Formulario di Statistica con



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R version 2.7.0 (2008-04-22)

Work in progress!

6 settembre 2008

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Indice

In	dice	iii
I	Matematica ed algebra lineare	vii
1	Background 1.1 Operatori matematici 1.2 Operatori relazionali 1.3 Operatori logici 1.4 Funzioni di base 1.5 Funzioni insiemistiche 1.6 Funzioni indice 1.7 Funzioni combinatorie 1.8 Funzioni trigonometriche dirette 1.9 Funzioni trigonometriche inverse 1.10 Funzioni iperboliche dirette 1.11 Funzioni iperboliche inverse 1.12 Funzioni iperboliche inverse 1.12 Funzioni di successione 1.14 Funzioni di ordinamento 1.15 Funzioni di troncamento e di arrotondamento 1.16 Funzioni sui numeri complessi 1.18 Funzioni cumulate 1.19 Funzioni in parallelo 1.20 Funzioni di analisi numerica 1.21 Costanti 1.22 Miscellaneous	1 1 5 7 9 12 15 17 19 21 22 24 25 29 33 36 39 47 50 52 53 59 62
2	Vettori, Matrici ed Arrays 2.1 Creazione di Vettori	75 75 84 99 135 143
3	Misure ed indici statistici 3.1 Minimo e massimo 3.2 Campo di variazione e midrange 3.3 Media aritmetica, geometrica ed armonica 3.4 Mediana e quantili 3.5 Differenza interquartile e deviazione assoluta dalla mediana 3.6 Asimmetria e curtosi 3.7 Coefficiente di variazione 3.8 Scarto quadratico medio e deviazione standard 3.9 Errore standard 3.10 Varianza e devianza 3.11 Covarianza e codevianza 3.12 Matrice di varianza e covarianza 3.13 Correlazione di Pearson, Spearman e Kendall	147 149 150 153 155 158 159 164 166 167 168 170 172

	3.14 Media e varianza pesate	
	3.15 Momenti centrati e non centrati	
	3.16Connessione e dipendenza in media	
	3.17 Sintesi di dati	
	3.18 Distribuzione di frequenza	
	3.20 Variabili casuali discrete	
	3.21 Variabili casuali continue	
	3.22 Logit	
	3.23 Serie storiche	
	3.24 Valori mancanti	. 252
	3.25 Miscellaneous	. 254
4	Analisi Componenti Principali (ACP)	261
	4.1 ACP con matrice di covarianza di popolazione	26 1
	4.2 ACP con matrice di covarianza campionaria	. 264
	4.3 ACP con matrice di correlazione di popolazione	
	4.4 ACP con matrice di correlazione campionaria	. 273
5	Analisi dei Gruppi	281
	5.1 Indici di distanza	
	5.2 Criteri di Raggruppamento	
Π	I Statistica Inferenziale	291
		201
6		293
	6.1 Test di ipotesi sulla media con uno o due campioni	
	6.2 Test di ipotesi sulla media con uno o due campioni (summarized data)	
	6.3 Test di ipotesi sulla varianza con uno o due campioni	
	6.4 Test di ipotesi su proporzioni	
	6.5 Test di ipotesi sull'omogeneità delle varianze	. 348
7	Analisi della varianza (Anova)	351
	7.1 Simbologia	. 351
	7.2 Modelli di analisi della varianza	
	7.3 Comandi utili in analisi della varianza	. 357
Q	Confronti multipli	373
0	8.1 Simbologia	
	8.2 Metodo di Tukey	
	8.3 Metodo di Bonferroni	
	8.4 Metodo di Student	
_		00=
9	****	385
	9.1 Test di ipotesi sulla correlazione lineare9.2 Test di ipotesi sulla autocorrelazione	
	9.2 Test di ipotesi sulla autocorrelazione	. 402
10	O Test di ipotesi non parametrici	409
	10.1 Simbologia	
	10.2Test di ipotesi sulla mediana con uno o due campioni	
	10.3Test di ipotesi sulla mediana con più campioni	
	10.4Test di ipotesi sull'omogeneità delle varianze	
	10.5 Anova non parametrica a due fattori senza interazione	
	10.6Test di ipotesi su una proporzione	
	10.7Test di ipotesi sul ciclo di casualità	
	10.8Test di ipotesi sulla differenza tra parametri di scala	. 450
1:	l Tabelle di contingenza	453
	11.1 Simbologia	
	11.2Test di ipotesi per tabelle di contingenza 2 righe per 2 colonne	. 100
	11.2 Test di ipotesi per tabelle di contingenza 2 righe per k colonne	. 466

	Test di ipotesi sull'adattamento 12.1Test di ipotesi sulla distribuzione normale	495
IV	Modelli Lineari	503
13	Regressione lineare semplice	505
	13.1 Simbologia	
	13.2Stima	
	13.3 Adattamento	
	13.4 Diagnostica	525
14	Regressione lineare multipla	537
	14.1 Simbologia	
	14.2Stima	
	14.3Adattamento	
	14.4 Diagnostica	580
15	Regressione lineare semplice pesata	599
	15.1 Simbologia	5 99
	15.2 Stima	
	15.3Adattamento	
	15.4 Diagnostica	621
16	Regressione lineare multipla pesata	633
	16.1 Simbologia	633
	16.2Stima	634
	16.3 Adattamento	
	16.4 Diagnostica	666
V		685
17	Regressione Logit	687
17	Regressione Logit 17.1 Simbologia	687 687
17	Regressione Logit 17.1 Simbologia	687 688
17	Regressione Logit 17.1 Simbologia	687 687 688 700
17	Regressione Logit 17.1 Simbologia	687 687 688 700 707
17	Regressione Logit 17.1 Simbologia	687 687 688 700 707
17	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia	687 688 700 707 721 721
17	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima	687 687 688 700 707 721 721 722
17	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento	687 688 700 707 721 721 722 734
17	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica	687 688 700 707 721 721 722 734 741
17: 18:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare	687 688 700 707 721 721 722 734 741 755
17:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia	687 688 700 707 721 721 722 734 741 755 755
177	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima	687 688 700 707 721 721 722 734 741 755 756
17:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento	687 688 700 707 721 721 722 734 741 755 756 769
17:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima	687 688 700 707 721 721 722 734 741 755 756 769
17: 18: 19:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy	687 688 700 707 721 721 722 734 741 755 756 769 776
177 183 193	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia	687 688 700 707 721 721 721 722 734 741 755 756 769 776
171 181 191	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima	687 688 700 707 721 721 722 734 741 755 756 769 776 789 789
171 181 191	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima 20.3 Adattamento	687 688 700 707 721 721 722 734 741 755 756 769 776 789 789 790 802
171 181 191	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima	687 688 700 707 721 721 722 734 741 755 756 769 776 789 789 790 802
17: 18: 19:	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima 20.3 Adattamento	687 688 700 707 721 721 722 734 741 755 756 769 776 789 789 790 802
171 181 191 201	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima 20.3 Adattamento 20.4 Diagnostica Regressione di Poisson 21.1 Simbologia	687 688 700 707 721 721 721 722 734 741 755 756 769 776 789 789 780 802 809 823 823
171 181 191 201	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima 20.3 Adattamento 20.4 Diagnostica Regressione di Poisson 21.1 Simbologia 21.2 Stima	687 688 700 707 721 721 721 722 734 741 755 756 769 776 789 789 790 802 809 823 823 824
17 18 19 20	Regressione Logit 17.1 Simbologia 17.2 Stima 17.3 Adattamento 17.4 Diagnostica Regressione Probit 18.1 Simbologia 18.2 Stima 18.3 Adattamento 18.4 Diagnostica Regressione Log-log complementare 19.1 Simbologia 19.2 Stima 19.3 Adattamento 19.4 Diagnostica Regressione di Cauchy 20.1 Simbologia 20.2 Stima 20.3 Adattamento 20.4 Diagnostica Regressione di Poisson 21.1 Simbologia	687 688 700 707 721 721 721 722 734 741 755 756 769 776 789 789 790 802 809 823 823 824 836

00 D. day and an a Community	055
22 Regressione Gamma	855
22.1 Simbologia	
22.2 Stima	856
22.3 Adattamento	867
22.4 Diagnostica	871
23 Regressione di Wald	879
23.1 Simbologia	879
23.2 Stima	
23.3 Adattamento	
23.4 Diagnostica	
VI Appendice	903
A Packages	905
B Links	907
Bibliografia	909
Indice analitico	911

Parte I Matematica ed algebra lineare

Capitolo 1

Background

1.1 Operatori matematici

```
+
```

Package: base
Description: addizione
Example:

> 1 + 2

[1] 3

> x <- c(1, 2, 3, 4, 5)

> y <- c(1.2, 3.4, 5.2, 3.5, 7.8)

> x + y

[1] 2.2 5.4 8.2 7.5 12.8

> x <- c(1, 2, 3, 4, 5)

> x + 10

[1] 11 12 13 14 15

-

• Package: base

• Example:

> x - 10

[1] -9 -8 -7 -6 -5

• **Description**: sottrazione

> 1.2 - 6.7
[1] -5.5

> x <- c(1, 2, 3, 4, 5)
> y <- c(1.2, 3.4, 5.2, 3.5, 7.8)
> x - y

[1] -0.2 -1.4 -2.2 0.5 -2.8

> x <- c(1, 2, 3, 4, 5)</pre>

```
> Inf - Inf
 [1] NaN
 > --3
 [1] 3
• Package: base
• Description: moltiplicazione
• Example:
 > 2.3 * 4
 [1] 9.2
 > x < -c(1.2, 3.4, 5.6, 7.8, 0, 9.8)
 > 3 * x
 [1] 3.6 10.2 16.8 23.4 0.0 29.4
 > x < -c(1, 2, 3, 4, 5, 6, 7)
 > y <- c(-3.2, -2.2, -1.2, -0.2, 0.8, 1.8, 2.8)
 > x * y
 [1] -3.2 -4.4 -3.6 -0.8 4.0 10.8 19.6
• Package: base
• Description: rapporto
• Example:
 > 21/7
 [1] 3
 > x < -c(1.2, 3.4, 5.6, 7.8, 0, 9.8)
 [1] 0.6 1.7 2.8 3.9 0.0 4.9
 > 2/0
 [1] Inf
 > -1/0
 [1] -Inf
 > 0/0
```

*

**

```
[1] NaN
 > Inf/Inf
 [1] NaN
 > Inf/0
 [1] Inf
 > -Inf/0
 [1] -Inf
 > x < -c(1, 2, 3, 4, 5, 6, 7)
 > y < -c(-3.2, -2.2, -1.2, -0.2, 0.8, 1.8, 2.8)
 > y/x
 [1] -3.20 -1.10 -0.40 -0.05 0.16 0.30 0.40
• Package: base
• Description: elevamento a potenza
• Example:
 > 2**4
 [1] 16
 > x < -c(1.2, 3.4, 5.6, 7.8, 0.0, 9.8)
 > x**2
 [1] 1.44 11.56 31.36 60.84 0.00 96.04
 > x < -c(1, 2, 3, 4)
 > y < -c(-3.2, -2.2, -1.2, -0.2)
 > y**x
 [1] -3.2000 4.8400 -1.7280 0.0016
```



• Package: base

• **Description:** elevamento a potenza

• Example:

```
> 2^4
[1] 16
> x <- c(1.2, 3.4, 5.6, 7.8, 0, 9.8)
> x^2
[1] 1.44 11.56 31.36 60.84 0.00 96.04
> x <- c(1, 2, 3, 4)
> y <- c(-3.2, -2.2, -1.2, -0.2)
> y^x
[1] -3.2000 4.8400 -1.7280 0.0016
```

%/%

• Package: base

• **Description:** quoziente intero della divisione

• Example:

```
> 22.6%/%3.4
[1] 6
> 23%/%3
[1] 7
```

%%

• Package: base

• **Description:** resto della divisione (modulo)

```
> 22.6%%3.4
[1] 2.2
> 23%%3
[1] 2
```

• Package: base

1.2 Operatori relazionali

```
<
```

• **Description:** minore • Example: > 1 < 2 [1] TRUE > x < -c(0.11, 1.2, 2.3, 4.5)> x < 2.4[1] TRUE TRUE TRUE FALSE > • Package: base • **Description:** maggiore • Example: > 3 > 1.2 [1] TRUE > x < -c(0.11, 1.2, 2.3, 4.5)> x > 2.4[1] FALSE FALSE FALSE TRUE <= • Package: base • **Description:** minore od uguale • Example: > 3.4 <= 8.5 [1] TRUE > x < -c(0.11, 1.2, 2.3, 4.5)> x <= 2.4[1] TRUE TRUE TRUE FALSE

>=

• Package: base

• **Description:** maggiore od uguale • Example: > 3.4 >= 5.4 [1] FALSE > x < -c(0.11, 1.2, 2.3, 5.4)> x >= 5.4[1] FALSE FALSE FALSE TRUE != • Package: base • **Description:** diverso • Example: > 2 != 3 [1] TRUE > x < -c(0.11, 1.2, 2.3, 5.4)> x != 5.4[1] TRUE TRUE TRUE FALSE == • Package: base • **Description:** uguale • Example: > 4 == 4 [1] TRUE > x < -c(0.11, 1.2, 2.3, 5.4)> x == 5.4[1] FALSE FALSE FALSE TRUE > TRUE == 1 [1] TRUE > FALSE == 0

[1] TRUE

1.3 Operatori logici

```
&
```

• Package: base

• **Description:** AND termine a termine

• Example:

```
> 1 & 5
[1] TRUE
> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> x & 3
[1] TRUE TRUE TRUE TRUE FALSE
```

&&

• Package: base

• Description: AND si arresta al primo elemento che soddisfa la condizione

• Example:

```
> 1 && 5
[1] TRUE
> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> x && 3
[1] TRUE
> x <- c(0, 1.2, 2.3, 4.5, 0)
> x && 3
[1] FALSE
```

I

• Package: base

• **Description:** OR termine a termine

```
> 5 | 0
[1] TRUE

> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> x | 0

[1] TRUE TRUE TRUE TRUE FALSE
```

11

```
• Package: base
```

• **Description:** OR si arresta al primo elemento che soddisfa la condizione

• Example:

```
> 5 || 0
[1] TRUE
> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> x || 3
[1] TRUE
> x <- c(0, 1.2, 2.3, 4.5, 0)
> x || 0
[1] FALSE
```

xor()

• Package: base

• **Description:** EXCLUSIVE OR termine a termine

• Example:

```
> xor(4, 5)
[1] FALSE
> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> xor(x, 3)
[1] FALSE FALSE FALSE FALSE TRUE
```

!

• Package: base

• Description: NOT

```
> !8
[1] FALSE
> x <- c(0.11, 1.2, 2.3, 4.5, 0)
> !x
[1] FALSE FALSE FALSE FALSE TRUE
```

1.4 Funzioni di base

sum()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• **Description:** somma

• Formula:

$$\sum_{i=1}^{n} x_i$$

• Example:

```
> x <- c(1.2, 2, 3)
> 1.2 + 2 + 3

[1] 6.2
> sum(x)

[1] 6.2
> x <- c(1.2, 3.4, 5.1, 5.6, 7.8)
> 1.2 + 3.4 + 5.1 + 5.6 + 7.8

[1] 23.1
> sum(x)

[1] 23.1
```

prod()

• Package: base

• Input:

 \times vettore numerico di dimensione n

• **Description:** prodotto

• Formula:

$$\prod_{i=1}^{n} x_i$$

```
> x <- c(1, 2, 3.2)
> 1 * 2 * 3.2

[1] 6.4
> prod(x)

[1] 6.4
> x <- c(1.2, 3.4, 5.1, 5.6, 7.8)
> 1.2 * 3.4 * 5.1 * 5.6 * 7.8
```

```
[1] 908.8934
> prod(x)
[1] 908.8934
```

abs()

• Package: base

• Input:

x valore numerico

• **Description:** valore assoluto

• Formula:

$$|x| = \begin{cases} x & \text{se } x > 0 \\ 0 & \text{se } x = 0 \\ -x & \text{se } x < 0 \end{cases}$$

• Example:

```
> abs(x = 1.3)

[1] 1.3

> abs(x = 0)

[1] 0

> abs(x = -2.3)

[1] 2.3

> abs(x = 3 + 4i)

[1] 5

> Mod(x = 3 + 4i)
```

• **Note:** Equivale alla funzione Mod().

```
sign()
```

• Package: base

• Input:

x valore numerico

• **Description:** segno

• Formula:

$$sign(x) = \begin{cases} 1 & \text{se } x > 0 \\ 0 & \text{se } x = 0 \\ -1 & \text{se } x < 0 \end{cases}$$

• Example:

```
> sign(x = 1.2)
[1] 1
> sign(x = 0)
[1] 0
> sign(x = -1.2)
[1] -1
```

sqrt()

• Package: base

• Input:

 \times valore numerico tale che x > 0

• **Description:** radice quadrata

• Formula:

 \sqrt{x}

```
> sqrt(x = 2)
[1] 1.414214
> sqrt(x = 3.5)
[1] 1.870829
> sqrt(x = -9)
[1] NaN
> sqrt(x = -9 + 0i)
[1] 0+3i
```

1.5 Funzioni insiemistiche

```
union()
```

```
• Package: base
```

• Input:

- x vettore alfanumerico di dimensione n
- ${\bf y}~$ vettore alfanumerico di dimensione m
- **Description:** unione
- Formula:

 $x \cup y$

• Example:

```
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> y <- c(1, 2, 6, 11)
> union(x, y)

[1] 1 2 3 4 5 6 7 8 9 10 11

> x <- c("a", "b", "c", "d", "e", "f", "g")
> y <- c("a", "e", "f", "h")
> union(x, y)

[1] "a" "b" "c" "d" "e" "f" "g" "h"
```

intersect()

- Package: base
- Input:
 - \mathbf{x} vettore alfanumerico di dimensione n
 - y vettore alfanumerico di dimensione m
- **Description:** intersezione
- Formula:

 $x \cap y$

```
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> y <- c(1, 2, 6, 11)
> intersect(x, y)

[1] 1 2 6

> x <- c("a", "b", "c", "d", "e", "f", "g")
> y <- c("a", "e", "f", "h")
> intersect(x, y)

[1] "a" "e" "f"
```

setdiff()

```
• Package: base
```

• Input:

- \times vettore alfanumerico di dimensione n
- y vettore alfanumerico di dimensione m
- Description: differenza
- Formula:

 $x \setminus y$

• Example:

```
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> y <- c(1, 2, 6, 11)
> setdiff(x, y)

[1]  3  4  5  7  8  9  10

> x <- c("a", "b", "c", "d", "e", "f", "g")
> y <- c("a", "e", "f", "h")
> setdiff(x, y)

[1]  "b" "c" "d" "g"
```

is.element()

• Package: base

• Input:

el valore x alfanumerico set vettore y alfanumerico di dimensione n

- Description: appartenenza di \boldsymbol{x} all'insieme \boldsymbol{y}
- Formula:

 $x \in y$

```
> x <- 2
> y < -c(1, 2, 6, 11)
> is.element(el = x, set = y)
[1] TRUE
> x < - 3
> y < -c(1, 2, 6, 11)
> is.element(el = x, set = y)
[1] FALSE
> x <- "d"
> y <- c("a", "b", "c", "d", "e", "f", "g")</pre>
> is.element(el = x, set = y)
[1] TRUE
> x <- "h"
> y <- c("a", "b", "c", "d", "e", "f", "g")
> is.element(el = x, set = y)
[1] FALSE
```

```
%in%
```

• Package: base

• Input:

x valore alfanumerico

y vettore alfanumerico di dimensione n

• **Description:** appartenenza di x all'insieme y

• Formula:

 $x \in y$

• Example:

```
> x <- 2
> y < -c(1, 2, 6, 11)
> x %in% y
[1] TRUE
> x <- 3
> y < -c(1, 2, 6, 11)
> x %in% y
[1] FALSE
> x <- "d"
> y <- c("a", "b", "c", "d", "e", "f", "g")</pre>
> x %in% y
[1] TRUE
> x <- "h"
> y <- c("a", "b", "c", "d", "e", "f", "g")</pre>
> x %in% y
[1] FALSE
```

setequal()

• Package: base

• Input:

x vettore alfanumerico di dimensione n

y vettore alfanumerico di dimensione \boldsymbol{m}

• **Description:** uguaglianza

• Formula:

$$x = y \Leftrightarrow \left\{ \begin{array}{l} x \subseteq y \\ y \subseteq x \end{array} \right.$$

```
> x <- c(1, 4, 5, 6, 8, 77)
> y <- c(1, 1, 1, 4, 5, 6, 8, 77)
> setequal(x, y)

[1] TRUE
```

```
> x <- c("a", "b")
> y <- c("a", "b", "a", "b", "a", "b", "a")
> setequal(x, y)
[1] TRUE
```

1.6 Funzioni indice

which()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- ullet Description: indici degli elementi di x che soddisfano ad una condizione fissata
- Example:

```
> x <- c(1.2, 4.5, -1.3, 4.5)
> which(x > 2)

[1] 2 4

> x <- c(1.2, 4.5, -1.3, 4.5)
> which((x >= -1) & (x < 5))

[1] 1 2 4

> x <- c(1.2, 4.5, -1.3, 4.5)
> which((x >= 3.6) | (x < -1.6))

[1] 2 4

> x <- c(1.2, 4.5, -1.3, 4.5)
> x | x < 4]

[1] 1.2 -1.3

> x [which(x < 4)]

[1] 1.2 -1.3</pre>
```

which.min()

• Package: base

• Input:

 \times vettore numerico di dimensione n

- ullet **Description:** indice del primo elemento minimo di x
- Example:

```
> x <- c(1.2, 1, 2.3, 4, 1, 4)
> min(x)
```

```
> which(x == min(x))[1]

[1] 2
> which.min(x)

[1] 2
> x <- c(1.2, 4.5, -1.3, 4.5)
> min(x)

[1] -1.3
> which(x == min(x))[1]

[1] 3
> which.min(x)

[1] 3
```

which.max()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• **Description:** indice del primo elemento massimo di \boldsymbol{x}

```
> x <- c(1.2, 1, 2.3, 4, 1, 4)
> max(x)

[1] 4
> which(x == max(x))[1]

[1] 4
> which.max(x)

[1] 4
> x <- c(1.2, 4.5, -1.3, 4.5)
> max(x)

[1] 4.5
> which(x == max(x))[1]

[1] 2
> which.max(x)
```

1.7 Funzioni combinatorie

choose()

• Package: base

• Input:

n valore naturale

k valore naturale tale che $0 \le k \le n$

• **Description:** coefficiente binomiale

• Formula:

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

• Example:

```
> n <- 10
> k <- 3
> prod(1:n) / (prod(1:k) * prod(1:(n - k)))

[1] 120
> choose(n = 10, k = 3)

[1] 120
> n <- 8
> k <- 5
> prod(1:n) / (prod(1:k) * prod(1:(n - k)))

[1] 56
> choose(n = 8, k = 5)

[1] 56
```

lchoose()

• Package: base

• Input:

n valore naturale

k valore naturale tale che $0 \le k \le n$

- Description: logaritmo naturale del coefficiente binomiale
- Formula:

$$\log \binom{n}{k}$$

```
> n <- 10
> k <- 3
> log(prod(1:n)/(prod(1:k) * prod(1:(n - k))))

[1] 4.787492
> lchoose(n = 10, k = 3)
```

```
[1] 4.787492
   > n <- 8
    > k <- 5
    > \log(prod(1:n)/(prod(1:k) * prod(1:(n - k))))
    [1] 4.025352
    > lchoose(n = 8, k = 5)
    [1] 4.025352
factorial()
  • Package: base
  • Input:
       x valore naturale
  • Description: fattoriale
  • Formula:
                                               x!
  • Example:
   > x <- 4
    > prod(1:x)
    [1] 24
   > factorial(x = 4)
    [1] 24
    > x <- 6
    > prod(1:x)
    [1] 720
   > factorial(x = 6)
```

[1] 720

```
lfactorial()
```

• Package: base

```
• Input:
    x valore naturale
• Formula:
                                    \log(x!)
• Example:
 > x <- 4
 > log(prod(1:x))
 [1] 3.178054
 > lfactorial(x = 4)
 [1] 3.178054
 > x <- 6
 > log(prod(1:x))
 [1] 6.579251
 > lfactorial(x = 6)
 [1] 6.579251
```

1.8 Funzioni trigonometriche dirette

sin()

- Package: base
- Input:
 - x valore numerico
- **Description:** seno
- Formula:

 $\sin(x)$

```
> \sin(x = 1.2)

[1] 0.932039

> \sin(x = pi)

[1] 1.224606e-16
```

cos()

• Package: base

• Input:

x valore numerico

• **Description:** coseno

• Formula:

 $\cos(x)$

• Example:

```
 > \cos(x = 1.2) 
[1] 0.3623578
 > \cos(x = pi/2) 
[1] 6.123032e-17
```

tan()

• Package: base

• Input:

x valore numerico

• **Description:** tangente

• Formula:

$$\tan(x) = \frac{\sin(x)}{\cos(x)}$$

```
> tan(x = 1.2)
[1] 2.572152
> tan(x = pi)
[1] -1.224606e-16
> tan(x = 2.3)
[1] -1.119214
> sin(x = 2.3)/cos(x = 2.3)
[1] -1.119214
```

> atan(x = 0.9)

1.9 Funzioni trigonometriche inverse

```
asin()
  • Package: base
  • Input:
        \times valore numerico tale che |x| \leq 1
  • Description: arcoseno di x, espresso in radianti nell'intervallo tra -\pi/2 e \pi/2
  • Formula:
                                                   \arcsin(x)
  • Example:
    > asin(x = 0.9)
    [1] 1.119770
    > asin(x = -1)
    [1] -1.570796
acos()
  • Package: base
  • Input:
         x valore numerico tale che |x| \leq 1
  • Description: arcocoseno di x, espresso in radianti nell'intervallo tra 0 e \pi
  • Formula:
                                                   \arccos(x)
  • Example:
    > acos(x = 0.9)
    [1] 0.4510268
    > acos (x = -1)
    [1] 3.141593
atan()
  • Package: base
  • Input:
         x valore numerico
  • Description: arcotangente di x, espressa in radianti nell'intervallo tra -\pi/2 e \pi/2
  • Formula:
                                                   \arctan(x)
  • Example:
```

```
[1] 0.7328151
> atan(x = -34)
[1] -1.541393
```

atan2()

• Package: base

• Input:

- y valore numerico di ordinata
- x valore numerico di ascissa
- **Description:** arcotangente in radianti dalle coordinate x e y specificate, nell'intervallo tra $-\pi$ e π
- Formula:

 $\arctan(x)$

• Example:

```
> atan2 (y = -2, x = 0.9)

[1] -1.147942

> atan2 (y = -1, x = -1)

[1] -2.356194
```

1.10 Funzioni iperboliche dirette

sinh()

• Package: base

• Input:

x valore numerico

• **Description:** seno iperbolico

• Formula:

$$\sinh(x) = \frac{e^x - e^{-x}}{2}$$

```
> x <- 2.45
> (exp(x) - exp(-x))/2

[1] 5.751027
> sinh(x = 2.45)

[1] 5.751027
> x <- 3.7
> (exp(x) - exp(-x))/2

[1] 20.21129
> sinh(x = 3.7)

[1] 20.21129
```

cosh()

• Package: base

• Input:

x valore numerico

• **Description:** coseno iperbolico

• Formula:

$$\cosh(x) = \frac{e^x + e^{-x}}{2}$$

• Example:

```
> x <- 2.45
> (exp(x) + exp(-x))/2

[1] 5.83732
> cosh(x = 2.45)

[1] 5.83732
> x <- 3.7
> (exp(x) + exp(-x))/2

[1] 20.23601
> cosh(x = 3.7)

[1] 20.23601
```

tanh()

• Package: base

• Input:

x valore numerico

• **Description:** tangente iperbolica

• Formula:

$$\tanh(x) = \frac{\sinh(x)}{\cosh(x)} = \frac{e^{2x} - 1}{e^{2x} + 1}$$

```
> x < -2.45

> (exp(2 * x) - 1)/(exp(2 * x) + 1)

[1] 0.985217

> tanh(x = 2.45)

[1] 0.985217

> x < -3.7

> (exp(2 * x) - 1)/(exp(2 * x) + 1)

[1] 0.9987782
```

```
> tanh(x = 3.7)
    [1] 0.9987782
    > tanh(x = 2.3)
    [1] 0.9800964
    > sinh(x = 2.3)/cosh(x = 2.3)
    [1] 0.9800964
        Funzioni iperboliche inverse
1.11
asinh()
  • Package: base
  • Input:
        x valore numerico
  • Description: inversa seno iperbolico
  • Formula:
                                             arcsinh(x)
  • Example:
    > asinh(x = 2.45)
    [1] 1.628500
    > asinh(x = 3.7)
    [1] 2.019261
acosh()
  • Package: base
  • Input:
        	imes valore numerico tale che x \ge 1
  • Description: inversa coseno iperbolico
  • Formula:
                                             arccosh(x)
  • Example:
    > acosh(x = 2.45)
```

[1] 1.544713

[1] 1.982697

> acosh(x = 3.7)

atanh()

```
• Package: base
```

• Input:

 \mathbf{x} valore numerico tale che |x| < 1

• Description: inversa tangente iperbolica

• Formula:

$$\operatorname{arctanh}(x) = \frac{1}{2} \log \left(\frac{1+x}{1-x} \right)$$

• Example:

```
> x <- 0.45
> 0.5 * log((1 + x)/(1 - x))

[1] 0.4847003

> atanh(x = 0.45)

[1] 0.4847003

> x <- 0.7
> 0.5 * log((1 + x)/(1 - x))

[1] 0.8673005

> atanh(x = 0.7)

[1] 0.8673005
```

1.12 Funzioni esponenziali e logaritmiche

exp()

• Package: base

• Input:

x valore numerico

• **Description:** esponenziale

• Formula:

 e^x

```
> \exp(x = 1.2)
[1] 3.320117
> \exp(x = 0)
[1] 1
```

```
expm1()
```

```
• Package: base
  • Input:
        x valore numerico
  • Description: esponenziale
  • Formula:
                                                 e^x - 1
  • Example:
    > x <- 1.2
    > \exp(x) - 1
    [1] 2.320117
    > expm1(x = 1.2)
    [1] 2.320117
    > x <- 0
    > \exp(x) - 1
    [1] 0
    > expm1(x = 0)
    [1] 0
log2()
  • Package: base
  • Input:
        \times valore numerico tale che x > 0
  • Description: logaritmo di x in base 2
  • Formula:
                                                 \log_2(x)
  • Example:
     > log2(x = 1.2) 
    [1] 0.2630344
     > log2(x = 8) 
    [1] 3
     > log2(x = -1.2) 
    [1] NaN
```

```
log10()
```

```
• Package: base
  • Input:
         \times valore numerico tale che x > 0
  • Description: logaritmo di x in base 10
  • Formula:
                                                  \log_{10}(x)
  • Example:
     > log10(x = 1.2) 
    [1] 0.07918125
     > log10(x = 1000) 
    [1] 3
     > log10(x = -6.4) 
    [1] NaN
log()
  • Package: base
  • Input:
         \times valore numerico tale che x > 0
         base il valore b tale che b>0
  • Description: logaritmo di x in base b
  • Formula:
                                                   \log_b(x)
  • Example:
     > \log(x = 2, base = 4) 
    [1] 0.5
    > log(x = 8, base = 2)
    [1] 3
     > \log(x = 0, base = 10) 
    [1] -Inf
     > \log(x = 100, base = -10) 
    [1] NaN
```

logb()

```
• Package: base
  • Input:
        \times valore numerico tale che x > 0
        base il valore b tale che b>0
  • Description: logaritmo di x in base b
  • Formula:
                                                 \log_b(x)
  • Example:
    > logb(x = 2, base = 4)
    [1] 0.5
    > logb(x = 8, base = 2)
    [1] 3
    > logb(x = -1.2, base = 2)
    [1] NaN
log1p()
  • Package: base
  • Input:
        \times valore numerico tale che x > -1
  • Description: logaritmo di x in base e
  • Formula:
                                                \log(x+1)
  • Example:
    > x < -2.3
    > log(x + 1)
    [1] 1.193922
     > log1p(x = 2.3) 
    [1] 1.193922
    > x <- 8
    > log(x + 1)
    [1] 2.197225
    > log1p(x = 8)
    [1] 2.197225
    > log1p(x = -1)
    [1] -Inf
     > log1p(x = -1.2) 
    [1] NaN
```

• Package: base

• Example: > 1:10

1.13 Funzioni di successione

• **Description:** successione con intervallo unitario

```
:
```

[1] 1 2 3 4 5 6 7 8 9 10 > 1:10.2 [1] 1 2 3 4 5 6 7 8 9 10 > 1.1:10.2 [1] 1.1 2.1 3.1 4.1 5.1 6.1 7.1 8.1 9.1 10.1 > 1:5 + 1 [1] 2 3 4 5 6 > 1: (5 + 1)[1] 1 2 3 4 5 6 rep() • Package: base • Input: \times vettore alfanumerico di dimensione ntimes ogni elemento del vettore viene ripetuto lo stesso numero times di volte length.out dimensione del vettore risultato each ogni elemento del vettore viene ripetuto each volte • **Description:** replicazioni • Example: > rep(x = 2, times = 5) [1] 2 2 2 2 2 > rep(x = c(1, 2, 3), times = 5)[1] 1 2 3 1 2 3 1 2 3 1 2 3 1 2 3 > rep(x = c(8.1, 6.7, 10.2), times = c(1, 2, 3))[1] 8.1 6.7 6.7 10.2 10.2 10.2 > rep(x = c(1, 2, 3), each = 2)[1] 1 1 2 2 3 3

```
> rep(x = c(1, 2, 3), length.out = 7)

[1] 1 2 3 1 2 3 1

> rep(x = TRUE, times = 5)

[1] TRUE TRUE TRUE TRUE

> rep(x = c(1, 2, 3, 4), each = 3, times = 2)

[1] 1 1 1 2 2 2 3 3 3 4 4 4 1 1 1 2 2 2 3 3 3 4 4 4
```

• Note: Il parametro each ha precedenza sul parametro times.

rep.int()

• Package: base

• Input:

 ${\tt x}$ vettore alfanumerico di dimensione n times ogni elemento del vettore viene ripetuto lo stesso numero times di volte

• **Description:** replicazioni

```
> rep.int(x = 2, times = 5)

[1] 2 2 2 2 2

> rep.int(x = c(1, 2, 3), times = 5)

[1] 1 2 3 1 2 3 1 2 3 1 2 3 1 2 3

> rep.int(x = c(1, 2, 3), times = c(1, 2, 3))

[1] 1 2 2 3 3 3

> rep.int(x = TRUE, times = 5)

[1] TRUE TRUE TRUE TRUE TRUE
```

sequence()

- Package: base
- Input:

nvec vettore numerico x di valori naturali di dimensione n

• **Description:** serie di sequenze di interi dove ciascuna sequenza termina con i numeri naturali passati come argomento

• Example:

```
> n1 <- 2
> n2 <- 5
> c(1:n1, 1:n2)

[1] 1 2 1 2 3 4 5

> sequence(nvec = c(2, 5))

[1] 1 2 1 2 3 4 5

> n1 <- 6
> n2 <- 3
> c(1:n1, 1:n2)

[1] 1 2 3 4 5 6 1 2 3

> sequence(nvec = c(6, 3))

[1] 1 2 3 4 5 6 1 2 3
```

seq()

- Package: base
- Input:

```
from punto di partenza to punto di arrivo by passo length.out dimensione along.with vettore di dimensione n per creare la sequenza di valori naturali 1,\,2,\,\ldots,\,n
```

- **Description:** successione
- Example:

```
> seq(from = 1, to = 3.4, by = 0.4)
[1] 1.0 1.4 1.8 2.2 2.6 3.0 3.4
> seq(from = 1, to = 3.4, length.out = 5)
[1] 1.0 1.6 2.2 2.8 3.4
> seq(from = 3.4, to = 1, length.out = 5)
[1] 3.4 2.8 2.2 1.6 1.0
```

```
> x < -c(1.5, 6.4, 9.6, 8.8)
   > n <- 4
   > 1:n
   [1] 1 2 3 4
   > seq(along.with = x)
    [1] 1 2 3 4
   > x < -c(1.5, 6.4, 9.6, 8.8)
   > seq(from = 88, to = 50, along.with = x)
    [1] 88.00000 75.33333 62.66667 50.00000
   > seq(from = 88, to = 50, length.out = length(x))
    [1] 88.00000 75.33333 62.66667 50.00000
   > seq(from = 5, by = -1, along.with = 1:6)
   [1] 5 4 3 2 1 0
   > seq(from = 8)
    [1] 1 2 3 4 5 6 7 8
   > seq(from = -8)
     [1] 1 0 -1 -2 -3 -4 -5 -6 -7 -8
seq_along()
  • Package: base
  • Input:
        along.with \mbox{vettore numerico}\ x\ \mbox{di dimensione}\ n
  • Description: sequenza di valori naturali 1, 2, \ldots, n
  • Example:
   > x < -c(1.2, 2.3, 3.4, 4.5, 5.6, 6.7)
   > n < -6
   > seq_along(along.with = x)
    [1] 1 2 3 4 5 6
```

> x < -c(1.5, 6.4, 9.6, 8.8)

> seq_along(along.with = x)

> n < -4

[1] 1 2 3 4

```
seq_len()
```

```
• Package: base
```

• Input:

length.out valore n naturale

- **Description:** sequenza di valori naturali $1, 2, \ldots, n$
- Example:

```
> n <- 6
> seq_len(length.out = 6)

[1] 1 2 3 4 5 6

> n <- 4
> seq_len(length.out = 4)

[1] 1 2 3 4
```

1.14 Funzioni di ordinamento

sort()

• Package: base

• Input:

```
{\tt x} vettore numerico di dimensione n decreasing = TRUE / FALSE decremento oppure incremento index.return = TRUE / FALSE vettore indici ordinati
```

- Description: ordinamento crescente oppure decrescente
- Output:

```
x vettore ordinato
```

ix vettore indici ordinati

• Formula:

Х

$$\begin{array}{c} \texttt{decreasing = TRUE} \\ \\ x_{(n)}, \, x_{(n-1)}, \, \ldots, \, x_{(1)} \\ \\ \texttt{decreasing = FALSE} \\ \\ x_{(1)}, \, x_{(2)}, \, \ldots, \, x_{(n)} \\ \end{array}$$

```
> x <- c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
> sort(x, decreasing = TRUE, index.return = FALSE)

[1] 4.21 3.40 2.30 2.10 1.20 0.00

> x <- c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
> res <- sort(x, decreasing = TRUE, index.return = TRUE)
> res$x

[1] 4.21 3.40 2.30 2.10 1.20 0.00
```

```
> res$ix
    [1] 3 6 2 5 1 4
   > x[res$ix]
    [1] 4.21 3.40 2.30 2.10 1.20 0.00
   > x \leftarrow c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
   > sort(x, decreasing = FALSE, index.return = FALSE)
    [1] 0.00 1.20 2.10 2.30 3.40 4.21
   > x \leftarrow c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
   > res <- sort(x, decreasing = FALSE, index.return = TRUE)</pre>
    > res$x
    [1] 0.00 1.20 2.10 2.30 3.40 4.21
   > res$ix
    [1] 4 1 5 2 6 3
   > x[res$ix]
    [1] 0.00 1.20 2.10 2.30 3.40 4.21
   > x < -c(1.2, 4.2, 4.5, -5.6, 6.5, 1.2)
   > sort(x, decreasing = TRUE)
    [1] 6.5 4.5 4.2 1.2 1.2 -5.6
   > rev(sort(x))
    [1] 6.5 4.5 4.2 1.2 1.2 -5.6
  • Note: Equivale alla funzione order() quando index.return = TRUE.
rev()
  • Package: base
  • Input:
       	imes vettore numerico di dimensione n
  • Description: elementi di un vettore in ordine invertito
  • Formula:
                                          x_n, x_{n-1}, \ldots, x_1
  • Example:
   > x \leftarrow c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
   > rev(x)
    [1] 3.40 2.10 0.00 4.21 2.30 1.20
   > x < -c(1.2, 4.2, 4.5, -5.6, 6.5, 1.2)
   > rev(x)
    [1] 1.2 6.5 -5.6 4.5 4.2 1.2
```

order()

- Package: base
- Input:

```
{\tt x} vettore numerico di dimensione n decreasing = TRUE / FALSE decremento oppure incremento
```

- **Description:** restituisce la posizione di ogni elemento di x se questo fosse ordinato in maniera decrescente oppure crescente
- Example:

```
> x <- c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
> order(x, decreasing = FALSE)

[1] 4 1 5 2 6 3

> x <- c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
> order(x, decreasing = TRUE)

[1] 3 6 2 5 1 4

> x <- c(1.6, 6.8, 7.7, 7.2, 5.4, 7.9, 8, 8, 3.4, 12)
> sort(x, decreasing = FALSE)

[1] 1.6 3.4 5.4 6.8 7.2 7.7 7.9 8.0 8.0 12.0

> x[order(x, decreasing = FALSE)]

[1] 1.6 3.4 5.4 6.8 7.2 7.7 7.9 8.0 8.0 12.0
```

rank()

- Package: base
- Input:

```
x vettore numerico di dimensione n ties.method = "average" / "first" / "random" / "max" / "min" metodo da utilizzare in presenza di ties
```

- **Description:** rango di x ossia viene associato ad ogni elemento del vettore x il posto occupato nello stesso vettore ordinato in modo crescente
- Example:

```
> x <- c(1.2, 2.3, 4.5, 2.3, 4.5, 6.6, 1.2, 3.4)
> rank(x, ties.method = "average")

[1] 1.5 3.5 6.5 3.5 6.5 8.0 1.5 5.0

> x <- c(1.2, 2.3, 4.21, 0, 2.1, 3.4)
> rank(x, ties.method = "average")

[1] 2 4 6 1 3 5

> x <- c(1.2, 4.2, 4.5, -5.6, 6.5, 1.2)
> rank(x, ties.method = "first")

[1] 2 4 5 1 6 3
```

• Note: Solo per ties.method = "average" e ties.method = "first" la somma del vettore finale rimane uguale a n(n+1)/2.

1.15 Funzioni di troncamento e di arrotondamento

trunc()

• Package: base

• Input:

x valore numerico

• **Description:** tronca la parte decimale

• Formula:

[x]

• Example:

```
> trunc(x = 2)
[1] 2
> trunc(x = 2.999)
[1] 2
> trunc(x = -2.01)
[1] -2
```

floor()

• Package: base

• Input:

x valore numerico

- **Description:** arrotonda all'intero inferiore
- Formula:

$$\lfloor x \rfloor = \begin{cases} x & \text{se } x \text{ è intero} \\ [x] & \text{se } x \text{ è positivo non intero} \\ [x] - 1 & \text{se } x \text{ è negativo non intero} \end{cases}$$

```
> floor(x = 2)
[1] 2
> floor(x = 2.99)
[1] 2
> floor(x = -2.01)
[1] -3
```

ceiling()

• Package: base

• Input:

x valore numerico

• Description: arrotonda all'intero superiore

• Formula:

$$\lceil x \rceil = \left\{ \begin{array}{ll} x & \text{se } x \text{ è intero} \\ [x] + 1 & \text{se } x \text{ è positivo non intero} \\ [x] & \text{se } x \text{ è negativo non intero} \end{array} \right.$$

• Example:

```
> ceiling(x = 2)
[1] 2
> ceiling(x = 2.001)
[1] 3
> ceiling(x = -2.01)
[1] -2
```

round()

• Package: base

• Input:

imes valore numerico digits valore naturale n

 Description: arrotonda al numero di cifre specificato da \boldsymbol{n}

```
> pi
[1] 3.141593
> round(x = pi, digits = 4)
[1] 3.1416
> exp(1)
[1] 2.718282
> round(x = exp(1), digits = 3)
[1] 2.718
```

signif()

```
    Package: base
    Input:

            x valore numerico
            digits valore naturale n

    Description: arrotonda al numero di cifre significative specificate da n
    Example:
```

```
> pi
[1] 3.141593
> signif(x = pi, digits = 4)
[1] 3.142
> exp(1)
[1] 2.718282
> signif(x = exp(1), digits = 3)
[1] 2.72
```

fractions()

- Package: MASS
- Input:
 - x oggetto numerico
- Description: trasforma un valore decimale in frazionario
- Example:

rational()

```
• Package: MASS
```

• Input:

x oggetto numerico

• **Description:** approssimazione razionale

• Example:

1.16 Funzioni avanzate

gamma()

• Package: base

• Input:

 \times valore numerico tale che x > 0

• **Description:** funzione gamma

• Formula:

$$\Gamma(x) = \int_0^{+\infty} u^{x-1} e^{-u} du$$

```
> gamma(x = 3.45)
[1] 3.146312
> gamma(x = 5)
[1] 24
```

lgamma()

• Package: base

• Input:

 \times valore numerico tale che x > 0

• **Description:** logaritmo naturale della funzione gamma

• Formula:

 $\log (\Gamma(x))$

• Example:

```
> log(gamma(x = 3.45))
[1] 1.146231
> lgamma(x = 3.45)

[1] 1.146231
> log(gamma(x = 5))

[1] 3.178054
> lgamma(x = 5)
```

digamma()

• Package: base

• Input:

 \times valore numerico tale che x > 0

• **Description:** funzione digamma

• Formula:

$$\Psi(x) \, = \, \frac{d}{dx} \, \log \left(\Gamma(x) \right)$$

```
> digamma(x = 2.45)

[1] 0.6783387

> digamma(x = 5.3)

[1] 1.570411
```

trigamma()

• Package: base

• Input:

 \times valore numerico tale che x > 0

• Description: derivata prima della funzione digamma

• Formula:

$$\frac{d}{dx}\,\Psi(x)$$

• Example:

```
> trigamma(x = 2.45)

[1] 0.5024545

> trigamma(x = 5.3)

[1] 0.2075909
```

psigamma()

• Package: base

• Input:

```
\mathbf{x} valore numerico tale che x>0 deriv valore naturale n
```

- **Description:** derivata *n*-esima della funzione digamma
- Formula:

$$\frac{d^n}{dx}\,\Psi(x)$$

```
> psigamma(x = 2.45, deriv = 0)
[1] 0.6783387
> digamma(x = 2.45)
[1] 0.6783387
> psigamma(x = 5.3, deriv = 1)
[1] 0.2075909
> trigamma(x = 5.3)
[1] 0.2075909
```

beta()

• Package: base

• Input:

a valore numerico tale che a > 0

b valore numerico tale che b > 0

• Description: funzione beta

• Formula:

$$B(a, b) = \frac{\Gamma(a) \Gamma(b)}{\Gamma(a+b)} = \int_0^1 u^{a-1} (1-u)^{b-1} du$$

• Example:

```
> a <- 3.45
> b <- 2.3
> gamma(a) * gamma(b)/gamma(a + b)

[1] 0.04659344
> beta(a = 3.45, b = 2.3)

[1] 0.04659344
> a <- 5
> b <- 4
> gamma(a) * gamma(b)/gamma(a + b)

[1] 0.003571429
> beta(a = 5, b = 4)

[1] 0.003571429
```

lbeta()

• Package: base

• Input:

a valore numerico tale che a > 0

b valore numerico tale che b>0

- Description: logaritmo naturale della funzione beta
- Formula:

log(B(a, b))

```
> a <- 3.45
> b <- 2.3
> log(gamma(a) * gamma(b)/gamma(a + b))

[1] -3.066296

> lbeta(a = 3.45, b = 2.3)

[1] -3.066296
```

```
> a <- 5
> b <- 4
> log(gamma(a) * gamma(b)/gamma(a + b))

[1] -5.63479

> lbeta(a = 5, b = 4)

[1] -5.63479
```

fbeta()

• Package: MASS

• Input:

```
\times valore numerico tale che x > 0 e x < 1
```

- a valore numerico tale che a>0
- b valore numerico tale che b>0
- Description: funzione beta
- Formula:

$$x^{a-1} (1-x)^{b-1}$$

```
> x <- 0.67
> a <- 3.45
> b <- 2.3
> x^(a - 1) * (1 - x)^(b - 1)

[1] 0.08870567

> fbeta(x = 0.67, a = 3.45, b = 2.3)

[1] 0.08870567

> x <- 0.12
> a <- 5
> b <- 4
> x^(a - 1) * (1 - x)^(b - 1)

[1] 0.0001413100

> fbeta(x = 0.12, a = 5, b = 4)
```

sigmoid()

• Package: e1071

• Input:

x valore numerico

• Description: funzione sigmoide

• Formula:

$$S(x) = (1 + e^{-x})^{-1} = \frac{e^x}{1 + e^x}$$

• Example:

```
> x <- 3.45
> (1 + exp(-x))^(-1)
[1] 0.9692311
> sigmoid(x = 3.45)
[1] 0.9692311
> x <- -1.7
> (1 + exp(-x))^(-1)
[1] 0.1544653
> sigmoid(x = -1.7)
[1] 0.1544653
```

dsigmoid()

• Package: e1071

• Input:

x valore numerico

- Description: derivata prima della funzione sigmoide
- Formula:

$$\frac{d}{dx} S(x) = \frac{e^x}{(1+e^x)^2} = \frac{e^x}{1+e^x} \left(1 - \frac{e^x}{1+e^x}\right) = S(x) \left(1 - S(x)\right)$$

```
> x <- 3.45
> exp(x)/(1 + exp(x))^2
[1] 0.02982214
> dsigmoid(x = 3.45)
[1] 0.02982214
> x <- -1.7
> exp(x)/(1 + exp(x))^2
[1] 0.1306057
> dsigmoid(x = -1.7)
[1] 0.1306057
```

d2sigmoid()

• Package: e1071

• Input:

x valore numerico

• Description: derivata seconda della funzione sigmoide

• Formula:

$$\frac{d^2}{dx}S(x) = \frac{e^x (1 - e^x)}{(1 + e^x)^3} = \frac{e^x}{1 + e^x} \left(1 - \frac{e^x}{1 + e^x} \right) \left(\frac{1}{1 + e^x} - \frac{e^x}{1 + e^x} \right) = S^2(x) (1 - S(x)) (e^{-x} - 1)$$

• Example:

```
> x <- 3.45
> (exp(x) * (1 - exp(x)))/(1 + exp(x))^3

[1] -0.02798695

> d2sigmoid(x = 3.45)

[1] -0.02798695

> x <- -1.7
> (exp(x) * (1 - exp(x)))/(1 + exp(x))^3

[1] 0.09025764

> d2sigmoid(x = -1.7)

[1] 0.09025764
```

besselI()

• Package: base

• Input:

 ${\bf x}~$ valore numerico tale che x>0 nu valore naturale

• **Description:** funzione BesselI

```
> besselI(x = 2.3, nu = 3)
[1] 0.3492232
> besselI(x = 1.6, nu = 2)
[1] 0.3939673
```

```
besselJ()
  • Package: base
  • Input:
        \times valore numerico tale che x > 0
        nu valore naturale
  • Description: funzione BesselJ
  • Example:
    > besselJ(x = 2.3, nu = 3)
    [1] 0.1799789
    > besselJ(x = 1.6, nu = 2)
    [1] 0.2569678
```

besselK()

- Package: base
- Input:
 - \times valore numerico tale che x > 0nu valore naturale
- **Description:** funzione BesselK
- Example:

```
> besselK(x = 2.3, nu = 3)
[1] 0.3762579
> besselK(x = 1.6, nu = 2)
[1] 0.4887471
```

besselY()

- Package: base
- Input:
 - \times valore numerico tale che x > 0nu valore naturale
- **Description:** funzione BesselY
- Example:

```
> besselY(x = 2.3, nu = 3)
[1] -0.8742197
> besselY(x = 1.6, nu = 2)
[1] -0.8548994
```

1.17 Funzioni sui numeri complessi

complex()

```
• Package: base
```

• Input:

```
real parte reale \alpha imaginary parte immaginaria \beta modulus modulo r argument argomento \phi
```

- **Description:** numero complesso
- Formula:

```
\begin{array}{rcl} \alpha + i \, \beta & = & r \left( \cos(\phi) + i \, \sin(\phi) \right) \\ & \alpha & = & r \, \cos(\phi) \\ & \beta & = & r \, \sin(\phi) \\ & r & = & \sqrt{\alpha^2 + \beta^2} \\ & \phi & = & \arctan\left(\frac{\beta}{\alpha}\right) \end{array}
```

• Example:

```
> complex(real = 1, imaginary = 3)
[1] 1+3i
> complex(modulus = Mod(1 + 3i), argument = Arg(1 + 3i))
[1] 1+3i
> complex(real = -3, imaginary = 4)
[1] -3+4i
> complex(modulus = Mod(-3 + 4i), argument = Arg(-3 + 4i))
[1] -3+4i
```

Re()

• Package: base

• Input:

x numero complesso

- **Description**: parte reale
- Formula:

 α

```
> Re(x = 2 + 3i)

[1] 2

> Re(x = -3 + 4i)

[1] -3
```

Im()

• Package: base

• Input:

 \times numero complesso

• **Description:** parte immaginaria

• Formula:

 β

• Example:

```
> Im(x = -2 + 3i)
[1] 3
> Im(x = 3 - 4i)
[1] -4
```

Mod()

• Package: base

• Input:

x numero complesso

• **Description:** modulo

• Formula:

$$r = \sqrt{\alpha^2 + \beta^2}$$

```
> x <- 2 + 3i
> sqrt(2^2 + 3^2)

[1] 3.605551

> Mod(x = 2 + 3i)

[1] 3.605551

> x <- -3 + 4i
> sqrt((-3)^2 + 4^2)

[1] 5

> Mod(x = -3 + 4i)

[1] 5

> x <- 3 + 4i
> sqrt(3^2 + 4^2)

[1] 5

> Mod(x = 3 + 4i)
```

1.17 Funzioni sui numeri complessi

```
[1] 5 > abs(x = 3 + 4i) [1] 5
```

• **Note:** Equivale alla funzione abs().

Arg()

- Package: base
- Input:
 - x numero complesso
- **Description:** argomento
- Formula:

$$\phi = \arctan\left(\frac{\beta}{\alpha}\right)$$

• Example:

```
> x <- 2 + 3i
> atan(3/2)

[1] 0.9827937

> Arg(x = 2 + 3i)

[1] 0.9827937

> x <- 4 + 5i
> atan(5/4)

[1] 0.8960554

> Arg(x = 4 + 5i)

[1] 0.8960554
```

Conj()

• Package: base

• Input:

x numero complesso

- Description: conjugato
- Formula:

 $\alpha - i\beta$

```
> Conj(x = 2 + 3i)

[1] 2-3i

> Conj(x = -3 + 4i)

[1] -3-4i
```

is.real()

```
• Package: base
```

• Input:

x valore numerico

• **Description:** segnalazione di valore numerico reale

• Example:

```
> is.real(x = 2 + 3i)
[1] FALSE
> is.real(x = 4)
[1] TRUE
```

is.complex()

• Package: base

• Input:

x valore numerico

- **Description:** segnalazione di valore numerico complesso
- Example:

```
> is.complex(x = 2 + 3i)
[1] TRUE
> is.complex(x = 4)
[1] FALSE
```

1.18 Funzioni cumulate

cumsum()

```
• Package: base
```

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- **Description:** somma cumulata
- Formula:

$$\sum_{j=1}^{i} x_j \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(1, 2, 4, 3, 5, 6)
> cumsum(x)

[1] 1 3 7 10 15 21
```

```
> x <- c(1, 2.3, 4.5, 6.7, 2.1)
> cumsum(x)

[1] 1.0 3.3 7.8 14.5 16.6
```

cumprod()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• **Description:** prodotto cumulato

• Formula:

$$\prod_{j=1}^{i} x_j \quad \forall i = 1, 2, \dots, n$$

• Example:

cummin()

• Package: base

• Input:

 \times vettore numerico di dimensione n

• **Description:** minimo cumulato

• Formula:

$$\min(x_1, x_2, ..., x_i) \quad \forall i = 1, 2, ..., n$$

```
> x <- c(3, 4, 3, 2, 4, 1)
> cummin(x)

[1] 3 3 3 2 2 1

> x <- c(1, 3, 2, 4, 5, 1)
> cummin(x)
[1] 1 1 1 1 1 1
```

cummax()

• Package: base

• Input:

 \times vettore numerico di dimensione n

• **Description:** massimo cumulato

• Formula:

$$\max(x_1, x_2, ..., x_i) \quad \forall i = 1, 2, ..., n$$

• Example:

```
> x <- c(1, 3, 2, 4, 5, 1)
> cummax(x)

[1] 1 3 3 4 5 5

> x <- c(1, 3, 2, 4, 5, 1)
> cummax(x)

[1] 1 3 3 4 5 5
```

1.19 Funzioni in parallelo

pmin()

• Package: base

• Input:

- ${\bf x}~$ vettore numerico di dimensione n
- ${\bf y}~$ vettore numerico di dimensione n
- **Description:** minimo in parallelo
- Formula:

$$\min(x_i, y_i) \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(1.2, 2.3, 0.11, 4.5)
> y <- c(1.1, 2.1, 1.3, 4.4)
> pmin(x, y)

[1] 1.10 2.10 0.11 4.40

> x <- c(1.2, 2.3, 0.11, 4.5)
> y <- c(1.1, 2.1, 1.1, 2.1)
> pmin(x, y)

[1] 1.10 2.10 0.11 2.10
```

pmax()

```
• Package: base
```

• Input:

```
\mathbf{x} vettore numerico di dimensione n vettore numerico di dimensione n
```

- Description: massimo in parallelo
- Formula:

$$\max(x_i, y_i) \quad \forall i = 1, 2, \dots, n$$

• Example:

```
> x <- c(1.2, 2.3, 0.11, 4.5)
> y <- c(1.1, 2.1, 1.3, 4.4)
> pmax(x, y)

[1] 1.2 2.3 1.3 4.5

> x <- c(1.2, 2.3, 0.11, 4.5)
> y <- c(1.1, 2.1, 1.1, 2.1)
> pmax(x, y)

[1] 1.2 2.3 1.1 4.5
```

1.20 Funzioni di analisi numerica

optimize()

• Package: stats

• Input:

```
f funzione f(x) lower estremo inferiore upper estremo superiore maximum = TRUE / FALSE massimo oppure minimo tol tolleranza
```

- Description: ricerca di un massimo oppure di un minimo
- Output:

```
minimum punto di minimo
maximum punto di massimo
objective valore assunto dalla funzione nel punto individuato
```

• Formula:

$$\max_{x} f(x)$$

$$\max_{x} f(x)$$

$$\max_{x} f(x)$$

$$\min_{x} f(x)$$

```
> f <- function(x) x * exp(-x^3) - (log(x))^2
   > optimize(f, lower = 0.3, upper = 1.5, maximum = TRUE, tol = 1e-04)
   $maximum
   [1] 0.8374697
   $objective
    [1] 0.4339975
   > f <- function(x) (x - 0.1)^2
   > optimize(f, lower = 0, upper = 1, maximum = FALSE, tol = 1e-04)
   $minimum
   [1] 0.1
   $objective
   [1] 7.70372e-34
   > f <- function(x) dchisq(x, df = 8)
   > optimize(f, lower = 0, upper = 10, maximum = TRUE, tol = 1e-04)
   $maximum
   [1] 5.999999
   $objective
    [1] 0.1120209
optim()
  • Package: stats
  • Input:
       par valore di partenza
       fn funzione f(x)
       method = "Nelder-Mead" / "BFGS" / "CG" / "L-BFGS-B" / "SANN" metodo di ottimizzazio-
  • Description: ottimizzazione
  • Output:
       par punto di ottimo
       value valore assunto dalla funzione nel punto individuato
  • Example:
   > f <- function(x) x * exp(-x^3) - (log(x))^2
   > optim(par = 1, fn = f, method = "BFGS")$par
   [1] 20804.91
   > optim(par = 1, fn = f, method = "BFGS")$value
   [1] -98.86214
   > f \leftarrow function(x) (x - 0.1)^2
```

> optim(par = 1, fn = f, method = "BFGS")\$par

[1] 0.1

```
> optim(par = 1, fn = f, method = "BFGS")$value
   [1] 7.70372e-34
   > f \leftarrow function(x) dchisq(x, df = 8)
   > optim(par = 1, fn = f, method = "BFGS")$par
   [1] 0.0003649698
   > optim(par = 1, fn = f, method = "BFGS")$value
    [1] 5.063142e-13
   > nLL <- function(mu, x) {</pre>
         z <- mu * x
         lz \leftarrow log(z)
         L1 <- sum(lz)
         L2 \leftarrow mu/2
          LL <- (L1 - L2)
   +
   + }
   > x <- c(1.2, 3.4, 5.6, 6.1, 7.8, 8.6, 10.7, 12, 13.7, 14.7)
   > optim(par = 10000, fn = nLL, method = "CG", x = x)$par
   [1] 9950.6
   > optim(par = 10000, fn = nLL, method = "CG", x = x)$value
    [1] 4863.693
uniroot()
  • Package: stats
  • Input:
       f funzione f(x)
       lower estremo inferiore
       upper estremo superiore
       tol tolleranza
       maxiter mumero massimo di iterazioni
```

• Description: ricerca di uno zero

• Output:

root radice
f.root valore assunto dalla funzione nel punto individuato
iter numero di iterazioni
estim.prec tolleranza

• Formula:

$$f(x) = 0$$

```
> f <- function(x) exp(-x) - x 
> uniroot(f, lower = 0, upper = 1, tol = 1e-04, maxiter = 1000)
```

```
$root
[1] 0.5671439
$f.root
[1] -9.448109e-07
$iter
[1] 3
$estim.prec
[1] 7.425e-05
> f <- function(x) log10(x) + x
> uniroot(f, lower = 0.1, upper = 1, tol = 1e-04, maxiter = 1000)
$root
[1] 0.3990136
$f.root
[1] 1.279136e-06
$iter
[1] 5
$estim.prec
[1] 5e-05
```

polyroot()

• Package: stats

• Input:

a vettore dei k coefficienti di un polinomio di ordine k-1

- Description: ricerca di uno zero in un polinomio
- Formula:

$$a_1 + a_2 x + a_3 x^2 + \dots + a_k x^{k-1} = 0$$

```
> k <- 3
> a1 <- 3
> a2 <- -2
> a3 <- 2
> a <- c(a1, a2, a3)
> polyroot(a)

[1] 0.5+1.118034i 0.5-1.118034i

> radice1 <- 0.5 + (0+1.118034i)
> a1 + a2 * radice1 + a3 * radice1^2

[1] -5.0312e-08+0i

> radice2 <- 0.5 - (0+1.118034i)
> a1 + a2 * radice2 + a3 * radice2^2
```

```
> k <- 4
> a1 <- 3
> a2 < -2
> a3 <- 2
> a4 <- -1
> a <- c(a1, a2, a3, a4)
> polyroot(a)
[1] 0.094732+1.283742i 0.094732-1.283742i 1.810536+0.000000i
> radice1 <- 0.09473214 + (0+1.283742i)</pre>
> a1 + a2 * radice1 + a3 * radice1^2 + a4 * radice1^3
[1] 7.477461e-07-5.808714e-07i
> radice2 <- 0.09473214 - (0+1.283742i)</pre>
> a1 + a2 * radice2 + a3 * radice2^2 + a4 * radice2^3
[1] 7.477461e-07+5.808714e-07i
> radice3 <- 1.81053571 + (0+0i)</pre>
> a1 + a2 * radice3 + a3 * radice3^2 + a4 * radice3^3
[1] 1.729401e-08+0i
```

D()

• Package: stats

• Input:

expr espressione contenente la funzione f(x) da derivare name variabile x di derivazione

- Description: derivata simbolica al primo ordine
- Formula:

$$\frac{d}{dx} f(x)$$

```
> D(expr = expression(exp(-x) - x), name = "x")
-(exp(-x) + 1)
> D(expr = expression(x * exp(-a)), name = "x")
exp(-a)
```

DD()

- Package:
- Input:

```
expr espressione contenente la funzione f(x) da derivare name variabile x di derivazione order il valore k dell'ordine di derivazione
```

- **Description:** derivata simbolica al k-esimo ordine
- Formula:

$$\frac{d^k}{d^k x} f(x)$$

• Example:

```
> DD(expr = expression(exp(-x) - x), name = "x", order = 1)
> DD(expr = expression(x \star exp(-a)), name = "a", order = 2)
```

integrate()

- Package: stats
- Input:

```
f funzione f(x) lower estremo inferiore a di integrazione upper estremo superiore b di integrazione subdivisions mumero di suddivisioni dell'intervallo di integrazione
```

- **Description:** integrazione numerica
- Output:

value integrale definito

• Formula:

$$\int_{a}^{b} f(x) dx$$

```
> f <- function(x) exp(-x)
> integrate(f, lower = 1.2, upper = 2.3, subdivisions = 150)

0.2009354 with absolute error < 2.2e-15

> f <- function(x) sqrt(x)
> integrate(f, lower = 2.1, upper = 4.5, subdivisions = 150)

4.335168 with absolute error < 4.8e-14

> f <- function(x) dnorm(x)
> integrate(f, lower = -1.96, upper = 1.96, subdivisions = 150)

0.9500042 with absolute error < 1.0e-11</pre>
```

1.21 Costanti

```
pi
```

Package: baseDescription: pi greco

• Formula:

• Example:

```
> pi
[1] 3.141593
> 2 * pi
[1] 6.283185
```

Inf

• Package:

• **Description:** infinito

• Formula:

 $\pm \infty$

 π

• Example:

```
> 2/0
[1] Inf
> -2/0
[1] -Inf
> 0^Inf
[1] 0
> exp(-Inf)
[1] 0
> 0/Inf
[1] 0
> Inf - Inf
[1] NaN
> Inf/Inf
[1] NaN
> exp(Inf)
```

[1] Inf

NaN

- Package:
- **Description:** not a number
- Example:

```
> Inf - Inf
[1] NaN
> 0/0
[1] NaN
```

NA

- Package:
- **Description:** not available
- Example:

```
> x <- c(1.2, 3.4, 5.6, NA)
> mean(x)

[1] NA
> mean(x, na.rm = TRUE)

[1] 3.4
```

NULL

- Package:
- **Description:** oggetto nullo
- Example:

```
> x <- c(1.2, 3.4, 5.6)
> names(x) <- c("a", "b", "c")
> names(x) <- NULL
> x

[1] 1.2 3.4 5.6
```

TRUE

- Package:
- Description: vero
- Example:

```
> TRUE | TRUE
[1] TRUE
> TRUE & TRUE
[1] TRUE
```

T

• Package: base

• **Description:** vero

• Example:

> T

[1] TRUE

> T & T

[1] TRUE

FALSE

• Package:

• **Description:** falso

• Example:

```
> FALSE | TRUE
```

[1] TRUE

> FALSE & TRUE

[1] FALSE

F

• Package: base

• **Description:** falso

• Example:

> F

[1] FALSE

> F | T

[1] TRUE

1.22 Miscellaneous

list()

• Package: base

• Description: creazione di un oggetto lista

```
> x < -c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1)
> y < -c(4.5, 5.4, 6.1, 6.1, 5.4)
> lista <- list(x = x, y = y)
> lista
$x
[1] 7.8 6.6 6.5 7.4 7.3 7.0 6.4 7.1
$y
[1] 4.5 5.4 6.1 6.1 5.4
> lista[1]
$x
[1] 7.8 6.6 6.5 7.4 7.3 7.0 6.4 7.1
> lista$x
[1] 7.8 6.6 6.5 7.4 7.3 7.0 6.4 7.1
> lista[[1]]
[1] 7.8 6.6 6.5 7.4 7.3 7.0 6.4 7.1
> lista[[1]][1]
[1] 7.8
> lista[2]
[1] 4.5 5.4 6.1 6.1 5.4
> lista$y
[1] 4.5 5.4 6.1 6.1 5.4
> lista[[2]]
[1] 4.5 5.4 6.1 6.1 5.4
> lista[[2]][1]
[1] 4.5
> x < -c(1, 2.3, 4.5, 6.7, 8.9)
> y < -c(154, 109, 137, 115, 140)
> z <- c(108, 115, 126, 92, 146)
> lista <- list(x = x, y = y, z = z)
> lista
```

```
[1] 1.0 2.3 4.5 6.7 8.9
[1] 154 109 137 115 140
$z
[1] 108 115 126 92 146
> lista[1]
$x
[1] 1.0 2.3 4.5 6.7 8.9
> lista$x
[1] 1.0 2.3 4.5 6.7 8.9
> lista[[1]]
[1] 1.0 2.3 4.5 6.7 8.9
> lista[[1]][1]
[1] 1
> lista[2]
$у
[1] 154 109 137 115 140
> lista$y
[1] 154 109 137 115 140
> lista[[2]]
[1] 154 109 137 115 140
> lista[[2]][1]
[1] 154
> lista[3]
[1] 108 115 126 92 146
> lista$z
[1] 108 115 126 92 146
> lista[[3]]
[1] 108 115 126 92 146
> lista[[3]][1]
```

```
[1] 108
    > x < -c(1, 2, 3)
    > y <- c(11, 12, 13, 14, 15)
    > lista <- list(x, y)</pre>
    > lista
    [[1]]
    [1] 1 2 3
    [[2]]
    [1] 11 12 13 14 15
    > names(lista)
   NULL
   > x < -c(1, 2, 3)
    > y <- c(11, 12, 13, 14, 15)
    > lista <- list(A = x, B = y)
    > lista
    $A
    [1] 1 2 3
    [1] 11 12 13 14 15
   > names(lista)
    [1] "A" "B"
lapply()
  • Package: base
  • Input:
        x oggetto lista
        FUN funzione
  • Description: applica la funzione FUN ad ogni elemento di lista
  • Example:
    > vec1 <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1)
    > mean(vec1)
    [1] 7.0125
    > vec2 <- c(4.5, 5.4, 6.1, 6.1, 5.4)
    > mean (vec2)
    [1] 5.5
    > x <- list(vec1 = vec1, vec2 = vec2)
```

> lapply(x, FUN = mean)

```
$vec1
[1] 7.0125
$vec2
[1] 5.5
> vec1 <- c(1, 2.3, 4.5, 6.7, 8.9)
> sd(vec1)
[1] 3.206556
> vec2 <- c(154, 109, 137, 115, 140)
> sd(vec2)
[1] 18.61451
> vec3 <- c(108, 115, 126, 92, 146)
> sd(vec3)
[1] 20.19406
> x <- list(vec1 = vec1, vec2 = vec2, vec3 = vec3)
> lapply(x, FUN = sd)
$vec1
[1] 3.206556
$vec2
[1] 18.61451
$vec3
[1] 20.19406
```

.Last.value

• Package: base

• **Description:** ultimo valore calcolato

• Example:

identical()

```
• Package: base
```

• Description: uguaglianza tra due oggetti

• Example:

```
> u <- c(1, 2, 3)
> v <- c(1, 2, 4)
> if (identical(u, v)) print("uguali") else print("non uguali")

[1] "non uguali"

> u <- c(1, 2, 3)
> v <- c(1, 3, 2)
> identical(u, v)
[1] FALSE
```

any()

• Package: base

• Input:

imes vettore numerico di dimensione n

- Description: restituisce TRUE se almeno un elemento del vettore soddisfa ad una condizione fissata
- Example:

```
> x <- c(3, 4, 3, 2, 4, 1)
> x < 2

[1] FALSE FALSE FALSE FALSE FALSE TRUE

> any(x < 2)

[1] TRUE

> x <- c(1, 2, 3, 4, 5, 6, 7, 8)
> x > 4

[1] FALSE FALSE FALSE FALSE TRUE TRUE TRUE

> any(x > 4)

[1] TRUE
```

all()

- Package: base
- Input:
 - \times vettore numerico di dimensione n
- Description: restituisce TRUE se tutti gli elementi del vettore soddisfano ad una condizione fissata
- Example:

```
> x <- c(3, 4, 3, 2, 4, 1)
> x < 2

[1] FALSE FALSE FALSE FALSE FALSE TRUE

> all(x < 2)

[1] FALSE

> x <- c(1, 2, 3, 4, 5, 6, 7, 8)
> x > 4

[1] FALSE FALSE FALSE FALSE TRUE TRUE TRUE

> all(x > 4)

[1] FALSE
```

match()

- Package: base
- Input:
- ullet Description: per ogni elemento di x restituisce la posizione della prima occorrenza in y
- Example:

```
> x <- c(1, 1, 1, 2, 2, 2, 3, 3, 3, 4, 4, 4, 5, 5, 5)
> match(x, table = c(2, 4), nomatch = 0)

[1] 0 0 0 1 1 1 0 0 0 2 2 2 0 0 0

> x <- c(1, 1, 1, 2, 2, 2, 3, 3, 3, 4, 4, 4, 5, 5, 5)
> match(x, table = c(2, 4), nomatch = NA)

[1] NA NA NA 1 1 1 NA NA NA 2 2 2 NA NA NA

> match(x = c(-3, 3), table = c(5, 33, 3, 6, -3, -4, 3, 5, -3), nomatch = NA)

[1] 5 3
```

outer()

```
• Package: base
```

• Input:

```
X vettore numerico x di dimensione n Y vettore numerico y di dimensione m FUN funzione f(x,\,y)
```

- **Description:** applica la funzione FUN ad ogni coppia ordinata costituita da un elemento di x ed uno di y
- Formula:

$$f(x_i, y_i) \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., m$$

• Example:

```
> outer(X = c(1, 2, 2, 4), Y = c(1.2, 2.3), FUN = "+")

[,1] [,2]
[1,] 2.2 3.3
[2,] 3.2 4.3
[3,] 3.2 4.3
[4,] 5.2 6.3

> outer(X = c(1, 2, 2, 4), Y = c(1.2, 2.3), FUN = "*")

[,1] [,2]
[1,] 1.2 2.3
[2,] 2.4 4.6
[3,] 2.4 4.6
[4,] 4.8 9.2
```

expression()

- Package: base
- Input:
 - x oggetto
- Description: crea una espressione simbolica
- Example:

```
> u <- c(4.3, 5.5, 6.8, 8)
> w <- c(4, 5, 6, 7)
> z <- expression(x = u/w)
> z

expression(x = u/w)

> u <- c(1.2, 3.4, 4.5)
> w <- c(1, 2, 44)
> z <- expression(x = u * w)
> z

expression(x = u * w)
```

eval()

```
• Package: base
```

• Input:

expr espressione simbolica

- Description: valuta una espressione simbolica
- Example:

```
> u <- c(4.3, 5.5, 6.8, 8)
> w <- c(4, 5, 6, 7)
> z <- expression(x = u/w)
> eval(expr = z)

[1] 1.075000 1.100000 1.133333 1.142857

> u <- c(1.2, 3.4, 4.5)
> w <- c(1, 2, 44)
> z <- expression(expr = u * w)
> eval(z)

[1] 1.2 6.8 198.0
```

replace()

• Package: base

• Input:

 ${\tt x}$ vettore numerico di dimensione n list indice dell'elemento da rimpiazzare values valore da inserire

- **Description:** rimpiazza un elemento del vettore x
- Example 1:

```
> x <- c(1, 2, 3, 4, 5, 6, 7, 8)
> replace(x, list = 1, values = 10)

[1] 10  2  3  4  5  6  7  8
> x

[1] 1 2 3 4 5 6 7 8
```

• Example 2:

```
> x <- c(1.2, 3.4, 5.6, 7.8)
> replace(x, list = 3, values = 8.9)

[1] 1.2 3.4 8.9 7.8
> x

[1] 1.2 3.4 5.6 7.8
```

• **Note:** Il vettore *x* rimane invariato.

```
e
```

• Package: base • Description: scrittura rapida di un valore numerico potenza di 10 • Example: > 1e3 [1] 1000 > -2e-2 [1] -0.02> 1e-2 [1] 0.01 > 3e4 [1] 30000 even() • Package: gtools • Input: x valore naturale • **Description:** verifica numero pari • Example: > even(x = 22) [1] TRUE > even(x = 7) [1] FALSE odd() • Package: gtools • Input: x valore naturale • **Description:** verifica numero dispari

- Example:

```
> odd (x = 22)
[1] FALSE
> odd(x = 7)
[1] TRUE
```

• Package: base

• Example:

• Description: notazione polacca inversa (RPN)

,

```
> 1 + 2
    [1] 3
    > 3 * 4.2
    [1] 12.6
  • Note: RPN = Reverse Polish Notation.
gcd()
  • Package: schoolmath
  • Input:
        x valore naturale
        y valore naturale
  • Description: massimo comun divisore
  • Example:
     > \gcd(x = 6, y = 26) 
    [1] 2
     > \gcd(x = 8, y = 36) 
    [1] 4
scm()
  • Package: schoolmath
  • Input:
        x valore naturale
        y valore naturale
  • Description: minimo comune multiplo
  • Example:
    > scm(6, 14)
    [1] 42
    > scm(12, 16)
    [1] 48
```

```
is.vector()
```

```
• Package: base
• Input:
     x oggetto
• Description: oggetto di tipo vettore
• Example 1:
 > x < -c(1.2, 2.34, 4.5, 6.7, 8.9)
 > is.vector(x)
 [1] TRUE
 > is.matrix(x)
 [1] FALSE
• Example 2:
 > x <- matrix(data = 1:12, nrow = 3, ncol = 4)
  [,1] [,2] [,3] [,4]
 [1,] 1 4 7 10
[2,] 2 5 8 11
 [3,] 3 6 9
                      12
 > is.vector(x)
 [1] FALSE
 > is.matrix(x)
 [1] TRUE
• Example 3:
 > x <- matrix(data = 1:12, nrow = 3, ncol = 4)
    [,1] [,2] [,3] [,4]
 [1,] 1 4 7 10
[2,] 2 5 8 11
 [3,] 3 6
                9 12
 > is.vector(x)
 [1] FALSE
 > is.matrix(x)
 [1] TRUE
```

```
is.matrix()
```

```
• Package: base
• Input:
     x oggetto
• Description: oggetto di tipo matrice
• Example 1:
 > x < -c(1.2, 2.34, 4.5, 6.7, 8.9)
 > is.vector(x)
 [1] TRUE
 > is.matrix(x)
 [1] FALSE
• Example 2:
 > x <- matrix(data = 1:12, nrow = 3, ncol = 4)
 > x
     [,1] [,2] [,3] [,4]
 [1,] 1 4 7 10
[2,] 2 5 8 11
 [3,]
      3
           6
                     12
 > is.vector(x)
 [1] FALSE
 > is.matrix(x)
 [1] TRUE
• Example 3:
 > x <- matrix(data = 1:12, nrow = 3, ncol = 4)
      [,1] [,2] [,3] [,4]
 [1,] 1 4 7 10
      2 5
 [2,]
                8
                     11
 [3,]
        3
             6
                       12
 > is.vector(x)
 [1] FALSE
 > is.matrix(x)
 [1] TRUE
```

Capitolo 2

• Package: base

> x

Vettori, Matrici ed Arrays

2.1 Creazione di Vettori

```
c()
```

```
• Input:
     ... oggetti da concatenare
     recursive = TRUE / FALSE concatenazione per oggetti di tipo list()
• Description: funzione di concatenazione
• Example:
 > x < -c(1.2, 3.4, 5.6, 7.8)
 [1] 1.2 3.4 5.6 7.8
 > x < -c(x, 9.9)
 [1] 1.2 3.4 5.6 7.8 9.9
 > x < -c(1.2, 3.4, 5.6, 7.8)
 [1] 1.2 3.4 5.6 7.8
 > x[5] < -9.9
 [1] 1.2 3.4 5.6 7.8 9.9
 > x <- c("a", "b")</pre>
 [1] "a" "b"
 > x <- c("a", "b")
 [1] "a" "b"
 > x <- c("a", "b", "a", "a", "b")
```

```
[1] "a" "b" "a" "a" "b"
> x < -c(x, "a")
> x
[1] "a" "b" "a" "a" "b" "a"
> x <- c("a", "b", "a", "a", "b")
[1] "a" "b" "a" "a" "b"
> x[6] <- "a"
> x
[1] "a" "b" "a" "a" "b" "a"
> x < -c("a", 1)
> x
[1] "a" "1"
> x < -c(x, 2)
[1] "a" "1" "2"
> lista <- list(primo = c(1, 2, 3), secondo = c(1.2, 5.6))
> lista
$primo
[1] 1 2 3
$secondo
[1] 1.2 5.6
> vettore <- c(lista, recursive = TRUE)</pre>
> vettore
  primo1 primo2 primo3 secondo1 secondo2
           2.0 3.0 1.2 5.6
    1.0
> y <- 1.2
> z < - y[-1]
> z
numeric(0)
```

- Note 1: Se il vettore è molto lungo, conviene utilizzare la funzione scan().
- Note 2: I vettori alfanumerici possono essere definiti usando " oppure '.

scan()

```
• Package: base
```

• Input:

```
what = double(0) / "character" tipo dei dati numerico oppure carattere
```

- **Description:** creazione di un vettore
- Example:

```
> x <- scan(what = double(0))
> x <- scan(what = "character")</pre>
```

[]

- Package: base
- Input:
 - ${\bf x}~$ vettore alfanumerico di dimensione n
- **Description:** estrazione di elementi da un vettore
- Example:

```
> x < -c(1.2, 3.4, 5.6, 7.8, 9, 9.9)
[1] 1.2 3.4 5.6 7.8 9.0 9.9
> x[2]
[1] 3.4
> x[c(1, 3, 4)]
[1] 1.2 5.6 7.8
> x[1:3]
[1] 1.2 3.4 5.6
> x[-c(1:3)]
[1] 7.8 9.0 9.9
> x[-(1:3)]
[1] 7.8 9.0 9.9
> x[x %in% c(1.2, 7.8)]
[1] 1.2 7.8
> x[x > 6.3]
[1] 7.8 9.0 9.9
```

> x[x > 6.3 & x < 9.7]

names()

- Package: base
- Input:
 - \times vettore numerico di dimensione n
- Description: assegnazioni di nomi agli elementi di un vettore
- Example:

```
> x < -c(1.2, 3.4, 5.6)
> names(x)
NULL
> names(x) <- c("primo", "secondo", "terzo")</pre>
  primo secondo terzo
    1.2
        3.4 5.6
> names(x)
[1] "primo" "secondo" "terzo"
> x[c("primo", "terzo")]
primo terzo
  1.2
      5.6
> names(x) <- NULL</pre>
> names(x)
NULL
```

vector()

```
• Package: base
```

• Input:

```
mode = "numeric" / "complex" / "logical" tipo di oggetto length valore <math>n della dimensione
```

- Example:

```
> x <- vector(mode = "numeric", length = 5)
> x

[1] 0 0 0 0 0

> x <- vector(mode = "complex", length = 3)
> x

[1] 0+0i 0+0i 0+0i

> x <- vector(mode = "logical", length = 4)
> x

[1] FALSE FALSE FALSE FALSE
```

numeric()

- Package: base
- Input:

length dimensione

- Example:

```
> x <- numeric(length = 5)
> x

[1] 0 0 0 0 0

> x <- numeric(length = 4)
> x

[1] 0 0 0 0
```

complex()

```
• Package: base
```

• Input:

length dimensione

- Example:

```
> x <- complex(length = 5)
> x

[1] 0+0i 0+0i 0+0i 0+0i 0+0i
> x <- complex(length = 4)
> x

[1] 0+0i 0+0i 0+0i 0+0i
```

logical()

• Package: base

• Input:

length dimensione

- **Description:** inizializzazione di un vettore logico di dimensione n
- Example:

```
> x <- logical(length = 5)
> x

[1] FALSE FALSE FALSE FALSE FALSE
> x <- logical(length = 4)
> x

[1] FALSE FALSE FALSE FALSE
```

head()

```
• Package: utils
```

- Input:
 - imes vettore numerico di dimensione m
 - n numero di elementi
- **Description:** seleziona i primi n elementi
- Example:

```
> x <- c(1.2, 3.2, 3.3, 2.5, 5, 5.6)
> head(x, n = 2)

[1] 1.2 3.2
> x <- c(4.5, 6.7, 8.9, 7.7, 11.2)
> head(x, n = 3)

[1] 4.5 6.7 8.9
```

tail()

- Package: utils
- Input:
 - ${\bf x}~$ vettore numerico di dimensione m
 - n numero di elementi
- **Description:** seleziona gli ultimi n elementi
- Example:

```
> x <- c(1.2, 3.2, 3.3, 2.5, 5, 5.6)
> tail(x, n = 3)

[1] 2.5 5.0 5.6
> x <- c(4.5, 6.7, 8.9, 7.7, 11.2)
> tail(x, n = 2)

[1] 7.7 11.2
```

%**o**%

- Package: base
- Input:
 - imes vettore numerico di dimensione n
 - y vettore numerico di dimensione m
- **Description:** prodotto esterno
- Formula:

$$x_i y_i \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., m$$

• Example:

```
> x < -c(1, 2, 3, 4)
> n <- 4
> y < -c(1.2, 3.4)
> m <- 2
> x %o% y
     [,1] [,2]
[1,] 1.2 3.4
[2,] 2.4 6.8
[3,] 3.6 10.2
[4,] 4.8 13.6
> x < -c(3, 4, 7)
> n < -3
> y < -c(1.1, 2.2, 3.3)
> m < - 3
> x %0% y
    [,1] [,2] [,3]
[1,] 3.3 6.6 9.9
[2,] 4.4 8.8 13.2
[3,] 7.7 15.4 23.1
```

append()

```
• Package: base
```

• Input:

```
{\tt x} vettore numerico di dimensione n values valore v numerico after valore j naturale
```

- **Description:** aggiunge un elemento ad un vettore
- Formula:

```
 \begin{array}{c} \texttt{after} \leq 0 \\ \\ v, \, x_1, \, x_2, \, \ldots, \, x_n \\ \\ \hline \texttt{after} \geq n \\ \\ x_1, \, x_2, \, \ldots, \, x_n, \, v \\ \\ \hline \hline 1 \leq \texttt{after} \leq n-1 \\ \\ x_1, \, x_2, \, \ldots, \, x_j, \, v, \, x_{j+1}, \, x_{j+2}, \, \ldots, \, x_n \end{array}
```

• Example:

```
> x <- c(1.2, 3.4, 5.6)
> append(x, values = 6, after = -2)

[1] 6.0 1.2 3.4 5.6

> x <- c(1.2, 3.4, 5.6)
> append(x, values = 6, after = 2)

[1] 1.2 3.4 6.0 5.6

> x <- c(1.2, 3.4, 5.6)
> append(x, values = 6, after = 7)

[1] 1.2 3.4 5.6 6.0
```

sapply()

• Package: base

• Input:

 ${\tt X}$ vettore numerico di dimensione n ${\tt FUN}$ funzione scelta

- ullet **Description:** applica FUN ad ogni elemento del vettore X
- Example:

```
> sapply(X = c(1.2, 3.2, 4.5, 6.7), FUN = sin)

[1] 0.93203909 -0.05837414 -0.97753012 0.40484992

> sapply(X = c(1.2, 3.2, 4.5, 6.7), FUN = log)
```

```
[1] 0.1823216 1.1631508 1.5040774 1.9021075
> a < -c(2, 4, 7, 3, 5, 2, 9, 0)
> X < -c(2, 4, 6)
> myfun <- function(x) which(a > x)
> sapply(X, FUN = myfun)
[[1]]
[1] 2 3 4 5 7
[[2]]
[1] 3 5 7
[[3]]
[1] 3 7
> x < -c(1.5, 6.4, 9.6, 8.8, 7.7, 2.2, 4.8)
> sapply(X = 1:5, FUN = function(i) sample(x, size = 3, replace = FALSE))
     [,1] [,2] [,3] [,4] [,5]
[1,] 9.6 8.8 2.2 1.5 7.7
[2,] 1.5 9.6 9.6
                    7.7 9.6
[3,] 8.8 6.4 7.7 9.6 6.4
> x <- matrix(data = c(2, 3, 4, 5, 5, 4, 1, 3, 4, 7, 6, 5, 12,
     13, 4, 11, 21, 10, 9, 7), nrow = 4, ncol = 5)
    [,1] [,2] [,3] [,4] [,5]
          5
[1,]
                4 12
      2
      3
                 7
                    13
            4
                         10
[2,]
[3,]
      4
           1
                 6
                     4
[4,]
       5
            3
                 5
                     11
                          7
> fattore <- factor(c(1, 2, 2, 1), labels = letters[1:2])
> fattore
[1] a b b a
Levels: a b
> sapply(X = 1:ncol(x), FUN = function(i) tapply(x[, i], INDEX = fattore,
    FUN = mean))
 [,1] [,2] [,3] [,4] [,5]
a 3.5 4.0 4.5 11.5 14.0
b 3.5 2.5 6.5 8.5 9.5
> myfun <- function(x) prod(1:x)</pre>
> sapply(X = 1:5, myfun)
[1] 1 2 6 24 120
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> sumsq <- function(b, xv = x, yv = y) {
    yhat <- 1.2 + b * xv
+
     sum((yv - yhat)^2)
+ }
> b <- seq(0, 2, by = 0.05)
> sapply(X = b, FUN = sumsq)
```

```
[1] 367.20560 339.53785 313.06340 287.78225 263.69440 240.79985 219.09860 [8] 198.59065 179.27600 161.15465 144.22660 128.49185 113.95040 100.60225 [15] 88.44740 77.48585 67.71760 59.14265 51.76100 45.57265 40.57760 [22] 36.77585 34.16740 32.75225 32.53040 33.50185 35.66660 39.02465 [29] 43.57600 49.32065 56.25860 64.38985 73.71440 84.23225 95.94340 [36] 108.84785 122.94560 138.23665 154.72100 172.39865 191.26960
```

subset()

• Package: base

• Input:

```
{\bf x}~ vettore numerico di dimensione n subset selezione
```

- **Description:** sottoinsieme del vettore x
- Example 1:

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> subset(x, subset = x > 7.5)
[1] 7.8 7.6
```

• Example 2:

```
> x <- c(7.8, 6.6, 6.5, 6.6)
> subset(x, subset = x == 6.6)

[1] 6.6 6.6
```

2.2 Creazione di Matrici

matrix()

• Package: base

• Input:

```
data vettore numerico di dimensione n\,m nrow numero n di righe ncol numero m di colonne byrow = TRUE / FALSE elementi disposti per riga oppure per colonna dimnames etichette di riga e di colonna
```

- Description: definizione di una matrice
- Example:

```
> n < -3
> m < - 2
> x < -c(1, -0.2, 3, 4, 5.6, 6.7)
> A <- matrix(data = x, nrow = n, ncol = m, byrow = FALSE)
> A
   [,1] [,2]
[1,] 1.0 4.0
[2,] -0.2 5.6
[3,] 3.0 6.7
> n <- 2
> m < - 3
> x <- 0
> A <- matrix(data = x, nrow = n, ncol = m)
> A
   [,1] [,2] [,3]
[1,] 0 0 0
[2,] 0 0 0
> n <- 2
> m < - 3
> x <- 1
> A <- matrix(data = x, nrow = n, ncol = m)
   [,1] [,2] [,3]
[1,] 1 1
[2,] 1 1
> n <- 3
> m <- 3
> x <- 1:9
> riga <- c("r1", "r2", "r3")</pre>
> colonna <- c("c1", "c2", "c3")</pre>
> A <- matrix(data = x, nrow = n, ncol = m, byrow = FALSE, dimnames = list(riga,
+ colonna))
> A
  c1 c2 c3
r1 1 4 7
r2 2 5 8
r3 3 6 9
```

dim()

• Package: base

• Input:

 \times vettore numerico di dimensione nm

- **Description:** dimensione
- Example:

```
> n <- 3
> m <- 3
> x <- 1:9
> dim(x) <- c(n, m)
> x
```

```
[,1] [,2] [,3]
[1,] 1 4 7
[2,] 2 5 8
[3,] 3 6 9

> n <- 1
> m <- 5
> x <- 1:5
> dim(x) <- c(n, m)
> x

[,1] [,2] [,3] [,4] [,5]
[1,] 1 2 3 4 5
```

rownames()

• Package: base

• Input:

 \times matrice di dimensione $n \times m$

- Description: etichette di riga
- Example:

```
> x < -matrix(data = c(1, 3, 5, 2, 4, 1), nrow = 2, ncol = 3, byrow = TRUE)
> x
   [,1] [,2] [,3]
[1,] 1 3 5
[2,]
      2 4
                1
> rownames(x)
NULL
> rownames(x) <- c("r1", "r2")</pre>
  [,1] [,2] [,3]
r1 1 3 5
r2 2 4 1
> rownames(x)
[1] "r1" "r2"
> x <- matrix(data = c(1, 4, 2, 3, 3, 2, 4, 1, 3.4, 4.3, 4.56,
+ 11.1), nrow = 3, ncol = 4)
> x
   [,1] [,2] [,3] [,4]
[1,] 1 3 4.0 4.30
[2,] 4 3 1.0 4.56
      2 2 3.4 11.10
[3,]
> rownames(x)
NULL
```

NULL

```
> rownames(x) <- c("r1", "r2", "r3")</pre>
      [,1] [,2] [,3] [,4]
        1 3 4.0 4.30
   r1
            3 1.0 4.56
   r2
        4
   r3
        2
             2 3.4 11.10
   > rownames(x)
   [1] "r1" "r2" "r3"
colnames()
  • Package: base
  • Input:
       \times matrice di dimensione n \times m
  • Description: etichette di colonna
  • Example:
   > x <- matrix(data = c(1, 3, 5, 2, 4, 1), nrow = 2, ncol = 3, byrow = TRUE)
       [,1] [,2] [,3]
   [1,] 1 3 5
[2,] 2 4 1
   > colnames(x)
   NULL
   > colnames(x) <- c("c1", "c2", "c3")</pre>
   > x
       c1 c2 c3
   [1,] 1 3 5
   [2,] 2 4 1
   > colnames(x)
   [1] "c1" "c2" "c3"
   > x <- matrix(data = c(1, 4, 2, 3, 3, 2, 4, 1, 3.4, 4.3, 4.56,
   + 11.1), nrow = 3, ncol = 4)
   > x
       [,1] [,2] [,3] [,4]
   [1,] 1 3 4.0 4.30
   [2,]
          4 3 1.0 4.56
             2 3.4 11.10
   [3,]
         2
   > colnames(x)
```

```
> colnames(x) <- c("c1", "c2", "c3", "c4")
       c1 c2 c3 c4
   [1,] 1 3 4.0 4.30 [2,] 4 3 1.0 4.56
   [3,] 2 2 3.4 11.10
   > colnames(x)
   [1] "c1" "c2" "c3" "c4"
dimnames()
  • Package: base
  • Input:
       {\bf x} matrice di dimensione n \times m
  • Description: etichette di riga e di colonna
  • Example:
   > x <- matrix(data = 1:9, nrow = 3, ncol = 3)
       [,1] [,2] [,3]
   [1,] 1 4 7
           2
                 5
   [2,]
   [3,]
           3 6
   > dimnames(x)
   NULL
   > dimnames(x) <- list(c("r1", "r2", "r3"), c("c1", "c2", "c3"))</pre>
   > x
      c1 c2 c3
   r1 1 4 7
   r2 2 5 8
   r3 3 6 9
   > dimnames(x)
   [[1]]
   [1] "r1" "r2" "r3"
```

[[2]]

[1] "c1" "c2" "c3"

```
[]
```

- Package: base
- Input:

A matrice di dimensione $n \times m$

- Description: estrazione di elementi da una matrice
- Example:

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)
> dimnames(A) <- list(c("r1", "r2", "r3"), c("c1", "c2", "c3"))
> n <- 3
> m < - 3
> A[2, 3]
[1] 8
> A[1, ]
c1 c2 c3
1 4 7
> A["r1", ]
c1 c2 c3
1 4 7
> A[, 3]
r1 r2 r3
7 8 9
> A[, "c3"]
r1 r2 r3
7 8 9
> A[c(1, 2), ]
  c1 c2 c3
r1 1 4 7
r2 2 5 8
> A[c("r1", "r2"), ]
  c1 c2 c3
r1 1 4 7
r2 2 5 8
> A[, c(2, 3)]
  c2 c3
r1 4 7
r2 5 8
r3 6 9
> A[, c("c2", "c3")]
```

```
c2 c3
r1 4 7
r2 5 8
r3 6 9
> A[-1, ]
 c1 c2 c3
r2 2 5 8
r3 3 6 9
> A[, -3]
  c1 c2
r1 1 4
r2 2 5
r3 3 6
> A[A[, "c2"] > 4.1, ]
  c1 c2 c3
r2 2 5 8
r3 3 6 9
> x[x > 3]
[1] 4 5 6 7 8 9
> A <- matrix(data = c(1.2, 3.4, 5.6, 7.8, 9.1), nrow = 1, ncol = 5)
> is.matrix(A)
[1] TRUE
> myvec <- A[1, ]</pre>
> is.vector(myvec)
[1] TRUE
> myvec2 <- A[, 1]
> is.vector(myvec2)
[1] TRUE
> myvec3 <- A[1, , drop = FALSE]</pre>
> is.vector(myvec3)
[1] FALSE
> is.matrix(myvec3)
[1] TRUE
```

col()

- Package: base
- Input:

data $\mbox{matrice di dimensione } n \times m$

- Description: colonna di appartenenza di ogni elemento
- Example:

```
> x <- matrix(data = 1:9, nrow = 3, ncol = 3)
    [,1] [,2] [,3]
[1,] 1 4
      2
           5
                8
[2,]
[3,]
    3 6
> n <- 3
> m <- 3
> col(x)
    [,1] [,2] [,3]
[1,] 1 2 3
      1
          2
                3
[2,]
      1 2
               3
[3,]
> x <- matrix(data = c(1.1, 2.3, 4.5, 6.7, 8.8, 6.1), nrow = 2,
+ ncol = 3)
> x
    [,1] [,2] [,3]
[1,] 1.1 4.5 8.8
[2,] 2.3 6.7 6.1
> n < -2
> m < - 3
> col(x)
    [,1] [,2] [,3]
[1,] 1 2 3
[2,] 1 2 3
```

row()

- Package: base
- Input:

data matrice di dimensione $n \times m$

- Description: riga di appartenenza di ogni elemento
- Example:

```
> n < -3
 > m <- 3
 > row(x)
     [,1] [,2] [,3]
 [1,] 1 1 1
[2,] 2 2 2
                  2
 [3,]
        3 3 3
 > x <- matrix(data = c(1.1, 2.3, 4.5, 6.7, 8.8, 6.1), nrow = 2,
 > n <- 2
 > m < - 3
 > row(x)
     [,1] [,2] [,3]
 [1,] 1 1 1
[2,] 2 2 2
• Package: utils
• Input:
```

head()

data matrice di dimensione $k \times m$ n numero di righe

- **Description:** seleziona le prime n righe
- Example:

```
> x <- matrix(data = 1:9, nrow = 3, ncol = 3)
> x
    [,1] [,2] [,3]
[1,] 1 4 7
[2,] 2 5 8
[3,] 3 6 9
> k <- 3
> m < - 3
> head(x, n = 2)
    [,1] [,2] [,3]
[1,] 1 4 7
[2,] 2 5 8
> x <- matrix(data = 1:9, nrow = 3, ncol = 3, byrow = TRUE)
> x
 [,1] [,2] [,3]
[1,] 1 2 3
[2,] 4 5 6
[3,] 7 8 9
> k <- 3
> m < - 3
> head(x, n = 2)
    [,1] [,2] [,3]
[1,] 1 2 3
[2,] 4 5 6
```

tail()

```
• Package: utils
```

• Input:

```
data \mbox{ matrice di dimensione } k \times m n \mbox{ numero di righe}
```

- **Description:** seleziona le ultime n righe
- Example:

```
> x <- matrix(data = 1:9, nrow = 3, ncol = 3)
    [,1] [,2] [,3]
[1,] 1 4
      2
           5
                8
[2,]
    3 6 9
[3,]
> k < - 3
> m <- 3
> tail(x, n = 2)
   [,1] [,2] [,3]
[2,] 2 5 8
[3,]
      3 6
                9
> x <- matrix(data = 1:9, nrow = 3, ncol = 3, byrow = TRUE)</pre>
> k < - 3
> m < - 3
> tail(x, n = 2)
   [,1] [,2] [,3]
[2,] 4 5 6
[3,] 7 8 9
```

vech()

- Package: fUtilities
- Input:

```
\times matrice di dimensione m \times n
```

- Description: seleziona gli elementi della sezione triangolare inferiore di una matrice simmetrica
- Example:

```
> x <- matrix(data = c(1, 2, 3, 4, 2, 4, 5, 6, 3, 5, 7, 8, 4, 6,
+ 8, 9), nrow = , <math>ncol = 4)
> x
    [,1] [,2] [,3] [,4]
[1,] 1 2
                3
      2
           4
                5
                     6
[2,]
     3
         5
[3,]
                7
                    8
         6
              8
      4
[4,]
> vech(x)
```

```
[,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
[1,] 1
                        4 5
                   4
> x < -matrix(data = c(11, 12, 13, 12, 14, 15, 13, 15, 16), nrow = 3,
+ ncol = 3)
> x
    [,1] [,2] [,3]
[1,] 11
         12
              13
[2,] 12
         14
             15
[3,] 13
          15
              16
> vech(x)
    [,1] [,2] [,3] [,4] [,5] [,6]
[1,] 11 12
              13
                  14
                       15
```

xpnd()

- Package: MCMCpack
- Input:

```
x vettore numerico di dimensione n\left(n+1\right)/2 nrow numero n di righe
```

- **Description:** crea una matrice simmetrica a partire da un vettore
- Example:

```
> xpnd(x = c(1, 2, 3, 4, 4, 5, 6, 7, 8, 9), nrow = 4)
    [,1] [,2] [,3] [,4]
[1,]
     1
           2
                3
[2,]
      2
            4
                 5
                     6
[3,]
     3
            5
                7
                    8
[4,]
      4
            6
                8
                    9
> xpnd(x = c(11, 12, 13, 14, 15, 16), nrow = 3)
    [,1] [,2] [,3]
[1,] 11 12 13
              15
          14
[2,]
      12
           15
               16
[3,]
     13
```

length()

- Package: base
- Input:

A matrice di dimensione $n \times m$

- Description: numero di elementi
- Formula:

 $n\,m$

• Example:

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)
     [,1] [,2] [,3]
     1 2
[1,]
          4
             5
                 8
[2,]
[3,]
       3
             6
                  9
> n <- 3
> m <- 3
> n * m
[1] 9
> length(A)
[1] 9
> A <- matrix(data = c(1.2, 4.5, 2.3, 3.1), nrow = 2, ncol = 2)
> A
     [,1] [,2]
[1,] 1.2 2.3 [2,] 4.5 3.1
> n <- 2
> m < - 2
> n * m
[1] 4
> length(A)
[1] 4
```

cbind()

• Package: base

• Input:

A matrice di dimensione $n \times m$

 ${\tt B}\;$ matrice di dimensione $n\times k$

- **Description:** unisce due matrici accostandole per colonna
- Example:

```
> A <- matrix(data = c(9.9, 1, 12), nrow = 3, ncol = 1)
> A

[,1]
[1,] 9.9
[2,] 1.0
[3,] 12.0

> B <- matrix(data = 1:3, nrow = 3, ncol = 1)
> B
```

```
[,1]
[1,] 1
[2,] 2
     3
[3,]
> n <- 3
> m < -1
> k <- 1
> cbind(A, B)
     [,1] [,2]
[1,] 9.9
[2,] 1.0
           1
2
[3,] 12.0
> A <- matrix(data = 1:2, nrow = 2, ncol = 1)
> A
   [,1]
[1,] 1
[2,] 2
> B <- matrix(data = 3:4, nrow = 2, ncol = 1)
> B
    [,1]
[1,] 3
[2,] 4
> n < -2
> m <- 1
> k <- 1
> cbind(A, B)
    [,1] [,2]
[1,] 1 3
[2,] 2 4
```

rbind()

- Package: base
- Input:
 - A matrice di dimensione $n \times m$
 - ${\,{\,}^{\,}\!{\,}^{\,}}$ matrice di dimensione $k\times m$
- Description: unisce due matrici accostandole per riga
- Example:

```
[,1] [,2] [,3]
[1,] 1 2 3
> n <- 1
> m < - 3
> k <- 1
> rbind(A, B)
    [,1] [,2] [,3]
[1,] 9.9 1 12
[2,] 1.0 2 3
> A <- matrix(data = 1:2, nrow = 2, ncol = 1)
    [,1]
[1,] 1
[2,] 2
> B <- matrix(data = 3:4, nrow = 2, ncol = 1)
> B
   [,1]
[1,] 3
[2,] 4
> n < -2
> m < -1
> k <- 2
> rbind(A, B)
   [,1]
[1,] 1
[2,] 2
[2,]
       3
[3,]
     4
[4,]
```

toeplitz()

• Package: stats

• Input:

data $\,$ vettore numerico di dimensione n

- Description: matrice simmetrica di Toeplitz di dimensione $n \times n$
- Example:

hilbert()

• Package: fUtilities

• Input:

n valore n naturale

- Description: matrice di Hilbert
- Formula:

$$1/(i+j-1) \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., n$$

[3,] 0.3333333 0.2500000 0.2000000 0.1666667 0.1428571 0.12500000 0.11111111 [4,] 0.2500000 0.2000000 0.1666667 0.1428571 0.1250000 0.11111111 0.10000000 [5,] 0.2000000 0.1666667 0.1428571 0.1250000 0.1111111 0.10000000 0.09090909 [6,] 0.1666667 0.1428571 0.1250000 0.1111111 0.1000000 0.09090909 0.08333333 [7,] 0.1428571 0.1250000 0.1111111 0.1000000 0.0909091 0.08333333 0.07692308

• Example:

> n < -5

```
> hilbert(n)
                    [,2]
                              [,3]
                                         [,4]
          [,1]
[1,] 1.0000000 0.5000000 0.3333333 0.2500000 0.2000000
[2,] 0.5000000 0.3333333 0.2500000 0.2000000 0.1666667
[3,] 0.3333333 0.2500000 0.2000000 0.1666667 0.1428571
[4,] 0.2500000 0.2000000 0.1666667 0.1428571 0.1250000
[5,] 0.2000000 0.1666667 0.1428571 0.1250000 0.1111111
> n < -7
> hilbert(n)
                    [,2]
                               [,3]
                                         [,4]
                                                   [,5]
[1,] 1.0000000 0.5000000 0.3333333 0.2500000 0.2000000 0.16666667 0.14285714
[2,] 0.5000000 0.3333333 0.2500000 0.2000000 0.1666667 0.14285714 0.12500000
```

• Package: futilities

pascal()

```
• Input:

n valore n naturale
```

• Description: matrice di Pascal

• Example:

```
> n <- 5
> pascal(n)
    [,1] [,2] [,3] [,4] [,5]
[1,]
     1 1
             1
                  1
         2
              3
                      5
[2,]
     1
                  4
[3,]
     1
         3
              6
                 10
                      15
     1
         4 10
                  20
                      35
[4,]
         5 15
     1
                  35
                      70
[5,]
> n < -7
> pascal(n)
    [,1] [,2] [,3] [,4] [,5] [,6] [,7]
[1,]
      1
         1
              1
                  1
                      1
                           1
          2
              3
                      5
                               7
[2,]
      1
                  4
                           6
         3
[3,]
      1
              6
                 10
                      15
                          21
                              28
     1 4 10 20
1 5 15 35
                     35
                         56
                             84
[4,]
                     70 126 210
[5,]
[6,]
     1 6 21 56 126 252 462
    1 7 28 84 210 462 924
[7,]
```

2.3 Operazioni sulle Matrici

rk()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times n$

- Description: rango cioé il numero di righe (colonne) linearmente indipendenti
- Example:

```
[,1] [,2] [,3]
    [1,] 1.2 6.5 2.3
[2,] 2.3 7.6 4.5
[3,] 4.5 1.1 6.7
    > n <- 3
    > rk(A)
    [1] 3
det()
  • Package: base
  • Input:
        A matrice di dimensione n \times n
  • Description: determinante
  • Formula:
                                                 det(A)
  • Example:
    > A <- matrix(data = c(1, 4, -0.2, 5.6), nrow = 2, ncol = 2)
    > A
         [,1] [,2]
    [1,] 1 -0.2
[2,] 4 5.6
    > n < -2
    > det(A)
    [1] 6.4
    > A <- matrix(data = c(1.2, 2.3, 4.5, 6.5, 7.6, 1.1, 2.3, 4.5,
    + 6.7), nrow = 3, ncol = 3)
    > A
         [,1] [,2] [,3]
    [1,] 1.2 6.5 2.3 [2,] 2.3 7.6 4.5
    [3,] 4.5 1.1 6.7
    > n < -3
    > det(A)
```

[1] 13.783

determinant()

```
• Package: base
```

• Input:

A matrice di dimensione $n \times n$

logarithm = TRUE / FALSE logaritmo naturale del modulo del determinante

- **Description:** determinante
- Output:

```
modulus modulo sign segno
```

• Formula:

```
\begin{array}{c} & & & \\ & & & \\ & & & \\ & & & \\ & & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ & & \\ &
```

```
> A <- matrix(data = c(1, 4, -0.2, 5.6), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1 -0.2
       4 5.6
[2,]
> n <- 2
> abs(det(A))
[1] 6.4
> determinant(A, logarithm = FALSE)$modulus
[1] 6.4
attr(,"logarithm")
[1] FALSE
> sign(det(A))
[1] 1
> determinant(A, logarithm = FALSE)$sign
[1] 1
> A <- matrix(data = c(1.2, 4.5, 6.7, 8.9, 4.5, 6.6, 7.8, 7.5,
     3.3), \text{ nrow} = 3, \text{ ncol} = 3)
> A
```

```
[,1] [,2] [,3]
[1,] 1.2 8.9 7.8
[2,] 4.5 4.5 7.5
[3,] 6.7 6.6 3.3

> n <- 3
> abs(det(A))

[1] 269.97

> determinant(A, logarithm = FALSE)$modulus

[1] 269.97
attr(,"logarithm")
[1] FALSE

> sign(det(A))

[1] 1

> determinant(A, logarithm = FALSE)$sign

[1] 1
```

determinant.matrix()

- Package: base
- Input:

 ${\tt A} \ \ {\bf matrice} \ {\bf di} \ {\bf dimensione} \ n \times n$

logarithm = TRUE / FALSE logaritmo naturale del modulo del determinante

- **Description:** determinante
- Output:

modulus modulo sign segno

• Formula:

```
\begin{array}{c} & & \\ & \log \operatorname{rithm} = \operatorname{TRUE} \\ \\ & \log \left( | \det(A)| \right) \\ \\ & \operatorname{sign} \left( \det(A) \right) \\ \\ & \log \operatorname{rithm} = \operatorname{FALSE} \\ \\ & \operatorname{modulus} \\ & | \det(A)| \\ \\ & \operatorname{sign} \left( \det(A) \right) \end{array}
```

```
> A <- matrix(data = c(1, 4, -0.2, 5.6), nrow = 2, ncol = 2) > A
```

```
[,1] [,2]
[1,] 1 -0.2
[2,] 4 5.6
> n <- 2
> abs(det(A))
[1] 6.4
> determinant.matrix(A, logarithm = FALSE)$modulus
[1] 6.4
attr(,"logarithm")
[1] FALSE
> sign(det(A))
[1] 1
> determinant.matrix(A, logarithm = FALSE)$sign
[1] 1
> A <- matrix(data = c(1.2, 4.5, 6.7, 8.9, 4.5, 6.6, 7.8, 7.5,
    3.3), nrow = 3, ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.2 8.9 7.8
[2,] 4.5 4.5 7.5
[3,] 6.7 6.6 3.3
> n < -3
> abs(det(A))
[1] 269.97
> determinant.matrix(A, logarithm = FALSE)$modulus
[1] 269.97
attr(,"logarithm")
[1] FALSE
> sign(det(A))
[1] 1
> determinant.matrix(A, logarithm = FALSE)$sign
[1] 1
```

tr()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times n$

• **Description:** traccia

• Formula:

$$\sum_{i=1}^{n} a_{i,i}$$

• Example:

```
> A <- matrix(data = c(1, 4, 2, 8), nrow = 2, ncol = 2)
   [,1] [,2]
[1,] 1 2
[2,] 4 8
> n <- 2
> tr(A)
[1] 9
> A <- matrix(data = c(1.2, 2.3, 4.5, 6.5, 7.6, 1.1, 2.3, 4.5,
+ 6.7), nrow = 3, ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.2 6.5 2.3
[2,] 2.3 7.6 4.5
[3,] 4.5 1.1 6.7
> n <- 3
> tr(A)
[1] 15.5
```

norm()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times m$

p = 1 / 2 / Inf massima somma assoluta di colonna, radice quadrata del massimo autovalore della matrice A^T A, massima somma assoluta di riga

- **Description:** norma
- Formula:

$$\max \left(\sum_{i=1}^{n} |a_{i,j}| \right) \quad \forall j = 1, 2, \dots, m$$

$$p = 2$$

$$\max_{i} (\lambda_i) \quad \forall i = 1, 2, \dots, m$$

$$\max\left(\sum_{j=1}^{m} |a_{i,j}|\right) \quad \forall i = 1, 2, \dots, n$$

```
> n <- 2
> m < - 2
> A <- matrix(data = c(2.2, 3.4, 0.2, -1.2), nrow = 2, ncol = 2,
+ byrow = FALSE)
> A
    [,1] [,2]
[1,] 2.2 0.2
[2,] 3.4 -1.2
> \max(abs(2.2) + abs(3.4), abs(0.2) + abs(-1.2))
[1] 5.6
> norm(A, p = 1)
[1] 5.6
> autovalori <- eigen(t(A) %*% A)$values</pre>
> sqrt(max(autovalori))
[1] 4.152189
> norm(A, p = 2)
[1] 4.152189
> \max(abs(2.2) + abs(0.2), abs(3.4) + abs(-1.2))
[1] 4.6
> norm(A, p = Inf)
[1] 4.6
```

isPositiveDefinite()

```
• Package: fUtilities
```

• Input:

 \times matrice di dimensione $n \times n$

• **Description:** matrice definita positiva

```
> A <- matrix(data = c(1, 4, -0.2, 5.6), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1 -0.2
[2,] 4 5.6
> n < -2
> isPositiveDefinite(A)
[1] TRUE
> A <- matrix(data = c(1.2, 2.3, 4.5, 6.5, 7.6, 1.1, 2.3, 4.5,
     6.7), nrow = 3, ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.2 6.5 2.3
[2,] 2.3 7.6 4.5
[3,] 4.5 1.1 6.7
> n <- 3
> isPositiveDefinite(A)
[1] TRUE
> A <- matrix(data = c(-1, 1, 1, -1), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] -1 1
[2,] 1 -1
> n <- 2
> isPositiveDefinite(A)
[1] FALSE
```

as.vector()

- Package: base
- Input:
 - A matrice di dimensione $n \times m$
- **Description:** trasforma la matrice in vettore di dimensione nm seguendo l'ordine delle colonne
- Example:

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)</pre>
     [,1] [,2] [,3]
[1,] 1 4
[2,] 2 5
[3,] 3 6
                   8
> n < - 3
> m < - 3
> as.vector(A)
[1] 1 2 3 4 5 6 7 8 9
> A <- matrix(data = c(1.2, 2.3, 6.5, 7.6), nrow = 2, ncol = 2)
     [,1] [,2]
[1,] 1.2 6.5
[2,] 2.3 7.6
> n < -2
> m < - 2
> as.vector(A)
[1] 1.2 2.3 6.5 7.6
```

solve()

- Package: base
- Input:
 - A matrice invertibile di dimensione $n \times n$
 - ${\mathtt B}$ matrice di dimensione $n \times k$
- Description: matrice inversa oppure soluzione di un sistema quadrato lineare
- Formula:

$$A^{-1}$$
 $A^{-1} B$

```
> A <- matrix(data = c(1, -0.2, 4, 5.6), nrow = 2, ncol = 2)
> A

[,1] [,2]
[1,] 1.0 4.0
[2,] -0.2 5.6
```

```
> n <- 2
> invA <- solve(A)</pre>
> A %*% invA
             [,1] [,2]
[1,] 1.000000e+00 0
[2,] 1.109952e-17 1
[1,] 1.000000e+00
> invA %*% A
            [,1]
                         [,2]
[1,] 1.00000e+00 2.220446e-16
[2,] 5.20417e-18 1.000000e+00
> A <- matrix(data = c(1, -0.2, 4, 5.6), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1.0 4.0
[2,] -0.2 5.6
> B <- c(11, -2)
[1] 11 -2
> n <- 2
> k <- 1
> solve(A, B)
[1] 10.87500 0.03125
> solve(A) %*% B
         [,1]
[1,] 10.87500
[2,] 0.03125
> A <- matrix(data = c(1, -0.2, 4, 5.6), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1.0 4.0
[2,] -0.2 5.6
> B <- matrix(data = c(11, -2, 13, 4.1), nrow = 2, ncol = 2)
> B
    [,1] [,2]
[1,] 11 13.0
[2,] -2 4.1
> n < -2
> k <- 2
> solve(A, B)
         [,1] [,2]
[1,] 10.87500 8.812500
[2,] 0.03125 1.046875
```

eigen()

• Package: base

• Input:

A matrice simmetrica di dimensione $n \times n$ only.values = TRUE / FALSE calcola i soli autovalori

• Description: autovalori ed autovettori

nrow = 3, ncol = 3)

• Output:

values la diagonale della matrice D degli autovalori di dimensione $n \times n$ vectors matrice ortogonale Γ degli autovettori di dimensione $n \times n$

> A <- matrix(data = c(1.2, 3, 5.6, 3, 4, 6.7, 5.6, 6.7, 9.8),

• Formula:

$$A = \Gamma D \Gamma^T$$

dove
$$\Gamma^T \Gamma = I_n = \Gamma \Gamma^T$$
 e $D = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_n)$

```
> A
     [,1] [,2] [,3]
[1,] 1.2 3.0 5.6
[2,] 3.0 4.0 6.7
[3,] 5.6 6.7 9.8
> n < -3
> D <- diag(eigen(A)$values)</pre>
        [,1]
                   [,2]
                             [,3]
[1,] 16.77455 0.0000000 0.000000
[2,] 0.00000 -0.1731794 0.000000
[3,] 0.00000 0.0000000 -1.601373
> GAMMA <- eigen(A) $vectors
> GAMMA
                     [,2]
          [,1]
                                [,3]
[1,] -0.3767594 0.3675643 0.8502640
[2,] -0.4980954 -0.8542951 0.1485966
[3,] -0.7809951 0.3675274 -0.5049458
> GAMMA %*% D %*% t(GAMMA)
    [,1] [,2] [,3]
[1,] 1.2 3.0 5.6
[2,] 3.0 4.0 6.7
[3,] 5.6 6.7 9.8
> A <- matrix(data = c(1.2, 2.3, 2.3, 2.2), nrow = 2, ncol = 2)
    [,1] [,2]
[1,] 1.2 2.3
[2,] 2.3 2.2
```

crossprod()

• Package: base

• Input:

- A matrice di dimensione $n \times m$
- B matrice di dimensione $n \times k$
- Description: prodotto scalare

[2,] 53.12 69.89 109.26 [3,] 81.70 109.26 172.29

• Formula:

$$A^T A \qquad A^T B$$

```
> A <- matrix(data = c(1.2, 3, 5.6, 3, 4, 6.7, 5.6, 6.7, 9.8),
+ nrow = 3, ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.2 3.0 5.6
[2,] 3.0 4.0 6.7
[3,] 5.6 6.7 9.8
> n <- 3
> m < - 3
> t(A) %*% A
     [,1] [,2] [,3]
[1,] 41.80 53.12 81.70
[2,] 53.12 69.89 109.26
[3,] 81.70 109.26 172.29
> crossprod(A)
     [,1]
          [,2] [,3]
[1,] 41.80 53.12 81.70
```

```
> A <- matrix(data = c(1.2, 3, 5.6, 3, 4, 6.7, 5.6, 6.7, 9.8),
+ nrow = 3, ncol = 3
> A
    [,1] [,2] [,3]
[1,] 1.2 3.0 5.6
[2,] 3.0 4.0 6.7
[3,] 5.6 6.7 9.8
> B <- matrix(data = c(11, -2, 3.4, 4.1, 5, 7), nrow = 3, ncol = 2)
> B
    [,1] [,2]
[1,] 11.0 4.1
[2,] -2.0 5.0
[3,] 3.4 7.0
> n < -3
> m < - 3
> k <- 2
> t(A) %*% B
           [,2]
      [,1]
          59.12
[1,] 26.24
[2,] 47.78 79.20
[3,] 81.52 125.06
> crossprod(A, B)
           [,2]
     [,1]
[1,] 26.24 59.12
[2,] 47.78 79.20
[3,] 81.52 125.06
```

tcrossprod()

• Package: base

• Input:

A matrice di dimensione $n \times m$

 ${\tt B}$ matrice di dimensione $k \times m$

• **Description:** prodotto scalare

• Formula:

$$A\,A^T \qquad A\,B^T$$

```
> n < -3
> m <- 3
> A % * % t(A)
      [,1]
           [,2] [,3]
[1,] 41.80 53.12 81.70
[2,] 53.12 69.89 109.26
[3,] 81.70 109.26 172.29
> tcrossprod(A)
[,1] [,2] [,3]
[1,] 41.80 53.12 81.70
[2,] 53.12 69.89 109.26
[3,] 81.70 109.26 172.29
> A <- matrix(data = c(1.2, 3, 5.6, 3, 4, 6.7, 5.6, 6.7, 9.8),
+ nrow = 3, ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.2 3.0 5.6
[2,] 3.0 4.0 6.7
[3,] 5.6 6.7 9.8
> B <- matrix(data = c(11, 4.1, -2, 5, 3.4, 7), nrow = 2, ncol = 3)
> B
    [,1] [,2] [,3]
[1,] 11.0 -2 3.4 [2,] 4.1 5 7.0
> n <- 3
> m <- 3
> k < - 2
> A %*% t(B)
      [,1] [,2]
[1,] 26.24 59.12
[2,] 47.78 79.20
[3,] 81.52 125.06
> tcrossprod(A, B)
      [,1] [,2]
[1,] 26.24 59.12
[2,] 47.78 79.20
[3,] 81.52 125.06
```



- Package: base
- Input:
 - A matrice di dimensione $n \times m$
 - B matrice di dimensione $n \times m$
- Description: prodotto di Hadamard
- Formula:

$$x_i y_j \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., m$$

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)</pre>
   [,1] [,2] [,3]
    1
          4
[1,]
      2
            5
[2,]
                 8
[3,]
     3
            6
                 9
> B <- matrix(data = c(4.1, 2.3, 4.1, 5.4, 4.6, 4.2, 2.1, 3.2,
+ 4.3), nrow = 3, ncol = 3)
    [,1] [,2] [,3]
[1,] 4.1 5.4 2.1
[2,] 2.3 4.6 3.2
[3,] 4.1 4.2 4.3
> n < -3
> m < - 3
> A * B
   [,1] [,2] [,3]
[1,] 4.1 21.6 14.7
[2,] 4.6 23.0 25.6
[3,] 12.3 25.2 38.7
> A <- matrix(data = c(1, 2, 3, 5), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1
      2
[2,]
> B <- matrix(data = c(1.1, 2.3, 4.5, 6.7), nrow = 2, ncol = 2)
> B
    [,1] [,2]
[1,] 1.1 4.5
[2,] 2.3 6.7
> n <- 2
> m < - 2
> A * B
    [,1] [,2]
[1,] 1.1 13.5
[2,] 4.6 33.5
```

```
%*%
```

- Package: base
- Input:
 - A matrice di dimensione $n \times m$
 - ${\mathbb B}$ matrice di dimensione $m \times k$
- **Description:** prodotto scalare
- Formula:

AB

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> B <- matrix(data = c(11, -1, 3.4, 4.1, 5, 7), nrow = 3, ncol = 2)
    [,1] [,2]
[1,] 11.0 4.1
[2,] -1.0 5.0
[3,] 3.4 7.0
> n <- 3
> m < - 3
> k < - 2
> A %*% B
           [,2]
     [,1]
[1,] 40.66 93.40
[2,] -4.40 34.18
[3,] 66.00 135.30
> A <- matrix(data = 1:2, nrow = 1, ncol = 2)
> A
    [,1] [,2]
[1,] 1 2
> B <- matrix(data = 3:4, nrow = 2, ncol = 1)
> B
    [,1]
[1,] 3
[2,] 4
> n <- 1
> m < - 2
> k <- 1
> A %*% B
    [,1]
[1,] 11
```

kronecker()

- Package: base
- Input:
 - A matrice di dimensione $n \times m$
 - ${\tt B}$ matrice di dimensione $h \times k$
- Description: prodotto di Kronecker
- Formula:

$$A \otimes B = \left(\begin{array}{ccc} a_{1,1}B & \cdots & a_{1,m}B \\ \vdots & \vdots & \vdots \\ a_{n,1}B & \cdots & a_{n,m}B \end{array}\right)$$

```
> A <- matrix(data = 1:3, nrow = 3, ncol = 1)</pre>
> A
   [,1]
[1,] 1
[2,] 2
[2,]
[3,]
> B <- matrix(data = 7:9, nrow = 1, ncol = 3)
[,1] [,2] [,3]
[1,] 7 8 9
> n <- 3
> m <- 1
> h <- 1
> k <- 3
> kronecker(A, B)
    [,1] [,2] [,3]
[1,] 7 8 9
[2,]
     14
            16
                 18
[3,]
     21
            24
                 27
> A <- matrix(data = 1:2, nrow = 1, ncol = 2)</pre>
> A
    [,1] [,2]
[1,] 1 2
> B <- matrix(data = 3:4, nrow = 2, ncol = 1)
   [,1]
[1,] 3
[2,] 4
> n <- 1
> m < - 2
> h <- 2
> k <- 1
> kronecker(A, B)
    [,1] [,2]
[1,] 3 6
[2,] 4 8
```



- Package: base
- Input:
 - A matrice di dimensione $n \times m$
 - B matrice di dimensione $h \times k$
- **Description:** prodotto di *Kronecker*
- Formula:

$$A \otimes B = \left(\begin{array}{ccc} a_{1,1}B & \cdots & a_{1,m}B \\ \vdots & \vdots & \vdots \\ a_{n,1}B & \cdots & a_{n,m}B \end{array}\right)$$

```
> A <- matrix(data = 1:3, nrow = 3, ncol = 1)
   [,1]
[1,] 1
[2,] 2
[3,]
> B <- matrix(data = 7:9, nrow = 1, ncol = 3)
> B
    [,1] [,2] [,3]
[1,] 7 8 9
> n < -3
> m <- 1
> h <- 1
> k <- 3
> A %x% B
   [,1] [,2] [,3]
[1,]
      7 8 9
           16 18
[2,]
     14
[3,]
     21
           24
               27
> A <- matrix(data = 1:2, nrow = 1, ncol = 2)</pre>
   [,1] [,2]
[1,] 1 2
> B <- matrix(data = 3:4, nrow = 2, ncol = 1)
   [,1]
[1,] 3
[2,] 4
> n <- 1
> m <- 2
> h <- 2
> k <- 1
> A %x% B
    [,1] [,2]
[1,] 3 6
[2,] 4 8
```

diag()

• Package: base

• Input:

A matrice di dimensione $n \times n$

imes vettore numerico di dimensione n

- h valore naturale
- Description: estrae gli elementi diagonali o crea una matrice diagonale
- Example:

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)</pre>
> A
   [,1] [,2] [,3]
    1
         4
[1,]
     2
          5 8
[2,]
[3,]
    3
         6
              9
> n < -3
> diag(A)
[1] 1 5 9
> x <- 1:3
> diag(x)
   [,1] [,2] [,3]
         0 0
[1,] 1
[2,] 0
          2
               0
         0
[3,]
    0
               3
> h <- 2
> diag(h)
   [,1] [,2]
[1,] 1 0
[2,] 0 1
```

t()

• Package: base

• Input:

A matrice di dimensione $n \times m$

• **Description:** trasposta

• Formula:

 A^T

```
[,1] [,2] [,3]
    [1,] 1.20 1.0 4.60
   [2,] 3.40 2.0 7.80
   [3,] 4.23 3.4 9.88
   > n < - 3
   > m < - 3
   > t(A)
        [,1] [,2] [,3]
   [1,] 1.2 3.4 4.23
   [2,] 1.0 2.0 3.40
   [3,] 4.6 7.8 9.88
   > A <- matrix(data = 1:2, nrow = 1, ncol = 2)
        [,1] [,2]
   [1,] 1 2
   > n <- 1
   > m < - 2
   > t(A)
       [,1]
   [1,] 1
[2,] 2
aperm()
  • Package: base
  • Input:
       A matrice di dimensione n \times m
  • Description: trasposta
  • Formula:
                                             A^T
  • Example:
   > A <- matrix(data = c(1.2, 3.4, 4.23, 1, 2, 3.4, 4.6, 7.8, 9.88),
   + nrow = 3, ncol = 3
   > A
        [,1] [,2] [,3]
   [1,] 1.20 1.0 4.60
    [2,] 3.40 2.0 7.80
   [3,] 4.23 3.4 9.88
   > n <- 3
   > m < - 3
   > aperm(A)
        [,1] [,2] [,3]
    [1,] 1.2 3.4 4.23
    [2,] 1.0 2.0 3.40
```

[3,] 4.6 7.8 9.88

```
> A <- matrix(data = 1:2, nrow = 1, ncol = 2)
> A

[,1] [,2]
[1,] 1 2

> n <- 1
> m <- 2
> t(A)

[,1]
[1,] 1
[2,] 2
```

dim()

• Package: base

• Input:

A matrice di dimensione $n \times m$

- **Description:** numero di righe e di colonne
- Formula:

n - m

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
   [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> dim(A)
[1] 3 3
> A <- matrix(data = c(1.2, 2.3, 6.5, 7.6), nrow = 2, ncol = 2)
> A
   [,1] [,2]
[1,] 1.2 6.5
[2,] 2.3 7.6
> n < -2
> m <- 2
> dim(A)
[1] 2 2
```

```
nrow()
```

```
• Package: base
```

• Input:

A matrice di dimensione $n \times m$

• Description: numero di righe

• Formula:

n

• Example:

NROW()

• Package: base

• Input:

A matrice di dimensione $n \times m$

- Description: numero di righe
- Formula:

n

[1] 2

```
[1] 3
   > A <- matrix(data = c(1.2, 2.3, 6.5, 7.6), nrow = 2, ncol = 2)
    > A
        [,1] [,2]
    [1,] 1.2 6.5
[2,] 2.3 7.6
   > NROW(A)
    [1] 2
ncol()
  • Package: base
  • Input:
        A matrice di dimensione n \times m
  • Description: numero di colonne
  • Formula:
                                                m
  • Example:
    > A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
   + ncol = 3)
   > A
        [,1] [,2] [,3]
    [1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
    [3,] 3.0 7.8 12.0
    > ncol(A)
    [1] 3
    > A <- matrix(data = 1:2, nrow = 1, ncol = 2)
        [,1] [,2]
    [1,] 1 2
   > ncol(A)
```

NCOL()

• Package: base

• Input:

A matrice di dimensione $n \times m$

• Description: numero di colonne

• Formula:

m

• Example:

rowSums()

• Package: futilities

• Input:

A matrice di dimensione $n \times m$

- Description: somme di riga
- Formula:

$$\sum_{i=1}^{m} x_{ij} \quad \forall i = 1, 2, \dots, n$$

```
> n <- 3
> m <- 3
> rowSums(A)

[1] 14.9 6.4 22.8

> A <- matrix(data = c(1.2, 3.4, 4.5, 5.6), nrow = 2, ncol = 2)
> A

        [,1] [,2]
[1,] 1.2 4.5
[2,] 3.4 5.6

> n <- 2
> m <- 2
> rowSums(A)

[1] 5.7 9.0
```

rowMeans()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times m$

- Description: medie di riga
- Formula:

$$\frac{1}{m} \sum_{i=1}^{m} x_{ij} \quad \forall i = 1, 2, \dots, n$$

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
> A
   [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> n <- 3
> m <- 3
> rowMeans(A)
[1] 4.966667 2.133333 7.600000
> A <- matrix(data = c(1.2, 3.4, 4.5, 5.6), nrow = 2, ncol = 2)
   [,1] [,2]
[1,] 1.2 4.5
[2,] 3.4 5.6
> n < -2
> m <- 2
> rowMeans(A)
[1] 2.85 4.50
```

colSums()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times m$

• Description: somme di colonna

• Formula:

$$\sum_{i=1}^{n} x_{ij} \quad \forall j = 1, 2, \dots, m$$

• Example:

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
> A
[1,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> n < -3
> m < - 3
> colSums(A)
[1] 3.8 17.4 22.9
> A <- matrix(data = c(1.2, 3.4, 4.5, 5.6), nrow = 2, ncol = 2)
> A
    [,1] [,2]
[1,] 1.2 4.5
[2,] 3.4 5.6
> n < -2
> m <- 2
> colSums(A)
[1] 4.6 10.1
```

colMeans()

• Package: fUtilities

• Input:

A matrice di dimensione $n \times m$

• Description: medie di colonna

• Formula:

$$\frac{1}{n}\sum_{i=1}^{n}x_{ij} \quad \forall j = 1, 2, \dots, m$$

```
[,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> n < -3
> m < - 3
> colMeans(A)
[1] 1.266667 5.800000 7.633333
> A <- matrix(data = c(1.2, 3.4, 4.5, 5.6), nrow = 2, ncol = 2)
    [,1] [,2]
[1,] 1.2 4.5
[2,] 3.4 5.6
> n <- 2
> m <- 2
> colMeans(A)
[1] 2.30 5.05
```

rowsum()

- Package: base
- Input:

A matrice di dimensione $n \times m$ group fattore f a k livelli di dimensione n

- \bullet **Description:** applica la funzione somma ad ogni gruppo di elementi in ciascuna colonna di A definito dai livelli di f
- Example 1:

```
> A <- matrix(data = c(1.2, 2.3, 4.3, 4.2, 4.2, 2.1, 2.2, 4), nrow = 4,
+
     ncol = 2)
> A
     [,1] [,2]
[1,] 1.2 4.2
[2,] 2.3 2.1
[3,] 4.3 2.2
[4,] 4.2 4.0
> n < -4
> m < - 2
> f <- factor(rep(1:2, times = 2))</pre>
> k <- nlevels(f)</pre>
> k
[1] 2
> rowsum(A, f)
  [,1] [,2]
1 5.5 6.4
2 6.5 6.1
```

• Example 2:

```
> A <- matrix(data = c(1, 2, 3, 4, 7, 8, 9, 8), nrow = 4, ncol = 2)
> A
     [,1] [,2]
[1,]
       1 7
      2
[2,]
             8
[3,]
      3
4
       3
             9
[4,]
> n < -4
> m < - 2
> k <- nlevels(f)</pre>
> k
[1] 2
> rowsum(A, f)
 [,1] [,2]
  4 16
2
   6 16
```

apply()

- Package: base
- Input:

```
A matrice di dimensione n \times m 
MARGIN = 1 / 2 riga o colonna 
FUN funzione scelta
```

- **Description:** applica FUN ad ogni riga o colonna della matrice A
- Example 1:

```
[,1] [,2] [,3]
 [1,] 1.0 4.0 9.9
 [2,] -0.2 5.6 1.0
 [3,] 3.0 7.8 12.0
 > n < -3
 > m < - 3
 > apply(A, MARGIN = 2, FUN = mean)
 [1] 1.266667 5.800000 7.633333
• Example 3:
 > A <- matrix(data = c(2, -1, -10.2, 1, -1, 5, 5.8, 3, 1, 3, 3.1,
      4), nrow = 4, ncol = 3)
 > A
       [,1] [,2] [,3]
 [1,]
       2.0 -1.0 1.0
 [2,] -1.0 5.0 3.0
 [3,] -10.2 5.8
                 3.1
 [4,] 1.0 3.0 4.0
 > n <- 4
 > m < - 3
 > apply(A, MARGIN = 2, FUN = sort)
       [,1] [,2] [,3]
 [1,] -10.2 -1.0 1.0
 [2,] -1.0
            3.0 3.0
       1.0 5.0 3.1
 [3,]
       2.0 5.8 4.0
 [4,]
• Example 4:
 > A <- matrix(data = c(2, -1, -10.2, 1, -1, 5, 5.8, 3, 1, 3, 3.1,
 + 4), nrow = 4, ncol = 3)
 > A
       [,1] [,2] [,3]
 [1,] 2.0 -1.0 1.0
 [2,] -1.0 5.0 3.0
 [3,] -10.2 5.8 3.1
 [4,] 1.0 3.0 4.0
 > n <- 4
 > m < - 3
 > apply(A, MARGIN = 2, FUN = function(x) {
      sort(x, decreasing = TRUE)
 + })
       [,1] [,2] [,3]
 [1,]
      2.0 5.8 4.0
      1.0 5.0 3.1
 [2,]
 [3,] -1.0 3.0 3.0
 [4,] -10.2 -1.0 1.0
• Example 5:
 > A <- matrix(data = c(1, 10, 100, 2, 20, 200, 3, 30, 300), nrow = 3,
 + ncol = 3)
 > A
```

```
[,1] [,2] [,3]
[1,] 1
          2
     10
         20
[2,]
[3,] 100 200 300
> n < -3
> m < - 3
> apply(A, MARGIN = 2, FUN = cumsum)
    [,1] [,2] [,3]
[1,] 1 2
[2,] 11 22
               33
[3,] 111 222 333
> t(apply(A, MARGIN = 1, FUN = cumsum))
    [,1] [,2] [,3]
[1,] 1
[2,] 10
[2,]
          30
              60
[3,] 100 300 600
```

solveCrossprod()

