

Introduction to Quantile Regression and Applications

Cliff J. Huang
Emeritus Professor
Department of Economics
Vanderbilt University

Yung Lieh. Yang
Emeritus Professor
Department of Finance
Ling Tung University

November 2015

Outline

- I. Mean, Quantiles, and Distribution
- II. Quantile Regression
- III. Fundamentals of Quantile Estimation
- IV. Computation of Quantile Regression: R-Program
- V. Conclusion Remarks

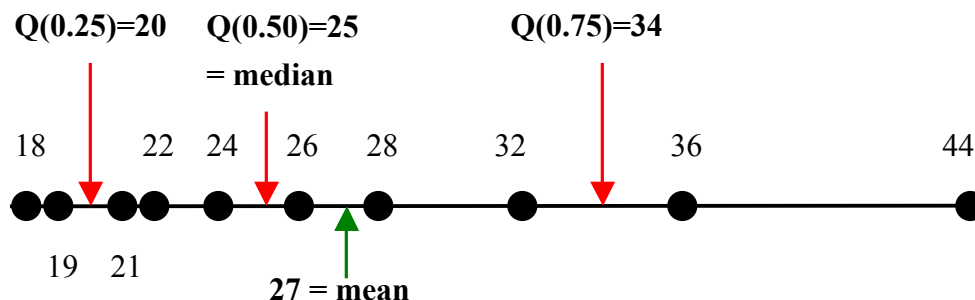
I. Mean, Quantiles, and Distribution

1. Distribution of a Random Variable

- Example: Starting Salary of 10 Employees

$Y = 18, 19, 21, 22, 24, 26, 28, 32, 36, 44$ (in thousands of NT\$)

❖ Mean = 27



❖ Quartiles:

First Quartile (25 Percentile) = 20

Second Quartile (50 Percentile) = 25 = Median

Third Quartile (75 Percentile) = 34

❖ Distribution Skewness

Symmetric (No Skewness):

Median = Mean

Median – First Quartile = Third Quartile – Median

Positive Skewness (Skew to the Right)

Median < Mean

Median – First Quartile < Third Quartile – Median

Negative Skewness (Skew to the Left)

Median > Mean

Median – First Quartile > Third Quartile – Median

- Quantiles of order τ : $Q(\tau)$, $0 \leq \tau \leq 1$

❖ $Q(0.20) = 19$; $Q(0.25) = 20$; $Q(0.50) = 25$

$Q(0.75) = 34$; $Q(0.90) = 36$; $Q(1.00) = ?$

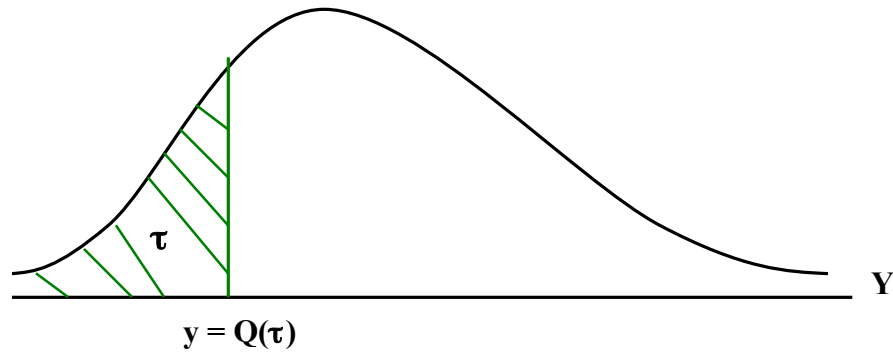
❖ Quantile as a Ranking:

$Q(\tau) < Q(\lambda)$ if $\tau < \lambda$

2. Quantile and Distribution Function

- Distribution Function

$$F(y) = P(Y \leq y)$$



- Quantiles

❖ Quantiles of order τ :

$$F(y) = P(Y \leq y) = \tau \text{ implies } y = F^{-1}(\tau) = Q(\tau)$$

Examples:

$$25 = F^{-1}(0.5) = Q(0.5)$$

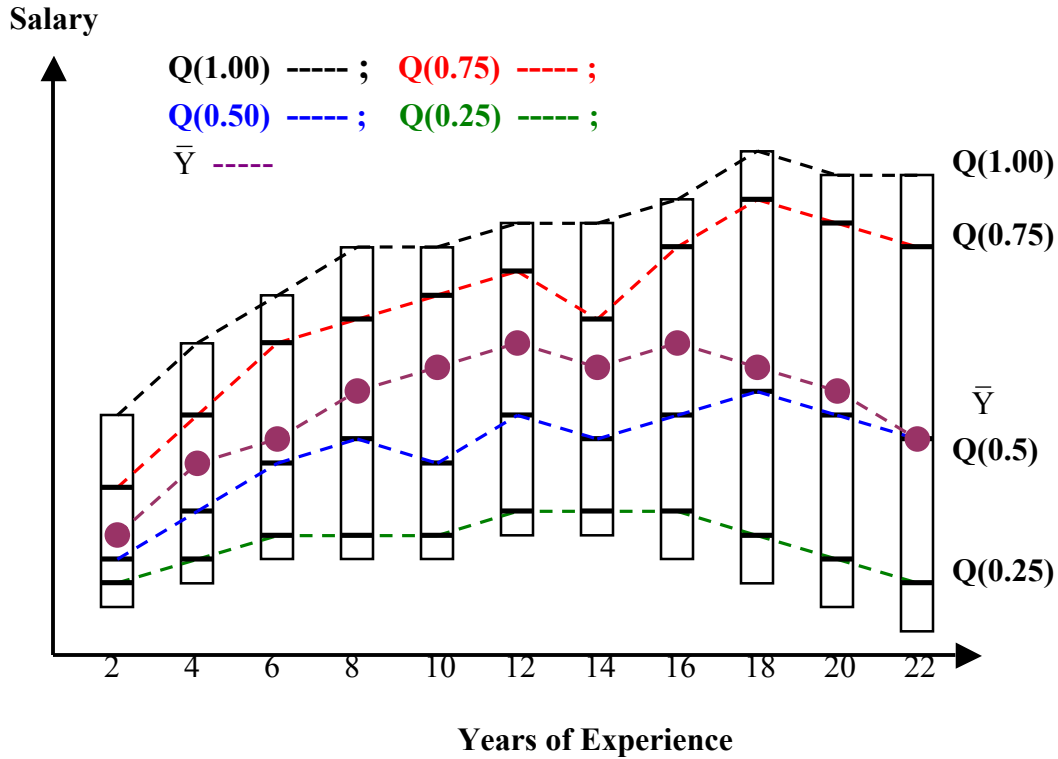
$$34 = F^{-1}(0.75) = Q(0.75)$$

II. Quantile Regression

1. Classical Regression versus Quantile Regression

- Example:

$y = \text{salary}; x = \text{years of experience}$



- Classical Regression

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

- ❖ Conditional Mean Regression

$$E(y_i | x_i) = \beta_0 + \beta_1 x_i \equiv \mu_i$$

- Quantile Regression of order $\tau \in [0, 1]$

$$y_i = \beta_0^\tau + \beta_1^\tau x_i + \varepsilon_i$$

- ❖ Conditional Quantile Regression

$$Q(\tau | x_i) = \beta_0^\tau + \beta_1^\tau x_i$$

2. Example: Quantile Regressions of Wage Determination in Taiwan

- Data: 29,133 plant workers in 1996

wage = hourly wage; experience = months on current job

- Regressions:

❖ Conditional Mean Regression

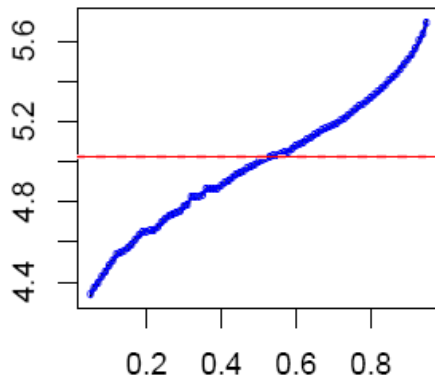
$$\ln(\text{wage}) = \beta_0 + \beta_1 (\text{experience}) + \varepsilon$$

