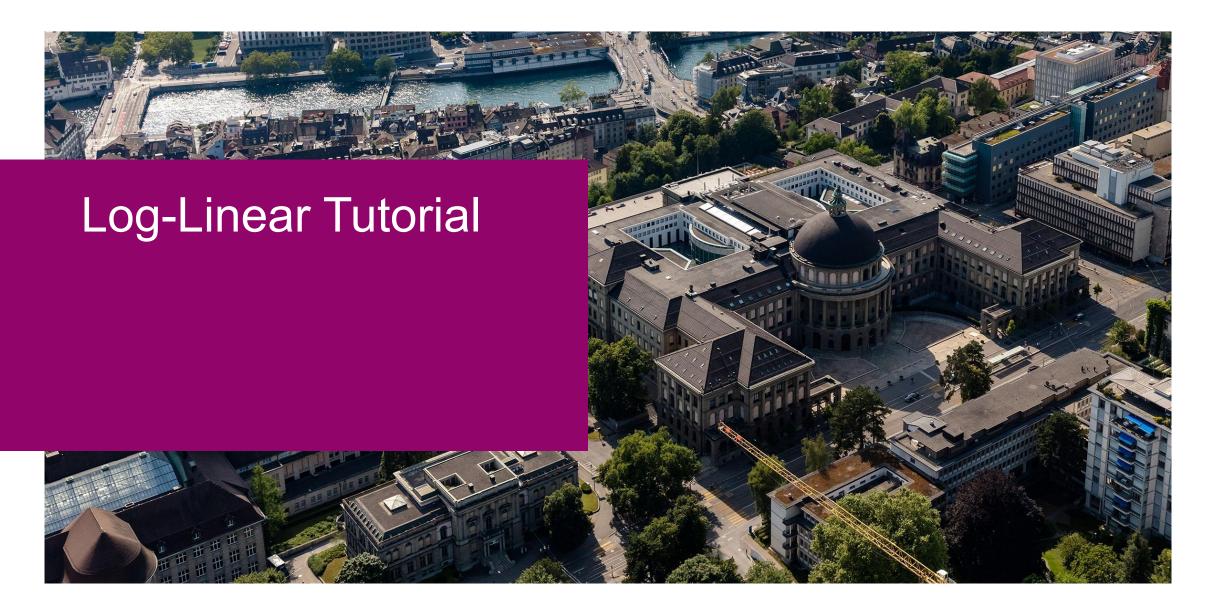
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- 1 Examples on random variables
- The gradient of the log linear model
- The exponential family
- Interactive visualization





Examples on Random Variables

What is a random variable?

- A Random Variable X is a function that maps outcomes of random experiments to a set of properties. (We will dig more into this later.)
- A Probability Distribution p(X=x) is a function that measures the probability that outcomes with the particular property x will occur

Example: rolling two dices.

Outcome: {2, 3}

Property: (1) sum of the numbers; (2) product of the numbers; ...

Value/measure: (1) 5; (2) 6; ...

$$6 \stackrel{\text{random variable: product}}{\longleftarrow} \{2, 3\} \stackrel{\text{random variable: sum}}{\longrightarrow} 5$$

adapted from A. Aldo Faisal, Cheng Soon Ong, and Marc Peter Deisenroth Mathematics for Machine Learning



1 Why do we need Random Variables?

- Random variables are fundamentally about interactions between different properties
 of elements of the sample space
- Independence and correlation are properties of random variables and not of the probability spaces

Example: rolling two dices.

- 1. The probability that (the sum is smaller than 5) while (the product is larger than 5)
- 2. ...





Finding the partial derivative

$$L(oldsymbol{ heta}) = -\sum_{n=1}^N \log p(y_n \mid x_n, oldsymbol{ heta})$$

$$\sum_{j=1}^{m} \log \Pr_{\vec{\theta}}(y_j^* \mid x_j)$$

$$p(y \mid x, oldsymbol{ heta}) = rac{\exp(oldsymbol{ heta} \cdot \mathbf{f}(x, y))}{\sum_{y' \in \mathcal{Y}} \exp(oldsymbol{ heta} \cdot \mathbf{f}(x, y'))}$$



$$\Pr_{\vec{\theta}}(y \mid x) = \frac{\exp\left(\vec{\theta} \cdot \vec{f}(x, y)\right)}{\sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x, y')\right)}$$

