

# Optimal Urban Transportation Policy: Evidence from Chicago<sup>†</sup>

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## Abstract

We characterize optimal urban transportation policies in the presence of congestion and environmental externalities and evaluate their welfare and distributional effects. We present a framework of a municipal government that implements different transportation equilibria through its choice of public transit policies—prices and frequencies—as well as road pricing. The government faces a budget constraint that introduces monopoly-like distortions. We apply this framework to Chicago, for which we construct a new dataset that comprehensively captures transportation choices. We find that road pricing alone leads to large welfare gains by reducing externalities, but at the expense of consumers (travelers), whose surplus falls even if road pricing revenues are fully rebated. The largest losses are borne by middle income consumers, who are most reliant on cars. We find that the optimal price of public transit is close to zero and goes along with a reduction in the frequency of buses and an increase in the frequency of trains. Combining these transit policies with road pricing eliminates budget constraints. This allows the government to implement higher transit frequencies and even lower prices, in which case consumer surplus increases after rebates.

**JEL classification:** L91, L5, L13, H23, R41, R48.

**Keywords:** Urban transportation, public transit subsidy design, road pricing, spatial equilibrium, Ramsey pricing.

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# 1 Introduction

Since the 1950s, urban transportation in the U.S. has been characterized by the overwhelming use of personal cars, while public transit accounts for only 3.4% of the 850 million city trips that Americans undertake every day.<sup>1</sup> This heavy reliance on cars poses significant challenges for city governments: the economic costs of road congestion in the U.S. alone are estimated at \$87 billion per year, and car usage causes significant environmental impacts through emissions of carbon and other pollutants.<sup>2,3</sup>

Cities' efforts to mitigate their effect on the climate and reduce inequality have led to a renewed discussion about the right mix of urban transportation policies.<sup>4</sup> Some argue that public transit should be cheaper, and indeed several municipalities have recently introduced free public transit.<sup>5</sup> Others suggest that instead of reducing fares, cities should provide more frequent, higher-quality public transit.<sup>6</sup> Although these proposals appeal to different types of riders, it may not be feasible to pursue both because of stressed municipal budgets. By contrast, some cities have recently started taxing private forms of transportation. For example, London enacted a £15 congestion surcharge during the day-time and New York City recently approved a cordon tax below 60th street in Manhattan.<sup>7</sup> These forms of road pricing not only steer consumers away from private transportation towards public alternatives but also help financially strained municipalities collect tax revenues. Importantly, all of the above policy levers cannot be evaluated in isolation — they are linked through mode substitution on the demand side, technologically through road congestion, and fiscally due to municipal budget constraints.

Given these interactions, what is the right mix of urban transportation policies? Should cities aim to increase the use of public transit, discourage the private use of roads, or some combination of the two? Are these two policy measures complements or substitutes in achieving cities' objectives? In this paper, we characterize the optimal mix of policies for a budget-constrained municipal government. Given prevailing difficulties to build new

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<sup>1</sup> This is despite generous subsidies for public transit, which account, on average, for 3/4 of marginal trip costs. See <https://www.bts.dot.gov>. and [newgeography.com](http://newgeography.com)

<sup>2</sup> See [World Economic Forum—US Traffic Congestion Cost in 2018](#).

<sup>3</sup> A private car emits 0.96 pounds of CO<sub>2</sub> per passenger-mile whereas public transit emits 0.45, even with low utilization rates.

<sup>4</sup> See [Brookings—U.S. Transportation policy](#) and [HKS—Free Public Transit](#)

<sup>5</sup> See [NYT—“Should Public Transit Be Free? More Cities Say, Why Not?”](#)

<sup>6</sup> See [The Conversation—Low-cost, high-quality public transportation.](#)

<sup>7</sup> For a comprehensive list of congestion pricing policies see <https://ops.fhwa.dot.gov>.

transit infrastructure in the US (Brooks and Liscow, 2021), we focus on three interventions that do not require new infrastructure: changing the fares of public transit, adjusting their service frequency, and road pricing.<sup>8</sup>

In order to measure the welfare effects of these interventions, we must account for several ingredients. First, we need to understand how people substitute across transportation modes as prices and travel times change. Second, we need to account for important forces that arise in transportation markets—traffic and environmental externalities as well as scale economies. Third, we have to account for the interaction of travel decisions across different locations, which are linked through roads that can become congested. To achieve this, we contribute to the literature both theoretically and empirically.

We first formulate a framework of a municipal government who wants to maximize welfare. The government chooses the prices and quality (in terms of frequencies) of different modes of transportation, subject to a budget constraint that accounts for operational costs and revenues from fares and road pricing. Given these government choices, travelers in the market adjust and reach an equilibrium. We first find that an unconstrained social planner would set price minus marginal cost equal to the marginal externality—a canonical result in the theory of optimal taxation. However, our social planner deviates from this solution because budget considerations introduce two monopoly-like distortions. First, the planner charges markups that downwards-distort quantities. Second, given the ability to choose the service frequency of public transit, the planner directly affects the quality that travelers experience. Budget considerations distort quality towards the marginal consumer, as in Spence (1975).

Next, we move to an empirical application of this framework to Chicago, home to one of the largest public transit system in the US.<sup>9</sup> Chicago serves as an ideal setting for our purposes for several reasons. First, public and private transportation are both important, with 69% of trips made by car and 23% by public transit.<sup>10</sup> Therefore, both changes in public transit provision and road pricing could be important policy levers. Second, Chicago is a city with large economic disparities, which means that it is important

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<sup>8</sup> We focus on short-run adjustments, keeping residents' and firms' locations as well as infrastructure fixed. We thus measure welfare gains in the short run, before firms and residents readjust.

<sup>9</sup> In Chicago, during 2019, the number of unlinked bus trips was 237,276,500 and the number of unlinked heavy rail trips was 218,467,000, which sum up to a total of 455,743,500 public transit trips. Source: [American Public Transportation Association \(APTA\), Public Transportation Ridership Report, Fourth Quarter 2019](#).

<sup>10</sup> The remaining trips are taxi and ride-hailing trips.

to account for the distributional effects of those policies.