- Package: strucchange
- Input:

```
A matrice di dimensione n \times k di rango k = \min(n, k) method = qr / chol / solve algoritmo risolutivo
```

- **Description:** inversa del prodotto incrociato di X
- Formula:

$$(A^T A)^{-1}$$

• Example 1:

```
> A <- matrix(data = c(11, -2, 3.4, 4.1, 5, 7), nrow = 3, ncol = 2)
> A
    [,1] [,2]
[1,] 11.0 4.1
[2,] -2.0 5.0
[3,] 3.4 7.0
> n < -3
> k < - 2
> solve(t(A) %*% A)
             [,1]
                         [,2]
[1,] 0.010167039 -0.006594413
[2,] -0.006594413 0.015289185
> solveCrossprod(A, method = "qr")
             [,1]
[1,] 0.010167039 -0.006594413
[2,] -0.006594413 0.015289185
```

```
> A <- matrix(data = c(1, 2, 3, 4, 7, 8, 9, 8), nrow = 4, ncol = 2)
     [,1] [,2]
[1,]
      1 7
            8
[2,]
       2
       3
           9
[3,]
      4
[4,]
> n < -4
> k <- 2
> solve(t(A) %*% A)
            [,1]
                        [,2]
      0.25393701 -0.08070866
[2,] -0.08070866 0.02952756
> solveCrossprod(A, method = "qr")
            [,1]
                       [,2]
[1,] 0.25393701 -0.08070866
[2,] -0.08070866 0.02952756
```

model.matrix()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** matrice del modello di regressione lineare di dimensione $n \times k$
- Formula:

$$X = \begin{pmatrix} 1 & x_{1,1} & \dots & x_{1,k-1} \\ 1 & x_{2,1} & \dots & x_{2,k-1} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_{n,1} & \dots & x_{n,k-1} \end{pmatrix}$$

```
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3)
> k <- 4
> n <- 8
> X <- model.matrix(object = modello)</pre>
  (Intercept) x1 x2
     1 1.1 1.2 1.40
1
             1 2.3 3.4 5.60
3
             1 4.5 5.6 7.56
             1 6.7 7.5 6.00
             1 8.9 7.5 5.40
5
             1 3.4 6.7 6.60
             1 5.6 8.6 8.70
              1 6.7 7.6 8.70
attr(,"assign")
[1] 0 1 2 3
```

kappa()

- Package: base
- Input:

```
A matrice di dimensione n \times m exact = TRUE
```

- ullet Description: calcola il ConditionNumber come rapporto tra il maggiore ed il minore valore singolare non nullo della matrice diagonale D
- Formula:

$$\frac{\max\left(\operatorname{diag}(D)\right)}{\min\left(\operatorname{diag}(D)\right)}$$

```
\mbox{dove} \quad A \,=\, U\,D\,V^T \quad \mbox{e} \quad U^T\,U \,=\, I_m \,=\, V^T\,V \,=\, V\,V^T \label{eq:constraints}
```

• Example 1:

• Example 2:

```
> A <- matrix(data = c(1, 2, 3, 4, 7, 8, 9, 8), nrow = 4, ncol = 2)
    [,1] [,2]
[1,]
      1
      2
             8
[2,]
[3,]
       3
             9
[4,]
> n <- 4
> m < - 2
> D <- diag(svd(A)$d)
> max(diag(D))/min(diag(D))
[1] 8.923297
> kappa(A, exact = TRUE)
[1] 8.923297
```

• **Note:** Calcola il *Condition Number* con la funzione svd().

lower.tri()

- Package: base
- Input:

A matrice di dimensione $n \times n$

- **Description:** matrice triangolare inferiore di dimensione $n \times n$ a partire dalla matrice A
- Example 1:

```
> A <- matrix(data = 1:9, nrow = 3, ncol = 3)
    [,1] [,2] [,3]
    1
2
[1,]
         5
[2,]
     3 6
[3,]
> n < -3
> A[t(lower.tri(A, diag = FALSE))] <- 0</pre>
> A
    [,1] [,2] [,3]
    1 0
[1,]
       2
           5
[2,]
[3,]
     3 6
```

• Example 2:

upper.tri()

- Package: base
- Input:

A matrice di dimensione $n \times n$

- **Description:** matrice triangolare superiore di dimensione $n \times n$ a partire dalla matrice A
- Example 1:

• Example 2:

backsolve()

- Package: base
- Input:

```
r matrice A dei coefficienti di dimensione n \times n data matrice b dei termini noti di dimensione 1 \times n upper.tri = TRUE / FALSE sistema triangolare superiore od inferiore transpose = TRUE / FALSE matrice dei coefficienti trasposta
```

- **Description:** soluzione di un sistema triangolare di dimensione $n \times n$
- Formula:

```
upper.tri = FALSE AND transpose = FALSE
```

```
\begin{pmatrix}
a_{1,1} & 0 & \dots & \dots & 0 & b_1 \\
a_{2,1} & a_{2,2} & 0 & \dots & 0 & b_2 \\
\vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\
a_{n-1,1} & a_{n-1,2} & \dots & \ddots & 0 & \vdots \\
a_{n,1} & a_{n,2} & \dots & \dots & a_{n,n} & b_n
\end{pmatrix}
```

• Example 1:

forwardsolve()

• Package: base

• Input:

```
1 matrice A dei coefficienti di dimensione n \times n
```

imes matrice b dei termini noti di dimensione $1 \times n$

upper.tri = TRUE / FALSE sistema triangolare superiore od inferiore
transpose = TRUE / FALSE matrice dei coefficienti trasposta

- **Description:** soluzione di un sistema triangolare di dimensione $n \times n$
- Formula:

upper.tri = TRUE AND transpose = FALSE

$$\begin{pmatrix} a_{1,1} & a_{1,2} & \dots & a_{1,n-1} & a_{1,n} & b_1 \\ 0 & a_{2,2} & \dots & a_{2,n-1} & a_{2,n} & b_2 \\ \vdots & 0 & \ddots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ 0 & 0 & \dots & 0 & a_{n,n} & b_n \end{pmatrix}$$

upper.tri = FALSE AND transpose = TRUE

$$\begin{pmatrix} a_{1,1} & a_{2,1} & \dots & a_{n-1,1} & a_{n,1} & b_1 \\ 0 & a_{2,2} & \dots & a_{n-1,2} & a_{n,2} & b_2 \\ \vdots & 0 & \ddots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ 0 & 0 & \dots & 0 & a_{n,n} & b_n \end{pmatrix}$$

upper.tri = FALSE AND transpose = FALSE

$$\begin{pmatrix} a_{1,1} & 0 & \dots & 0 & b_1 \\ a_{2,1} & a_{2,2} & 0 & \dots & 0 & b_2 \\ \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ a_{n-1,1} & a_{n-1,2} & \dots & \ddots & 0 & \vdots \\ a_{n,1} & a_{n,2} & \dots & \dots & a_{n,n} & b_n \end{pmatrix}$$

```
[1] 8 4 2

> forwardsolve(1 = A, x = b, upper.tri = TRUE, transpose = TRUE)

[1] 8.000000 -5.000000 -6.016667

• Example 2:

> A <- matrix(data = c(1.2, 0.34, 7.7, 4.5), nrow = 2, ncol = 2, tolon tol
```

2.4 Fattorizzazioni di Matrici

svd()

- Package: base
- Input:
 - A matrice di dimensione $n \times m$
- Description: fattorizzazione ai valori singolari
- Output:
 - d diagonale della matrice D dei valori singolari di dimensione $m \times m$
 - u matrice U di dimensione $n \times m$
 - v matrice ortogonale V di dimensione $m \times m$
- Formula:

$$A\,=\,U\,D\,V^T$$

$$dove \quad U^T U = I_m = V^T V = V V^T$$

```
> A <- matrix(data = c(11, -2, 3.4, 4.1, 5, 7), nrow = 3, ncol = 2)
> A

[,1] [,2]
[1,] 11.0   4.1
[2,] -2.0   5.0
[3,]  3.4   7.0
```

```
> n < -3
 > m <- 2
 > D <- diag(svd(A)$d)
         [,1] [,2]
 [1,] 13.29929 0.000000
 [2,] 0.00000 7.106262
 > U <- svd(A)$u
 > U
            [,1] [,2]
 [1,] -0.8566792 0.3981302
 [2,] -0.0882360 -0.7395948
 [3,] -0.5082471 -0.5426710
 > t(U) %*% U
              [,1]
 [1,] 1.000000e+00 -3.762182e-17
 [2,] -3.762182e-17 1.000000e+00
 > V <- svd(A)$v
            [,1] [,2]
 [1,] -0.8252352 0.5647893
 [2,] -0.5647893 -0.8252352
 > t(V) %*% V
                       [,2]
              [,1]
 [1,] 1.000000e+00 -2.222614e-18
 [2,] -2.222614e-18 1.000000e+00
 > V %*% t(V)
              [,1] [,2]
 [1,] 1.000000e+00 2.222614e-18
 [2,] 2.222614e-18 1.000000e+00
 > U %*% D %*% t(V)
     [,1] [,2]
 [1,] 11.0 4.1
 [2, ] -2.0 5.0
 [3,] 3.4 7.0
• Example 2:
 > A <- matrix(data = c(1, 2, 3.45, 7.8), nrow = 2, ncol = 2)
 > A
    [,1] [,2]
 [1,] 1 3.45
[2,] 2 7.80
```

```
> n <- 2
> m <- 2
> D <- diag(svd(A)$d)
       [,1] [,2]
[1,] 8.81658 0.0000000
[2,] 0.00000 0.1020804
> U <- svd(A)$u
[,1] [,2]
[1,] -0.4072775 -0.9133044
[2,] -0.9133044 0.4072775
> t(U) %*% U
             [,1]
                           [,2]
[1,] 1.000000e+00 -2.201201e-16
[2,] -2.201201e-16 1.000000e+00
> V <- svd(A)$v
> V
          [,1] [,2]
[1,] -0.2533734 -0.9673686
[2,] -0.9673686 0.2533734
> t(V) %*% V
            [,1] [,2]
[1,] 1.000000e+00 1.585646e-18
[2,] 1.585646e-18 1.000000e+00
> V %*% t(V)
            [,1]
                        [,2]
[1,] 1.000000e+00 1.585646e-18
[2,] 1.585646e-18 1.000000e+00
> U %*% D %*% t(V)
   [,1] [,2]
[1,] 1 3.45
[2,] 2 7.80
```

qr.Q()

- Package: base
- Input:

A matrice di rango pieno di dimensione $n \times m$

- **Description:** matrice Q di dimensione $n \times m$
- Formula:

$$A = QR$$

dove
$$Q^T Q = I_m$$

• Example 1:

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> n < -3
> m < - 3
> Q <- qr.Q(qr(A))
> Q
                       [,2]
            [,1]
[1,] -0.31559720 -0.220214186 -0.9229865
[2,] 0.06311944 -0.975415572 0.2111407
[3,] -0.94679160 0.008377024 0.3217382
> t(Q) %*% Q
                           [,2]
             [,1]
                                         [,3]
[1,] 1.000000e+00 -1.690678e-17 -4.214836e-17
[2,] -1.690678e-17 1.000000e+00 3.281046e-17
[3,] -4.214836e-17 3.281046e-17 1.000000e+00
```

• Example 2:

```
> A <- matrix(data = c(1, 2, 3.45, 7.8), nrow = 2, ncol = 2)
> A
   [,1] [,2]
[1,] 1 3.45
[2,] 2 7.80
> n < -2
> m < - 2
> Q <- qr.Q(qr(A))
> Q
           [,1]
                      [,2]
[1,] -0.4472136 -0.8944272
[2,] -0.8944272 0.4472136
> t(Q) %*% Q
              [,1]
                             [,2]
[1,] 1.000000e+00 -1.260385e-17
[2,] -1.260385e-17 1.000000e+00
```

qr.R()

- Package: base
- Input:

A matrice di rango pieno di dimensione $n \times m$

- **Description:** matrice R triangolare superiore di dimensione $m \times m$
- Formula:

$$A = QR$$

• Example 1:

```
> A <- matrix(data = c(1, -0.2, 3, 4, 5.6, 7.8, 9.9, 1, 12), nrow = 3,
+ ncol = 3)
> A
    [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
> n < -3
> m <- 3
> R <- qr.R(qr(A))
> R
                 [,2] [,3]
         [,1]
[1,] -3.168596 -8.293894 -14.422792
[2,] 0.000000 -6.277843 -3.055012
[3,] 0.000000 0.000000 -5.065567
> Q <- qr.Q(qr(A))
> Q
           [,1]
                      [,2] [,3]
[1,] -0.31559720 -0.220214186 -0.9229865
[2,] 0.06311944 -0.975415572 0.2111407
[3,] -0.94679160 0.008377024 0.3217382
> Q %*% R
    [,1] [,2] [,3]
[1,] 1.0 4.0 9.9
[2,] -0.2 5.6 1.0
[3,] 3.0 7.8 12.0
```

• Example 2:

```
[,1] [,2]
[1,] -2.236068 -8.5194190
[2,] 0.000000 0.4024922

> Q <- qr.Q(qr(A))
> Q

[,1] [,2]
[1,] -0.4472136 -0.8944272
[2,] -0.8944272 0.4472136

> Q *** R

[,1] [,2]
[1,] 1 3.45
[2,] 2 7.80
```

chol()

• Package: base

• Input:

A matrice simmetrica definita positiva di dimensione $n \times n$

- **Description:** matrice P triangolare superiore di dimensione $n \times n$
- Formula:

$$A = P^T P$$

• Example 1:

• Example 2:

```
> A <- matrix(data = c(1.2, 3.4, 3.4, 11.2), nrow = 2, ncol = 2)
> A

[,1] [,2]
[1,] 1.2 3.4
[2,] 3.4 11.2
```

chol2inv()

• Package: base

• Input:

 ${\mathbb P}\$ matrice P triangolare superiore di dimensione $n\times n$

• **Description:** funzione inversa di chol ()

[1,] 0.21428571 -0.07142857 [2,] -0.07142857 0.35714286

• Formula:

$$(P^T P)^{-1}$$

```
> A <- matrix(data = c(5, 1, 1, 3), nrow = 2, ncol = 2)
   [,1] [,2]
[1,] 5 1
[2,] 1 3
> n <- 2
> P <- chol(A)
> P
        [,1] [,2]
[1,] 2.236068 0.4472136
[2,] 0.000000 1.6733201
> t(P) %*% P
    [,1] [,2]
[1,] 5 1
[2,] 1 3
> chol2inv(P)
           [,1] [,2]
[1,] 0.21428571 -0.07142857
[2,] -0.07142857 0.35714286
> solve(A)
           [,1] [,2]
```

• Example 2:

```
> A <- matrix(data = c(1.2, 3.4, 3.4, 11.2), nrow = 2, ncol = 2)
    [,1] [,2]
[1,] 1.2 3.4
[2,] 3.4 11.2
> n <- 2
> P <- chol(A)
> P
        [,1] [,2]
[1,] 1.095445 3.103761
[2,] 0.000000 1.251666
> t(P) %*% P
   [,1] [,2]
[1,] 1.2 3.4
[2,] 3.4 11.2
> chol2inv(P)
         [,1]
               [,2]
[1,] 5.957447 -1.8085106
[2,] -1.808511 0.6382979
> solve(A)
         [,1]
               [,2]
[1,] 5.957447 -1.8085106
[2,] -1.808511 0.6382979
```

ginv()

- Package: MASS
- Input:

A matrice di dimensione $n \times m$

- Description: inversa generalizzata A_g di dimensione $m \times n$
- Formula:

$$A = A A_q A$$

• Example 1:

```
[,1]
                        [,2]
 [1,] 0.007783879 -0.4266172 0.302297558
 [2,] 0.035078001 0.1553743 -0.001334379
 > A %*% Ag %*% A
      [,1] [,2]
 [1,] 1.0 4.0
 [2,] -0.2 5.6
 [3,] 3.0 7.8
• Example 2:
 > A <- matrix(data = c(1.2, 3.4, 3.4, 11.2), nrow = 2, ncol = 2)
 > A
     [,1] [,2]
 [1,] 1.2 3.4
 [2,] 3.4 11.2
 > n <- 2
 > m <- 2
 > Ag <- ginv(A)</pre>
 > Ag
                 [,2]
           [,1]
 [1,] 5.957447 -1.8085106
 [2,] -1.808511 0.6382979
 > A %*% Ag %*% A
     [,1] [,2]
 [1,] 1.2 3.4
 [2,] 3.4 11.2
```

2.5 Creazione di Arrays

array()

• Package: base

• Input:

data vettore numerico
dim dimensione
dimnames etichette di dimensione

• **Description:** creazione

```
> etichette <- list(c("A", "B"), c("a", "b"), c("X", "Y"))
> myarray <- array(data = 1:8, dim = c(2, 2, 2), dimnames = etichette)
> myarray
```

```
, , X
 a b
A 1 3
B 2 4
, , Y
 a b
A 5 7
B 6 8
> etichette <- list(c("A", "B"), c("a", "b"))</pre>
> x <- array(data = 1:8, dim = c(2, 2), dimnames = etichette)
 a b
A 1 3
B 2 4
> x < - seq(1:12)
> dim(x) <- c(3, 2, 2)
> x
, , 1
    [,1] [,2]
[1,] 1 4
[2,] 2 5
[3,] 3 6
, , 2
 [,1] [,2]
[1,] 7 10
[2,] 8 11
     9 12
[3,]
> array(data = 1, dim = c(4, 5))
   [,1] [,2] [,3] [,4] [,5]
[1,] 1 1 1 1 1
[2,] 1 1 1 1 1
[2,]
     \begin{matrix}1&&1\\1&&1\end{matrix}
[3,]
                  1
                       1
                             1
[4,]
                  1
                       1
                             1
```

dim()

• Package: base

• Input:

x array

• **Description:** dimensione

```
> n <- 3
> m <- 3
> x <- 1:9
> dim(x) <- c(n, m)
> x
```

```
[,1] [,2] [,3]
[1,] 1 4 7
[2,] 2 5 8
[3,] 3 6 9

> x <- seq(1:12)
> dim(x) <- c(3, 2, 2)
> x

, , 1

[,1] [,2]
[1,] 1 4
[2,] 2 5
[3,] 3 6

, , 2

[,1] [,2]
[1,] 7 10
[2,] 8 11
[3,] 9 12
```

[]

- Package: base
- Input:

x array

- **Description:** estrazione di elementi
- Example:

dimnames()

• Package: base

• Input:

x array

• **Description:** etichette di dimensione

```
> x
, , 1
[,1] [,2] [,3]
[1,] 1 3 5
[2,] 2 4 6
, , 2
 [,1] [,2] [,3]
[1,] 7 9 11
[2,] 8 10 12
> dimnames(x) <- list(letters[1:2], LETTERS[1:3], c("primo", "secondo"))</pre>
> x
, , primo
 АВС
a 1 3 5
b 2 4 6
, , secondo
 A B C
a 7 9 11
b 8 10 12
```

Parte II Statistica Descrittiva

Capitolo 3

> max(x)

[1] 6.4

Misure ed indici statistici

3.1 Minimo e massimo

```
min()
  • Package: base
  • Input:
        	imes vettore numerico di dimensione n
  • Description: minimo
  • Formula:
                                                  x_{(1)}
  • Examples:
    > x < -c(4.5, 3.4, 8.7, 3.6)
    > \min(x)
    [1] 3.4
    > x < -c(1.1, 3.4, 4.5, 6.4, 4, 3, 4)
    > \min(x)
    [1] 1.1
max()
  • Package: base
  • Input:
        \times vettore numerico di dimensione n
  • Description: massimo
  • Formula:
                                                  x_{(n)}
  • Examples:
    > x < -c(1.2, 2.3, 4.5, 6.5)
    > max(x)
    [1] 6.5
    > x < -c(1.1, 3.4, 4.5, 6.4, 4, 3, 4)
```

3.2 Campo di variazione e midrange

```
range()
  • Package: base
  • Input:
        	imes vettore numerico di dimensione n
  • Description: minimo e massimo
  • Formula:
                                               x_{(1)}
                                                    x_{(n)}
  • Examples:
    > x < -c(1, 1.2, 3.4, 0.8)
    > min(x)
    [1] 0.8
    > max(x)
    [1] 3.4
    > range(x)
    [1] 0.8 3.4
    > x < -c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
    > min(x)
    [1] 1.2
    > \max(x)
    [1] 6.4
    > range(x)
    [1] 1.2 6.4
range2()
  • Package: sigma2tools
  • Input:
        {\bf x}~ vettore numerico di dimensione n
  • Description: campo di variazione
  • Formula:
                                                x_{(n)} - x_{(1)}
  • Examples:
```

> x < -c(1, 1.2, 3.4, 0.8)

> min(x)

[1] 0.8

```
> \max(x)
    [1] 3.4
    > max(x) - min(x)
    [1] 2.6
    > range2(x)
    [1] 2.6
    > x \leftarrow c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
    > min(x)
    [1] 1.2
   > max(x)
    [1] 6.4
   > max(x) - min(x)
    [1] 5.2
    > range2(x)
    [1] 5.2
midrange()
  • Package: sigma2tools
  • Input:
        	imes vettore numerico di dimensione n
  • Description: midrange
  • Formula:
                                            (x_{(1)} + x_{(n)}) / 2
  • Examples:
    > x < -c(1, 1.2, 3.4, 0.8, 1.77, 7.8)
    > min(x)
    [1] 0.8
    > \max(x)
    [1] 7.8
    > (min(x) + max(x))/2
    [1] 4.3
```

> midrange(x)

```
[1] 4.3
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> min(x)

[1] 1.2
> max(x)

[1] 6.4
> (min(x) + max(x))/2

[1] 3.8
> midrange(x)
```

extendrange()

• Package: grDevices

- Input:
 - \times vettore numerico di dimensione n
 - f percentuale di estensione α del campo di variazione
- **Description:** campo di variazione
- Formula:

$$x_{(1)} - \alpha (x_{(n)} - x_{(1)})$$
 $x_{(n)} + \alpha (x_{(n)} - x_{(1)})$

```
> x <- c(1, 1.2, 3.4, 0.8)
> alpha <- 0.05
> min(x)

[1] 0.8

> max(x)

[1] 3.4

> min(x) - alpha * (max(x) - min(x))

[1] 0.67

> max(x) + alpha * (max(x) - min(x))

[1] 3.53

> extendrange(x, f = 0.05)

[1] 0.67 3.53
```

```
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> alpha <- 0.05
> min(x)

[1] 1.2

> max(x)

[1] 6.4

> min(x) - alpha * (max(x) - min(x))

[1] 0.94

> max(x) + alpha * (max(x) - min(x))

[1] 6.66

> extendrange(x, f = 0.05)
```

3.3 Media aritmetica, geometrica ed armonica

mean()

• Package: base

• Input:

 \times vettore numerico di dimensione n

trim il valore di α con $0 \le \alpha \le 0.5$ che rappresenta la percentuale di osservazioni più basse e più alte che deve essere esclusa dal calcolo della media aritmetica

- **Description:** media α -trimmed
- Formula:

$$\bar{x}_{\alpha} = \begin{cases} \bar{x} & \text{se } \alpha = 0 \\ \frac{1}{n-2 \lfloor n \alpha \rfloor} \sum_{i=\lfloor n \alpha \rfloor+1}^{n-\lfloor n \alpha \rfloor} x_{(i)} & \text{se } 0 < \alpha < 0.5 \\ Q_{0.5}(x) & \text{se } \alpha = 0.5 \end{cases}$$

```
> x <- c(1, 1.2, 3.4, 0.8, 10.2, 9.3, 7.34)
> n <- 7
> sum(x)/n

[1] 4.748571

> mean(x, trim = 0)

[1] 4.748571

> x <- c(1, 1.2, 3.4, 0.8, 10.2, 9.3, 7.34)
> x <- sort(x)
> x
```

mean.g()

- Package: labstatR
- Input:
 - ${\bf x}~$ vettore numerico di elementi positivi di dimensione n
- Description: media geometrica
- Formula:

$$\bar{x}_G = \left(\prod_{i=1}^n x_i\right)^{1/n} = \exp\left(\frac{1}{n}\sum_{i=1}^n \log(x_i)\right)$$

```
> x <- c(1.2, 2.3, 4.5, 6.5)
> n <- 4
> prod(x)^(1/n)

[1] 2.997497
> mean.g(x)

[1] 2.997497
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> n <- 7
> prod(x)^(1/n)

[1] 3.434782
> mean.g(x)

[1] 3.434782
```

mean.a()

• Package: labstatR

• Input:

 ${\bf x}~$ vettore numerico di elementi non nulli di dimensione n

• Description: media armonica

• Formula:

$$\bar{x}_A = \left(\frac{1}{n} \sum_{i=1}^n \frac{1}{x_i}\right)^{-1}$$

• Examples:

```
> x <- c(1.2, 2.3, 4.5, 6.5)
> 1/mean(1/x)

[1] 2.432817
> mean.a(x)

[1] 2.432817
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> 1/mean(1/x)

[1] 2.992404
> mean.a(x)

[1] 2.992404
```

3.4 Mediana e quantili

median()

• Package: stats

• Input:

 \times vettore numerico di dimensione n

• Description: mediana

• Formula:

$$Q_{0.5}(x) \,=\, \left\{ \begin{array}{ll} x_{(\frac{n+1}{2})} & \text{se } n \text{ \`e dispari} \\ \\ 0.5 \left(x_{(\frac{n}{2})} + x_{(\frac{n}{2}+1)} \right) & \text{se } n \text{ \`e pari} \end{array} \right.$$

```
> x <- c(1.2, 0.34, 5.6, 7.4, 2.1, 3.2, 9.87, 10.1)
> x <- sort(x)
> x

[1] 0.34 1.20 2.10 3.20 5.60 7.40 9.87 10.10
> n <- 8
> 0.5 * (x[n/2] + x[n/2 + 1])
```

```
[1] 4.4
 > median(x)
 [1] 4.4
 > x \leftarrow c(1.2, 0.34, 5.6, 7.4, 2.1, 3.2, 9.87)
 > x <- sort(x)
 > x
 [1] 0.34 1.20 2.10 3.20 5.60 7.40 9.87
 > n <- 7
 > x[(n + 1)/2]
 [1] 3.2
 > median(x)
 [1] 3.2
• Note: Equivale alla funzione quantile() quando questa è calcolata in probs = 0.5.
```

quantile()

• Package: stats

• Input:

 \times vettore numerico di dimensione nprobs valore p di probabilità

- **Description:** quantile al (100 p)%
- Formula:

$$Q_p(x) \,=\, \left\{ \begin{array}{ll} x_{(\alpha)} & \text{se } \alpha \text{ è intero} \\ \\ x_{(\lfloor \alpha \rfloor)} + (\alpha - \lfloor \alpha \rfloor) \, \left(x_{(\lfloor \alpha \rfloor + 1)} - x_{(\lfloor \alpha \rfloor)} \right) & \text{se } \alpha \text{ non è intero} \end{array} \right.$$

dove
$$\alpha = 1 + (n-1)p$$

```
> x < -c(1.2, 2.3, 0.11, 4.5, 2.3, 4.55, 7.8, 6.6, 9.9)
> x <- sort(x)
> x
[1] 0.11 1.20 2.30 2.30 4.50 4.55 6.60 7.80 9.90
> n < -9
> p <- 0.25
> alpha <-1+(n-1)*p
> alpha
[1] 3
> x[alpha]
[1] 2.3
```

```
> quantile(x, probs = 0.25)
 25%
 2.3
 > x < -c(1.2, 2.3, 0.11, 4.5)
 > x <- sort(x)
 > x
 [1] 0.11 1.20 2.30 4.50
 > n <- 4
 > p < -0.34
 > alpha <- 1 + (n - 1) * p
 > alpha
 [1] 2.02
 > x[floor(alpha)] + (alpha - floor(alpha)) * (x[floor(alpha) +
     1] - x[floor(alpha)])
 [1] 1.222
 > quantile(x, probs = 0.34)
   34%
 1.222
 > x < -c(1.2, 4.2, 4.5, -5.6, 6.5, 1.2)
 > x <- sort(x)
 > n < -6
 > p < -0.68
 > alpha <- 1 + (n - 1) * p
 > alpha
 [1] 4.4
 > x[floor(alpha)] + (alpha - floor(alpha)) * (x[floor(alpha) +
 + 1] - x[floor(alpha)])
 [1] 4.32
 > quantile(x, probs = 0.68)
  68%
 4.32
• Note 1: Equivale alla funzione median() quando probs = 0.5.
• Note 2: Equivale alla funzione min() quando probs = 0.
• Note 3: Equivale alla funzione max() quando probs = 1.
```

3.5 Differenza interquartile e deviazione assoluta dalla mediana

```
IQR()
```

```
• Package: stats
```

• Input:

x vettore numerico di dimensione n

• **Description:** differenza interquartile

• Formula:

$$IQR(x) = Q_{0.75}(x) - Q_{0.25}(x)$$

• Examples:

```
> x <- c(1, 1.2, 3.4, 0.8, 10.2, 9.3, 7.34)
> diff(quantile(x, probs = c(0.25, 0.75)))

75%
7.22
> IQR(x)

[1] 7.22
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> diff(quantile(x, probs = c(0.25, 0.75)))

75%
1.05
> IQR(x)
[1] 1.05
```

• **Note:** Calcola i quartili con la funzione quantile().

mad()

• Package: stats

• Input:

x vettore numerico di dimensione n center parametro rispetto al quale si effettuano gli scarti constant il valore α della costante positiva

• Description: deviazione assoluta dalla mediana

• Formula:

$$\alpha Q_{0.5}(|x - \mathbf{center}(x)|)$$

```
> x <- c(1.2, 3.4, 4.5, 6.4, 4)
> alpha <- 1.23
> alpha * median(abs(x - median(x)))

[1] 0.738

> mad(x, center = median(x), constant = 1.23)
```

```
[1] 0.738

> x <- c(1, 1.2, 3.4, 0.8, 10.2, 9.3, 7.34)
> alpha <- 1.55
> alpha * median(abs(x - mean(x)))

[1] 5.810286

> mad(x, center = mean(x), constant = 1.55)

[1] 5.810286

> x <- c(1.2, 4.2, 4.5, -5.6, 6.5, 1.2)
> alpha <- 2.42
> alpha * median(abs(x - mean(x)))

[1] 5.687

> mad(x, center = mean(x), constant = 2.42)

[1] 5.687

• Note: Per default vale constant = 1.4826 = 1/Φ<sup>-1</sup>(0.75) e center = median(x).
```

3.6 Asimmetria e curtosi

skew()

• Package: labstatR

• Input:

x vettore numerico di dimensione n

• Description: asimmetria nella popolazione

• Formula:

$$\gamma_3 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{\sigma_x} \right)^3$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean((x - mean(x))^3/sigmax^3)

[1] 0.1701538
> skew(x)

[1] 0.1701538

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean((x - mean(x))^3/sigmax^3)

[1] -0.5845336
> skew(x)

[1] -0.5845336
```

skewness()

• Package: fBasics

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• Description: asimmetria campionaria

• Formula:

$$\hat{\gamma}_3 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right)^3$$

• Examples:

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> mean((x - mean(x))^3/sd(x)^3)

[1] 0.1217521

> skewness(x)

[1] 0.1217521

attr(,"method")
[1] "moment"

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> mean((x - mean(x))^3/sd(x)^3)

[1] -0.4182582

> skewness(x)

[1] -0.4182582

attr(,"method")
[1] "moment"
```

skewness()

• Package: e1071

• Input:

 \times vettore numerico di dimensione n

• **Description:** asimmetria campionaria

• Formula:

$$\hat{\gamma}_3 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right)^3$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> mean((x - mean(x))^3/sd(x)^3)

[1] 0.1217521
> skewness(x)
```

```
[1] 0.1217521
attr(,"method")
[1] "moment"

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> mean((x - mean(x))^3/sd(x)^3)

[1] -0.4182582

> skewness(x)

[1] -0.4182582
attr(,"method")
[1] "moment"
```

kurt()

• Package: labstatR

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- Description: kurtosi nella popolazione
- Formula:

$$\gamma_4 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{\sigma_x} \right)^4$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean((x - mean(x))^4/sigmax^4)

[1] 1.623612

> kurt(x)

[1] 1.623612

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean((x - mean(x))^4/sigmax^4)

[1] 2.312941

> kurt(x)

[1] 2.312941
```

kurtosis()

• Package: fBasics

• Input:

 \times vettore numerico di dimensione n

• Description: kurtosi campionaria

• Formula:

$$\hat{\gamma}_4 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right)^4 - 3$$

• Examples:

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> mean((x - mean(x))^4/sd(x)^4) - 3

[1] -1.960889

> kurtosis(x)

[1] -1.960889
attr(,"method")
[1] "excess"

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> mean((x - mean(x))^4/sd(x)^4) - 3

[1] -1.519718

> kurtosis(x)

[1] -1.519718
attr(,"method")
[1] "excess"
```

kurtosis()

• Package: e1071

• Input:

 \times vettore numerico di dimensione n

• Description: kurtosi campionaria

• Formula:

$$\hat{\gamma}_4 = \frac{1}{n} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right)^4 - 3$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> mean((x - mean(x))^4/sd(x)^4) - 3
[1] -1.960889
> kurtosis(x)
```

```
[1] -1.960889
attr(,"method")
[1] "excess"

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> mean((x - mean(x))^4/sd(x)^4) - 3

[1] -1.519718

> kurtosis(x)

[1] -1.519718
attr(,"method")
[1] "excess"
```

geary()

- Package:
- Input:
 - ${\bf x}~$ vettore numerico di dimensione n
- **Description:** kurtosi secondo *Geary*
- Formula:

$$\gamma_4^G = \frac{1}{n} \sum_{i=1}^n \frac{|x_i - \bar{x}|}{\sigma_x}$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean(abs(x - mean(x))/sigmax)

[1] 0.8702836

> geary(x)

[1] 0.8702836

> x <- c(1.2, 3.4, 5.2, 3.4, 4.4)
> sigmax <- sqrt(mean((x - mean(x))^2))
> mean(abs(x - mean(x))/sigmax)

[1] 0.7629055

> geary(x)

[1] 0.7629055
```

3.7 Coefficiente di variazione

var.coeff()

• Package: ineq

• Input:

```
x vettore numerico di dimensione n square = TRUE / FALSE quadrato
```

- Description: coefficiente di variazione nella popolazione
- Formula:

square = FALSE
$$CV_x = \sigma_x \, / \, \bar{x}$$

$$CV_x^2 = (\sigma_x / \bar{x})^2$$

• Examples:

```
> x <- c(1, 1.2, 3.4, 0.8)
> sigmax <- sqrt(mean((x - mean(x))^2))
> sigmax/mean(x)

[1] 0.6555055

> var.coeff(x, square = FALSE)

[1] 0.6555055

> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> sigmax <- sqrt(mean((x - mean(x))^2))
> (sigmax/mean(x))^2

[1] 0.1484087

> var.coeff(x, square = TRUE)

[1] 0.1484087
```

cv()

• Package: labstatR

• Input:

 \times vettore numerico di dimensione n

- Description: coefficiente di variazione nella popolazione
- Formula:

$$CV_x = \sigma_x / |\bar{x}| = \sqrt{\frac{n-1}{n}} cv_x$$

```
> x <- c(1, 1.2, 3.4, 0.8)
> sigmax <- sqrt(mean((x - mean(x))^2))
> sigmax/abs(mean(x))

[1] 0.6555055

> cv(x)

[1] 0.6555055

> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> sigmax <- sqrt(mean((x - mean(x))^2))
> sigmax/abs(mean(x))

[1] 0.3852385

> cv(x)

[1] 0.3852385
```

cv2()

• Package: sigma2tools

• Input:

 \times vettore numerico di dimensione n

- **Description:** coefficiente di variazione campionario
- Formula:

$$cv_x = s_x / |\bar{x}| = \sqrt{\frac{n}{n-1}} CV_x$$

```
> x <- c(1, 1.2, 3.4, 0.8)
> sd(x)/abs(mean(x))

[1] 0.7569126

> cv2(x)

[1] 0.7569126

> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> sd(x)/abs(mean(x))

[1] 0.4161051

> cv2(x)

[1] 0.4161051
```

3.8 Scarto quadratico medio e deviazione standard

sigma()

• Package: sigma2tools

• Input:

x vettore numerico di dimensione n

• Description: scarto quadratico medio

• Formula:

$$\sigma_x = \left(\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2\right)^{1/2} = \sqrt{\frac{1}{n} s s_x} = \sqrt{\frac{n-1}{n}} s_x$$

• Examples:

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> sqrt(mean((x - mean(x))^2))

[1] 2.868031
> sigma(x)

[1] 2.868031
> x <- c(1.2, 2.3, 4.5, 6.5)
> sqrt(mean((x - mean(x))^2))

[1] 2.041292
> sigma(x)

[1] 2.041292
```

sd()

• Package: stats

• Input:

 \times vettore numerico di dimensione n

• Description: deviazione standard

• Formula:

$$s_x = \left(\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2\right)^{1/2} = \sqrt{\frac{1}{n-1} s s_x} = \sqrt{\frac{n}{n-1}} \sigma_x$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> n <- 5
> sqrt(sum((x - mean(x))^2)/(n - 1))

[1] 3.206556
> sd(x)

[1] 3.206556
```

```
> x <- c(1.3, 4.2, 3.3, 8.7)
> n <- 4
> sqrt(sum((x - mean(x))^2)/(n - 1))

[1] 3.127699
> sd(x)

[1] 3.127699
```

3.9 Errore standard

popstderror()

• Package: sigma2tools

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- **Description:** errore standard nella popolazione
- Formula:

$$SE_x = \sigma_x / \sqrt{n} = \sqrt{\frac{n-1}{n}} se_x$$

```
> x <- c(1, 1.2, 3.4, 0.8)
> n <- 4
> sigmax <- sqrt(sum((x - mean(x))^2)/n)
> sigmax/sqrt(n)

[1] 0.5244044

> popstderror(x)

[1] 0.5244044

> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> n <- 7
> sigmax <- sqrt(sum((x - mean(x))^2)/n)
> sigmax/sqrt(n)

[1] 0.5512245

> popstderror(x)
```

stderror()

• Package: sigma2tools

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• Description: errore standard campionario

• Formula:

$$se_x = s_x / \sqrt{n} = \sqrt{\frac{n}{n-1}} SE_x$$

• Examples:

```
> x <- c(1, 1.2, 3.4, 0.8)
> n <- 4
> sd(x)/sqrt(n)

[1] 0.6055301
> stderror(x)

[1] 0.6055301
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> n <- 7
> sd(x)/sqrt(n)

[1] 0.5953905
> stderror(x)

[1] 0.5953905
```

3.10 Varianza e devianza

sigma2()

• Package: labstatR

• Input:

imes vettore numerico di dimensione n

• Description: varianza nella popolazione

• Formula:

$$\sigma_x^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{x}^2 = \frac{1}{n} s s_x = \frac{n-1}{n} s_x^2$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> mean((x - mean(x))^2)

[1] 8.2256
> sigma2(x)

[1] 8.2256
```

```
> x <- c(1.2, 2.3, 4.5, 6.5)
> mean((x - mean(x))^2)

[1] 4.166875

> sigma2(x)

[1] 4.166875
```

var()

• Package: fUtilities

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- Description: varianza campionaria
- Formula:

$$s_x^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 = \frac{1}{n-1} \sum_{i=1}^n x_i^2 - \frac{n}{n-1} \bar{x}^2 = \frac{1}{n-1} s s_x = \frac{n}{n-1} \sigma_x^2$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> n <- 5
> sum((x - mean(x))^2)/(n - 1)

[1] 10.282

> var(x)

[1] 10.282

> x <- c(1.2, 3.4, 5.6, 3.7, 7.8, 8.5)
> n <- 6
> sum((x - mean(x))^2)/(n - 1)

[1] 7.826667

> var(x)
```

ssdev()

• Package: sigma2tools

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• **Description**: devianza

• Formula:

$$ss_x = \sum_{i=1}^n (x_i - \bar{x})^2 = \sum_{i=1}^n x_i^2 - n \,\bar{x}^2 = (n-1) \,s_x^2 = n \,\sigma_x^2$$

• Examples:

```
> x <- c(1, 1.2, 3.4, 0.8)
> sum((x - mean(x))^2)

[1] 4.4
> ssdev(x)

[1] 4.4
> x <- c(1.2, 2.3, 4.5, 6.5)
> sum((x - mean(x))^2)

[1] 16.6675
> ssdev(x)

[1] 16.6675
```

3.11 Covarianza e codevianza

COV()

• Package: labstatR

• Input:

x vettore numerico di dimensione n

y vettore numerico di dimensione n

• Description: covarianza nella popolazione

• Formula:

$$\sigma_{xy} = \frac{1}{n} \sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y}) = \frac{1}{n} \sum_{i=1}^{n} x_i y_i - \bar{x} \bar{y} = \frac{1}{n} s s_{xy} = \frac{n-1}{n} s_{xy}$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> y <- c(1.2, 3.4, 4.5, 6.4, 4)
> mean((x - mean(x)) * (y - mean(y)))

[1] 3.298
> COV(x, y)
[1] 3.298
```

```
> x <- c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
> y <- c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
> mean((x - mean(x)) * (y - mean(y)))

[1] 4.442222

> COV(x, y)
```

cov()

• Package: fUtilities

• Input:

- \times vettore numerico di dimensione n
- y vettore numerico di dimensione n
- Description: covarianza campionaria
- Formula:

$$s_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y}) = \frac{1}{n-1} \sum_{i=1}^{n} x_i y_i - \frac{n}{n-1} \bar{x} \bar{y} = \frac{1}{n-1} s s_{xy} = \frac{n}{n-1} \sigma_{xy}$$

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> y <- c(1.3, 4.2, 3.3, 8.7, 3.7)
> n <- 5
> sum((x - mean(x)) * (y - mean(y)))/(n - 1)

[1] 4.4535

> cov(x, y)

[1] 4.4535

> x <- c(1.5, 6.4, 6.3, 6.7, 7.5, 4.5, 4.2, 7.8)
> y <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4, 3.4)
> n <- 8
> sum((x - mean(x)) * (y - mean(y)))/(n - 1)

[1] 1.970893

> cov(x, y)

[1] 1.970893
```

codev()

• Package: sigma2tools

• Input:

x vettore numerico di dimensione n

y vettore numerico di dimensione n

• **Description:** codevianza

• Formula:

$$ss_{xy} = \sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y}) = \sum_{i=1}^{n} x_i y_i - n \bar{x} \bar{y} = (n-1) s_{xy} = n \sigma_{xy}$$

• Examples:

```
> x <- c(1.5, 6.4, 6.3, 6.7, 7.5)
> y <- c(1.2, 3.4, 4.5, 6.4, 4)
> sum((x - mean(x)) * (y - mean(y)))

[1] 14.03
> codev(x, y)

[1] 14.03
> x <- c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
> y <- c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
> sum((x - mean(x)) * (y - mean(y)))

[1] 26.65333
> codev(x, y)
[1] 26.65333
```

3.12 Matrice di varianza e covarianza

sigma2m()

• Package: sigma2tools

• Input:

x matrice di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici $x_1,\,x_2,\,\ldots,\,x_k$

- Description: matrice di covarianza non corretta
- Formula:

$$s_{x_i x_j} = \frac{1}{n} (x_i - \bar{x}_i)^T (x_j - \bar{x}_j) \quad \forall i, j = 1, 2, \dots, k$$

```
> k <- 2
> x1 <- c(1.3, 4.6, 7.7, 8.4, 12.4)
> x2 <- c(1.2, 3.4, 4.5, 6.4, 4)
> n <- 5
> (n - 1) * var(x1)/n
[1] 13.9576
```

```
> (n - 1) * var(x2)/n
 [1] 2.832
 > (n - 1) * cov(x1, x2)/n
 [1] 4.21
 > x <- cbind(x1, x2)
 > sigma2m(x)
          x1 x2
 x1 13.9576 4.210
 x2 4.2100 2.832
 > k <- 3
 > x1 <- c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
 > x2 <- c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
 > x3 < -c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
 > n < -7
 > (n - 1) * var(x1)/n
 [1] 7.670612
 > (n - 1) * var(x2)/n
 [1] 2.380869
 > (n - 1) * var(x3)/n
 [1] 1042.793
 > (n - 1) * cov(x1, x2)/n
 [1] 0.5416122
 > (n - 1) \star cov(x1, x3)/n
 [1] 56.06959
 > (n - 1) * cov(x2, x3)/n
 [1] 11.56516
 > x \leftarrow cbind(x1, x2, x3)
 > sigma2m(x)
                        x2
             x1
 x1 7.6706122 0.5416122
                              56.06959
 x2 0.5416122 2.3808694
                              11.56516
 x3 56.0695918 11.5651633 1042.79265
• Note: Naturalmente vale che s_{x_i x_i} = s_{x_i}^2 \quad \forall i = 1, 2, ..., k.
```

Var()

- Package: car
- Input:

x matrice di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici x_1, x_2, \ldots, x_k diag = TRUE / FALSE varianze campionarie o matrice di covarianza

- Description: matrice di covarianza
- Formula:

$$\begin{aligned} s_{x_i}^2 &= \frac{1}{n-1} \left(x_i - \bar{x}_i \right)^T \left(x_i - \bar{x}_i \right) & \forall i = 1, 2, \dots, k \\ \\ & \boxed{\text{diag = FALSE}} \\ s_{x_i x_j} &= \frac{1}{n-1} \left(x_i - \bar{x}_i \right)^T \left(x_j - \bar{x}_j \right) & \forall i, j = 1, 2, \dots, k \end{aligned}$$

• Examples:

[1] 7.717

```
> k <- 2
> x1 <- c(0.5, -0.1, 0.2, -1.9, 1.9, 0.7, -1.5, 0, -2.5, 1.6, 0.2,
> x2 \leftarrow c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 6.5, 2, 1.2, 3.4)
> n <- 12
> var(x1)
[1] 1.734545
> var(x2)
[1] 12.89295
> cov(x1, x2)
[1] -1.070909
> x \leftarrow cbind(x1, x2)
> Var(x, diag = TRUE)
       x1
 1.734545 12.892955
> Var(x, diag = FALSE)
          x1
x1 1.734545 -1.070909
x2 -1.070909 12.892955
> k < - 3
> x1 <- c(1.2, 3.4, 5.6, 7.5, 7.7)
> x2 <- c(1.1, 2.1, 4.2, 5.3, 3.3)
> x3 \leftarrow c(1, 2.6, 7.6, 7.7, 7.7)
> n < -5
> var(x1)
```

```
> var(x2)
[1] 2.76
> var(x3)
[1] 10.647
> cov(x1, x2)
[1] 3.965
> cov(x1, x3)
[1] 8.628
> cov(x2, x3)
[1] 4.895
> x \leftarrow cbind(x1, x2, x3)
> Var(x, diag = TRUE)
   x1
        x2 x3
 7.717 2.760 10.647
> Var(x, diag = FALSE)
          x2 x3
      x1
x1 7.717 3.965 8.628
x2 3.965 2.760 4.895
x3 8.628 4.895 10.647
```

• Note: Naturalmente vale che $s_{x_ix_i}=s_{x_i}^2 \quad \forall i=1,2,\ldots,k.$

3.13 Correlazione di Pearson, Spearman e Kendall

```
cor()
```

```
• Package: fUtilities
```

• Input:

```
{\tt x} vettore numerico di dimensione n y vettore numerico di dimensione n method = "pearson" / "spearman" / "kendall" tipo di coefficiente
```

- Description: coefficiente di correlazione
- Formula:

$$r_{xy} = \frac{\sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y})}{\left(\sum_{i=1}^{n} (x_i - \bar{x})^2\right)^{1/2} \left(\sum_{i=1}^{n} (y_i - \bar{y})^2\right)^{1/2}} = \frac{\sum_{i=1}^{n} x_i y_i - n \bar{x} \bar{y}}{\left(\sum_{i=1}^{n} x_i^2 - n \bar{x}^2\right)^{1/2} \left(\sum_{i=1}^{n} y_i^2 - n \bar{y}^2\right)^{1/2}}$$

method = "pearson"

method = "spearman"

$$r_{xy}^{S} = \frac{\sum_{i=1}^{n} (a_{i} - \bar{a}) (b_{i} - \bar{b})}{\left(\sum_{i=1}^{n} (a_{i} - \bar{a})^{2}\right)^{1/2} \left(\sum_{i=1}^{n} (b_{i} - \bar{b})^{2}\right)^{1/2}} = \frac{\sum_{i=1}^{n} a_{i} b_{i} - n ((n+1)/2)^{2}}{\left(\sum_{i=1}^{n} a_{i}^{2} - n ((n+1)/2)^{2}\right)^{1/2} \left(\sum_{i=1}^{n} b_{i}^{2} - n ((n+1)/2)^{2}\right)^{1/2}}$$

dove a, b sono i ranghi di x ed y rispettivamente.

$$r_{xy}^{K} = \frac{2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n} \operatorname{sign}((x_{j} - x_{i}) (y_{j} - y_{i}))}{\left(n\left(n-1\right) - \sum_{i=1}^{g} t_{i} (t_{i} - 1)\right)^{1/2}\left(n\left(n-1\right) - \sum_{j=1}^{h} u_{j} (u_{j} - 1)\right)^{1/2}}$$

dove t, u sono i ties di x ed y rispettivamente.

```
> x <- c(1, 2, 2, 4, 3, 3)
> y <- c(6, 6, 7, 7, 7, 9)
> cov(x, y)/(sd(x) * sd(y))
[1] 0.522233
> cor(x, y, method = "pearson")
[1] 0.522233
> x < -c(1, 2, 3, 5.6, 7.6, 2.3, 1)
> y < -c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)
> a <- rank(x)
> b <- rank(y)
> rhoS <- cov(a, b)/(sd(a) * sd(b))
> rhoS
[1] 0.9908674
> cor(x, y, method = "spearman")
[1] 0.9908674
> x < -c(1, 2, 2, 4, 3, 3)
> y <- c(6, 6, 7, 7, 7, 9)
> matrice <- matrix(0, nrow = n - 1, ncol = n, byrow = FALSE)
> for (i in 1:(n - 1)) for (j in (i + 1):n) matrice[i, j] <- sign((x[j] -
+ x[i]) * (y[j] - y[i]))
> table(rank(x))
  1 2.5 4.5 6
  1 2 2 1
> g < - 2
> t1 <- 2
> t2 <- 2
> t <- c(t1, t2)
> t
```

```
[1] 2 2
> table(rank(y))
1.5 4 6
        1
 2
    3
> h <- 2
> u1 <- 2
> u2 <- 3
> u <- c(u1, u2)
> u
[1] 2 3
> rhoK <- (2 * sum(matrice))/((n * (n - 1) - sum(t * (t - 1)))^0.5 *
+ (n * (n - 1) - sum(u * (u - 1)))^0.5)
> rhoK
[1] 0.5853694
> cor(x, y, method = "kendall")
[1] 0.5853694
> x < -c(1, 2, 3, 5.6, 7.6, 2.3, 1)
> y <- c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)
> cov(x, y)/(sd(x) * sd(y))
[1] 0.8790885
> cor(x, y, method = "pearson")
[1] 0.8790885
> x < -c(1, 2, 2, 4, 3, 3)
> y < -c(6, 6, 7, 7, 7, 9)
> a <- rank(x)
> b <- rank(y)
> rhoS <- cov(a, b)/(sd(a) * sd(b))
> rhoS
[1] 0.6833149
> cor(x, y, method = "spearman")
[1] 0.6833149
> x < -c(1, 2, 3, 5.6, 7.6, 2.3, 1)
> y <- c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)
> n < -7
> matrice <- matrix(0, nrow = n - 1, ncol = n, byrow = FALSE)
> for (i in 1:(n - 1)) for (j in (i + 1):n) matrice[i, j] <- sign((x[j] -
     x[i]) * (y[j] - y[i]))
> table(rank(x))
1.5 3 4 5 6 7
    1 1 1 1
 2
```

```
> g < -1
> t <- 2
> table(rank(y))
1.5 3.5 5 6 7
 2 2 1 1 1
> h <- 2
> u1 <- 2
> u2 <- 2
> u <- c(u1, u2)
> u
[1] 2 2
> rhoK <- (2 * sum(matrice))/((n * (n - 1) - sum(t * (t - 1)))^0.5 *
+ (n * (n - 1) - sum(u * (u - 1)))^0.5)
> rhoK
[1] 0.9746794
> cor(x, y, method = "kendall")
[1] 0.9746794
```

cov2cor()

• Package: stats

• Input:

V matrice di covarianza di dimensione $k \times k$ relativa ai vettori numerici x_1, x_2, \ldots, x_k

• Description: converte la matrice di covarianza nella matrice di correlazione

• Formula:

$$r_{x_i x_j} \, = \, \frac{\sigma_{x_i x_j}}{\sigma_{x_i} \, \sigma_{x_j}} \, = \, \frac{s_{x_i x_j}}{s_{x_i} \, s_{x_j}} \, = \, \frac{s s_{x_i x_j}}{\sqrt{s s_{x_i} \, s s_{x_j}}} \quad \forall \, i, j \, = \, 1, \, 2, \, \ldots, \, k$$

• Examples:

```
> x2 <- c(1, 2, 3, 5, 6, 7.3)
> dati <- cbind(x1, x2)
> dati
      x1 x2
[1,] -1.2 1.0
[2,] -1.3 2.0
[3,] -6.7 3.0
[4,] 0.8 5.0
[5,] -7.6 6.0
[6,] -5.6 7.3
> n <- 6
> k < - 2
> V <- cov(dati)
> V
      x1 x2
x1 12.004 -3.780
x2 - 3.780 5.975
```

> x1 < -c(-1.2, -1.3, -6.7, 0.8, -7.6, -5.6)

```
> cor(dati)
            x1 x2
 x1 1.0000000 -0.4463339
 x2 -0.4463339 1.0000000
 > cov2cor(V)
            x1
 x1 1.0000000 -0.4463339
 x2 -0.4463339 1.0000000
 > x1 \leftarrow c(1, 2, 4.5, 1.2, 1.23)
 > x2 \leftarrow c(2.7, -7.8, 8.8, 4.5, 5.21)
 > x3 < -c(1, 4.77, 8.9, 7.8, 0.8)
 > dati <- cbind(x1, x2, x3)
 > dati
       x1 x2 x3
 [1,] 1.00 2.70 1.00
 [2,] 2.00 -7.80 4.77
 [3,] 4.50 8.80 8.90
 [4,] 1.20 4.50 7.80
 [5,] 1.23 5.21 0.80
 > n <- 5
 > k < - 3
 > V <- cov(dati)
 > V
         x1
               x2 x3
 x1 2.120480 2.969010 3.679945
 x2 2.969010 39.249620 5.167965
 x3 3.679945 5.167965 14.036080
 > cor(dati)
                    x2
           x1
 x1 1.0000000 0.3254444 0.6745301
 x2 0.3254444 1.0000000 0.2201805
 x3 0.6745301 0.2201805 1.0000000
 > cov2cor(V)
                x2 x3
           x1
 x1 1.0000000 0.3254444 0.6745301
 x2 0.3254444 1.0000000 0.2201805
 x3 0.6745301 0.2201805 1.0000000
• Note: Naturalmente vale che s_{x_i x_i} = s_{x_i}^2 \quad \forall i = 1, 2, ..., k.
```

179

cancor()

• Package: stats

• Input:

x vettore numerico di dimensione n
y vettore numerico di dimensione n
xcenter = TRUE / FALSE parametro di posizione
ycenter = TRUE / FALSE parametro di posizione

- Description: correlazione canonica
- Output:

cor coefficiente di correlazione xcenter parametro di locazione ycenter parametro di locazione

• Formula:

cor

$$xcenter = TRUE \ \ AND \ \ ycenter = TRUE$$

$$r_{xy} = \frac{\sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y})}{\left(\sum_{i=1}^{n} (x_i - \bar{x})^2\right)^{1/2} \left(\sum_{i=1}^{n} (y_i - \bar{y})^2\right)^{1/2}}$$

$$xcenter = TRUE \ \ AND \ \ ycenter = FALSE$$

$$r_{xy} = \frac{\sum_{i=1}^{n} (x_i - \bar{x}) y_i}{\left(\sum_{i=1}^{n} (x_i - \bar{x})^2\right)^{1/2} \left(\sum_{i=1}^{n} y_i^2\right)^{1/2}}$$

$$xcenter = FALSE \ \ AND \ \ ycenter = TRUE$$

$$r_{xy} = \frac{\sum_{i=1}^{n} x_i (y_i - \bar{y})}{\left(\sum_{i=1}^{n} x_i^2\right)^{1/2} \left(\sum_{i=1}^{n} (y_i - \bar{y})^2\right)^{1/2}}$$

$$xcenter = FALSE \ \ AND \ \ ycenter = FALSE$$

$$r_{xy} = \frac{\sum_{i=1}^{n} x_i y_i}{\left(\sum_{i=1}^{n} x_i^2\right)^{1/2} \left(\sum_{i=1}^{n} y_i^2\right)^{1/2}}$$

$$xcenter = TRUE$$

$$\bar{x}$$

$$xcenter = TRUE$$

$$\bar{y}$$

$$ycenter = TRUE$$

$$\bar{y}$$

$$ycenter = FALSE$$

• Examples:

xcenter

ycenter

```
> x <- c(1, 2, 3, 5.6, 7.6, 2.3, 1)

> y <- c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)

> n <- 7

> sum((x - mean(x)) * (y - mean(y)))/(sum((x - mean(x))^2)^0.5 *

+ sum((y - mean(y))^2)^0.5)
```

```
[1] 0.8790885
> cancor(x, y, xcenter = TRUE, ycenter = TRUE)$cor
[1] 0.8790885
> mean(x)
[1] 3.214286
> cancor(x, y, xcenter = TRUE, ycenter = TRUE)$xcenter
[1] 3.214286
> mean(y)
[1] 13.85714
> cancor(x, y, xcenter = TRUE, ycenter = TRUE) $ycenter
[1] 13.85714
> sum((x - mean(x)) * y)/(sum((x - mean(x))^2)^0.5 * sum(y^2)^0.5)
[1] 0.7616638
> cancor(x, y, xcenter = TRUE, ycenter = FALSE)$cor
[1] 0.7616638
> mean(x)
[1] 3.214286
> cancor(x, y, xcenter = TRUE, ycenter = FALSE)$xcenter
[1] 3.214286
> cancor(x, y, xcenter = TRUE, ycenter = FALSE) $ycenter
[1] 0
> sum(x * (y - mean(y)))/(sum(x^2)^0.5 * sum((y - mean(y))^2)^0.5)
[1] 0.5118281
> cancor(x, y, xcenter = FALSE, ycenter = TRUE)$cor
[1] 0.5118281
> cancor(x, y, xcenter = FALSE, ycenter = TRUE) $xcenter
[1] 0
> mean(y)
[1] 13.85714
```

```
> cancor(x, y, xcenter = FALSE, ycenter = TRUE)$ycenter
[1] 13.85714
> sum(x * y)/(sum(x^2)^0.5 * sum(y^2)^0.5)
[1] 0.8494115
> cancor(x, y, xcenter = FALSE, ycenter = FALSE)$cor
[1] 0.8494115
> cancor(x, y, xcenter = FALSE, ycenter = FALSE)$xcenter
[1] 0
> cancor(x, y, xcenter = FALSE, ycenter = FALSE) $ycenter
[1] 0
> x < -c(1.2, 2.3, 4.5, 3.2, 4.7)
> y < -c(1.8, 9.87, 7.5, 6.6, 7.7)
> n < -5
> sum((x - mean(x)) * (y - mean(y)))/(sum((x - mean(x))^2)^0.5 *
  sum((y - mean(y))^2)^0.5)
[1] 0.536735
> cancor(x, y, xcenter = TRUE, ycenter = TRUE)$cor
[1] 0.536735
> mean(x)
[1] 3.18
> cancor(x, y, xcenter = TRUE, ycenter = TRUE) $xcenter
[1] 3.18
> mean(y)
[1] 6.694
> cancor(x, y, xcenter = TRUE, ycenter = TRUE) $ycenter
[1] 6.694
> sum((x - mean(x)) * y)/(sum((x - mean(x))^2)^0.5 * sum(y^2)^0.5)
[1] 0.1990048
> cancor(x, y, xcenter = TRUE, ycenter = FALSE)$cor
[1] 0.1990048
> mean(x)
```

```
[1] 3.18
> cancor(x, y, xcenter = TRUE, ycenter = FALSE) $xcenter
[1] 3.18
> cancor(x, y, xcenter = TRUE, ycenter = FALSE) $ycenter
[1] 0
> sum(x * (y - mean(y)))/(sum(x^2)^0.5 * sum((y - mean(y))^2)^0.5)
[1] 0.2061343
> cancor(x, y, xcenter = FALSE, ycenter = TRUE) $cor
[1] 0.2061343
> cancor(x, y, xcenter = FALSE, ycenter = TRUE) $xcenter
[1] 0
> mean(y)
[1] 6.694
> cancor(x, y, xcenter = FALSE, ycenter = TRUE)$ycenter
[1] 6.694
> sum(x * y)/(sum(x^2)^0.5 * sum(y^2)^0.5)
[1] 0.9339306
> cancor(x, y, xcenter = FALSE, ycenter = FALSE)$cor
[1] 0.9339306
> cancor(x, y, xcenter = FALSE, ycenter = FALSE)$xcenter
[1] 0
> cancor(x, y, xcenter = FALSE, ycenter = FALSE)$ycenter
[1] 0
```

partial.cor()

- Package: Rcmdr
- Input:

X matrice di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici x_1, x_2, \ldots, x_k

- **Description:** correlazione parziale
- Formula:

$$r_{x_i x_j|.} = -\frac{R_{i,j}^{-1}}{\sqrt{R_{i,i}^{-1} R_{j,j}^{-1}}} \quad \forall i \neq j = 1, 2, \dots, k$$

dove R è la matrice di correlazione tra i k vettori

```
> k < - 3
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> X <- cbind(x1, x2, x3)
     x1 x2 x3
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
[8,] 6.7 7.6 8.70
> n <- 8
> R \leftarrow cor(X)
> RI <- solve(R)
> D <- 1/sqrt(diag(RI))
> mat <- -RI * (D %o% D)
> diag(mat) <- 0
> mat
                 x2
x1 0.0000000 0.8221398 -0.4883764
x2 0.8221398 0.0000000 0.8022181
x3 -0.4883764 0.8022181 0.0000000
> partial.cor(X)
                 x2
           x1
x1 0.0000000 0.8221398 -0.4883764
x2 0.8221398 0.0000000 0.8022181
x3 -0.4883764 0.8022181 0.0000000
> k <- 2
> x1 <- c(-1.2, -1.3, -6.7, 0.8, -7.6, -5.6)
> x2 < -c(1, 2, 3, 5, 6, 7.3)
> X \leftarrow cbind(x1, x2)
> X
```

```
x1 x2
[1,] -1.2 1.0
[2,] -1.3 2.0
[3,] -6.7 3.0
[4,] 0.8 5.0
[5,] -7.6 6.0
[6,] -5.6 7.3
> n < -6
> R < - cor(X)
> RI <- solve(R)
> D <- 1/sqrt(diag(RI))</pre>
> mat <- -RI * (D %o% D)
> diag(mat) <- 0
> mat
           x1
                      x2
x1 0.0000000 -0.4463339
x2 -0.4463339 0.0000000
> partial.cor(X)
           x1
x1 0.0000000 -0.4463339
x2 -0.4463339 0.0000000
```

cor2pcor()

- Package: corpcor
- Input:

m matrice di covarianza o di correlazione di dimensione $n \times k$ dei vettori numerici x_1, x_2, \ldots, x_k

- **Description:** correlazione parziale
- Formula:

$$r_{x_i x_j \mid \cdot} = -\frac{R_{i,j}^{-1}}{\sqrt{R_{i,i}^{-1} R_{j,j}^{-1}}} \quad \forall i, j = 1, 2, \dots, k$$

dove R è la matrice di correlazione tra i k vettori

• Example 1:

```
> n < - 8
 > R <- cor(X)
 > RI <- solve(R)</pre>
 > D <- 1/sqrt(diag(RI))</pre>
 > mat <- -RI * (D %o% D)
 > diag(mat) <- 1</pre>
 > mat
            x1
                     x2
 x1 1.0000000 0.8221398 -0.4883764
 x2 0.8221398 1.0000000 0.8022181
 x3 -0.4883764 0.8022181 1.0000000
 > cor2pcor(m = cor(X))
                   [,2]
            [,1]
                                 [,3]
 [1,] 1.0000000 0.8221398 -0.4883764
 [2,] 0.8221398 1.0000000 0.8022181
 [3,] -0.4883764 0.8022181 1.0000000
 > cor2pcor(m = cov(X))
            [,1]
                      [,2]
                                  [,3]
 [1,] 1.0000000 0.8221398 -0.4883764
 [2,] 0.8221398 1.0000000 0.8022181
 [3,] -0.4883764 0.8022181 1.0000000
• Example 2:
 > k < - 2
 > x1 <- c(-1.2, -1.3, -6.7, 0.8, -7.6, -5.6)
 > x2 < -c(1, 2, 3, 5, 6, 7.3)
 > X < - cbind(x1, x2)
 > X
       x1 x2
 [1,] -1.2 1.0
 [2,] -1.3 2.0
 [3,] -6.7 3.0
 [4,] 0.8 5.0
 [5,] -7.6 6.0
 [6,] -5.6 7.3
 > n < -6
 > R <- cor(X)
 > RI <- solve(R)</pre>
 > D <- 1/sqrt(diag(RI))
 > mat <- -RI * (D %o% D)
 > diag(mat) <- 1</pre>
 > mat
            x1
 x1 1.0000000 -0.4463339
 x2 -0.4463339 1.0000000
 > cor2pcor(m = cor(X))
            [,1]
                      [,2]
 [1,] 1.0000000 -0.4463339
 [2,] -0.4463339 1.0000000
```

pcor2cor()

- Package: corpcor
- Input:

m matrice di correlazione parziale di dimensione $k \times k$ dei vettori numerici x_1, x_2, \ldots, x_k

- **Description:** correlazione parziale
- Formula:

$$r_{x_i x_j} \, = \, \frac{\sigma_{x_i x_j}}{\sigma_{x_i} \, \sigma_{x_j}} \, = \, \frac{s_{x_i x_j}}{s_{x_i} \, s_{x_j}} \, = \, \frac{s s_{x_i x_j}}{\sqrt{s s_{x_i} \, s s_{x_j}}} \quad \forall \, i,j \, = \, 1, \, 2, \, \ldots, \, k$$

```
> k < - 3
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> X \leftarrow cbind(x1, x2, x3)
> X
      x1 x2 x3
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
[8,] 6.7 7.6 8.70
> n < - 8
> cor(X)
          x1
                    x2
x1 1.0000000 0.8260355 0.5035850
x2 0.8260355 1.0000000 0.8066075
x3 0.5035850 0.8066075 1.0000000
> mat <- cor2pcor(cor(X))</pre>
> mat
           [,1]
                     [,2]
[1,] 1.0000000 0.8221398 -0.4883764
[2,] 0.8221398 1.0000000 0.8022181
[3,] -0.4883764 0.8022181 1.0000000
> pcor2cor(m = mat)
                    [,2]
          [,1]
                            [,3]
[1,] 1.0000000 0.8260355 0.5035850
[2,] 0.8260355 1.0000000 0.8066075
[3,] 0.5035850 0.8066075 1.0000000
```

```
> x1 < -c(-1.2, -1.3, -6.7, 0.8, -7.6, -5.6)
> x2 <- c(1, 2, 3, 5, 6, 7.3)
> X \leftarrow cbind(x1, x2)
> X
      x1 x2
[1,] -1.2 1.0
[2,] -1.3 2.0
[3,] -6.7 3.0
[4,] 0.8 5.0
[5,] -7.6 6.0
[6,] -5.6 7.3
> n <- 6
> cor(X)
x1 1.0000000 -0.4463339
x2 -0.4463339 1.0000000
> mat <- cor2pcor(m = cor(X))</pre>
> cor2pcor(m = mat)
           [,1]
                 [,2]
[1,] 1.0000000 -0.4463339
[2,] -0.4463339 1.0000000
```

3.14 Media e varianza pesate

weighted.mean()

- Input:
- Package: stats
 - $\times\,$ vettore numerico di dimensione n
 - $\ensuremath{\mathbf{w}}$ vettore numerico w di pesi di dimensione n
- Description: media pesata
- Formula:

$$\bar{x}_W = \frac{\sum_{i=1}^n x_i w_i}{\sum_{j=1}^n w_j}$$

```
> x <- c(3.7, 3.3, 3.5, 2.8)
> w <- c(5, 5, 4, 1)
> sum(w)

[1] 15
> sum(x * w)/sum(w)

[1] 3.453333
> weighted.mean(x, w)

[1] 3.453333
```

```
> x <- c(3.7, 3.3, 3.5, 2.8)
> w <- c(0.16, 0.34, 0.28, 0.22)
> sum(w)

[1] 1
> sum(x * w)

[1] 3.31
> weighted.mean(x, w)
```

wt.var()

- Input:
- Package: corpcor

xvec vettore numerico di dimensione n

- ${\tt w}\$ vettore numerico w di pesi a somma unitaria di dimensione n
- Description: varianza pesata
- Formula:

$$s_x^2 = (1 - w^T w)^{-1} (x - \bar{x}_W)^T W^{-1} (x - \bar{x}_W)$$

```
> x < -c(3.7, 3.3, 3.5, 2.8)
> w < -c(5, 5, 4, 1)
> w <- w/sum(w)
> xW <- sum(x * w)
> W <- diag(1/w)
> as.numeric(1/(1 - t(w) %*% w) * t(x - xW) %*% solve(W) %*% (x - xW) %*% (x - xW) %*% solve(W) %*% (x - xW) %*% (x - xW
                                 xW))
[1] 0.0813924
> wt.var(xvec = x, w)
[1] 0.0813924
> x < -c(3.7, 3.3, 3.5, 2.8)
> w < -c(0.16, 0.34, 0.28, 0.22)
> xW <- sum(x * w)
> W <- diag(1/w)
> as.numeric(1/(1 - t(w) %*% w) * t(x - xW) %*% solve(W) %*% (x -
                              xW))
[1] 0.1252732
> wt.var(xvec = x, w)
[1] 0.1252732
```

wt.moments()

- Package: corpcor
- Input:
 - x matrice di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici x_1, x_2, \ldots, x_k
 - ${\tt w}\$ vettore numerico w di pesi a somma unitaria di dimensione n
- Description: media e varinza pesate pesata
- Output:

```
mean medie pesate var varianze pesate
```

• Formula:

mean

$$\bar{x}_{iW} \quad \forall i = 1, 2, \ldots, k$$

var

$$s_{x_i}^2 = (1 - w^T w)^{-1} (x_i - \bar{x}_{iW})^T W^{-1} (x_i - \bar{x}_{iW}) \quad \forall i = 1, 2, ..., k$$

• Examples 1:

```
> k <- 2
> x1 <- c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
> x2 < -c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
> x \leftarrow cbind(x1, x2)
> n <- 6
> w < -c(0.16, 0.34, 0.28, 0.12, 0.08, 0.02)
> xW1 <- sum(x1 * w)
> xW2 <- sum(x2 * w)
> c(xW1, xW2)
[1] 4.588 3.208
> wt.moments(x, w)$mean
             x1
                                   x2
4.588 3.208
> W <- diag(1/w)
> var1 <- as.numeric(1/(1 - t(w) %*% w) * t(x1 - xW1) %*% solve(W) %*%
                         (x1 - xW1))
> var2 <- as.numeric(1/(1 - t(w) %*% w) * t(x2 - xW2) %*% solve(W) %*% w) * t(x2 - xW2) %*% w) * t(x2 - xW2) %*% solve(W) %*% w) * t(x2 - xW2) %*% w) * t(x
                         (x2 - xW2))
> c(var1, var2)
[1] 6.061454 3.200126
> wt.moments(x, w)$var
                                                                x2
                            x1
6.061454 3.200126
```

```
> k < - 3
          > x1 \leftarrow c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
          > x2 < -c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
          > x3 < -c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
          > x \leftarrow cbind(x1, x2, x3)
          > n < -7
          > w \leftarrow c(0.16, 0.34, 0.15, 0.12, 0.08, 0.03, 0.12)
          > xW1 <- sum(x1 * w)
          > xW2 <- sum(x2 * w)
          > xW3 < - sum(x3 * w)
          > c(xW1, xW2, xW3)
          [1] 4.7940 6.0606 14.0310
          > wt.moments(x, w)$mean
                                               x2
                        \times 1
             4.7940 6.0606 14.0310
          > W <- diag(1/w)
          > var1 <- as.numeric(1/(1 - t(w) %*% w) * t(x1 - xW1) %*% solve(W) %*%
                            (x1 - xW1))
          > var2 <- as.numeric(1/(1 - t(w) %*% w) * t(x2 - xW2) %*% solve(W) %*%
                           (x2 - xW2))
          > var3 <- as.numeric(1/(1 - t(w) %*% w) * t(x3 - xW3) %*% solve(W) %*% w) * t(x3 - xW3) %*% w) * t
                           (x3 - xW3))
          > c(var1, var2, var3)
           [1]
                           8.159415
                                                       3.336630 781.977429
          > wt.moments(x, w)$var
                                  x1
                                                                 x2
                                                                                                  x3
                8.159415 3.336630 781.977429
cov.wt()
     • Package: stats
     • Input:
                     x matrice di dimensione n \times k le cui colonne corrispondono ai vettori numerici x_1, x_2, \ldots, x_k
                     {\sf wt}\; vettore numerico w di pesi a somma unitaria di dimensione n
                     center = TRUE / FALSE parametro di posizione
                     cor = TRUE / FALSE correlazione pesata
     • Description: matrice di covarianza e correlazione pesata
     • Output:
                     cov matrice di covarianza pesata
```

center media pesata

n.obs dimensione campionaria

wt vettore numerico w

cor matrice di correlazione pesata

• Formula:

COV

(x2 - x2W))

```
[1] 5.330667
 > z \leftarrow cbind(x1, x2)
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE)$cov
          x1
 x1 7.406667 5.330667
 x2 5.330667 7.185667
 > as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% x1)
 [1] 44.148
 > as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*% x2)
 [1] 27.194
 > as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% x2)
 [1] 32.444
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE)$cov
        x1
              x2
 x1 44.148 32.444
 x2 32.444 27.194
• Examples 2:
 > k <- 2
 > x1 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
 > x2 \leftarrow c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
 > n < -6
 > w < - rep(1/n, times = n)
 > sum(w)
 [1] 1
 > x1W <- sum(x1 * w)
 > x2W <- sum(x2 * w)
 > W <- diag(1/w)
 > c(x1W, x2W)
 [1] 5.533333 4.083333
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE)$center
       x1
 5.533333 4.083333
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE) $center
 [1] 0
```

```
> k < - 2
 > x1 <- c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
 > x2 < -c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
 > n < -6
 > w < - rep(1/n, times = n)
 > sum(w)
 [1] 1
 > n
 [1] 6
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE)$n.obs
 [1] 6
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE)$n.obs
 [1] 6
• Example 4:
 > x1 <- c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
 > x2 \leftarrow c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
 > n <- 6
 > w < - rep(1/n, times = n)
 > sum(w)
 [1] 1
 > w
 [1] 0.1666667 0.1666667 0.1666667 0.1666667 0.1666667
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE) $wt
 [1] 0.1666667 0.1666667 0.1666667 0.1666667 0.1666667
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE) $wt
 [1] 0.1666667 0.1666667 0.1666667 0.1666667 0.1666667
• Example 5:
 > k <- 2
 > x1 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.7, 7.8)
 > x2 <- c(1.1, 2.3, 4.4, 5.1, 2.9, 8.7)
 > n < -6
 > w <- rep(1/n, times = n)
 > sum(w)
 [1] 1
 > x1W <- sum(x1 * w)
 > x2W <- sum(x2 * w)
 > W <- diag(1/w)
 > covx1x2 < -1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
      (x2 - x2W)
 > covx1x2 <- as.numeric(covx1x2)</pre>
 > covx1x2
```

```
[1] 5.330667
> sx1 < - sqrt(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
          (x1 - x1W))
> sx1 <- as.numeric(sx1)</pre>
> sx1
[1] 2.721519
> sx2 < - sqrt(1/(1 - t(w) %*% w) * t(x2 - x2W) %*% solve(W) %*%
+ (x2 - x2W))
> sx2 <- as.numeric(sx2)</pre>
> sx2
[1] 2.680609
> rx1x2 <- covx1x2/(sx1 * sx2)
> rx1x2
[1] 0.7306958
> cov.wt(z, wt = w, center = TRUE, cor = TRUE)$cor
                               x1
x1 1.0000000 0.7306958
x2 0.7306958 1.0000000
> covx1x2 <- as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*%
+ x2)
> covx1x2
[1] 32.444
> sx1 <- sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% w) * t(x1) %*% w) *
+ x1))
> sx1
[1] 6.644396
> sx2 < - sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*%
+ x2))
> sx2
[1] 5.214787
> rx1x2 <- covx1x2/(sx1 * sx2)
> rx1x2
[1] 0.9363589
> cov.wt(z, wt = w, center = FALSE, cor = TRUE) $cor
                                 x1
x1 1.0000000 0.9363589
x2 0.9363589 1.0000000
```

• Example 6:

```
> k < - 3
> x1 \leftarrow c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
> x2 <- c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
> x3 \leftarrow c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
> n < -7
> w < rep(1/n, times = n)
> sum(w)
[1] 1
> x1W <- sum(x1 * w)
> x2W <- sum(x2 * w)
> x3W <- sum(x3 * w)
> W <- diag(1/w)
> as.numeric(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
     (x1 - x1W))
[1] 8.949048
> as.numeric(1/(1 - t(w) %*% w) * t(x2 - x2W) %*% solve(W) %*%
+ (x2 - x2W))
[1] 2.777681
> as.numeric(1/(1 - t(w) %*% w) * t(x3 - x3W) %*% solve(W) %*%
+ (x3 - x3W))
[1] 1216.591
> as.numeric(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
+ (x2 - x2W))
[1] 0.631881
> as.numeric(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
+ (x3 - x3W))
[1] 65.41452
> as.numeric(1/(1 - t(w) %*% w) * t(x2 - x2W) %*% solve(W) %*%
+ (x3 - x3W))
[1] 13.49269
> z <- cbind(x1, x2, x3)
> cov.wt(z, wt = w, center = TRUE, cor = TRUE)$cov
                   x2
          x1
                               x3
x1 8.949048 0.631881
                         65.41452
x2 0.631881 2.777681
                       13.49269
x3 65.414524 13.492690 1216.59143
> as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% x1)
[1] 47.235
> as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*% x2)
[1] 39.34568
```

```
> as.numeric(1/(1 - t(w) %*% w) * t(x3) %*% solve(W) %*% x3)
 [1] 1665.432
 > as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% x2)
 [1] 38.049
 > as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% x3)
 [1] 196.5033
 > as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*% x3)
 [1] 141.6067
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE)$cov
           x1
                     x2
                               x3
 x1 47.2350 38.04900 196.5033
 x2 38.0490 39.34568 141.6067
 x3 196.5033 141.60667 1665.4317
• Example 7:
 > k < - 3
 > x1 <- c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
 > x2 <- c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
 > x3 \leftarrow c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
 > n < -7
 > w < - rep(1/n, times = n)
 > sum(w)
 [1] 1
 > c(x1W, x2W, x3W)
 [1] 5.728571 5.598571 19.614286
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE) $center
        x1
                   x2
  5.728571 5.598571 19.614286
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE) $center
 [1] 0
• Example 8:
 > k < - 3
 > x1 \leftarrow c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
 > x2 \leftarrow c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
 > x3 <- c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
 > n < -7
 > w <- rep(1/n, times = n)
 > sum(w)
 [1] 1
```

```
> n
 [1] 7
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE) $n.obs
 [1] 7
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE)$n.obs
 [1] 7
• Example 9:
 > k < - 3
 > x1 \leftarrow c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
 > x2 <- c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
 > x3 < -c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
 > n < -7
 > w <- rep(1/n, times = n)
 > sum(w)
 [1] 1
 > w
 [1] 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571
 > cov.wt(z, wt = w, center = TRUE, cor = TRUE)$wt
 [1] 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571
 > cov.wt(z, wt = w, center = FALSE, cor = TRUE) $wt
 [1] 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571 0.1428571
• Example 10:
 > k < - 3
 > x1 <- c(1.1, 3.6, 7.4, 6.8, 9.8, 7.6, 3.8)
 > x2 <- c(5.6, 7.54, 7.3, 3.5, 6.45, 5.4, 3.4)
 > x3 < -c(2.8, 8.5, 6.4, 7.8, 98.6, 7.5, 5.7)
 > n < -7
 > w < rep(1/n, times = n)
 > sum(w)
 [1] 1
 > x1W <- sum(x1 * w)
 > x2W <- sum(x2 * w)
 > x3W <- sum(x3 * w)
 > W <- diag(1/w)
 > covx1x2 < -1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
       (x2 - x2W)
 > covx1x2 <- as.numeric(covx1x2)</pre>
 > covx1x2
 [1] 0.631881
```

```
> covx1x3 < -1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
+ (x3 - x3W)
> covx1x3 <- as.numeric(covx1x3)</pre>
> covx1x3
[1] 65.41452
> covx2x3 < -1/(1 - t(w) %*% w) * t(x2 - x2W) %*% solve(W) %*%
+ (x3 - x3W)
> covx2x3 <- as.numeric(covx2x3)</pre>
> covx2x3
[1] 13.49269
> sx1 < - sqrt(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*%
+ (x1 - x1W))
> sx1 <- as.numeric(sx1)</pre>
> sx1
[1] 2.991496
> sx2 < - sqrt(1/(1 - t(w) %*% w) * t(x2 - x2W) %*% solve(W) %*%
+ (x2 - x2W))
> sx2 <- as.numeric(sx2)</pre>
> sx2
[1] 1.666638
> sx3 < - sqrt(1/(1 - t(w) %*% w) * t(x3 - x3W) %*% solve(W) %*%
+ (x3 - x3W))
> sx3 <- as.numeric(sx3)</pre>
> sx3
[1] 34.87967
> rx1x2 <- covx1x2/(sx1 * sx2)
> rx1x2
[1] 0.1267377
> rx1x3 <- covx1x3/(sx1 * sx3)
> rx1x3
[1] 0.6269218
> rx2x3 \leftarrow covx2x3/(sx2 * sx3)
> rx2x3
[1] 0.2321053
> cov.wt(z, wt = w, center = TRUE, cor = TRUE)$cor
          x1
                    x2
                              x3
x1 1.0000000 0.1267377 0.6269218
x2 0.1267377 1.0000000 0.2321053
x3 0.6269218 0.2321053 1.0000000
> covx1x2 < - as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*%
     x2)
> covx1x2
```

```
[1] 38.049
    > covx1x3 < - as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*%
    > covx1x3
    [1] 196.5033
    > covx2x3 < - as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*%
            x3)
    > covx2x3
    [1] 141.6067
    > sx1 < - sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x1 - x1W) %*% solve(W) %*% w) * t(x1 - x1W) %*% w) * t(x1 - x1
    + (x1 - x1W)))
    > sx1
    [1] 2.991496
    > sx1 <- sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x1) %*% solve(W) %*% )
    + x1))
    > sx1
    [1] 6.872772
    > sx2 < - sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x2) %*% solve(W) %*%
    + x2))
    > sx2
    [1] 6.272614
    > sx3 < - sqrt(as.numeric(1/(1 - t(w) %*% w) * t(x3) %*% solve(W) %*%
                  x3))
    > sx3
    [1] 40.8097
    > rx1x2 <- covx1x2/(sx1 * sx2)
    > rx1x2
    [1] 0.8825976
    > rx1x3 <- covx1x3/(sx1 * sx3)
    > rx1x3
    [1] 0.7006071
    > rx2x3 <- covx2x3/(sx2 * sx3)
    > rx2x3
    [1] 0.5531867
    > cov.wt(z, wt = w, center = FALSE, cor = TRUE) $cor
                                                  x2
                                 x1
    x1 1.0000000 0.8825976 0.7006071
    x2 0.8825976 1.0000000 0.5531867
    x3 0.7006071 0.5531867 1.0000000
• Note 1: W è la matrice diagonale definita positiva di dimensione n \times n tale che W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})
• Note 2: Naturalmente vale che s_{x_i x_i} = s_{x_i}^2 \quad \forall i = 1, 2, ..., k.
```

corr()

- Package: boot
- Input:
 - d matrice di dimensione $n \times 2$ le cui colonne corrispondono ai vettori numerici x ed y
 - ${\tt w}\$ vettore numerico w di pesi a somma unitaria di dimensione n
- Description: correlazione pesata
- Formula:

$$r_{xy} = \frac{(x - \bar{x}_W)^T W^{-1} (y - \bar{y}_W)}{((x - \bar{x}_W)^T W^{-1} (x - \bar{x}_W))^{1/2} ((y - \bar{y}_W)^T W^{-1} (y - \bar{y}_W))^{1/2}}$$

```
> x < -c(1.2, 2.3, 3.4, 4.5, 5.6, 6.7)
> y < -c(1, 2, 3, 5, 6, 7.3)
> d <- as.matrix(cbind(x, y))</pre>
> n <- 6
> w <- abs(rnorm(n))
> w <- w/sum(w)
> sum(w)
[1] 1
> mxw <- weighted.mean(x, w)</pre>
> myw <- weighted.mean(y, w)</pre>
> W <- diag(1/w)
> num <- as.numeric(t(x - mxw) %*% solve(W) %*% (y - myw))
> den <- as.numeric(sqrt(t(x - mxw) %*% solve(W) %*% (x - mxw) *
+ t(y - myw) %*% solve(W) %*% (y - myw)))
> rho <- num/den
> rho
[1] 0.9988987
> corr(d, w)
[1] 0.9988987
> x < -c(1, 2, 3, 5.6, 7.6, 2.3, 1)
> y <- c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)
> d <- as.matrix(cbind(x, y))</pre>
> n < -7
> w <- abs(rnorm(n))</pre>
> w <- w/sum(w)
> sum(w)
[1] 1
> mxw <- weighted.mean(x, w)</pre>
> myw <- weighted.mean(y, w)</pre>
> W <- diag(1/w)
> num <- as.numeric(t(x - mxw) %*% solve(W) %*% (y - myw))
> den <- as.numeric(sqrt(t(x - mxw) %*% solve(W) %*% (x - mxw) *
+ t(y - myw) %*% solve(W) %*% (y - myw)))
> rho <- num/den
> rho
[1] 0.9095326
```

```
> corr(d, w)
[1] 0.9095326
> x < -c(1.1, 2.3, 4.5, 6.7, 8.9)
> y < -c(2.3, 4.5, 6.7, 8.9, 10.2)
> d <- as.matrix(cbind(x, y))</pre>
> n < -5
> w <- rep(1/n, times = n)
> sum(w)
[1] 1
> mxw <- weighted.mean(x, w)</pre>
> myw <- weighted.mean(y, w)</pre>
> W <- diag(1/w)
> num <- as.numeric(t(x - mxw) %*% solve(W) %*% (y - myw))
> den <- as.numeric(sqrt(t(x - mxw) %*% solve(W) %*% (x - mxw) *
  t(y - myw) %*% solve(W) %*% (y - myw)))
> rho <- num/den</pre>
> rho
[1] 0.9866942
> corr(d, w)
[1] 0.9866942
```

• Note: W è la matrice diagonale definita positiva di dimensione $n \times n$ tale che $W = \mathrm{diag}(w_1^{-1},\,w_2^{-1},\,\ldots,\,w_n^{-1})$

3.15 Momenti centrati e non centrati

moment()

- Package: moments
- Input:

```
x vettore numerico di dimensione n
order il valore k dell'ordine
central = TRUE / FALSE parametro di posizione
absolute = TRUE / FALSE modulo
```

- ullet **Description:** momento centrato e non centrato di ordine k
- Formula:

	absolute = TRUE	absolute = FALSE
central = TRUE	$\sum_{i=1}^{n} x_i - \bar{x} ^k / n$	$\sum_{i=1}^{n} (x_i - \bar{x})^k / n$
central = FALSE	$\sum_{i=1}^{n} x_i ^k / n$	$\sum_{i=1}^{n} x_i^k / n$

```
> x <- c(-1.2, 1.2, 3.4, 4.2, 12.4, 13.4, 17.3, 18.1)

> n <- 8

> k <- 5

> mean(abs(x - mean(x))^k)
```

```
[1] 31074.24
> moment(x, central = TRUE, absolute = TRUE, order = 5)
[1] 31074.24
> mean((x - mean(x))^k)
[1] 1565.904
> moment(x, central = TRUE, absolute = FALSE, order = 5)
[1] 1565.904
> mean(abs(x)^k)
[1] 527406.3
> moment(x, central = FALSE, absolute = TRUE, order = 5)
[1] 527406.3
> mean(x^k)
[1] 527405.6
> moment(x, central = FALSE, absolute = FALSE, order = 5)
[1] 527405.6
> x < -c(1.2, 4.5, 6.7, 7.8, 9.8)
> n < -5
> k < - 3
> mean(abs(x - mean(x))^k)
[1] 35.0028
> moment(x, central = TRUE, absolute = TRUE, order = 3)
[1] 35.0028
> mean((x - mean(x))^k)
[1] -10.584
> moment(x, central = TRUE, absolute = FALSE, order = 3)
[1] -10.584
> mean(abs(x)^k)
[1] 361.872
> moment(x, central = FALSE, absolute = TRUE, order = 3)
[1] 361.872
> mean(x^k)
[1] 361.872
> moment(x, central = FALSE, absolute = FALSE, order = 3)
[1] 361.872
```

scale()

• Package: base

• Input:

```
{\tt x} vettore numerico di dimensione n center = TRUE / FALSE parametro di posizione scale = TRUE / FALSE parametro di scala
```

- **Description:** centratura o normalizzazione
- Formula:

	scale = TRUE	scale = FALSE
center = TRUE	$(x-\bar{x})/s_x$	$x-ar{x}$
center = FALSE	$x / \left(\frac{1}{n-1} \sum_{i=1}^{n} x_i^2\right)^{1/2}$	x

```
> x \leftarrow c(1.2, 3.4, 4.2, 12.4, 13.4, 17.3, 18.1)
> n < -7
> (x - mean(x))/sd(x)
 \begin{smallmatrix} 1 \end{smallmatrix} \end{bmatrix} - 1.2639104 - 0.9479328 - 0.8330319 \quad 0.3447028 \quad 0.4883290 \quad 1.0484712 \quad 1.1633721 
> as.numeric(scale(x, center = TRUE, scale = TRUE))
 \begin{smallmatrix} [1] \end{smallmatrix} - 1.2639104 - 0.9479328 - 0.8330319 & 0.3447028 & 0.4883290 & 1.0484712 & 1.1633721 \\ \end{smallmatrix} 
> x - mean(x)
[1] -8.8 -6.6 -5.8 2.4 3.4 7.3 8.1
> as.numeric(scale(x, center = TRUE, scale = FALSE))
[1] -8.8 -6.6 -5.8 2.4 3.4 7.3 8.1
> x/sqrt(sum(x^2)/(n-1))
[1] 0.09337932 0.26457475 0.32682763 0.96491968 1.04273578 1.34621858 1.40847146
> as.numeric(scale(x, center = FALSE, scale = TRUE))
[1] \quad 0.09337932 \quad 0.26457475 \quad 0.32682763 \quad 0.96491968 \quad 1.04273578 \quad 1.34621858 \quad 1.40847146 \quad 1.04273578 \quad 1.40847146 \quad 1.04847146 \quad 1.04847146
> x \leftarrow c(1.2, 3.4, 4.2, 12.4, 13.4, 17.3, 18.1)
> as.numeric(scale(x, center = FALSE, scale = FALSE))
[1] 1.2 3.4 4.2 12.4 13.4 17.3 18.1
> x < -c(1.2, 4.5, 6.7, 7.8, 9.8)
> n <- 5
> (x - mean(x))/sd(x)
[1] -1.4562179 -0.4550681 0.2123651 0.5460817 1.1528392
> as.numeric(scale(x, center = TRUE, scale = TRUE))
```

```
[1] -1.4562179 -0.4550681  0.2123651  0.5460817  1.1528392
> x - mean(x)
[1] -4.8 -1.5  0.7  1.8  3.8
> as.numeric(scale(x, center = TRUE, scale = FALSE))
[1] -4.8 -1.5  0.7  1.8  3.8
> x/sqrt(sum(x^2)/(n - 1))
[1] 0.1605504  0.6020639  0.8964063  1.0435775  1.3111615
> as.numeric(scale(x, center = FALSE, scale = TRUE))
[1] 0.1605504  0.6020639  0.8964063  1.0435775  1.3111615
> x <- c(1.2, 4.5, 6.7, 7.8, 9.8)
> as.numeric(scale(x, center = FALSE, scale = FALSE))
[1] 1.2 4.5 6.7 7.8 9.8
```

cum3()

- Package: boot
- Input:
 - a vettore numerico x di dimensione n
 - b vettore numerico y di dimensione n
 - ${\tt c}\;$ vettore numerico z di dimensione n

unbiased = TRUE / FALSE distorsione

- **Description:** momento terzo centrato
- Formula:

$$\frac{n}{(n-1)(n-2)} \sum_{i=1}^{n} (x_i - \bar{x}) (y_i - \bar{y}) (z_i - \bar{z})$$

unbiased = FALSE

$$\frac{1}{n}\sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)\left(y_{i}-\bar{y}\right)\left(z_{i}-\bar{z}\right)$$

```
> x <- c(-3, -2, -1, 0, 1, 2)

> y <- c(1.2, 2.3, 2, 3.1, 3.55, 6.7)

> z <- c(2, 3.45, 2.6, 3.11, 3.5, 6.2)

> n <- 6

> (n/((n - 1) * (n - 2))) * sum((x - mean(x)) * (y - mean(y)) *

+ (z - mean(z)))
```

```
[1] 4.96385

> cum3(a = x, b = y, c = z, unbiased = TRUE)

[1] 4.96385

> x <- c(-3, -2, -1, 0, 1, 2)
> y <- c(1.2, 2.3, 2, 3.1, 3.55, 6.7)
> z <- c(2, 3.45, 2.6, 3.11, 3.5, 6.2)
> n <- 6
> (1/n) * sum((x - mean(x)) * (y - mean(y)) * (z - mean(z)))

[1] 2.757694

> cum3(a = x, b = y, c = z, unbiased = FALSE)

[1] 2.757694
```

emm()

• Package: actuar

• Input:

 \mathbf{x} vettore numerico di dimensione n order il valore k dell'ordine

- **Description:** momento non centrato di ordine k
- Formula:

$$\frac{1}{n} \sum_{i=1}^{n} x_i^k$$

```
> x <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> k <- 3
> mean(x^3)

[1] 534.2372

> emm(x, order = 3)

[1] 534.2372

> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> n <- 5
> k <- 4
> mean(x^4)

[1] 1745.677

> emm(x, order = 4)
```

3.16 Connessione e dipendenza in media

eta()

• Package: labstatR

• Input:

- y vettore numerico di dimensione n
- f fattore a k livelli di dimensione n
- **Description:** rapporto di correlazione $\eta^2_{u|f}$
- Formula:

$$\eta_{y|f}^2 = \frac{\sum_{j=1}^k (\bar{y}_j - \bar{y})^2 n_j}{\sum_{i=1}^n (\bar{y}_i - \bar{y})^2}$$

```
> y <- c(1, 1.2, 2.1, 3.4, 5.4, 5.6, 7.2, 3.2, 3, 1, 2.3)
> f <- factor(c("a", "b", "c", "b", "a", "c", "a", "b", "b", "c",</pre>
     "a"))
> f
[1] abcbacabbca
Levels: a b c
> k <- 3
> n <- 11
> table(f)
a b c
4 4 3
> enne <- tapply(y, f, FUN = length)</pre>
> enne
a b c
4 4 3
> ymedio <- tapply(y, f, FUN = mean)</pre>
> sum((ymedio - mean(y))^2 * enne)/sum((y - mean(y))^2)
[1] 0.08657807
> eta(f, y)
[1] 0.08657807
> y <- c(1.2, 3.4, 55.6, 5.1, 7.8, 8.4, 8.7, 9.8)
> f <- factor(c("a", "b", "b", "b", "b", "a", "a", "b"))</pre>
[1] abbbbaab
Levels: a b
> k <- 2
> n < - 8
> table(f)
```

```
f
a b
3 5

> enne <- tapply(y, f, FUN = length)
> enne

a b
3 5

> ymedio <- tapply(y, f, FUN = mean)
> sum((ymedio - mean(y))^2 * enne)/sum((y - mean(y))^2)

[1] 0.0900426

> eta(f, y)

[1] 0.0900426
```

Gini()

- Package: ineq
- Input:
 - x vettore numerico di dimensione n
- Description: rapporto di concentrazione di Gini
- Formula:

$$\frac{n-1}{n}\,G$$

dove
$$G = \frac{2}{n-1} \sum_{i=1}^{n-1} (p_i - q_i) = 1 - \frac{2}{n-1} \sum_{i=1}^{n-1} q_i = \frac{1}{n(n-1)\bar{x}} \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} (x_{(j)} - x_{(i)})$$

```
> x <- c(1.2, 3.4, 55.6, 5.1, 7.8, 8.4, 8.7, 9.8)
> x <- sort(x)
> x

[1] 1.2 3.4 5.1 7.8 8.4 8.7 9.8 55.6

> n <- 8
> q <- cumsum(x[1:(n - 1)])/sum(x)
> G <- 2/(n - 1) * sum((1:(n - 1))/n - q)
> G

[1] 0.606

> R <- (n - 1)/n * G
> R

[1] 0.53025

> Gini(x)

[1] 0.53025
```

gini()

• Package: labstatR

• Input:

 $\label{eq:posterior} \mbox{y vettore numerico di dimensione } n \\ \mbox{plot = FALSE}$

- Description: indici di concentrazione
- Output:
 - G indice di Gini
 - R rapporto di concentrazione di Gini
 - P proporzioni
 - Q somme cumulate
- Formula:

$$G = \frac{2}{n-1} \sum_{i=1}^{n-1} (p_i - q_i) = 1 - \frac{2}{n-1} \sum_{i=1}^{n-1} q_i = \frac{1}{n (n-1) \bar{y}} \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} (y_{(j)} - y_{(i)})$$

$$\operatorname{dove} \quad p_i = i/n \quad \forall i = 1, 2, \dots, n$$

$$q_i = \sum_{j=1}^{i} y_{(j)} / \sum_{j=1}^{n} y_j \quad \forall i = 1, 2, \dots, n$$

$$\frac{n-1}{n} G$$

$$0, p_i \quad \forall i = 1, 2, \dots, n$$

$$0, q_i \quad \forall i = 1, 2, \dots, n$$

```
> y <- c(1, 1, 1, 4, 4, 5, 7, 10)
> y <- sort(y)
> y
[1] 1 1 1 4 4 5 7 10
> n <- 8
> q <- cumsum(y[1:(n - 1)])/sum(y)
> G <- 2/(n - 1) * sum((1:(n - 1))/n - q)
[1] 0.4545455
> gini(y, plot = FALSE)$G
[1] 0.4545455
> R <- (n - 1)/n * G
[1] 0.3977273
> gini(y, plot = FALSE)$R
[1] 0.3977273
> P <- c(0, (1:n)/n)
> P
[1] 0.000 0.125 0.250 0.375 0.500 0.625 0.750 0.875 1.000
> gini(y, plot = FALSE)$P
[1] 0.000 0.125 0.250 0.375 0.500 0.625 0.750 0.875 1.000
> Q <- c(0, cumsum(y)/sum(y))
> 0
[1] 0.00000000 0.03030303 0.06060606 0.09090909 0.21212121 0.33333333 0.48484848
[8] 0.69696970 1.00000000
> gini(y, plot = FALSE)$Q
[1] 0.00000000 0.03030303 0.06060606 0.09090909 0.21212121 0.33333333 0.48484848
[8] 0.69696970 1.00000000
> y < -c(1.2, 3.4, 55.6, 5.1, 7.8, 8.4, 8.7, 9.8)
> y <- sort(y)
> y
[1] 1.2 3.4 5.1 7.8 8.4 8.7 9.8 55.6
> n <- 8
> q <- cumsum(y[1:(n - 1)])/sum(y)
> G <- 2/(n - 1) * sum((1:(n - 1))/n - q)
> G
[1] 0.606
```

```
> gini(y, plot = FALSE)$G

[1] 0.606

> R <- (n - 1)/n * G
> R

[1] 0.53025

> gini(y, plot = FALSE)$R

[1] 0.53025

> P <- c(0, (1:n)/n)
> P

[1] 0.000 0.125 0.250 0.375 0.500 0.625 0.750 0.875 1.000

> gini(y, plot = FALSE)$P

[1] 0.000 0.125 0.250 0.375 0.500 0.625 0.750 0.875 1.000

> Q <- c(0, cumsum(y)/sum(y))
> Q

[1] 0.000 0.012 0.046 0.097 0.175 0.259 0.346 0.444 1.000

> gini(y, plot = FALSE)$Q

[1] 0.000 0.012 0.046 0.097 0.175 0.259 0.346 0.444 1.000
```

RS()

- Package: ineq
- Input:
 - imes vettore numerico di dimensione n
- **Description:** coefficiente di disuguaglianza di *Ricci Schutz*
- Formula:

$$\frac{1}{2 n \bar{x}} \sum_{i=1}^{n} |x_i - \bar{x}|$$

```
> x <- c(1, 1.2, 3.4, 0.8)
> mean(abs(x - mean(x)))/(2 * mean(x))
[1] 0.28125
> RS(x)
[1] 0.28125
> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> mean(abs(x - mean(x)))/(2 * mean(x))
[1] 0.1417790
> RS(x)
[1] 0.1417790
```

chi2()

- Package: labstatR
- Input:
 - f fattore a k livelli
 - g fattore a h livelli
- **Description:** quadrato dell'indice di connessione $\tilde{\chi}^2$ di *Cramer*
- Formula:

$$\tilde{\chi}^{2} = \frac{\chi^{2}}{\chi_{\max}^{2}} = \frac{\sum_{i=1}^{k} \sum_{j=1}^{h} \frac{(n_{ij} - \hat{n}_{ij})^{2}}{\hat{n}_{ij}}}{n_{..} \min(k-1, h-1)} = \frac{\sum_{i=1}^{h} \sum_{j=1}^{k} \frac{n_{ij}^{2}}{\hat{n}_{ij}} - n_{..}}{n_{..} \min(k-1, h-1)} = \frac{\sum_{i=1}^{k} \sum_{j=1}^{h} \frac{n_{ij}^{2}}{n_{i.} n_{.j}} - 1}{\min(k-1, h-1)}$$

$$\text{dove} \qquad \hat{n}_{ij} = \frac{n_{i.} n_{.j}}{n_{..}} \quad \forall i = 1, 2, \dots, k \quad \forall j = 1, 2, \dots, h$$

$$n_{..} = \sum_{i=1}^{k} \sum_{j=1}^{h} n_{ij} = \sum_{i=1}^{k} \sum_{j=1}^{h} \hat{n}_{ij}$$

```
> f <- factor(c("a", "b", "c", "b", "a", "c", "a", "b", "b", "c",</pre>
     "a"))
> f
[1] abcbacabbca
Levels: a b c
> k <- nlevels(f)</pre>
> g <- factor(c("0", "P", "W", "P", "P", "0", "0", "W", "W", "P",
    "P"))
Levels: O P W
> h <- nlevels(g)</pre>
> table(f, g)
 g
f OPW
 a 2 2 0
 b 0 2 2
 c 1 1 1
> n.. <- sum(table(f, g))
> chi2(f, g)
[1] 0.1777778
> f <- factor(c("a", "b", "b", "b", "b", "a", "a", "b"))</pre>
[1] abbbbaab
Levels: a b
> k <- nlevels(f)</pre>
> g <- factor(c("A", "B", "B", "B", "A", "A", "B", "A"))
```

```
[1] A B B B A A B A
Levels: A B

> h <- nlevels(g)
> table(f, g)

g
f A B
    a 2 1
    b 2 3

> n.. <- sum(table(f, g))
> chi2(f, g)

[1] 0.06666667
```

E()

• Package: labstatR

• Input:

f fattore a k livelli di dimensione n

- Description: indice di eterogeneità di Gini
- Formula:

$$E = \frac{k}{k-1} \left(1 - \frac{1}{n^2} \sum_{i=1}^{k} n_i^2 \right)$$

```
> f <- factor(c("a", "b", "c", "b", "a", "c", "a", "b", "b", "c",</pre>
     "a"))
 [1] abcbacabbca
Levels: a b c
> k <- 3
> n <- 11
> enne <- table(f)</pre>
> enne
f
a b c
4 4 3
> E <- k/(k - 1) * (1 - 1/n^2 * sum(enne^2))
> E
[1] 0.9917355
> E(f)
[1] 0.9917355
> f <- factor(c("A", "B", "B", "B", "A", "A", "B", "A"))</pre>
> f
```

```
[1] A B B B A A B A
Levels: A B

> k <- 2
> n <- 8
> enne <- table(f)
> enne

f
A B
4 4

> E <- k/(k - 1) * (1 - 1/n^2 * sum(enne^2))
> E

[1] 1

> E(g)
[1] 1
```

3.17 Sintesi di dati

summary()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• **Description:** statistiche descrittive

• Output:

Min. minimo
1st Qu. primo quartile
Median mediana
Mean media aritmetica
3rd Qu. terzo quartile
Max. massimo

• Formula:

Min. $x_{(1)}$ 1st Qu. $Q_{0.25}(x)$ Median $Q_{0.5}(x)$ Mean \bar{x} 3rd Qu. $Q_{0.75}(x)$ Max. $x_{(n)}$

```
> x < -c(1, 2.3, 5, 6.7, 8)
> min(x)
[1] 1
> quantile(x, probs = 0.25)
25%
2.3
> median(x)
[1] 5
> mean(x)
[1] 4.6
> quantile(x, probs = 0.75)
75%
6.7
> \max(x)
[1] 8
> summary(x)
  Min. 1st Qu. Median
                         Mean 3rd Qu.
   1.0 2.3 5.0
                          4.6
                                  6.7
                                          8.0
> x <- c(1.2, 2.2, 3, 15.6, 71.6, 2.2, 1.2)
> min(x)
[1] 1.2
> quantile(x, probs = 0.25)
25%
1.7
> median(x)
[1] 2.2
> mean(x)
[1] 13.85714
> quantile(x, probs = 0.75)
75%
9.3
> max(x)
[1] 71.6
> summary(x)
  Min. 1st Qu. Median
                         Mean 3rd Qu.
                                         Max.
                                9.30 71.60
        1.70
                  2.20
                        13.86
   1.20
```

• **Note:** Calcola i quartili con la funzione quantile().

fivenum()

```
• Package: stats
```

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• Description: cinque numeri di Tukey

• Formula:

```
x_{(1)}
0.5 \left(x_{\lfloor \lfloor (n+3)/2 \rfloor/2 \rfloor} + x_{\lceil \lfloor (n+3)/2 \rfloor/2 \rceil}\right)
Q_{0.5}(x)
0.5 \left(x_{\lfloor n+1-\lfloor (n+3)/2 \rfloor/2 \rfloor} + x_{\lceil n+1-\lfloor (n+3)/2 \rfloor/2 \rceil}\right)
x_{(n)}
```

```
> x < -c(1, 2.3, 5, 6.7, 8)
> n < -5
> min(x)
[1] 1
> 0.5 * (x[floor(floor((n + 3)/2)/2)] + x[ceiling(floor((n + 3)/2)/2)])
[1] 2.3
> median(x)
[1] 5
> 0.5 * (x[n + 1 - floor(floor((n + 3)/2)/2)] + x[n + 1 - ceiling(floor((n + 3)/2)/2)]
+ 3)/2)/2)])
[1] 6.7
> \max(x)
[1] 8
> fivenum(x)
[1] 1.0 2.3 5.0 6.7 8.0
> x < -c(1.2, 1.2, 2.2, 2.2, 3, 15.6, 71.6)
> n <- 7
> \min(x)
[1] 1.2
> 0.5 * (x[floor(floor((n + 3)/2)/2)] + x[ceiling(floor((n + 3)/2)/2)])
[1] 1.7
```

```
> median(x)
[1] 2.2
> 0.5 * (x[n + 1 - floor(floor((n + 3)/2)/2)] + x[n + 1 - ceiling(floor((n + 3)/2)/2)]
+ 3)/2)/2)])
[1] 9.3
> \max(x)
[1] 71.6
> fivenum(x)
[1] 1.2 1.7 2.2 9.3 71.6
> x < -c(1.44, 5.76, 21.16, 60.84)
> n <- 4
> min(x)
[1] 1.44
> 0.5 * (x[floor(floor((n + 3)/2)/2)] + x[ceiling(floor((n + 3)/2)/2)])
[1] 3.6
> median(x)
[1] 13.46
> 0.5 * (x[n + 1 - floor(floor((n + 3)/2)/2)] + x[n + 1 - ceiling(floor((n + 3)/2)/2)]
     3)/2)/2)])
[1] 41
> max(x)
[1] 60.84
> fivenum(x)
[1] 1.44 3.60 13.46 41.00 60.84
```

basicStats()

• Package: fBasics

• Input:

 \times vettore numerico di dimensione n

ci livello di confidenza $1-\alpha$

• **Description:** statistiche riassuntive

• Output:

nobs dimensione campionaria

NAs numero di valori NA oppure NaN

Minimum minimo

Maximum massimo

1. Quartile primo quartile

3. Quartile terzo quartile

Mean media aritmetica

Median mediana

Sum somma

SE Mean errore standard della media

LCL Mean estremo inferiore dell'intervallo di confidenza a livello $1-\alpha$ per la media incognita

UCL Mean estremo superiore dell'intervallo di confidenza a livello $1-\alpha$ per la media incognita

Variance varianza campionaria

Stdev deviazione standard

Skewness asimmetria campionaria

Kurtosis kurtosi campionaria

• Formula:

nobs

n

NAs

NA + # NaN

Minimum

 $x_{(1)}$

Maximum

 $x_{(m)}$

1. Quartile

 $Q_{0.25}(x)$

3. Quartile

 $Q_{0.75}(x)$

Mean

 \bar{x}

Median

 $Q_{0.5}(x)$

Sum

 $\sum_{i=1}^{m} x_i$

SE Mean

 s_x / \sqrt{m}

LCL Mean

 $\bar{x} - t_{1-\alpha/2, m-1} s_x / \sqrt{m}$

UCL Mean $\bar{x} + t_{1-\alpha/2,m-1} \, s_x / \sqrt{m}$ Variance s_x^2 Stdev s_x Skewness $\frac{1}{m} \sum_{i=1}^m \left(\frac{x_i - \bar{x}}{s_x}\right)^3$ Kurtosis $\frac{1}{m} \sum_{i=1}^m \left(\frac{x_i - \bar{x}}{s_x}\right)^4 - 3$

```
> x < -c(1, 2.3, 5, 6.7, 8)
> length(x)
[1] 5
> sum(is.na(x))
[1] 0
> min(x)
[1] 1
> \max(x)
[1] 8
> quantile(x, probs = 0.25)
25%
2.3
> quantile(x, probs = 0.75)
75%
6.7
> mean(x)
[1] 4.6
> median(x)
[1] 5
> sum(x)
[1] 23
> sd(x)/sqrt(length(x))
[1] 1.311106
```

```
> alpha <- 0.05
> mean(x) - qt(1 - alpha/2, length(x) - 1) * sd(x)/sqrt(length(x))
[1] 0.959785
> mean(x) + qt(1 - alpha/2, length(x) - 1) \star sd(x)/sqrt(length(x))
[1] 8.240215
> var(x)
[1] 8.595
> sd(x)
[1] 2.931723
> mean((x - mean(x))^3/sd(x)^3)
[1] -0.08091067
> mean((x - mean(x))^4/sd(x)^4) - 3
[1] -2.055005
> basicStats(x, ci = 0.95)
            round.ans..digits...6.
nobs
                           5.000000
                           0.000000
NAs
Minimum
                           1.000000
Maximum
                           8.000000
1. Quartile
                           2.300000
3. Quartile
                           6.700000
Mean
                           4.600000
Median
                           5.000000
Sum
                          23.000000
SE Mean
                          1.311106
LCL Mean
                           0.959785
UCL Mean
                           8.240215
Variance
                           8.595000
Stdev
                           2.931723
Skewness
                          -0.113076
Kurtosis
                           1.476555
> x < -c(1.3, NaN, 2, 3.4, 3.4, 5.7, NA, 3.8, 0, 9, 0)
> n <- 11
> m <- 11 - sum(is.na(x))
> m
[1] 9
> sum(is.na(x))
[1] 2
> min(x, na.rm = TRUE)
[1] 0
```

```
> max(x, na.rm = TRUE)
[1] 9
> quantile(x, probs = 0.25, na.rm = TRUE)
25%
1.3
> quantile(x, probs = 0.75, na.rm = TRUE)
75%
3.8
> mean(x, na.rm = TRUE)
[1] 3.177778
> median(x, na.rm = TRUE)
[1] 3.4
> sum(x, na.rm = TRUE)
[1] 28.6
> sd(x, na.rm = TRUE)/sqrt(m)
[1] 0.9563788
> alpha <- 0.05</pre>
> mean(x, na.rm = TRUE) - qt(1 - alpha/2, m - 1) * sd(x, na.rm = TRUE)/sqrt(m)
[1] 0.9723642
> mean(x, na.rm = TRUE) + qt(1 - alpha/2, m - 1) * sd(x, na.rm = TRUE)/sqrt(m)
[1] 5.383191
> var(x, na.rm = TRUE)
[1] 8.231944
> sd(x, na.rm = TRUE)
[1] 2.869137
> mean((x - mean(x, na.rm = TRUE))^3/sd(x, na.rm = TRUE)^3, na.rm = TRUE)
[1] 0.6644322
> mean((x - mean(x, na.rm = TRUE))^4/sd(x, na.rm = TRUE)^4, na.rm = TRUE) -
[1] -0.6913239
> basicStats(x, ci = 0.95)
```

```
round.ans..digits...6.
                          11.000000
nobs
NAs
                           2.000000
                           0.000000
Minimum
Maximum
                           9.000000
1. Quartile
                          1.300000
3. Quartile
                          3.800000
                          3.177778
Mean
                           3.400000
Median
                          28.600000
Sum
SE Mean
                           0.956379
LCL Mean
                           0.972364
UCL Mean
                           5.383191
Variance
                           8.231944
Stdev
                           2.869137
Skewness
                           0.792829
                           2.921918
Kurtosis
```

- Note 1: Calcola le statistiche descrittive utilizzando x privato dei valori NA e NaN.
- Note 2: Vale la relazione m = n (#NA + #NaN).
- Note 3: Calcola i quartili con la funzione quantile ().

stat.desc()

- Package: pastecs
- Input:
 - \times vettore numerico di dimensione n
 - p livello di confidenza $1-\alpha$
- Description: statistiche descrittive
- Output:

```
nbr.val dimensione campionaria m di x privato dei valori NA e NaN nbr.null numero di valori nulli nbr.na numero di valori NA e NaN min minimo max massimo range campo di variazione sum somma median mediana mean media aritmetica SE.mean errore standard della media CI.mean.p ampiezza dell'intervallo di confidenza a livello 1-\alpha var varianza campionaria std.dev deviazione standard coef.var coefficiente di variazione campionario
```

• Formula:

```
nbr.val $m$ nbr.null $\#\,0$ nbr.na $\#\,\mathrm{NA}\ +\ \#\,\mathrm{NaN}\ ]
```

min $x_{(1)}$ max $x_{(m)}$ range $x_{(m)} - x_{(1)}$ sum $\sum_{i=1}^{m} x_i$ median $Q_{0.5}(x)$ mean \bar{x} SE.mean $s_x \, / \, \sqrt{m}$ CI.mean.p $t_{1-\alpha/2,\,m-1}\,s_x/\sqrt{m}$ var s_x^2 std.dev s_x coef.var s_x / \bar{x}

• Examples:

[1] 23

```
> x <- c(1, 2.3, 5, 6.7, 8)
> length(x)

[1] 5
> sum(x == 0)

[1] 0
> sum(is.na(x))

[1] 0
> min(x)

[1] 1
> max(x)

[1] 8
> max(x) - min(x)

[1] 7
> sum(x)
```

```
> median(x)
[1] 5
> mean(x)
[1] 4.6
> sd(x)/sqrt(length(x))
[1] 1.311106
> alpha <- 0.05</pre>
> qt(1 - alpha/2, df = length(x) - 1) * sd(x)/sqrt(length(x))
[1] 3.640215
> var(x)
[1] 8.595
> sd(x)
[1] 2.931723
> sd(x)/mean(x)
[1] 0.6373311
> stat.desc(x, p = 0.95)
    nbr.val
                nbr.null
                               nbr.na
                                                min
                                                             max
                                                                         range
   5.0000000
               0.0000000
                           0.0000000
                                         1.0000000 8.0000000
                                                                     7.0000000
                                           SE.mean CI.mean.0.95
                   median
                                  mean
                                                                           var
         sum
               5.0000000
  23.0000000
                             4.6000000
                                         1.3111064
                                                       3.6402150
                                                                     8.5950000
    std.dev
                coef.var
   2.9317230
                0.6373311
> x < -c(1.3, NaN, 2, 3.4, 3.4, 5.7, NA, 3.8, 0, 9, 0)
> n <- 11
> m <- 11 - sum(is.na(x))
> m
[1] 9
> sum(x == 0, na.rm = TRUE)
[1] 2
> sum(is.na(x))
[1] 2
> min(x, na.rm = TRUE)
[1] 0
> max(x, na.rm = TRUE)
```

```
[1] 9
> max(x, na.rm = TRUE) - min(x, na.rm = TRUE)
[1] 9
> sum(x, na.rm = TRUE)
[1] 28.6
> median(x, na.rm = TRUE)
[1] 3.4
> mean(x, na.rm = TRUE)
[1] 3.177778
> sd(x, na.rm = TRUE)/sqrt(m)
[1] 0.9563788
> alpha <- 0.05
> qt(1 - alpha/2, df = m - 1) * sd(x, na.rm = TRUE)/sqrt(m)
[1] 2.205414
> var(x, na.rm = TRUE)
[1] 8.231944
> sd(x, na.rm = TRUE)
[1] 2.869137
> sd(x, na.rm = TRUE)/mean(x, na.rm = TRUE)
[1] 0.9028751
> stat.desc(x, p = 0.95)
     nbr.val
                nbr.null
                               nbr.na
                                                min
                                                              max
                                                                         range
                             nbr.na min
2.0000000 0.0000000
   9.000000
                2.0000000
                                                       9.0000000
                                                                     9.0000000
                                            SE.mean CI.mean.0.95
                   median
                                  mean
                                                                           var
         sum
  28.6000000
                3.4000000
                             3.1777778 0.9563788
                                                       2.2054136
                                                                     8.2319444
     std.dev
                coef.var
   2.8691365
                0.9028751
```

- Note 1: Calcola le statistiche descrittive utilizzando x privato dei valori NA e NaN.
- Note 2: Vale la relazione m = n (#NA + #NaN).
- Note 3: Calcola i quartili con la funzione quantile ().

boxplot.stats()

```
• Package: grDevices
```

• Input:

```
{\bf x} vettore numerico di dimensione n coef valore c positivo
```

- **Description:** statistiche necessarie per il boxplot
- Output:

```
stats cinque numeri di Tukey n dimensione del vettore x conf intervallo di notch out valori di x esterni all'intervallo tra i baffi
```

• Formula:

stats
$$x_{(1)} \qquad Q_{0.5}\left(x_i\left|_{x_i \leq Q_{0.5}(x)}\right.\right) \qquad Q_{0.5}(x) \qquad Q_{0.5}\left(x_i\left|_{x_i \geq Q_{0.5}(x)}\right.\right) \qquad x_{(n)}$$
 n
$$n$$
 conf
$$Q_{0.5}(x) \mp 1.58 \cdot IQR(x) / \sqrt{n}$$
 out
$$x_i < Q_{0.25}(x) - c \cdot IQR(x) \quad OR \quad x_i > Q_{0.75}(x) + c \cdot IQR(x)$$

```
> x < -c(1.2, 1.2, 2.2, 3, 15.6, 71.6)
> c <- 1.4
> fn <- fivenum(x)</pre>
> fn
[1] 1.2 1.2 2.6 15.6 71.6
> boxplot.stats(x, coef = 1.4)$stats
[1] 1.2 1.2 2.6 15.6 15.6
> n <- 6
> boxplot.stats(x, coef = 1.4)n
[1] 6
> median(x) + c(-1, 1) * 1.58 * (fn[4] - fn[2])/sqrt(n)
[1] -6.688465 11.888465
> boxplot.stats(x, coef = 1.4)$conf
[1] -6.688465 11.888465
> x[x < fn[2] - c * (fn[4] - fn[2]) | x > fn[4] + c * (fn[4] -
   fn[2])]
[1] 71.6
> boxplot.stats(x, coef = 1.4)$out
```

```
[1] 71.6
> x < -c(1, 2.3, 5, 6.7, 8)
> c <- 2.6
> fn <- fivenum(x)</pre>
[1] 1.0 2.3 5.0 6.7 8.0
> boxplot.stats(x, coef = 2.6)$stats
[1] 1.0 2.3 5.0 6.7 8.0
> n <- 5
> boxplot.stats(x, coef = 2.6)$n
[1] 5
> median(x) + c(-1, 1) * 1.58 * (fn[4] - fn[2])/sqrt(n)
[1] 1.890971 8.109029
> boxplot.stats(x, coef = 2.6)$conf
[1] 1.890971 8.109029
> x[x < fn[2] - c * (fn[4] - fn[2]) | x > fn[4] + c * (fn[4] -
   fn[2])]
numeric(0)
> boxplot.stats(x, coef = 2.6)$out
numeric(0)
```

.18 Distribuzione di frequenza

• Note: Calcola i quartili con la funzione fivenum().

tabulate()

- Package: base
- Input:

bin vettore di valori naturali di dimensione n

- **Description:** distribuzione di frequenza per i valori naturali $1, 2, \ldots, \max(\texttt{bin})$
- Examples:

```
> tabulate(bin = c(2, 3, 5))
[1] 0 1 1 0 1
> tabulate(bin = c(2, 3, 3, 5))
[1] 0 1 2 0 1
> tabulate(bin = c(-2, 0, 2, 3, 3, 5))
[1] 0 1 2 0 1
```

table()

```
• Package: base
```

• Input:

 \times vettore alfanumerico di dimensione n

• **Description:** distribuzione di frequenza

```
> x <- c("a", "a", "b", "c", "a", "c")
> table(x)
X
a b c
3 1 2
> table(x)/length(x)
                b
0.5000000 0.1666667 0.3333333
> f <- factor(c("a", "b", "c", "b", "a", "c", "a", "b", "b", "c",</pre>
+ "a"))
> f
[1] abcbacabbca
Levels: a b c
> g <- factor(c("A", "S", "A", "S", "S", "S", "A", "S", "A",
+ "A"))
> g
[1] A S A S S S A S S A A
Levels: A S
> table(f, g)
  g
f AS
 a 3 1
 b 0 4
 c 2 1
> x \leftarrow c(1, 2, 3, 2, 1, 3, 1, 1, 2, 3)
> table(x)
1 2 3
4 3 3
```

unique()

• Package: base

```
• Input:
       {\bf x}~ vettore alfanumerico di dimensione n
  • Description: supporto (valori distinti di x)
  • Examples:
   > x <- c("a", "a", "b", "c", "a", "c")
   > unique(x)
   [1] "a" "b" "c"
   > x <- c(1, 2, 3, 2, 1, 3, 1, 1, 2, 3)
   > unique(x)
   [1] 1 2 3
   > x < -c(12, -3, 7, 12, 4, -3, 12, 7, -3)
   > x[!duplicated(x)]
   [1] 12 -3 7 4
   > unique(x)
   [1] 12 -3 7 4
duplicated()
  • Package: base
  • Input:
       \times vettore numerico di dimensione n
  • Description: segnalazione di valori duplicati
  • Examples:
   > x < -c(1, 2, 1, 3, 2, 2, 4)
   > duplicated(x)
   [1] FALSE FALSE TRUE FALSE TRUE TRUE FALSE
   > x < -c(1, 2, 1, 2, 1, 2)
   > duplicated(x)
   [1] FALSE FALSE TRUE TRUE TRUE TRUE
   > x < -c(12, -3, 7, 12, 4, -3, 12, 7, -3)
   > unique(x[duplicated(x)])
   [1] 12 -3 7
```

3.19 Istogramma

hist()

```
• Package: graphics
```

• Input:

```
x vettore numerico di dimensione n breaks estremi delle classi di ampiezza b_i right = TRUE / FALSE classi chiuse a destra \left(a_{(i)},\,a_{(i+1)}\right] oppure a sinistra \left[a_{(i)},\,a_{(i+1)}\right) include.lowest = TRUE / FALSE estremo incluso plot = FALSE
```

• Description: istogramma

• Output:

breaks estremi delle classi counts frequenze assolute density densità di frequenza mids punti centrali delle classi

• Formula:

breaks
$$a_{(i)} \quad \forall i=1,2,\ldots,m$$
 counts
$$n_i \quad \forall i=1,2,\ldots,m-1$$
 density
$$\frac{n_i}{n\,b_i} \quad \forall i=1,2,\ldots,m-1$$
 mids
$$\frac{a_{(i)}+a_{(i+1)}}{2} \quad \forall i=1,2,\ldots,m-1$$

```
> x \leftarrow c(51.1, 52.3, 66.7, 77.1, 77.15, 77.17)
> n < -6
> m < -4
> a1 <- 50
> a2 <- 65
> a3 < -70
> a4 < -85
> a <- c(a1, a2, a3, a4)
> b1 <- 65 - 50
> b2 <- 70 - 65
> b3 <- 85 - 70
> b <- c(b1, b2, b3)
[1] 15 5 15
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE) $breaks
[1] 50 65 70 85
> count <- numeric(m - 1)</pre>
> count[1] <- sum(x >= a1 & x < a2)
> count[2] <- sum(x >= a2 & x < a3)
> count[3] <- sum(x >= a3 & x < a4)
> count
```

```
[1] 2 1 3
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE)$counts
[1] 2 1 3
> count/(n * b)
[1] 0.02222222 0.03333333 0.03333333
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE)$density
[1] 0.02222222 0.03333333 0.03333333
> (a[-m] + a[-1])/2
[1] 57.5 67.5 77.5
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE) $mids
[1] 57.5 67.5 77.5
> x \leftarrow c(1, 1.2, 2.2, 2.3, 3, 5, 6.7, 8, 15.6)
> n < -9
> m < -5
> a1 <- 0
> a2 < -5
> a3 <- 10
> a4 <- 15
> a5 <- 20
> a <- c(a1, a2, a3, a4, a5)
[1] 0 5 10 15 20
> b1 <- a2 - a1
> b2 <- a3 - a2
> b3 <- a4 - a3
> b4 <- a5 - a4
> b <- c(b1, b2, b3, b4)
> b
[1] 5 5 5 5
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE) $breaks
[1] 0 5 10 15 20
> count <- numeric(m - 1)</pre>
> count[1] <- sum(x >= a1 & x < a2)
> count[2] <- sum(x >= a2 & x < a3)
> count[3] <- sum(x >= a3 & x < a4)
> count[4] <- sum(x >= a4 & x < a5)
> count
[1] 5 3 0 1
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE)$counts
```

```
[1] 5 3 0 1
> count/(n * b)

[1] 0.11111111 0.06666667 0.00000000 0.02222222
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE)$density
[1] 0.11111111 0.06666667 0.00000000 0.02222222
> (a[-m] + a[-1])/2
[1] 2.5 7.5 12.5 17.5
> hist(x, breaks = a, right = FALSE, include.lowest = FALSE, plot = FALSE)$mids
[1] 2.5 7.5 12.5 17.5
```

n.bins()

- Package: car
- Input:

```
{\tt x} vettore numerico di dimensione n rule = "freedman.diaconis" / "sturges" / "scott" / "simple" algoritmo
```

- Description: algoritmo di calcolo per il numero di classi di un istogramma
- Formula:

$$n_c = \left\lceil \frac{x_{(n)} - x_{(1)}}{2 \, IQR(x) \, n^{-1/3}} \right\rceil$$

$$\boxed{\text{rule = "sturges"}}$$

$$n_c = \left\lceil \log_2(n) + 1 \right\rceil$$

$$\boxed{\text{rule = "scott"}}$$

$$n_c = \left\lceil \frac{x_{(n)} - x_{(1)}}{3.5 \, s_x \, n^{-1/3}} \right\rceil$$

$$\boxed{\text{rule = "simple"}}$$

$$n_c = \left\{ \begin{array}{l} \left\lfloor 2 \sqrt{n} \right\rfloor & \text{se } n \leq 100 \\ \left\lfloor 10 \, \log_{10}(n) \right\rfloor & \text{se } n > 100 \end{array} \right.$$

```
> x <- c(2.3, 1, 5, 6.7, 8)
> x <- sort(x)
> x

[1] 1.0 2.3 5.0 6.7 8.0
```

```
> n <- 5
> nc <- ceiling((x[n] - x[1])/(2 * IQR(x) * n^(-1/3)))
[1] 2
> n.bins(x, rule = "freedman.diaconis")
[1] 2
> x < -c(2.3, 1, 5, 6.7, 8)
> n <- 5
> nc <- ceiling(log2(n) + 1)
> nc
[1] 4
> n.bins(x, rule = "sturges")
[1] 4
> x < -c(2.3, 1, 5, 6.7, 8)
> x <- sort(x)
> x
[1] 1.0 2.3 5.0 6.7 8.0
> n <- 5
> sx <- sd(x)
> nc <- ceiling((x[n] - x[1])/(3.5 * sx * n^(-1/3)))
> nc
[1] 2
> n.bins(x, rule = "scott")
[1] 2
> x < -c(2.3, 1, 5, 6.7, 8)
> n <- 5
> nc <- floor(2 * sqrt(n))
> nc
[1] 4
> n.bins(x, rule = "simple")
[1] 4
```

 \bullet $\,$ Note: Calcola i quartili con la funzione ${\tt quantile}\,()$.

nclass.FD()

- Package: grDevices
- Input:
 - \times vettore numerico di dimensione n
- Description: numero di classi di un istogramma secondo Freedman Diaconis
- Formula:

$$n_c = \left[\frac{x_{(n)} - x_{(1)}}{2 \, IQR(x) \, n^{-1/3}} \right]$$

• Examples:

```
> x < -c(2.3, 1, 5, 6.7, 8)
> x <- sort(x)
> X
[1] 1.0 2.3 5.0 6.7 8.0
> n < -5
> nc <- ceiling((x[n] - x[1])/(2 * IQR(x) * n^(-1/3)))
> nc
[1] 2
> nclass.FD(x)
[1] 2
> x < -c(3.4, 5.52, 6.4, 7.56, 8.7, 8.6, 5.4, 5.5)
> x <- sort(x)
> x < -c(3.4, 5.4, 5.5, 5.52, 6.4, 7.56, 8.6, 8.7)
> nc <- ceiling((x[n] - x[1])/(2 * IQR(x) * n^{(-1/3)}))
[1] 3
> nclass.FD(x)
[1] 3
```

• **Note:** Calcola i quartili con la funzione quantile().

nclass.Sturges()

- Package: grDevices
- Input:
 - \times vettore numerico di dimensione n
- Description: numero di classi di un istogramma secondo Sturges
- Formula:

$$n_c = \lceil \log_2(n) + 1 \rceil$$

```
> x <- c(1, 2.3, 5, 6.7, 8)
> n <- 5
> nc <- ceiling(log2(n) + 1)
> nc

[1] 4
> nclass.Sturges(x)

[1] 4
> x <- c(3.4, 5.4, 5.5, 5.52, 6.4, 7.56, 8.6, 8.7)
> n <- 8
> nc <- ceiling(log2(n) + 1)
> nc

[1] 4
> nclass.Sturges(x)

[1] 4
```

nclass.scott()

• Package: grDevices

• Input:

imes vettore numerico di dimensione n

- **Description:** numero di classi di un istogramma secondo *Scott*
- Formula:

$$n_c = \left\lceil \frac{x_{(n)} - x_{(1)}}{3.5 \, s_x \, n^{-1/3}} \right\rceil$$

```
> x <- c(2.3, 1, 5, 6.7, 8)
> x <- sort(x)
> x

[1] 1.0 2.3 5.0 6.7 8.0

> n <- 5
> sx <- sd(x)
> nc <- ceiling((x[n] - x[1])/(3.5 * sx * n^(-1/3)))
> nc

[1] 2
> nclass.scott(x)

[1] 2
> x <- c(3.4, 5.4, 5.5, 5.52, 6.4, 7.56, 8.6, 8.7)
> x <- sort(x)
> x

[1] 3.40 5.40 5.50 5.52 6.40 7.56 8.60 8.70
```

```
> n <- 8
> sx <- sd(x)
> nc <- ceiling((x[n] - x[1])/(3.5 * sx * n^(-1/3)))
> nc

[1] 2
> nclass.scott(x)
```

3.20 Variabili casuali discrete

Bernoulli

$$p_X(x) = p^x (1-p)^{1-x} \quad x = 0, 1, \quad 0
$$\mu_X = p$$

$$\sigma_X^2 = p (1-p)$$$$

Binomiale

$$p_X(x) = {m \choose x} p^x (1-p)^{m-x} \quad x = 0, 1, 2, ..., m, \quad m \in \mathbb{N} / \{0\}, \quad 0
$$\mu_X = m p$$

$$\sigma_X^2 = m p (1-p)$$$$

Binomiale Negativa

$$p_X(x) = \binom{r+x-1}{x} p^r (1-p)^x = \binom{r+x-1}{r-1} p^r (1-p)^x \quad x \in \mathbb{N}, \quad r \in \mathbb{N} \setminus \{0\}, \quad 0
$$\mu_X = r (1-p) / p$$

$$\sigma_X^2 = r (1-p) / p^2$$$$

Geometrica

$$p_X(x) = p(1-p)^x$$
 $x \in \mathbb{N}$, $0
 $\mu_X = (1-p)/p$
 $\sigma_X^2 = (1-p)/p^2$$

Geometrica 2

$$p_X(x) \,=\, p\,(1-p)^{x-1} \quad x \in \mathbb{N}\backslash\{0\}, \quad 0
$$\mu_X \,=\, 1\,/\,p$$

$$\sigma_X^2 \,=\, (1-p)\,/\,p^2$$$$

Ipergeometrica

$$p_X(x) = {M \choose x} {N-M \choose k-x} / {N \choose k}$$
$$x = 0, 1, 2, \dots, k$$
$$N \in \mathbb{N} \setminus \{0\}$$
$$k = 1, 2, \dots, N$$

$$M = 0, 1, 2, ..., N - 1$$

 $\mu_X = k (M/N)$
 $\sigma_X^2 = k (M/N) (1 - M/N) (N - k) / (N - 1)$

Multinomiale

$$\begin{split} p_{X_1,\,X_2,\,\dots,\,X_k}(x_1,\,x_2,\,\dots,\,x_k) &= \frac{m\,!}{x_1\,!\,x_2\,!\cdots x_k\,!}\,\prod_{i=1}^k\,p_i^{x_i}\\ x_i &= 0,\,1,\,2,\,\dots,\,m\quad\forall i=1,\,2,\,\dots,\,k\\ 0 &< p_i &< 1\quad\forall i=1,\,2,\,\dots,\,k\\ \sum_{i=1}^k\,x_i &= m\\ \sum_{i=1}^k\,p_i &= 1\\ \mu_{X_i} &= m\,p_i\quad\forall i=1,\,2,\,\dots,\,k\\ \sigma_{X_i}^2 &= m\,p_i\,(1-p_i)\quad\forall i=1,\,2,\,\dots,\,k\\ \sigma_{X_iX_j} &= -m\,p_i\,p_j\quad\forall i\neq j=1,\,2,\,\dots,\,k \end{split}$$

