

The determinants of the differential exposure to COVID-19 in New York City and their evolution over time*

Milena Almagro[†] Angelo Orane-Hutchinson[‡]

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Abstract

In this paper, we explore different channels to explain the disparities in COVID-19 incidence across New York City neighborhoods. To do so, we estimate several regression models to assess the statistical relevance of different variables such as neighborhood characteristics and occupations. Our results suggest occupations are crucial for explaining the observed patterns, with those with a high degree of human interaction being more likely to be exposed to the virus. Moreover, after controlling for occupations, commuting patterns no longer play a significant role. The relevance of occupations is robust to the inclusion of demographics, with some of them, such as income or the share of Asians, having no statistical significance. On the other hand, racial disparities still persist for Blacks and Hispanics compared to Whites, although their magnitudes are economically small. In our temporal analysis, although the coefficients of many occupations and demographics decrease over time, we find results consistent with higher intra-household contagion, in line with the progression of the pandemic as well as with the health policies that have been set in place.

* Any errors or omissions are our own.

[†]Department of Economics, New York University. Email: m.almagro@nyu.edu

[‡]Department of Economics, New York University. Email: aoh227@nyu.edu

1 Introduction

Extensive evidence shows large differences in the incidence of COVID-19 across demographic groups. For example, in the US, Blacks are being hit the hardest by the pandemic, which creates an additional source of worry for this already vulnerable socio-economic group.¹ In the case of New York City (NYC), on April 6, the Department of Health and Mental Hygiene (DOH) released a report with race-specific data revealing that Blacks and Hispanics have a 47% and 36% higher crude rate, number of deaths per 100,000 population, respectively, compared to Whites.² These patterns repeat themselves in other areas of the US; however, the disparities across racial groups are less pronounced than in NYC.³

The fact that more vulnerable groups and minorities are being hit harder by the pandemic garnered the attention of many economists and policy makers. For example, Borjas (2020) and Schmitt-Grohé et al. (2020) show the incidence disparities of testing and positive rates across NYC neighborhoods is explained by a selection on demographics. Given that COVID-19 does not intrinsically discriminate across demographic groups, the reason for such disparities still remains an open question. Hence, our goal is to assess the importance of a set of observable channels, such as population density, commuting patterns, and occupations, that explain the existing differences across groups.

To understand the relevance of different mechanisms, we use data on the number of tests and positives across NYC zip codes provided by DOH.⁴ We combine these data with demographic data provided by the American Community Survey (ACS), also at the zip code level.

Because we focus on pointing out observable channels that are likely to explain why some demographic groups have a higher incidence of COVID-19 exposure, we estimate different specifications highlighting their importance at each stage. We start by including a small set of neighborhood controls, such as commuting patterns, population density, and health controls. In all of our specifications, we also include the share of the population being tested, or equivalently, tests per capita. The inclusion of this variable is motivated by the limited availability of tests in NYC that have been mostly administered to those showing symptoms.⁵ Therefore, we can interpret this variable as a proxy for the rate of infection *within* the neighborhood and thus the estimates of the other variables should be interpreted related to this rate. We find that when the number of tests per capita increases, the share of positives also increases, a result which is consistent with our interpretation. Not surprisingly, we find that

¹For a recent article see <https://nyti.ms/2UUI5je>

²For the latest data release visit: www1.nyc.gov/site/doh/covid/covid-19-data.page

³For a summary see: <https://nyti.ms/2wu9jCK>

⁴Unfortunately, at the time of this analysis, there is no data available with the number of deaths by zip code.

⁵As a matter of fact, at earlier dates tests were only performed only on those who required hospitalization.

as tests become more available to a larger share of the population, the magnitude of tests per capita decreases over time.

We continue by analyzing the role of occupations, motivated by the fact that they vary in their degree of human interaction, with those with a high level of human contact being more likely to be exposed to the virus, and vice versa.⁶ We do so by including the share of workers in each zip code for 13 categories constructed from ACS according to their degree of human interaction showing that, indeed, occupations are a key component in explaining the observed differences across NYC areas. For example, in our preferred specification including demographics and borough fixed effects, we find that a one-percentage-point increase in the number of workers employed in transportation, an occupation that has been declared essential and with a high exposure to human interaction, increases the share of positives by 2% and 1% percent for April 1 and April 20, respectively. Moreover, we show that after controlling for occupations, length of commute and the use of public transport are not significant.⁷

Additionally, these results are robust to the inclusion of demographics as well as borough fixed effects.⁸ Including demographics leads to several striking patterns. Whereas simple correlations show that wealthier neighborhoods have a lower rate of positives, we show that income is not significant when occupations are included. However, we still see positive significant effects on positive rates for minorities. These results could be due to minorities being less likely to get tested or having to be in worse conditions compared to Whites.⁹ However, it is arguable that this racial disparities are economically important, even becoming less relevant over time as more testing becomes available, with Asians showing no statistical significance at the end of our sample. For example, for April 1 and April 20, we find that a one-percentage point increase in the population of Blacks correlates with an increase of 0.34% and 0.15% in the share of positives for an average number of 51% and 54% of positive cases, respectively. For Hispanics, the disparity is larger, where a one-percentage-point increase in their population corresponds to an increase of 0.38% increase and 0.23% in the rate of positives, respectively. On the other hand, the estimate corresponding to average household size is positive throughout our sample and becomes larger over time.

Our daily analysis also reveals that as the stay-at-home orders starts to be effective, the magnitude of multiple occupations decrease as days go by. However, we still find a rather stable coefficient of household size over time, which is consistent with the stay-at-home order being more helpful at mitigating contagion at work or

⁶Michaels et al. (2019) show that interactive occupations have become more important in larger metros over time. A recent paper by Barbieri et al. (2020) shows evidence of this mechanism for workers in Italy.

⁷Harries (2020) argues that the NYC subway was crucial for spreading the pandemic in NYC. More recently, Furth (2020) shows that “local infections are negatively correlated with subway use.”

⁸We use similar controls to those in Borjas (2020) for comparability purposes.

⁹Some evidence that this is plausible mechanism can be found in www.modernhealthcare.com/safety-quality/long-standing-racial-and-income-disparities-seen-creeping-covid-19-care

in public spaces than within the household.

We conclude that much of the disparity in the rates of positives can be explained by different demographic groups being more or less representative across different occupations. In particular, a key channel is the differences in exposure to human contact across jobs. An immediate implication for policy is to target first these more sensitive groups regarding the distribution of protective gear, testing, and vaccination should target first, not only considering their risk of exposure but also the potential spillovers that it may have on the rest of the population.

