

MA215 Probability Theory

Assignment 14

1. Suppose that the m.g.f. of X is $M_X(t) = \frac{2}{\sqrt{4-t}}$, ($t < 4$).

(i) Find $E(X)$, $E(X^2)$ and $\text{Var}(X)$.

(ii) Suppose that X and Y are independent and both with this m.g.f. (i.e., $M_X(t) = M_Y(t) = \frac{2}{\sqrt{4-t}}$). Find the m.g.f. of $X + Y$ and identify the distribution of $X + Y$.

$$(i) \quad E(X) = \left. \frac{d}{dt} E(e^{tx}) \right|_{t=0} = \left. \frac{d}{dt} M_X(t) \right|_{t=0} = \left. \frac{d}{dt} \frac{2}{\sqrt{4-t}} \right|_{t=0} = \left. \frac{1}{(4-t)^{\frac{3}{2}}} \right|_{t=0} = \frac{1}{8}$$

$$E(X^2) = \left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} = \left. \frac{d}{dt} \frac{1}{(4-t)^{\frac{3}{2}}} \right|_{t=0} = \left. \frac{3}{2} (4-t)^{-\frac{5}{2}} \right|_{t=0} = \frac{3}{64}$$

$$\text{Var}(X) = E(X^2) - (E(X))^2 = \frac{3}{64} - \frac{1}{64} = \frac{1}{32}$$

$$(ii) \quad \because X, Y \text{ are independent} \quad \therefore M_{X+Y}(t) = M_X(t) \cdot M_Y(t) = \frac{4}{4-t}$$

$\therefore X+Y$ follows Gamma distribution, and $\lambda = \frac{1}{2}$, $k=2$

$$\therefore X+Y \sim \text{Gamma}(2, \frac{1}{2})$$

2. (a) If the m.g.f. of X is $M_X(t) = \frac{a^2}{a^2-t^2}$, then find the k th moment $E(X^k)$, $k \in \mathbb{N}_+$.

(b) Suppose the m.g.f. of X can be expressed as a power series

$$M_X(t) = \sum_{k=0}^{\infty} a_k t^k = a_0 + a_1 t + a_2 t^2 + a_3 t^3 + \dots$$

and assume further that $a_0 = 1$, $a_1 = 3$ and $a_2 = 7$. Find $E(X)$ and $\text{Var}(X)$.

$$(a) \quad E(X^k) = \left. \frac{d^k}{dt^k} M_X(t) \right|_{t=0} = \left. \frac{d^k}{dt^k} \frac{a^2}{a^2-t^2} \right|_{t=0} = 0$$

$$(b) \quad E(X) = \left. \frac{d}{dt} M_X(t) \right|_{t=0} = a_1 = 3$$

$$E(X^2) = \left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} = 2a_2 = 14$$

$$\text{Var}(X) = E(X^2) - (E(X))^2 = 5$$

3. Suppose the m.g.f. of X has the Maclaurin series

$$M_X(t) = 1 + a_1 t + a_2 t^2 + a_3 t^3 + \dots$$

Find the variance and the third central moment $E[(X - E(X))^3]$ of X in terms of a_1 , a_2 , and a_3 .

$$E[(X - E(X))^3] = E(X^3 - 3X^2 E(X) + 3X E(X)^2 - E(X)^3) = E(X^3) - 3E(X^2)E(X) + 2E(X)^3$$

$$E(X) = \left. \frac{d}{dt} M_X(t) \right|_{t=0} = a_1 \quad E(X^2) = \left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} = 2a_2$$

$$E(X^3) = \left. \frac{d^3}{dt^3} M_X(t) \right|_{t=0} = 6a_3$$

$$\therefore E[(X - E(X))^3] = 6a_3 - 6a_1 a_2 + 2a_1^3$$

4. Let X_1, X_2, \dots, X_n be i.i.d., each having the normal distribution with parameters μ and σ^2 .

(i) Find the m.g.f.s of the sample sum $S_n = \sum_{i=1}^n X_i$ and sample average $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$.