Poisson

$$p_X(x) \,=\, \lambda^x\,e^{-\lambda}\,/\,x\,! \quad x\in\mathbb{N}, \quad \lambda>0$$

$$\mu_X \,=\, \lambda$$

$$\sigma_X^2 \,=\, \lambda$$

Tavola argomenti comandi R

Variabile Casuale	Suffisso	Parametri	Package
Bernoulli	binom	size, prob	stats
Binomiale	binom	size, prob	stats
Binomiale Negativa	nbinom	size, prob	stats
Geometrica	geom	prob	stats
Geometrica 2	geomet	р	distributions
Ipergeometrica	hyper	m, n, k	stats
Multinomiale	multinom	size, prob	stats
Poisson	pois	lambda	stats

Tavola esempi comandi R

Variabile Casuale	Oggetto	Comando in R
Bernoulli	Densità	dbinom(x=x, size=1, prob=p)
	Ripartizione	pbinom(q=x, size=1, prob=p)
	Quantile	$qbinom(p=\alpha, size=1, prob=p)$
	Random	rbinom(n,size=1,prob=p)
Binomiale	Densità	dbinom(x=x, size=m, prob=p)
	Ripartizione	pbinom(q=x,size=m,prob=p)
	Quantile	$qbinom(p=\alpha,size=m,prob=p)$
	Random	rbinom(n, size=m, prob=p)
Binomiale Negativa	Densità	dnbinom(x=x,size=r,prob=p)
	Ripartizione	pnbinom(q=x,size=r,prob=p)
	Quantile	$qnbinom(p=\alpha,size=r,prob=p)$
	Random	rnbinom(n, size=r, prob=p)
Geometrica	Densità	dgeom(x=x,prob=p)
	Ripartizione	pgeom(q=x,prob=p)
	Quantile	$qgeom(p=\alpha,prob=p)$
	Random	rgeom(n,prob=p)

Geometrica 2	Densità	geometpdf ($p=p$, $x=x$)	
	Ripartizione	geometcdf(p= p , x= x)	
Ipergeometrica	Densità	dhyper($x=x$, $m=M$, $n=N-M$, $k=k$)	
	Ripartizione	phyper($q=x$, $m=M$, $n=N-M$, $k=k$)	
	Quantile	qhyper(p= α , m= M , n= $N-M$, k= k)	
	Random	rhyper(nn, m= M , n= $N-M$, k= k)	
Multinomiale	Densità	dmultinom(x=c(x_1,\ldots,x_k),prob=c(p_1,\ldots,p_k))	
	Random	rmultinom(n,size= m ,prob=c(p_1,\ldots,p_k))	
Poisson	Densità	dpois($x=x$, lambda= λ)	
	Ripartizione	ppois($q=x$, lambda= λ)	
	Quantile	qpois($p=\alpha$, lambda= λ)	
	Random	rpois(n,lambda= λ)	

3.21 Variabili casuali continue

Beta

$$\begin{split} f_X(x) &= \frac{\Gamma(\theta + \lambda)}{\Gamma(\theta) \, \Gamma(\lambda)} \, x^{\theta - 1} \, (1 - x)^{\lambda - 1} \quad 0 < x < 1, \quad \theta > 0, \quad \lambda > 0 \\ \mu_X &= \theta \, / \, (\theta + \lambda) \\ \sigma_X^2 &= \theta \, \lambda \, / \, \left[(\theta + \lambda + 1) \, (\theta + \lambda)^2 \right] \end{split}$$

Beta NC

$$\frac{\chi_{\theta}^2(\delta)}{\chi_{\theta}^2(\delta) + \chi_{\lambda}^2} \quad 0 < x < 1, \quad \theta > 0, \quad \lambda > 0, \quad \delta > 0$$

Burr

$$\begin{split} f_X(x) &= \frac{\theta \, \mu \, (x \, / \, \lambda)^{\theta}}{x \, \left(1 + (x \, / \, \lambda)^{\theta}\right)^{\mu + 1}} \quad x > 0, \quad \theta > 0, \quad \mu > 0, \quad \lambda > 0 \\ \\ \mu_X &= \lambda \, \Gamma(1 - 1 \, / \, \theta) \, \Gamma(1 \, / \, \theta + \mu) \, / \, \Gamma(\mu) \\ \\ \sigma_X^2 &= \left[\Gamma(\mu) \, \Gamma(1 - 2 \, / \, \theta) \, \Gamma(2 \, / \, \theta + \mu) - \Gamma^2(1 - 1 \, / \, \theta) \, \Gamma(1 \, / \, \theta + \mu) \right] \, \lambda^2 \, / \, \Gamma^2(\mu) \quad \text{per } \theta > 2 \end{split}$$

Cauchy

$$f_X(x) = (\pi \lambda)^{-1} \left[1 + ((x - \theta) / \lambda)^2 \right]^{-1} \quad x \in \mathbb{R}, \quad \theta \in \mathbb{R}, \quad \lambda > 0$$

$$\mu_X = \mathbb{A}$$

$$\sigma_X^2 = \mathbb{A}$$

Chi - Quadrato

$$f_X(x) = \frac{2^{-k/2}}{\Gamma(k/2)} x^{(k-2)/2} e^{-x/2} \quad x > 0, \quad k > 0$$

$$\mu_X = k$$

$$\sigma_X^2 = 2 k$$

Chi - Quadrato NC

$$\begin{split} f_X(x) &= \exp\left(-(x+\delta) \, / \, 2\right) \, \sum_{i=0}^{\infty} \, \frac{(\delta \, / \, 2)^i \, x^{k \, / \, 2+i-1}}{2^{k \, / \, 2+i} \, \Gamma(k \, / \, 2+i) \, i!} \quad x > 0, \quad k > 0, \quad \delta > 0 \\ \mu_X &= k + \delta \\ \sigma_X^2 &= 2 \, (k+2 \, \delta) \end{split}$$

Dirichlet

$$f_{X_1,X_2,\dots,X_k}(x_1,x_2,\dots,x_k) = \frac{\Gamma(\alpha_1+\alpha_2+\dots+\alpha_k)}{\Gamma(\alpha_1)\Gamma(\alpha_2)\dots\Gamma(\alpha_k)} \prod_{i=1}^k x_i^{\alpha_i-1}$$

$$x_i > 0 \quad \forall i = 1, 2, \dots, k$$

$$\alpha_i > 0 \quad \forall i = 1, 2, \dots, k$$

$$\sum_{i=1}^k x_i = 1$$

$$\sum_{i=1}^k \alpha_i = \alpha$$

$$\mu_{X_i} = \frac{\alpha_i}{\alpha} \quad \forall i = 1, 2, \dots, k$$

$$\sigma_{X_i}^2 = \frac{\alpha_i(\alpha-\alpha_i)}{\alpha^2(\alpha+1)} \quad \forall i = 1, 2, \dots, k$$

$$\sigma_{X_iX_j} = -\frac{\alpha_i \alpha_j}{\alpha^2(\alpha+1)} \quad \forall i \neq j = 1, 2, \dots, k$$

Esponenziale

$$f_X(x) = \lambda e^{-\lambda x}$$
 $x > 0$, $\lambda > 0$
 $\mu_X = 1/\lambda$
 $\sigma_X^2 = 1/\lambda^2$

Fisher

$$\begin{split} f_X(x) &= \frac{\Gamma((n_1+n_2)/2)}{\Gamma(n_1/2)\,\Gamma(n_2/2)} \, \left(\frac{n_1}{n_2}\right)^{n_1/2} \, x^{(n_1-2)/2} \, \left(1+\frac{n_1}{n_2}\,x\right)^{-(n_1+n_2)/2} & x,\, n_1,\, n_2>0 \\ \mu_X &= \frac{n_2}{n_2-2} \quad \text{per} \, n_2>2 \\ \sigma_X^2 &= \frac{2\,n_2^2\,(n_1+n_2-2)}{n_1\,(n_2-2)^2\,(n_2-4)} \quad \text{per} \, n_2>4 \end{split}$$

Fisher NC

$$\begin{split} f_X(x) &= \tfrac{n_1^{n_1/2} n_2^{n_2/2}}{\exp{(\delta/2)}} \tfrac{x^{n_1/2-1}}{(n_1\,x + n_2)^{(n_1+n_2)/2}} \sum_{i=0}^\infty \tfrac{(\delta/2)^i}{i!} \tfrac{\Gamma(n_1/2 + n_2/2 + i)}{\Gamma(n_1/2 + i) \Gamma(n_2/2)} \left(\tfrac{n_1\,x}{n_1\,x + n_2} \right)^i \quad x, \, n_1, \, n_2, \, \delta > 0 \\ \mu_X &= \tfrac{n_2\,(n_1 + \delta)}{n_1\,(n_2 - 2)} \quad \text{per} \, n_2 > 2 \\ \sigma_X^2 &= 2 \left(\tfrac{n_2}{n_1} \right)^2 \tfrac{(n_1 + \delta)^2 + (n_1 + 2\,\delta)\,(n_2 - 2)}{(n_2 - 2)^2\,(n_2 - 4)} \quad \text{per} \, n_2 > 4 \end{split}$$

Friedman

$$x > 0$$
 $r \in \mathbb{N} / \{0, 1\}, N \in \mathbb{N} / \{0, 1\}$

Gamma

$$f_X(x) = \frac{\lambda^{\theta}}{\Gamma(\theta)} x^{\theta-1} e^{-\lambda x} \quad x > 0, \quad \theta > 0, \quad \lambda > 0$$

$$\mu_X = \theta / \lambda$$

$$\sigma_X^2 = \theta / \lambda^2$$

Gamma 2

$$\begin{split} f_X(x) &= \frac{1}{\lambda^{\theta} \, \Gamma(\theta)} \, x^{\theta-1} \, e^{-x \, / \, \lambda} \quad x > 0, \quad \theta > 0, \quad \lambda > 0 \\ \mu_X &= \theta \, \lambda \\ \sigma_X^2 &= \theta \, \lambda^2 \end{split}$$

Gamma inversa

$$\begin{split} f_X(x) &= \frac{\lambda^{\theta}}{\Gamma(\theta)} \, x^{-\,(\theta+1)} \, e^{-\lambda\,/\,x} \quad x > 0, \quad \theta > 0, \quad \lambda > 0 \\ \mu_X &= \lambda\,/\,(\theta-1) \quad \text{per } \theta > 1 \\ \sigma_X^2 &= \lambda^2\,/\,[(\theta-1)^2\,(\theta-2)] \quad \text{per } \theta > 2 \end{split}$$

Gamma inversa 2

$$\begin{split} f_X(x) &= \frac{1}{\lambda^\theta \, \Gamma(\theta)} \, x^{-\, (\theta+1)} \, e^{-1\, /\, (\lambda\, x)} \quad x > 0, \quad \theta > 0, \quad \lambda > 0 \\ \mu_X &= 1\, /\, [\lambda\, (\theta-1)] \quad \text{per } \theta > 1 \\ \sigma_X^2 &= 1\, /\, [\lambda^2\, (\theta-1)^2\, (\theta-2)] \quad \text{per } \theta > 2 \end{split}$$

Laplace

$$f_X(x) = \frac{1}{2} \lambda^{-1} \exp\left(-\frac{|x-\theta|}{\lambda}\right) \quad x \in \mathbb{R}, \quad \theta \in \mathbb{R}, \quad \lambda > 0$$

$$\mu_X = \theta$$

$$\sigma_X^2 = 2 \lambda^2$$

Logistica

$$f_X(x) = \lambda^{-1} \exp((x - \theta) / \lambda) (1 + \exp((x - \theta) / \lambda))^{-2} \quad x \in \mathbb{R}, \quad \theta \in \mathbb{R}, \quad \lambda > 0$$

$$\mu_X = \theta$$

$$\sigma_X^2 = (\pi \lambda)^2 / 3$$

LogLogistica

$$\begin{split} f_X(x) &= \frac{\theta \left(x/\lambda \right)^{\theta}}{x \left(1 + \left(x/\lambda \right)^{\theta} \right)^2} \quad x > 0, \quad \theta > 0, \quad \lambda > 0 \\ \\ \mu_X &= \lambda \, \Gamma(1 - 1/\theta) \, \Gamma(1/\theta + 1) \\ \\ \sigma_X^2 &= \left[\Gamma(1 - 2/\theta) \, \Gamma(2/\theta + 1) - \Gamma^2(1 - 1/\theta) \, \Gamma(1/\theta + 1) \right] \, \lambda^2 \quad \text{per } \theta > 2 \end{split}$$

LogNormale

$$\begin{split} f_X(x) &= \left(\sigma \, x \, \sqrt{2 \, \pi}\right)^{-1} \, \exp\left(-(\log(x) - \mu)^2 \, / \, (2 \, \sigma^2)\right) \quad x > 0, \quad \mu \in \mathbb{R}, \, \sigma > 0 \\ \mu_X &= \exp\left(\mu + \sigma^2 \, / \, 2\right) \\ \sigma_X^2 &= \exp\left(2 \, \mu + \sigma^2\right) \, \left(\exp\left(\sigma^2\right) - 1\right) \end{split}$$

Mann - Whitney

$$\begin{split} &0 \leq x \leq n_x \, n_y, \quad n_x \in \mathbb{N} \, / \, \{0\}, \quad n_y \in \mathbb{N} \, / \, \{0\} \\ &\mu_X \, = \, n_x \, n_y \, / \, 2 \\ &\sigma_X^2 \, = \, n_x \, n_y \, (n_x + n_y + 1) \, / \, 12 \end{split}$$

Normale

$$\begin{split} f_X(x) &= \left(2\,\pi\,\sigma^2\right)^{-1\,/\,2}\,\exp\left(-(x-\mu)^2\,/\,(2\,\sigma^2)\right) \quad x\in\mathbb{R},\quad \mu\in\mathbb{R},\quad \sigma>0\\ \mu_X &= \mu\\ \sigma_X^2 &= \sigma^2 \end{split}$$

Normale doppia

$$f_{X_1,X_2}(x_1,x_2) \,=\, \frac{1}{2\,\pi\,\sqrt{\sigma_{11}\,\sigma_{22}\,(1-\rho^2)}}\,\exp\left(-\frac{1}{2\,(1-\rho^2)}\left[\left(\frac{x_1-\mu_1}{\sqrt{\sigma_{11}}}\right)^2 - 2\,\rho\,\frac{x_1-\mu_1}{\sqrt{\sigma_{11}}}\,\frac{x_2-\mu_2}{\sqrt{\sigma_{22}}} + \left(\frac{x_2-\mu_2}{\sqrt{\sigma_{22}}}\right)^2\right]\right)$$

$$x_i \in \mathbb{R} \quad \forall i = 1, 2$$

$$\mu_i \in \mathbb{R} \quad \forall i = 1, 2$$

$$\rho \, = \, \sigma_{12} \, / \, \sqrt{\sigma_{11} \, \sigma_{22}} \, = \, \sigma_{21} \, / \, \sqrt{\sigma_{11} \, \sigma_{22}} \in (0, \, 1)$$

$$V_2 = \left(egin{array}{cc} \sigma_{11} & \sigma_{12} \ \sigma_{21} & \sigma_{22} \end{array}
ight) \quad ext{definita positiva}$$

$$\sigma_{ii} > 0 \quad \forall i = 1, 2$$

$$\mu_{X_i} = \mu_i \quad \forall i = 1, 2$$

$$\sigma_{X_i}^2 = \sigma_{ii} \quad \forall i = 1, 2$$

$$\sigma_{X_1X_2} = \sigma_{12} = \sigma_{21}$$

Normale multipla

$$f_{X_1,X_2,\dots,X_k}(x_1,x_2,\dots,x_k) = \frac{1}{(2\pi)^{k/2}\sqrt{\det(V_k)}} \exp\left(-\frac{1}{2}(x_1-\mu_1,x_2-\mu_2,\dots,x_k-\mu_k)^T V_k^{-1}(x_1-\mu_1,x_2-\mu_2,\dots,x_k-\mu_k)\right)$$

$$x_i \in \mathbb{R} \quad \forall i = 1, 2, \dots, k$$

$$\mu_i \in \mathbb{R} \quad \forall i = 1, 2, \dots, k$$

$$V_k = \left(egin{array}{cccc} \sigma_{11} & \sigma_{12} & \dots & \sigma_{1k} \\ \sigma_{21} & \sigma_{22} & \dots & \sigma_{2k} \\ \vdots & \vdots & \vdots & \vdots \\ \sigma_{k1} & \sigma_{k2} & \dots & \sigma_{kk} \end{array}
ight)$$
 definita positiva

$$\sigma_{ii} > 0 \quad \forall i = 1, 2, \dots, k$$

$$\mu_{X_i} = \mu_i \quad \forall i = 1, 2, \dots, k$$

$$\sigma_{X_i}^2 = \sigma_{ii} \quad \forall i = 1, 2, \dots, k$$

$$\sigma_{X_i X_i} = \sigma_{ij} = \sigma_{ji} \quad \forall i \neq j = 1, 2, \dots, k$$

Pareto

$$f_X(x) = \frac{\theta \lambda^{\theta}}{x^{\theta+1}} \quad x > \lambda, \quad \theta > 0, \quad \lambda > 0$$

$$\mu_X = \theta \lambda / (\lambda - 1)$$

$$\sigma_X^2 \,=\, \theta\,\lambda^2\,/\,\left((\theta-2)\,(\theta-1)^2\right) \quad \text{per }\theta>2$$

Student

$$f_X(x) = \frac{\Gamma((k+1)/2)}{\Gamma(k/2)} (k\pi)^{-1/2} (1+x^2/k)^{-(k+1)/2} \quad x \in \mathbb{R}, \quad k > 0$$

$$\mu_X = 0 \quad \text{per } k > 1$$

$$\sigma_X^2 = k / (k - 2)$$
 per $k > 2$

Student NC

$$f_X(x) = \frac{k^{k/2} \exp\left(-\delta^2/2\right)}{\sqrt{\pi} \Gamma(n/2) (k+x^2)^{(k+1)/2}} \sum_{i=0}^{\infty} \frac{\Gamma((k+i+1)/2) \delta^i}{i!} \left(\frac{2x^2}{k+x^2}\right)^{i/2} \quad x \in \mathbb{R}, \quad k > 0, \quad \delta \in \mathbb{R}$$

$$\mu_X = \sqrt{k/2} \delta \Gamma ((k-1)/2) / \Gamma (k/2)$$
 per $k > 1$

$$\sigma_X^2 = k(1+\delta^2)/(k-2) - \delta(k/2) (\Gamma((k-1)/2)/\Gamma(k/2))^2$$
 per $k > 2$

Tukey

$$x > 0$$
, $n \in \mathbb{N} / \{0, 1, 2\}$, $p \in \mathbb{N} / \{0, 1\}$

Uniforme

$$f_X(x) = 1/(b-a)$$
 $a < x < b$, $a \in \mathbb{R}$, $b \in \mathbb{R}$, $a < b$
 $\mu_X = (a+b)/2$
 $\sigma_X^2 = (b-a)^2/12$

Wald

$$f_X(x) = (\lambda/(2\pi x^3))^{1/2} \exp\left(-\lambda(x-\theta)^2/(2\theta^2 x)\right) \quad x>0, \quad \theta>0, \quad \lambda>0$$

$$\mu_X = \theta$$

$$\sigma_X^2 = \theta^3/\lambda$$

Weibull

$$f_X(x) = (\theta/\lambda) (x/\lambda)^{\theta-1} \exp\left(-(x/\lambda)^{\theta}\right) \quad x > 0, \quad \theta > 0, \quad \lambda > 0$$

$$\mu_X = \lambda \Gamma((\theta+1)/\theta)$$

$$\sigma_X^2 = \lambda^2 \left[\Gamma((\theta+2)/\theta) - \Gamma^2((\theta+1)/\theta)\right]$$

Wilcoxon signed rank

$$0 \le x \le n(n+1)/2, \quad n \in \mathbb{N}/\{0\}$$

 $\mu_X = n(n+1)/4$
 $\sigma_X^2 = n(n+1)(2n+1)/24$

Tavola argomenti comandi R

Variabile Casuale	Suffisso	Parametri	Package
Beta	beta	shape1, shape2	stats
Beta NC	beta	shape1, shape2, ncp	stats
Burr	burr	shape1, shape2, scale, rate	actuar
Cauchy	cauchy	location, scale	stats
Chi - Quadrato	chisq	df	stats
Chi - Quadrato NC	chisq	df, ncp	stats
Dirichlet	dirichlet	alpha	MCMCpack
Esponenziale	exp	rate	stats
Fisher	f	df1, df2	stats
Fisher NC	f	df1, df2, ncp	stats
Friedman	Friedman	r, N	SuppDists
Gamma	gamma	shape, scale, rate	stats
Gamma 2	gamma	shape, scale, rate	stats
Gamma inversa	invgamma	shape, scale	MCMCpack
Gamma inversa 2	invgamma	shape, scale	MCMCpack
Laplace	laplace	m, s	formularioR
Logistica	logis	location, scale	stats
LogLogistica	llogis	shape, scale, rate	actuar
LogNormale	lnorm	meanlog, sdlog	stats
Mann - Whitney	wilcox	m, n	stats
Normale	norm	mean, sd	stats
Normale doppia	mvnorm	mean, sigma	mvtnorm
Normale multipla	mvnorm	mean, sigma	mvtnorm
Pareto	pareto1	shape, min	actuar

Student	t	df	stats
Student NC	t	df, ncp	stats
Tukey	tukey	nmeans, df	stats
Uniforme	unif	min, max	stats
Wald	invGauss	nu, lambda	SuppDists
Weibull	weibull	shape, scale	stats
Wilcoxon signed rank	signrank	n	stats

Tavola esempi comandi R

Variabile Casuale	Oggetto	Comando in R	
Beta	Densità	dbeta(x= x , shape1= θ , shape2= λ)	
	Ripartizione	pbeta (q= x , shape1= θ , shape2= λ)	
	Quantile	qbeta (p= α , shape1= θ , shape2= λ)	
	Random	rbeta(n, shape1= θ , shape2= λ)	
Beta NC	Densità	dbeta(x= x , shape1= θ , shape2= λ , ncp= δ)	
	Ripartizione	pbeta (q= x , shape1= θ , shape2= λ , ncp= δ)	
	Quantile	qbeta (p= α , shape1= θ , shape2= λ , ncp= δ)	
	Random	rbeta(n, shape1= θ , shape2= λ , ncp= δ)	
Burr	Densità	dburr(x= x , shape1= μ , shape2= θ , scale= λ)	
		dburr(x= x , shape1= μ , shape2= θ , rate=1/ λ)	
	Ripartizione	pburr(q= x , shape1= μ , shape2= θ , scale= λ)	
		pburr(q= x , shape1= μ , shape2= θ , rate= $1/\lambda$)	
	Quantile	qburr(p= α , shape1= μ , shape2= θ , scale= λ)	
		qburr(p= α , shape1= μ , shape2= θ , rate= $1/\lambda$)	
	Random	rburr(n, shape1= μ , shape2= θ , scale= λ)	
		rburr(n,shape1= μ ,shape2= θ ,rate= $1/\lambda$)	
Cauchy	Densità	dcauchy (x= x , location= θ , scale= λ)	
	Ripartizione	pcauchy (q= x , location= θ , scale= λ)	
	Quantile	qcauchy (p= α , location= θ , scale= λ)	
	Random	rcauchy (n, location= θ , scale= λ)	
Chi - Quadrato	Densità	dchisq(x=x,df=k)	
	Ripartizione	pchisq(q=x,df=k)	
	Quantile	$qchisq(p=\alpha, df=k)$	
	Random	rchisq(n,df=k)	
Chi - Quadrato NC	Densità	$dchisq(x=x, df=k, ncp=\delta)$	
	Ripartizione	pchisq (q= x , df= k , ncp= δ)	
	Quantile Random	qchisq(p= α , df= k , ncp= δ)	
Dirichlet	Densità	rchisq(n,df= k ,ncp= δ) ddirichlet(x=c($x_1,,x_k$),alpha=c($\alpha_1,,\alpha_k$))	
Dirichlet	Random		
Esponenziale	Densità	rdirichlet (n, alpha=c $(\alpha_1, \ldots, \alpha_k)$) $\text{dexp}(x=x, \text{rate}=\lambda)$	
Esponenziale	Ripartizione	$(\text{dexp}(x=x, \text{rate}=\lambda))$ $(\text{pexp}(q=x, \text{rate}=\lambda))$	
	Quantile	$ \text{qexp}(q-x, \text{rate}-\lambda) $ $ \text{qexp}(p=\alpha, \text{rate}=\lambda) $	
	Random	$rexp(p-\alpha, rate=\lambda)$	
Fisher	Densità	$df(x=x,df1=n_1,df2=n_2)$	
risiici		$pf(q=x,df1=n_1,df2=n_2)$	
	Quantile	$\begin{array}{c} \text{pr}(q x, \text{dif } n_1, \text{dif } n_2) \\ \text{qf}(\text{p=}\alpha, \text{df1=}n_1, \text{df2=}n_2) \end{array}$	
	Random	$\begin{array}{c} \text{qf}(p, \text{df} = n_1, \text{df} = n_2) \\ \text{rf}(n, \text{df} = n_1, \text{df} = n_2) \end{array}$	
Fisher NC	Densità	$df(x=x,df1=n_1,df2=n_2,ncp=\delta)$	
	Ripartizione	$pf(q=x, df1=n_1, df2=n_2, ncp=\delta)$	
	Quantile	$ qf(p=\alpha,df1=n_1,df2=n_2,ncp=\delta) $	
	Random	$rf(n, df1=n_1, df2=n_2, ncp=\delta)$	
Friedman	Densità	dFriedman ($x=x$, $r=r$, $N=N$)	
	Ripartizione	pFriedman $(q=x, r=r, N=N)$	
	Quantile	qFriedman (p= α , r= r , N= N)	
	Random	rFriedman($n, r=r, N=N$)	
Gamma	Densità	dgamma($x=x$, shape= θ , rate= λ)	
		dgamma (x=x, shape= θ , scale= $1/\lambda$)	
	Ripartizione	pgamma ($q=x$, shape= θ , rate= λ)	
	-	pgamma ($q=x$, shape= θ , scale= $1/\lambda$)	
	Quantile	qgamma ($p=\alpha$, shape= θ , rate= λ)	
	1 =		

		qgamma(p= α , shape= θ , scale= $1/\lambda$)
	Random	qgamma($p=\alpha$, snape= θ , scare= $1/\lambda$) rgamma(n , shape= θ , rate= λ)
	Kanuom	rgamma (n, shape= θ , rate= λ) rgamma (n, shape= θ , scale= $1/\lambda$)
Gamma 2	Densità	dgamma (x=x, shape= θ , rate= $1/\lambda$)
Gamma 2	Delisita	dgamma($x=x$, $shape=\theta$, $scale=\lambda$)
	Ripartizione	$\begin{array}{l} \text{dgamma}(x=x, \text{shape}=v, \text{scale}=\lambda) \\ \text{pgamma}(\text{q}=x, \text{shape}=\theta, \text{rate}=1/\lambda) \end{array}$
	Ripartizione	pgamma($q=x$, shape= θ , scale= λ)
	Quantile	$ qqamma(q=x,shape=\theta,rate=1/\lambda) $
	guunin	qqamma (p= α , shape= θ , scale= λ)
	Random	rgamma (n, shape= θ , rate= $1/\lambda$)
		rgamma (n, shape= θ , scale= λ)
Gamma inversa	Densità	dinvgamma (x= x , shape= θ , scale= $1/\lambda$)
	Random	rinvgamma(n, shape= θ , scale= λ)
Gamma inversa 2	Densità	dinvgamma (x= x , shape= θ , scale= λ)
	Random	rinvgamma(n, shape= θ , scale= $1/\lambda$)
Laplace	Densità	dlaplace $(x=x, m=\theta, s=\lambda)$
•	Ripartizione	plaplace $(q=x, m=\theta, s=\lambda)$
	Quantile	qlaplace (p= α , m= θ , s= λ)
	Random	rlaplace($n, m=\theta, s=\lambda$)
Logistica	Densità	dlogis (x= x , location= θ , scale= λ)
-	Ripartizione	plogis (q= x , location= θ , scale= λ)
	Quantile	qlogis (p= α , location= θ , scale= λ)
	Random	rlogis(n,location= θ ,scale= λ)
LogLogistica	Densità	dllogis(x= x , shape= θ , scale= λ)
		dllogis(x= x , shape= $ heta$, rate= $1/\lambda$)
	Ripartizione	pllogis(q= x , shape= θ , scale= λ)
		pllogis(q= x ,shape= $ heta$,rate= $1/\lambda$)
	Quantile	qllogis (p= α , shape= θ , scale= λ)
		qllogis (p= α , shape= θ , rate= $1/\lambda$)
	Random	rllogis (n, shape= θ , scale= λ)
V	D 143	rllogis (n, shape= θ , rate= $1/\lambda$)
LogNormale	Densità	dlnorm(x= x , meanlog= μ , sdlog= σ)
	Ripartizione Quantile	plnorm($q=x$, meanlog= μ , sdlog= σ) qlnorm($p=\alpha$, meanlog= μ , sdlog= σ)
	Random	
Mann - Whitney	Densità	rlnorm(n, meanlog= μ , sdlog= σ) dwilcox(x= x , m= n_x , n= n_y)
Maiiii - Williney	Ripartizione	
	Quantile	qwilcox ($q=x$, $m=n_x$, $n=n_y$) qwilcox ($p=\alpha$, $m=n_x$, $n=n_y$)
	Random	rwilcox (nn, m= n_x , n= n_y)
Normale	Densità	$dnorm(x=x, mean=\mu, sd=\sigma)$
1101111110	Ripartizione	pnorm (q= x , mean= μ , sd= σ)
	Quantile	$q = \alpha$, $m = \alpha = \mu$, $s = \sigma$
	Random	$r_{n,mean} = \mu, sd = \sigma,$
Normale doppia	Densità	dmvnorm(x=c(x_1,x_2),mean=c(μ_1,μ_2),sigma= V_2)
	Ripartizione	pmvnorm(u=c(x_1,x_2), mean=c(μ_1,μ_2), sigma= V_2)
	Random	rmvnorm(n, mean=c(μ_1,μ_2), sigma= V_2)
Normale multipla	Densità	dmvnorm(x=c(x_1,x_2,\ldots,x_k), mean=c(μ_1,μ_2,\ldots,μ_k), sigma= V_k)
_	Ripartizione	pmvnorm(u=c(x_1,x_2,\ldots,x_k), mean=c(μ_1,μ_2,\ldots,μ_k), sigma= V_k)
	Random	rmvnorm(n,mean=c(μ_1,μ_2,\ldots,μ_k),sigma= V_k)
Pareto	Densità	dpareto1(x= x , shape= θ , min= λ)
	Ripartizione	ppareto1(q= x , shape= θ , min= λ)
	Quantile	qpareto1(p= $lpha$, shape= $ heta$, min= λ)
	Random	rparetol(n, shape= θ , min= λ)
Student	Densità	dt(x=x, df=k)
	Ripartizione	pt(q=x,df=k)
	Quantile	$qt(p=\alpha, df=k)$
A. 4	Random	rt (n, df=k)
Student NC	Densità	$dt (x=x, df=k, ncp=\delta)$
	Ripartizione	pt (q= x , df= k , ncp= δ)
	Quantile	qt (p= α , df= k , ncp= δ)
M. 1	Random	rt (n, df= k , ncp= δ)
Tukey	Ripartizione	ptukey(q=x,nmeans=p,df=n)
	Quantile	qtukey(p= α ,nmeans= p ,df= n)

Uniforme	Densità	dunif(x=x, min=a, max=b)		
	Ripartizione	punif $(q=x, min=a, max=b)$		
	Quantile	qunif (p= α , min= a , max= b)		
	Random	runif(n,min=a,max=b)		
Wald	Densità	dinvGauss($x=x$, $nu=\theta$, $lambda=\lambda$)		
	Ripartizione	pinvGauss(q= x , nu= θ , lambda= λ)		
	Quantile	qinvGauss(p= α , nu= θ , lambda= λ)		
	Random	rinvGauss(n,nu= θ ,lambda= λ)		
Weibull	Densità	dweibull($x=x$, shape= θ , scale= λ)		
	Ripartizione	pweibull($q=x$, shape= θ , scale= λ)		
	Quantile	qweibull(p= α , shape= θ , scale= λ)		
	Random	rweibull(n,shape= θ ,scale= λ)		
Wilcoxon signed rank	Densità	dsignrank ($x=x$, $n=n$)		
	Ripartizione	psignrank ($q=x$, $n=n$)		
	Quantile	qsignrank (p= α , n= n)		
	Random	rsignrank(nn, n=n)		

3.22 Logit

logit()

• Package: faraway

• Input:

 $\times\,$ vettore numerico di probabilità di dimensione n

- **Description:** transformazione logit
- Formula:

$$\log\left(\frac{x_i}{1-x_i}\right) \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(0.2, 0.34, 0.54, 0.65, 0.11)
> log(x/(1 - x))

[1] -1.3862944 -0.6632942  0.1603427  0.6190392 -2.0907411

> logit(x)

[1] -1.3862944 -0.6632942  0.1603427  0.6190392 -2.0907411

> x <- c(0.23, 0.45, 0.67, 0.89, 0.11)
> log(x/(1 - x))

[1] -1.2083112 -0.2006707  0.7081851  2.0907411 -2.0907411

> logit(x)
```

ilogit()

• Package: faraway

• Input:

x vettore numerico di dimensione n

• Description: trasformazione logit inversa

• Formula:

$$\frac{e^{x_i}}{1 + e^{x_i}} \, = \, \frac{1}{1 + e^{-x_i}} \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$$

• Examples:

```
> x <- c(1, 2, 3, 5, -6)
> exp(x)/(1 + exp(x))

[1] 0.731058579 0.880797078 0.952574127 0.993307149 0.002472623
> ilogit(x)

[1] 0.731058579 0.880797078 0.952574127 0.993307149 0.002472623
> x <- c(2.3, 4.5, 6.7, 7.8, 12)
> exp(x)/(1 + exp(x))

[1] 0.9088770 0.9890131 0.9987706 0.9995904 0.9999939
> ilogit(x)

[1] 0.9088770 0.9890131 0.9987706 0.9995904 0.9999939
```

inv.logit()

• Package: boot

• Input:

 \times vettore numerico di dimensione n

- Description: trasformazione logit inversa
- Formula:

$$\frac{e^{x_i}}{1 + e^{x_i}} = \frac{1}{1 + e^{-x_i}} \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(1, 2, 3, 5, -6)
> exp(x)/(1 + exp(x))

[1] 0.731058579 0.880797078 0.952574127 0.993307149 0.002472623
> inv.logit(x)

[1] 0.731058579 0.880797078 0.952574127 0.993307149 0.002472623
> x <- c(2.3, 4.5, 6.7, 7.8, 12)
> exp(x)/(1 + exp(x))

[1] 0.9088770 0.9890131 0.9987706 0.9995904 0.9999939
> ilogit(x)

[1] 0.9088770 0.9890131 0.9987706 0.9995904 0.9999939
```

3.23 Serie storiche

length()

• Package: base

• Input:

x vettore numerico di dimensione n

- **Description:** dimensione campionaria
- Formula:

n

• Examples:

```
> x <- c(1.2, 2.3, 4.5, 6.5)
> length(x)

[1] 4

> x <- c(1.2, 3.4, 4.5, 6.4, 4, 3, 4)
> length(x)
[1] 7
```

diff()

• Package: base

• Input:

 ${\tt x}$ vettore numerico di dimensione n lag il valore d del ritardo differences il valore k dell'ordine delle differenze

- Description: differenze in una serie storica
- Formula:

$$(1-B^d)^k x_t \quad \forall t = dk+1, dk+2, \dots, n$$

dove
$$(1 - B^d)^k = \sum_{i=0}^k {k \choose j} (-1)^j B^{jd}$$
 $B^h x_t = x_{t-h}$

```
> x < -c(1, 2, 4, 3, 5, 6, -9)
> n < -7
> d < - 2
> k <- 2
> x[(k * d + 1):n] - 2 * x[(k * d + 1 - d):(n - d)] + x[(k * d + d)]
     1 - k * d): (n - k * d)
[1] -2 2 -15
> diff(x, lag = 2, differences = 2)
[1] -2 2 -15
> x <- c(2, 6, 10, 9, 9, 8, 9, 9, 10, 12)
> n <- 10
> d < - 2
> k <- 3
> x[(k * d + 1):n] - 3 * x[(k * d + 1 - d):(n - d)] + 3 * x[(k * d + 1 - d):(n - d)]
    d + 1 - 2 * d): (n - 2 * d)] - x[(k * d + 1 - k * d): (n - 2 * d)]
    k * d)
[1] 10 6 0 0
> diff(x, lag = 2, differences = 3)
[1] 10 6 0 0
```

diffinv()

- Package: stats
- Input:
 - ${\bf x}~$ vettore numerico di dimensione n
 - lag il valore d del ritardo
 - differences il valore k dell'ordine delle differenze
 - xi valore necessari a ricostruire la serie storica di partenza
- **Description:** operazione inversa del comando diff()
- Examples:

```
[1] -2 2 -15
> diffinv(diff(x, lag = 2, differences = 2), lag = 2, differences = 2,
   xi = c(1, 2, 4, 3))
[1] 1 2 4 3 5 6 -9
> x \leftarrow c(2, 6, 10, 9, 9, 8, 9, 9, 10, 12)
> n <- 10
> d < - 2
> k <- 3
> diff(x, lag = 2, differences = 3)
[1] 10 6 0 0
> diffinv(diff(x, lag = 2, differences = 3), lag = 2, differences = 3,
  xi = c(2, 6, 10, 9, 9, 8))
 [1] 2 6 10 9 9 8 9 9 10 12
```

acf()

• Package: stats

• Input:

```
\times vettore numerico di dimensione n
lag.max il valore d del ritardo
type = "correlation" / "covariance" / "partial" tipo di legame
demean = TRUE / FALSE centratura
plot = FALSE
```

- **Description:** autocovarianza oppure autocorrelazione
- Output:

acf autocovarianza oppure autocorrelazione n.used dimensione campionaria lag il valore d del ritardo

• Formula:

acf

type = "correlation" AND demean = TRUE
$$\hat{\rho}(k) = \frac{\sum_{t=1}^{n-k} (x_t - \bar{x}) (x_{t+k} - \bar{x})}{\sum_{t=1}^{n} (x_t - \bar{x})^2} \quad \forall k = 0, 1, 2, \dots, d$$
 type = "correlation" AND demean = FALSE
$$\hat{\rho}(k) = \frac{\sum_{t=1}^{n-k} x_t x_{t+k}}{\sum_{t=1}^{n} x_t^2} \quad \forall k = 0, 1, 2, \dots, d$$
 type = "covariance" AND demean = TRUE
$$\hat{\gamma}(k) = \frac{1}{n} \sum_{t=1}^{n-k} (x_t - \bar{x}) (x_{t+k} - \bar{x}) \quad \forall k = 0, 1, 2, \dots, d$$
 type = "covariance" AND demean = FALSE

$$\hat{\gamma}(k) = \frac{1}{n} \sum_{t=1}^{n-k} x_t x_{t+k} \quad \forall k = 0, 1, 2, \dots, d$$

n.used

lag

d

n

```
> x <- c(1, 2, 7, 3, 5, 2, 0, 1, 4, 5)
> n <- 10
> d < - 4
> sum((x[1:(n - d)] - mean(x)) * (x[(d + 1):n] - mean(x)))/((n - d))
+ 1) * var(x))
[1] -0.3409091
> acf(x, lag.max = d, type = "correlation", demean = TRUE, plot = FALSE) $acf[d +
+ 1]
[1] -0.3409091
> x < -c(1, 2, 7, 3, 5, 2, 0, 1, 4, 5)
> n < -10
> d < -4
> sum((x[1:(n - d)]) * (x[(d + 1):n]))/(sum(x^2))
[1] 0.3134328
> acf(x, lag.max = d, type = "correlation", demean = FALSE, plot = FALSE) $acf[d +
+ 1]
[1] 0.3134328
> x < -c(1, 2, 7, 3, 5, 2, 0, 1, 4, 5)
> n < -10
> sum((x[1:(n - d)] - mean(x)) * (x[(d + 1):n] - mean(x)))/n
[1] -1.5
> acf(x, lag.max = d, type = "covariance", demean = TRUE, plot = FALSE) $acf[d +
    11
[1] -1.5
```

```
> x <- c(1, 2, 7, 3, 5, 2, 0, 1, 4, 5)
> n <- 10
> d <- 4
> sum((x[1:(n - d)]) * (x[(d + 1):n]))/n

[1] 4.2
> acf(x, lag.max = d, type = "covariance", demean = FALSE, plot = FALSE)$acf[d + 1]

[1] 4.2
```

pacf()

• Package: stats

• Input:

 ${\tt x}$ vettore numerico di dimensione n lag.max il valore d del ritardo demean = TRUE / FALSE centratura plot = FALSE

- **Description:** autocorrelazione parziale
- Output:

acf autocorrelazione parziale ${\tt n.used}$ dimensione campionaria ${\tt lag}$ il valore d del ritardo

• Formula:

acf

$$\hat{\pi}(k) = \frac{\begin{vmatrix} 1 & \hat{\rho}(1) & \hat{\rho}(2) & \dots & \hat{\rho}(1) \\ \hat{\rho}(1) & 1 & \hat{\rho}(1) & \dots & \hat{\rho}(2) \\ \hat{\rho}(2) & \hat{\rho}(1) & 1 & \dots & \hat{\rho}(3) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \hat{\rho}(k-1) & \hat{\rho}(k-2) & \hat{\rho}(k-3) & \dots & \hat{\rho}(k) \end{vmatrix}}{\begin{vmatrix} 1 & \hat{\rho}(1) & \hat{\rho}(2) & \dots & \hat{\rho}(k-1) \\ \hat{\rho}(1) & 1 & \hat{\rho}(1) & \dots & \hat{\rho}(k-2) \\ \hat{\rho}(2) & \hat{\rho}(1) & 1 & \dots & \hat{\rho}(k-3) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \hat{\rho}(k-1) & \hat{\rho}(k-2) & \hat{\rho}(k-3) & \dots & 1 \end{vmatrix}}$$

$$\hat{\rho}(k) = \frac{\sum_{t=1}^{n-k} (x_t - \bar{x}) (x_{t+k} - \bar{x})}{\sum_{t=1}^{n} (x_t - \bar{x})^2} \quad \forall k = 0, 1, 2, \dots, d$$

$$\frac{\text{demean = FALSE}}{\hat{\rho}(k)}$$

$$\hat{\rho}(k) = \frac{\sum_{t=1}^{n-k} x_t x_{t+k}}{\sum_{t=1}^{n} x_t^2} \quad \forall k = 0, 1, 2, \dots, d$$

n.used

n

lag

d

• Examples:

3.24 Valori mancanti

is.na()

• Package: base

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

- Description: rileva la presenza di valori NA e NaN
- Examples:

```
> x <- c(1.3, 1, 2, 3.4, 3.4, 5.7, NA, 3.8)
> is.na(x)

[1] FALSE FALSE FALSE FALSE FALSE FALSE TRUE FALSE

> x <- c(1.3, NaN, 2, 3.4, 3.4, 5.7, NA, 3.8)
> is.na(x)

[1] FALSE TRUE FALSE FALSE FALSE FALSE TRUE FALSE

> x <- c(1, 2, NA, 4, 5.6, NaN, 1.2, 4, 4.4)
> x[!is.na(x)]

[1] 1.0 2.0 4.0 5.6 1.2 4.0 4.4

> x <- c(3, 4, NA, 5)
> mean(x)

[1] NA

> mean(x[!is.na(x)])

[1] 4
```

is.nan()

```
• Package: base
```

• Input:

 ${\bf x}~$ vettore numerico di dimensione n

• Description: rileva la presenza di valori NaN

• Examples:

```
> x <- c(1.3, 1, 2, 3.4, 3.4, 5.7, NA, 3.8)
> is.nan(x)

[1] FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE
> x <- c(1.3, NaN, 2, 3.4, 3.4, 5.7, NA, 3.8)
> is.nan(x)

[1] FALSE TRUE FALSE FALSE FALSE FALSE FALSE FALSE
> x <- c(1, 2, NA, 4, 5.6, NaN, 1.2, 4, 4.4)
> x[!is.nan(x)]

[1] 1.0 2.0 NA 4.0 5.6 1.2 4.0 4.4
```

na.omit()

• Package: stats

• Input:

x vettore numerico di dimensione n

• Description: elimina i valori NA e NaN

```
> x <- c(1.3, 1, 2, 3.4, 3.4, 5.7, NA, 3.8)
> na.omit(x)

[1] 1.3 1.0 2.0 3.4 3.4 5.7 3.8
attr(,"na.action")
[1] 7
attr(,"class")
[1] "omit"

> x <- c(1.3, NaN, 2, 3.4, 3.4, 5.7, NA, 3.8)
> na.omit(x)

[1] 1.3 2.0 3.4 3.4 5.7 3.8
attr(,"na.action")
[1] 2 7
attr(,"class")
[1] "omit"
```

3.25 Miscellaneous

• Package: fUtilities

```
sample()
```

```
    Input:
        x vettore alfanumerico di dimensione n
        size ampiezza campionaria
        replace = TRUE / FALSE estrazione con oppure senza ripetizione
        prob vettore di probabilità
```

- Description: estrazione campionaria
- Examples:

nsize()

• Package: BSDA

• Input:

```
b valore del margine di errore E sigma valore dello scarto quadratico medio \sigma_x p valore della proporzione campionaria p conf.level livello di confidenza 1-\alpha type = "mu" / "pi" media nella popolazione oppure proporzione campionaria
```

- **Description:** dimensione campionaria dato il margine di errore E
- Formula:

$$type = "mu"$$

$$n = \left\lceil \left(z_{1-\alpha/2}\sigma_x\right)/E\right)^2 \right\rceil$$

$$type = "pi"$$

$$n = \left\lceil p\left(1-p\right)\left(z_{1-\alpha/2}/E\right)^2 \right\rceil$$

• Examples:

```
> nsize(b = 0.15, sigma = 0.31, conf.level = 0.95, type = "mu")
The required sample size (n) to estimate the population
mean with a 0.95 confidence interval so that the margin
of error is no more than 0.15 is 17 .

> nsize(b = 0.03, p = 0.77, conf.level = 0.95, type = "pi")
The required sample size (n) to estimate the population
proportion of successes with a 0.95 confidence interval
so that the margin of error is no more than 0.03 is 756 .
```

ic.var()

- Package: labstatR
- Input:
 - x vettore numerico di dimensione n conf.level livello di confidenza $1-\alpha$
- Description: intervallo di confidenza Chi-Quadrato per la varianza incognita
- Formula:

$$\frac{(n-1) s_x^2}{\chi_{1-\alpha/2, n-1}^2} \quad \frac{(n-1) s_x^2}{\chi_{\alpha/2, n-1}^2}$$

```
> x < -c(1.2, 3.4, 4.2, 12.4, 13.4, 17.3, 18.1)
> n < -7
> alpha <- 0.05
> lower <- (n - 1) * var(x)/qchisq(1 - alpha/2, df = n - 1)
> upper <- (n - 1) * var(x)/qchisq(alpha/2, df = n - 1)
> c(lower, upper)
[1] 20.12959 235.06797
> ic.var(x, conf.level = 0.95)
[1] 20.12959 235.06797
> x \leftarrow c(1, 2, 3, 4, 5.6, 7.4, 1.2, 4, 4.4)
> n < -9
> alpha <- 0.05
> lower <- (n - 1) * var(x)/qchisq(1 - alpha/2, df = n - 1)
> upper <- (n - 1) * var(x)/qchisq(alpha/2, df = n - 1)
> c(lower, upper)
[1] 1.986681 15.981587
> ic.var(x, conf.level = 0.95)
[1] 1.986681 15.981587
```

sweep()

```
• Package: base
```

• Input:

- **Description:** operazioni da compiere su ogni riga (colonna) della matrice x
- Examples:

```
> X1 <- c(1.2, 3.4, 5.6)
> X2 <- c(7.5, 6.7, 8.4)
> X3 \leftarrow c(4.3, 3.2, 3.2)
> x \leftarrow cbind(X1, X2, X3)
> mediecolonna <- apply(x, MARGIN = 2, FUN = mean)</pre>
> mediecolonna
                        Х3
      X1
           X2
3.400000 7.533333 3.566667
> sweep(x, MARGIN = 2, STATS = mediecolonna, FUN = "-")
       X1
                   X2
[1,] -2.2 -0.03333333 0.7333333
[2,] 0.0 -0.83333333 -0.3666667
[3,] 2.2 0.86666667 -0.3666667
> X1 < -c(1.2, 3.4, 5.6)
> X2 <- c(7.5, 6.7, 8.4)
> X3 < -c(4.3, 3.2, 3.2)
> x \leftarrow cbind(X1, X2, X3)
> medieriga <- apply(x, MARGIN = 1, FUN = mean)</pre>
> medieriga
[1] 4.333333 4.433333 5.733333
> sweep(x, MARGIN = 1, STATS = medieriga, FUN = "-")
             X1
                     X2
[1,] -3.1333333 3.166667 -0.03333333
[2,] -1.0333333 2.266667 -1.23333333
[3,] -0.1333333 2.666667 -2.53333333
```

set.seed()

• Package: base

• Input:

seed **seme**

- Description: fissa un seme per rendere riproducibili i risultati di un'estrazione
- Examples:

```
> set.seed(seed = 100)
> rnorm(1)
```

```
[1] -0.5021924
> rnorm(1)
[1] 0.1315312
> rnorm(1)
[1] -0.07891709
> rnorm(1)
[1] 0.8867848
> set.seed(seed = 100)
> rnorm(1)
[1] -0.5021924
> rnorm(1)
[1] 0.1315312
```

simple.z.test()

• Package: UsingR

• Input:

```
x vettore numerico di dimensione n sigma valore di \sigma_x conf.level livello di confidenza 1-\alpha
```

- **Description:** intervallo di confidenza per la media incognita a livello $1-\alpha$
- Formula:

$$\bar{x} \mp z_{1-\alpha/2} \, \sigma_x / \sqrt{n}$$

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> xmedio <- mean(x)
> xmedio

[1] 7.018182

> sigmax <- 1.2
> alpha <- 0.05
> n <- 11
> lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> c(lower, upper)

[1] 6.309040 7.727323

> simple.z.test(x, sigma = 1.2, conf.level = 0.95)

[1] 6.309040 7.727323
```

```
> x <- c(1, 2.3, 4.5, 6.7, 8.9)
> xmedio <- mean(x)
> xmedio

[1] 4.68

> sigmax <- 1.45
> alpha <- 0.05
> n <- 5
> lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> c(lower, upper)

[1] 3.409042 5.950958

> simple.z.test(x, sigma = 1.45, conf.level = 0.95)

[1] 3.409042 5.950958
```

median.test()

- Package: formularioR
- Input:
 - ${\bf x}~$ vettore numerico di dimensione n ${\bf m0}~$ valore $Q_{0.5}(x)$ della mediana
- Description: verifica di ipotesi per la mediana
- Formula:

$$2 \min (P(X \le v), P(X \ge v))$$

dove $X \sim Binomiale(n, p_0)$ $v = \#(x_i < Q_{0.5}(x) \ \forall i = 1, 2, ..., n)$

```
> x < -c(1, 2, 8, 12, 12, 17, 25, 52)
> n < - 8
> m0 < 12
> v <- sum(x < 12)
> V
[1] 3
> 2 * min(pbinom(q = v, size = 8, prob = 0.5), 1 - pbinom(q = v -
      1, size = 8, prob = 0.5)
[1] 0.7265625
> median.test(x, m0 = 12)
[1] 0.7265625
> x < -c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> n <- 11
> m0 < -6.6
> v <- sum(x < 6.6)
> v
```

Capitolo 4

Analisi Componenti Principali (ACP)

4.1 ACP con matrice di covarianza di popolazione

Simbologia

- matrice dei dati di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici w_1, w_2, \ldots, w_k : W
- media di colonna della matrice dei dati: $\bar{w}_i \quad \forall j=1,2,\ldots,k$
- matrice dei dati centrata di dimensione $n \times k$: Z
- elemento di riga i e colonna j della matrice dei dati centrata: $z_{ij}=w_{ij}-\bar{w}_j \quad \forall i=1,2,\ldots,n \quad \forall j=1,2,\ldots,k$
- matrice di covarianza di dimensione $k \times k$: $S = \frac{Z^T Z}{n} = \Gamma D \Gamma^T$
- matrice ortogonale degli autovettori di dimensione $k \times k$: Γ
- *j*-esima colonna della matrice Γ : $\Gamma^j \quad \forall j = 1, 2, ..., k$
- matrice diagonale degli autovalori di dimensione $k \times k$: $D = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_k)$
- componente principale j-esima: $x_j = Z \Gamma^j \quad \forall j = 1, 2, ..., k$
- scarto quadratico medio della j-esima componente principale: $\sigma_{x_j}=\sqrt{\lambda_{(k-j+1)}} \quad \forall j=1,2,\ldots,k$
- problema di ottimo vincolato:

$$x_{j} = Z \gamma_{j} \quad \forall j = 1, 2, ..., k$$

$$\sigma_{x_{j}}^{2} = \frac{x_{j}^{T} x_{j}}{n} = \frac{(Z \gamma_{j})^{T} (Z \gamma_{j})}{n} = \gamma_{j}^{T} \frac{Z^{T} Z}{n} \gamma_{j} = \gamma_{j}^{T} S \gamma_{j} \quad \forall j = 1, 2, ..., k$$

$$\max_{\gamma_{i}^{T} \gamma_{j} = 1} \sigma_{x_{j}}^{2} = \max_{\gamma_{i}^{T} \gamma_{j} = 1} \gamma_{j}^{T} S \gamma_{j} = \lambda_{(k-j+1)} \quad \forall j = 1, 2, ..., k$$

princomp()

- Package: stats
- Input:

w matrice dei dati

• Output:

sdev scarto quadratico medio delle componenti principali center media di colonna della matrice W n.obs dimensione campionaria scores componenti principali

• Formula:

sdev

$$\sigma_{x_j} \quad \forall j = 1, 2, \ldots, k$$

```
center
                                                              \bar{w}_i \quad \forall j = 1, 2, \ldots, k
n.obs
                                                                             n
scores
                                                              x_i \quad \forall j = 1, 2, \ldots, k
```

```
• Examples:
 > w1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > w2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > w3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > W <- cbind(w1, w2, w3)
 > W
       w1 w2 w3
 [1,] 1.1 1.2 1.40
 [2,] 2.3 3.4 5.60
 [3,] 4.5 5.6 7.56
 [4,] 6.7 7.5 6.00
 [5,] 8.9 7.5 5.40
 [6,] 3.4 6.7 6.60
 [7,] 5.6 8.6 8.70
 [8,] 6.7 7.6 8.70
 > res <- princomp(W)</pre>
 > n <- 8
 > k < - 3
 > Z <- scale(W, scale = FALSE)</pre>
 > colnames(Z) <- c("z1", "z2", "z3")</pre>
        z1
              z2
 [1,] -3.8 -4.8125 -4.845
 [2,] -2.6 -2.6125 -0.645
 [3,] -0.4 -0.4125 1.315
 [4,] 1.8 1.4875 -0.245
 [5,] 4.0 1.4875 -0.845
 [6,] -1.5 0.6875 0.355
 [7,] 0.7 2.5875 2.455
 [8,] 1.8 1.5875 2.455
 attr(, "scaled:center")
     w1
          w2
                 w3
 4.9000 6.0125 6.2450
 > S < (1/n) * t(Z) %*% Z
 > dimnames(S) <- list(NULL, NULL)</pre>
 > S
                 [,2]
                           [,3]
          [,1]
 [1,] 5.82250 4.688750 2.668250
 [2,] 4.68875 5.533594 4.166437
 [3,] 2.66825 4.166437 4.821675
 > sdev <- sqrt(eigen(S)$values)</pre>
 > names(sdev) <- c("Comp.1", "Comp.2", "Comp.3")</pre>
 > sdev
             Comp.2
                        Comp.3
 3.6303620 1.6179210 0.6169052
 > res$sdev
```

```
Comp.1
             Comp.2
                       Comp.3
3.6303620 1.6179210 0.6169052
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
    w1
         w2
                  w3
4.9000 6.0125 6.2450
> res$center
   w1
          w2
4.9000 6.0125 6.2450
> n
[1] 8
> res$n.obs
[1] 8
> D <- diag(eigen(S)$values)</pre>
         [,1]
                  [,2]
                            [,3]
[1,] 13.17953 0.000000 0.0000000
[2,] 0.00000 2.617668 0.0000000
[3,] 0.00000 0.000000 0.3805721
> GAMMA <- eigen(S)$vectors</pre>
> GAMMA
          [,1]
                      [,2]
[1,] 0.5867813 0.68021602 0.4393107
[2,] 0.6341906 -0.04872184 -0.7716401
[3,] 0.5034779 -0.73139069 0.4599757
> scores <- Z %*% GAMMA
> colnames(scores) <- c("Comp.1", "Comp.2", "Comp.3")</pre>
> scores
         Comp.1
                   Comp.2
[1,] -7.7211617 1.1932409 -0.1844450
[2,] -3.5071975 -1.1695288 0.5770175
[3,] 0.1657573 -1.2137674 0.7474453
[4,] 1.8762127
                1.3311058 -0.4697494
[5,] 2.8650447
                3.2664155 0.2207489
[6,] -0.2654312 -1.3134640 -1.0261773
[7,] 3.2877534 -1.4454807 -0.5598609
[8,] 3.2990222 -0.6485212 0.6950210
> res$scores
         Comp.1
                    Comp.2
                              Comp.3
     7.7211617 1.1932409 -0.1844450
[1,]
[2,] 3.5071975 -1.1695288 0.5770175
[3,] -0.1657573 -1.2137674 0.7474453
[4,] -1.8762127 1.3311058 -0.4697494
[5,] -2.8650447 3.2664155 0.2207489
[6,] 0.2654312 -1.3134640 -1.0261773
[7,] -3.2877534 -1.4454807 -0.5598609
[8,] -3.2990222 -0.6485212 0.6950210
```

4.2 ACP con matrice di covarianza campionaria

Simbologia

- matrice dei dati di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici w_1, w_2, \ldots, w_k : W
- media di colonna della matrice dei dati: $\bar{w}_j \quad \forall j = 1, 2, \ldots, k$
- matrice dei dati centrata di dimensione $n \times k$: Z
- elemento di riga i e colonna j della matrice dei dati centrata: $z_{ij}=w_{ij}-\bar{w}_j \quad \forall i=1,2,\ldots,n \quad \forall j=1,2,\ldots,k$
- matrice di covarianza di dimensione $k \times k$: $S = \frac{Z^T Z}{n-1} = \Gamma D \Gamma^T$
- matrice ortogonale degli autovettori di dimensione $k \times k$: Γ
- *j*-esima colonna della matrice Γ : $\Gamma^j \quad \forall j = 1, 2, ..., k$
- matrice diagonale degli autovalori di dimensione $k \times k$: $D = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_k)$
- componente principale j-esima: $x_j = Z \Gamma^j \quad \forall j = 1, 2, ..., k$
- deviazione standard della j-esima componente principale: $s_{x_j}=\sqrt{\lambda_{(k-j+1)}} \quad \forall j=1,\,2,\,\ldots,\,k$
- problema di ottimo vincolato:

$$x_{j} = Z \gamma_{j} \quad \forall j = 1, 2, \dots, k$$

$$s_{x_{j}}^{2} = \frac{x_{j}^{T} x_{j}}{n-1} = \frac{(Z \gamma_{j})^{T} (Z \gamma_{j})}{n-1} = \gamma_{j}^{T} \frac{Z^{T} Z}{n-1} \gamma_{j} = \gamma_{j}^{T} S \gamma_{j} \quad \forall j = 1, 2, \dots, k$$

$$\max_{\gamma_{j}^{T} \gamma_{j} = 1} s_{x_{j}}^{2} = \max_{\gamma_{j}^{T} \gamma_{j} = 1} \gamma_{j}^{T} S \gamma_{j} = \lambda_{(k-j+1)} \quad \forall j = 1, 2, \dots, k$$

prcomp()

- Package: stats
- Input:

w matrice dei dati

• Output:

sdev deviazione standard delle componenti principali rotation matrice ortogonale degli autovettori center media di colonna della matrice W x componenti principali

• Formula:

sdev
$$s_{x_j} \quad \forall j=1,2,\ldots,k$$
 rotation
$$\Gamma$$
 center
$$\bar{w}_j \quad \forall j=1,2,\ldots,k$$
 x
$$x_j \quad \forall j=1,2,\ldots,k$$

```
> w1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> w2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> w3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> W <- cbind(w1, w2, w3)
> W
```

```
w1 w2
               w3
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
[8,] 6.7 7.6 8.70
> res <- prcomp(W)</pre>
> n < - 8
> k <- 3
> Z <- scale(W, scale = FALSE)</pre>
> colnames(Z) <- c("z1", "z2", "z3")</pre>
             z2
       z1
                     z3
[1,] -3.8 -4.8125 -4.845
[2,] -2.6 -2.6125 -0.645
[3,] -0.4 -0.4125 1.315
[4,] 1.8 1.4875 -0.245
[5,] 4.0 1.4875 -0.845
[6,] -1.5 0.6875 0.355
[7,] 0.7 2.5875 2.455
[8,] 1.8 1.5875 2.455
attr(, "scaled:center")
   w1
         w2
                 w3
4.9000 6.0125 6.2450
> S <- (1/(n-1)) * t(Z) %*% Z
> dimnames(S) <- list(NULL, NULL)</pre>
> S
                  [,2]
         [,1]
                            [,3]
[1,] 6.654286 5.358571 3.049429
[2,] 5.358571 6.324107 4.761643
[3,] 3.049429 4.761643 5.510486
> sdev <- sqrt(eigen(S)$values)</pre>
> sdev
[1] 3.8810202 1.7296303 0.6594994
> res$sdev
[1] 3.8810202 1.7296303 0.6594994
> GAMMA <- eigen(S)$vectors</pre>
> dimnames(GAMMA) <- list(c("w1", "w2", "w3"), c("PC1", "PC2",</pre>
      "PC3"))
> GAMMA
          PC1
                       PC2
w1 -0.5867813 -0.68021602 0.4393107
w2 -0.6341906 0.04872184 -0.7716401
w3 -0.5034779 0.73139069 0.4599757
> res$rotation
```

```
PC1
                     PC2
w1 0.5867813 0.68021602 -0.4393107
w2 0.6341906 -0.04872184 0.7716401
w3 0.5034779 -0.73139069 -0.4599757
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
         w2
    w1
                  w3
4.9000 6.0125 6.2450
> res$center
   w1
          w2
4.9000 6.0125 6.2450
> D <- diag(eigen(S)$values)</pre>
> D
         [,1]
                 [,2]
[1,] 15.06232 0.000000 0.0000000
[2,] 0.00000 2.991621 0.0000000
[3,] 0.00000 0.000000 0.4349395
> scores <- Z %*% GAMMA
> colnames(scores) <- c("PC1", "PC2", "PC3")</pre>
> scores
            PC1
                       PC2
                                 PC3
[1,] 7.7211617 -1.1932409 -0.1844450
[2,] 3.5071975 1.1695288 0.5770175
[3,] -0.1657573 1.2137674 0.7474453
[4,] -1.8762127 -1.3311058 -0.4697494
[5,] -2.8650447 -3.2664155 0.2207489
[6,] 0.2654312 1.3134640 -1.0261773
[7,] -3.2877534 1.4454807 -0.5598609
[8,] -3.2990222 0.6485212 0.6950210
> res$x
                      PC2
            PC1
                                  PC3
[1,] -7.7211617 1.1932409 0.1844450
[2,] -3.5071975 -1.1695288 -0.5770175
[3,] 0.1657573 -1.2137674 -0.7474453
[4,] 1.8762127 1.3311058 0.4697494
[5,] 2.8650447 3.2664155 -0.2207489
[6,] -0.2654312 -1.3134640 1.0261773
[7,] 3.2877534 -1.4454807 0.5598609
[8,] 3.2990222 -0.6485212 -0.6950210
```

summary()

• Package: base

• Input:

object oggetto di tipo prcomp()

• Output:

sdev deviazione standard delle componenti principali

rotation matrice ortogonale degli autovettori

center $\$ media di colonna della $\$ matrice $\ W$

x componenti principali

importance deviazione standard delle componenti principali, quota di varianza spiegata da ciascuna componente principale e quota di varianza spiegata dalle prime l componenti principali (l = 1, 2, ..., k)

• Formula:

sdev
$$s_{x_j} \quad \forall j=1,2,\ldots,k$$
 rotation
$$\Gamma$$
 center
$$\bar{w}_j \quad \forall j=1,2,\ldots,k$$

$$\mathbf{x}$$

$$x_j \quad \forall j=1,2,\ldots,k$$
 importance
$$s_{x_j} \quad \frac{\lambda_{(k-j+1)}}{\sum_{i=1}^k \lambda_i} \quad \frac{\sum_{j=1}^l \lambda_{(k-j+1)}}{\sum_{i=1}^k \lambda_i} \quad \forall j,l=1,2,\ldots,k$$

```
> w1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> w2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> w3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> W \leftarrow cbind(w1, w2, w3)
> W
      w1 w2
             w3
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
[8,] 6.7 7.6 8.70
> res <- summary(object = prcomp(W))</pre>
> n < - 8
> k < - 3
> Z <- scale(W, scale = FALSE)</pre>
> colnames(Z) <- c("z1", "z2", "z3")</pre>
       z1
             z2
                     z3
[1,] -3.8 -4.8125 -4.845
[2,] -2.6 -2.6125 -0.645
[3,] -0.4 -0.4125 1.315
[4,] 1.8 1.4875 -0.245
[5,] 4.0 1.4875 -0.845
          0.6875 0.355
[6,] -1.5
[7,] 0.7
          2.5875 2.455
[8,] 1.8 1.5875 2.455
attr(, "scaled:center")
   w1 w2 w3
4.9000 6.0125 6.2450
```

```
> S <- (1/(n - 1)) * t(Z) %*% Z
> dimnames(S) <- list(NULL, NULL)</pre>
         [,1]
                [,2] [,3]
[1,] 6.654286 5.358571 3.049429
[2,] 5.358571 6.324107 4.761643
[3,] 3.049429 4.761643 5.510486
> sdev <- sqrt(eigen(S)$values)</pre>
> sdev
[1] 3.8810202 1.7296303 0.6594994
> res$sdev
[1] 3.8810202 1.7296303 0.6594994
> GAMMA <- eigen(S)$vectors</pre>
> GAMMA
           [,1]
                       [,2]
                                  [,3]
[1,] -0.5867813 -0.68021602 0.4393107
[2,] -0.6341906 0.04872184 -0.7716401
[3,] -0.5034779 0.73139069 0.4599757
> res$rotation
         PC1
               PC2
w1 0.5867813 0.68021602 -0.4393107
w2 0.6341906 -0.04872184 0.7716401
w3 0.5034779 -0.73139069 -0.4599757
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
    w1
         w2
4.9000 6.0125 6.2450
> res$center
    w1
         w2
4.9000 6.0125 6.2450
> D <- diag(eigen(S)$values)</pre>
> D
         [,1]
                 [,2]
                           [,3]
[1,] 15.06232 0.000000 0.0000000
[2,] 0.00000 2.991621 0.0000000
[3,] 0.00000 0.000000 0.4349395
> x <- Z %*% GAMMA
> colnames(x) <- c("PC1", "PC2", "PC3")</pre>
> x
```

```
PC1
                      PC2
[1,] 7.7211617 -1.1932409 -0.1844450
[2,] 3.5071975 1.1695288 0.5770175
[3,] -0.1657573 1.2137674 0.7474453
[4,] -1.8762127 -1.3311058 -0.4697494
[5,] -2.8650447 -3.2664155 0.2207489
[6,] 0.2654312 1.3134640 -1.0261773
[7,] -3.2877534 1.4454807 -0.5598609
[8,] -3.2990222 0.6485212 0.6950210
> res$x
            PC1
                      PC2
[1,] -7.7211617 1.1932409 0.1844450
[2,] -3.5071975 -1.1695288 -0.5770175
[3,] 0.1657573 -1.2137674 -0.7474453
[4,] 1.8762127 1.3311058 0.4697494
[5,] 2.8650447 3.2664155 -0.2207489
[6,] -0.2654312 -1.3134640 1.0261773
[7,] 3.2877534 -1.4454807 0.5598609
[8,] 3.2990222 -0.6485212 -0.6950210
> lambda <- sdev^2
> importance <- rbind(sdev, lambda/sum(lambda), cumsum(lambda)/sum(lambda))</pre>
> dimnames(importance) <- list(c("Standard deviation", "Proportion of Variance",</pre>
+ "Cumulative Proportion"), c("PC1", "PC2", "PC3"))
> importance
                            PC1
                                      PC2
Standard deviation 3.8810202 1.7296303 0.65949942
Proportion of Variance 0.8146691 0.1618065 0.02352438
Cumulative Proportion 0.8146691 0.9764756 1.00000000
> res$importance
                          PC1
                                  PC2
                                            PC3
Standard deviation 3.88102 1.729630 0.6594994
Proportion of Variance 0.81467 0.161810 0.0235200
Cumulative Proportion 0.81467 0.976480 1.0000000
```

4.3 ACP con matrice di correlazione di popolazione

Simbologia

- matrice dei dati di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici w_1, w_2, \ldots, w_k : W
- media di colonna della matrice dei dati: $\bar{w}_j \quad \forall j = 1, 2, \dots, k$
- varianza campionaria di colonna della matrice dei dati: $\sigma_{w_j}^2 = n^{-1} \left(w_j \bar{w}_j \right)^T \left(w_j \bar{w}_j \right) \ \ \, \forall j=1,\,2,\,\ldots,\,k$
- matrice dei dati standardizzata di dimensione $n \times k$: Z
- elemento di riga i e colonna j della matrice dei dati standardizzata: $z_{ij} = (w_{ij} \bar{w}_j) / \sigma_{w_j} \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2, \dots, k$
- matrice di correlazione di dimensione $k \times k$: $R = \frac{Z^T Z}{n} = \Gamma D \Gamma^T$
- matrice ortogonale degli autovettori di dimensione $k \times k$: Γ
- *j*-esima colonna della matrice Γ : $\Gamma^j \quad \forall j = 1, 2, ..., k$

- matrice diagonale degli autovalori di dimensione $k \times k$: $D = diag(\lambda_1, \lambda_2, \dots, \lambda_k)$
- componente principale j-esima: $x_j = Z \Gamma^j \quad \forall j = 1, 2, ..., k$
- scarto quadratico medio della *j*-esima componente principale: $\sigma_{x_j} = \sqrt{\lambda_{(k-j+1)}} \quad \forall j = 1, 2, \dots, k$
- problema di ottimo vincolato:

$$x_j = Z\gamma_j \quad \forall j = 1, 2, \dots, k$$

$$\sigma^2 = \frac{x_j^T x_j}{\sigma^2} = \frac{(Z\gamma_j)^T (Z\gamma_j)}{\sigma^2} = \gamma_j^T \frac{Z^T Z}{\sigma^2} \gamma_j$$

$$\sigma_{x_{j}}^{2} = \frac{x_{j}^{T} x_{j}}{n} = \frac{(Z \gamma_{j})^{T} (Z \gamma_{j})}{n} = \gamma_{j}^{T} \frac{Z^{T} Z}{n} \gamma_{j} = \gamma_{j}^{T} R \gamma_{j} \quad \forall j = 1, 2, ..., k$$

$$\max_{\gamma_{j}^{T} \gamma_{j} = 1} \sigma_{x_{j}}^{2} = \max_{\gamma_{j}^{T} \gamma_{j} = 1} \gamma_{j}^{T} R \gamma_{j} = \lambda_{(k-j+1)} \quad \forall j = 1, 2, ..., k$$

princomp()

- Package: stats
- Input:

W matrice dei dati

cor = TRUE matrice di correlazione

• Output:

sdev scarto quadratico medio delle componenti principali center $\,$ media di colonna della $\,$ matrice Wscale $\,$ scarto quadratico medio di colonna della matrice Wn.obs dimensione campionaria scores componenti principali

• Formula:

sdev
$$\sigma_{x_j} \quad \forall j=1,2,\ldots,k$$
 center
$$\bar{w}_j \quad \forall j=1,2,\ldots,k$$
 scale
$$\sigma_{w_j} \quad \forall j=1,2,\ldots,k$$
 n.obs
$$n$$

 $x_i \quad \forall i = 1, 2, \ldots, k$

• Examples:

[8,] 6.7 7.6 8.70

```
> w1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> w2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> w3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> W <- cbind(w1, w2, w3)
> W
      w1 w2
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
```

```
> res <- princomp(W, cor = TRUE)</pre>
> n <- 8
> k < - 3
> sigma <- function(x) sqrt((length(x) - 1) * var(x)/length(x))
> Z <- sweep(W, 2, apply(W, MARGIN = 2, FUN = mean)) %*% diag(1/apply(W,
     MARGIN = 2, FUN = sigma))
> colnames(Z) <- c("z1", "z2", "z3")</pre>
> Z
                         7.2
             z.1
[1,] -1.5748125 -2.0458185 -2.2064537
[2,] -1.0775033 -1.1105872 -0.2937384
[3,] -0.1657697 -0.1753559 0.5988620
[4,] 0.7459638 0.6323439 -0.1115751
[5,] 1.6576973 0.6323439 -0.3848201
[6,] -0.6216365 0.2922598 0.1616700
[7,] 0.2900970 1.0999596 1.1180276
[8,] 0.7459638 0.6748544 1.1180276
> R <- (1/n) * t(Z) %*% Z
> dimnames(R) <- list(NULL, NULL)</pre>
          [,1]
                    [,2]
                              [,3]
[1,] 1.0000000 0.8260355 0.5035850
[2,] 0.8260355 1.0000000 0.8066075
[3,] 0.5035850 0.8066075 1.0000000
> sdev <- sqrt(eigen(R)$values)</pre>
> names(sdev) <- c("Comp.1", "Comp.2", "Comp.3")</pre>
> sdev
   Comp.1
            Comp.2
                       Comp.3
1.5599434 0.7047305 0.2644457
> res$sdev
            Comp.2
   Comp.1
                       Comp.3
1.5599434 0.7047305 0.2644457
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
          w2
                 w3
   w1
4.9000 6.0125 6.2450
> res$center
         w2
    w 1
                 w3
4.9000 6.0125 6.2450
> scale <- apply(W, MARGIN = 2, FUN = sigma)</pre>
> scale
      w1
              w2
2.412986 2.352359 2.195831
> res$scale
```

```
w1
              w2
                        w3
2.412986 2.352359 2.195831
> n
[1] 8
> res$n.obs
[1] 8
> D <- diag(eigen(R)$values)</pre>
                   [,2]
         [,1]
                             [,3]
[1,] 2.433423 0.0000000 0.0000000
[2,] 0.000000 0.4966451 0.0000000
[3,] 0.000000 0.0000000 0.0699315
> GAMMA <- eigen(R)$vectors</pre>
> GAMMA
           [,1]
                       [,2]
[1,] -0.5538345 -0.69330367 0.4610828
[2,] -0.6272670 -0.01674325 -0.7786242
[3,] -0.5475431 0.72045103 0.4256136
> scores <- Z %*% GAMMA
> colnames(scores) <- c("Comp.1", "Comp.2", "Comp.3")</pre>
> scores
          Comp.1
                   Comp.2
                                Comp.3
[1,] 3.36358843 -0.4635649 -0.07229172
[2,] 1.45422766 0.5540077 0.24289279
[3,] -0.12609881 0.5493156 0.31498656
[4,] -0.74869682 -0.6081513 -0.19589504
[5,] -1.10403287 -1.4371192 0.10819286
[6,] 0.07243752 0.5425648 -0.44537755
[7,] -1.46280241 0.5859419 -0.24684871
[8,] -1.44862269 0.2770054 0.29434081
> res$scores
          Comp.1
                   Comp.2
                            Comp.3
[1,] 3.36358843 -0.4635649 -0.07229172
[2,] 1.45422766 0.5540077 0.24289279
[3,] -0.12609881 0.5493156 0.31498656
[4,] -0.74869682 -0.6081513 -0.19589504
[5,] -1.10403287 -1.4371192 0.10819286
[6,] 0.07243752 0.5425648 -0.44537755
[7,] -1.46280241 0.5859419 -0.24684871
[8,] -1.44862269 0.2770054 0.29434081
```

4.4 ACP con matrice di correlazione campionaria

Simbologia

- matrice dei dati di dimensione $n \times k$ le cui colonne corrispondono ai vettori numerici w_1, w_2, \ldots, w_k : W
- media di colonna della matrice dei dati: $\bar{w}_i \quad \forall j = 1, 2, \dots, k$
- varianza campionaria di colonna della matrice dei dati: $s_{w_i}^2 = (n-1)^{-1} (w_j \bar{w}_j)^T (w_j \bar{w}_j) \quad \forall j=1,2,\ldots,k$
- matrice dei dati standardizzata di dimensione $n \times k$: Z
- elemento di riga i e colonna j della matrice dei dati standardizzata: $z_{ij} = (w_{ij} \bar{w}_j)/s_{w_i} \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., k$
- matrice di correlazione di dimensione $k \times k$: $R = \frac{Z^T Z}{n-1} = \Gamma D \Gamma^T$
- matrice ortogonale degli autovettori di dimensione $k \times k$: Γ
- *j*-esima colonna della matrice Γ : $\Gamma^j \quad \forall j = 1, 2, ..., k$
- matrice diagonale degli autovalori di dimensione $k \times k$: $D = diag(\lambda_1, \lambda_2, \dots, \lambda_k)$
- componente principale j-esima: $x_j = Z \Gamma^j \quad \forall j = 1, 2, ..., k$
- deviazione standard della j-esima componente principale: $s_{x_j}=\sqrt{\lambda_{(k-j+1)}} \quad \forall j=1,\,2,\,\ldots,\,k$
- problema di ottimo vincolato:

$$x_{j} = Z \gamma_{j} \quad \forall j = 1, 2, ..., k$$

$$s_{x_{j}}^{2} = \frac{x_{j}^{T} x_{j}}{n-1} = \frac{(Z \gamma_{j})^{T} (Z \gamma_{j})}{n-1} = \gamma_{j}^{T} \frac{Z^{T} Z}{n-1} \gamma_{j} = \gamma_{j}^{T} R \gamma_{j} \quad \forall j = 1, 2, ..., k$$

$$\max_{\gamma_{j}^{T} \gamma_{j} = 1} s_{x_{j}}^{2} = \max_{\gamma_{j}^{T} \gamma_{j} = 1} \gamma_{j}^{T} R \gamma_{j} = \lambda_{(k-j+1)} \quad \forall j = 1, 2, ..., k$$

prcomp()

- Package: stats
- Input:

W matrice dei dati scale. = TRUE matrice di correlazione

• Output:

sdev deviazione standard delle componenti principali rotation matrice ortogonale degli autovettori center media di colonna della matrice W scale deviazione standard di colonna della matrice W x componenti principali

• Formula:

sdev
$$s_{x_j} \quad \forall j=1,2,\ldots,k$$
 rotation
$$\Gamma$$
 center
$$\bar{w}_j \quad \forall j=1,2,\ldots,k$$
 scale
$$s_{w_j} \quad \forall j=1,2,\ldots,k$$
 x
$$x_j \quad \forall j=1,2,\ldots,k$$

```
> w1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> w2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> w3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> W <- cbind(w1, w2, w3)
> 1/7
      w1 w2 w3
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,] 3.4 6.7 6.60
[7,] 5.6 8.6 8.70
[8,] 6.7 7.6 8.70
> res <- prcomp(W, scale. = TRUE)</pre>
> n <- 8
> k < - 3
> Z <- scale(W, scale = TRUE)</pre>
> colnames(Z) <- c("z1", "z2", "z3")</pre>
             z1
                         z2
[1,] -1.4731022 -1.9136880 -2.0639484
[2,] -1.0079120 -1.0388592 -0.2747671
[3,] -0.1550634 -0.1640304 0.5601841
[4,] 0.6977852 0.5915036 -0.1043689
[5,] 1.5506339 0.5915036 -0.3599662
[6,] -0.5814877 0.2733840 0.1512284
[7,] 0.2713609 1.0289180 1.0458191
[8,] 0.6977852 0.6312685 1.0458191
attr(, "scaled:center")
    w1
        w2
4.9000 6.0125 6.2450
attr(, "scaled:scale")
           w2
                         wЗ
      w1
2.579590 2.514778 2.347442
> R < (1/(n - 1)) * t(Z) %*% Z
> dimnames(R) <- list(NULL, NULL)</pre>
> R
                     [,2]
          [,1]
                               [,3]
[1,] 1.0000000 0.8260355 0.5035850
[2,] 0.8260355 1.0000000 0.8066075
[3,] 0.5035850 0.8066075 1.0000000
> sdev <- sqrt(eigen(R)$values)</pre>
> sdev
[1] 1.5599434 0.7047305 0.2644457
> res$sdev
[1] 1.5599434 0.7047305 0.2644457
> D <- diag(eigen(R)$values)</pre>
> D
```

```
[,1]
                   [,2]
[1,] 2.433423 0.0000000 0.0000000
[2,] 0.000000 0.4966451 0.0000000
[3,] 0.000000 0.0000000 0.0699315
> GAMMA <- eigen(R)$vectors</pre>
> dimnames(GAMMA) <- list(c("w1", "w2", "w3"), c("PC1", "PC2",</pre>
+ "PC3"))
> GAMMA
         PC1
                     PC2
w1 0.5538345 0.69330367 0.4610828
w2 0.6272670 0.01674325 -0.7786242
w3 0.5475431 -0.72045103 0.4256136
> res$rotation
         PC1
                     PC2
w1 0.5538345 0.69330367 -0.4610828
w2 0.6272670 0.01674325 0.7786242
w3 0.5475431 -0.72045103 -0.4256136
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
   w1
         w2
4.9000 6.0125 6.2450
> res$center
         w2
   w 1
                 w3
4.9000 6.0125 6.2450
> scale <- apply(W, MARGIN = 2, FUN = sigma)</pre>
> scale
              w2
                       w3
      w1
2.412986 2.352359 2.195831
> res$scale
              w2
      w1
2.579590 2.514778 2.347442
> x <- Z %*% GAMMA
> colnames(x) <- c("PC1", "PC2", "PC3")</pre>
> x
                       PC2
             PC1
[1,] -3.14634887 0.4336252 -0.06762271
[2,] -1.36030541 -0.5182267 0.22720540
[3,] 0.11795463 -0.5138377 0.29464294
[4,] 0.70034175 0.5688735 -0.18324303
[5,] 1.03272818 1.3443019 0.10120515
[6,] -0.06775909 -0.5075229 -0.41661255
[7,] 1.36832636 -0.5480985 -0.23090583
[8,] 1.35506245 -0.2591149 0.27533061
> res$x
```

```
PC1 PC2 PC3
[1,] -3.14634887 0.4336252 0.06762271
[2,] -1.36030541 -0.5182267 -0.22720540
[3,] 0.11795463 -0.5138377 -0.29464294
[4,] 0.70034175 0.5688735 0.18324303
[5,] 1.03272818 1.3443019 -0.10120515
[6,] -0.06775909 -0.5075229 0.41661255
[7,] 1.36832636 -0.5480985 0.23090583
[8,] 1.35506245 -0.2591149 -0.27533061
```

summary()

• Package: base

• Input:

object oggetto di tipo prcomp()

• Output:

sdev deviazione standard delle componenti principali rotation matrice ortogonale degli autovettori center media di colonna della matrice W scale deviazione standard di colonna della matrice W x componenti principali

importance deviazione standard delle componenti principali, quota di varianza spiegata da ciascuna componente principale e quota di varianza spiegata dalle prime l componenti principali (l = 1, 2, ..., k)

• Formula:

sdev
$$s_{x_j} \quad \forall j=1,2,\ldots,k$$
 rotation
$$\Gamma$$
 center
$$\bar{w}_j \quad \forall j=1,2,\ldots,k$$
 scale
$$s_{w_j} \quad \forall j=1,2,\ldots,k$$

$$\mathbf{x}$$

$$x_j \quad \forall j=1,2,\ldots,k$$
 importance
$$s_{x_j} \quad \frac{\lambda_{(k-j+1)}}{k} \qquad \frac{1}{k} \sum_{j=1}^l \lambda_{(k-j+1)} \quad \forall j,l=1,2,\ldots,k$$

```
> res <- summary(object = prcomp(W, scale. = TRUE))</pre>
> n <- 8
> k < - 3
> Z <- scale(W, scale = TRUE)</pre>
> colnames(Z) <- c("z1", "z2", "z3")</pre>
> Z
             z1
                         z2
                                    z3
[1,] -1.4731022 -1.9136880 -2.0639484
[2,] -1.0079120 -1.0388592 -0.2747671
[3,] -0.1550634 -0.1640304 0.5601841
[4,] 0.6977852 0.5915036 -0.1043689
[5,]
     1.5506339
                 0.5915036 -0.3599662
[6,] -0.5814877 0.2733840 0.1512284
[7,] 0.2713609 1.0289180 1.0458191
[8,] 0.6977852 0.6312685 1.0458191
attr(, "scaled:center")
   w1
         w2
                 w3
4.9000 6.0125 6.2450
attr(, "scaled:scale")
     w1
              w2
2.579590 2.514778 2.347442
> R <- (1/(n - 1)) * t(Z) %*% Z
> dimnames(R) <- list(NULL, NULL)</pre>
> R
          [,1]
                    [,2]
                              [,3]
[1,] 1.0000000 0.8260355 0.5035850
[2,] 0.8260355 1.0000000 0.8066075
[3,] 0.5035850 0.8066075 1.0000000
> sdev <- sqrt(eigen(R)$values)</pre>
> sdev
[1] 1.5599434 0.7047305 0.2644457
> res$sdev
[1] 1.5599434 0.7047305 0.2644457
> GAMMA <- eigen(R)$vectors</pre>
> dimnames(GAMMA) <- list(c("w1", "w2", "w3"), c("PC1", "PC2",</pre>
+ "PC3"))
> GAMMA
                      PC2
                                 PC3
w1 0.5538345 0.69330367 0.4610828
w2 0.6272670 0.01674325 -0.7786242
w3 0.5475431 -0.72045103 0.4256136
> res$rotation
         PC1
                      PC2
                                 PC3
w1 0.5538345 0.69330367 -0.4610828
w2 0.6272670 0.01674325 0.7786242
w3 0.5475431 -0.72045103 -0.4256136
> center <- apply(W, MARGIN = 2, FUN = mean)</pre>
> center
```

```
w1
         w2
4.9000 6.0125 6.2450
> res$center
         w2
                w3
   w1
4.9000 6.0125 6.2450
> scale <- apply(W, MARGIN = 2, FUN = sd)
> scale
      w1
              w2
2.579590 2.514778 2.347442
> res$scale
             w2
2.579590 2.514778 2.347442
> D <- diag(eigen(S)$values)</pre>
         [,1]
                 [,2]
                            [,3]
[1,] 15.06232 0.000000 0.0000000
[2,] 0.00000 2.991621 0.0000000
[3,] 0.00000 0.000000 0.4349395
> x <- Z %*% GAMMA
> colnames(x) <- c("PC1", "PC2", "PC3")</pre>
            PC1
                       PC2
                                   PC3
[1,] -3.14634887 0.4336252 -0.06762271
[2,] -1.36030541 -0.5182267 0.22720540
[3,] 0.11795463 -0.5138377 0.29464294
[4,] 0.70034175 0.5688735 -0.18324303
[5,] 1.03272818 1.3443019 0.10120515
[6,] -0.06775909 -0.5075229 -0.41661255
[7,] 1.36832636 -0.5480985 -0.23090583
[8,] 1.35506245 -0.2591149 0.27533061
> res$x
             PC1
                       PC2
[1,] -3.14634887 0.4336252 0.06762271
[2,] -1.36030541 -0.5182267 -0.22720540
[3,] 0.11795463 -0.5138377 -0.29464294
[4,] 0.70034175 0.5688735 0.18324303
[5,] 1.03272818 1.3443019 -0.10120515
[6,] -0.06775909 -0.5075229 0.41661255
[7,] 1.36832636 -0.5480985 0.23090583
[8,] 1.35506245 -0.2591149 -0.27533061
> lambda <- sdev^2</pre>
> importance <- rbind(sdev, lambda/k, cumsum(lambda)/k)</pre>
> dimnames(importance) <- list(c("Standard deviation", "Proportion of Variance",</pre>
+ "Cumulative Proportion"), c("PC1", "PC2", "PC3"))
> importance
```

	PC1	PC2	PC3
Standard deviation	1.5599434	0.7047305	0.2644457
Proportion of Variance	0.8111411	0.1655484	0.0233105
Cumulative Proportion	0.8111411	0.9766895	1.0000000

> res\$importance

	PC1	PC2	PC3
Standard deviation	1.559943	0.7047305	0.2644457
Proportion of Variance	0.811140	0.1655500	0.0233100
Cumulative Proportion	0.811140	0.9766900	1.0000000

Capitolo 5

Analisi dei Gruppi

5.1 Indici di distanza

```
dist()
```

• Package: stats

• Input:

```
x matrice di dimensione n \times k le cui righe corrispondono ai vettori numerici x_1, x_2, \ldots, x_n method = "euclidean" / "maximum" / "manhattan" / "canberra" / "binary" / "minkowski" indice di distanza per la distanza di Minkowski upper = TRUE diag = TRUE
```

- **Description:** matrice di distanza o di dissimilarità per gli n vettori di dimensione $n \times n$
- Formula:

$$d_{x_ix_j} = \left(\sum_{h=1}^k (x_{ih} - x_{jh})^2\right)^{1/2} \quad \forall i,j = 1, 2, \dots, n$$

$$\boxed{\text{method} = \text{"maximum"}}$$

$$d_{x_ix_j} = \max_h |x_{ih} - x_{jh}| \quad \forall i,j = 1, 2, \dots, n$$

$$\boxed{\text{method} = \text{"manhattan"}}$$

$$d_{x_ix_j} = \sum_{h=1}^k |x_{ih} - x_{jh}| \quad \forall i,j = 1, 2, \dots, n$$

$$\boxed{\text{method} = \text{"canberra"}}$$

$$d_{x_ix_j} = \sum_{h=1}^k \frac{x_{ih} - x_{jh}}{x_{ih} + x_{jh}} \quad \forall i,j = 1, 2, \dots, n$$

$$\boxed{\text{method} = \text{"binary"}}$$

$$d_{x_ix_j} = 1 - \frac{n_{11}}{n_{01} + n_{10} + n_{11}} \quad \forall i,j = 1, 2, \dots, n$$

$$\boxed{\text{method} = \text{"minkowski"}}$$

$$d_{x_i x_j} = \left(\sum_{h=1}^k |x_{ih} - x_{jh}|^p\right)^{1/p} \quad \forall i, j = 1, 2, \dots, n$$

• Examples:

```
> x < -matrix(data = rnorm(n = 30), nrow = 10, ncol = 3, byrow = FALSE)
> k < - 3
> n <- 10
> dist(x, method = "euclidean", upper = TRUE, diag = TRUE)
  0.0000000 1.5948359 1.6080407 1.5836525 2.2113048 3.0581815 2.3820407
  1.5948359 0.0000000 1.4765220 1.5084132 0.9847730 2.9608231 0.8150047
  1.6080407 1.4765220 0.0000000 1.8622265 2.3977451 1.7540114 1.9745533
  1.5836525 1.5084132 1.8622265 0.0000000 1.6478362 2.6834204 2.1774463
  2.2113048 0.9847730 2.3977451 1.6478362 0.0000000 3.6618122 1.0875239
 6 \quad 3.0581815 \quad 2.9608231 \quad 1.7540114 \quad 2.6834204 \quad 3.6618122 \quad 0.0000000 \quad 3.3142664 
  2.3820407 0.8150047 1.9745533 2.1774463 1.0875239 3.3142664 0.0000000
  3.4274432 2.2298585 2.1613885 3.3445427 2.8214454 2.8972571 1.7918570
9 1.2371199 2.3024300 2.7601394 1.8380083 2.4297830 4.0248341 3.0452671
10 3.6159883 2.4770211 2.3594738 2.7396964 2.7641401 2.1990887 2.2918994
   3.4274432 1.2371199 3.6159883
  2.2298585 2.3024300 2.4770211
  2.1613885 2.7601394 2.3594738
  3.3445427 1.8380083 2.7396964
  2.8214454 2.4297830 2.7641401
  2.8972571 4.0248341 2.1990887
  1.7918570 3.0452671 2.2918994
  0.0000000 4.4430280 1.8632088
  4.4430280 0.0000000 4.4151604
10 1.8632088 4.4151604 0.0000000
> dist(x, method = "minkowski", p = 1, upper = TRUE, diag = TRUE)
                                      4
  0.000000 2.511879 2.548073 2.084588 3.795046 5.216133 3.593517 4.051206
  2.511879 0.000000 1.680889 2.443684 1.416056 3.923327 1.081638 3.134763
  2.548073 1.680889 0.000000 3.218951 2.964057 2.668059 2.762527 2.681157
  2.084588 2.443684 3.218951 0.000000 2.707806 3.603471 3.501799 4.819033
  3.795046 1.416056 2.964057 2.707806 0.000000 4.320338 1.832726 4.550819
6 5.216133 3.923327 2.668059 3.603471 4.320338 0.000000 4.704210 4.925776
   3.593517 1.081638 2.762527 3.501799 1.832726 4.704210 0.000000 2.718093
  4.051206 3.134763 2.681157 4.819033 4.550819 4.925776 2.718093 0.000000
  1.984456 2.705089 3.960357 3.037213 3.622008 6.628417 3.420478 5.463490
10 5.547416 4.254610 3.611224 3.922487 4.651621 3.572303 3.814418 2.523997
          9
  1.984456 5.547416
  2.705089 4.254610
  3.960357 3.611224
   3.037213 3.922487
  3.622008 4.651621
  6.628417 3.572303
  3.420478 3.814418
  5.463490 2.523997
  0.000000 6.959700
10 6.959700 0.000000
```

- Note 1: Possiamo ottenere le variabili standardizzate se applichiamo il comando scale() alla matrice x.
- **Note 2:** La distanza di dissimilarità calcolata con method = "binary" corrisponde al complemento ad uno dell'indice di *Jaccard*.

as.dist()

• Package: stats

• Input:

```
m matrice simmetrica con elementi nulli sulla diagonale di dimensione n \times n upper = TRUE / FALSE matrice triangolare superiore diag = TRUE / FALSE elementi nulli sulla diagonale
```

• **Description:** oggetto di tipo dist ()

• Examples:

```
> m <- matrix(data = c(0, 1, 5, 1, 0, 3, 5, 3, 0), nrow = 3, ncol = 3,
    byrow = TRUE)
> m
   [,1] [,2] [,3]
    0
         1
[1,]
                 3
           0
      1
[2,]
[3,]
    5
           3
                 0
> n < -3
> as.dist(m, upper = TRUE, diag = TRUE)
 1 2 3
1 0 1 5
2 1 0 3
3 5 3 0
> as.dist(m, upper = TRUE, diag = FALSE)
 1 2 3
1 1 5
2 1
    3
3 5 3
> as.dist(m, upper = FALSE, diag = TRUE)
 1 2 3
1 0
2 1 0
3 5 3 0
> as.dist(m, upper = FALSE, diag = FALSE)
 1 2
2 1
3 5 3
```

mahalanobis()

```
• Package: stats
```

• Input:

```
x vettore numerico di dimensione k center vettore numerico \bar{x} delle medie di dimensione k cov matrice S di covarianza di dimensione k \times k
```

- Description: quadrato della distanza di Mahalanobis
- Formula:

$$MD^2 = (x - \bar{x})^T S^{-1} (x - \bar{x})$$

• Example 1:

```
> X <- matrix(data = c(1.1, 1.2, 1.4, 2.3, 3.4, 5.6, 4.5, 5.6,
     7.56, 6.7, 7.5, 6, 8.9, 7.5, 5.4, 3.4, 6.7, 6.6, 5.6, 8.6,
     8.7, 6.7, 7.6, 8.7), nrow = 8, ncol = 3, byrow = TRUE)
> X
     [,1] [,2] [,3]
[1,] 1.1 1.2 1.40
[2,] 2.3 3.4 5.60
[3,] 4.5 5.6 7.56
[4,] 6.7 7.5 6.00
[5,] 8.9 7.5 5.40
[6,]
     3.4 6.7 6.60
[7,] 5.6 8.6 8.70
          7.6 8.70
[8,] 6.7
> k < - 3
> medie <- apply(X, MARGIN = 2, FUN = mean)</pre>
> S <- cov(X)
> x < -c(1.2, 3.4, 5.7)
> as.numeric(t(x - medie) %*% solve(S) %*% (x - medie))
[1] 2.487141
> mahalanobis(x, center = medie, cov = S)
[1] 2.487141
```

```
> X <- matrix(data = c(1.1, 3.4, 2.3, 5.6, 4.5, 6.7, 6.7, 6.7,
+ 8.9, 8.6), nrow = 5, ncol = 2, byrow = FALSE)
> X

      [,1] [,2]
[1,] 1.1 6.7
[2,] 3.4 6.7
[3,] 2.3 6.7
[4,] 5.6 8.9
[5,] 4.5 8.6

> k <- 2
> medie <- apply(X, MARGIN = 2, FUN = mean)
> S <- cov(X)
> x <- c(1.4, 6.7)
> as.numeric(t(x - medie) %*% solve(S) %*% (x - medie))
[1] 1.530355
```

```
> mahalanobis(x, center = medie, cov = S)
 [1] 1.530355
• Example 3:
 > X < -matrix(data = c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7,
       1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6, 1.4, 5.6, 7.56, 6,
       5.4, 6.6, 8.7, 8.7, 1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6),
       nrow = 8, ncol = 4, byrow = TRUE)
      [,1] [,2] [,3] [,4]
 [1,] 1.10 2.3 4.50 6.7
 [2,] 8.90 3.4 5.60 6.7
 [3,] 1.20 3.4 5.60
                      7.5
 [4,] 7.50 6.7 8.60
                      7.6
 [5,] 1.40 5.6 7.56 6.0
 [6,] 5.40 6.6 8.70 8.7
 [7,] 1.50 6.4 9.60 8.8
 [8,] 8.86 7.8 8.60 8.6
 > k <- 4
 > medie <- apply(X, MARGIN = 2, FUN = mean)</pre>
 > S <- cov(X)
 > x < -c(1.1, 2.4, 10.4, 7.8)
 > as.numeric(t(x - medie) %*% solve(S) %*% (x - medie))
 [1] 114.4839
 > mahalanobis(x, center = medie, cov = S)
 [1] 114.4839
```

5.2 Criteri di Raggruppamento

hclust()

```
• Package: stats
```

• Input:

```
d oggetto di tipo dist()
method = "ward" / "single" / "complete" / "average" / "mcquitty" / "median" /
"centroid" criterio di Ward, Legame Singolo, Legame Completo, Legame Medio, McQuitty, Mediana
e Centroide
```

- **Description:** analisi dei gruppi per gli n vettori di dimensione k
- Output:

merge matrice di dimensione $(n-1)\times 2$ le cui righe descrivono le aggregazioni avvenute a ciascun passo dell'intero procedimento. Gli elementi negativi indicano singole unità, mentre quelli positivi indicano gruppi già formati

height vettore di n-1 valori numerici non decrescenti che indicano i livelli di dissomiglianza ai quali avvengono le aggregazioni

```
order permutazioni delle osservazioni originali
labels vettore delle etichette delle osservazioni
method criterio di aggregazione utilizzato
dist.method criterio di distanza utilizzato
```

• Formula:

$$d_{(xy)z} = \frac{(n_x + n_z) \, d_{xz} + (n_y + n_z) \, d_{yz} - n_z \, d_{(xy)}}{n_{xy} + n_z}$$

$$\text{method = "single"}$$

$$d_{(xy)z} = \min(d_{xz}, d_{yz})$$

$$\text{method = "complete"}$$

$$d_{(xy)z} = \max(d_{xz}, d_{yz})$$

$$\text{method = "average"}$$

$$d_{(xy)z} = \frac{n_x \, d_{xz} + n_y \, d_{yz}}{n_{(xy)}}$$

$$\text{method = "mequitty"}$$

$$d_{(xy)z} = \frac{d_{xz} + d_{yz}}{2}$$

$$\text{method = "median"}$$

$$d_{(xy)z} = \frac{d_{xz} + d_{yz}}{2} - \frac{d_{(xy)}}{4}$$

$$\text{method = "centroid"}$$

$$d_{(xy)z} = \frac{n_x \, d_{xz} + n_y \, d_{yz}}{n_{(xy)}} - \frac{n_x \, n_y \, d_{xy}}{n_{(xy)}^2}$$

```
> x <- matrix(data = rnorm(n = 30), nrow = 3, ncol = 10, byrow = FALSE)
> k <- 3
> n <- 10
> d <- dist(x, method = "euclidean", upper = TRUE, diag = TRUE)</pre>
> hclust(d = d, method = "single")
Call:
hclust(d = d, method = "single")
Cluster method : single
Distance
           : euclidean
Number of objects: 3
> res <- hclust(d = d, method = "single")</pre>
> res$merge
   [,1] [,2]
[1,] -2 -3 [2,] -1 1
> res$height
[1] 2.985362 3.761878
```

```
> res$order
 [1] 1 2 3
 > res$labels
 NULL
 > res$method
 [1] "single"
 > res$dist.method
 [1] "euclidean"
• Example 2:
 > x < -matrix(data = rnorm(n = 100), nrow = 20, ncol = 5, byrow = FALSE)
 > k < - 3
 > n < -10
 > d <- dist(x, method = "euclidean", upper = TRUE, diag = TRUE)</pre>
 > hclust(d = d, method = "median")
 Call:
 hclust(d = d, method = "median")
 Cluster method : median
 Distance : euclidean
 Number of objects: 20
 > res <- hclust(d = d, method = "median")</pre>
 > res$merge
      [,1] [,2]
  [1,] -6 -16
            1
  [2,]
        -2
  [3,] -14
              2
  [4,] -12 -20
  [5,] -19 4
        3
  [6,]
       -15 6
  [7,]
       -13 -18
  [8,]
            8
  [9,] -10
              9
 [10,] -11
       7 10
 [11,]
 [12,]
       -4 -17
 [13,]
       11 12
        -5 13
 [14,]
        -7 14
 [15,]
 [16,]
        -1
             -8
        15
            16
 [17,]
        -3 17
 [18,]
        -9 18
 [19,]
 > res$height
  [1] 1.129097 1.070475 1.196478 1.351082 1.274444 1.390697 1.335846 1.440786
  [9] 1.606760 1.559425 1.650469 1.819976 1.762757 1.643485 2.162323 2.422278
 [17] 2.680234 2.464257 2.140949
```

```
> res$order
    [1] 9 3 7 5 15 14 2 6 16 19 12 20 11 10 13 18 4 17 1 8
   > res$labels
   NULL
   > res$method
   [1] "median"
   > res$dist.method
   [1] "euclidean"
kmeans()
  • Package: stats
  • Input:
       x matrice di dimensione n \times k le cui righe corrispondono ai vettori numerici x_1, x_2, \ldots, x_n
       centers scalare che indica il numero di gruppi
       iter.max massimo numero di iterazioni concesse al criterio di ottimizzazione
  • Description: analisi di ragguppamento non gerarchica con il metodo k-means
  • Output:
       cluster gruppo di appartenenza di ciascuna osservazione
       centers centroidi dei gruppi ottenuti
       withinss devianza di ciascun gruppo
       size numero di osservazioni in ciascun gruppo
  • Example 1:
   > x < -matrix(data = rnorm(n = 100, mean = 0, sd = 0.3), nrow = 50,
   + ncol = 2, byrow = FALSE)
   > kmeans(x, centers = 2, iter.max = 10)
   K-means clustering with 2 clusters of sizes 29, 21
   Cluster means:
            [,1]
                        [,2]
   1 -0.05916688 -0.1945814
   2 0.04105267 0.2989030
   Clustering vector:
    [39] 2 1 1 1 2 2 1 1 1 2 2 1
   Within cluster sum of squares by cluster:
   [1] 2.771814 2.263145
   Available components:
   [1] "cluster" "centers" "withinss" "size"
   > res <- kmeans(x, centers = 2, iter.max = 10)</pre>
```

> res\$cluster

```
[39] 2 2 2 2 2 2 1 2 1 2 1 2
 > res$centers
          [,1]
                    [,2]
 1 0.07741224 -0.2356923
 2 -0.10429336 0.2419507
 > res$withinss
 [1] 2.079959 2.784218
 > res$size
 [1] 24 26
• Example 2:
 > x < matrix(data = rnorm(n = 80, mean = 0, sd = 0.3), nrow = 40,
 + ncol = 2, byrow = FALSE)
 > kmeans(x, centers = 5, iter.max = 15)
 K-means clustering with 5 clusters of sizes 5, 5, 7, 13, 10
 Cluster means:
         [,1]
                     [,2]
 1 -0.2826432 0.37367857
 2 -0.4721982 -0.53828582
 3 0.2601737 0.14589161
 4 -0.2726225 -0.07709169
 5 0.2381249 -0.14376129
 Clustering vector:
  [1] 4 4 3 4 5 5 5 4 5 1 1 4 4 3 2 1 4 2 2 4 5 3 1 4 4 5 4 3 4 5 3 1 3 5 2 5 3 5
 [39] 2 4
 Within cluster sum of squares by cluster:
 [1] 0.2127299 0.2585805 0.1444599 0.4426205 0.2739510
 Available components:
 [1] "cluster" "centers" "withinss" "size"
 > res <- kmeans(x, centers = 5, iter.max = 15)
 > res$cluster
  [1] \ 2 \ 3 \ 5 \ 3 \ 5 \ 5 \ 2 \ 3 \ 2 \ 1 \ 1 \ 3 \ 3 \ 5 \ 4 \ 1 \ 2 \ 4 \ 4 \ 3 \ 2 \ 5 \ 1 \ 3 \ 3 \ 2 \ 3 \ 5 \ 5 \ 5 \ 5 \ 5 \ 4 \ 5 \ 2 \ 2
 [39] 4 3
 > res$centers
          [,1]
                      [,2]
 1 -0.28264316 0.37367857
 2 0.06019474 -0.09067425
 3 -0.30619549 -0.08337684
 4 -0.47219821 -0.53828582
 5 0.32226949 0.02036143
 > res$withinss
 [1] 0.2127299 0.2084292 0.3159412 0.2585805 0.4271144
```

```
> res$size
```

[1] 5 8 11 5 11

Parte III Statistica Inferenziale

Capitolo 6

Test di ipotesi parametrici

6.1 Test di ipotesi sulla media con uno o due campioni

Test Z con un campione

```
• Package: BSDA
```

• **Sintassi:** z.test()

• Input:

```
x vettore numerico di dimensione n sigma.x valore di \sigma_x mu valore di \mu_0 alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica Z p.value p-value conf.int intervallo di confidenza per la media incognita a livello 1-\alpha estimate media campionaria null.value valore di \mu_0 alternative ipotesi alternativa
```

• Formula:

statistic $z \, = \, \frac{\bar{x} - \mu_0}{\sigma_x \, / \, \sqrt{n}} \label{eq:z}$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(- z)$

conf.int
$$\bar{x}\mp z_{1-\alpha/2}\,\sigma_x\,/\,\sqrt{n}$$
 estimate
$$\bar{x}$$
 null.value
$$\mu_0$$

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> xmedio <- 7.018182</pre>
> sigmax <- 1.2</pre>
> n <- 11
> mu0 < - 6.5
> z <- (xmedio - mu0)/(sigmax/sqrt(n))</pre>
> z
[1] 1.432179
> res <- z.test(x, sigma.x = 1.2, mu = 6.5, alternative = "two.sided",
+ conf.level = 0.95)
> res$statistic
1.432179
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 0.1520925
> res$p.value
[1] 0.1520926
> alpha <- 0.05</pre>
> lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> c(lower, upper)
[1] 6.309040 7.727324
> res$conf.int
[1] 6.309040 7.727323
attr(,"conf.level")
[1] 0.95
> xmedio
[1] 7.018182
> res$estimate
mean of x
 7.018182
> mu0
[1] 6.5
> res$null.value
mean
 6.5
> res$alternative
```

[1] "two.sided"

```
[1] "two.sided"
• Example 2:
 > x < -c(1, 2.3, 4.5, 6.7, 8.9)
 > xmedio <- 4.68
 > sigmax <- 1.45
 > n < -5
 > mu0 < - 5.2
 > z <- (xmedio - mu0)/(sigmax/sqrt(n))</pre>
 [1] -0.8019002
 > res <- z.test(x, sigma.x = 1.45, mu = 5.2, alternative = "two.sided",
 + conf.level = 0.95)
 > res$statistic
 -0.8019002
 > p.value <-2 * pnorm(-abs(z))
 > p.value
 [1] 0.4226107
 > res$p.value
 [1] 0.4226107
 > alpha <- 0.05
 > lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
 > upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
 > c(lower, upper)
 [1] 3.409042 5.950958
 > res$conf.int
 [1] 3.409042 5.950958
 attr(,"conf.level")
 [1] 0.95
 > xmedio
 [1] 4.68
 > res$estimate
 mean of x
      4.68
 > mu0
 [1] 5.2
 > res$null.value
 mean
  5.2
 > res$alternative
```

Test di Student con un campione

```
• Package: stats
```

• Sintassi: t.test()

• Input:

```
x vettore numerico di dimensione n mu valore di \mu_0 alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica t parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la media incognita a livello 1-\alpha estimate media campionaria null.value valore di \mu_0 alternative ipotesi alternativa
```

• Formula:

statistic
$$t = \frac{\bar{x} - \mu_0}{s_x / \sqrt{n}}$$
 parameter
$$df = n - 1$$
 p.value

alternative	less	greater	two.sided
p.value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \le - t)$

```
conf.int \bar{x}\mp t_{1-\alpha/2,\,\mathrm{df}}\,s_x\,/\,\sqrt{n} estimate \bar{x} null.value \mu_0
```

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> xmedio <- 7.018182
> sx <- 0.4643666
> n <- 11
> mu0 <- 6.5
> t <- (xmedio - mu0)/(sx/sqrt(n))
> t

[1] 3.700988
> res <- t.test(x, mu = 6.5, alternative = "two.sided", conf.level = 0.95)
> res$statistic
t
3.700987
```

```
> parameter <- n - 1
> parameter
[1] 10
> res$parameter
df
10
> p.value < 2 * pt(-abs(t), df = n - 1)
> p.value
[1] 0.004101807
> res$p.value
[1] 0.004101817
> alpha <- 0.05
> lower <- xmedio - qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
> upper <- xmedio + qt(1 - 0.05/2, df = n - 1) \star sx/sqrt(n)
> c(lower, upper)
[1] 6.706216 7.330148
> res$conf.int
[1] 6.706216 7.330148
attr(,"conf.level")
[1] 0.95
> xmedio
[1] 7.018182
> res$estimate
mean of x
7.018182
> mu0
[1] 6.5
> res$null.value
mean
 6.5
> res$alternative
[1] "two.sided"
```

```
> x < -c(1, 2.3, 4.5, 6.7, 8.9)
> xmedio <- 4.68
> sx <- 3.206556
> n <- 5
> mu0 < - 5.2
> t <- (xmedio - mu0)/(sx/sqrt(n))
> t
[1] -0.3626181
> res <- t.test(x, mu = 5.2, alternative = "two.sided", conf.level = 0.95)</pre>
> res$statistic
-0.3626182
> parameter <- n - 1
> parameter
[1] 4
> res$parameter
df
 4
> p.value < 2 * pt(-abs(t), df = n - 1)
> p.value
[1] 0.7352382
> res$p.value
[1] 0.7352382
> alpha <- 0.05</pre>
> lower <- xmedio - qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
> upper <- xmedio + qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
> c(lower, upper)
[1] 0.6985349 8.6614651
> res$conf.int
[1] 0.6985351 8.6614649
attr(,"conf.level")
[1] 0.95
> mean(x)
[1] 4.68
> res$estimate
mean of x
     4.68
> mu0
```

```
[1] 5.2
> res$null.value
mean
   5.2
> res$alternative
[1] "two.sided"
```

Test Z con due campioni indipendenti

```
• Package: BSDA
```

• **Sintassi:** z.test()

• Input:

```
x vettore numerico di dimensione n_x y vettore numerico di dimensione n_y sigma.x valore di \sigma_x sigma.y valore di \sigma_y mu valore di (\mu_x - \mu_y)_{|H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica Z p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello 1-\alpha estimate medie campionarie null.value valore di (\mu_x-\mu_y)_{|H_0} alternative ipotesi alternativa
```

• Formula:

statistic

$$z = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{\sqrt{\sigma_x^2 / n_x + \sigma_y^2 / n_y}}$$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z)$

conf.int
$$\bar{x}-\bar{y}\mp z_{1-\alpha/2}\sqrt{{\sigma_x^2/n_x+\sigma_y^2/n_y}}$$
 estimate
$$\bar{x}=\bar{y}$$
 null.value
$$(\mu_x-\mu_y)_{|H_0}$$

```
> x < -c(154, 109, 137, 115, 140)
> xmedio <- 131
> sigmax <- 15.5
> nx <- 5
> y <- c(108, 115, 126, 92, 146)
> ymedio <- 117.4
> sigmay <- 13.5</pre>
> ny <- 5
> mu0 <- 10
> z <- (xmedio - ymedio - mu0)/sqrt(sigmax^2/nx + sigmay^2/ny)
[1] 0.3916284
> res <- z.test(x, y, sigma.x = 15.5, sigma.y = 13.5, mu = 10,
+ alternative = "two.sided", conf.level = 0.95)
> res$statistic
0.3916284
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 0.6953328
> res$p.value
[1] 0.6953328
> alpha <- 0.05</pre>
> lower <- (xmedio - ymedio) - qnorm(1 - 0.05/2) * sqrt(sigmax^2/nx +
     sigmay^2/ny)
> upper <- (xmedio - ymedio) + qnorm(1 - 0.05/2) * sqrt(sigmax^2/nx +
     sigmay^2/ny)
> c(lower, upper)
[1] -4.41675 31.61675
> res$conf.int
[1] -4.41675 31.61675
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean of x mean of y
   131.0 117.4
> mu0
[1] 10
> res$null.value
```

```
difference in means
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x \leftarrow c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
 > xmedio <- 7.018182
 > sigmax <- 0.5
 > nx <- 11
 > y <- c(4.5, 5.4, 6.1, 6.1, 5.4, 5, 4.1, 5.5)
 > ymedio <- mean(y)</pre>
 > ymedio
 [1] 5.2625
 > sigmay <- 0.8
 > ny <- length(y)</pre>
 > ny
 [1] 8
 > mu0 <- 1.2
 > z <- (xmedio - ymedio - mu0)/sqrt(sigmax^2/nx + sigmay^2/ny)</pre>
 > res <- z.test(x, y, sigma.x = 0.5, sigma.y = 0.8, mu = 1.2, alternative = "two.sided",
      conf.level = 0.95)
 > res$statistic
 1.733737
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.0829646
 > res$p.value
 [1] 0.0829647
 > alpha <- 0.05
 > lower <- (xmedio - ymedio) - qnorm(1 - 0.05/2) * sqrt(sigmax^2/nx +
       sigmay^2/ny)
 > upper <- (xmedio - ymedio) + qnorm(1 - 0.05/2) * sqrt(sigmax^2/nx +
       sigmay^2/ny)
 > c(lower, upper)
 [1] 1.127492 2.383872
 > res$conf.int
 [1] 1.127492 2.383872
 attr(,"conf.level")
 [1] 0.95
 > c(xmedio, ymedio)
```

Test di Student con due campioni indipendenti con varianze non note e supposte uguali

```
• Sintassi: t.test()
• Input:

x vettore numerico di dimensione n_x
y vettore numerico di dimensione n_y
mu valore di (\mu_x - \mu_y)_{|H_0}
alternative = "less" / "greater" / "two.sided" ipotesi alternativa
```

conf.level livello di confidenza $1-\alpha$

var.equal = TRUE

• Output:

• Package: stats

statistic valore empirico della statistica t parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello $1-\alpha$ estimate medie campionarie null.value valore di $(\mu_x-\mu_y)_{|H_0}$ alternative ipotesi alternativa

• Formula:

statistic $t=\frac{(\bar x-\bar y)-(\mu_x-\mu_y)_{|\,H_0}}{s_P\sqrt{1/n_x+1/n_y}}$ dove $s_P^2=\frac{(n_x-1)\,s_x^2+(n_y-1)\,s_y^2}{n_x+n_y-2}$ parameter $d\!f=n_x+n_y-2$ p.value

alternative	less	greater	two.sided
p.value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \leq - t)$

```
conf.int \bar{x}-\bar{y}\mp t_{1-\alpha/2,\,df}\,s_P\,\sqrt{1\left/\,n_x+1\right/\,n_y} estimate \bar{x}\quad\bar{y} null.value (\mu_x-\mu_y)_{|\,H_0}
```

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> xmedio <- 7.018182</pre>
> sx <- 0.4643666
> nx <- 11
> y <- c(4.5, 5.4, 6.1, 6.1, 5.4, 5, 4.1, 5.5)
> ymedio <- 5.2625
> sy <- 0.7069805
> ny <- 8
> mu0 <- 1.2
> Sp < - sqrt(((nx - 1) * sx^2 + (ny - 1) * sy^2)/(nx + ny - 2))
[1] 0.5767614
> t <- (xmedio - ymedio - mu0)/(Sp * sqrt(1/nx + 1/ny))
[1] 2.073455
> res <- t.test(x, y, mu = 1.2, alternative = "two.sided", conf.level = 0.95,
+ var.equal = TRUE)
> res$statistic
2.073455
> parameter <- nx + ny - 2</pre>
> parameter
[1] 17
> res$parameter
df
17
> p.value <-2 * pt(-abs(t), df = nx + ny - 2)
> p.value
[1] 0.05364035
> res$p.value
[1] 0.05364043
```

```
> alpha <- 0.05
 > lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = nx + ny - 2) *
       Sp * sqrt(1/nx + 1/ny)
 > upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = nx + ny - 2) \star
       Sp * sqrt(1/nx + 1/ny)
 > c(lower, upper)
 [1] 1.190256 2.321108
 > res$conf.int
 [1] 1.190255 2.321108
 attr(,"conf.level")
 [1] 0.95
 > c(xmedio, ymedio)
 [1] 7.018182 5.262500
 > res$estimate
 mean of x mean of y
  7.018182 5.262500
 > mu0
 [1] 1.2
 > res$null.value
 difference in means
                  1.2
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < -c(154, 109, 137, 115, 140)
 > xmedio <- 131
 > sx <- 18.61451
 > nx < -5
 > y <- c(108, 115, 126, 92, 146)
 > ymedio <- 117.4
 > sy <- 20.19406
 > ny <- 5
 > mu0 <- 10
 > Sp <- sqrt(((nx - 1) * sx^2 + (ny - 1) * sy^2)/(nx + ny - 2))
 > Sp
 [1] 19.42035
 > t <- (xmedio - ymedio - mu0)/(Sp * sqrt(1/nx + 1/ny))
 > t
 [1] 0.2930997
```

```
> res <- t.test(x, y, mu = 10, alternative = "two.sided", conf.level = 0.95,
+ var.equal = TRUE)
> res$statistic
0.2930998
> parameter <- nx + ny - 2
> parameter
[1] 8
> res$parameter
df
 8
> p.value < 2 * pt(-abs(t), df = nx + ny - 2)
> p.value
[1] 0.7769049
> res$p.value
[1] 0.7769049
> alpha <- 0.05</pre>
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = nx + ny - 2) *
     Sp * sqrt(1/nx + 1/ny)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = nx + ny - 2) *
      Sp * sqrt(1/nx + 1/ny)
> c(lower, upper)
[1] -14.72351 41.92351
> res$conf.int
[1] -14.72351 41.92351
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean\ of\ x\ mean\ of\ y
   131.0 117.4
> mu0
[1] 10
> res$null.value
difference in means
                 10
> res$alternative
[1] "two.sided"
```

Test di Student con due campioni indipendenti con varianze non note e supposte diverse

• Package: stats

• Sintassi: t.test()

• Input:

x vettore numerico di dimensione n_x y vettore numerico di dimensione n_y mu valore di $(\mu_x - \mu_y)_{|H_0}$ alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza $1-\alpha$

• Output:

statistic valore empirico della statistica t parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello $1-\alpha$ estimate medie campionarie null.value valore di $(\mu_x-\mu_y)_{|H_0}$ alternative ipotesi alternativa

• Formula:

statistic

$$t = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{\sqrt{s_x^2 / n_x + s_y^2 / n_y}}$$

parameter

$$df = \frac{\left(s_x^2 / n_x + s_y^2 / n_y\right)^2}{s_x^4 / \left(n_x^2 \left(n_x - 1\right)\right) + s_y^4 / \left(n_y^2 \left(n_y - 1\right)\right)} = \left(\frac{1}{n_x - 1}C^2 + \frac{1}{n_y - 1}(1 - C)^2\right)^{-1}$$

dove
$$C = \frac{s_x^2 / n_x}{s_x^2 / n_x + s_y^2 / n_y}$$

p.value

alternative	less	greater	two.sided
p.value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \le - t)$

conf.int

$$\bar{x} - \bar{y} \mp t_{1-\alpha/2, df} \sqrt{s_x^2/n_x + s_y^2/n_y}$$

estimate

$$\bar{x}$$
 i

null.value

$$(\mu_x - \mu_y)_{\mid H_0}$$

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> xmedio <- 7.018182
> sx <- 0.4643666
> nx <- 11
> y <- c(4.5, 5.4, 6.1, 6.1, 5.4, 5, 4.1, 5.5)
> ymedio <- 5.2625
> sy <- 0.7069805
> ny <- 8
> mu0 <- 1.2
> t <- (xmedio - ymedio - mu0)/sqrt(sx^2/nx + sy^2/ny)
> t
```

```
[1] 1.939568
> res <- t.test(x, y, mu = 1.2, alternative = "two.sided", conf.level = 0.95)</pre>
> res$statistic
1.939568
> g1 < (sx^2/nx + sy^2/ny)^2/(sx^4/(nx^2 * (nx - 1)) + sy^4/(ny^2 * (nx - 1)) + sy^4/(nx^2 * (nx - 1)) + sy^4/(nx^2 * (nx - 1)) + sy^4/(nx^2 * (
+ (ny - 1)))
> gl
[1] 11.30292
> C <- (sx^2/nx)/(sx^2/nx + sy^2/ny)
> gl <- as.numeric(solve(solve(nx - 1) * C^2 + solve(ny - 1) *
                 (1 - C)^2)
> al
[1] 11.30292
> res$parameter
                     df
11.30292
> p.value < 2 * pt(-abs(t), df = gl)
> p.value
[1] 0.0777921
> res$p.value
[1] 0.07779219
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
                   sy^2/ny)
> c(lower, upper)
[1] 1.127160 2.384204
> res$conf.int
[1] 1.127160 2.384203
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 7.018182 5.262500
> res$estimate
\hbox{mean of } x \hbox{ mean of } y
   7.018182 5.262500
```

```
> mu0
     [1] 1.2
     > res$null.value
     difference in means
                                                               1.2
     > res$alternative
     [1] "two.sided"
• Example 2:
     > x < -c(154, 109, 137, 115, 140)
     > xmedio <- 131
     > sx <- 18.61451
     > nx <- 5
     > y < -c(108, 115, 126, 92, 146)
     > ymedio <- 117.4
     > sy <- 20.19406
     > ny < -5
     > mu0 <- 10
     > t <- (xmedio - ymedio - mu0)/sqrt(sx^2/nx + sy^2/ny)
     > t
     [1] 0.2930997
     > res <- t.test(x, y, mu = 10, alternative = "two.sided", conf.level = 0.95)
     > res$statistic
     0.2930998
     > g1 < (sx^2/nx + sy^2/ny)^2/(sx^4/(nx^2 * (nx - 1)) + sy^4/(ny^2 * (nx - 1)) + sy^4/(nx^2 * (nx - 1)) + sy^4/(nx^2 * (nx - 1)) + sy^4/(nx^2 * (
                       (ny - 1))
     > q1
     [1] 7.947511
     > C <- (sx^2/nx)/(sx^2/nx + sy^2/ny)
     > gl <- as.numeric(solve(solve(nx - 1) * C^2 + solve(ny - 1) *
                          (1 - C)^2)
     > gl
     [1] 7.947511
     > res$parameter
     7.947512
     > p.value < 2 * pt(-abs(t), df = gl)
     > p.value
     [1] 0.7769531
     > res$p.value
```

```
[1] 0.7769531
> alpha <- 0.05
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
     sy^2/ny)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
     sy^2/ny)
> c(lower, upper)
[1] -14.75611 41.95611
> res$conf.int
[1] -14.75611 41.95611
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean of x mean of y
          117.4
   131.0
> mu0
[1] 10
> res$null.value
difference in means
> res$alternative
[1] "two.sided"
```

Test di Student per dati appaiati

parameter gradi di libertà

```
• Package: stats
• Sintassi: t.test()
• Input:

x vettore numerico di dimensione n
y vettore numerico di dimensione n
mu valore di (\mu_x - \mu_y)_{|H_0}
alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
paired = TRUE
• Output:
statistic valore empirico della statistica t
```

```
p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello 1-\alpha estimate differenza tra le medie campionarie null.value valore di (\mu_x-\mu_y)_{|H_0} alternative ipotesi alternativa
```

• Formula:

statistic $t = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{s_{x-y} / \sqrt{n}}$ $\text{dove } s_{x-y}^2 = \frac{1}{n-1} \sum_{i=1}^n \left((x_i - y_i) - (\bar{x} - \bar{y}) \right)^2 = s_x^2 + s_y^2 - 2 \, s_{xy}$ parameter df = n-1 p.value $\boxed{ \text{alternative less greater two.sided} } \boxed{ \text{p.value } P(t_{df} \leq t) \mid 1 - P(t_{df} \leq t) \mid 2 \, P(t_{df} \leq - |t|) }$

$$ar{x}-ar{y}\mp t_{1-lpha/2,\,df}\,s_{x-y}\,/\,\sqrt{n}$$
 $ar{x}-ar{y}$

null.value

conf.int

estimate

$$(\mu_x - \mu_y)_{\mid H_0}$$

• Example 1:

7

```
> x < -c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1)
> xmedio <- 7.0125
> y < -c(4.5, 5.4, 6.1, 6.1, 5.4, 5, 4.1, 5.5)
> ymedio <- 5.2625</pre>
> n < - 8
> mu0 <- 1.2
> t <- (xmedio - ymedio - mu0)/(sd(x - y)/sqrt(n))
[1] 1.815412
> res <- t.test(x, y, mu = 1.2, alternative = "two.sided", conf.level = 0.95,
    paired = TRUE)
> res$statistic
1.815412
> parameter <- n - 1
> parameter
[1] 7
> res$parameter
df
```

```
> p.value < 2 * pt(-abs(t), df = n - 1)
 > p.value
 [1] 0.1123210
 > res$p.value
 [1] 0.1123210
 > alpha <- 0.05</pre>
 > lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = n - 1) * sd(x -
       y)/sqrt(n)
 > upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = n - 1) * sd(x - 1)
       y)/sqrt(n)
 > c(lower, upper)
 [1] 1.033610 2.466390
 > res$conf.int
 [1] 1.033610 2.466390
 attr(,"conf.level")
 [1] 0.95
 > xmedio - ymedio
 [1] 1.75
 > res$estimate
 mean of the differences
                     1.75
 > mu0
 [1] 1.2
 > res$null.value
 difference in means
                  1.2
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x <- c(154, 109, 137, 115, 140)
 > xmedio <- 131
 > y <- c(108, 115, 126, 92, 146)
 > ymedio <- 117.4
 > n < -5
 > mu0 <- 10
 > t <- (xmedio - ymedio - mu0)/(sd(x - y)/sqrt(n))
 [1] 0.3680758
```

```
> res <- t.test(x, y, mu = 10, alternative = "two.sided", conf.level = 0.95,
+ paired = TRUE)
> res$statistic
0.3680758
> parameter <- n - 1
> parameter
[1] 4
> res$parameter
df
 4
> p.value <- 2 * pt(-abs(t), df = n - 1)
> p.value
[1] 0.7314674
> res$p.value
[1] 0.7314674
> alpha <- 0.05</pre>
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = n - 1) * sd(x -
     y)/sqrt(n)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = n - 1) * sd(x -
      y)/sqrt(n)
> c(lower, upper)
[1] -13.55528 40.75528
> res$conf.int
[1] -13.55528 40.75528
attr(,"conf.level")
[1] 0.95
> xmedio - ymedio
[1] 13.6
> res$estimate
mean of the differences
                   13.6
> mu0
[1] 10
> res$null.value
difference in means
                 10
> res$alternative
[1] "two.sided"
```

Test di Fisher con k campioni indipendenti

```
• Package: stats
```

• Sintassi: oneway.test()

• Input:

formula modello di regressione lineare con una variabile esplicativa fattore f a k livelli ed n unità var.equal = TRUE

• Output:

```
statistic valore empirico della statistica F parameter gradi di libertà p.value p-value
```

• Formula:

statistic $Fvalue \,=\, \frac{\left[\,\sum_{j=1}^k\,n_j\,(\bar{y}_j-\bar{y})^2\,\right]/\left(k-1\right)}{\left[\,\sum_{j=1}^k\,\sum_{i=1}^{n_j}\,(y_{ij}-\bar{y}_j)^2\,\right]/\left(n-k\right)}$ parameter

f	k-1
Residuals	n-k

p.value

$$P(F_{k-1, n-k} \ge Fvalue)$$

• Examples:

6.2 Test di ipotesi sulla media con uno o due campioni (summarized data)

Test Z con un campione

```
    Package: BSDA
    Sintassi: zsum.test()
    Input:
        mean.x valore di x̄
        sigma.x valore di σ<sub>x</sub>
```

```
n.x valore di n mu valore di \mu_0 alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica Z p.value p-value conf.int intervallo di confidenza per la media incognita a livello 1-\alpha estimate media campionaria null.value valore di \mu_0 alternative ipotesi alternativa
```

• Formula:

statistic $z = \frac{\bar{x} - \mu_0}{\sigma_x \, / \, \sqrt{n}} \label{eq:z}$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z))$

conf.int $\bar{x}\mp z_{1-\alpha/2}\,\sigma_x\,/\,\sqrt{n}$ estimate \bar{x} null.value μ_0

```
> xmedio <- 7.018182
> sigmax <- 1.2
> n <- 11
> mu0 <- 6.5
> z <- (xmedio - mu0)/(sigmax/sqrt(n))
> z

[1] 1.432179

> res <- zsum.test(mean.x = 7.018182, sigma.x = 1.2, n.x = 11,
+ mu = 6.5, alternative = "two.sided", conf.level = 0.95)
> res$statistic

2
1.432179

> p.value <- 2 * pnorm(-abs(z))
> p.value

[1] 0.1520925

> res$p.value

[1] 0.1520925
```

```
> alpha <- 0.05</pre>
 > lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
 > upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
 > c(lower, upper)
 [1] 6.309040 7.727324
 > res$conf.int
 [1] 6.309040 7.727324
 attr(,"conf.level")
 [1] 0.95
 > xmedio
 [1] 7.018182
 > res$estimate
 mean\ of\ x
  7.018182
 > mu0
 [1] 6.5
 > res$null.value
 mean
  6.5
 > res$alternative
 [1] "two.sided"
• Example 2:
 > xmedio <- 4.68
 > sigmax <- 1.45</pre>
 > n <- 5
 > mu0 < - 5.2
 > z <- (xmedio - mu0)/(sigmax/sqrt(n))</pre>
 > z
 [1] -0.8019002
 > res <- zsum.test(mean.x = 4.68, sigma.x = 1.45, n.x = 5, mu = 5.2,
 + alternative = "two.sided", conf.level = 0.95)
 > res$statistic
 -0.8019002
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.4226107
 > res$p.value
```

```
[1] 0.4226107
> alpha <- 0.05
> lower <- xmedio - qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> upper <- xmedio + qnorm(1 - 0.05/2) * sigmax/sqrt(n)
> c(lower, upper)
[1] 3.409042 5.950958
> res$conf.int
[1] 3.409042 5.950958
attr(,"conf.level")
[1] 0.95
> xmedio
[1] 4.68
> res$estimate
mean\ of\ x
     4.68
> mu0
[1] 5.2
> res$null.value
mean
5.2
> res$alternative
[1] "two.sided"
```

Test di Student con un campione

parameter gradi di libertà

p.value p-value

```
Package: BSDA
Sintassi: tsum.test()
Input:

mean.x valore di x

s.x valore di s<sub>x</sub>

n.x valore di n

mu valore di μ<sub>0</sub>

alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1 – α
Output:

statistic valore empirico della statistica t
```

```
conf.int intervallo di confidenza per la media incognita a livello 1-\alpha estimate media campionaria null.value valore di \mu_0 alternative ipotesi alternativa
```

• Formula:

statistic $t = \frac{\bar{x} - \mu_0}{s_x / \sqrt{n}}$ parameter df = n - 1 p.value

conf.int $\bar{x} \mp t_{1-\alpha/2,\,df}\,s_x\,/\,\sqrt{n}$ estimate \bar{x} null.value μ_0

```
> xmedio <- 7.018182</pre>
> sx <- 1.2
> n <- 11
> mu0 < - 6.5
> t <- (xmedio - mu0)/(sx/sqrt(n))
> t
[1] 1.432179
> res <- tsum.test(mean.x = 7.018182, s.x = 1.2, n.x = 11, mu = 6.5,
     alternative = "two.sided", conf.level = 0.95)
> res$statistic
1.432179
> parameter <- n - 1
> parameter
[1] 10
> res$parameter
df
10
> p.value < 2 * pt(-abs(t), df = n - 1)
> p.value
[1] 0.1826001
> res$p.value
```

```
[1] 0.1826001
 > alpha <- 0.05</pre>
 > lower <- xmedio - qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
 > upper <- xmedio + qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
 > c(lower, upper)
 [1] 6.212011 7.824353
 > res$conf.int
 [1] 6.212011 7.824353
 attr(,"conf.level")
 [1] 0.95
 > xmedio
 [1] 7.018182
 > res$estimate
 mean of x
  7.018182
 > mu0
 [1] 6.5
 > res$null.value
 mean
  6.5
 > res$alternative
 [1] "two.sided"
• Example 2:
 > xmedio <- 4.68
 > sx <- 1.45
 > n <- 5
 > mu0 < -5.2
 > t <- (xmedio - mu0)/(sx/sqrt(n))
 [1] -0.8019002
 > res <- tsum.test(mean.x = 4.68, s.x = 1.45, n.x = 5, mu = 5.2,
 + alternative = "two.sided", conf.level = 0.95)
 > res$statistic
 -0.8019002
 > parameter <- n - 1
 > parameter
 [1] 4
```

```
> res$parameter
df
 4
> p.value < 2 * pt(-abs(t), df = n - 1)
> p.value
[1] 0.4675446
> res$p.value
[1] 0.4675446
> alpha <- 0.05
> lower <- xmedio - qt(1 - 0.05/2, df = n - 1) * sx/sqrt(n)
> upper <- xmedio + qt(1 - 0.05/2, df = n - 1) \star sx/sqrt(n)
> c(lower, upper)
[1] 2.879587 6.480413
> res$conf.int
[1] 2.879587 6.480413
attr(, "conf.level")
[1] 0.95
> xmedio
[1] 4.68
> res$estimate
mean of x
     4.68
> mu0
[1] 5.2
> res$null.value
mean
 5.2
> res$alternative
[1] "two.sided"
```

Test Z con due campioni indipendenti

• Package: BSDA

• Sintassi: zsum.test()

• Input:

```
mean.x valore di \bar{x} sigma.x valore di \sigma_x n.x valore di n_x mean.y valore di \bar{y} sigma.y valore di \sigma_y n.y valore di n_y mu valore di (\mu_x - \mu_y)_{|H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica Z p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello 1-\alpha estimate medie campionarie null.value valore di (\mu_x-\mu_y)_{|H_0} alternative ipotesi alternativa
```

• Formula:

statistic

$$z = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{\sqrt{\sigma_x^2 / n_x + \sigma_y^2 / n_y}}$$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z)$

```
conf.int \bar{x}-\bar{y}\mp z_{1-\alpha/2}\sqrt{{\sigma_x^2/n_x+\sigma_y^2/n_y}} estimate \bar{x}=\bar{y} null.value (\mu_x-\mu_y)_{|\,H_0}
```