2 Data description and patterns

Our source of incidence rates of COVID-19 and the number of tests performed is the NYC DOH data release. The DOH releases (almost) daily data on the cumulative count of COVID-19 cases and the total number of residents that have been tested, divided by the zip code of residence. We have collected data from April 1 to April 24, with only April 2 and April 6 missing from our sample.¹⁰

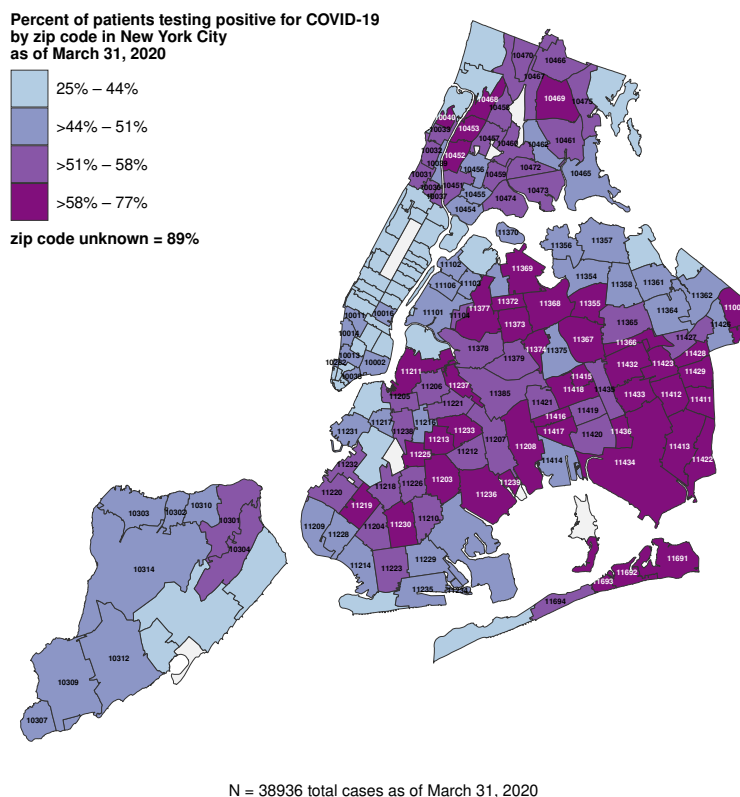


Figure 1 gives an example of differences in the rates of positives across NYC zip codes. From simple inspection, zip codes with the highest rates belong to the boroughs of Bronx, Brooklyn, and Queens. These boroughs are also home to the majority of Blacks and Hispanics living in NYC.¹¹ We obtain demographic and occupation data at the zip code level from the American Community Survey (ACS). The demographic characteristics we include are zip code median income, average age, racial breakdown, and health insurance status. We also include commuting-related variables: average commute time to work as well as means of transportation. Moreover, the zip codes with the lowest income levels are also located in those three boroughs. We plot a simple correlation between the share of positives and demographics. We see that shares of Blacks and Hispanics are positively correlated with rate of positives, whereas Figure 2 shows a flat relationship for the share of Asians. We observe a negative relationship for income, a slightly positive one for the share of males, and a positive and significant one for household size, as shown in Figure 3.

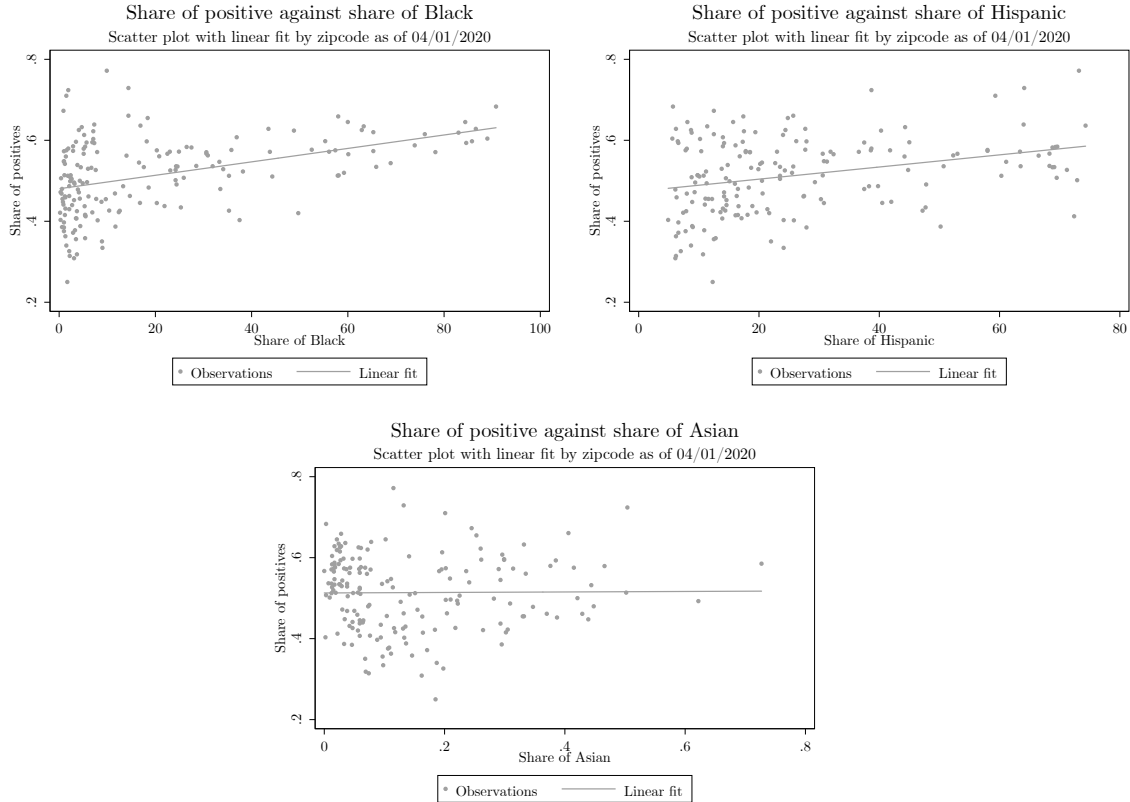


Figure 2: Share of positives against income, share of males, and household size by zip code

¹¹These groups compose 29 % and 56 %, respectively, of all Bronx residents, 31 % and 19 % for Brooklyn, and 17 % and 28 % in the case of Queens.

