## Formula Sheet LEC0101/0201 Notation

| Simple Linear Regression  | Multiple Linear Regression  |
|---|---|
| $y_i = \beta_0 + \beta_1 x_i + \epsilon_i, i = 1, \dots, n$   | $\mathbf{Y} = \mathbf{X}oldsymbol{eta} + oldsymbol{\epsilon}$   |
| $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$   | $\hat{oldsymbol{eta}} = (\mathbf{X^TX})^{-1}\mathbf{X}^T\mathbf{Y}$   |
| $\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n x_i y_i - n\bar{x}\bar{y}}{\sum_{i=1}^n x_i^2 - n\bar{x}^2}$ |   |
| $s^{2} = \hat{\sigma}^{2} = \frac{\sum_{i=1}^{n} \hat{e}_{i}^{2}}{n-2} = \frac{\sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}}{n-2}$  | $s^2 = \hat{\sigma}^2 = \frac{\hat{\mathbf{e}}^T \hat{\mathbf{e}}}{n - p - 1}$  |
| $Var(\hat{\beta}_0) = \sigma^2 \left( \frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right)$   | $Cov(\hat{\boldsymbol{\beta}}) = \sigma^2(\mathbf{X}^T\mathbf{X})^{-1}$   |
| $Var(\hat{\beta}_1) = \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}$  |   |
| $Var(\hat{y}_0 \mid x_0, X) = \sigma^2 \left( \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right)$   | $Var(\hat{y}_0 \mid \mathbf{X}, \mathbf{x}_0) = \sigma^2 \mathbf{x}_0^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_0$                              |
| $\hat{\beta}_j \pm t_{\frac{\alpha}{2}, n-2} \sqrt{\widehat{Var}(\hat{\beta}_j)}$   | $\hat{\boldsymbol{\beta}}_j \pm t_{\frac{\alpha}{2}, n-p-1} s \sqrt{(\mathbf{X}^T \mathbf{X})_{(j+1, j+1)}^{-1}}$                                 |
| $\hat{\beta}_0 + \hat{\beta}_1 x_0 \pm t_{\frac{\alpha}{2}, n-2} \sqrt{\widehat{Var}(\hat{y}_0)}$   | $\mathbf{x}_0^T \hat{\boldsymbol{\beta}} \pm t_{\frac{\alpha}{2}, n-p-1} s \sqrt{\mathbf{x}_0^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_0}$     |
| $\hat{\beta}_0 + \hat{\beta}_1 x_0 \pm t_{\frac{\alpha}{2}, n-2} s \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$                                 | $\mathbf{x}_0^T \hat{\boldsymbol{\beta}} \pm t_{\frac{\alpha}{2}, n-p-1} s \sqrt{1 + \mathbf{x}_0^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_0}$ |
| $T = \frac{\hat{\beta}_j - \beta_j^0}{\sqrt{\widehat{Var}(\hat{\beta}_j)}} \sim T_{n-2}$  | $T = \frac{\hat{\beta}_j - \beta_j^0}{s\sqrt{(\mathbf{X}^T \mathbf{X})_{(j+1,j+1)}^{-1}}} \sim T_{n-p-1}$   |

$$\mathbf{X}^{\mathbf{T}}\mathbf{X} = \begin{pmatrix} n & \sum x_{i1} & \sum x_{i2} & \dots & \sum x_{ip} \\ \sum x_{i1} & \sum x_{i1}^{2} & \sum x_{i1}x_{i2} & \dots & \sum x_{i1}x_{ip} \\ \sum x_{i2} & \sum x_{i1}x_{i2} & \sum x_{i2}^{2} & \dots & \sum x_{i2}x_{ip} \\ \vdots & \vdots & \vdots & & \vdots \\ \sum x_{ip} & \sum x_{i1}x_{ip} & \sum x_{i2}x_{ip} & \dots & \sum x_{ip}^{2} \end{pmatrix}$$