Beyond these economic considerations, Chicago is also particularly good in terms of data availability. We combine several data sources to construct a rich dataset of travel flows, travel times, and prices for all relevant modes at a high temporal and spatial resolution. We have access to the near-universe of public transit through records from the CTA and the universe of ride-hailing trips, which are made public by the city of Chicago. One challenge we face is that there are no official records of private car trips. To overcome this problem, we determine mobility patterns from individual cellphone location records, which we use to construct the total number of trips in the city. We then recover the number of car trips by subtracting public transit and livery-vehicle (ride-hail and taxi) trips from the total number of trips.

Our data reveals that both mode choice and access to public transportation vary by travelers' income levels and location. Car ownership is less common for both high and low income travelers relative to those with median incomes. These patterns together motivate a mode choice (demand) model that can account for these sources of heterogeneity.

The richness of our data allows us to designate granular transportation markets—people traveling from one community area (CA) to another during one hour of the week—and still conduct our analysis with aggregate market shares (Berry *et al.*, 1995).<sup>11</sup> This approach has the advantage that we can use standard inversion techniques to address endogeneity concerns. While the cost of operating a car and public transit prices are invariant to demand shocks, both the travel times of road-based modes and the prices of ride-hail are endogenous.

We instrument endogenous travel times using free-flow traffic speeds and argue that those capture long-run cross-sectional differences in infrastructure that are independent of the demand shocks that would bias our estimated model coefficients. To address the endogeneity of ride-hailing prices, we use price variation introduced by a surcharge that is levied on ride-hailing trips that either start or end downtown between 6AM and 10PM.<sup>12</sup> We construct indirect-inference moments that capture the deviation between the model-predicted demand response and the (difference in difference) treatment effect of the surcharge policy. The resulting model allows for heterogeneous substitution patterns across locations, income, and car ownership. Our estimates reveal substantial hetero-

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<sup>11</sup> There are 77 community areas in the City of Chicago, with an average population of 35,600.

<sup>12</sup> The details of this policy are laid out [on the website of the City of Chicago](#).

geneity in the value of time across travelers and time of day, ranging from \$4 to \$45 per hour.<sup>13</sup>

Another key component of our model is a congestion technology that allows us to map traffic flows into speeds. To estimate the model, we exploit variation across hours of the day in travel speeds and in the number of vehicles traveling between adjacent Community Areas, following Akbar and Duranton (2017) and Kreindler (2023). We find an average elasticity of travel speed with respect to traffic flows between -0.12 to -0.19, comparable to existing estimates in the literature (Akbar and Duranton, 2017; Couture *et al.*, 2018). We find substantial heterogeneity across different areas of the city: traffic has a larger effect on travel times in central areas than in more peripheral areas.

We model wait times for public transit as a function of their frequency (the “fleet size” of buses and trains), taking into account the empirical frequency of delays and resulting changes in headway.

With our estimates of preferences and the congestion technology, we proceed to quantify the welfare effects of different transportation policies. We first explore the cases where the government only adjusts transit prices and frequencies or only implements road pricing. To explore the interactions between these two we then compute a counterfactual where the government does both.

We find that a planner who only controls public transit would like to reduce prices of both buses and trains by slightly more than fifty percent. The changes in frequency are different for buses and trains. The frequency of the former is reduced by 28% while that of trains increases by 10%. This is because trains serve higher value of time consumers and buses have low utilization. Together, these price and frequency changes lead to a weekly increase in overall welfare of \$1.45 million (\$0.54 per resident), which comes from a \$0.74 million weekly increase in consumer welfare and a reduction in environmental externalities of \$0.71 million per week. The effects on consumer welfare are progressive: lower-income travelers gain from lower prices, despite higher waiting times, whereas the highest income travelers are worse off due to their higher value of time. These welfare effects are strongly affected by budget considerations. A planner who did not face a budget constraint would reduce prices further and would set bus and train frequencies around 10% higher. This would lead to much larger consumer welfare gains for travelers

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<sup>13</sup>For comparison, the average hourly wage in Chicago’s metropolitan area is \$30. See [wage statistics from Bureau of Labor Statistics for the Chicago region](#).

of around \$10 million per week.

The optimal road price, when it is the only instrument of the planner, is 37 cents per kilometer, which implies a total price of \$7.4 for the average car commuter. This per-kilometer price leads to overall welfare gains of \$4.19 million per week, almost three times as much as with optimal transit policies. Most of these gains are due to a reduction in externalities of \$3.04 million per week. In the absence of rebates, consumer surplus would fall by \$25.1 million per week, which amounts to \$9.30 per resident. Thus, while road pricing is beneficial on efficiency grounds, it hurts consumers. Middle income consumers, who rely the most on cars, experience the largest welfare losses. Most of the consumer welfare losses would be offset if the government were to fully rebate the \$24.25 million in weekly revenue generated from road pricing, but the average consumer would still be worse off by \$0.31 per week—amounting to a total consumer surplus loss of \$0.85 million per week.

We then consider road pricing and public transit policies in conjunction. In this case, the government's budget constraint is no longer binding because of the revenue collected from road pricing. Transit policies are thus close to what they would be in isolation and without a budget constraint, as described above. Public transit becomes virtually free: the optimal fares are \$0.07 for buses and \$0.10 for trains. Compared to the status quo, train frequencies are around 20% higher and bus frequencies are around 20% lower. Overall welfare under this policy increases by \$5.60 million per week. Without rebates, consumer surplus decreases by \$16.20 million dollars per week (\$6 per resident), but it can increase by up to \$1.20 million per week if the government rebates its full surplus. These findings highlight the importance of jointly considering public transit policies and road pricing: road pricing can help achieve large reductions in externalities, but it only benefits consumers if resulting revenues are used to cross-subsidize public transit and to rebate travelers.

**Related Literature** Our work relates to several strands of the literature in transportation economics and industrial organization.

Some papers analyze transportation markets based on spatial equilibrium models. These studies are closely linked to theoretical work by Arnott (1996), which shows that taxis should be subsidized because of increasing returns to scale, and Lagos (2003), who formulates a spatial search and matching model that he calibrates to the New York City

taxi market. More recent empirical work includes Frechette *et al.* (2019) and Buchholz (2021), who also both model the New York City taxi market, as well as Brancaccio *et al.* (2020a), who model the dry bulk shipping industry. Castillo (2023), Rosaia (2023), and Gaineddenova (2022) study ride-hailing platforms. Kreindler (2023) combines a structural approach with experimental evidence to study the effects of congestion on the well-being of travelers. Fuchs and Wong (2022) also study a multi-modal transportation model in the context of freight transportation. Like Brancaccio *et al.* (2020b), who derive optimal policies for frictional transportation markets, we derive them for urban transportation markets with a budget-constrained social planner.

