

The correlation function of X and Y is defined

as

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}} \\ = \frac{\mathbb{E}[(X - \mathbb{E}(X))(Y - \mathbb{E}(Y))]}{\sqrt{\mathbb{E}[(X - \mathbb{E}(X))^2] \mathbb{E}[(Y - \mathbb{E}(Y))^2]}}.$$

Conditional moments.

As conditional distribution function $F_{X|Y}(x)$ is a random variable (randomness is induced by the random variable Y), it is natural to imagine that conditional moments are also random. Consider a function (continuous if the bivariate random vector is of continuous type) $g: \mathbb{R}^2 \rightarrow \mathbb{R}$. Then we define the random variable (of continuous type)

$$\mathbb{E}(g(x, y) | Y) : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}, \mathcal{B}(\mathbb{R})).$$

such that $\mathbb{E}(g(x, y) | Y)$ takes the value

$$\mathbb{E}_{Y=y} (g(x, y) | Y) = \mathbb{E}(g(x, y) | y) \\ = \mathbb{E}(g(x, y) | Y=y).$$

$$= \begin{cases} \int_{-\infty}^{\infty} dx f_{x|y}(x|y) g(x,y) & (x,y) \text{ is a random vector of continuous type} \\ \sum_{i=1}^{\infty} g(x_i, y) p_{x|y}(x_i|y) & (x,y) \text{ is a random vector of discrete type and } p_{x|y}(x_i|y) \neq 0, \\ & \text{that is, } p_y(y) \neq 0. \end{cases}$$

with conditional probability density function $f_{x|y}(x|y)$ and mass function $p_{x|y}(x_i|y)$.

As we have seen that there is a relation between conditional probability density/mass function, marginal probability density/mass function and joint probability density/mass function, it is natural to see imagine that those relations can be used to obtain some formulae which relate the moments of (x,y) to the conditional moments.

We can show that $E(g(x,y))$ exists that is,

$E(|g(x,y)|) < \infty$, then

$$E(g(x,y)) = E[E(g(x,y)|y)].$$

$$= \int_{-\infty}^{\infty} dy f_y(y) \mathbb{E}(g(x,y) | y=y)$$

$$= \left\{ \int_{-\infty}^{\infty} dy f_y(y) \left[\int_{-\infty}^{\infty} dz f_{x|y}(z|y) g(x,y) \right] \right.$$

$$\left. \sum_{j=1}^{\infty} p_y(y_j) \sum_{i=1}^{\infty} p_{x|y}(x_i|y_j) g(x_i, y_j) \right\}$$

Similarly, we can use the following formula

$$\mathbb{E}(g(x,y)) = \mathbb{E}[\mathbb{E}(g(x,y) | x)].$$

Properties of conditional expectation

i) For any constant c , $E(c|Y) = c$ almost surely.

" $X = c$ almost surely" means " $P(X = c) = 1$ ".

ii) Linearity.

$$E[a_1 g_1(x) + a_2 g_2(x) | Y] = a_1 E(g_1(x) | Y) + a_2 E(g_2(x) | Y)$$

almost surely (with probability one).

iii) If $P(X \geq 0) = 1$ then $E(X|Y) \geq 0$ with probability 1.

iv) If $P(g(x,Y) \geq 0) = 1$, then $E(g(x,Y) | Y) \geq 0$ with probability 1.

v) If $\phi: \mathbb{R}^2 \rightarrow \mathbb{R}$ and $E(|\phi(x,Y)|) < \infty$, then

$$E[\phi(x,Y) | Y=y] = E[\phi(x,y) | Y=y]$$

vi) $E[\psi(x) \phi(x,Y) | X] = \psi(x) E(\phi(x,Y) | X)$ with probability 1.

vii) The n th conditional moment ~~of X given~~ of X given Y is defined to be $E(X^n | Y)$.

viii) Let $E(h(x))$ exists that is, $E(|h(x)|) < \infty$, then $E(h(x)) = E[E(h(x) | Y)]$.

$\forall x \in \mathbb{R}$ if $\mathbb{E}(x^2) < \infty$, then

$$\text{Var}(x) = \mathbb{V}\text{ar}(\mathbb{E}(x|Y)) + \mathbb{E}(\mathbb{V}\text{ar}(x|Y)).$$

Lecture - XI Independent random vectors, moments
and moment generating functions.

Consider two distribution functions $F_1 : \mathbb{R} \rightarrow [0,1]$ and $F_2 : \mathbb{R} \rightarrow [0,1]$. Define a function

$$F : \mathbb{R} \times \mathbb{R} \rightarrow [0,1] \text{ such that } F(x,y) = F_1(x) F_2(y)$$

for all $(x,y) \in \mathbb{R}^2$.

Exercise Check that $F : \mathbb{R}^2 \rightarrow [0,1]$ is a two-dimensional distribution function.

According to this exercise, there exists a bivariate random vector (X,Y) such that

$$\mathbb{P}(\{X \leq x\} \cap \{Y \leq y\}) = F(x,y) = F_1(x) F_2(y).$$

for all $(x,y) \in \mathbb{R}^2$. This means that there exists

a probability space (Ω, \mathcal{F}, P) such that

$(X,Y) : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$ is a bivariate random vector.