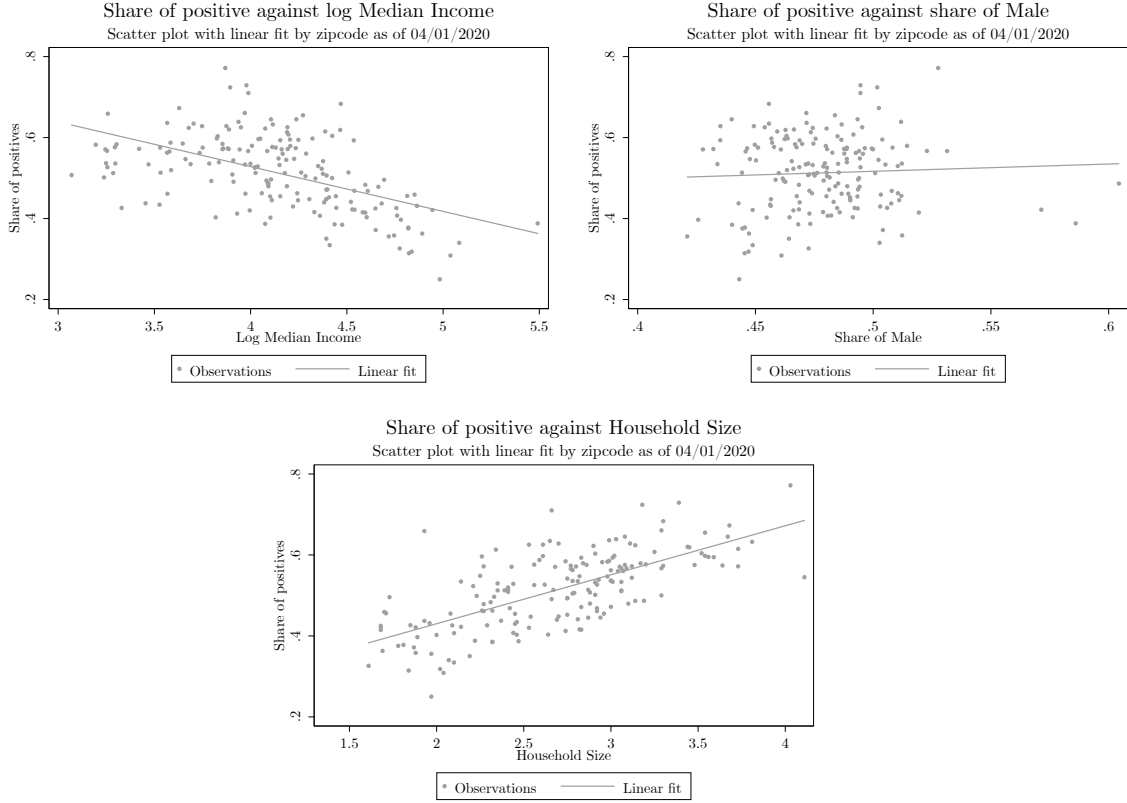


Figure 3: Share of positives against income, share of males, and household size by zip code

We also construct the shares of the working-age population employed at different occupations. We then categorize them according to the groups listed in column 2 of Table 1, and the ACS provides the number of workers in each one occupation, by zip code of residence. Taking into account the share of essential occupations, we divide the occupations into 13 categories for our empirical exercise, according to their essential definition, correlations between them, and similarity in work environments and social exposure. Table 1 shows the occupation groups that we use in the estimating regressions.¹² Summary statistics for all variables included in our empirical analysis can be found in Table 2

¹²For robustness we perform our analysis with two alternative classifications for occupations. First, we divided occupations between essential and non-essential as declared by the US government. Second, we used the four categories defined in Kaplan et al. (2020). In both cases, the high level of aggregation lead to non-significant estimates or results that were hard to reconcile with observational evidence.

Table 1: Occupation categories

Category	ACS Occupations
(1) Essential - Professional	Management, Business, Finance
(2) Non essential - Professional	Computer and Mathematical, Architecture and Engineering, Sales and Related, Community and Social Services, Education, Training, and Library, Arts, Design, Entertainment, Sports, and Media Administrative and Office Support
(3) Science fields	Life, Physical, and Social Science
(4) Law and related	Legal
(5) Health practitioners	Health practitioners
(6) Other health	Health technologists, technicians, and Healthcare Support
(7) Firefighting	Firefighting and prevention
(8) Law enforcement	Law enforcement
(9) Essential - Service	Food Preparation and Serving, Building and Grounds Cleaning and Maintenance
(10) Non essential - Service	Personal Care and Service
(11) Industrial, Natural resources and Construction	Construction and Extraction, Material Moving, Farming, Fishing, and Forestry, Production
(12) Essential - Technical	Installation, Maintenance, and Repair
(13) Transportation	Transportation

Finally, Table 2 presents the summary statistics of all the variables that are used in our analysis.

Table 2: Summary statistics

Variable	Mean	Std. Dev.	p10	Median	p90
Share of positives	0.563	0.085	0.438	0.583	0.645
Tests per Capita	0.018	0.006	0.012	0.017	0.026
Median Income (in 000's)	68.604	31.878	34.122	62.202	115.084
Share ≥ 20 , ≤ 40	0.323	0.084	0.246	0.308	0.433
Share ≥ 40 , ≤ 60	0.258	0.033	0.220	0.261	0.296
Share ≥ 60	0.200	0.079	0.132	0.190	0.276
Share Male	0.477	0.029	0.446	0.479	0.508
Household Size	2.683	0.537	1.930	2.750	3.300
% Black	0.200	0.240	0.010	0.076	0.600
% Hispanic	0.263	0.195	0.078	0.189	0.634
% Asian	0.144	0.139	0.017	0.094	0.335
Density (in 000's)	43.380	31.045	10.784	36.639	90.075
% Public Transport	0.532	0.150	0.312	0.543	0.712
Commuting Time (in mins)	40.647	7.054	27.200	42.100	48.100
% Uninsured	0.089	0.043	0.042	0.084	0.143
% Essential - Professional	0.126	0.089	0.046	0.092	0.285
% Non ess. - Professional	0.279	0.075	0.195	0.271	0.359
% Science fields	0.006	0.007	0.001	0.004	0.015
% Law and related	0.018	0.026	0.003	0.008	0.049
% Health practitioners	0.029	0.018	0.009	0.026	0.050
% Other health	0.038	0.024	0.010	0.035	0.073
% Firefighting	0.012	0.009	0.003	0.012	0.023
% Law enforcement	0.007	0.007	0.001	0.006	0.014
% Essential - Service	0.065	0.033	0.035	0.060	0.107
% Non ess. - Service	0.032	0.013	0.016	0.032	0.047
% Ind. and Construction	0.054	0.027	0.014	0.056	0.090
% Essential - Technical	0.014	0.009	0.004	0.013	0.022
% Transportation	0.029	0.016	0.004	0.032	0.048