```
> xmedio <- 131
> sigmax <- 15.5
> nx <- 5
> ymedio <- 117.4
> sigmay <- 13.5
> ny <- 5
> mu0 <- 10
> z <- (xmedio - ymedio - mu0)/sqrt(sigmax^2/nx + sigmay^2/ny)
> z
[1] 0.3916284
```

```
> res <- zsum.test(mean.x = 131, sigma.x = 15.5, n.x = 5, mean.y = 117.4,
      sigma.y = 13.5, n.y = 5, mu = 10, alternative = "two.sided",
      conf.level = 0.95)
> res$statistic
0.3916284
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 0.6953328
> res$p.value
[1] 0.6953328
> alpha <- 0.05</pre>
> lower <- xmedio - ymedio - qnorm(1 - 0.05/2) \star sqrt(sigmax^2/nx +
     sigmay^2/ny)
> upper <- xmedio - ymedio + qnorm(1 - 0.05/2) \star sqrt(sigmax^2/nx +
     sigmay^2/ny)
> c(lower, upper)
[1] -4.41675 31.61675
> res$conf.int
[1] -4.41675 31.61675
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean of x mean of y
    131.0 117.4
> mu0
[1] 10
> res$null.value
difference in means
                 10
> res$alternative
[1] "two.sided"
```

```
> xmedio <- 7.018182</pre>
> sigmax <- 0.5
> nx <- 11
> ymedio <- 5.2625</pre>
> sigmay <- 0.8
> ny <- 8
> mu0 < -1.2
> z <- (xmedio - ymedio - mu0)/sqrt(sigmax^2/nx + sigmay^2/ny)
[1] 1.733738
> res <- zsum.test(mean.x = 7.018182, sigma.x = 0.5, n.x = 11,
+ mean.y = 5.2625, sigma.y = 0.8, n.y = 8, mu = 1.2, alternative = "two.sided",
     conf.level = 0.95)
> res$statistic
1.733738
> p.value <- 2 * pnorm(-abs(z))</pre>
> p.value
[1] 0.0829646
> res$p.value
[1] 0.0829646
> alpha <- 0.05</pre>
> lower <- xmedio - ymedio - qnorm(1 - 0.05/2) \star sqrt(sigmax^2/nx +
     sigmay^2/ny)
> upper <- xmedio - ymedio + qnorm(1 - 0.05/2) \star sqrt(sigmax^2/nx +
     sigmay^2/ny)
> c(lower, upper)
[1] 1.127492 2.383872
> res$conf.int
[1] 1.127492 2.383872
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 7.018182 5.262500
> res$estimate
mean of x mean of y
7.018182 5.262500
> mu0
[1] 1.2
> res$null.value
```

Test di Student con due campioni indipendenti con varianze non note e supposte uguali

```
• Package: BSDA
```

• Sintassi: tsum.test()

• Input:

```
mean.x valore di \bar{x} s.x valore di s_x n.x valore di n_x mean.y valore di \bar{y} s.y valore di s_y n.y valore di n_y mu valore di (\mu_x - \mu_y)_{|H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha var.equal = TRUE
```

• Output:

```
statistic valore empirico della statistica t parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello 1-\alpha estimate medie campionarie null.value valore di (\mu_x-\mu_y)_{|H_0} alternative ipotesi alternativa
```

• Formula:

statistic

$$t = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{s_P \sqrt{1/n_x + 1/n_y}}$$

dove
$$s_P^2 = \frac{(n_x - 1) s_x^2 + (n_y - 1) s_y^2}{n_x + n_y - 2}$$

parameter

$$df = n_x + n_y - 2$$

p.value

alt	ernative	less	greater	two.sided
r	value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \le - t)$

conf.int

$$\bar{x} - \bar{y} \mp t_{1-\alpha/2, df} s_P \sqrt{1/n_x + 1/n_y}$$

```
estimate \bar{x} = \bar{y} null.value (\left. \mu_x - \mu_y \right.)_{\mid H_0}
```

```
> xmedio <- 7.018182</pre>
> sx < -0.5
> nx <- 11
> ymedio <- 5.2625
> sy <- 0.8
> ny <- 8
> mu0 <- 1.2
> Sp <- sqrt(((nx - 1) * sx^2 + (ny - 1) * sy^2)/(nx + ny - 2))
> Sp
[1] 0.6407716
> t <- (xmedio - ymedio - mu0)/(Sp * sqrt(1/nx + 1/ny))
> res <- tsum.test(mean.x = 7.018182, s.x = 0.5, n.x = 11, mean.y = 5.2625,
     s.y = 0.8, n.y = 8, mu0 < -1.2, alternative = "two.sided",
     conf.level = 0.95)
> res$statistic
1.866326
> parameter <- nx + ny - 2
> parameter
[1] 17
> res$parameter
df
17
> p.value <-2 * pt(-abs(t), df = nx + ny - 2)
> p.value
[1] 0.07934364
> res$p.value
[1] 0.07934364
> alpha <- 0.05
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = nx + ny - 2) *
     Sp * sqrt(1/nx + 1/ny)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = nx + ny - 2) *
      Sp * sqrt(1/nx + 1/ny)
> c(lower, upper)
[1] 1.127503 2.383861
> res$conf.int
[1] 1.127503 2.383861
attr(,"conf.level")
[1] 0.95
```

```
> c(xmedio, ymedio)
 [1] 7.018182 5.262500
 > res$estimate
 mean of x mean of y
  7.018182 5.262500
 > mu0
 [1] 1.2
 > res$null.value
 difference in means
                 1.2
 > res$alternative
 [1] "two.sided"
• Example 2:
 > xmedio <- 131
 > sx <- 15.5
 > nx <- 5
 > ymedio <- 117.4
 > sy <- 13.5
 > ny <- 5
 > mu0 <- 10
 > Sp <- sqrt(((nx - 1) * sx^2 + (ny - 1) * sy^2)/(nx + ny - 2))
 > Sp
 [1] 14.53444
 > t <- (xmedio - ymedio - mu0)/(Sp * sqrt(1/nx + 1/ny))
 [1] 0.3916284
 > res <- tsum.test(mean.x = 131, s.x = 15.5, n.x = 5, mean.y = 117.4,
       s.y = 13.5, n.y = 5, mu = 10, alternative = "two.sided",
       conf.level = 0.95, var.equal = TRUE)
 > res$statistic
 0.3916284
 > parameter <- nx + ny - 2
 > parameter
 [1] 8
 > res$parameter
 df
  8
```

```
> p.value < 2 * pt(-abs(t), df = nx + ny - 2)
> p.value
[1] 0.705558
> res$p.value
[1] 0.705558
> alpha <- 0.05</pre>
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = nx + ny - 2) *
     Sp * sqrt(1/nx + 1/ny)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = nx + ny - 2) \star
      Sp * sqrt(1/nx + 1/ny)
> c(lower, upper)
[1] -7.597685 34.797685
> res$conf.int
[1] -7.597685 34.797685
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean of x mean of y
   131.0 117.4
> mu0
[1] 10
> res$null.value
difference in means
                 10
> res$alternative
[1] "two.sided"
```

Test di Student con due campioni indipendenti con varianze non note e supposte diverse

• Package: BSDA

• **Sintassi:** tsum.test()

• Input:

```
mean.x valore di \bar{x} s.x valore di s_x n.x valore di n_x mean.y valore di \bar{y} s.y valore di s_y n.y valore di n_y mu valore di (\mu_x - \mu_y)_{|H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha var.equal = FALSE
```

• Output:

statistic valore empirico della statistica t parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la differenza tra le medie incognite a livello $1-\alpha$ estimate medie campionarie null.value valore di $(\mu_x-\mu_y)_{|H_0}$ alternative ipotesi alternativa

• Formula:

statistic

$$t = \frac{(\bar{x} - \bar{y}) - (\mu_x - \mu_y)_{|H_0}}{\sqrt{s_x^2 / n_x + s_y^2 / n_y}}$$

parameter

$$df = \frac{\left(s_x^2 / n_x + s_y^2 / n_y\right)^2}{s_x^4 / \left(n_x^2 \left(n_x - 1\right)\right) + s_y^4 / \left(n_y^2 \left(n_y - 1\right)\right)} = \left(\frac{1}{n_x - 1}C^2 + \frac{1}{n_y - 1}(1 - C)^2\right)^{-1}$$

dove
$$C = \frac{s_x^2 / n_x}{s_x^2 / n_x + s_y^2 / n_y}$$

p.value

alternative	less	greater	two.sided
p.value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \le - t)$

conf.int

$$\bar{x} - \bar{y} \mp t_{1-\alpha/2, df} \sqrt{s_x^2/n_x + s_y^2/n_y}$$

estimate

$$\bar{x}$$
 \bar{y}

null.value

$$(\mu_x - \mu_y)_{\mid H_0}$$

```
> xmedio <- 7.018182</pre>
> sx <- 0.5
> nx <- 11
> ymedio <- 5.2625</pre>
> sy <- 0.8
> ny <- 8
> mu0 < -1.2
> t <- (xmedio - ymedio - mu0)/sqrt(sx^2/nx + sy^2/ny)
> t
[1] 1.733738
> res <- tsum.test(mean.x = 7.018182, s.x = 0.5, n.x = 11, mean.y = 5.2625,
+ s.y = 0.8, n.y = 8, mu = 1.2, alternative = "two.sided",
                conf.level = 0.95, var.equal = FALSE)
> res$statistic
                  t
1.733738
> g1 < (sx^2/nx + sy^2/ny)^2/(sx^4/(nx^2 * (nx - 1)) + sy^4/(ny^2 * (nx - 1)) + sy^4/(nx^2 * (
               (ny - 1))
> gl
[1] 10.92501
> C <- (sx^2/nx)/(sx^2/nx + sy^2/ny)
> gl <- as.numeric(solve(solve(nx - 1) * C^2 + solve(ny - 1) *
              (1 - C)^2)
> q1
[1] 10.92501
> res$parameter
                  df
10.92501
> p.value < 2 * pt(-abs(t), df = gl)
> p.value
[1] 0.1110536
> res$p.value
[1] 0.1110536
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
                  sy^2/ny)
> c(lower, upper)
[1] 1.049651 2.461713
> res$conf.int
[1] 1.049651 2.461713
attr(,"conf.level")
[1] 0.95
```

```
> c(xmedio, ymedio)
 [1] 7.018182 5.262500
 > res$estimate
 mean of x mean of y
  7.018182 5.262500
 > mu0
 [1] 1.2
 > res$null.value
 difference in means
                  1.2
 > res$alternative
 [1] "two.sided"
• Example 2:
 > xmedio <- 131
 > sx < -15.5
 > nx <- 5
 > ymedio <- 117.4
 > sy <- 13.5
 > ny <- 5
 > mu0 <- 10
 > t <- (xmedio - ymedio - mu0)/sqrt(sx^2/nx + sy^2/ny)
 > t
 [1] 0.3916284
 > res <- tsum.test(mean.x = 131, s.x = 15.5, n.x = 5, mean.y = 117.4,
 + s.y = 13.5, n.y = 5, mu = 10, alternative = "two.sided",
       conf.level = 0.95, var.equal = FALSE)
 > res$statistic
 0.3916284
 > g1 < (sx^2/nx + sy^2/ny)^2/(sx^4/(nx^2 * (nx - 1)) + sy^4/(ny^2 * (nx^2 + nx^2)) + sy^4/(ny^2 * (nx^2 + nx^2))
       (ny - 1))
 > gl
 [1] 7.852026
 > C <- (sx^2/nx)/(sx^2/nx + sy^2/ny)
 > gl <- as.numeric(solve(solve(nx - 1) * C^2 + solve(ny - 1) *
      (1 - C)^2)
 > gl
 [1] 7.852026
 > res$parameter
```

```
df
7.852026
> p.value < 2 * pt(-abs(t), df = gl)
> p.value
[1] 0.7057463
> res$p.value
[1] 0.7057463
> lower <- (xmedio - ymedio) - qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx +
     sy^2/ny)
> upper <- (xmedio - ymedio) + qt(1 - 0.05/2, df = gl) * sqrt(sx^2/nx + c
+ sy^2/ny
> c(lower, upper)
[1] -7.667421 34.867421
> res$conf.int
[1] -7.667421 34.867421
attr(,"conf.level")
[1] 0.95
> c(xmedio, ymedio)
[1] 131.0 117.4
> res$estimate
mean of x mean of y
   131.0 117.4
> mu0
[1] 10
> res$null.value
difference in means
> res$alternative
[1] "two.sided"
```

6.3 Test di ipotesi sulla varianza con uno o due campioni

Test Chi-Quadrato con un campione

```
Package: sigma2toolsSintassi: sigma2.test()
```

• Input:

```
x vettore numerico di dimensione n var0 valore di \sigma_0^2 alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha
```

• Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la media incognita a livello 1-\alpha estimate varianza campionaria null.value valore di \sigma_0^2 alternative ipotesi alternativa
```

• Formula:

statistic
$$c \, = \, \frac{(n-1) \, s_x^2}{\sigma_0^2} \label{eq:constraint}$$
 parameter
$$d f \, = \, n-1 \label{eq:constraint}$$
 p.value

alternative	less	greater	two.sided
p.value	$P(\chi_{df}^2 \le c)$	$P(\chi_{df}^2 \ge c)$	$2 \min \left(P(\chi_{df}^2 \le c), P(\chi_{df}^2 \ge c) \right)$

```
conf.int \frac{(n-1)\,s_x^2}{\chi_{1-\alpha/2,\,df}^2} - \frac{(n-1)\,s_x^2}{\chi_{\alpha/2,\,df}^2} estimate s_x^2 null.value \sigma_0^2
```

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> sx <- 0.4643666
> n <- 11
> var0 <- 0.5
> c <- (n - 1) * sx^2/var0
> c

[1] 4.312727

> res <- sigma2.test(x, var0 = 0.5, alternative = "two.sided", conf.level = 0.95)
> res$statistic
```

```
X-squared
 4.312727
> parameter <- n - 1
> parameter
[1] 10
> res$parameter
df
10
> p.value < 2 * min(pchisq(c, df = n - 1), 1 - pchisq(c, df = n -
+ 1))
> p.value
[1] 0.1357228
> res$p.value
[1] 0.1357229
> alpha <- 0.05</pre>
> lower <- (n - 1) * sx^2/qchisq(1 - alpha/2, df = n - 1)
> upper <- (n - 1) * sx^2/qchisq(alpha/2, df = n - 1)
> c(lower, upper)
[1] 0.1052748 0.6641150
> res$conf.int
[1] 0.1052749 0.6641151
attr(,"conf.level")
[1] 0.95
> sx^2
[1] 0.2156363
> res$estimate
var of x
0.2156364
> var0
[1] 0.5
> res$null.value
variance
     0.5
> res$alternative
[1] "two.sided"
```

```
> x < -c(1, 2.3, 4.5, 6.7, 8.9)
> sx <- 3.206556
> n <- 5
> var0 <- 12
> c <- (n - 1) * sx^2/var0
[1] 3.427334
> res <- sigma2.test(x, var0 = 12, alternative = "two.sided", conf.level = 0.95)
> res$statistic
X-squared
3.427333
> parameter <- n - 1
> parameter
[1] 4
> res$parameter
df
 4
> p.value < 2 * min(pchisq(c, df = n - 1), 1 - pchisq(c, df = n -
     1))
> p.value
[1] 0.9780261
> res$p.value
[1] 0.9780263
> alpha <- 0.05
> lower <- (n - 1) * sx^2/qchisq(1 - alpha/2, df = n - 1)
> upper <- (n - 1) * sx^2/qchisq(alpha/2, df = n - 1)
> c(lower, upper)
[1] 3.690833 84.901796
> res$conf.int
[1] 3.690832 84.901785
attr(,"conf.level")
[1] 0.95
> sx^2
[1] 10.28200
> res$estimate
var of x
  10.282
```

Test di Fisher con due campioni

• Package: stats

• Sintassi: var.test()

• Input:

x vettore numerico di dimensione n_x y vettore numerico di dimensione n_y ratio il valore di $\frac{\sigma_x^2}{\sigma_y^2} \left| H_0 \right|$ alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza $1-\alpha$

• Output:

statistic valore empirico della statistica F parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per il rapporto tra le varianze incognite al livello $1-\alpha$ estimate rapporto tra le varianze campionarie null.value valore di $\frac{\sigma_x^2}{\sigma_y^2} \left| H_0 \right|$ alternative ipotesi alternativa

• Formula:

statistic

$$Fval = \frac{s_x^2}{s_y^2} \frac{1}{\frac{\sigma_x^2}{\sigma_y^2} \mid H_0}$$

parameter

$$df_1 = n_x - 1 \qquad df_2 = n_y - 1$$

p.value

alternative	less	greater	two.sided
p.value	$P(F_{df_1,df_2} \leq Fval)$	$P(F_{df_1,df_2} \geq Fval)$	$2 \min (P(F_{df_1,df_2} \leq Fval), P(F_{df_1,df_2} \geq Fval))$

conf.int $\frac{1}{F_{1-\frac{\alpha}{2},df_1,df_2}}\frac{s_x^2}{s_y^2} - \frac{1}{F_{\frac{\alpha}{2},df_1,df_2}}\frac{s_x^2}{s_y^2}$ estimate $\frac{s_x^2}{s_y^2}$

null.value

 $\left| \frac{\sigma_x^2}{\sigma_y^2} \right| H_0$

```
> x < -c(7, -4, 18, 17, -3, -5, 1, 10, 11, -2, -3)
> nx <- 11
> y < -c(-1, 12, -1, -3, 3, -5, 5, 2, -11, -1, -3)
> ny <- 11
> ratio <- 1.3
> Fval <- sd(x)^2/sd(y)^2 * (1/ratio)
> Fval
[1] 1.648524
> res <- var.test(x, y, ratio = 1.3, alternative = "two.sided",
+ conf.level = 0.95)
> res$statistic
1.648524
> c(nx - 1, ny - 1)
[1] 10 10
> res$parameter
  num df denom df
     10
          10
> p.value <-2 * min(pf(Fval, df1 = nx - 1, df2 = ny - 1), 1 -
     pf(Fval, df1 = nx - 1, df2 = ny - 1))
> p.value
[1] 0.4430561
> res$p.value
[1] 0.4430561
> alpha <- 0.05</pre>
> lower <- (1/qf(1 - 0.05/2, df1 = nx - 1, df2 = ny - 1)) * sd(x)^2/sd(y)^2
> upper <- (1/qf(0.05/2, df1 = nx - 1, df2 = ny - 1)) * sd(x)^2/sd(y)^2
> c(lower, upper)
[1] 0.5765943 7.9653858
> res$conf.int
[1] 0.5765943 7.9653858
attr(,"conf.level")
[1] 0.95
> sd(x)^2/sd(y)^2
[1] 2.143081
> res$estimate
```

```
ratio of variances
           2.143081
 > ratio
 [1] 1.3
 > res$null.value
 ratio of variances
                1.3
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < -c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
 > nx <- 11
 > y <- c(4.5, 5.4, 6.1, 6.1, 5.4, 5, 4.1, 5.5)
 > ny <- 8
 > ratio <- 1.1
 > Fval <- sd(x)^2/sd(y)^2 * (1/ratio)
 > Fval
 [1] 0.3922062
 > res <- var.test(x, y, ratio = 1.1, alternative = "two.sided",</pre>
      conf.level = 0.95)
 > res$statistic
 0.3922062
 > c(nx - 1, ny - 1)
 [1] 10 7
 > res$parameter
   num df denom df
       10
 > p.value < 2 * min(pf(Fval, df1 = nx - 1, df2 = ny - 1), 1 -
 + pf(Fval, df1 = nx - 1, df2 = ny - 1))
 > p.value
 [1] 0.1744655
 > res$p.value
 [1] 0.1744655
 > alpha <- 0.05</pre>
 > lower <- (1/qf(1 - 0.05/2, df1 = nx - 1, df2 = ny - 1)) * sd(x)^2/sd(y)^2
 > upper <- (1/qf(0.05/2, df1 = nx - 1, df2 = ny - 1)) * sd(x)^2/sd(y)^2
 > c(lower, upper)
```

[1] 0.09061463 1.70405999

```
> res$conf.int
    [1] 0.09061463 1.70405999
    attr(, "conf.level")
    [1] 0.95
    > sd(x)^2/sd(y)^2
    [1] 0.4314268
    > res$estimate
    ratio of variances
              0.4314268
    > ratio
    [1] 1.1
    > res$null.value
    ratio of variances
    > res$alternative
    [1] "two.sided"
      Test di ipotesi su proporzioni
Test con un campione
   • Package: stats
   • Sintassi: prop.test()
   • Input:
        x numero di successi
        n dimensione campionaria
        p il valore di p_0
        alternative = "less" / "greater" / "two.sided" ipotesi alternativa
        conf.level livello di confidenza 1-\alpha
        correct = FALSE
   • Output:
        statistic valore empirico della statistica \chi^2
        parameter gradi di libertà
        p.value p-value
```

conf.int intervallo di confidenza per la proporzione incognita al livello $1-\alpha$

estimate proporzione calcolata sulla base del campione

null.value il valore di p_0

alternative ipotesi alternativa

• Formula:

statistic $z^2 = \left(\frac{\frac{x}{n} - p_0}{\sqrt{\frac{p_0 \left(1 - p_0\right)}{n}}}\right)$ parameter 1

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$P(\chi_1^2 \ge z^2)$

conf.int $\frac{\left(2\,x+z_{1-\alpha/2}^2\right)\mp\sqrt{\left(2\,x+z_{1-\alpha/2}^2\right)^2-4\,\left(n+z_{1-\alpha/2}^2\right)\,x^2/n}}{2\,\left(n+z_{1-\alpha/2}^2\right)}$ estimate $\frac{x}{n}$ null.value

```
> x <- 10
> n < -23
> p0 < - 0.45
> z <- (x/n - p0)/sqrt(p0 * (1 - p0)/n)
[1] -0.1466954
> z^2
[1] 0.02151954
> res <- prop.test(x = 10, n = 23, p = 0.45, alternative = "two.sided",
+ conf.level = 0.95, correct = FALSE)
> res$statistic
X-squared
0.02151954
> res$parameter
df
1
> p.value < 1 - pchisq(z^2, df = 1)
> p.value
[1] 0.8833724
> res$p.value
[1] 0.8833724
```

```
> alpha <- 0.05
 > zc <- qnorm(1 - 0.05/2)
 > lower <- ((2 * x + zc^2) - sqrt((2 * x + zc^2)^2 - 4 * (n + zc^2) *
       x^2/n)/(2 * (n + zc^2))
 > upper <- ((2 * x + zc^2) + sqrt((2 * x + zc^2)^2 - 4 * (n + zc^2) *
       x^2/n)/(2 * (n + zc^2))
 > c(lower, upper)
 [1] 0.2563464 0.6318862
 > res$conf.int
 [1] 0.2563464 0.6318862
 attr(,"conf.level")
 [1] 0.95
 > x/n
 [1] 0.4347826
 > res$estimate
 0.4347826
 > p0
 [1] 0.45
 > res$null.value
    p
 0.45
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x <- 18
 > n < -30
 > p0 < -0.55
 > z <- (x/n - p0)/sqrt(p0 * (1 - p0)/n)
 > z
 [1] 0.5504819
 > z^2
 [1] 0.3030303
 > res <- prop.test(x = 18, n = 30, p = 0.55, alternative = "two.sided",
 + conf.level = 0.95, correct = FALSE)
 > res$statistic
 X-squared
 0.3030303
```

```
> res$parameter
df
1
> p.value < 1 - pchisq(z^2, df = 1)
> p.value
[1] 0.5819889
> res$p.value
[1] 0.5819889
> alpha <- 0.05
> zc <- qnorm(1 - 0.05/2)
> lower <- (zc^2/(2 * n) + x/n - zc * sqrt(zc^2/(4 * n^2) + x/n *
+ (1 - x/n)/n))/(1 + zc^2/n)
> upper <- (zc<sup>2</sup>/(2 * n) + x/n + zc * sqrt(zc<sup>2</sup>/(4 * n<sup>2</sup>) + x/n *
+ (1 - x/n)/n))/(1 + zc^2/n)
> c(lower, upper)
[1] 0.4232036 0.7540937
> res$conf.int
[1] 0.4232036 0.7540937
attr(,"conf.level")
[1] 0.95
> x/n
[1] 0.6
> res$estimate
 p
0.6
> p0
[1] 0.55
> res$null.value
  р
0.55
> res$alternative
[1] "two.sided"
```

Potenza nel Test con un campione

• Package: stats

• Sintassi: power.prop.test()

• Input:

n il valore n della dimensione di ciascun campione

p1 valore p_1 della proporzione sotto ipotesi nulla

p2 il valore p_2 della proporzione sotto l'ipotesi alternativa

sig.level livello di significatività α

power potenza $1 - \beta$

alternative $pu\`{o}$ essere cambiata in one.sided, two.sided a seconda del numero di code che interessano

• Output:

p1 il valore p_1 della proporzione sotto l'ipotesi nulla

p2 il valore p_2 della proporzione sotto l'ipotesi alternativa

 ${\tt n}\;$ il valore n della dimensione di ciascun campione

sig.level livello di significatività α

power potenza $1 - \beta$

alternative ipotesi alternativa

• Formula:

$$\xi = \sqrt{p_1 (1 - p_1) + p_2 (1 - p_2)}$$

$$\delta = \sqrt{(p_1 + p_2) (1 - (p_1 + p_2)/2)}$$

$$\gamma = |p_1 - p_2|$$
 alternative = one.sided

p1

 p_1

р2

 p_2

n

$$n = [(\xi/\gamma) \Phi^{-1}(1-\beta) + (\delta/\gamma) \Phi^{-1}(1-\alpha)]^{2}$$

sig.level

$$\alpha = 1 - \Phi \left((\gamma / \delta) \sqrt{n} - (\xi / \delta) \Phi^{-1} (1 - \beta) \right)$$

power

$$1 - \beta = \Phi \left(\left(\gamma / \xi \right) \sqrt{n} - \left(\delta / \xi \right) \Phi^{-1} (1 - \alpha) \right)$$

p1

 p_1

p2

 p_2

n

$$n \, = \, \left[(\xi \, / \, \gamma) \, \Phi^{-1} (1 - \beta) + (\delta \, / \, \gamma) \, \Phi^{-1} (1 - \alpha \, / \, 2) \right]^2$$

sig.level

$$\alpha = 2 \left[1 - \Phi \left((\gamma / \delta) \sqrt{n} - (\xi / \delta) \Phi^{-1} (1 - \beta) \right) \right]$$

power

$$1 - \beta = \Phi\left(\left(\gamma/\xi\right)\sqrt{n} - \left(\delta/\xi\right)\Phi^{-1}(1 - \alpha/2)\right)$$

```
> n < -23
    > p1 <- 0.23
    > p2 < -0.31
    > power.prop.test(n, p1, p2, sig.level = NULL, power = 0.9, alternative = "one.sided")
         Two-sample comparison of proportions power calculation
                  n = 23
                 p1 = 0.23
                 p2 = 0.31
          sig.level = 0.7470593
              power = 0.9
        alternative = one.sided
     NOTE: n is number in *each* group
  • Example 2:
    > p1 < -0.23
    > p2 <- 0.31
    > power.prop.test(n = NULL, p1, p2, sig.level = 0.05, power = 0.9,
          alternative = "one.sided")
         Two-sample comparison of proportions power calculation
                  n = 525.6022
                 p1 = 0.23
                 p2 = 0.31
          sig.level = 0.05
              power = 0.9
        alternative = one.sided
     NOTE: n is number in *each* group
  • Example 3:
    > n < -23
    > p1 <- 0.23
    > p2 <- 0.31
    > power.prop.test(n, p1, p2, sig.level = 0.05, power = NULL, alternative = "one.sided")
         Two-sample comparison of proportions power calculation
                  n = 23
                 p1 = 0.23
                 p2 = 0.31
          sig.level = 0.05
              power = 0.1496353
        alternative = one.sided
     NOTE: n is number in *each* group
Test con due campioni indipendenti
  • Package: stats
```

```
• Sintassi: prop.test()
• Input:
      x numero di successi nei due campioni
      n dimensione dei due campioni
```

```
alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza 1-\alpha correct = FALSE
```

• Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value conf.int intervallo di confidenza per la differenza tra le proporzioni incognite al livello $1-\alpha$ estimate proporzioni calcolate sulla base dei campioni alternative ipotesi alternativa

• Formula:

statistic

$$z^{2} = \left(\frac{\left|\frac{x_{1}}{n_{1}} - \frac{x_{2}}{n_{2}}\right| - 0.5\left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)}{\sqrt{\frac{x_{1} + x_{2}}{n_{1} + n_{2}}\left(1 - \frac{x_{1} + x_{2}}{n_{1} + n_{2}}\right)\left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)}}\right)^{2}$$

$$z^{2} = \left(\frac{\frac{x_{1}}{n_{1}} - \frac{x_{2}}{n_{2}}}{\sqrt{\frac{x_{1} + x_{2}}{n_{1} + n_{2}} \left(1 - \frac{x_{1} + x_{2}}{n_{1} + n_{2}}\right) \left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)}}\right)^{2}$$

parameter

1

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$1 - P(\chi_1^2 \le z^2)$

conf.int

$$\left| \frac{x_1}{n_1} - \frac{x_2}{n_2} \right| \mp 0.5 \left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mp z_{1-\alpha/2} \sqrt{\frac{\frac{x_1}{n_1} \left(1 - \frac{x_1}{n_1} \right)}{n_1} + \frac{\frac{x_2}{n_2} \left(1 - \frac{x_2}{n_2} \right)}{n_2}} \right|$$

$$\boxed{\text{correct = FALSE}}$$

$$\frac{x_1}{n_1} - \frac{x_2}{n_2} \mp z_{1-\alpha/2} \sqrt{\frac{\frac{x_1}{n_1} \left(1 - \frac{x_1}{n_1}\right)}{n_1} + \frac{\frac{x_2}{n_2} \left(1 - \frac{x_2}{n_2}\right)}{n_2}}$$

estimate

$$\frac{x_1}{n_1}$$
 $\frac{x_2}{n_2}$

```
> x <- c(9, 11)

> n <- c(23, 32)

> x1 <- 9

> x2 <- 11

> n1 <- 23

> n2 <- 32

> z <- (x1/n1 - x2/n2)/sqrt((x1 + x2)/(n1 + n2) * (1 - (x1 + x2)/(n1 + n2)) * (1/n1 + 1/n2))

> z^2
```

```
[1] 0.1307745
 > res <- prop.test(x = c(9, 11), n = c(23, 32), alternative = "two.sided",
      conf.level = 0.95, correct = FALSE)
 > res$statistic
 X-squared
 0.1307745
 > res$parameter
 df
 > p.value < 1 - pchisq(z^2, df = 1)
 > p.value
 [1] 0.7176304
 > res$p.value
 [1] 0.7176304
 > lower <- (x1/n1 - x2/n2) - qnorm(1 - 0.05/2) * sqrt(x1/n1 * (1 -
 + x1/n1)/n1 + x2/n2 * (1 - x2/n2)/n2)
 > upper <- (x1/n1 - x2/n2) + qnorm(1 - 0.05/2) * sqrt(x1/n1 * (1 -
      x1/n1)/n1 + x2/n2 * (1 - x2/n2)/n2)
 > c(lower, upper)
 [1] -0.2110231 0.3061318
 > res$conf.int
 [1] -0.2110231 0.3061318
 attr(,"conf.level")
 [1] 0.95
 > c(x1/n1, x2/n2)
 [1] 0.3913043 0.3437500
 > res$estimate
    prop 1
             prop 2
 0.3913043 0.3437500
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < - c(4, 11)
 > n < -c(20, 24)
 > x1 <- 4
 > x2 <- 11
 > n1 <- 20
 > n2 < -24
 > z <- (x1/n1 - x2/n2)/sqrt((x1 + x2)/(n1 + n2) * (1 - (x1 + x2)/(n1 + n2))
      n2)) * (1/n1 + 1/n2))
 > z^2
```

```
[1] 3.240153
> res <- prop.test(x = c(4, 11), n = c(20, 24), alternative = "two.sided",
+ conf.level = 0.95, correct = FALSE)
> res$statistic
X-squared
 3.240153
> res$parameter
df
1
> p.value < 1 - pchisq(z^2, df = 1)
> p.value
[1] 0.07185392
> res$p.value
[1] 0.07185392
> lower <- (x1/n1 - x2/n2) - qnorm(1 - 0.05/2) * sqrt(x1/n1 * (1 -
+ x1/n1)/n1 + x2/n2 * (1 - x2/n2)/n2)
> upper <- (x1/n1 - x2/n2) + qnorm(1 - 0.05/2) * sqrt(x1/n1 * (1 -
     x1/n1)/n1 + x2/n2 * (1 - x2/n2)/n2)
> c(lower, upper)
[1] -0.523793280 0.007126613
> res$conf.int
[1] -0.523793280 0.007126613
attr(,"conf.level")
[1] 0.95
> c(x1/n1, x2/n2)
[1] 0.2000000 0.4583333
> res$estimate
   prop 1
           prop 2
0.2000000 0.4583333
> res$alternative
[1] "two.sided"
```

Test con k campioni indipendenti

Package: stats
Sintassi: prop.test()
Input:

x numero di successi nei k campioni
n dimensione dei k campioni
correct = FALSE

• Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value estimate proporzioni calcolate sulla base dei k campioni

• Formula:

statistic $c = \sum_{i=1}^k \left(\frac{\frac{x_i}{n_i} - \hat{p}}{\sqrt{\hat{p}\left(1 - \hat{p}\right)/n_i}}\right)^2$ $\text{dove} \quad \hat{p} = \frac{\sum_{j=1}^k x_j}{\sum_{j=1}^k n_j}$ $parameter \\ df = k-1$ $p.value \\ P(\chi_{df}^2 \geq c)$ estimate $\frac{x_i}{n_i} \quad \forall i=1,2,\ldots,k$

```
> k <- 3
> x < -c(10, 21, 32)
> n <- c(23, 55, 81)
> phat <- sum(x)/sum(n)
> statistic <- sum(((x/n - phat)/sqrt(phat * (1 - phat)/n))^2)
> statistic
[1] 0.1911084
> prop.test(x, n, correct = FALSE)$statistic
X-squared
0.1911084
> parameter <- k - 1
> parameter
[1] 2
> prop.test(x, n, correct = FALSE)$parameter
df
 2
```

```
> p.value < 1 - pchisq(statistic, df = k - 1)
 > p.value
 [1] 0.9088691
 > prop.test(x, n, correct = FALSE) $p.value
 [1] 0.9088691
 > estimate <- x/n
 > estimate
 [1] 0.4347826 0.3818182 0.3950617
 > prop.test(x, n, correct = FALSE)$estimate
             prop 2
    prop 1
                       prop 3
 0.4347826 0.3818182 0.3950617
• Example 2:
 > k <- 4
 > x < -c(17, 14, 21, 34)
 > n <- c(26, 22, 33, 45)
 > phat <- sum(x)/sum(n)
 > statistic <- sum(((x/n - phat)/sqrt(phat * (1 - phat)/n))^2)
 > statistic
 [1] 1.747228
 > prop.test(x, n, correct = FALSE) $statistic
 X-squared
  1.747228
 > parameter <- k - 1
 > parameter
 [1] 3
 > prop.test(x, n, correct = FALSE)$parameter
 df
  3
 > p.value <- 1 - pchisq(statistic, df = k - 1)</pre>
 > p.value
 [1] 0.6264855
 > prop.test(x, n, correct = FALSE)$p.value
 [1] 0.6264855
 > estimate <- x/n
 > estimate
 [1] 0.6538462 0.6363636 0.6363636 0.7555556
 > prop.test(x, n, correct = FALSE) $estimate
    prop 1
             prop 2
                        prop 3
                                  prop 4
 0.6538462 0.6363636 0.6363636 0.7555556
```

6.5 Test di ipotesi sull'omogeneità delle varianze

Test di Bartlett

• Package: stats

```
Sintassi: bartlett.test()
Input:

x vettore numerico di dimensione n
g fattore a k livelli di dimensione n
```

• Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p\text{-value}
```

• Formula:

statistic
$$c=\frac{(n-k)\log{(s_P^2)}-\sum_{j=1}^k{(n_j-1)\log{(s_j^2)}}}{1+\frac{1}{3(k-1)}\left(\sum_{j=1}^k{\frac{1}{n_j-1}}-\frac{1}{n-k}\right)}$$
 dove
$$s_P^2=\frac{\sum_{j=1}^k{(n_j-1)\,s_j^2}}{n-k}$$
 parameter
$$df=k-1$$
 p.value
$$P(\chi_{df}^2\geq c)$$

```
> x <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 16.5, 22, 1.2, 3.4)
> g <- factor(rep(1:4, each = 3))</pre>
> g
 [1] 1 1 1 2 2 2 3 3 3 4 4 4
Levels: 1 2 3 4
> n < -12
> k <- 4
> s2 <- tapply(x, g, var)
 21.000000 3.103333 16.470000 130.573333
> enne <- tapply(x, g, length)</pre>
> enne
1 2 3 4
3 3 3 3
> Sp2 <- sum((enne - 1) * s2/(n - k))
> Sp2
[1] 42.78667
> c <- ((n - k) * log(Sp2) - sum((enne - 1) * log(s2)))/(1 + 1/(3 *
     (k-1)) * (sum(1/(enne-1)) - 1/(n-k)))
> c
```

```
[1] 5.254231
     > res <- bartlett.test(x, g)</pre>
      > res$statistic
      Bartlett's K-squared
                                                      5.254231
      > parameter <- k - 1
      > parameter
     [1] 3
     > res$parameter
      df
          3
      > p.value < 1 - pchisq(c, df = k - 1)
      > p.value
      [1] 0.1541
     > res$p.value
      [1] 0.1541
• Example 2:
      > x < -c(0.7, -1.6, -0.2, -1.2, -0.1, 3.4, 3.7, 0.8, 0, 2, 1.9,
      > g <- factor(rep(1:2, c(8, 4)))
         [1] 1 1 1 1 1 1 1 1 2 2 2 2
     Levels: 1 2
     > n <- 12
      > k <- 2
      > s2 <- tapply(x, g, var)
      > s2
      3.8069643 0.9091667
     > enne <- tapply(x, g, length)</pre>
      > enne
      1 2
      8 4
     > Sp2 <- sum((enne - 1) * s2/(n - k))
      > Sp2
      [1] 2.937625
      > c <- ((n - k) * log(Sp2) - sum((enne - 1) * log(s2)))/(1 + 1/(3 * (enne - 1) * log(s2))/(1 + 1/(3 * (
                          (k-1)) * (sum(1/(enne - 1)) - 1/(n - k)))
      > C
```

```
[1] 1.514017
> res <- bartlett.test(x, g)</pre>
> res$statistic
Bartlett's K-squared
            1.514017
> parameter <- k - 1
> parameter
[1] 1
> res$parameter
df
1
> p.value <- 1 - pchisq(c, df = k - 1)
> p.value
[1] 0.2185271
> res$p.value
[1] 0.2185271
```

Capitolo 7

Analisi della varianza (Anova)

7.1 Simbologia

• numero di livelli dei fattori di colonna e di riga:

Anova	f (colonna)	g (riga)
ad un fattore	k	/
a due fattori senza interazione	k	h
a due fattori con interazione	k	h

• dimensione campionaria di colonna, di riga e di cella:

Anova	j-esima colonna	<i>i</i> -esima riga	<i>ij</i> -esima cella
ad un fattore	n_j	/	/
a due fattori senza interazione	hl	kl	l
a due fattori con interazione	hl	kl	l

• medie campionarie di colonna, di riga e di cella:

Anova	j-esima colonna	<i>i</i> -esima riga	<i>ij</i> -esima cella
ad un fattore	\bar{y}_j	/	/
a due fattori senza interazione	$ar{y}_{\cdot j}$.	$\bar{y}_{i\cdots}$	\bar{y}_{ij} .
a due fattori con interazione	$ar{y}_{\cdot j}$.	\bar{y}_{i}	$ar{y}_{ij}$.

• media campionaria generale: \bar{y}

7.2 Modelli di analisi della varianza

Anova ad un fattore

• Package: stats

• Sintassi: anova()

• Input:

y vettore numerico di dimensione n

f fattore a k livelli di dimensione n

• Output:

Df gradi di libertà

Sum Sq somma dei quadrati

Mean Sq media dei quadrati

 ${ t F}$ value valore empirico della statistica F

Pr(>F) p-value

• Formula:

Df

f	k-1
Residuals	n-k

Sum Sq

f	$\sum_{j=1}^{k} n_j (\bar{y}_j - \bar{y})^2$
Residuals	$\sum_{j=1}^{k} \sum_{i=1}^{n_j} (y_{ij} - \bar{y}_j)^2$

Mean Sq

F value
$$Fvalue = \frac{\left[\sum_{j=1}^{k} n_j \left(\bar{y}_j - \bar{y}\right)^2\right]/(k-1)}{\left[\sum_{j=1}^{k} \sum_{i=1}^{n_j} \left(y_{ij} - \bar{y}_j\right)^2\right]/\left(n-k\right)}$$
 Pr(>F)
$$P(F_{k-1,\,n-k} \geq Fvalue)$$

• Examples:

```
> y <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 16.5, 22, 1.2, 3.4)
> f <- factor(rep(letters[1:4], each = 3))</pre>
 [1] a a a b b b c c c d d d
Levels: a b c d
> n < -12
> k <- 4
> modello <- lm(formula = y \sim f)
> anova(modello)
Analysis of Variance Table
Response: y
          Df Sum Sq Mean Sq F value Pr(>F)
           3 136.03 45.34 1.0597 0.4184
Residuals 8 342.29
                     42.79
> res <- anova(object = modello)</pre>
> res$Df
[1] 3 8
> res$"Sum Sq"
[1] 136.0292 342.2933
> res$"Mean Sq"
[1] 45.34306 42.78667
> res$"F value"
```

[1] 1.059747 NA
> res\$"Pr(>F)"

[1] 0.4183517 NA

Anova a due fattori senza interazione

• Package: stats

• Sintassi: anova()

• Input:

y vettore numerico di dimensione khl

f fattore a k livelli di dimensione khl

g fattore a h livelli di dimensione khl

• Output:

Df gradi di libertà

Sum Sq somma dei quadrati

Mean Sq media dei quadrati

 ${\mathbb F}$ value valore empirico della statistica F

Pr(>F) *p*-value

• Formula:

Df

f	k-1
g	h-1
Residuals	k h l - (k+h-1)

Sum Sq

	f	$hl \sum_{j=1}^{k} (\bar{y}_{\cdot j} - \bar{y})^2$
	g	$kl \sum_{i=1}^{h} (\bar{y}_{i\cdots} - \bar{y})^2$
ĺ	Residuals	

Mean Sq

f	$\left[hl \sum_{j=1}^{k} (\bar{y}_{\cdot j} - \bar{y})^{2}\right] / (k-1)$
g	$\left[kl \sum_{i=1}^{h} (\bar{y}_{i} - \bar{y})^{2}\right] / (h - 1)$
Residuals	$\frac{\left[l\sum_{j=1}^{k}\sum_{i=1}^{h}(\bar{y}_{ij}\bar{y}_{i}\bar{y}_{\cdot j}.+\bar{y})^{2}+\sum_{j=1}^{k}\sum_{i=1}^{h}\sum_{m=1}^{l}(y_{ijm}-\bar{y}_{ij\cdot})^{2}\right]}{[khl-(k\!+\!h\!-\!1)]}$

F value

$$F_{f}value = \frac{\left[h \, l \, \sum_{j=1}^{k} \left(\bar{y}_{\cdot j \cdot } - \bar{y}\right)^{2}\right] / \left(k-1\right)}{\frac{\left[l \, \sum_{j=1}^{k} \, \sum_{i=1}^{h} \left(\bar{y}_{i j \cdot } - \bar{y}_{i \cdot } - \bar{y}_{\cdot j \cdot } + \bar{y}\right)^{2} + \sum_{j=1}^{k} \, \sum_{i=1}^{h} \, \sum_{m=1}^{l} \left(y_{i j m} - \bar{y}_{i j \cdot }\right)^{2}\right]}{\left[k \, h \, l - \left(k + h - 1\right)\right]}}$$

$$F_gvalue = \frac{\left[kl\sum_{i=1}^{h}(\bar{y}_{i\cdot\cdot} - \bar{y})^2\right]/(h-1)}{\frac{\left[l\sum_{j=1}^{k}\sum_{i=1}^{h}(\bar{y}_{ij\cdot} - \bar{y}_{i\cdot\cdot} - \bar{y}_{\cdot j\cdot} + \bar{y})^2 + \sum_{j=1}^{k}\sum_{i=1}^{h}\sum_{m=1}^{l}(y_{ijm} - \bar{y}_{ij\cdot})^2\right]}{[k\,h\,l - (k+h-1)]}$$

```
Pr(>F)
                                              P(F_{k-1,\,k\,h\,l-(k+h-1)} \ge F_f value)
                                              P(F_{h-1, k h l-(k+h-1)}) \ge F_g value)
```

```
> y <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 6.5, 2, 1.2, 3.4)
> f <- factor(rep(letters[1:2], each = 6))</pre>
> f
 [1] a a a a a a b b b b b b
Levels: a b
> g <- factor(rep(LETTERS[2:1], times = 6))</pre>
 [1] B A B A B A B A B A B A
Levels: A B
> table(f, g)
  g
f AB
 a 3 3
 b 3 3
> n <- 12
> k <- 2
> h <- 2
> 1 <- 3
> 1
[1] 3
> modello <- lm(formula = y \sim f + g)
> anova(object = modello)
Analysis of Variance Table
Response: y
          Df Sum Sq Mean Sq F value Pr(>F)
          1 4.441 4.441 0.2913 0.6025
              0.188
                      0.188 0.0123 0.9141
Residuals 9 137.194 15.244
> res <- anova(object = modello)</pre>
> res$Df
[1] 1 1 9
> res$"Sum Sq"
[1]
    4.440833 0.187500 137.194167
> res$"Mean Sq"
[1] 4.440833 0.187500 15.243796
> res$"F value"
```

• Note: Il numero di replicazioni per cella l deve essere maggiore od uguale ad uno.

NA

Anova a due fattori con interazione

• Package: stats

• Sintassi: anova()

• Input:

y vettore numerico di dimensione khl

f fattore a k livelli di dimensione khl

g fattore a h livelli di dimensione khl

• Output:

Df gradi di libertà

Sum Sq somma dei quadrati

Mean Sq media dei quadrati

 ${\mathbb F}$ value valore empirico della statistica F

Pr(>F) p-value

• Formula:

Df

\int	k-1
\parallel g	h-1
f:g	(k-1)(h-1)
Residuals	k h (l-1)

Sum Sq

f	$hl \sum_{j=1}^k (\bar{y}_{\cdot j} - \bar{y})^2$
g	$kl \sum_{i=1}^{h} (\bar{y}_{i\cdots} - \bar{y})^2$
f:g	$l \sum_{j=1}^{k} \sum_{i=1}^{h} (\bar{y}_{ij.} - \bar{y}_{i} - \bar{y}_{.j.} + \bar{y})^2$
Residuals	$\sum_{j=1}^{k} \sum_{i=1}^{h} \sum_{m=1}^{l} (y_{ijm} - \bar{y}_{ij.})^2$

Mean Sq

f	$\left[hl \sum_{j=1}^{k} (\bar{y}_{\cdot j} - \bar{y})^{2} \right] / (k-1)$
g	$\left[kl \sum_{i=1}^{h} (\bar{y}_{i} - \bar{y})^{2}\right] / (h - 1)$
f:g	$ \left[l \sum_{j=1}^{k} \sum_{i=1}^{h} (\bar{y}_{ij.} - \bar{y}_{i} - \bar{y}_{.j.} + \bar{y})^{2} \right] / \left[(k-1)(h-1) \right] $
Residuals	$\left[\sum_{j=1}^{k} \sum_{i=1}^{h} \sum_{m=1}^{l} (y_{ijm} - \bar{y}_{ij})^{2}\right] / \left[k h (l-1)\right]$

```
F value
                                     F_{f}value = \frac{\left[h l \sum_{j=1}^{k} (\bar{y}_{\cdot j \cdot - \bar{y}})^{2}\right] / (k-1)}{\left[\sum_{i=1}^{k} \sum_{i=1}^{h} \sum_{m=1}^{l} (y_{ijm} - \bar{y}_{ij \cdot})^{2}\right] / [k h (l-1)]}
                                     F_gvalue = \frac{\left[kl \sum_{i=1}^{h} (\bar{y}_{i\cdot\cdot} - \bar{y})^2\right] / (h-1)}{\left[\sum_{j=1}^{k} \sum_{i=1}^{h} \sum_{m=1}^{l} (y_{ijm} - \bar{y}_{ij\cdot})^2\right] / [k h (l-1)]}
                                F_{f:g}value = \frac{\left[l \sum_{j=1}^{k} \sum_{i=1}^{h} (\bar{y}_{ij} - \bar{y}_{i} - \bar{y}_{ij} + \bar{y})^{2}\right] / \left[(k-1)(h-1)\right]}{\left[\sum_{j=1}^{k} \sum_{i=1}^{h} \sum_{m=1}^{l} (y_{ijm} - \bar{y}_{ij})^{2}\right] / \left[kh(l-1)\right]}
        Pr(>F)
                                                          P(F_{k-1,kh(l-1)} \ge F_f value)
                                                          P(F_{h-1, k h (l-1)} \ge F_g value)
                                                    P(F_{(k-1)(h-1), k h (l-1)}) \ge F_{f:g} value)
• Examples:
  > y <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 6.5, 2, 1.2, 3.4)
  > f <- factor(rep(letters[1:2], each = 6))</pre>
   [1] a a a a a a b b b b b b
  Levels: a b
  > g <- factor(rep(LETTERS[2:1], times = 6))</pre>
  > g
   [1] B A B A B A B A B A B A
  Levels: A B
  > table(f, q)
  f AB
     a 3 3
     b 3 3
  > n <- 12
  > k <- 2
  > h <- 2
  > 1 <- 3
  > modello <- lm(formula = y \sim f + g + f:g)
  > anova(object = modello)
  Analysis of Variance Table
  Response: y
                   Df Sum Sq Mean Sq F value Pr(>F)
                         4.441 4.441 0.2616 0.6228
                                      0.188 0.0110 0.9189
                         0.188
  g
  f:q
                          1.401
                                       1.401
                                                   0.0825 0.7812
                    1
  Residuals 8 135.793 16.974
  > res <- anova(object = modello)</pre>
  > res$Df
  [1] 1 1 1 8
  > res$"Sum Sq"
            4.440833 0.187500 1.400833 135.793333
  [1]
```

```
> res$"Mean Sq"

[1] 4.440833 0.187500 1.400833 16.974167

> res$"F value"

[1] 0.26162305 0.01104620 0.08252737 NA

> res$"Pr(>F)"

[1] 0.6228225 0.9188831 0.7812018 NA
```

• **Note:** Il numero di replicazioni per cella *l* deve essere maggiore di uno.

7.3 Comandi utili in analisi della varianza

factor()

• Package: base

• Input:

```
x vettore alfanumerico di dimensione n
levels etichette di livello
labels etichette di livello
ordered = TRUE / FALSE livelli su scala ordinale
```

- **Description:** crea un fattore
- Examples:

```
> factor(x = rep(c("U", "D"), each = 4), levels = c("U", "D"))
Levels: U D
> factor(x = rep(c("U", "D"), each = 4), levels = c("D", "U"))
[1] U U U U D D D
Levels: D U
> factor(x = rep(1:2, each = 4), labels = c("U", "D"))
[1] U U U U D D D
Levels: U D
> factor(x = rep(1:2, each = 4), labels = c("D", "U"))
[1] D D D D U U U U
Levels: D U
> factor(x = rep(1:2, each = 4), labels = c("U", "D"), ordered = TRUE)
Levels: U < D
> factor(x = rep(1:2, each = 4), labels = c("D", "U"), ordered = TRUE)
Levels: D < U
```

as.factor()

• Package: base

• Input:

 \times vettore alfanumerico di dimensione n

- Description: creazione di un fattore
- Examples:

relevel()

• Package: stats

• Input:

 \mathbf{x} fattore a k livelli ref livello di riferimento

- **Description:** ricodificazione dei livelli di un fattore
- Examples:

```
> x <- factor(c("a", "b", "c", "a", "b", "b", "c", "c", "a", "b"))
> x

[1] a b c a b b c c a b
Levels: a b c
> relevel(x, ref = "b")
```

```
[1] abcabbccab
    Levels: b a c
    > relevel(x, ref = "c")
    [1] abcabbccab
    Levels: c a b
levels()
  • Package: base
  • Input:
        f fattore a k livelli
  • Description: nome dei livelli
  • Examples:
    > f <- factor(rep(1:2, each = 5))
    > f
     [1] 1 1 1 1 1 2 2 2 2 2 2
    Levels: 1 2
    > levels(f)
    [1] "1" "2"
    > f <- factor(rep(c("U", "D"), each = 4))
    > f
    [1] U U U U D D D D
    Levels: D U
    > levels(f)
    [1] "D" "U"
nlevels()
  • Package: base
  • Input:
        f fattore a k livelli
  • Description: numero di livelli
  • Examples:
    > f <- factor(rep(1:2, each = 5))</pre>
    > f
    [1] 1 1 1 1 1 2 2 2 2 2
    Levels: 1 2
    > nlevels(f)
```

```
[1] 2
   > f <- factor(c("A", "A", "A", "B", "B", "B", "B", "C", "C"))
   > f
    [1] A A A A B B B B C C
   Levels: A B C
   > nlevels(f)
   [1] 3
ordered()
 • Package: base
 • Input:
       \times vettore alfanumerico di dimensione n
       levels etichette dei livelli
 • Description: fattore con livelli su scala ordinale
 • Examples:
   >  ordered(x = c(rep("U", 5), rep("D", 5)), levels = c("U", "D"))
    Levels: U < D
   >  ordered(x = c(rep("U", 5), rep("D", 5)), levels = c("D", "U"))
    Levels: D < U
   > fattore <- ordered(x = c("a", "b", "c", "a", "b", "b", "c", "c",</pre>
        "a", "b"), levels = c("a", "b", "c"))
   > fattore
    [1] abcabbccab
   Levels: a < b < c
   > fattore < "b"</pre>
```

[1] TRUE FALSE FALSE TRUE FALSE FALSE FALSE TRUE FALSE

as.ordered()

```
• Package: base
```

• Input:

x vettore alfanumerico di dimensione n

• Description: fattore con livelli su scala ordinale

• Examples:

```
> as.ordered(x = c(rep("U", 5), rep("D", 5)))

[1] U U U U U D D D D D
Levels: D < U
> as.ordered(x = c(rep("U", 5), rep("D", 5)))

[1] U U U U D D D D D
Levels: D < U
> as.ordered(x = c("a", "b", "c", "a", "b", "b", "c", "c", "a", "b"))

[1] a b c a b b c c a b
Levels: a < b < c</pre>
```

letters[]

• Package: base

• **Description:** lettere minuscole

• Examples:

```
> letters[1:6]
[1] "a" "b" "c" "d" "e" "f"
> letters[c(3, 5, 6, 26)]
[1] "c" "e" "f" "z"
```

LETTERS[]

• Package: base

• Description: lettere maiuscole

```
> LETTERS[1:6]

[1] "A" "B" "C" "D" "E" "F"

> LETTERS[c(3, 5, 6, 26)]

[1] "C" "E" "F" "Z"
```

as.numeric()

```
• Package: base
```

• Input:

x fattore a k livelli

• Description: codici dei livelli

• Examples:

as.integer()

• Package: base

• Input:

 \times fattore a k livelli

• Description: codici dei livelli

unclass()

```
• Package: base
```

• Input:

 \times fattore a k livelli

• Description: codici dei livelli

• Examples:

```
> x < - factor(c(2, 3, 1, 1, 1, 3, 4, 4, 1, 2), labels = c("A",
+ "B", "C", "D"))
 [1] B C A A A C D D A B
Levels: A B C D
> unclass(x)
[1] 2 3 1 1 1 3 4 4 1 2
attr(,"levels")
[1] "A" "B" "C" "D"
> x <- factor(c("M", "F", "M", "F", "M", "F", "F", "M"), levels = c("M",</pre>
     "F"))
> x
[1] M F M F M F F M
Levels: M F
> unclass(x)
[1] 1 2 1 2 1 2 2 1
attr(,"levels")
[1] "M" "F"
```

by()

• Package: base

• Input:

data vettore numerico y di dimensione n INDICES fattore f a k livelli FUN funzione

- Description: applica FUN ad ogni vettore numerico per livello del fattore
- Example 1:

```
> y <- c(1.2, 2.3, 5.6, 3.5, 2.5, 3.8, 6.8, 5.7, 3.7, 6.4)
> f <- factor(c("a", "b", "c", "a", "b", "c", "c", "a", "b"))
> f

[1] a b c a b b c c a b
Levels: a b c
> by(data = y, INDICES = f, FUN = mean)
```

```
f: a
 [1] 2.8
 f: b
 [1] 3.75
 f: c
 [1] 6.033333
• Example 2:
 > y <- c(1.2, 2.3, 5.6, 3.5, 2.5, 3.8, 6.8, 5.7, 3.7, 6.4)
 > g <- factor(c("alto", "medio", "basso", "alto", "medio", "basso",</pre>
      "medio", "alto", "alto", "basso"))
 > g
  [1] alto medio basso alto medio basso medio alto basso
 Levels: alto basso medio
 > by(data = y, INDICES = g, FUN = mean)
 g: alto
 [1] 3.525
 g: basso
 [1] 5.266667
 g: medio
 [1] 3.866667
• Example 3:
 > y < -c(1.2, 2.3, 5.6, 3.5, 2.5, 3.8, 6.8, 5.7, 3.7, 6.4)
 > f <- factor(c("a", "b", "c", "a", "b", "b", "c", "c", "a", "b"))</pre>
 > f
  [1] abcabbccab
 Levels: a b c
 > g <- factor(c("alto", "medio", "basso", "alto", "medio", "basso",</pre>
      "medio", "alto", "alto", "basso"))
 > g
  [1] alto medio basso alto medio basso medio alto basso
 Levels: alto basso medio
 > by(data = y, INDICES = list(f, g), FUN = mean)
 : a
 : alto
 [1] 2.8
 : b
 : alto
 [1] NA
 : C
 : alto
 [1] 5.7
 : a
 : basso
 [1] NA
```

```
: b
: basso
[1] 5.1

: c
: basso
[1] 5.6

: a
: medio
[1] NA

: b
: medio
[1] 2.4

: c
: medio
[1] 6.8
```

tapply()

• Package: base

• Input:

```
X vettore numerico x di dimensione n INDEX fattore f a k livelli FUN funzione
```

- ullet Description: applica la funzione FUN $\,$ ad ogni gruppo di elementi di x definito dai livelli di f
- Examples:

```
> X \leftarrow c(1.2, 2.3, 5.6, 3.5, 2.5, 3.8, 6.8, 5.7, 3.7, 6.4)
> f <- factor(c("a", "b", "c", "a", "b", "b", "c", "c", "a", "b"))</pre>
[1] abcabbccab
Levels: a b c
> g <- factor(c("alto", "medio", "basso", "alto", "medio", "basso",</pre>
     "medio", "alto", "alto", "basso"))
> g
[1] alto medio basso alto medio basso medio alto basso
Levels: alto basso medio
> tapply(X, INDEX = f, FUN = mean)
              b
2.800000 3.750000 6.033333
> tapply(X, INDEX = list(f, g), FUN = mean)
 alto basso medio
a 2.8 NA NA
  NA
        5.1
             2.4
c 5.7 5.6 6.8
```

g1()

```
• Package: base
  • Input:
       n numero dei livelli
       k numero delle replicazioni
       length dimensione del fattore risultato
       labels nomi dei livelli
       ordered = TRUE / FALSE fattore ordinato
  • Description: crea un fattore
  • Examples:
   > gl(n = 2, k = 5, labels = c("M", "F"))
    [1] MMMMMFFFFF
   Levels: M F
   > gl(n = 2, k = 1, length = 10, labels = c("A", "B"))
    [1] A B A B A B A B A B
   Levels: A B
   > gl(n = 2, k = 8, labels = c("Control", "Treat"), ordered = TRUE)
    [1] Control Control Control Control Control Control Control Treat
    [10] Treat Treat Treat Treat
                                          Treat
                                                  Treat
                                                          Treat
   Levels: Control < Treat
ave()
  • Package: stats
  • Input:
       x vettore numerico di dimensione n
```

- f fattore a k livelli di dimensione n
- FUN funzione
- **Description:** applica e replica la funzione FUN ad ogni gruppo di elementi di x definito dai livelli di f
- Examples:

```
> x < -c(1, 2, 3, 4, 5, 6, 7, 8)
> f <- factor(rep(letters[1:2], each = 4))</pre>
> f
[1] a a a a b b b b
Levels: a b
> mean(x[f == "a"])
[1] 2.5
> mean(x[f == "b"])
[1] 6.5
```

```
> ave(x, f, FUN = mean)
[1] 2.5 2.5 2.5 2.5 6.5 6.5 6.5 6.5
> x < -c(1, 2, 3, 4, 5, 6, 7, 8)
> f <- factor(rep(letters[1:2], each = 4))</pre>
[1] a a a a b b b b
Levels: a b
> sum(x[f == "a"])
[1] 10
> sum(x[f == "b"])
[1] 26
> ave(x, f, FUN = sum)
[1] 10 10 10 10 26 26 26 26
> x < -c(1, 2, 3, 4, 5, 6, 7, 8)
> f <- factor(rep(letters[1:2], each = 4))</pre>
> f
[1] a a a a b b b b
Levels: a b
> mean(x[f == "a"])
[1] 2.5
> mean(x[f == "b"])
[1] 6.5
> ave(x, f, FUN = function(x) mean(x, trim = 0.1))
[1] 2.5 2.5 2.5 2.5 6.5 6.5 6.5 6.5
```

cut()

```
• Package: base
```

• Input:

```
x vettore numerico di dimensione n breaks estremi delle classi di ampiezza b_i right = TRUE / FALSE classi chiuse a destra \left(a_{(i)},\,a_{(i+1)}\right] oppure a sinistra \left[a_{(i)},\,a_{(i+1)}\right) include.lowest = TRUE / FALSE estremo incluso labels etichette ordered_result = TRUE / FALSE fattore ordinato
```

- **Description:** raggruppamento in classi
- Examples:

```
> x < -c(1.2, 2.3, 4.5, 5.4, 3.4, 5.4, 2.3, 2.1, 1.23, 4.3, 0.3)
> n <- 11
> cut(x, breaks = c(0, 4, 6), right = TRUE, include.lowest = FALSE,
     labels = c("0-4", "4-6"))
[1] 0-4 0-4 4-6 4-6 0-4 4-6 0-4 0-4 0-4 4-6 0-4
Levels: 0-4 4-6
> x < -c(1, 2, 3, 4, 5.6, 7.4, 1.2, 4, 4.4)
> n < -9
> cut(x, breaks = c(0, 4, 8), right = TRUE, include.lowest = FALSE,
  labels = c("0-4", "4-8"))
[1] 0-4 0-4 0-4 0-4 4-8 4-8 0-4 0-4 4-8
Levels: 0-4 4-8
> x < -c(1, 2, 3, 4, 5.6, 7.4, 1.2, 4, 4.4)
> cut(x, breaks = c(0, 4, 8), right = TRUE, include.lowest = FALSE,
     labels = c("0-4", "4-8"), ordered_result = TRUE)
[1] 0-4 0-4 0-4 0-4 4-8 4-8 0-4 0-4 4-8
Levels: 0-4 < 4-8
```

summary()

• Package: base

• Input:

object fattore a k livelli di dimensione n

- **Description:** distribuzione di frequenza assoluta
- Examples:

```
a b c
   4 4 4
   > f <- factor(c("ALTO", "ALTO", "BASSO", "MEDIO", "ALTO", "BASSO",
        "MEDIO", "BASSO"))
   > f
    [1] ALTO ALTO BASSO MEDIO ALTO BASSO MEDIO BASSO
   Levels: ALTO BASSO MEDIO
   > summary(object = f)
    ALTO BASSO MEDIO
      3 3 2
interaction()
  • Package: base
  • Input:
        ... fattori su cui eseguire l'interazione
  • Description: interazione tra fattori
  • Example 1:
   > a <- factor(rep(1:2, each = 4))
    [1] 1 1 1 1 2 2 2 2
   Levels: 1 2
   > b <- factor(rep(c("ctrl", "treat"), times = 2, each = 2))</pre>
    [1] ctrl ctrl treat treat ctrl ctrl treat treat
   Levels: ctrl treat
   > interaction(a, b)
    [1] 1.ctrl 1.ctrl 1.treat 1.treat 2.ctrl 2.ctrl 2.treat 2.treat
   Levels: 1.ctrl 2.ctrl 1.treat 2.treat
  • Example 2:
   > a <- factor(rep(1:2, each = 4))
   > a
   [1] 1 1 1 1 2 2 2 2
   Levels: 1 2
   > b <- factor(rep(c("M", "F"), times = 4))</pre>
   [1] M F M F M F M F
   Levels: F M
   > interaction(a, b)
```

```
[1] 1.M 1.F 1.M 1.F 2.M 2.F 2.M 2.F Levels: 1.F 2.F 1.M 2.M
```

• Example 3:

```
> a <- factor(rep(c("M", "F"), times = 4))
> a

[1] M F M F M F M F
Levels: F M

> b <- factor(rep(c("M", "F"), times = 4))
> b

[1] M F M F M F M F
Levels: F M

> interaction(a, b)

[1] M.M F.F M.M F.F M.M F.F M.M F.F
Levels: F.F M.F F.M M.M
```

expand.grid()

- Package: base
- Input:
 - ... vettori numerici o fattori
- Description: creazione di un data frame da tutte le combinazioni di vettori numerici o fattori
- Example 1:

```
> height <- c(60, 80)
> weight <- c(100, 300, 500)
> sex <- factor(c("Male", "Female"))</pre>
> mydf <- expand.grid(height = height, weight = weight, sex = sex)
> mydf
  height weight
                  sex
                Male
1
      60
          100
2
      80
            100
                 Male
3
      60
           300
                 Male
     80
4
           300 Male
5
     60 500 Male
6
     80
          500 Male
7
     60
           100 Female
8
           100 Female
     80
9
      60
            300 Female
10
      80
            300 Female
11
      60
            500 Female
      80
           500 Female
12
> is.data.frame(mydf)
[1] TRUE
```

• Example 2:

```
> Sex <- factor(c("Women", "Men"), levels = c("Women", "Men"))</pre>
 > Age <- factor(c("18-23", "24-40", ">40"), levels = c("18-23",
       "24-40", ">40"))
 > Response <- factor(c("little importance", "importance", "very importance"),</pre>
       levels = c("little importance", "importance", "very importance"))
 > mydf <- expand.grid(Sex = Sex, Age = Age, Response = Response)</pre>
 > Freq <- c(26, 40, 9, 17, 5, 8, 12, 17, 21, 15, 14, 15, 7, 8,
       15, 12, 41, 18)
 > mydf <- cbind(mydf, Freq)</pre>
 > mydf
            Age
                         Response Freq
    Women 18-23 little importance
    Men 18-23 little importance
                                    40
    Women 24-40 little importance
                                    9
    Men 24-40 little importance
                                    17
 5
    Women >40 little importance
                                    5
 6
    Men >40 little importance
                                    8
 7 Women 18-23
                      importance
                                    12
    Men 18-23
                                    17
                      importance
 9 Women 24-40
                      importance
                                    21
 10 Men 24-40
                      importance
                                    15
 11 Women
          >40
                       importance
                                    14
 12 Men
            >40
                      importance
                                    15
 13 Women 18-23 very importance
                                    7
 Men 18-23 very importance
                                    8
 15 Women 24-40 very importance
                                    15
 16 Men 24-40 very importance
                                    12
 17 Women >40 very importance
                                    41
            >40 very importance
 18
    Men
                                    18
 > is.data.frame(mydf)
 [1] TRUE
• Example 3:
 > x <- LETTERS[1:3]
 > y <- 1:2
 > z <- letters[1:2]</pre>
 > mydf <- expand.grid(x = x, y = y, z = z)
    х у г
 1
   A 1 a
 2 B 1 a
 3 C 1 a
 4 A 2 a
   B 2 a
 5
   C 2 a
 6
 7
    A 1 b
    B 1 b
 9 C 1 b
 10 A 2 b
 11 B 2 b
 12 C 2 b
 > is.data.frame(mydf)
 [1] TRUE
```

Capitolo 8

Confronti multipli

8.1 Simbologia

• numero di livelli dei fattori di colonna e di riga:

Anova	f (colonna)	g (riga)
ad un fattore	k	/
a due fattori senza interazione	k	h
a due fattori con interazione	k	h

• dimensione campionaria di colonna, di riga e di cella:

Anova	j-esima colonna	<i>i</i> -esima riga	<i>ij</i> -esima cella
ad un fattore	n_j	/	/
a due fattori senza interazione	hl	kl	/
a due fattori con interazione	hl	kl	l

• medie campionarie di colonna, di riga e di cella:

Anova	j-esima colonna	<i>i</i> -esima riga	<i>ij-</i> esima cella
ad un fattore	\bar{y}_j	/	/
a due fattori senza interazione	$ar{y}_{\cdot j}$.	$\bar{y}_{i\cdots}$	\bar{y}_{ij} .
a due fattori con interazione	$ar{y}_{\cdot j}$.	$\bar{y}_{i\cdots}$	$ar{y}_{ij}$.

ullet media campionaria generale: $ar{y}$

8.2 Metodo di Tukey

Applicazione in Anova ad un fattore

• Package: stats

• Sintassi: TukeyHSD()

• Input:

y vettore numerico di dimensione n

f fattore con livelli $1, 2, \ldots, k$

conf.level livello di confidenza $1-\alpha$

- Output:
 - f intervallo di confidenza a livello $1-\alpha$ per il fattore f
- Formula:

f

$$\bar{y}_i - \bar{y}_j \quad \forall i > j = 1, 2, \dots, k$$

$$\bar{y}_i - \bar{y}_j \mp q_{1-\alpha, k, n-k} s_P \sqrt{1/(2n_i) + 1/(2n_j)} \quad \forall i > j = 1, 2, \dots, k$$

dove
$$s_P^2 = \sum_{i=1}^k \sum_{j=1}^{n_j} (y_{ij} - \bar{y}_j)^2 / (n - k)$$

```
> y <- c(19, 24, 24, 27, 20, 24, 22, 21, 22, 29, 18, 17)
> f <- factor(rep(1:3, times = 4))
 [1] 1 2 3 1 2 3 1 2 3 1 2 3
Levels: 1 2 3
> n <- 12
> k < - 3
> alpha <- 0.05
> qTUKEY <- qtukey(0.95, nmeans = k, df = n - k)
> qTUKEY
[1] 3.948492
> TukeyHSD(aov(formula = y \sim f), conf.level = 0.95)
  Tukey multiple comparisons of means
    95% family-wise confidence level
Fit: aov(formula = y \sim f)
$f
   diff
                lwr
                        upr
                                p adj
2-1 -3.5 -10.534094 3.534094 0.3860664
3-1 -2.5 -9.534094 4.534094 0.5996130
3-2 1.0 -6.034094 8.034094 0.9175944
> res <- TukeyHSD(aov(formula = y ~ f), conf.level = 0.95)
> y1m <- mean(y[f == "1"])
> y1m
[1] 24.25
> y2m <- mean(y[f == "2"])
> y2m
[1] 20.75
> y3m <- mean(y[f == "3"])
> y3m
[1] 21.75
> differ <- c(y2m - y1m, y3m - y1m, y3m - y2m)
> n1 <- length(y[f == "1"])
> n1
[1] 4
> n2 <- length(y[f == "2"])
> n2
```

```
[1] 4
> n3 <- length(y[f == "3"])
> n3
[1] 4
> Sp2 <- anova(lm(formula = y ~ f))$"Mean Sq"[2]
> stderror <- sqrt(Sp2) * sqrt(c(1/(2 * n2) + 1/(2 * n1), 1/(2 *
     n3) + 1/(2 * n1), 1/(2 * n3) + 1/(2 * n2)))
> lower <- differ - qTUKEY * stderror
> upper <- differ + qTUKEY * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 3, ncol = 3,
     dimnames = list(c("2-1", "3-1", "3-2"), c("diff", "lwr",
          "upr")))
   diff
                lwr
2-1 -3.5 -10.534094 3.534094
3-1 -2.5 -9.534094 4.534094
3-2 1.0 -6.034094 8.034094
> res$f
               lwr
                       upr
                               p adj
2-1 -3.5 -10.534094 3.534094 0.3860664
3-1 -2.5 -9.534094 4.534094 0.5996130
3-2 1.0 -6.034094 8.034094 0.9175944
```

• **Note:** Il numero di confronti è pari a $\binom{k}{2}$ per il fattore f.

Applicazione in Anova a due fattori senza interazione

```
• Package: stats
• Sintassi: TukeyHSD()
• Input:
       y vettore numerico di dimensione khl
       f fattore con livelli 1, 2, \ldots, k
       g fattore con livelli 1, 2, \ldots, h
       conf.level livello di confidenza 1-\alpha
```

• Output:

```
f intervallo di confidenza a livello 1-\alpha per il fattore f
g intervallo di confidenza a livello 1-\alpha per il fattore g
```

• Formula:

f $\bar{y}_{\cdot i}$ $-\bar{y}_{\cdot j}$ $\forall i > j = 1, 2, \ldots, k$ $\bar{y}_{\cdot i} - \bar{y}_{\cdot j} \mp q_{1-\alpha, k, k h l-(k+h-1)} s_P / \sqrt{h l} \quad \forall i > j = 1, 2, \dots, k$ dove $s_P^2 = \frac{l \sum_{j=1}^k \sum_{i=1}^h (\bar{y}_{ij\cdot} - \bar{y}_{i\cdot\cdot} - \bar{y}_{\cdot j\cdot} + \bar{y})^2 + \sum_{j=1}^k \sum_{i=1}^h \sum_{m=1}^l (y_{ijm} - \bar{y}_{ij\cdot})^2}{k h l - (k+h-1)}$

Levels: a b > g <- factor(rep(LETTERS[2:1], times = 6))</pre> [1] B A B A B A B A B A B A Levels: A B > table(f, g) g f AB a 3 3 b 3 3 > n <- 12 > k <- 2 > h <- 2 > 1 <- 3 > alpha <- 0.05 > qTUKEYf <- qtukey(0.95, nmeans = k, df = k * h * 1 - (k + h -1)) > qTUKEYf [1] 3.199173 > qTUKEYg <- qtukey(0.95, nmeans = h, df = k * h * 1 - (k + h -1)) > qTUKEYg [1] 3.199173 > TukeyHSD(aov(formula = $y \sim f + g$), conf.level = 0.95) Tukey multiple comparisons of means 95% family-wise confidence level Fit: $aov(formula = y \sim f + g)$ \$f diff lwr upr p adj b-a 6.216667 -2.001707 14.43504 0.1212097 \$9 diff lwr upr p adj

B-A -1.416667 -9.63504 6.801707 0.7056442

```
> res <- TukeyHSD(aov(formula = y ~ f + g), conf.level = 0.95)
> y.1.m <- mean(y[f == "a"])
> y.1.m
[1] 4.366667
> y.2.m <- mean(y[f == "b"])
> y.2.m
[1] 10.58333
> differ <- y.2.m - y.1.m
> Sp2 <- anova(lm(formula = y ~ f + g))$"Mean Sq"[3]
> stderror <- sqrt(Sp2)/sqrt(h * 1)</pre>
> lower <- differ - qTUKEYf * stderror
> upper <- differ + qTUKEYf * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 1, ncol = 3,
     dimnames = list("b-a", c("diff", "lwr", "upr")))
        diff
                 lwr
                          upr
b-a 6.216667 -2.001707 14.43504
> res$f
               lwr
                        upr padj
b-a 6.216667 -2.001707 14.43504 0.1212097
> y1..m <- mean(y[g == "A"])
> y1..m
[1] 8.183333
> y2..m <- mean(y[g == "B"])
> y2..m
[1] 6.766667
> differ <- y2..m - y1..m
> Sp2 <- anova(lm(formula = y ~ f + g))$"Mean Sq"[3]
> stderror <- sqrt(Sp2)/sqrt(k * 1)</pre>
> lower <- differ - qTUKEYg * stderror
> upper <- differ + qTUKEYg * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 1, ncol = 3,
      dimnames = list("B-A", c("diff", "lwr", "upr")))
         diff
                  lwr
B-A -1.416667 -9.63504 6.801707
> res$g
         diff
                  lwr
                           upr
B-A -1.416667 -9.63504 6.801707 0.7056442
```

- Note 1: Il numero di replicazioni per cella l deve essere maggiore od uguale ad uno.
- **Note 2:** Il numero di confronti è pari a $\binom{k}{2}$ per il fattore f.
- **Note 3:** Il numero di confronti è pari a $\binom{h}{2}$ per il fattore g.

Applicazione in Anova a due fattori con interazione

```
Package: stats
Sintassi: TukeyHSD()
Input:

y vettore numerico di dimensione khl
f fattore con livelli 1, 2, ..., k
g fattore con livelli 1, 2, ..., h
conf.level livello di confidenza 1 - α
```

• Output:

- f intervallo di confidenza a livello $1-\alpha$ per il fattore f q intervallo di confidenza a livello $1-\alpha$ per il fattore q
- f:g intervallo di confidenza a livello $1-\alpha$ per l'interazione f:g

• Formula:

$$\begin{split} \bar{y}_{\cdot i \cdot} - \bar{y}_{\cdot j \cdot} &\quad \forall i > j = 1, \, 2, \, \dots, \, k \\ \bar{y}_{\cdot i \cdot} - \bar{y}_{\cdot j \cdot} &\quad \mp \, q_{1-\alpha, \, k, \, k \, h \, (l-1)} \, s_P \, / \sqrt{h \, l} \quad \forall i > j = 1, \, 2, \, \dots, \, k \end{split}$$

$$\begin{aligned} \operatorname{dove} \quad s_P^2 &= \sum_{j=1}^k \sum_{i=1}^h \sum_{m=1}^l \left(y_{ijm} - \bar{y}_{ij \cdot} \right)^2 / \left[k \, h \, (l-1) \right] \\ \bar{y}_{i \cdot} - \bar{y}_{j \cdot} &\quad \forall i > j = 1, \, 2, \, \dots, \, h \\ \bar{y}_{i \cdot} - \bar{y}_{j \cdot} &\quad \mp \, q_{1-\alpha, \, h, \, k \, h \, (l-1)} \, s_P \, / \sqrt{k \, l} \quad \forall i > j = 1, \, 2, \, \dots, \, h \end{aligned}$$

$$\begin{aligned} \operatorname{dove} \quad s_P^2 &= \sum_{j=1}^k \sum_{i=1}^h \sum_{m=1}^l \left(y_{ijm} - \bar{y}_{ij \cdot} \right)^2 / \left[k \, h \, (l-1) \right] \\ \bar{t} : \mathsf{g} \\ \bar{y}_{ij \cdot} - \bar{y}_{uw} \quad \forall i, \, u = 1, \, 2, \, \dots, \, h \quad \forall j, \, w = 1, \, 2, \, \dots, \, k \\ \bar{y}_{ij \cdot} - \bar{y}_{uw} \cdot \mp \, q_{1-\alpha, \, k \, h, \, k \, h \, (l-1)} \, s_P \, / \sqrt{l} \quad \forall i, \, u = 1, \, 2, \, \dots, \, h \quad \forall j, \, w = 1, \, 2, \, \dots, \, k \end{aligned}$$

dove
$$s_P^2 = \sum_{i=1}^k \sum_{i=1}^h \sum_{m=1}^l (y_{ijm} - \bar{y}_{ij.})^2 / [k h (l-1)]$$

```
> y <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 16.5, 22, 1.2, 3.4)
> f <- factor(rep(letters[1:2], each = 6))
> f

[1] a a a a a a b b b b b
Levels: a b

> g <- factor(rep(LETTERS[1:2], times = 6))
> g

[1] A B A B A B A B A B A B B
Levels: A B
> table(f, g)
```

```
g
   АВ
  a 3 3
 b 3 3
> n < -12
> k < - 2
> h <- 2
> 1 <- 3
> alpha <- 0.05
> qTUKEYf <- qtukey(0.95, nmeans = k, df = k * h * (1 - 1))
> qTUKEYf
[1] 3.261182
> qTUKEYg <- qtukey(0.95, nmeans = h, df = k * h * (1 - 1))
> qTUKEYg
[1] 3.261182
> qTUKEYfg <- qtukey(0.95, nmeans = k * h, df = k * h * (1 - 1))
> qTUKEYfg
[1] 4.52881
> TukeyHSD(aov(y \sim f + g + f:g), conf.level = 0.95)
  Tukey multiple comparisons of means
    95% family-wise confidence level
Fit: aov(formula = y \sim f + g + f:g)
$f
        diff
                   lwr
                            upr
                                    p adj
b-a 6.216667 -2.460179 14.89351 0.1371018
$g
        diff
                 lwr
                           upr
                                   p adj
B-A 1.416667 -7.26018 10.09351 0.7163341
$`f:g`
              diff
                          lwr
                                    upr
                                            p adj
b:A-a:A 3.8666667 -13.173972 20.90731 0.8838028
a:B-a:A -0.9333333 -17.973972 16.10731 0.9979198
b:B-a:A 7.6333333 -9.407306 24.67397 0.5144007
a:B-b:A -4.8000000 -21.840639 12.24064 0.8043752
b:B-b:A 3.7666667 -13.273972 20.80731 0.8912420
b:B-a:B 8.5666667 -8.473972 25.60731 0.4251472
> res <- TukeyHSD(aov(y ~ f + g + f:g), conf.level = 0.95)
> y.1.m <- mean(y[f == "a"])
> y.1.m
[1] 4.366667
> y.2.m \leftarrow mean(y[f == "b"])
> y.2.m
[1] 10.58333
```

```
> differ <- y.2.m - y.1.m</pre>
> Sp2 <- anova(lm(formula = y ~ f + g))$"Mean Sq"[4]
> stderror <- sqrt(Sp2)/sqrt(h * 1)</pre>
> lower <- differ - qTUKEYf * stderror</pre>
> upper <- differ + qTUKEYf * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 1, ncol = 3,
      dimnames = list("b-a", c("diff", "lwr", "upr")))
        diff lwr upr
b-a 6.216667 NA NA
> res$f
        diff
                  lwr
                            upr
                                   p adj
b-a 6.216667 -2.460179 14.89351 0.1371018
> y1..m <- mean(y[g == "A"])
> y1..m
[1] 6.766667
> y2..m <- mean(y[g == "B"])
> y2..m
[1] 8.183333
> differ <- y2..m - y1..m</pre>
> Sp2 <- anova(lm(formula = y ~ f + g))$"Mean Sq"[3]
> stderror <- sqrt(Sp2)/sqrt(k * 1)</pre>
> lower <- differ - qTUKEYg * stderror</pre>
> upper <- differ + qTUKEYg * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 1, ncol = 3,
      dimnames = list("B-A", c("diff", "lwr", "upr")))
        diff
                  lwr
B-A 1.416667 -6.961002 9.794335
> res$g
        diff
                 lwr
                          upr
                                  p adj
B-A 1.416667 -7.26018 10.09351 0.7163341
> y11.m <- mean(y[f == "a" & g == "A"])
> y11.m
[1] 4.833333
> y12.m <- mean(y[f == "b" & g == "A"])
> y12.m
[1] 8.7
> y21.m <- mean(y[f == "a" & g == "B"])
> y21.m
[1] 3.9
> y22.m <- mean(y[f == "b" & g == "B"])
> y22.m
```

```
[1] 12.46667
> differ <- c(y12.m - y11.m, y21.m - y11.m, y22.m - y11.m, y21.m -
      y12.m, y22.m - y12.m, y22.m - y21.m)
> Sp2 <- anova(lm(formula = y ~ f * g))$"Mean Sq"[4]
> stderror <- rep(sqrt(Sp2)/sqrt(1), times = 6)</pre>
> lower <- differ - qTUKEYfg * stderror
> upper <- differ + qTUKEYfg * stderror
> matrix(data = cbind(differ, lower, upper), nrow = 6, ncol = 3,
     dimnames = list(c("b:A-a:A", "a:B-a:A", "b:B-a:A", "a:B-b:A",
          "b:B-b:A", "b:B-a:B"), c("diff", "lwr", "upr")))
              diff
                          lwr
b:A-a:A 3.8666667 -13.173972 20.90731
a:B-a:A -0.9333333 -17.973972 16.10731
b:B-a:A 7.6333333 -9.407306 24.67397
a:B-b:A -4.8000000 -21.840639 12.24064
b:B-b:A 3.7666667 -13.273972 20.80731
b:B-a:B 8.5666667 -8.473972 25.60731
> res$"f:q"
                         lwr
                                   upr
                                          p adj
b:A-a:A 3.8666667 -13.173972 20.90731 0.8838028
a:B-a:A -0.9333333 -17.973972 16.10731 0.9979198
b:B-a:A 7.6333333 -9.407306 24.67397 0.5144007
a:B-b:A -4.8000000 -21.840639 12.24064 0.8043752
b:B-b:A 3.7666667 -13.273972 20.80731 0.8912420
b:B-a:B 8.5666667 -8.473972 25.60731 0.4251472
```

- Note 1: Il numero di replicazioni per cella l deve essere maggiore di uno.
- Note 2: Il numero di confronti è pari a $\binom{k}{2}$ per il fattore f.
- **Note 3:** Il numero di confronti è pari a $\binom{h}{2}$ per il fattore g.
- Note 4: Il numero di confronti è pari a $\binom{kh}{2}$ per l'interazione f:g.

8.3 Metodo di Bonferroni

p.value p-value

Applicazione in Anova ad un fattore

```
Package: stats
Sintassi: pairwise.t.test()
Input:

y vettore numerico di dimensione n
f fattore con livelli 1, 2, ..., k livelli di dimensione n
p.adjust.method = "bonferroni"
Output:
```

• Formula:

```
p.value
                               2\binom{k}{2} P(t_{n-k} \le -|t|) = k(k-1) P(t_{n-k} \le -|t|)
                              \mbox{dove} \quad t \, = \, \frac{\bar{y}_i - \bar{y}_j}{s_P \, \sqrt{1 \, / \, n_i + 1 \, / \, n_j}} \quad \forall \, i > j \, = \, 1, \, 2, \, \ldots, \, k \label{eq:dove}
                                     con s_P^2 = \sum_{i=1}^k \sum_{i=1}^{n_j} (y_{ij} - \bar{y}_j)^2 / (n - k)
• Examples:
 > y < -c(1, 14, 1, 12.1, 3.5, 5.6, 18.4, 12, 1.65, 22, 1.2, 1.34)
 > f \leftarrow factor(rep(1:3, times = 4))
 > f
   [1] 1 2 3 1 2 3 1 2 3 1 2 3
 Levels: 1 2 3
 > n <- 12
 > k <- 3
 > m.1 <- mean(y[f == "1"])
 > m.2 <- mean(y[f == "2"])
 > m.3 <- mean(y[f == "3"])
 > n1 <- length(y[f == "1"])
 > n2 <- length(y[f == "2"])
 > n3 <- length(y[f == "3"])
 > s2 <- anova(lm(formula = y ~ f))$"Mean Sq"[2]
 > s <- sqrt(s2)
 > t12 <- (m.2 - m.1)/(s * sqrt(1/n1 + 1/n2))
 > t13 <- (m.3 - m.1)/(s * sqrt(1/n3 + 1/n1))
 > t23 <- (m.3 - m.2)/(s * sqrt(1/n3 + 1/n2))
 > p12 <- k * (k - 1) * pt(-abs(t12), df = n - k)
 > p13 < -k * (k - 1) * pt(-abs(t13), df = n - k)
 > p23 <- k * (k - 1) * pt(-abs(t23), df = n - k)
 > matrix(data = c(p12, p13, NA, p23), dimnames = list(c("2", "3"),
        c("1", "2")), nrow = 2, ncol = 2)
 2 0.7493036
 3 0.1258454 0.8521961
 > pairwise.t.test(y, f, p.adjust.method = "bonferroni")
           Pairwise comparisons using t tests with pooled SD
 data: y and f
    1
 20.75 -
 3 0.13 0.85
 P value adjustment method: bonferroni
 > res <- pairwise.t.test(y, f, p.adjust.method = "bonferroni")</pre>
 > res$p.value
```

2 0.7493036

3 0.1258454 0.8521961

Metodo di Student

```
Applicazione in Anova ad un fattore
   • Package: stats
   • Sintassi: pairwise.t.test()
   • Input:
           \vee vettore numerico di dimensione n
           f fattore con livelli 1, 2, \ldots, k di dimensione n
           p.adjust.method = "none"
   • Output:
           p.value p-value
   • Formula:
           p.value
                                                         2P(t_{n-k} \le -|t|)
                                      dove t = \frac{\bar{y}_i - \bar{y}_j}{s_P \sqrt{1/n_i + 1/n_j}} \quad \forall i > j = 1, 2, ..., k
                                              con s_P^2 = \sum_{i=1}^k \sum_{j=1}^{n_j} (y_{ij} - \bar{y}_j)^2 / (n - k)
   • Examples:
     > y <- c(19, 24, 24, 27, 20, 24, 22, 21, 22, 29, 18, 17)
     > f <- factor(rep(1:3, times = 4))</pre>
```

```
[1] 1 2 3 1 2 3 1 2 3 1 2 3
Levels: 1 2 3
> n <- 12
> k < - 3
> m.1 <- mean(y[f == "1"])
> m.2 <- mean(y[f == "2"])
> m.3 <- mean(y[f == "3"])
> n1 <- length(y[f == "1"])
> n2 <- length(y[f == "2"])
> n3 <- length(y[f == "3"])
> s2 <- anova(lm(formula = y ~ f))$"Mean Sq"[2]
> s <- sqrt(s2)
> t12 <- (m.2 - m.1)/(s * sqrt(1/n1 + 1/n2))
> t13 <- (m.3 - m.1)/(s * sqrt(1/n3 + 1/n1))
> t23 <- (m.3 - m.2)/(s * sqrt(1/n3 + 1/n2))
> p12 <- 2 * pt(-abs(t12), df = n - k)
> p13 <- 2 * pt(-abs(t13), df = n - k)
> p23 <- 2 * pt(-abs(t23), df = n - k)
> matrix(data = c(p12, p13, NA, p23), dimnames = list(c("2", "3"),
     c("1", "2")), nrow = 2, ncol = 2)
2 0.1981691
3 0.3469732 0.7006709
> pairwise.t.test(y, f, p.adjust.method = "none")
```

```
Pairwise comparisons using t tests with pooled SD

data: y and f

1 2
2 0.20 -
3 0.35 0.70

P value adjustment method: none

> res <- pairwise.t.test(y, f, p.adjust.method = "none")
> res$p.value

1 2
2 0.1981691 NA
3 0.3469732 0.7006709
```

Capitolo 9

Test di ipotesi su correlazione ed autocorrelazione

9.1 Test di ipotesi sulla correlazione lineare

Test di Pearson

```
• Package: stats
```

• Sintassi: cor.test()

• Input:

```
\times vettore numerico di dimensione n
```

y vettore numerico di dimensione \boldsymbol{n}

alternative = "less" / "greater" / "two.sided" ipotesi alternativa conf.level livello di confidenza $1-\alpha$

• Output:

statistic valore empirico della statistica t

parameter gradi di libertà

p.value p-value

conf.int intervallo di confidenza a livello $1-\alpha$ ottenuto con la trasformazione Z di Fisher

estimate coefficiente di correlazione campionario

alternative ipotesi alternativa

• Formula:

statistic

$$t = r_{xy} \sqrt{\frac{n-2}{1-r_{xy}^2}} = \frac{\hat{\beta}_2}{s/\sqrt{ss_x}}$$

dove
$$r_{xy} = \frac{s_{xy}}{s_x s_y} = \hat{\beta}_2 \frac{s_x}{s_y}$$

parameter

$$df = n - 2$$

p.value

alternative	less	greater	two.sided
p.value	$P(t_{df} \leq t)$	$1 - P(t_{df} \le t)$	$2P(t_{df} \le - t)$

conf.int

$$\tanh\left(\frac{1}{2}\log\left(\frac{1+r_{xy}}{1-r_{xy}}\right) \mp z_{1-\alpha/2} \frac{1}{\sqrt{n-3}}\right)$$

dove
$$\tanh(x) = \frac{\sinh(x)}{\cosh(x)} = \frac{e^{2x} - 1}{e^{2x} + 1}$$

estimate

 r_{xy}

```
• Example 1:
```

```
> x < -c(1, 2, 2, 4, 3, 3)
> y < -c(6, 6, 7, 7, 7, 9)
> n < -6
> r <- cov(x, y)/(sd(x) * sd(y))
[1] 0.522233
> t <- r * sqrt((n - 2)/(1 - r^2))
[1] 1.224745
> res <- cor.test(x, y, alternative = "two.sided", conf.level = 0.95)</pre>
> res$statistic
1.224745
> parameter <- n - 2</pre>
> parameter
[1] 4
> res$parameter
df
4
> p.value < 2 * pt(-abs(t), df = n - 2)
> p.value
[1] 0.2878641
> res$p.value
[1] 0.2878641
> lower <- tanh(0.5 * log((1 + r)/(1 - r)) - qnorm(1 - 0.05/2)/sqrt(n -
> upper <- tanh(0.5 * log((1 + r)/(1 - r)) + qnorm(1 - 0.05/2)/sqrt(n -
     3))
> c(lower, upper)
[1] -0.5021527 0.9367690
> res$conf.int
[1] -0.5021527 0.9367690
attr(,"conf.level")
[1] 0.95
> r
[1] 0.522233
```

```
> res$estimate
      cor
 0.522233
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < -c(1.2, 1.2, 3.4, 3.4, 4.5, 5.5, 5.5, 5.6, 6.6, 6.6)
 > y <- c(1.3, 1.3, 1.3, 4.5, 5.6, 6.7, 6.7, 6.7, 8.8, 8.8, 9)
 > n <- 11
 > r <- cov(x, y)/(sd(x) * sd(y))
 > r
 [1] 0.9527265
 > t <- r * sqrt((n - 2)/(1 - r^2))
 [1] 9.40719
 > res <- cor.test(x, y, alternative = "two.sided", conf.level = 0.95)</pre>
 > res$statistic
       t
 9.40719
 > parameter <- n - 2
 > parameter
 [1] 9
 > res$parameter
 df
  9
 > p.value < 2 * pt(-abs(t), df = n - 2)
 > p.value
 [1] 5.936572e-06
 > res$p.value
 [1] 5.936572e-06
 > lower <- tanh(0.5 * log((1 + r)/(1 - r)) - qnorm(1 - 0.05/2)/sqrt(n -
 + 3))
 > upper <- tanh(0.5 * log((1 + r)/(1 - r)) + qnorm(1 - 0.05/2)/sqrt(n -
       3))
 > c(lower, upper)
 [1] 0.8234897 0.9879637
 > res$conf.int
```

```
[1] 0.8234897 0.9879637
 attr(, "conf.level")
 [1] 0.95
 > r
 [1] 0.9527265
 > res$estimate
       cor
 0.9527265
 > res$alternative
 [1] "two.sided"
• Example 3:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > r <- cov(x, y)/(sd(x) * sd(y))
 [1] 0.740661
 > t <- r * sqrt((n - 2)/(1 - r^2))
 [1] 2.700251
 > res <- cor.test(x, y, alternative = "two.sided", conf.level = 0.95)</pre>
 > res$statistic
 2.700251
 > parameter <- n - 2
 > parameter
 [1] 6
 > res$parameter
 df
  6
 > p.value < 2 * pt(-abs(t), df = n - 2)
 > p.value
 [1] 0.03556412
 > res$p.value
 [1] 0.03556412
```

```
> lower <- tanh(0.5 * log((1 + r)/(1 - r)) - qnorm(1 - 0.05/2)/sqrt(n -
> upper <- tanh(0.5 * log((1 + r)/(1 - r)) + qnorm(1 - 0.05/2)/sqrt(n -
      3))
> c(lower, upper)
[1] 0.07527696 0.94967566
> res$conf.int
[1] 0.07527696 0.94967566
attr(,"conf.level")
[1] 0.95
> r
[1] 0.740661
> res$estimate
     cor
0.740661
> res$alternative
[1] "two.sided"
```

Test di Kendall

```
Package: statsSintassi: cor.test()Input:
```

```
x vettore numerico di dimensione n y vettore numerico di dimensione n alternative = "less" / "greater" / "two.sided" ipotesi alternativa method = "kendall" exact = F
```

• Output:

statistic valore empirico della statistica Z p.value p-value estimate coefficiente di correlazione campionario alternative ipotesi alternativa

• Formula:

statistic

$$z = \frac{1}{\sigma_K} \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} sign((x_j - x_i)(y_j - y_i))$$

dove

$$\begin{split} \sigma_{K}^{2} &= \frac{n\left(n-1\right)\left(2\,n+5\right)}{18} + \\ &- \frac{\sum_{i=1}^{g} t_{i}\left(t_{i}-1\right)\left(2\,t_{i}+5\right) + \sum_{j=1}^{h} u_{j}\left(u_{j}-1\right)\left(2\,u_{j}+5\right)}{18} + \\ &+ \frac{\left[\sum_{i=1}^{g} t_{i}\left(t_{i}-1\right)\left(t_{i}-2\right)\right]\left[\sum_{j=1}^{h} u_{j}\left(u_{j}-1\right)\left(u_{j}-2\right)\right]}{9\,n\left(n-1\right)\left(n-2\right)} + \\ &+ \frac{\left[\sum_{i=1}^{g} t_{i}\left(t_{i}-1\right)\right]\left[\sum_{j=1}^{h} u_{j}\left(u_{j}-1\right)\right]}{2\,n\left(n-1\right)} \end{split}$$

e t, u sono i ties di x ed y rispettivamente.

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(- z))$

estimate

$$r_{xy}^{K} = \frac{2\sum_{i=1}^{n-1}\sum_{j=i+1}^{n} \operatorname{sign}((x_{j} - x_{i})(y_{j} - y_{i}))}{\left(n(n-1) - \sum_{i=1}^{g} t_{i}(t_{i} - 1)\right)^{1/2} \left(n(n-1) - \sum_{j=1}^{h} u_{j}(u_{j} - 1)\right)^{1/2}}$$

• Example 1:

> u

```
> x <- c(1, 2, 2, 4, 3, 3)
> y <- c(6, 6, 7, 7, 7, 9)
> matrice <- matrix(data = 0, nrow = n - 1, ncol = n, byrow = F)
> for (i in 1:(n - 1)) for (j in (i + 1):n) matrice[i, j] <- sign((x[j] -
+ x[i]) * (y[j] - y[i]))
> num <- sum(matrice)</pre>
> num
[1] 7
> table(x)
1 2 3 4
1 2 2 1
> g <- 2
> t1 <- 2
> t2 <- 2
> t <- c(t1, t2)
> t
[1] 2 2
> table(y)
6 7 9
2 3 1
> h <- 2
> u1 <- 2
> u2 <- 3
> u <- c(u1, u2)
```

```
[1] 2 3
 > sigmaK <- sqrt(n * (n - 1) * (2 * n + 5)/18 - (sum(t * (t - 1) *
       (2 * t + 5)) + sum(u * (u - 1) * (2 * u + 5)))/18 + (sum(t * 1))
       (t-1) * (t-2)) * sum(u * (u-1) * (u-2)))/(9 * n *
       (n-1) * (n-2)) + (sum(t * (t-1)) * sum(u * (u-1)))/(2 *
       n * (n - 1)))
 > sigmaK
 [1] 4.711688
 > z <- num/sigmaK</pre>
 > z
 [1] 1.485667
 > res <- cor.test(x, y, alternative = "two.sided", method = "kendall",</pre>
 + exact = F)
 > res$statistic
 1.485667
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.1373672
 > res$p.value
 [1] 0.1373672
 > cor(x, y, method = "kendall")
 [1] 0.5853694
 > res$estimate
       tau
 0.5853694
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < -c(1.2, 1.2, 3.4, 3.4, 4.5, 5.5, 5.5, 5.6, 6.6, 6.6)
 > y <- c(1.3, 1.3, 1.3, 4.5, 5.6, 6.7, 6.7, 6.7, 8.8, 8.8, 9)
 > matrice <- matrix(data = 0, nrow = n - 1, ncol = n, byrow = F)
 > for (i in 1:(n - 1)) for (j in (i + 1):n) matrice[i, j] <- sign((x[j] -
       x[i]) * (y[j] - y[i]))
 > num <- sum(matrice)</pre>
 > num
 [1] 45
 > table(x)
```

```
1.2 3.4 4.5 5 5.5 6.6
      2 2 1 1 2 3
> g < - 4
> t1 <- 2
> t2 <- 2
> t3 <- 2
> t4 <- 3
> t <- c(t1, t2, t3, t4)
[1] 2 2 2 3
> table(y)
1.3 4.5 5.6 6.7 8.8
    3 1 1 3 2
> h <- 3
> u1 <- 3
> u2 <- 3
> u3 <- 2
> u <- c(u1, u2, u3)
[1] 3 3 2
> sigmaK < - sqrt(n * (n - 1) * (2 * n + 5)/18 - (sum(t * (t - 1) * (n - 1
                    (2 * t + 5)) + sum(u * (u - 1) * (2 * u + 5)))/18 + (sum(t * 1))
                     (t-1) * (t-2)) * sum(u * (u-1) * (u-2)))/(9 * n *
                    (n-1) * (n-2)) + (sum(t * (t-1)) * sum(u * (u-1)))/(2 *
                   n * (n - 1))
> sigmaK
[1] 12.27891
> z <- num/sigmaK
> z
[1] 3.664819
> res <- cor.test(x, y, alternative = "two.sided", method = "kendall",</pre>
+ exact = F)
> res$statistic
3.664819
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 0.0002475132
> res$p.value
[1] 0.0002475132
```

```
> cor(x, y, method = "kendall")
     [1] 0.9278844
    > res$estimate
                       t.au
    0.9278844
    > res$alternative
     [1] "two.sided"
• Example 3:
    > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
    > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
    > n <- 8
    > matrice < - matrix(data = 0, nrow = n - 1, ncol = n, byrow = F)
    > for (i in 1:(n - 1)) for (j in (i + 1):n) matrice[i, j] <- sign((x[j] -
                   x[i]) * (y[j] - y[i]))
    > num <- sum(matrice)</pre>
    > num
    [1] 18
    > table(x)
    1.1 2.3 3.4 4.5 5.6 6.7 8.9
        1 1 1 1 1 2 1
    > g <- 1
    > t1 <- 2
    > t <- c(t1)
    > t
    [1] 2
    > table(y)
       1.5 6.4 7.8 8.6 8.8 8.86 9.6
          1 1 1 2 1 1
    > h <- 1
    > u1 <- 2
    > u <- c(u1)
    > u
    [1] 2
    > sigmaK <- sqrt(n * (n - 1) * (2 * n + 5)/18 - (sum(t * (t - 1) *
                    (2 * t + 5)) + sum(u * (u - 1) * (2 * u + 5)))/18 + (sum(t * 2 * 4 * 4 * 5)))/18 + (sum(t * 4 * 4 * 4 * 5)))/18 + (sum(t * 4 * 6 * 6))/18 + (sum(t * 4 * 6))/18 + (sum(t * 6 * 6
                        (t-1) * (t-2)) * sum(u * (u-1) * (u-2)))/(9 * n *
                       (n-1) * (n-2)) + (sum(t * (t-1)) * sum(u * (u-1)))/(2 *
                     n * (n - 1))
    > sigmaK
     [1] 7.960468
```

```
> z <- num/sigmaK</pre>
[1] 2.261174
> res <- cor.test(x, y, alternative = "two.sided", method = "kendall",</pre>
+ exact = F)
> res$statistic
2.261174
> p.value <- 2 * pnorm(-abs(z))</pre>
> p.value
[1] 0.02374851
> res$p.value
[1] 0.02374851
> cor(x, y, method = "kendall")
[1] 0.6666667
> res$estimate
      tau
0.6666667
> res$alternative
[1] "two.sided"
```

Test Z con una retta di regressione

estimate coefficiente di correlazione

alternative ipotesi alternativa

null.value valore di ρ_0

• Formula:

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z))$

 $z = \frac{\arctan h(r_{xy}) - \arctan h(\rho_0)}{\frac{1}{\sqrt{n-3}}}$ $\operatorname{dove} \ \operatorname{arctanh}(x) = \frac{1}{2} \log \left(\frac{1+x}{1-x}\right)$ $\operatorname{p.value}$ $\operatorname{conf.int}$ $\tanh \left(\frac{1}{2} \log \left(\frac{1+r_{xy}}{1-r_{xy}}\right) \mp z_{1-\alpha/2} \frac{1}{\sqrt{n-3}}\right)$ $\operatorname{dove} \ \tanh(x) = \frac{\sinh(x)}{\cosh(x)} = \frac{e^{2x}-1}{e^{2x}+1}$ estimate r_{xy} $\operatorname{null.value}$

```
> x \leftarrow c(1.2, 3.4, 5.6, 7.4, 3.2, 5.6, 7.8, 8.9)
> y <- c(1.5, 6.7, 8.5, 4.2, 3.7, 8.8, 9.1, 10.2)
> n <- 8
> r <- cor(x, y)
> r
[1] 0.7354548
> res <- cor2.test(r1 = r, n1 = n, alternative = "two.sided", rho = 0.8,
+ conf.level = 0.95)
> rho0 <- 0.8
> z \leftarrow (atanh(r) - atanh(rho0))/(1/sqrt(n - 3))
[1] -0.3535357
> res$statistic
-0.3535357
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 0.7236869
> res$p.value
[1] 0.7236869
> lower <- tanh(0.5 * log((1 + r)/(1 - r)) - qnorm(1 - 0.05/2)/sqrt(n -
+ 3))
> upper <- tanh(0.5 * log((1 + r)/(1 - r)) + qnorm(1 - 0.05/2)/sqrt(n -
     3))
> c(lower, upper)
```

```
[1] 0.0638966 0.9485413
 > res$conf.int
 [1] 0.0638966 0.9485413
 attr(,"conf.level")
 [1] 0.95
 > r
 [1] 0.7354548
 > res$estimate
 0.7354548
 > rho0
 [1] 0.8
 > res$null.value
 corr coef
       0.8
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x < -c(1, 2, 2, 4, 3, 3)
 > y < -c(6, 6, 7, 7, 7, 9)
 > n < -6
 > r <- cor(x, y)
 > res <- cor2.test(r1 = r, n1 = n, alternative = "two.sided", rho = 0.6,
 + conf.level = 0.95)
 > rho0 <- 0.6
 > z <- (atanh(r) - atanh(rho0))/(1/sqrt(n - 3))
 > z
 [1] -0.1970069
 > res$statistic
 -0.1970069
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.8438221
 > res$p.value
 [1] 0.8438221
```

```
> lower <- tanh(atanh(r) - qnorm(1 - 0.05/2)/sqrt(n - 3))
 > upper <- tanh(atanh(r) + qnorm(1 - 0.05/2)/sqrt(n - 3))
 > c(lower, upper)
 [1] -0.5021527 0.9367690
 > res$conf.int
 [1] -0.5021527 0.9367690
 attr(,"conf.level")
 [1] 0.95
 > r
 [1] 0.522233
 > res$estimate
        r
 0.522233
 > rho0
 [1] 0.6
 > res$null.value
 corr coef
       0.6
 > res$alternative
 [1] "two.sided"
• Example 3:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > r < -cor(x, y)
 > res <- cor2.test(r1 = r, n1 = n, alternative = "two.sided", rho = 0.77,
       conf.level = 0.95)
 > rho0 <- 0.77
 > z <- (atanh(r) - atanh(rho0))/(1/sqrt(n - 3))
 > z
 [1] -0.1529148
 > res$statistic
 -0.1529148
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.8784655
 > res$p.value
```

```
[1] 0.8784655
> lower <- tanh(atanh(r) - qnorm(1 - 0.05/2)/sqrt(n - 3))
> upper <- tanh(atanh(r) + qnorm(1 - 0.05/2)/sqrt(n - 3))
> c(lower, upper)
[1] 0.07527696 0.94967566
> res$conf.int
[1] 0.07527696 0.94967566
attr(,"conf.level")
[1] 0.95
> r
[1] 0.740661
> res$estimate
0.740661
> rho0
[1] 0.77
> res$null.value
corr coef
     0.77
> res$alternative
[1] "two.sided"
```

Test Z con due rette di regressione

p.value p-value

livello $1 - \alpha$

conf.int intervallo di confidenza per la differenza tra i coefficienti di correlazione incogniti a

statistic valore empirico della statistica Z

estimate coefficienti di correlazione alternative ipotesi alternativa

• Formula:

statistic

$$z \, = \, \frac{ \operatorname{arctanh}(r_{x_1 y_1}) - \operatorname{arctanh}(r_{x_2 y_2})}{\sqrt{\frac{1}{n_1 - 3} + \frac{1}{n_2 - 3}}}$$

$$\mathbf{dove} \quad \mathrm{arctanh}(x) \, = \, \frac{1}{2} \, \log \left(\frac{1+x}{1-x} \right)$$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z))$

conf.int

$$\tanh\left(\frac{1}{2}\log\left(\frac{1+r_{x_1y_1}}{1-r_{x_1y_1}}\right) - \frac{1}{2}\log\left(\frac{1+r_{x_2y_2}}{1-r_{x_2y_2}}\right) \mp z_{1-\alpha/2}\sqrt{\frac{1}{n_1-3} + \frac{1}{n_2-3}}\right)$$

dove
$$\tanh(x) = \frac{\sinh(x)}{\cosh(x)} = \frac{e^{2x} - 1}{e^{2x} + 1}$$

estimate

[1] 0.6209827

$$r_{x_1y_1} \qquad r_{x_2y_2}$$

```
> x1 <- c(1.2, 3.4, 5.6, 7.4, 3.2, 5.6, 7.8, 8.9)
> y1 <- c(1.5, 6.7, 8.5, 4.2, 3.7, 8.8, 9.1, 10.2)
> n1 <- 8
> r1 <- cor(x1, y1)
> r1
[1] 0.7354548
> x2 < -c(1, 2, 2, 4, 3, 3)
> y2 <- c(6, 6, 7, 7, 7, 9)
> n2 < -6
> r2 <- cor(x2, y2)
> r2
[1] 0.522233
> res <- cor2.test(r1, n1, r2, n2, alternative = "two.sided", conf.level = 0.95)
> z <- (atanh(r1) - atanh(r2))/sqrt(1/(n1 - 3) + 1/(n2 - 3))
> z
[1] 0.4944581
> res$statistic
0.4944581
> p.value <-2 * pnorm(-abs(z))
> p.value
```

```
> res$p.value
 [1] 0.6209827
 > lower <- tanh(atanh(r1) - atanh(r2) - qnorm(1 - 0.05/2) * sqrt(1/(n1 -
 + 3) + 1/(n2 - 3)))
 > upper <- tanh(atanh(r1) - atanh(r2) + qnorm(1 - 0.05/2) * sqrt(1/(n1 -
       3) + 1/(n2 - 3))
 > c(lower, upper)
 [1] -0.7895570 0.9460192
 > res$conf.int
 [1] -0.7895570 0.9460192
 attr(,"conf.level")
 [1] 0.95
 > c(r1, r2)
 [1] 0.7354548 0.5222330
 > res$estimate
 0.7354548 0.5222330
 > res$alternative
 [1] "two.sided"
• Example 2:
 > x1 < -c(1.2, 5.6, 7.4, 6.78, 6.3, 7.8, 8.9)
 > y1 <- c(2.4, 6.4, 8.4, 8.5, 8.54, 8.7, 9.7)
 > n1 <- 7
 > r1 <- cor(x1, y1)
 > r1
 [1] 0.9755886
 > x2 < -c(3.7, 8.6, 9.9, 10.4)
 > y2 <- c(5.8, 9.7, 12.4, 15.8)
 > n2 <- 4
 > r2 <- cor(x2, y2)
 > r2
 [1] 0.9211733
 > res <- cor2.test(r1, n1, r2, n2, alternative = "two.sided", conf.level = 0.95)
 > z < - (atanh(r1) - atanh(r2))/sqrt(1/(n1 - 3) + 1/(n2 - 3))
 [1] 0.5367157
 > res$statistic
 0.5367157
```

```
> p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.591464
 > res$p.value
 [1] 0.591464
 > lower <- tanh(atanh(r1) - atanh(r2) - qnorm(1 - 0.05/2) * sqrt(1/(n1 -
 + 3) + 1/(n2 - 3)))
 > upper <- tanh(atanh(r1) - atanh(r2) + qnorm(1 - 0.05/2) * sqrt(1/(n1 -
       3) + 1/(n2 - 3))
 > c(lower, upper)
 [1] -0.9203392 0.9925038
 > res$conf.int
 [1] -0.9203392 0.9925038
 attr(,"conf.level")
 [1] 0.95
 > c(r1, r2)
 [1] 0.9755886 0.9211733
 > res$estimate
        r1
                 r2
 0.9755886 0.9211733
 > res$alternative
 [1] "two.sided"
• Example 3:
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y1 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > n1 <- 8
 > r1 <- cor(x1, y1)
 > r1
 [1] 0.8260355
 > x2 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y2 <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n2 < - 8
 > r2 <- cor(x2, y2)
 > r2
 [1] 0.8531061
 > res <- cor2.test(r1, n1, r2, n2, alternative = "two.sided", conf.level = 0.95)
 > z <- (atanh(r1) - atanh(r2))/sqrt(1/(n1 - 3) + 1/(n2 - 3))
 [1] -0.1453518
```

```
> res$statistic
-0.1453518
> p.value <- 2 * pnorm(-abs(z))</pre>
> p.value
[1] 0.8844331
> res$p.value
[1] 0.8844331
> lower <- tanh(atanh(r1) - atanh(r2) - qnorm(1 - 0.05/2) * sqrt(1/(n1 -
+ 3) + 1/(n2 - 3)))
> upper <- tanh(atanh(r1) - atanh(r2) + qnorm(1 - 0.05/2) * sqrt(1/(n1 -
     3) + 1/(n2 - 3))
> c(lower, upper)
[1] -0.8696200 0.8169779
> res$conf.int
[1] -0.8696200 0.8169779
attr(,"conf.level")
[1] 0.95
> c(r1, r2)
[1] 0.8260355 0.8531061
> res$estimate
      r1
0.8260355 0.8531061
> res$alternative
[1] "two.sided"
```

9.2 Test di ipotesi sulla autocorrelazione

statistic valore empirico della statistica χ^2

Test di Box - Pierce

```
Package: stats
Sintassi: Box.test()
Input:

x vettore numerico di dimensione n
lag il valore d del ritardo

Output:
```

parameter gradi di libertà

p.value p-value

• Formula:

statistic
$$c=n\sum_{k=1}^d\hat\rho^2(k)$$

$${\rm dove}\quad \hat\rho(k)=\frac{\sum_{t=1}^{n-k}(x_t-\bar x)\,(x_{t+k}-\bar x)}{\sum_{t=1}^n(x_t-\bar x)^2}\quad\forall\,k=1,2,\ldots,d$$
 parameter
$$df=d$$
 p.value
$$P(\chi^2_{df}\geq c)$$

• Example 1:

```
> x < -c(1.2, 3.4, 5.6, 7.4, 3.2, 5.6, 7.8, 8.9)
> n <- 8
> d < - 3
> autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])</pre>
> autocorr <- autocorr[-1]</pre>
> autocorr
[1] 0.2562830 -0.1947304 -0.1413042
> c <- n * sum(autocorr^2)</pre>
[1] 0.9885422
> Box.test(x, lag = d)$statistic
X-squared
0.9885422
> d
[1] 3
> Box.test(x, lag = d)$parameter
df
3
> p.value < 1 - pchisq(c, df = d)
> p.value
[1] 0.8040244
> Box.test(x, lag = d)$p.value
[1] 0.8040244
```

```
> x < -c(1.2, 2.6, 3.8, 4.4, 5.2)
 > n <- 5
 > d < - 2
 > autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])</pre>
 > autocorr <- autocorr[-1]</pre>
 > autocorr
 [1] 0.36612642 -0.09918963
 > c <- n * sum(autocorr^2)</pre>
 [1] 0.7194357
 > Box.test(x, lag = d)$statistic
 X-squared
 0.7194357
 > d
 [1] 2
 > Box.test(x, lag = d)$parameter
 df
  2
 > p.value <- 1 - pchisq(c, df = d)
 > p.value
 [1] 0.6978732
 > Box.test(x, lag = d)$p.value
 [1] 0.6978732
• Example 3:
 > x < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > n <- 8
 > d < - 2
 > autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])</pre>
 > autocorr <- autocorr[-1]</pre>
 > autocorr
 [1] 0.2271066 -0.2233210
 > c <- n * sum(autocorr^2)</pre>
 > c
 [1] 0.8115975
 > Box.test(x, lag = d)$statistic
 X-squared
 0.8115975
 > d
```

```
[1] 2
> Box.test(x, lag = d) $parameter

df
2
> p.value <- 1 - pchisq(c, df = d)
> p.value

[1] 0.6664443
> Box.test(x, lag = d) $p.value

[1] 0.6664443
```

Test di Ljung - Box

• Package: stats

• Sintassi: Box.test()

• Input:

x vettore numerico di dimensione n lag il valore d del ritardo type = "Ljung-Box"

• Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value

• Formula:

statistic $c=n\left(n+2\right)\sum_{k=1}^d\frac{1}{n-k}\hat{\rho}^2(k)$ $\text{dove}\quad \hat{\rho}(k)=\frac{\sum_{t=1}^{n-k}(x_t-\bar{x})\left(x_{t+k}-\bar{x}\right)}{\sum_{t=1}^n(x_t-\bar{x})^2}\quad\forall\,k=1,\,2,\,\ldots,\,d$ parameter df=d p.value $P(\chi_{df}^2\geq c)$

```
> x <- c(1.2, 3.4, 5.6, 7.4, 3.2, 5.6, 7.8, 8.9)
> n <- 8
> d <- 3
> autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])
> autocorr
<- autocorr[-1]
> autocorr
[1] 0.2562830 -0.1947304 -0.1413042
```

```
> c <- n * (n + 2) * sum(autocorr^2/(n - 1:d))
 [1] 1.575709
 > Box.test(x, lag = d, type = "Ljung-Box")$statistic
 X-squared
  1.575709
 > d
 [1] 3
 > Box.test(x, lag = d, type = "Ljung-Box")$parameter
 df
  3
 > p.value < 1 - pchisq(c, df = d)
 > p.value
 [1] 0.6649102
 > Box.test(x, lag = d, type = "Ljung-Box")$p.value
 [1] 0.6649102
• Example 2:
 > x < -c(1.2, 2.6, 3.8, 4.4, 5.2)
 > n <- 5
 > d < - 2
 > autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])</pre>
 > autocorr <- autocorr[-1]</pre>
 > autocorr
 [1] 0.36612642 -0.09918963
 > c <- n * (n + 2) * sum(autocorr^2/(n - 1:d))
 > C
 [1] 1.287708
 > Box.test(x, lag = d, type = "Ljung-Box")$statistic
 X-squared
  1.287708
 > d
 [1] 2
 > Box.test(x, lag = d, type = "Ljung-Box")$parameter
 df
  2
```

```
> p.value < 1 - pchisq(c, df = d)
 > p.value
 [1] 0.5252641
 > Box.test(x, lag = d, type = "Ljung-Box")$p.value
 [1] 0.5252641
• Example 3:
 > x < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > n <- 8
 > d < - 2
 > autocorr <- as.vector(acf(x, lag.max = d, plot = F)[[1]])</pre>
 > autocorr <- autocorr[-1]</pre>
 > autocorr
 [1] 0.2271066 -0.2233210
 > c <- n * (n + 2) * sum(autocorr^2/(n - 1:d))
 > C
 [1] 1.254420
 > Box.test(x, lag = d, type = "Ljung-Box")$statistic
 X-squared
  1.254420
 > d
 [1] 2
 > Box.test(x, lag = d, type = "Ljung-Box")$parameter
 df
  2
 > p.value < 1 - pchisq(c, df = d)
 > p.value
 [1] 0.5340799
 > Box.test(x, lag = d, type = "Ljung-Box") $p.value
 [1] 0.5340799
```

Capitolo 10

Test di ipotesi non parametrici

10.1 Simbologia

- dimensione del campione j-esimo: $n_j \quad \forall j = 1, 2, ..., k$
- media aritmetica del campione j-esimo: $\bar{x}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} x_{ij} \quad \forall j = 1, 2, \dots, k$
- varianza nel campione j-esimo: $s_j^2 = \frac{1}{n_j-1} \sum_{i=1}^{n_j} (x_{ij} \bar{x}_j)^2 \quad \forall j = 1, 2, \ldots, k$
- varianza pooled: $s_P^2 = \sum_{j=1}^k \left(n_j 1\right) s_j^2 / \left(n k\right)$
- somma dei ranghi nel campione j-esimo: $R_j \quad \forall j=1,\,2,\,\ldots,\,k$
- media dei ranghi nel campione j-esimo: $\bar{R}_j \quad \forall j = 1, 2, ..., k$
- media dei ranghi nel campione di dimensione n: \bar{R}
- ties nel campione di dimensione n: $t_j \quad \forall j=1,2,\ldots,g \qquad \sum_{j=1}^g t_j=n \qquad 1\leq g\leq n$

10.2 Test di ipotesi sulla mediana con uno o due campioni

Test esatto Wilcoxon signed rank

- Package: stats
- Sintassi: wilcox.test()
- Input:

```
x vettore numerico di dimensione n mu il valore di Q_{0.5}(x)_{\mid H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa exact = TRUE
```

• Output:

```
statistic valore empirico della statistica V p.value p-value null.value il valore di Q_{0.5}(x)_{\mid H_0} alternative ipotesi alternativa
```

• Formula:

statistic v

alternative	less	greater	two.sided
p.value	$P(V \le v)$	$P(V \ge v)$	$2 \min (P(V \le v), P(V \ge v))$

null.value

 $Q_{0.5}(x)_{|H_0}$

```
• Example 1:
```

```
> x < -c(-0.1, -0.2, 0.7, 0.8, -1.2, -1.6, 2, 3.4, 3.7)
> n < -9
> mu <- 3.3
> x - mu
[1] -3.4 -3.5 -2.6 -2.5 -4.5 -4.9 -1.3 0.1 0.4
> xx <- rank(abs(x - mu)) * sign(x - mu)
> xx
[1] -6 -7 -5 -4 -8 -9 -3 1 2
> v \leftarrow sum(xx[xx > 0])
> v
[1] 3
> res1 <- wilcox.test(x, mu = 3.3, alternative = "less", exact = TRUE)
> res1$statistic
۲,7
3
> p.value.less <- psignrank(v, n)</pre>
> p.value.less
[1] 0.009765625
> res1$p.value
[1] 0.009765625
> p.value.greater <- 1 - psignrank(v - 1, n)</pre>
> p.value.greater
[1] 0.9941406
> res2 <- wilcox.test(x, mu = 3.3, alternative = "greater", exact = TRUE)
> res2$p.value
[1] 0.9941406
> p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.two.sided
[1] 0.01953125
> res3 <- wilcox.test(x, mu = 3.3, alternative = "two.sided", exact = TRUE)
> res3$p.value
[1] 0.01953125
```

```
> x <- c(3.8, 5.6, 1.8, 5, 2.4, 4.2, 7.3, 8.6, 9.1, 5.2)
 > n < -10
 > mu < - 6.3
 > x - mu
  [1] -2.5 -0.7 -4.5 -1.3 -3.9 -2.1 1.0 2.3 2.8 -1.1
 > xx <- rank(abs(x - mu)) * sign(x - mu)
 > xx
  [1] -7 -1 -10 -4 -9 -5 2 6 8 -3
 > v \leftarrow sum(xx[xx > 0])
 > 77
 [1] 16
 > res1 <- wilcox.test(x, mu = 6.3, alternative = "less", exact = TRUE)
 > res1$statistic
  V
 16
 > p.value.less <- psignrank(v, n)</pre>
 > p.value.less
 [1] 0.1376953
 > res1$p.value
 [1] 0.1376953
 > p.value.greater <- 1 - psignrank(v - 1, n)</pre>
 > p.value.greater
 [1] 0.883789
 > res2 <- wilcox.test(x, mu = 6.3, alternative = "greater", exact = TRUE)
 > res2$p.value
 [1] 0.883789
 > p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
 > p.value.two.sided
 [1] 0.2753906
 > res3 <- wilcox.test(x, mu = 6.3, alternative = "two.sided", exact = TRUE)
 > res3$p.value
 [1] 0.2753906
• Example 3:
 > x \leftarrow c(1.2, 3.4, 4.5, 6.4, 3, 4, 2.3, 8.8, 9.87, 12.34)
 > n <- 10
 > mu < - 2.7
 > xx <- rank(abs(x - mu)) * sign(x - mu)
 > xx
```

```
[1] -5 3 6 7 1 4 -2 8 9 10
> v \leftarrow sum(xx[xx > 0])
> v
[1] 48
> res1 <- wilcox.test(x, mu = 2.7, alternative = "less", exact = TRUE)
> res1$statistic
V
48
> p.value.less <- psignrank(v, n)</pre>
> p.value.less
[1] 0.9863281
> res1$p.value
[1] 0.9863281
> p.value.greater <- 1 - psignrank(v - 1, n)</pre>
> p.value.greater
[1] 0.01855469
> res2 <- wilcox.test(x, mu = 2.7, alternative = "greater", exact = TRUE)
> res2$p.value
[1] 0.01855469
> p.value.twosided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.twosided
[1] 0.03710938
> res3 <- wilcox.test(x, mu = 2.7, alternative = "two.sided", exact = TRUE)</pre>
> res3$p.value
[1] 0.03710938
```

• Note: Il vettore abs (x-mu) non deve contenere valori duplicati o nulli.

Test asintotico Wilcoxon signed rank

```
• Sintassi: wilcox.test()
```

• Package: stats

• Input:

```
\times vettore numerico di dimensione n
mu il valore di Q_{0.5}(x)_{\mid H_0}
alternative = "less" / "greater" / "two.sided" ipotesi alternativa
correct = TRUE / FALSE correzione di continuità di Yates
exact = FALSE
```

• Output:

```
statistic valore empirico della statistica V
p.value p-value
null.value il valore di Q_{0.5}(x)_{\mid H_0}
alternative ipotesi alternativa
```

• Formula:

p.value

statistic

less two.sided alternative greater

alternative less greater two.sided
$$\text{p.value} \qquad \Phi(z) \qquad 1 - \Phi(z) \qquad 2\,\Phi(-\,|\,z\,|)$$

$$z = \frac{v - \frac{m(m+1)}{4} + 0.5}{\left[\frac{1}{24} \left(m(m+1)(2m+1) - \frac{1}{2} \sum_{j=1}^{g} t_j(t_j^2 - 1)\right)\right]^{1/2}}$$

correct = TRUE

v

$$z = \frac{v - \frac{m(m+1)}{4}}{\left[\frac{1}{24} \left(m(m+1)(2m+1) - \frac{1}{2} \sum_{j=1}^{g} t_j(t_j^2 - 1)\right)\right]^{1/2}}$$

null.value

$$Q_{0.5}(x)_{|H_0}$$

```
> x <- c(4, 3, 4, 5, 2, 3, 4, 5, 4, 4, 5, 5, 4, 5, 4, 4, 3, 4,
    2, 4, 5, 5, 4, 4)
> mu <- 4
> xx <- (x - mu)[(x - mu) != 0]
 [1] -1 1 -2 -1 1 1 1 1 -1 -2 1 1
> m <- length(xx)
> m
[1] 12
> xx <- rank(abs(xx)) * sign(xx)
> xx
```

```
> v \leftarrow sum(xx[xx > 0])
 > v
 [1] 38.5
 > res <- wilcox.test(x, mu = 4, alternative = "less", correct = FALSE,
 + exact = FALSE)
 > res$statistic
   V
 38.5
 > table(rank(abs(xx)))
 5.5 11.5
  10 2
 > g <- 2
 > t1 <- 10
 > t2 <- 2
 > t <- c(t1, t2)
 > num <- v - m \star (m + 1)/4
 > den <- sqrt((m * (m + 1) * (2 * m + 1) - 0.5 * sum(t * (t^2 -
    1)))/24)
 > z <- num/den
 > p.value <- pnorm(z)</pre>
 > p.value
[1] 0.4832509
 > res$p.value
 [1] 0.4832509
• Example 2:
 > x <- c(4, 3, 4, 5, 2, 3, 4, 5, 4, 4, 5, 5, 4, 5, 4, 4, 3, 4,
 + 2, 4, 5, 5, 4, 4)
 > n < -24
 > mu <- 3
 > xx <- (x - mu) [(x - mu) != 0]
 > xx
 [1] 1 1 2 -1 1 2 1 1 2 2 1 2 1 1 1 -1 1 2 2 1 1
 > m <- length(xx)
 > m
 [1] 21
 > xx <- rank(abs(xx)) * sign(xx)
 > XX
  [1] 7.5 7.5 18.0 -7.5 7.5 18.0 7.5 7.5 18.0 18.0 7.5 18.0 7.5 7.5
 [16] -7.5 7.5 18.0 18.0 7.5 7.5
 > v \leftarrow sum(xx[xx > 0])
 > V
```

```
[1] 216
 > res <- wilcox.test(x, mu = 3, alternative = "less", correct = TRUE,
 + exact = FALSE)
 > res$statistic
  V
 216
 > table(rank(abs(xx)))
 7.5 18
  14 7
 > g <- 2
 > t1 <- 14
 > t2 <- 7
 > t <- c(t1, t2)
 > num <- v - m * (m + 1)/4 + 0.5
 > den <- sqrt((m * (m + 1) * (2 * m + 1) - 0.5 * sum(t * (t^2 -
       1)))/24)
 > z <- num/den
 > p.value <- pnorm(z)</pre>
 > p.value
 [1] 0.999871
 > res$p.value
 [1] 0.999871
• Example 3:
 > x < -c(1.2, 3.4, 4.5, 6.4, 3, 4, 2.3, 8.8, 9.87, 12.34)
 > n <- 10
 > mu <- 2.7
 > xx <- (x - mu)[(x - mu) != 0]
 > xx <- c(-1.5, 0.7, 1.8, 3.7, 0.3, 1.3, -0.4, 6.1, 7.17, 9.64)
 > m <- length(xx)
 > m
 [1] 10
 > xx <- rank(abs(xx)) * sign(xx)
 > xx
  [1] -5 3 6 7 1 4 -2 8 9 10
 > v \leftarrow sum(xx[xx > 0])
 > V
 [1] 48
 > res <- wilcox.test(x, mu = 2.7, alternative = "less", correct = TRUE,
 + exact = FALSE)
 > res$statistic
  V
 48
```

```
1 2 3 4 5 6 7 8 9 10
        1 1 1 1 1 1 1 1
    > q < -10
    > t1 <- 1
    > t2 <- 1
    > t3 <- 1
    > t4 <- 1
    > t5 <- 1
    > t6 <- 1
    > t7 <- 1
    > t8 <- 1
    > t9 <- 1
    > t10 <- 1
    > t <- c(t1, t2, t3, t4, t5, t6, t7, t8, t9, t10)
    > num < -v - m * (m + 1)/4 + 0.5
    > den <- sqrt((m * (m + 1) * (2 * m + 1) - 0.5 * sum(t * (t^2 -
           1)))/24)
    > z <- num/den
    > p.value <- pnorm(z)</pre>
     > p.value
    [1] 0.9838435
    > res$p.value
     [1] 0.9838435
Test esatto di Mann - Whitney
   • Package: stats
   • Sintassi: wilcox.test()
   • Input:
         	imes vettore numerico di dimensione n_x
         y vettore numerico di dimensione n_y
         mu il valore di \left(\,Q_{0.5}(x)-Q_{0.5}(y)\,
ight)_{\mid\,H_0}
         alternative = "less" / "greater" / "two.sided" ipotesi alternativa
         exact = TRUE
   • Output:
         statistic valore empirico della statistica {\it W}
         p.value p-value
         null.value il valore di (Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}
```

alternative ipotesi alternativa

• Formula:

statistic

> table(rank(abs(xx)))

w

p.value

alternative	less	greater	two.sided
p.value	$P(W \le w)$	$P(W \ge w)$	$2 \min (P(W \le w), P(W \ge w))$

null.value

$$(Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}$$

```
> x < -c(1.2, 3.4, 5.4, -5.6, 7.3, 2.1)
> nx <- 6
> y < -c(-1.1, -0.1, 0.9, 1.9, 2.9, 3.9, 4.99)
> ny < -7
> mu < -2.1
> c(x, y + mu)
[1] 1.20 3.40 5.40 -5.60 7.30 2.10 -3.20 -2.20 -1.20 -0.20 0.80 1.80
[13] 2.89
> Rx <- sum(rank(c(x, y + mu))[1:nx])
> Rx
[1] 53
> w <- Rx - nx * (nx + 1)/2
> W
[1] 32
> res1 <- wilcox.test(x, y, mu = -2.1, alternative = "less", exact = TRUE)
> res1$statistic
W
32
> p.value.less <- pwilcox(w, nx, ny)</pre>
> p.value.less
[1] 0.9493007
> res1$p.value
[1] 0.9493007
> p.value.greater <- 1 - pwilcox(w - 1, nx, ny)</pre>
> p.value.greater
[1] 0.06876457
> res2 <- wilcox.test(x, y, mu = -2.1, alternative = "greater",
+ exact = TRUE)
> res2$p.value
[1] 0.06876457
> p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.two.sided
```

```
> x <- c(33.3, 30.1, 38.62, 38.94, 42.63, 41.96, 46.3, 43.25)
> y <- c(31.62, 46.33, 31.82, 40.21, 45.72, 39.8, 45.6, 41.25)
> ny <- 8
> mu <- 1.1
> c(x, y + mu)
[1] 33.30 30.10 38.62 38.94 42.63 41.96 46.30 43.25 32.72 47.43 32.92 41.31
[13] 46.82 40.90 46.70 42.35
> Rx <- sum(rank(c(x, y + mu))[1:nx])
> Rx
[1] 61
> w <- Rx - nx \star (nx + 1)/2
> W
[1] 25
> res1 <- wilcox.test(x, y, mu = 1.1, alternative = "less", exact = TRUE)
> res1$statistic
W
25
> p.value.less <- pwilcox(w, nx, ny)</pre>
> p.value.less
[1] 0.2526807
> res1$p.value
[1] 0.2526807
> p.value.greater <- 1 - pwilcox(w - 1, nx, ny)</pre>
> p.value.greater
[1] 0.7790987
> res2 <- wilcox.test(x, y, mu = 1.1, alternative = "greater",</pre>
+ exact = TRUE)
> res2$p.value
[1] 0.7790987
> p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.two.sided
[1] 0.5053613
> res3 <- wilcox.test(x, y, mu = 1.1, alternative = "two.sided",</pre>
+ exact = TRUE)
> res3$p.value
[1] 0.5053613
```

• Example 3:

```
> x < -c(4, 2.3, 8.8, 9.87, 12.34, 1.4)
> nx <- 6
> y <- c(6.4, 9.6, 8.86, 7.8, 8.6, 8.7, 1.1)
> ny < -7
> mu <- 2.3
> c(x, y + mu)
[1] 4.00 2.30 8.80 9.87 12.34 1.40 8.70 11.90 11.16 10.10 10.90 11.00
[13] 3.40
> Rx <- sum(rank(c(x, y + mu))[1:nx])
> Rx
[1] 33
> w <- Rx - nx * (nx + 1)/2
> w
[1] 12
> res1 <- wilcox.test(x, y, mu = 2.3, alternative = "less", exact = TRUE)
> res1$statistic
W
12
> p.value.less <- pwilcox(w, nx, ny)</pre>
> p.value.less
[1] 0.1171329
> res1$p.value
[1] 0.1171329
> p.value.greater <- 1 - pwilcox(w - 1, nx, ny)</pre>
> p.value.greater
[1] 0.9096737
> res2 <- wilcox.test(x, y, mu = 2.3, alternative = "greater",
+ exact = TRUE)
> res2$p.value
[1] 0.9096737
> p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.two.sided
[1] 0.2342657
> res3 <- wilcox.test(x, y, mu = 2.3, alternative = "two.sided",</pre>
+ exact = TRUE)
> res3$p.value
[1] 0.2342657
```

• Note: Il vettore c (x, y+mu) non deve contenere valori duplicati.