As the distribution function uniquely characterizes the law of a random vector, it is immediate to see that for any $A \times B \in \mathcal{B}(\mathbb{R}^2)$, we should have

$$\mathbb{P}(\{X \in A\} \cap \{Y \in B\}) = \mathbb{P}(X \in A) \mathbb{P}(Y \in B). \quad \rightarrow (*)$$

(2)

Let $(X, Y) : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$ be a bivariate random vector. Then we call X and Y to be ~~independent~~ if for every $A \times B \in \mathcal{B}(\mathbb{R}^2)$

$$P(\{(X \in A) \cap (Y \in B)\}) = P(X \in A) P(Y \in B) \rightarrow (*)_1$$

The components of an n -dimensional random vector $(X_1, X_2, \dots, X_n) : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ are called independent if for every $A_1 \times A_2 \times \dots \times A_n \in \mathcal{B}(\mathbb{R}^n)$, we have

$$P\left(\bigcap_{i=1}^n \{X_i \in A_i\}\right) = \prod_{i=1}^n P(X_i \in A_i). \rightarrow (*)_2$$

Let $F_{X_1, \dots, X_n} : \mathbb{R}^n \rightarrow [0, 1]$ be the distribution function of the random vector (X_1, X_2, \dots, X_n) . Then it follows from $(*)_2$ that

$$F_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \prod_{i=1}^n F_{X_i}(x_i). \rightarrow (*)_3$$

where $F_{X_i} : \mathbb{R} \rightarrow [0, 1]$ is the distribution function of X_i for all $i = 1, 2, \dots, n$.

Theorem

Consider n distribution functions $(F_i : 1 \leq i \leq n)$. Define a function $F : \mathbb{R}^n \rightarrow [0, 1]$ such that

$$F(x_1, x_2, \dots, x_n) = \prod_{i=1}^n F_i(x_i) \text{ for every } (x_1, \dots, x_n) \in \mathbb{R}^n. \quad (3)$$

Then there exists an n -dimensional random vector $(X_1, \dots, X_n): (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ with distribution function given by F .

Probability mass/density function of independent random variables.

Let (X, Y) be a random vector of discrete type.

Then

$$\begin{aligned} P(X=x_i, Y=y_j) &= P(X=x_i) P(Y=y_j) \\ &= p_x(x_i) p_y(y_j) \end{aligned}$$

Let (X, Y) be a random vector of continuous type with density function $f_{X,Y}(x, y)$. Then it follows

that

$$f_{X,Y}(x, y) = f_x(x) f_y(y) \quad \forall (x, y) \in \mathbb{R}^2$$

where f_x and f_y are the density functions of X and Y respectively.

Joint mass/density function is the product of marginal mass/density functions if the random variables are independent.

Pairwise independent random variables.

Let $(X_1, X_2, \dots, X_n) : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ be an n -dimensional random vector. We call the random variables $(X_1, \dots, X_n) = (X_i : 1 \leq i \leq n)$ to be pairwise independent if ~~independent~~ for every $i \neq j$, we have

$$\mathbb{P}(X_i \in A \cap X_j \in B) = \mathbb{P}(X_i \in A) \mathbb{P}(X_j \in B)$$

for every $A, B \in \mathcal{B}(\mathbb{R}^2)$.

Remark Pairwise independent random variables might not be independent completely.

Theorem

Let $(X, Y) : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$ be a random vector. Suppose that $f : \mathbb{R} \rightarrow \mathbb{R}$ and $g : \mathbb{R} \rightarrow \mathbb{R}$ are Borel measurable maps. Then $f(X)$ and $g(Y)$ are also independent random variables.

Identically distributed random variables

Let X and Y be two random variables (may not be defined on the same probability space). Then we say X and Y are identically distributed if for every $A \in \mathcal{B}(\mathbb{R})$, $\mathbb{P}(X \in A) = \mathbb{P}(Y \in A)$.

Independently and identically distributed random variables

We call components of an n -dimensional random vector $(X_1, \dots, X_n) : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ to be ~~independently~~ ~~independently~~ independently and identically distributed if the components are independent and identically distributed.

Independent random vectors.

Let $(\underline{X}, \underline{Y}) = (X_1, \dots, X_m, Y_1, Y_2, \dots, Y_n) : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^{m+n}, \mathcal{B}(\mathbb{R}^{m+n}))$ be random vector. We call the random vectors $\underline{X} = (X_1, \dots, X_m)$ and $\underline{Y} = (Y_1, \dots, Y_n)$ to be independent if for every $A \in \mathcal{B}(\mathbb{R}^m)$ and $B \in \mathcal{B}(\mathbb{R}^n)$, we have

$$\begin{aligned} P((\underline{X}, \underline{Y}) \in A \times B) &= P(\{\underline{X} \in A\} \cap \{\underline{Y} \in B\}) \\ &= P(\underline{X} \in A) P(\underline{Y} \in B). \end{aligned}$$

Some properties of i.i.d. random variables.

Suppose that (X_1, X_2, \dots, X_n) be i.i.d. random variables. Then we can see that

$$F_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \prod_{i=1}^n F_{X_i}(x_i) = F_{X_{i_1}, \dots, X_{i_n}}(x_{i_1}, x_{i_2}, \dots, x_{i_n})$$

that is, $(X_1, \dots, X_n) \stackrel{d}{=} (X_{i_1}, X_{i_2}, \dots, X_{i_n})$