Within this strand of literature, Durrmeyer and Martínez (2023) and Kreindler *et al.* (2023) are most closely related to our work. Durrmeyer and Martínez estimate an equilibrium model of mode substitution and investigate the welfare effects of private car restrictions and road pricing. We depart from Durrmeyer and Martínez mainly in two ways. First, we introduce a framework to design optimal government policies that allows for different modes of transportation to interact through various channels: consumer choices, the congestion function, and the budget constraint of the planner. Second, we also consider public transit policies, their potential interaction with road pricing, as well as the distributional implications of such interventions—for which our granular data is particularly well suited. Kreindler *et al.* (2023) study optimal transit policies but focus on the optimal network configuration for buses. While our policy simulations are less granular in terms of network planning, we incorporate the trade-off that the social planner faces when setting policies for both public transit and private modes of transportation through congestion surcharges.

There are several papers that investigate the long run effects of urban transportation policy, such as new rail infrastructure and interactions with residential location choices. Severen (2023) measures the effects on commuter welfare and productivity of the Los Angeles Commuter Rail. Tsivanidis (2023) quantifies the equilibrium effects of new public transit infrastructure in Bogotá. Brinkman and Lin (2022) study the welfare implications of highway construction during the mid-1900s. Fajgelbaum and Schaal (2020), Allen and Arkolakis (2022), and Bordeu (2023) characterize the optimal allocation of road infrastructure. These papers focus only on one mode of transportation and abstract away from incorporating rich heterogeneity in substitution patterns. We depart from their work by abstracting away from the residential and firm location choice models but allowing for

rich demand substitution patterns across modes and heterogeneity, which is a key ingredient to understand distributional effects of transit policies. More recently, Barwick *et al.* (2021) explores the interaction between mode choices and residential location choices and Herzog (2021) quantifies the equilibrium effects of London’s downtown tolls. Whereas these papers focus on individual policies—road pricing, driving restrictions, and subway expansion—we allow the government to choose the optimal combination among a portfolio of policies: public transit prices and frequencies, and road pricing.

We also build on a classic theoretical literature in transportation economics that develop models that capture the interaction between schedule constraints and congestion (Small, 1982; Arnott *et al.*, 1990, 1993; Small *et al.*, 2005). We enrich these models by combining a congestion model with the demand approach used in industrial organization (Berry, 1994; Berry *et al.*, 1995), which allows us to model rider heterogeneity and account for the endogeneity of travel times and prices.

Finally, our work relates to several other works in empirical transportation economics. The estimation strategy we follow for the traffic congestion technology builds on the literature that measures the effects of traffic congestion (Akbar and Duranton, 2017; Akbar *et al.*, 2023; Couture *et al.*, 2018; Kreindler, 2023). This paper also relates to works that analyze different forms of road pricing. Hall (2018) shows theoretically that pricing some highway lanes (“Lexus lanes”) can lead to Pareto improvements. Cook and Li (2023) empirically explore the distributional effects of dynamic tolling, and Yang *et al.* (2020) exploit the variation that is induced by driving restrictions in Beijing to derive the optimal road congestion surcharge. The last two papers abstract away from mode substitution and the interaction between public and private transportation. Parry and Small (2009) derive theoretical expressions for the optimal prices of public transit. We extend their results to account for the joint effect of prices and quality improvements and for the distortions introduced by budget considerations. Furthermore, we model the resulting equilibrium adjustments by taking into account the linkages across the many thousands of markets in a city. Lastly, Leccese (2021) and Leccese (2022) investigate the pass-through and the distributional consequences of the ride-hailing surcharge in downtown Chicago that we use to identify price elasticities.

## 2 Background and Data

### 2.1 Background

Chicago is the third largest city in the U.S. and its public transit system, which is operated by the Chicago Transit Authority (henceforth CTA), is one of the largest in the nation. It includes a bus network of 127 routes with almost two thousand buses, and a train rapid transit system, which is known as the “Chicago L,” that has eight routes and 145 stations. Prices are per ride and independent of distance. The full fare for bus and the L are \$2.25 and \$2.50 respectively.<sup>14</sup> The CTA has a history of budget shortfalls, which suggests that it is important to account for budget considerations (see, for instance, Ramos (2019)). Lastly, passengers may travel by private for-hire-vehicles (FHV) in the form of taxis and ride hail. Taxis have a regulated rate of \$2.25 per mile or \$0.2 per 36 seconds as well as a \$3.25 base fare.<sup>15</sup> Ride hailing companies adjust prices dynamically according to market conditions.

### 2.2 Data description

We define a market as an origin-destination-time tuple. Chicago’s Community Areas (CAs) serve as our origin and destination locations. In total there are 77 CAs, with an average size of three square miles and an average population of 36,000 people.<sup>16</sup> We define a unit of time as an hour of the day, distinguishing between weekdays and weekends. Thus, we have 48 time periods. Our main dataset consists of travel flows, prices, and travel times for every mode in every market during January 2020.

To construct this dataset, we rely on a variety of raw data sources. First, we use administrative public transit microdata from the CTA. We have access to the near-universe of individual public transit trips for both buses and the Chicago L train. We observe the station or bus stop of origin, the time when the passenger tapped in, and an inferred drop-off L station or bus stop (Zhao *et al.*, 2007).<sup>17</sup>

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<sup>14</sup> Reduced fares exist for students and seniors and a monthly pass for unlimited rides can be obtained at \$75. There are also daily, 3-day, weekly, and monthly passes. See [CTA Fares](#) for additional details.

<sup>15</sup> See [Chicago Taxi Fare Regulation](#).

<sup>16</sup> See [Community Areas in Chicago](#).

<sup>17</sup> Unfortunately, we do not have data on Metra rides as this is managed by the Regional Transit Authority (RTA) rather than by the CTA. There were 74 million Metra rides in 2019 in Cook, DuPage, Kane, Kendall, Lake, McHenry, and Will counties combined, corresponding to less than 1% of trips in the region before the pandemic. See [My Daily Travel](#) and [Annual/Monthly Ridership](#) for details.