⁽²⁾ The Gradient of a Log-Linear Model

$$\sum_{j=1}^{m} \log \frac{\exp\left(\vec{\theta} \cdot \vec{f}(x,y)\right)}{\sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x,y')\right)}$$

$$\mathcal{L}\left(\vec{\theta}\right) = \left(\sum_{j=1}^{m} \vec{\theta} \cdot \vec{f}(x_j, y_j^*)\right) - \sum_{j=1}^{m} \log \sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x_j, y')\right)$$

$$\sum_{j=1}^{m} \frac{\exp\left(\vec{\theta} \cdot \vec{f}(x,y)\right)}{\sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x,y')\right)}$$

$$\mathcal{L}\left(\vec{\theta}\right) = \left(\sum_{j=1}^{m} \vec{\theta} \cdot \vec{f}(x_j, y_j^*)\right) - \sum_{j=1}^{m} \left(\log \sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x_j, y')\right)\right)$$



$$\mathcal{L}\left(\vec{\theta}\right) = \left(\sum_{j=1}^{m} \vec{\theta} \cdot \vec{f}(x_j, y_j^*)\right) - \sum_{j=1}^{m} \log \sum_{y'} \exp\left(\vec{\theta} \cdot \vec{f}(x_j, y')\right)$$

$$\left(\sum_{j=1}^{m} \sum_{k} \theta_k f_k(x_j, y_j^*)\right) - \sum_{j=1}^{m} \log \sum_{y'} \exp \left(\sum_{k} \theta_k f_k(x_j, y')\right)$$



$$\mathcal{L}\left(\vec{\theta}\right) = \left(\sum_{j=1}^{m} \sum_{k} \theta_k f_k(x_j, y_j^*)\right) - \sum_{j=1}^{m} \log \sum_{y'} \exp\left(\sum_{k} \theta_k f_k(x_j, y')\right)$$

Take the **derivative** with respect to $heta_{\ell}$, we have

$$\frac{\partial \mathcal{L}}{\partial \theta_{\ell}} = \left(\sum_{j=1}^{m} f_{\ell}(x_j, y_j^*)\right) - \sum_{j=1}^{m} \frac{\sum_{y'} \left(\exp \sum_{k} \theta_k f_k(x_j, y')\right) f_{\ell}(x_j, y')}{\sum_{y'} \exp \sum_{k} \theta_k f_k(x_j, y')}$$



The Gradient of a Log-Linear Model $\frac{\partial \mathcal{L}}{\partial \theta_{\theta}}$

$$\sum_{j=1}^{m} \log \left(\sum_{y'} \exp \left(\sum_{k} \theta_k f_k(x_j, y') \right) \right)$$

$$(\ln x)' = rac{1}{x} \qquad \sum_{j=1}^m rac{\sum_{y'} \exp \sum_k heta_k f_k(x_j, y')}$$

The Gradient of a Log-Linear Model $\frac{\partial \mathcal{L}}{\partial \theta_{\theta}}$

$$\sum_{j=1}^{m} \log \sum_{y'} \left(\exp \left(\sum_{k} \theta_{k} f_{k}(x_{j}, y') \right) \right)$$

$$(e^{x})' = e^{x} \sum_{j=1}^{m} \frac{\sum_{y'} (\exp \sum_{k} \theta_{k} f_{k}(x_{j}, y'))}{\sum_{y'} \exp \sum_{k} \theta_{k} f_{k}(x_{j}, y')}$$

The Gradient of a Log-Linear Model $\frac{\partial \mathcal{L}}{\partial \theta}$

$$\sum_{j=1}^{m} \log \sum_{y'} \exp \left(\sum_{k} \theta_k f_k(x_j, y') \right)$$

$$\sum_{j=1}^{m} \frac{\sum_{y'} (\exp \sum_{k} \theta_{k} f_{k}(x_{j}, y')) f_{\ell}(x_{j}, y')}{\sum_{y'} \exp \sum_{k} \theta_{k} f_{k}(x_{j}, y')}$$

(2)

