

binomial-gamma-hurdle

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0.1 Binomial-Gamma Hurdle Models

0.1.1 Model description

Dependent variable, a number of animals observed per minute (y) is semi-continuous (i.e. a point mass in a single value and a continuous distribution elsewhere). The data generating process for this type of data can be modelled using a gamma distribution. The main problem is however that response variable has a high proportion of zeros (96%), which is more than expected from a gamma distribution with, therefore it cannot be readily applied.

Lets consider the two common methods for dealing with zero-inflated data:

- (1) Modelling a zero-inflation parameter that represents the probability a given 0 comes from the main distribution (say the negative binomial distribution) or is an excess 0;
- (2) Modelling the zero and non-zero data with one model and then modelling the non-zero data with another. This is often called a hurdle model.

In (1), the response variable is modelled as a mixture of a Bernoulli distribution (a point mass at zero) and a Poisson distribution (or any other count distribution supported on non-negative integers). In (2), the basic idea is that a Bernoulli probability governs the binary outcome of whether a variable has a zero or positive realization. If the realization is positive, the hurdle is crossed, and the conditional distribution of the positives is governed by a truncated-at-zero model. Hurdle models model the zeros and non-zeros as two separate processes and can be useful in that they allow you to model the zeros and non-zeros with different predictors or different roles of the same predictors.

Zero-inflation models may be more elegant and informative if the same predictors are thought to contribute to the extra and real zeros.

Hurdle models can be useful in that they allow you to model the zeros and non-zeros with different predictors or different roles of the same predictors. Maybe one process leads to the zero/non-zero data and another leads to the non-zero magnitude.

Here we shall focus on (2) and model the zeros separately from the non-zeros in a binomial-Gamma hurdle model.

0.1.2 Load libraries

```
In [25]: # library(R.utils)
library(ggplot2)
# library(GGally)
# library(lmtest)
```

```
# library(tidyverse)
library(lme4)
library(effects)
library(optimx)
```

```
In [26]: set.seed(4322)
         Sys.time()
```

```
[1] "2019-03-19 15:19:04 GMT"
```

```
In [15]: # require(devtools)
         # install_version("effects", version = "4.0-0")
```

0.1.3 Read in data

```
In [31]: dat <- read.csv(file = 'data.csv', row.names=1)
         # sunfish <- read.csv('ignore/sunfish.csv')
```

Variable *y* is a response variable, variables *x1* and *x2* are explanatory variables. Variable *x1* represent a number of observers, variable *x2* represent an environmental variable (such as sea surface temperature).

```
In [29]: head(dat)
```

y	x1	x2	year
0	1	10.40875	1971
0	1	10.40875	1971
0	1	10.40875	1971
0	1	10.40875	1971
0	1	10.40875	1971
0	1	10.40875	1971
0	1	10.40875	1971

0.1.4 Scale data

```
In [39]: # select variables to scale
         cols = c("x1", "x2")
         # scale variables and add to a df
         dat[, paste0(cols, "_", "sc")] <- scale(dat[, cols])
         summary(dat)
```

y	x1	x2	year
Min. :0.000000	Min. : 1.000	Min. : 9.207	Min. :1971
1st Qu.:0.000000	1st Qu.: 2.000	1st Qu.:13.169	1st Qu.:1984
Median :0.000000	Median : 4.000	Median :15.261	Median :1998
Mean :0.002676	Mean : 4.824	Mean :14.621	Mean :1996
3rd Qu.:0.000000	3rd Qu.: 6.000	3rd Qu.:16.288	3rd Qu.:2009
Max. :0.132941	Max. :40.000	Max. :18.168	Max. :2017
	NA's :1485	NA's :61	

x1_sc	x2_sc
Min. : -0.9718	Min. : -2.5651
1st Qu.: -0.7176	1st Qu.: -0.6882
Median : -0.2093	Median : 0.3032
Mean : 0.0000	Mean : 0.0000
3rd Qu.: 0.2990	3rd Qu.: 0.7894
Max. : 8.9408	Max. : 1.6802
NA's : 1485	NA's : 61

0.1.5 Binomial model

When relating the sightings to temperature what we are interested in detecting are annual trends over and above seasonal fluctuations that we would expect. So we would expect that within each year as temperature increases during spring and summer and zooplankton blooms occur, sunfish sightings will increase. What we want to know is -- in a year when zooplankton abundance and temperatures are high are sunfish sightings also high.

```
In [43]: summary(glm(iffelse(dat$y>0,1,0) ~
                        x1_sc +
                        x2_sc +
                        year,
                        data = dat,
                        family = binomial(link = logit)))
```

Call:

```
glm(formula = iffelse(dat$y > 0, 1, 0) ~ x1_sc + x2_sc + year,
     family = binomial(link = logit), data = dat)
```

Deviance Residuals:

Min	1Q	Median	3Q	Max
-0.3255	-0.2907	-0.2742	-0.2629	2.6607

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	17.079047	13.167616	1.297	0.195
x1_sc	0.018503	0.089106	0.208	0.836
x2_sc	0.051264	0.094037	0.545	0.586
year	-0.010179	0.006605	-1.541	0.123

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 1136.3 on 3487 degrees of freedom
 Residual deviance: 1133.6 on 3484 degrees of freedom
 (1487 observations deleted due to missingness)
 AIC: 1141.6

Number of Fisher Scoring iterations: 6

We see that observer related variables (x1) is highly significant. We try to isolate its effect for each year. We apply a mixed-effect modeling framework and fit a varying intercept model with lmer. This approach is useful when we are interested explicitly in variation among and by groups. Group level variables are specified using a special syntax: (1|year) to fit a linear model with a varying-intercept group effect using the variable year.

We include 'year' as random effect with noise variables.

```
In [45]: m.bin.full.re <- glmer(ifelse(dat$y>0,1,0) ~
      x1_sc +
      (1|year) ,
      data = dat,
      # control = glmerControl(optimizer='optimx', optCtrl=list(method='nlopt',
      family = binomial(link = logit))
```

```
In [46]: summary(m.bin.full.re)
```

Generalized linear mixed model fit by maximum likelihood (Laplace

Approximation) [glmerMod]

Family: binomial (logit)

Formula: ifelse(dat\$y > 0, 1, 0) ~ x1_sc + (1 | year)

Data: dat

AIC	BIC	logLik	deviance	df.resid
1141.4	1159.9	-567.7	1135.4	3487

Scaled residuals:

Min	1Q	Median	3Q	Max
-0.2532	-0.2034	-0.1948	-0.1898	5.5795

Random effects:

Groups	Name	Variance	Std.Dev.
--------	------	----------	----------

year	(Intercept)	0.05928	0.2435
------	-------------	---------	--------

Number of obs: 3490, groups: year, 38

Fixed effects:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-3.24648	0.10363	-31.327	<2e-16 ***
x1_sc	0.04109	0.08446	0.486	0.627

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

(Intr)
x1_sc -0.044