The second data source, which is published by the City of Chicago, contains the universe of de-identified taxi and ride-hailing trips.<sup>18</sup> It includes prices, drop-off and pick-up locations, trip length, trip duration, and the number of pooled riders in the case that the trip is a pool ride.

The third source is based on Veraset mobile-phone location data, which keeps track of people's movements over time.<sup>19</sup> This dataset records a device ID and a sequence of GPS coordinates and timestamps, for approximately 40% of active cellphone devices in the US. We infer all motorized trips from the sequence of GPS coordinates for each individual device. Appendix A.3.1 describes this process in detail. The frequency with which records are generated depends on the applications installed by the user, so we restrict our analysis to devices with frequent location information. As we explain at the end of this section, this restriction results in a sample of trips and travelers that is representative across many dimensions. We thus multiply the cellphone trips by a common inflation factor to arrive at the total number of trips.

We combine these three data sources to construct the near-universe of market flows across all modes: private car, taxi, ride-hailing, and public transit (buses and subway). While the CTA data allow us to observe the total number of trips for buses, trains, taxis, and ride-hailing, we do not have official records to directly measure car trips. Given that the cellphone data covers all motorized trips, we can recover car trips by subtracting public transit, taxi, and ride-hailing trips from the cellphone trips.

To observe the counterfactual travel times and routes that would have taken place if travelers had chosen a different mode, we use Google Maps data. We query all 30,796,848 possible trips by (origin census tract, destination census tract, hour of the week) triple, which we then combine with our flows data.

Our comprehensive dataset has several advantages over survey data. First, while survey data is designed to be representative at the city level, it often falls short of being representative at a more granular spatial or temporal resolution. Even worse, the coverage of surveys diminishes at higher resolutions, leading to sparse data. We show that over 60% of the origin-destination CA pairs have zero trips—a problem that is exacerbated when we further break down the data by time period (see Appendix A.5). This means that, although surveys are useful to understand aggregate movement patterns, they are

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<sup>18</sup>Source: Chicago Data Portal, Transportation Network Providers - Trips

<sup>19</sup>Source: Veraset: Location Data Provider

ill-suited to studying fine-grained movement patterns or policy interventions that are heterogeneous across space and across time. Our detailed data also allows us to estimate the relationship between vehicle flow and traffic speed throughout the city. Lastly, the richness of our data also has econometric advantages. We can invert market shares (Berry (1994), Berry *et al.* (1995)) at a very granular level, which allows us to construct moment conditions based on high-resolution instruments to address the endogeneity of prices and travel times.

We add demographic information to our main dataset using census data from the 2016-2020 American Community Survey (ACS) sample. We use information on income and car ownership rates at the census tract level. We match devices to census tracts by inferring a device user's home location based on the modal GPS location during nighttime hours. We partition devices into two groups: those who spend at least three nights of the month in their modal night location, which we call *residents*, and those who spent at most two nights, which we call *visitors*.<sup>20</sup> Trips made by residents account for 93.3% of all cellphone trips. For residents, we impute their income and car ownership probability as the median income and car ownership rates of their home census tract. We are thus able to construct the distribution of travelers' income and car ownership for every market. Lastly, for validation of our data construction we use travel survey data from the Chicago Metropolitan Agency for Planning (CMAP) 2019 Household Travel Survey.<sup>21</sup>

Since the cellphone data only covers forty percent of devices, and among those we select the ones that provide sufficient location data, one may wonder whether some demographic groups are over- or underrepresented. We test the representativeness of our data based on census covariates assigned to devices by our inferred home location. We find that income groups are equally represented in the data (see Appendix A.1.3). We also test whether our cellphone trips are representative of travel patterns by comparing the distribution of the time and distance travelled in the cellphone data and in the survey data, and find that both aggregate distributions align quite well (Appendix A.1.4).

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<sup>20</sup>Figure 13 in Appendix A.1.3 shows the average share of visitors by origin location. As expected, visitors are highly concentrated in the city center and in the airports.

<sup>21</sup>[Source: My Daily Travel survey \(website\)](#)

## 2.3 Summary Statistics and Descriptive Results

We present descriptive evidence in four parts: First, we explore the characteristics and usage of different transportation modes.<sup>22</sup> Next, we analyze rider characteristics and how they correlate with mode choice. We then document evidence on the low utilization rates of buses. Finally, we present evidence of traffic congestion in the raw data.