The Gradient of a Log-Linear Model

$$\left(\sum_{j=1}^{m} f_{\ell}(x_{j}, y_{j}^{*})\right) - \sum_{j=1}^{m} \frac{\sum_{y'} \left(\exp \sum_{k} \theta_{k} f_{k}(x_{j}, y')\right) f_{\ell}(x_{j}, y')}{\sum_{y'} \exp \sum_{k} \theta_{k} f_{k}(x_{j}, y')}$$



$$\sum_{j=1}^{m} \left(f_{\ell}(x_j, y_j^*) - \frac{\sum_{y'} (\exp \sum_{k} \theta_k f_k(x_j, y')) f_{\ell}(x_j, y')}{\sum_{y'} \exp \sum_{k} \theta_k f_k(x_j, y')} \right)$$



$$\sum_{j=1}^{m} \left(f_{\ell}(x_j, y_j^*) - \frac{\sum_{y'} \left(\exp \sum_{k} \theta_k f_k(x_j, y') \right) f_{\ell}(x_j, y')}{\sum_{y'} \exp \sum_{k} \theta_k f_k(x_j, y')} \right)$$

$$\sum_{j=1}^{m} \left(f_{\ell}(x_j, y_j^*) - \sum_{y'} \underbrace{\Pr_{\vec{\theta}} \left(y' \mid x_j \right)} f_{\ell}(x_j, y') \right)$$



$$\sum_{j=1}^{m} \left(f_{\ell}(x_j, y_j^*) - \left[\sum_{y'} \Pr_{\vec{\theta}} \left(y' \mid x_j \right) f_{\ell}(x_j, y') \right) \right)$$

observed feature "counts"

expected feature "count"





The Exponential Family

Bernoulli

Gaussian

Why "The Exponential Family"?

• The **exponential family** is a family of probability distributions over $x \in X$, parameterized by some θ , of the form

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

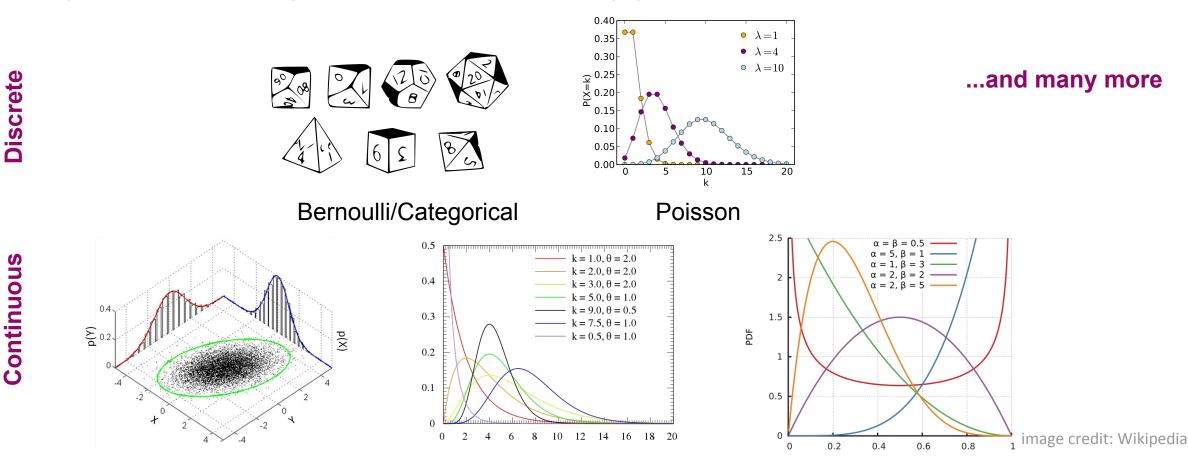
where

- $Z(\theta)$ is the partition function
- h(x) determines the support (exact zeros in the model)
- θ are the canonical parameters
- $\phi(x)$ are the sufficient statistics
 - This is the same as a feature function! Just different terminology between statistics and NLP!

3

Why care about the Exponential Family?

- This is just one of the many ways to define the joint distribution between x and θ , why should you care?
- If you prove something about the exponential family, you've proven it about a lot of distributions at once!



Gamma

Gaussian

Beta

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

Bernoulli Distribution: Standard formulation

The Bernoulli for $x \in \{0, 1\}$

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$$(x|\mu) = \mu^x (1-\mu)^{1-x}$$

 $= \exp \log \mu^x (1-\mu)^{1-x}$
 $= \exp[\log \mu^x + \log(1-\mu)^{1-x}]$
 $= \exp[x \log(\mu) + (1-x) \log(1-\mu)]$

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

Bernoulli Distribution: Standard formulation

$$= \exp[x \log(\mu) + (1 - x) \log(1 - \mu)]$$

$$=\exp[\boldsymbol{\phi}(x)^T\boldsymbol{\theta}]$$

where
$$\phi(x) = [\mathbb{I}(x=0), \mathbb{I}(x=1)]$$
 and $\boldsymbol{\theta} = [\log(\mu), \log(1-\mu)]$.