Test asintotico di Mann - Whitney

```
• Package: stats
```

• **Sintassi:** wilcox.test()

• Input:

```
x vettore numerico di dimensione n_x y vettore numerico di dimensione n_y mu il valore di (Q_{0.5}(x)-Q_{0.5}(y))_{|H_0} alternative = "less" / "greater" / "two.sided" ipotesi alternativa correct = TRUE / FALSE correzione di continuità di Yates exact = FALSE
```

• Output:

```
statistic valore empirico della statistica W p.value p-value null.value il valore di (Q_{0.5}(x)-Q_{0.5}(y))_{|H_0} alternative ipotesi alternativa
```

• Formula:

statistic

w

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(- z)$

$$z = \frac{w - \frac{n_x n_y}{2} + 0.5}{\left[\frac{n_x n_y}{12} \left(n_x + n_y + 1 - \frac{\sum_{j=1}^g t_j (t_j^2 - 1)}{(n_x + n_y) (n_x + n_y - 1)}\right)\right]^{1/2}}$$

$$z = \frac{w - \frac{n_x n_y}{2}}{\left[\frac{n_x n_y}{12} \left(n_x + n_y + 1 - \frac{\sum_{j=1}^g t_j (t_j^2 - 1)}{(n_x + n_y) (n_x + n_y - 1)}\right)\right]^{1/2}}$$

null.value

$$(Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}$$

```
> x <- c(-1, 1, -2, -1, 1, 1, 1, 1, -1, -2, 1, 1)
> nx <- 12
> y <- c(1, 1, 2, 3, 4, 5, 3, 2, 1)
> ny <- 9
> mu <- -4
> Rx <- sum(rank(c(x, y + mu))[1:nx])
> Rx

[1] 163.5

> w <- Rx - nx * (nx + 1)/2
> w

[1] 85.5
```

```
> res <- wilcox.test(x, y, mu = -4, alternative = "less", correct = TRUE,
 + exact = FALSE)
 > res$statistic
 85.5
 > table(rank(c(x, y + mu)))
    2 5.5 10 13 17.5
            5 1 8
    3 4
 > g <- 4
 > t1 <- 3
 > t2 <- 4
 > t3 <- 5
 > t4 <- 8
 > t <- c(t1, t2, t3, t4)
 > num <- w - nx * ny/2 + 0.5
 > den <- sqrt(nx * ny/12 * (nx + ny + 1 - sum(t * (t^2 - 1))/((nx +
 + ny) * (nx + ny - 1))))
 > z <- num/den
 > p.value <- pnorm(z)</pre>
 > p.value
 [1] 0.9910242
 > res$p.value
 [1] 0.9910242
• Example 2:
 > x < -c(33.3, 30.1, 38.62, 38.94, 42.63, 41.96, 46.3, 43.25)
 > y <- c(31.62, 46.33, 31.82, 40.21, 45.72, 39.8, 45.6, 41.25)
 > ny <- 8
 > mu <- 4
 > Rx <- sum(rank(c(x, y + mu))[1:nx])
 [1] 51
 > w < - Rx - nx * (nx + 1)/2
 [1] 15
 > res <- wilcox.test(x, y, mu = 4, alternative = "less", correct = FALSE,
 + exact = FALSE)
 > res$statistic
  W
 15
 > table(rank(x, y + mu))
 1 2 3 4 5 6 7 8
 1 1 1 1 1 1 1 1
```

```
> g <- 8
> t1 <- 1
> t2 <- 1
> t3 <- 1
> t4 <- 1
> t5 <- 1
> t6 <- 1
> t7 <- 1
> t8 <- 1
> t <- c(t1, t2, t3, t4, t5, t6, t7, t8)
> num <- w - nx * ny/2
> den <- sqrt(nx * ny/12 * (nx + ny + 1 - sum(t * (t^2 - 1))/((nx +
+ ny) * (nx + ny - 1))))
> z <- num/den
> p.value <- pnorm(z)</pre>
> p.value
[1] 0.03710171
> res$p.value
[1] 0.03710171
```

```
> x < -c(4, 2.3, 8.8, 9.87, 12.34, 1.4)
> nx <- 6
> y < -c(6.4, 9.6, 8.86, 7.8, 8.6, 8.7, 1.1)
> ny <- 7
> mu <- 2.3
> Rx <- sum(rank(c(x, y + mu))[1:nx])
> Rx
[1] 33
> w <- Rx - nx * (nx + 1)/2
[1] 12
> res <- wilcox.test(x, y, mu = 2.3, alternative = "less", correct = TRUE,
     exact = FALSE)
> res$statistic
W
12
> table(rank(c(x, y + mu)))
 1 2 3 4 5 6 7 8 9 10 11 12 13
   1 1 1 1 1 1 1 1
                           1 1
                                 1
> q < -13
> t1 <- 1
> t2 <- 1
> t3 <- 1
> t4 <- 1
> t5 <- 1
> t6 <- 1
> t7 <- 1
> t8 <- 1
> t9 <- 1
> t10 <- 1
> t11 <- 1
> t12 <- 1
> t13 <- 1
> t <- c(t1, t2, t3, t4, t5, t6, t7, t8, t9, t10, t11, t12, t13)
> num <- w - nx * ny/2 + 0.5
> den <- sqrt(nx * ny/12 * (nx + ny + 1 - sum(t * (t^2 - 1))/((nx +
     ny) * (nx + ny - 1)))
> z <- num/den
> p.value <- pnorm(z)</pre>
> p.value
[1] 0.1123193
> res$p.value
[1] 0.1123193
```

Test esatto Wilcoxon signed rank per dati appaiati

```
• Sintassi: wilcox.test()
• Input:

x vettore numerico di dimensione n
y vettore numerico di dimensione n
mu il valore di (Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}
alternative = "less" / "greater" / "two.sided" ipotesi alternativa exact = TRUE
paired = TRUE
```

• Output:

• Package: stats

```
statistic valore empirico della statistica V p.value p-value null.value il valore di (Q_{0.5}(x)-Q_{0.5}(y))_{|H_0} alternative ipotesi alternativa
```

• Formula:

```
statistic p.value
```

alternative	less	greater	two.sided
p.value	$P(V \le v)$	$P(V \ge v)$	$2 \min (P(V \le v), P(V \ge v))$

null.value

$$(Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}$$

v

```
> x < -c(-0.1, -0.2, 0.7, 0.8, -1.2, -1.6, 2, 3.4, 3.7)
> n < -9
> y < -c(1, 2, 3, 4, 5, 6, 7, 8, 9)
> mu < - -4
> x - y - mu
[1] 2.9 1.8 1.7 0.8 -2.2 -3.6 -1.0 -0.6 -1.3
> xy <- rank(abs(x - y - mu)) * sign(x - y - mu)
> xy
[1] 8 6 5 2 -7 -9 -3 -1 -4
> v \leftarrow sum(xy[xy > 0])
> V
[1] 21
> res1 <- wilcox.test(x, y, mu = -4, alternative = "less", exact = TRUE,
     paired = TRUE)
> res1$statistic
V
21
```

```
> p.value.less <- psignrank(v, n)</pre>
 > p.value.less
 [1] 0.4550781
 > res1$p.value
 [1] 0.4550781
 > p.value.greater <- 1 - psignrank(v - 1, n)</pre>
 > p.value.greater
 [1] 0.5898438
 > res2 <- wilcox.test(x, y, mu = -4, alternative = "greater", paired = TRUE,
 + exact = TRUE)
 > res2$p.value
 [1] 0.5898438
 > p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
 > p.value.two.sided
 [1] 0.9101562
 > res3 <- wilcox.test(x, y, mu = -4, alternative = "two.sided",</pre>
 + paired = TRUE, exact = TRUE)
 > res3$p.value
 [1] 0.9101562
• Example 2:
 > x <- c(33.3, 30.1, 38.62, 38.94, 42.63, 41.96, 46.3, 43.25)
 > n <- 8
 > y <- c(31.62, 46.33, 31.82, 40.21, 45.72, 39.8, 45.6, 41.25)
 > mu <- 1.1
 > x - y - mu
 \begin{bmatrix} 1 \end{bmatrix} 0.58 -17.33 5.70 -2.37 -4.19 1.06 -0.40 0.90
 > xy <- rank(abs(x - y - mu)) * sign(x - y - mu)
 > xy
 [1] 2 -8 7 -5 -6 4 -1 3
 > v \leftarrow sum(xy[xy > 0])
 [1] 16
 > res1 <- wilcox.test(x, y, mu = 1.1, alternative = "less", exact = TRUE,
 + paired = TRUE)
 > res1$statistic
  V
 16
```

```
> p.value.less <- psignrank(v, n)</pre>
 > p.value.less
 [1] 0.421875
 > res1$p.value
 [1] 0.421875
 > p.value.greater <- 1 - psignrank(v - 1, n)</pre>
 > p.value.greater
 [1] 0.6289062
 > res2 <- wilcox.test(x, y, mu = 1.1, alternative = "greater",
 + exact = TRUE, paired = TRUE)
 > res2$p.value
 [1] 0.6289062
 > p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
 > p.value.two.sided
 [1] 0.84375
 > res3 <- wilcox.test(x, y, mu = 1.1, alternative = "two.sided",</pre>
 + exact = TRUE, paired = TRUE)
 > res3$p.value
 [1] 0.84375
Example 3:
 > x < -c(4, 2.3, 8.8, 9.87, 12.34, 1.4)
 > n < -6
 > y <- c(6.4, 9.6, 8.86, 7.8, 8.6, 8.8)
 > mu < - 2.3
 > x - y - mu
 [1] -4.70 -9.60 -2.36 -0.23 1.44 -9.70
 > xy <- rank(abs(x - y - mu)) * sign(x - y - mu)
 > xy
 [1] -4 -5 -3 -1 2 -6
 > v \leftarrow sum(xy[xy > 0])
 > v
 [1] 2
 > res1 <- wilcox.test(x, y, mu = 2.3, alternative = "less", exact = TRUE,
 + paired = TRUE)
 > res1$statistic
 V
 2
```

```
> p.value.less <- psignrank(v, n)</pre>
> p.value.less
[1] 0.046875
> res2 <- wilcox.test(x, y, mu = 2.3, alternative = "less", exact = TRUE,
     paired = TRUE)
> res2$p.value
[1] 0.046875
> p.value.greater <- 1 - psignrank(v - 1, n)</pre>
> p.value.greater
[1] 0.96875
> res2$p.value
[1] 0.046875
> p.value.two.sided <- 2 * min(p.value.less, p.value.greater)</pre>
> p.value.two.sided
[1] 0.09375
> res3 <- wilcox.test(x, y, mu = 2.3, alternative = "two.sided",</pre>
     exact = TRUE, paired = TRUE)
> res3$p.value
[1] 0.09375
```

• Note: Il vettore abs (x-y-mu) non deve contenere valori duplicati o nulli.

Test asintotico Wilcoxon signed rank per dati appaiati

```
• Package: stats
• Sintassi: wilcox.test()
• Input:

    x vettore numerico di dimensione n
    y vettore numerico di dimensione n
    mu il valore di (Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}
    alternative = "less" / "greater" / "two.sided" ipotesi alternativa correct = TRUE / FALSE correzione di continuità di Yates
    exact = FALSE
    paired = TRUE

• Output:

    statistic valore empirico della statistica V
    p.value p-value
    null.value il valore di (Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}
    alternative ipotesi alternativa
```

• Formula:

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(- z)$

 ${\tt statistic} \\ v$

p.value

$$z = \frac{v - \frac{m\left(m+1\right)}{4} + 0.5}{\left[\frac{1}{24}\left(m\left(m+1\right)\left(2\,m+1\right) - \frac{1}{2}\,\sum_{j=1}^{g}\,t_{j}\left(t_{j}^{2} - 1\right)\right)\right]^{1/2}}$$

$$\boxed{\text{correct = FALSE}}$$

$$z = \frac{v - \frac{m\left(m+1\right)}{4}}{\left[\frac{1}{24}\left(m\left(m+1\right)\left(2\,m+1\right) - \frac{1}{2}\,\sum_{j=1}^{g}\,t_{j}\left(t_{j}^{2} - 1\right)\right)\right]^{1/2}}$$

null.value

$$(Q_{0.5}(x) - Q_{0.5}(y))_{|H_0}$$

```
> x \leftarrow c(4, 4, 3, 4, 2, 4, 5, 5, 4, 3.3)
> n < -10
> y <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> mu < -2
> xy <- (x - y - mu)[(x - y - mu) != 0]
> xy
[1] 5.0 4.0 2.0 2.0 -1.0 -1.0 -3.0 -4.7
> m <- length(xy)</pre>
> m
[1] 8
> xy <- rank(abs(xy)) * sign(xy)</pre>
> xy
[1] 8.0 6.0 3.5 3.5 -1.5 -1.5 -5.0 -7.0
> v \leftarrow sum(xy[xy > 0])
> V
[1] 21
> res <- wilcox.test(x, y, mu = -2, alternative = "less", correct = TRUE,
+ exact = FALSE, paired = TRUE)
> res$statistic
V
21
> table(rank(abs(xy)))
1.5 3.5 5 6 7
 2 2 1 1 1
                    1
```

```
> g <- 2
 > t1 <- 2
 > t2 <- 2
 > t <- c(t1, t2)
 > num <- v - m * (m + 1)/4 + 0.5
 > den <- sqrt(1/24 * (m * (m + 1) * (2 * m + 1) - 0.5 * sum(t *
      (t^2 - 1))
 > z <- num/den
 > p.value <- pnorm(z)</pre>
 > p.value
 [1] 0.6883942
 > res$p.value
 [1] 0.6883942
• Example 2:
 > x <- c(33.3, 30.1, 38.62, 38.94, 42.63, 41.96, 46.3, 43.25)
 > y <- c(31.62, 46.33, 31.82, 40.21, 45.72, 39.8, 45.6, 41.25)
 > mu <- 2
 > xy <- (x - y - mu)[(x - y - mu) != 0]
 > xy
 [1] -0.32 -18.23 4.80 -3.27 -5.09 0.16 -1.30
 > m <- length(xy)</pre>
 > m
 [1] 7
 > xy <- rank(abs(xy)) * sign(xy)</pre>
 > xy
 [1] -2 -7 5 -4 -6 1 -3
 > v <- sum(xy[xy > 0])
 > V
 [1] 6
 > res <- wilcox.test(x, y, mu = 2, alternative = "less", correct = FALSE,
 + exact = FALSE, paired = TRUE)
 > res$statistic
 V
 6
 > table(rank(abs(xy)))
 1 2 3 4 5 6 7
 1 1 1 1 1 1 1
```

```
> g < - 7
 > t1 <- 1
 > t2 <- 1
 > t3 <- 1
 > t4 <- 1
 > t5 <- 1
 > t6 <- 1
 > t7 <- 1
 > t <- c(t1, t2, t3, t4, t5, t6, t7)
 > \text{num} < - \text{v} - \text{m} * (\text{m} + 1)/4
 > den <- sqrt(1/24 * (m * (m + 1) * (2 * m + 1) - 0.5 * sum(t *
       (t^2 - 1)))
 > z <- num/den
 > p.value <- pnorm(z)</pre>
 > p.value
 [1] 0.08814819
 > res$p.value
 [1] 0.08814819
• Example 3:
 > x < -c(4.5, 6.4, 3, 4, 2.3, 8.8, 9.87, 12.34)
 > n <- 8
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > mu <- 2.3
 > xy <- (x - y - mu)[(x - y - mu) != 0]
 > xy
 [1] 0.70 -2.30 -8.90 -7.10 -8.86 -1.30 -1.03 1.44
 > m <- length(xy)</pre>
 > m
 [1] 8
 > xy <- rank(abs(xy)) * sign(xy)</pre>
 > xy
 [1] 1 -5 -8 -6 -7 -3 -2 4
 > v \leftarrow sum(xy[xy > 0])
 > V
 [1] 5
 > res <- wilcox.test(x, y, mu = 2.3, alternative = "less", correct = TRUE,
       exact = FALSE, paired = TRUE)
 > res$statistic
 V
 5
 > table(rank(abs(xy)))
 1 2 3 4 5 6 7 8
 1 1 1 1 1 1 1 1
```

```
> g < - 8
> t1 <- 1
> t2 <- 1
> t3 <- 1
> t4 <- 1
> t5 <- 1
> t6 <- 1
> t7 <- 1
> t8 <- 1
> t <- c(t1, t2, t3, t4, t5, t6, t7, t8)
> num <- v - m * (m + 1)/4 + 0.5
> den <- sqrt(1/24 * (m * (m + 1) * (2 * m + 1) - 0.5 * sum(t *
+ (t^2 - 1)))
> z <- num/den
> p.value <- pnorm(z)</pre>
> p.value
[1] 0.04002896
> res$p.value
[1] 0.04002896
```

10.3 Test di ipotesi sulla mediana con più campioni

Test di Kruskal - Wallis

• Package: stats

• Sintassi: kruskal.test()

• Input:

 \times vettore numerico di dimensione n

g fattore a k livelli di dimensione n

• Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value

• Formula:

statistic
$$c = \frac{1}{C} \frac{12}{n(n+1)} \sum_{i=1}^k n_i \left(\bar{R}_i - \bar{R} \right)^2 = \frac{1}{C} \frac{12}{n(n+1)} \sum_{i=1}^k \frac{R_i^2}{n_i} - 3(n+1)$$

$$\text{dove } C = 1 - \frac{\sum_{i=1}^h t_i (t_i^2 - 1)}{n(n^2 - 1)} \quad \text{e} \quad \bar{R} = \frac{1}{n} \sum_{i=1}^k R_i = \frac{1}{n} \sum_{i=1}^k n_i \bar{R}_i = \frac{n+1}{2}$$
 parameter
$$df = k-1$$

$$p.\text{value}$$

$$P(\chi_{df}^2 \geq c)$$

```
> x <- c(2.1, 3, 2.1, 5.3, 5.3, 2.1, 5.6, 7.5, 2.1, 5.3, 2.1, 7.5)
> g <- factor(rep(letters[1:4], each = 3))
> g
```

```
[1] a a a b b b c c c d d d
Levels: a b c d
> n <- 12
> k < - 4
> R1 <- sum(rank(x)[g == "a"])
> R2 <- sum(rank(x)[g == "b"])
> R3 <- sum(rank(x)[g == "c"])
> R4 <- sum(rank(x)[g == "d"])
> R < - c(R1, R2, R3, R4)
[1] 12.0 19.0 24.5 22.5
> table(rank(x))
   3
       6
            8 10 11.5
            3 1 2
       1
> h <- 3
> t1 <- 5
> t2 <- 3
> t3 <- 2
> t <- c(t1, t2, t3)
> tapply(x, g, FUN = "length")
a b c d
3 3 3 3
> n1 <- 3
> n2 <- 3
> n3 <- 3
> n4 < -3
> enne <- c(n1, n2, n3, n4)
> C <-1 - sum(t * (t^2 - 1))/(n * (n^2 - 1))
> statistic <- (12/(n * (n + 1)) * sum(R^2/enne) - 3 * (n + 1))/C
> statistic
[1] 2.542784
> res <- kruskal.test(x, g)</pre>
> res$statistic
Kruskal-Wallis chi-squared
                  2.542784
> parameter <- k - 1
> parameter
[1] 3
> res$parameter
df
 3
> p.value <- 1 - pchisq(statistic, df = parameter)</pre>
> p.value
[1] 0.4676086
```

```
> res$p.value
 [1] 0.4676086
• Example 2:
 > x < -c(0.7, 1.6, 0.2, 1.2, 0.1, 3.4, 3.7, 0.8, 0, 2, 1.9, 0.8,
      1.1, 0.1, 0.1, 4.4, 5.5, 1.6, 4.6, 3.4)
 > g <- factor(rep(letters[1:2], each = 10))</pre>
 > g
  [1] a a a a a a a a a b b b b b b b b b
 Levels: a b
 > n < -20
 > k <- 2
 > R1 <- sum(rank(x)[g == "a"])
 > R2 <- sum(rank(x)[g == "b"])
 > R < - c(R1, R2)
 > R
 [1] 90.5 119.5
 > table(rank(x))
             5 6 7.5 9 10 11.5 13 14 15.5 17 18 19
1 1 2 1 1 2 1 1 2 1 1 1
                                                                           20
    1
 > h <- 4
 > t1 <- 3
 > t2 <- 2
 > t3 <- 2
 > t4 <- 2
 > t <- c(t1, t2, t3, t4)
 > tapply(x, g, FUN = "length")
  a b
 10 10
 > n1 <- 10
 > n2 < -10
 > enne <- c(n1, n2)
 > C <-1 - sum(t * (t^2 - 1))/(n * (n^2 - 1))
 > statistic <- (12/(n * (n + 1)) * sum(R^2/enne) - 3 * (n + 1))/C
 > statistic
 [1] 1.207785
 > res <- kruskal.test(x, g)</pre>
 > res$statistic
 Kruskal-Wallis chi-squared
                   1.207785
 > parameter <- k - 1
 > parameter
 [1] 1
 > res$parameter
```

```
df
  1
 > p.value <- 1 - pchisq(statistic, df = parameter)</pre>
 > p.value
 [1] 0.2717712
 > res$p.value
 [1] 0.2717712
• Example 3:
 > x <- c(4, 2.3, 8.8, 9.87, 12.34, 1.4, 6.4, 9.6, 8.86, 7.8, 8.6,
      8.8, 2, 0.3)
 > g \leftarrow factor(rep(c("Ctl", "Trt"), times = c(10, 4)))
 > g
  [1] Ctl Ctl Ctl Ctl Ctl Ctl Ctl Ctl Ctl Trt Trt Trt
 Levels: Ctl Trt
 > n <- 14
 > k < - 2
 > R1 <- sum(rank(x)[g == "Ctl"])
 > R2 <- sum(rank(x)[g == "Trt"])
 > R <- c(R1, R2)
 > R
 [1] 83.5 21.5
 > table(rank(x))
     2 3 4 5 6 7 8 9.5 11 12 13 14
   1
         1 1 1 1 1
                             1 2
                                     1 1
                                             1 1
 > h <- 1
 > t1 <- 2
 > t <- c(t1)
 > tapply(x, g, FUN = "length")
 Ctl Trt
 10 4
 > n1 <- 10
 > n2 < -4
 > enne <- c(n1, n2)
 > C <- 1 - sum(t * (t^2 - 1))/(n * (n^2 - 1))
 > statistic <- (12/(n * (n + 1)) * sum(R^2/enne) - 3 * (n + 1))/C
 > statistic
 [1] 1.448183
 > res <- kruskal.test(x, g)</pre>
 > res$statistic
 Kruskal-Wallis chi-squared
                   1.448183
```

```
> parameter <- k - 1
> parameter

[1] 1
> res$parameter

df
    1
> p.value <- 1 - pchisq(statistic, df = parameter)
> p.value

[1] 0.2288198
> res$p.value

[1] 0.2288198
```

10.4 Test di ipotesi sull'omogeneità delle varianze

Test di Levene

• Package: car

• Sintassi: levene.test()

• Input:

y vettore numerico di dimensione n group fattore f a k livelli di dimensione n

• Output:

Df gradi di libertà $\begin{tabular}{ll} F & value & valore empirico della statistica F \\ Pr(>F) & p-value \end{tabular}$

• Formula:

Df

$$F \text{ value} \\ Fvalue = \frac{\left[\sum_{j=1}^{k}\sum_{i=1}^{n_{j}}\left(xij-\bar{x}_{j}\right)^{2}\right]/(k-1)}{\left[\sum_{j=1}^{k}\left(n_{j}-1\right)s_{j}^{2}\right]/\left(n-k\right)} \\ \text{dove} \quad x_{ij} = \left|y_{ij}-Q_{0.5}\left(\left\{y_{1j},\, \ldots,\, y_{n_{j}j}\right\}\right)\right| \quad \forall j=1,\,2,\, \ldots,\, k \quad \forall i=1,\,2,\, \ldots,\, n_{j} \\ P(F_{k-1,\,n-k} \geq Fvalue) \\ \end{cases}$$

```
> y <- c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 16.5, 22, 1.2, 3.4)
 > f <- factor(rep(letters[1:4], each = 3))</pre>
 > n < -12
 > k <- 4
 > Df <- c(k - 1, n - k)
 > Df
 [1] 3 8
 > res <- levene.test(y, group = f)</pre>
 > res$Df
 [1] 3 8
 > x <- abs(y - ave(y, f, FUN = "median"))
 > Fvalue <- anova(lm(formula = x \sim f))$F
 > Fvalue
 [1] 0.608269
                     NA
 > res$"F value"
 [1] 0.608269
                     NA
 > p.value < 1 - pf(Fvalue, df1 = k - 1, df2 = n - k)
 > p.value
 [1] 0.6281414
                      NA
 > res$"Pr(>F)"
 [1] 0.6281414
                      NA
• Example 2:
 > y \leftarrow c(1.2, 3.4, 4.5, 6.4, 4, 3, 4, 3.4)
 > f <- factor(c("A", "B", "B", "B", "A", "A", "B", "A"))</pre>
 > n < - 8
 > k <- 2
 > Df <- c(k - 1, n - k)
 > Df
 [1] 1 6
 > res <- levene.test(y, group = f)</pre>
 > res$Df
 [1] 1 6
 > x <- abs(y - ave(y, f, FUN = "median"))
 > Fvalue <- anova(lm(formula = x \sim f))$F
 > Fvalue
 [1] 0.01477833
                          NA
 > res$"F value"
 [1] 0.01477833
                          NA
```

```
> p.value <- 1 - pf(Fvalue, df1 = k - 1, df2 = n - k)
 > p.value
 [1] 0.9072118
               NA
 > res$"Pr(>F)"
 [1] 0.9072118 NA
• Example 3:
 > y <- c(4, 2.3, 8.8, 9.87, 12.34, 1.4, 6.4, 9.6, 8.86, 7.8, 8.6,
      8.8, 2, 0.3)
 > f <- factor(rep(c("Ctl", "Trt"), times = c(10, 4)))
 > f
  [1] Ctl Ctl Ctl Ctl Ctl Ctl Ctl Ctl Ctl Trt Trt Trt
 Levels: Ctl Trt
 > n <- 14
 > k <- 2
 > Df <- c(k - 1, n - k)
 > Df
 [1] 1 12
 > res <- levene.test(y, group = f)</pre>
 > res$Df
 [1] 1 12
 > x <- abs(y - ave(y, f, FUN = "median"))
 > Fvalue <- anova(lm(formula = x \sim f))$F
 > Fvalue
 [1] 0.6701819
                    NA
 > res$"F value"
 [1] 0.6701819 NA
 > p.value < 1 - pf(Fvalue, df1 = k - 1, df2 = n - k)
 > p.value
 [1] 0.4289462
                    NA
 > res$"Pr(>F)"
 [1] 0.4289462
                    NA
```

10.5 Anova non parametrica a due fattori senza interazione

Test di Friedman

```
Package: stats
Sintassi: friedman.test()
Input:

x matrice di dimensione n × k
```

• Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p\text{-value}
```

• Formula:

statistic
$$c = \frac{12}{n\,k\,(k+1)} \sum_{j=1}^k \,R_j^2 - 3\,n\,(k+1)$$
 parameter
$$df = k-1$$
 p.value
$$P(\chi_{df}^2 \ge c)$$

```
> x <- matrix(c(6, 15, 8, 26, 29, 56, 60, 52, 20), nrow = 3, ncol = 3,
+ dimnames = list(NULL, c("X1", "X2", "X3")))
> x
    X1 X2 X3
[1,] 6 26 60
[2,] 15 29 52
[3,] 8 56 20
> n < -3
> k <- 3
> matrice <- t(apply(x, MARGIN = 1, FUN = "rank"))</pre>
> matrice
    X1 X2 X3
[1,] 1 2 3
[2,] 1 2 3
[3,] 1 3 2
> colSums(matrice)
X1 X2 X3
3 7 8
> R1 <- colSums(matrice)[1]</pre>
> R2 <- colSums(matrice)[2]</pre>
> R3 <- colSums(matrice)[3]</pre>
> R <- c(R1, R2, R3)
> R
X1 X2 X3
 3 7 8
```

```
> statistic <- 12/(n * k * (k + 1)) * sum(R^2) - 3 * n * (k + 1)
 > statistic
 [1] 4.666667
 > res <- friedman.test(x)</pre>
 > res$statistic
 Friedman chi-squared
             4.666667
 > parameter <- k - 1
 > parameter
 [1] 2
 > res$parameter
 df
 > p.value <- 1 - pchisq(statistic, df = parameter)</pre>
 > p.value
 [1] 0.09697197
 > res$p.value
 [1] 0.09697197
• Example 2:
 > x <- matrix(c(1, 3, 1, 3, 2, 2, 2, 3, 2, 3, 3, 1, 2, 1, 1), nrow = 5,
 + ncol = 3, dimnames = list(NULL, c("X1", "X2", "X3")))
 > x
      X1 X2 X3
 [1,] 1 2 3
 [2,] 3 2 1
 [3,] 1 3 2
 [4,] 3 2 1
 [5,] 2 3 1
 > n <- 5
 > k <- 3
 > matrice <- t(apply(x, MARGIN = 1, FUN = "rank"))</pre>
 > matrice
     X1 X2 X3
 [1,] 1 2 3
 [2,] 3 2 1
 [3,] 1 3 2
 [4,] 3 2 1
 [5,] 2 3 1
 > colSums(matrice)
 X1 X2 X3
 10 12 8
```

```
> R1 <- colSums(matrice)[1]</pre>
 > R2 <- colSums(matrice)[2]</pre>
 > R3 <- colSums(matrice)[3]</pre>
 > R < - c(R1, R2, R3)
 > R
 X1 X2 X3
 10 12 8
 > statistic <- 12/(n * k * (k + 1)) * sum(R^2) - 3 * n * (k + 1)
 > statistic
 [1] 1.6
 > res <- friedman.test(x)</pre>
 > res$statistic
 Friedman chi-squared
                  1.6
 > parameter <- k - 1
 > parameter
 [1] 2
 > res$parameter
 df
  2
 > p.value <- 1 - pchisq(statistic, df = parameter)</pre>
 > p.value
 [1] 0.449329
 > res$p.value
 [1] 0.449329
• Example 3:
 > x <- matrix(0, nrow = 10, ncol = 6, byrow = TRUE, dimnames = list(NULL,
 + c("X1", "X2", "X3", "X4", "X5", "X6")))
 > for (i in 1:10) x[i, ] \leftarrow sample(1:6)
 > x
       X1 X2 X3 X4 X5 X6
  [1,] 5 3 4 2 6 1
  [2,]
       3
           1
             4 2
                    6 5
  [3,]
       1
           4 5
                3 2 6
  [4,]
        3
           1 6 2 5 4
           2 5
                    3 1
  [5,]
       6
                 4
           4 5
  [6,]
        6
                 2
  [7,]
        1
           4
              2
                 3
                    5
                       6
              3
                 2
  [8,] 1
           6
                    5
                       4
           2 1
                 5
                       3
  [9,] 6
                    4
                 5
 [10,] 2
           3 1
```

```
> n <- 10
> k <- 6
> matrice <- t(apply(x, MARGIN = 1, FUN = "rank"))</pre>
> matrice
      X1 X2 X3 X4 X5 X6
 [1,] 5 3 4 2 6 1
 [2,] 3 1 4 2 6 5
         4 5 3 2 6
 [3,] 1
 [4,] 3
         1 6 2 5
 [5,] 6
               4 3 1
         2
            5
 [6,] 6
         4
            5
               2
                   3
                      1
 [7,] 1
         4 2
               3 5 6
         6 3 2 5 4
 [8,] 1
               5 4 3
 [9,] 6 2 1
         3 1
                5
[10,] 2
> colSums(matrice)
X1 X2 X3 X4 X5 X6
34 30 36 30 45 35
> R1 <- colSums(matrice)[1]</pre>
> R2 <- colSums(matrice)[2]</pre>
> R3 <- colSums(matrice)[3]</pre>
> R4 <- colSums(matrice)[4]</pre>
> R5 <- colSums(matrice)[5]</pre>
> R6 <- colSums(matrice)[6]
> R <- c(R1, R2, R3, R4, R5, R6)
X1 X2 X3 X4 X5 X6
34 30 36 30 45 35
> statistic <- 12/(n * k * (k + 1)) * sum(R^2) - 3 * n * (k + 1)
> statistic
[1] 4.342857
> res <- friedman.test(x)</pre>
> res$statistic
Friedman chi-squared
            4.342857
> parameter <- k - 1</pre>
> parameter
[1] 5
> res$parameter
df
 5
> p.value <- 1 - pchisq(statistic, df = parameter)</pre>
> p.value
[1] 0.5011797
> res$p.value
[1] 0.5011797
```

10.6 Test di ipotesi su una proporzione

Test di Bernoulli

• Package: stats

• Sintassi: binom.test()

• Input:

x numero di successi

n dimensione campionaria

p valore di p_0

alternative = "less" / "greater" / "two.sided" ipotesi alternativa

conf.level livello di confidenza $1-\alpha$

• Output:

statistic numero di successi parameter dimensione campionaria p.value p-value conf.int intervallo di confidenza per la proporzione incognita a livello $1-\alpha$ estimate proporzione campionaria null.value valore di p_0 alternative ipotesi alternativa

• Formula:

statistic
$$x$$
 parameter
$$n$$
 p.value
$$\frac{\text{alternative = "less"}}{\text{p.value}}$$

$$\text{p.value} = \sum_{i=0}^{x} \binom{n}{i} p_0^i \, (1-p_0)^{n-i}$$

$$\frac{\text{alternative = "greater"}}{\text{p.value}}$$

$$\text{p.value} = 1 - \sum_{i=0}^{x-1} \binom{n}{i} p_0^i \, (1-p_0)^{n-i}$$

$$\begin{array}{|c|c|c|c|c|} \textbf{Caso} & \textbf{p.value} \\ \hline x = n \, p_0 & 1 \\ \hline x < n \, p_0 & F_X(x) - F_X(n-y) + 1 & y = \# \left(p_X(k) \le p_X(x) \quad \forall \, k = \lceil n \, p_0 \rceil, \, \dots, \, n \right) \\ \hline x > n \, p_0 & F_X(y-1) - F_X(x-1) + 1 & y = \# \left(p_X(k) \le p_X(x) \quad \forall \, k = 0, \, \dots, \, \lfloor n \, p_0 \rfloor \right) \\ \hline \end{array}$$

$$X \sim Binomiale(n, p_0)$$

$$p_X(x) = \binom{n}{x} p_0^x (1 - p_0)^{n-x} \quad \forall x = 0, 1, ..., n$$

$$F_X(x) = \sum_{i=0}^x \binom{n}{i} p_0^i (1 - p_0)^{n-i} \quad \forall x = 0, 1, ..., n$$

alternative = "two.sided"

```
conf.int
                                   F_U^{-1}(\alpha/2) F_H^{-1}(1-\alpha/2)
                       dove U \sim Beta(x, n-x+1) e H \sim Beta(x+1, n-x)
     estimate
                                              n
     null.value
                                             p_0
• Example 1:
 > x < -682
 > n < -925
 > p0 < -0.75
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
       conf.level = 0.95)$statistic
 number of successes
                  682
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
 + conf.level = 0.95) $parameter
 number of trials
              925
 > n * p0
 [1] 693.75
 > y <- sum(dbinom(ceiling(n * p0):n, n, p0) <= dbinom(x, n, p0))
 > y
 [1] 220
 > p.value <- pbinom(x, n, p0) - pbinom(n - y, n, p0) + 1
 > p.value
 [1] 0.3824916
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
 + conf.level = 0.95)$p.value
 [1] 0.3824916
 > lower <- qbeta(0.025, x, n - x + 1)
 > upper <- qbeta(0.975, x + 1, n - x)
 > c(lower, upper)
 [1] 0.7076683 0.7654066
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
 + conf.level = 0.95)$conf.int
 [1] 0.7076683 0.7654066
 attr(, "conf.level")
```

[1] 0.95

```
> x/n
 [1] 0.7372973
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
       conf.level = 0.95)$estimate
 probability of success
              0.7372973
 > p0
 [1] 0.75
 > binom.test(x = 682, n = 925, p = 0.75, alternative = "two.sided",
     conf.level = 0.95)$null.value
 probability of success
                   0.75
• Example 2:
 > x < -682
 > n <- 925
 > p0 < -0.63
 > binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
       conf.level = 0.95)$statistic
 number of successes
 > binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
 + conf.level = 0.95) $parameter
 number of trials
              925
 > n * p0
 [1] 582.75
 > y <- sum(dbinom(0:floor(n * p0), n, p0) <= dbinom(x, n, p0))
 [1] 480
 > p.value <- pbinom(y - 1, n, p0) - pbinom(x - 1, n, p0) + 1
 > p.value
 [1] 4.925171e-12
 > binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
     conf.level = 0.95)$p.value
 [1] 4.925209e-12
 > ower <- qbeta(0.025, x, n - x + 1)
 > upper <- qbeta(0.975, x + 1, n - x)
 > c(lower, upper)
```

```
[1] 0.7076683 0.7654066
> binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
+ conf.level = 0.95)$conf.int
[1] 0.7076683 0.7654066
attr(,"conf.level")
[1] 0.95
> x/n
[1] 0.7372973
> binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
     conf.level = 0.95)$estimate
probability of success
             0.7372973
> p0
[1] 0.63
> binom.test(x = 682, n = 925, p = 0.63, alternative = "two.sided",
     conf.level = 0.95)$null.value
probability of success
                  0.63
```

10.7 Test di ipotesi sul ciclo di casualità

Test dei Runs

```
    Package: tseries
    Sintassi: runs.test()
    Input:
        x fattore a 2 livelli di dimensione n
        alternative = "less" / "greater" / "two.sided" ipotesi alternativa
    Output:
```

statistic valore empirico della statistica Z p.value $p ext{-} ext{value}$ alternative ipotesi alternativa

• Formula:

statistic
$$z=\frac{V-\frac{n_1+2\,n_1\,n_2+n_2}{n_1+n_2}}{\sqrt{\frac{2\,n_1\,n_2\,(2\,n_1\,n_2-n_1-n_2)}{(n_1+n_2)^2\,(n_1+n_2-1)}}}$$
 p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z)$

```
> x <- factor(c("HIGH", "LOW", "LOW", "HIGH", "LOW", "HIGH", "HIGH", "HIGH", "LOW", "HIGH", "LOW", "LOW", "HIGH", "LOW", "LOW", "HIGH", "LOW", "LOW", "HIGH", "LOW", "LOW
                   "HIGH", "LOW", "HIGH", "HIGH", "LOW", "HIGH", "LOW",
                    "HIGH", "LOW", "HIGH", "HIGH", "LOW", "HIGH", "LOW"))
> X
  [1] HIGH LOW LOW HIGH LOW HIGH HIGH LOW HIGH HIGH LOW LOW HIGH LOW
[16] HIGH LOW HIGH HIGH LOW HIGH LOW HIGH LOW HIGH HIGH LOW HIGH LOW
Levels: HIGH LOW
> n <- 30
> V <-1 + sum(as.numeric(x[-1] != x[-n]))
[1] 22
> n1 <- length(x[x == "HIGH"])
[1] 16
> n2 <- length(x[x == "LOW"])
[1] 14
> media <- (n1 + 2 * n1 * n2 + n2)/(n1 + n2)
> media
[1] 15.93333
> varianza <- (2 * n1 * n2 * (2 * n1 * n2 - n1 - n2))/((n1 + n2)^2 *
+ (n1 + n2 - 1))
> varianza
[1] 7.174866
> z <- (V - media)/sqrt(varianza)</pre>
> z
[1] 2.26487
> runs.test(x, alternative = "less")$statistic
Standard Normal
                         2.26487
> p.value <- pnorm(z)</pre>
> p.value
[1] 0.9882397
> runs.test(x, alternative = "less")$p.value
```

[1] 0.9882397

```
• Example 2:
```

```
> x
[1] abbbabbabbaabbaabbab
Levels: a b
> n < -22
> V <-1 + sum(as.numeric(x[-1] != x[-n]))
[1] 12
> n1 <- length(x[x == "a"])
> n1
[1] 8
> n2 <- length(x[x == "b"])
> n2
[1] 14
> media <- (n1 + 2 * n1 * n2 + n2)/(n1 + n2)
> media
[1] 11.18182
> varianza <- (2 * n1 * n2 * (2 * n1 * n2 - n1 - n2))/((n1 + n2)^2 *
+ (n1 + n2 - 1))
> varianza
[1] 4.451791
> z <- (V - media)/sqrt(varianza)</pre>
[1] 0.3877774
> runs.test(x, alternative = "two.sided")$statistic
Standard Normal
    0.3877774
> p.value <- 2 * pnorm(-abs(z))</pre>
> p.value
[1] 0.6981808
> runs.test(x, alternative = "two.sided")$p.value
[1] 0.6981808
```

```
> x \leftarrow factor(rep(1:2, each = 10))
Levels: 1 2
> n <- 20
> V <- 1 + sum(as.numeric(x[-1] != x[-n]))
> V
[1] 2
> n1 <- length(x[x == "1"])
> n1
[1] 10
> n2 <- length(x[x == "2"])
> n2
[1] 10
> media <- (n1 + 2 * n1 * n2 + n2)/(n1 + n2)
> media
[1] 11
> varianza <- (2 * n1 * n2 * (2 * n1 * n2 - n1 - n2))/((n1 + n2)^2 *
+ (n1 + n2 - 1))
> varianza
[1] 4.736842
> z <- (V - media)/sqrt(varianza)</pre>
> z
[1] -4.135215
> runs.test(x, alternative = "two.sided")$statistic
Standard Normal
     -4.135215
> p.value <- 2 * pnorm(-abs(z))
> p.value
[1] 3.546230e-05
> runs.test(x, alternative = "two.sided")$p.value
[1] 3.546230e-05
```

10.8 Test di ipotesi sulla differenza tra parametri di scala

Test di Mood

```
• Package: stats
• Sintassi: mood.test()
• Input:

    x vettore numerico di dimensione n_x
    y vettore numerico di dimensione n_y
    alternative = "less" / "greater" / "two.sided" ipotesi alternativa
```

• Output:

```
statistic valore empirico della statistica Z p.value p	ext{-}\mathrm{value} alternative ipotesi alternativa
```

• Formula:

statistic
$$z = \frac{V - \frac{n_x \left(n_x + n_y + 1\right) \left(n_x + n_y - 1\right)}{12}}{\sqrt{\frac{n_x \, n_y \left(n_x + n_y + 1\right) \left(n_x + n_y + 2\right) \left(n_x + n_y - 2\right)}{180}}}$$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(-\mid z\mid)$

• Example 1:

```
> x < -c(-1, 1, -2, -1, 1, 1, 1, 1, -1, -2, 1, 1)
> y <- c(1, 2, 3, 4, 5, 6, 7, 8, 9)
> nx <- 12
> ny <- 9
> Rx <- rank(c(x, y))[1:nx]
> V <- sum((Rx - (nx + ny + 1)/2)^2)
> media <- nx * (nx + ny + 1) * (nx + ny - 1)/12
> varianza <- nx * ny * (nx + ny + 1) * (nx + ny + 2) * (nx + ny -
+ 2)/180
> z <- (V - media)/sqrt(varianza)</pre>
[1] -1.273865
> mood.test(x, y, alternative = "less")$statistic
-1.273865
> p.value <- pnorm(z)</pre>
> p.value
[1] 0.1013557
> mood.test(x, y, alternative = "less")$p.value
[1] 0.1013557
```

```
> x < -c(1, 4.5, 6.78, 9.8, 7.7)
 > y < -c(1, 4, 10, 2.1, 3.5, 5.6, 8.4, 12, 16.5, 22, 1.2, 3.4)
 > nx < -5
 > ny <- 12
 > Rx <- rank(c(x, y))[1:nx]
 > V <- sum((Rx - (nx + ny + 1)/2)^2)
 > media <- nx * (nx + ny + 1) * (nx + ny - 1)/12
 > media
 [1] 120
 > varianza <- nx * ny * (nx + ny + 1) * (nx + ny + 2) * (nx + ny -
       2)/180
 > varianza
 [1] 1710
 > z <- (V - media)/sqrt(varianza)</pre>
 [1] -1.009621
 > mood.test(x, y, alternative = "two.sided")$statistic
 -1.009621
 > p.value <-2 * pnorm(-abs(z))
 > p.value
 [1] 0.3126768
 > mood.test(x, y, alternative = "two.sided")$p.value
 [1] 0.3126768
• Example 3:
 > x \leftarrow c(1, 1.2, 3.4, 0.8, 10.2, 9.3, 7.34)
 > y <- c(-3.4, 0.2, 1.2, 2.1, 2.2, 2.2, 2.3, 3.1, 3.2, 4.2, 4.3,
      5.43)
 > nx < -7
 > ny <- 12
 > Rx <- rank(c(x, y))[1:nx]
 > V <- sum((Rx - (nx + ny + 1)/2)^2)
 > media <- nx * (nx + ny + 1) * (nx + ny - 1)/12
 > media
 [1] 210
 > varianza < - nx * ny * (nx + ny + 1) * (nx + ny + 2) * (nx + ny -
 + 2)/180
 > varianza
 [1] 3332
 > z <- (V - media)/sqrt(varianza)</pre>
 [1] 1.702080
```

```
> mood.test(x, y, alternative = "two.sided")$statistic

Z
1.702080
> p.value <- 2 * pnorm(-abs(z))
> p.value

[1] 0.0887403
> mood.test(x, y, alternative = "two.sided")$p.value

[1] 0.0887403
```

Capitolo 11

Tabelle di contingenza

11.1 Simbologia

- frequenze osservate: $n_{ij} \quad \forall i = 1, 2, ..., h \quad \forall j = 1, 2, ..., k$
- frequenze osservate nella m-esima tabella di contingenza 2×2 : $n_{ijm} \quad \forall i,j=1,2 \quad \forall m=1,2,\ldots,l$
- frequenze marginali di riga: $n_{i.} = \sum_{j=1}^{k} n_{ij} \quad \forall i = 1, 2, ..., h$
- frequenze marginali di riga nella m-esima tabella di contingenza 2×2 : $n_{i \cdot m} = \sum_{i=1}^{2} n_{ijm} \quad \forall i = 1, 2 \quad \forall m = 1, 2, \dots, l$
- frequenze marginali di colonna: $n_{i,j} = \sum_{i=1}^{h} n_{i,j} \quad \forall j = 1, 2, \dots, k$
- $\bullet\,$ frequenze marginali di colonna nella m-esimatabella di contingenza 2×2 :

$$n_{\cdot jm} = \sum_{i=1}^{2} n_{ijm} \quad \forall j = 1, 2 \quad \forall m = 1, 2, \dots, l$$

- frequenze attese: $\hat{n}_{ij} = n_i \cdot n_{ij} / n_{ij}$ $\forall i = 1, 2, ..., h \quad \forall j = 1, 2, ..., k$
- frequenze attese nella m-esima tabella di contingenza 2×2 :

$$\hat{n}_{ijm} = n_{i \cdot m} \, n_{\cdot jm} \, / \, n_{\cdot \cdot m} \quad \forall i, j = 1, 2 \quad \forall m = 1, 2, \dots, l$$

- totale frequenze assolute: $n_{\cdot \cdot \cdot} = \sum_{i=1}^h \sum_{j=1}^k n_{ij} = \sum_{i=1}^h \sum_{j=1}^k \hat{n}_{ij}$
- totale frequenze assolute nella m-esima tabella di contingenza 2×2 : $n_{\cdots m} = \sum_{i=1}^2 \sum_{j=1}^2 n_{ijm} = \sum_{i=1}^2 \sum_{j=1}^2 \hat{n}_{ijm} \quad \forall \, m=1,\,2,\,\ldots,\,l$

11.2 Test di ipotesi per tabelle di contingenza 2 righe per 2 colonne

Test Chi - Quadrato di indipendenza

- Package: stats
- Sintassi: chisq.test()
- Input:
 - x matrice di dimensione 2×2 contenente frequenze assolute correct = TRUE / FALSE correzione di *Yates*
- Output:
 - statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value observed frequenze osservate expected frequenze attese residuals residui di Pearson
- Formula:

statistic

parameter

p.value

observed

expected

residuals

$$c = \sum_{i=1}^{2} \sum_{j=1}^{2} \frac{(|n_{ij} - \hat{n}_{ij}| - 1/2)^{2}}{\hat{n}_{ij}} = \frac{n ... (|n_{11} n_{22} - n_{12} n_{21}| - n .../2)^{2}}{n_{1}..n_{2}..n_{1} n_{\cdot 2}}$$

$$c = \sum_{i=1}^{2} \sum_{j=1}^{2} \frac{(n_{ij} - \hat{n}_{ij})^{2}}{\hat{n}_{ij}} = \frac{n ... (n_{11} n_{22} - n_{12} n_{21})^{2}}{n_{1}..n_{2}..n_{\cdot 1} n_{\cdot 2}}$$

$$df = 1$$

$$P(\chi_{df}^{2} \ge c)$$

$$n_{ij} \quad \forall i, j = 1, 2$$

$$\frac{n_{ij} - \hat{n}_{ij}}{\sqrt{\hat{n}_{ij}}} \quad \forall i, j = 1, 2$$

```
> x < -matrix(data = c(2, 10, 23, 21), nrow = 2, ncol = 2, byrow = FALSE)
> riga <- c("A", "B")</pre>
> colonna <- c("A", "B")</pre>
> dimnames(x) <- list(riga, colonna)</pre>
> x
 A B
A 2 23
B 10 21
> chisq.test(x, correct = FALSE)
        Pearson's Chi-squared test
X-squared = 4.8369, df = 1, p-value = 0.02786
> res <- chisq.test(x, correct = FALSE)</pre>
> res$statistic
X-squared
4.836911
> res$parameter
df
1
> res$p.value
[1] 0.02785675
> res$observed
```

```
А В
 A 2 23
 B 10 21
 > res$expected
          Α
 A 5.357143 19.64286
 B 6.642857 24.35714
 > res$residuals
 A -1.450451 0.7574736
 B 1.302544 -0.6802314
• Example 2:
 > x <- matrix(data = c(2, 10, 23, 21), nrow = 2, ncol = 2, byrow = FALSE)
 > riga <- c("A", "B")</pre>
 > colonna <- c("A", "B")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
    A B
 A 2 23
 B 10 21
 > chisq.test(x, correct = TRUE)
         Pearson's Chi-squared test with Yates' continuity correction
 X-squared = 3.5034, df = 1, p-value = 0.06124
 > res <- chisq.test(x, correct = TRUE)</pre>
 > res$statistic
 X-squared
  3.503421
 > res$parameter
 df
  1
 > res$p.value
 [1] 0.06124219
 > res$observed
    A B
 A 2 23
 B 10 21
 > res$expected
```

```
A 5.357143 19.64286
 B 6.642857 24.35714
 > res$residuals
           Α
 A -1.450451 0.7574736
 B 1.302544 -0.6802314
• Example 3:
 > x <- matrix(data = c(12, 5, 7, 7), nrow = 2, ncol = 2, byrow = FALSE)
 > riga <- c("A", "B")</pre>
 > colonna <- c("A", "B")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
   АВ
 A 12 7
 B 5 7
 > chisq.test(x, correct = TRUE)
         Pearson's Chi-squared test with Yates' continuity correction
 data: x
 X-squared = 0.6411, df = 1, p-value = 0.4233
 > res <- chisq.test(x, correct = TRUE)</pre>
 > res$statistic
 X-squared
 0.6411203
 > res$parameter
 df
  1
 > res$p.value
 [1] 0.4233054
 > res$observed
   АВ
 A 12 7
 B 5 7
 > res$expected
           Α
 A 10.419355 8.580645
 B 6.580645 5.419355
 > res$residuals
            Α
 A 0.4896818 -0.5396031
 B -0.6161694 0.6789856
```

Test di McNemar

```
    Package: stats
    Sintassi: mcnemar.test()
    Input:
        x matrice di dimensione 2 × 2 contenente frequenze assolute correct = TRUE / FALSE correzione di Yates
```

• Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p\text{-value}
```

• Formula:

statistic
$$c = \frac{(|n_{12} - n_{21}| - 1)^2}{n_{12} + n_{21}}$$

$$c = \frac{(|n_{12} - n_{21}| - 1)^2}{n_{12} + n_{21}}$$

$$c = \frac{(n_{12} - n_{21})^2}{n_{12} + n_{21}}$$
 parameter
$$df = 1$$
 p.value
$$P(\chi_{df}^2 \ge c)$$

```
> x < -matrix(data = c(2, 10, 23, 21), nrow = 2, ncol = 2, byrow = FALSE)
> riga <- c("A", "B")</pre>
> colonna <- c("A", "B")</pre>
> dimnames(x) <- list(riga, colonna)</pre>
> x
  A B
A 2 23
B 10 21
> mcnemar.test(x, correct = FALSE)
        McNemar's Chi-squared test
data:
McNemar's chi-squared = 5.1212, df = 1, p-value = 0.02364
> res <- mcnemar.test(x, correct = FALSE)</pre>
> res$statistic
McNemar's chi-squared
              5.121212
> res$parameter
df
 1
```

```
> res$p.value
 [1] 0.0236351
• Example 2:
 > x <- matrix(data = c(2, 10, 23, 21), nrow = 2, ncol = 2, byrow = FALSE)
 > riga <- c("A", "B")</pre>
 > colonna <- c("A", "B")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
    A B
 A 2 23
 B 10 21
 > mcnemar.test(x, correct = TRUE)
         McNemar's Chi-squared test with continuity correction
 data: x
 McNemar's chi-squared = 4.3636, df = 1, p-value = 0.03671
 > res <- mcnemar.test(x, correct = TRUE)</pre>
 > res$statistic
 McNemar's chi-squared
               4.363636
 > res$parameter
 df
  1
 > res$p.value
 [1] 0.03671386
• Example 3:
 > x <- matrix(data = c(12, 5, 7, 7), nrow = 2, ncol = 2, byrow = FALSE)
 > riga <- c("A", "B")</pre>
 > colonna <- c("A", "B")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
    АВ
 A 12 7
 B 5 7
 > mcnemar.test(x, correct = TRUE)
         McNemar's Chi-squared test with continuity correction
 data: x
 McNemar's chi-squared = 0.0833, df = 1, p-value = 0.7728
 > res <- mcnemar.test(x, correct = TRUE)</pre>
 > res$statistic
```

```
McNemar's chi-squared
0.08333333

> res$parameter

df
1

> res$p.value

[1] 0.77283
```

Test esatto di Fisher

```
• Package: stats
```

• Sintassi: fisher.test()

• Input:

• Output:

```
p.value p-value
alternative ipotesi alternativa
```

• Formula:

p.value

alternative	p.value		
less	$\sum_{i=0}^{n_{11}} p(i)$		
greater	$1 - \sum_{i=0}^{n_{11}-1} p(i)$		
two.sided	$\sum_{i=0}^{n_{11}} p(i) + \sum_{p(i) \le p(n_{11})} p(i) \forall i = n_{11} + 1, \dots, \min(n_1, n_{11})$		

$$p(i) = \frac{\max(n_{1.}, n_{.1})C_{i}}{nC_{\min(n_{1.}, n_{.1})}} \frac{C_{\min(n_{1.}, n_{.1})-i}}{C_{\min(n_{1.}, n_{.1})}} \quad \forall i = 0, 1, ..., \min(n_{1.}, n_{.1})$$

```
> x <- matrix(data = c(2, 9, 5, 4), nrow = 2, ncol = 2, byrow = FALSE)
> riga <- c("A", "B")
> colonna <- c("A", "B")
> dimnames(x) <- list(riga, colonna)
> x

A B
A 2 5
B 9 4

> n11 <- 2
> n1. <- 2 + 5
> n.1 <- 2 + 9
> n.. <- 2 + 5 + 9 + 4
> n..
[1] 20
```

```
> minimo <- min(n1., n.1)</pre>
> minimo
[1] 7
> massimo <- max(n1., n.1)</pre>
> massimo
[1] 11
> p <- function(i) dhyper(i, massimo, n.. - massimo, minimo)</pre>
> p.value.less <- 0</pre>
> for (i in 0:n11) p.value.less <- p.value.less + p(i)</pre>
> p.value.less
[1] 0.1017802
> fisher.test(x, alternative = "less")$p.value
[1] 0.1017802
> p.value.greater <- 0</pre>
> for (i in 0:(n11 - 1)) p.value.greater <- p.value.greater + p(i)</pre>
> p.value.greater <- 1 - p.value.greater</pre>
> p.value.greater
[1] 0.9876161
> fisher.test(x, alternative = "greater")$p.value
[1] 0.9876161
> p.value1 <- 0
> for (i in 0:n11) p.value1 <- p.value1 + p(i)</pre>
> p.value1
[1] 0.1017802
> p.value2 <- 0
> for (i in (n11 + 1):minimo) {
     if (p(i) \le p(n11))
          p.value2 <- p.value2 + p(i)</pre>
+ }
> p.value2
[1] 0.05789474
> p.value.two.sided <- p.value1 + p.value2</pre>
> p.value.two.sided
[1] 0.1596749
> fisher.test(x, alternative = "two.sided")$p.value
[1] 0.1596749
```

```
> x <- matrix(data = c(3, 7, 6, 5), nrow = 2, ncol = 2, byrow = FALSE)
> riga <- c("A", "B")</pre>
> colonna <- c("A", "B")</pre>
> dimnames(x) <- list(riga, colonna)</pre>
> X
 АВ
A 3 6
B 7 5
> n11 <- 3
> n1. < -3 + 6
> n.1 <- 3 + 7
> n.. <- 3 + 6 + 7 + 5
> n..
[1] 21
> minimo <- min(n1., n.1)</pre>
> minimo
[1] 9
> massimo <- max(n1., n.1)</pre>
> massimo
[1] 10
> p <- function(i) dhyper(i, massimo, n.. - massimo, minimo)</pre>
> p.value.less <- 0</pre>
> for (i in 0:n11) p.value.less <- p.value.less + p(i)</pre>
> p.value.less
[1] 0.2449393
> fisher.test(x, alternative = "less")$p.value
[1] 0.2449393
> p.value.greater <- 0</pre>
> for (i in 0:(n11 - 1)) p.value.greater <- p.value.greater + p(i)</pre>
> p.value.greater <- 1 - p.value.greater</pre>
> p.value.greater
[1] 0.943677
> fisher.test(x, alternative = "greater")$p.value
[1] 0.943677
> p.value1 <- 0
> for (i in 0:n11) p.value1 <- p.value1 + p(i)</pre>
> p.value1
[1] 0.2449393
```

```
> p.value2 <- 0
 > for (i in (n11 + 1):minimo) {
       if (p(i) \le p(n11))
           p.value2 <- p.value2 + p(i)</pre>
 + }
 > p.value2
 [1] 0.1420576
 > p.value.two.sided <- p.value1 + p.value2</pre>
 > p.value.two.sided
 [1] 0.3869969
 > fisher.test(x, alternative = "two.sided")$p.value
 [1] 0.3869969
• Example 3:
 > x <- matrix(c(2, 9, 3, 4), nrow = 2, ncol = 2, byrow = FALSE)
 > riga <- c("A", "B")
 > colonna <- c("A", "B")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
  AΒ
 A 2 3
 B 9 4
 > n11 <- 2
 > n1. < -2 + 3
 > n.1 < -2 + 9
 > n.. < -2 + 3 + 9 + 4
 > n..
 [1] 18
 > minimo <- min(n1., n.1)</pre>
 > minimo
 [1] 5
 > massimo <- max(n1., n.1)</pre>
 > massimo
 [1] 11
 > p <- function(i) dhyper(i, massimo, n.. - massimo, minimo)
 > p.value.less <- 0</pre>
 > for (i in 0:n11) p.value.less <- p.value.less + p(i)
 > p.value.less
 [1] 0.2720588
 > fisher.test(x, alternative = "less")$p.value
 [1] 0.2720588
```

```
> p.value.greater <- 0</pre>
> for (i in 0:(n11 - 1)) p.value.greater <- p.value.greater + p(i)</pre>
> p.value.greater <- 1 - p.value.greater</pre>
> p.value.greater
[1] 0.9526144
> fisher.test(x, alternative = "greater")$p.value
[1] 0.9526144
> p.value1 <- 0
> for (i in 0:n11) p.value1 <- p.value1 + p(i)</pre>
> p.value1
[1] 0.2720588
> p.value2 <- 0
> for (i in (n11 + 1):minimo) {
      if (p(i) \le p(n11))
         p.value2 <- p.value2 + p(i)</pre>
+ }
> p.value2
[1] 0.05392157
> p.value.two.sided <- p.value1 + p.value2</pre>
> p.value.two.sided
[1] 0.3259804
> fisher.test(x, alternative = "two.sided")$p.value
[1] 0.3259804
```

Test di Mantel - Haenszel

```
Package: stats
Sintassi: mantelhaen.test()
Input:

x array di dimensione 2 × 2 × l contenente l tabelle di contingenza 2 × 2 conf.level livello di confidenza 1 - α correct = FALSE
Output:

statistic valore empirico della statistica χ² parameter gradi di libertà p.value p-value estimate stima campionaria del comune OR conf.int intervallo di confidenza a livello 1 - α
```

statistic

$$c = \frac{\left[\sum_{m=1}^{l} (n_{11m} - \hat{n}_{11m})\right]^2}{\sum_{m=1}^{l} \hat{\sigma}_{n_{11m}}^2}$$

dove
$$\hat{\sigma}_{n_{11m}}^2 = \frac{n_{1 \cdot m} \, n_{2 \cdot m} \, n_{\cdot 1m} \, n_{\cdot 2m}}{n_{\cdot m}^2 \, (n_{\cdot m} - 1)} \quad \forall \, m = 1, \, 2, \, \dots, \, l$$

parameter

$$df = 1$$

p.value

$$P(\chi_{df}^2 \ge c)$$

estimate

$$\hat{\theta}_{MH} = \frac{\sum_{m=1}^{l} n_{11m} n_{22m} / n_{\cdots m}}{\sum_{m=1}^{l} n_{12m} n_{21m} / n_{\cdots m}} = \frac{\sum_{m=1}^{l} R_m}{\sum_{m=1}^{l} S_m} = \frac{R}{S}$$

conf.int

$$\hat{\theta}_{MH} e^{-z_{1-lpha/2} \hat{\sigma}_{\log(\hat{ heta}_{MH})}} \quad \hat{ heta}_{MH} e^{z_{1-lpha/2} \hat{\sigma}_{\log(\hat{ heta}_{MH})}}$$

dove

$$\hat{\sigma}_{\log(\hat{\theta}_{MH})}^{2} = \frac{1}{R^{2}} \sum_{m=1}^{l} \frac{(n_{11m} + n_{22m}) R_{m}}{n_{\cdots m}} + \frac{1}{S^{2}} \sum_{m=1}^{l} \frac{(n_{12m} + n_{21m}) S_{m}}{n_{\cdots m}} + \frac{1}{2RS} \sum_{m=1}^{l} \frac{(n_{11m} + n_{22m}) S_{m} + (n_{12m} + n_{21m}) R_{m}}{n_{\cdots m}}$$

• Examples:

Response

Treatment Success Failure
Drug 11 25
Control 10 27

, , Center = 2

Response

Treatment Success Failure
Drug 16 4
Control 22 10

, , Center = 3

Response

Treatment Success Failure
Drug 14 5
Control 7 12

, , Center = 4

Response

Treatment Success Failure
Drug 2 14
Control 1 16

```
, , Center = 5
        Response
Treatment Success Failure
                  11
 Drug
              6
 Control
              0
                     12
, , Center = 6
        Response
Treatment Success Failure
         1 10
 Drug
             0
                     10
 Control
, , Center = 7
        Response
Treatment Success Failure
           1
 Drug
              1
 Control
, , Center = 8
        Response
Treatment Success Failure
 Drug
               4
 Control
               6
                       1
> mantelhaen.test(x, conf.level = 0.95, correct = FALSE)
       Mantel-Haenszel chi-squared test without continuity correction
data: x
Mantel-Haenszel X-squared = 6.3841, df = 1, p-value = 0.01151
alternative hypothesis: true common odds ratio is not equal to 1
95 percent confidence interval:
1.177590 3.869174
sample estimates:
common odds ratio
        2.134549
> res <- mantelhaen.test(x, conf.level = 0.95, correct = FALSE)
> res$statistic
Mantel-Haenszel X-squared
                6.384113
> res$parameter
df
1
> res$p.value
[1] 0.01151463
> res$estimate
common odds ratio
        2.134549
```

```
> res$conf.int
[1] 1.177590 3.869174
attr(,"conf.level")
[1] 0.95
```

11.3 Test di ipotesi per tabelle di contingenza n righe per k colonne

Test Chi - Quadrato di indipendenza

Package: statsSintassi: chisq.test()

• Input:

x matrice di dimensione $h \times k$ contenente frequenze assolute

• Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p-value observed frequenze osservate expected frequenze attese residuals residui di Pearson
```

• Formula:

statistic

$$c = \sum_{i=1}^{h} \sum_{j=1}^{k} \frac{(n_{ij} - \hat{n}_{ij})^2}{\hat{n}_{ij}} = \sum_{i=1}^{h} \sum_{j=1}^{k} \frac{n_{ij}^2}{\hat{n}_{ij}} - n... = n... \left(\sum_{i=1}^{h} \sum_{j=1}^{k} \frac{n_{ij}^2}{n_i \cdot n_{\cdot j}} - 1 \right)$$

parameter

$$df = (h-1)(k-1)$$

p.value

$$P(\chi_{df}^2 \ge c)$$

observed

$$n_{ij} \quad \forall i = 1, 2, ..., h \quad \forall j = 1, 2, ..., k$$

expected

$$\hat{n}_{ij} \quad \forall i = 1, 2, \dots, h \quad \forall j = 1, 2, \dots, k$$

residuals

$$\frac{n_{ij} - \hat{n}_{ij}}{\sqrt{\hat{n}_{ij}}} \quad \forall i = 1, 2, ..., h \quad \forall j = 1, 2, ..., k$$

• Examples:

```
> h <- 3
> k <- 3
> chisq.test(x)
        Pearson's Chi-squared test
data: x
X-squared = 22.9907, df = 4, p-value = 0.0001272
> res <- chisq.test(x)</pre>
> res$statistic
X-squared
22.99074
> res$parameter
df
> res$p.value
[1] 0.0001271668
> res$observed
  A B C
A 2 21 43
B 10 11 32
C 23 12 30
> res$expected
             В
A 12.55435 15.78261 37.66304
B 10.08152 12.67391 30.24457
C 12.36413 15.54348 37.09239
> res$residuals
                  В
           Α
A -2.97875184 1.3133002 0.8696329
B -0.02567500 -0.4701945 0.3191986
C 3.02476204 -0.8987847 -1.1645289
```

Test di McNemar

Package: stats
Sintassi: mcnemar.test()
Input:

matrice di dimensione n × n contenente frequenze assolute

• Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà

```
p.value p-value
```

• Formula:

statistic
$$c=\sum_{i=1}^n\sum_{j=i+1}^n\frac{(n_{ij}-n_{ji})^2}{n_{ij}+n_{ji}}$$
 parameter
$$d\!f=n\left(n-1\right)/2$$
 p.value
$$P(\chi^2_{d\!f}\geq c)$$

• Examples:

```
> x < -matrix(data = c(2, 10, 23, 21, 11, 12, 43, 32, 30), nrow = 3,
+ ncol = 3)
> riga <- c("A", "B", "C")</pre>
> colonna <- c("A", "B", "C")</pre>
> dimnames(x) <- list(riga, colonna)</pre>
> x
 A B C
A 2 21 43
B 10 11 32
C 23 12 30
> n < -3
> mcnemar.test(x)
        McNemar's Chi-squared test
data: x
McNemar's chi-squared = 19.0547, df = 3, p-value = 0.0002664
> res <- mcnemar.test(x)</pre>
> res$statistic
McNemar's chi-squared
             19.05474
> res$parameter
df
 3
> res$p.value
[1] 0.0002663652
```

11.4 Comandi utili per le tabelle di contingenza

margin.table()

```
• Package: base
```

• Input:

```
x matrice di dimensione h \times k contenente frequenze assolute margin = NULL / 1 / 2 marginale assoluto totale, di riga o di colonna
```

- **Description:** distribuzione marginale assoluta
- Formula:

```
\begin{array}{c} \text{margin = NULL} \\ \\ n.. \\ \\ \hline margin = 1 \\ \\ n_i. \quad \forall i = 1, 2, \ldots, h \\ \\ \hline \text{margin = 2} \\ \\ n._j \quad \forall j = 1, 2, \ldots, k \\ \end{array}
```

• Example 1:

• Example 3:

```
> x < -matrix(data = c(1, 3, 0, 1, 3, 2, 2, 1, 2), nrow = 3, ncol = 3,
     byrow = TRUE)
> riga <- c("a", "b", "c")</pre>
> colonna <- c("A", "B", "C")</pre>
> dimnames(x) <- list(riga, colonna)</pre>
 A B C
a 1 3 0
b 1 3 2
c 2 1 2
> h < - 3
> k <- 3
> margin.table(x, margin = 1)
a b c
4 6 5
> margin.table(x, margin = 2)
A B C
4 7 4
```

prop.table()

- Package: base
- Input:
 - x matrice di dimensione $h \times k$ contenente frequenze assolute margin = NULL / 1 / 2 frequenza relativa totale, di riga o di colonna
- Description: distribuzione relativa
- Formula:

```
АВС
 a 1 3 0
 b 1 3 2
 c 2 1 2
 > h <- 3
 > k <- 3
 > prop.table(x, margin = NULL)
                        В
 a 0.06666667 0.20000000 0.0000000
 b 0.06666667 0.20000000 0.1333333
 c 0.13333333 0.06666667 0.1333333
• Example 2:
 > x < -matrix(data = c(1, 3, 0, 1, 3, 2, 2, 1, 2), nrow = 3, ncol = 3,
 + byrow = TRUE)
 > riga <- c("a", "b", "c")</pre>
 > colonna <- c("A", "B", "C")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
   A B C
 a 1 3 0
 b 1 3 2
 c 2 1 2
 > h <- 3
 > k < - 3
 > prop.table(x, margin = 1)
                В
            Α
 a 0.2500000 0.75 0.0000000
 b 0.1666667 0.50 0.3333333
 c 0.4000000 0.20 0.4000000
• Example 3:
 > x < -matrix(data = c(1, 3, 0, 1, 3, 2, 2, 1, 2), nrow = 3, ncol = 3,
 + byrow = TRUE)
 > riga <- c("a", "b", "c")</pre>
 > colonna <- c("A", "B", "C")</pre>
 > dimnames(x) <- list(riga, colonna)</pre>
 > x
   A B C
 a 1 3 0
 b 1 3 2
 c 2 1 2
 > h <- 3
 > k < - 3
 > prop.table(x, margin = 2)
      Α
                в с
 a 0.25 0.4285714 0.0
 b 0.25 0.4285714 0.5
 c 0.50 0.1428571 0.5
```

xtabs()

```
• Package: stats
```

• Input:

```
y vettore numerico di dimensione n
```

- f fattore a k livelli
- g fattore a h livelli
- Description: costruzione di una tabella di contingenza a partire da un dataframe
- Examples:

```
> y <- c(1.2, 2.1, 1.1, 2.3, 5.4, 4.3, 3.1, 2.3, 4.3, 5.4, 5.5,
> f <- factor(rep(letters[1:2], each = 6))</pre>
Levels: a b
> g <- factor(rep(LETTERS[2:1], times = 6))</pre>
[1] B A B A B A B A B A B A
Levels: A B
> data.frame(f, g, y)
  f g y
1 a B 1.2
2 a A 2.1
3
  a B 1.1
  a A 2.3
4
  a B 5.4
6 a A 4.3
7 b B 3.1
8 b A 2.3
9 b B 4.3
10 b A 5.4
11 b B 5.5
12 b A 5.7
> xtabs(y \sim f + g)
     Α
 a 8.7 7.7
 b 13.4 12.9
```

ftable()

• Package: stats

• Input:

```
x oggetto di tipo table contenente frequenze assolute
row.vars variabili di riga
col.vars variabili di colonna
```

• **Description:** costruzione di flat tables

• Examples:

```
> Titanic
, , Age = Child, Survived = No
     Sex
Class Male Female
      0
 1st
 2nd
        0
              0
             17
 3rd 35
 Crew 0
              0
, , Age = Adult, Survived = No
     Sex
Class Male Female
 1st 118 4
 2nd 154
             13
 3rd 387
             89
 Crew 670
              3
, , Age = Child, Survived = Yes
     Sex
Class Male Female
 1st 5 1
 2nd
       11
             13
      13
 3rd
             14
 Crew 0
             0
, , Age = Adult, Survived = Yes
     Sex
Class Male Female
      57 140
       14
             80
 2nd
       75
             76
 3rd
 Crew 192
              20
> ftable(x = Titanic, row.vars = c("Class", "Sex", "Age"), col.vars = c("Survived"))
                Survived No Yes
Class Sex
           Age
    Male Child
                         0
1st
                         118
                            57
           Adult
                         0 1
     Female Child
          Adult
                         4 140
2nd
    Male Child
                         0 11
           Adult
                        154 14
                         0 13
     Female Child
                        13 80
           Adult
3rd
     Male Child
                         35
                             13
                             75
           Adult
                        387
     Female Child
                         17 14
                        89 76
          Adult
    Male Child
                         0 0
Crew
           Adult
                        670 192
     Female Child
                         0 0
                          3 20
           Adult
> ftable(x = Titanic, row.vars = c("Age"), col.vars = c("Sex"))
```

```
Sex Male Female
Age
Child 64 45
Adult 1667 425
```

summary()

• Package: base

• Input:

x oggetto di tipo table di dimensione $h \times k$ contenente frequenze assolute

• **Description:** test χ^2 di indipendenza

• Output:

```
n.cases totale frequenze statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p\text{-value}
```

• Formula:

```
n.cases n.. statistic c=\sum_{i=1}^h\sum_{j=1}^k\frac{(n_{ij}-\hat{n}_{ij})^2}{\hat{n}_{ij}}=n..\left(\sum_{i=1}^h\sum_{j=1}^k\frac{n_{ij}^2}{n_{i\cdot}n_{\cdot j}}-1\right) parameter df=(h-1)\left(k-1\right) p.value P(\chi^2_{df}\geq c)
```

```
> f <- factor(c("a", "b", "c", "b", "a", "c", "a", "b", "b", "c",</pre>
     "a"))
[1] abcbacabbca
Levels: a b c
> g <- factor(c("A", "S", "A", "S", "S", "S", "A", "S", "A",
     "A"))
> q
[1] A S A S S S A S S A A
Levels: A S
> x < - table(f, g)
> x
  g
f AS
 a 3 1
 b 0 4
 c 2 1
```

```
> h <- 3
 > k <- 2
 > summary(x)
 Number of cases in table: 11
 Number of factors: 2
 Test for independence of all factors:
        Chisq = 5.286, df = 2, p-value = 0.07114
        Chi-squared approximation may be incorrect
 > res <- summary(x)</pre>
 > res$n.cases
 [1] 11
 > res$statistic
 [1] 5.286111
 > res$parameter
 [1] 2
 > res$p.value
 [1] 0.07114355
• Example 2:
 "a"))
  [1] ababaabbaba
 Levels: a b
 > g <- factor(c("A", "S", "A", "S", "S", "S", "A", "S", "A",
      "A"))
  [1] A S A S S S A S S A A
 Levels: A S
 > x < - table(f, g)
 > x
   g
 f AS
   a 3 3
  b 2 3
 > h <- 2
 > k < - 2
 > summary(x)
 Number of cases in table: 11
 Number of factors: 2
 Test for independence of all factors:
        Chisq = 0.11, df = 1, p-value = 0.7401
        Chi-squared approximation may be incorrect
```

```
> res <- summary(x)
> res$n.cases

[1] 11
> res$statistic

[1] 0.11
> res$parameter

[1] 1
> res$p.value

[1] 0.7401441
```

Capitolo 12

Test di ipotesi sull'adattamento

12.1 Test di ipotesi sulla distribuzione normale

Test di Kolmogorov - Smirnov

```
• Package: stats
```

• Sintassi: ks.test()

• Input:

x vettore numerico di n valori distinti

- **Description:** test di ipotesi per $H_0: F_0(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$ contro $H_1: F_0(x) \neq \Phi\left(\frac{x-\mu}{\sigma}\right)$
- Output:

statistic valore empirico della statistica D

• Formula:

statistic
$$d = \max_{1 \leq i \leq n} \left\{ \max \left[\frac{i}{n} - F_0\left(x_{(i)}\right), \, F_0\left(x_{(i)}\right) - \frac{i-1}{n} \right] \right\}$$

$$\mbox{dove} \quad F_0\left(x_{(i)}\right) \,=\, \Phi\left(\frac{x_{(i)}-\mu}{\sigma}\right) \quad \forall \, i \,=\, 1,\, 2,\, \ldots,\, n$$

• Example 1:

```
> x <- c(0.1, 2.3, 4.3, 4.2, 5.6, 7.21, 8.2)
> n <- 7
> x <- sort(x)
> x

[1] 0.10 2.30 4.20 4.30 5.60 7.21 8.20

> Fo <- pnorm(x, mean = 3.3, sd = 1.2)
> vettore1 <- (1:n)/n - Fo
> vettore2 <- Fo - ((1:n) - 1)/n
> d <- max(pmax(vettore1, vettore2))
> d

[1] 0.4876584

> ks.test(x, "pnorm", 3.3, 1.2)$statistic
D
0.4876584
```

```
> x < -c(1.1, 3.4, 5.6, 7.8, 2.3, 4.5, 1.2, 2.2)
    > n <- 8
    > x <- sort(x)
    [1] 1.1 1.2 2.2 2.3 3.4 4.5 5.6 7.8
    > Fo <- pnorm(x, mean = 4.1, sd = 2.3)
    > vettore1 <- (1:n)/n - Fo
    > vettore2 <- Fo - ((1:n) - 1)/n
    > d <- max(pmax(vettore1, vettore2))</pre>
    > d
    [1] 0.2830715
    > ks.test(x, "pnorm", 4.1, 2.3)$statistic
             D
    0.2830715
  • Example 3:
    > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.8)
    > n <- 8
    > x <- sort(x)
    [1] 1.1 2.3 3.4 4.5 5.6 6.7 6.8 8.9
    > Fo <- pnorm(x, mean = 6.3, sd = 1.1)
    > vettore1 <- (1:n)/n - Fo
    > vettore2 <- Fo - ((1:n) - 1)/n
    > d <- max(pmax(vettore1, vettore2))</pre>
    > d
    [1] 0.4491182
    > ks.test(x, "pnorm", 6.3, 1.1)$statistic
    0.4491182
Test di Jarque - Bera
  • Package: tseries
  • Sintassi: jarque.bera.test()
  • Input:
        \times vettore numerico di dimensione n
  • Output:
        statistic valore empirico della statistica \chi^2
        parameter gradi di libertà
        p.value p-value
```

• Formula:

statistic
$$c=\frac{n}{6}\left(\frac{m_3}{m_2^{3/2}}\right)^2+\frac{n}{24}\left(\frac{m_4}{m_2^2}-3\right)^2$$

$${\rm dove}\ m_k=\frac{1}{n}\sum_{i=1}^n\left(x_i-\bar x\right)^k\ \ \forall\,k=2,\,3,\,4$$
 parameter
$$d\!f=2$$
 p.value
$$P(\chi^2_{d\!f}\geq c)$$

• Example 1:

• Example 2:

```
> x < -c(0.1, 2.3, 4.3, 4.2, 5.6, 7.21, 8.2)
> m2 <- mean((x - mean(x))^2)
> m2
[1] 6.650012
> m3 <- mean((x - mean(x))^3)
> m3
[1] -4.594487
> m4 <- mean((x - mean(x))^4)
> m4
[1] 92.51966
> c <- (n/6) * (m3/m2^(3/2))^2 + (n/24) * (m4/m2^2 - 3)^2
> C
[1] 0.3241426
> jarque.bera.test(x)$statistic
X-squared
0.3241426
> jarque.bera.test(x)$parameter
df
> p.value < 1 - pchisq(c, df = 2)
> p.value
[1] 0.8503806
> jarque.bera.test(x)$p.value
X-squared
0.8503806
```

479

```
> x \leftarrow c(1.1, 3.4, 5.6, 7.8, 2.3, 4.5, 1.2, 2.2, 1.1)
 > n <- 9
 > m2 <- mean((x - mean(x))^2)
 [1] 4.806914
 > m3 <- mean((x - mean(x))^3)
 > m3
 [1] 8.816102
 > m4 <- mean((x - mean(x))^4)
 > m4
 [1] 58.41274
 > c <- (n/6) * (m3/m2^(3/2))^2 + (n/24) * (m4/m2^2 - 3)^2
 [1] 1.133201
 > jarque.bera.test(x)$statistic
 X-squared
  1.133201
 > jarque.bera.test(x)$parameter
 df
  2
 > p.value < 1 - pchisq(c, df = 2)
 > p.value
 [1] 0.5674513
 > jarque.bera.test(x)$p.value
 X-squared
 0.5674513
• Example 3:
 > x \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > n <- 8
 > m2 <- mean((x - mean(x))^2)
 > m2
 [1] 5.8225
 > m3 <- mean((x - mean(x))^3)
 > m3
 [1] 0.015
 > m4 <- mean((x - mean(x))^4)
 > m4
```

```
[1] 67.06683

> c <- (n/6) * (m3/m2^(3/2))^2 + (n/24) * (m4/m2^2 - 3)^2
> c

[1] 0.347969

> jarque.bera.test(x)$statistic

X-squared
    0.347969

> jarque.bera.test(x)$parameter

df
    2

> p.value <- 1 - pchisq(c, df = 2)
> p.value

[1] 0.8403099

> jarque.bera.test(x)$p.value

X-squared
    0.8403099
```

Test di Cramer - von Mises

• Package: nortest

• Sintassi: cvm.test()

• Input:

 \mathbf{x} vettore numerico di dimensione $n \geq 7$

• Output:

statistic valore empirico della statistica ${\it Z}$ p.value ${\it p}$ -value

• Formula:

statistic

$$W = \frac{1}{12n} + \sum_{i=1}^{n} \left[\Phi\left(\frac{x_{(i)} - \bar{x}}{s_x}\right) - \frac{2i - 1}{2n} \right]^2$$

p.value

$$WW = (1 + 0.5 / n) W$$

ww	< 0.0275	$\geq 0.0275 \text{ AND} < 0.051$	
p.value	$1 - e^{-13.953 + 775.5 WW - 12542.61 WW^2}$	$1 - e^{-5.903 + 179.546 WW - 1515.29 WW^2}$	
ww	$\geq 0.051~\text{AND} < 0.092$	≥ 0.092	
p.value	$e^{0.886-31.62WW+10.897WW^2}$	$e^{1.111-34.242WW+12.832WW^2}$	

```
> x \leftarrow c(1.1, 1.2, 2.2, 2.3, 3.4, 4.5, 5.6, 7.8)
 > n <- 8
 > x <- sort(x)
 > W <- 1/(12 * n) + sum((pnorm((x - mean(x)))/sd(x)) - (2 * (1:n) -
      1)/(2 * n))^2
 > W
 [1] 0.04611184
 > cvm.test(x)$statistic
 0.04611184
 > WW <- (1 + 0.5/n) * W
 > WW
 [1] 0.04899383
 > p.value < 1 - exp(-5.903 + 179.546 * WW - 1515.29 * WW^2)
 > p.value
 [1] 0.5246239
 > cvm.test(x)$p.value
 [1] 0.5246239
• Example 2:
 > x <- c(80, 96.19, 98.07, 99.7, 99.79, 99.81, 101.14, 101.6, 103.44,
 > n < -10
 > x <- sort(x)
 > W <- (1/(12 * n)) + sum((pnorm((x - mean(x)))/sd(x)) - (2 * (1:n) -
      1)/(2 * n))^2
 > W
 [1] 0.2296694
 > cvm.test(x)$statistic
 0.2296694
 > WW <- (1 + 0.5/n) * W
 > WW
 [1] 0.2411529
 > p.value <- exp(1.111 - 34.242 * WW + 12.832 * WW^2)
 > p.value
 [1] 0.001661032
 > cvm.test(x)$p.value
 [1] 0.001661032
```

• Example 3:

Test di Anderson - Darlin

```
Package: nortest
Sintassi: ad.test()
Input:

x vettore numerico di dimensione n ≥ 7
```

• Output:

statistic valore empirico della statistica ${\it Z}$ p.value ${\it p}$ -value

• Formula:

statistic

$$A = -n - \frac{1}{n} \sum_{i=1}^{n} (2i - 1) \left[\log \left(\Phi \left(\frac{x_{(i)} - \bar{x}}{s_x} \right) \right) + \log \left(1 - \Phi \left(\frac{x_{(n-i+1)} - \bar{x}}{s_x} \right) \right) \right]$$

p.value

$$AA = (1 + 0.75 / n + 2.25 / n^2) A$$

AA	< 0.2	$\geq 0.2 \text{ AND} < 0.34$
p.value	$1 - e^{-13.436 + 101.14 AA - 223.73 AA^2}$	$1 - e^{-8.318 + 42.796 AA - 59.938 AA^2}$
AA	$\geq 0.34 \text{ AND} < 0.6$	≥ 0.6
p.value	$e^{0.9177-4.279AA-1.38AA^2}$	$e^{1.2937-5.709AA+0.0186AA^2}$

```
> x <- c(99.7, 99.79, 101.14, 99.32, 99.27, 101.29, 100.3, 102.4,
 + 105.2)
 > n <- 9
 > x <- sort(x)
 > A <- -n - mean((2 * (1:n) - 1) * (log(pnorm((x - mean(x))/sd(x))) +
       log(1 - pnorm((rev(x) - mean(x))/sd(x))))
 [1] 0.5914851
 > ad.test(x)$statistic
         Α
 0.5914851
 > AA <- (1 + 0.75/n + 2.25/n^2) * A
 > AA
 [1] 0.6572057
 > p.value <- exp(1.2937 - 5.709 * AA + 0.0186 * AA^2)
 > p.value
 [1] 0.08627171
 > ad.test(x)$p.value
 [1] 0.08627171
• Example 2:
 > x \leftarrow c(1.1, 1.2, 2.2, 2.3, 3.4, 4.5, 5.6, 7.8)
 > n <- 8
 > x <- sort(x)
 > A <- -n - mean((2 * (1:n) - 1) * (log(pnorm((x - mean(x))/sd(x))) +
      log(1 - pnorm((rev(x) - mean(x))/sd(x))))
 > A
 [1] 0.3073346
 > ad.test(x)$statistic
 0.3073346
 > AA <- (1 + 0.75/n + 2.25/n^2) * A
 > AA
 [1] 0.346952
 > p.value <- exp(0.9177 - 4.279 * AA - 1.38 * AA^2)
 > p.value
```

```
[1] 0.480453
 > ad.test(x)$p.value
 [1] 0.480453
• Example 3:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > n <- 8
 > x <- sort(x)
 > A <- -n - mean((2 * (1:n) - 1) * (log(pnorm((x - mean(x))/sd(x))) +
      log(1 - pnorm((rev(x) - mean(x))/sd(x))))
 > A
 [1] 0.1546968
 > ad.test(x)$statistic
 0.1546968
 > AA <- (1 + 0.75/n + 2.25/n^2) * A
 > AA
 [1] 0.1746381
 > p.value < 1 - exp(-13.436 + 101.14 * AA - 223.73 * AA^2)
 > p.value
 [1] 0.9254678
 > ad.test(x)$p.value
 [1] 0.9254678
```

Test di Shapiro - Francia

```
• Package: nortest
```

• Sintassi: sf.test()

• Input:

x vettore numerico di dimensione $5 \le n \le 5000$

• Output:

statistic valore empirico della statistica ${\it Z}$ p.value ${\it p}\text{-}{\rm value}$

• Formula:

statistic

$$W = \frac{\left(\sum_{i=1}^{n} x_{(i)} y_{i} - n \bar{x} \bar{y}\right)^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2} \sum_{i=1}^{n} (y_{i} - \bar{y})^{2}}$$

dove
$$y_i = \Phi^{-1}\left(\frac{i-3/8}{n+1/4}\right) \, \forall i = 1, 2, ..., n$$

p.value

$$1 - \Phi(z)$$

```
dove z = \frac{\log(1-W) - [-1.2725 + 1.0521 [\log(\log(n)) - \log(n)]]}{1.0308 - 0.26758 [\log(\log(n)) + 2/\log(n)]}
```

```
• Example 1:
```

```
> x < -c(7.7, 5.6, 4.3, 3.2, 3.1, 2.2, 1.2, 1)
> n <- 8
> x <- sort(x)
> y <- qnorm(((1:n) - 3/8)/(n + 1/4))
> W <- cor(x, y)^2
> W
[1] 0.9420059
> sf.test(x)$statistic
0.9420059
> z <- (\log(1 - W) - (-1.2725 + 1.0521 * (\log(\log(n)) - \log(n))))/(1.0308 - (-1.2725 + 1.0521))
+ 0.26758 * (log(log(n)) + 2/log(n)))
> z
[1] -0.2724882
> p.value <- 1 - pnorm(z)
> p.value
[1] 0.6073767
> sf.test(x)$p.value
[1] 0.6073767
```

```
> p.value <- 1 - pnorm(z)</pre>
 > p.value
 [1] 0.1292269
 > sf.test(x)$p.value
 [1] 0.1292269
• Example 3:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > n <- 8
 > x <- sort(x)
 > y <- qnorm(((1:n) - 3/8)/(n + 1/4))
 > W < -cor(x, y)^2
 > W
 [1] 0.9838034
 > sf.test(x)$statistic
 0.9838034
 > z <- (log(1 - W) - (-1.2725 + 1.0521 * (log(log(n)) - log(n))))/(1.0308 - (log(n)) - log(n))))
 + 0.26758 * (log(log(n)) + 2/log(n)))
 > z
 [1] -2.48103
 > p.value <- 1 - pnorm(z)</pre>
 > p.value
 [1] 0.9934498
 > sf.test(x)$p.value
 [1] 0.9934498
```

Test di Lilliefors

```
    Package: nortest
    Sintassi: lillie.test()
    Input:

            x vettore numerico di dimensione n ≥ 5

    Output:

            statistic valore empirico della statistica Z
            p.value p-value
```

• Formula:

n	$n \le 100$	n > 100
Kd	D	$(n/100)^{0.49} D$
nd	n	100

statistic

$$D = \max(a, b)$$

$$\begin{array}{ll} \text{dove} & \quad \textbf{a} \, = \, \max \left\{ \frac{i}{n} - \Phi \left(\frac{x_{(i)} - \bar{x}}{s_x} \right) \right\}_{i \, = \, 1, \, 2, \, \dots, \, n} \\ \\ \textbf{b} \, = \, \max \left\{ \Phi \left(\frac{x_{(i)} - \bar{x}}{s_x} \right) - \frac{i - 1}{n} \right\}_{i \, = \, 1, \, 2, \, \dots, \, n} \end{array}$$

p.value

$$pvalue = e^{-7.01256\,Kd^2\,(nd+2.78019)\,+2.99587\,Kd\,\sqrt{nd+2.78019}-0.122119+\frac{0.974598}{\sqrt{nd}}+\frac{1.67997}{nd}}$$

$$pvalue \leq 0.1$$

$$p.value = pvalue$$

$$pvalue > 0.1$$

$$kk = (\sqrt{n} - 0.01 + 0.85\,/\,\sqrt{n})\,D$$

kk	p.value
≤ 0.302	1
≤ 0.5	$2.76773 - 19.828315kk + 80.709644kk^2 - 138.55152kk^3 + 81.218052kk^4$
≤ 0.9	$-4.901232 + 40.662806 kk - 97.490286 kk^2 + 94.029866 kk^3 - 32.355711 kk^4$
≤ 1.31	$6.198765 - 19.558097 kk + 23.186922 kk^2 - 12.234627 kk^3 + 2.423045 kk^4$
> 1.31	0

```
> x <- c(1.1, 1.2, 2.2, 2.3, 3.4, 4.5, 5.6, 7.8)
> n <- 8
> x <- sort(x)
> a <- max((1:n)/n - pnorm((x - mean(x))/sd(x)))
> a

[1] 0.1983969
> b <- max(pnorm((x - mean(x))/sd(x)) - ((1:n) - 1)/n)
> b

[1] 0.1505139
> D <- max(a, b)
> D

[1] 0.1983969
> lillie.test(x)$statistic
D
0.1983969
```

```
> Kd <- D
 > nd <- n
 > pvalue <- exp(-7.01256 * Kd^2 * (nd + 2.78019) + 2.99587 * Kd *
       sqrt(nd + 2.78019) - 0.122119 + 0.974598/sqrt(nd) + 1.67997/nd)
 > pvalue
 [1] 0.5534262
 > kk <- (sqrt(n) - 0.01 + 0.85/sqrt(n)) * D
 > kk
 [1] 0.6187895
 > p.value <- -4.901232 + 40.662806 * kk - 97.490286 * kk^2 + 94.029866 *
 + kk^3 - 32.355711 * kk^4
 > p.value
 [1] 0.4665968
 > lillie.test(x)$p.value
 [1] 0.4665968
• Example 2:
 > x \leftarrow c(42.3, 31.4, 11.2, 9, 8.5, 7.5, 5.6, 2.3)
 > n <- 8
 > x <- sort(x)
 > a <- max((1:n)/n - pnorm((x - mean(x))/sd(x)))
 [1] 0.3479997
 > b <- max(pnorm((x - mean(x))/sd(x)) - ((1:n) - 1)/n)
 > b
 [1] 0.1908506
 > D <- max(a, b)
 > D
 [1] 0.3479997
 > lillie.test(x)$statistic
 0.3479997
 > Kd <- D
 > nd <- n
 > pvalue <- exp(-7.01256 * Kd^2 * (nd + 2.78019) + 2.99587 * Kd *
       sqrt(nd + 2.78019) - 0.122119 + 0.974598/sqrt(nd) + 1.67997/nd)
 > pvalue
 [1] 0.004993897
 > p.value <- pvalue
 > p.value
 [1] 0.004993897
```

```
> lillie.test(x)$p.value
 [1] 0.004993897
• Example 3:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > n <- 8
 > x <- sort(x)
 > a <- max((1:n)/n - pnorm((x - mean(x))/sd(x)))
 [1] 0.1176558
 > b <- max(pnorm((x - mean(x))/sd(x)) - ((1:n) - 1)/n)
 > b
 [1] 0.1323442
 > D <- max(a, b)
 > D
 [1] 0.1323442
 > lillie.test(x)$statistic
         D
 0.1323442
 > Kd <- D
 > nd <- n
 > pvalue <- exp(-7.01256 * Kd^2 * (nd + 2.78019) + 2.99587 * Kd *
 + sqrt (nd + 2.78019) - 0.122119 + 0.974598/sqrt (nd) + 1.67997/nd)
 > pvalue
 [1] 1.507065
 > kk <- (sqrt(n) - 0.01 + 0.85/sqrt(n)) * D
 > kk
 [1] 0.4127748
 > p.value <- 2.76773 - 19.828315 \star kk + 80.709644 \star kk^2 - 138.55152 \star
       kk^3 + 81.218052 * kk^4
 > p.value
 [1] 0.9481423
 > lillie.test(x)$p.value
 [1] 0.9481423
```

Test di Anscombe - Glynn

• Package: moments

• Sintassi: anscombe.test()

• Input:

 ${\tt x}$ vettore numerico di dimensione n alternative = "less" / "greater" / "two.sided" ipotesi alternativa

• Output:

statistic valore empirico della statistica Z p.value p-value alternative ipotesi alternativa

• Formula:

statistic

$$z \, = \, \frac{1 - \frac{2}{9 \, a} - \left(\frac{1 - 2 \, / \, a}{1 + xx \sqrt{2 \, / \, (a - 4)}}\right)^{1 \, / \, 3}}{\sqrt{\frac{2}{9 \, a}}}$$

dove

$$b = \frac{1}{n} \sum_{i=1}^{n} \left(\frac{x_i - \bar{x}}{\sigma_x}\right)^4$$

$$eb2 = \frac{3(n-1)}{(n+1)}$$

$$vb2 = \frac{24 n (n-2) (n-3)}{(n+1)^2 (n+3) (n+5)}$$

$$m3 = \frac{6(n^2 - 5 n + 2)}{(n+7) (n+9)} \sqrt{\frac{6(n+3) (n+5)}{n (n-2) (n-3)}}$$

$$a = 6 + \frac{8}{m3} \left(\frac{2}{m3} + \sqrt{1 + \frac{4}{m3}}\right)$$

$$xx = (b - eb2) / \sqrt{vb2}$$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1 - \Phi(z)$	$2\Phi(-\mid z\mid)$

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> n <- length(x)
> b <- n * sum((x - mean(x))^4)/(sum((x - mean(x))^2)^2)
> eb2 <- 3 * (n - 1)/(n + 1)
> vb2 <- 24 * n * (n - 2) * (n - 3)/((n + 1)^2 * (n + 3) * (n + 5))
> m3 <- (6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9))) * sqrt((6 * (n + 3) * (n + 5))/(n * (n - 2) * (n - 3)))
> a <- 6 + (8/m3) * (2/m3 + sqrt(1 + 4/m3))
> xx <- (b - eb2)/sqrt(vb2)
> res <- anscombe.test(x, alternative = "two.sided")
> z <- (1 - 2/(9 * a) - ((1 - 2/a)/(1 + xx * sqrt(2/(a - 4))))^(1/3))/sqrt(2/(9 * a))
> c(b, z)
[1] 1.8382073 -0.9304068
```

```
> res$statistic
                  kurt
      1.8382073 -0.9304068
   > p.value <- 2 * pnorm(-abs(z))
   > p.value
   [1] 0.3521605
   > res$p.value
    [1] 0.3521605
• Example 2:
   > x < -c(1, 2.3, 4.5, 6.7, 8.9)
   > n <- length(x)</pre>
   > b <- n * sum((x - mean(x))^4)/(sum((x - mean(x))^2)^2)
   > eb2 <- 3 * (n - 1)/(n + 1)
   > vb2 < -24 * n * (n - 2) * (n - 3)/((n + 1)^2 * (n + 3) * (n + 3)
                  5))
   > m3 < -(6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9)))) * sqrt((6 * (n + 7) * (n + 9))
                  (n + 3) * (n + 5))/(n * (n - 2) * (n - 3)))
   > a <- 6 + (8/m3) * (2/m3 + sqrt(1 + 4/m3))
   > xx <- (b - eb2)/sqrt(vb2)
   > res <- anscombe.test(x, alternative = "two.sided")</pre>
   > z <- (1 - 2/(9 * a) - ((1 - 2/a)/(1 + xx * sqrt(2/(a - 4))))^(1/3))/sqrt(2/(9 * a))
               a))
   > c(b, z)
    [1] 1.623612 -0.734540
   > res$statistic
                kurt
      1.623612 -0.734540
   > p.value <- 2 * pnorm(-abs(z))
   > p.value
   [1] 0.4626197
   > res$p.value
    [1] 0.4626197
• Example 3:
   > x < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
   > n <- length(x)
   > b <- n * sum((x - mean(x))^4)/(sum((x - mean(x))^2)^2)
   > eb2 <- 3 * (n - 1)/(n + 1)
   > vb2 < -24 * n * (n - 2) * (n - 3)/((n + 1)^2 * (n + 3) * (n + 3))
                  5))
    > m3 <- (6 * (n^2 - 5 * n + 2)/((n + 7) * (n + 9))) * sqrt((6 *
                 (n + 3) * (n + 5))/(n * (n - 2) * (n - 3)))
   > a <- 6 + (8/m3) * (2/m3 + sqrt(1 + 4/m3))
   > xx <- (b - eb2)/sqrt(vb2)
   > res <- anscombe.test(x, alternative = "two.sided")</pre>
   > z <- (1 - 2/(9 * a) - ((1 - 2/a)/(1 + xx * sqrt(2/(a - 4))))^(1/3))/sqrt(2/(9 * a))
                a))
   > c(b, z)
```

Test di Bonett - Seier

• Package: moments

• Sintassi: bonett.test()

• Input:

 $\begin{tabular}{ll} $\tt x$ & vettore & numerico & di & dimensione & n \\ & alternative & = "less" & / "greater" & / "two.sided" & ipotesi & alternativa \\ \end{tabular}$

• Output:

statistic valore empirico della statistica Z p.value $p ext{-}\mathrm{value}$ alternative ipotesi alternativa

• Formula:

statistic

$$z\,=\,\sqrt{n+2}\,\left(13.29\,\log\left(\rho\,/\,\tau\right)-3\right)\,/\,3.54$$

dove
$$\rho = \sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_i - \bar{x})^2}$$
 e $\tau = \frac{1}{n} \sum_{i=1}^{n} |x_i - \bar{x}|$

p.value

alternative	less	greater	two.sided
p.value	$\Phi(z)$	$1-\Phi(z)$	$2\Phi(- z)$

```
> x <- c(7.8, 6.6, 6.5, 7.4, 7.3, 7, 6.4, 7.1, 6.7, 7.6, 6.8)
> n <- length(x)
> rho <- sqrt((n - 1) * var(x)/n)
> tau <- mean(abs(x - mean(x)))
> res <- bonett.test(x, alternative = "two.sided")
> z <- sqrt(n + 2) * (13.29 * log(rho/tau) - 3)/3.54
> c(tau, z)
[1] 0.3834711 -1.1096692
```

```
> res$statistic
        tau
  0.3834711 -1.1096692
 > p.value <-2 * pnorm(-abs(z))
 > p.value
 [1] 0.2671416
 > res$p.value
 [1] 0.2671416
• Example 2:
 > x < -c(1, 2.3, 4.5, 6.7, 8.9)
 > n <- length(x)
 > rho <- sqrt((n - 1) * var(x)/n)
 > tau <- mean(abs(x - mean(x)))
 > res <- bonett.test(x, alternative = "two.sided")</pre>
 > z <- sqrt(n + 2) * (13.29 * log(rho/tau) - 3)/3.54
 > c(tau, z)
 [1] 2.49600 -0.86214
 > res$statistic
      tau
  2.49600 -0.86214
 > p.value <- 2 * pnorm(-abs(z))
 > p.value
 [1] 0.3886105
 > res$p.value
 [1] 0.3886105
• Example 3:
 > x <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- length(x)</pre>
 > rho <- sqrt((n - 1) * var(x)/n)
 > tau <- mean(abs(x - mean(x)))
 > res <- bonett.test(x, alternative = "two.sided")</pre>
 > z <- sqrt(n + 2) * (13.29 * log(rho/tau) - 3)/3.54
 > c(tau, z)
 [1] 1.785000 1.035715
 > res$statistic
      t.au
 1.785000 1.035715
 > p.value <-2 * pnorm(-abs(z))
 > p.value
 [1] 0.3003353
 > res$p.value
 [1] 0.3003353
```

12.2 Funzioni di adattamento normale

```
qqnorm()
```

```
• Package: stats
```

• Input:

```
y vettore numerico di dimensione n ordinato in maniera crescente plot.it = FALSE
```

- Description: quantili teorici e campionari per QQ-Norm
- Output:
 - x quantili teorici
 - y quantili campionari
- Formula:

```
\begin{cases} \Phi^{-1}\left(\left(8\,i-3\right)/\left(8\,n+2\right)\right) & \forall\,i\,=\,1,\,2,\,\ldots,\,n & \text{se }n\leq10 \\ \\ \Phi^{-1}\left(\left(i-1/2\right)/n\right) & \forall\,i\,=\,1,\,2,\,\ldots,\,n & \text{se }n>10 \end{cases} 
 y_{(i)} \quad \forall\,i\,=\,1,\,2,\,\ldots,\,n
```

• Example 1:

```
> y <- c(3.2, 1.4, 4.2, 12.4, 13.4, 17.3, 18.1)
> y <- sort(y)
> y

[1]  1.4  3.2  4.2  12.4  13.4  17.3  18.1

> n <- 7
> qqnorm(y, plot.it = FALSE)$y

[1]  1.4  3.2  4.2  12.4  13.4  17.3  18.1

> qnorm((8 * (1:n) - 3)/(8 * n + 2))

[1]  -1.3644887 -0.7582926 -0.3529340  0.0000000  0.3529340  0.7582926  1.3644887

> qqnorm(y, plot.it = FALSE)$x

[1]  -1.3644887 -0.7582926 -0.3529340  0.0000000  0.3529340  0.7582926  1.3644887
```

```
 \begin{smallmatrix} [1] & -1.7316644 & -1.1503494 & -0.8122178 & -0.5485223 & -0.3186394 & -0.1046335 \end{smallmatrix} 
   [7] 0.1046335 0.3186394 0.5485223 0.8122178 1.1503494 1.7316644
 > qqnorm(y, plot.it = FALSE)$x
    \begin{smallmatrix} [1] & -1.7316644 & -1.1503494 & -0.8122178 & -0.5485223 & -0.3186394 & -0.1046335 \end{smallmatrix} 
   [7] 0.1046335 0.3186394 0.5485223 0.8122178 1.1503494 1.7316644
• Example 3:
 > y < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- sort(y)
 [1] 1.1 2.3 3.4 4.5 5.6 6.7 6.7 8.9
 > n < - 8
 > qqnorm(y, plot.it = FALSE)$y
 [1] 1.1 2.3 3.4 4.5 5.6 6.7 6.7 8.9
 > qnorm((8 * (1:n) - 3)/(8 * n + 2))
   \begin{smallmatrix} 1 \end{smallmatrix} \end{bmatrix} - 1.4342002 - 0.8524950 - 0.4727891 - 0.1525060 0.1525060 0.4727891 0.8524950 
  [8] 1.4342002
 > qqnorm(y, plot.it = FALSE)$x
  \lceil 1 \rceil -1.4342002 -0.8524950 -0.4727891 -0.1525060 0.1525060 0.4727891 0.8524950
  [8] 1.4342002
```

ppoints()

- Package: stats
- Input:
 - n valore naturale
- **Description:** rapporti per QQ-Norm
- Formula:

$$\begin{cases} (8i-3)/(8n+2) & \forall i = 1, 2, ..., n & \text{se } n \le 10 \\ (i-1/2)/n & \forall i = 1, 2, ..., n & \text{se } n > 10 \end{cases}$$

• Example 1:

```
> n <- 5
> (8 * (1:n) - 3)/(8 * n + 2)

[1] 0.1190476 0.3095238 0.5000000 0.6904762 0.8809524
> ppoints(n = 5)

[1] 0.1190476 0.3095238 0.5000000 0.6904762 0.8809524
```

```
> n <- 12
> ((1:n) - 1/2)/n
```

```
[1] 0.04166667 0.12500000 0.20833333 0.29166667 0.37500000 0.45833333 [7] 0.54166667 0.62500000 0.70833333 0.79166667 0.87500000 0.95833333  
> ppoints(n = 12)

[1] 0.04166667 0.12500000 0.20833333 0.29166667 0.37500000 0.45833333 [7] 0.54166667 0.62500000 0.70833333 0.79166667 0.87500000 0.95833333  
• Example 3:

> n <- 15

> ((1:n) - 1/2)/n

[1] 0.03333333 0.10000000 0.16666667 0.23333333 0.30000000 0.36666667 [7] 0.43333333 0.9000000 0.96666667  
> ppoints(n = 15)

[1] 0.03333333 0.10000000 0.16666667 0.23333333 0.30000000 0.36666667 [7] 0.43333333 0.50000000 0.96666667  
| O.03333333 0.10000000 0.16666667 0.23333333 0.30000000 0.36666667 [7] 0.43333333 0.50000000 0.96666667 0.63333333 0.70000000 0.76666667 [7] 0.43333333 0.50000000 0.56666667 0.63333333 0.70000000 0.76666667 [13] 0.83333333 0.90000000 0.96666667
```

12.3 Test di ipotesi su una distribuzione generica

Test Chi - Quadrato GOF

- Package: stats
- Sintassi: chisq.test()
- Input:
 - ${\bf x}~$ vettore di frequenze assolute a somma n di dimensione k
 - p vettore p di probabilità a somma unitaria di dimensione k
- Output:

```
statistic valore empirico della statistica \chi^2 parameter gradi di libertà p.value p-value observed valori osservati expected valori attesi residuals residui di Pearson
```

• Formula:

statistic
$$c=\sum_{i=1}^k\frac{(n_i-\hat{n}_i)^2}{\hat{n}_i}=\sum_{i=1}^k\frac{n_i^2}{\hat{n}_i}-n$$

$$\text{dove}\quad \hat{n}_i=n\,p_i\quad\forall\,i=1,\,2,\,\ldots,\,k$$
 parameter
$$df=k-1$$

$$p.\text{value}$$

$$P(\chi^2_{df}\geq c)$$
 observed
$$n_i\quad\forall\,i=1,\,2,\,\ldots,\,k$$

```
expected \hat{n}_i = n\,p_i \quad \forall\, i=1,\,2,\,\ldots,\,k residuals \frac{n_i-\hat{n}_i}{\sqrt{\hat{n}_i}} \quad \forall\, i=1,\,2,\,\ldots,\,k
```

```
> x < -c(100, 110, 80, 55, 14)
> n <- sum(x)
> n
[1] 359
> prob < c(0.29, 0.21, 0.17, 0.17, 0.16)
> k <- 5
> osservati <- x
> attesi <- n * prob</pre>
> c <- sum((osservati - attesi)^2/attesi)</pre>
> C
[1] 55.3955
> chisq.test(x, p = prob)$statistic
X-squared
  55.3955
> parameter <- k - 1
> parameter
[1] 4
> chisq.test(x, p = prob)$parameter
df
 4
> p.value <- 1 - pchisq(c, df = parameter)</pre>
> p.value
[1] 2.684530e-11
> chisq.test(x, p = prob)$p.value
[1] 2.684534e-11
> osservati
[1] 100 110 80 55 14
> chisq.test(x, p = prob)$observed
[1] 100 110 80 55 14
> attesi
[1] 104.11 75.39 61.03 61.03 57.44
```

```
> chisq.test(x, p = prob)$expected
 [1] 104.11 75.39 61.03 61.03 57.44
 > residui <- (osservati - attesi)/sqrt(attesi)</pre>
 > residui
 [1] -0.4028057 3.9860682 2.4282626 -0.7718726 -5.7316888
 > chisq.test(x, p = prob)$residuals
 [1] -0.4028057 3.9860682 2.4282626 -0.7718726 -5.7316888
• Example 2:
 > x < -c(89, 37, 30, 28, 2)
 > n <- sum(x)
 > n
 [1] 186
 > prob <- c(0.4, 0.2, 0.2, 0.15, 0.05)
 > k <- 5
 > osservati <- x</pre>
 > attesi <- n * prob</pre>
 > c <- sum((osservati - attesi)^2/attesi)</pre>
 > C
 [1] 9.990143
 > chisq.test(x, p = prob)$statistic
 X-squared
  9.990143
 > parameter <- k - 1
 > parameter
 [1] 4
 > chisq.test(x, p = prob)$parameter
 df
  4
 > p.value <- 1 - pchisq(c, df = parameter)</pre>
 > p.value
 [1] 0.04059404
 > chisq.test(x, p = prob)$p.value
 [1] 0.04059404
 > osservati
 [1] 89 37 30 28 2
 > chisq.test(x, p = prob)$observed
```

```
[1] 89 37 30 28 2
 > attesi
 [1] 74.4 37.2 37.2 27.9 9.3
 > chisq.test(x, p = prob)$expected
 [1] 74.4 37.2 37.2 27.9 9.3
 > residui <- (osservati - attesi)/sqrt(attesi)</pre>
 > residui
 [1] 1.69264697 -0.03279129 -1.18048650 0.01893206 -2.39376430
 > chisq.test(x, p = prob)$residuals
 [1] 1.69264697 -0.03279129 -1.18048650 0.01893206 -2.39376430
• Example 3:
 > x < -c(54, 29, 5)
 > n <- sum(x)
 [1] 88
 > prob <- c(0.5, 0.25, 0.25)
 > k <- 3
 > osservati <- x</pre>
 > attesi <- n * prob</pre>
 > c <- sum((osservati - attesi)^2/attesi)</pre>
 [1] 17.63636
 > chisq.test(x, p = prob)$statistic
 X-squared
  17.63636
 > parameter <- k - 1
 > parameter
 [1] 2
 > chisq.test(x, p = prob)$parameter
 df
 > p.value <- 1 - pchisq(c, df = parameter)</pre>
 > p.value
 [1] 0.0001480172
 > chisq.test(x, p = prob)$p.value
 [1] 0.0001480172
```

```
> osservati
[1] 54 29 5
> chisq.test(x, p = prob)$observed
[1] 54 29 5
> attesi
[1] 44 22 22
> chisq.test(x, p = prob)$expected
[1] 44 22 22
> residui <- (osservati - attesi)/sqrt(attesi) > residui
[1] 1.507557 1.492405 -3.624412
> chisq.test(x, p = prob)$residuals
[1] 1.507557 1.492405 -3.624412
```

Parte IV Modelli Lineari

Capitolo 13

Regressione lineare semplice

13.1 Simbologia

$$y_i = \beta_1 + \beta_2 \ x_i + \varepsilon_i \quad \forall i = 1, 2, ..., n \qquad \varepsilon \sim N(0, \sigma^2 I_n)$$

- variabile dipendente: y
- matrice del modello di dimensione $n \times 2$: X
- numero di parametri da stimare e rango della matrice del modello: 2
- ullet numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_i) \quad \forall i = 1, 2, ..., n$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^TX)^{-1}X^T$
- matrice identità di dimensione $n \times n$: I_n
- devianza residua: $RSS = \sum_{i=1}^{n} e_i^2 = y^T e = y^T (I_n H) y$
- stima di σ^2 : $s^2 = RSS/(n-2)$
- gradi di libertà della devianza residua: n-2
- stima di σ^2 tolta la i-esima unità: $s_{-i}^2 = s^2 \left(1 + \frac{1 r s t a d a r d_i^2}{n-3}\right) = s^2 \left(1 + \frac{r s t u d e n t_i^2 1}{n-2}\right)^{-1} \quad \forall i = 1, 2, \ldots, n$
- codevianza tra x ed y: $ss_{xy} = \sum_{i=1}^{n} (x_i \bar{x}) (y_i \bar{y})$
- devianza di x: $ss_x = \sum_{i=1}^n (x_i \bar{x})^2$
- devianza di y: $ss_y = \sum_{i=1}^n (y_i \bar{y})^2$
- stime OLS: $\hat{\beta} = (X^T X)^{-1} X^T y$
- stima OLS intercetta: $\hat{\beta}_1 = \bar{y} \bar{x} \, s s_{xy} / s s_x$
- stima OLS coefficiente angolare: $\ \hat{eta}_2 = s s_{xy} \, / \, s s_x$
- standard error delle stime OLS: $s_{\hat{\beta}} = s \sqrt{\operatorname{diag}((X^T X)^{-1})}$
- standard error della stima OLS intercetta: $s_{\hat{\beta}_1} = s \sqrt{\sum_{i=1}^n x_i^2 / (n s s_x)}$
- standard error della stima OLS coefficiente angolare: $s_{\hat{\beta}_2} = s / \sqrt{s s_x}$
- covarianza tra le stime OLS: $s_{\hat{\beta}_1\,\hat{\beta}_2} = -\bar{x}\,s^2/ss_x$
- t-values delle stime OLS: $t_{\hat{eta}} = \hat{eta} \, / \, s_{\hat{eta}}$
- residui: $e = (I_n H) y$
- residui standard: $rstandard_i = \frac{e_i}{s\sqrt{1-h_i}} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = \frac{e_i}{s_{-i}\sqrt{1-h_i}} = rstandard_i\sqrt{\frac{n-3}{n-2-rstandard_i^2}} \quad \forall i=1,2,\ldots,n$
- valori adattati: $\hat{y} = Hy$

- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- stime OLS tolta la i-esima unità: $\hat{\beta}_{(-i)} \quad \forall i=1,2,\ldots,n$
- correlazione tra le stime OLS: $r_{\hat{\beta}_1\,\hat{\beta}_2}=rac{s_{\hat{\beta}_1\,\hat{\beta}_2}}{s_{\hat{\beta}_1}\,s_{\hat{\beta}_2}}$
- devianza residua modello nullo: $RSS_{nullo} = \sum_{i=1}^{n} (y_i \bar{y})^2 = (y \bar{y})^T (y \bar{y})$
- indice di determinazione: $R^2=1-RSS/RSS_{nullo}=1-\left(1-R_{adj}^2\right)\left(n-2\right)/\left(n-1\right)=r_{xy}^2$
- indice di determinazione aggiustato: $R_{adj}^2 = 1 \frac{RSS / (n-2)}{RSS_{nullo} / (n-1)} = 1 \left(1 R^2\right) \left(n 1\right) / \left(n 2\right)$
- valore noto del regressore per la previsione: x_0
- log-verosimiglianza normale: $\hat{\ell} = -n \left(\log(2\pi) + \log(RSS/n) + 1 \right) / 2$
- distanza di Cook: $cd_i=\frac{h_i\,rstandard_i^2}{2\,(1-h_i)}=\frac{e_i^2}{2\,s^2}\,\frac{h_i}{(1-h_i)^2}\quad\forall\,i=1,\,2,\,\ldots,\,n$
- covratio: $cr_i = (1 h_i)^{-1} \left(1 + \frac{rstudent_i^2 1}{n 2}\right)^{-2} = (1 h_i)^{-1} \left(\frac{s_{-i}}{s}\right)^4 \quad \forall i = 1, 2, \dots, n$

13.2 Stima

lm()

- Package: stats
- Input:

formula modello di regressione lineare con una variabile esplicativa ed n unità

- x = TRUE matrice del modello
- y = TRUE variabile dipendente
- Description: analisi di regressione lineare
- Output:

coefficients stime OLS

residuals **residui**

rank rango della matrice del modello

fitted.values valori adattati

df.residual gradi di libertà della devianza residua

- x matrice del modello
- y variabile dipendente
- Formula:

coefficients

$$\hat{\beta}_i \quad \forall j = 1, 2$$

residuals

$$e_i \quad \forall i = 1, 2, \dots, n$$

rank

2

fitted.values

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

df.residual

$$n-2$$

Х

X

У

y

```
> x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x, x = TRUE, y = TRUE)
 > modello$coefficients
 (Intercept)
   3.8486818 0.7492486
 > modello$residuals
           1
                                   3
 -3.17285530 0.82804637 2.37969944 -0.06864749 -1.65699442 1.40387291
  0.55552598 -0.26864749
 > modello$rank
 [1] 2
 > modello$fitted.values
                            3
                                      4
  4.672855 5.571954 7.220301 8.868647 10.516994 6.396127 8.044474 8.868647
 > modello$df.residual
 [1] 6
 > modello$x
  (Intercept) x
            1 1.1
             1 2.3
 3
             1 4.5
             1 6.7
 4
 5
             1 8.9
 6
             1 3.4
 7
             1 5.6
             1 6.7
 attr(, "assign")
 [1] 0 1
 > modello$y
        2 3 4 5 6 7 8
 1.50 6.40 9.60 8.80 8.86 7.80 8.60 8.60
• Note 1: Il modello nullo si ottiene con lm(formula = y ~ 1).
• Note 2: L'istruzione lm(formula = y ~ x) è equivalente a lm(formula = y ~ X - 1).
• Note 3: L'istruzione lm(formula = y \sim x) è equivalente a lm(formula = y \sim 1 + x).
```

summary.lm()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità correlation = TRUE correlazione tra le stime OLS

- Description: analisi di regressione lineare
- Output:

residuals residui coefficients stima puntuale, standard error, t-value, p-value sigma stima di σ r.squared indice di determinazione adj.r.squared indice di determinazione aggiustato fstatistic valore empirico della statistica F, df numeratore, df denominatore cov.unscaled matrice di covarianza delle stime OLS non scalata per σ^2 correlation matrice di correlazione tra le stime OLS

• Formula:

residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 coefficients
$$\hat{\beta}_j \qquad s_{\hat{\beta}_j} \qquad t_{\hat{\beta}_j} \qquad p\text{-value} = 2\,P(t_{n-2} \le -\,|\,t_{\hat{\beta}_j}\,|) \qquad \forall j=1,2$$
 sigma
$$s$$
 r.squared
$$R^2$$
 adj.r.squared
$$R^2_{adj}$$
 fstatistic
$$Fvalue = \frac{RSS_{nullo} - RSS}{RSS/(n-2)} = t_{\hat{\beta}_2}^2 \qquad 1 \qquad n-2$$
 cov.unscaled
$$(X^T\,X)^{-1}$$
 correlation
$$r_{\hat{\beta}_1\,\hat{\beta}_2}$$

```
Estimate Std. Error t value
                                      Pr(>|t|)
(Intercept) 3.8486818 1.5155372 2.539484 0.04411163
          > res$sigma
[1] 1.893745
> res$r.squared
[1] 0.5485788
> res$adj.r.squared
[1] 0.4733419
> res$fstatistic
  value numdf dendf
7.291356 1.000000 6.000000
> res$cov.unscaled
          (Intercept)
(Intercept) 0.6404573 -0.10519536
           -0.1051954 0.02146844
> res$correlation
          (Intercept)
(Intercept) 1.0000000 -0.8971215
          -0.8971215 1.0000000
```

vcov()

• Package: stats

• Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- Description: matrice di covarianza delle stime OLS
- Formula:

$$s^2 (X^T X)^{-1}$$

lm.fit()

- Package: stats
- Input:
 - x matrice del modello
 - y variabile dipendente
- Description: analisi di regressione lineare
- Output:

```
coefficients stime OLS
residuals residui
rank rango della matrice del modello
fitted.values valori adattati
df.residual gradi di libertà della devianza residua
```

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall j=1,2$$
 residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 rank
$$2$$
 fitted.values
$$\hat{y}_i \quad \forall i=1,2,\dots,n$$
 df.residual
$$n-2$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x)
> X <- model.matrix(object = modello)</pre>
> res <- lm.fit(x = X, y)
> res$coefficients
(Intercept)
 3.8486818 0.7492486
> res$residuals
[7] 0.55552598 -0.26864749
> res$rank
[1] 2
> res$fitted.values
[1] 4.672855 5.571954 7.220301 8.868647 10.516994 6.396127 8.044474
[8] 8.868647
> res$df.residual
[1] 6
```

lsfit()

```
• Package: stats
```

• Input:

```
x matrice del modello
y variabile dipendente
intercept = FALSE
```

- Description: analisi di regressione lineare
- Output:

```
coefficients stime OLS residuals residui
```

• Formula:

```
coefficients \hat{\beta}_j \quad \forall \, j \, = \, 1, \, 2 residuals e_i \quad \forall \, i \, = \, 1, \, 2, \, \ldots, \, n
```

• Examples:

confint()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- Description: intervallo di confidenza per le stime OLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-2} s_{\hat{\beta}_i} \quad \forall j = 1, 2$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> confint(object = modello, parm = c(1, 2), level = 0.95)
```

```
2.5 % 97.5 % (Intercept) 0.14029581 7.557068 x 0.07029498 1.428202
```

coef()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** stime OLS
- Formula:

$$\hat{\beta}_j \quad \forall j = 1, 2$$

• Examples:

boxcox()

• Package: MASS

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità lambda parametro di trasformazione λ plotit = FALSE

- **Description:** modello trasformato secondo *Box–Cox*
- Output:
 - \times valore del parametro λ
 - y funzione di verosimiglianza $L(\lambda)$ da minimizzare in λ
- Formula:

y
$$L(\lambda) = -\frac{n}{2} \log \left(RSS_{t_{\lambda}(y)}\right) + (\lambda - 1) \sum_{i=1}^{n} \log(y_i)$$

$$\text{dove} \quad t_{\lambda}(y) = \begin{cases} \frac{y^{\lambda} - 1}{\lambda} & \text{se } \lambda \neq 0 \\ \log(y) & \text{se } \lambda = 0 \end{cases}$$

 $RSS_{t_{\lambda}(y)}$ rappresenta il valore di RSS per il modello che presenta $t_{\lambda}(y)$ come variabile dipendente.

• Example 1:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> res <- boxcox(object = modello, lambda = 1.2, plotit = FALSE)
> res$x

[1] 1.2

**Example 2:

** x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
** y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> res <- boxcox(object = modello, lambda = 4.1, plotit = FALSE)
> res$x
```

[1] 4.1

> res\$y

[1] -11.30996

fitted()

• Package: stats

• Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

predict.lm()

• Package: stats

• Input:

```
object modello di regressione lineare con una variabile esplicativa ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- Description: intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

fit

$$\hat{\beta}_1+\hat{\beta}_2\,x_0 \qquad \hat{\beta}_1+\hat{\beta}_2\,x_0 \mp t_{1-\alpha/2,\,df}\,s^*\,\sqrt{\frac{1}{n}}+\frac{(x_0-\bar{x})^2}{\sum_{i=1}^n{(x_i-\bar{x})^2}}$$

$$\boxed{\text{interval = "prediction"}}$$

$$\hat{\beta}_1+\hat{\beta}_2\,x_0 \qquad \hat{\beta}_1+\hat{\beta}_2\,x_0 \mp t_{1-\alpha/2,\,df}\,s^*\,\sqrt{1+\frac{1}{n}}+\frac{(x_0-\bar{x})^2}{\sum_{i=1}^n{(x_i-\bar{x})^2}}$$
 se.fit
$$s^*\,\sqrt{\frac{1}{n}+\frac{(x_0-\bar{x})^2}{\sum_{i=1}^n{(x_i-\bar{x})^2}}}$$
 df
$$df=n-2$$
 residual.scale

• Example 1:

> c(yhat, lower, upper)

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> x0 <- c(1, 1.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))
> yhat

[1] 4.822705

> new <- data.frame(x = 1.3)
> s <- summary.lm(object = modello) $sigma
> X <- model.matrix(object = modello)
> lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*% + X) %*% x0)
> upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*% + X) %*% x0)</pre>
```

```
[1] 4.822705 2.465776 7.179634
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
      scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
        fit
               lwr
                         upr
 1 4.822705 2.465776 7.179634
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
      x0))
 > se.fit
 [1] 1.202537
 > res$se.fit
 [1] 1.202537
 > s
 [1] 1.893745
 > res$residual.scale
 [1] 1.893745
• Example 2:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x)
 > x0 <- c(1, 1.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 4.822705
 > new <- data.frame(x = 1.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > lower <- yhat - qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% X) %*% x0)
 > upper <- yhat + qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 4.8227050 -0.6664366 10.3118467
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
 + interval = "prediction", level = 0.95)
 > res$fit
                   lwr
 1 4.822705 -0.6664366 10.31185
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
 + x0))
 > se.fit
```

```
[1] 1.202537
> res$se.fit
[1] 1.202537
> s
[1] 1.893745
> res$residual.scale
[1] 1.893745
```

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n 2 e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

predict()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di σ df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza $1-\alpha$

- **Description:** intervallo di confidenza o di previsione
- Output:

fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di σ

• Formula:

fit

$$\hat{\beta}_{1}+\hat{\beta}_{2}\,x_{0} \qquad \hat{\beta}_{1}+\hat{\beta}_{2}\,x_{0}\mp t_{1-\alpha/2,\,df}\,s^{*}\,\sqrt{\frac{1}{n}+\frac{(x_{0}-\bar{x})^{2}}{\sum_{i=1}^{n}\,(x_{i}-\bar{x})^{2}}}$$

$$\boxed{\text{interval = "prediction"}}$$

$$\hat{\beta}_{1}+\hat{\beta}_{2}\,x_{0} \qquad \hat{\beta}_{1}+\hat{\beta}_{2}\,x_{0}\mp t_{1-\alpha/2,\,df}\,s^{*}\,\sqrt{1+\frac{1}{n}+\frac{(x_{0}-\bar{x})^{2}}{\sum_{i=1}^{n}\,(x_{i}-\bar{x})^{2}}}}$$
 se.fit
$$s^{*}\,\sqrt{\frac{1}{n}+\frac{(x_{0}-\bar{x})^{2}}{\sum_{i=1}^{n}\,(x_{i}-\bar{x})^{2}}}$$

```
df
                                           df = n - 2
     residual.scale
                                              s^*
• Example 1:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x)
 > x0 < -c(1, 1.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 4.822705
 > new <- data.frame(x = 1.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
       X) %*% x0)
 > upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
       X) %*% x0)
 > c(yhat, lower, upper)
 [1] 4.822705 2.465776 7.179634
 > res <- predict(object = modello, newdata = new, se.fit = TRUE,
       scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
                 lwr
        fit
                          upr
 1 4.822705 2.465776 7.179634
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
      \times 0)
 > se.fit
 [1] 1.202537
 > res$se.fit
 [1] 1.202537
 > s
 [1] 1.893745
 > res$residual.scale
 [1] 1.893745
• Example 2:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x)
 > x0 < -c(1, 1.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
```

```
[1] 4.822705
> new <- data.frame(x = 1.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)</pre>
> lower <- yhat - qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
     solve(t(X) %*% X) %*% x0)
> upper <- yhat + qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
     solve(t(X) %*% X) %*% x0)
> c(yhat, lower, upper)
[1] 4.8227050 -0.6664366 10.3118467
> res <- predict(object = modello, newdata = new, se.fit = TRUE,
     interval = "prediction", level = 0.95)
> res$fit
                 lwr
       fit.
                           upr
1 4.822705 -0.6664366 10.31185
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
+ x0))
> se.fit
[1] 1.202537
> res$se.fit
[1] 1.202537
> s
[1] 1.893745
> res$residual.scale
[1] 1.893745
```

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n 2 e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

cov2cor()

• Package: stats

• Input:

 \forall matrice di covarianza delle stime OLS di dimensione 2×2

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i, j = 1, 2$$

13.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** log-verosimiglianza normale
- Formula:

 $\hat{\ell}$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> logLik(object = modello)
'log Lik.' -15.30923 (df=3)
```

durbin.watson()

• Package: car

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / RSS$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y \sim x)

> durbin.watson(model = modello)$dw

[1] 1.75205
```

AIC()

- Package: stats
- Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** indice AIC
- Formula:

$$-2\hat{\ell} + 6$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> AIC(object = modello)

[1] 36.61846
```

extractAIC()

- Package: stats
- Input:

fit modello di regressione lineare con una variabile esplicativa ed <math>n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$2 \qquad n \log(RSS/n) + 4$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> extractAIC(fit = modello)
[1] 2.00000 11.91545
```

deviance()

- Package: stats
- Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- Description: devianza residua
- Formula:

RSS

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> deviance(object = modello)

[1] 21.51762
```

PRESS()

• Package: MPV

• Input:

 ${f x}~$ modello di regressione lineare con una variabile esplicativa ed n unità

• **Description:** PRESS

• Formula:

$$\sum_{i=1}^{n} e_i^2 / (1 - h_i)^2$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> PRESS(x = modello)
[1] 53.41271
```

anova()

• Package: stats

• Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- Description: anova di regressione
- Output:

Df gradi di libertà Sum Sq devianze residue Mean Sq quadrati medi F value valore empirico della statistica F Pr(>F) p-value

• Formula:

Df
$$1 n-2$$
 Sum Sq
$$RSS_{nullo}-RSS RSS$$
 Mean Sq
$$RSS_{nullo}-RSS RSS/(n-2)$$
 F value
$$F_{value}=\frac{RSS_{nullo}-RSS}{RSS/(n-2)}=t_{\hat{\beta}_2}^2$$
 Pr(>F)
$$P(F_{1,n-2}\geq F_{value})$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> anova(object = modello)
```

```
Analysis of Variance Table

Response: y

Df Sum Sq Mean Sq F value Pr(>F)

x 1 26.1488 26.1488 7.2914 0.03556 *

Residuals 6 21.5176 3.5863

---

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

drop1()

• Package: stats

• Input:

object modello di regressione lineare con una variabile esplicativa ed n unità scale selezione indice AIC oppure Cp test = "F"

• **Description:** submodels

• Output:

Df differenza tra gradi di libertà Sum of Sq differenza tra devianze residue RSS devianza residua AIC indice AIC Cp indice Cp F value valore empirico della statistica F Pr(F) p-value

• Formula:

Df $RSS_{nullo} - RSS$ RSS, RSS_{nullo} AIC $\boxed{\text{scale = 0}}$ $n \log (RSS/n) + 4, n \log (RSS_{nullo}/n) + 2$ Cp $\boxed{\text{scale = s}^2}$ $2, \frac{RSS_{nullo}}{RSS/(n-2)} + 2 - n$ F value $F_{value} = \frac{RSS_{nullo} - RSS}{RSS/(n-2)} = t_{\hat{\beta}_2}^2$ Pr(F)

 $P(F_{1, n-2} \ge F_{value})$

• Example 1:

Pr(F) p-value

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
   > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
   > n < - 8
   > modello <- lm(formula = y \sim x)
   > drop1(object = modello, scale = 0, test = "F")
   Single term deletions
   Model:
   y ~ x
          Df Sum of Sq RSS AIC F value Pr(F)
                21.518 11.915
   <none>
                26.149 47.666 16.278 7.2914 0.03556 *
   Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
  • Example 2:
   > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
   > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
   > modello <- lm(formula = y \sim x)
   > s <- summary.lm(object = modello)$sigma</pre>
   > drop1(object = modello, scale = s^2, test = "F")
   Single term deletions
   Model:
   у ~ х
   scale: 3.586271
          Df Sum of Sq RSS Cp F value Pr(F)
                        21.518 2.0000
                 26.149 47.666 7.2914 7.2914 0.03556 *
   X
   Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
add1()
  • Package: stats
  • Input:
       object modello nullo di regressione lineare semplice
       scope \,modello di regressione lineare con una variabile esplicativa ed n unità
       scale selezione indice AIC oppure Cp
       test = "F"
  • Description: submodels
  • Output:
       Df differenza tra gradi di libertà
       Sum of Sq differenza tra devianze residue
       RSS devianza residua
       AIC indice AIC
       Cp indice Cp
       {\mathbb F} value valore empirico della statistica F
```

• Formula:

```
Df
                                                   1
      Sum of Sq
                                             RSS_{nullo} - RSS
      RSS
                                              RSS_{nullo}, RSS
      AIC
                                              scale = 0
                                 n \log (RSS_{nullo}/n) + 2, n \log (RSS/n) + 4
      Ср
                                              scale = s^2
                                          \frac{RSS_{nullo}}{RSS \, / \, (n-2)} + 2 - n, \, 2
      F value
                                     F_{value} = \frac{RSS_{nullo} - RSS}{RSS / (n-2)} = t_{\hat{\beta}_2}^2
      Pr(F)
                                            P(F_{1,n-2} \geq F_{value})
• Example 1:
 > x \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x)
 > nullo <- lm(formula = y \sim 1)
 > add1(object = nullo, scope = modello, scale = 0, test = "F")
 Single term additions
 Model:
         Df Sum of Sq RSS AIC F value Pr(F)
                         47.666 16.278
 <none>
               26.149 21.518 11.915 7.2914 0.03556 *
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
• Example 2:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x)
 > nullo <- lm(formula = y \sim 1)
 > s <- summary.lm(object = modello)$sigma</pre>
 > add1(object = nullo, scope = modello, scale = s^2, test = "F")
 Single term additions
 Model:
 y ~ 1
 scale: 3.586271
```

Df Sum of Sq

RSS

Cp F value Pr(F)

13.4 Diagnostica

ls.diag()

• Package: stats

• Input:

ls.out modello di regressione lineare con una variabile eplicativa ed n unità

• Description: analisi di regressione lineare

• Output:

```
std.dev stima di \sigma hat valori di leva std.res residui standard stud.res residui studentizzati cooks distanza di Cook dfits dfits correlation matrice di correlazione tra le stime OLS std.err standard error delle stime OLS cov.scaled matrice di covarianza delle stime OLS non scalata per \sigma^2
```

• Formula:

std.dev
$$s$$
 hat
$$h_i \quad \forall i=1,2,\dots,n$$
 std.res
$$rstandard_i \quad \forall i=1,2,\dots,n$$
 stud.res
$$rstudent_i \quad \forall i=1,2,\dots,n$$
 cooks
$$cd_i \quad \forall i=1,2,\dots,n$$
 dfits
$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i=1,2,\dots,n$$
 correlation
$$r_{\beta_1\beta_2}$$
 std.err
$$s_{\beta_j} \quad \forall j=1,2$$
 cov.scaled
$$s^2 (X^TX)^{-1}$$
 cov.unscaled
$$(X^TX)^{-1}$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x)
> res <- ls.diag(ls.out = modello)</pre>
> res$std.dev
[1] 1.893745
> res$hat
[1] 0.4350043 0.2701267 0.1284350 0.1945578 0.4684951 0.1733040 0.1355195
[8] 0.1945578
> res$std.res
 \begin{smallmatrix} [1] & -2.22897996 & 0.51181072 & 1.34601741 & -0.04039112 & -1.20017856 & 0.81532985 \end{smallmatrix} 
[7] 0.31550428 -0.15806803
> res$stud.res
[1] -4.90710471 0.47776268 1.47068630 -0.03687690 -1.25680777 0.78929887
[7] 0.29043398 -0.14459710
> res$cooks
[1] 1.9126289653 0.0484739848 0.1334918569 0.0001970407 0.6348329327
[6] 0.0696786009 0.0078023824 0.0030176734
> res$dfits
 \begin{smallmatrix} 1 \end{smallmatrix} \end{smallmatrix} -4.30575707 \quad 0.29065126 \quad 0.56456215 \quad -0.01812431 \quad -1.17996116 \quad 0.36138726 
[7] 0.11499284 -0.07106678
> res$correlation
              (Intercept)
              1.0000000 -0.8971215
(Intercept)
              -0.8971215 1.0000000
> res$std.err
(Intercept) 1.5155372
             0.2774737
> res$cov.scaled
              (Intercept)
              2.2968531 -0.37725904
(Intercept)
               -0.3772590 0.07699164
Х
> res$cov.unscaled
             (Intercept)
              0.6404573 -0.10519536
(Intercept)
Х
              -0.1051954 0.02146844
```

cooks.distance()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

rstandard()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

• Examples:

rstandard.lm()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

• Examples:

rstudent()

- Package: stats
- Input:

model modello di regressione lineare con una variabile esplicativa ed <math>n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, ..., n$$

• Examples:

rstudent.lm()

- Package: stats
- Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \dots, n$$

lmwork()

• Package: MASS

• Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

• **Description:** diagnostica di regressione

• Output:

```
stdedv stima di \sigma stdres residui standard studres residui studentizzati
```

• Formula:

```
stdedv s stdres rstandard_i \  \  \, \forall i=1,2,\ldots,n studres rstudent_i \  \  \, \forall i=1,2,\ldots,n
```

• Examples:

dffits()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- **Description**: dffits
- Formula:

$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i = 1, 2, \dots, n$$

• Examples:

covratio()

- Package: stats
- Input:

modello di regressione lineare con una variabile esplicativa ed <math>n unità

- **Description:** covratio
- Formula:

$$cr_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

lm.influence()

- Package: stats
- Input:

 ${\tt modell}$ modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** diagnostica di regressione
- Output:

```
hat valori di leva coefficients differenza tra le stime OLS eliminando una unità sigma stima di \sigma eliminando una unità wt.res residui
```

• Formula:

hat $h_i \quad \forall \, i=1,\,2,\,\ldots,\,n$ coefficients $\hat{\beta}_j - \hat{\beta}_{j\,(-i)} \,=\, e_i\,(1-h_i)^{-1}\,(X^T\,X)_j^{-1}\,X_i^T \quad \forall i=1,\,2,\,\ldots,\,n \quad \forall j=1,\,2$

```
sigma s_{-i} \quad \forall i = 1, 2, \ldots, n wt.res e_i \quad \forall i = 1, 2, \ldots, n
```

• Examples:

```
> x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x)
> res <- lm.influence(model = modello)</pre>
> res$hat
                                  4
0.4350043 \ 0.2701267 \ 0.1284350 \ 0.1945578 \ 0.4684951 \ 0.1733040 \ 0.1355195 \ 0.1945578
> res$coefficients
   (Intercept)
1 -2.946804056 0.458130527
2 0.452110031 -0.063325849
3 0.456185994 -0.023446758
4 0.005484663 -0.003293542
5 0.922114131 -0.267715952
6 0.480231536 -0.054685694
7 0.033006665 0.009657123
8 0.021463873 -0.012889065
> res$sigma
                            3
                                       4
0.8602058\ 2.0287040\ 1.7332139\ 2.0742118\ 1.8084168\ 1.9562006\ 2.0572134\ 2.0701700
> res$wt.res
          1
                      2
                                   3
                                                4
-3.17285530 \quad 0.82804637 \quad 2.37969944 \quad -0.06864749 \quad -1.65699442 \quad 1.40387291
         7
 0.55552598 - 0.26864749
```

residuals.lm()

- Package: stats
- Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- **Description:** residui
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> residuals.lm(object = modello)
```

```
1 2 3 4 5 6

-3.17285530 0.82804637 2.37969944 -0.06864749 -1.65699442 1.40387291

7 8

0.555552598 -0.26864749
```

df.residual()

• Package: stats

• Input:

object $\,$ modello di regressione lineare con una variabile esplicativa ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-2

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x)
> df.residual(object = modello)
[1] 6
```

hatvalues()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

dfbeta()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

• **Description:** dfbeta

• Formula:

$$\hat{\beta}_j - \hat{\beta}_{j(-i)} = e_i (1 - h_i)^{-1} (X^T X)_i^{-1} X_i^T \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2$$

• Examples:

dfbetas()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- **Description**: dfbetas
- Formula:

$$\frac{\hat{\beta}_{j} - \hat{\beta}_{j (-i)}}{s_{\hat{\beta}_{j} - \hat{\beta}_{j (-i)}}} = \frac{e_{i} (1 - h_{i})^{-1} (X^{T} X)_{j}^{-1} X_{i}^{T}}{s_{-i} \sqrt{(X^{T} X)_{j, j}^{-1}}} \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2$$

outlier.test()

- Package:
- Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: test sugli outliers
- Output:

test massimo residuo studentizzato assoluto, gradi di libertà, p-value

• Formula:

```
test t = \max_i(\,|\,rstudent_i\,|) \quad n-3 \quad p\text{-value} = 2\,P(\,t_{n-3} \leq -|\,t\,|) \qquad \forall\, i\,=\,1,\,2,\,\ldots\,,n
```

• Examples:

influence.measures()

• Package: stats

• Input:

model modello di regressione lineare con una variabile esplicativa ed n unità

- Description: dfbetas, dffits, covratio, distanza di Cook, valori di leva
- Output:

```
infmat misure di influenza di dimensione n\times 6 is .inf matrice di influenza con valori logici di dimensione n\times 6
```

• Formula:

infmat

$$DFBETAS_{ij} = \frac{e_{i} (1-h_{i})^{-1} (X^{T} X)_{j}^{-1} X_{i}^{T}}{s_{-i} \sqrt{(X^{T} X)_{j,j}^{-1}}} \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2$$

$$DFFITS_{i} = rstudent_{i} \sqrt{\frac{h_{i}}{1-h_{i}}} \quad \forall i = 1, 2, ..., n$$

$$COVRATIO_{i} = (1-h_{i})^{-1} \left(1 + \frac{rstudent_{i}^{2}-1}{n-2}\right)^{-2} \quad \forall i = 1, 2, ..., n$$

$$COOKD_{i} = \frac{h_{i} rstandard_{i}^{2}}{2(1-h_{i})} \quad \forall i = 1, 2, ..., n$$

$$HAT_{i} = h_{i} \quad \forall i = 1, 2, ..., n$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8</pre>
```

```
> modello <- lm(formula = y \sim x)
 > res <- influence.measures(model = modello)</pre>
 > res
 Influence measures of
           lm(formula = y \sim x):
     dfb.1_
               dfb.x
                      dffit cov.r
                                        cook.d
                                                  hat inf
 1 -4.28059 3.6349 -4.3058 0.0753 1.912629 0.435
 2 0.27847 -0.2130 0.2907 1.8044 0.048474 0.270
 3 0.32889 -0.0923 0.5646 0.8050 0.133492 0.128
 4 0.00330 -0.0108 -0.0181 1.7869 0.000197 0.195
    0.63715 -1.0104 -1.1800 1.5646 0.634833 0.468
 6 0.30676 -0.1908
                      0.3614 1.3773 0.069679 0.173
    0.02005 0.0320 0.1150 1.6109 0.007802 0.136
    0.01296 -0.0425 -0.0711 1.7730 0.003018 0.195
 > res$infmat
          dfb.1_
                        dfb.x
                                     dffit
                                                 cov.r
                                                              cook.d
 1 -4.280591734
                  3.63485094 -4.30575707 0.07534912 1.9126289653 0.4350043
 2\quad 0.278471258 \ -0.21304046 \quad 0.29065126 \ 1.80443448 \ 0.0484739848 \ 0.2701267
 3\quad 0.328885485\ -0.09232735\quad 0.56456215\ 0.80504974\ 0.1334918569\ 0.1284350
 4 \quad 0.003304089 \quad -0.01083702 \quad -0.01812431 \quad 1.78686556 \quad 0.0001970407 \quad 0.1945578
 5 \quad 0.637149075 \quad -1.01035839 \quad -1.17996116 \quad 1.56459066 \quad 0.6348329327 \quad 0.4684951
 0.020048284 0.03203820 0.11499284 1.61092794 0.0078023824 0.1355195
    0.012955584 -0.04249278 -0.07106678 1.77297867 0.0030176734 0.1945578
 > res$is.inf
   dfb.1_ dfb.x dffit cov.r cook.d
     TRUE TRUE TRUE FALSE
                               TRUE FALSE
   FALSE FALSE FALSE FALSE FALSE
 3 FALSE FALSE FALSE FALSE FALSE
 4 FALSE FALSE FALSE FALSE FALSE
 5 FALSE TRUE FALSE FALSE FALSE
 6 FALSE FALSE FALSE FALSE FALSE
    FALSE FALSE FALSE FALSE FALSE
    FALSE FALSE FALSE FALSE FALSE
• Note 1: Il caso i-esimo è influente se |DFBETAS_{ij}| > 1 \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2
• Note 2: Il caso i-esimo è influente se |DFFITS_i| > 3\sqrt{2/(n-2)} \forall i = 1, 2, ..., n
• Note 3: Il caso i-esimo è influente se |1 - COVRATIO_i| > 6 / (n-2) \forall i = 1, 2, ..., n
• Note 4: Il caso i-esimo è influente se P(F_{2,n-2} \geq COOKD_i) > 0.5 \quad \forall i = 1, 2, ..., n
• Note 5: Il caso i-esimo è influente se HAT_i > 6/n \quad \forall i = 1, 2, ..., n
• Note 6: I casi influenti rispetto ad almeno una tra queste misure sono marcati con un asterisco.
```

Corrispondentemente la stessa riga della matrice is.inf riporterà almeno un simbolo TRUE.