$$RSS = \sum_{i=1}^{n} (y_{i} - \hat{E}(Y_{i} \mid \mathbf{X}))^{2}$$

## Formula Sheet LEC5101 Notation

| Simple Linear Regression  | Multiple Linear Regression  |
|---|---|
| $Y_i = \beta_0 + \beta_1 x_i + e_i, i = 1, \dots, n$  | $\mathbf{Y} = \mathbf{X}oldsymbol{eta} + e$   |
| $\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{x}$ $\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(Y_i - \bar{Y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n x_i Y_i - n\bar{x}\bar{Y}}{\sum_{i=1}^n x_i^2 - n\bar{x}^2}$ | $\hat{oldsymbol{eta}} = (\mathbf{X^TX})^{-1}\mathbf{X}^T\mathbf{Y}$   |
| $\hat{\sigma}^{1} - \sum_{i=1}^{n} (x_{i} - \bar{x})^{2} - \sum_{i=1}^{n} x_{i}^{2} - n\bar{x}^{2}$ $\hat{\sigma}^{2} = \frac{\sum_{i=1}^{n} \hat{e}_{i}^{2}}{n-2} = \frac{\sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})^{2}}{n-2}$              | $\hat{\sigma}^2 = \frac{\hat{\mathbf{e}}^T \hat{\mathbf{e}}}{n-p-1}$  |
| $Var(\hat{\beta}_0) = \sigma^2 \left( \frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right)$   | $Cov(\hat{\boldsymbol{eta}}) = \sigma^2(\mathbf{X}^T\mathbf{X})^{-1}$   |
| $Var(\hat{\beta}_{1}) = \frac{\sigma^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}}$ $Var(\hat{Y}^{*} \mid x^{*}, X) = \sigma^{2} \left( \frac{1}{n} + \frac{(x^{*} - \bar{x})^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} \right)$          | $Var(\hat{Y}^* \mid \mathbf{X}, \mathbf{x}^*) = \sigma^2 \mathbf{x}^{*T} (\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}^*$                                     |
| $\hat{\beta}_j \pm t_{1-\frac{\alpha}{2},n-2} \hat{se}(\hat{\beta}_j)$  | $\hat{\beta}_{j} \pm t_{1-\frac{\alpha}{2},n-p-1} \hat{\sigma} \sqrt{(\mathbf{X}^{T}\mathbf{X})_{(j+1,j+1)}^{-1}}$                                      |
| $\hat{\beta}_0 + \hat{\beta}_1 x^* \pm t_{1-\frac{\alpha}{2}, n-2} \hat{se}(\hat{Y}^*)$   | $\mathbf{x}^{*T}\hat{\boldsymbol{\beta}} \pm t_{1-\frac{\alpha}{2},n-p-1}\hat{\sigma}\sqrt{\mathbf{x}^{*T}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{x}^*}$   |
| $\hat{\beta}_0 + \hat{\beta}_1 x^* \pm t_{1-\frac{\alpha}{2},n-2} \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$   | $\mathbf{x}^{*T}\hat{\boldsymbol{\beta}} \pm t_{1-\frac{\alpha}{2},n-p-1}\hat{\sigma}\sqrt{1+\mathbf{x}^{*T}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{x}^*}$ |
| $\frac{\hat{\beta}_j - \beta_j^0}{\hat{se}(\hat{\beta}_j)} \sim t_{n-2}$  | $\frac{\hat{\beta}_{j} - \beta_{j}^{0}}{\hat{\sigma}\sqrt{(\mathbf{X}^{T}\mathbf{X})_{(j+1,j+1)}^{-1}}} \sim t_{n-p-1}$                                 |

$$\mathbf{X}^{\mathbf{T}}\mathbf{X} = \begin{pmatrix} n & \sum x_{i1} & \sum x_{i2} & \dots & \sum x_{ip} \\ \sum x_{i1} & \sum x_{i1}^{2} & \sum x_{i1}x_{i2} & \dots & \sum x_{i1}x_{ip} \\ \sum x_{i2} & \sum x_{i1}x_{i2} & \sum x_{i2}^{2} & \dots & \sum x_{i2}x_{ip} \\ \vdots & \vdots & & \vdots & & \vdots \\ \sum x_{ip} & \sum x_{i1}x_{ip} & \sum x_{i2}x_{ip} & \dots & \sum x_{ip}^{2} \end{pmatrix}$$

$$RSS = \sum_{i=1}^{n} (Y_{i} - \hat{E}(Y_{i} \mid \mathbf{X}))^{2}$$