

For the eqn.

$$(\lambda + \mu)P_n = \lambda P_{n-1} + \mu P_{n+1}$$

Substitute by the following in the RHS

$$P_n = \rho P_0$$

$$\therefore \lambda \rho^{n-1} P_0 + \mu \rho^{n+1} P_0 = (\lambda \rho^{n-1} + \mu \rho^{n+1}) P_0$$

$$= P_0 \rho^n (\lambda \rho^{-1} + \mu \rho) = P_0 \rho^n (\mu + \lambda)$$

= L.H.S.

$\therefore P_n = \rho P_0$ satisfies the equation.

For the eqn.

$$(\lambda + \mu)P_n = \lambda P_{n-1} + \mu P_{n+1}$$

substitute by

$$P_{n+1} = \rho P_n \text{ i.e. } (P_n = P_{n+1} / \rho) \text{ in the R.H.S. of the equation}$$

$$\therefore \lambda P_n / \rho + \mu \rho P_n = \mu P_n + \lambda P_n = P_n (\lambda + \mu) = \text{L.H.S.}$$

$\therefore P_{n+1} = \rho P_n$ satisfies the equation.

For the M/M/1 queue

$$P_0 = 1 - \sum_{n=1}^N P_n$$

$$= 1 - \frac{P_1 (1 - (\frac{P_n}{P_{n-1}})^N)}{1 - (\frac{P_n}{P_{n-1}})}$$

$$= 1 - \frac{P_0 \rho (1 - \rho^N)}{1 - \rho}$$

$$\therefore P_0 [1 + \frac{(\rho - \rho^{N+1})}{1 - \rho}] = 1$$

$$\therefore P_0 = \frac{1 - \rho}{1 - \rho^{N+1}}$$

$$\text{and } \therefore P_n = P_0 \rho^n = \frac{(1 - \rho) \rho^n}{1 - \rho^{N+1}}$$

Problem 2 (Schwartz 2.9)

For finite m/m/1 queue

$$\gamma = \lambda(1 - P_B) = \mu(1 - P_0)$$

$$\therefore 1 - P_B = \frac{\mu}{\lambda}(1 - P_0)$$

$$\text{so } P_B = 1 - \frac{\mu}{\lambda}(1 - P_0)$$

$$\therefore P_0 = \frac{1 - \rho}{1 - \rho^{N+1}}$$

$$\begin{aligned} \therefore P_B &= 1 - \frac{1}{\rho} + \frac{1}{\rho} \frac{1 - \rho}{1 - \rho^{N+1}} = (1 - \rho) \rho^N \left[\frac{-1}{\rho^{N+1}} + \frac{1}{1 - \rho^{N+1}} \right] \\ &= (1 - \rho) \rho^N \left[\frac{\rho^{N+1} - 1 + 1}{\rho^{N+1}(1 - \rho^{N+1})} \right] = \frac{(1 - \rho) \rho^N}{1 - \rho^{N+1}} \\ &= P_N \end{aligned}$$

Problem 3 (Schwartz 2.10)

$$P_B = \frac{(1 - \rho) \rho^N}{1 - \rho^{N+1}}$$

1. For $\rho = 0.5$, $P_B = 10^{-3}$
 - a. We can get the length of the buffer by trial and error
N= 9 customers
 - b. For $\rho = 0.5$, $P_B = 10^{-6}$
N=19 customers
2. For $\rho = 0.8$,
 - a. $P_B = 10^{-3}$
N=23 customers
 - b. For $\rho = 0.8$, $P_B = 10^{-6}$
N= 29 customers

From the results we conclude that, in order to decrease the blocking probability of the system or to increase the utilization of the system we have to use a longer buffer.