Capitolo 14

Regressione lineare multipla

14.1 Simbologia

$$y_i = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} + \varepsilon_i \quad \forall i = 1, 2, \dots, n \qquad \varepsilon \sim N(0, \sigma^2 I_n)$$

- variabile dipendente: y
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^TX)^{-1}X^T$
- matrice identità di dimensione $n \times n$: I_n
- devianza residua: $RSS = \sum_{i=1}^{n} e_i^2 = y^T e = y^T (I_n H) y$
- stima di σ^2 : $s^2 = RSS/(n-k)$
- gradi di libertà della devianza residua: n-k
- stima di σ^2 tolta la i-esima unità: $s_{-i}^2 = s^2 \left(1 + \frac{1 rstandard_i^2}{n k 1}\right) = s^2 \left(1 + \frac{rstudent_i^2 1}{n k}\right)^{-1} \quad \forall i = 1, 2, \ldots, n$
- stime OLS: $\hat{\beta} = (X^T X)^{-1} X^T y$
- standard error delle stime OLS: $s_{\hat{\beta}} = s \sqrt{\mathrm{diag}((X^TX)^{-1})}$
- t-values delle stime OLS: $t_{\hat{eta}} = \hat{eta} / s_{\hat{eta}}$
- residui: $e = (I_n H) y$
- residui standard: $rstandard_i = \frac{e_i}{s\sqrt{1-h_i}} \quad \forall i=1,2,\ldots,n$
- residui studentizzati: $rstudent_i = \frac{e_i}{s_{-i}\sqrt{1-h_i}} = rstandard_i\sqrt{\frac{n-k-1}{n-k-rstandard_i^2}} \quad \forall i=1,2,\ldots,n$
- valori adattati: $\hat{y} = Hy$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \dots, n$
- stime OLS tolta la *i*-esima unità: $\hat{\beta}_{(-i)} \quad \forall i = 1, 2, ..., n$
- correlazione tra le stime OLS: $r_{\hat{\beta}_i\,\hat{\beta}_j}=rac{s^2\,(X^T\,X)_{i,\,j}^{-1}}{s_{\hat{\beta}_i}\,s_{\hat{\beta}_j}}\quad \forall\,i,j\,=\,1,\,2,\,\ldots,\,k$
- devianza residua modello nullo: $RSS_{nullo} = \sum_{i=1}^n (y_i \bar{y})^2 = (y \bar{y})^T (y \bar{y})$
- indice di determinazione: $R^2=1-RSS/RSS_{nullo}=1-\left(1-R_{adj}^2\right)\left(n-k\right)/\left(n-1\right)$
- indice di determinazione aggiustato: $R_{adj}^2 = 1 \frac{RSS/(n-k)}{RSS_{nullo}/(n-1)} = 1 \left(1 R^2\right) \left(n 1\right)/\left(n k\right)$
- valore noto dei regressori per la previsione: $x_0^T = (1, x_{01}, x_{02}, \dots, x_{0k-1})$
- log-verosimiglianza normale: $\hat{\ell} = -n (\log(2\pi) + \log(RSS/n) + 1)/2$

- distanza di *Cook*: $cd_i=rac{h_i\,rstandard_i^2}{k\,(1-h_i)}=rac{e_i^2}{k\,s^2}\,rac{h_i}{(1-h_i)^2}\quad \forall\,i\,=\,1,\,2,\,\ldots,\,n$
- covratio: $cr_i = (1 h_i)^{-1} \left(1 + \frac{rstudent_i^2 1}{n k} \right)^{-k} = (1 h_i)^{-1} \left(\frac{s_{-i}}{s} \right)^{2k} \quad \forall i = 1, 2, \dots, n$

14.2 Stima

lm()

• Package: stats

• Input:

formula modello di regressione lineare con k-1 variabili esplicative ed n unità $\mathbf{x} = \mathtt{TRUE}$ matrice del modello

y = TRUE variabile dipendente

• Description: analisi di regressione lineare

• Output:

coefficients stime OLS
residuals residui
rank rango della matrice del modello
fitted.values valori adattati
df.residual gradi di libertà della devianza residua
x matrice del modello
y variabile dipendente

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall j=1,2,\dots,k$$
 residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 rank
$$k$$
 fitted.values
$$\hat{y}_i \quad \forall i=1,2,\dots,n$$
 df.residual
$$n-k$$
 X

• Examples:

> modello\$residuals

```
0.4358424 \quad 1.3067117 \quad 0.6974820 \quad 0.2575634 \quad 0.6607787 \quad -0.9691173
-0.9536382
-1.4356227
> modello$rank
[1] 4
> modello$fitted.values
 2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
> modello$df.residual
[1] 4
> modello$x
  (Intercept) x1 x2 x3
            1 1.1 1.2 1.40
2
            1 2.3 3.4 5.60
            1 4.5 5.6 7.56
            1 6.7 7.5 6.00
            1 8.9 7.5 5.40
5
            1 3.4 6.7 6.60
            1 5.6 8.6 8.70
            1 6.7 7.6 8.70
attr(,"assign")
[1] 0 1 2 3
> modello$v
                                 7
            3
                 4
                      5
                            6
1.50 6.40 9.60 8.80 8.86 7.80 8.60 8.60
```

- **Note 1:** Il modello nullo si ottiene con lm (formula = y ~ 1).
- Note 2: L'istruzione update (object = $y \sim x1 + x2$, formula = . ~ . + x3) è esattamente equivalente a lm (formula = $y \sim x1 + x2 + x3$).
- Note 3: In seguito ad una modifica come ad esempio x1[3] < -1.2, conviene adoperare il comando update (modello) anziché ripetere modello $< -lm(formula = y \sim x1 + x2 + x3)$.
- Note 4: L'operatore I () permette di poter modellare regressioni lineari polinomiali. Per un polinomio di terzo grado occorre scrivere $lm(formula = y \sim x + I(x^2) + I(x^3))$.
- **Note 5:** Per regressioni polinomiali occorre usare il comando poly(). Per un polinomio di quarto grado occorre scrivere lm(formula = y ~ poly(x, degree = 4, raw = TRUE)).
- **Note 6:** Per regressioni polinomiali ortogonali occorre usare il comando poly(). Per un polinomio ortogonale di quarto grado occorre scrivere lm(formula = y ~ poly(x,degree = 4)).
- Note 7: Il comando lm (formula = y ~ x1 + x2) è equivalente a lm (formula = y ~ X-1).
- Note 8: Il comando $lm (formula = y \sim x1 + x2)$ è equivalente a $lm (formula = y \sim 1 + x1 + x2)$.

summary.lm()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità correlation = TRUE correlazione tra le stime OLS

- Description: analisi di regressione lineare
- Output:

residuals residui coefficients stima puntuale, standard error, t-value, p-value sigma stima di σ r.squared indice di determinazione adj.r.squared indice di determinazione aggiustato fstatistic valore empirico della statistica F, df numeratore, df denominatore cov.unscaled matrice di covarianza delle stime OLS non scalata per σ^2 correlation matrice di correlazione tra le stime OLS

• Formula:

residuals
$$e_i \quad \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j \quad s_{\hat{\beta}_j} \quad t_{\hat{\beta}_j} \quad p\text{-value} = 2\,P(t_{n-k} \le -\,|\,t_{\hat{\beta}_j}\,|) \quad \forall j=1,2,\ldots,k$$
 sigma
$$s$$
 r.squared
$$R^2$$
 adj.r.squared
$$R^2_{adj}$$
 fstatistic
$$Fvalue = \frac{(RSS_{nullo} - RSS) \ / \ (k-1)}{RSS \ / \ (n-k)} \qquad k-1 \qquad n-k$$
 cov.unscaled
$$(X^T \, X)^{-1}$$
 correlation
$$r_{\hat{\beta}_i,\hat{\beta}_i} \quad \forall i,j=1,2,\ldots,k$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> res <- summary.lm(object = modello, correlation = TRUE)
> res$residuals
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

> res\$coefficients

```
Estimate Std. Error t value Pr(>|t|) (Intercept) 0.988514333 1.4292308 0.691640822 0.5272118 x1 0.422516384 0.3883267 1.088043731 0.3377443 x2 -0.001737381 0.5822146 -0.002984091 0.9977619 x3 0.716029046 0.4068987 1.759723294 0.1532663
```

- > res\$sigma
- [1] 1.303508
- > res\$r.squared
- [1] 0.8574147
- > res\$adj.r.squared
- [1] 0.7504757
- > res\$fstatistic

```
value numdf dendf
8.017793 3.000000 4.000000
```

> res\$cov.unscaled

```
    (Intercept)
    x1
    x2
    x3

    (Intercept)
    1.20220217 -0.06075872 0.0350553 -0.15856757

    x1
    -0.06075872 0.08874976 -0.1093953 0.04541621

    x2
    0.03505530 -0.10939532 0.1994982 -0.11184964

    x3
    -0.15856757 0.04541621 -0.1118496 0.09744180
```

> res\$correlation

```
(Intercept) x1 x2 x3

(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900

x1 -0.18600997 1.0000000 -0.82213982 0.4883764

x2 0.07158062 -0.8221398 1.00000000 -0.8022181

x3 -0.46329002 0.4883764 -0.80221810 1.0000000
```

vcov()

- Package: stats
- Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime OLS
- Formula:

$$s^2 (X^T X)^{-1}$$

```
> k <- 4
> x1 < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> vcov(object = modello)
            (Intercept)
                                 x1
                                              x2
(Intercept) 2.04270054 -0.10323710 0.05956359 -0.26942727
            -0.10323710 0.15079759 -0.18587712 0.07716815
x1
x2
            0.05956359 -0.18587712 0.33897378 -0.19004733
            -0.26942727 0.07716815 -0.19004733 0.16556652
x3
```

lm.fit()

• Package: stats

- Input:
 - x matrice del modello
 - y variabile dipendente
- Description: analisi di regressione lineare
- Output:

coefficients stime OLS
residuals residui
rank rango della matrice del modello
fitted.values valori adattati
df.residual gradi di libertà della devianza residua

• Formula:

coefficients
$$\hat{\beta_j} \quad \forall j=1,2,\dots,k$$
 residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 rank
$$k$$
 fitted.values
$$\hat{y_i} \quad \forall i=1,2,\dots,n$$
 df.residual
$$n-k$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> X <- model.matrix(object = modello)
> res <- lm.fit(x = X, y)
> res$coefficients
```

```
(Intercept)
                                x1
                                               x2
     > res$residuals
     \begin{smallmatrix} 1 \end{smallmatrix} \end{bmatrix} - 0.9536382 \quad 0.4358424 \quad 1.3067117 \quad 0.6974820 \quad 0.2575634 \quad 0.6607787 \quad -0.9691173 
    [8] -1.4356227
    > res$rank
    [1] 4
    > res$fitted.values
    [1] 2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117
    [8] 10.035623
    > res$df.residual
    [1] 4
lsfit()
  • Package: stats
  • Input:
        x matrice del modello
        y variabile dipendente
        intercept = FALSE
  • Description: analisi di regressione lineare
  • Output:
        coefficients stime OLS
        residuals residui
  • Formula:
        coefficients
                                               \hat{\beta}_i \quad \forall j = 1, 2, \ldots, k
        residuals
                                               e_i \quad \forall i = 1, 2, \ldots, n
  • Examples:
    > k <- 4
    > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
    > x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
    > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
    > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
    > modello <- lm(formula = y \sim x1 + x2 + x3)
    > X <- model.matrix(object = modello)</pre>
    > res <- lsfit(x = X, y, intercept = FALSE)</pre>
    > res$coefficients
     (Intercept)
                                x1
     0.988514333 0.422516384 -0.001737381 0.716029046
    > res$residuals
     \begin{smallmatrix} 1 \end{smallmatrix} \rbrack -0.9536382 \quad 0.4358424 \quad 1.3067117 \quad 0.6974820 \quad 0.2575634 \quad 0.6607787 \quad -0.9691173 
    [8] -1.4356227
```

confint()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- Description: intervallo di confidenza per le stime OLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-k} s_{\hat{\beta}_i} \quad \forall j = 1, 2, ..., k$$

• Example 1:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> confint (object = modello, parm = c(1, 2, 3, 4), level = 0.95)
                 2.5 % 97.5 %
(Intercept) -2.9796664 4.956695
            -0.6556513 1.500684
x1
x2
            -1.6182241 1.614749
x3
            -0.4137027 1.845761
```

• Example 2:

Confint()

• Package: Rcmdr

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- Description: intervallo di confidenza per le stime OLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-k} s_{\hat{\beta}_j} \quad \forall j = 1, 2, \dots, k$$

• Example 2:

coef()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** stime OLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** stime OLS
- Formula:

$$\hat{\beta}_i \quad \forall \ i = 1, 2, \dots, k$$

• Examples:

coeftest()

- Package: lmtest
- Input:

 ${\tt x}$ modello di regressione lineare con k-1 variabili esplicative ed n unità ${\tt df}={\tt NULL}$ / ${\tt Inf}$ significatività delle stime effettuata con la variabile casuale t oppure Z

- **Description:** stima puntuale, standard error, t-value, p-value
- Formula:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3)
> coeftest(x = modello, df = NULL)
t test of coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.9885143 1.4292308 0.6916 0.5272
x1
            0.4225164 0.3883267 1.0880
                                          0.3377
            -0.0017374 0.5822146 -0.0030
                                           0.9978
x2
             0.7160290 0.4068987 1.7597
                                          0.1533
x3
```

• Example 2:

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> coeftest(x = modello, df = Inf)
z test of coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.9885143 1.4292308 0.6916 0.48916
             0.4225164 0.3883267 1.0880 0.27658
            -0.0017374 0.5822146 -0.0030 0.99762
x2
x3
             0.7160290 0.4068987 1.7597 0.07845
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

• Note: Naturalmente vale che $t_{\hat{\beta}_i} = z_{\hat{\beta}_i} \quad \forall j = 1, 2, ..., k$.

boxcox()

• Package: MASS

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità lambda parametro di trasformazione λ plotit = FALSE

- Description: modello trasformato secondo Box-Cox
- Output:
 - \times valore del parametro λ
 - y funzione di verosimiglianza $L(\lambda)$ da minimizzare in λ
- Formula:

y
$$L(\lambda) = -\frac{n}{2}\log\left(RSS_{t_{\lambda}(y)}\right) + (\lambda - 1)\sum_{i=1}^{n}\log(y_{i})$$

$$\text{dove} \quad t_{\lambda}(y) = \begin{cases} \frac{y^{\lambda} - 1}{\lambda} & \text{se } \lambda \neq 0 \\ \log(y) & \text{se } \lambda = 0 \end{cases}$$

 $RSS_{t_{\lambda}(y)}$ rappresenta il valore di RSS per il modello che presenta $t_{\lambda}(y)$ come variabile dipendente.

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> res <- boxcox(object = modello, lambda = 1.2, plotit = FALSE)

> res$x
```

```
[1] 1.2

> res$y

[1] -7.185995

• Example 2:

> k <- 4

> x1 <- c(1.1)
```

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> res <- boxcox(object = modello, lambda = 4.1, plotit = FALSE)
> res$x

[1] 4.1
> res$y
```

box.cox()

- Package: car
- Input:
 - y vettore numerico positivo di dimensione n
 - p parametro di trasformazione λ
- **Description:** variabile y trasformata secondo Box-Cox
- Formula:

$$t_{\lambda}(y) = \left\{ egin{array}{ll} rac{y^{\lambda}-1}{\lambda} & ext{se } \lambda
eq 0 \\ \log(y) & ext{se } \lambda = 0 \end{array}
ight.$$

• Example 1:

```
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> box.cox(y, p = 0.5)

[1] 0.4494897 3.0596443 4.1967734 3.9329588 3.9531504 3.5856960 3.8651513
[8] 3.8651513
```

```
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> box.cox(y, p = 2)

[1] 0.6250 19.9800 45.5800 38.2200 38.7498 29.9200 36.4800 36.4800
```

box.cox.var()

• Package: car

• Input:

y vettore numerico positivo di dimensione n

- **Description:** variabile y trasformata secondo *Box–Cox*
- Formula:

$$y_i (\log (y_i / \bar{y}_G) - 1) \quad \forall i = 1, 2, ..., n$$

dove
$$\bar{y}_G = \left(\prod_{i=1}^n y_i\right)^{1/n} = \exp\left(\frac{1}{n}\sum_{i=1}^n \log(y_i)\right)$$

• Examples:

```
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> box.cox.var(y)

[1] -3.748828 -6.709671 -6.172042 -6.423405 -6.406997 -6.634371 -6.475128
[8] -6.475128
```

bc()

- Package: car
- Input:
 - y vettore numerico positivo di dimensione n
 - p parametro di trasformazione λ
- **Description:** variabile y trasformata secondo *Box–Cox*
- Formula:

$$t_{\lambda}(y) = \begin{cases} rac{y^{\lambda} - 1}{\lambda} & \text{se } \lambda \neq 0 \\ \log(y) & \text{se } \lambda = 0 \end{cases}$$

• Example 1:

```
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> bc(y, p = 0.5)

[1] 0.4494897 3.0596443 4.1967734 3.9329588 3.9531504 3.5856960 3.8651513
[8] 3.8651513
```

```
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> bc(y, p = 2)

[1] 0.6250 19.9800 45.5800 38.2200 38.7498 29.9200 36.4800 36.4800
```

fitted()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

• Description: valori adattati

• Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> fitted(object = modello)
1 2 3 4 5 6 7 8
2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
```

fitted.values()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> fitted.values(object = modello)
1 2 3 4 5 6 7 8
2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
```

predict.lm()

• Package: stats

• Input:

```
object modello di regressione lineare con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- Description: intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

fit $x_0^T \hat{\beta} = \text{"confidence"}$ $x_0^T \hat{\beta} = x_0^T \hat{\beta} \mp t_{1-\alpha/2,df} s^* \sqrt{x_0^T (X^T X)^{-1} x_0}$ interval = "prediction" $x_0^T \hat{\beta} = x_0^T \hat{\beta} \mp t_{1-\alpha/2,df} s^* \sqrt{1 + x_0^T (X^T X)^{-1} x_0}$ se.fit $s^* \sqrt{x_0^T (X^T X)^{-1} x_0}$ df = n - k residual.scale s^*

```
> k < -4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> x0 \leftarrow c(1, 1.3, 2.1, 2.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 3.181004
> new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)</pre>
> lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
      X) %*% x0)
> upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%)
      X) %*% x0)
> c(yhat, lower, upper)
```

```
[1] 3.181004 1.200204 5.161803
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
       scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
               lwr
                         upr
 1 3.181004 1.200204 5.161803
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
 + x0))
 > se.fit
 [1] 1.010631
 > res$se.fit
 [1] 1.010631
 > s
 [1] 1.303508
 > res$residual.scale
 [1] 1.303508
• Example 2:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > x0 \leftarrow c(1, 1.3, 2.1, 2.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 3.181004
 > new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > lower <- yhat - qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
        solve(t(X) %*% X) %*% x0)
 > upper <- yhat + qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 3.181004 -1.398453 7.760461
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
      interval = "prediction", level = 0.95)
 > res$fit
                   lwr
                            upr
 1 3.181004 -1.398453 7.760461
```

```
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
+ x0))
> se.fit

[1] 1.010631
> res$se.fit

[1] 1.010631
> s

[1] 1.303508
> res$residual.scale

[1] 1.303508
```

- **Note 1:** Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n k e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

predict()

- Package: stats
- Input:

```
object modello di regressione lineare con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- Description: intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

fit

```
se.fit s^* \, \sqrt{x_0^T \, (X^T \, X)^{-1} \, x_0} df df \, = n - k residual.scale s^*
```

• Example 1:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> x0 \leftarrow c(1, 1.3, 2.1, 2.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 3.181004
> new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)</pre>
> lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
      X) %*% x0)
> upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
     X) %*% x0)
> c(yhat, lower, upper)
[1] 3.181004 1.200204 5.161803
> res <- predict(object = modello, newdata = new, se.fit = TRUE,
     scale = s, df = Inf, interval = "confidence", level = 0.95)
> res$fit
               lwr
1 3.181004 1.200204 5.161803
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
+ x0))
> se.fit
[1] 1.010631
> res$se.fit
[1] 1.010631
> s
[1] 1.303508
> res$residual.scale
[1] 1.303508
```

```
> k < - 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> x0 <- c(1, 1.3, 2.1, 2.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 3.181004
> new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)</pre>
> lower <- yhat - qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
      solve(t(X) %*% X) %*% x0)
> upper <- yhat + qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
     solve(t(X) %*% X) %*% x0)
> c(yhat, lower, upper)
[1] 3.181004 -1.398453 7.760461
> res <- predict(object = modello, newdata = new, se.fit = TRUE,
      interval = "prediction", level = 0.95)
> res$fit
       fit
              lwr
                         upr
1 3.181004 -1.398453 7.760461
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% X) %*%
+ x0))
> se.fit
[1] 1.010631
> res$se.fit
[1] 1.010631
> s
[1] 1.303508
> res$residual.scale
[1] 1.303508
```

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n k e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

linear.hypothesis()

• Package: car

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità hypothesis.matrix matrice C di dimensione $q \times k$ e rango pari a $q = \min(q, k)$ rhs vettore b della previsione lineare di dimensione q

• **Description:** test di ipotesi per H_0 : $C\beta = b$ contro H_1 : $C\beta \neq b$ dove C e b sono così definiti:

$$C = \begin{pmatrix} c_{1,1} & c_{1,2} & \dots & c_{1,k} \\ c_{2,1} & c_{2,2} & \dots & c_{2,k} \\ \vdots & \vdots & \vdots & \vdots \\ c_{q,1} & c_{q,2} & \dots & c_{q,k} \end{pmatrix} \quad b = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_q \end{pmatrix}$$

• Output:

Res.Df gradi di libertà della devianza residua RSS devianza residua Df gradi di libertà della devianza relativa all'ipotesi nulla H_0 Sum of Sq devianza relativa all'ipotesi nulla H_0

 \mathbb{F} valore empirico della statistica F

Pr(>F) p-value

• Formula:

Res.Df
$$n-k \qquad n-k+q$$
 RSS
$$RSS + \left(b-C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b-C\,\hat{\beta}\right)$$
 Df
$$-q$$
 Sum of Sq
$$-\left(b-C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b-C\,\hat{\beta}\right)$$
 F
$$Fvalue = \frac{\left[\left(b-C\,\hat{\beta}\right)^T\,\left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b-C\,\hat{\beta}\right)\right]/q}{RSS/(n-k)}$$
 Pr (>F)
$$P(F_{a,n-k} \geq Fvalue)$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> C <- matrix(data = c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3), nrow = 2, ncol = 4, byrow = TRUE)
> C

    [,1] [,2] [,3] [,4]
[1,] 1 3 5.0 2.3
[2,] 2 4 1.1 4.3
```

```
> b <- c(1.1, 2.3)
[1] 1.1 2.3
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b)
Linear hypothesis test
Hypothesis:
 (Intercept) + 3 x1 + 5 x2 + 2.3 x3 = 1.1
2 (Intercept) + 4 \times 1 + .1 \times 2 + 4.3 \times 3 = 2.3
Model 1: y \sim x1 + x2 + x3
Model 2: restricted model
     Res.Df
                                           RSS Df Sum of Sq F Pr(>F)
                               6.7965
            4
                      6 17.9679 -2 -11.1713 3.2874 0.1431
> res <- linear.hypothesis(model = modello, hypothesis.matrix = C,</pre>
                rhs = b)
> q < - 2
> c(n - k, n - k + q)
[1] 4 6
> res$Res.Df
[1] 4 6
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(residuals(object = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*% beta) %*% solve(t(X) %*% solve(t(X) %*% solve(t(X) %*% beta) %*% solve(t(X) %*% solve(
                 X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
[1] 6.796529 17.967863
> res$RSS
[1] 6.796529 17.967863
> -q
[1] -2
> res$Df
[1] NA -2
> -CSS
[1] -11.17133
> res$"Sum of Sq"
 [1]
                                    NA -11.17133
```

```
> Fvalue <- (CSS/q)/(RSS/(n - k))
 > Fvalue
 [1] 3.287364
 > res$F
 [1] NA 3.287364
 > 1 - pf(Fvalue, df1 = q, df2 = n - k)
 [1] 0.1430808
 > res$"Pr(>F)"
 [1]
           NA 0.1430808
• Example 2:
 > k <- 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > C <- matrix(data = c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3, 12.3, 3.4,
       4.5, 6.9), nrow = 3, ncol = 4, byrow = TRUE)
      [,1] [,2] [,3] [,4]
 [1,] 1.0 3.0 5.0 2.3
 [2,] 2.0 4.0 1.1 4.3
 [3,] 12.3 3.4 4.5 6.9
 > b <- c(1.1, 2.3, 5.6)
 > b
 [1] 1.1 2.3 5.6
 > linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b)
 Linear hypothesis test
 Hypothesis:
 (Intercept) + 3 x1 + 5 x2 + 2.3 x3 = 1.1
 2 (Intercept) + 4 \times 1 + .1 \times 2 + 4.3 \times 3 = 2.3
 2.3 (Intercept) + 3.4 \times 1 + 4.5 \times 2 + 6.9 \times 3 = 5.6
 Model 1: y \sim x1 + x2 + x3
 Model 2: restricted model
  Res.Df
             RSS Df Sum of Sq F Pr(>F)
            6.797
 1 4
       7 109.041 -3 -102.244 20.058 0.007131 **
 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
 > res <- linear.hypothesis(model = modello, hypothesis.matrix = C,
 + rhs = b)
 > q < - 3
 > c(n - k, n - k + q)
```

```
[1] 4 7
> res$Res.Df
[1] 4 7
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(residuals(object = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*%
     X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
[1] 6.796529 109.040699
> res$RSS
[1] 6.796529 109.040699
> -q
[1] -3
> res$Df
[1] NA -3
> -CSS
[1] -102.2442
> res$"Sum of Sq"
[1]
         NA -102.2442
> Fvalue <- (CSS/q)/(RSS/(n - k))
> Fvalue
[1] 20.05811
> res$F
[1] NA 20.05811
> 1 - pf(Fvalue, df1 = q, df2 = n - k)
[1] 0.007131315
> res$"Pr(>F)"
[1]
       NA 0.007131315
```

lht()

• Package: car

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità hypothesis.matrix matrice C di dimensione $q \times k$ e rango pari a $q = \min(q, k)$ rhs vettore b della previsione lineare di dimensione q

• **Description:** test di ipotesi per $H_0: C\beta = b$ contro $H_1: C\beta \neq b$ dove C e b sono così definiti:

$$C = \begin{pmatrix} c_{1,1} & c_{1,2} & \dots & c_{1,k} \\ c_{2,1} & c_{2,2} & \dots & c_{2,k} \\ \vdots & \vdots & \vdots & \vdots \\ c_{q,1} & c_{q,2} & \dots & c_{q,k} \end{pmatrix} \quad b = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_q \end{pmatrix}$$

• Output:

Res.Df gradi di libertà della devianza residua

RSS devianza residua

 ${\tt Df}\;$ gradi di libertà della devianza relativa all'ipotesi nulla H_0

Sum of Sq devianza relativa all'ipotesi nulla \mathcal{H}_0

 \mathbb{F} valore empirico della statistica F

Pr(>F) p-value

• Formula:

Res.Df
$$n-k \qquad n-k+q$$
 RSS
$$RSS + \left(b-C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1} \,\left(b-C\,\hat{\beta}\right)$$
 Df
$$-q$$
 Sum of Sq
$$-\left(b-C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1} \,\left(b-C\,\hat{\beta}\right)$$
 F
$$Fvalue = \frac{\left[\left(b-C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,X\right)^{-1}\,C^T\right]^{-1} \,\left(b-C\,\hat{\beta}\right)\right]/q}{RSS/(n-k)}$$
 Pr (>F)
$$P(F_{a,n-k} \geq Fvalue)$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> C <- matrix(data = c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3), nrow = 2, ncol = 4, byrow = TRUE)
> C
[,1] [,2] [,3] [,4]
[1,] 1 3 5.0 2.3
[2,] 2 4 1.1 4.3
```

```
> b <- c(1.1, 2.3)
[1] 1.1 2.3
> lht(model = modello, hypothesis.matrix = C, rhs = b)
Linear hypothesis test
Hypothesis:
(Intercept) + 3 \times 1 + 5 \times 2 + 2.3 \times 3 = 1.1
2 (Intercept) + 4 x1 + .1 x2 + 4.3 x3 = 2.3
Model 1: y \sim x1 + x2 + x3
Model 2: restricted model
 Res.Df
             RSS Df Sum of Sq F Pr(>F)
          6.7965
     4
       6 17.9679 -2 -11.1713 3.2874 0.1431
> res <- lht(model = modello, hypothesis.matrix = C, rhs = b)
> q < - 2
> c(n - k, n - k + q)
[1] 4 6
> res$Res.Df
[1] 4 6
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(residuals(object = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*%
+ X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
[1] 6.796529 17.967863
> res$RSS
[1] 6.796529 17.967863
> -q
[1] -2
> res$Df
[1] NA -2
> -CSS
[1] -11.17133
> res$"Sum of Sq"
[1]
           NA -11.17133
```

```
> Fvalue <- (CSS/q)/(RSS/(n - k))
 > Fvalue
 [1] 3.287364
 > res$F
 [1]
          NA 3.287364
 > 1 - pf(Fvalue, df1 = q, df2 = n - k)
 [1] 0.1430808
 > res$"Pr(>F)"
           NA 0.1430808
 [1]
• Example 2:
 > k < - 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > C <- matrix(data = c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3, 12.3, 3.4,
       4.5, 6.9), nrow = 3, ncol = 4, byrow = TRUE)
 > C
       [,1] [,2] [,3] [,4]
 [1,] 1.0 3.0 5.0 2.3
 [2,] 2.0 4.0 1.1 4.3
 [3,] 12.3 3.4 4.5 6.9
 > b < -c(1.1, 2.3, 5.6)
 > h
 [1] 1.1 2.3 5.6
 > lht(model = modello, hypothesis.matrix = C, rhs = b)
 Linear hypothesis test
 Hypothesis:
 (Intercept) + 3 x1 + 5 x2 + 2.3 x3 = 1.1
 2 (Intercept) + 4 \times 1 + .1 \times 2 + 4.3 \times 3 = 2.3
 2.3 (Intercept) + 3.4 \times 1 + 4.5 \times 2 + 6.9 \times 3 = 5.6
 Model 1: y \sim x1 + x2 + x3
 Model 2: restricted model
  Res.Df
              RSS Df Sum of Sq F Pr(>F)
             6.797
 1 4
        7 109.041 -3 -102.244 20.058 0.007131 **
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 > res <- lht(model = modello, hypothesis.matrix = C, rhs = b)</pre>
 > q < - 3
 > c(n - k, n - k + q)
```

```
[1] 4 7
> res$Res.Df
[1] 4 7
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(residuals(object = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*%
     X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
[1] 6.796529 109.040699
> res$RSS
[1] 6.796529 109.040699
> -q
[1] -3
> res$Df
[1] NA -3
> -CSS
[1] -102.2442
> res$"Sum of Sq"
[1]
         NA -102.2442
> Fvalue <- (CSS/q)/(RSS/(n - k))
> Fvalue
[1] 20.05811
> res$F
[1] NA 20.05811
> 1 - pf(Fvalue, df1 = q, df2 = n - k)
[1] 0.007131315
> res$"Pr(>F)"
[1]
       NA 0.007131315
```

lm.ridge()

• Package: MASS

• Input:

formula modello di regressione lineare con k-1 variabili esplicative ed n unità lambda valore del parametro λ

• Description: Ridge-Regression

• Output:

coef stime scales scarto quadratico medio delle k-1 variabili esplicative lambda λ ym media della variabile dipendente xm media delle k-1 variabili esplicative GCV i valori di λ e GCV kHKB kHKB kLW kLW

• Formula:

coef
$$V\left(D^{2} + \lambda I_{k-1}\right)^{-1}DU^{T}\left(y - \bar{y}\right)$$
 scales
$$\sigma_{x_{j}} \quad \forall j = 1, 2, \ldots, k-1$$
 lambda
$$\lambda$$
 ym
$$\bar{y}$$
 xm
$$\bar{x}_{j} \quad \forall j = 1, 2, \ldots, k-1$$
 GCV
$$\left(\frac{(y - \bar{y})^{T}\left(I_{n} - UD\left(D^{2} + \lambda I_{k-1}\right)^{-1}DU^{T}\right)^{2}\left(y - \bar{y}\right)}{\left(n - \sum_{i=1}^{k-1} \frac{D_{i,i}^{2}}{\lambda + D_{i,i}^{2}}\right)^{2}}\right)$$
 kHKB
$$\frac{k - 3}{n - k} \frac{(y - \bar{y})^{T}\left(I_{n} - UU^{T}\right)(y - \bar{y})}{(y - \bar{y})^{T}UD^{-2}U^{T}\left(y - \bar{y}\right)}$$
 kLW
$$\frac{n(k - 3)}{n - k} \frac{(y - \bar{y})^{T}\left(I_{n} - UU^{T}\right)(y - \bar{y})}{(y - \bar{y})^{T}UU^{T}\left(y - \bar{y}\right)}$$

• Example 1:

> res\$scales

```
x1
               x2
 2.412986 2.352359 2.195831
 > res$lambda
 [1] 1.2
 > res$ym
 [1] 7.52
 > res$xm
          x2
     x1
                 x3
 4.9000 6.0125 6.2450
 > res$GCV
       1.2
 0.2049004
 > res$kHKB
 [1] 0.483875
 > res$kLW
 [1] 0.3325936
• Example 2:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > res <- lm.ridge(formula = modello, lambda = 3.78)</pre>
 > res$coef
                  x2
 0.5765168 0.6291156 0.8724114
 > res$scales
                x2
       x1
 2.412986 2.352359 2.195831
 > res$lambda
 [1] 3.78
 > res$ym
 [1] 7.52
 > res$xm
```

- Note 1: La matrice del modello X viene privata della prima colonna (intercetta) e poi trasformata nella matrice standardizzata Z. Successivamente viene applicata la fattorizzazione ai valori singolari $Z = UDV^T$ mediante il comando $\operatorname{svd}()$.
- Note 2: I parametri stimati sono k-1 e non k (modello senza intercetta).

cov2cor()

• Package: stats

• Input:

 ${\tt V}\;$ matrice di covarianza delle stime OLS di dimensione $k\times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$$

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> V <- vcov(object = modello)
> cov2cor(V)
             (Intercept)
                                 x1
                                              x2
(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900
            -0.18600997 1.0000000 -0.82213982
                                                 0.4883764
x1
                                     1.00000000 -0.8022181
             0.07158062 -0.8221398
x2
x3
            -0.46329002 0.4883764 -0.80221810
```

14.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** log-verosimiglianza normale
- Formula:

 $\hat{\ell}$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> logLik(object = modello)
'log Lik.' -10.69939 (df=5)
```

durbin.watson()

• Package: car

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / RSS$$

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> res <- durbin.watson(model = modello)

> res$dw
```

AIC()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

• **Description:** indice AIC

• Formula:

$$-2\hat{\ell} + 2(k+1)$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> AIC(object = modello)

[1] 31.39878
```

BIC()

• Package: nlme

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

• **Description:** indice *BIC*

• Formula:

$$-2\,\hat{\ell} + (k+1)\,\log(n)$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> BIC(object = modello)

[1] 31.79599
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k = n \log(RSS/n) + 2k$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> extractAIC(fit = modello)

[1] 4.000000 6.695764
```

deviance()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

RSS

• Examples:

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> deviance(object = modello)
```

PRESS()

• Package: MPV

• Input:

 \times modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description**: PRESS
- Formula:

$$\sum_{i=1}^{n} e_i^2 / (1 - h_i)^2$$

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> PRESS(x = modello)

[1] 35.00228
```

drop1()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità scale selezione indice AIC oppure Cp test = "F"

• **Description:** submodels

• Output:

Df differenza tra gradi di libertà Sum of Sq differenza tra devianze residue RSS devianza residua AIC indice AIC Cp indice Cp F value valore empirico della statistica F Pr (F) p-value

• Formula:

Df

$$\underbrace{1, 1, \ldots, 1}_{k-1 \text{ volte}}$$

Sum of Sq

$$RSS_{-x_i} - RSS \quad \forall j = 1, 2, \dots, k-1$$

dove RSS_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

RSS

$$RSS, RSS_{-x_{j}} \quad \forall j = 1, 2, ..., k-1$$

AIC

$$n \log (RSS/n) + 2k, n \log (RSS_{-x_i}/n) + 2(k-1) \quad \forall j = 1, 2, ..., k-1$$

Ср

scale =
$$s^2$$

$$k, \frac{RSS_{-x_{j}}}{RSS/(n-k)} + 2(k-1) - n \quad \forall j = 1, 2, ..., k-1$$

F value

$$F_j = \frac{RSS_{-x_j} - RSS}{RSS/(n-k)} \quad \forall j = 1, 2, ..., k-1$$

Pr(F)

$$P(F_{1,n-k} \ge F_j) \quad \forall j = 1, 2, \dots, k-1$$

• Example 1:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y ~ x1 + x2 + x3)
> drop1(object = modello, scale = 0, test = "F")
```

```
Single term deletions
 Model:
 y \sim x1 + x2 + x3
                        RSS AIC F value Pr(F)
        Df Sum of Sq
                      6.7965 6.6958
             2.0115 8.8080 6.7698
                                        1.1838 0.3377
        1 1.513e-05 6.7965 4.6958 8.905e-06 0.9978
 x2
             5.2616 12.0581 9.2824
                                        3.0966 0.1533
 x3
        1
 > res <- drop1(object = modello, scale = 0, test = "F")</pre>
 > res$Df
 [1] NA 1 1 1
 > res$"Sum of Sq"
              NA 2.011499e+00 1.513044e-05 5.261577e+00
 [1]
 > res$RSS
 [1] 6.796529 8.808029 6.796544 12.058107
 > res$AIC
 [1] 6.695764 6.769777 4.695782 9.282365
 > res$"F value"
              NA 1.183839e+00 8.904801e-06 3.096626e+00
 [1]
 > res$"Pr(F)"
           NA 0.3377443 0.9977619 0.1532663
• Example 2:
 > k <- 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > s
 [1] 1.303508
 > drop1(object = modello, scale = s^2, test = "F")
 Single term deletions
 Model:
 y \sim x1 + x2 + x3
 scale: 1.699132
        Df Sum of Sq
                        RSS
                               Cp F value Pr(F)
                      6.7965 4.0000
 x1
             2.0115 8.8080 3.1838
                                      1.1838 0.3377
         1 1.513e-05 6.7965 2.0000 8.905e-06 0.9978
 x2
             5.2616 12.0581 5.0966
                                      3.0966 0.1533
 x3
```

```
> res <- drop1(object = modello, scale = s^2, test = "F")</pre>
    > res$Df
    [1] NA 1 1 1
    > res$"Sum of Sq"
                     NA 2.011499e+00 1.513044e-05 5.261577e+00
    [1]
    > res$RSS
    [1] 6.796529 8.808029 6.796544 12.058107
    > res$Cp
    [1] 4.000000 3.183839 2.000009 5.096626
    > res$"F value"
                     NA 1.183839e+00 8.904801e-06 3.096626e+00
    [1]
    > res$"Pr(F)"
    [1]
                 NA 0.3377443 0.9977619 0.1532663
add1()
  • Package: stats
  • Input:
         object modello nullo di regressione lineare
         scope modello di regressione lineare con k-1 variabili esplicative ed n unità
         scale selezione indice AIC oppure Cp
         test = "F"
  • Description: submodels
  • Output:
         Df differenza tra gradi di libertà
         Sum of Sq differenza tra devianze residue
         RSS devianza residua
         AIC indice AIC
         {\it Cp}\ \ {\it indice}\ {\it Cp}
         {\mathbb F} value valore empirico della statistica F
         Pr(F) p-value
  • Formula:
         Df
                                                     \underbrace{1, 1, \dots, 1}_{k-1 \text{ volte}}
         Sum of Sq
                                       RSS_{nullo} - RSS_{x_i} \quad \forall j = 1, 2, \ldots, k-1
```

dove RSS_{x_j} rappresenta la devianza residua del modello con la sola variabile esplicativa x_j .

```
RSS
                                    RSS_{nullo}, RSS_{x_i} \quad \forall j = 1, 2, ..., k-1
      AIC
                                                scale = 0
                       n \log (RSS_{nullo}/n) + 2, n \log (RSS_{x_i}/n) + 4 \quad \forall j = 1, 2, ..., k-1
      Ср
                                              scale = s^2
                       \frac{RSS_{nullo}}{RSS/(n-k)} + 2 - n, \, \frac{RSS_{x_j}}{RSS/(n-k)} + 4 - n \quad \forall j = 1, 2, \dots, k-1
      F value
                                F_j = \frac{RSS_{nullo} - RSS_{x_j}}{RSS_{x_i} / (n-2)} \quad \forall j = 1, 2, \dots, k-1
      Pr(F)
                                    P(F_{1,n-2} \ge F_j) \quad \forall j = 1, 2, \dots, k-1
• Example 1:
 > k < - 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > nullo <- lm(formula = y \sim 1)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > add1(object = nullo, scope = modello, scale = 0, test = "F")
 Single term additions
 Model:
 y ~ 1
          Df Sum of Sq
                           RSS AIC F value
                         47.666 16.278
               26.149 21.518 11.915 7.2914 0.035564 *
                35.492 12.175 7.359 17.4911 0.005799 **
 x2
                34.691 12.975 7.869 16.0418 0.007077 **
 x3
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 > res <- add1(object = nullo, scope = modello, scale = 0, test = "F")
 > res$Df
 [1] NA 1 1 1
 > res$"Sum of Sq"
            NA 26.14878 35.49165 34.69113
  [1]
 > res$RSS
  [1] 47.66640 21.51762 12.17475 12.97527
 > res$AIC
  [1] 16.278282 11.915446 7.359380 7.868828
```

```
> res$"F value"
 [1] NA 7.291356 17.491113 16.041811
 > res$"Pr(F)"
            NA 0.035564122 0.005799048 0.007076764
 [1]
• Example 2:
 > k < - 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > nullo <- lm(formula = y \sim 1)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > s
 [1] 1.303508
 > add1(object = nullo, scope = modello, scale = s^2, test = "F")
 Single term additions
 Model:
 y ~ 1
 scale: 1.699132
        Df Sum of Sq RSS Cp F value
                     47.666 22.0534
 <none>
             26.149 21.518 8.6639 7.2914 0.035564 *
 x1
             35.492 12.175 3.1653 17.4911 0.005799 **
 x2
        1
 x3
             34.691 12.975 3.6364 16.0418 0.007077 **
 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
 > res <- add1(object = nullo, scope = modello, scale = s^2, test = "F")</pre>
 > res$Df
 [1] NA 1 1 1
 > res$"Sum of Sq"
         NA 26.14878 35.49165 34.69113
 > res$RSS
 [1] 47.66640 21.51762 12.17475 12.97527
 > res$Cp
 [1] 22.053378 8.663889 3.165274 3.636408
 > res$"F value"
 [1]
          NA 7.291356 17.491113 16.041811
 > res$"Pr(F)"
            NA 0.035564122 0.005799048 0.007076764
 [1]
```

leaps()

• Package: leaps

• Input:

 $\times\,$ matrice del modello priva della prima colonna (intercetta) di dimensione $n\times(h-1)$

y variabile dipendente

method = "r2" / "adjr2" / "Cp" indice
$$R^2,\,R_{adj}^2,\,C_p$$
 nbest = 1

• Description: Best Subsets

• Output:

which variabili selezionate size numero di parametri r2 / adjr2 / Cp indice $R^2,\,R^2_{adj},\,C_p$

• Formula:

size

$$k_i \quad \forall j = 1, 2, \dots, h-1$$

Numero di esplicative	Numero di parametri	Numero di Subsets
1	$k_1 = 2$	$\binom{h-1}{1}$
2	$k_2 = 3$	$\binom{h-1}{2}$
	•	•
	•	•
j	$k_j = j + 1$	$\binom{h-1}{j}$
	•	
h-1	$k_{h-1} = h$	$\binom{h-1}{h-1}$

r2

$$R_i^2 \quad \forall j = 1, 2, \dots, h-1$$

 R_j^2 rappresenta il massimo R^2 tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

adjr2

$$R_{adj j}^{2} = 1 - \frac{RSS / (n - k_{j})}{RSS_{nullo} / (n - 1)}$$

$$= \frac{1 - k_{j}}{n - k_{j}} + \frac{n - 1}{n - k_{j}} R_{j}^{2} \quad \forall j = 1, 2, ..., h - 1$$

 $R^2_{adj\,j}$ rappresenta il massimo R^2_{adj} tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

Ср

$$Cp_{j} = (n - k_{h-1}) \frac{1 - R_{j}^{2}}{1 - R_{h-1}^{2}} + 2 k_{j} - n$$

$$= \left(\frac{n - k_{h-1}}{1 - R_{h-1}^{2}} + 2 k_{j} - n\right) - \frac{n - k_{h-1}}{1 - R_{h-1}^{2}} R_{j}^{2} \qquad \forall j = 1, 2, \dots, h - 1$$

 Cp_j rappresenta il minimo Cp tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

• Example 1: > k <- 4

```
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > X <- model.matrix(object = modello)
 > A <- X[, -1]
 > leaps(x = A, y, method = "r2", nbest = 1)
 $which
      1 2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 $label
 [1] "(Intercept)" "1"
                                 "2"
 $size
 [1] 2 3 4
 [1] 0.7445843 0.8574144 0.8574147
 > res <- leaps(x = A, y, method = "r2", nbest = 1)
 > res$which
         2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 > res$size
 [1] 2 3 4
 > res$r2
 [1] 0.7445843 0.8574144 0.8574147
• Example 2:
 > k <- 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > X <- model.matrix(object = modello)
 > A \leftarrow X[, -1]
 > leaps(x = A, y, method = "adjr2", nbest = 1)
```

```
$which
             2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 $label
 [1] "(Intercept)" "1"
                                 "2"
                                                "3"
 $size
 [1] 2 3 4
 $adjr2
 [1] 0.7020150 0.8003801 0.7504757
 > res <- leaps(x = A, y, method = "adjr2", nbest = 1)
 > res$which
             2
       1
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 > res$size
 [1] 2 3 4
 > res$adjr2
 [1] 0.7020150 0.8003801 0.7504757
• Example 3:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > X <- model.matrix(object = modello)</pre>
 > A <- X[, -1]
 > leaps(x = A, y, method = "Cp", nbest = 1)
 $which
      1
             2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 $label
 [1] "(Intercept)" "1"
                                  "2"
                                                "3"
 $size
 [1] 2 3 4
 $Cp
 [1] 3.165274 2.000009 4.000000
 > res <- leaps(x = A, y, method = "Cp", nbest = 1)
 > res$which
```

```
1 2 3
1 FALSE TRUE FALSE
2 TRUE FALSE TRUE
3 TRUE TRUE TRUE
> res$size

[1] 2 3 4
> res$Cp

[1] 3.165274 2.000009 4.000000
```

- Note 1: Tutti i modelli contengono l'intercetta.
- Note 2: R_{adij}^2 è una trasformazione lineare crescente di $R_i^2 \quad \forall j = 1, 2, ..., h-1$.
- Note 3: Cp_j è una trasformazione lineare decrescente di $R_j^2 \quad \forall j = 1, 2, \ldots, h-1$.

bptest()

- Package: lmtest
- Input:

formula modello di regressione lineare con k-1 variabili esplicative ed n unità studentize = TRUE / FALSE metodo di Koenker

- Description: test di Breusch-Pagan per l'omoschedasticità dei residui
- Output:

statistic valore empirico della statistica χ^2 parameter gradi di libertà p.value p-value

• Formula:

statistic

parameter

p.value

$$v_i = e_i^2 - RSS/n \quad \forall i = 1, 2, \dots, n$$

$$c = n \frac{v^T H v}{v^T v}$$

$$\texttt{studentize} = \texttt{FALSE}$$

$$v_i = n e_i^2 / RSS - 1 \quad \forall i = 1, 2, \dots, n$$

$$c = \frac{1}{2} v^T H v$$

$$df = k - 1$$

$$P(\chi^2_{df} \geq c)$$

• Example 1:

```
> k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > bptest(formula = modello, studentize = TRUE)
          studentized Breusch-Pagan test
 data: modello
 BP = 3.2311, df = 3, p-value = 0.3574
 > res <- bptest(formula = modello, studentize = TRUE)</pre>
 > res$statistic
        BP
 3.231074
 > res$parameter
 df
  3
 > res$p.value
 0.3573517
• Example 2:
 > k <- 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > bptest(formula = modello, studentize = FALSE)
          Breusch-Pagan test
 data: modello
 BP = 0.9978, df = 3, p-value = 0.8018
 > res <- bptest(formula = modello, studentize = FALSE)</pre>
 > res$statistic
         ВP
 0.9977698
 > res$parameter
 df
  3
 > res$p.value
         BP
 0.8017916
```

14.4 Diagnostica

ls.diag()

• Package: stats

• Input:

ls.out modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: analisi di regressione lineare
- Output:

```
std.dev stima di \sigma hat valori di leva std.res residui standard stud.res residui studentizzati cooks distanza di Cook dfits dfits correlation matrice di correlazione tra le stime OLS std.err standard error delle stime OLS cov.scaled matrice di covarianza delle stime OLS non scalata per \sigma^2
```

• Formula:

hat
$$h_i \quad \forall i=1,2,\dots,n$$
 std.res
$$rstandard_i \quad \forall i=1,2,\dots,n$$
 stud.res
$$rstudent_i \quad \forall i=1,2,\dots,n$$
 cooks
$$cd_i \quad \forall i=1,2,\dots,n$$
 dfits
$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i=1,2,\dots,n$$
 correlation
$$r_{\hat{\beta}_i \hat{\beta}_j} \quad \forall i,j=1,2,\dots,k$$
 std.err
$$s_{\hat{\beta}_j} \quad \forall j=1,2,\dots,k$$
 cov.scaled
$$s^2 (X^T X)^{-1}$$
 cov.unscaled
$$(X^T X)^{-1}$$

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> res <- ls.diag(ls.out = modello)

> res$std.dev
```

x2

x3

```
[1] 1.303508
> res$hat
 \begin{smallmatrix} 1 \end{smallmatrix} ] \hspace{0.1cm} 0.7695906 \hspace{0.1cm} 0.4163361 \hspace{0.1cm} 0.3791092 \hspace{0.1cm} 0.3154744 \hspace{0.1cm} 0.7283511 \hspace{0.1cm} 0.5539241 \hspace{0.1cm} 0.4302463 
[8] 0.4069682
> res$std.res
\begin{bmatrix} 1 \end{bmatrix} -1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
[8] -1.4301703
> res$stud.res
[1] -2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
[8] -1.7718134
> res$cooks
[1] 1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
[8] 0.35091186
> res$dfits
\begin{bmatrix} 1 \end{bmatrix} -3.7255223 0.3280660 1.1157578 0.4018144 0.5475321 0.7916935 -0.8516950
[8] -1.4677742
> res$correlation
             (Intercept)
                                    x1
                                                  x2
(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900
             -0.18600997 1.0000000 -0.82213982 0.4883764
x1
              0.07158062 -0.8221398 1.00000000 -0.8022181
x2
x3
             -0.46329002 0.4883764 -0.80221810 1.0000000
> res$std.err
                   [,1]
(Intercept) 1.4292308
             0.3883267
x1
x2
             0.5822146
             0.4068987
x3
> res$cov.scaled
                                      x1
                                                   x2
              (Intercept)
(Intercept) 2.04270054 -0.10323710 0.05956359 -0.26942727
             -0.10323710 0.15079759 -0.18587712 0.07716815
x1
              0.05956359 -0.18587712 0.33897378 -0.19004733
x2
             -0.26942727 0.07716815 -0.19004733 0.16556652
x3
> res$cov.unscaled
             (Intercept)
                                      \times 1
                                                  x2
(Intercept) 1.20220217 -0.06075872 0.0350553 -0.15856757
             -0.06075872 0.08874976 -0.1093953 0.04541621
x1
```

 $0.03505530 - 0.10939532 \quad 0.1994982 - 0.11184964$

cooks.distance()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> cooks.distance(model = modello)
1 2 3 4 5 6 7
1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
8 0.35091186
```

cookd()

• Package: car

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> cookd(model = modello)
1 2 3 4 5 6 7
1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
8 0.35091186
```

rstandard()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

rstandard.lm()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

stdres()

• Package: MASS

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

rstudent()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \dots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> rstudent(model = modello)
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8
-1.7718134
```

rstudent.lm()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> rstudent.lm(model = modello)
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8
-1.7718134
```

studres()

• Package: MASS

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \dots, n$$

lmwork()

```
• Package: MASS
```

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: diagnostica di regressione
- Output:

```
stdedv stima di \sigma stdres residui standard studres residui studentizzati
```

• Formula:

```
stdedv s stdres rstandard_i \  \  \forall i=1,2,\ldots,n studres rstudent_i \  \  \forall i=1,2,\ldots,n
```

• Examples:

> res\$studres

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> lmwork(object = modello)
$stdedv
[1] 1.303508
$stdres
                      2
                                   3
                                                4
-1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
-1.4301703
$studres
                     2
                                   3
        1
                                                4
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
-1.7718134
> res <- lmwork(object = modello)</pre>
> res$stdedv
[1] 1.303508
> res$stdres
-1.5241225 \quad 0.4376576 \quad 1.2722093 \quad 0.6467323 \quad 0.3791111 \quad 0.7589935 \quad -0.9849613
-1.4301703
```

```
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8 -1.7718134
```

dffits()

• Package: stats

• Input:

modeln modello di regressione lineare con n-1 variabili esplicative ed n unità

- **Description**: dffits
- Formula:

$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i = 1, 2, \dots, n$$

• Examples:

covratio()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description**: covratio
- Formula:

$$cr_i \quad \forall i = 1, 2, \ldots, n$$

lm.influence()

- Package: stats
- Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** diagnostica di regressione
- Output:

```
hat valori di leva coefficients differenza tra le stime OLS eliminando una unità sigma stima di \sigma eliminando una unità wt.res residui
```

• Formula:

hat
$$h_i \quad \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j-\hat{\beta}_{j\,(-i)}=e_i\,(1-h_i)^{-1}\,(X^T\,X)_j^{-1}\,X_i^T\quad \forall i=1,2,\ldots,n\quad \forall j=1,2,\ldots,k$$
 sigma
$$s_{-i} \quad \forall i=1,2,\ldots,n$$
 wt.res
$$e_i \quad \forall i=1,2,\ldots,n$$

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> lm.influence(model = modello)
$hat
                           3
                                     4
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
$coefficients
  (Intercept)
                      x1
                                  x2
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 -0.15505077 0.14594807 -0.39064531 0.32853997
  0.10864633 -0.01436987 0.12965355 -0.11055404
  0.27248353 -0.28472521 0.38742501 -0.16358023
  0.36758841 0.18614884 -0.28071294 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
$sigma
                                              5
                           3
                                     4
0.9745992 1.4686808 1.1613865 1.4242946 1.4778725 1.3925645 1.3099769 1.0521638
$wt.res
                   2
                                                                         7
        1
           0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
-0.9536382
-1.4356227
```

```
> res <- lm.influence(model = modello)</pre>
> res$hat
                             3
                                        4
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
> res$coefficients
  (Intercept)
                        x1
                                    x2
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 -0.15505077 0.14594807 -0.39064531 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
5 \quad 0.06456839 \quad 0.14591697 \quad -0.04391330 \quad -0.06357315
0.27248353 - 0.28472521 0.38742501 - 0.16358023
7 \quad 0.36758841 \quad 0.18614884 \quad -0.28071294 \quad 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
> res$sigma
                            3
                                       4
                                                  5
                                                             6
0.9745992 1.4686808 1.1613865 1.4242946 1.4778725 1.3925645 1.3099769 1.0521638
> res$wt.res
                                3
                                           4
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
-1.4356227
```

influence()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: diagnostica di regressione
- Output:

hat valori di leva coefficients differenza tra le stime OLS eliminando una unità sigma stima di σ eliminando una unità wt.res residui

• Formula:

hat
$$h_i \quad \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j-\hat{\beta}_{j\,(-i)}=e_i\,(1-h_i)^{-1}\,(X^T\,X)_j^{-1}\,X_i^T\quad \forall i=1,2,\ldots,n\quad \forall j=1,2,\ldots,k$$
 sigma
$$s_{-i}\quad \forall i=1,2,\ldots,n$$
 wt.res
$$e_i\quad \forall i=1,2,\ldots,n$$

```
> k < -4
> x1 < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3)
> influence(model = modello)
$hat
                     3
                                    4
0.7695906\ 0.4163361\ 0.3791092\ 0.3154744\ 0.7283511\ 0.5539241\ 0.4302463\ 0.4069682
$coefficients
                     x1
 (Intercept)
                                 x2
1 - 3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 - 0.15505077 \quad 0.14594807 \quad -0.39064531 \quad 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
5 0.06456839 0.14591697 -0.04391330 -0.06357315
 6 \quad 0.27248353 \ -0.28472521 \quad 0.38742501 \ -0.16358023 \\
7 \quad 0.36758841 \quad 0.18614884 \quad -0.28071294 \quad 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
$sigma
                          3 4 5
0.9745992\ 1.4686808\ 1.1613865\ 1.4242946\ 1.4778725\ 1.3925645\ 1.3099769\ 1.0521638
$wt.res
              2
                             3
                                         4
                                                               6
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
-1.4356227
> res <- influence(model = modello)</pre>
> res$hat
                                4
                 2
                            3
0.7695906\ 0.4163361\ 0.3791092\ 0.3154744\ 0.7283511\ 0.5539241\ 0.4302463\ 0.4069682
> res$coefficients
                      x1
                                  x2
 (Intercept)
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 - 0.15505077 \quad 0.14594807 - 0.39064531 \quad 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
5 0.06456839 0.14591697 -0.04391330 -0.06357315
0.27248353 - 0.28472521 0.38742501 - 0.16358023
7 0.36758841 0.18614884 -0.28071294 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
> res$sigma
                                4
                                          5
                            3
0.9745992 1.4686808 1.1613865 1.4242946 1.4778725 1.3925645 1.3099769 1.0521638
> res$wt.res
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
-1.4356227
```

residuals()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

• **Description:** residui

• Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> residuals(object = modello)
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

residuals.lm()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> residuals.lm(object = modello)
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

residuals.default()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

• Description: residui

• Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> residuals.default(object = modello)
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

resid()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: residui
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> resid(object = modello)
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

df.residual()

• Package: stats

• Input:

object modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-k

• Examples:

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3)

> df.residual(object = modello)
```

hatvalues()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> hatvalues(model = modello)
1 2 3 4 5 6 7 8
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
```

hat()

• Package: stats

• Input:

x matrice del modello

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3)
> X <- model.matrix(object = modello)
> hat(x = X)

[1] 0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463
[8] 0.4069682
```

dfbeta()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** dfbeta
- Formula:

$$\hat{\beta}_j - \hat{\beta}_{j(-i)} = e_i (1 - h_i)^{-1} (X^T X)_i^{-1} X_i^T \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., k$$

• **Examples:**> k <- 4

```
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < -8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> dfbeta(model = modello)
  (Intercept)
                         x1
                                      x2
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 - 0.15505077 \quad 0.14594807 \quad -0.39064531 \quad 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
5 \quad 0.06456839 \quad 0.14591697 \quad -0.04391330 \quad -0.06357315
0.27248353 - 0.28472521 0.38742501 - 0.16358023
7 \quad 0.36758841 \quad 0.18614884 \quad -0.28071294 \quad 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
```

dfbetas()

• Package: stats

• Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description**: dfbetas
- Formula:

$$\frac{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}{s_{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}} = \frac{e_{i}\,(1 - h_{i})^{-1}\,(X^{T}\,X)_{j}^{-1}\,X_{i}^{T}}{s_{-i}\,\sqrt{(X^{T}\,X)_{j,\,j}^{-1}}} \quad \forall i = 1,\,2,\,\ldots,\,n \quad \forall j = 1,\,2,\,\ldots,\,k$$

• Examples:

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3)
> dfbetas(model = modello)
 (Intercept)
                 x1
1 -3.70059595 0.43942641 0.02349647 1.44767218
 3 -0.12176106  0.42183052 -0.75307182  0.90623075
 0.06957072 -0.03386642 0.20380513 -0.24865783
 0.66729165 - 0.75363662 \quad 0.79740312 - 1.05700791
```

vif()

- Package: car
- Input:

mod modello di regressione lineare con k-1 variabili esplicative ed n unità

- **Description:** variance inflation factors
- Formula:

$$\left(1 - R_{x_j}^2\right)^{-1} \quad \forall j = 1, 2, \dots, k - 1$$

 $R_{x_j}^2$ rappresenta il valore di R^2 per il modello che presenta il regressore j-esimo come variabile dipendente.

• Examples:

outlier.test()

- Package: car
- Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: test sugli outliers
- Output:

test massimo residuo studentizzato assoluto, gradi di libertà, p-value

• Formula:

test

$$t = \max(|rstudent_i|)$$
 $n - k - 1$ p -value $= 2P(t_{n-k-1} \le -|t|)$ $\forall i = 1, 2, ..., n$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3)
> outlier.test(model = modello)
max|rstudent| = 2.038485, degrees of freedom = 3,
unadjusted p = 0.1342423, Bonferroni p > 1
Observation: 1
> res <- outlier.test(model = modello)</pre>
> res$test
max|rstudent|
                         df unadjusted p Bonferroni p
    2.0384846 3.0000000 0.1342423
```

influence.measures()

- Package: stats
- Input:

model modello di regressione lineare con k-1 variabili esplicative ed n unità

- Description: dfbetas, dffits, covratio, distanza di Cook, valori di leva
- Output:

```
infmat misure di influenza di dimensione n \times (k+4) is .inf matrice di influenza con valori logici di dimensione n \times (k+4)
```

• Formula:

infmat

$$DFBETAS_{ij} = \frac{e_{i} (1-h_{i})^{-1} (X^{T} X)_{j}^{-1} X_{i}^{T}}{s_{-i} \sqrt{(X^{T} X)_{j,j}^{-1}}} \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., k$$

$$DFFITS_{i} = rstudent_{i} \sqrt{\frac{h_{i}}{1-h_{i}}} \quad \forall i = 1, 2, ..., n$$

$$COVRATIO_{i} = (1-h_{i})^{-1} \left(1 + \frac{rstudent_{i}^{2}-1}{n-k}\right)^{-k} \quad \forall i = 1, 2, ..., n$$

$$COOKD_{i} = \frac{h_{i} rstandard_{i}^{2}}{k(1-h_{i})} \quad \forall i = 1, 2, ..., n$$

$$HAT_{i} = h_{i} \quad \forall i = 1, 2, ..., n$$

```
> k < - 4
> x1 < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
```

```
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3)
 > res <- influence.measures(model = modello)</pre>
 > res
 Influence measures of
          lm(formula = y \sim x1 + x2 + x3):
    dfb.1_
            dfb.x1 dfb.x2 dfb.x3 dffit
                                          cov.r cook.d
 1 -3.7006 0.4394
                   0.0235
                           1.4477 -3.726
                                          0.424 1.9397 0.770
   0.1362 0.0440 -0.1874 0.1812 0.328
                                          4.450 0.0342 0.416
 1.116 0.640 0.2471 0.379
 4 0.0696 -0.0339 0.2038 -0.2487 0.402
                                         2.968 0.0482 0.315
  0.1785 -0.6863 0.6229 -0.3763 0.792
                                         3.804 0.1788 0.554
 7
   0.2559 0.4770 -0.4798 0.0765 -0.852
                                         1.826 0.1832 0.430
    0.6673 -0.7536  0.7974 -1.0570 -1.468  0.304  0.3509  0.407
 > res$infmat
        dfb.1_
                    dfb.x1
                               dfb.x2
                                           dfb.x3
                                                       dffit
                                                                  cov.r
               0.43942641 0.02349647 1.44767218 -3.7255223
 1 - 3.70059595
                                                             0.4238374
 2 \quad 0.13617748 \quad 0.04395152 \quad -0.18739044 \quad 0.18124433 \quad 0.3280660
                                                             4.4498753
 3 -0.12176106  0.42183052 -0.75307182  0.90623075
                                                   1.1157578
                                                             0.6395729
   0.06957072 -0.03386642 0.20380513 -0.24865783
                                                   0.4018144
                                                             2.9682483
   0.5475321 10.0502975
   0.7916935
                                                              3.8036903
               0.47699422 -0.47976587 0.07653668 -0.8516950
    0.25592307
                                                              1.8260516
   0.66729165 - 0.75363662 0.79740312 - 1.05700791 - 1.4677742
                                                             0.3038647
       cook.d
                   hat.
 1 1.93972080 0.7695906
 2 0.03415783 0.4163361
 3 0.24706215 0.3791092
 4 0.04819074 0.3154744
 5 0.09633983 0.7283511
 6 0.17883712 0.5539241
 7 0.18315058 0.4302463
 8 0.35091186 0.4069682
 > res$is.inf
   dfb.1_ dfb.x1 dfb.x2 dfb.x3 dffit cov.r cook.d
     TRUE FALSE FALSE
                        TRUE TRUE FALSE
                                           TRUE FALSE
   FALSE FALSE FALSE FALSE
 2
                                    TRUE FALSE FALSE
    FALSE FALSE
                 FALSE FALSE FALSE
                                           FALSE FALSE
          FALSE
                 FALSE
                        FALSE FALSE FALSE
                                           FALSE FALSE
    FALSE
    FALSE
          FALSE
                 FALSE
                        FALSE FALSE
                                     TRUE
                                           FALSE FALSE
   FALSE FALSE FALSE FALSE FALSE FALSE
 7
    FALSE FALSE FALSE FALSE FALSE FALSE
    FALSE FALSE FALSE
                         TRUE FALSE FALSE FALSE
• Note 1: Il caso i-esimo è influente se |DFBETAS_{ij}| > 1 \forall i = 1, 2, ..., n \forall j = 1, 2, ..., k
• Note 2: Il caso i-esimo è influente se |DFFITS_i| > 3\sqrt{k/(n-k)} \forall i = 1, 2, ..., n
• Note 3: Il caso i-esimo è influente se |1 - COVRATIO_i| > 3k/(n-k) \forall i = 1, 2, ..., n
• Note 4: Il caso i-esimo è influente se P(F_{k,n-k} \geq COOKD_i) > 0.5 \quad \forall i = 1, 2, ..., n
• Note 5: Il caso i-esimo è influente se HAT_i > 3k/n \quad \forall i = 1, 2, ..., n
• Note 6: I casi influenti rispetto ad almeno una tra queste misure sono marcati con un asterisco.
```

Corrispondentemente la stessa riga della matrice is .inf riporterà almeno un simbolo TRUE.

Capitolo 15

Regressione lineare semplice pesata

15.1 Simbologia

$$y_i = \beta_1 + \beta_2 \ x_i + \varepsilon_i \quad \forall i = 1, 2, ..., n \qquad \varepsilon \sim N(0, \sigma^2 W)$$

- variabile dipendente: y
- matrice del modello di dimensione $n \times 2$: X
- numero di parametri da stimare e rango della matrice del modello: 2
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_i) \quad \forall i = 1, 2, ..., n$
- vettore numerico positivo dei pesi WLS: $w = (w_1, w_2, \dots, w_n)$
- matrice diagonale definita positiva di dimensione $n \times n$: $W = \mathrm{diag}(w_1^{-1},\,w_2^{-1},\,\ldots,\,w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X (X^T W^{-1} X)^{-1} X^T W^{-1}$
- matrice identità di dimensione $n \times n$: I_n
- devianza residua: $RSS = \sum_{i=1}^n w_i \, e_i^2 = y^T \, W^{-1} \, e = y^T \, W^{-1} \, (I_n H) \, y$
- stima di σ^2 : $s^2 = RSS/(n-2)$
- gradi di libertà della devianza residua: n-2
- stima di σ^2 tolta la i-esima unità: $s_{-i}^2 = s^2 \left(1 + \frac{1 rstandard_i^2}{n-3}\right) = s^2 \left(1 + \frac{rstudent_i^2 1}{n-2}\right)^{-1} \quad \forall i = 1, 2, \ldots, n$
- codevianza pesata tra x ed y: $ss_{xy} = \sum_{i=1}^n w_i (x_i \bar{x}_W) (y_i \bar{y}_W)$
- devianza pesata di x: $ss_x = \sum_{i=1}^n w_i (x_i \bar{x}_W)^2$
- devianza pesata di y: $ss_y = \sum_{i=1}^n w_i \left(y_i \bar{y}_W\right)^2$
- stime WLS: $\hat{\beta} = (X^T W^{-1} X)^{-1} X^T W^{-1} y$
- stima WLS intercetta: $\hat{\beta}_1 = \bar{y}_W \bar{x}_W s s_{xy} / s s_x$
- stima WLS coefficiente angolare: $\hat{\beta}_2 = s s_{xy} / s s_x$
- standard error delle stime WLS: $s_{\hat{\beta}} = s \sqrt{\mathrm{diag}((X^T W^{-1} X)^{-1})}$
- standard error della stima WLS intercetta: $s_{\hat{\beta}_1} = s \sqrt{\sum_{i=1}^n w_i x_i^2 / (s s_x \sum_{i=1}^n w_i)}$
- standard error della stima WLS coefficiente angolare: $s_{\hat{eta}_2} = s / \sqrt{s s_x}$
- covarianza tra le stime WLS: $s_{\hat{\beta}_1 \, \hat{\beta}_2} = -\bar{x}_W \, s^2 \, / \, s s_x$
- t-values delle stime WLS: $t_{\hat{\beta}} = \hat{\beta} / s_{\hat{\beta}}$
- residui: $e = (I_n H) y$
- residui pesati: $\sqrt{w_i} e_i \quad \forall i = 1, 2, ..., n$

- residui standard: $rstandard_i = \frac{e_i}{s\sqrt{(1-h_i)/w_i}} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = \frac{e_i}{s_{-i}\sqrt{(1-h_i)/w_i}} = rstandard_i\sqrt{\frac{n-3}{n-2-rstandard_i^2}} \quad \forall i=1,2,\ldots,n$
- valori adattati: $\hat{y} = H y$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- stime WLS tolta la i-esima unità: $\hat{\beta}_{(-i)} \quad \forall i = 1, 2, ..., n$
- correlazione delle stime WLS: $r_{\hat{\beta}_i\,\hat{\beta}_j}=rac{s^2\,(X^T\,W^{-1}\,X)_{(i,\,j)}^{-1}}{s_{\hat{\beta}_i}\,s_{\hat{\beta}_j}}\quad \forall\,i,j\,=\,1,\,2$
- devianza residua modello nullo: $RSS_{nullo} = \sum_{i=1}^n w_i (y_i \bar{y}_W)^2 = (y \bar{y}_W)^T W^{-1} (y \bar{y}_W)$
- indice di determinazione: $R^2=1-RSS/RSS_{nullo}=1-\left(1-R_{adj}^2\right)\left(n-2\right)/\left(n-1\right)=r_{xy}^2$
- indice di determinazione aggiustato: $R_{adj}^2 = 1 \frac{RSS / (n-2)}{RSS_{nullo} / (n-1)} = 1 \left(1 R^2\right) \left(n 1\right) / \left(n 2\right)$
- valore noto dei regressori per la previsione: x_0
- log-verosimiglianza normale: $\hat{\ell} = -n \left(\log(2\pi) + \log\left(RSS/n\right) + 1 \sum_{i=1}^{n} \log(w_i)/n\right)/2$
- distanza di Cook: $cd_i=\frac{h_i\,rstandard_i^2}{2\,(1-h_i)}=\frac{e_i^2}{2\,s^2}\,\frac{h_i}{(1-h_i)^2}\quad \forall\,i=1,\,2,\,\ldots,\,n$
- covratio: $cr_i = (1 h_i)^{-1} \left(1 + \frac{rstudent_i^2 1}{n 2} \right)^{-2} = (1 h_i)^{-1} \left(\frac{s_{-i}}{s} \right)^4 \quad \forall i = 1, 2, \dots, n$

15.2 Stima

lm()

- Package: stats
- Input:

formula modello di regressione lineare pesata con una variabile esplicativa ed n unità weights pesi

x = TRUE matrice del modello

y = TRUE variabile dipendente

- Description: analisi di regressione lineare pesata
- Output:

coefficients stime WLS

residuals **residui**

fitted.values valori adattati

weights pesi

rank rango della matrice del modello

df.residual gradi di libertà della devianza residua

x matrice del modello

y variabile dipendente

• Formula:

coefficients

$$\hat{\beta}_i \quad \forall i = 1, 2$$

residuals

$$e_i \quad \forall i = 1, 2, \ldots, n$$

fitted.values

$$\hat{y}_i \quad \forall i = 1, 2, \dots, n$$

```
weights
                                     w_i \quad \forall i = 1, 2, \ldots, n
     rank
                                            2
     df.residual
                                           n-2
     Х
                                            X
     У
                                            y
• Examples:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x, weights = rep(1/n, n), x = TRUE,
 + y = TRUE)
 > modello$coefficients
 (Intercept)
   3.8486818 0.7492486
 > modello$residuals
                                  3
 -3.17285530 0.82804637 2.37969944 -0.06864749 -1.65699442 1.40387291
  0.55552598 -0.26864749
 > modello$fitted.values
                            3
                                      4
                                                5
  4.672855 5.571954 7.220301 8.868647 10.516994 6.396127 8.044474 8.868647
 > modello$weights
 [1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125
 > modello$rank
 [1] 2
 > modello$df.residual
 [1] 6
 > modello$x
  (Intercept) x
          1 1.1
            1 2.3
 2
 3
             1 4.5
 4
             1 6.7
 5
             1 8.9
 6
             1 3.4
 7
            1 5.6
 attr(, "assign")
 [1] 0 1
```

> modello\$y

- Note 1: Il modello nullo si ottiene attraverso con lm(formula = y ~ 1, weights = w).
- Note 2: L'istruzione $lm (formula = y \sim x, weights = w)$ è equivalente $a lm (formula = y \sim X 1, weight$
- Note 3: L'istruzione lm (formula = y ~ x, weights = w) è equivalente a lm (formula = y ~ 1 + x, weight

summary.lm()

- Package: stats
- Input:

object modello di regressione lineare pesata con una variabile esplicativa ed n unità correlation = TRUE correlazione delle stime WLS

- Description: analisi di regressione lineare pesata
- Output:

residuals residui coefficients stima puntuale, standard error, t-value, p-value sigma stima di σ r.squared indice di determinazione adj.r.squared indice di determinazione aggiustato fstatistic valore empirico della statistica F, df numeratore, df denominatore cov.unscaled matrice di covarianza delle stime WLS non scalata per σ^2 correlation matrice di correlazione delle stime WLS

• Formula:

residuals
$$e_i \quad \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j \quad s_{\hat{\beta}_j} \quad t_{\hat{\beta}_j} \quad p\text{-value} = 2P(t_{n-2} \leq -|t_{\hat{\beta}_j}|) \quad \forall j=1,2$$
 sigma
$$s$$
 r.squared
$$R^2$$
 adj.r.squared
$$R^2_{adj}$$
 fstatistic
$$Fvalue = \frac{RSS_{nullo} - RSS}{RSS/(n-2)} = t_{\hat{\beta}_2}^2 \qquad 1 \qquad n-2$$
 cov.unscaled
$$(X^TW^{-1}X)^{-1}$$
 correlation
$$r_{\hat{\beta}_i,\hat{\beta}_i} \quad \forall i,j=1,2$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x, weights = rep(1/n, n))
> res <- summary.lm(object = modello, correlation = TRUE)</pre>
> res$residuals
                      2
                                  3
                                              4
-1.12177375 0.29275860 0.84135081 -0.02427055 -0.58583599 0.49634403
 0.19640809 -0.09498123
> res$coefficients
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.8486818 1.5155372 2.539484 0.04411163
           0.7492486  0.2774737  2.700251  0.03556412
> res$sigma
[1] 0.66954
> res$r.squared
[1] 0.5485788
> res$adj.r.squared
[1] 0.4733419
> res$fstatistic
  value
          numdf dendf
7.291356 1.000000 6.000000
> res$cov.unscaled
            (Intercept)
(Intercept)
            5.1236582 -0.8415629
            -0.8415629 0.1717475
> res$correlation
            (Intercept)
(Intercept) 1.0000000 -0.8971215
            -0.8971215 1.0000000
```

vcov()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: matrice di covarianza delle stime WLS
- Formula:

$$s^2 (X^T W^{-1} X)^{-1}$$

• Examples:

lm.wfit()

• Package: stats

• Input:

x matrice del modello

y variabile dipendente

w pesi

• Description: analisi di regressione lineare pesata

• Output:

```
coefficients stime WLS
residuals residui
fitted.values valori adattati
weights pesi
rank rango della matrice del modello
df.residual gradi di libertà della devianza residua
```

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall j=1,2,\dots,k$$
 residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 fitted.values
$$\hat{y}_i \quad \forall i=1,2,\dots,n$$
 weights
$$w_i \quad \forall i=1,2,\dots,n$$
 rank
$$k$$
 df.residual
$$n-k$$

• Examples:

```
> x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x, weights = rep(1/n, n))
> X <- model.matrix(object = modello)</pre>
> res <- lm.wfit(x = X, y, w = rep(1/n, n))
> res$coefficients
(Intercept)
 3.8486818 0.7492486
> res$residuals
[7] 0.55552598 -0.26864749
> res$fitted.values
[1] 4.672855 5.571954 7.220301 8.868647 10.516994 6.396127 8.044474
[8] 8.868647
> res$weights
[1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125
> res$rank
[1] 2
> res$df.residual
[1] 6
```

lsfit()

• Package: stats

• Input:

```
x matrice del modello
y variabile dipendente
wt pesi
intercept = FALSE
```

- Description: analisi di regressione lineare pesata
- Output:

```
coefficients stime WLS
residuals residui
wt pesi
```

• Formula:

coefficients

$$\hat{\beta}_i \quad \forall j = 1, 2$$

```
residuals e_i \quad \forall i = 1, 2, \ldots, n wt w_i \quad \forall i = 1, 2, \ldots, n
```

• Examples:

confint()

• Package: stats

• Input:

object modello di regressione lineare pesata con una variabile esplicativa ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- Description: intervallo di confidenza per le stime WLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-2} s_{\hat{\beta}_j} \quad \forall j = 1, 2$$

coef()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** stime WLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2$$

• Examples:

fitted()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

predict.lm()

• Package: stats

• Input:

```
object modello di regressione lineare pesata con una variabile esplicativa ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

• Description: intervallo di confidenza o di previsione

• Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

```
fit x_0^T \hat{\beta} \qquad x_0^T \hat{\beta} \mp t_{1-\alpha/2,\,df} \, s^* \, \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} \qquad \qquad \text{interval = "prediction"} x_0^T \hat{\beta} \qquad x_0^T \hat{\beta} \mp t_{1-\alpha/2,\,df} \, s^* \, \sqrt{1 + x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} se.fit s^* \, \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} \text{df} \qquad \qquad df = n-2 residual.scale s^*
```

• Example 1:

```
> x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x, weights = rep(1/n, n))
> x0 <- c(1, 1.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 4.822705
> new <- data.frame(x = 1.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)</pre>
> W <- diag(1/rep(1/n, n))
> lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
     solve(W) %*% X) %*% x0)
> upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
     solve(W) %*% X) %*% x0)
> c(yhat, lower, upper)
[1] 4.822705 2.465776 7.179634
> res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
+ scale = s, df = Inf, interval = "confidence", level = 0.95)
> res$fit
       fit
                lwr
1 4.822705 2.465776 7.179634
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
+ X) %*% x0))
> se.fit
```

```
[1] 1.202537
 > res$se.fit
 [1] 1.202537
 > s
 [1] 0.66954
 > res$residual.scale
 [1] 0.66954
• Example 2:
 > x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x, weights = rep(1/n, n))
 > x0 <- c(1, 1.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 4.822705
 > new <- data.frame(x = 1.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > W <- diag(1/rep(1/n, n))</pre>
 > lower <- yhat - qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% solve(W) %*% X) %*% x0)
 > upper <- yhat + qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% solve(W) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 4.822705 1.454862 8.190548
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
 + interval = "prediction", level = 0.95)
 > res$fit
                 lwr
        fit
 1 4.822705 1.454862 8.190548
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
 + X) %*% x0))
 > se.fit
 [1] 1.202537
 > res$se.fit
 [1] 1.202537
 > s
 [1] 0.66954
 > res$residual.scale
```

```
[1] 0.66954
```

- **Note 1:** Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n 2 e scale = summary.lm(object = modello)\$sigma.
- **Note 2:** Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

predict()

- Package: stats
- Input:

```
object modello di regressione lineare pesata con una variabile esplicativa ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- Description: intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

fit $x_0^T \hat{\beta} = \frac{\text{"confidence"}}{x_0^T \hat{\beta}}$ $x_0^T \hat{\beta} + t_{1-\alpha/2, df} s^* \sqrt{x_0^T (X^T W^{-1} X)^{-1} x_0}$ interval = "prediction" $x_0^T \hat{\beta} = x_0^T \hat{\beta} + t_{1-\alpha/2, df} s^* \sqrt{1 + x_0^T (X^T W^{-1} X)^{-1} x_0}$ se.fit $s^* \sqrt{x_0^T (X^T W^{-1} X)^{-1} x_0}$ df df = n-2 residual.scale s^*

• Example 1:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> x0 <- c(1, 1.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))
> yhat
[1] 4.822705
```

```
> new <- data.frame(x = 1.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > W <- diag(1/rep(1/n, n))</pre>
 > lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
       solve(W) %*% X) %*% x0)
 > upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
       solve(W) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 4.822705 2.465776 7.179634
 > res <- predict(object = modello, newdata = new, se.fit = TRUE,
      scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
        fit
                 lwr
                           upr
 1 4.822705 2.465776 7.179634
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
 + X) % * % x 0))
 > se.fit
 [1] 1.202537
 > res$se.fit
 [1] 1.202537
 > s
 [1] 0.66954
 > res$residual.scale
 [1] 0.66954
• Example 2:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x, weights = rep(1/n, n))
 > x0 < -c(1, 1.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 4.822705
 > new <- data.frame(x = 1.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > W <- diag(1/rep(1/n, n))</pre>
 > lower <- yhat - qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% solve(W) %*% X) %*% x0)
 > upper <- yhat + qt(1 - 0.05/2, df = n - 2) * s * sqrt(1 + t(x0) %*%
       solve(t(X) %*% solve(W) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 4.822705 1.454862 8.190548
```

611

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n 2 e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

cov2cor()

• Package: stats

• Input:

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2$$

15.3 Adattamento

logLik()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** log-verosimiglianza normale
- Formula:

 $\hat{\ell}$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> logLik(object = modello)
'log Lik.' -15.30923 (df=3)
```

durbin.watson()

• Package: car

• Input:

 ${\tt model}$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / RSS$$

AIC()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** indice AIC
- Formula:

$$-2\hat{\ell} + 6$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> AIC(object = modello)
[1] 36.61846
```

extractAIC()

- Package: stats
- Input:

fit modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$2 \qquad n \log(RSS/n) + 4$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> extractAIC(fit = modello)

[1] 2.000000 -4.720086
```

deviance()

- Package: tt stats
- Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: devianza residua
- Formula:

RSS

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> deviance(object = modello)
[1] 2.689703
```

PRESS()

• Package: MPV

• Input:

imes modello di regressione lineare pesata con una variabile esplicativa ed n unità

• **Description:** PRESS

• Formula:

$$\sum_{i=1}^{n} e_i^2 / (1 - h_i)^2$$

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> PRESS(x = modello)
[1] 53.41271
```

anova()

• Package: stats

• Input:

object modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: anova di regressione
- Output:

Df gradi di libertà
Sum Sq devianze residue
Mean Sq quadrati medi
F value valore empirico della statistica F
Pr(>F) p-value

• Formula:

Df
$$1 n-2$$
 Sum Sq
$$RSS_{nullo}-RSS RSS$$
 Mean Sq
$$RSS_{nullo}-RSS RSS/(n-2)$$
 F value
$$F_{value}=\frac{RSS_{nullo}-RSS}{RSS/(n-2)}=t_{\hat{\beta}_2}^2$$
 Pr(>F)
$$P(F_{1,n-2}\geq F_{value})$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x, weights = rep(1/n, n))

> anova(object = modello)
```

```
Analysis of Variance Table

Response: y

Df Sum Sq Mean Sq F value Pr(>F)

x 1 3.2686 3.2686 7.2914 0.03556 *

Residuals 6 2.6897 0.4483

---

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

drop1()

• Package: stats

• Input:

object modello di regressione lineare pesata con una variabile esplicativa ed n unità scale selezione indice AIC oppure Cp test = "F"

• **Description:** submodels

• Output:

Df differenza tra gradi di libertà Sum of Sq differenza tra devianze residue RSS devianza residua AIC indice AIC Cp indice Cp F value valore empirico della statistica F Pr (F) p-value

• Formula:

Df $RSS_{nullo} - RSS$ $RSS_{nullo} - RSS$ RSS_{nullo} AIC $\frac{\text{scale} = 0}{n \log (RSS/n) + 4, n \log (RSS_{nullo}/n) + 2}$ Cp $\frac{\text{scale} = s^2}{2}$ $2, \frac{RSS_{nullo}}{RSS/(n-2)} + 2 - n$ F value $F_{value} = \frac{RSS_{nullo} - RSS}{RSS/(n-2)} = t_{\hat{\beta}_2}^2$ Pr(F)

 $P(F_{1, n-2} \ge F_{value})$

• Example 1:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x, weights = rep(1/n, n))
 > drop1(object = modello, scale = 0, test = "F")
 Single term deletions
 Model:
 y ~ x
        Df Sum of Sq
                        RSS AIC F value Pr(F)
                      2.6897 -4.7201
 <none>
              3.2686 5.9583 -0.3573 7.2914 0.03556 *
 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
 > res <- drop1(object = modello, scale = 0, test = "F")</pre>
 > res$Df
 [1] NA 1
 > res$"Sum of Sq"
         NA 3.268597
 [1]
 > res$RSS
 [1] 2.689703 5.958300
 > res$AIC
 [1] -4.7200862 -0.3572507
 > res$"F value"
 [1]
         NA 7.291356
 > res$"Pr(F)"
 [1]
            NA 0.03556412
• Example 2:
 > x < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x, weights = rep(1/n, n))
 > s <- summary.lm(object = modello)$sigma</pre>
 > drop1(object = modello, scale = s^2, test = "F")
 Single term deletions
 Model:
 y ~ x
 scale: 0.4482838
        Df Sum of Sq
                       RSS Cp F value
                                            Pr(F)
 <none>
                     2.6897 2.0000
              3.2686 5.9583 7.2914 7.2914 0.03556 *
 X
 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
```

```
> res <- drop1(object = modello, scale = s^2, test = "F")</pre>
    > res$Df
    [1] NA 1
    > res$"Sum of Sq"
    [1]
               NA 3.268597
    > res$RSS
    [1] 2.689703 5.958300
    > res$Cp
    [1] 2.000000 7.291356
    > res$"F value"
               NA 7.291356
    [1]
    > res$"Pr(F)"
    [1]
                NA 0.03556412
add1()
  • Package: stats
  • Input:
        object modello nullo di regressione lineare semplice pesata
        scope \, modello di regressione lineare pesata con una variabile esplicativa ed n unità
        scale selezione indice AIC oppure Cp
        test = "F"
  • Description: submodels
  • Output:
        Df differenza tra gradi di libertà
        Sum of Sq differenza tra devianze residue
        RSS devianza residua
        AIC indice AIC
        Cp indice Cp
        {\mathbb F} value valore empirico della statistica F
        Pr(F) p-value
  • Formula:
        Df
                                                      1
        Sum of Sq
                                               RSS_{nullo} - RSS
        RSS
                                                RSS_{nullo}, RSS
```

AIC

scale = 0

 $n \log (RSS_{nullo}/n) + 2$, $n \log (RSS/n) + 4$

Ср

scale =
$$s^2$$

$$\frac{RSS_{nullo}}{RSS/(n-2)} + 2 - n, 2$$

F value

$$F_{value} \, = \, \frac{RSS_{nullo} - RSS}{RSS \, / \, (n-2)} \, = \, t_{\hat{\beta}_2}^2 \label{eq:Fvalue}$$

Pr(F)

$$P(F_{1, n-2} \ge F_{value})$$

• Example 1:

```
> x \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> nullo <- lm(formula = y \sim 1, weights = rep(1/n, n))
> modello <- lm(formula = y \sim x, weights = rep(1/n, n))
> add1(object = nullo, scope = modello, scale = 0, test = "F")
Single term additions
Model:
y ~ 1
      Df Sum of Sq
                       RSS
                               AIC F value Pr(F)
                     5.9583 -0.3573
<none>
             3.2686 2.6897 -4.7201 7.2914 0.03556 *
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- add1(object = nullo, scope = modello, scale = 0, test = "F")
> res$Df
[1] NA 1
> res$"Sum of Sq"
        NA 3.268597
[1]
> res$RSS
[1] 5.958300 2.689703
> res$AIC
[1] -0.3572507 -4.7200862
> res$"F value"
        NA 7.291356
[1]
> res$"Pr(F)"
[1]
          NA 0.03556412
```

• Example 2:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> nullo <- lm(formula = y \sim 1, weights = rep(1/n, n))
> modello <- lm(formula = y \sim x, weights = rep(1/n, n))
> s <- summary.lm(object = modello)$sigma</pre>
> add1(object = nullo, scope = modello, scale = s^2, test = "F")
Single term additions
Model:
y ~ 1
scale: 0.4482838
    Df Sum of Sq RSS Cp F value
                    5.9583 7.2914
<none>
           3.2686 2.6897 2.0000 7.2914 0.03556 *
X
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- add1(object = nullo, scope = modello, scale = s^2, test = "F")</pre>
> res$Df
[1] NA 1
> res$"Sum of Sq"
[1] NA 3.268597
> res$RSS
[1] 5.958300 2.689703
> res$Cp
[1] 7.291356 2.000000
> res$"F value"
[1] NA 7.291356
> res$"Pr(F)"
[1]
          NA 0.03556412
```

15.4 Diagnostica

ls.diag()

• Package: stats

• Input:

ls.out modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: analisi di regressione lineare pesata
- Output:

```
std.dev stima di \sigma hat valori di leva std.res residui standard stud.res residui studentizzati cooks distanza di Cook dfits dfits correlation matrice di correlazione delle stime WLS std.err standard error delle stime WLS cov.scaled matrice di covarianza delle stime WLS non scalata per \sigma^2
```

• Formula:

std.dev
$$s$$
 hat
$$h_i \quad \forall i=1,2,\ldots,n$$
 std.res
$$rstandard_i \quad \forall i=1,2,\ldots,n$$
 stud.res
$$rstudent_i \quad \forall i=1,2,\ldots,n$$
 cooks
$$cd_i \quad \forall i=1,2,\ldots,n$$
 dfits
$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i=1,2,\ldots,n$$
 correlation
$$r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i,j=1,2$$
 std.err
$$s_{\hat{\beta}_j} \quad \forall j=1,2$$
 cov.scaled
$$s^2 \, (X^T \, W^{-1} \, X)^{-1}$$
 cov.unscaled

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> res <- ls.diag(ls.out = modello)
> res$std.dev
[1] 1.893745
```

> res\$hat [1] 0.4350043 0.2701267 0.1284350 0.1945578 0.4684951 0.1733040 0.1355195 [8] 0.1945578 > res\$std.res $\begin{bmatrix} 1 \end{bmatrix}$ -2.22897996 0.51181072 1.34601741 -0.04039112 -1.20017856 0.81532985 [7] 0.31550428 -0.15806803 > res\$stud.res [1] -4.90710471 0.47776268 1.47068630 -0.03687690 -1.25680777 0.78929887 [7] 0.29043398 -0.14459710 > res\$cooks [1] 1.9126289653 0.0484739848 0.1334918569 0.0001970407 0.6348329327 [6] 0.0696786009 0.0078023824 0.0030176734 > res\$dfits $\begin{smallmatrix} 1 \end{smallmatrix} \end{smallmatrix} -4.30575707 \quad 0.29065126 \quad 0.56456215 \quad -0.01812431 \quad -1.17996116 \quad 0.36138726$ [7] 0.11499284 -0.07106678 > res\$correlation (Intercept) 1.0000000 -0.8971215 (Intercept) -0.8971215 1.0000000 > res\$std.err [,1] (Intercept) 4.286587 0.784814 > res\$cov.scaled (Intercept) 18.374825 -3.0180723 (Intercept) -3.018072 0.6159331

> res\$cov.unscaled

(Intercept)

(Intercept)

5.1236582 -0.8415629

-0.8415629 0.1717475

cooks.distance()

• Package: stats

• Input:

modell modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

rstandard()

• Package: stats

• Input:

model modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

• Examples:

rstandard.lm()

• Package: stats

• Input:

model modello di regressione lineare pesata con una variabile esplicativa ed <math>n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

rstudent.lm()

- Package: stats
- Input:

modell modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, ..., n$$

• Examples:

lmwork()

- Package: MASS
- Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** diagnostica di regressione
- Output:

```
stdedv stima di \sigma stdres residui standard studres residui studentizzati
```

• Formula:

```
stdedv s stdres rstandard_i \  \, \forall i\,=\,1,\,2,\,\ldots,\,n
```

studres

$$rstudent_i \quad \forall i = 1, 2, ..., n$$

• Examples:

dffits()

• Package: stats

• Input:

model modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description**: dffits
- Formula:

$$rstudent_i \sqrt{\frac{h_i}{1 - h_i}} \quad \forall i = 1, 2, \dots, n$$

covratio()

• Package: stats

• Input:

model modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description**: covratio
- Formula:

$$cr_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

lm.influence()

• Package: stats

• Input:

 ${\tt model}$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** diagnostica di regressione
- Output:

```
hat valori di leva coefficients differenza tra le stime WLS eliminando una unità sigma stima di \sigma eliminando una unità wt.res residui pesati
```

• Formula:

```
hat h_i \quad \forall \, i=1,2,\ldots,n coefficients \hat{\beta}_j - \hat{\beta}_{j\,(-i)} = w_i\,e_i\,(1-h_i)^{-1}\,(X^T\,W^{-1}\,X)_j^{-1}\,X_i^T \quad \forall i=1,2,\ldots,n \quad \forall j=1,2 sigma s_{-i} \quad \forall \, i=1,2,\ldots,n wt.res \sqrt{w_i}\,e_i \quad \forall \, i=1,2,\ldots,n
```

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> res <- lm.influence(model = modello)
> res$hat
```

weights()

• Package: stats

0.19640809 -0.09498123

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- **Description:** pesi
- Formula:

$$w_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> weights(object = modello)
[1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125
```

weighted.residuals()

• Package: stats

• Input:

obj modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: residui pesati
- Formula:

$$\sqrt{w_i} e_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

residuals.lm()

• Package: stats

• Input:

object modello di regressione lineare pesata con una variabile esplicativa ed n unità type = "response" / "pearson" tipo di residuo

- **Description:** residui
- Formula:

type = "response"
$$e_i \quad \forall i = 1, 2, \dots, n$$

$$type = "pearson"
$$\sqrt{w_i} \, e_i \quad \forall i = 1, 2, \dots, n$$$$

df.residual()

• Package: stats

• Input:

object $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-2

• Examples:

```
> x <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x, weights = rep(1/n, n))
> df.residual(object = modello)
[1] 6
```

hatvalues()

• Package: stats

• Input:

modell modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

dfbeta()

- Package: stats
- Input:

formula $\,$ modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: dfbeta
- Formula:

$$\hat{\beta}_j - \hat{\beta}_{j(-i)} = w_i e_i (1 - h_i)^{-1} (X^T W^{-1} X)_j^{-1} X_i^T \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2$$

dfbetas()

• Package: stats

• Input:

formula modello di regressione lineare pesata con una variabile esplicativa ed n unità

• **Description:** dfbetas

• Formula:

$$\frac{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}{s_{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}} = \frac{w_{i} e_{i} (1 - h_{i})^{-1} (X^{T} W^{-1} X)_{j}^{-1} X_{i}^{T}}{s_{-i} \sqrt{(X^{T} W^{-1} X)_{j,j}^{-1}}} \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2$$

• Examples:

outlier.test()

• Package: car

• Input:

model modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: test sugli outliers
- Output:

test massimo residuo studentizzato assoluto, gradi di libertà, p-value

• Formula:

```
test t = \max_i ( \, | \, rstudent_i \, | ) \quad n-3 \quad p\text{-value} = 2 \, P( \, t_{n-3} \, \leq \, -| \, t \, | ) \qquad \forall \, i \, = \, 1, \, 2, \, \ldots \, , n
```

• Examples:

influence.measures()

• Package: stats

• Input:

modell modello di regressione lineare pesata con una variabile esplicativa ed n unità

- Description: dfbetas, dffits, covratio, distanza di Cook, valori di leva
- Output:

infmat misure di influenza di dimensione $n\times 6$ is .inf matrice di influenza con valori logici di dimensione $n\times 6$

• Formula:

$$DFBETAS_{ij} = \frac{w_i e_i (1-h_i)^{-1} (X^T W^{-1} X)_j^{-1} X_i^T}{s_{-i} \sqrt{(X^T W^{-1} X)_j^{-1}}} \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2$$

$$DFFITS_i = rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i = 1, 2, ..., n$$

$$COVRATIO_i = (1-h_i)^{-1} \left(1 + \frac{rstudent_i^2 - 1}{n-2}\right)^{-2} \quad \forall i = 1, 2, ..., n$$

$$COOKD_i = \frac{h_i rstandard_i^2}{2(1-h_i)} \quad \forall i = 1, 2, ..., n$$

$$HAT_i = h_i \quad \forall i = 1, 2, ..., n$$

```
2 0.278471258 -0.21304046 0.29065126 1.80443448 0.0484739848 0.2701267  
3 0.328885485 -0.09232735 0.56456215 0.80504974 0.1334918569 0.1284350  
4 0.003304089 -0.01083702 -0.01812431 1.78686556 0.0001970407 0.1945578  
5 0.637149075 -1.01035839 -1.17996116 1.56459066 0.6348329327 0.4684951  
6 0.306755388 -0.19079196 0.36138726 1.37727804 0.0696786009 0.1733040  
7 0.020048284 0.03203820 0.11499284 1.61092794 0.0078023824 0.1355195  
8 0.012955584 -0.04249278 -0.07106678 1.77297867 0.0030176734 0.1945578
```

> res\$is.inf

```
dfb.1_ dfb.x dffit cov.r cook.d hat

TRUE TRUE TRUE FALSE TRUE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE

FALSE TRUE FALSE FALSE FALSE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE

FALSE FALSE FALSE FALSE FALSE FALSE
```

- Note 1: Il caso *i*-esimo è influente se $|DFBETAS_{ij}| > 1 \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2$
- Note 2: Il caso *i*-esimo è influente se $|DFFITS_i| > 3\sqrt{2/(n-2)}$ $\forall i = 1, 2, ..., n$
- Note 3: Il caso *i*-esimo è influente se $|1 COVRATIO_i| > 6 / (n-2)$ $\forall i = 1, 2, ..., n$
- Note 4: Il caso *i*-esimo è influente se $P(F_{2,n-2} \ge COOKD_i) > 0.5 \quad \forall i = 1, 2, ..., n$
- Note 5: Il caso *i*-esimo è influente se $HAT_i > 6 / n \quad \forall i = 1, 2, ..., n$
- **Note 6:** I casi influenti rispetto ad almeno una tra queste misure sono marcati con un asterisco. Corrispondentemente la stessa riga della matrice is.inf riporterà almeno un simbolo TRUE.

Capitolo 16

Regressione lineare multipla pesata

16.1 Simbologia

$$y_i = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} + \varepsilon_i \quad \forall i = 1, 2, \dots, n \qquad \varepsilon \sim N(0, \sigma^2 W)$$

- variabile dipendente: y
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi WLS: $w=(w_1,\,w_2,\,\ldots,\,w_n)$
- matrice diagonale definita positiva di dimensione $n \times n$: $W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- matrice identità di dimensione $n \times n$: I_n
- devianza residua: $RSS = \sum_{i=1}^n w_i \, e_i^2 = y^T \, W^{-1} \, e = y^T \, W^{-1} \, (I_n H) \, y$
- stima di σ^2 : $s^2 = RSS/(n-k)$
- gradi di libertà della devianza residua: n-k
- stima di σ^2 tolta la i-esima unità: $s_{-i}^2 = s^2 \left(1 + \frac{1 rstandard_i^2}{n k 1}\right) = s^2 \left(1 + \frac{rstudent_i^2 1}{n k}\right)^{-1} \quad \forall i = 1, 2, \ldots, n$
- stime WLS: $\hat{\beta} = (X^T W^{-1} X)^{-1} X^T W^{-1} y$
- standard error delle stime WLS: $s_{\hat{\beta}} = s \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- t-values delle stime WLS: $t_{\hat{eta}} = \hat{eta} / s_{\hat{eta}}$
- residui: $e = (I_n H) y$
- residui pesati: $\sqrt{w_i} e_i \quad \forall i = 1, 2, ..., n$
- residui standard: $rstandard_i = \frac{e_i}{s\sqrt{(1-h_i)/w_i}} \quad \forall i=1,2,\ldots,n$
- residui studentizzati: $rstudent_i = \frac{e_i}{s_{-i}\sqrt{(1-h_i)/w_i}} = rstandard_i\sqrt{\frac{n-k-1}{n-k-rstandard_i^2}} \quad \forall i=1,2,\ldots,n$
- valori adattati: $\hat{y} = Hy$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, ..., n$
- stime WLS tolta la *i*-esima unità: $\hat{\beta}_{(-i)} \quad \forall i = 1, 2, ..., n$
- correlazione delle stime WLS: $r_{\hat{\beta}_i\,\hat{\beta}_j}=rac{s^2\,(X^T\,W^{-1}\,X)_{i,\,j}^{-1}}{s_{\hat{\beta}_i}\,s_{\hat{\beta}_j}}\quad \forall\,i,j\,=\,1,\,2,\,\ldots,\,k$
- devianza residua modello nullo: $RSS_{nullo} = \sum_{i=1}^n w_i (y_i \bar{y}_W)^2 = (y \bar{y}_W)^T W^{-1} (y \bar{y}_W)$
- indice di determinazione: $R^2=1-RSS/RSS_{nullo}=1-\left(1-R_{adj}^2\right)\left(n-k\right)/\left(n-1\right)$

- indice di determinazione aggiustato: $R_{adj}^{2}=1-\frac{RSS/(n-k)}{RSS_{nullo}/(n-1)}=1-\left(1-R^{2}\right)\left(n-1\right)/\left(n-k\right)$
- valore noto dei regressori per la previsione: $x_0^T = (1, x_{01}, x_{02}, \dots, x_{0k-1})$
- log-verosimiglianza normale: $\hat{\ell} = -n \left(\log(2\pi) + \log\left(RSS/n\right) + 1 \sum_{i=1}^{n} \log(w_i)/n\right)/2$
- distanza di Cook: $cd_i=rac{h_i\,rstandard_i^2}{k\,(1-h_i)}=rac{e_i^2}{k\,s^2}\,rac{h_i}{(1-h_i)^2}\quad \forall\,i\,=\,1,\,2,\,\ldots,\,n$
- covratio: $cr_i = (1 h_i)^{-1} \left(1 + \frac{rstudent_i^2 1}{n k} \right)^{-k} = (1 h_i)^{-1} \left(\frac{s_{-i}}{s} \right)^{2k} \quad \forall i = 1, 2, \dots, n$

16.2 Stima

lm()

- Package: stats
- Input:

formula modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

weights pesi

x = TRUE matrice del modello

y = TRUE variabile dipendente

- Description: analisi di regressione lineare pesata
- Output:

coefficients stime WLS

residuals **residui**

fitted.values valori adattati

weights pesi

rank rango della matrice del modello

df.residual gradi di libertà della devianza residua

- x matrice del modello
- y variabile dipendente

• Formula:

coefficients

 $\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$

residuals

 $e_i \quad \forall i = 1, 2, \ldots, n$

fitted.values

 $\hat{y}_i \quad \forall i = 1, 2, \dots, n$

weights

 $w_i \quad \forall i = 1, 2, \ldots, n$

rank

k

df.residual

n-k

Х

X

У

y

```
> k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
 + n), x = TRUE, y = TRUE)
 > modello$coefficients
  (Intercept)
                      x1
                                   x2
  > modello$residuals
         1
                   2
                              3
                                        4
 -0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
 -1.4356227
 > modello$fitted.values
                                4
  2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
 > modello$weights
 [1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125
 > modello$rank
 [1] 4
 > modello$df.residual
 [1] 4
 > modello$x
   (Intercept) x1 x2 x3
            1 1.1 1.2 1.40
            1 2.3 3.4 5.60
 3
            1 4.5 5.6 7.56
            1 6.7 7.5 6.00
 5
            1 8.9 7.5 5.40
 6
            1 3.4 6.7 6.60
 7
            1 5.6 8.6 8.70
            1 6.7 7.6 8.70
 attr(, "assign")
 [1] 0 1 2 3
 > modello$y
                 4
                      5
 1.50 6.40 9.60 8.80 8.86 7.80 8.60 8.60
• Note 1: Il modello nullo si ottiene con lm(formula = y ~ 1, weights = w).
```

• Note 2: L'istruzione update (object = $y \sim x1 + x2$, formula = . $\sim . + x3$) è esattamente equi-

valente a $lm(formula = y \sim x1 + x2 + x3, weights = w)$.

- Note 3: In seguito ad una modifica come ad esempio x1[3] <- 1.2, conviene adoperare il comando update (modello) anziché ripetere modello <- lm(formula = y ~ x1 + x2 + x3, weights = w).
- **Note 4:** L'operatore I () permette di poter modellare regressioni lineari polinomiali. Per un polinomio di terzo grado occorre scrivere $lm(formula = y \sim x + I(x^2) + I(x^3), weights = w)$.
- **Note 5:** Per regressioni polinomiali occorre usare il comando poly(). Per un polinomio di quarto grado occorre scrivere lm(formula = y ~ poly(x, degree = 4, raw = TRUE), weights = w).
- **Note 6:** Per regressioni polinomiali ortogonali occorre usare il comando poly(). Per un polinomio ortogonale di quarto grado occorre scrivere lm(formula = y ~ poly(x,degree = 4),weights = w).
- Note 7: Il comando uzione lm(formula = y ~ x1 + x2 + x3, weights=w) è esattamente equivalente a lm(formula = y ~ X-1, weights = w).
- Note 8: Il comando $lm(formula = y \sim x1 + x2 + x3, weights = w)$ è esattamente equivalente a $lm(formula = y \sim 1 + x1 + x2 + x3, weights = w)$.

summary.lm()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime WLS

- Description: analisi di regressione lineare pesata
- Output:

residuals residui

coefficients stima puntuale, standard error, t-value, p-value

sigma \sin a di σ

r.squared indice di determinazione

adj.r.squared indice di determinazione aggiustato

 $\verb|fstatistic| valore empirico della statistica| \textit{F, df numeratore, df denominatore}|$

cov.unscaled matrice di covarianza delle stime WLS non scalata per σ^2

correlation matrice di correlazione delle stime WLS

• Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

coefficients

$$\hat{\beta}_j$$
 $s_{\hat{\beta}_j}$ $t_{\hat{\beta}_j}$ p -value = $2P(t_{n-k} \le -|t_{\hat{\beta}_j}|)$ $\forall j = 1, 2, ..., k$

sigma

s

r.squared

 R^2

adj.r.squared

 R^2 .

fstatistic

$$Fvalue = \frac{\left(RSS_{nullo} - RSS\right) / (k-1)}{RSS / (n-k)} \qquad k-1 \qquad n-k$$

cov.unscaled

$$(X^T W^{-1} X)^{-1}$$

correlation

> res\$correlation

x1

x2

x3

(Intercept)

x1

-0.18600997 1.0000000 -0.82213982 0.4883764

0.07158062 -0.8221398 1.00000000 -0.8022181

-0.46329002 0.4883764 -0.80221810 1.0000000

(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900

```
r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k
```

```
• Examples:
```

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> res <- summary.lm(object = modello, correlation = TRUE)</pre>
> res$residuals
                                                                    2
                                                                                                        3
                                                                                                                                             4
                                                                                                                                                                                  5
                                                                                                                                                                                                                         6
-0.3371620 \quad 0.1540936 \quad 0.4619923 \quad 0.2465971 \quad 0.0910624 \quad 0.2336206 \quad -0.3426347 \quad 0.0910624 \quad 0.
-0.5075693
> res$coefficients
                                                      Estimate Std. Error t value Pr(>|t|)
 (Intercept) 0.988514333 1.4292308 0.691640822 0.5272118
                                           0.422516384 0.3883267 1.088043731 0.3377443
x1
x2
                                        -0.001737381 0.5822146 -0.002984091 0.9977619
x3
                                        0.716029046 0.4068987 1.759723294 0.1532663
> res$sigma
[1] 0.4608596
> res$r.squared
[1] 0.8574147
> res$adj.r.squared
[1] 0.7504757
> res$fstatistic
          value numdf
                                                                  dendf
8.017793 3.000000 4.000000
> res$cov.unscaled
                                                                                  x1
                                         (Intercept)
                                                                                                                                                x2
 (Intercept) 9.6176174 -0.4860697 0.2804424 -1.2685405
x1
                                           -0.4860697 0.7099981 -0.8751626 0.3633297
x2
                                            0.2804424 -0.8751626 1.5959854 -0.8947971
                                           -1.2685405 0.3633297 -0.8947971 0.7795344
x3
```

vcov()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime WLS
- Formula:

$$s^2 (X^T W^{-1} X)^{-1}$$

• Examples:

```
> k <- 4
> x1 < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
     n))
> vcov(object = modello)
            (Intercept)
                                 x1
                                              x2
                                                          x3
(Intercept) 2.04270054 -0.10323710 0.05956359 -0.26942727
            -0.10323710 0.15079759 -0.18587712 0.07716815
x1
             0.05956359 - 0.18587712 \quad 0.33897378 - 0.19004733
x2
x3
            -0.26942727 0.07716815 -0.19004733 0.16556652
```

lm.wfit()

• Package: stats

• Input:

- x matrice del modello
- y variabile dipendente
- w pesi
- Description: analisi di regressione lineare pesata
- Output:

```
coefficients stime WLS
residuals residui
fitted.values valori adattati
weights pesi
rank rango della matrice del modello
df.residual gradi di libertà della devianza residua
```

• Formula:

coefficients
$$\hat{\beta_j} \quad \forall j=1,2,\dots,k$$
 residuals
$$e_i \quad \forall i=1,2,\dots,n$$
 fitted.values
$$\hat{y_i} \quad \forall i=1,2,\dots,n$$
 weights
$$w_i \quad \forall i=1,2,\dots,n$$

```
rank
                                                     k
        df.residual
                                                    n-k
  • Examples:
    > k <- 4
    > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
    > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
    > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
    > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
    > n < - 8
    > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
         n))
    > X <- model.matrix(object = modello)</pre>
    > res <- lm.wfit(x = X, y, w = rep(1/n, n))
    > res$coefficients
     (Intercept)
                              x1
                                             x2
                                                             x3
     0.988514333 \quad 0.422516384 \quad -0.001737381 \quad 0.716029046
    > res$residuals
     \begin{smallmatrix} 1 \end{smallmatrix} \end{bmatrix} - 0.9536382 \quad 0.4358424 \quad 1.3067117 \quad 0.6974820 \quad 0.2575634 \quad 0.6607787 \quad -0.9691173 
    [8] -1.4356227
    > res$fitted.values
    [1] 2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117
    [8] 10.035623
    > res$weights
    [1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125
    > res$rank
    [1] 4
    > res$df.residual
    [1] 4
lsfit()
  • Package: stats
  • Input:
        x matrice del modello
        y variabile dipendente
        wt pesi
        intercept = FALSE
```

coefficients stime WLS

• Output:

• **Description:** analisi di regressione lineare pesata

```
residuals residui
wt pesi
```

• Formula:

```
coefficients \hat{\beta_j} \quad \forall \, j \, = \, 1, \, 2, \, \ldots, \, k residuals e_i \quad \forall \, i \, = \, 1, \, 2, \, \ldots, \, n wt w_i \quad \forall \, i \, = \, 1, \, 2, \, \ldots, \, n
```

• Examples:

```
> k < - 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> X <- model.matrix(object = modello)</pre>
> res <- lsfit(x = X, y, wt = rep(1/n, n), intercept = FALSE)
> res$coefficients
 (Intercept)
                          x1
                                         x2
 0.988514333 0.422516384 -0.001737381 0.716029046
> res$residuals
 \begin{smallmatrix} 1 \end{smallmatrix} \rbrack -0.9536382 \quad 0.4358424 \quad 1.3067117 \quad 0.6974820 \quad 0.2575634 \quad 0.6607787 \quad -0.9691173 
[8] -1.4356227
> res$wt
[1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125
```

confint()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- Description: intervallo di confidenza per le stime WLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-k} s_{\hat{\beta}_j} \quad \forall j = 1, 2, ..., k$$

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))

> confint(object = modello, parm = c(1, 2, 3), level = 0.95)
```

```
2.5 % 97.5 % (Intercept) -2.9796664 4.956695 x1 -0.6556513 1.500684 x2 -1.6182241 1.614749
```

Confint()

• Package: Rcmdr

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità parm parametri del modello su cui calcolare l'intervallo di confidenza level livello di confidenza $1-\alpha$

- **Description:** intervallo di confidenza per le stime WLS
- Formula:

$$\hat{\beta}_j \mp t_{1-\alpha/2, n-k} s_{\hat{\beta}_j} \quad \forall j = 1, 2, ..., k$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
     n))
> Confint (object = modello, parm = c(1, 2, 3), level = 0.95)
                 2.5 %
                         97.5 %
(Intercept) -2.9796664 4.956695
           -0.6556513 1.500684
x1
           -1.6182241 1.614749
x2
```

coef()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** stime WLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** stime WLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

coeftest()

• Package: lmtest

• Input:

imes modello di regressione lineare pesata con k-1 variabili esplicative ed n unità df = NULL / Inf significatività delle stime effettuata con la variabile casuale t oppure Z

- **Description:** stima puntuale, standard error, t-value, p-value
- Formula:

• Example 1:

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))

> coeftest(x = modello, df = NULL)
```

```
t test of coefficients:
```

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.9885143 1.4292308 0.6916 0.5272
x1 0.4225164 0.3883267 1.0880 0.3377
x2 -0.0017374 0.5822146 -0.0030 0.9978
x3 0.7160290 0.4068987 1.7597 0.1533
```

• Example 2:

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> coeftest(x = modello, df = Inf)
z test of coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.9885143
                       1.4292308 0.6916 0.48916
x1
            0.4225164
                       0.3883267 1.0880 0.27658
x2
            -0.0017374 0.5822146 -0.0030 0.99762
x3
             0.7160290 0.4068987 1.7597 0.07845 .
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

• Note: Naturalmente vale che $t_{\hat{\beta}_i} = z_{\hat{\beta}_i} \quad \forall j = 1, 2, ..., k$.

fitted()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \dots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> fitted(object = modello)
1 2 3 4 5 6 7 8
2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
```

fitted.values()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{y}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> fitted.values(object = modello)
1 2 3 4 5 6 7 8
2.453638 5.964158 8.293288 8.102518 8.602437 7.139221 9.569117 10.035623
```

predict.lm()

• Package: stats

• Input:

```
object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- Description: intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

fit

[1] 0.4608596

```
se.fit
                                     s^* \sqrt{x_0^T (X^T W^{-1} X)^{-1} x_0}
     df
                                           df = n - k
     residual.scale
                                               s^*
• Example 1:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
 + n))
 > x0 \leftarrow c(1, 1.3, 2.1, 2.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 3.181004
 > new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > W <- diag(1/rep(1/n, n))</pre>
 > lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
       solve(W) %*% X) %*% x0)
 > upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
 + solve(W) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 3.181004 1.200204 5.161803
 > res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
 + scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
                 lwr
                          upr
 1 3.181004 1.200204 5.161803
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
 + X) %*% x0))
 > se.fit
 [1] 1.010631
 > res$se.fit
 [1] 1.010631
 > s
 [1] 0.4608596
 > res$residual.scale
```

• Example 2:

```
> k < - 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
+ n))
> x0 <- c(1, 1.3, 2.1, 2.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 3.181004
> new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
> s <- summary.lm(object = modello)$sigma</pre>
> X <- model.matrix(object = modello)
> W <- diag(1/rep(1/n, n))
> lower <- yhat - qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
     solve(t(X) %*% solve(W) %*% X) %*% x0)
> upper <- yhat + qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%
     solve(t(X) %*% solve(W) %*% X) %*% x0)
> c(yhat, lower, upper)
[1] 3.18100394 0.09706736 6.26494051
> res <- predict.lm(object = modello, newdata = new, se.fit = TRUE,
+ interval = "prediction", level = 0.95)
> res$fit
       fit
                  lwr
                         upr
1 3.181004 0.09706736 6.26494
> se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
+ X) %*% x0))
> se.fit
[1] 1.010631
> res$se.fit
[1] 1.010631
> s
[1] 0.4608596
> res$residual.scale
[1] 0.4608596
```

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n k e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

predict()

• Package: stats

• Input:

```
object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime scale stima s^* di \sigma df il valore df dei gradi di libertà interval = "confidence" / "prediction" intervallo di confidenza o previsione level livello di confidenza 1-\alpha
```

- **Description:** intervallo di confidenza o di previsione
- Output:

```
fit valore previsto ed intervallo di confidenza se.fit standard error delle stime df il valore df dei gradi di libertà residual.scale stima s^* di \sigma
```

• Formula:

```
\inf x_0^T \hat{\beta} \qquad \inf x_0^T \hat{\beta} = \operatorname{"confidence"} x_0^T \hat{\beta} = t_{1-\alpha/2,df} \, s^* \, \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} \inf x_0^T \hat{\beta} \qquad x_0^T \, \hat{\beta} \mp t_{1-\alpha/2,df} \, s^* \, \sqrt{1 + x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} se.fit s^* \, \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0} df df = n - k residual.scale
```

• Example 1:

```
> k < -4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
     n))
> x0 <- c(1, 1.3, 2.1, 2.3)
> yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
> yhat
[1] 3.181004
> new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
> s <- summary.lm(object = modello)$sigma
> X <- model.matrix(object = modello)</pre>
> W <- diag(1/rep(1/n, n))</pre>
> lower <- yhat - qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
     solve(W) %*% X) %*% x0)
> upper <- yhat + qnorm(1 - 0.05/2) * s * sqrt(t(x0) %*% solve(t(X) %*%
      solve(W) %*% X) %*% x0)
> c(yhat, lower, upper)
```

```
[1] 3.181004 1.200204 5.161803
 > res <- predict(object = modello, newdata = new, se.fit = TRUE,
       scale = s, df = Inf, interval = "confidence", level = 0.95)
 > res$fit
        fit
                  lwr
 1 3.181004 1.200204 5.161803
 > se.fit <- as.numeric(s * sqrt(t(x0) %*% solve(t(X) %*% solve(W) %*%
 + X) %*% x0))
 > se.fit
 [1] 1.010631
 > res$se.fit
 [1] 1.010631
 > s
 [1] 0.4608596
 > res$residual.scale
 [1] 0.4608596
• Example 2:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
 + n))
 > x0 <- c(1, 1.3, 2.1, 2.3)
 > yhat <- as.numeric(t(x0) %*% coef(object = modello))</pre>
 > yhat
 [1] 3.181004
 > new <- data.frame(x1 = 1.3, x2 = 2.1, x3 = 2.3)
 > s <- summary.lm(object = modello)$sigma</pre>
 > X <- model.matrix(object = modello)</pre>
 > W <- diag(1/rep(1/n, n))</pre>
 > lower <- yhat - qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0) %*%)
      solve(t(X) %*% solve(W) %*% X) %*% x0)
 > upper <- yhat + qt(1 - 0.05/2, df = n - k) * s * sqrt(1 + t(x0)) %*%
       solve(t(X) %*% solve(W) %*% X) %*% x0)
 > c(yhat, lower, upper)
 [1] 3.18100394 0.09706736 6.26494051
 > res <- predict(object = modello, newdata = new, se.fit = TRUE,
 + interval = "prediction", level = 0.95)
 > res$fit
```

- Note 1: Per il calcolo dell'intervallo classico di confidenza o previsione impostare i parametri df = n k e scale = summary.lm(object = modello)\$sigma.
- Note 2: Per il calcolo dell'intervallo asintotico di confidenza o previsione impostare i parametri df = Inf e scale = summary.lm(object = modello)\$sigma.

linear.hypothesis()

- Package: car
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità hypothesis.matrix matrice C di dimensione $q \times k$ e rango pari a $q = \min(q, k)$ rhs vettore b della previsione lineare di dimensione q

• **Description:** test di ipotesi per $H_0: C\beta = b$ contro $H_1: C\beta \neq b$ dove C e b sono così definiti:

$$C = \begin{pmatrix} c_{1,1} & c_{1,2} & \dots & c_{1,k} \\ c_{2,1} & c_{2,2} & \dots & c_{2,k} \\ \vdots & \vdots & \vdots & \vdots \\ c_{q,1} & c_{q,2} & \dots & c_{q,k} \end{pmatrix} \quad b = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_q \end{pmatrix}$$

• Output:

Res.Df gradi di libertà della devianza residua

RSS devianza residua

 ${\tt Df}\;$ gradi di libertà della devianza relativa all'ipotesi nulla H_0

Sum of Sq devianza relativa all'ipotesi nulla \mathcal{H}_0

 \mathbb{F} valore empirico della statistica F

Pr(>F) p-value

• Formula:

Res.Df

$$n-k$$
 $n-k+q$

```
RSS \qquad RSS + \left(b - C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1} \left(b - C\,\hat{\beta}\right) Df -q Sum of Sq -\left(b - C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1} \left(b - C\,\hat{\beta}\right) F Fvalue = \frac{\left[\left(b - C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1} \left(b - C\,\hat{\beta}\right)\right]/q}{RSS/(n-k)} Pr (>F) P(F_{q,\,n-k} \geq Fvalue)
```

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
+ n))
> W <- diag(1/rep(1/n, n))</pre>
> C < -matrix(c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3), nrow = 2, ncol = 4,
     byrow = TRUE)
> C
    [,1] [,2] [,3] [,4]
      1 3 5.0 2.3
[1,]
      2 4 1.1 4.3
[2,]
> b < -c(1.1, 2.3)
> b
[1] 1.1 2.3
> q < -2
> c(n - k, n - k + q)
[1] 4 6
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b) $Res.Df
[1] 4 6
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(weighted.residuals(obj = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*%
     solve(W) %*% X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
[1] 0.8495662 2.2459829
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b)$RSS
[1] 0.8495662 2.2459829
```

```
> -q
[1] -2
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b) $Df
[1] NA -2
> -CSS
[1] -1.396417
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b) $"Sum of Sq"
           NA -1.396417
[1]
> Fvalue <- (CSS/q)/(RSS/(n - k))
> Fvalue
[1] 3.287364
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b) $F
          NA 3.287364
[1]
> 1 - pf(Fvalue, df1 = q, df2 = n - k)
[1] 0.1430808
> linear.hypothesis(model = modello, hypothesis.matrix = C, rhs = b)$"Pr(>F)"
          NA 0.1430808
[1]
```

lht()

- Package: car
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità hypothesis.matrix matrice C di dimensione $q \times k$ e rango pari a $q = \min(q, k)$ rhs vettore b della previsione lineare di dimensione q

• **Description:** test di ipotesi per $H_0: C\beta = b$ contro $H_1: C\beta \neq b$ dove C e b sono così definiti:

$$C = \begin{pmatrix} c_{1,1} & c_{1,2} & \dots & c_{1,k} \\ c_{2,1} & c_{2,2} & \dots & c_{2,k} \\ \vdots & \vdots & \vdots & \vdots \\ c_{q,1} & c_{q,2} & \dots & c_{q,k} \end{pmatrix} \quad b = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_q \end{pmatrix}$$

• Output:

Res.Df gradi di libertà della devianza residua RSS devianza residua Df gradi di libertà della devianza relativa all'ipotesi nulla H_0 Sum of Sq devianza relativa all'ipotesi nulla H_0 F valore empirico della statistica F

Pr(>F) p-value

• Formula:

RSS
$$RSS + \left(b - C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b - C\,\hat{\beta}\right)$$
 Df
$$-q$$
 Sum of Sq
$$-\left(b - C\,\hat{\beta}\right)^T \left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b - C\,\hat{\beta}\right)$$
 F
$$Fvalue = \frac{\left[\left(b - C\,\hat{\beta}\right)^T\left[C\,\left(X^T\,W^{-1}\,X\right)^{-1}\,C^T\right]^{-1}\,\left(b - C\,\hat{\beta}\right)\right]/q}{RSS/\left(n - k\right)}$$
 Pr (>F)
$$P(F_{a,\,n-k} \geq Fvalue)$$

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> W <- diag(1/rep(1/n, n))</pre>
> C <- matrix(c(1, 3, 5, 2.3, 2, 4, 1.1, 4.3), nrow = 2, ncol = 4,
      byrow = TRUE)
    [,1] [,2] [,3] [,4]
     1 3 5.0 2.3
2 4 1.1 4.3
[1,]
[2,]
> b <- c(1.1, 2.3)
> b
[1] 1.1 2.3
> q < -2
> c(n - k, n - k + q)
[1] 4 6
> lht(model = modello, hypothesis.matrix = C, rhs = b)$Res.Df
[1] 4 6
> X <- model.matrix(object = modello)</pre>
> RSS <- sum(weighted.residuals(obj = modello)^2)</pre>
> beta <- coefficients(object = modello)</pre>
> CSS <- as.numeric(t(b - C %*% beta) %*% solve(C %*% solve(t(X) %*%
+ solve(W) %*% X) %*% t(C)) %*% (b - C %*% beta))
> c(RSS, RSS + CSS)
```

```
[1] 0.8495662 2.2459829
> lht(model = modello, hypothesis.matrix = C, rhs = b)$RSS
[1] 0.8495662 2.2459829
> -q
[1] -2
> lht(model = modello, hypothesis.matrix = C, rhs = b)$Df
[1] NA -2
> -CSS
[1] -1.396417
> 1ht(model = modello, hypothesis.matrix = C, rhs = b) $"Sum of Sq"
[1]
         NA -1.396417
> Fvalue <- (CSS/q)/(RSS/(n - k))
> Fvalue
[1] 3.287364
> lht(model = modello, hypothesis.matrix = C, rhs = b)$F
        NA 3.287364
[1]
> 1 - pf(Fvalue, df1 = q, df2 = n - k)
[1] 0.1430808
> lht(model = modello, hypothesis.matrix = C, rhs = b) $"Pr(>F)"
[1]
         NA 0.1430808
```

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime WLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i, j \, = \, 1, \, 2, \, \ldots, \, k$$

```
> k < -4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
     n))
> V <- vcov(object = modello)
> cov2cor(V)
            (Intercept)
                                 x1
                                             x2
(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900
            -0.18600997 1.0000000 -0.82213982 0.4883764
x1
            0.07158062 -0.8221398 1.00000000 -0.8022181
x2
            -0.46329002 0.4883764 -0.80221810 1.0000000
x3
```

16.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** log-verosimiglianza normale
- Formula:

 $\hat{\ell}$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> logLik(object = modello)
'log Lik.' -10.69939 (df=5)
```

durbin.watson()

- Package: car
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / RSS$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> durbin.watson(model = modello)$dw
[1] 0.9255503
```

AIC()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

$$-2\,\hat{\ell} + 2\,(k+1)$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> AIC(object = modello)
[1] 31.39878
```

BIC()

• Package: nlme

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** indice *BIC*
- Formula:

$$-2\,\hat{\ell} + (k+1)\,\log(n)$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> BIC(object = modello)
[1] 31.79599
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k \qquad n \log(RSS/n) + 2k$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> extractAIC(fit = modello)
[1] 4.000000 -9.939768
```

deviance()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

RSS

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> deviance(object = modello)
[1] 0.8495662
```

PRESS()

• Package: MPV

• Input:

x modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

• Description: PRESS

• Formula:

$$\sum_{i=1}^{n} e_i^2 / (1 - h_i)^2$$

• Examples:

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n, n))

> PRESS(x = modello)
```

drop1()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità scale selezione indice AIC oppure Cp test = "F"

• **Description:** submodels

• Output:

Df differenza tra gradi di libertà Sum of Sq differenza tra devianze residue RSS devianza residua AIC indice AIC Cp indice Cp F value valore empirico della statistica F Pr (F) p-value

• Formula:

Df $\underbrace{1,\,1,\,\ldots,\,1}_{k-1\,\mathrm{volte}}$ Sum of Sq $RSS_{-x_j}-RSS \ \ \forall\,j\,=\,1,\,2,\,\ldots,\,k-1$

dove RSS_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

 $RSS, RSS_{-x_{j}} \quad \forall j = 1, 2, ..., k-1$

```
AIC
                                              scale = 0
                   n \log (RSS/n) + 2k, n \log (RSS_{-x_i}/n) + 2(k-1) \quad \forall j = 1, 2, ..., k-1
      Ср
                                             scale = s^2
                            k, \frac{RSS_{-x_j}}{RSS/(n-k)} + 2(k-1) - n \quad \forall j = 1, 2, ..., k-1
     F value
                                F_j = \frac{RSS_{-x_j} - RSS}{RSS/(n-k)} \quad \forall j = 1, 2, ..., k-1
     Pr(F)
                                   P(F_{1,n-k} \ge F_i) \quad \forall j = 1, 2, ..., k-1
• Example 1:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
 > drop1(object = modello, scale = 0, test = "F")
 Single term deletions
 Model:
 y \sim x1 + x2 + x3
                          RSS AIC 0.8496 -9.9398
         Df Sum of Sq
                                       AIC
                                              F value Pr(F)
 <none>
                                                1.1838 0.3377
               0.2514
                         1.1010 -9.8658
 x1
          1
 x2
         1 1.891e-06
                          0.8496 -11.9398 8.905e-06 0.9978
 x3
          1
                0.6577
                          1.5073 -7.3532
                                                3.0966 0.1533
 > res <- drop1(object = modello, scale = 0, test = "F")</pre>
 > res$Df
 [1] NA 1 1 1
 > res$"Sum of Sq"
  [1]
                 NA 2.514374e-01 1.891304e-06 6.576972e-01
 > res$RSS
 [1] 0.8495662 1.1010036 0.8495680 1.5072633
 > res$AIC
 [1] -9.939768 -9.865756 -11.939750 -7.353167
 > res$"F value"
  [1]
                 NA 1.183839e+00 8.904801e-06 3.096626e+00
 > res$"Pr(F)"
```

```
[1] NA 0.3377443 0.9977619 0.1532663
```

• Example 2:

```
> k < - 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
+ n))
> s <- summary.lm(object = modello)$sigma</pre>
> drop1(object = modello, scale = s^2, test = "F")
Single term deletions
Model:
y \sim x1 + x2 + x3
scale: 0.2123915
       Df Sum of Sq
                      RSS Cp F value Pr(F)
                    0.84957 4.0000
       1 0.25144 1.10100 3.1838
                                     1.1838 0.3377
x1
       1 1.891e-06 0.84957 2.0000 8.905e-06 0.9978
x2
          0.65770 1.50726 5.0966
                                      3.0966 0.1533
> res <- drop1(object = modello, scale = s^2, test = "F")</pre>
> res$Df
[1] NA 1 1 1
> res$"Sum of Sq"
[1]
            NA 2.514374e-01 1.891304e-06 6.576972e-01
> res$RSS
[1] 0.8495662 1.1010036 0.8495680 1.5072633
> res$Cp
[1] 4.000000 3.183839 2.000009 5.096626
> res$"F value"
             NA 1.183839e+00 8.904801e-06 3.096626e+00
[1]
> res$"Pr(F)"
[1]
         NA 0.3377443 0.9977619 0.1532663
```

add1()

• Package: stats

• Input:

object modello nullo di regressione lineare pesata scope modello di regressione lineare pesata con k-1 variabili esplicative ed n unità scale selezione indice AIC oppure Cp test = "F"

• **Description:** submodels

• Output:

Df differenza tra gradi di libertà $\begin{array}{lll} \text{Sum of Sq differenza tra devianze residue} \\ \text{RSS devianza residua} \\ \text{AIC indice } AIC \\ \text{Cp indice } Cp \\ \text{F value valore empirico della statistica } F \\ \text{Pr}\left(\mathbb{F}\right) \ p\text{-value} \\ \end{array}$

• Formula:

Df $\underbrace{\frac{1,\,1,\,\ldots,\,1}{k-1\,\mathrm{volte}}}$ Sum of Sq $RSS_{nullo}-RSS_{x_j} \ \, \forall\,j\,=\,1,\,2,\,\ldots,\,k-1$

dove RSS_{x_j} rappresenta la devianza residua del modello con la sola variabile esplicativa x_j .

