

# Generalizable Recommendation to a Target Population by Leveraging Randomized and Observational Studies

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Lili Wu & Shu Yang

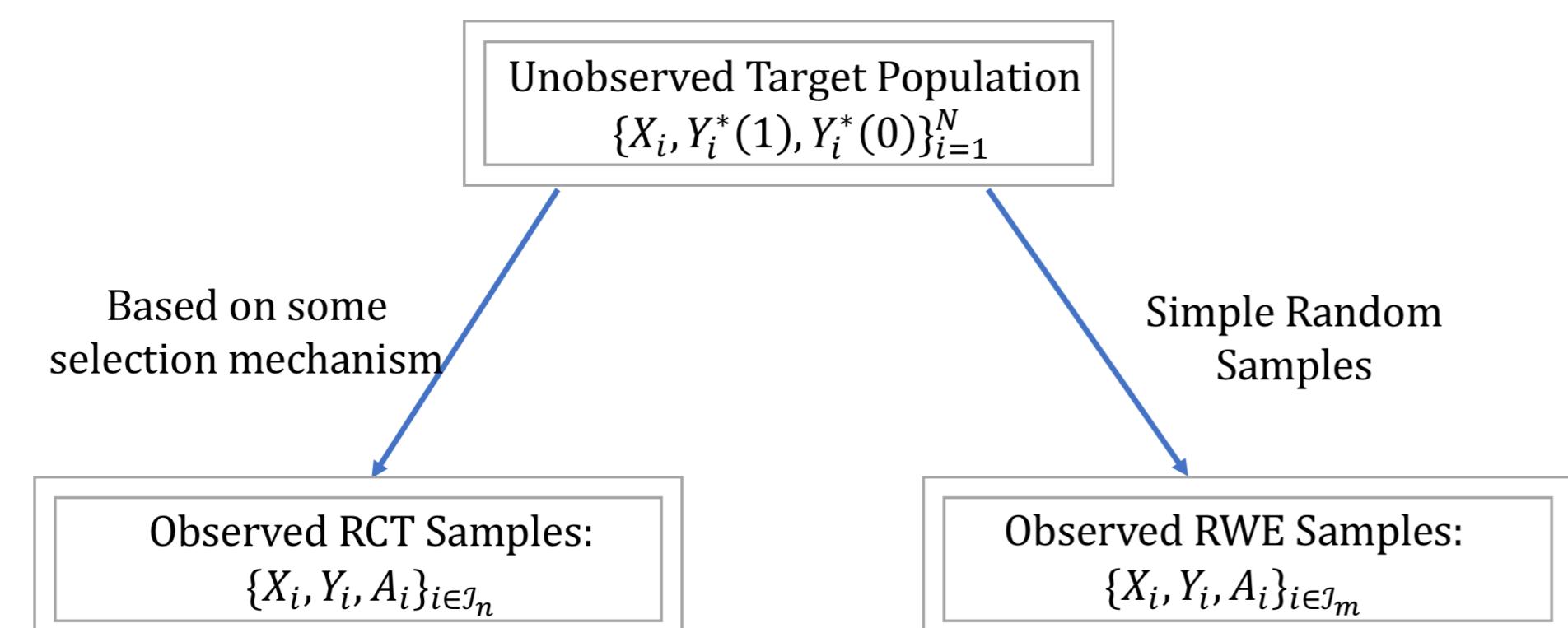
Department of Statistics, North Carolina State University

{lwu9, syang24}@ncsu.edu

## Abstract

- Personalized medicine focuses on the recommendation of treatment to patients based on their individual characteristics. The golden standard is to use randomized clinical trials (RCT).
- Comparing with the complex non-linear decision rules, clinicians are interested in parsimonious decision rules (e.g., linear or decision tree rules) which are more interpretable.
- This work is to give generalizable recommendations to the target population by use the information from RCT and real world evidence (RWE).

