1) Sigmoid function derivatives $\sigma(\eta)$

The sigmoid function is written as $\sigma(\eta) = \frac{1}{1+e^{-\eta}} = \frac{e^{\eta}}{1+e^{\eta}}$, where $0 < \sigma(\eta) < 1$. Show that $\frac{d\sigma(\eta)}{d\eta} = \sigma(\eta) \left[1 - \sigma(\eta)\right]$ and $\frac{d\log\sigma(\eta)}{d\eta} = 1 - \sigma(\eta)$.

Solution

$$\begin{split} &\sigma(\eta) = \frac{1}{1 + e^{-\eta}} = \frac{e^{\eta}}{1 + e^{\eta}} \ , \ \ 0 < \sigma(\eta) < 1 \\ &\frac{d\sigma(\eta)}{d\eta} = -\frac{-e^{-n}}{(1 + e^{-n})^2} = \frac{e^{-n}}{(1 + e^{-n})^2} = \frac{1}{1 + e^{-n}} \left(\frac{e^{-n}}{1 + e^{-n}} \right) = \frac{1}{1 + e^{-n}} \left(1 - \frac{1}{1 + e^{-n}} \right) = \sigma(\eta) \left[1 - \sigma(\eta) \right] \\ &\frac{d \log \sigma(\eta)}{d\eta} = \frac{1}{\sigma(\eta)} \cdot \frac{d\sigma(\eta)}{d\eta} = 1 - \sigma(\eta) \end{split}$$

2) Logistic Regression Likelihood & Cross-Entropy

Let $\mathcal{D} = \{(\mathbf{x_1}, y_1), \dots, (\mathbf{x_N}, y_N)\}$, where $\mathbf{x_n} \in \mathbb{R}^D$ and $y_n \in \mathbb{R}$, be the training data of a binary logistic regression model with weights $\mathbf{w} \in \mathbb{R}^D$. The probability of sample $(\mathbf{x_n}, y_n)$ belonging to class 1 is $p(y = 1 | \mathbf{x}, \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x})$, while the probability of belonging to class 0 is $p(y = 0 | \mathbf{x}, \mathbf{w}) = 1 - \sigma(\mathbf{w}^T \mathbf{x})$. Compute the likelihood $\mathcal{L}(\mathcal{D}|\mathbf{w})$ of data \mathcal{D} given the model parameters \mathbf{w} , as well as the cross-entropy error $\mathcal{E}(\mathbf{w}) = -\log \mathcal{L}(\mathcal{D}|\mathbf{w})$.

Solution

Input: $\mathbf{x} \in \mathbb{R}^D$

Output: $y \in \{0, 1\}$

Training data: $\mathcal{D} = \{(\mathbf{x_1}, y_1), \dots, (\mathbf{x_N}, y_N)\}\$

Model:

$$\begin{split} p(y = 1 | \mathbf{x}, \mathbf{w}) &= \sigma(\mathbf{w}^T \mathbf{x}) \\ p(y = 0 | \mathbf{x}, \mathbf{w}) &= 1 - \sigma(\mathbf{w}^T \mathbf{x}), \ \sigma(\eta) = \frac{1}{1 + e^{-\eta}} \\ f(\mathbf{x}) &: \mathbf{x} \to y, \ f(\mathbf{x}) = \left\{ \begin{array}{l} 1, & p(y = 1 | \mathbf{x}, \mathbf{w}) > 0.5 \\ 0, & \text{otherwise} \end{array} \right. = \left. \left\{ \begin{array}{l} 1, & \sigma(\mathbf{w}^T \mathbf{x}) > 0.5 \\ 0, & \text{otherwise} \end{array} \right. \end{split}$$

Model parameters: Weights $\mathbf{w} \in \mathbb{R}^D$ (to be learned)

Data likelihood for 1 training sample:

$$p(y_n|\mathbf{x_n}, \mathbf{w}) = \left\{ \begin{array}{ll} \sigma(\mathbf{w}^T \mathbf{x_n}), & y_n = 1 \\ 1 - \sigma(\mathbf{w}^T \mathbf{x_n}), & y_n = 0 \end{array} \right\} = \left[\sigma(\mathbf{w}^T \mathbf{x_n}) \right]^{y_n} \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right]^{1 - y_n}$$

Data likelihood for all training data:

$$L(\mathcal{D}|\mathbf{w}) = \prod_{n=1}^{N} p(y_n|\mathbf{x_n}, \mathbf{w}) = \prod_{n=1}^{N} \left[\sigma(\mathbf{w}^T \mathbf{x_n}) \right]^{y_n} \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right]^{1 - y_n}$$

Log-likelihood for all training data:

$$l(\mathcal{D}|\mathbf{w}) = \sum_{n=1}^{N} \left\{ y_n \log \left[\sigma(\mathbf{w}^T \mathbf{x_n}) \right] + (1 - y_n) \log \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right] \right\}$$