RSS

$$RSS_{nullo}, RSS_{x_j} \quad \forall j = 1, 2, \dots, k-1$$

AIC

$$n \log (RSS_{nullo}/n) + 2$$
, $n \log (RSS_{x_i}/n) + 4 \quad \forall j = 1, 2, ..., k-1$

Ср

scale =
$$s^2$$

$$\frac{RSS_{nullo}}{RSS/(n-k)} + 2 - n, \frac{RSS_{x_j}}{RSS/(n-k)} + 4 - n \quad \forall j = 1, 2, \dots, k-1$$

F value

$$F_{j} = \frac{RSS_{nullo} - RSS_{x_{j}}}{RSS_{x_{j}} / (n-2)} \quad \forall j = 1, 2, ..., k-1$$

Pr(F)

$$P(F_{1,n-2} \ge F_j) \quad \forall j = 1, 2, \dots, k-1$$

• Example 1:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> nullo <- lm(formula = y ~ 1, weights = rep(1/n, n))
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> add1(object = nullo, scope = modello, scale = 0, test = "F")
```

```
Single term additions
 Model:
 y ~ 1
        Df Sum of Sq
                       RSS AIC F value Pr(F)
 <none>
                      5.9583 -0.3573
             3.2686 2.6897 -4.7201 7.2914 0.035564 *
             4.4365 1.5218 -9.2762 17.4911 0.005799 **
 x2
        1
             4.3364 1.6219 -8.7667 16.0418 0.007077 **
 x3
         1
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 > res <- add1(object = nullo, scope = modello, scale = 0, test = "F")</pre>
 > res$Df
 [1] NA 1 1 1
 > res$"Sum of Sq"
 [1] NA 3.268597 4.436456 4.336392
 > res$RSS
 [1] 5.958300 2.689703 1.521844 1.621908
 > res$AIC
 [1] -0.3572507 -4.7200862 -9.2761525 -8.7667043
 > res$"F value"
       NA 7.291356 17.491113 16.041811
 [1]
 > res$"Pr(F)"
 [1]
         NA 0.035564122 0.005799048 0.007076764
• Example 2:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n < - 8
 > nullo <- lm(formula = y \sim 1, weights = rep(1/n, n))
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
      n))
 > s <- summary.lm(object = modello)$sigma</pre>
 > add1(object = nullo, scope = modello, scale = s^2, test = "F")
 Single term additions
 Model:
 y ~ 1
 scale: 0.2123915
        Df Sum of Sq
                       RSS Cp F value Pr(F)
                     5.9583 22.0534
 <none>
             3.2686 2.6897 8.6639 7.2914 0.035564 *
 x1
 x2
             4.4365 1.5218 3.1653 17.4911 0.005799 **
 x3
        1
             4.3364 1.6219 3.6364 16.0418 0.007077 **
 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
```

```
> res <- add1(object = nullo, scope = modello, scale = s^2, test = "F")
    > res$Df
    [1] NA 1 1 1
    > res$"Sum of Sq"
               NA 3.268597 4.436456 4.336392
    [1]
    > res$RSS
    [1] 5.958300 2.689703 1.521844 1.621908
    > res$Cp
    [1] 22.053378 8.663889 3.165274 3.636408
    > res$"F value"
               NA 7.291356 17.491113 16.041811
    [1]
    > res$"Pr(F)"
                  NA 0.035564122 0.005799048 0.007076764
    [1]
leaps()
  • Package: leaps
  • Input:
        \times matrice del modello priva della prima colonna (intercetta) di dimensione n \times (h-1)
        y variabile dipendente
        wt vettore positivo dei pesi di dimensione n
        method = "r2" / "adjr2" / "Cp" indice R^2, R_{adi}^2, C_p
        nbest = 1
  • Description: Best Subsets
  • Output:
        which variabili selezionate
        size numero di parametri
        r2 / adjr2 / Cp indice R^2,\,R^2_{adj},\,C_p
  • Formula:
        size
                                            k_i \quad \forall j = 1, 2, \ldots, h-1
        r2
                                               method = "r2"
                                            R_i^2 \quad \forall j = 1, 2, \dots, h-1
```

 R_j^2 rappresenta il massimo R^2 tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

adjr2

Numero di esplicative	Numero di parametri	Numero di Subsets
1	$k_1 = 2$	$\binom{h-1}{1}$
2	$k_2 = 3$	$\binom{h-1}{2}$
	•	
	•	
j	$k_j = j + 1$	$\binom{h-1}{j}$
	•	
	•	
h-1	$k_{h-1} = h$	$\binom{h-1}{h-1}$

$$R_{adj j}^{2} = 1 - \frac{RSS / (n - k_{j})}{RSS_{nullo} / (n - 1)}$$

$$= \frac{1 - k_{j}}{n - k_{j}} + \frac{n - 1}{n - k_{j}} R_{j}^{2} \quad \forall j = 1, 2, ..., h - 1$$

 $R_{adj\,j}^2$ rappresenta il massimo R_{adj}^2 tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

Ср

$$Cp_{j} = (n - k_{h-1}) \frac{1 - R_{j}^{2}}{1 - R_{h-1}^{2}} + 2 k_{j} - n$$

$$= \left(\frac{n - k_{h-1}}{1 - R_{h-1}^{2}} + 2 k_{j} - n\right) - \frac{n - k_{h-1}}{1 - R_{h-1}^{2}} R_{j}^{2} \qquad \forall j = 1, 2, \dots, h - 1$$

 Cp_j rappresenta il minimo Cp tra i $\binom{h-1}{j}$ modelli di regressione con j variabili esplicative oppure k_j parametri.

• Example 1:

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> X <- model.matrix(object = modello)
> A <- X[, -1]
> leaps(x = A, y, wt = rep(1/n, n), method = "r2", nbest = 1)
$which
          2
     1
1 FALSE TRUE FALSE
2 TRUE FALSE TRUE
3 TRUE TRUE TRUE
$label
                                "2"
[1] "(Intercept)" "1"
                                               "3"
$size
[1] 2 3 4
```

```
$r2
 [1] 0.7445843 0.8574144 0.8574147
 > res < leaps(x = A, y, wt = rep(1/n, n), method = "r2", nbest = 1)
 > res$which
           2
       1
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 > res$size
 [1] 2 3 4
 > res$r2
 [1] 0.7445843 0.8574144 0.8574147
• Example 2:
 > k <- 4
 > x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
 > x2 \leftarrow c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > n <- 8
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
       n))
 > X <- model.matrix(object = modello)</pre>
 > A <- X[, -1]
 > leaps(x = A, y, wt = rep(1/n, n), method = "adjr2", nbest = 1)
 $which
            2
      1
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 $label
                               "2"
                                               "3"
 [1] "(Intercept)" "1"
 $size
 [1] 2 3 4
 $adjr2
 [1] 0.7020150 0.8003801 0.7504757
 > res < leaps(x = A, y, wt = rep(1/n, n), method = "adjr2", nbest = 1)
 > res$which
             2
       1
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 > res$size
 [1] 2 3 4
```

> res\$adjr2

```
[1] 0.7020150 0.8003801 0.7504757
• Example 3:
 > k <- 4
 > x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
 > x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
 > y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
 > modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
       n))
 > X <- model.matrix(object = modello)</pre>
 > A <- X[, -1]
 > leaps(x = A, y, wt = rep(1/n, n), method = "Cp", nbest = 1)
 $which
             2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 $label
 [1] "(Intercept)" "1"
                                  "2"
                                                   "3"
 $size
 [1] 2 3 4
 $Cp
 [1] 3.165274 2.000009 4.000000
 > res < leaps(x = A, y, wt = rep(1/n, n), method = "Cp", nbest = 1)
 > res$which
            2
 1 FALSE TRUE FALSE
 2 TRUE FALSE TRUE
 3 TRUE TRUE TRUE
 > res$size
 [1] 2 3 4
 > res$Cp
 [1] 3.165274 2.000009 4.000000
```

- Note 1: Tutti i modelli contengono l'intercetta.
- Note 2: R_{adj}^2 è una trasformazione lineare crescente di R_j^2 $\forall j = 1, 2, ..., h-1$.
- Note 3: Cp_j è una trasformazione lineare decrescente di $R_i^2 \quad \forall j = 1, 2, ..., h-1$.

16.4 Diagnostica

ls.diag()

• Package: stats

• Input:

ls.out modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: analisi di regressione lineare pesata
- Output:

```
std.dev stima di \sigma hat valori di leva std.res residui standard stud.res residui studentizzati cooks distanza di Cook dfits dfits correlation matrice di correlazione delle stime WLS std.err standard error delle stime WLS cov.scaled matrice di covarianza delle stime WLS non scalata per \sigma^2
```

• Formula:

hat
$$h_i \quad \forall i=1,2,\ldots,n$$
 std.res
$$rstandard_i \quad \forall i=1,2,\ldots,n$$
 stud.res
$$rstudent_i \quad \forall i=1,2,\ldots,n$$
 cooks
$$cd_i \quad \forall i=1,2,\ldots,n$$
 dfits
$$rstudent_i \sqrt{\frac{h_i}{1-h_i}} \quad \forall i=1,2,\ldots,n$$
 correlation
$$r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i,j=1,2,\ldots,k$$
 std.err
$$s_{\hat{\beta}_j} \quad \forall j=1,2,\ldots,k$$
 cov.scaled
$$s^2 \, (X^T \, W^{-1} \, X)^{-1}$$
 cov.unscaled
$$(X^T \, W^{-1} \, X)^{-1}$$

```
> k <- 4

> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n <- 8

> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))

> res <- ls.diag(ls.out = modello)

> res$std.dev
```

x3

```
[1] 1.303508
> res$hat
 \begin{smallmatrix} 1 \end{smallmatrix} \rbrack \quad 0.7695906 \quad 0.4163361 \quad 0.3791092 \quad 0.3154744 \quad 0.7283511 \quad 0.5539241 \quad 0.4302463 
[8] 0.4069682
> res$std.res
\begin{bmatrix} 1 \end{bmatrix} -1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
[8] -1.4301703
> res$stud.res
[1] -2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
[8] -1.7718134
> res$cooks
[1] 1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
[8] 0.35091186
> res$dfits
\begin{bmatrix} 1 \end{bmatrix} -3.7255223 0.3280660 1.1157578 0.4018144 0.5475321 0.7916935 -0.8516950
[8] -1.4677742
> res$correlation
             (Intercept)
                                  x1
                                                x2
(Intercept) 1.00000000 -0.1860100 0.07158062 -0.4632900
             -0.18600997 1.0000000 -0.82213982 0.4883764
x1
             0.07158062 -0.8221398 1.00000000 -0.8022181
x2
x3
             -0.46329002 0.4883764 -0.80221810 1.0000000
> res$std.err
                 [,1]
(Intercept) 4.042475
             1.098354
x1
x2
             1.646751
            1.150883
x3
> res$cov.scaled
             (Intercept)
                                               x2
                                  x1
(Intercept) 16.3416044 -0.8258968 0.4765087 -2.1554182
              -0.8258968 1.2063807 -1.4870170 0.6173452
x1
               0.4765087 -1.4870170 2.7117903 -1.5203786
x2
              -2.1554182 0.6173452 -1.5203786 1.3245321
x3
> res$cov.unscaled
             (Intercept)
                                   x1
                                               x2
              9.6176174 -0.4860697 0.2804424 -1.2685405
(Intercept)
              -0.4860697 0.7099981 -0.8751626 0.3633297
x1
              0.2804424 -0.8751626 1.5959854 -0.8947971
x2
```

-1.2685405 0.3633297 -0.8947971 0.7795344

cooks.distance()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> cooks.distance(model = modello)
1 2 3 4 5 6 7
1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
8
0.35091186
```

cookd()

- Package: car
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> cookd(model = modello)
1 2 3 4 5 6 7
1.93972080 0.03415783 0.24706215 0.04819074 0.09633983 0.17883712 0.18315058
8
0.35091186
```

rstandard()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, \ldots, n
```

• Examples:

rstandard.lm()

- Package: stats
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> rstandard.lm(model = modello)
1 2 3 4 5 6 7
-1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
8
-1.4301703
```

stdres()

• Package: MASS

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> stdres(object = modello)
1 2 3 4 5 6 7
-1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
8
-1.4301703
```

rstudent()

- Package: stats
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> rstudent(model = modello)
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8
-1.7718134
```

rstudent.lm()

• Package: stats

• Input:

 ${\tt model}$ modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, \ldots, n
```

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> rstudent.lm(model = modello)
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8
-1.7718134
```

studres()

• Package: MASS

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> studres(object = modello)
1 2 3 4 5 6 7
-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972
8
-1.7718134
```

lmwork()

• Package: MASS

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: diagnostica di regressione
- Output:

```
stdedv stima di \sigma stdres residui standard studres residui studentizzati
```

• Formula:

```
stdedv s stdres rstandard_i \  \  \forall i=1,\,2,\,\ldots,\,n studres rstudent_i \  \  \forall i=1,\,2,\,\ldots,\,n
```

• Examples:

-1.7718134

```
> k < - 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> res <- lmwork(object = modello)</pre>
> res$stdedv
[1] 0.4608596
> res$stdres
         1
                    2
                                3
                                           4
-1.5241225 0.4376576 1.2722093 0.6467323 0.3791111 0.7589935 -0.9849613
-1.4301703
> res$studres
```

-2.0384846 0.3884371 1.4278921 0.5918863 0.3343822 0.7104546 -0.9800972

dffits()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

• **Description:** dffits

• Formula:

$$rstudent_i \sqrt{\frac{h_i}{1 - h_i}} \quad \forall i = 1, 2, \dots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> dffits(model = modello)
1 2 3 4 5 6 7
-3.7255223 0.3280660 1.1157578 0.4018144 0.5475321 0.7916935 -0.8516950
8
-1.4677742
```

covratio()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** covratio
- Formula:

$$cr_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> covratio(model = modello)
1 2 3 4 5 6 7
0.4238374 4.4498753 0.6395729 2.9682483 10.0502975 3.8036903 1.8260516
8
0.3038647
```

lm.influence()

- Package: stats
- Input:

modell modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: diagnostica di regressione
- Output:

hat valori di leva coefficients differenza tra le stime WLS eliminando una unità sigma stima di σ eliminando una unità wt.res residui pesati

• Formula:

hat
$$h_i \quad \forall \, i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j - \hat{\beta}_{j\,(-i)} = w_i\,e_i\,(1-h_i)^{-1}\,(X^T\,W^{-1}\,X)_j^{-1}\,X_i^T \quad \forall i=1,2,\ldots,n \quad \forall j=1,2,\ldots,k$$
 sigma
$$s_{-i} \quad \forall \, i=1,2,\ldots,n$$
 wt.res
$$\sqrt{w_i}\,e_i \quad \forall \, i=1,2,\ldots,n$$

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> lm.influence(model = modello)
$hat
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
$coefficients
  (Intercept)
                       x1
                                    x2
                                                 x3
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 - 0.15505077 0.14594807 - 0.39064531 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
   6 \quad 0.27248353 \quad -0.28472521 \quad 0.38742501 \quad -0.16358023 
7 0.36758841 0.18614884 -0.28071294 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
$sigma
0.3445728 \ 0.5192571 \ 0.4106121 \ 0.5035642 \ 0.5225068 \ 0.4923459 \ 0.4631468 \ 0.3719961
$wt.res
                    2
                                3
                                                                   6
         1
                                            4
-0.3371620 \quad 0.1540936 \quad 0.4619923 \quad 0.2465971 \quad 0.0910624 \quad 0.2336206 \quad -0.3426347
-0.5075693
```

influence()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: diagnostica di regressione
- Output:

hat valori di leva coefficients differenza tra le stime WLS eliminando una unità sigma stima di σ eliminando una unità wt.res residui pesati

• Formula:

hat
$$h_i \quad \forall \, i=1,\,2,\,\ldots,\,n$$
 coefficients
$$\hat{\beta}_j - \hat{\beta}_{j\,(-i)} \,=\, w_i\,e_i\,(1-h_i)^{-1}\,(X^T\,W^{-1}\,X)_j^{-1}\,X_i^T \quad \forall i=1,\,2,\,\ldots,\,n \quad \forall j=1,\,2,\,\ldots,\,k$$
 sigma
$$s_{-i} \quad \forall \, i=1,\,2,\,\ldots,\,n$$
 wt.res
$$\sqrt{w_i}\,e_i \quad \forall \, i=1,\,2,\,\ldots,\,n$$

```
> k <- 4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> influence(model = modello)
$hat
                             3
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
$coefficients
  (Intercept)
                        x1
                                    x2
                                                 x3
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 - 0.15505077 0.14594807 - 0.39064531 0.32853997
4 0.10864633 -0.01436987 0.12965355 -0.11055404
   6 \quad 0.27248353 \quad -0.28472521 \quad 0.38742501 \quad -0.16358023 
7 0.36758841 0.18614884 -0.28071294 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
$sigma
0.3445728 \ 0.5192571 \ 0.4106121 \ 0.5035642 \ 0.5225068 \ 0.4923459 \ 0.4631468 \ 0.3719961
$wt.res
                     2
                                3
                                                                   6
         1
                                            4
-0.3371620 \quad 0.1540936 \quad 0.4619923 \quad 0.2465971 \quad 0.0910624 \quad 0.2336206 \quad -0.3426347
-0.5075693
```

weights()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: pesi
- Formula:

$$w_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> weights(object = modello)
[1] 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125 0.125
```

weighted.residuals()

• Package: stats

• Input:

obj modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: residui pesati
- Formula:

$$\sqrt{w_i} e_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> weighted.residuals(obj = modello)
1 2 3 4 5 6 7
-0.3371620 0.1540936 0.4619923 0.2465971 0.0910624 0.2336206 -0.3426347
8
-0.5075693
```

residuals()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità type = "response" / "pearson" tipo di residuo

- **Description:** residui
- Formula:

```
type = "response" e_i \quad \forall i = 1, 2, \dots, n type = "pearson" <math display="block">\sqrt{w_i} \, e_i \quad \forall i = 1, 2, \dots, n
```

• Example 1:

• Example 2:

residuals.lm()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

• **Description:** residui

• Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

residuals.default()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** residui
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

resid()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** residui
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> resid(object = modello)
1 2 3 4 5 6 7
-0.9536382 0.4358424 1.3067117 0.6974820 0.2575634 0.6607787 -0.9691173
8
-1.4356227
```

df.residual()

• Package: stats

• Input:

object modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

$$n-k$$

• Examples:

```
> k < -4

> x1 < -c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)

> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)

> x3 < -c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)

> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)

> n < -8

> modello < -lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))

> df.residual(object = modello)
```

hatvalues()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> hatvalues(model = modello)
1 2 3 4 5 6 7 8
0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463 0.4069682
```

hat()

- Package: stats
- Input:
 - x matrice del modello
- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, + n))
> X <- model.matrix(object = modello)
> hat(x = X)
[1] 0.7695906 0.4163361 0.3791092 0.3154744 0.7283511 0.5539241 0.4302463
[8] 0.4069682
```

dfbeta()

- Package: stats
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** dfbeta
- Formula:

$$\hat{\beta}_j - \hat{\beta}_{j(-i)} = w_i e_i (1 - h_i)^{-1} (X^T W^{-1} X)_j^{-1} X_i^T \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2, \dots, k$$

```
> k < -4
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
     n))
> dfbeta(model = modello)
  (Intercept)
                        x1
                                     x2
1 -3.95445343 0.12758388 0.01022818 0.44042192
2 0.21929134 0.01923025 -0.12292616 0.08309302
3 -0.15505077 0.14594807 -0.39064531 0.32853997
  0.10864633 -0.01436987 0.12965355 -0.11055404
5 \quad 0.06456839 \quad 0.14591697 \quad -0.04391330 \quad -0.06357315
 6 \quad 0.27248353 \quad -0.28472521 \quad 0.38742501 \quad -0.16358023 
7 0.36758841 0.18614884 -0.28071294 0.03129723
8 0.76981755 -0.23622669 0.37474061 -0.34716366
```

dfbetas()

• Package: stats

• Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** dfbetas
- Formula:

$$\frac{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}{s_{\hat{\beta}_{j} - \hat{\beta}_{j\,(-i)}}} = \frac{w_{i} e_{i} (1 - h_{i})^{-1} (X^{T} W^{-1} X)_{j}^{-1} X_{i}^{T}}{s_{-i} \sqrt{(X^{T} W^{-1} X)_{j,j}^{-1}}} \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2, \dots, k$$

```
> k < - 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 < -c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
      n))
> dfbetas(model = modello)
  (Intercept)
                         x1
                                      x2
1 -3.70059595 0.43942641 0.02349647 1.44767218
2 \quad 0.13617748 \quad 0.04395152 \quad -0.18739044 \quad 0.18124433
3 - 0.12176106 \quad 0.42183052 \quad -0.75307182 \quad 0.90623075
4 \quad 0.06957072 \quad -0.03386642 \quad 0.20380513 \quad -0.24865783
5 0.03984687 0.33142498 -0.06652573 -0.13780473
6 0.17845806 -0.68632053 0.62287782 -0.37630746
7 0.25592307 0.47699422 -0.47976587 0.07653668
8 0.66729165 -0.75363662 0.79740312 -1.05700791
```

vif()

- Package: car
- Input:

mod modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** variance inflation factors
- Formula:

$$\left(1 - R_{x_j}^2\right)^{-1} \quad \forall j = 1, 2, \dots, k - 1$$

 $R_{x_j}^2$ rappresenta il valore di R^2 per il modello che presenta il regressore j-esimo come variabile dipendente

• Examples:

outlier.test()

- Package: car
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- Description: test sugli outliers
- Output:

test massimo residuo studentizzato assoluto, gradi di libertà, p-value

• Formula:

test

```
t = \max_{i}(|rstudent_{i}|)  n-k-1 p-value = 2P(t_{n-k-1} \le -|t|) \forall i = 1, 2, \dots, n
```

```
> k <- 4
> x1 <- c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 <- c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y <- c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n <- 8
> modello <- lm(formula = y ~ x1 + x2 + x3, weights = rep(1/n, n))
> outlier.test(model = modello)

max|rstudent| = 2.038485, degrees of freedom = 3, unadjusted p = 0.1342423, Bonferroni p > 1
Observation: 1
```

influence.measures()

- Package: stats
- Input:

model modello di regressione lineare pesata con k-1 variabili esplicative ed n unità

- **Description:** dfbetas, dffits, covratio, distanza di *Cook*, valori di leva
- Output:

```
infmat misure di influenza di dimensione n \times (k+4) is.inf matrice di influenza con valori logici di dimensione n \times (k+4)
```

• Formula:

infmat

$$DFBETAS_{ij} = \frac{w_i \, e_i \, (1 - h_i)^{-1} \, (X^T \, W^{-1} \, X)_j^{-1} \, X_i^T}{s_{-i} \, \sqrt{(X^T \, W^{-1} \, X)_{j,j}^{-1}}} \quad \forall i = 1, 2, \dots, n \quad \forall j = 1, 2, \dots, k$$

$$DFFITS_i = rstudent_i \, \sqrt{\frac{h_i}{1 - h_i}} \quad \forall i = 1, 2, \dots, n$$

$$COVRATIO_i = (1 - h_i)^{-1} \, \left(1 + \frac{rstudent_i^2 - 1}{n - k}\right)^{-k} \quad \forall i = 1, 2, \dots, n$$

$$COOKD_i = \frac{h_i \, rstandard_i^2}{k \, (1 - h_i)} \quad \forall i = 1, 2, \dots, n$$

$$HAT_i = h_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

```
> x1 \leftarrow c(1.1, 2.3, 4.5, 6.7, 8.9, 3.4, 5.6, 6.7)
> x2 <- c(1.2, 3.4, 5.6, 7.5, 7.5, 6.7, 8.6, 7.6)
> x3 \leftarrow c(1.4, 5.6, 7.56, 6, 5.4, 6.6, 8.7, 8.7)
> y < -c(1.5, 6.4, 9.6, 8.8, 8.86, 7.8, 8.6, 8.6)
> n < - 8
> modello <- lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n,
> res <- influence.measures(model = modello)</pre>
> res
Influence measures of
          lm(formula = y \sim x1 + x2 + x3, weights = rep(1/n, n)):
   dfb.1_ dfb.x1 dfb.x2 dfb.x3 dffit cov.r cook.d hat inf
1 - 3.7006 \quad 0.4394 \quad 0.0235 \quad 1.4477 \quad -3.726 \quad 0.424 \quad 1.9397 \quad 0.770
2 0.1362 0.0440 -0.1874 0.1812 0.328 4.450 0.0342 0.416
3 - 0.1218 \quad 0.4218 \quad -0.7531 \quad 0.9062 \quad 1.116 \quad 0.640 \quad 0.2471 \quad 0.379
4 0.0696 -0.0339 0.2038 -0.2487 0.402 2.968 0.0482 0.315
5 \quad 0.0398 \quad 0.3314 \quad -0.0665 \quad -0.1378 \quad 0.548 \quad 10.050 \quad 0.0963 \quad 0.728
6 0.1785 -0.6863 0.6229 -0.3763 0.792 3.804 0.1788 0.554
  0.2559   0.4770   -0.4798   0.0765   -0.852   1.826   0.1832   0.430
  0.6673 -0.7536  0.7974 -1.0570 -1.468  0.304  0.3509  0.407
```

> res\$infmat

```
dfb.1_{-}
                   dfb.x1
                               dfb.x2
                                           dfb.x3
                                                        dffit
                                                                   cov.r
1 - 3.70059595 0.43942641 0.02349647 1.44767218 -3.7255223
                                                               0.4238374
2 0.13617748 0.04395152 -0.18739044 0.18124433
                                                   0.3280660
                                                               4.4498753
3 -0.12176106  0.42183052 -0.75307182  0.90623075
                                                   1.1157578
                                                               0.6395729
4 \quad 0.06957072 \quad -0.03386642 \quad 0.20380513 \quad -0.24865783 \quad 0.4018144 \quad 2.9682483
5 \quad 0.03984687 \quad 0.33142498 \quad -0.06652573 \quad -0.13780473 \quad 0.5475321 \quad 10.0502975
3.8036903
  0.25592307 0.47699422 -0.47976587 0.07653668 -0.8516950
                                                              1.8260516
8 \quad 0.66729165 \quad -0.75363662 \quad 0.79740312 \quad -1.05700791 \quad -1.4677742 \quad 0.3038647
      cook.d
                   hat
1 1.93972080 0.7695906
2 0.03415783 0.4163361
3 0.24706215 0.3791092
4 0.04819074 0.3154744
5 0.09633983 0.7283511
6 0.17883712 0.5539241
7 0.18315058 0.4302463
8 0.35091186 0.4069682
> res$is.inf
  dfb.1_ dfb.x1 dfb.x2 dfb.x3 dffit cov.r cook.d
   TRUE
         FALSE
                FALSE
                         TRUE
                              TRUE FALSE
                                            TRUE FALSE
2
  FALSE
         FALSE
                FALSE
                       FALSE FALSE
                                     TRUE
                                           FALSE FALSE
3
  FALSE FALSE FALSE FALSE FALSE
                                           FALSE FALSE
4 FALSE FALSE FALSE FALSE FALSE
                                           FALSE FALSE
5 FALSE FALSE FALSE
                        FALSE FALSE
                                     TRUE
                                           FALSE FALSE
6 FALSE
         FALSE
                FALSE
                        FALSE FALSE FALSE
                                           FALSE FALSE
7
  FALSE FALSE
               FALSE
                        FALSE FALSE FALSE
                                           FALSE FALSE
```

• Note 1: Il caso *i*-esimo è influente se $|DFBETAS_{ij}| > 1 \quad \forall i = 1, 2, ..., n \quad \forall j = 1, 2, ..., k$

TRUE FALSE FALSE FALSE

- Note 2: Il caso *i*-esimo è influente se $|DFFITS_i| > 3\sqrt{k/(n-k)}$ $\forall i = 1, 2, ..., n$
- Note 3: Il caso *i*-esimo è influente se $|1 COVRATIO_i| > 3 k / (n k)$ $\forall i = 1, 2, ..., n$
- Note 4: Il caso *i*-esimo è influente se $P(F_{k,n-k} \geq COOKD_i) > 0.5 \quad \forall i = 1, 2, ..., n$
- Note 5: Il caso *i*-esimo è influente se $HAT_i > 3k/n \quad \forall i = 1, 2, ..., n$

8 FALSE FALSE FALSE

• **Note 6:** I casi influenti rispetto ad almeno una tra queste misure sono marcati con un asterisco. Corrispondentemente la stessa riga della matrice is.inf riporterà almeno un simbolo TRUE.

Parte V Modelli Lineari Generalizzati

Capitolo 17

Regressione Logit

17.1 Simbologia

$$\log\left(\frac{\pi_i}{1-\pi_i}\right) = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} \qquad Y_i \sim \text{Bin}(\pi_i, n_i) \quad \forall i = 1, 2, \dots, n$$

- numero di successi: $y_i \quad \forall i = 1, 2, ..., n$
- numero di prove: $n_i \quad \forall i = 1, 2, \ldots, n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w=(w_1,w_2,\ldots,w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \mathrm{diag}(w_1^{-1},\,w_2^{-1},\,\dots,\,w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, ..., n$
- distanza di Cook: $cd_i=\left(e_i^P\right)^2 rac{h_i}{k\left(1-h_i\right)^2} \quad \forall \, i=1,\,2,\,\ldots,\,n$
- stime IWLS: $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{\beta}} = \hat{\beta} / s_{\hat{\beta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \; \hat{\beta}_j} = \frac{(X^T \, W^{-1} \, X)_{i, \, j}^{-1}}{s_{\hat{\beta}_i} \; s_{\hat{\beta}_j}} \quad \forall \, i, j \, = \, 1, \, 2, \, \ldots, \, k$
- residui di devianza: $e_i = \operatorname{sign}(y_i \hat{y}_i) \sqrt{2 \left[y_i \log \left(\frac{y_i}{\hat{y}_i} + C_{i1} \right) + (n_i y_i) \log \left(\frac{n_i y_i}{n_i \hat{y}_i} + C_{i2} \right) \right]}$ $\forall i = 1, 2, \dots, n$ dove $C_{i1} = 0.5 \left(1 \operatorname{sign}(y_i) \right) / \hat{y}_i$ e $C_{i2} = 0.5 \left(1 \operatorname{sign}(n_i y_i) \right) / (n_i \hat{y}_i)$
- residui standard: $rstandard_i = e_i / \sqrt{1 h_i} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = sign(y_i \hat{y}_i) \sqrt{e_i^2 + h_i (e_i^P)^2 / (1 h_i)} \quad \forall i = 1, 2, ..., n$
- residui di *Pearson*: $e_i^P = \frac{y_i n_i \, \hat{\pi}_i}{\sqrt{n_i \, \hat{\pi}_i \, (1 \hat{\pi}_i)}} \quad \forall i = 1, 2, \ldots, n$
- residui di lavoro: $e_i^W=rac{y_i-n_i\,\hat{\pi}_i}{n_i\,\hat{\pi}_i\,(1-\hat{\pi}_i)} \quad \forall i=1,\,2,\,\ldots,\,n$
- residui di riposta: $e_i^R = y_i / n_i \hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$
- log-verosimiglianza binomiale: $\hat{\ell} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{\hat{y}_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{\hat{y}_i}{n_i} \right) \right]$
- valori adattati: $\hat{\pi}_i = \frac{\exp(X_i \hat{\beta})}{1 + \exp(X_i \hat{\beta})} \quad \forall i = 1, 2, ..., n$

- numero di successi attesi: $\hat{y}_i = n_i \hat{\pi}_i \quad \forall i = 1, 2, ..., n$
- log-verosimiglianza binomiale modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{y_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{y_i}{n_i} \right) \right]$
- devianza residua: $D=2\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^n e_i^2$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza binomiale modello nullo: $\hat{\ell}_{nullo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log (\hat{\pi}) + (n_i y_i) \log (1 \hat{\pi}) \right]$
- valori adattati modello nullo: $\hat{\pi} = \sum_{j=1}^n y_j / \sum_{j=1}^n n_j \quad \forall i = 1, 2, ..., n$
- numero di successi attesi modello nullo: $\hat{y}_i = n_i \hat{\pi} \quad \forall i = 1, 2, ..., n$
- devianza residua modello nullo: $D_{nullo} = 2 \left(\hat{\ell}_{saturo} \hat{\ell}_{nullo} \right)$
- gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{eta}_{nullo} = \log\left(\frac{\hat{\pi}}{1-\hat{\pi}}\right)$

17.2 Stima

glm()

- Package: stats
- Input:

formula modello di regressione logit con k-1 variabili esplicative ed n unità family = binomial(link="logit") famiglia e link del modello x = TRUE matrice del modello

- Description: analisi di regressione logit
- Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y proporzione di successi
x matrice del modello

• Formula:

coefficients $\hat{\beta}_j \quad \forall j=1,2,\ldots,k$ residuals $e^W_i \quad \forall i=1,2,\ldots,n$ fitted.values $\hat{\pi}_i \quad \forall i=1,2,\ldots,n$

```
rank
                                                   k
linear.predictors
                                                 X \hat{\beta}
deviance
                                                  D
aic
                                               -2\hat{\ell}+2k
null.deviance
                                                D_{nullo}
weights
                                         w_i \quad \forall i = 1, 2, \ldots, n
prior.weights
                                         n_i \quad \forall i = 1, 2, \ldots, n
df.residual
                                                n-k
df.null
                                                 n-1
У
                                       y_i / n_i \quad \forall i = 1, 2, \ldots, n
                                                  X
 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
  14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
  88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"),</pre>
+ x = TRUE)
> modello$coefficients
(Intercept)
-21.226395 1.631968
> modello$residuals
                                 3
-1.00203763 \ -1.01042031 \ -1.01905988 \ -0.41336424 \ -0.48212701 \ -0.07089826
         7
                     8
                                 9
                                            10
                                                         11
```

22

17

23

 $0.07938086 \quad 0.22704866 \quad -0.13926878 \quad 0.33629857 \quad 0.25835047 \quad 0.17881393$

 $0.19514514 \ -0.43506531 \ -0.25760272 \ -0.64783388 \ -0.44626460 \ -0.78405425$

14 15 16

21

-0.22141017 0.01336452 0.26283804 -0.24965088 -0.36552096

20

> modello\$fitted.values

13

19

1.00057358

```
1 2 3 4 5 6
0.002033490 0.010312851 0.018703394 0.027863526 0.041320994 0.060871141
7 8 9 10 11 12
0.088814107 0.127838223 0.180610428 0.248949062 0.332647930 0.428434554
13 14 15 16 17 18
0.529902047 0.628956590 0.718237396 0.793102235 0.852169542 0.896572801
19 20 21 22 23 24
0.928753893 0.951463983 0.967190831 0.977939948 0.985221193 0.990123427
25
0.999426746
```

> modello\$rank

[1] 2

> modello\$linear.predictors

```
-6.1959664 -4.5639981 -3.9601698 -3.5521777 -3.1441856 -2.7361935 -2.3282014
                   10
                             11
                                       12
                9
                                                      13
-1.9202093 -1.5122173 -1.1042252 -0.6962331 -0.2882410 0.1197511 0.5277432
      15
          16
                   17
                             18
                                       19
                                                      20
0.9357353 \quad 1.3437274 \quad 1.7517194 \quad 2.1597115 \quad 2.5677036 \quad 2.9756957 \quad 3.3836878
          23
                    24
                              25
      2.2
3.7916799 4.1996720 4.6076640 7.4636087
```

> modello\$deviance

[1] 26.70345

> modello\$aic

[1] 114.7553

> modello\$null.deviance

[1] 3693.884

> modello\$weights

```
2
                          3
                                   4
0.7630428 \quad 2.0413099 \quad 1.7068902 \quad 3.2504707 \quad 3.5652333 \quad 5.0306085 \quad 8.4972661
                9
                    10
                             11
                                       12
                                                     13
12.3760338 14.7990471 17.3885402 22.1993347 26.4468672 24.6614810 24.7372446
      15 16 17
                             18
                                       19
                                                20
21.2491158 19.1986735 12.3457255 8.9948289 7.9404319 4.7104022 3.8714069
               23 24
2.3946581 1.3686835 1.1148148 0.6010036
```

> modello\$prior.weights

```
3
              5
                      7 8 9
1
           4
                  6
                                10 11 12 13 14 15 16
         120
              90
                  88 105 111 100
                                93 100 108 99 106 105 117
          20
              21
                  22
                     23
                            25
17
   18
      19
                        24
   97 120 102 122 111
                     94 114 1049
98
```

> modello\$df.residual

[1] 23

```
> modello$df.null
```

[1] 24

> modello\$y

> modello\$x

```
(Intercept)
            1 9.21
1
            1 10.21
3
            1 10.58
4
            1 10.83
5
           1 11.08
6
           1 11.33
7
           1 11.58
8
           1 11.83
9
            1 12.08
10
            1 12.33
            1 12.58
11
           1 12.83
12
           1 13.08
13
14
           1 13.33
15
           1 13.58
           1 13.83
16
17
           1 14.08
18
            1 14.33
            1 14.58
19
           1 14.83
20
           1 15.08
21
22
           1 15.33
23
           1 15.58
24
           1 15.83
25
            1 17.58
attr(,"assign")
[1] 0 1
```

summary.glm()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- **Description:** analisi di regressione logit
- Output:

deviance devianza residua aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS

• Formula:

deviance
$$D$$
 aic
$$-2\hat{\ell}+2k$$
 df.residual
$$n-k$$
 null.deviance
$$D_{nullo}$$
 df.null
$$n-1$$
 deviance.resid
$$e_i \ \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j \ s_{\hat{\beta}_j} \ z_{\hat{\beta}_j} \ p\text{-value} = 2\Phi(-|z_{\hat{\beta}_j}|) \quad \forall j=1,2,\ldots,k$$
 cov.unscaled
$$(X^TW^{-1}X)^{-1}$$
 cov.scaled
$$(X^TW^{-1}X)^{-1}$$
 correlation
$$r_{\hat{\beta},\hat{\beta}_i} \ \forall i,j=1,2,\ldots,k$$

```
> res$null.deviance
   [1] 3693.884
   > res$df.null
   [1] 24
   > res$deviance.resid
                      2
                                3
                                    4
   -1.2372312 -2.0363101 -1.8739732 -0.8043827 -0.9953320 -0.1607163 0.2289532
           8 9 10 11 12 13
    0.7780252 \; -0.5441548 \quad 1.3675388 \quad 1.2016944 \quad 0.9162826 \; -1.0982255 \quad 0.0665090
              16 17
                                             19 20 21
                                    18
    1.2375553 \ -1.0695134 \ -1.2358120 \ 1.0633044 \ 0.5665503 \ -0.8912577 \ -0.4883964
              23
                         24
                                    25
   -0.9195743 -0.4900070 -0.7461893 1.0968278
   > res$coefficients
                Estimate Std. Error z value
                                                Pr(>|z|)
   (Intercept) -21.226395 0.77068466 -27.54226 5.479038e-167
                1.631968 0.05895308 27.68249 1.134448e-168
   > res$cov.unscaled
              (Intercept)
   (Intercept) 0.59395485 -0.045281754
              -0.04528175 0.003475466
   > res$cov.scaled
              (Intercept)
   (Intercept) 0.59395485 -0.045281754
              -0.04528175 0.003475466
   > res$correlation
              (Intercept)
   (Intercept)
                1.000000 -0.996644
                -0.996644 1.000000
glm.fit()
 • Package: stats
 • Input:
      x matrice del modello
      y proporzione di successi
      weights numero di prove
      family = binomial(link="logit") famiglia e link del modello
 • Description: analisi di regressione logit
```

coefficients stime IWLS

• Output:

residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua
df.null gradi di libertà devianza residua modello nullo
y proporzione di successi

• Formula:

coefficients $\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$ residuals $e_i^W \quad \forall i = 1, 2, \dots, n$ fitted.values $\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$ rank klinear.predictors $X \hat{\beta}$ deviance Daic $-2\hat{\ell}+2k$ null.deviance D_{nullo} weights $w_i \quad \forall i = 1, 2, \ldots, n$ prior.weights $n_i \quad \forall i = 1, 2, \ldots, n$ df.residual n-kdf.null n-1У $y_i / n_i \quad \forall i = 1, 2, \ldots, n$

• Examples:

```
(Intercept)
 -21.226395 1.631968
> res$residuals
  [1] \quad -1.00203763 \quad -1.01042031 \quad -1.01905988 \quad -0.41336424 \quad -0.48212701 \quad -0.07089826 
 [7] 0.07938086 0.22704866 -0.13926878 0.33629857 0.25835047 0.17881393
[13] \ -0.22141017 \ \ 0.01336452 \ \ 0.26283804 \ -0.24965088 \ -0.36552096 \ \ 0.33713195
[19] \quad 0.19514514 \quad -0.43506531 \quad -0.25760272 \quad -0.64783388 \quad -0.44626460 \quad -0.78405425
[25] 1.00057358
> res$fitted.values
 [1] 0.002033490 0.010312851 0.018703394 0.027863526 0.041320994 0.060871141
 [7] 0.088814107 0.127838223 0.180610428 0.248949062 0.332647930 0.428434554
[13] 0.529902047 0.628956590 0.718237396 0.793102235 0.852169542 0.896572801
[19] 0.928753893 0.951463983 0.967190831 0.977939948 0.985221193 0.990123427
[25] 0.999426746
> res$rank
[1] 2
> res$linear.predictors
 [1] -6.1959664 -4.5639981 -3.9601698 -3.5521777 -3.1441856 -2.7361935
  \lceil 7 \rceil \ -2.3282014 \ -1.9202093 \ -1.5122173 \ -1.1042252 \ -0.6962331 \ -0.2882410 
     0.1197511 0.5277432 0.9357353 1.3437274 1.7517194
                                                               2.1597115
[19] 2.5677036 2.9756957 3.3836878 3.7916799 4.1996720
[25] 7.4636087
> res$deviance
[1] 26.70345
> res$aic
[1] 114.7553
> res$null.deviance
[1] 3693.884
> res$weights
 [1] 0.7630428 2.0413099 1.7068902 3.2504707 3.5652333 5.0306085
     8.4972661 12.3760338 14.7990471 17.3885402 22.1993347 26.4468672
[13] 24.6614810 24.7372446 21.2491158 19.1986735 12.3457255
                                                               8.9948289
[19] 7.9404319 4.7104022 3.8714069 2.3946581 1.3686835 1.1148148
[25] 0.6010036
> res$prior.weights
[1] 376 200
                 93 120
                           90
                                88 105 111
                                               100
                                                     93 100 108 99 106 105
[16] 117
           98
                 97 120 102 122
                                    111
                                          94
                                               114 1049
> res$df.residual
[1] 23
```

```
> res$df.null

[1] 24

> res$y

[1] 0.00000000 0.00000000 0.00000000 0.01666667 0.02222222 0.05681818
[7] 0.09523810 0.15315315 0.16000000 0.31182796 0.39000000 0.47222222
[13] 0.47474747 0.63207547 0.77142857 0.75213675 0.80612245 0.92783505
[19] 0.94166667 0.93137255 0.95901639 0.96396396 0.97872340 0.98245614
[25] 1.00000000
```

vcov()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$(X^T W^{-1} X)^{-1}$$

• Examples:

coef()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_j \quad \forall j = 1, 2, \ldots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

Package: stats

• Input:

```
object modello di regressione logit con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime
```

- **Description**: previsione
- Output:

```
fit valore previsto
se.fit standard error delle stime
```

• Formula:

```
fit x_0^T \, \hat{\beta} se.fit \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}
```

• Examples:

predict()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

```
1
-19.10484
> res$se.fit
[1] 0.6943312
```

fitted()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))
> fitted(object = modello)
0.002033490 \ 0.010312851 \ 0.018703394 \ 0.027863526 \ 0.041320994 \ 0.060871141
                                              1.0
                                                          11
0.088814107 0.127838223 0.180610428 0.248949062 0.332647930 0.428434554
         13
                     14
                                 15
                                             16
                                                          17
0.529902047 \ 0.628956590 \ 0.718237396 \ 0.793102235 \ 0.852169542 \ 0.896572801
                     20
                                       22
                                                          23
        19
                                 21
0.928753893 0.951463983 0.967190831 0.977939948 0.985221193 0.990123427
         25
0.999426746
```

fitted.values()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> fitted.values(object = modello)
0.002033490 \ 0.010312851 \ 0.018703394 \ 0.027863526 \ 0.041320994 \ 0.060871141
                8
                                 9
                                            10
                                                         11
0.088814107 0.127838223 0.180610428 0.248949062 0.332647930 0.428434554
        13
                    14
                                 15
                                             16
                                                         17
0.529902047 \ 0.628956590 \ 0.718237396 \ 0.793102235 \ 0.852169542 \ 0.896572801
                     20
                                21
                                            22
                                                         23
        19
0.928753893 0.951463983 0.967190831 0.977939948 0.985221193 0.990123427
0.999426746
```

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$$

• Examples:

17.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** log-verosimiglianza binomiale
- Formula:

 $\hat{\ell}$

• Examples:

AIC()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

 $-2\hat{\ell}+2k$

• Examples:

durbin.watson()

• Package: car

• Input:

model modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> durbin.watson(model = modello)
 lag Autocorrelation D-W Statistic p-value
  1
           0.3440895
                         1.209446 0.034
 Alternative hypothesis: rho != 0
> res <- durbin.watson(model = modello)</pre>
> res$dw
[1] 1.209446
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2k$$

deviance()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

• Examples:

[1] 26.70345

anova()

• Package: stats

• Input:

nullo modello nullo di regressione logit con n unità modello modello di regressione logit con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
         12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
          14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
   > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
          88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
   > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
          108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
          1049)
   > nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "logit"))
   > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
   > anova(nullo, modello, test = "Chisq")
   Analysis of Deviance Table
   Model 1: cbind(y, Total - y) ~ 1
   Model 2: cbind(y, Total - y) \sim x
     Resid. Df Resid. Dev Df Deviance P(>|Chi|)
             24
                  3693.9
   2
             23
                     26.7 1 3667.2
                                             0.0
   > res <- anova(nullo, modello, test = "Chisq")</pre>
   > res$"Resid. Df"
   [1] 24 23
   > res$"Resid. Dev"
   [1] 3693.88357 26.70345
   > res$Df
   [1] NA 1
   > res$Deviance
   [1]
           NA 3667.18
   > res$"P(>|Chi|)"
   [1] NA 0
drop1()
  • Package: stats
  • Input:
       object modello di regressione logit con k-1 variabili esplicative ed n unità
       test = "Chisq"
  • Description: submodels
  • Output:
       Df differenza tra gradi di libertà
       Deviance differenza tra devianze residue
```

AIC indice AIC

Pr(Chi) p-value

LRT valore empirico della statistica χ^2

• Formula:

Deviance $\underbrace{1,1,\ldots,1}_{k-1\,\mathrm{volte}}$ Deviance $D,\,D_{-x_j}\ \, \forall\,j\,=\,1,\,2,\,\ldots,\,k-1$

dove D_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

AIC

$$-2\hat{\ell} + 2k, -2\hat{\ell}_{-x_i} + 2(k-1) \quad \forall j = 1, 2, \dots, k-1$$

dove $\hat{\ell}_{-x_j}$ rappresenta la log-verosimiglianza binomiale del modello eliminata la variabile esplicativa x_j . LRT

 $c_j = D_{-x_j} - D \quad \forall j = 1, 2, \dots, k - 1$

Pr(Chi)

$$P(\chi_1^2 \ge c_j) \quad \forall j = 1, 2, \dots, k-1$$

• Examples:

[1] NA 0

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> drop1(object = modello, test = "Chisq")
Single term deletions
Model:
cbind(y, Total - y) \sim x
                    AIC
       Df Deviance
                                    Pr(Chi)
                             LRT
              26.7 114.8
            3693.9 3779.9 3667.2 < 2.2e-16 ***
Х
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- drop1(object = modello, test = "Chisq")</pre>
> res$Df
[1] NA 1
> res$Deviance
      26.70345 3693.88357
[1]
> res$AIC
[1] 114.7553 3779.9354
> res$LRT
[1]
        NA 3667.18
> res$"Pr(Chi)"
```

add1()

• Package: stats

• Input:

object modello nullo di regressione logit scope modello di regressione logit con k-1 variabili esplicative ed n unità test = "Chisq"

- Description: submodels
- Output:

Df differenza tra gradi di libertà Deviance differenza tra devianze residue AIC indice AIC LRT valore empirico della statistica χ^2 Pr(Chi) p-value

• Formula:

Df

$$\underbrace{1, 1, \ldots, 1}_{k-1 \text{ volte}}$$

Deviance

$$D_{nullo}, D_{x_j} \quad \forall j = 1, 2, \dots, k-1$$

dove D_{x_i} rappresenta la devianza residua del modello con la sola variabile esplicativa x_i .

AIC

$$-2\,\hat{\ell}_{nullo} + 2, \, -2\,\hat{\ell}_{x_j} + 4 \quad \forall j = 1, \, 2, \, \dots, \, k-1$$

dove $\hat{\ell}_{x_j}$ rappresenta la log-verosimiglianza binomiale del modello con la sola variabile esplicativa x_j .

LRT

$$c_j = D_{nullo} - D_{x_j} \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "logit"))</pre>
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> add1(object = nullo, scope = modello, test = "Chisq")
Single term additions
Model:
cbind(y, Total - y) \sim 1
      Df Deviance
                     AIC
                            LRT Pr(Chi)
            3693.9 3779.9
<none>
             26.7 114.8 3667.2 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

17.4 Diagnostica

rstandard()

• Package: stats

• Input:

model modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

 $rstandard_i \quad \forall i = 1, 2, \ldots, n$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> rstandard(model = modello)
                                 3
-1.26387269 -2.10534096 -1.91498313 -0.83301527 -1.02729335 -0.16669886
         7
                    8
                                9
                                           1.0
                                                        11
 0.24077974   0.82521025   -0.57526008   1.44049872   1.26945542   0.97065728
        13
                               15
                                           16
                                                       17
-1.15658902 0.07035119 1.30959757 -1.13960327 -1.30015928 1.11385953
        19
                    20
                        21
                                   22
                                                23
 0.59653144 \ -0.92511157 \ -0.50699153 \ -0.94525426 \ -0.49917710 \ -0.75953595
 1.12275650
```

rstandard.glm()

• Package: stats

• Input:

modeln modello di regressione logit con n-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> rstandard.glm(model = modello)
                                   3
                                               4
-1.26387269 -2.10534096 -1.91498313 -0.83301527 -1.02729335 -0.16669886
                      8
                                              10
                                                           11
 0.24077974 \quad 0.82521025 \quad -0.57526008 \quad 1.44049872 \quad 1.26945542 \quad 0.97065728
         13
                     14
                                  15
                                              16
                                                           17
-1.15658902 0.07035119 1.30959757 -1.13960327 -1.30015928
                                                              1.11385953
        19
                                                     23
                                       22
                    2.0
                                2.1
 0.59653144 \ -0.92511157 \ -0.50699153 \ -0.94525426 \ -0.49917710 \ -0.75953595
         25
 1.12275650
```

rstudent()

• Package: stats

• Input:

modell modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, ..., n
```

rstudent.glm()

• Package: stats

• Input:

modeln modello di regressione logit con n-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))
> rstudent.glm(model = modello)
                               3
                                          4
-1.25063645 -2.07129265 -1.89478391 -0.82902073 -1.02213647 -0.16657527
             8 9 10 11
 0.24102704 \quad 0.82768067 \quad -0.57433275 \quad 1.44416053 \quad 1.27117259 \quad 0.97103803
            14 15 16 17
       13
-1.15672425 0.07034687 1.30668616 -1.14272936 -1.30517189 1.10911742
                                  22
                       21
                                                     23
 0.59483577 - 0.92917154 - 0.50839548 - 0.95001692 - 0.50040422 - 0.76258344
 1.10987159
```

residuals.default()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals.default(object = modello)
                                 3
         1
                     2
-1.00203763 -1.01042031 -1.01905988 -0.41336424 -0.48212701 -0.07089826
         7
              8 9 10
                                                       11
0.07938086 \quad 0.22704866 \quad -0.13926878 \quad 0.33629857 \quad 0.25835047 \quad 0.17881393
                        15
        13
             14
                                           16
                                                       17
-0.22141017 0.01336452 0.26283804 -0.24965088 -0.36552096 0.33713195
                               21
                                           22
        19
                    2.0
                                                       2.3
0.19514514 \ -0.43506531 \ -0.25760272 \ -0.64783388 \ -0.44626460 \ -0.78405425
1.00057358
```

residuals()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{l} \text{type = "deviance"} \\ e_i \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \end{array}$$

• Example 1:

• Example 2:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals(object = modello, type = "pearson")
                    2
                              3
                                         4
-0.87529996 -1.44362837 -1.33137848 -0.74525548 -0.91034225 -0.15901761
            8 9 10 11 12
 0.23139551 \quad 0.79874716 \quad -0.53576012 \quad 1.40235004 \quad 1.21724831 \quad 0.91957777
       13 14 15 16 17
-1.09953015 0.06647053 1.21159801 -1.09387707 -1.28431127 1.01110426
           20 21 22 23 24
        19
 0.54989436 \ -0.94424085 \ -0.50685539 \ -1.00250029 \ -0.52208706 \ -0.82783987
 0.77568558
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals(object = modello, type = "working")
                               3
-1.00203763 \ -1.01042031 \ -1.01905988 \ -0.41336424 \ -0.48212701 \ -0.07089826
                       9
                    8
                                   10
                                              11
 0.07938086 \quad 0.22704866 \quad -0.13926878 \quad 0.33629857 \quad 0.25835047 \quad 0.17881393
        1.3
             14 15 16 17
-0.22141017 0.01336452 0.26283804 -0.24965088 -0.36552096 0.33713195
                   20 21 22 23 24
 0.19514514 \ -0.43506531 \ -0.25760272 \ -0.64783388 \ -0.44626460 \ -0.78405425
        25
 1.00057358
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
+ 88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)</pre>
```

```
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "logit"))
> residuals(object = modello, type = "response")
-0.0020334895 \ -0.0103128513 \ -0.0187033936 \ -0.0111968589 \ -0.0190987716
                     7
           6
                                      8
                                                     9
-0.0040529588 \quad 0.0064239884 \quad 0.0253149298 \quad -0.0206104280 \quad 0.0628788951
           11
                         12
                                        13
                                                       14
 0.0573520700 0.0437876678 -0.0551545725 0.0031188816 0.0531911753
                         17
           16
                                        18
                                                       19
-0.0409654825 \ -0.0460470931 \ \ 0.0312622502 \ \ 0.0129127734 \ -0.0200914343
           21
                         22
                                        23
                                                       24
-0.0081744371 -0.0139759836 -0.0064977884 -0.0076672869 0.0005732538
```

residuals.glm()

• Package: stats

• Input:

```
object modello di regressione logit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- Description: residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

• Example 2:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals.glm(object = modello, type = "pearson")
                    2
                               3
                                          4
-0.87529996 -1.44362837 -1.33137848 -0.74525548 -0.91034225 -0.15901761
            8 9 10 11 12
 0.23139551 \quad 0.79874716 \quad -0.53576012 \quad 1.40235004 \quad 1.21724831 \quad 0.91957777
       13 14 15 16 17
-1.09953015 0.06647053 1.21159801 -1.09387707 -1.28431127 1.01110426
           20 21 22 23 24
        19
 0.54989436 \ -0.94424085 \ -0.50685539 \ -1.00250029 \ -0.52208706 \ -0.82783987
 0.77568558
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals.glm(object = modello, type = "working")
                                3
-1.00203763 \ -1.01042031 \ -1.01905988 \ -0.41336424 \ -0.48212701 \ -0.07089826
                        9
                    8
                                   10
                                               11
 0.07938086 \quad 0.22704866 \quad -0.13926878 \quad 0.33629857 \quad 0.25835047 \quad 0.17881393
        1.3
                  14 15 16 17
-0.22141017 0.01336452 0.26283804 -0.24965088 -0.36552096 0.33713195
                   20 21 22 23 24
 0.19514514 \ -0.43506531 \ -0.25760272 \ -0.64783388 \ -0.44626460 \ -0.78405425
        25
 1.00057358
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
+ 88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)</pre>
```

```
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> residuals.glm(object = modello, type = "response")
            1
                           2
-0.0020334895 \ -0.0103128513 \ -0.0187033936 \ -0.0111968589 \ -0.0190987716
                       7
           6
                                        8
                                                      9
-0.0040529588 \quad 0.0064239884 \quad 0.0253149298 \quad -0.0206104280 \quad 0.0628788951
           11
                          12
                                        13
                                                       14
 0.0573520700 0.0437876678 -0.0551545725 0.0031188816 0.0531911753
           16
                          17
                                         18
                                                       19
-0.0409654825 \ -0.0460470931 \ \ 0.0312622502 \ \ 0.0129127734 \ -0.0200914343
           21
                          22
                                         23
                                                       24
-0.0081744371 -0.0139759836 -0.0064977884 -0.0076672869 0.0005732538
```

resid()

• Package: stats

• Input:

```
object modello di regressione logit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- Description: residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \hline \end{array}$$

• Example 1:

• Example 2:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> resid(object = modello, type = "pearson")
                    2
                              3
                                          4
-0.87529996 -1.44362837 -1.33137848 -0.74525548 -0.91034225 -0.15901761
             8 9 10 11 12
 0.23139551 \quad 0.79874716 \quad -0.53576012 \quad 1.40235004 \quad 1.21724831 \quad 0.91957777
       13 14 15 16 17
-1.09953015 0.06647053 1.21159801 -1.09387707 -1.28431127 1.01110426
           20 21 22 23 24
        19
 0.54989436 \ -0.94424085 \ -0.50685539 \ -1.00250029 \ -0.52208706 \ -0.82783987
 0.77568558
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> resid(object = modello, type = "working")
                               3
-1.00203763 \ -1.01042031 \ -1.01905988 \ -0.41336424 \ -0.48212701 \ -0.07089826
                        9
                    8
                                   10
                                               11
 0.07938086 \quad 0.22704866 \quad -0.13926878 \quad 0.33629857 \quad 0.25835047 \quad 0.17881393
        13
                  14 15 16 17
-0.22141017 0.01336452 0.26283804 -0.24965088 -0.36552096 0.33713195
                   20 21 22 23 24
 0.19514514 \ -0.43506531 \ -0.25760272 \ -0.64783388 \ -0.44626460 \ -0.78405425
        25
 1.00057358
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
+ 88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)</pre>
```

```
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))
> resid(object = modello, type = "response")
-0.0020334895 -0.0103128513 -0.0187033936 -0.0111968589 -0.0190987716
                             8
                                          9
                       7
-0.0040529588 \quad 0.0064239884 \quad 0.0253149298 \quad -0.0206104280 \quad 0.0628788951
         11
               12
                           13
0.0573520700 0.0437876678 -0.0551545725 0.0031188816 0.0531911753
              17
                           18
                                                19
         16
-0.0409654825 -0.0460470931 0.0312622502 0.0129127734 -0.0200914343
              22
                                    23
-0.0081744371 \ -0.0139759836 \ -0.0064977884 \ -0.0076672869 \ \ 0.0005732538
```

weighted.residuals()

• Package: stats

• Input:

obj modello di regressione logit con k-1 variabili esplicative ed n unità

• **Description:** residui pesati

• Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "logit"))
> weighted.residuals(obj = modello)
                                                   5
                              3
                                        4
-1.2372312 \ -2.0363101 \ -1.8739732 \ -0.8043827 \ -0.9953320 \ -0.1607163 \ \ 0.2289532
                  9
                            10
                                       11
                                                  12
 0.7780252 \; -0.5441548 \quad 1.3675388 \quad 1.2016944 \quad 0.9162826 \; -1.0982255 \quad 0.0665090
       15 16 17 18 19
                                                      2.0
1.2375553 - 1.0695134 - 1.2358120  1.0633044  0.5665503 - 0.8912577 - 0.4883964
       22 23 24
                                  25
-0.9195743 -0.4900070 -0.7461893 1.0968278
```

weights()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

• Description: pesi iniziali

• Formula:

$$n_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> weights(object = modello)
  1
            3
                 4
                     5
                           6
                                7
                                     8
                                          9
                                              10
                                                   11
                                                        12
                                                             13
                                                                  14
                                                                       15
                                                                            16
 376
     200
           93 120
                     90
                          88 105 111
                                        100
                                              93 100
                                                       108
                                                             99 106 105 117
           19
                     21
                           22
                               23
                                   24
  17
      18
               20
                                         25
  98
       97 120 102 122 111
                              94 114 1049
```

df.residual()

• Package: stats

• Input:

object modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-k

• Examples:

hatvalues()

• Package: stats

• Input:

model modello di regressione logit con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> hatvalues(model = modello)
                               3
                                          4
0.04171418 \ 0.06450180 \ 0.04237196 \ 0.06756306 \ 0.06125644 \ 0.07048903 \ 0.09582267
                   9
                             10
                                         11
                                                    12
                                                               13
0.11108936\ 0.10521957\ 0.09873284\ 0.10390681\ 0.10889885\ 0.09837709\ 0.10624609
       1.5
                             17
                                        18
                                                    19
                                                               20
                  16
0.10699575 0.11922484 0.09653421 0.08871474 0.09799217 0.07184963 0.07200939
                  23
                             24
                                         25
0.05359644 0.03640349 0.03483536 0.04565424
```

cooks.distance()

• Package: stats

• Input:

modeln modello di regressione logit con n-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> cooks.distance(model = modello)
0.0174011270 \ 0.0768009809 \ 0.0409503781 \ 0.0215799628 \ 0.0288029684 \ 0.0010315088
                        8
                                      9
                                                  10
                                                                11
0.0031379129 0.0448481919 0.0188614178 0.1195191319 0.0958663105 0.0579850735
                                     15
          13
                       14
                                                   16
                                                                17
0.0731523657 \ 0.0002938362 \ 0.0984796718 \ 0.0919482890 \ 0.0975367746 \ 0.0546070811
                        20
                                     21
                                                   22
                                                                23
0.0182095530\ 0.0371812046\ 0.0107408856\ 0.0300692243\ 0.0053432866\ 0.0128138673
0.0150803356
```

cookd()

• Package: car

• Input:

model modello di regressione logit con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "logit"))</pre>
> cookd(model = modello)
                                       3
0.0174011270 \ 0.0768009809 \ 0.0409503781 \ 0.0215799628 \ 0.0288029684 \ 0.0010315088
                                      9
                        8
                                                   10
                                                                 11
0.0031379129\ 0.0448481919\ 0.0188614178\ 0.1195191319\ 0.0958663105\ 0.0579850735
          13
                        14
                                     15
                                                   16
                                                                 17
0.0731523657 \ 0.0002938362 \ 0.0984796718 \ 0.0919482890 \ 0.0975367746 \ 0.0546070811
          19
                        20
                                     21
                                                   22
                                                                 23
0.0182095530 \ 0.0371812046 \ 0.0107408856 \ 0.0300692243 \ 0.0053432866 \ 0.0128138673
0.0150803356
```

Capitolo 18

Regressione Probit

18.1 Simbologia

$$\Phi^{-1}(\pi_i) = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} \qquad Y_i \sim \text{Bin}(\pi_i, n_i) \quad \forall i = 1, 2, \dots, n$$

- numero di successi: $y_i \quad \forall i = 1, 2, ..., n$
- numero di prove: $n_i \quad \forall i = 1, 2, \ldots, n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: $\,k\,$
- ullet numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w=(w_1,\,w_2,\,\ldots,\,w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \mathrm{diag}(w_1^{-1},\,w_2^{-1},\,\ldots,\,w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- distanza di Cook: $cd_i = \left(e_i^P\right)^2 \frac{h_i}{k\left(1-h_i\right)^2} \quad \forall i=1,2,\ldots,n$
- stime IWLS: $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{\beta}} = \hat{\beta} / s_{\hat{\beta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \; \hat{\beta}_j} = \frac{(X^T \, W^{-1} \, X)_{i, \, j}^{-1}}{s_{\hat{\beta}_i} \; s_{\hat{\beta}_j}} \quad \forall \, i, j \, = \, 1, \, 2, \, \ldots, \, k$
- residui di devianza: $e_i = \operatorname{sign}(y_i \hat{y}_i) \sqrt{2 \left[y_i \log \left(\frac{y_i}{\hat{y}_i} + C_{i1} \right) + (n_i y_i) \log \left(\frac{n_i y_i}{n_i \hat{y}_i} + C_{i2} \right) \right]}$ $\forall i = 1, 2, \dots, n$ dove $C_{i1} = 0.5 \left(1 \operatorname{sign}(y_i) \right) / \hat{y}_i$ e $C_{i2} = 0.5 \left(1 \operatorname{sign}(n_i y_i) \right) / (n_i \hat{y}_i)$
- residui standard: $rstandard_i = e_i / \sqrt{1 h_i} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = sign(y_i \hat{y}_i) \sqrt{e_i^2 + h_i (e_i^P)^2 / (1 h_i)} \quad \forall i = 1, 2, ..., n$
- residui di *Pearson*: $e_i^P = \frac{y_i n_i \, \hat{\pi}_i}{\sqrt{n_i \, \hat{\pi}_i \, (1 \hat{\pi}_i)}} \quad \forall i = 1, 2, \ldots, n$
- \bullet residui di lavoro: $e^W_i=rac{y_i-n_i\,\hat{\pi}_i}{n_i\,\hat{\pi}_i\,(1-\hat{\pi}_i)} \ \ \forall i=1,\,2,\,\ldots,\,n$
- residui di riposta: $e_i^R = y_i / n_i \hat{\pi}_i \quad \forall i = 1, 2, \dots, n$
- log-verosimiglianza binomiale: $\hat{\ell} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{\hat{y}_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{\hat{y}_i}{n_i} \right) \right]$
- valori adattati: $\hat{\pi}_i = \Phi\left(X_i\,\hat{eta}\right) \quad \forall\, i\,=\,1,\,2,\,\ldots,\,n$

- numero di successi attesi: $\hat{y}_i = n_i \, \hat{\pi}_i \quad \forall \, i = 1, \, 2, \, \ldots, \, n$
- log-verosimiglianza binomiale modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{y_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{y_i}{n_i} \right) \right]$
- devianza residua: $D=2\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^{n}e_{i}^{2}$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza binomiale modello nullo: $\hat{\ell}_{nullo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log (\hat{\pi}) + (n_i y_i) \log (1 \hat{\pi}) \right]$
- valori adattati modello nullo: $\hat{\pi} = \sum_{j=1}^n y_j / \sum_{j=1}^n n_j \quad \forall i = 1, 2, ..., n$
- numero di successi attesi modello nullo: $\hat{y}_i = n_i \hat{\pi} \quad \forall i = 1, 2, ..., n$
- devianza residua modello nullo: $D_{nullo} = 2 \left(\hat{\ell}_{saturo} \hat{\ell}_{nullo} \right)$
- gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = \Phi^{-1}(\hat{\pi})$

18.2 Stima

glm()

- Package: stats
- Input:

formula modello di regressione probit con k-1 variabili esplicative ed n unità family = binomial(link="probit") famiglia e link del modello x = TRUE matrice del modello

- Description: analisi di regressione probit
- Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y proporzione di successi
x matrice del modello

• Formula:

coefficients $\hat{\beta}_j \quad \forall \, j \, = \, 1, \, 2, \, \dots, \, k$ residuals $e^W_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$ fitted.values $\hat{\pi}_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$

```
rank
                                                              k
linear.predictors
                                                            X \hat{\beta}
deviance
                                                             D
aic
                                                         -2\hat{\ell}+2k
null.deviance
                                                           D_{nullo}
weights
                                                  w_i \quad \forall i = 1, 2, \ldots, n
prior.weights
                                                  n_i \quad \forall i = 1, 2, \ldots, n
df.residual
                                                           n-k
df.null
                                                           n-1
У
                                               y_i / n_i \quad \forall i = 1, 2, \ldots, n
                                                             X
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"),</pre>
  x = TRUE)
> modello$coefficients
(Intercept)
 -11.818942 0.907823
> modello$residuals
                                     3
-0.269418259 \ -0.348625023 \ -0.389983219 \ -0.122461411 \ -0.200141756 \ -0.046955683
```

10

22

16

11

23

17

9

21

15

 $-0.118836507 \quad 0.054563070 \quad 0.218884846 \quad -0.056123202 \quad -0.104260350$

 $-0.002815914 \quad 0.058111915 \ -0.133324114 \quad 0.140220542 \quad 0.121793589 \quad 0.102604272$

 $0.136088873 \ -0.179601128 \ -0.148819712 \ -0.409392515 \ -0.420317445 \ -0.792660540$

> modello\$fitted.values

1.3

19

0.229368032

8

20

14

> modello\$rank

[1] 2

> modello\$linear.predictors

```
4
-3.4578913 -2.5500682 -2.2141737 -1.9872179 -1.7602621 -1.5333064 -1.3063506
                   10
                            11
                                      12
              9
                                                    13
-1.0793948 -0.8524391 -0.6254833 -0.3985275 -0.1715718 0.0553840 0.2823398
      15
          16
                   17
                            18
                                      19
                                               20
0.5092955 \quad 0.7362513 \quad 0.9632071 \quad 1.1901628 \quad 1.4171186 \quad 1.6440744 \quad 1.8710301
      22
          23 24
                              25
2.0979859 2.3249417 2.5518974 4.1405878
```

> modello\$deviance

[1] 22.88743

> modello\$aic

[1] 110.9392

> modello\$null.deviance

[1] 3693.884

> modello\$weights

> modello\$prior.weights

```
3
              5
                      7 8 9
1
           4
                  6
                                10 11 12 13 14 15 16
         120
              90
                  88 105 111 100
                                93 100 108 99 106 105 117
          20
              21
                  22
                     23
17
   18
      19
                         24
                            2.5
   97 120 102 122 111
                     94 114 1049
98
```

> modello\$df.residual

[1] 23

```
> modello$df.null
```

> modello\$y

[1] 24

> modello\$x

```
(Intercept)
            1 9.21
1
            1 10.21
3
            1 10.58
4
            1 10.83
5
           1 11.08
6
           1 11.33
7
           1 11.58
8
           1 11.83
9
           1 12.08
10
            1 12.33
            1 12.58
11
           1 12.83
12
           1 13.08
13
14
           1 13.33
15
           1 13.58
           1 13.83
16
17
           1 14.08
18
           1 14.33
            1 14.58
19
           1 14.83
20
           1 15.08
21
22
           1 15.33
23
           1 15.58
24
           1 15.83
25
            1 17.58
attr(,"assign")
[1] 0 1
```

summary.glm()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- **Description:** analisi di regressione probit
- Output:

deviance devianza residua aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS

• Formula:

deviance Daic $-2\hat{\ell}+2k$ df.residual n-knull.deviance D_{nullo} df.null n-1deviance.resid $e_i \quad \forall i = 1, 2, \ldots, n$ coefficients $\hat{\beta}_{j}$ $s_{\hat{\beta}_{i}}$ $z_{\hat{\beta}_{i}}$ p-value = $2\Phi(-|z_{\hat{\beta}_{i}}|)$ $\forall j = 1, 2, ..., k$ cov.unscaled $(X^T W^{-1} X)^{-1}$ cov.scaled $(X^T W^{-1} X)^{-1}$ correlation $r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i, j = 1, 2, \dots, k$

• Examples:

[1] 23

```
> res$null.deviance
   [1] 3693.884
   > res$df.null
   [1] 24
   > res$deviance.resid
   -0.45247119 -1.46964542 -1.58456196 -0.51743600 -0.90056726 -0.22725786
                                      9
                        8
                                                 10
                                                              11
   -0.01668127 \quad 0.38801751 \quad -0.95408459 \quad 0.98731872 \quad 0.93524092 \quad 0.84356724
            13
                        14
                                     15
                                                 16
                                                              17
                                                                          18
   -0.94228925 0.44328398 1.75392860 -0.43468903 -0.67959504 1.46607128
            19
                        20
                                    21
                                         22
                                                     23
    0.84691681 \ -0.81514441 \ -0.62908579 \ -1.26364877 \ -0.95089420 \ -1.40845258
    0.19062911
   > res$coefficients
                  Estimate Std. Error z value
                                                    Pr(>|z|)
    (Intercept) -11.818942 0.38701607 -30.53863 8.004674e-205
                 0.907823 0.02955339 30.71807 3.265395e-207
   > res$cov.unscaled
                (Intercept)
    (Intercept) 0.14978143 -0.0113907885
               -0.01139079 0.0008734026
   > res$cov.scaled
                (Intercept)
    (Intercept) 0.14978143 -0.0113907885
               -0.01139079 0.0008734026
   > res$correlation
               (Intercept)
   (Intercept) 1.0000000 -0.9959042
                -0.9959042 1.0000000
glm.fit()
  • Package: stats
  • Input:
       x matrice del modello
       y proporzione di successi
       weights numero di prove
       family = binomial(link="probit") famiglia e link del modello
```

• Description: analisi di regressione probit

• Output:

727

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua
df.null gradi di libertà devianza residua modello nullo
y proporzione di successi

• Formula:

coefficients	$\hat{\beta}_j \forall j = 1, 2, \dots, k$
residuals	$e_i^W \forall j = 1, 2, \dots, n$
fitted.values	$\hat{\pi}_i \forall i = 1, 2, \dots, n$
rank	k
linear.predictors	$X\hat{eta}$
deviance	D
aic	$-2\hat\ell+2k$
null.deviance	D_{nullo}
weights	$w_i \forall i = 1, 2, \dots, n$
prior.weights	$n_i orall i = 1, 2, \ldots, n$
df.residual	n-k
df.null	n-1
У	$y_i / n_i \forall i = 1, 2, \ldots, n$

```
(Intercept)
 -11.818942 0.907823
> res$residuals
 [1] -0.269418259 -0.348625023 -0.389983219 -0.122461411 -0.200141756
  [6] \quad -0.046955683 \quad -0.002815914 \quad 0.058111915 \quad -0.133324114 \quad 0.140220542 
[11] \quad 0.121793589 \quad 0.102604272 \quad -0.118836507 \quad 0.054563070 \quad 0.218884846
[16] -0.056123202 -0.104260350 0.228143827 0.136088873 -0.179601128
[21] -0.148819712 -0.409392515 -0.420317445 -0.792660540 0.229368032
> res$fitted.values
 [1] 0.0002722105 0.0053850922 0.0134084170 0.0234491271 0.0391816851
 [6] 0.0626001924 0.0957166773 0.1402058751 0.1969852207 0.2658269508
[11] 0.3451206813 0.4318871004 0.5220837266 0.6111585001 0.6947274541
[16] 0.7692111098 0.8322781892 0.8830088002 0.9217758718 0.9499195786
[21] 0.9693295476 0.9820468044 0.9899624601 0.9946430973 0.9999826792
> res$rank
[1] 2
> res$linear.predictors
 [1] -3.4578913 -2.5500682 -2.2141737 -1.9872179 -1.7602621 -1.5333064
  \lceil 7 \rceil \ -1.3063506 \ -1.0793948 \ -0.8524391 \ -0.6254833 \ -0.3985275 \ -0.1715718 
                                                              1.1901628
     0.0553840 0.2823398 0.5092955 0.7362513 0.9632071
[13]
     1.4171186 1.6440744 1.8710301 2.0979859 2.3249417
                                                              2.5518974
[19]
[25] 4.1405878
> res$deviance
[1] 22.88743
> res$aic
[1] 110.9392
> res$null.deviance
[1] 3693.884
> res$weights
     1.4104551 8.9094789 8.3105953 16.0744621 17.1659357 22.7386165
 [7] 35.0406005 45.7076709 48.6499031 51.2857797 60.0774428 68.0228376
[13] 62.9551408 65.5510152 60.7937719 60.9999288 44.1838731 36.2494196
[19] 35.5528528 22.8652682 19.7074642 12.2829626 6.7637482 5.0575577
[25] 0.3453737
> res$prior.weights
[1] 376 200
                 93 120
                          90
                               88 105 111
                                              100
                                                    93 100 108 99 106 105
[16] 117
           98
                 97 120 102 122
                                    111
                                          94
                                               114 1049
> res$df.residual
[1] 23
```

```
> res$df.null

[1] 24

> res$y

[1] 0.00000000 0.00000000 0.00000000 0.01666667 0.02222222 0.05681818
[7] 0.09523810 0.15315315 0.16000000 0.31182796 0.39000000 0.47222222
[13] 0.47474747 0.63207547 0.77142857 0.75213675 0.80612245 0.92783505
[19] 0.94166667 0.93137255 0.95901639 0.96396396 0.97872340 0.98245614
[25] 1.00000000
```

vcov()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$(X^T W^{-1} X)^{-1}$$

• Examples:

coef()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_j \quad \forall j = 1, 2, \ldots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description**: previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

```
fit x_0^T \, \hat{\beta} se.fit \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}
```

• Examples:

predict()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto
se.fit standard error delle stime

• Formula:

fit $x_0^T \, \hat{\beta}$ se.fit $\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$

```
1
-10.63877
> res$se.fit
[1] 0.3487713
```

fitted()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> fitted(object = modello)
0.0002722105 \ 0.0053850922 \ 0.0134084170 \ 0.0234491271 \ 0.0391816851 \ 0.0626001924
                                                  1.0
                                                               11
0.0957166773 \ 0.1402058751 \ 0.1969852207 \ 0.2658269508 \ 0.3451206813 \ 0.4318871004
         13
                       14
                                    15
                                                  16
                                                               17
0.5220837266 0.6111585001 0.6947274541 0.7692111098 0.8322781892 0.8830088002
         19
                       20
                                          22
                                    21
                                                              2.3
0.9217758718 0.9499195786 0.9693295476 0.9820468044 0.9899624601 0.9946430973
0.9999826792
```

fitted.values()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> fitted.values(object = modello)
0.0002722105 \ 0.0053850922 \ 0.0134084170 \ 0.0234491271 \ 0.0391816851 \ 0.0626001924
                                    9
                                                 10
                                                               11
0.0957166773 \ 0.1402058751 \ 0.1969852207 \ 0.2658269508 \ 0.3451206813 \ 0.4318871004
          13
                       14
                                    15
                                                  16
                                                               17
0.5220837266 0.6111585001 0.6947274541 0.7692111098 0.8322781892 0.8830088002
                       20
                                    21
                                                  22
          19
                                                               23
0.9217758718 0.9499195786 0.9693295476 0.9820468044 0.9899624601 0.9946430973
0.9999826792
```

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$$

• Examples:

18.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** log-verosimiglianza binomiale
- Formula:

 $\hat{\ell}$

• Examples:

AIC()

- Package: stats
- Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

 $-2\hat{\ell}+2k$

• Examples:

durbin.watson()

- Package: car
- Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> durbin.watson(model = modello)
 lag Autocorrelation D-W Statistic p-value
  1
           0.3108564
                         1.367754 0.07
 Alternative hypothesis: rho != 0
> res <- durbin.watson(model = modello)</pre>
> res$dw
[1] 1.367754
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2k$$

• Examples:

[1]

2.0000 110.9392

deviance()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

• Examples:

[1] 22.88743

anova()

• Package: stats

• Input:

nullo modello nullo di regressione probit con n unità modello modello di regressione probit con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
          12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
          14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
   > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
          88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
   > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
          108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
          1049)
   > nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "probit"))</pre>
   > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
   > anova(nullo, modello, test = "Chisq")
   Analysis of Deviance Table
   Model 1: cbind(y, Total - y) ~ 1
   Model 2: cbind(y, Total - y) \sim x
     Resid. Df Resid. Dev Df Deviance P(>|Chi|)
             24
                  3693.9
   2
             23
                     22.9 1 3671.0
                                             0.0
   > res <- anova(nullo, modello, test = "Chisq")</pre>
   > res$"Resid. Df"
   [1] 24 23
   > res$"Resid. Dev"
    [1] 3693.88357 22.88743
   > res$Df
   [1] NA 1
   > res$Deviance
   [1]
            NA 3670.996
   > res$"P(>|Chi|)"
    [1] NA 0
drop1()
  • Package: stats
  • Input:
       object modello di regressione probit con k-1 variabili esplicative ed n unità
       test = "Chisq"
  • Description: submodels
  • Output:
       Df differenza tra gradi di libertà
       Deviance differenza tra devianze residue
       AIC indice AIC
       LRT valore empirico della statistica \chi^2
```

Pr(Chi) p-value

• Formula:

Df $\underbrace{\frac{1,\,1,\,\ldots,\,1}{k-1\,\mathrm{volte}}}_{}$ Deviance $D,\,D_{-x_j}\ \, \forall\,j\,=\,1,\,2,\,\ldots,\,k-1$

dove D_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

AIC

$$-2\hat{\ell} + 2k, -2\hat{\ell}_{-x_i} + 2(k-1) \quad \forall j = 1, 2, \dots, k-1$$

dove $\hat{\ell}_{-x_j}$ rappresenta la log-verosimiglianza binomiale del modello eliminata la variabile esplicativa x_j . LRT

 $c_j = D_{-x_j} - D \quad \forall j = 1, 2, \dots, k - 1$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))
> drop1(object = modello, test = "Chisq")
Single term deletions
Model:
cbind(y, Total - y) \sim x
                   AIC
       Df Deviance
                                   Pr(Chi)
                             LRT
              22.9 110.9
            3693.9 3779.9 3671.0 < 2.2e-16 ***
Х
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- drop1(object = modello, test = "Chisq")</pre>
> res$Df
[1] NA 1
> res$Deviance
      22.88743 3693.88357
[1]
> res$AIC
[1] 110.9392 3779.9354
> res$LRT
[1]
         NA 3670.996
> res$"Pr(Chi)"
[1] NA 0
```

add1()

• Package: stats

• Input:

object modello nullo di regressione probit scope modello di regressione probit con k-1 variabili esplicative ed n unità test = "Chisq"

- Description: submodels
- Output:

Df differenza tra gradi di libertà Deviance differenza tra devianze residue AIC indice AIC LRT valore empirico della statistica χ^2 Pr (Chi) p-value

• Formula:

Df

$$\underbrace{1, 1, \ldots, 1}_{k-1 \text{ volte}}$$

Deviance

$$D_{nullo}, D_{x_j} \quad \forall j = 1, 2, \ldots, k-1$$

dove D_{x_i} rappresenta la devianza residua del modello con la sola variabile esplicativa x_i .

AIC

$$-2\,\hat{\ell}_{nullo} + 2, \, -2\,\hat{\ell}_{x_j} + 4 \quad \forall j = 1, \, 2, \, \dots, \, k-1$$

dove $\hat{\ell}_{x_j}$ rappresenta la log-verosimiglianza binomiale del modello con la sola variabile esplicativa x_j .

LRT

$$c_j = D_{nullo} - D_{x_j} \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "probit"))</pre>
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> add1(object = nullo, scope = modello, test = "Chisq")
Single term additions
Model:
cbind(y, Total - y) \sim 1
      Df Deviance
                     AIC
                            LRT Pr(Chi)
            3693.9 3779.9
<none>
             22.9 110.9 3671.0 < 2.2e-16 ***
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

18.4 Diagnostica

rstandard()

• Package: stats

• Input:

modello di regressione probit con <math>k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

 $rstandard_i \quad \forall i = 1, 2, \ldots, n$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> rstandard(model = modello)
                                 3
-0.45702180 -1.52667261 -1.62930398 -0.54193441 -0.93825575 -0.23771437
         7
                    8
                                9
                                           1.0
                                                        11
-0.01766532 0.41236338 -1.00506815 1.03243853 0.97758496 0.88234046
                               15
                                           16 17
-0.98089408 0.46342071 1.83843010 -0.46019719 -0.71464732 1.54273708
        19
                    20
                                21
                                    22
                                               23
 0.90128028 \; -0.85537455 \; -0.66151138 \; -1.31119403 \; -0.97372238 \; -1.43789404
 0.19126471
```

rstandard.glm()

• Package: stats

• Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> rstandard.glm(model = modello)
                                   3
-0.45702180 \ -1.52667261 \ -1.62930398 \ -0.54193441 \ -0.93825575 \ -0.23771437
                                  9
                                                          11
                      8
                                              10
-0.01766532 0.41236338 -1.00506815 1.03243853 0.97758496 0.88234046
         13
                                 15
                                              16
                                                          17
-0.98089408 0.46342071 1.83843010 -0.46019719 -0.71464732
                                                              1.54273708
                                21
                                                          23
                     20
                                     22
        19
 0.90128028 \; -0.85537455 \; -0.66151138 \; -1.31119403 \; -0.97372238 \; -1.43789404
```

rstudent()

• Package: stats

0.19126471

• Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, ..., n
```

rstudent.glm()

• Package: stats

• Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> rstudent.glm(model = modello)
                              3
                                         4
-0.45475250 \ -1.49850744 \ -1.60724034 \ -0.53954353 \ -0.93261903 \ -0.23741494
        7
            8 9 10 11
-0.01766390 0.41295880 -1.00258075 1.03395739 0.97836584 0.88258097
       13
            14 15 16 17
-0.98094312 0.46328566 1.83403420 -0.46061490 -0.71601113 1.53357601
                      21
                                 22
 0.89694597 - 0.85968513 - 0.66475785 - 1.32462729 - 0.98094946 - 1.45532717
 0.19094718
```

residuals.default()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals.default(object = modello)
                                    3
-0.269418259 -0.348625023 -0.389983219 -0.122461411 -0.200141756 -0.046955683
          7
                      8
                              9
                                               10 11
-0.002815914 \quad 0.058111915 \quad -0.133324114 \quad 0.140220542 \quad 0.121793589 \quad 0.102604272
         13
                      14
                          15
                                       16
                                                    17
-0.118836507 0.054563070 0.218884846 -0.056123202 -0.104260350 0.228143827
                                                   23
                                               22
         19
                      2.0
                                  2.1
0.136088873 \ -0.179601128 \ -0.148819712 \ -0.409392515 \ -0.420317445 \ -0.792660540
0.229368032
```

residuals()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \hline e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \hline \end{array}$$

• Example 1:

```
1 2 3 4 5 6
-0.45247119 -1.46964542 -1.58456196 -0.51743600 -0.90056726 -0.22725786
7 8 9 10 11 12
-0.01668127 0.38801751 -0.95408459 0.98731872 0.93524092 0.84356724
13 14 15 16 17 18
-0.94228925 0.44328398 1.75392860 -0.43468903 -0.67959504 1.46607128
19 20 21 22 23 24
0.84691681 -0.81514441 -0.62908579 -1.26364877 -0.95089420 -1.40845258
25
0.19062911
```

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals(object = modello, type = "pearson")
                   2
                             3
-0.31996722 -1.04060064 -1.12424645 -0.49098375 -0.82922265 -0.22390818
               8 9 10 11
-0.01666883 0.39287973 -0.92992864 1.00417656 0.94401767 0.84623856
           14 15 16 17
       13
-0.94289966 0.44176215 1.70665302 -0.43833594 -0.69302839
                     21 22 23
        19
                  20
 0.81144619 - 0.85880990 - 0.66065634 - 1.43479933 - 1.09312733 - 1.78261348
 0.13479572
```

• Example 3:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals(object = modello, type = "working")
                                    3
-0.269418259 \ -0.348625023 \ -0.389983219 \ -0.122461411 \ -0.200141756 \ -0.046955683
                       8
                                    9
                                                10
                                                     11
-0.002815914
             0.058111915 - 0.133324114 \quad 0.140220542 \quad 0.121793589 \quad 0.102604272
         13
                      14
                                  15 16 17
-0.118836507
             0.054563070 \quad 0.218884846 \quad -0.056123202 \quad -0.104260350 \quad 0.228143827
                       20 21 22 23
          19
 0.136088873 \ -0.179601128 \ -0.148819712 \ -0.409392515 \ -0.420317445 \ -0.792660540
 0.229368032
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals(object = modello, type = "response")
                         2
-2.722105e-04 -5.385092e-03 -1.340842e-02 -6.782460e-03 -1.695946e-02
                        7
                               8
-5.782011e-03 -4.785821e-04 1.294728e-02 -3.698522e-02 4.600101e-02
                                     13
          11
                        12
                                           14
 4.487932e-02 4.033512e-02 -4.733625e-02 2.091697e-02
                                                        7.670112e-02
                                      18
                        17
          16
                                                   19
                                                                  20
-1.707436e-02 -2.615574e-02 4.482625e-02 1.989079e-02 -1.854703e-02
                       22
                                     23
                                                   24
-1.031315e-02 -1.808284e-02 -1.123906e-02 -1.218696e-02 1.732085e-05
```

residuals.glm()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

• **Description:** residui

• Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \\ e_i^R & \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals.glm(object = modello, type = "pearson")
                   2
                             3
-0.31996722 -1.04060064 -1.12424645 -0.49098375 -0.82922265 -0.22390818
              8 9 10 11
-0.01666883 0.39287973 -0.92992864 1.00417656 0.94401767 0.84623856
           14 15 16
                                           17
       13
-0.94289966 0.44176215 1.70665302 -0.43833594 -0.69302839
                     21 22 23
        19
                  20
 0.81144619 - 0.85880990 - 0.66065634 - 1.43479933 - 1.09312733 - 1.78261348
 0.13479572
```

• Example 3:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals(object = modello, type = "working")
                                    3
-0.269418259 \ -0.348625023 \ -0.389983219 \ -0.122461411 \ -0.200141756 \ -0.046955683
                      8
                                    9
                                                10
                                                     11
-0.002815914
             0.058111915 - 0.133324114 \quad 0.140220542 \quad 0.121793589 \quad 0.102604272
         13
                      14
                                  15 16 17
-0.118836507
             0.054563070 \quad 0.218884846 \quad -0.056123202 \quad -0.104260350 \quad 0.228143827
                      20 21 22 23
          19
 0.136088873 \ -0.179601128 \ -0.148819712 \ -0.409392515 \ -0.420317445 \ -0.792660540
 0.229368032
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals.glm(object = modello, type = "response")
                         2
-2.722105e-04 -5.385092e-03 -1.340842e-02 -6.782460e-03 -1.695946e-02
                         7
                                      8
-5.782011e-03 -4.785821e-04 1.294728e-02 -3.698522e-02 4.600101e-02
                                       13
           11
                        12
                                                    14
 4.487932e-02 4.033512e-02 -4.733625e-02 2.091697e-02
                                                         7.670112e-02
                         17
                                       18
           16
                                                     19
                                                                   20
-1.707436e-02 -2.615574e-02 4.482625e-02 1.989079e-02 -1.854703e-02
                        22
                                       23
                                                    24
-1.031315e-02 -1.808284e-02 -1.123906e-02 -1.218696e-02 1.732085e-05
```

resid()

- Package: stats
- Input:

```
object modello di regressione probit con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \hline \end{array}$$

• Example 1:

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> resid(object = modello, type = "pearson")
                   2
                             3
-0.31996722 -1.04060064 -1.12424645 -0.49098375 -0.82922265 -0.22390818
            8 9 10 11
-0.01666883 0.39287973 -0.92992864 1.00417656 0.94401767 0.84623856
           14 15 16
                                           17
       13
-0.94289966 0.44176215 1.70665302 -0.43833594 -0.69302839
                     21 22 23
       19
                  20
 0.81144619 - 0.85880990 - 0.66065634 - 1.43479933 - 1.09312733 - 1.78261348
 0.13479572
```

• Example 3:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> residuals(object = modello, type = "working")
                                     3
-0.269418259 \ -0.348625023 \ -0.389983219 \ -0.122461411 \ -0.200141756 \ -0.046955683
                       8
                                    9
                                                10
                                                     11
-0.002815914
             0.058111915 - 0.133324114 \quad 0.140220542 \quad 0.121793589 \quad 0.102604272
         13
                      14
                                   15 16 17
-0.118836507
             0.054563070 \quad 0.218884846 \quad -0.056123202 \quad -0.104260350 \quad 0.228143827
                       20 21 22 23
          19
 0.136088873 \ -0.179601128 \ -0.148819712 \ -0.409392515 \ -0.420317445 \ -0.792660540
 0.229368032
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> resid(object = modello, type = "response")
-2.722105e-04 -5.385092e-03 -1.340842e-02 -6.782460e-03 -1.695946e-02
                                  8
                                               9
-5.782011e-03 -4.785821e-04 1.294728e-02 -3.698522e-02 4.600101e-02
         11
               12
                           13
                                       14
 4.487932e-02 4.033512e-02 -4.733625e-02 2.091697e-02 7.670112e-02
             17 18
                                      19
        16
-1.707436e-02 -2.615574e-02 4.482625e-02 1.989079e-02 -1.854703e-02
         21 22 23 24
-1.031315e-02 -1.808284e-02 -1.123906e-02 -1.218696e-02 1.732085e-05
```

weighted.residuals()

• Package: stats

• Input:

obj modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> weighted.residuals(obj = modello)
                     2
                                3
-0.45247119 -1.46964542 -1.58456196 -0.51743600 -0.90056726 -0.22725786
               8 9
         7
                                                       11
                                           10
-0.01668127 \quad 0.38801751 \ -0.95408459 \quad 0.98731872 \quad 0.93524092 \quad 0.84356724
                                   16
                                               17
                        15
        13
                    14
-0.94228925 0.44328398 1.75392860 -0.43468903 -0.67959504
                                                           1.46607128
        19
                    20
                                2.1
                                           2.2
                                                       23
 0.84691681 \ -0.81514441 \ -0.62908579 \ -1.26364877 \ -0.95089420 \ -1.40845258
 0.19062911
```

weights()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: pesi iniziali
- Formula:

$$n_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> weights(object = modello)
             3
                                7
                                           9
   1
                      5
                           6
                                      8
                                               10
                                                   11
                                                         12
                                                              13
                                                                  14
                                                                        15
                                                                             16
 376
      200
            93 120
                      90
                           88
                              105
                                    111
                                         100
                                               93 100
                                                       108
                                                              99 106 105 117
  17
      18
           19
                20
                      21
                           22
                                23
                                    24
                                          25
  98
      97 120 102 122 111
                               94 114 1049
```

df.residual()

• Package: stats

• Input:

object modello di regressione probit con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

$$n-k$$

• Examples:

[1] 23

hatvalues()

• Package: stats

• Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

• Description: valori di leva

• Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> hatvalues(model = modello)
                                   3
                                                4
0.019815055 0.073312514 0.054167532 0.088367447 0.078723832 0.086040497
                                   9
                                               10
                                                           11
0.108307417 \ \ 0.114593994 \ \ 0.098879759 \ \ 0.085494466 \ \ 0.084753718 \ \ 0.085956150
         13
                     14
                                  15
                                               16
                                                           17
0.077164589 0.085016631 0.089815211 0.107785168 0.095690966 0.096919770
        19
                     20
                                  21
                                               22
                                                           23
0.116997841 \ 0.091852356 \ 0.095632164 \ 0.071207217 \ 0.046338837 \ 0.040531561
         25
0.006635307
```

cooks.distance()

• Package: stats

• Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

• Description: distanza di Cook

• Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
1 2 3 4 5 6
1.055748e-03 4.622210e-02 3.826517e-02 1.281613e-02 3.188885e-02 2.582016e-03
7 8 9 10 11 12
1.892378e-05 1.128148e-02 5.265155e-02 5.154131e-02 4.508303e-02 3.683821e-02
13 14 15 16 17 18
4.027824e-02 9.908879e-03 1.578888e-01 1.300781e-02 2.810019e-02 1.121110e-01
19 20 21 22 23 24
4.940191e-02 4.107159e-02 2.551732e-02 8.496473e-02 3.044167e-02 6.995461e-02
25
6.108938e-05
```

cookd()

- Package: car
- Input:

model modello di regressione probit con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x < -c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "probit"))</pre>
> cookd(model = modello)
                                     3
1.055748e-03 4.622210e-02 3.826517e-02 1.281613e-02 3.188885e-02 2.582016e-03
                                    9
                       8
                                                 1.0
                                                             11
1.892378e-05 1.128148e-02 5.265155e-02 5.154131e-02 4.508303e-02 3.683821e-02
          13
                       14
                                    15
                                                 16
                                                              17
4.027824e-02 9.908879e-03 1.578888e-01 1.300781e-02 2.810019e-02 1.121110e-01
          19
                      20
                                    21
                                                 22
                                                             23
4.940191e-02 4.107159e-02 2.551732e-02 8.496473e-02 3.044167e-02 6.995461e-02
6.108938e-05
```

Capitolo 19

Regressione Log-log complementare

19.1 Simbologia

$$\log(-\log(1-\pi_i)) = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} \qquad Y_i \sim \text{Bin}(\pi_i, n_i) \quad \forall i = 1, 2, \dots, n$$

- numero di successi: $y_i \quad \forall i = 1, 2, ..., n$
- numero di prove: $n_i \quad \forall i = 1, 2, \ldots, n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- ullet numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w=(w_1,\,w_2,\,\ldots,\,w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X (X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, ..., n$
- distanza di Cook: $cd_i = \left(e_i^P\right)^2 \frac{h_i}{k\left(1-h_i\right)^2} \quad \forall i=1,2,\ldots,n$
- stime IWLS: $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{eta}} = \hat{eta} \, / \, s_{\hat{eta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \, \hat{\beta}_j} = \frac{(X^T \, W^{-1} \, X)_{i,\,j}^{-1}}{s_{\hat{\beta}_i} \, s_{\hat{\beta}_j}} \quad \forall i,j = 1,\,2,\,\ldots,\,k$
- residui di devianza: $e_i = \operatorname{sign}(y_i \hat{y}_i) \sqrt{2 \left[y_i \log \left(\frac{y_i}{\hat{y}_i} + C_{i1} \right) + (n_i y_i) \log \left(\frac{n_i y_i}{n_i \hat{y}_i} + C_{i2} \right) \right]}$ $\forall i = 1, 2, \dots, n \qquad \text{dove} \quad C_{i1} = 0.5 \left(1 \operatorname{sign}(y_i) \right) / \hat{y}_i \quad \text{e} \quad C_{i2} = 0.5 \left(1 \operatorname{sign}(n_i y_i) \right) / \left(n_i \hat{y}_i \right)$
- residui standard: $rstandard_i = e_i / \sqrt{1 h_i} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = sign(y_i \hat{y}_i) \sqrt{e_i^2 + h_i (e_i^P)^2 / (1 h_i)} \quad \forall i = 1, 2, ..., n$
- residui di *Pearson*: $e_i^P = \frac{y_i n_i \, \hat{\pi}_i}{\sqrt{n_i \, \hat{\pi}_i \, (1 \hat{\pi}_i)}} \quad \forall i = 1, 2, \ldots, n$
- residui di lavoro: $e_i^W=\frac{y_i-n_i\,\hat{\pi}_i}{n_i\,\hat{\pi}_i\,(1-\hat{\pi}_i)} \quad \forall i=1,\,2,\,\ldots,\,n$
- residui di riposta: $e_i^R = y_i / n_i \hat{\pi}_i \quad \forall i = 1, 2, \dots, n$
- log-verosimiglianza binomiale: $\hat{\ell} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{\hat{y}_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{\hat{y}_i}{n_i} \right) \right]$
- valori adattati: $\hat{\pi}_i = 1 \exp\left(-\exp\left(X_i\,\hat{\beta}\right)\right) \quad \forall i = 1,\,2,\,\ldots,\,n$

- numero di successi attesi: $\hat{y}_i = n_i \, \hat{\pi}_i \quad \forall i = 1, 2, \dots, n$
- log-verosimiglianza binomiale modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{y_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{y_i}{n_i} \right) \right]$
- devianza residua: $D=2\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^{n}e_{i}^{2}$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza binomiale modello nullo: $\hat{\ell}_{nullo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log (\hat{\pi}) + (n_i y_i) \log (1 \hat{\pi}) \right]$
- valori adattati modello nullo: $\hat{\pi} = \sum_{j=1}^n y_j / \sum_{j=1}^n n_j \quad \forall i = 1, 2, ..., n$
- numero di successi attesi modello nullo: $\hat{y}_i = n_i \hat{\pi} \quad \forall i = 1, 2, ..., n$
- devianza residua modello nullo: $D_{nullo} = 2 \left(\hat{\ell}_{saturo} \hat{\ell}_{nullo} \right)$
- gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = \log(-\log(1-\hat{\pi}))$

19.2 Stima

glm()

- Package: stats
- Input:

formula modello di regressione log-log complementare con k-1 variabili esplicative ed n unità family = binomial(link="cloglog") famiglia e link del modello x = TRUE matrice del modello

- **Description:** analisi di regressione log-log complementare
- Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y proporzione di successi
x matrice del modello

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall \, j \, = \, 1, \, 2, \, \dots, \, k$$
 residuals
$$e^W_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$$
 fitted.values
$$\hat{\pi}_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$$

```
rank
                                                              k
linear.predictors
                                                            X \hat{\beta}
deviance
                                                             D
aic
                                                         -2\hat{\ell}+2k
null.deviance
                                                           D_{nullo}
weights
                                                 w_i \quad \forall i = 1, 2, \ldots, n
prior.weights
                                                  n_i \quad \forall i = 1, 2, \ldots, n
df.residual
                                                           n-k
df.null
                                                           n-1
У
                                               y_i / n_i \quad \forall i = 1, 2, \ldots, n
                                                             X
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"),</pre>
+ x = TRUE)
> modello$coefficients
(Intercept)
-12.9851164 0.9530076
> modello$residuals
```

```
3
-1.00747570 \ -1.01954272 \ -1.02795778 \ -0.77915832 \ -0.77094233 \ -0.49169111
        7
                   8
                               9
                                          10
                                                      11
-0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224
       13 14 15
                                  16
                                                     17
0.08077108 0.28820279 0.42232719 0.13526781 0.06070359
       19
                   20
                               21
                                          22
                                                      23
0.12113911 \ -0.19177587 \ -0.30930043 \ -0.93966307 \ -1.91670214 \ -7.49366104
1.0000000
```

> modello\$fitted.values

> modello\$rank

[1] 2

> modello\$linear.predictors

```
1
                  2
                           3
-4.20791595 -3.25490830 -2.90229547 -2.66404356 -2.42579164 -2.18753973
                           9
                                    10
                 8
                                               11
-1.94928782 \ -1.71103591 \ -1.47278400 \ -1.23453209 \ -0.99628017 \ -0.75802826
      13
          14
                    15
                                     16
                                         17
-0.51977635 -0.28152444 -0.04327253 0.19497939 0.43323130 0.67148321
          20 21 22 23
      19
0.90973512 1.14798703 1.38623894 1.62449086 1.86274277 2.10099468
      25
3.76875806
```

> modello\$deviance

[1] 118.8208

> modello\$aic

[1] 206.8726

> modello\$null.deviance

[1] 3693.884

> modello\$weights

```
1 2 3 4 5 6
5.551912e+00 7.568498e+00 4.966316e+00 8.071724e+00 7.609886e+00 9.329133e+00
7 8 9 10 11 12
1.391005e+01 1.829764e+01 2.040002e+01 2.331378e+01 3.052613e+01 3.967311e+01
13 14 15 16 17 18
4.309158e+01 5.356986e+01 5.997599e+01 7.287294e+01 6.342595e+01 6.111898e+01
19 20 21 22 23 24
6.738325e+01 4.527553e+01 3.641982e+01 1.797138e+01 6.226026e+00 2.146377e+00
25
2.329248e-13
```

> modello\$prior.weights

```
9 10 11
                  6 7 8
 1
    2
       3
           4
              5
                                       12 13 14 15 16
       93 120
              90
                 88 105 111 100
                               93 100 108 99 106 105 117
376
  200
          20
             21
                  22 23 24
                            25
    97 120 102 122 111
                     94 114 1049
```

> modello\$df.residual

```
[1] 23
> modello$df.null
[1] 24
> modello$y
                              3
                                         4
0.00000000 \ 0.00000000 \ 0.000000000 \ 0.01666667 \ 0.02222222 \ 0.05681818 \ 0.09523810
                       10 11
                                                 12
                                                            13
0.15315315 \ \ 0.16000000 \ \ 0.31182796 \ \ 0.39000000 \ \ 0.47222222 \ \ 0.47474747 \ \ 0.63207547 
                     17 18 19 20
                 16
0.77142857 \ 0.75213675 \ 0.80612245 \ 0.92783505 \ 0.94166667 \ 0.93137255 \ 0.95901639
       22 23 24
0.96396396 0.97872340 0.98245614 1.00000000
> modello$x
   (Intercept)
            1 9.21
2
            1 10.21
3
            1 10.58
4
            1 10.83
5
            1 11.08
6
            1 11.33
7
            1 11.58
8
            1 11.83
9
            1 12.08
10
            1 12.33
            1 12.58
11
12
            1 12.83
13
            1 13.08
14
            1 13.33
15
           1 13.58
           1 13.83
16
17
            1 14.08
18
            1 14.33
19
            1 14.58
20
            1 14.83
21
           1 15.08
22
           1 15.33
23
           1 15.58
24
            1 15.83
25
            1 17.58
attr(,"assign")
```

summary.glm()

[1] 0 1

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- Description: analisi di regressione log-log complementare
- Output:

deviance devianza residua
aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS

• Formula:

deviance Daic $-2\hat{\ell}+2k$ df.residual n-knull.deviance D_{nullo} df.null n-1deviance.resid $e_i \quad \forall i = 1, 2, \ldots, n$ coefficients $\hat{\beta}_j$ $s_{\hat{\beta}_i}$ $z_{\hat{\beta}_i}$ p-value = $2\Phi(-|z_{\hat{\beta}_i}|)$ $\forall j = 1, 2, ..., k$ cov.unscaled $(X^T W^{-1} X)^{-1}$ cov.scaled $(X^T W^{-1} X)^{-1}$ correlation $r_{\hat{\beta}_i \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$

• Examples:

> res\$df.residual

```
[1] 23
```

> res\$null.deviance

```
[1] 3693.884
```

> res\$df.null

[1] 24

> res\$deviance.resid

```
2
                     3
-3.344811e+00 -3.928580e+00 -3.195443e+00 -2.625263e+00 -2.501326e+00
          7 8 9 10
       6
-1.632697e+00 -1.183466e+00 -3.479272e-01 -1.146176e+00 1.287445e+00
                    13
       11
          12
                              14
1.722479e+00 2.078066e+00 5.293632e-01 2.125777e+00 3.393960e+00
       16
           17
                           18
                                    19
                                               20
1.175000e+00 4.892018e-01 2.127667e+00 1.046796e+00 -1.190182e+00
      21 22 23 24 25
-1.608195e+00 -2.739982e+00 -2.588698e+00 -3.552944e+00 6.825317e-07
```

> res\$coefficients

```
Estimate Std. Error z value Pr(>|z|)
(Intercept) -12.9851164 0.42631012 -30.45932 9.016015e-204
x 0.9530076 0.03133172 30.41671 3.303275e-203
```

> res\$cov.unscaled

```
(Intercept) x
(Intercept) 0.1817403 -0.0133057991
x -0.0133058 0.0009816765
```

> res\$cov.scaled

```
(Intercept) x
(Intercept) 0.1817403 -0.0133057991
x -0.0133058 0.0009816765
```

> res\$correlation

```
(Intercept) x
(Intercept) 1.0000000 -0.9961646
x -0.9961646 1.0000000
```

glm.fit()

• Package: stats

• Input:

x matrice del modello

y proporzione di successi

weights numero di prove

family = binomial(link="cloglog") famiglia e link del modello

• Description: analisi di regressione log-log complementare

• Output:

coefficients stime IWLS

residuals residui di lavoro

fitted.values valori adattati

rank rango della matrice del modello

linear.predictors predittori lineari

deviance devianza residua

aic indice AIC

null.deviance devianza residua modello nullo

weights pesi IWLS

prior.weights pesi iniziali

df.residual gradi di libertà devianza residua

df.null gradi di libertà devianza residua modello nullo

y proporzione di successi

• Formula:

coefficients

 $\hat{\beta}_j \quad \forall j = 1, 2, \ldots, k$

residuals

 $e_i^W \quad \forall i = 1, 2, \dots, n$

fitted.values

 $\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$

rank

k

linear.predictors

 $X \hat{\beta}$

deviance

D

aic

 $-2\hat{\ell}+2k$

null.deviance

 D_{nullo}

weights

 $w_i \quad \forall i = 1, 2, \ldots, n$

prior.weights

 $n_i \quad \forall i = 1, 2, \ldots, n$

df.residual

n-k

df.null

n-1

У

```
y_i / n_i \quad \forall i = 1, 2, \ldots, n
```

```
• Examples:
```

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> X <- model.matrix(object = modello)</pre>
> res <- glm.fit(x = X, y/Total, weights = Total, family = binomial(link = "cloglog"))</pre>
> res$coefficients
(Intercept)
-12.9851164
             0.9530076
> res$residuals
 [1] -1.00747570 -1.01954272 -1.02795778 -0.77915832 -0.77094233 -0.49169111
  \lceil 7 \rceil \ -0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224 
[13] \quad 0.08077108 \quad 0.28820279 \quad 0.42232719 \quad 0.13526781 \quad 0.06070359 \quad 0.24992698
[19] \quad 0.12113911 \quad -0.19177587 \quad -0.30930043 \quad -0.93966307 \quad -1.91670214 \quad -7.49366104
[25] 1.00000000
> res$fitted.values
 [1] 0.01476722 0.03784946 0.05341742 0.06729466 0.08461277 0.10612777
 [7] 0.13270442 0.16529635 0.20489911 0.25246255 0.30874773 0.37411551
[13] \quad 0.44824630 \quad 0.52981661 \quad 0.61620640 \quad 0.70337481 \quad 0.78609705 \quad 0.85873787
[19] 0.91656310 0.95722673 0.98168030 0.99375413 0.99840579 0.99971820
[25] 1.00000000
> res$rank
[1] 2
> res$linear.predictors
 [1] -4.20791595 -3.25490830 -2.90229547 -2.66404356 -2.42579164 -2.18753973
 [7] -1.94928782 -1.71103591 -1.47278400 -1.23453209 -0.99628017 -0.75802826
[13] \ -0.51977635 \ -0.28152444 \ -0.04327253 \ \ 0.19497939 \ \ 0.43323130 \ \ 0.67148321
     0.90973512 1.14798703 1.38623894 1.62449086 1.86274277 2.10099468
[19]
[25] 3.76875806
> res$deviance
[1] 118.8208
> res$aic
[1] 206.8726
> res$null.deviance
[1] 3693.884
```

```
> res$weights
 [1] 5.551912e+00 7.568498e+00 4.966316e+00 8.071724e+00 7.609886e+00
 [6] 9.329133e+00 1.391005e+01 1.829764e+01 2.040002e+01 2.331378e+01
[11] 3.052613e+01 3.967311e+01 4.309158e+01 5.356986e+01 5.997599e+01
[16] 7.287294e+01 6.342595e+01 6.111898e+01 6.738325e+01 4.527553e+01
[21] 3.641982e+01 1.797138e+01 6.226026e+00 2.146377e+00 2.329248e-13
> res$prior.weights
                                                       100 108
[1] 376 200
                93 120
                         90
                              88 105
                                             100
                                                   93
                                                                 99 106 105
                                        111
                97 120 102 122 111
                                        94
                                             114 1049
[16]
     117
            98
> res$df.residual
[1] 23
> res$df.null
[1] 24
> res$y
 [1] 0.00000000 0.00000000 0.00000000 0.01666667 0.02222222 0.05681818
 [7] 0.09523810 0.15315315 0.16000000 0.31182796 0.39000000 0.47222222
[13] 0.47474747 0.63207547 0.77142857 0.75213675 0.80612245 0.92783505
[19] 0.94166667 0.93137255 0.95901639 0.96396396 0.97872340 0.98245614
```

vcov()

• Package: stats

[25] 1.00000000

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$(X^T W^{-1} X)^{-1}$$

coef()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

coefficients()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

predict.glm()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> predict.glm(object = modello, newdata = data.frame(x = 1.3),
     se.fit = TRUE)
$fit
-11.74621
$se.fit
[1] 0.3857516
$residual.scale
[1] 1
> res <- predict.glm(object = modello, newdata = data.frame(x = 1.3),
     se.fit = TRUE)
> res$fit
-11.74621
> res$se.fit
[1] 0.3857516
```

predict()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))
> predict.glm(object = modello, newdata = data.frame(x = 1.3),
     se.fit = TRUE)
$fit
-11.74621
$se.fit
[1] 0.3857516
$residual.scale
[1] 1
> res <- predict.glm(object = modello, newdata = data.frame(x = 1.3),
     se.fit = TRUE)
> res$fit
-11.74621
> res$se.fit
[1] 0.3857516
```

fitted()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x < -c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> fitted(object = modello)
                    2
                                           4
0.01476722\ 0.03784946\ 0.05341742\ 0.06729466\ 0.08461277\ 0.10612777\ 0.13270442
                              10
                                          11
                                                     12
0.16529635 0.20489911 0.25246255 0.30874773 0.37411551 0.44824630 0.52981661
                              17
                                          18
                                                     19
                                                                20
                   16
0.61620640 0.70337481 0.78609705 0.85873787 0.91656310 0.95722673 0.98168030
                   2.3
                              24
0.99375413 0.99840579 0.99971820 1.00000000
```

fitted.values()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

```
0.01476722\ 0.03784946\ 0.05341742\ 0.06729466\ 0.08461277\ 0.10612777\ 0.13270442
         9
                         10
                                   11
                                      12
                                                       13
0.16529635 0.20489911 0.25246255 0.30874773 0.37411551 0.44824630 0.52981661
                         17
      15
               16
                                   18
                                             19
                                                       20
0.61620640\ 0.70337481\ 0.78609705\ 0.85873787\ 0.91656310\ 0.95722673\ 0.98168030
      22 23 24
0.99375413 0.99840579 0.99971820 1.00000000
```

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \dots, k$$

• Examples:

19.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** log-verosimiglianza binomiale
- Formula:

 $\hat{\ell}$

AIC()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

$$-2\hat{\ell}+2k$$

• Examples:

durbin.watson()

• Package: car

• Input:

model modello di regressione cloglog con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> durbin.watson(model = modello)
 lag Autocorrelation D-W Statistic p-value
  1
           0.7610921 0.3836592
 Alternative hypothesis: rho != 0
> res <- durbin.watson(model = modello)</pre>
> res$dw
[1] 0.3836592
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2k$$

• Examples:

deviance()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

anova()

• Package: stats

• Input:

nullo modello nullo di regressione log-log complementare con n unità modello di regressione log-log complementare con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi_{df}^2 \geq c)$$

Deviance

```
Analysis of Deviance Table
    Model 1: cbind(y, Total - y) \sim 1
    Model 2: cbind(y, Total - y) \sim x
      Resid. Df Resid. Dev Df Deviance P(>|Chi|)
             24
                     3693.9
    2
              23
                      118.8 1 3575.1
                                                 0.0
    > res <- anova(nullo, modello, test = "Chisq")</pre>
    > res$"Resid. Df"
    [1] 24 23
    > res$"Resid. Dev"
    [1] 3693.8836 118.8208
    > res$Df
    [1] NA 1
    > res$Deviance
               NA 3575.063
    [1]
    > res$"P(>|Chi|)"
    [1] NA 0
drop1()
  • Package: stats
  • Input:
        object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità
        test = "Chisq"
  • Description: submodels
  • Output:
        Df differenza tra gradi di libertà
        Deviance differenza tra devianze residue
        AIC indice AIC
        LRT valore empirico della statistica \chi^2
        Pr(Chi) p-value
  • Formula:
        Df
```

dove D_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

 $D, D_{-x_i} \quad \forall j = 1, 2, \ldots, k-1$

AIC

$$-2\hat{\ell} + 2k, -2\hat{\ell}_{-x_i} + 2(k-1) \quad \forall j = 1, 2, \dots, k-1$$

dove $\hat{\ell}_{-x_j}$ rappresenta la log-verosimiglianza binomiale del modello eliminata la variabile esplicativa x_j .

LRT

$$c_j = D_{-x_j} - D \quad \forall j = 1, 2, \dots, k-1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> drop1(object = modello, test = "Chisq")
Single term deletions
Model:
cbind(y, Total - y) \sim x
      Df Deviance
                   AIC
                             LRT Pr(Chi)
            118.8 206.9
<none>
            3693.9 3779.9 3575.1 < 2.2e-16 ***
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- drop1(object = modello, test = "Chisq")</pre>
> res$Df
[1] NA 1
> res$Deviance
[1] 118.8208 3693.8836
> res$AIC
[1] 206.8726 3779.9354
> res$LRT
[1]
         NA 3575.063
> res$"Pr(Chi)"
[1] NA 0
```

add1()

• Package: stats

• Input:

object modello nullo di regressione log-log complementare scope modello di regressione log-log complementare con k-1 variabili esplicative ed n unità test = "Chisq"

- Description: submodels
- Output:

Df differenza tra gradi di libertà Deviance differenza tra devianze residue AIC indice AIC LRT valore empirico della statistica χ^2 Pr (Chi) p-value

• Formula:

Df

$$\underbrace{1,\,1,\,\ldots,\,1}_{k-1\,\text{volte}}$$

Deviance

$$D_{nullo}, D_{x_i} \quad \forall j = 1, 2, \dots, k-1$$

dove D_{x_i} rappresenta la devianza residua del modello con la sola variabile esplicativa x_i .

AIC

$$-2\,\hat{\ell}_{nullo} + 2, \, -2\,\hat{\ell}_{x_j} + 4 \quad \forall j = 1, \, 2, \, \dots, \, k-1$$

dove $\hat{\ell}_{x_j}$ rappresenta la log-verosimiglianza binomiale del modello con la sola variabile esplicativa x_j .