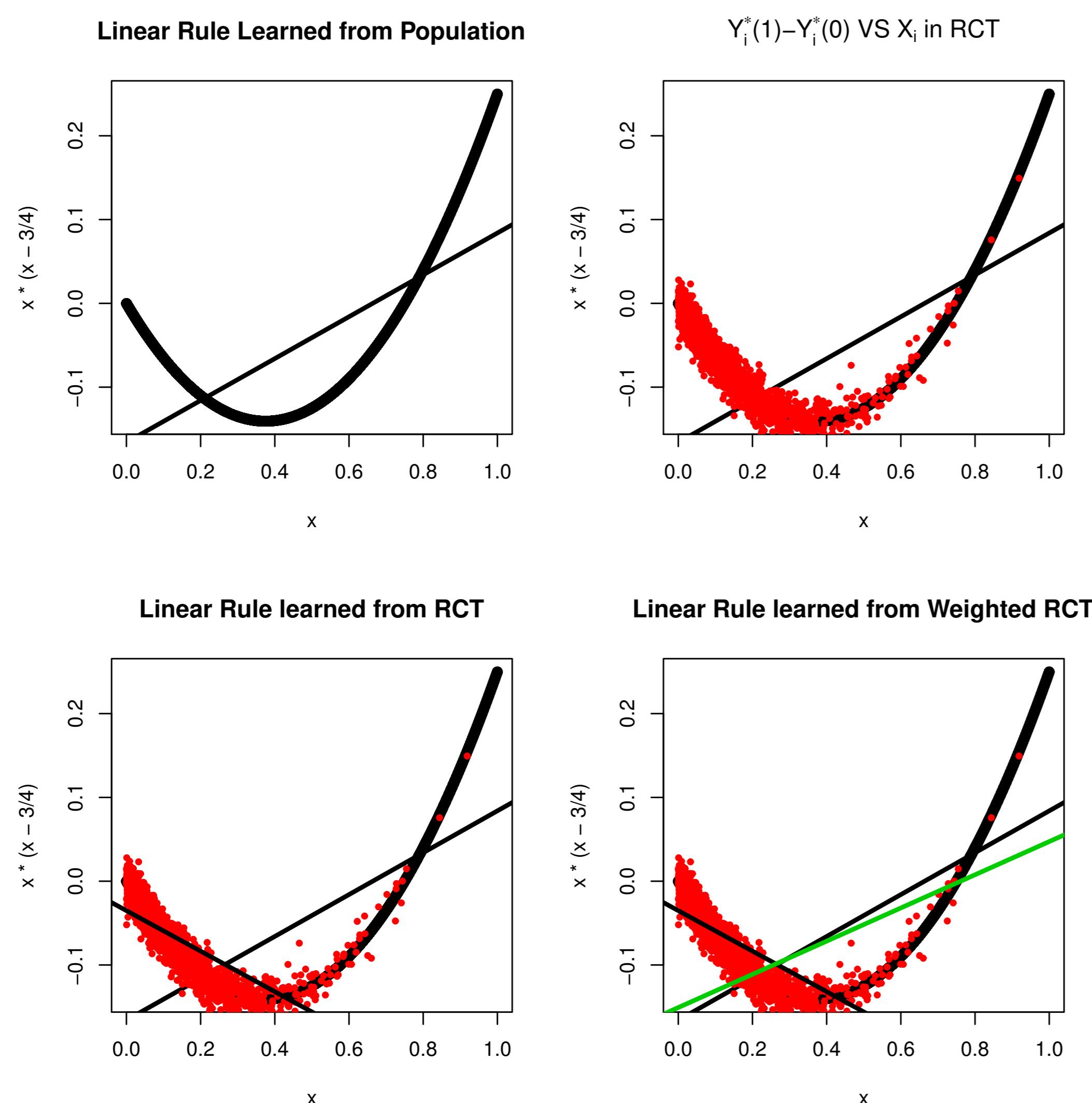
## Data Sources Illustration



## Issues

- There may exist unmeasured confounders affecting treatment and outcomes in RWE.
- There is selection bias in RCT (i.e., people's characteristics distribution in RCT may be very different from the target population), can a parsimonious optimal rule derived from it really be the optimal rule to the target population?

## Selection Bias in RCT



## Main Objectives

- Estimate transferring weights to correct covariate distribution in RCT.
- Use the weights obtained above to learn interpretable decision rules which have better generalizability to the target population.

## Methods

- Notations: treatment  $A \in \mathcal{A} = \{0, 1\}$ , covariates  $X \in \mathbf{R}^p$ , potential outcomes  $Y^*(0), Y^*(1) \in \mathbf{R}$ , target population size  $N$ , index set of RCT samples  $\mathcal{I}_n$  with size  $n$ , index set of RWE samples  $\mathcal{I}_m$  with size  $m$ .  $\delta_i = 1$  means the  $i$ th individual is in RWE, and  $\tilde{\delta}_i = 1$  means the  $i$ th individual is in RCT.