Problem 4 (Shwartz 2-11)

$$\begin{aligned}
 E[n] &= \sum_{n=0}^{\infty} nP_n = \sum_{n=0}^{\infty} nP_0 \rho^n = \sum_{n=0}^{\infty} n(1-\rho)\rho^n = (1-\rho) \sum_{n=0}^{\infty} n\rho^n = (1-\rho)\rho \sum_{n=1}^{\infty} n\rho^{n-1} \\
 &\because \sum_{n=0}^{\infty} \rho^n = \frac{1}{1-\rho} \\
 &\therefore \frac{d}{dn} \sum_{n=1}^{\infty} \rho^n = \frac{1}{(1-\rho)^2} \\
 \text{and so, } E[n] &= \frac{\rho}{1-\rho}
 \end{aligned}$$

Problem 5

1. For M / M / ∞
 $\lambda_n = \lambda$, $\mu = n\mu$

$$\therefore P_n = P_0 \frac{\rho^n}{n!}$$

From the property of probability, we get $\sum_{i=0}^{\infty} P_i = 1$

$$\begin{aligned}
 \therefore \sum_{n=0}^{\infty} P_0 \frac{\rho^n}{n!} &= 1 \\
 P_0 &= e^{-\rho}
 \end{aligned}$$

$$\therefore P_n = e^{-\rho} \frac{\rho^n}{n!} \text{ which is the Poisson's distribution}$$

$$E[n] = \rho; \quad E[T] = \frac{\rho}{\lambda} = \frac{1}{\mu}$$

$$E[w] = E[T] - \frac{1}{\mu} = 0; \quad E[q] = 0$$

2. For queue with discouragement

$$\lambda_n = \frac{\lambda}{n+1}, \quad \mu_n = \mu$$

$$\therefore P_n = \frac{\lambda^n P_0}{n! \mu^n} = \frac{P_0 \rho^n}{n!}$$

Thus we get the value of $P_0 = e^{-\rho}$ which is the same as the M/M/ ∞ case

Similarly $P_n = \frac{e^{-\rho} \rho^n}{n!}$ is the Poisson's distribution

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Problem 6 (Galleger 3.8)

- a. Given $\lambda=10$ packets/sec
 $\tau(\text{Transmission Time})=20\text{msec}=0.02\text{sec}$

$$\therefore \text{Probability [Inter-arrival time } < \text{ Transmission Time]} = 1 - \text{Prob. [no arrival of Packets]} \\ = 1 - e^{-\lambda\tau}$$

$$\therefore \text{Probability of no collision with the predecessor} = 1 - (1 - e^{-\lambda\tau})$$

$$\begin{aligned} \text{Also Probability of no collision with the predecessor and successor} \\ = e^{-\lambda\tau} * e^{-\lambda\tau} \\ = e^{-(0.2+0.2)} = e^{-0.4} = 0.67 \end{aligned}$$

Also Prob. that a packet does-not collide with another Packet will be the same = 0.67

Problem 7 (Galleger 3.9)

$$\begin{aligned} \lambda &= 150 \text{ packets/min/session} \\ \mu &= 50 * 10^3 \text{ bits/sec} / 1000 \text{ bits} = 50 \text{ packets / sec} = 3000 \text{ packets / min} \\ &= 300 \text{ packets/min/session} \\ \rho &= \lambda/\mu = 150/300 = 0.5 \end{aligned}$$

- a) For TDM with 10 channels ($m=10$)
 $E[n] = m\rho/(1-\rho) = 10*0.5/(1-0.5) = 10$ packets
 $E[T] = m / (\mu-\lambda) = 10 / (300-150) = 6.67 * 10^{-3} \text{ min} = 0.4 \text{ sec}$
 $E[q] = E[n] - m\rho = 10 - 10*0.5 = 5$ packets

For Statistical Multiplexing

$$\begin{aligned} E[n] &= \rho/(1-\rho) = 0.5 / (1-0.5) = 1 \text{ packet} \\ E[T] &= 1 / (\mu-\lambda) = 1/(3000-1500) = 6.67 * 10^{-4} \text{ min} = 0.04 \text{ sec} \\ E[q] &= E[n] - \rho = 1 - 0.5 = 0.5 \text{ packet} \end{aligned}$$

