# Week 8: Joint, Marginal, Conditional Continuous Distributions, Expectations, Covariance and Correlations

MATH697

Sahir Bhatnagar

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McGill University

# **Joint Continuous Distributions**

#### Introduction

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- We simply make the now-familiar substitutions of integrals for sums and PDFs for PMFs
- Remembering that the probability of any individual point is now
   0

# Joint Continuous CDF

 In order for X and Y to have a continuous joint distribution, we require that the joint CDF

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- The partial derivative with respect to x and y is called the joint PDF.
- The joint PDF determines the joint distribution, as does the joint CDF.

# Joint Continuous PDF

#### Definition 1 (Joint Continuous PDF)

If X and Y are continuous with joint CDF  $F_{X,Y}$ , their **joint PDF** is the derivative of the joint CDF with respect to x and y:

$$f_{X,Y}(x,y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y)$$
 (1)

We require  $f_{X,Y}(x,y) \ge 0$  for all x and y and

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1$$

# Joint Continuous PDF

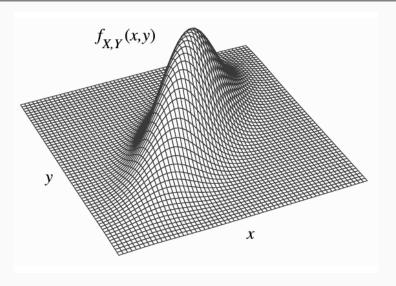


Figure 1

# Joint Continuous PDF Example

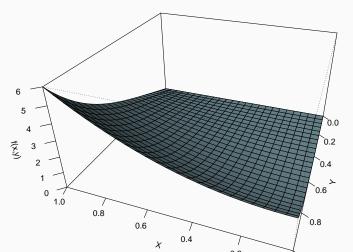
#### Example 2 (Joint Continuous PDF)

Let X and Y be continuous with joint PDF f given by

$$f_{X,Y}(x,y) = \begin{cases} 4x^2y + 2y^5 & 0 \le x \le 1, 0 \le y \le 1\\ 0 & \text{else} \end{cases}$$

- 1. Verify f is indeed a density function
- 2. Compute  $P(0.5 \le X \le 0.7, 0.2 \le Y \le 0.9)$

# Joint Continuous PDF Example



# Marginal Continous PDF

#### Definition 3 (Marginal Continuous PDF)

If X and Y are continuous RVs with joint PDF  $f_{X,Y}(x,y)$ , the marginal PDF of X is:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy, \quad x \in \mathbb{R}$$
 (2)

Similarly the marginal PDF of Y is:

$$f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx, \quad y \in \mathbb{R}$$
 (3)

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- · We have mainly been looking at the joint distribution of two RVs
- However marginalization works analogously with any number of variables
- For example, if we have the joint PDF of X, Y, Z, W but want the joint PDF of X, W, we just have to integrate over all possible values of Y and Z:

$$f_{X,W}(x,w) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y,Z,W}(x,y,z,w) dy dz$$

# Marginal Continous PDF (Rule of Thumb)

Integrate over the unwanted variables to get the joint PDF of the wanted variables

# Marginal Continuous PDF Example (continued)

#### Example 4 (Joint Continuous PDF)

Let X and Y be continuous with joint PDF f given by

$$f_{X,Y}(x,y) = \begin{cases} 4x^2y + 2y^5 & 0 \le x \le 1, 0 \le y \le 1\\ 0 & \text{else} \end{cases}$$

- 1. Find the marginals of X and Y
- 2. Are X and Y independent?

# Continuous PDF Example

#### Example 5 (Joint Continuous PDF)

Let X and Y be continuous with joint PDF f given by

$$f_{X,Y}(x,y) = \begin{cases} 120x^3y & x \ge 0, y \ge 0, x + y \le 1\\ 0 & \text{else} \end{cases}$$

- 1. Verify that f is indeed a valid joint PDF
- 2. Find the marginals of X and Y

#### Conditional Continuous PDF

#### Definition 6 (Conditional Continuous PDF)

If *X* and *Y* are continuous with joint PDF  $f_{X,Y}$ , the **conditional PDF** of *Y* given X = x is:

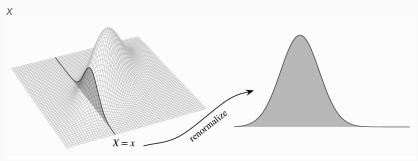
$$f_{Y|X}(y \mid x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$
 (4)