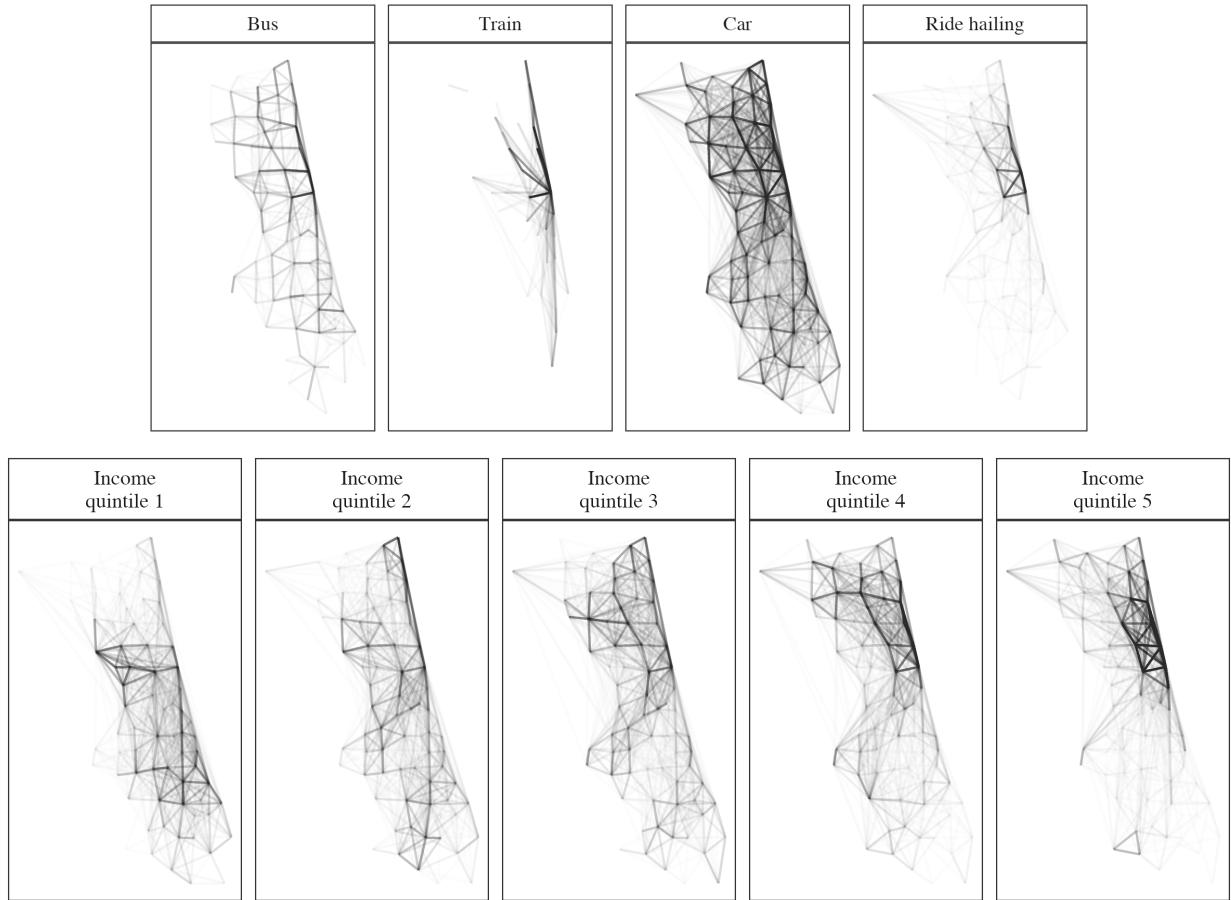


Figure 1: Trips by mode and by income

*Notes:* These figures show a random sample of 20,000 trips. A line connects the origin and destination of every trip. The top figure splits trips by mode. The bottom figure splits them by the income quintile of the traveler.

Although Chicago has one of the most extensive public transit systems in the US, about 69% of trips are taken by car. Public transit accounts for 23% of trips, with buses

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<sup>22</sup>We exclude biking but these modes only represent 1.8% and 2% of the overall trips.

taking slightly more than half of this share. Ride-hailing accounts for 6% of trips. Lastly, taxis account for only 0.9% of trips—and, hence, we exclude them for most of our analysis. Figure 1 shows that the characteristics of trips differ across modes. Bus and car trips are spread throughout the city. Train trips are concentrated along the corridors connected by the L lines. Ride hailing mostly accounts for short trips downtown or north of downtown, along the coast of Lake Michigan, as well as for trips to and from the two major airports in Chicago: O’Hare to the northwest and Midway to the southwest.

Chicago has stark income differences that are reflected in distinct travel patterns. The bottom panel of Figure 1 shows that low income travelers mostly stay in the south and the west parts of the city. The highest income travelers mostly stay downtown and to the north, along the coast of Lake Michigan. Trips of intermediate income travelers are more evenly spread throughout the city.

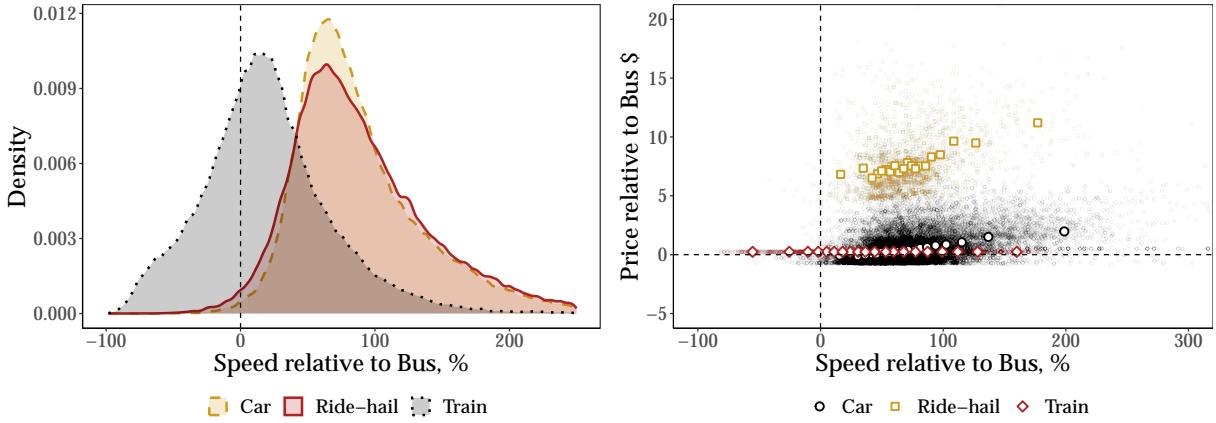


Figure 2: Speed and price differences across modes of transportation

*Notes:* The left panel reports density of speeds across four modes of transportation. The right panel presents scatter and binscatter plots of prices and speed by mode. Observations are at the market level, weighted by the total number of trips in the market.

Figure 2 shows the differences in speed, travel time, and prices across modes. The left panel shows the distributions of the speed of all modes relative to the speed of buses. Trains are, on average, 10% faster than buses, and cars and ride-hailing are on average almost twice as fast as buses. The right panel shows that buses are also the cheapest option for most trips, but car prices are lower for some trips, which are mostly short, inexpensive trips with a low price. We can also see that almost all of the comparisons with respect to buses lie on the upper-right quadrant, indicating that modes that are faster

also tend to be more expensive relative to travelling by bus.<sup>23</sup> This price-speed trade-off will therefore be a key margin of decision for our travellers when choosing their preferred mode of transportation.

Figure 3 shows patterns in car ownership across income levels. We observe an inverted U-shape: car ownership first increases in income and then declines again at the top of the income distribution. Our demand estimation incorporates car ownership, as otherwise our estimates would conflate preferences for non-car modes of transportation with the actual possibility of travelling by car.

