Week 9: Multidimensional Change of Variable

MATH697

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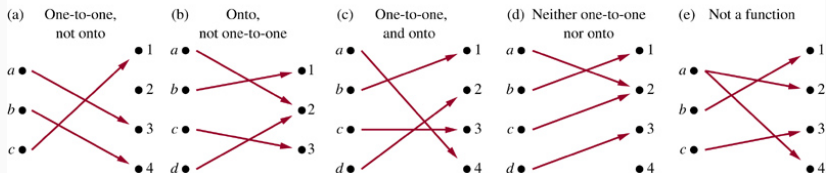McGill University

First Some Terminology

Figure 1

Multidimensional Change of Variable

- **Univariate Transformations**: We previously discussed the problem of starting with a single random variable X, forming some function of X, such as X^2 or e^X, to obtain a new random variable $Y = h(X)$

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- Multivariate Transformations: We now generalize this scenario by starting with more than a single random variable

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    1. The total lifetime $X + Y$
    2. The ratio of lifetimes X/Y; e.g. if the ratio is 2, the original component lasted twice as long as its replacement
    3. The ratio $X/(X + Y)$; the proportion of system lifetime during which the original component operated

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- We now focus on finding the joint distribution of these two new continuous variables (U, V)

Recall, $U = g_1(X, Y)$ and $V = g_2(X, Y)$.

Joint Distribution of Two New Random Variables

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- $g_1(\cdot)$ and $g_2(\cdot)$ are functions that express the **new variables in terms of the original ones**
- The general result presumes that these functions can be inverted to solve for the original variables in terms of the new ones:

$$X = h_1(U, V) \qquad Y = h_2(U, V)$$

- For example, if

$$u = g_1(x, y) = x + y \qquad v = g_2(x, y) = x - y$$

then

$$x = h_1(u, v) = \frac{1}{2}(u + v) \qquad y = h_2(u, v) = \frac{1}{2}(u - v)$$

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- $\mathcal{B}$ is the **image** of $\mathcal{A}$ under the transformation

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- Define the **Jacobian** as

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\begin{aligned}
\mathcal{J} &= \det \begin{bmatrix} \frac{\partial h_1(u,v)}{\partial u} & \frac{\partial h_1(u,v)}{\partial v} \\ \frac{\partial h_2(u,v)}{\partial u} & \frac{\partial h_2(u,v)}{\partial v} \end{bmatrix} \\
&= \frac{\partial h_1(u, v)}{\partial u} \frac{\partial h_2(u, v)}{\partial v} - \frac{\partial h_1(u, v)}{\partial v} \frac{\partial h_2(u, v)}{\partial u}
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- Then

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f_{U,V}(u, v) = \begin{cases} f_{X,Y}(h_1(u, v), h_2(u, v)) \times |\mathcal{J}| & u, v \in \mathcal{B} \\ 0 & else \end{cases}
$$

Theorem 1 (Multivariate Transformations (2 random variables))

Let X and Y be jointly continuous, with PDF $f_{X,Y}(x, y)$ on the set
$\mathcal{A} = \{(x, y) : f_{X,Y}(x, y) > 0\}$. *Let $U = g_1(X, Y)$ and $V = g_2(X, Y)$ on the set*
$\mathcal{B} = \{(u, v) : u = g_1(x, y), v = g_2(x, y) \text{ for some } (x, y) \in \mathcal{A}\}$, *where*
$g_1, g_2 : \mathbb{R}^2 \to \mathbb{R}$ *are differentiable functions. Assume g_1, g_2 is one-to-one on $\mathcal{A}$,*
i.e., we can solve the equations $u = g_1(x, y), v = g_2(x, y)$ for x and y denoted by
$x = h_1(u, v)$ *and $y = h_2(u, v)$. Then U and V are also jointly continuous with PDF*

$$f_{U,V}(u, v) = \begin{cases} f_{X,Y}(h_1(u, v), h_2(u, v)) \times |\mathcal{J}| & u, v \in \mathcal{B} \\ 0 & else \end{cases} \tag{1}$$

where

$$\mathcal{J} = det \begin{bmatrix} \frac{\partial h_1(u,v)}{\partial u} & \frac{\partial h_1(u,v)}{\partial v} \\ \frac{\partial h_2(u,v)}{\partial u} & \frac{\partial h_2(u,v)}{\partial v} \end{bmatrix}$$

Example 2 (Uniform Square Transformation)

Let X and Y be independent $Uniform(0, 1)$ random variables.

Consider the two random variables $U = X + Y$ and $V = X - Y$.