- Assume the covariate distribution in RWE is the same as that in the target population. Thus  $\frac{1}{m} \sum_{i \in \mathcal{I}_m} X_i$  would be an unbiased estimator of  $E(X)$ , the covariate expectation in the target population.
- Note that  $E\left\{\frac{\tilde{\delta}_i X_i}{P(\tilde{\delta}_i = 1|X_i)}\right\} = E\left\{E\left\{\frac{\tilde{\delta}_i X_i}{P(\tilde{\delta}_i = 1|X_i)} \mid X_i\right\}\right\} = E(X_i)$ ,
- Goal: to find transferring weights  $\tilde{w}_i := \frac{1}{P(\tilde{\delta}_i = 1|X_i)}$ .

## Two Approaches to estimate Weights

- Method of propensity score: Assume  $\pi(X_i; \alpha) := P(\tilde{\delta}_i = 1|X_i; \alpha) = \frac{\exp(\alpha^\top X_i)}{1 + \exp(\alpha^\top X_i)}$ , where  $\alpha \in \mathbf{R}^p$ , so we can get the maximum likelihood estimator

$$\begin{aligned} \hat{\alpha} &= \underset{\alpha}{\operatorname{argmax}} \sum_{i=1}^N \{\tilde{\delta}_i \log \pi(x_i; \alpha) + (1 - \tilde{\delta}_i) \log(1 - \pi(x_i; \alpha))\} \\ &= \underset{\alpha}{\operatorname{argmax}} \sum_{i=1}^N \{\tilde{\delta}_i (\alpha^\top x_i) - \log(1 + \exp(\alpha^\top x_i))\} \\ &\approx \underset{\alpha}{\operatorname{argmax}} \left\{ \sum_{i=1}^N \tilde{\delta}_i (\alpha^\top x_i) - \sum_{i \in \mathcal{I}_m} \frac{N}{m} \log(1 + \exp(\alpha^\top x_i)) \right\}. \end{aligned}$$

- Method of estimating equations:

$$\frac{1}{N} \sum_{i=1}^N \frac{\tilde{\delta}_i X_i}{P(\tilde{\delta}_i = 1|X_i; \alpha)} = \frac{1}{m} \sum_{i \in \mathcal{I}_m} X_i,$$

where  $P(\tilde{\delta}_i = 1|X_i; \alpha)$  is modeled same as in the above (1). So we can solve  $p$  equations to get  $\hat{\alpha}$ .

## Decision Rules Larning with the Weights

- Linear Decision Rules:

- Target population parameter:  $\beta_N^* = \arg \min_{\beta} \sum_{i=1}^N \{Y_i^*(1) - Y_i^*(0) - X_i^\top \beta\}^2$
- Add  $\tilde{w}_i$  into least square loss function:

$$\hat{\beta}^* = \arg \min_{\beta} \sum_{i \in \mathcal{I}_n} \tilde{w}_i \{\hat{Y}_i^* - X_i^\top \beta\}^2,$$