- b) In case of five sessions have rates 250 packets/min and other five have rates of 50 packets/min

$$\begin{aligned} \text{For the 1}^{\text{st}} \text{ five sessions} \\ \lambda &= 250 * 5 = 1250 \text{ packets/min} \\ \mu &= 3000 * 5 / 10 = 1500 \text{ packets/min} \\ &= 3000 * 5 / 10 = 1500 \text{ packets/min} \\ \rho &= \lambda/\mu = 1250/1500 = 0.833 \\ E[n] &= 5 * 0.83 / (1 - 0.83) = 24.9 \text{ packets} \\ &0.166) = 1.1 \text{ packets} \end{aligned}$$

$$\begin{aligned} \text{For the 2}^{\text{nd}} \text{ five sessions} \\ \lambda &= 50 * 5 = 250 \text{ packets/min} \\ \mu & \\ \rho &= \lambda/\mu = 1250/1500 = 0.166 \\ E[n] &= 5 * 0.166 / (1 - \end{aligned}$$

$$E[n] \text{ total} = 24.9 + 1.1 = 26 \text{ packets}$$

$$E[T] = m / (\mu - \lambda) = 5 / (1500 - 1250) = 0.02 \text{ min} \quad E[T] = 5 / (1500 - 250) = 4 \times 10^{-3} \text{ min}$$

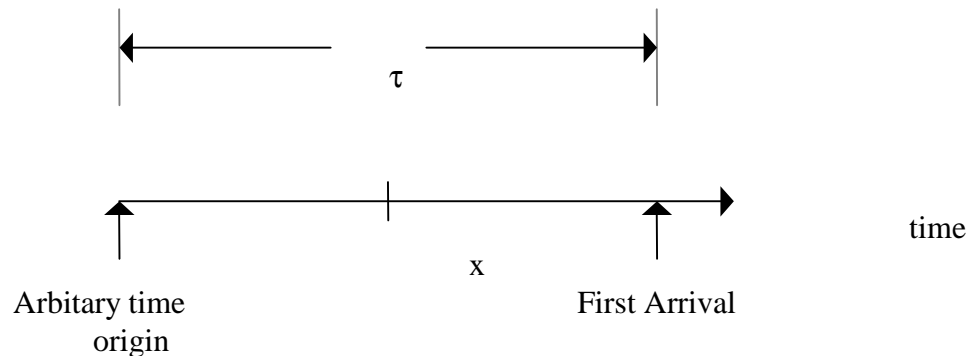
$$E[T] \text{ total} = (0.02 + 0.004) \times 60 = 1.44 \text{ sec}$$

$$E[q] = E[n] - m\rho = 24.9 - 5 \times 0.83 = 20.735 \text{ packets} \quad E[q] = 1.1 - 5 \times 0.16 = 0.27 \text{ packets}$$

$$E[q] \text{ total} = 20.735 + 0.27 = 21 \text{ packets}$$

Problem 8 (Galleger 3.10)

Consider the time diagram shown



- a. Let τ representing the time the first arrival after some arbitrary time origin.
 Take any value x .
 No arrival occur in the interval $(0, x)$ if and only if $\tau > x$.
 The probability that no arrivals occur in $(0, x)$; i.e.
 $P(\tau > x) = \text{prob.}(\text{no arrivals in } (0, x))$

For Poisson's distribution

$$p(k) = (\lambda t)^k e^{-\lambda t} / k!$$

here $k=0$

$$\therefore p(\tau > x) = e^{-\lambda x}$$

Then the prob. that $\tau \leq x = 1 - e^{-\lambda x}$

On the basis, the probability of packet arrivals in the small time interval

$(t, t + \delta t)$ is just $\lambda \delta t + o(\delta t)$

Let $N^{(i)}(t, t + \delta t)$ be the number of events in Poisson process I , $I=1, 2, \dots, m$ in the interval $(t, t + \delta t)$.