LRT

$$c_j = D_{nullo} - D_{x_j} \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "cloglog"))</pre>
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> add1(object = nullo, scope = modello, test = "Chisq")
Single term additions
Model:
cbind(y, Total - y) \sim 1
      Df Deviance
                     AIC
                            LRT Pr(Chi)
            3693.9 3779.9
<none>
            118.8 206.9 3575.1 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
> res <- add1(object = nullo, scope = modello, test = "Chisq")
> res$Df

[1] NA     1
> res$Deviance
[1] 3693.8836     118.8208
> res$AIC

[1] 3779.9354     206.8726
> res$LRT

[1]      NA     3575.063
> res$"Pr(Chi)"

[1] NA     0
```

19.4 Diagnostica

rstandard()

• Package: stats

• Input:

modell modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, \ldots, n
```

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
               12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
                14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
                 88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
                 108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))
> rstandard(model = modello)
-3.546647e + 00 \quad -4.126490e + 00 \quad -3.278516e + 00 \quad -2.722320e + 00 \quad -2.574884e + 000e +
                                                                        7
                                                                                                                  8
                                                                                                                                                          9
                                 6
-1.682464e+00 -1.228898e+00 -3.625140e-01 -1.189748e+00 1.332682e+00
                                               12
                                                                                       13
  1.787005e+00 2.161401e+00 5.487673e-01 2.212887e+00 3.545180e+00
                                                                                                                                 19
                               16
                                                                                       18
                                               17
   1.243292e+00 5.172376e-01 2.269593e+00 1.144446e+00 -1.279947e+00
                                                                        22
                                                                                                                23
                                                                                                                                24
```

rstandard.glm()

• Package: stats

• Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))
> rstandard.glm(model = modello)
                                      3
-3.546647e+00 -4.126490e+00 -3.278516e+00 -2.722320e+00 -2.574884e+00
                        7
                                     8
-1.682464e+00 -1.228898e+00 -3.625140e-01 -1.189748e+00 1.332682e+00
          11
                       12
                                     13
                                                  14
 1.787005e+00 2.161401e+00 5.487673e-01 2.212887e+00
                                                     3.545180e+00
                              18
                                                 19
          16
                 17
 1.243292e+00 5.172376e-01 2.269593e+00 1.144446e+00 -1.279947e+00
          2.1
                22
                            23
                                         24
-1.728057e+00 -2.857626e+00 -2.633515e+00 -3.577897e+00 6.825317e-07
```

rstudent()

• Package: stats

• Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, ..., n
```

```
-3.447960e + 00 \quad -4.030684e + 00 \quad -3.238407e + 00 \quad -2.694633e + 00 \quad -2.554716e + 00 \quad -2.694633e + 00 \quad -2.694636e + 00 \quad -2.694666e + 00 \quad -2.69466e + 00 \quad -2.694666e + 00 \quad -2.69666e + 00 \quad -2.69666e + 00 \quad -2.69666e + 00 \quad -2.69666e + 00 \quad -2
                                                                                                                                                                                  8
                                                                                                                                                                                                                                                                                                                   9
-1.674902e+00 -1.225072e+00 -3.622277e-01 -1.187261e+00 1.334804e+00
                                                                11
                                                                                                         12
                                                                                                                                                                                  13
                                                                                                                                                                                                                                                                                                                          14
      1.789702e+00 2.163690e+00 5.488287e-01 2.211575e+00 3.534607e+00
                                                                                                                                                                                    18
                                                                16
                                                                                                      17
                                                                                                                                                                                                                                                                                                                          19
      1.241017e+00 5.165991e-01 2.247950e+00 1.135287e+00 -1.295065e+00
                                                                                                                                                                                                                                                                     24
                                                                                                  22
                                                                                                                                                                                  23
                                                                  2.1
-1.767784e+00 -2.983221e+00 -2.738686e+00 -3.784579e+00 6.825317e-07
```

rstudent.glm()

• Package: stats

• Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> rstudent.glm(model = modello)
-3.447960e+00 -4.030684e+00 -3.238407e+00 -2.694633e+00 -2.554716e+00
          6
               7
                                8 9
-1.674902e+00 -1.225072e+00 -3.622277e-01 -1.187261e+00 1.334804e+00
         11
              12 13 14
 1.789702e+00 2.163690e+00 5.488287e-01 2.211575e+00 3.534607e+00
              17
                           18
                                       19
         16
 1.241017e+00 5.165991e-01 2.247950e+00 1.135287e+00 -1.295065e+00
                                                24
          2.1
                      2.2
                                   2.3
-1.767784e+00 -2.983221e+00 -2.738686e+00 -3.784579e+00 6.825317e-07
```

residuals.default()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals.default(object = modello)
                                 3
                                              4
-1.00747570 -1.01954272 -1.02795778 -0.77915832 -0.77094233 -0.49169111
         7
             8 9 10
                                                        11
-0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224
                        15
       13
            14
                                    16
                                                        17
 0.08077108 \quad 0.28820279 \quad 0.42232719 \quad 0.13526781 \quad 0.06070359 \quad 0.24992698
        19
                    2.0
                               2.1
                                            2.2
                                                        2.3
 0.12113911 \ -0.19177587 \ -0.30930043 \ -0.93966307 \ -1.91670214 \ -7.49366104
 1.00000000
```

residuals()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{l} \texttt{type} = \texttt{"deviance"} \\ \\ e_i \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \texttt{type} = \texttt{"pearson"} \\ \\ e_i^P \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \texttt{type} = \texttt{"working"} \\ \\ e_i^W \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \texttt{type} = \texttt{"response"} \\ \\ e_i^R \quad \forall i=1,\,2,\,\ldots,\,n \end{array}$$

• Example 1:

```
1 2 3 4 5
-3.344811e+00 -3.928580e+00 -3.195443e+00 -2.625263e+00 -2.501326e+00
6 7 8 9 10
-1.632697e+00 -1.183466e+00 -3.479272e-01 -1.146176e+00 1.287445e+00
11 12 13 14 15
1.722479e+00 2.078066e+00 5.293632e-01 2.125777e+00 3.393960e+00
16 17 18 19 20
1.175000e+00 4.892018e-01 2.127667e+00 1.046796e+00 -1.190182e+00
21 22 23 24 25
-1.608195e+00 -2.739982e+00 -2.588698e+00 -3.552944e+00 6.825317e-07
```

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals(object = modello, type = "pearson")
                      2
-2.373963e+00 -2.804939e+00 -2.290887e+00 -2.213700e+00 -2.126766e+00
              7 8 9
         6
-1.501829e+00 -1.131643e+00 -3.444267e-01 -1.112389e+00 1.317832e+00
                                    14
                        13
         11
             12
1.758796e+00 2.106981e+00 5.302147e-01 2.109393e+00 3.270668e+00
                         18
                                     19
              17
         16
1.154719e+00 4.834456e-01 1.953903e+00 9.944108e-01 -1.290438e+00
             22 23
                                    24 25
         21
-1.866683e+00 -3.983806e+00 -4.783173e+00 -1.098075e+01 4.826228e-07
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals(object = modello, type = "working")
                                 3
-1.00747570 -1.01954272 -1.02795778 -0.77915832 -0.77094233 -0.49169111
         7
                    8
                                9
                                    10
                                                       11
-0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224
             14 15 16
                                                      17
 0.08077108 \quad 0.28820279 \quad 0.42232719 \quad 0.13526781 \quad 0.06070359 \quad 0.24992698
        19
                       21 22 23
                    20
 0.12113911 \ -0.19177587 \ -0.30930043 \ -0.93966307 \ -1.91670214 \ -7.49366104
 1.00000000
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals(object = modello, type = "response")
                          2
-1.476722e-02 -3.784946e-02 -5.341742e-02 -5.062800e-02 -6.239055e-02
                         7
           6
                                      8
-4.930959e-02 -3.746632e-02 -1.214320e-02 -4.489911e-02 5.936540e-02
                                       13
          11
                        12
                                                    14
 8.125227e-02 9.810671e-02 2.650118e-02 1.022589e-01
                                                        1.552222e-01
                         17
                                       18
                                                     19
          16
 4.876194e-02 2.002539e-02 6.909718e-02 2.510357e-02 -2.585418e-02
          21
                        22
                                      23
                                                     24
-2.266391e-02 -2.979016e-02 -1.968239e-02 -1.726206e-02 2.220446e-16
```

residuals.glm()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ e_i & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \hline \\ \text{type = "response"} \\ \\ \hline e_i^R & \forall i=1,2,\ldots,n \\ \\ \hline \end{array}$$

• Example 1:

```
-3.344811e+00 -3.928580e+00 -3.195443e+00 -2.625263e+00 -2.501326e+00
           7
                     8
                                9
-1.632697e+00 -1.183466e+00 -3.479272e-01 -1.146176e+00 1.287445e+00
        11
                 12 13 14
1.722479e+00 2.078066e+00 5.293632e-01 2.125777e+00 3.393960e+00
                 17
                                 19
                             18
1.175000e+00 4.892018e-01 2.127667e+00 1.046796e+00 -1.190182e+00
            22
                      23
                                24
-1.608195e+00 -2.739982e+00 -2.588698e+00 -3.552944e+00 6.825317e-07
```

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals.glm(object = modello, type = "pearson")
                      2
                                  3
-2.373963e+00 -2.804939e+00 -2.290887e+00 -2.213700e+00 -2.126766e+00
               7 8 9
         6
-1.501829e+00 -1.131643e+00 -3.444267e-01 -1.112389e+00 1.317832e+00
                                     14
                        13
         11
             12
1.758796e+00 2.106981e+00 5.302147e-01 2.109393e+00 3.270668e+00
                                      19
              17
                                 18
         16
1.154719e+00 4.834456e-01 1.953903e+00 9.944108e-01 -1.290438e+00
             22 23
                                    24 25
         21
-1.866683e+00 -3.983806e+00 -4.783173e+00 -1.098075e+01 4.826228e-07
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals.glm(object = modello, type = "working")
-1.00747570 -1.01954272 -1.02795778 -0.77915832 -0.77094233 -0.49169111
         7
                    8
                                9
                                    10
                                                       11
-0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224
             14 15 16
                                                      17
 0.08077108 \quad 0.28820279 \quad 0.42232719 \quad 0.13526781 \quad 0.06070359 \quad 0.24992698
        19
                       21 22 23
                    20
 0.12113911 \ -0.19177587 \ -0.30930043 \ -0.93966307 \ -1.91670214 \ -7.49366104
 1.00000000
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> residuals.glm(object = modello, type = "response")
-1.476722e-02 -3.784946e-02 -5.341742e-02 -5.062800e-02 -6.239055e-02
                         7
           6
                                       8
-4.930959e-02 -3.746632e-02 -1.214320e-02 -4.489911e-02 5.936540e-02
                                       13
           11
                         12
                                                     14
 8.125227e-02 9.810671e-02 2.650118e-02
                                          1.022589e-01
                                                         1.552222e-01
                         17
                                       18
                                                     19
          16
 4.876194e-02 2.002539e-02 6.909718e-02 2.510357e-02 -2.585418e-02
                         22
                                       23
                                                     24
-2.266391e-02 -2.979016e-02 -1.968239e-02 -1.726206e-02 2.220446e-16
```

resid()

- Package: stats
- Input:

```
object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \hline \\ \text{type = "response"} \\ \\ \\ e_i^R & \forall i=1,2,\ldots,n \\ \\ \hline \end{array}$$

• Example 1:

```
-3.344811e+00 -3.928580e+00 -3.195443e+00 -2.625263e+00 -2.501326e+00
             7
                      8
                                 9
-1.632697e+00 -1.183466e+00 -3.479272e-01 -1.146176e+00 1.287445e+00
        11
                   12 13
                                 14
1.722479e+00 2.078066e+00 5.293632e-01 2.125777e+00 3.393960e+00
                  17
                              18
                                         19
1.175000e+00 4.892018e-01 2.127667e+00 1.046796e+00 -1.190182e+00
                       23
                                  24
              22
-1.608195e+00 -2.739982e+00 -2.588698e+00 -3.552944e+00 6.825317e-07
```

> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,

• Example 2:

```
12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
    14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> resid(object = modello, type = "pearson")
                       2
                                   3
-2.373963e+00 -2.804939e+00 -2.290887e+00 -2.213700e+00 -2.126766e+00
               7 8
          6
-1.501829e+00 -1.131643e+00 -3.444267e-01 -1.112389e+00 1.317832e+00
                                      14
                         13
         11
             12
 1.758796e+00 2.106981e+00 5.302147e-01 2.109393e+00 3.270668e+00
                                       19
                     17
                                  18
         16
 1.154719e+00 4.834456e-01 1.953903e+00 9.944108e-01 -1.290438e+00
                  22 23
                                      24 25
         21
-1.866683e+00 -3.983806e+00 -4.783173e+00 -1.098075e+01 4.826228e-07
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> resid(object = modello, type = "working")
                                3
-1.00747570 -1.01954272 -1.02795778 -0.77915832 -0.77094233 -0.49169111
         7
                   8
                               9
                                   10
                                                      11
-0.30341626 \ -0.08051823 \ -0.24628470 \ \ 0.27292979 \ \ 0.31833027 \ \ 0.33451224
             14 15 16
                                                     17
 0.08077108 0.28820279 0.42232719 0.13526781 0.06070359
                                              23
        19
                       21 22
                    20
 0.12113911 \ -0.19177587 \ -0.30930043 \ -0.93966307 \ -1.91670214 \ -7.49366104
 1.00000000
```

• Example 4:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
+ 12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
+ 14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
```

```
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> resid(object = modello, type = "response")
-1.476722e-02 -3.784946e-02 -5.341742e-02 -5.062800e-02 -6.239055e-02
                                   8
-4.930959e-02 -3.746632e-02 -1.214320e-02 -4.489911e-02 5.936540e-02
         11
                12
                                13
                                             14
 8.125227e-02 9.810671e-02 2.650118e-02 1.022589e-01 1.552222e-01
               17
                          18
                                       19
         16
 4.876194e-02 2.002539e-02 6.909718e-02 2.510357e-02 -2.585418e-02
                          23
          21
               22
                                       24
-2.266391e-02 -2.979016e-02 -1.968239e-02 -1.726206e-02 2.220446e-16
```

weighted.residuals()

- Package: stats
- Input:

obj modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> weighted.residuals(obj = modello)
                         2
-3.344811e+00 -3.928580e+00 -3.195443e+00 -2.625263e+00 -2.501326e+00
                         7
           6
                                       8
-1.632697e+00 -1.183466e+00 -3.479272e-01 -1.146176e+00 1.287445e+00
                                      13
          11
                        12
                                                    1 4
 1.722479e+00 2.078066e+00 5.293632e-01 2.125777e+00 3.393960e+00
          16
                       17
                                      18
                                                    19
 1.175000e+00 4.892018e-01 2.127667e+00 1.046796e+00 -1.190182e+00
                        22
                                      23
                                                    24
-1.608195e+00 -2.739982e+00 -2.588698e+00 -3.552944e+00 6.825317e-07
```

weights()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: pesi iniziali
- Formula:

$$n_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> weights(object = modello)
             3
                                7
                                           9
   1
                      5
                           6
                                      8
                                               10
                                                   11
                                                         12
                                                              13
                                                                  14
                                                                        15
                                                                             16
 376
      200
            93 120
                      90
                           88
                              105
                                    111
                                         100
                                               93 100
                                                       108
                                                              99 106 105 117
  17
      18
           19
                20
                      21
                           22
                                23
                                    24
                                          25
  98
      97 120 102 122 111
                               94 114 1049
```

df.residual()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

$$n-k$$

hatvalues()

• Package: stats

• Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> hatvalues(model = modello)
           1
                                      3
                                                    4
                                                                  5
1.105792e-01 9.362145e-02 5.003535e-02 7.003405e-02 5.631849e-02 5.828511e-02
                                      9
                                                   10
                                                                11
                        8
7.257287e - 02 \ 7.885661e - 02 \ 7.190461e - 02 \ 6.673601e - 02 \ 7.091234e - 02 \ 7.562508e - 02
          13
                       14
                                     15
                                                   16
                                                                17
6.946860e-02 7.717999e-02 8.349045e-02 1.068393e-01 1.054680e-01 1.211568e-01
          19
                       20
                                     21
                                                  22
                                                                23
1.633692e-01 1.353446e-01 1.339136e-01 8.064188e-02 3.374658e-02 1.389985e-02
          25
4.030027e-15
```

cooks.distance()

• Package: stats

• Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

cookd()

- Package: car
- Input:

model modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x < -c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cloglog"))</pre>
> cookd(model = modello)
3.938916e-01 4.483042e-01 1.454921e-01 1.984188e-01 1.430242e-01 7.411901e-02
                                         9
                          8
                                                       1.0
                                                                      11
5.402610e-02 5.512482e-03 5.164813e-02 6.653361e-02 1.270601e-01 1.964540e-01
           13
                          14
                                         15
                                                       16
                                                                       17
1.127717e-02 2.016302e-01 5.316254e-01 8.928832e-02 1.540260e-02 2.994339e-01
           19
                          20
                                         21
                                                       22
                                                                      23
1.153996 {\text{e}} {\text{-}} 01 \ 1.507299 {\text{e}} {\text{-}} 01 \ 3.110377 {\text{e}} {\text{-}} 01 \ 7.571077 {\text{e}} {\text{-}} 01 \ 4.134756 {\text{e}} {\text{-}} 01 \ 8.617915 {\text{e}} {\text{-}} 01
4.693465e-28
```

Capitolo 20

Regressione di Cauchy

20.1 Simbologia

$$F_U^{-1}(\pi_i) = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1}$$
 $Y_i \sim \text{Bin}(\pi_i, n_i) \ \forall i = 1, 2, \dots, n$ $U \sim \text{Cauchy}(0, 1)$

- numero di successi: $y_i \quad \forall i = 1, 2, \ldots, n$
- numero di prove: $n_i \quad \forall i = 1, 2, ..., n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- ullet numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w = (w_1, w_2, \dots, w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, ..., n$
- distanza di Cook: $cd_i=\left(e_i^P\right)^2 \frac{h_i}{k\left(1-h_i\right)^2} \quad \forall\, i=1,\,2,\,\ldots,\,n$
- stime IWLS: $\hat{\beta} = \left(\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k\right)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{eta}} = \hat{eta} / s_{\hat{eta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \, \hat{\beta}_j} = \frac{(X^T \, W^{-1} \, X)_{i,\,j}^{-1}}{s_{\hat{\beta}_i} \, s_{\hat{\beta}_j}} \quad \forall \, i,j \, = \, 1,\,2,\,\ldots,\,k$
- residui di devianza: $e_i = \operatorname{sign}(y_i \hat{y}_i) \sqrt{2 \left[y_i \log \left(\frac{y_i}{\hat{y}_i} + C_{i1} \right) + (n_i y_i) \log \left(\frac{n_i y_i}{n_i \hat{y}_i} + C_{i2} \right) \right]}$ $\forall i = 1, 2, \ldots, n$ dove $C_{i1} = 0.5 \left(1 \operatorname{sign}(y_i) \right) / \hat{y}_i$ e $C_{i2} = 0.5 \left(1 \operatorname{sign}(n_i y_i) \right) / (n_i \hat{y}_i)$
- residui standard: $rstandard_i = e_i / \sqrt{1 h_i} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = sign\left(y_i \hat{y}_i\right) \sqrt{e_i^2 + h_i \left(e_i^P\right)^2 / (1 h_i)} \quad \forall i = 1, 2, \dots, n$
- residui di *Pearson*: $e_i^P = \frac{y_i n_i \, \hat{\pi}_i}{\sqrt{n_i \, \hat{\pi}_i \, (1 \hat{\pi}_i)}} \quad \forall i = 1, 2, \ldots, n$
- residui di lavoro: $e^W_i=rac{y_i-n_i\,\hat{\pi}_i}{n_i\,\hat{\pi}_i\,(1-\hat{\pi}_i)} \ \ \forall i=1,\,2,\,\ldots,\,n$
- residui di riposta: $e_i^R = y_i \hat{\mu}_i \quad \forall i = 1, 2, \dots, n$
- log-verosimiglianza binomiale: $\hat{\ell} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{\hat{y}_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{\hat{y}_i}{n_i} \right) \right]$
- valori adattati: $\hat{\pi}_i = F_U\left(X_i\,\hat{\beta}\right) \quad \forall\, i=1,\,2,\,\ldots,\,n$

- numero di successi attesi: $\hat{y}_i = n_i \, \hat{\pi}_i \quad \forall \, i = 1, \, 2, \, \ldots, \, n$
- log-verosimiglianza binomiale modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log \left(\frac{y_i}{n_i} \right) + (n_i y_i) \log \left(1 \frac{y_i}{n_i} \right) \right]$
- devianza residua: $D=2\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^{n}e_{i}^{2}$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza binomiale modello nullo: $\hat{\ell}_{nullo} = \sum_{i=1}^{n} \left[\log \binom{n_i}{y_i} + y_i \log (\hat{\pi}) + (n_i y_i) \log (1 \hat{\pi}) \right]$
- valori adattati modello nullo: $\hat{\pi} = \sum_{j=1}^n y_j / \sum_{j=1}^n n_j \quad \forall i = 1, 2, ..., n$
- numero di successi attesi modello nullo: $\hat{y}_i = n_i \hat{\pi} \quad \forall i = 1, 2, ..., n$
- devianza residua modello nullo: $D_{nullo} = 2 \left(\hat{\ell}_{saturo} \hat{\ell}_{nullo} \right)$
- gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = F_{II}^{-1}(\hat{\pi})$

20.2 Stima

glm()

- Package: stats
- Input:

formula modello di regressione di Cauchy con k-1 variabili esplicative ed n unità family = binomial(link="cauchit") famiglia e link del modello x = TRUE matrice del modello

- Description: analisi di regressione di Cauchy
- Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y proporzione di successi
x matrice del modello

• Formula:

coefficients $\hat{\beta}_j \quad \forall \, j \, = \, 1, \, 2, \, \dots, \, k$ residuals $e^W_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$ fitted.values $\hat{\pi}_i \quad \forall \, i \, = \, 1, \, 2, \, \dots, \, n$

```
rank
                                                             k
linear.predictors
                                                            X \hat{\beta}
deviance
                                                             D
aic
                                                        -2\hat{\ell}+2k
null.deviance
                                                          D_{nullo}
weights
                                                 w_i \quad \forall i = 1, 2, \ldots, n
prior.weights
                                                 n_i \quad \forall i = 1, 2, \ldots, n
df.residual
                                                           n-k
df.null
                                                           n-1
                                               y_i / n_i \quad \forall i = 1, 2, \ldots, n
Х
                                                             X
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"),</pre>
+ x = TRUE)
> modello$coefficients
(Intercept)
-33.544126 2.583834
> modello$residuals
                   2
                              3
                                         4
-9.8152648 -7.2558854 -6.3140094 -4.0086223 -3.2932991 -0.9917917 0.4226277
```

12

19

13

20

11

 $1.5498952 \quad 0.6272238 \quad 1.7058520 \quad 0.9553468 \quad 0.3321975 \quad -0.3474066 \quad -0.5728429$

 $-0.4855652 \ -2.0313711 \ -2.4430322 \ \ 0.6948164 \ \ 0.9814772 \ -0.2170523 \ \ 1.6310583$

18

> modello\$fitted.values

15

9

23

16 17

1.8963437 3.7327336 4.4091809 11.9357223

10

24

> modello\$rank

[1] 2

> modello\$linear.predictors

```
5
               2
                      3
                               4
-9.7470111 -7.1631766 -6.2071579 -5.5611993 -4.9152406 -4.2692820 -3.6233234
          9
                 10 11 12 13
-2.9773648 -2.3314062 -1.6854476 -1.0394890 -0.3935303 0.2524283 0.8983869
                         18
                 17
         16
                                  19
                                           20
     15
1.5443455 2.1903041 2.8362627 3.4822213 4.1281800 4.7741386 5.4200972
      22
             23
                      24
6.0660558 6.7120144 7.3579730 11.8796833
```

> modello\$deviance

[1] 180.8584

> modello\$aic

[1] 268.9102

> modello\$null.deviance

[1] 3693.884

> modello\$weights

```
3
0.13128604 0.17547429 0.12496388 0.22326973 0.24087950 0.35536805
                     8
                                 9
                                             1.0
                                                         11
0.68009289 \quad 1.24943550 \quad 2.17782383 \quad 4.51791817 \quad 12.69591273 \quad 34.80291036
        13
              14
                          15
                                      16
                                                         17
36.35987656 16.80244939 6.21201298 2.99536877 1.26102284 0.70343728
       19
                   2.0
                              21
                                            22
                                                         23
0.53414690 \quad 0.29731270 \quad 0.24487355 \quad 0.15967458 \quad 0.10010712 \quad 0.09232367
       2.5
0.20223732
```

> modello\$prior.weights

```
3
               5
                      7
1
           4
                  6
                          8 9 10 11
                                        12 13
                                                14 15 16
         120
               90
                  88 105
                        111 100
                                93 100 108 99 106 105 117
          20
               21
                  22
                     23
                            25
17
   18
       19
                         24
98
   97 120 102 122 111
                     94 114 1049
```

> modello\$df.residual

[1] 23

```
> modello$df.null
[1] 24
> modello$y
                                         5
               2
                        3
                                4
9
                  10 11 12 13
0.15315315 \ \ 0.160000000 \ \ 0.31182796 \ \ 0.390000000 \ \ 0.47222222 \ \ 0.47474747 \ \ 0.63207547
     15 16 17 18 19 20
0.77142857 \ 0.75213675 \ 0.80612245 \ 0.92783505 \ 0.94166667 \ 0.93137255 \ 0.95901639
     22 23 24 25
0.96396396 0.97872340 0.98245614 1.00000000
> modello$x
  (Intercept)
          1 9.21
1
          1 10.21
3
          1 10.58
4
          1 10.83
5
         1 11.08
6
         1 11.33
7
         1 11.58
8
         1 11.83
9
         1 12.08
10
          1 12.33
         1 12.58
11
         1 12.83
12
         1 13.08
13
14
         1 13.33
```

summary.glm()

[1] 0 1

• Package: stats

attr(,"assign")

• Input:

15

16 17

18

19

20

21 22

23

24

25

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

• Description: analisi di regressione di Cauchy

1 13.58 1 13.83

1 14.08

1 14.33 1 14.58

1 14.83

1 15.08

1 15.33

1 15.58

1 15.83

1 17.58

• Output:

deviance devianza residua aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS

• Formula:

deviance
$$D$$
 aic
$$-2\,\hat{\ell} + 2\,k$$
 df.residual
$$n-k$$
 null.deviance
$$D_{nullo}$$
 df.null
$$n-1$$
 deviance.resid
$$e_i \quad \forall i=1,2,\ldots,n$$
 coefficients
$$\hat{\beta}_j \quad s_{\hat{\beta}_j} \quad z_{\hat{\beta}_j} \quad p\text{-value} = 2\,\Phi(-\,|z_{\hat{\beta}_j}|) \qquad \forall j=1,2,\ldots,k$$
 cov.unscaled
$$(X^T\,W^{-1}\,X)^{-1}$$
 cov.scaled
$$(X^T\,W^{-1}\,X)^{-1}$$
 correlation
$$r_{\hat{\beta}_i\,\hat{\beta}_j} \quad \forall i,j=1,2,\ldots,k$$

• Examples:

[1] 23

• Output:

coefficients stime IWLS

```
> res$null.deviance
   [1] 3693.884
   > res$df.null
   [1] 24
   > res$deviance.resid
                    2
                        3
                                  4
   -4.9879493 -4.2499874 -3.1154320 -2.2134735 -1.8547635 -0.6138012 0.3429411
          8 9 10 11 12 13
   1.6292015 0.8969607 3.3340955 3.2290861 1.9359119 -2.0794099 -2.2707637
             16 17 18 19 20 21
   -1.1752053 \ -3.2150141 \ -2.5014455 \ \ 0.6008633 \ \ 0.7452777 \ -0.1175573 \ \ 0.8498527
             23
                       24
                                 2.5
    0.8002034 1.3186785 1.5146367 7.5396162
   > res$coefficients
               Estimate Std. Error z value
                                            Pr(>|z|)
   (Intercept) -33.544126 2.1690507 -15.46489 5.987702e-54
               > res$cov.unscaled
             (Intercept)
             4.7047808 -0.36150385
   (Intercept)
              -0.3615038 0.02782502
   > res$cov.scaled
             (Intercept)
             4.7047808 -0.36150385
   (Intercept)
              -0.3615038 0.02782502
   > res$correlation
             (Intercept)
               1.000000 -0.999138
   (Intercept)
              -0.999138 1.000000
glm.fit()
 • Package: stats
 • Input:
      x matrice del modello
      y proporzione di successi
      weights numero di prove
      family = binomial(link="cauchit") famiglia e link del modello
 • Description: analisi di regressione di Cauchy
```

795

residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua
df.null gradi di libertà devianza residua modello nullo
y proporzione di successi

• Formula:

coefficients $\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$ residuals $e_i^W \quad \forall i = 1, 2, \dots, n$ fitted.values $\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$ rank klinear.predictors $X \hat{\beta}$ deviance Daic $-2\hat{\ell}+2k$ null.deviance D_{nullo} weights $w_i \quad \forall i = 1, 2, \ldots, n$ prior.weights $n_i \quad \forall i = 1, 2, \ldots, n$ df.residual n-kdf.null n-1У $y_i / n_i \quad \forall i = 1, 2, \ldots, n$

• Examples:

```
(Intercept)
 -21.226395 1.631968
> res$residuals
  [1] \quad -1.00203763 \quad -1.01042031 \quad -1.01905988 \quad -0.41336424 \quad -0.48212701 \quad -0.07089826 
 [7] 0.07938086 0.22704866 -0.13926878 0.33629857 0.25835047 0.17881393
[13] \ -0.22141017 \ \ 0.01336452 \ \ 0.26283804 \ -0.24965088 \ -0.36552096 \ \ 0.33713195
[19] \quad 0.19514514 \quad -0.43506531 \quad -0.25760272 \quad -0.64783388 \quad -0.44626460 \quad -0.78405425
[25] 1.00057358
> res$fitted.values
 [1] 0.002033490 0.010312851 0.018703394 0.027863526 0.041320994 0.060871141
 [7] 0.088814107 0.127838223 0.180610428 0.248949062 0.332647930 0.428434554
[13] 0.529902047 0.628956590 0.718237396 0.793102235 0.852169542 0.896572801
[19] 0.928753893 0.951463983 0.967190831 0.977939948 0.985221193 0.990123427
[25] 0.999426746
> res$rank
[1] 2
> res$linear.predictors
 [1] -6.1959664 -4.5639981 -3.9601698 -3.5521777 -3.1441856 -2.7361935
  \lceil 7 \rceil \ -2.3282014 \ -1.9202093 \ -1.5122173 \ -1.1042252 \ -0.6962331 \ -0.2882410 
     0.1197511 0.5277432 0.9357353 1.3437274 1.7517194
                                                               2.1597115
[19] 2.5677036 2.9756957 3.3836878 3.7916799 4.1996720
[25] 7.4636087
> res$deviance
[1] 26.70345
> res$aic
[1] 114.7553
> res$null.deviance
[1] 3693.884
> res$weights
 [1] 0.7630428 2.0413099 1.7068902 3.2504707 3.5652333 5.0306085
     8.4972661 12.3760338 14.7990471 17.3885402 22.1993347 26.4468672
[13] 24.6614810 24.7372446 21.2491158 19.1986735 12.3457255
                                                               8.9948289
[19] 7.9404319 4.7104022 3.8714069 2.3946581 1.3686835 1.1148148
[25] 0.6010036
> res$prior.weights
 [1] 376 200
                 93 120
                           90
                                88 105 111 100
                                                     93 100 108 99 106 105
[16] 117
           98
                 97 120 102 122
                                    111
                                          94
                                               114 1049
> res$df.residual
[1] 23
```

```
> res$df.null

[1] 24

> res$y

[1] 0.00000000 0.00000000 0.00000000 0.01666667 0.02222222 0.05681818
[7] 0.09523810 0.15315315 0.16000000 0.31182796 0.39000000 0.47222222
[13] 0.47474747 0.63207547 0.77142857 0.75213675 0.80612245 0.92783505
[19] 0.94166667 0.93137255 0.95901639 0.96396396 0.97872340 0.98245614
[25] 1.00000000
```

vcov()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$(X^T W^{-1} X)^{-1}$$

• Examples:

coef()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description**: previsione
- Output:

```
fit valore previsto
se.fit standard error delle stime
```

• Formula:

```
fit x_0^T \, \hat{\beta} se.fit \sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}
```

• Examples:

predict()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

```
1
-30.18514

> res$se.fit

[1] 1.952408
```

fitted()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x < -c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> fitted(object = modello)
        1
                    2
                               3
                                          4
                                                     5
0.03254332\ 0.04415163\ 0.05084422\ 0.05663242\ 0.06388783\ 0.07323785\ 0.08571643
                   9
                             10 11
                                                   12
                                                               13
0.10314181 \ \ 0.12897631 \ \ 0.17045144 \ \ 0.24383760 \ \ 0.38066032 \ \ 0.57870619 \ \ 0.73297838
                                 18
       15
                  16
                             17
                                                    19
                                                               2.0
0.81708886\ 0.86366984\ 0.89210300\ 0.91098535\ 0.92435062\ 0.93427641\ 0.94192536
                   23
0.94799380 0.95292239 0.95700290 0.97326854
```

fitted.values()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\pi}_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
                  12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
                  14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
                  88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
                  108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> fitted.values(object = modello)
0.03254332\ 0.04415163\ 0.05084422\ 0.05663242\ 0.06388783\ 0.07323785\ 0.08571643
                                                         9 10 11
                                                                                                                                          12
                                                                                                                                                                                                  13
0.10314181 \ \ 0.12897631 \ \ 0.17045144 \ \ 0.24383760 \ \ 0.38066032 \ \ 0.57870619 \ \ 0.73297838
                                                         16
                                                                                          17
                                                                                                                             18
                        1.5
                                                                                                                                                                 19
                                                                                                                                                                                                   2.0
0.81708886 \ 0.86366984 \ 0.89210300 \ 0.91098535 \ 0.92435062 \ 0.93427641 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.941925368 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0.94192548 \ 0
                        22 23
                                                                                             24
                                                                                                                               2.5
0.94799380 0.95292239 0.95700290 0.97326854
```

cov2cor()

• Package: stats

• Input:

imes matrice di covarianza delle stime IWLS di dimensione k imes k

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$$

• Examples:

20.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

• **Description:** log-verosimiglianza binomiale

• Formula:

 $\hat{\ell}$

• Examples:

AIC()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

 $-2\hat{\ell} + 2k$

• Examples:

durbin.watson()

• Package: car

• Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> durbin.watson(model = modello)
 lag Autocorrelation D-W Statistic p-value
          0.5390491
                       0.4700264
Alternative hypothesis: rho != 0
> res <- durbin.watson(model = modello)</pre>
> res$dw
[1] 0.4700264
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2k$$

• Examples:

[1] 2.0000 268.9102

deviance()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

• Examples:

anova()

• Package: stats

• Input:

nullo modello nullo di regressione di Cauchy con n unità modello modello di regressione di Cauchy con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

```
Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value
```

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
          12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
          14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
   > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
          88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
   > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
          108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
          1049)
   > nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "cauchit"))
   > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
   > anova(nullo, modello, test = "Chisq")
   Analysis of Deviance Table
   Model 1: cbind(y, Total - y) \sim 1
   Model 2: cbind(y, Total - y) \sim x
     Resid. Df Resid. Dev Df Deviance P(>|Chi|)
             24
                    3693.9
                    180.9 1 3513.0
   2
             23
                                             0.0
   > res <- anova(nullo, modello, test = "Chisq")</pre>
   > res$"Resid. Df"
   [1] 24 23
   > res$"Resid. Dev"
    [1] 3693.8836 180.8584
   > res$Df
   [1] NA 1
   > res$Deviance
    [1]
             NA 3513.025
   > res$"P(>|Chi|)"
    [1] NA 0
drop1()
  • Package: stats
  • Input:
       object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità
       test = "Chisq"
  • Description: submodels
  • Output:
       Df differenza tra gradi di libertà
       Deviance differenza tra devianze residue
       AIC indice AIC
       LRT valore empirico della statistica \chi^2
```

Pr(Chi) p-value

• Formula:

Define $\underbrace{\frac{1,\,1,\,\ldots,\,1}{k-1\,\mathrm{volte}}}_{}$ Deviance $D,\,D_{-x_j}\ \ \forall\,j\,=\,1,\,2,\,\ldots,\,k-1$

dove D_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

AIC

$$-2\hat{\ell} + 2k, -2\hat{\ell}_{-x_i} + 2(k-1) \quad \forall j = 1, 2, \dots, k-1$$

dove $\hat{\ell}_{-x_j}$ rappresenta la log-verosimiglianza binomiale del modello eliminata la variabile esplicativa x_j .

LRT

$$c_j = D_{-x_j} - D \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
      12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
      14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> drop1(object = modello, test = "Chisq")
Single term deletions
Model:
cbind(y, Total - y) \sim x
                   AIC
       Df Deviance
                                    Pr(Chi)
                             LRT
             180.9 268.9
            3693.9 3779.9 3513.0 < 2.2e-16 ***
Х
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> res <- drop1(object = modello, test = "Chisq")</pre>
> res$Df
[1] NA 1
> res$Deviance
[1] 180.8584 3693.8836
> res$AIC
[1] 268.9102 3779.9354
> res$LRT
[1]
         NA 3513.025
> res$"Pr(Chi)"
[1] NA 0
```

add1()

• Package: stats

• Input:

object modello nullo di regressione di Cauchy scope modello di regressione di Cauchy con k-1 variabili esplicative ed n unità test = "Chisq"

- **Description:** submodels
- Output:

Df differenza tra gradi di libertà Deviance differenza tra devianze residue AIC indice AIC LRT valore empirico della statistica χ^2 Pr(Chi) p-value

• Formula:

Df

$$\underbrace{1, 1, \ldots, 1}_{k-1 \text{ volte}}$$

Deviance

$$D_{nullo}, D_{x_i} \quad \forall j = 1, 2, ..., k-1$$

dove D_{x_i} rappresenta la devianza residua del modello con la sola variabile esplicativa x_i .

AIC

$$-2\,\hat{\ell}_{nullo} + 2, \, -2\,\hat{\ell}_{x_j} + 4 \quad \forall j = 1, \, 2, \, \dots, \, k-1$$

dove $\hat{\ell}_{x_j}$ rappresenta la log-verosimiglianza binomiale del modello con la sola variabile esplicativa x_j .

LRT

$$c_j = D_{nullo} - D_{x_j} \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
      108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> nullo <- glm(formula = cbind(y, Total - y) ~ 1, family = binomial(link = "cauchit"))</pre>
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> add1(object = nullo, scope = modello, test = "Chisq")
Single term additions
Model:
cbind(y, Total - y) \sim 1
      Df Deviance
                     AIC
                            LRT Pr(Chi)
            3693.9 3779.9
<none>
            180.9 268.9 3513.0 < 2.2e-16 ***
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

20.4 Diagnostica

rstandard()

- Package: stats
- Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y < -c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> rstandard(model = modello)
                              3
                                         4
-5.1264853 -4.3358475 -3.1490590 -2.2484272 -1.8797967 -0.6232837
                                                                  0.3506059
        8
                   9
                            10
                                        11
                                                   12
                                                             13
1.6777851 \quad 0.9291382 \quad 3.4984066 \quad 3.5293420 \quad 2.3265176 \quad -2.4900358 \quad -2.5224910
       15
                 16
                            17
                                  18
                                             19
                                                       20
-1.2457978 -3.3570127 -2.5688041 0.6134906 0.7613634 -0.1193833 0.8636473
       22 23
                             24
                                        25
 0.8106387 1.3317047 1.5311383 8.0376682
```

rstandard.glm()

- Package: stats
- Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, \ldots, n
```

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> rstandard.glm(model = modello)
                            3
-5.1264853 -4.3358475 -3.1490590 -2.2484272 -1.8797967 -0.6232837
                                                             0.3506059
                          10
                                     11
                                               12
                                                         13
15
                 16
                          17
                                     18
                                               19
                                                         20
-1.2457978 -3.3570127 -2.5688041 0.6134906 0.7613634 -0.1193833 0.8636473
       22
                23
                           24
                                     2.5
 0.8106387 1.3317047 1.5311383 8.0376682
```

rstudent()

- Package: stats
- Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

```
rstudent_i \quad \forall i = 1, 2, ..., n
```

rstudent.glm()

• Package: stats

• Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- **Description:** residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, ..., n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
      1049)
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "cauchit"))
> rstudent.glm(model = modello)
                              3
                                         4
-5.0588500 \ -4.2941160 \ -3.1327370 \ -2.2391220 \ -1.8738045 \ -0.6226038 \ \ 0.3508547
                      10
                                                        13
        8 9
                                 11 12
 1.6840319 \quad 0.9311874 \quad 3.5275840 \quad 3.5611698 \quad 2.3353549 \quad -2.4956524 \quad -2.5390300
                                  18
       15
                 16
                             17
                                                    19
                                                               2.0
-1.2499439 \ -3.3841296 \ -2.5822550 \ \ 0.6127486 \ \ 0.7601912 \ -0.1194079 \ \ 0.8623051
                  23
                              24
                                        25
 0.8095676 1.3291375 1.5275625 7.7960241
```

residuals.default()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- **Description:** residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> residuals.default(object = modello)
-9.8152648 \ -7.2558854 \ -6.3140094 \ -4.0086223 \ -3.2932991 \ -0.9917917 \ \ 0.4226277
                               11 12
                                                    13
           9
                     10
          0.6272238 1.7058520 0.9553468 0.3321975 -0.3474066 -0.5728429
1.5498952
                                                19 20
                           17
                                     18
       15
                16
                                                                      2.1
-0.4855652 -2.0313711 -2.4430322 0.6948164 0.9814772 -0.2170523 1.6310583
       22 23
                            24
                                      25
1.8963437 3.7327336 4.4091809 11.9357223
```

residuals()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- Description: residui
- Formula:

$$\begin{array}{l} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

1049)

> residuals(object = modello, type = "response")

```
-4.9879493 \ -4.2499874 \ -3.1154320 \ -2.2134735 \ -1.8547635 \ -0.6138012 \ \ 0.3429411
                                                           11
                                                                             12 13
                     9
                                        10
    1.6292015 0.8969607 3.3340955 3.2290861 1.9359119 -2.0794099 -2.2707637
                                                                       18 19 20
                15
                                 16 17
  -1.1752053 \ -3.2150141 \ -2.5014455 \ \ 0.6008633 \ \ 0.7452777 \ -0.1175573 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.84985279
                22 23 24 25
    0.8002034 1.3186785 1.5146367 7.5396162
• Example 2:
  > x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
           12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
            14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
  > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
           88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
  > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
            108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
  > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
  > residuals(object = modello, type = "pearson")
                                    2
                                                       3
                                                                                            5
  -3.5563874 \ -3.0394419 \ -2.2319966 \ -1.8941117 \ -1.6163149 \ -0.5912262 \ \ 0.3485259
                                  9
                                                                      11
                                                   10
                                                                                        12
                                                                                                           1.3
    1.7324103 0.9256002 3.6257473 3.4039079 1.9597174 -2.0948691 -2.3482148
                                                   17
                                                                                        19
                                                                                                  20
               1.5
                                 16
                                                                      18
  -1.2102597 -3.5158214 -2.7434754 0.5827626 0.7173290 -0.1183527 0.8071359
                                23
                                         24
    0.7577756 1.1810403 1.3397363 5.3676317
• Example 3:
  > x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
            12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
            14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
  > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
            88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
  > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
            108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
            1049)
  > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
  > residuals(object = modello, type = "working")
                                                       3
                                                                         4
  -9.8152648 -7.2558854 -6.3140094 -4.0086223 -3.2932991 -0.9917917
                                  9
                                                           11
                                                                                         12
                                                                                                 13
                                         10
    1.5498952 \quad 0.6272238 \quad 1.7058520 \quad 0.9553468 \quad 0.3321975 \quad -0.3474066 \quad -0.5728429
               15
                     16 17 18
                                                                                        19
                                                                                                 20
  -0.4855652 -2.0313711 -2.4430322 0.6948164 0.9814772 -0.2170523 1.6310583
                                 23
                                                   24
    1.8963437 3.7327336 4.4091809 11.9357223
• Example 4:
  > x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
           12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
            14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
  > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
            88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
  > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
            108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
```

> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>

residuals.glm()

• Package: stats

• Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R & \forall i=1,2,\ldots,n \\ \\ \end{array}$$

• Example 1:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> residuals.glm(object = modello, type = "deviance")
                              3
                                        4
-4.9879493 -4.2499874 -3.1154320 -2.2134735 -1.8547635 -0.6138012 0.3429411
                  9
                      10
                                 11
                                                 12
                                                      13
1.6292015 \quad 0.8969607 \quad 3.3340955 \quad 3.2290861 \quad 1.9359119 \quad -2.0794099 \quad -2.2707637
           16 17 18 19 20
       15
-1.1752053 -3.2150141 -2.5014455 0.6008633 0.7452777 -0.1175573 0.8498527
```

22 23 24 25

0.8002034 1.3186785 1.5146367 7.5396162

• Example 2:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> residuals.glm(object = modello, type = "pearson")
                              3
                                        4
-3.5563874 -3.0394419 -2.2319966 -1.8941117 -1.6163149 -0.5912262 0.3485259
                 9 10 11 12 13
 1.7324103 \quad 0.9256002 \quad 3.6257473 \quad 3.4039079 \quad 1.9597174 \quad -2.0948691 \quad -2.3482148
                                           19 20
                           17
                                 18
       15
                 16
-1.2102597 -3.5158214 -2.7434754 0.5827626 0.7173290 -0.1183527 0.8071359
       22
                 23
                             24
                                       25
 0.7577756 1.1810403 1.3397363 5.3676317
```

• Example 3:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "cauchit"))
> residuals.glm(object = modello, type = "working")
                            3
                                      4
-9.8152648 -7.2558854 -6.3140094 -4.0086223 -3.2932991 -0.9917917
                               11
                 9
                          10
                                               12
                                                    13
 1.5498952
          0.6272238 1.7058520 0.9553468 0.3321975 -0.3474066 -0.5728429
                                          19
                                                    20
       1.5
                16 17
                                18
-0.4855652 -2.0313711 -2.4430322 0.6948164 0.9814772 -0.2170523 1.6310583
       22 23
                         24
 1.8963437 3.7327336 4.4091809 11.9357223
```

• Example 4:

19

2.0

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> residuals.glm(object = modello, type = "response")
                                                         5
                      2
                                  3
                                             4
-0.032543316 \ -0.044151625 \ -0.050844224 \ -0.039965753 \ -0.041665609 \ -0.016419665
                                 9
                     8
                                            10
                                                         11
 14
                                15
                                    16
                                                         17
-0.103958715 -0.100902908 -0.045660287 -0.111533087 -0.085980550 0.016849703
```

22

23

21

24

resid()

• Package: stats

• Input:

```
object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

$$\begin{array}{l} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \\ \text{type = "response"} \\ \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \end{array}$$

• Example 1:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
     12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "cauchit"))
> resid(object = modello, type = "deviance")
                                        4
                                                   5
-4.9879493 -4.2499874 -3.1154320 -2.2134735 -1.8547635 -0.6138012 0.3429411
                                                  12
                            10
                                       11
                                                             13
1.6292015 0.8969607 3.3340955 3.2290861 1.9359119 -2.0794099 -2.2707637
       15
                  16
                          17
                                  18
                                                  19
                                                             20
-1.1752053 -3.2150141 -2.5014455 0.6008633 0.7452777 -0.1175573 0.8498527
       2.2
               23
                             2.4
```

• Example 2:

0.8002034 1.3186785 1.5146367 7.5396162

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
       12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
       14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
 > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
       88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
       108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
 > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
 > resid(object = modello, type = "pearson")
                              3
                    2
                                        4
                                                   5
 -3.5563874 -3.0394419 -2.2319966 -1.8941117 -1.6163149 -0.5912262 0.3485259
              9 10 11 12 13
  17
                                 18
                                           19 20
                  16
 -1.2102597 -3.5158214 -2.7434754 0.5827626 0.7173290 -0.1183527 0.8071359
                  2.3
                      24
                                        25
  0.7577756 1.1810403 1.3397363 5.3676317
• Example 3:
 > x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
       12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
       14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
 > y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
       108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
       1049)
 > modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "cauchit"))
 > resid(object = modello, type = "working")
                                                   5
 -9.8152648 -7.2558854 -6.3140094 -4.0086223 -3.2932991 -0.9917917
                            10 11 12 13
                   9
           0.6272238 1.7058520 0.9553468 0.3321975 -0.3474066 -0.5728429
  1.5498952
                                 18
                                                      20
                            17
                  16
                                                  19
 -0.4855652 -2.0313711 -2.4430322 0.6948164 0.9814772 -0.2170523 1.6310583
                  23
                             2.4
  1.8963437 3.7327336 4.4091809 11.9357223
• Example 4:
 > x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
       12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
       14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
  y < -c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
      88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
 > Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
       108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
       1049)
 > modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
 > resid(object = modello, type = "response")
                                    3
 -0.032543316 -0.044151625 -0.050844224 -0.039965753 -0.041665609 -0.016419665
                       8
                          9
                                               10
                                                   11
  0.009521665 \quad 0.050011345 \quad 0.031023688 \quad 0.141376522 \quad 0.146162404 \quad 0.091561906
          13
                       14
                                   1.5
                                                            17
                                               16
 -0.103958715 \ -0.100902908 \ -0.045660287 \ -0.111533087 \ -0.085980550 \ \ 0.016849703
                                   21
                                              22
                       20
  0.017316049 \; -0.002903864 \quad 0.017091031 \quad 0.015970168 \quad 0.025801013 \quad 0.025453243
          2.5
  0.026731456
```

weighted.residuals()

• Package: stats

• Input:

obj modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
               12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
                14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
              88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
                108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
                 1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> weighted.residuals(obj = modello)
                                                          2
                                                                                          3
                                                                                                                         4
-4.9879493 \ -4.2499874 \ -3.1154320 \ -2.2134735 \ -1.8547635 \ -0.6138012 \ \ 0.3429411
                                      9
                                                                 10 11 12 13
  15 16 17 18 19 20
-1.1752053 \ -3.2150141 \ -2.5014455 \ \ 0.6008633 \ \ 0.7452777 \ -0.1175573 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.849852799 \ \ 0.84985279
                       22
                                  23
                                                                                       24
  0.8002034 1.3186785 1.5146367 7.5396162
```

weights()

• Package: stats

• Input:

object modello di regressione log-log complementare con k-1 variabili esplicative ed n unità

- Description: pesi iniziali
- Formula:

$$n_i \quad \forall i = 1, 2, \ldots, n$$

```
2
           3
                     5
                         6
                               7
                                    8
                                         9
                                             10
                                                            13
 1
                                                 11
                                                      12
                                                               14
                                                                      15
                                                                           16
                    90
                                                           99 106 105 117
376
          93
              120
                         88
                             105
                                             93 100 108
    200
                                  111
                                       100
17
     18
          19
              20
                    21
                         22
                              23
                                  24
                                        25
     97 120 102 122 111
                              94 114 1049
```

df.residual()

- Package: stats
- Input:

object modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

$$n-k$$

• Examples:

hatvalues()

[1] 23

• Package: stats

• Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

cooks.distance()

- Package:
- Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
    88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
     1049)
> modello <- glm(formula = cbind(y, Total - y) ~ x, family = binomial(link = "cauchit"))</pre>
> cooks.distance(model = modello)
                                     3
                                                 4
0.3762214804 \ \ 0.1962136349 \ \ 0.0552357880 \ \ 0.0589188486 \ \ 0.0364623856 \ \ 0.0056112386
                      8
                                   9
                                                10
                                                            11
0.0028692913 \ 0.0963310836 \ 0.0335706735 \ 0.7308700108 \ 1.3468893627 \ 1.2320350055
                     14 15
                                                16 17
         13
1.3653510505 \ 0.7961188111 \ 0.1018405155 \ 0.6083887972 \ 0.2166167590 \ 0.0075183418
                                       22
                       20
                                    21
                                                              23
0.0117156580 \ 0.0002261279 \ 0.0110091368 \ 0.0077349710 \ 0.0141216419 \ 0.0200921981
2.2344212321
```

cookd()

- Package: car
- Input:

model modello di regressione di Cauchy con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(9.21, 10.21, 10.58, 10.83, 11.08, 11.33, 11.58, 11.83,
    12.08, 12.33, 12.58, 12.83, 13.08, 13.33, 13.58, 13.83, 14.08,
     14.33, 14.58, 14.83, 15.08, 15.33, 15.58, 15.83, 17.58)
> y <- c(0, 0, 0, 2, 2, 5, 10, 17, 16, 29, 39, 51, 47, 67, 81,
     88, 79, 90, 113, 95, 117, 107, 92, 112, 1049)
> Total <- c(376, 200, 93, 120, 90, 88, 105, 111, 100, 93, 100,
     108, 99, 106, 105, 117, 98, 97, 120, 102, 122, 111, 94, 114,
> modello <- glm(formula = cbind(y, Total - y) \sim x, family = binomial(link = "cauchit"))
> cookd(model = modello)
                                 3
0.3762214804 \ 0.1962136349 \ 0.0552357880 \ 0.0589188486 \ 0.0364623856 \ 0.0056112386
                     8
                                9 10 11
0.0028692913 \ 0.0963310836 \ 0.0335706735 \ 0.7308700108 \ 1.3468893627 \ 1.2320350055
         13
                   14 15 16 17
1.3653510505 0.7961188111 0.1018405155 0.6083887972 0.2166167590 0.0075183418
         19 20 21 22 23
0.0117156580 \ 0.0002261279 \ 0.0110091368 \ 0.0077349710 \ 0.0141216419 \ 0.0200921981
2.2344212321
```

Capitolo 21

Regressione di Poisson

21.1 Simbologia

$$\log(\mu_i) = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1} \qquad Y_i \sim \text{Poisson}(\mu_i) \quad \forall i = 1, 2, \dots, n$$

- numero di conteggi: $y_i \quad \forall i = 1, 2, ..., n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w = (w_1, w_2, \dots, w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- distanza di Cook: $cd_i = \left(e_i^P\right)^2 \frac{h_i}{k \left(1-h_i\right)^2} \quad \forall i=1,2,\ldots,n$
- stime IWLS: $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \sqrt{\operatorname{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{eta}} = \hat{eta} / s_{\hat{eta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \, \hat{\beta}_j} = \frac{(X^T \, W^{-1} \, X)_{i,\,j}^{-1}}{s_{\hat{\beta}_i} \, s_{\hat{\beta}_j}} \quad \forall \, i,j \, = \, 1,\, 2,\, \ldots,\, k$
- residui di devianza: $e_i = \operatorname{sign}(y_i \hat{\mu}_i) \sqrt{2 \left(y_i \log \left(\frac{y_i}{\hat{\mu}_i} + C_i\right) (y_i \hat{\mu}_i)\right)}$ dove $C_i = 0.5 \left(1 \operatorname{sign}(y_i)\right) / \hat{\mu}_i \quad \forall i = 1, 2, \dots, n$
- residui standard: $rstandard_i = e_i / \sqrt{1 h_i} \quad \forall i = 1, 2, ..., n$
- residui studentizzati: $rstudent_i = sign\left(y_i \hat{\mu}_i\right) \sqrt{e_i^2 + h_i \left(e_i^P\right)^2 / (1 h_i)} \quad \forall i = 1, 2, \dots, n$
- residui di *Pearson*: $e_i^P = (y_i \hat{\mu}_i) / \sqrt{\hat{\mu}_i} \quad \forall i = 1, 2, \dots, n$
- residui di lavoro: $e_i^W = \left(y_i \hat{\mu}_i\right) / \hat{\mu}_i \quad \forall i = 1, 2, \dots, n$
- residui di riposta: $e_i^R = y_i \hat{\mu}_i \quad \forall i = 1, \, 2, \, \ldots, \, n$
- log-verosimiglianza di *Poisson*: $\hat{\ell} = \sum_{i=1}^n \left[y_i \log (\hat{\mu}_i) \hat{\mu}_i \log(y_i!) \right]$
- valori adattati: $\hat{\mu}_i = \exp\left(X_i\,\hat{\beta}\right) \ \ \forall\, i\,=\,1,\,2,\,\ldots,\,n$
- log-verosimiglianza di *Poisson* modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} [y_i \log(y_i) y_i \log(y_i!)]$

- devianza residua: $D=2\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^n e_i^2=2\sum_{i=1}^n y_i\log\left(\frac{y_i}{\hat{\mu}_i}+C_i\right)$ dove $C_i=0.5\left(1-\mathrm{sign}(y_i)\right)/\hat{\mu}_i \quad \forall i=1,2,\ldots,n$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza di *Poisson* modello nullo: $\hat{\ell}_{nullo} = \sum_{i=1}^{n} [y_i \log(\bar{y}) \bar{y} \log(y_i!)]$
- valori adattati modello nullo: $\hat{\mu} = \bar{y} \quad \forall i = 1, 2, ..., n$
- devianza residua modello nullo: $D_{nullo} = 2 \left(\hat{\ell}_{saturo} \hat{\ell}_{nullo} \right)$
- ullet gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = \log{(\hat{\mu})}$

21.2 Stima

glm()

- Package: stats
- Input:

```
formula modello di regressione di Poisson con k-1 variabili esplicative ed n unità family = poisson(link="log") famiglia e link del modello x = TRUE matrice del modello
```

- **Description:** analisi di regressione di *Poisson*
- Output:

```
coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y numero di conteggi
x matrice del modello
```

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall j=1,2,\dots,k$$
 residuals
$$e^W_i \quad \forall i=1,2,\dots,n$$
 fitted.values
$$\hat{\mu}_i \quad \forall i=1,2,\dots,n$$
 rank
$$k$$
 linear.predictors
$$X\,\hat{\beta}$$

```
deviance
                                                           D
aic
                                                       -2\hat{\ell}+2k
null.deviance
                                                        D_{nullo}
weights
                                                w_i \quad \forall i = 1, 2, \ldots, n
prior.weights
                                                      1, 1, \ldots, 1
                                                         n volte
df.residual
                                                         n-k
df.null
                                                         n-1
У
                                                y_i \quad \forall i = 1, 2, \ldots, n
                                                           X
```

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8, 
+ 9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"),
+ x = TRUE)
> modello$coefficients
(Intercept)
0.916392046 0.001997418
```

> modello\$residuals

1	2	3	4	5	6
-0.20165148	-0.56413249	0.29042202	0.70199431	0.34247005	-0.43487568
7	8	9	10	11	12
0.16386402	-0.20455645	0.04999536	0.33172955	-0.33831611	0.32602805
13	14	15	16	17	18
0.87408986	-0.35912141	0.10943462	-0.40119990	0.08161077	-0.33034568
19	20	21	22	23	24
0.50898714	0.21924503	-0.15404144	-0.68653798	-0.03098119	-0.37430000
25	26	27	28	29	30
-0.17573412	1.66878447	0.56630428	-0.10405228	0.04163966	-0.71290188
31	32				
-0.46243717	-0.65221412				

> modello\$fitted.values

```
7.515515 9.177101 13.173985 5.287914 10.428538 14.156177 4.296035 8.800122
                  11
                                12 13
                                                       15
      9
              10
                                                   14
6.666696 5.256322 9.067774 6.033055 14.940586 6.241432
                                                        9.013600 6.680026
      17 18
                                20 21
                                                    22
                                                        23
                  19
7.396376 13.439770 15.242012
                            7.381617 7.092546
                                              3.190179
                                                        9.287745
      2.5
              26
                    27
                                 28
                                           29
                                                    30
                                                             31
10.918807 \quad 5.245834 \quad 10.853574 \quad 11.161366 \quad 6.720174 \quad 10.449389 \quad 16.742229
                                                                5.750665
```

> modello\$rank

[1] 2

> modello\$linear.predictors

```
3
                                    5
                                                    7
     1
                            4
2.016970 2.216711 2.578244 1.665424 2.344546 2.650151 1.457692 2.174766
            10
                    11 12
                                   13
                                            14
                                                   15
1.897124 1.659432 2.204727 1.797253 2.704081 1.831210 2.198735 1.899122
           18
                                   21
    17
                    19
                       20
                                           22
                                                23
2.000990 2.598218 2.724056 1.998993 1.959044 1.160077 2.228696 1.855179
          26
                   27
                         28
                                   29
                                         30
                                                 31
2.390487 1.657434 2.384494 2.412458 1.905114 2.346544 2.817934 1.749315
```

- > modello\$deviance
- [1] 62.8054
- > modello\$aic
- [1] 190.1035
- > modello\$null.deviance
- [1] 103.7138
- > modello\$weights

```
2
                                        5
                       3
                               4
                                                6
7.515661 9.177255 13.174144 5.288041 10.428696 14.156336 4.296149 8.800275
             10
                     11
                              12
                                      13
                                                14
                                                        15
6.666836 5.256449 9.067928 6.033189 14.940742 6.241568 9.013754 6.680166
                               20
                                       21
                                                         23
     17
              18
                      19
                                                22
7.396521 13.439929 15.242168 7.381762 7.092689
                                           3.190277
                                                   9.287900 6.392978
             26
                  27
                          28
                                   29
                                           30
                                                    31
10.918966 5.245960 10.853733 11.161525 6.720315 10.449547 16.742380 5.750797
```

> modello\$prior.weights

- > modello\$df.residual
- [1] 30
- > modello\$df.null
- [1] 31
- > modello\$y

```
1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26 6 4 17 9 14 8 5 7 7 7 6 8 28 4 10 4 8 9 23 9 6 1 9 4 9 14 27 28 29 30 31 32 17 10 7 3 9 2
```

> modello\$x

```
(Intercept) x
            1 551
1
2
            1 651
3
            1 832
            1 375
4
5
            1 715
            1 868
7
            1 271
8
            1 630
9
            1 491
10
            1 372
11
            1 645
            1 441
12
13
            1 895
14
            1 458
15
            1 642
            1 492
16
            1 543
17
18
            1 842
19
            1 905
20
            1 542
21
            1 522
22
            1 122
23
            1 657
24
            1 470
25
            1 738
            1 371
26
27
            1 735
28
            1 749
29
            1 495
30
            1 716
           1 952
31
32
            1 417
attr(,"assign")
[1] 0 1
```

summary.glm()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- Description: analisi di regressione di Poisson
- Output:

```
deviance devianza residua
aic indice AIC
df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
```

correlation matrice di correlazione delle stime IWLS

• Formula:

deviance Daic $-2\hat{\ell}+2k$ df.residual n-knull.deviance D_{nullo} df.null n-1deviance.resid $e_i \quad \forall i = 1, 2, \ldots, n$ coefficients $\hat{\beta}_j$ $s_{\hat{\beta}_i}$ $z_{\hat{\beta}_i}$ p-value = $2\Phi(-|z_{\hat{\beta}_i}|)$ $\forall j = 1, 2, ..., k$ cov.unscaled $(X^T W^{-1} X)^{-1}$ cov.scaled $(X^T W^{-1} X)^{-1}$ correlation $r_{\hat{\beta}_i \, \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$

• Examples:

> res\$deviance.resid

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> res <- summary.glm(object = modello, correlation = TRUE)</pre>
> res$deviance
[1] 62.8054
> res$aic
[1] 190.1035
> res$df.residual
[1] 30
> res$null.deviance
[1] 103.7138
> res$df.null
[1] 31
```

```
-0.5731569 -1.9263607 1.0084275 1.4656879 1.0504241 -1.7835363 0.3309445
                  10
                            11 12
      8 9
                                              13
-0.6294980 0.1280339 0.7234253 -1.0862504 0.7623113 3.0093299 -0.9610107
         16 17 18 19 20 21
      15
0.3228171 \ -1.1213526 \ 0.2190303 \ -1.2890517 \ 1.8466732 \ 0.5756799 \ -0.4215129
      22 23 24 25 26 27 28
-1.4353411 -0.0949116 -1.0171558 -0.5990789 3.1586571 1.7215083 -0.3539304
         30 31
                           32
      29
0.1072073 - 2.7223502 - 2.0764597 - 1.8101537
> res$coefficients
                                         Pr(>|z|)
            Estimate Std. Error z value
(Intercept) 0.916392046 0.2215541099 4.136200 3.531049e-05
         0.001997418 0.0003184551 6.272213 3.559532e-10
> res$cov.unscaled
           (Intercept)
(Intercept) 4.908622e-02 -6.797742e-05
         -6.797742e-05 1.014137e-07
> res$cov.scaled
           (Intercept)
(Intercept) 4.908622e-02 -6.797742e-05
         -6.797742e-05 1.014137e-07
> res$correlation
         (Intercept)
(Intercept) 1.0000000 -0.9634665
          -0.9634665 1.0000000
```

glm.fit()

- Package: stats
- Input:
 - x matrice del modello
 - y numero di conteggi

family = poisson(link="log") famiglia e link del modello

- Description: analisi di regressione di Poisson
- Output:

```
coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
```

```
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua
df.null gradi di libertà devianza residua modello nullo
y numero di conteggi
```

• Formula:

coefficients	$\hat{\beta}_j \forall j = 1, 2, \dots, k$
residuals	$e_i^W \forall i = 1, 2, \dots, n$
fitted.values	$\hat{\mu}_i \forall i = 1, 2, \ldots, n$
rank	k
linear.predictors	$X\hat{eta}$
deviance	D
aic	$-2\hat\ell+2k$
null.deviance	D_{nullo}
weights	$w_i \forall i = 1, 2, \dots, n$
prior.weights	$\underbrace{1, 1, \ldots, 1}_{n \text{ volte}}$
df.residual	n-k
df.null	n-1
У	$y_i \forall i = 1, 2, \ldots, n$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
      9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> X <- model.matrix(object = modello)
> res <- glm.fit(x = X, y, family = poisson(link = "log"))
> res$coefficients
(Intercept)
0.916392046 0.001997418
> res$residuals
 [7] \quad 0.16386402 \quad -0.20455645 \quad 0.04999536 \quad 0.33172955 \quad -0.33831611 \quad 0.32602805
[13] \quad 0.87408986 \quad -0.35912141 \quad 0.10943462 \quad -0.40119990 \quad 0.08161077 \quad -0.33034568
[19] \quad 0.50898714 \quad 0.21924503 \quad -0.15404144 \quad -0.68653798 \quad -0.03098119 \quad -0.37430000
[25] -0.17573412 1.66878447
                              0.56630428 -0.10405228 0.04163966 -0.71290188
[31] -0.46243717 -0.65221412
```

```
> res$fitted.values
     7.515515 9.177101 13.173985 5.287914 10.428538 14.156177 4.296035
     8.800122 6.666696 5.256322 9.067774 6.033055 14.940586 6.241432
 [8]
     9.013600
              6.680026 7.396376 13.439770 15.242012 7.381617
[15]
     3.190179 9.287745 6.392840 10.918807 5.245834 10.853574 11.161366
[22]
[29] 6.720174 10.449389 16.742229 5.750665
> res$rank
[1] 2
> res$linear.predictors
 [1] 2.016970 2.216711 2.578244 1.665424 2.344546 2.650151 1.457692 2.174766
 [9] 1.897124 1.659432 2.204727 1.797253 2.704081 1.831210 2.198735 1.899122
[17] 2.000990 2.598218 2.724056 1.998993 1.959044 1.160077 2.228696 1.855179
[25] 2.390487 1.657434 2.384494 2.412458 1.905114 2.346544 2.817934 1.749315
> res$deviance
[1] 62.8054
> res$aic
[1] 190.1035
> res$null.deviance
[1] 103.7138
> res$weights
 [1]
    7.515661 9.177255 13.174144 5.288041 10.428696 14.156336 4.296149
 [8] 8.800275 6.666836 5.256449 9.067928 6.033189 14.940742 6.241568
[15] 9.013754 6.680166 7.396521 13.439929 15.242168 7.381762 7.092689
[22] 3.190277 9.287900 6.392978 10.918966 5.245960 10.853733 11.161525
    6.720315 10.449547 16.742380 5.750797
[29]
> res$prior.weights
> res$df.residual
[1] 30
> res$df.null
[1] 31
> res$y
 [1] 6 4 17 9 14 8 5 7 7 7 6 8 28 4 10 4 8 9 23 9 6 1 9 4 9
[26] 14 17 10 7 3 9
```

vcov()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$(X^T W^{-1} X)^{-1}$$

• Examples:

coef()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall i = 1, 2, \dots, k$$

• Examples:

coefficients()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

predict()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime

• Formula:

fit
$$x_0^T \, \hat{\beta}$$
 se.fit
$$\sqrt{x_0^T \, (X^T \, W^{-1} \, X)^{-1} \, x_0}$$

• Examples:

fitted()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> fitted(object = modello)
```

```
7.515515
          9.177101 13.173985
                               5.287914 10.428538 14.156177
                                                             4.296035
                10
                          11
                                     12
                                               13
                                                                   15
                                                         14
6.666696
          5.256322
                    9.067774
                               6.033055 14.940586
                                                   6.241432
                                                             9.013600
                                                                       6.680026
      17
                18
                          19
                                     20
                                               21
                                                         2.2
                                                                   23
7.396376 13.439770 15.242012
                               7.381617
                                        7.092546
                                                   3.190179
                                                             9.287745
                                                                       6.392840
      25
                26
                          27
                                     28
                                               29
                                                         30
                                                                   31
                                                                             32
10.918807 5.245834 10.853574 11.161366 6.720174 10.449389 16.742229
```

fitted.values()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
      9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> fitted.values(object = modello)
                            3
                                       4
                                                 5
                                                           6
 7.515515
           9.177101 13.173985 5.287914 10.428538 14.156177
                                                              4.296035
                 10
                           11
                                     12
                                               13
                                                          14
                                                                    1.5
 6.666696 5.256322 9.067774
                               6.033055 14.940586
                                                   6.241432
                                                              9.013600
                                                                         6.680026
                                                                    23
       17
                 18
                           19
                                     20
                                                21
                                                          22
 7.396376 13.439770 15.242012
                               7.381617
                                         7.092546
                                                    3.190179
                                                              9.287745
                                                                         6.392840
                 26
                           27
                                     28
                                                29
                                                          30
                                                                    31
10.918807 5.245834 10.853574 11.161366 6.720174 10.449389 16.742229
```

cov2cor()

- Package: stats
- Input:

 \forall matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \, \hat{\beta}_j} \quad \forall i, j = 1, 2, \dots, k$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> V <- vcov(object = modello)
> cov2cor(V)
```

```
(Intercept) x
(Intercept) 1.0000000 -0.9634665
x -0.9634665 1.0000000
```

21.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: log-verosimiglianza di Poisson
- Formula:

 $\hat{\ell}$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> logLik(object = modello)
'log Lik.' -93.05175 (df=2)
```

AIC()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

 $-2\hat{\ell} + 2k$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> AIC(object = modello)
[1] 190.1035
```

durbin.watson()

• Package: car

• Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D-W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

• Examples:

extractAIC()

• Package: stats

• Input:

fit modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2k$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> extractAIC(fit = modello)
[1] 2.0000 190.1035
```

deviance()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> deviance(object = modello)
[1] 62.8054
```

anova()

• Package: stats

• Input:

nullo modello nullo di regressione di Poisson con n unità modello di regressione di Poisson con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

• Formula:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
          441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
          470, 738, 371, 735, 749, 495, 716, 952, 417)
   > y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
          9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
   > nullo <- glm(formula = y ~ 1, family = poisson(link = "log"))
   > modello <- glm(formula = y \sim x, family = poisson(link = "log"))
   > anova(nullo, modello, test = "Chisq")
   Analysis of Deviance Table
   Model 1: y ~ 1
   Model 2: y \sim x
     Resid. Df Resid. Dev Df Deviance P(>|Chi|)
             31
                   103.714
                    62.805 1 40.908 1.595e-10
   2
             30
   > res <- anova(nullo, modello, test = "Chisq")</pre>
   > res$"Resid. Df"
   [1] 31 30
   > res$"Resid. Dev"
    [1] 103.7138 62.8054
   > res$Df
    [1] NA 1
   > res$Deviance
   [1]
             NA 40.90836
   > res$"P(>|Chi|)"
                 NA 1.595374e-10
    [1]
drop1()
  • Package: stats
  • Input:
       object modello di regressione di Poisson con k-1 variabili esplicative ed n unità
       test = "Chisq"
  • Description: submodels
  • Output:
       Df differenza tra gradi di libertà
       Deviance differenza tra devianze residue
       AIC indice AIC
       LRT valore empirico della statistica \chi^2
       Pr(Chi) p-value
```

839

Df

$$\underbrace{1, 1, \ldots, 1}_{k-1 \text{ volte}}$$

Deviance

$$D, D_{-x_i} \quad \forall j = 1, 2, \dots, k-1$$

dove D_{-x_j} rappresenta la devianza residua del modello eliminata la variabile esplicativa x_j .

AIC

$$-2\hat{\ell} + 2k, -2\hat{\ell}_{-x_i} + 2(k-1) \quad \forall j = 1, 2, \dots, k-1$$

dove $\hat{\ell}_{-x_j}$ rappresenta la log-verosimiglianza di *Poisson* del modello eliminata la variabile esplicativa x_j .

LRT

$$c_j = D_{-x_j} - D \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_j) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
      9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> drop1(object = modello, test = "Chisq")
Single term deletions
Model:
       Df Deviance
                     AIC
                            LRT
                                     Pr(Chi)
           62.805 190.104
       1 103.714 229.012 40.908 1.595e-10 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> res <- drop1(object = modello, test = "Chisq")</pre>
> res$Df
[1] NA 1
> res$Deviance
[1] 62.8054 103.7138
> res$AIC
[1] 190.1035 229.0119
> res$LRT
[1]
         NA 40.90836
> res$"Pr(Chi)"
[1]
             NA 1.595374e-10
```

add1()

• Package: stats

• Input:

object modello nullo di regressione di Poisson scope modello di regressione di Poisson con k-1 variabili esplicative ed n unità test = "Chisq"

- Description: submodels
- Output:

Df differenza tra gradi di libertà Deviance differenza tra devianze residue AIC indice AIC LRT valore empirico della statistica χ^2 Pr(Chi) p-value

• Formula:

Df

$$\underbrace{1,\,1,\,\ldots,\,1}_{k-1\,\text{volte}}$$

Deviance

$$D_{nullo}, D_{x_j} \quad \forall j = 1, 2, \dots, k-1$$

dove D_{x_j} rappresenta la devianza residua del modello con la sola variabile esplicativa x_j .

AIC

$$-2\,\hat{\ell}_{nullo} + 2,\, -2\,\hat{\ell}_{x_j} + 4 \quad \forall \, j \, = \, 1,\, 2,\, \ldots,\, k-1$$

dove $\hat{\ell}_{x_j}$ rappresenta la log-verosimiglianza di *Poisson* del modello con la sola variabile esplicativa x_j .

LRT

$$c_j = D_{nullo} - D_{x_j} \quad \forall j = 1, 2, \dots, k - 1$$

Pr(Chi)

$$P(\chi_1^2 \ge c_i) \quad \forall j = 1, 2, \dots, k-1$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
      9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> nullo <- glm(formula = y ~ 1, family = poisson(link = "log"))</pre>
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> add1(object = nullo, scope = modello, test = "Chisq")
Single term additions
Model:
y ~ 1
       Df Deviance
                     AIC
                               LRT
                                   Pr(Chi)
          103.714 229.012
          62.805 190.104 40.908 1.595e-10 ***
Х
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
> res <- add1(object = nullo, scope = modello, test = "Chisq")
> res$Df
```

```
[1] NA 1
> res$Deviance
[1] 103.7138 62.8054
> res$AIC
[1] 229.0119 190.1035
> res$LRT
[1] NA 40.90836
> res$"Pr(Chi)"
[1] NA 1.595374e-10
```

21.4 Diagnostica

rstandard()

• Package: stats

• Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
      441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8, 
+ 9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> rstandard(model = modello)
                                    3
                         1.05211402 1.51608947 1.07143385 -1.88626732
-0.58415822 -1.95861072
                      8
                                   9
                                               10
                                                            11
 0.34589794 \ -0.63996238 \ 0.13103010 \ 0.74852597 \ -1.10435414 \ 0.78352354
         13
                     14
                                                            17
                                  15
                                               16
 3.22469291 - 0.98623876 0.32818923 - 1.14750260 0.22333743 - 1.34944537
         19
                     20
                                  21
                                              22
                                                            23
 1.98995067 0.58703566 -0.43038260 -1.52017691 -0.09651101 -1.04276847
         25
                26
                                  27
                                       28
                                                           29
-0.61255699 3.26857905 1.75959764 -0.36242210 0.10968144 -2.77705113
         31
-2.31245034 -1.86471908
```

rstandard.glm()

• Package: stats

• Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

```
rstandard_i \quad \forall i = 1, 2, ..., n
```

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> rstandard.glm(model = modello)
-0.58415822 -1.95861072 1.05211402 1.51608947 1.07143385 -1.88626732
                8
                                      10
                                                         11
 0.34589794 \ -0.63996238 \ 0.13103010 \ 0.74852597 \ -1.10435414 \ 0.78352354
        13
             14
                                 15
                                             16
                                                          17
 3.22469291 - 0.98623876 \ 0.32818923 - 1.14750260 \ 0.223333743 - 1.34944537
                                                          23
        19
                     2.0
                                 2.1
                                             2.2
 1.98995067 \quad 0.58703566 \quad -0.43038260 \quad -1.52017691 \quad -0.09651101 \quad -1.04276847
         25
                    26
                                 27
                                             28
                                                         29
-0.61255699 3.26857905 1.75959764 -0.36242210 0.10968144 -2.77705113
         31
                     32
-2.31245034 -1.86471908
```

rstudent()

• Package: stats

• Input:

model modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> rstudent(model = modello)
                                                           5
                                  3
                                               4
-0.58339795 \ -1.95178717 \ 1.05607073 \ 1.52661113 \ 1.07368887 \ -1.87037216
              8
                                  9
                                             10
                                                         11
 0.34667588 - 0.63922752 \ 0.13107905 \ 0.75111918 - 1.10219023
        13
                                 15
                                             16
                                                          17
                    14
 3.27847151 - 0.98303536 0.32838016 - 1.14375042 0.22345192 - 1.34249887
```

rstudent.glm()

• Package: stats

• Input:

model modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: residui studentizzati
- Formula:

$$rstudent_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
    9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> rstudent.glm(model = modello)
                               3
                                          4
-0.58339795 \ -1.95178717 \ 1.05607073 \ 1.52661113 \ 1.07368887 \ -1.87037216
            8
        7
                       9
                                 10
                                             11
 0.34667588 - 0.63922752 \quad 0.13107905 \quad 0.75111918 - 1.10219023 \quad 0.78568685
       1.3
           14
                      15
                                 16
                                                    17
 3.27847151 - 0.98303536 0.32838016 - 1.14375042 0.22345192 - 1.34249887
           20
                      21
                                 22
                                            23
        19
 2.01164323 \quad 0.58782968 \quad -0.42991912 \quad -1.49773238 \quad -0.09649454 \quad -1.03936493
                                 28
                      27
                                            29
            26
        2.5
-0.61175065 3.31837107 1.76616018 -0.36212559 0.10971516 -2.76165762
        31
-2.27414465 -1.85104246
```

residuals.default()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \dots, n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> residuals.default(object = modello)
```

```
-0.20165148 -0.56413249 0.29042202 0.70199431 0.34247005 -0.43487568
                             10
                                       11
          8
                   9
0.16386402 - 0.20455645 \ 0.04999536 \ 0.33172955 - 0.33831611 \ 0.32602805
      13 14 15 16 17
0.87408986 \ -0.35912141 \quad 0.10943462 \ -0.40119990 \quad 0.08161077 \ -0.33034568
          20 21 22 23 24
0.50898714 \quad 0.21924503 \quad -0.15404144 \quad -0.68653798 \quad -0.03098119 \quad -0.37430000
           26
                   27 28 29 30
       2.5
-0.17573412 1.66878447 0.56630428 -0.10405228 0.04163966 -0.71290188
       31
-0.46243717 -0.65221412
```

residuals()

• Package: stats

• Input:

object modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{l} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

```
441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> residuals(object = modello, type = "deviance")
                                              5
-0.5731569 \ -1.9263607 \ 1.0084275 \ 1.4656879 \ 1.0504241 \ -1.7835363 \ 0.3309445
                             11
                                        12
                                                      13
                9
                         10
-0.6294980 0.1280339 0.7234253 -1.0862504 0.7623113 3.0093299 -0.9610107
      15
           16
                    17
                             18 19
                                                      20
0.3228171 \ -1.1213526 \ 0.2190303 \ -1.2890517 \ 1.8466732 \ 0.5756799 \ -0.4215129
       22 23 24 25 26 27 28
-1.4353411 -0.0949116 -1.0171558 -0.5990789 3.1586571 1.7215083 -0.3539304
       29 30 31 32
0.1072073 - 2.7223502 - 2.0764597 - 1.8101537
```

> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,

• Example 2:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
    441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
    9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> residuals(object = modello, type = "pearson")
-0.55281621 -1.70896773 1.05411532 1.61426859 1.10594698 -1.63620653
        7 8 9 10 11
0.33963895 \; -0.60681668 \quad 0.12908774 \quad 0.76054544 \; -1.01876268 \quad 0.80079916
       13 14 15 16 17 18
3.37862422 - 0.89718790 \quad 0.32855181 - 1.03693106 \quad 0.22195094 - 1.21105688
                                            23
       19
            20
                             21
                                        22
1.98713767 \quad 0.59566971 \quad -0.41024061 \quad -1.22623047 \quad -0.09441767 \quad -0.94638261
              26
                      27
                                 28
                                            29
-0.58068913 3.82214815 1.86567606 -0.34762443 0.10794374 -2.30449201
       31
-1.89216663 -1.56404492
```

• Example 3:

```
441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
    9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> residuals(object = modello, type = "working")
-0.20165148 \ -0.56413249 \ \ 0.29042202 \ \ 0.70199431 \ \ 0.34247005 \ -0.43487568
                      9
                                 10 11 12
           8
 0.16386402 \; -0.20455645 \quad 0.04999536 \quad 0.33172955 \; -0.33831611 \quad 0.32602805
                  14 15 16 17
        13
 0.87408986 \ -0.35912141 \quad 0.10943462 \ -0.40119990 \quad 0.08161077 \ -0.33034568
                             21 22 23
                   20
 0.50898714 \quad 0.21924503 \quad -0.15404144 \quad -0.68653798 \quad -0.03098119 \quad -0.37430000
                                  28 29
                       27
             26
-0.17573412 1.66878447 0.56630428 -0.10405228 0.04163966 -0.71290188
        31
                  32
-0.46243717 -0.65221412
```

> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,

• Example 4:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
    441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+ 9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> residuals(object = modello, type = "response")
-1.5155146 \ -5.1771007 \quad 3.8260153 \quad 3.7120857 \quad 3.5714619 \ -6.1561773 \quad 0.7039655
      8 9
                   10
                             11 12 13
-1.8001216 0.3333039 1.7436775 -3.0677741 1.9669451 13.0594144 -2.2414318
       15
                    17 18 19 20
               16
 0.9863999 \ -2.6800256 \quad 0.6036240 \ -4.4397699 \quad 7.7579880 \quad 1.6183829 \ -1.0925460
       22 23 24 25 26 27
-2.1901791 -0.2877454 -2.3928401 -1.9188070 8.7541661 6.1464257 -1.1613656
       29 30 31
 0.2798258 - 7.4493890 - 7.7422291 - 3.7506647
```

residuals.glm()

• Package: stats

• Input:

object modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,\,2,\,\ldots,\,n \\ \\ \end{array}$$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
    441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> residuals.glm(object = modello, type = "deviance")
                                      4
-0.5731569 \ -1.9263607 \ 1.0084275 \ 1.4656879 \ 1.0504241 \ -1.7835363 \ 0.3309445
       8 9 10 11 12 13
-0.6294980 \quad 0.1280339 \quad 0.7234253 \quad -1.0862504 \quad 0.7623113 \quad 3.0093299 \quad -0.9610107
      15 16
                           17 18 19 20
 0.3228171 - 1.1213526 \quad 0.2190303 - 1.2890517 \quad 1.8466732 \quad 0.5756799 - 0.4215129
       22 23 24 25
                                          26
                                                          27
-1.4353411 -0.0949116 -1.0171558 -0.5990789 3.1586571 1.7215083 -0.3539304
                 30
                            31
 0.1072073 - 2.7223502 - 2.0764597 - 1.8101537
```

• Example 2:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> residuals.glm(object = modello, type = "pearson")
```

• Example 3:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
    441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> residuals.glm(object = modello, type = "working")
-0.20165148 -0.56413249 0.29042202 0.70199431 0.34247005 -0.43487568
       7
           8 9
                                10 11 12
0.16386402 \; -0.20455645 \quad 0.04999536 \quad 0.33172955 \; -0.33831611 \quad 0.32602805
      13 14 15 16 17 18
0.87408986 - 0.35912141 \quad 0.10943462 - 0.40119990 \quad 0.08161077 - 0.33034568
       19 20 21 22 23 24
0.50898714 \quad 0.21924503 \ -0.15404144 \ -0.68653798 \ -0.03098119 \ -0.37430000
                     27
                                28 29
           26
-0.17573412 1.66878447 0.56630428 -0.10405228 0.04163966 -0.71290188
       31
-0.46243717 -0.65221412
```

• Example 4:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
    441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+ 9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> residuals.glm(object = modello, type = "response")
                           3
                                    4
-1.5155146 \ -5.1771007 \ \ 3.8260153 \ \ 3.7120857 \ \ 3.5714619 \ -6.1561773 \ \ 0.7039655
       8 9 10 11 12 13
-1.8001216 0.3333039 1.7436775 -3.0677741 1.9669451 13.0594144 -2.2414318
      15 16 17 18 19 20 21
 0.9863999 \ -2.6800256 \quad 0.6036240 \ -4.4397699 \quad 7.7579880 \quad 1.6183829 \ -1.0925460
                                        26
                   24
                             25
                                                 27 28
       22
          23
-2.1901791 -0.2877454 -2.3928401 -1.9188070 8.7541661 6.1464257 -1.1613656
       29
                    31
          30
 0.2798258 - 7.4493890 - 7.7422291 - 3.7506647
```

resid()

• Package: stats

• Input:

object modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

type = "deviance"
$$e_i \quad \forall i = 1, 2, \dots, n$$

$$type = "pearson"$$

$$e_i^P \quad \forall i = 1, 2, \dots, n$$

$$type = "working"$$

$$e_i^W \quad \forall i = 1, 2, \dots, n$$

$$type = "response"$$

$$e_i^R \quad \forall i = 1, 2, \dots, n$$

• Example 1:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
 441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+ 9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> resid(object = modello, type = "deviance")
                         3
-0.5731569 -1.9263607 1.0084275 1.4656879 1.0504241 -1.7835363 0.3309445
      8 9 10 11 12 13 14
-0.6294980 0.1280339 0.7234253 -1.0862504 0.7623113 3.0093299 -0.9610107
         16
                  17 18
                                     19 20 21
      15
0.3228171 \ -1.1213526 \ \ 0.2190303 \ -1.2890517 \ \ 1.8466732 \ \ 0.5756799 \ -0.4215129
      22 23 24 25
                                     26 27
-1.4353411 -0.0949116 -1.0171558 -0.5990789 3.1586571 1.7215083 -0.3539304
      29 30 31 32
0.1072073 - 2.7223502 - 2.0764597 - 1.8101537
```

• Example 2:

```
470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
   9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> resid(object = modello, type = "pearson")
                            3
-0.55281621 -1.70896773 1.05411532 1.61426859 1.10594698 -1.63620653
       7 8 9 10 11 12
0.33963895 \; -0.60681668 \quad 0.12908774 \quad 0.76054544 \; -1.01876268 \quad 0.80079916
          14
                     15
                               16
                                          17
       13
3.37862422 -0.89718790 0.32855181 -1.03693106 0.22195094 -1.21105688
           20 21 22 23 24
       19
1.98713767 \quad 0.59566971 \quad -0.41024061 \quad -1.22623047 \quad -0.09441767 \quad -0.94638261
           26 27 28 29
-0.58068913 3.82214815 1.86567606 -0.34762443 0.10794374 -2.30449201
       31
-1.89216663 -1.56404492
```

> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,

441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,

• Example 3:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> resid(object = modello, type = "working")
-0.20165148 -0.56413249 0.29042202 0.70199431 0.34247005 -0.43487568
         7
            8
                                9
                                          10 11
0.16386402 \; -0.20455645 \quad 0.04999536 \quad 0.33172955 \; -0.33831611 \quad 0.32602805
        13 14 15 16
                                                       17
0.87408986 - 0.35912141 \quad 0.10943462 - 0.40119990 \quad 0.08161077 - 0.33034568
        19
                   20
                                           22
                                                       2.3
                               21
0.50898714 \quad 0.21924503 \quad -0.15404144 \quad -0.68653798 \quad -0.03098119 \quad -0.37430000
               26
                        27
                                    28
                                               29
-0.17573412 \quad 1.66878447 \quad 0.56630428 \quad -0.10405228 \quad 0.04163966 \quad -0.71290188
        31
-0.46243717 -0.65221412
```

• Example 4:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
    470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
    9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> resid(object = modello, type = "response")
                           3
-1.5155146 \ -5.1771007 \quad 3.8260153 \quad 3.7120857 \quad 3.5714619 \ -6.1561773 \quad 0.7039655
          9
                    10
                             11
                                       12 13
-1.8001216 0.3333039 1.7436775 -3.0677741 1.9669451 13.0594144 -2.2414318
                   17 18 19 20
      15 16
0.9863999 -2.6800256 0.6036240 -4.4397699 7.7579880 1.6183829 -1.0925460
      22 23 24 25 26 27
-2.1901791 -0.2877454 -2.3928401 -1.9188070 8.7541661 6.1464257 -1.1613656
          30
                   31
0.2798258 - 7.4493890 - 7.7422291 - 3.7506647
```

weighted.residuals()

- Package: stats
- Input:

obj modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- **Description:** residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> weighted.residuals(obj = modello)
```

```
-0.5731569 -1.9263607 1.0084275 1.4656879
                                             1.0504241 -1.7835363
        8
                   9
                             10
                                         11
                                                    12
                                                               13
-0.6294980 0.1280339 0.7234253 -1.0862504 0.7623113
                                                        3.0093299 -0.9610107
       15
                  16
                              17
                                         18
                                                    19
                                                               20
                                                                          2.1
 0.3228171 - 1.1213526 \quad 0.2190303 - 1.2890517 \quad 1.8466732
                                                        0.5756799 - 0.4215129
                 23
                              24
                                         25
                                                    26
                                                               27
-1.4353411 -0.0949116 -1.0171558 -0.5990789 3.1586571 1.7215083 -0.3539304
       29
           30
                              31
                                         32
 0.1072073 - 2.7223502 - 2.0764597 - 1.8101537
```

weights()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- **Description:** pesi iniziali
- Formula:

$$\underbrace{1,\,1,\,\ldots,\,1}_{n\,\mathrm{volte}}$$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> weights(object = modello)
1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26
              1
        1 1
                  1 1 1
                          1 1 1
                                   1 1 1 1
                                              1 1 1 1
                                                          1 1
27 28 29 30 31 32
1 1 1 1 1
```

df.residual()

• Package: stats

• Input:

object modello di regressione di Poisson con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

$$n-k$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
+     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
+     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
+     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))
> df.residual(object = modello)
[1] 30
```

hatvalues()

• Package: stats

• Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

• **Description:** valori di leva

• Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
     9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> hatvalues(model = modello)
                                                   5
        1
                   2
                              3
                                        4
0.03731074\ 0.03266037\ 0.08132102\ 0.06538376\ 0.03883352\ 0.10595899\ 0.08459283
                  9 10 11 12
                                                             1.3
0.03243571 \ \ 0.04520986 \ \ 0.06594243 \ \ 0.03251736 \ \ 0.05341286 \ \ 0.12911084 \ \ 0.05050580
       15
                 16 17 18 19 20
0.03247008 \ 0.04505800 \ 0.03819908 \ 0.08750591 \ 0.13881691 \ 0.03831420 \ 0.04079290
       22
                  23
                             24
                                        25
                                                   26
                                                              27
0.10849868 \ \ 0.03286992 \ \ 0.04852097 \ \ 0.04352190 \ \ 0.06612878 \ \ 0.04282468 \ \ 0.04631162 
       29
                  30
                             31
0.04460584 0.03900696 0.19368977 0.05766771
```

cooks.distance()

• Package: stats

• Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
     470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
    9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y ~ x, family = poisson(link = "log"))</pre>
> cooks.distance(model = modello)
0.0061516720 0.0509683838 0.0535329887 0.0975269911 0.0257068065 0.1774472070
                        8
                                      9
                                                                11
                                                  1.0
0.0058225056 \ 0.0063789436 \ 0.0004131972 \ 0.0218593896 \ 0.0180278945 \ 0.0191135734
          13
                       14
                                    15
                                                  16
                                                                17
0.9715982423 0.0225472435 0.0018721138 0.0265636449 0.0010171067 0.0770683993
                       20
                                     2.1
                                                   2.2
                                                                23
0.3695534723 \ 0.0073497811 \ 0.0037308438 \ 0.1026348110 \ 0.0001566410 \ 0.0240012884
```

```
25 26 27 28 29 30
0.0080207542 0.5538620110 0.0813492551 0.0030765755 0.0002847026 0.1121558914
31 32
0.5333239875 0.0794315456
```

cookd()

- Package: car
- Input:

model modello di regressione di *Poisson* con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

```
> x <- c(551, 651, 832, 375, 715, 868, 271, 630, 491, 372, 645,
     441, 895, 458, 642, 492, 543, 842, 905, 542, 522, 122, 657,
      470, 738, 371, 735, 749, 495, 716, 952, 417)
> y <- c(6, 4, 17, 9, 14, 8, 5, 7, 7, 7, 6, 8, 28, 4, 10, 4, 8,
      9, 23, 9, 6, 1, 9, 4, 9, 14, 17, 10, 7, 3, 9, 2)
> modello <- glm(formula = y \sim x, family = poisson(link = "log"))
> cookd(model = modello)
                         2
                                       3
                                                    4
                                                                  5
           1
0.0061516720 \ 0.0509683838 \ 0.0535329887 \ 0.0975269911 \ 0.0257068065 \ 0.1774472070
           7
                        8
                                     9
                                                  10
                                                                11
0.0058225056 \ 0.0063789436 \ 0.0004131972 \ 0.0218593896 \ 0.0180278945 \ 0.0191135734
          13
                        14
                                     15
                                                   16
                                                                 17
0.9715982423 \ 0.0225472435 \ 0.0018721138 \ 0.0265636449 \ 0.0010171067 \ 0.0770683993
                        20
                                     21
                                                   22
0.3695534723 \ 0.0073497811 \ 0.0037308438 \ 0.1026348110 \ 0.0001566410 \ 0.0240012884
                        26
                                     27
                                                   28
                                                                 29
0.0080207542\ 0.5538620110\ 0.0813492551\ 0.0030765755\ 0.0002847026\ 0.1121558914
          31
0.5333239875 0.0794315456
```

Capitolo 22

Regressione Gamma

22.1 Simbologia

$$1/\mu_i = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1}$$
 $Y_i \sim \text{Gamma}(\omega, \omega / \mu_i) \quad \forall i = 1, 2, \dots, n$

- valori osservati: $y_i \quad \forall i = 1, 2, ..., n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w=(w_1,\,w_2,\,\ldots,\,w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \operatorname{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X(X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- distanza di Cook: $cd_i=\left(e_i^P\right)^2\frac{h_i}{\hat{\phi}^2\,k\,(1-h_i)^2}\,\,\,\,\forall\,i=1,\,2,\,\ldots,\,n$
- stime IWLS: $\hat{\beta} = \left(\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k\right)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \hat{\phi} \sqrt{\mathrm{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{\beta}} = \hat{\beta} / s_{\hat{\beta}}$
- correlazione delle stime IWLS: $r_{\hat{\beta}_i \; \hat{\beta}_j} = \frac{\hat{\phi}^2 \, (X^T \, W^{-1} \, X)_{i, \, j}^{-1}}{s_{\hat{\beta}_i \; s_{\hat{\beta}_j}}} \quad \forall \, i, j \, = \, 1, \, 2, \, \ldots, \, k$
- stima del parametro di dispersione: $\hat{\phi}^2=rac{1}{n-k}\sum_{i=1}^n\left(e_i^P
 ight)^2=rac{1}{n-k}\sum_{i=1}^n\left(y_i-\hat{\mu}_i
 ight)^2/\hat{\mu}_i^2$
- residui di devianza: $e_i = \operatorname{sign}\left(y_i \hat{\mu}_i\right) \sqrt{2\left(\left(y_i \hat{\mu}_i\right) / \hat{\mu}_i \log\left(y_i / \hat{\mu}_i\right)\right)} \quad \forall i = 1, 2, \dots, n$
- residui standard: $rstandard_i = \frac{e_i}{\hat{\phi}\sqrt{1-h_i}} \quad \forall i=1,2,\ldots,n$
- residui di *Pearson*: $e_i^P = (y_i \hat{\mu}_i) / \hat{\mu}_i \quad \forall i = 1, 2, ..., n$
- residui di lavoro: $e_i^W = -(y_i \hat{\mu}_i) / \hat{\mu}_i^2 \quad \forall i = 1, 2, \dots, n$
- residui di riposta: $e_i^R = y_i \hat{\mu}_i \quad \forall i=1,\,2,\,\ldots,\,n$
- log-verosimiglianza gamma: $\hat{\ell} = \sum_{i=1}^{n} \left[\hat{\omega} \left(-y_i / \hat{\mu}_i \log\left(\hat{\mu}_i\right) \right) + (\hat{\omega} 1) \log\left(y_i\right) + \hat{\omega} \log\left(\hat{\omega}\right) \log\left(\Gamma\left(\hat{\omega}\right) \right) \right]$
- stima del parametro ω della distribuzione Gamma: $\hat{\omega} = n \, / \, D$
- valori adattati: $\hat{\mu}_i = \left(X_i \, \hat{\beta}\right)^{-1} \quad \forall i = 1, 2, \ldots, n$
- log-verosimiglianza gamma modello saturo: $\hat{\ell}_{saturo} = \sum_{i=1}^{n} \left[\hat{\omega} \, \left(-1 \log \left(y_i \right) \right) + \left(\hat{\omega} 1 \right) \, \log \left(y_i \right) + \hat{\omega} \, \log \left(\hat{\omega} \right) \log \left(\Gamma \left(\hat{\omega} \right) \right) \right]$

- devianza residua: $D=2\,\hat{\omega}^{-1}\,\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=2\,\sum_{i=1}^{n}\,\left[\left(y_{i}-\hat{\mu}_{i}\right)/\,\hat{\mu}_{i}-\log\left(y_{i}/\,\hat{\mu}_{i}\right)\right]=\sum_{i=1}^{n}\,e_{i}^{2}$
- gradi di libertà della devianza residua: n-k
- log-verosimiglianza gamma modello nullo:

$$\hat{\ell}_{nullo} = \sum_{i=1}^{n} \left[\hat{\omega} \left(-y_i / \bar{y} - \log\left(\bar{y}\right) \right) + \left(\hat{\omega} - 1 \right) \log\left(y_i\right) + \hat{\omega} \log\left(\hat{\omega}\right) - \log\left(\Gamma\left(\hat{\omega}\right)\right) \right]$$

- valori adattati modello nullo: $\hat{\mu} = \bar{y} \quad \forall i = 1, 2, \dots, n$
- devianza residua modello nullo: $D_{nullo} = 2\,\hat{\omega}^{-1}\,\left(\hat{\ell}_{saturo} \hat{\ell}_{nullo}\right)$
- $\bullet\,$ gradi di libertà della devianza residua modello nullo: $\,n-1\,$
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = 1/\bar{y}$

22.2 Stima

glm()

- Package: stats
- Input:

```
formula modello di regressione gamma con k-1 variabili esplicative ed n unità family = Gamma (link="inverse") famiglia e link del modello \mathbf{x} = \mathtt{TRUE} matrice del modello
```

- Description: analisi di regressione gamma
- Output:

```
coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y valori osservati
x matrice del modello
```

• Formula:

coefficients
$$\hat{\beta}_j \quad \forall j=1,2,\ldots,k$$
 residuals
$$e^W_i \quad \forall i=1,2,\ldots,n$$
 fitted.values
$$\hat{\mu}_i \quad \forall i=1,2,\ldots,n$$
 rank
$$k$$
 linear.predictors
$$X\,\hat{\beta}$$

```
deviance
                                             D
     aic
                                        -2\hat{\ell} + 2(k+1)
     null.deviance
                                            D_{nullo}
     weights
                                      w_i \quad \forall i = 1, 2, \ldots, n
     prior.weights
     df.residual
                                            n-k
     df.null
                                            n-1
     У
                                      y_i \quad \forall i = 1, 2, \ldots, n
     X
                                             X
• Examples:
 > x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
 + 4.094345, 4.382027, 4.60517)
 > y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
 > modello <- glm(formula = y \sim x, family = Gamma(link = "inverse"),
 + x = TRUE)
 > modello$coefficients
 (Intercept)
 -0.01655439 0.01534312
 > modello$residuals
                           2
                                          3
  3.219110e-04 -1.669382e-03 -1.245097e-03 -8.626330e-04 1.353051e-03
                           7
                                   8
 -4.456480e-05 1.314954e-03 1.879616e-03 1.414317e-03
 > modello$fitted.values
 122.85903 53.26389 40.00713 34.00264 28.06578 24.97221 21.61432 19.73182
  18.48317
 > modello$rank
 [1] 2
 > modello$linear.predictors
                           3
                     2
                                      4
 0.00813941 \ 0.01877444 \ 0.02499554 \ 0.02940948 \ 0.03563058 \ 0.04004452 \ 0.04626563
 0.05067957 0.05410327
```

```
> modello$deviance
[1] 0.01672967
> modello$aic
[1] 37.9899
> modello$null.deviance
[1] 3.512826
> modello$weights
                  2
                             3
                                      4
15094.6872 2837.0712 1600.5833 1156.1874 787.6926 623.6144 467.1808
       8
                  9
  389.3463 341.6289
> modello$prior.weights
1 2 3 4 5 6 7 8 9
1 1 1 1 1 1 1 1 1
> modello$df.residual
[1] 7
> modello$df.null
[1] 8
> modello$y
  1 2 3 4 5 6 7
118 58 42 35 27 25 21 19 18
> modello$x
 (Intercept) x
          1 1.609438
2
           1 2.302585
3
           1 2.708050
           1 2.995732
5
           1 3.401197
6
           1 3.688879
7
           1 4.094345
           1 4.382027
           1 4.605170
attr(,"assign")
[1] 0 1
```

summary.glm()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- Description: analisi di regressione gamma
- Output:

deviance devianza residua
aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo

df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS

• Formula:

aic
$$-2\,\hat{\ell}+2\,(k+1)$$
 df.residual
$$n-k$$
 null.deviance
$$D_{nullo}$$
 df.null
$$n-1$$
 deviance.resid
$$e_j \ \forall j=1,2,\ldots,k$$
 coefficients
$$\hat{\beta}_j \ s_{\hat{\beta}_j} \ z_{\hat{\beta}_j} \ p\text{-value} = 2\,\Phi(-\,|z_{\hat{\beta}_j}|) \quad \forall j=1,2,\ldots,k$$
 cov.unscaled
$$(X^TW^{-1}X)^{-1}$$
 cov.scaled
$$\hat{\phi}^2\,(X^TW^{-1}X)^{-1}$$
 correlation
$$r_{\hat{\beta}_i\,\hat{\beta}_j} \ \forall i,j=1,2,\ldots,k$$

```
> res$aic
[1] 37.9899
> res$df.residual
[1] 7
> res$null.deviance
[1] 3.512826
> res$df.null
[1] 8
> res$deviance.resid
                              3
-0.040083434 0.086411120 0.049008874 0.029049825 -0.038466050 0.001112469
                      8
-0.028695647 -0.037556945 -0.026372375
> res$coefficients
             Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.01655439 0.0009275454 -17.84752 4.279105e-07
           0.01534312 0.0004149591 36.97501 2.751164e-09
> res$cov.unscaled
             (Intercept)
(Intercept) 0.0003517261 -0.0001474395
           -0.0001474395 0.0000703955
> res$cov.scaled
             (Intercept)
(Intercept) 8.603405e-07 -3.606447e-07
          -3.606447e-07 1.721911e-07
> res$correlation
           (Intercept)
           1.000000 -0.936999
(Intercept)
            -0.936999 1.000000
```

glm.fit()

• Package: stats

• Input:

x matrice del modello

y valori osservati

family = Gamma(link="inverse") famiglia e link del modello

• Description: analisi di regressione gamma

• Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y valori osservati

• Formula:

coefficients $\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$ residuals $e_i^W \quad \forall i = 1, 2, \dots, n$ fitted.values $\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$ rank klinear.predictors $X \hat{\beta}$ deviance Daic $-2\hat{\ell} + 2(k+1)$ null.deviance D_{nullo} weights $w_i \quad \forall i = 1, 2, \ldots, n$ prior.weights df.residual n-kdf.null n-1

У

```
y_i \quad \forall i = 1, 2, \ldots, n
```

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
    4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> modello <- glm(formula = y ~ x, family = Gamma(link = "inverse"))
> X <- model.matrix(object = modello)</pre>
> res <- qlm.fit(x = X, y, family = Gamma(link = "inverse"))</pre>
> res$coefficients
(Intercept)
-0.01655439 0.01534312
> res$residuals
[1] 3.219110e-04 -1.669382e-03 -1.245097e-03 -8.626330e-04 1.353051e-03
[6] -4.456480e-05 1.314954e-03 1.879616e-03 1.414317e-03
> res$fitted.values
[1] 122.85903 53.26389 40.00713 34.00264 28.06578 24.97221 21.61432
[8] 19.73182 18.48317
> res$rank
[1] 2
> res$linear.predictors
[8] 0.05067957 0.05410327
> res$deviance
[1] 0.01672967
> res$aic
[1] 37.9899
> res$null.deviance
[1] 3.512826
> res$weights
[1] 15094.6872 2837.0712 1600.5833 1156.1874 787.6926 623.6144
                                                                  467.1808
[8] 389.3463 341.6289
> res$prior.weights
[1] 1 1 1 1 1 1 1 1 1
> res$df.residual
[1] 7
```

```
> res$df.null
[1] 8
> res$y
[1] 118 58 42 35 27 25 21 19 18
```

vcov()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$\hat{\phi}^2 (X^T W^{-1} X)^{-1}$$

• Examples:

coef()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

```
fit valore previsto
se.fit standard error delle stime
residual.scale radice quadrata della stima del parametro di dispersione
```

• Formula:

fit
$$x_0^T\,\hat{\beta}$$
 se.fit
$$\hat{\phi}\,\sqrt{x_0^T\,(X^T\,W^{-1}\,X)^{-1}\,x_0}$$
 residual.scale
$$\hat{\phi}$$

predict()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime residual.scale radice quadrata della stima del parametro di dispersione

• Formula:

fit
$$x_0^T\,\hat{\beta}$$
 se.fit
$$\hat{\phi}\,\sqrt{x_0^T\,(X^T\,W^{-1}\,X)^{-1}\,x_0}$$
 residual.scale
$$\hat{\phi}$$

fitted()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

fitted.values()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i,\hat{\beta}_i} \quad \forall i,j = 1, 2, \ldots, k$$

• Examples:

22.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

• Description: log-verosimiglianza gamma

• Formula:

 $\hat{\ell}$

• Examples:

AIC()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

$$-2\hat{\ell} + 2(k+1)$$

• Examples:

durbin.watson()

• Package: car

• Input:

model modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D–W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

```
lag Autocorrelation D-W Statistic p-value
    1     0.1835659     1.495257     0
Alternative hypothesis: rho != 0

> res <- durbin.watson(model = modello)
> res$dw

[1] 1.495257
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k - 2\hat{\ell} + 2(k+1)$$

• Examples:

deviance()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

anova()

• Package: stats

• Input:

nullo modello nullo di regressione gamma con n unità modello di regressione gamma con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

• Examples:

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
     4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> nullo <- glm(formula = y ~ 1, family = Gamma(link = "inverse"))</pre>
> modello <- glm(formula = y \sim x, family = Gamma(link = "inverse"))
> anova(nullo, modello, test = "Chisq")
Analysis of Deviance Table
Model 1: y ~ 1
Model 2: y \sim x
 Resid. Df Resid. Dev Df Deviance P(>|Chi|)
         8
                3.5128
                0.0167 1 3.4961 9.112e-313
> res <- anova(nullo, modello, test = "Chisq")</pre>
> res$"Resid. Df"
[1] 8 7
> res$"Resid. Dev"
[1] 3.51282626 0.01672967
```

> res\$Df

22.4 Diagnostica

rstandard()

• Package: stats

• Input:

model modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

rstandard.glm()

• Package: stats

• Input:

model modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

residuals.default()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \dots, n$$

• Examples:

residuals()

- Package: stats
- Input:

```
object modello di regressione gamma con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ e_i & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \hline e_i^R & \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
+     4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> modello <- glm(formula = y ~ x, family = Gamma(link = "inverse"))
> residuals(object = modello, type = "deviance")
```

• Example 2:

• Example 3:

• Example 4:

residuals.glm()

• Package: stats

• Input:

```
object modello di regressione gamma con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

```
type = "deviance" e_i \quad \forall i = 1, 2, \dots, n type = "pearson"
```

```
e_i^P \quad \forall i=1,2,\ldots,n \begin{tabular}{c|c} type = "working" \end{tabular} e_i^W \quad \forall i=1,2,\ldots,n \begin{tabular}{c|c} type = "response" \end{tabular} e_i^R \quad \forall i=1,2,\ldots,n
```

• Example 1:

• Example 2:

• Example 3:

• Example 4:

resid()

• Package: stats

• Input:

```
object modello di regressione gamma con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

```
\begin{array}{ll} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \hline \end{array}
```

• Example 1:

• Example 2:

• Example 3:

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
+          4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> modello <- glm(formula = y ~ x, family = Gamma(link = "inverse"))
> resid(object = modello, type = "working")
```

```
1 2 3 4 5
3.219110e-04 -1.669382e-03 -1.245097e-03 -8.626330e-04 1.353051e-03
6 7 8 9
-4.456480e-05 1.314954e-03 1.879616e-03 1.414317e-03
```

• Example 4:

weighted.residuals()

- Package: stats
- Input:

obj modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

weights()

- Package: stats
- Input:

object modello di regressione di gamma con k-1 variabili esplicative ed n unità

- **Description:** pesi iniziali
- Formula:

$$\underbrace{1, 1, \ldots, 1}_{n \text{ volte}}$$

df.residual()

• Package: stats

• Input:

object modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-k

• Examples:

hatvalues()

• Package: stats

• Input:

model modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

cooks.distance()

• Package: stats

• Input:

model modello di regressione gamma con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

cookd()

- Package: car
- Input:

 ${\tt modell}$ modello di regressione gamma con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

Capitolo 23

Regressione di Wald

23.1 Simbologia

$$1/\mu_i^2 = \beta_1 + \beta_2 \ x_{i1} + \beta_3 \ x_{i2} + \dots + \beta_k \ x_{ik-1}$$
 $Y_i \sim \text{Wald}(\mu_i, \omega) \ \forall i = 1, 2, \dots, n$

- valori osservati: $y_i \quad \forall i = 1, 2, ..., n$
- matrice del modello di dimensione $n \times k$: X
- ullet numero di parametri da stimare e rango della matrice del modello: k
- numero di unità: n
- *i*-esima riga della matrice del modello : $X_i = (1, x_{i1}, x_{i2}, \dots, x_{ik-1}) \quad \forall i = 1, 2, \dots, n$
- vettore numerico positivo dei pesi IWLS: $w = (w_1, w_2, \dots, w_n)$
- matrice diagonale dei pesi IWLS di dimensione $n \times n$: $W = \text{diag}(w_1^{-1}, w_2^{-1}, \dots, w_n^{-1})$
- matrice di proiezione di dimensione $n \times n$: $H = X (X^T W^{-1} X)^{-1} X^T W^{-1}$
- valori di leva: $h_i = H_{i,i} \quad \forall i = 1, 2, \ldots, n$
- distanza di Cook: $cd_i=\left(e_i^P\right)^2\frac{h_i}{\hat{\phi}^2\,k\,(1-h_i)^2}\quad \forall\,i=1,\,2,\,\ldots,\,n$
- stime IWLS: $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k)^T$
- standard error delle stime IWLS: $s_{\hat{\beta}} = \hat{\phi} \sqrt{\mathrm{diag}((X^T W^{-1} X)^{-1})}$
- z-values delle stime IWLS: $z_{\hat{\beta}} = \hat{\beta} \, / \, s_{\hat{\beta}}$
- $\bullet \ \ \text{correlazione delle stime IWLS:} \quad r_{\hat{\beta}_i \ \hat{\beta}_j} \ = \ \frac{\hat{\phi}^2 \ (X^T \ W^{-1} \ X)_{i, \, j}^{-1}}{s_{\hat{\beta}_i} \ s_{\hat{\beta}_j}} \quad \forall \, i, j \, = \, 1, \, 2, \, \ldots, \, k$
- stima del parametro di dispersione: $\hat{\phi}^2 = \frac{1}{n-k} \sum_{i=1}^n \left(e_i^P\right)^2 = \frac{1}{n-k} \sum_{i=1}^n \left(\left(y_i \hat{\mu}_i\right) / \hat{\mu}_i^{3/2}\right)^2$
- residui di devianza: $e_i = \text{sign}(y_i \hat{\mu}_i) \sqrt{(y_i \hat{\mu}_i)^2/(y_i \hat{\mu}_i^2)} \quad \forall i = 1, 2, \dots, n$
- residui standard: $rstandard_i = \frac{e_i}{\hat{\phi}\sqrt{1-h_i}} \quad \forall i=1,2,\ldots,n$
- residui di *Pearson*: $e_i^P = (y_i \hat{\mu}_i) / \hat{\mu}_i^{3/2} \quad \forall i = 1, 2, \ldots, n$
- residui di lavoro: $e_i^W = -2 (y_i \hat{\mu}_i) / \hat{\mu}_i^3 \quad \forall i = 1, 2, \ldots, n$
- residui di riposta: $e_i^R = y_i \hat{\mu}_i \quad \forall i = 1, 2, \dots, n$
- log-verosimiglianza normale inversa: $\hat{\ell} = \frac{n}{2} \log (\hat{\omega}) \frac{3}{2} \sum_{i=1}^{n} \log (2\pi y_i) \hat{\omega} \sum_{i=1}^{n} (y_i \hat{\mu}_i)^2 / (2y_i \hat{\mu}_i^2)$
- stima del parametro ω della distribuzione Wald: $\hat{\omega}=n/D$
- valori adattati: $\hat{\mu}_i = \left(X_i\,\hat{\beta}\right)^{-1\,/\,2} \quad \forall\, i\,=\,1,\,2,\,\ldots,\,n$
- log-verosimiglianza normale inversa modello saturo: $\hat{\ell}_{saturo} = \frac{n}{2} \log{(\hat{\omega})} \frac{3}{2} \sum_{i=1}^{n} \log{(2\pi y_i)}$

- devianza residua: $D=2\,\hat{\omega}^{-1}\,\left(\hat{\ell}_{saturo}-\hat{\ell}\right)=\sum_{i=1}^n\left(y_i-\hat{\mu}_i\right)^2/\left(y_i\,\hat{\mu}_i^2\right)=\sum_{i=1}^n\,e_i^2$
- gradi di libertà della devianza residua: n-k
- $\bullet\,$ log-verosimiglianza normale inversa modello nullo:

$$\hat{\ell}_{nullo} \, = \, \frac{n}{2} \, \log \left(\hat{\omega} \right) - \frac{3}{2} \, \sum_{i=1}^{n} \, \log \left(2 \, \pi \, y_{i} \right) - \hat{\omega} \, \sum_{i=1}^{n} \, \left(y_{i} - \bar{y} \right)^{2} / \left(2 \, y_{i} \, \bar{y}^{2} \right)$$

- valori adattati modello nullo: $\hat{\mu} = \bar{y} \quad \forall i = 1, 2, \dots, n$
- devianza residua modello nullo: $D_{nullo} = 2\,\hat{\omega}^{-1}\,\left(\hat{\ell}_{saturo} \hat{\ell}_{nullo}\right)$
- gradi di libertà della devianza residua modello nullo: n-1
- stima IWLS intercetta modello nullo: $\hat{\beta}_{nullo} = 1/\bar{y}^2$

23.2 Stima



- Package: stats
- Input:

formula modello di regressione normale inversa con k-1 variabili esplicative ed n unità family = inverse.gaussian(link="1/mu^2") famiglia e link del modello x = TRUE matrice del modello

- Description: analisi di regressione normale inversa
- Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua
df.null gradi di libertà devianza residua modello nullo
y valori osservati
x matrice del modello

• Formula:

coefficients $\hat{\beta}_j \quad \forall j=1,2,\dots,k$ residuals $e_i^W \quad \forall i=1,2,\dots,n$ fitted.values $\hat{\mu}_i \quad \forall i=1,2,\dots,n$ rank k linear.predictors $X\,\hat{\beta}$

```
deviance
                                              D
     aic
                                         -2\,\hat{\ell} + 2\,(k+1)
     null.deviance
                                            D_{nullo}
     weights
                                      w_i \quad \forall i = 1, 2, \ldots, n
     prior.weights
     df.residual
                                            n-k
     df.null
                                             n-1
     У
                                      y_i \quad \forall i = 1, 2, \ldots, n
     X
                                              X
• Examples:
 > x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
 + 4.094345, 4.382027, 4.60517)
 > y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
 > modello <- glm(formula = y ~ x, family = inverse.gaussian(link = "1/mu^2"),</pre>
 + x = TRUE)
 > modello$coefficients
  (Intercept)
 -0.001107977 0.000721914
 > modello$residuals
                           2
                                          3
  1.441199e-05 -4.052050e-04 -3.766423e-04 -2.882582e-04 2.402256e-05
                           7
                                   8
  4.397338e-05 3.595650e-04 5.697415e-04 6.762886e-04
 > modello$fitted.values
 136.21078 42.47477 34.36037 30.79207 27.24286 25.35854 23.26344 22.05690
  21.24028
 > modello$rank
 [1] 2
 > modello$linear.predictors
                                       3
                          2
                                                     4
 5.389855e-05 5.542911e-04 8.470019e-04 1.054684e-03 1.347394e-03 1.555076e-03
 1.847788e-03 2.055469e-03 2.216559e-03
```

```
> modello$deviance
[1] 0.006931123
> modello$aic
[1] 61.57485
> modello$null.deviance
[1] 0.08779963
> modello$weights
                  2
                             3
                                      4
632025.412 19157.982 10142.024 7299.044 5054.816 4076.798 3147.514
       8
  2682.741 2395.664
> modello$prior.weights
1 2 3 4 5 6 7 8 9
1 1 1 1 1 1 1 1 1
> modello$df.residual
[1] 7
> modello$df.null
[1] 8
> modello$y
  1 2 3
           4 5 6 7
118 58 42 35 27 25 21 19 18
> modello$x
  (Intercept) x
           1 1.609438
2
           1 2.302585
3
           1 2.708050
           1 2.995732
5
           1 3.401197
6
           1 3.688879
7
           1 4.094345
           1 4.382027
           1 4.605170
attr(,"assign")
[1] 0 1
```

summary.glm()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità correlation = TRUE correlazione delle stime IWLS

- Description: analisi di regressione normale inversa
- Output:

```
deviance devianza residua
aic indice AIC

df.residual gradi di libertà devianza residua
null.deviance devianza residua modello nullo
df.null gradi di libertà devianza residua modello nullo
deviance.resid residui di devianza
coefficients stima puntuale, standard error, z-value, p-value
cov.unscaled matrice di covarianza delle stime IWLS non scalata
cov.scaled matrice di covarianza delle stime IWLS scalata
correlation matrice di correlazione delle stime IWLS
```

• Formula:

aic
$$-2\,\hat{\ell}+2\,(k+1)$$
 df.residual
$$n-k$$
 null.deviance
$$D_{nullo}$$
 df.null
$$n-1$$
 deviance.resid
$$e_j \ \forall j=1,2,\ldots,k$$
 coefficients
$$\hat{\beta}_j \ s_{\hat{\beta}_j} \ z_{\hat{\beta}_j} \ p\text{-value} = 2\,\Phi(-\,|\,z_{\hat{\beta}_j}\,|) \quad \forall j=1,2,\ldots,k$$
 cov.unscaled
$$(X^T\,W^{-1}\,X)^{-1}$$
 cov.scaled
$$\hat{\phi}^2\,(X^T\,W^{-1}\,X)^{-1}$$
 correlation
$$r_{\hat{\beta}_i\,\hat{\beta}_j} \ \forall i,j=1,2,\ldots,k$$

```
> res$aic
[1] 61.57485
> res$df.residual
[1] 7
> res$null.deviance
[1] 0.08779963
> res$df.null
[1] 8
> res$deviance.resid
                                    3
-0.012307674 \quad 0.047994662 \quad 0.034307576 \quad 0.023099121 \quad -0.001715587 \quad -0.002827732
                       8
-0.021231743 -0.031795091 -0.035957248
> res$coefficients
               Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.001107977 1.675366e-04 -6.613343 0.0003005580
            0.000721914 9.468635e-05 7.624267 0.0001237599
> res$cov.unscaled
             (Intercept)
(Intercept) 2.549583e-05 -1.399142e-05
            -1.399142e-05 8.143748e-06
> res$cov.scaled
             (Intercept)
(Intercept) 2.806852e-08 -1.540325e-08
           -1.540325e-08 8.965505e-09
> res$correlation
            (Intercept)
            1.000000 -0.970991
(Intercept)
             -0.970991 1.000000
```

glm.fit()

• Package: stats

• Input:

x matrice del modello

y valori osservati

family = inverse.gaussian(link="1/mu^2") famiglia e link del modello

• Description: analisi di regressione normale inversa

• Output:

coefficients stime IWLS
residuals residui di lavoro
fitted.values valori adattati
rank rango della matrice del modello
linear.predictors predittori lineari
deviance devianza residua
aic indice AIC
null.deviance devianza residua modello nullo
weights pesi IWLS
prior.weights pesi iniziali
df.residual gradi di libertà devianza residua modello nullo
y valori osservati

• Formula:

coefficients $\hat{\beta}_i \quad \forall j = 1, 2, \ldots, k$ residuals $e_i^W \quad \forall i = 1, 2, \dots, n$ fitted.values $\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$ rank klinear.predictors $X \hat{\beta}$ deviance Daic $-2\hat{\ell} + 2(k+1)$ null.deviance D_{nullo} weights $w_i \quad \forall i = 1, 2, \ldots, n$ prior.weights df.residual n-kdf.null n-1

У

 $y_i \quad \forall i = 1, 2, \ldots, n$

```
• Examples:
```

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
    4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> modello <- glm(formula = y ~ x, family = inverse.gaussian(link = "1/mu^2"))
> X <- model.matrix(object = modello)</pre>
> res <- qlm.fit(x = X, y, family = Gamma(link = "inverse"))</pre>
> res$coefficients
(Intercept)
-0.01655439 0.01534312
> res$residuals
[1] 3.219110e-04 -1.669382e-03 -1.245097e-03 -8.626330e-04 1.353051e-03
[6] -4.456480e-05 1.314954e-03 1.879616e-03 1.414317e-03
> res$fitted.values
[1] 122.85903 53.26389 40.00713 34.00264 28.06578 24.97221 21.61432
[8] 19.73182 18.48317
> res$rank
[1] 2
> res$linear.predictors
[8] 0.05067957 0.05410327
> res$deviance
[1] 0.01672967
> res$aic
[1] 37.9899
> res$null.deviance
[1] 3.512826
> res$weights
[1] 15094.6872 2837.0712 1600.5833 1156.1874 787.6926 623.6144
                                                                  467.1808
[8] 389.3463 341.6289
> res$prior.weights
[1] 1 1 1 1 1 1 1 1 1
> res$df.residual
[1] 7
```

```
> res$df.null
[1] 8
> res$y
[1] 118 58 42 35 27 25 21 19 18
```

vcov()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: matrice di covarianza delle stime IWLS
- Formula:

$$\hat{\phi}^2 (X^T W^{-1} X)^{-1}$$

• Examples:

coef()

- Package: stats
- Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

coefficients()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** stime IWLS
- Formula:

$$\hat{\beta}_i \quad \forall j = 1, 2, \dots, k$$

• Examples:

predict.glm()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

```
fit valore previsto
se.fit standard error delle stime
residual.scale radice quadrata della stima del parametro di dispersione
```

• Formula:

fit
$$x_0^T\,\hat{\beta}$$
 se.fit
$$\hat{\phi}\,\sqrt{x_0^T\,(X^T\,W^{-1}\,X)^{-1}\,x_0}$$
 residual.scale
$$\hat{\phi}$$

predict()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità newdata il valore di x_0 se.fit = TRUE standard error delle stime

- **Description:** previsione
- Output:

fit valore previsto se.fit standard error delle stime residual.scale radice quadrata della stima del parametro di dispersione

• Formula:

fit
$$x_0^T\,\hat{\beta}$$
 se.fit
$$\hat{\phi}\,\sqrt{x_0^T\,(X^T\,W^{-1}\,X)^{-1}\,x_0}$$
 residual.scale
$$\hat{\phi}$$

fitted()

• Package: stats

• Input:

object $\,$ modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

fitted.values()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** valori adattati
- Formula:

$$\hat{\mu}_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

cov2cor()

• Package: stats

• Input:

 \lor matrice di covarianza delle stime IWLS di dimensione $k \times k$

- Description: converte la matrice di covarianza nella matrice di correlazione
- Formula:

$$r_{\hat{\beta}_i \hat{\beta}_i} \quad \forall i, j = 1, 2, \ldots, k$$

• Examples:

23.3 Adattamento

logLik()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

• Description: log-verosimiglianza normale inversa

• Formula:

 $\hat{\ell}$

• Examples:

AIC()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** indice AIC
- Formula:

$$-2\hat{\ell} + 2(k+1)$$

• Examples:

durbin.watson()

• Package: car

• Input:

model modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: test di Durbin-Watson per verificare la presenza di autocorrelazioni tra i residui
- Output:

dw valore empirico della statistica D–W

• Formula:

dw

$$\sum_{i=2}^{n} (e_i - e_{i-1})^2 / D$$

```
lag Autocorrelation D-W Statistic p-value
    1     0.5326615     0.7262834     0
Alternative hypothesis: rho != 0

> res <- durbin.watson(model = modello)
> res$dw

[1] 0.7262834
```

extractAIC()

• Package: stats

• Input:

fit modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: numero di parametri del modello ed indice AIC generalizzato
- Formula:

$$k \qquad -2\,\hat{\ell} + 2\,(k+1)$$

• Examples:

deviance()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: devianza residua
- Formula:

D

anova()

• Package: stats

• Input:

nullo modello nullo di regressione normale inversa con n unità modello di regressione normale inversa con k-1 variabili esplicative con n unità test = "Chisq"

- Description: anova di regressione
- Output:

Resid. Df gradi di libertà
Resid. Dev devianza residua
Df differenza dei gradi di libertà
Deviance differenza tra le devianze residue
P(>|Chi|) p-value

• Formula:

Resid. Df
$$n-1 \quad n-k$$
 Resid. Dev
$$D_{nullo} \quad D$$

$$df = k-1$$
 Deviance
$$c = D_{nullo} - D$$
 P(>|Chi|)
$$P(\chi^2_{df} \geq c)$$

• Examples:

> res\$Df

```
> x <- c(1.609438, 2.302585, 2.70805, 2.995732, 3.401197, 3.688879,
     4.094345, 4.382027, 4.60517)
> y <- c(118, 58, 42, 35, 27, 25, 21, 19, 18)
> nullo <- glm(formula = y ~ 1, family = inverse.gaussian(link = "1/mu^2"))</pre>
> modello <- glm(formula = y \sim x, family = inverse.gaussian(link = "1/mu^2"))
> anova(nullo, modello, test = "Chisq")
Analysis of Deviance Table
Model 1: y ~ 1
Model 2: y \sim x
 Resid. Df Resid. Dev Df Deviance P(>|Chi|)
         8 0.087800
              0.006931 1 0.080869 1.029e-17
> res <- anova(nullo, modello, test = "Chisq")</pre>
> res$"Resid. Df"
[1] 8 7
> res$"Resid. Dev"
[1] 0.087799631 0.006931123
```

23.4 Diagnostica

rstandard()

• Package: stats

• Input:

model modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

• Examples:

rstandard.glm()

• Package: stats

• Input:

modello di regressione normale inversa con <math>k-1 variabili esplicative ed n unità

- Description: residui standard
- Formula:

$$rstandard_i \quad \forall i = 1, 2, ..., n$$

residuals.default()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: residui di lavoro
- Formula:

$$e_i^W \quad \forall i = 1, 2, \dots, n$$

• Examples:

residuals()

- Package: stats
- Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

$$\begin{array}{ll} \text{type = "deviance"} \\ e_i & \forall i=1,2,\ldots,n \\ \\ \text{type = "pearson"} \\ \\ e_i^P & \forall i=1,2,\ldots,n \\ \\ \text{type = "working"} \\ \\ e_i^W & \forall i=1,2,\ldots,n \\ \\ \text{type = "response"} \\ \\ e_i^R & \forall i=1,2,\ldots,n \end{array}$$

• Example 1:

• Example 2:

• Example 3:

• Example 4:

residuals.glm()

• Package: stats

• Input:

```
object modello di regressione normale inversa con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo
```

- **Description:** residui
- Formula:

```
e_i^P \quad \forall i=1,2,\ldots,n \texttt{type} = \texttt{"working"} e_i^W \quad \forall i=1,2,\ldots,n \texttt{type} = \texttt{"response"} e_i^R \quad \forall i=1,2,\ldots,n
```

• Example 1:

• Example 2:

• Example 3:

• Example 4:

resid()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità type = "deviance" / "pearson" / "working" / "response" tipo di residuo

- **Description:** residui
- Formula:

```
\begin{array}{l} \text{type = "deviance"} \\ \\ e_i \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "pearson"} \\ \\ e_i^P \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "working"} \\ \\ e_i^W \quad \forall i=1,2,\ldots,n \\ \\ \hline \text{type = "response"} \\ \\ e_i^R \quad \forall i=1,2,\ldots,n \\ \\ \end{array}
```

• Example 1:

• Example 2:

• Example 3:

```
1 2 3 4 5
1.441199e-05 -4.052050e-04 -3.766423e-04 -2.882582e-04 2.402256e-05
6 7 8 9
4.397338e-05 3.595650e-04 5.697415e-04 6.762886e-04
```

• Example 4:

weighted.residuals()

• Package: stats

• Input:

ob j $\,$ modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** residui pesati
- Formula:

$$e_i \quad \forall i = 1, 2, \dots, n$$

• Examples:

weights()

• Package: stats

• Input:

object modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** pesi iniziali
- Formula:

$$\underbrace{1, 1, \ldots, 1}_{n \text{ volte}}$$

df.residual()

• Package: stats

• Input:

object $\,$ modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: gradi di libertà della devianza residua
- Formula:

n-k

• Examples:

hatvalues()

• Package: stats

• Input:

model modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** valori di leva
- Formula:

$$h_i \quad \forall i = 1, 2, \ldots, n$$

cooks.distance()

• Package: stats

• Input:

model modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- **Description:** distanza di *Cook*
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

• Examples:

cookd()

- Package: car
- Input:

 ${\tt modell}$ modello di regressione normale inversa con k-1 variabili esplicative ed n unità

- Description: distanza di Cook
- Formula:

$$cd_i \quad \forall i = 1, 2, \ldots, n$$

Parte VI Appendice

Appendice A

Packages

Package	Descrizione	Status	Versione
actuar	Actuarial functions	Not Installed	0.9-7
base	The R Base Package	Loaded	2.7.0
boot	Bootstrap R (S-Plus) Functions (Canty)	Not Loaded	1.2-32
BSDA	Basic Statistics and Data Analysis	Not Installed	0.1
car	Companion to Applied Regression	Not Installed	1.2-7
corpcor	Efficient Estimation of Covariance and (Partial) Correlation	Not Installed	1.4.7
datasets	The R Datasets Package	Loaded	2.7.0
distributions	Probability distributions based on TI-83 Plus	Not Installed	1.4
e1071	Misc Functions of the Department of Statistics (e1071), TU Wien	Not Installed	1.5-17
formularioR	Formulario di Statistica con R	Not Installed	1.0
faraway	Functions and datasets for books by Julian Faraway.	Not Installed	1.0.3
fBasics	Rmetrics - Markets and Basic Statistics	Not Installed	240.10068.1
foreign	Read Data Stored by Minitab, S, SAS, SPSS, Stata, Systat, dBase,	Not Loaded	0.8-25
fUtilities	Rmetrics - Rmetrics Function Utilities	Not Installed	270.73
graphics	The R Graphics Package	Loaded	2.7.0
grDevices	The R Graphics Devices and Support for Colours and Fonts	Loaded	2.7.0
gtools	Various R programming tools	Not Installed	2.4.0

ineq	Measuring inequality, concentration and poverty	Not Installed	0.2-8
labstatR	Libreria del Laboratorio di Statistica con R	Not Installed	1.0.4
leaps	regression subset selection	Not Installed	2.7
lmtest	Testing Linear Regression Models	Not Installed	0.9-21
MASS	Main Package of Venables and Ripley's MASS	Not Loaded	7.2-41
MCMCpack	Markov chain Monte Carlo (MCMC) Package	Not Installed	0.9-4
methods	Formal Methods and Classes	Loaded	2.7.0
moments	Moments, cumulants, skewness, kurtosis and related tests	Not Installed	0.11
MPV	Data Sets from Montgomery, Peck and Vining's Book	Not Installed	1.25
mvtnorm	Multivariate Normal and T Distribution	Not Installed	0.8-1
nlme	Linear and Nonlinear Mixed Effects Models	Not Loaded	3.1-88
nortest	Tests for Normality	Not Installed	1.0
pastecs	Package for Analysis of Space-Time Ecological Series	Not Installed	1.3-4
Rcmdr	R Commander	Not Installed	1.3-11
schoolmath	Functions and datasets for math used in school	Not Installed	0.2
sigma2tools	Test of hypothesis about sigma2	Not Installed	1.2.6
stats	The R Stats Package	Loaded	2.7.0
strucchange	Testing, Monitoring and Dating Structural Changes	Not Installed	1.3-2
SuppDists	Supplementary distributions	Not Installed	1.1-2
tseries	Time series analysis and computational finance	Not Installed	0.10-13
UsingR	Data sets for the text Using R for Introductory	Not Installed	0.1-8
	Statistics	Wot Instance	

Download Packages from CRAN site

Appendice B

Links

R site search

Site search http://finzi.psych.upenn.edu/search.html

Mailing list archives http://tolstoy.newcastle.edu.au/R/

Help for R (Jonathan Baron) http://finzi.psych.upenn.edu/

R information

CRAN http://cran.r-project.org/

Web site http://www.r-project.org/

News http://cran.r-project.org/doc/Rnews/

R Wiki http://wiki.r-project.org/

Bioconductor http://www.bioconductor.org/

R GUIs

Projects (CRAN) http://www.sciviews.org/_rgui/

R Commander http://socserv.socsci.mcmaster.ca/jfox/Misc/Rcmdr/index.html

SciViews http://www.sciviews.org/SciViews-R/

Tinn-R

SourceForge (main) http://sourceforge.net/projects/tinn-r

SciViews http://www.sciviews.org/Tinn-R

Statistics

Journal of Statistical Soft- http://www.jstatsoft.org/

ware

Electronic Textbook Sta- http://www.statsoftinc.com/textbook/stathome.html

tSoft

Processing

Miktex http://miktex.org/

Deplate http://deplate.sourceforge.net/index.php

Txt2tags
http://txt2tags.sourceforge.net/

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Indice analitico

```
%o%, 81
                                                    atan2, 22
%x%, 116
                                                    atanh, 25
*, 2, 113
                                                    ave, 366
**, 3
                                                    backsolve, 132
+, 1
                                                    bartlett.test, 348
-, 1
.Last.value, 65
                                                    basicStats, 218
                                                    bc, 549
/, 2
                                                    besselI, 45
:, 29
                                                    besselJ, 46
==, 6
                                                    besselK, 46
[ ], 77, 89, 145
                                                    bessely, 46
%*%, 114
%in%, 14
                                                    beta, 42
                                                    BIC, 568, 655
|, 7
                                                    binom.test, 443
| | |, 8
                                                    bonett.test, 493
!, 8
                                                    box.cox, 548
! = , 6
                                                    box.cox.var, 549
응응. 4
                                                    Box.test, 402, 405
응/응, 4
                                                    boxcox, 512, 547
&, 7
                                                    boxplot.stats, 226
&&, 7
<, 5
                                                    bptest, 578
                                                    by, 363
<=, 5
>, 5
                                                    c, 75
>=, 6
                                                    cancor, 180
^, 4
                                                    cbind, 95
abs, 10
                                                    ceiling, 37
acf, 249
                                                    chi2, 212
acos, 21
                                                    chisq.test, 453, 466, 497
acosh, 24
                                                    chol, 140
ad.test, 483
                                                    chol2inv, 141
add1, 523, 572, 618, 660, 706, 740, 775, 808, 841
                                                    choose, 17
AIC, 520, 568, 614, 655, 701, 735, 770, 803, 836,
                                                    codev, 172
                                                    coef, 512, 545, 607, 641, 696, 730, 765, 798, 832,
        868, 892
all, 67
                                                            863, 887
                                                    coefficients, 546, 642, 697, 731, 765, 799, 832,
anova, 351, 353, 355, 521, 615, 703, 737, 772, 805,
                                                            864, 888
       838, 870, 894
                                                    coeftest, 546, 642
anscombe.test, 491
any, 66
                                                    col, 91
aperm, 118
                                                    colMeans, 124
append, 82
                                                    colnames, 87
                                                    colSums, 124
apply, 126
Arg, 49
                                                    complex, 47, 80
array, 143
                                                    Confint, 544, 641
as.dist, 283
                                                    confint, 511, 544, 606, 640
as.factor, 358
                                                    Conj, 49
                                                    cookd, 582, 668, 719, 753, 788, 820, 853, 878, 902
as.integer, 362
                                                    cooks.distance, 527, 582, 623, 668, 718, 752, 787,
as.numeric, 362
as.ordered, 361
                                                            820, 852, 878, 902
as.vector, 107
                                                    cor, 175
asin, 21
                                                    cor.test, 385, 389
asinh, 24
                                                    cor2.test, 394, 398
atan, 21
                                                    cor2pcor, 185
```

```
corr, 201
                                                   drop1, 522, 570, 616, 657, 704, 738, 773, 806, 839
cos, 20
                                                   dsigmoid, 44
cosh, 23
                                                   dsignrank, 245
COV, 170
                                                   dt, 244
                                                   dunif, 245
cov, 171
                                                   duplicated, 229
cov.wt, 191
cov2cor, 178, 518, 566, 612, 653, 700, 734, 769,
                                                   durbin.watson, 519, 567, 613, 654, 701, 735, 770,
                                                           803, 837, 868, 892
       802, 835, 867, 891
covratio, 530, 587, 626, 673
                                                   dweibull, 245
crossprod, 110
                                                   dwilcox, 244
cum3, 205
                                                   E, 213
cummax, 52
                                                   e, 70
cummin, 51
                                                   eigen, 109
cumprod, 51
                                                   emm, 206
cumsum, 50
                                                   eta, 207
cut, 368
                                                   eval, 69
cv, 164
                                                   even, 70
cv2, 165
                                                   exp, 25
cvm.test, 481
                                                   expand.grid, 370
                                                   expm1, 26
D, 57
                                                   expression, 68
d2sigmoid, 45
                                                   extendrange, 152
dbeta, 243
                                                   extractAIC, 520, 568, 614, 656, 702, 736, 771, 804,
dbinom, 237
                                                           837, 869, 893
dburr, 243
dcauchy, 243
                                                   F, 61
dchisq, 243
                                                   factor, 357
DD, 58
                                                   factorial, 18
ddirichlet, 243
                                                   FALSE, 61
det, 100
                                                   fbeta, 43
determinant, 101
                                                   fisher.test, 459
determinant.matrix, 102
                                                   fitted, 513, 550, 607, 643, 699, 733, 768, 801,
deviance, 520, 569, 614, 656, 703, 737, 771, 805,
                                                           834, 866, 890
       838, 869, 893
                                                   fitted.values, 550, 644, 699, 733, 768, 801, 835,
dexp, 243
                                                           867, 891
df, 243
                                                   fivenum, 216
df.residual, 532, 593, 629, 679, 717, 751, 786,
                                                   floor, 36
       819, 851, 877, 901
                                                   forwardsolve, 134
dfbeta, 533, 594, 629, 680
                                                   fractions, 38
dfbetas, 533, 594, 630, 681
                                                   friedman.test, 439
dffits, 529, 587, 625, 673
                                                   ftable, 472
dFriedman, 243
dgamma, 243, 244
                                                   gamma, 39
dgeom, 237
                                                   gcd, 71
dhyper, 238
                                                   geary, 163
diag, 117
                                                   geometcdf, 238
diff, 247
                                                   geometpdf, 238
diffinv, 248
                                                   Gini, 208
digamma, 40
                                                   gini, 209
dim, 85, 119, 144
                                                   ginv, 142
dimnames, 88, 146
                                                   q1, 366
dinvgamma, 244
                                                   glm, 688, 722, 756, 790, 824, 856, 880
dinvGauss, 245
                                                   glm.fit, 693, 727, 762, 795, 829, 861, 885
dist, 281
dlaplace, 244
                                                   hat, 593, 680
dllogis, 244
                                                   hatvalues, 532, 593, 629, 679, 717, 752, 787, 819,
dlnorm, 244
                                                           852, 877, 901
dlogis, 244
                                                   hclust, 285
dmultinom, 238
                                                   head, 80, 92
dmvnorm, 244
                                                   hilbert, 98
dnbinom, 237
                                                   hist, 230
dnorm, 244
dpareto1, 244
                                                   ic.var, 255
dpois, 238
                                                   identical, 66
```

ilogit, 246	mantelhaen.test, 463
Im, 48	margin.table, 469
Inf, 59	match, 67
influence, 589, 675	matrix, 84
influence.measures, 534, 596, 631, 683	max, 149
integrate, <mark>58</mark>	mcnemar.test, 457 , 467
interaction, 369	mean, 153
intersect, 12	mean.a, 155
inv.logit, 246	mean.g, 154
IQR, 158	median, 155
is.complex, 50	median.test, 258
is.element, 13	midrange, 151
is.matrix, 73	min, 149
is.na, 252	Mod, 48
is.nan, 253	model.matrix, 129
is.real, 50	moment, 202
is.vector, 72	mood.test, 450
isPositiveDefinite, 106	
	n.bins, 232
jarque.bera.test, 478	NA, 60
	na.omit, 253
kappa, 130	names, 78
kmeans, 288	NaN, 60
kronecker, 115	nclass.FD, 234
kruskal.test, 432	nclass.scott, 235
ks.test, 477	nclass.Sturges, 234
kurt, 161	NCOL, 122
kurtosis, <mark>162</mark>	ncol, 121
	nlevels, 359
lapply, 64	norm, 104
lbeta, <mark>42</mark>	NROW, 120
lchoose, 17	nrow, 120
leaps, 575 , 662	nsize, 254
length, 94, 247	NULL, 60
LETTERS[], 361	
letters[], 361	numeric, 79
levels, 359	odd, 70
levene.test, 436	oneway.test, 313
lfactorial, 19	optim, 54
lgamma, 40	
-	optimize, 53
1ht, 560, 651	order, 35
lillie.test, 487	ordered, 360
linear.hypothesis, 556, 649	outer, 68
list, 62	outlier.test, 534, 595, 630, 682
lm, 506, 538, 600, 634	5 051
lm.fit, 510, 542	pacf, 251
lm.influence, 530, 588, 626, 674	pairwise.t.test, 381, 383
lm.ridge, 564	partial.cor, 184
lm.wfit, 604 , 638	pascal, 99
lmwork, 529, 586, 624, 672	pbeta, 243
log, <mark>27</mark>	pbinom, 237
log10, 27	psignrank, 245
log1p, 28	pburr, 243
log2, <mark>26</mark>	pcauchy, 243
logb, 28	pchisq, 243
logical, 80	pcor2cor, 187
	pexp, 243
logit, 245	pf, 243
logLik, 519, 567, 613, 654, 700, 734, 769, 802,	pFriedman, 243
836, 867, 891	-
lower.tri, 131	pgamma, 243, 244
ls.diag, 525, 580, 621, 666	pgeom, 237
lsfit, 511 , 543 , 605 , 639	phyper, 238
1 150	pi, 59
mad, 158	pinvGauss, 245
mahalanobis, 284	plaplace, 244

pllogis, 244	rank, 35
plnorm, 244	rational, 39
plogis, 244	rbeta, 243
pmax, 53	rbind, 96
pmin, 52	rbinom, 237
pmvnorm, 244	rburr, 243
pnbinom, 237	rcauchy, 243
pnorm, 244	rchisq, 243
polyroot, 56	rdirichlet, 243
popstderror, 167	Re, 47
power.prop.test, 341	relevel, 358
pparetol, 244	rep, 29
ppoints, 496	rep.int, 30
ppois, 238	replace, 69
prcomp, 264, 273	resid, 592, 678, 714, 748, 783, 816, 848, 875, 899
predict, 516, 553, 610, 647, 698, 732, 767, 800,	
-	residuals, 591, 677, 710, 744, 779, 812, 845, 872, 896
834, 865, 889	
predict.glm, 697, 731, 766, 799, 833, 864, 888	residuals.default, 592, 678, 709, 743, 778, 811,
predict.lm, 514, 551, 607, 644	844, 872, 896
PRESS, 521 , 569 , 615 , 657	residuals.glm, 712, 746, 781, 814, 847, 873, 897
princomp, 261, 270	residuals.lm, 531, 591, 628, 677
prod, 9	rev, 34
prop.table, 470	rexp, 243
prop.test, 337, 342, 346	rf, 243
psigamma, 41	rFriedman, 243
pt, 244	rgamma, 243 , 244
ptukey, 244	rgeom, 237
punif, 245	rhyper, 238
pweibull, 245	rinvgamma, 244
pwilcox, 244	rinvGauss, 245
	rk, <mark>99</mark>
qbeta, 243	rlaplace, 244
qbinom, 237	rllogis, 244
qburr, 243	rlnorm, 244
qcauchy, 243	rlogis, 244
qchisq, 243	rmultinom, 238
qexp, 243	rmvnorm, 244
qf, 243	rnbinom, 237
qFriedman, 243	rnorm, 244
qgamma, 243, 244	round, 37
qgeom, 237	row, 91
qhyper, 238	rowMeans, 123
qinyGauss, 245	
-	rownames, 86
qlaplace, 244	rowsum, 125
qllogis, 244	rowSums, 122
qlnorm, 244	rparetol, 244
qlogis, 244	rpois, 238
qnbinom, 237	RS, 211
qnorm, 244	rsignrank, 245
qparetol, 244	rstandard, 527, 583, 623, 669, 707, 741, 776, 809,
qpois, 238	842, 871, 895
qqnorm, 495	rstandard.glm, 708, 742, 777, 810, 843, 871, 895
qr.Q, 138	rstandard.lm, 527 , 583 , 623 , 669
qr.R, 139	rstudent, 528, 584, 670, 708, 742, 777, 810, 843
qsignrank, 245	rstudent.glm, 709, 743, 778, 811, 844
qt, 244	rstudent.lm, 528, 585, 624, 671
qtukey, 244	rt, 244
quantile, 156	runif, 245
qunif, 245	runs.test, 446
qweibull, 245	rweibull, 245
qwilcox, 244	rwilcox, 244
- -	
range, 150	sample, 254
range2, 150	sapply, 82
	± ± 4′

```
scale, 204
                                                   Var, 174
scan, 77
                                                   var, 169
scm, 71
                                                   var.coeff, 164
sd, 166
                                                   var.test, 334
                                                   vcov, 509, 541, 604, 638, 696, 730, 764, 798, 832,
seq, 31
seq_along, 32
                                                          863, 887
                                                   vech, 93
seq_len, 33
                                                   vector, 79
sequence, 31
                                                   vif, 595, 682
set.seed, 256
setdiff, 13
                                                   weighted.mean, 188
setequal, 14
                                                   weighted.residuals, 628, 676, 716, 750, 785, 818,
sf.test, 485
                                                          850, 876, 900
sigma, 166
                                                   weights, 627, 676, 716, 751, 786, 818, 851, 876,
sigma2, 168
sigma2.test, 331
                                                   which, 15
sigma2m, 172
                                                   which.max, 16
sigmoid, 44
                                                   which.min, 15
sign, 11
                                                   wilcox.test, 409, 413, 416, 421, 425, 428
signif, 38
                                                   wt.moments, 190
simple.z.test, 257
                                                   wt.var, 189
sin, 19
sinh, 22
                                                   xor, 8
skew, 159
                                                   xpnd, 94
skewness, 160
                                                   xtabs, 472
solve, 107
solveCrossprod, 128
                                                   z.test, 293, 299
sort, 33
                                                   zsum.test, 313, 320
sqrt, 11
ssdev, 170
stat.desc, 222
stderror, 168
stdres, 584, 670
studres, 585, 671
subset, 84
sum, 9
summary, 214, 266, 276, 368, 474
summary.glm, 691, 725, 759, 793, 827, 859, 883
summary.lm, 508, 540, 602, 636
svd, 135
sweep, 256
T, 61
t, 117
t.test, 296, 302, 306, 309
table, 228
tabulate, 227
tail, 81, 93
tan, 20
tanh, 23
tapply, 365
tcrossprod, 111
toeplitz, 97
tr, 104
trigamma, 41
TRUE, 60
trunc, 36
tsum.test, 316, 323, 327
TukeyHSD, 373, 375, 378
unclass, 363
union, 12
unique, 229
uniroot, 55
```

upper.tri, 131