

EX 8.51: H_0 vs H_1 , $W(X)$ test stat. Suppose $0 < \alpha \leq 1$, choose c_α s.t. $\{x: W(x) \geq c_\alpha\}$ is the rejection region of a size α test of H_0 . show that p -value $P(x)$ in (8.3.9) is the smallest α level at which we could reject H_0 , having observed the data x .

proof: ① if $\alpha < P(x)$, $\sup_{\theta \in \Theta_0} P_\theta(W(X) \geq c_\alpha) = \alpha < P(x) = \sup_{\theta \in \Theta_0} P(W(X) \geq W(x))$

, $\Rightarrow W(x) < c_\alpha$, couldn't reject H_0 at level α , with observation x .

② if $\alpha \geq P(x)$, $\sup_{\theta \in \Theta_0} P_\theta(W(X) \geq c_\alpha) = \alpha \geq P(x) = \sup_{\theta \in \Theta_0} P(W(X) \geq W(x))$

if $W(x) < c_\alpha$, let $c'_\alpha = W(x)$ then $\alpha = \sup_{\theta \in \Theta_0} P_\theta(W(X) \geq c'_\alpha) = P(x)$, reject H_0 at size α

if $W(x) \geq c_\alpha$, then reject H_0 at level α

By 1) 2) \Rightarrow reject H_0 at level α having observed x

EX 9.13. X a single observation from the $\text{Beta}(\theta, 1)$ pdf.

(a) $Y = -(\log X)^{-1}$. Evaluate the confidence coefficient of the set $[4/2, 4]$

(b) Find a pivotal quantity & use it to set up a confidence interval having the same confidence coefficient as the interval in part (a)

(c) Compare the two confidence intervals

Proof (a) $Y = -(\log X)^{-1}$, $X \sim \text{Beta}(\theta, 1) \Rightarrow f_X(x) = \frac{\theta}{x^{1+\theta}} e^{-\theta/x}$, $0 < x < \infty$

$$P\left(\frac{Y}{2} \leq \theta \leq Y\right) = P(\theta \leq Y \leq 2\theta) = \int_0^{2\theta} \frac{\theta}{y^{1+\theta}} e^{-\theta/y} dy = e^{-\theta/y} \Big|_0^{2\theta} = e^{-1/2} - e^{-1} = 0.239$$

(b) $f_X(x) = \theta x^{\theta-1}$, $0 < x < 1$, $Q(t, \theta) := t^\theta$, $g(Q(t, \theta)) \equiv 1$

$$f(t|\theta) = \theta t^{\theta-1} = 1 \cdot \frac{\partial}{\partial t} Q(t, \theta) = g(Q(t, \theta)) \cdot \left| \frac{\partial}{\partial t} Q(t, \theta) \right|$$

$\Rightarrow Q(X, \theta) := X^\theta$ is the Pivotal quantity, with $f_T(t|\theta) \equiv 1$ $\forall 0 < t < 1$

\Rightarrow Suppose CI: $C(X) = \{\theta : a \leq Q(X, \theta) \leq b\}$

$$\text{Then } P_\theta(\theta \in C(X)) = P(a \leq Q(X, \theta) \leq b) \stackrel{f_T(t|\theta) \equiv 1}{=} b - a$$

$$\begin{aligned} \text{Since } Q(t, \theta) = t^\theta &\Rightarrow C(X) = \left\{ \theta : \frac{\log b}{\log X} \leq \theta \leq \frac{\log a}{\log X} \right\} \\ \Rightarrow b - a &\stackrel{(a)}{=} 0.239, \Rightarrow \text{CI: } \left[\frac{\log(b+0.239)}{\log X}, \log a / \log X \right] \end{aligned}$$

(c) part (a): $C_1(X) = \{\theta : \frac{Y}{2} \leq \theta \leq Y\} = \{\theta : -\frac{1}{2\log X} \leq \theta \leq \frac{1}{\log X}\}$

By (a) $C_1(X)$ with confidence coefficient 0.239

part (b): a generalization CI of part (a)

$$C_2(X) = \left\{ \theta : \frac{\log b}{\log X} \leq \theta \leq \frac{\log a}{\log X} \right\}$$

Suppose $C_2(X)$ with confidence coefficient $1-\alpha$

then $b-\alpha = 1-\alpha$, $b = 1-\alpha+a$

$$\text{give } X, \min_{0 \leq a \leq \alpha} \frac{\log a - \log b}{\log X} \iff \min_{0 < X < 1} \log b - \log a \stackrel{b=1-\alpha+a}{\iff} \min \log\left(\frac{1-\alpha}{a} + 1\right)$$

$\Rightarrow a^* = \alpha$, $b^* = 1 \Rightarrow$ best $1-\alpha$ pivotal interval is $\theta \in \left[0, \frac{\log \alpha}{\log X}\right]$

which is better than $C_1(X)$ in part (a), by set $1-\alpha = 0.239$

the $\left[0, \frac{\log(1-0.239)}{\log X}\right]$ is better than $\left[-\frac{1}{2\log X}, \frac{1}{\log X}\right]$