3 Results

3.1 General Results

In this section, we present the main empirical results, for our four different specifications. Our unit of analysis is the zip code, and all models include the share of positive tests as the dependent variable. Additionally, we include tests per capita as a proxy for the rate of infection within the neighborhoods. The first model includes some widely discussed potential factors of the spread of COVID-19 in NYC: density and commuting patterns, specifically, log of population density, percentage of workers using public transport, and average commute time. We also include the percentage of uninsured to control for those who do not have access to healthcare insurance. Our second model expands by including our proposed mechanism, namely, the percentage of the working-age population employed in each of the 13 occupation categories defined in Table 1. The third specification adds demographic controls related to income, age, gender, household size, and race. Finally, we include borough fixed effects in our last model. Exploiting the fact that we have daily data over multiple days, we estimate a separate regression for each of them, allowing us to detect any time variation in the correlations. Therefore, in all of our specifications we run the following regression equations

$$\text{share of positives}_{it} = \alpha_t + \beta_t \text{tests per capita} + \gamma_t X_i + \epsilon_{it},$$

where the set of controls X_{it} vary according to the description below.

The first model shows the effect of the variables commonly used to explain the incidence of COVID-19 in NYC. Whereas Harries (2020) finds subway use was a major factor of the virus spread, we find it does not have a significant effect. This result could be due to the lack of cross-neighborhood variation to identify this effect, because most New Yorkers use public transportation in their daily commute. Nonetheless, commute time is a significant factor. For example, for April 1 a four-minute increase in commute time, a 10% increase on average, correlates with a 2.3% increase in the share of positives. We also find a positive and significant effect of the share of the uninsured population on the rate of positives. This result can be explained by patient selection, with uninsured patients being willing to be tested only under very acute symptoms, because they fear of medical charges. Although its magnitude decreases as we include other covariates, the estimate of this variable remains positive and significant with a one-percentage-point increase in the share of uninsured population being correlated with a 0.65 percentage-point.

Specification (2) shows the effect of different occupations in the share of positives. The variables are defined as the shares of the working-age population employed in these occupations, so the coefficients are relative to the working-age but not employed population. The coefficients can be read as the effect of a one-percentage-point increase in the population employed in the particular category on the share of positives.

We find some occupations can explain a significant part of the variation in COVID-19 incidence. On the one hand, an increase in the share of workers employed in non-essential - professional, other health (not health practitioners), and transportation occupations are all associated with a higher percentage of positive tests. On the other hand, higher shares of workers in the science fields category, legal occupations, and law enforcement have a negative correlation with the share of positives. These results are discussed further in the time-trends section.

Perhaps surprisingly, under this specification, commute time no longer has a significant effect. This result suggests commuting patterns are closely related to occupations, and most of the explanatory variation for commuting patterns comes through this channel. This result also implies the existence of within-city location and mobility patterns that are occupation specific.

We include demographic variables in the third model. Despite the strong correlation between the share of positives and demographic characteristics, the results for specification (3) show some of it can be explained through the occupation mechanism. Notably, the income effect disappears when we control for occupations, suggesting the previous correlation is due to income differences across jobs. Still, some demographic effects remain significant. For April 20, a one-percentage-point increase in the share of Blacks and Hispanics leads to a 0.15% and 0.23% increase in the rate of positives, an effect that is economically small. A plausible explanation for these patterns could be driven by a racial bias on the incidence of testing, as pointed out by Borjas (2020). Another explanation is differences in adherence to the shelter-in-place policy, as explored by Coven and Gupta (2020). We also find household size has positive correlation with test outcomes. Adding one extra person to the average household, a 37% increase, corresponds with a 7% increase in the percentage of positive tests. Although neighborhood density does not explain variation in the share of positives, density in households appears to do so, with increasing magnitude over time.

The tests-per-capita coefficient is positive and highly significant across almost all specifications. Because of the scarcity of tests, testing was only available at earlier dates to those who required hospitalizations, and later on to those who were already showing symptoms. Therefore, we interpret this variable as a proxy for the rate of infections within the neighborhood, as the fraction of patients showing symptoms or requiring hospitalization is rather fixed for COVID-19.¹³ Its magnitude decreases over time as testing becomes more available and accessible to the rest of the population. We also find that its estimate remains significant even after including borough fixed effects.

¹³A potential concern is large differences in age distribution across NYC zip codes. In the data, we find that the average age in NYC ranges from 27.5 to 45.5 across neighborhoods in NYC, with the exception of zip code 11005. It is a fairly small zip code with 1700 residents, an average age of 76, and mainly composed of retired immigrant women. Given such differences, we have excluded it from our analysis.

Table 3: Dependent variable - Share of positives as of April 1, 2020

	(1)		(2)		(3)		(4)	
	Nbhd Controls		+ Occupations		+ Demographics		+ Borough FE	
Tests-per-capita	9.017***	(2.879)	11.186***	(2.447)	10.773***	(2.249)	12.050***	(2.386)
Log Density	0.015	(0.014)	0.022*	(0.012)	0.015	(0.012)	0.032***	(0.011)
% Public Transport	-0.015	(0.072)	0.013	(0.068)	0.053	(0.070)	-0.059	(0.062)
Log Commuting Time	0.237***	(0.046)	-0.016	(0.083)	0.009	(0.075)	-0.054	(0.062)
% Uninsured	1.002***	(0.141)	0.662***	(0.246)	0.336	(0.215)	0.150	(0.180)
% Essential - Professional			0.156	(0.271)	0.695***	(0.238)	0.766***	(0.236)
% Non ess. - Professional			0.669***	(0.189)	0.615***	(0.181)	0.544**	(0.216)
% Science fields			-4.703***	(1.294)	-3.745***	(1.064)	-2.965***	(1.118)
% Law and related			-0.410	(0.801)	-0.875	(0.754)	-1.427**	(0.697)
% Health practitioners			-0.432	(0.421)	-0.167	(0.431)	-0.167	(0.386)
% Other health			0.947***	(0.321)	0.027	(0.412)	0.346	(0.402)
% Firefighting			2.743**	(1.072)	1.624	(1.109)	1.629*	(0.965)
% Law enforcement			-0.301	(1.215)	0.815	(1.089)	-0.223	(1.016)
% Essential - Service			-0.100	(0.354)	0.258	(0.347)	0.245	(0.300)
% Non ess. - Service			0.769	(0.561)	1.166**	(0.509)	1.154**	(0.483)
% Ind. and Construction			1.091**	(0.437)	1.208***	(0.402)	0.839**	(0.401)
% Essential - Technical			-2.025*	(1.133)	-0.457	(0.979)	-0.319	(0.881)
% Transportation			1.752***	(0.588)	1.718***	(0.527)	1.102**	(0.469)
Log Income					-0.008	(0.034)	-0.010	(0.033)
Share $\geq 20, \leq 40$					-0.346*	(0.176)	-0.357**	(0.173)
Share $\geq 40, \leq 60$					-0.855***	(0.222)	-0.611**	(0.237)
Share ≥ 60					-0.380**	(0.175)	-0.347*	(0.197)
Share Male					-0.050	(0.267)	-0.146	(0.264)
Log Household Size					0.076	(0.073)	0.037	(0.061)
% Black					0.149***	(0.039)	0.175***	(0.040)
% Hispanic					0.003	(0.050)	0.194***	(0.050)
% Asian					0.136**	(0.053)	0.141***	(0.050)
Bronx							-0.014	(0.023)
Brooklyn							0.086***	(0.022)
Queens							0.084***	(0.024)
Staten Island							0.083***	(0.027)
Constant	-0.671**	(0.264)	-0.149	(0.342)	0.110	(0.372)	0.196	(0.334)
Observations	174		174		174		174	
R^2	0.514		0.694		0.785		0.839	