The **conditional PDF** of *X* given Y = y is:

$$f_{X|Y}(x \mid y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}$$
 (5)

#### **Conditional Continuous PDF**

Conditional PDF of Y given X=x. The conditional PDF  $f_{Y\mid X}(y\mid x)$  is obtained by renormalizing the slice of the joint PDF at the fixed value

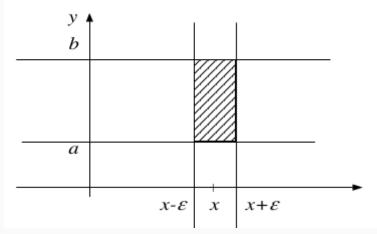


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- Fortunately, many important results such as Bayes' rule work in the continuous case exactly as one would hope

# Joint PDF from Conditional and Marginal

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Alternatively we have:

$$f_{X,Y}(x,y) = f_{Y|X}(y \mid x)f_X(x)$$

# Continuous form of Bayes' rule

# Definition 7 (Continuous form of Bayes' rule and Law of Total Probability)

If X and Y are continuous then:

$$f_{Y|X}(y \mid x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{f_{X|Y}(x \mid y)f_Y(y)}{f_X(x)}$$
(6)

and

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy = \int_{-\infty}^{\infty} f_{X|Y}(x \mid y) f_Y(y) dy$$
 (7)

# Continuous form of Bayes' rule - Remark

• In Equation (7), if we plugged in the other expression for  $f_{X,Y}(x,y) = f_{Y|X}(x \mid y)f_X(x)$  we get

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

$$= \int_{-\infty}^{\infty} f_{Y|X}(x \mid y) f_X(x) dy$$

$$= f_X(x) \int_{-\infty}^{\infty} f_{Y|X}(x \mid y) dy$$
(8)

Equation (8) implies that  $\int_{-\infty}^{\infty} f_{Y|X}(x \mid y) dy = 1$ , confirming the fact that conditional PDFs must integrate to 1.

# Conditional Continuous PDF Example (continued)

#### Example 8 (Joint Continuous PDF)

Let X and Y be continuous with joint PDF f given by

$$f_{X,Y}(x,y) = \begin{cases} 4x^2y + 2y^5 & 0 \le x \le 1, 0 \le y \le 1\\ 0 & \text{else} \end{cases}$$

- 1. Compute  $P(0.2 \le Y \le 0.3 | X = 0.8)$
- 2. Compute  $P(0.2 \le Y \le 0.3)$

# Summary of Bayes' Rule

	Y discrete	Y continuous
X discrete	$P(Y = y \mid X = x) = \frac{P(X = x \mid Y = y)P(Y = y)}{P(X = x)}$	$f_Y(y \mid X = x) = \frac{P(X = x \mid Y = y)f_Y(y)}{P(X = x)}$
	` ,	
X continuous	$P(Y = y \mid X = x) = \frac{f_X(x Y=y)P(Y=y)}{f_X(x)}$	$f_{Y X}(y \mid x) = \frac{f_{X Y}(x y)f_Y(y)}{f_X(x)}$

# Summary of Law of Total Probability

	Y discrete	Y continuous
X discrete	P(X = x) =	P(X = x) =
	$\sum_{y} P(X = x \mid Y = y) P(Y = y)$	$\int_{-\infty}^{\infty} P(X = x \mid Y = y) f_Y(y) dy$
X continuous	$f_X(x) = \sum_y f_X(x \mid Y = y) P(Y = y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X Y}(x \mid y) f_Y(y) dy$

# Independence of Continuous RVs

#### Definition 9 (Independence of Continuous RVs)

If X and Y are jointly continuous, then X and Y are independent if and only if their joint density function  $f_{X,Y}$  can be chosen to satisfy

$$f_{X,Y}(x,y) = f_X(x)f_Y(y), \qquad x,y \in \mathbb{R}$$
 (9)

### Independence of Continuous RVs - Example

#### Example 10 (Independence of Continuous RVs)

Let X and Y have joint density

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{8080} (12xy^2 + 6x + 4y^2 + 2) & 0 \le x \le 6, 3 \le y \le 5\\ 0 & else \end{cases}$$

- 1. Compute the marginal densities for X and Y
- 2. Show that X and Y are independent

### Independence of Continuous RVs - Example

#### Example 11 (Unit Disk)

Let X and Y be a completely random point in the unit disk  $\{(x,y): x^2+y^2\leq 1\}$  with joint PDF