Let $N(t, t + \delta t)$ be the total number of events from the whole stream.

$$\text{Then } \text{Prob.}[N(t, t + \delta t) = 0] = \prod_{i=1}^n \text{prob.}[N^{(i)}(t, t + \delta t) = 0]$$

Since the probability of no packet arrival in the time interval $(t, t + \delta t)$ is $(1 - \lambda_i \delta t + o(\delta t))$

$$\begin{aligned} \therefore \text{prob.}[N(t, t + \delta t) = 0] &= \prod_{i=1}^n (1 - \lambda_i \delta t + o(\delta t)) \\ &= 1 - \lambda \delta t + o(\delta t) + \text{higher powers of } (\delta t) \text{ that goes to zero} \end{aligned}$$

Where $\lambda = \sum_{i=1}^m \lambda_i$, since the individual process are independent

$$\begin{aligned} \text{Now } \text{prob.}[N(t, t + \delta t) = 1] &= \sum_{i=1}^n \lambda_i \delta t + o(\delta t) \\ &= \lambda \delta t + o(\delta t) + \text{higher powers of } (\delta t) \text{ which goes to zero} \end{aligned}$$

$$\begin{aligned} \text{prob.}[N(t, t + \delta t) > 2] \\ &= 1 - \text{prob.}[N(t, t + \delta t) = 0] - \text{prob.}[N(t, t + \delta t) = 1] \\ &= o(\delta t) \end{aligned}$$

- b. For a R.V. z with Poisson distribution $P_z[(N_1 + N_2) = n]$ is given by

$$\begin{aligned} P_z(n) &= \sum_{k=0}^n P[N_1 = k] P[N_2 = n - k] \\ &= \sum_{k=0}^n P_x(k) P_y(n - k) \\ &= \sum_{k=0}^n \frac{1}{k!} \frac{1}{(n - k)!} e^{-(\lambda_1 + \lambda_2)} \lambda_1^k \lambda_2^{n - k} \end{aligned}$$

Recall from the Binomial theorem

$$\sum_{k=0}^n \binom{n}{k} \lambda_1^k \lambda_2^{n - k} = (\lambda_1 + \lambda_2)^n$$

$$\begin{aligned} \text{Then } P_z(n) &= \frac{1}{n!} e^{-(\lambda_1 + \lambda_2)} \sum_{k=0}^n \binom{n}{k} \lambda_1^k \lambda_2^{n - k} \\ &= \frac{(\lambda_1 + \lambda_2)^n}{n!} e^{-(\lambda_1 + \lambda_2)} \text{ for } n \geq 0 \end{aligned}$$

Then the number of arrivals in the union of the intervals is Poisson distributed with parameter $\lambda_1 + \lambda_2$

- c. $P[1 \text{ arrival from } A_1 \text{ prior to } t | 1 \text{ occurred}]$

$$\begin{aligned} &= \frac{p[1 \text{ arrival from } A_1 \text{ prior to } t, 0 \text{ from } A_2]}{P[1 \text{ arrival occurred}]} \\ &= \frac{\lambda_1 t e^{-\lambda_1 t} e^{-\lambda_2 t}}{\lambda t e^{-\lambda t}} = \frac{\lambda_1}{\lambda} \end{aligned}$$

d. For all $s \in [t_1, t_2]$

$$\begin{aligned}
 &= \frac{P[1 \text{ arrival occurred in } [t_1, s], 0 \text{ arrivals occurred in } [s, t_2]]}{P[1 \text{ arrival occurred}]} \\
 &= \frac{\lambda(s - t_1)e^{-\lambda(s - t_1)}e^{-\lambda(t_2 - s)}}{\lambda(t_2 - t_1)e^{-\lambda(t_2 - t_1)}} = \frac{s - t_1}{t_2 - t_1}
 \end{aligned}$$

Hence conditional on the knowledge that only one arrival occurred, the time of this arrival is uniformly distributed in $[t_1, t_2]$.