Weighted OLS by population size. Robust standard errors in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, *** $p < 0.01$

Table 4: Dependent variable - Share of positives as of April 10, 2020

	(1)		(2)		(3)		(4)	
	Nbhd Controls		+ Occupations		+ Demographics		+ Borough FE	
Tests per Capita	1.913**	(0.832)	1.713*	(0.921)	1.795**	(0.763)	3.904***	(0.675)
Log Density	0.022*	(0.013)	0.018*	(0.010)	0.013	(0.008)	0.022***	(0.007)
% Public Transport	-0.001	(0.060)	0.012	(0.056)	0.095*	(0.056)	-0.004	(0.041)
Log Commuting Time	0.262***	(0.040)	0.019	(0.069)	0.022	(0.060)	-0.023	(0.045)
% Uninsured	1.038***	(0.104)	0.521***	(0.187)	0.316**	(0.136)	0.290***	(0.103)
% Essential - Professional			-0.003	(0.203)	0.579***	(0.177)	0.484***	(0.179)
% Non ess. - Professional			0.419**	(0.173)	0.354**	(0.166)	0.257*	(0.138)
% Science fields			-3.021***	(1.082)	-3.094***	(0.905)	-2.334***	(0.812)
% Law and related			-0.604	(0.525)	-1.050**	(0.480)	-1.293***	(0.422)
% Health practitioners			-0.248	(0.372)	-0.061	(0.379)	-0.124	(0.281)
% Other health			0.753***	(0.258)	-0.275	(0.302)	0.238	(0.231)
% Firefighting			1.282	(0.880)	-0.042	(0.869)	0.456	(0.570)
% Law enforcement			-1.859	(1.149)	-1.217	(0.892)	-1.323*	(0.751)
% Essential - Service			0.159	(0.262)	0.198	(0.280)	0.127	(0.199)
% Non ess. - Service			0.359	(0.471)	0.844**	(0.417)	0.781**	(0.350)
% Ind. and Construction			0.472	(0.332)	0.497*	(0.279)	0.101	(0.225)
% Essential - Technical			-0.729	(0.854)	-0.150	(0.718)	-0.474	(0.531)
% Transportation			1.824***	(0.419)	1.639***	(0.377)	0.831***	(0.299)
Log Income					-0.024	(0.023)	-0.027	(0.021)
Share $\geq 20, \leq 40$					-0.243**	(0.122)	-0.246**	(0.097)
Share $\geq 40, \leq 60$					-0.510***	(0.181)	-0.228	(0.149)
Share ≥ 60					0.127	(0.114)	-0.017	(0.115)
Share Male					0.453**	(0.180)	0.249	(0.171)
Log Household Size					0.167***	(0.054)	0.111***	(0.042)
% Black					0.165***	(0.030)	0.114***	(0.026)
% Hispanic					0.018	(0.041)	0.130***	(0.034)
% Asian					0.047	(0.043)	0.018	(0.030)
Bronx							-0.040***	(0.014)
Brooklyn							0.053***	(0.015)
Queens							0.058***	(0.016)
Staten Island							-0.022	(0.023)
Constant	-0.765***	(0.228)	0.010	(0.269)	-0.137	(0.258)	0.107	(0.213)
Observations	174		174		174		174	
R^2	0.674		0.801		0.871		0.921	

Weighted OLS by population size. Robust standard errors in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, *** $p < 0.01$

Table 5: Dependent variable - Share of positives as of April 20, 2020

	(1)		(2)		(3)		(4)	
	Nbhd Controls		+ Occupations		+ Demographics		+ Borough FE	
Tests-per-Capita	0.667	(0.485)	0.262	(0.560)	0.381	(0.497)	2.553***	(0.476)
Log Density	0.024**	(0.011)	0.015*	(0.009)	0.011	(0.008)	0.016***	(0.006)
% Public Transport	0.010	(0.055)	-0.001	(0.050)	0.080	(0.053)	-0.017	(0.040)
Log Commuting Time	0.232***	(0.034)	0.001	(0.060)	0.004	(0.055)	-0.008	(0.042)
% Uninsured	0.924***	(0.099)	0.417**	(0.171)	0.296**	(0.129)	0.351***	(0.098)
% Essential - Professional			-0.210	(0.168)	0.294*	(0.165)	0.235	(0.160)
% Non ess. - Professional			0.329**	(0.147)	0.274*	(0.152)	0.224*	(0.122)
% Science fields			-1.931*	(1.016)	-2.318***	(0.861)	-1.609**	(0.784)
% Law and related			-0.492	(0.456)	-0.851*	(0.460)	-0.898**	(0.397)
% Health practitioners			-0.155	(0.357)	0.010	(0.387)	-0.206	(0.278)
% Other health			0.815***	(0.232)	-0.053	(0.272)	0.365	(0.221)
% Firefighting			0.379	(0.829)	-0.765	(0.876)	-0.156	(0.556)
% Law enforcement			-1.970*	(1.049)	-1.472*	(0.820)	-1.344**	(0.655)
% Essential - Service			0.312	(0.229)	0.205	(0.242)	0.082	(0.171)
% Non ess. - Service			-0.046	(0.437)	0.455	(0.378)	0.578*	(0.296)
% Ind. and Construction			0.271	(0.317)	0.246	(0.271)	-0.079	(0.209)
% Essential - Technical			-0.785	(0.724)	-0.603	(0.617)	-0.908*	(0.487)
% Transportation			1.253***	(0.364)	1.083***	(0.327)	0.541*	(0.293)
Log Income					-0.021	(0.022)	-0.022	(0.019)
Share $\geq 20, \leq 40$					-0.169	(0.115)	-0.208**	(0.090)
Share $\geq 40, \leq 60$					-0.389**	(0.161)	-0.198	(0.126)
Share ≥ 60					0.248**	(0.108)	0.002	(0.104)
Share Male					0.540***	(0.166)	0.318**	(0.150)
Log Household Size					0.167***	(0.047)	0.099***	(0.036)
% Black					0.140***	(0.030)	0.081***	(0.026)
% Hispanic					0.027	(0.036)	0.125***	(0.033)
% Asian					0.015	(0.043)	0.012	(0.031)
Bronx							-0.062***	(0.014)
Brooklyn							0.034**	(0.014)
Queens							0.023	(0.015)
Staten Island							-0.064***	(0.022)
Constant	-0.682***	(0.195)	0.201	(0.232)	-0.076	(0.236)	0.111	(0.196)
Observations	174		174		174		174	
R^2	0.673		0.800		0.866		0.920	