Cross-entropy error (negative log-likelihood):

$$\mathcal{E}(\mathbf{w}) = -\sum_{n=1}^{N} \left\{ y_n \log \left[\sigma(\mathbf{w}^T \mathbf{x_n}) \right] + (1 - y_n) \log \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right] \right\}$$

Logistic Regression - Optimization

3a) Show that the first order derivative (i.e., gradient vector) of the cross-entropy function is $\nabla \mathcal{E}(\mathbf{w}) = \frac{\vartheta \mathcal{E}(\mathbf{w})}{\vartheta \mathbf{w}} = \sum_{n=1}^{N} \underbrace{\left(\sigma(\mathbf{w}^T \mathbf{x_n}) - y_n\right)}_{\mathbf{w}} \mathbf{x_n}$

Solution

$$\nabla \mathcal{E}(\mathbf{w}) = -\sum_{n=1}^{N} \left\{ y_n \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right] \mathbf{x_n} - (1 - y_n) \left[1 - \left(1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right) \right] \mathbf{x_n} \right\}$$

$$= -\sum_{n=1}^{N} \left\{ y_n \left[1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right] \mathbf{x_n} - (1 - y_n) \sigma(\mathbf{w}^T \mathbf{x_n}) \mathbf{x_n} \right\}$$

$$= -\sum_{n=1}^{N} \left[y_n - y_n \sigma(\mathbf{w}^T \mathbf{x_n}) - \sigma(\mathbf{w}^T \mathbf{x_n}) + y_n \sigma(\mathbf{w}^T \mathbf{x_n}) \right] \mathbf{x_n}$$

$$= \sum_{n=1}^{N} \underbrace{\left(\sigma(\mathbf{w}^T \mathbf{x_n}) - y_n \right)}_{\text{over a graph}} \mathbf{x_n}$$

No closed-form solution that minimizes the cross-entropy function.

We use an approximate method, e.g. gradient descent, so we need to compute $\nabla \mathcal{E}(\mathbf{w})$. Gradient descent update: $\mathbf{w_{k+1}} := \mathbf{w_k} - \alpha(k) \nabla \mathcal{E}(\mathbf{w})$

3b) Show that the Hessian of the cross-entropy function is $\mathbf{H} = \frac{\vartheta^2 \mathcal{E}(\mathbf{w})}{\vartheta^2 \mathbf{w}} = \nabla \left((\nabla \mathcal{E}(\mathbf{w}))^T \right) = \sum_{n=1}^N \sigma(\mathbf{w}^T \mathbf{x_n}) \cdot \left(1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right) \cdot \left(\mathbf{x_n} \cdot \mathbf{x_n}^T \right)$ and show that it is positive semi-definite.

$$\frac{\text{Solution}}{\mathbf{H}} = \frac{\vartheta^2 \mathcal{E}(\mathbf{w})}{\vartheta^2 \mathbf{w}} = \nabla \left((\nabla \mathcal{E}(\mathbf{w}))^T \right) = \nabla \left(\sum_{n=1}^N \left(\sigma(\mathbf{w}^T \mathbf{x_n}) - y_n \right) \mathbf{x_n}^T \right) \\
\mathbf{H} = \frac{\vartheta}{\vartheta \mathbf{w}} \left[\sum_{n=1}^N \left(\sigma(\mathbf{w}^T \mathbf{x_n}) \cdot \mathbf{x_n}^T - y_n \mathbf{x_n}^T \right) \right] \\
= \sum_{n=1}^N \frac{\vartheta}{\vartheta \mathbf{w}} \left[\sigma(\mathbf{w}^T \mathbf{x_n}) \right] \cdot \mathbf{x_n}^T \quad \text{(chain rule)} \\
= \sum_{n=1}^N \underbrace{\sigma(\mathbf{w}^T \mathbf{x_n})}_{\in [0,1]} \cdot \underbrace{\left(1 - \sigma(\mathbf{w}^T \mathbf{x_n}) \right)}_{\in [0,1]} \cdot \underbrace{\left(\mathbf{x_n} \cdot \mathbf{x_n}^T \right)}_{\in \mathcal{R}^{D \times D}}$$

For all $\mathbf{v} \in \mathbb{R}^D$, substituting $\mu_n = \sigma(\mathbf{w}^T \mathbf{x_n}) \left(1 - \sigma(\mathbf{w}^T \mathbf{x_n})\right) \ge 0$, we have:

$$\mathbf{v}^T \mathbf{H} \mathbf{v} = \mathbf{v}^T \left(\sum_{n=1}^N \mu_n \mathbf{x_n} \mathbf{x_n}^T \right) \mathbf{v} = \sum_{n=1}^N (\mu_n \mathbf{x_n}^T \mathbf{v})^T (\mathbf{x_n}^T \mathbf{v}) = \sum_{n=1}^N \mu_n \|\mathbf{x_n}^T \mathbf{v}\|_2^2 \ge 0$$