❖ Qunatile Regression

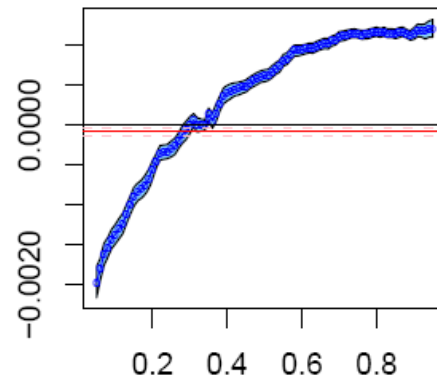
$$\ln(\text{wage}) = \beta_0^\tau + \beta_1^\tau (\text{experience}) + \varepsilon$$

τ	β_0^τ	β_1^τ
0.1	4.47987	-0.00137
0.2	4.65334	-0.00056
0.3	4.77952	0.00000
0.4	4.88014	0.00040
0.5	4.99839	0.00062
0.6	5.07736	0.00094
0.7	5.18110	0.00111
0.8	5.31419	0.00117
0.9	5.50766	0.00115
0.95	5.69128	0.00121
Conditional Mean	5.02400	-0.000086

(Intercept)



experience



III. Fundamentals of Quantile Regression Estimation

1. Estimation of Mean and Quantile

- Sample Observations: $y = (y_1, y_2, \dots, y_n)$
- Population Mean and Population Quantile

$$E(y_i) = \mu; \quad Q(\tau) = F^{-1}(\tau)$$

- Least-square Estimation of Population Mean

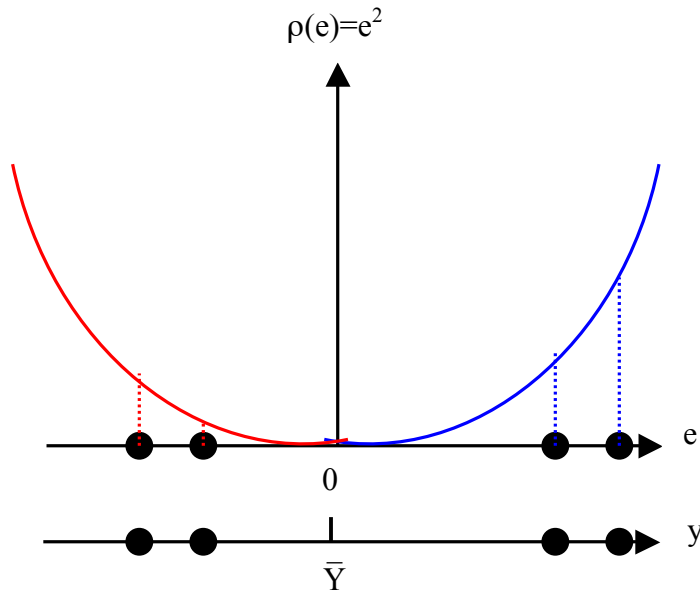
❖ Model: $y_i = \mu + \varepsilon_i = \bar{Y} + e_i$

Error: $e_i = y_i - \bar{Y}$

❖ Loss Function: $\rho(e_i) = e_i^2 = (y_i - \bar{Y})^2$

- ❖ Least-square method:

$$\min_{\bar{Y}} \sum_{i=1}^n \rho(e_i) = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i - \bar{Y})^2; \quad \bar{Y} = \frac{1}{n} \sum_{i=1}^n y_i$$