Weighted OLS by population size. Robust standard errors in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, *** $p < 0.01$

3.2 Daily comparison and time trends

We find some time-variant results that could provide insight on both the evolution of the pandemic effects as well as the health policies in place. Figures 4 to 6 show the time evolution of the coefficients for specification (4). The result for the tests-per-capita variable is particularly salient; we observe a strong correlation on the share of positive tests that becomes progressively smaller over time. This result could be reconciled with the fact that in the earlier days of the crisis, testing was severely limited. Zip codes with more tests implied a higher share of people at high risk of having the disease. So, a key takeaway from the results of our daily comparison is the importance of widespread testing, because it allows us to identify the mechanisms that explain demographic and occupational differences in COVID-19 exposure.

Notable time trends exist in the correlations associated with occupations. Higher shares of professional and non-essential - service categories were associated with higher percentage-point increases in the rate of positive tests in the early day. An additional percentage point in either occupation group implied an almost 0.6 percentage point increase in the positive rate. However, they eventually become not significant or trend toward zero, averaging closer to a 0.3-percentage-point increase. A plausible explanation is that these professions are either non-essential, or have the highest shares of remote workers. Although they were highly exposed to the virus in the beginning, once the workers shelter in place, their correlation with positive tests subsides. The opposite happens in science fields and law occupations — they are negatively correlated with COVID-19 incidence in the beginning, but the effect trends towards zero.

We find interesting patterns for the essential occupations as well. An additional percentage point in the share of transportation workers is associated with between a 0.5- and a one-percentage-point increase in the positives rate. The effect seems to decay over time, but at a slower rate than other occupations. This result could be due to its essential designation, but also due to its relatively high-exposure nature. The share of industrial, natural-resources, and construction occupations starts off with a positive correlation with COVID-19 incidence. However, a week after the general stay-at-home order, the governor determined construction was not essential, and this order could explain the eventual attenuation of the correlation. Law-enforcement-occupation shares have a consistently negative correlation on the share of positive, whereas firefighter shares have a declining trajectory toward zero. A plausible explanation for this difference could be the partnership between the NYPD and health care groups to provide free testing to its members.¹⁴

The share of the uninsured population increasingly predicts the variation in positive test results. We find that an additional percentage point in the share of uninsured predicts an almost 0.3-percentage-point increase in the share of positives. Although many health care providers are waiving COVID-19-related out-of-pocket costs, these

¹⁴www.nypost.com/2020/04/03/nypd-partners-with-health-care-groups-to-test-cops-for-covid-19/

fees remain very high for the uninsured, and so a higher incidence of COVID-19 in this group could imply a severe financial burden. Although still significant, the effect of neighborhood density declines over time, and the opposite occurs for household size. The stay-at-home order could mitigate part of the risk of high neighborhood density, while increasing the probability of within-household infections.

Finally, another outstanding time pattern is that the coefficients on racial composition decrease in magnitude as the selection of testing decreases. This result may suggest a stronger racial-selection component is at play among those in worse conditions at earlier dates. For example, an explanation for this pattern could be that Black citizens were less likely to be tested or had to be in worse conditions to access testing compared to White citizens.¹⁵

¹⁵Some evidence that this is plausible mechanism can be found in www.modernhealthcare.com/safety-quality/long-standing-racial-and-income-disparities-seen-creeping-covid-19-care