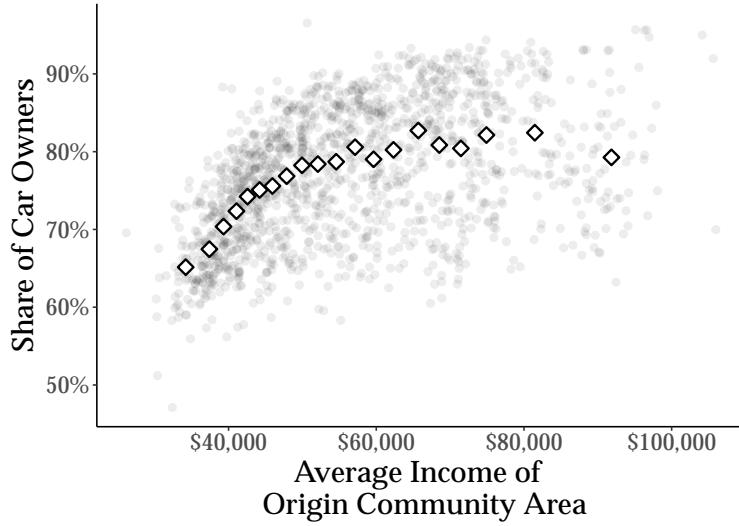


Figure 3: Car ownership by travelers' income

*Notes:* This figure plots a scatterplot and a binscatter of car ownership against median income. Our level of observation is a census tract.

Figure 4 shows how mode choices vary by travelers' income, which we proxy using the average income of the origin CA. Lower income travelers are more likely to travel by bus. Perhaps surprisingly, higher income people are more likely to use trains than lower income people (second panel): as we can see in Figure 1, train trips are concentrated along L lines, which tend to be located in higher income areas. Consistent with car ownership patterns, car usage follows an inverted-U shape: middle income people are most likely

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<sup>23</sup>The positive relationship between speed and price arises from the fact that longer trips—those that are more expensive by construction—are also more likely to travel through highways rather than secondary roads.

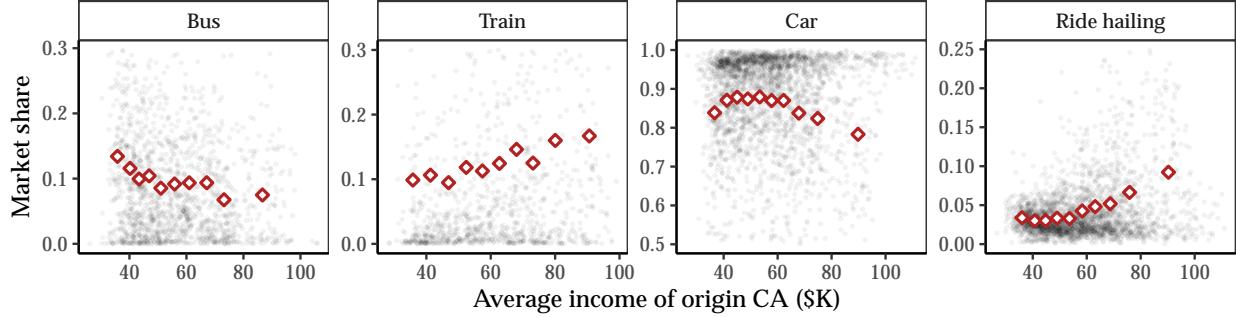


Figure 4: Mode market shares by travelers' income

*Notes:* Each one of these panels presents both a scatterplot and a binscatter of market shares against average income for each mode. Each observation represents trips going from an origin Community Area to a destination Community Area. Note that the vertical scale varies by mode.

to use cars—and, thus, they are likely to be hurt the most by road pricing. Finally, ride-hailing is overwhelmingly used by the highest income people.

We now turn to data on the occupancy of buses. Figure 5 shows that even though buses are the cheapest available mode of transportation, many of them are running empty. Even during the morning and afternoon rush hours, median utilization rates stay below 20%. If we look at the distribution within hour, we see that 90% of buses are at a utilization below 75% during the morning rush hour. Moreover, this graph also shows that, while certain buses may travel at full capacity at peak times, those events are extremely rare. Only the 99th percentile bus reaches full capacity at any time during the day, and only during the morning and afternoon rush hours. These low utilization rates suggest that resources may not be optimally allocated, implying that there is scope for the city government to optimize on the supply of public transit.

Lastly, we present raw data patterns that show how traffic congestion impacts travel times. We examine the relationship between the number of vehicles moving between pairs of adjacent CAs and travel times. The left panel of Figure 6 shows this relationship, after residualizing on market fixed effects. The right panel shows two separate markets. In both panels, the data exhibit a “hockey-stick” pattern: the trend is flat at lower vehicle counts and then, past a certain threshold, it begins to rise at a rate that is close to linear in log space. This means that when there are not many cars on the road, speeds are close to constant. However, beyond a certain traffic volume, additional vehicles decrease travel speeds with an approximately constant elasticity. These patterns guide our selection of

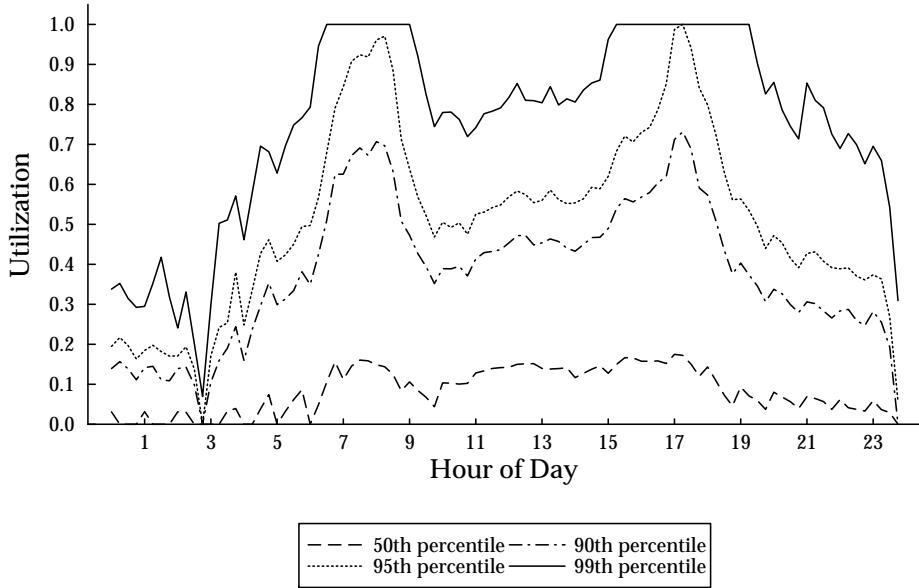


Figure 5: Bus utilization rates

*Notes:* This figure shows the 50th, 90th, 95th, and 99th percentile bus utilization rate over the course of the day, restricting to weekdays. We measure utilization for each bus every fifteen minutes by taking the number of riders on the bus divided by the capacity of the bus. We conservatively assume each bus has a capacity of 53, which is the smaller of the two bus sizes used by the CTA. If the number of observed riders is greater than the assumed capacity we set the utilization rate to 1.

the congestion function's form and the instrument that we use for travel times when estimating demand.