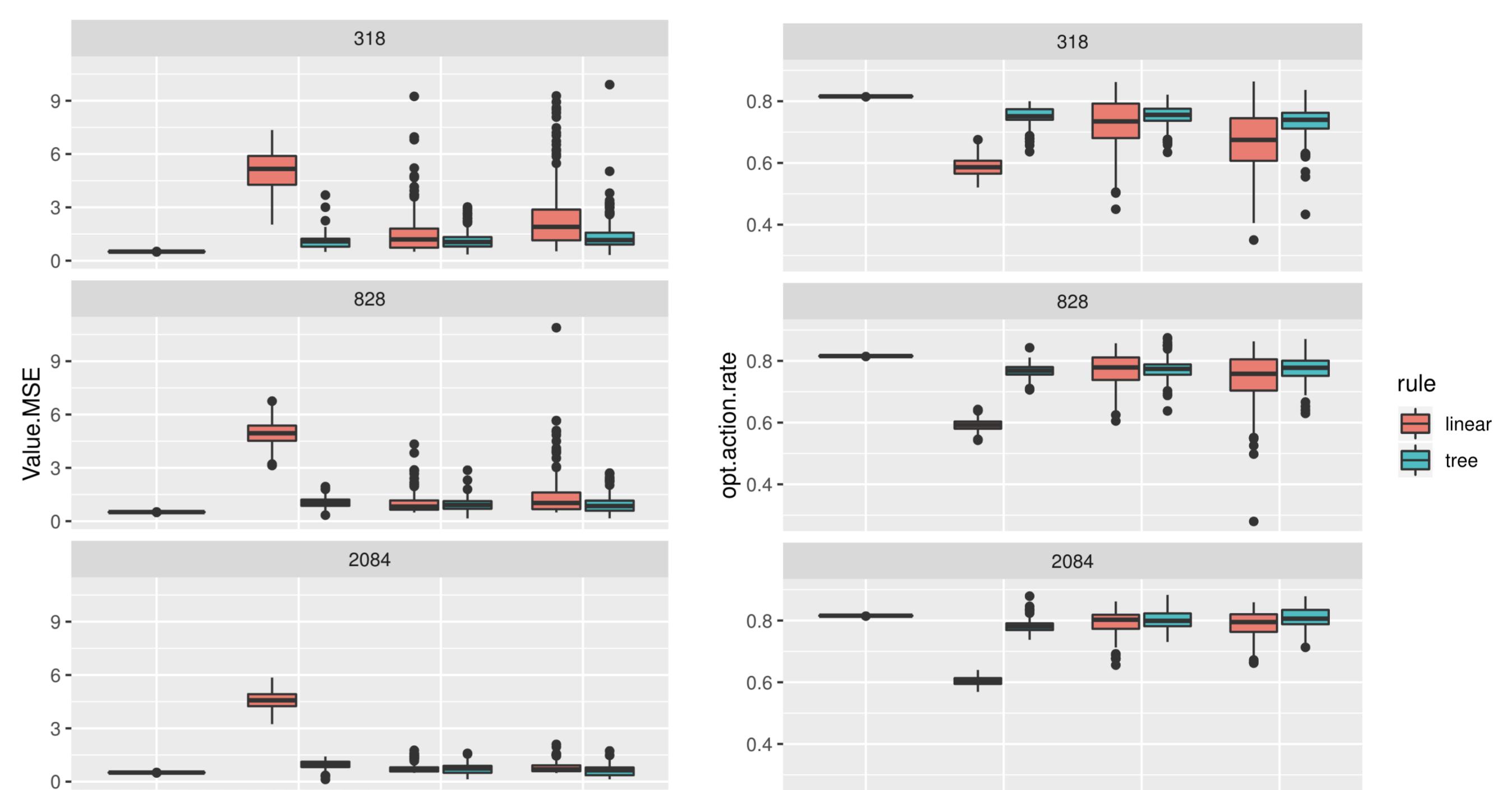
$$\text{where } \hat{Y}_i^* = \frac{A_i Y_i}{P(A_i = 1|X_i)} - \frac{(1-A_i) Y_i}{1 - P(A_i = 1|X_i)}.$$

- Decision Tree Rule:

- Define the policy  $D : \text{dom } X \rightarrow \{0, 1\}$  and the value of the policy  $D$  as  $V(D) = E_D(Y) = E[Y^* \{D(X)\}]$  over the target population. To maximize  $V(D)$  is equivalent to  $\min_{D} \sum_{i \in \mathcal{I}_n} \frac{|\tilde{R}_i|}{P(A_i, X_i)} I[\{(2A_i - 1)\text{sign}(\tilde{R}_i) + 1\}/2 \neq D(X_i)]$  [Wang et al., 2016], where  $\tilde{R}_i = Y_i - \hat{R}_i$ , and the predicted  $\hat{R}_i$  is fitted by outcomes  $Y_i$  only on  $X_i$ , and  $\pi(A_i, X_i)$  is the probability of taking action  $A_i$ .
- Add  $\tilde{w}_i$  into the loss function:  $\min_{D} \sum_{i \in \mathcal{I}_n} \frac{\tilde{w}_i |\tilde{R}_i|}{\pi(A_i, X_i)} I[\{(2A_i - 1)\text{sign}(\tilde{R}_i) + 1\}/2 \neq D(X_i)]$ .

## Simulation Results

- Generative model:  $Y_i = 1 + X_{i1} + A_i \tilde{X}_i + \epsilon_i$ , where  $N = 10^6$ ,  $\tilde{X}_i = (\cos(1) - \frac{1}{2}, \cos(X_{i1}) - \frac{1}{2}, \cos(X_{i2}) - \frac{1}{2}, \cos(X_{i3}) - \frac{1}{2})^\top$ ,  $X_i = (X_{i1}, X_{i2}, X_{i3})^\top \sim N(\mathbf{1}, I_3)$ ,  $A_i \sim \text{Bernoulli}(0.5)$ ,  $\beta = (1, 2, 3, 4)$ ,  $\epsilon_i \sim N(0, 0.5^2)$ ,  $i = 1, \dots, N$  independently. Let  $\alpha = (\alpha_1, 1, -2, 1)$ .



## References

- [Wang et al., 2016] Wang, Y., Wu, P., Liu, Y., Weng, C., and Zeng, D. (2016). Learning optimal individualized treatment rules from electronic health record data. In *2016 IEEE International Conference on Healthcare Informatics (ICHI)*, pages 65–71. IEEE.