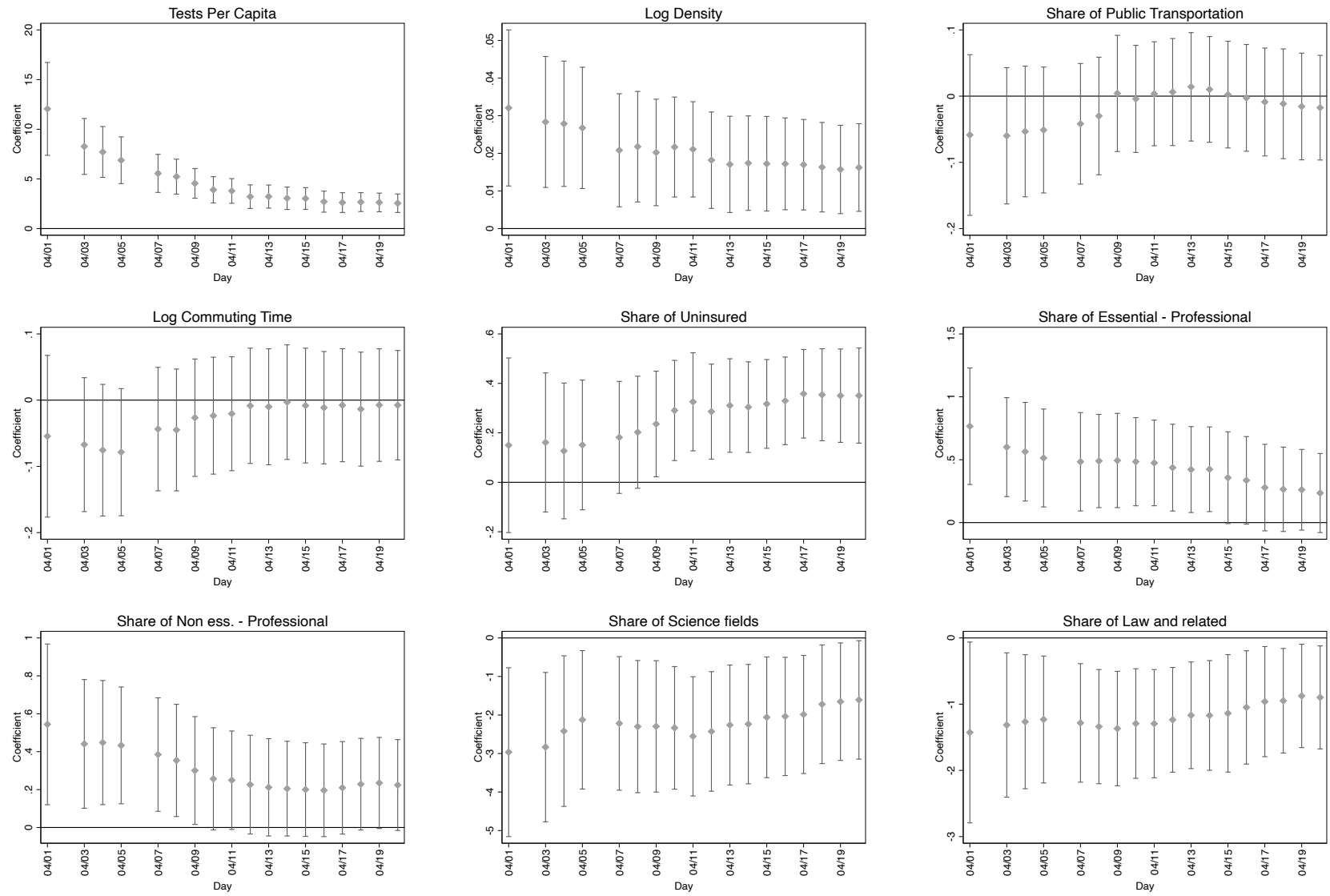


Figure 4: Regression coefficients of specification (4) over time

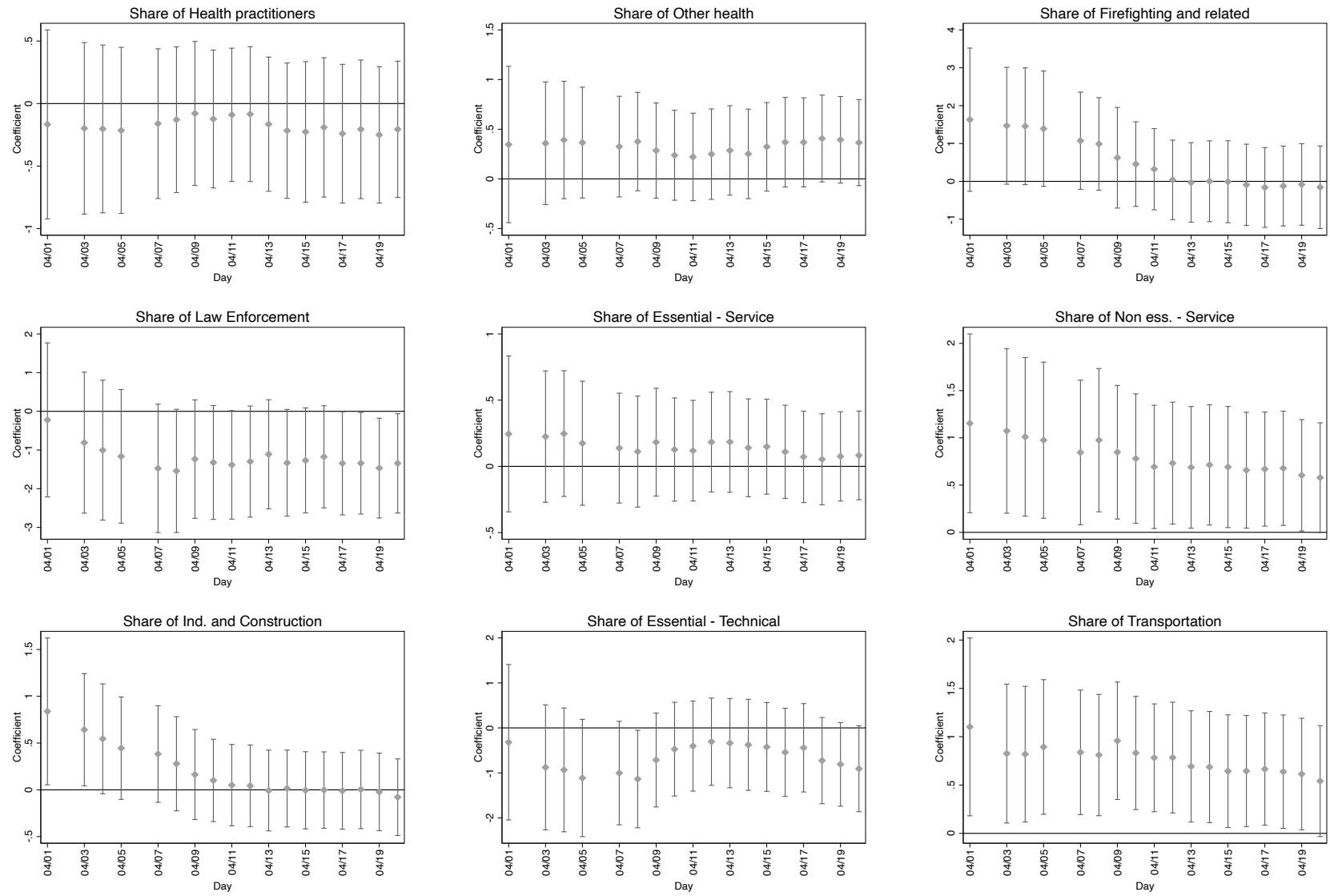


Figure 5: Regression coefficients of specification (4) over time

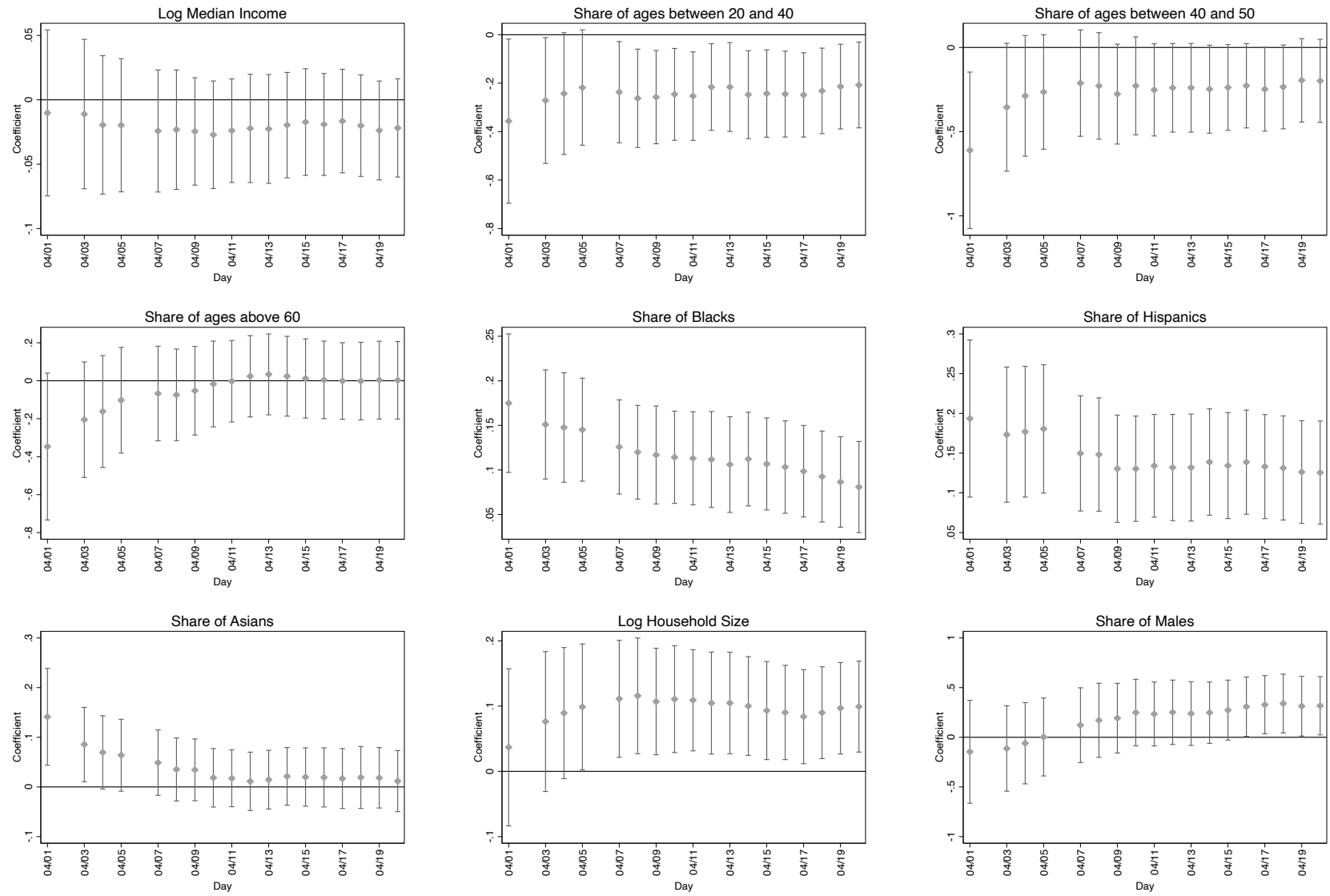


Figure 6: Regression coefficients of specification (4) over time

4 Conclusions and policy implications

In this paper, we present evidence showing that occupations are an important channel for explaining the difference in the rates of COVID-19 across neighborhoods. Using data from NYC at the zip code level, we study the relationship between the share of positive tests and the share of workers in different occupations. The DOH provides daily updates of COVID-19 test data, allowing us to study the aforementioned relationships over multiple days and to detect time variation in their magnitudes.

We begin by showing descriptive evidence of the heterogeneous incidence of positive cases across neighborhoods, income, race, gender, and household size. A zip code’s median income is negatively correlated with its share of positives. Conversely, we find a positive correlation between shares of Black and Hispanic residents as well as average household size and the share of positive tests. Highlighting these differences is important because this observations confirms that the disease has had more harmful effects on vulnerable communities, and finding an occupation mechanism that explains it could guide policy intended to alleviate its impact.

We estimate several models to explore the effects of occupations. Our first specification only includes neighborhood characteristics, such as the use of public transportation and the average length of daily commutes. Although commuting patterns have been put forth as a major factor in the spread of the disease in NYC, we show that after including occupation controls, they fail to significantly explain variation in share of positives.

We then augment the model by including the occupation shares. We find the strongest positive correlation on the share of positive tests with the share of workers in Transportation, Industrial, Natural-resources, Construction, and Non essential - Professional, with clear time trends in their estimated coefficients. For example, in the case of Transportation, a one-percentage-point increase in the share of workers in these occupations leads to a one- to two-percentage-point increase in thee rates of positive results. Although the other two have a significant effect in positive shares in the first days, the magnitude becomes insignificant by the end of the sample period. This trend could be a result of the stay-at-home order. Conversely, higher shares of workers in Science Fields and Law Enforcement reduce the number of positive rates, with Science Fields decreasing in magnitude over time.