(ii) What are the distributions of these two random variables?

(i) We already know that $X_1, \dots, X_n \sim N(\mu, \sigma^2)$, then

$$M_{X_1}(t) = \dots = M_{X_n}(t) = E(e^{tx_i}) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$$

$\therefore X_1, \dots, X_n$ are i.i.d.

$$\therefore M_{S_n}(t) = \prod_{i=1}^n M_{X_i}(t) = (e^{\mu t + \frac{1}{2}\sigma^2 t^2})^n = e^{\mu n t + \frac{1}{2}\sigma^2 n t^2}$$

$$M_{\bar{X}_n}(t) = E(e^{t \frac{S_n}{n}}) = M_{S_n}(\frac{t}{n}) = e^{\mu n \cdot \frac{t}{n} + \frac{1}{2}\sigma^2 n \cdot (\frac{t}{n})^2} = e^{\mu t + \frac{1}{2n}\sigma^2 t^2}$$

(ii) S_n, \bar{X}_n both follow the normal distribution

$$S_n \sim N(\mu n, n\sigma^2), \quad \bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$$

5. Suppose X is a discrete random variable taking values of non-negative integers (or subset of non-negative integers) with p.m.f $\{p_k; k \geq 0\}$. Define the probability generating function (p.g.f.) of X , denoted by $\Pi_X(t)$, as

$$\Pi_X(t) = E(t^X).$$

- Write down the form $\Pi_X(t)$ in terms of the p.m.f. $\{p_k; k \geq 0\}$.
- Investigate the problem as how to get $E(X)$ and $\text{Var}(X)$ by using $\Pi_X(t)$.
- Find the p.g.f. of the Binomial Random Variable X with parameter n and p .
- Find the p.g.f. of the Poisson Random Variable X with parameter λ .

$$(i) \quad \Pi_X(t) = E(t^X) = \sum_{k=0}^{\infty} p_k t^k$$

$$(ii) \quad \because \frac{d}{dt} t^X = X t^{X-1} \quad \therefore E(X) = \left. \frac{d}{dt} \Pi_X(t) \right|_{t=1} = \sum_{k=0}^{\infty} k p_k \cdot t^{k-1} \Big|_{t=1} = \sum_{k=0}^{\infty} k p_k = \Pi_X'(1)$$

$$E(X^2 - X) = \left. \frac{d^2}{dt^2} \Pi_X(t) \right|_{t=1} = \sum_{k=0}^{\infty} k(k-1) p_k \cdot t^{k-2} \Big|_{t=1} = \sum_{k=0}^{\infty} k(k-1) p_k = \Pi_X''(1)$$

$$\therefore E(X^2) = E(X^2 - X) + E(X) = \sum_{k=0}^{\infty} k^2 p_k$$

$$\text{Var}(X) = E(X^2) - (E(X))^2 = \sum_{k=0}^{\infty} k^2 p_k - \left(\sum_{k=0}^{\infty} k p_k \right)^2 = \Pi_X''(1) - (\Pi_X'(1))^2$$

$$(iii) \quad \text{p.m.f: } p_k = C_n^k p^k (1-p)^{n-k}, \quad k = 0, 1, 2, \dots, n$$

$$\text{p.g.f: } \Pi_X(t) = E(t^X) = \sum_{k=0}^n C_n^k p^k (1-p)^{n-k} t^k = (pt + (1-p))^n$$

$$(iv) \quad \text{p.m.f: } p_k = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, \dots$$

$$\text{p.g.f: } \Pi_X(t) = \sum_{k=0}^{\infty} p_k t^k = \sum_{k=0}^{\infty} e^{-\lambda} \cdot \frac{(\lambda t)^k}{k!} = e^{-\lambda(1-t)} \cdot \sum_{k=0}^{\infty} e^{-\lambda t} \cdot \frac{(\lambda t)^k}{k!} = e^{-\lambda(1-t)}$$