1. Determine the joint PDF of U and V

2. What is the marginal PDF of U and marginal PDF of V?

Example 3 (Distribution of the product of beta variables)

Let $X \sim Beta(\alpha, \beta)$ and $Y \sim Beta(\alpha + \beta, \gamma)$ be independent random variables. The joint PDF of (X, Y) is

$$f_{X,Y}(x, y) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1}(1-x)^{\beta-1} \frac{\Gamma(\alpha + \beta + \gamma)}{\Gamma(\alpha + \beta)\Gamma(\gamma)} y^{\alpha+\beta-1}(1-y)^{\gamma-1}$$

on the set $\mathcal{A} = \{(x, y) : 0 < x < 1, 0 < y < 1\}$ and 0 elsewhere. Consider the transformation $U = XY$ and $V = X$

1. Determine the joint PDF of U and V
2. What is the marginal PDF of U and marginal PDF of V?

Example 4 (Sum and difference of normal variables)

Let $X \sim N(0, 1)$ and $Y \sim N(0, 1)$ be independent standard normal random variables. The joint PDF of (X, Y) is

$$f_{X,Y}(x, y) = \frac{1}{2\pi} \exp\{-x^2/2\} \exp\{-y^2/2\}$$

on the set $\mathcal{A} = \{(x, y) : -\infty < x < \infty, -\infty < y < \infty\}$.
Consider the transformation $U = X + Y$ and $V = X - Y$

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- Then for each i, the inverse functions from $\mathcal{B}$ to $\mathcal{A}$ can be found.

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- Let $\mathcal{J}_i$ denote the Jacobian computed from the i the inverse, then

$$f_{U,V}(u, v) = \sum_{i=1}^{k} f_{X,Y}(h_{1i}(u, v), h_{2i}(u, v)) |\mathcal{J}_i| \qquad (2)$$

Multivariate Transformation Example

Example 5 (Distribution of the ratio of normal variables)

Let $X \sim N(0, 1)$ and $Y \sim N(0, 1)$ be independent standard normal random variables. The joint PDF of (X, Y) is

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on the set $\mathcal{A} = \{(x, y) : -\infty < x < \infty, -\infty < y < \infty\}$. Consider the transformation $U = X/Y$ and $V = |Y|$. Note that this transformation is not one-to-one since the points (x, y) and $(-x, -y)$ are both mapped into the same (u, v) point. But if we restrict consideration to either positive or negative values of y, then the transformation is one-to-one.

1. Determine the joint PDF of U and V
2. What is the marginal PDF of U ?

Hierarchical Models and Mixture Distributions

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- While in general, a RV can only have one distribution, it is often easier to model a situation by thinking of things in a hierarchy
- Perhaps the most classic hierarchical model is given in the following example

Example 6 (Binomial-Poisson hierarchy (revisited))

An chicken lays a large number of eggs, each surviving with probability p. On the average, how many eggs will survive? The large number of eggs laid is a random variable, often taken to be $Poisson(\lambda)$. Furthermore, if we assume that each egg's survival is independent, then we have Bernoulli trials. Therefore, if we let X=number of survivors and N=number of eggs laid, we have

$$X|N \sim Binomial(N, p) \quad \text{and} \quad N \sim Poisson(\lambda)$$

a hierarchical model.

- We showed in Example 25 (Week 7), that the marginal distribution of X is $Poisson(\lambda p)$ with Y playing no part at all

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- The answer to the original question is easy to compute $E(X) = \lambda p$. So on average λp eggs will survive
- If we were interested only in this mean and did not need the distribution, we could have used properties of **conditioal expectation**
- Sometimes, **calculations can be greatly simplified** by using the following theorem. Recal that $E(X|y)$ is a function of y and $E(X|Y)$ is a random variable whose value depends on the value of Y

Theorem 7 (Conditional Expectation)

If X and Y are any two random variables, then

$$E(X) = E(E(X|Y)) \tag{3}$$

provided that the expectations exist

Proof: on board

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- In general we can say that hierarchical models lead to mixture distributions

Example 9 (Binomial-Poisson-Exponential hierarchy)

Consider a generalization of Example 6 where instead of one mother chicken, there are a large number of mothers, and one mother is chosen at random. We are still interested in knowing the average number of survivors, but it is no longer clear that the number of eggs laid follows the same Poission distribution for each mother. The following three stage hierarchy may be more appropriate. Let X = number of survivors. Then

$$X|N \sim Binomial(N, p)$$
$$N|\Lambda \sim Poisson(\Lambda)$$
$$\Lambda \sim Exponential(\beta)$$

where the last stage of hierarchy accounts for the variability across different mothers.

Theorem 10 (Conditional Variance)

For any two random variables X and Y

$$Var(X) = E(Var(E|Y)) + Var(E(X|Y)) \tag{4}$$

Proof: on board

Example 11 (Beta-binomial hierarchy)

One generalization of the binomial is to allow the success probability to vary according to a distribution. A standard model for this situation is

$$X|P \sim Binomial(n, P)$$
$$P \sim beta(\alpha, \beta)$$

Session Info

```
devtools::session_info()
```

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##  version  R version 3.4.1 (2017-06-30)
##  system   x86_64, linux-gnu
##  ui       X11
##  language en_US
##  collate  en_US.UTF-8
##  tz       Canada/Eastern
##  date     2017-11-06
##
##  package     * version  date        source
##  abind         1.4-5    2016-07-21  cran (@1.4-5)
##  arm           1.9-3    2016-11-27  cran (@1.9-3)
##  assertthat    0.2.0    2017-04-11  CRAN (R 3.4.1)
##  backports     1.1.0    2017-05-22  cran (@1.1.0)
##  base        * 3.4.1    2017-07-08  local
##  bindr         0.1      2016-11-13  CRAN (R 3.4.1)
##  bindrcpp      0.2      2017-06-17  CRAN (R 3.4.1)
##  blme          1.0-4    2015-06-14  cran (@1.0-4)
##  broom         0.4.2    2017-02-13  CRAN (R 3.4.1)
```