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\pi} & x^2 + y^2 \le 1\\ 0 & else \end{cases}$$

- 1. Can  $f_{X,Y}(x,y) = f_X(x)f_Y(y)$  where  $f_X(x) = 1$  and  $f_Y(y) = 1/\pi$  for  $0 \le x \le 1, 0 \le y \le 1$ , demonstrating that X and Y are independent?
- 2. Try out the point (0.9,0.9) to support your answer above

### MGFs of Independent Continuous RVs

#### Theorem 12 (MGFs of Independent Continuous RVs)

Let X and Y be independent random variables with MGF  $M_X(t)$  and  $M_Y(t)$ , respectively. Then the MGF of the random variable Z = X + Y is given by

$$M_Z(t) = M_X(x)M_Y(y) \tag{10}$$

Proof: on board

• Let 
$$Z = X + Y$$

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#### Example 13 (Distribution of the sum of Normal RVs)

Let  $X \sim N(\mu, \sigma^2)$  and  $Y \sim N(\gamma, \tau^2)$  be independent normal random variables. From Assignment 3, we know that the MGFs of X and Y are

$$M_X(t) = \exp\{\mu t + \sigma^2 t^2 / 2\}$$
  $M_Y(t) = \exp\{\gamma t + \tau^2 t^2 / 2\}$ 

Thus from Theorem (12), the MGF of Z = X + Y is

$$M_Z(t) = M_X(t)M_Y(t) = \exp\{(\mu + \gamma)t + (\sigma^2 + \tau^2)t^2/2\}$$

MGF of a Normal RV with mean  $\mu + \gamma$  and variance  $\sigma^{2} + \tau^{2}$ 

# **Expectations**

### **Expectation of Continuous RVs**

### Theorem 14 (Expectation of Continuous RVs)

Let g be a function from  $\mathbb{R}^2 \to \mathbb{R}$ . If X and Y are jointly continuous, with joint PDF  $f_{X,Y}$ , then

$$E[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dx dy$$
 (11)

### Expectation of Continuous RVs - Example

### Example 15 (Expected distance between two Uniforms)

For independent  $X \sim \textit{Unif}(0,1)$  and  $Y \sim \textit{Unif}(0,1)$ , find E(|X-Y|)

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- Positive covariance between X and Y indicates that when X goes up, Y also tends to go up
- Negative covariance indicates that when X goes up, Y tends to go down.

### Definition 16 (Covariance)

The Covariance between RVs X and Y is

$$Cov(X, Y) = E((X - EX)(Y - EY))$$
(12)

$$= E((X - \mu_X)(Y - \mu_Y)) \tag{13}$$

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#### Proposition 17 (Alternative definition of Covariance)

$$Cov(X,Y) = E(XY) - E(X)E(Y)$$
(14)

Proof: Assignment 4

#### Definition 18 (Correlation)

The Correlation between RVs X and Y is

$$\rho_{XY} = R^{2} = Corr(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}}$$

$$= \frac{Cov(X, Y)}{\sigma_{X}\sigma_{Y}}$$
(15)

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- if X is in centimeters rather than meters, the covariance is multiplied a hundredfold

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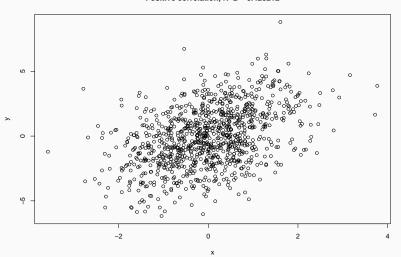
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- · Covariance depends on the units in which X and Y are measured
- if X is in centimeters rather than meters, the covariance is multiplied a hundredfold
- Correlation is easier to interpret since it is a unitless version of covariance

#### Visualization of Positive Correlation

```
x <- rnorm(1e3); y <- x + 2 * rnorm(1e3)
plot(x ,y, main = sprintf("Positive correlation, R^2 = %g",cor(x,y)))</pre>
```

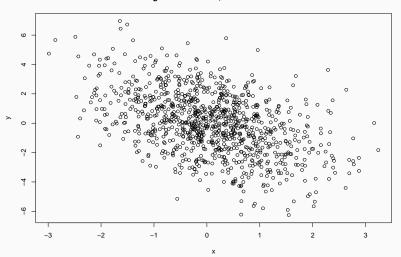
#### Positive correlation, R^2 = 0.420212



