLE for η is:

$$\hat{\eta} = \frac{2 \sum_{i=1}^n X_i}{n(n+1)}, \text{ unbiased for } \eta.$$

$$\Rightarrow MSE_{\eta}(\hat{\eta}) = Var_{\eta}(\hat{\eta}) = Var_{\lambda^2} \left(\frac{2 \sum_{i=1}^n X_i}{n(n+1)} \right).$$

$$= \left(\frac{2}{n(n+1)} \right)^2 \sum_{i=1}^n Var_{\lambda^2}(X_i).$$

$$= \left(\frac{2}{n(n+1)} \right)^2 \sum_{i=1}^n i \lambda^2 = \frac{2\lambda^2}{n(n+1)}.$$

$$(b). \text{CRLB: } \frac{\partial^2}{\partial \eta^2} \log L(\eta; x) = -\frac{1}{\eta^2} \sum_{i=1}^n X_i.$$

$$\Rightarrow E_{\eta} \left(-\frac{\partial^2}{\partial \eta^2} \log L(\eta; x) \right) = \frac{1}{\eta^2} \sum_{i=1}^n E(X_i) = \frac{n(n+1)}{2\eta}.$$

$$\Rightarrow CRLB = \frac{2\eta}{n(n+1)}. \quad \checkmark \text{ attained.}$$

(c). Attainment thm.

$$\begin{aligned}\frac{\partial}{\partial \eta} \log L(\eta; \underline{x}) &= \frac{1}{\eta} \sum_{i=1}^n x_i - \frac{n(n+1)}{2} \\ &= \frac{1}{\eta} \left(\sum_{i=1}^n x_i - \frac{n(n+1)}{2} \eta \right) \\ &= \frac{n(n+1)}{2\eta} (\hat{\eta} - \eta).\end{aligned}$$

Set $a(\eta) = \frac{n(n+1)}{2\eta} \Rightarrow$ CRLB is attainable.

Q2. $f(\theta | \underline{x}) \propto \pi(\theta) \prod_{i=1}^n f(x_i | \theta).$

$$\begin{aligned}&= \theta^{c-1} e^{-\lambda\theta} \cdot \theta^n e^{-\theta \sum_{i=1}^n x_i} \\ &= \theta^{c+n-1} e^{-\theta(\lambda + \sum_{i=1}^n x_i)}.\end{aligned}$$

\Rightarrow Posterior distⁿ: $\text{Gam}(c+n, \sum_{i=1}^n x_i + \lambda).$

① When $L(\theta, a) = (\theta - a)^2$, the Bayes risk is minimized at the posterior mean:

$$E(\theta | \underline{x}) = \frac{c+n}{\lambda + \sum_{i=1}^n x_i}.$$

② When $L(\theta, a) = \mathbb{1}_{\{|\theta - a| > \varepsilon\}}$, the Bayes risk is minimized at the posterior mode:

$$\begin{aligned}\frac{\partial}{\partial \theta} \log f(\theta | \underline{x}) &= \text{Constant} + \frac{\partial}{\partial \theta} \left((c+n-1) \log \theta - \theta \left(\lambda + \sum_{i=1}^n x_i \right) \right) \\ &= \frac{c+n-1}{\theta} - \left(\lambda + \sum_{i=1}^n x_i \right) \stackrel{\text{set}}{=} 0.\end{aligned}$$

$$\Rightarrow \hat{\theta} = \frac{c+n-1}{\lambda + \sum_{i=1}^n x_i}.$$

$$\frac{\partial^2}{\partial \theta^2} \log f(\theta | \underline{x}) = -\frac{c+n-1}{\theta^2} < 0. \Rightarrow \text{the posterior mode is } \frac{c+n-1}{\lambda + \sum_{i=1}^n x_i}.$$

Q3: (1). $\Theta_0 = \{0\}$. Set $\theta_0 = 0$.

\therefore Under H_0 , $IP_{\theta_0}(X_{(n)} \geq 1) = 0$.

$\Rightarrow \sup_{\theta \in \Theta_0} IP_{\theta}(X_{(1)} \geq k \text{ or } X_{(n)} \geq 1)$

$= IP_{\theta_0}(X_{(1)} \geq k) = (1-k)^n, 0 \leq k \leq 1$

Set $(1-k)^n = \alpha \Rightarrow k = 1 - \alpha^{1/n}$ to have a size α test.

(2). $L(\theta; x) = \theta e^{-\theta x}$

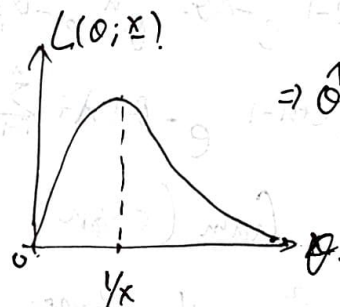
$\lambda(x) = \frac{\sup_{0 \leq \theta \leq 1} \theta e^{-\theta x}}{\sup_{\theta \geq 0} \theta e^{-\theta x}}$

$= \begin{cases} 1 & \text{if } \frac{1}{x} \leq 1 \\ x e^{-(x-1)} & \text{if } \frac{1}{x} > 1 \end{cases}$

$= \begin{cases} 1 & \text{if } x \geq 1 \\ x e^{-(x-1)} & \text{if } x < 1 \end{cases}$

Note that

$\frac{\partial L(\theta; x)}{\partial \theta} = (1 - \theta x) e^{-\theta x} = \begin{cases} \geq 0 & \text{if } \theta \leq 1/x \\ < 0 & \text{if } \theta > 1/x \end{cases}$



$\Rightarrow \hat{\theta}_{MLE} = \frac{1}{x} \mathbb{1}_{\{x \leq 1\}} + 1 \cdot \mathbb{1}_{\{x > 1\}}$

$\hat{\theta}_{MLE} = \frac{1}{x}$

\Rightarrow By CRT: Reject if $\lambda(x) \leq c$.

$\frac{d}{dx}(x e^{-(x-1)}) = (1-x) e^{-(x-1)} \geq 0$ for $\forall x \leq 1$.

$\Rightarrow \lambda(x)$ is \nearrow in x .

$\therefore \{x: \lambda(x) \leq c\} = \{x: x \leq c^*\}$

Set $\alpha = 0.05 = \sup_{\theta \leq 1} IP_{\theta}(X \leq c^*) = \sup_{\theta \leq 1} (1 - e^{-\theta \cdot c^*})$

$= 1 - e^{-c^*}$

$\Rightarrow c^* = -\log 0.95$

i.e. Reject if $x \leq -\log 0.95$.

Power fn: $\pi(\theta) = IP_{\theta}(X \leq -\log 0.95) = 1 - (0.95)^{\theta}$

Since $-\log 0.95 \approx 0.05$

then with $X_1 = 0.1$, $n=1$, we will NOT reject H_0 at $\alpha = 0.05$.