### 3 Model

Our model consists of three parts. First, there are travelers with fixed origin and destination who choose either one of the available modes or not to travel at all. Second, there is a transportation technology that captures the relationship between the number of people who use a mode and its travel time. Third, there is a social planner that maximizes welfare subject to a budget constraint.

#### 3.1 Setup and Equilibrium Definition

We first present a simple version of our model, in which we focus on only one market and describe traveler preferences and the transportation technology in the abstract. This

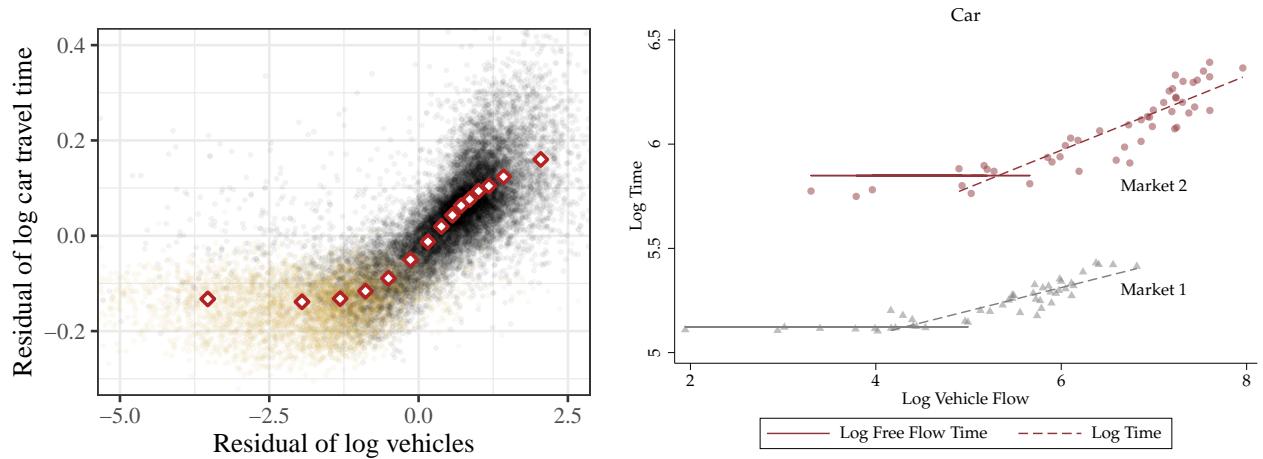


Figure 6: Congestion Graphs

*Notes:* These figures present the relationship between travel times and the number of vehicles, using data at the level of a pair of adjacent community areas during an hour of the week. Yellow points represent observations between midnight and 5 am. The left panel shows log-travel times against the logarithm of vehicles on the road, where both variables are residualized by market fixed effects. The right panel singles out two random markets and shows the same relationship.

will allow us to present theoretical results in Section 3.2 that highlight the general forces in the social planner's objective function. In Section 3.3 we present the empirical version of our model, which accounts for inter-temporal and spatial variation in supply and demand as well as for the spatial linkages across markets due to congestion and driver movements. In Sections 3.3.1 and 3.3.2 we provide the details on how we model traveler demand and the transportation technologies for different modes.

The first component of our model are travelers, who differ across two observable dimensions: their income and whether they own a car. We denote a traveler's type by  $\theta \in \mathbb{R}^n$ , with density  $f(\cdot)$ , which captures observable characteristics as well as unobservable preferences for modes.

A traveler of type  $\theta$  decides which transportation mode  $j$  to take to her destination. She can choose among the set  $\mathcal{J}(\theta)$ , which varies depending on whether public transit is easily accessible and on whether she owns a car. She can also choose the outside option of not taking a trip, which we denote by 0. Thus, her choice set is given by  $\mathcal{J}(\theta) \cup \{0\}$ . She gets utility  $u_j(t_j, \theta) - p_j$  if she takes transportation mode  $j$ , where  $p_j$  is the price and  $t_j$  is the travel time. This travel time includes the in-vehicle time, the waiting time before

the trip starts, and—for public transit—the walking time to the station or stop.<sup>24</sup> We normalize the utility of the outside option to zero. Therefore,  $u_j(t_j, \theta) - p_j$  is the utility measured relative to not taking a trip.

The traveler chooses the mode that maximizes her utility among her choice set:

$$j^*(\theta) = \operatorname{argmax}_{j \in \mathcal{J}(\theta) \cup \{0\}} u_j(t_j, \theta) - p_j \quad (1)$$

Given vectors of prices  $\mathbf{p}$  and total trip times  $\mathbf{t}$  for all modes, demand for mode  $j$  is given by

$$q_j = q_j(\mathbf{p}, \mathbf{t}) = \int_{\Theta_j(\mathbf{p}, \mathbf{t})} f(\theta) d\theta, \quad (2)$$

where  $\Theta_j(\mathbf{p}, \mathbf{t})$  is the set of traveler types choosing mode  $j$  at  $(\mathbf{p}, \mathbf{t})$ .

