

# Week 11: Convergence in Distribution and Central Limit Theorem

MATH697

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# Convergence in Distribution

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## Definition 1 (Convergence in Distribution)

Let  $X_1, X_2, \dots$  be a sequence of random variables. Then we say that the **sequence converges in distribution** to a random variable  $X$  if

$$\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$$

at all points  $x$  where  $F_X(x)$  is continuous and we write  $X_n \xrightarrow{D} X$

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- The importance of this, is that often the distribution of  $X_n$  is difficult to work with, while that of  $X$  is much simpler
- With  $X_n$  converging in distribution to  $X$ , however, we can approximate the distribution of  $X_n$  by that of  $X$ .

# Convergence in Probability Implies in Distribution

## Theorem 2 (Convergence in Probability Implies in Convergence in Distribution)

*If the sequence of random variables  $X_1, X_2, \dots$  converges to a random variable  $X$ , the sequence also converges in distribution to  $X$ . That is if  $X_n \xrightarrow{P} X$  then  $X_n \xrightarrow{D} X$*

# Poisson Approximation to the Binomial

## Example 3 (Poisson Approximation to the Binomial)

Suppose that  $X_n \sim \text{Binomial}(n, \lambda/n)$  and  $X \sim \text{Poisson}(\lambda)$ . We have previously seen that as  $n \rightarrow \infty$

$$P(X_n = j) = \binom{n}{j} \left(\frac{\lambda}{n}\right)^j \left(1 - \frac{\lambda}{n}\right)^{n-j} \rightarrow e^{-\lambda} \frac{\lambda^j}{j!}$$

This implies that  $F_{X_n}(x) \rightarrow F_X(x)$  and thus  $X_n \xrightarrow{D} X$



# Central Limit Theorem

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- Let  $X_1, X_2, \dots$ , be iid with mean  $\mu$  and finite variance  $\sigma^2$ . The law of large numbers says that as  $n \rightarrow \infty$ ,  $\bar{X}_n$  converges to the constant  $\mu$  (with probability 1).

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- But what is its distribution along the way to becoming a constant?
- This is addressed by the central limit theorem (CLT)

# Central Limit Theorem (CLT)

- The CLT states that for large  $n$ , the distribution of  $\bar{X}_n$  (after standardization) approaches a standard Normal distribution.
- By standardization, we mean that we subtract  $\mu$ , the expected value of  $\bar{X}_n$ , and divide by  $\sigma/\sqrt{n}$ , the standard deviation of  $\bar{X}_n$ .

## Theorem 4 (CLT)

As  $n \rightarrow \infty$

$$\sqrt{n} \left( \frac{\bar{X}_n - \mu}{\sigma} \right) \xrightarrow{D} \mathcal{N}(0, 1)$$

. From this we can also say that for large  $n$

$$\bar{X}_n \sim \mathcal{N}(\mu, \sigma^2/n)$$

*Proof:* on board

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- In words, the CDF of the left-hand side approaches  $\Phi$ , the CDF of a standard Normal distribution.
- Starting from virtually no assumptions (other than independence and finite variances), **we end up with normality**

## Running Proportion of Heads Revisited

Before, we used the law of large numbers to conclude that  $\bar{X}_n \rightarrow 1/2$  as  $n \rightarrow \infty$ . Now, using the central limit theorem, we can say more:  $E(\bar{X}_n) = 1/2$  and  $\text{Var}(\bar{X}_n) = 1/(4n)$ , so for large  $n$ ,

$$\bar{X}_n \sim \mathcal{N}\left(\frac{1}{2}, \frac{1}{4n}\right)$$



# Poisson Convergence to Normal

Let  $Y \sim \text{Pois}(n)$ . Using MGFs it can be shown that  $Y$  can be expressed as a sum of  $n$  iid  $\text{Poisson}(1)$  random variables. Therefore by CLT, for large  $n$ ,

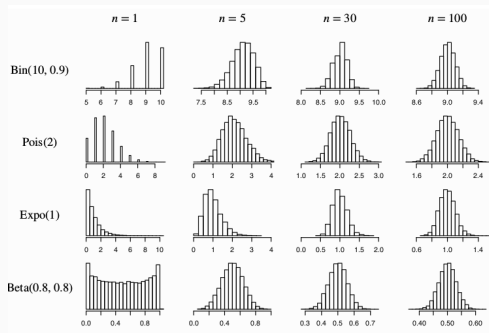
$$Y \sim \mathcal{N}(n, n)$$

# Binomial Convergence to Normal

Let  $Y \sim \text{Binomial}(n, p)$ . Using MGFs it can be shown that  $Y$  can be expressed as a sum of  $n$  iid Bernoulli( $p$ ) random variables. Therefore by CLT, for large  $n$ ,

$$Y \sim \mathcal{N}(np, np(1 - p))$$

Histograms of the distribution of  $\bar{X}_n$  for different starting distributions of the  $X_j$  (indicated by the rows) and increasing values of  $n$  (indicated by the columns). Each histogram is based on 10,000 simulated values of  $\bar{X}_n$ . Regardless of the starting distribution of the  $X_j$ , the distribution of  $\bar{X}_n$  approaches a Normal distribution as  $n$  grows.



## CLT Exercise

One way to visualize the CLT for a distribution of interest is to plot the distribution of  $\bar{X}_n$  for various values of  $n$ . To do this, we first have to generate iid  $X_1, \dots, X_n$  a bunch of times from our distribution of interest. For example, suppose that our distribution of interest is  $Unif(0, 1)$ , and we are interested in the distribution of  $\bar{X}_{12}$ , i.e., we set  $n = 12$ . In the following code, we create a matrix of iid standard Uniforms. The matrix has 12 columns, corresponding to  $X_1$  through  $X_{12}$ . Each row of the matrix is a different realization of  $X_1$  through  $X_{12}$ .

```
nsim <- 10^4
n <- 12
x <- matrix(runif(n*nsim), nrow=nsim, ncol=n)
xbar <- rowMeans(x)
hist(xbar)
# change runif for rexp. what do you notice?
```

# Session Info

```
devtools::session_info()
```

```
## setting value
## version R version 3.4.1 (2017-06-30)
## system x86_64, linux-gnu
## ui X11
## language en_US
## collate en_US.UTF-8
## tz Canada/Eastern
## date 2017-11-14
##
## package * version date source
## abind 1.4-5 2016-07-21 cran (@1.4-5)
## arm 1.9-3 2016-11-27 cran (@1.9-3)
## assertthat 0.2.0 2017-04-11 CRAN (R 3.4.1)
## backports 1.1.0 2017-05-22 cran (@1.1.0)
## base * 3.4.1 2017-07-08 local
## bindr 0.1 2016-11-13 CRAN (R 3.4.1)
## bindrcpp 0.2 2017-06-17 CRAN (R 3.4.1)
## blme 1.0-4 2015-06-14 cran (@1.0-4)
## broom 0.4.2 2017-02-13 CRAN (R 3.4.1)
```