### Visualization of Negative Correlation

```
x <- rnorm(1e3); y <- -x + 2 * rnorm(1e3)
plot(x ,y, main = sprintf("Negative correlation, R^2 = %g",cor(x,y)))</pre>
```

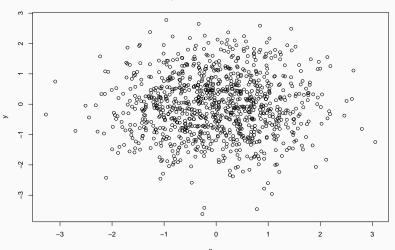
#### Negative correlation, R^2 = -0.463363



### Visualization of Independence

```
x <- rnorm(1e3); y <- rnorm(1e3)
plot(x ,y, main = sprintf("Independent, R^2 = %g",cor(x,y)))</pre>
```

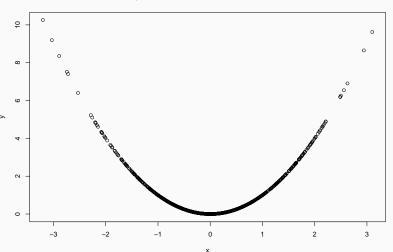
#### Independent, R^2 = 0.0218506



### Visualization of Dependent but Uncorrelated

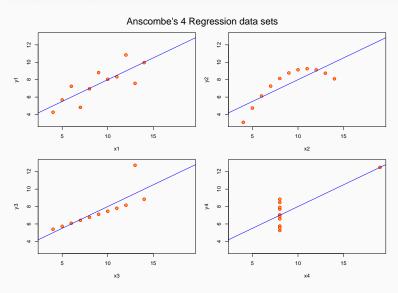
```
x <- rnorm(1e3); y <- x^2
plot(x ,y, main = sprintf("Independent but uncorrelated, R^2 = %g",cor(x,y,method = "s</pre>
```

Independent but uncorrelated, R^2 = 0.0476517



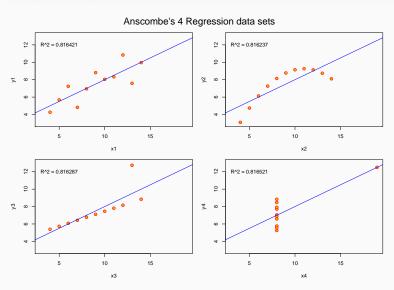
### Anscombes Quartet

data("anscombe")



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- If X and Y tend to move in opposite directions, then X EX and Y — EY will tend to have opposite signs, giving a negative covariance
- If X and Y are independent, then their covariance is zero. We say that RVs with zero covariance are uncorrelated.

### Theorem 19 (Independence and Not Correlated)

If X and Y are independent, then they are uncorrelated.

proof: on board

#### Theorem 19 (Independence and Not Correlated)

If X and Y are independent, then they are uncorrelated.

proof: on board

The converse of this theorem is false: just because X and Y are uncorrelated does not mean they are independent

### **Properties of Covariance**

- 1. Cov(X, X) = Var(X)
- 2. Cov(X, Y) = Cov(Y, X)
- 3. Cov(X, c) = 0 for any constant c
- 4. Cov(aX, Y) = aCov(X, Y) for any constant a
- 5. Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z)
- 6. Cov(X+Y,Z+W) = Cov(X,Z) + Cov(X,W) + Cov(Y,Z) + Cov(Y,W)
- 7.  $Var(aX + bY) = a^2Var(X) + b^2Var(Y) + 2abCov(X, Y)$

Proof: Assignment 4

### **Properties of Correlation**

#### Theorem 20 (Correlation Bounds and linear relationships)

For any random variables X and Y

- 1.  $-1 \le \rho_{XY} \le 1$
- 2.  $\rho_{XY}=\pm 1$  if and only if there exist numbers  $a\neq 0$  and b such that P(Y=aX+b)=1 (i.e.  $\rho_{XY}=\pm 1$  implies X and Y are linearly related, and if X and Y are linearly related, then  $\rho_{XY}=\pm 1$ )
- 3. If  $\rho_{\rm XY}=$  1, then a> 0, and if  $\rho_{\rm XY}=-$ 1 then a< 0 (i.e. a and  $\rho_{\rm XY}$  have the same sign)

proof: on board

### Correlation Example

#### Example 21 (Correlation)

Let the joint PDF of (X, Y) be

$$f(x,y) = \begin{cases} 1 & 0 < x < 1, x < y < x + 1 \\ 0 & else \end{cases}$$

Find the correlation of X and Y