EX3:

$$\begin{aligned}
 \text{proof: (a) type I error} &= P(P_n(X_1, \dots, X_n) \leq \alpha \mid \mu=0) \\
 &= P(1 - \Phi(\sqrt{n} \bar{X}) \leq \alpha \mid \mu=0) \\
 &= P(1 - \alpha \leq \Phi(\sqrt{n} \bar{X}) \mid \mu=0) \\
 &= P(\sqrt{n} \bar{X} \geq z_\alpha \mid \mu=0) \\
 &= P\left(\frac{\bar{X} - \mu}{\sqrt{n}} \geq \frac{z_\alpha}{n} - \frac{\mu}{\sqrt{n}} \mid \mu=0\right) \\
 &= 1 - \Phi\left(\frac{z_\alpha}{n} - \frac{\mu}{\sqrt{n}}\right) \stackrel{\mu=0}{=} 1 - \Phi\left(\frac{z_\alpha}{n}\right)
 \end{aligned}$$

$$\begin{aligned}
 \text{type II error} &= P(P_n(X_1, \dots, X_n) > \alpha \mid \mu=\mu_1) \\
 &= P(1 - \alpha > \Phi(\sqrt{n} \bar{X}) \mid \mu=\mu_1) \\
 &= P\left(\frac{\bar{X} - \mu}{\sqrt{n}} < \frac{z_\alpha}{n} - \frac{\mu}{\sqrt{n}} \mid \mu=\mu_1\right) \\
 &= \Phi\left(\frac{z_\alpha}{n} - \frac{\mu_1}{\sqrt{n}}\right)
 \end{aligned}$$

$$\begin{aligned}
 \text{(b) type I error} &= 1 - P(P_{10}(X_{10+t1}, \dots, X_{10+t10}) > \alpha, \forall t \in \{0, \dots, L-1\} \mid \mu=0) \\
 &= 1 - \prod_{t=0}^{L-1} P(1 - \Phi(\sqrt{n} \bar{X}_t) > \alpha \mid \mu=0) \\
 &= 1 - \prod_{t=0}^{L-1} \Phi\left(\frac{z_\alpha}{10}\right) \\
 &= 1 - \left[\Phi\left(\frac{z_\alpha}{10}\right)\right]^L
 \end{aligned}$$

$$\begin{aligned}
 \text{type II error} &= P(P_{10}(X_{10+t1}, \dots, X_{10+t10}) > \alpha \mid \mu=\mu_1) \\
 &= \prod_{t=0}^{L-1} P(1 - \Phi(\sqrt{n} \bar{X}_t) > \alpha \mid \mu=\mu_1) \\
 &= \prod_{t=0}^{L-1} P\left(\frac{\bar{X} - \mu}{\sqrt{10}} < \frac{z_\alpha}{10} - \frac{\mu}{\sqrt{10}} \mid \mu=\mu_1\right) \\
 &= \left[\Phi\left(\frac{z_\alpha}{10} - \frac{\mu_1}{\sqrt{10}}\right)\right]^L
 \end{aligned}$$

$$\text{Suppose } \hat{p} := P(P_{10}(X_1, \dots, X_{10}) \leq \alpha) \stackrel{X_1, \dots, X_{10} \text{ ind.}}{=} 1 - \Phi\left(\frac{z_\alpha}{10} - \frac{\mu}{\sqrt{10}}\right)$$

$$\begin{aligned}
 \text{Then } E[N] &= 10 \cdot \hat{p} + 2 \cdot 10 \cdot (1 - \hat{p}) \hat{p} + 3 \cdot 10 \cdot (1 - \hat{p})^2 \hat{p} + \dots + 1 \cdot 10 \cdot (1 - \hat{p})^{L-1} \hat{p} \\
 &\quad + 1 \cdot 10 (1 - \hat{p})^L = 10 \hat{p} [1 + 2(1 - \hat{p}) + 3(1 - \hat{p})^2 + \dots + 1(1 - \hat{p})^{L-1}] + 1 \cdot 10 (1 - \hat{p})^L
 \end{aligned}$$

$$\text{Suppose } S := 1 + 2(1 - \hat{p}) + 3(1 - \hat{p})^2 + \dots + 1(1 - \hat{p})^{L-1} \quad \textcircled{1}$$

$$\text{then } (1 - \hat{p})S = (1 - \hat{p}) + 2(1 - \hat{p})^2 + \dots + 1(1 - \hat{p})^L \quad \textcircled{2}$$

$$\textcircled{1} - \textcircled{2} = \hat{p} S = 1 + (1 - \hat{p}) + \dots + (1 - \hat{p})^{L-1} - 1(1 - \hat{p})^L = \frac{1 - (1 - \hat{p})^L}{\hat{p}} - 1(1 - \hat{p})^L$$

$$\Rightarrow E[N] = \frac{10}{\hat{p}} [1 - (1 - \hat{p})^L] = \frac{10}{1 - \Phi\left(\frac{z_\alpha}{10} - \frac{\mu}{\sqrt{10}}\right)} \cdot \left[1 - \left(\Phi\left(\frac{z_\alpha}{10} - \frac{\mu}{\sqrt{10}}\right)\right)^L\right]$$