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

Bernoulli Distribution: Standard formulation

$$= \exp[x \log(\mu) + (1 - x) \log(1 - \mu)]$$

$$=\exp[\boldsymbol{\phi}(x)^T\boldsymbol{\theta}]$$

where
$$\phi(x) = [\mathbb{I}(x=0), \mathbb{I}(x=1)]$$
 and $\boldsymbol{\theta} = [\log(\mu), \log(1-\mu)]$.

Is this representation good?

$$= \exp[x \log(\mu) + (1 - x) \log(1 - \mu)]$$

where
$$\phi(x) = [\mathbb{I}(x=0), \mathbb{I}(x=1)]$$
 and $\boldsymbol{\theta} = [\log(\mu), \log(1-\mu)]$.

there is a **linear dependence** between the features

$$\mathbf{1}^T \boldsymbol{\phi}(x) = \mathbb{I}(x=0) + \mathbb{I}(x=1) = 1$$

Consequently θ is not uniquely identifiable. It is common to require there is a unique θ associated with the distribution.

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

Bernoulli Distribution: Standard formulation

The Bernoulli for $x \in \{0, 1\}$

$$Ber(x|\mu) = \mu^x (1-\mu)^{1-x} = \mu^x (1-\mu)(1-\mu)^{-x}$$

$$= (1-\mu)(\frac{\mu}{1-\mu})^x = (1-\mu)\exp\log(\frac{\mu}{1-\mu})^x$$

$$= (1-\mu)\exp(x\log(\frac{\mu}{1-\mu}))$$

adapted from Kevin P. Murphy, Machine Learning: A Probabilistic Perspective, Chapter 9: Generalized linear models and the exponential family

The Gaussian is an Exponential Family Distribution

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

Gaussian Distribution: Standard formulation

$$\mathcal{N}(x|\mu,\sigma^{2}) = \frac{1}{(2\pi\sigma^{2})^{\frac{1}{2}}} \exp\left[-\frac{1}{2\sigma^{2}}(x-\mu)^{2}\right]$$

$$= \frac{1}{(2\pi\sigma^{2})^{\frac{1}{2}}} \exp\left[-\frac{1}{2\sigma^{2}}x^{2} + \frac{\mu}{\sigma^{2}}x \left[-\frac{1}{2\sigma^{2}}\mu^{2}\right]\right]$$

$$= \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left[-\frac{\mu^{2}}{2\sigma^{2}}\right] \exp\left[-\frac{1}{2\sigma^{2}}x^{2} + \frac{\mu}{\sigma^{2}}x\right]$$

The Gaussian is an Exponential Family Distribution

$$p(x \mid \boldsymbol{\theta}) = \frac{1}{Z(\boldsymbol{\theta})} h(x) \exp(\boldsymbol{\theta} \cdot \boldsymbol{\phi}(x))$$

$$\frac{1}{\sqrt{2\pi\sigma^2}} \exp[-\frac{\mu^2}{2\sigma^2}] \exp[-\frac{1}{2\sigma^2}x^2 + \frac{\mu}{\sigma^2}x]$$

$$Z(\theta) = \sqrt{2\pi\sigma^2} \exp(\frac{\mu^2}{2\sigma^2})$$

$$h(x) = 1$$

$$\theta = \left[-\frac{1}{2\sigma^2}, \frac{\mu}{\sigma^2} \right]^T$$

$$\phi(x) = [x^2, x]^T$$



Interactive Visualization

4 Interactive visualization



Welcome! This interactive visualization will help you understand the popular technique of log-linear modeling.

Try it out: The sliders below control the parameters ("weights") of a log-linear model. When you increase the circle weight, which filled shapes get bigger? Which ones get smaller?

One game is to try to match all 4 shapes to the **gray outlines**. You will need to use both sliders. A shape will turn gray if it matches well. It turns red if it is too small, blue if it is too big. *Note:* You may like to zoom in with your browser.

What the picture means: Your model defines a probability for each shape. You're adjusting these *model probabilities* by changing the weights. When the weights are 0, all 4 filled shapes have equal probability of ¼, as shown by their equal areas.