- Quantile Estimation of $Q(\tau)$

❖ Model: $y_i = Q(\tau) + \varepsilon_i = \hat{Q} + e_i$

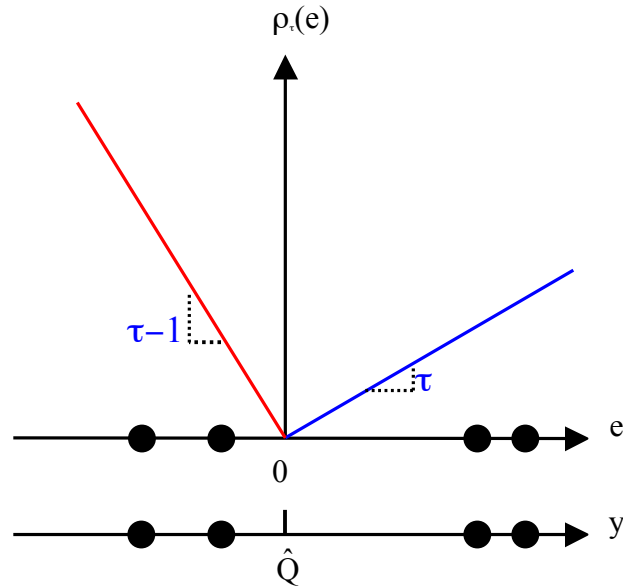
Error: $e_i = y_i - \hat{Q}$

Example: $\tau = 0.25$ implies 100 τ % of $e_i = y_i - \hat{Q} \leq 0$

100 (1- τ) % of $e_i = y_i - \hat{Q} \geq 0$

❖ Loss Function:

$$\begin{aligned}
\rho_{\tau}(e_i) &= \tau \times e_i = \tau \times (y_i - \hat{Q}) \quad \text{if } e_i \geq 0 \\
&= (\tau - 1) \times e_i = (\tau - 1) \times (y_i - \hat{Q}) \quad \text{if } e_i \leq 0 \\
\rho_{\tau}(e_i) &= (\tau - I(e_i < 0)) \times e_i = (\tau - I(y_i - \hat{Q} < 0)) \times (y_i - \hat{Q})
\end{aligned}$$



❖ Quantile Estimation Method

$$\min_{\hat{Q}} \sum_{i=1}^n \rho_{\tau}(e_i) = \sum_{i=1}^n (\tau - I(y_i - \hat{Q} < 0)) \times (y_i - \hat{Q})$$

Note: (1) For any i , there is the constraint

$$\text{If } (y_i - \hat{Q}) \begin{cases} \geq 0 \\ \leq 0 \end{cases}, \text{ then } \rho_{\tau}(e_i) = \begin{cases} \tau \times (y_i - \hat{Q}) \\ (1 - \tau) \times (\hat{Q} - y_i) \end{cases}$$

i.e, for any i ,

$$\hat{Q} + (y_i - \hat{Q}) - (\hat{Q} - y_i) = y_i$$

(2) Let $u_i = (y_i - \hat{Q}) \geq 0$, $v_i = (\hat{Q} - y_i) \geq 0$ be two “slack” variables,

then for any i ,

$$\rho_{\tau}(e_i) = \tau \times u_i + (1 - \tau) \times v_i$$

$$\text{subject to } \hat{Q} + u_i - v_i = y_i$$

❖ Linear Programming Estimation of Quantiles

$$\min_{\hat{Q}, u, v} \sum_{i=1}^n (\tau \times u_i + (1 - \tau) \times v_i)$$

$$\text{subject to } \hat{Q} + u_i - v_i = y_i, \quad i = 1, 2, \dots, n$$

$$u_i, v_i \geq 0$$

2. Estimation of Classical Regression and Quantile Regression

- Model: $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$
- Classical Regression

❖ Conditional Mean: $E(y_i | x_i) = \beta_0 + \beta_1 x_i$

❖ Least-square Method: $\min_{\hat{\beta}_0, \hat{\beta}_1} \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$

- Quantile Regression:

❖ Conditional Quantile: $Q(\tau | x_i) = \beta_0^\tau + \beta_1^\tau x_i$

❖ Quantile Method: $\min_{\hat{\beta}_0^\tau, \hat{\beta}_1^\tau} \sum_{i=1}^n \rho_\tau(e_i) = \sum_{i=1}^n (\tau - I(e_i < 0)) \times (e_i)$

$$e_i = y_i - \hat{\beta}_0^\tau - \hat{\beta}_1^\tau x_i$$

IV. Computation of Quantile Regression: R-Program

- Free Download of **R-Program** at
<http://www.r-project.org>
- Example on Wage Determination in Taiwan
 - ❖ Observations: 29,133 plant workers in 1996
 - ❖ Variables: wage = hourly wage; lnwage = $\ln(\text{wage})$; gender = 1 male, 0 female;
 age = year of age; married = 1 married, 0 unmarried;
 education = year of education;
 experience = months of working experience on current job

- Use of R-Program

- ❖ Data file: wage.dat

wage	lnwage	gender	age	married	education	experience
270.83	5.60	1	33	0	9	2
175	5.16	1	36	1	12	125
116.07	4.75	0	28	1	12	26
...