When adding demographic controls, we observe that racial patterns do persist, suggesting that the occupation mechanism does not fully explain all of the racial differences. However, their magnitude is small and arguably not economically relevant when compared to the statistical significance of the occupation magnitudes. Income and most age groups do not contribute to explaining the variation in positive tests, suggesting the occupation mechanism can explain to a greater extent the disparities along those demographics observed in the data.

In all of our regression models we include the number of tests per capita, and find that it is a strong predictor of the share of positive tests. However, its relative

importance declines over time, as tests become more widely available and therefore the number of tests per capita starts including other residents beyond those requiring hospitalization. Moreover, as this variable loses relevancy, more of the variation in COVID-19 incidence is explained through the occupation channel.

Our results suggest clear implications for policy all over the world, but in particular for NYC. First, as mentioned earlier, our results highlight the need of mass testing to enable the clean identification of the relevant channels that generate a greater risk of exposure. Second, once these channels are identified, policy-makers can target specific groups in the provision of protective gear, tests, and vaccinations. The purpose of this policy is twofold. On the one hand, for those who are more vulnerable it provides a form of insurance against contracting the disease. On the other hand, targeting specific workers or demographic groups may have spillovers on the rest of the population. For example, a policy that starts vaccinating and/or testing those workers with higher rates of human interactions affects not only those directly targeted by the policy, but also those who are likely to be in contact with them. We believe, given the high contagion rates of COVID-19, these type of spillovers from this policy may be substantial. Moreover, given that access to insurance plays a significant role that becomes more important over time, another policy recommendation is to incentivize testing for those without medical insurance such as complete coverage of out-of-pocket costs in relation to COVID-19, including testing. In this way, the authorities would be able to have a clearer image of potential focal points of contagion. Finally, our results also suggest the stay-at-home order could have mitigated contagion rates at work or in public spaces, while increasing the probability of intra-household infections. This contagion within a single home could be alleviated by a policy offering shelter to those families with a large number of members or with a limited shared space.

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