

MA215 Probability Theory

Assignment 10

1. Suppose a player plays the following gambling games which is known as the wheel of fortune. The player bets on one of the numbers 1 through 6. Three dice are then rolled, and if the number bet by the player appears i times, $i = 1, 2, 3$; then the player wins i units; on the other hand, if the number bet by the player does not appear on any of the dies, then the player loses 1 unit. Is this game fair to the player?

Suppose that the player get X units

$$X \in \{-1, 1, 2, 3\}$$

$$P(X=3) = \left(\frac{1}{6}\right)^3 = \frac{1}{216} \quad P(X=2) = C_3^2 \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right) = \frac{15}{216}$$

$$P(X=1) = C_3^1 \frac{1}{6} \left(\frac{5}{6}\right)^2 = \frac{75}{216} \quad P(X=-1) = \left(\frac{5}{6}\right)^3 = \frac{125}{216}$$

$$\therefore \begin{array}{c|cccc} X & -1 & 1 & 2 & 3 \\ \hline P & \frac{125}{216} & \frac{75}{216} & \frac{15}{216} & \frac{1}{216} \end{array}$$

$$\therefore E(X) = -1 \times \frac{125}{216} + 1 \times \frac{75}{216} + 2 \times \frac{15}{216} + 3 \times \frac{1}{216} = -\frac{17}{216} < 0$$

\therefore The game is not fair.

2. Suppose the r.v. X takes non-negative integer values only. Show that

$$E(X) = \sum_{n=0}^{\infty} P(X > n) = \sum_{n=1}^{\infty} P(X \geq n).$$

Proof. $\because X \in \mathbb{N}$.

$$\begin{aligned} \therefore \sum_{n=0}^{\infty} P(X > n) &= (P(X=1) + P(X=2) + \cdots + P(X=n) + \cdots) + (P(X=2) + P(X=3) + \cdots + P(X=n) + \cdots) \\ &\quad + \cdots + (P(X=n) + P(X=n+1) + \cdots) = P(X=1) + 2P(X=2) + \cdots + nP(X=n) + \cdots \\ &= \sum_{n=1}^{\infty} nP(X=n) = E(X) \end{aligned}$$

$$\begin{aligned} \text{同理, } \sum_{n=1}^{\infty} P(X \geq n) &= P(X \geq 1) + \cdots + P(X \geq n) + \cdots = (P(X=1) + P(X=2) + \cdots + P(X=n) + \cdots) + \\ &\quad (P(X=2) + P(X=3) + \cdots + P(X=n) + \cdots) + \cdots + (P(X=n) + P(X=n+1) + \cdots) \\ &= \sum_{n=1}^{\infty} nP(X=n) = E(X). \quad \square \end{aligned}$$

3. (a) Suppose the r.v. X obeys the uniformly distribution over $[a, b]$. Find $E(X)$.
 (b) Suppose the r.v. X obeys the general Γ distribution with parameters λ and α where $\lambda > 0, \alpha > 0$. Write down the p.d.f. of this general Γ random variable and the analytic form of the Γ function $\Gamma(\alpha)$ for $\alpha > 0$ and hence find the $E(X)$ of this general Γ random variable.
 (c) Suppose $Y = X^2$ where X is normally distributed with parameters μ and σ^2 . Obtain the p.d.f. of Y and then find $E(Y)$.

$$(a) \text{ p.d.f. } f(x) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{otherwise} \end{cases}$$

$$\therefore E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_a^b \frac{x}{b-a} dx = \frac{1}{2} \frac{x^2}{(b-a)} \Big|_a^b = \frac{a+b}{2}$$

$$(b) \text{ p.d.f. } f(x) = \begin{cases} \frac{\lambda^\alpha x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$$\Gamma(\alpha) = \int_0^\infty t^{\alpha-1} e^{-t} dt$$

$$\begin{aligned} E(X) &= \int_{-\infty}^{+\infty} x f(x) dx = \int_0^\infty \frac{\lambda^\alpha x^\alpha e^{-\lambda x}}{\Gamma(\alpha)} dx = \int_0^\infty \frac{(\lambda x)^\alpha e^{-\lambda x}}{\lambda \Gamma(\alpha)} d(\lambda x) \\ &= \frac{\Gamma(\alpha+1)}{\lambda \Gamma(\alpha)} = \frac{\alpha \Gamma(\alpha)}{\lambda \Gamma(\alpha)} = \frac{\alpha}{\lambda} \end{aligned}$$

$$(c) X \sim N(\mu, \sigma^2) \therefore \text{p.d.f. } f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < +\infty$$

$$\therefore Y = X^2 \therefore X = \ln Y, f_Y(y) = \left| \frac{1}{y} \right| \cdot f_X(\ln y) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma y} \cdot e^{-\frac{(\ln y - \mu)^2}{2\sigma^2}}, & y \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$$E(X^2) = \text{Var}(X) + E(X)^2 = \sigma^2 + \mu^2$$

$$\therefore E(Y) = \sigma^2 + \mu^2$$

4. (a) Suppose that the two discrete r.v.s X and Y have joint p.m.f. given by

X	$Y = 1$	$Y = 2$	$Y = 3$	$Y = 4$
$X = 1$	2/32	3/32	4/32	5/32
$X = 2$	3/32	4/32	5/32	6/32

Obtain $E(X)$ and $E(Y)$.

(b) Suppose that the two continuous r.v.s X and Y have joint p.d.f.

$$f(x, y) = \begin{cases} x + y, & 0 \leq x \leq 1, 0 \leq y \leq 1, \\ 0, & \text{otherwise.} \end{cases}$$

Find $E(X)$ and $E(Y)$.

$$(a) \begin{array}{c|cc} X & 1 & 2 \\ \hline P & \frac{7}{16} & \frac{9}{16} \end{array} \quad \begin{array}{c|cccc} Y & 1 & 2 & 3 & 4 \\ \hline P & \frac{5}{32} & \frac{7}{32} & \frac{9}{32} & \frac{11}{32} \end{array} \quad \frac{x^2}{2} + \frac{y^2}{4}$$

$$\therefore E(X) = 1 \times \frac{7}{16} + 2 \times \frac{9}{16} = \frac{25}{16}, \quad E(Y) = 1 \times \frac{5}{32} + 2 \times \frac{7}{32} + 3 \times \frac{9}{32} + 4 \times \frac{11}{32} = \frac{45}{16}$$

$$(b) f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^1 (x+y) dy = \left(xy + \frac{y^2}{2} \right) \Big|_0^1 = x + \frac{1}{2} \quad (0 \leq x \leq 1)$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = y + \frac{1}{2} \quad (0 \leq y \leq 1)$$

$$\therefore E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx = \int_0^1 \left(x^2 + \frac{x}{2} \right) dx = \frac{7}{12}, \quad E(Y) = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_0^1 \left(y^2 + \frac{y}{2} \right) dy = \frac{7}{12}$$