- ❖ Program file:

```
> library(quantreg)
## change directory
{
  dat=read.table("wage.dat",header=T)
  attach(dat)
  ls <- lm(lnwage~gender+married+education+experience)
  summary(ls)
  qt <- summary(rq(lnwage~gender+married+education+experience,
                   tau=5:95/100))
  postscript("wage.ps", horizontal = FALSE, width = 6.5, height = 3.5)
  plot(qt, nrow =1, ncol = 2)
  dev.off()
}
```

- ❖ Outputs:

- (1) Least-square outputs:

- Coefficients:

(Intercept)	gender	married	education	experience
4.078e+00	3.074e-01	2.242e-01	5.660e-02	-2.925e-05

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	4.078e+00	1.099e-02	370.912	<2e-16 ***
gender	3.074e-01	6.071e-03	50.641	<2e-16 ***
married	2.242e-01	6.344e-03	35.340	<2e-16 ***
education	5.660e-02	7.833e-04	72.260	<2e-16 ***
experience	-2.925e-05	2.814e-05	-1.039	0.299

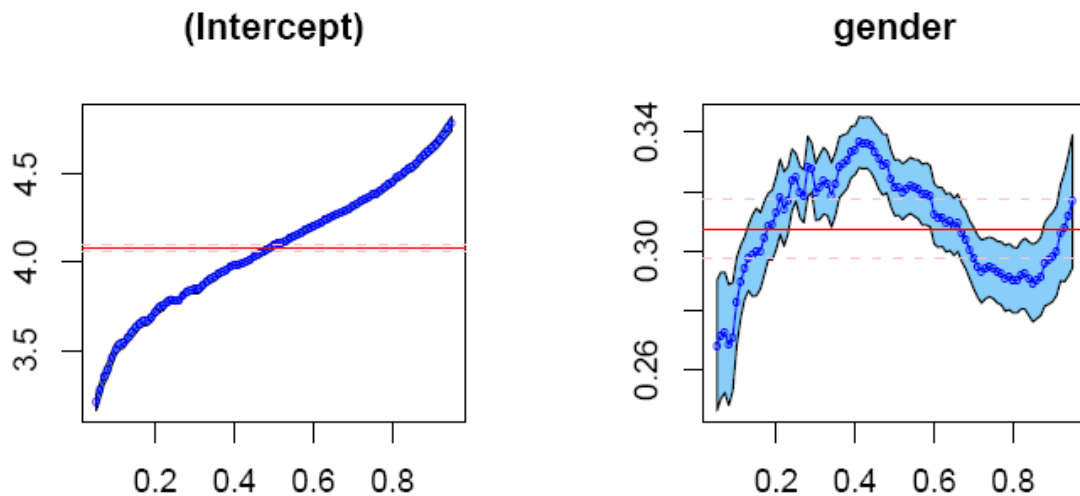
(2) Quantile Outputs:

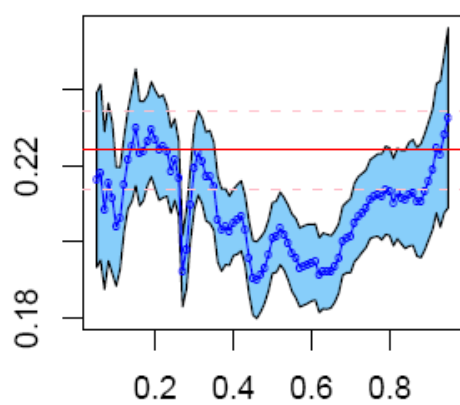
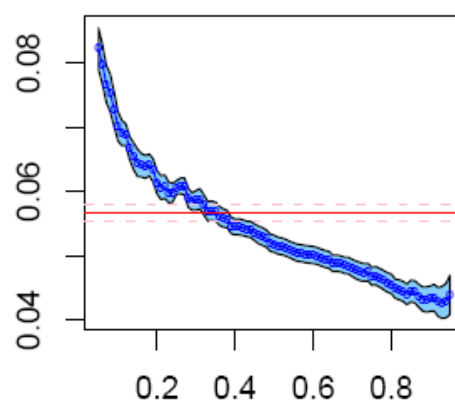
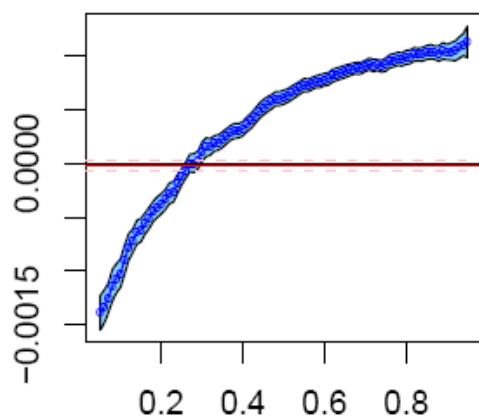
tau: [1] 0.05

Coefficients:

	Value	Std. Error	t value	Pr(> t)
(Intercept)	3.20542	0.02845	112.66064	0.00000
gender	0.26821	0.01336	20.08050	0.00000
married	0.21626	0.01395	15.49835	0.00000
education	0.08225	0.00195	42.17080	0.00000
experience	-0.00139	0.00010	-14.33464	0.00000

(3) Graphs of $y_i = \beta^\tau$



married**education****experience**

V. Conclusion Remarks on Quantile Regression

1. Regression:

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

- Homocedastic Case: $\varepsilon_i \sim \text{iid}$

$$\begin{aligned} Q_Y(\tau | x_i) &= \beta_0 + \beta_1 x_i + Q_\varepsilon(\tau) = (\beta_0 + Q_\varepsilon(\tau)) + \beta_1 x_i \\ &= \beta_0^\tau + \beta_1 x_i \end{aligned}$$

Intercept varies with τ , but not slope.

- Heteroscedastic Case: $\varepsilon_i = h(x_i) w_i = x_i w_i$

$$\begin{aligned} Q_Y(\tau | x_i) &= \beta_0 + \beta_1 x_i + x_i Q_w(\tau) = \beta_0 + (\beta_1 + Q_w(\tau)) x_i \\ &= \beta_0 + \beta_1^\tau x_i \end{aligned}$$

Slope varies with τ , but not intercept.

2. Interpreting Quantile Regression Coefficients:

- Classical Regression:

$$\diamond y_i = \beta_0 + \beta_1 x_i + \varepsilon_i; \quad E(y_i | x_i) = \beta_0 + \beta_1 x_i$$

$$\frac{\partial E(y_i | x_i)}{\partial x_i} = \beta_1$$

$$\diamond \text{Transformation: } \ln(y_i) = \beta_0 + \beta_1 x_i + \varepsilon_i; \quad E(\ln(y_i) | x_i) = \beta_0 + \beta_1 x_i;$$

$$\frac{\partial E(\ln(y_i) | x_i)}{\partial x_i} = \beta_1, \quad \text{but} \quad \frac{\partial E(y_i | x_i)}{\partial x_i} \neq e\beta_1$$

- Quantile Regression:

$$\diamond y_i = \beta_0 + \beta_1 x_i + \varepsilon_i; \quad Q_y(\tau | x_i) = \beta_0^\tau + \beta_1^\tau x_i$$

$$\frac{\partial Q_y(\tau | x_i)}{\partial x_i} = \beta_1^\tau$$

$$\diamond \text{Transformation: } \ln(y_i) = \beta_0 + \beta_1 x_i + \varepsilon_i; \quad Q_{\ln(y)}(\tau | x_i) = \beta_0^\tau + \beta_1^\tau x_i$$

$$\frac{\partial Q_{\ln(y)}(\tau | x_i)}{\partial x_i} = \beta_1^\tau \quad \text{and} \quad \frac{\partial Q_y(\tau | x_i)}{\partial x_i} = e\beta_1^\tau$$