We refer to the vector  $\mathbf{q}$  as trips. Gross consumer utility and consumer surplus are given by

$$U(\mathbf{p}, \mathbf{t}) = \sum_j \int_{\Theta_j(\mathbf{p}, \mathbf{t})} u(t_j, \theta) f(\theta) d\theta \quad \text{and} \quad CS(\mathbf{p}, \mathbf{t}) = \sum_j \int_{\Theta_j(\mathbf{p}, \mathbf{t})} (u(t_j, \theta) - p_j) f(\theta) d\theta. \quad (3)$$

Travel times are determined by a transportation technology that captures the dependence on the number of travelers choosing each mode as well as on the overall capacity of the fleet for each mode. The fleet size for public transit is a policy choice and determines the frequency at which buses and trains run. For ride-hailing, the fleet size is determined by the number of drivers. The transportation technology also captures the fact that the in-vehicle time for road-based modes of transportation depends on how congested roads are.

Accounting for all these considerations, we can compactly write the vector  $\mathbf{t}$  of travel times for all modes as

$$\mathbf{t} = T(\mathbf{q}, \mathbf{k}), \quad (4)$$

where  $\mathbf{k}$  is the vector of fleet sizes for all modes. For each mode  $j$  there is a cost  $C_j(q_j, k_j)$  to supply  $q_j$  rides with fleet size  $k_j$ . This cost function includes both labor costs and phys-

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<sup>24</sup>Observe that we can also incorporate heterogeneity in traveler's sensitivity to prices  $u_j(t_j, \theta) - \theta_p \cdot p_j$ . Note that a re-scaled version of utility, namely  $(u_j(t_j, \theta) - \theta_p \cdot p_j)/\theta_p$  leads to the same optimal choices. Using  $u_j(t_j, \theta) - p_j$  has the advantage of measuring utility directly in monetary terms for all consumers  $\theta \in \Theta$ , where the Value of Time (VOT) can be computed as  $\partial u_j(t_j, \theta)/\partial t_j$ .

ical costs, such as fuel and vehicle depreciation. Additionally, there is an environmental externality  $E_j(q_j, k_j)$  that is borne by society. We assume that both functions are increasing and convex in both arguments.

With this notation we can now define an equilibrium.

**Definition 1** (Transportation Equilibrium). *Given a set of prices  $\mathbf{p}$  and fleet sizes  $\mathbf{k}$ , a market equilibrium is a set of trips  $\mathbf{q}$  and travel times  $\mathbf{t}$ ,  $(\mathbf{q}^*(\mathbf{p}, \mathbf{k}), \mathbf{t}^*(\mathbf{p}, \mathbf{k}))$  such that (2) and (4) hold.*

Travel times and quantities are the equilibrium objects. Prices and fleet sizes are exogenous and the social planner is able to determine a subset of them. Therefore, for any given set of fleet size and prices, travel times adjust to bring the market into equilibrium.

## 3.2 The Social Planner's Problem

The city government chooses the prices and fleet sizes of buses and trains as well as the cost of traveling by car through a congestion surcharge. We denote these modes as  $\mathcal{J}_G$ . The city government's goal is to maximize welfare subject to a budget constraint.

To define welfare and the government's budget, it will be easier to think of the allocation  $(\mathbf{q}, \mathbf{k})$  that arises in equilibrium. The government's revenue is equal to the payments it obtains from travelers minus its costs:

$$\Pi(\mathbf{q}, \mathbf{k}) = \sum_{j \in \mathcal{J}_G} p_j(\mathbf{q}, T(\mathbf{q}, \mathbf{k})) q_j - C_j(q_j, k_j).$$

This revenue cannot fall below  $-B$ , where  $B$  is the transportation budget. Welfare is equal to gross consumer utility minus the cost of transportation provision and minus all externalities:

$$W(\mathbf{q}, \mathbf{k}) = U(\mathbf{q}, T(\mathbf{q}, \mathbf{k})) - C(\mathbf{q}, \mathbf{k}) - E(\mathbf{q}, \mathbf{k}).$$

If we let  $\mathbf{p}_G$  and  $\mathbf{k}_G$  be the vectors of prices and capacities set by the government, the objective function is:

$$\begin{aligned} & \max_{\mathbf{p}_G, \mathbf{k}_G} U(\mathbf{q}, T(\mathbf{q}, \mathbf{k})) - C(\mathbf{q}, \mathbf{k}) - E(\mathbf{q}, \mathbf{k}) \\ & \text{s.t. } \sum_j p_j(\mathbf{q}, T(\mathbf{q}, \mathbf{k})) \cdot q_j - C(\mathbf{q}, \mathbf{k}) \geq -B \end{aligned} \tag{5}$$

To state the optimality conditions, it will be useful to introduce notation for the derivatives of costs, externalities, travel times and utilities with respect to some quantity  $x$ . To do so, we will use superscripts. For example,  $C_j^q$  denotes the derivative of the cost of mode  $j$  with respect to the quantity of rides of mode  $j$ . Also, let  $\Omega_{lj}$  represent elements of the inverse Jacobian of  $q(p, T)$  with respect to  $p$ . And, finally, define  $D_{lj} = \frac{\partial q_l}{\partial p_j} / \frac{\partial q_l}{\partial p_j}$ , as the diversion ratio from  $j$  to  $l$  of a price increase for mode  $j$ .

With this notation in place we can now state the following proposition.

**Proposition 1.** *Prices under the solution of the social planner's problem (5) are given by:*

$$p_j = \overbrace{C_j^q + E_j^q}^{\text{Mg. cost and env. externality}} - \underbrace{\sum_l u_l^T \cdot T_{lj}^q}_{\text{Network effects}} + \overbrace{M_j^q}^{\text{Diversion}} + \frac{\lambda}{1+\lambda} \cdot \left( \underbrace{\sum_{k \in \mathcal{J}_G} q_k \cdot \Omega_{kj}}_{\text{Market power markup}} + \underbrace{\tilde{M}_j^q - M_j^q}_{\text{Diversion distortion}} - E_j - \underbrace{\sum_l (\tilde{u}_l^T - u_l^T) \cdot T_{lj}^q}_{\text{Spence distortion}} \right) \quad (6)$$

where  $\lambda$  is the Lagrange multiplier for the budget constraint,  $\tilde{u}_j^T$  is a weighted sum of the derivative of gross utility among marginal travelers with respect to the total time if pickup times for mode  $j$  increase by 1%, and  $M_j^q$  and  $\tilde{M}_j^q$  (which we discuss below) are defined as:

$$M_j^q \equiv \sum_{k \neq j} D_{kj} \left( C_k^q + E_k^q - \sum_l u_l^T \cdot T_{lk}^q - p_k \right) \quad (7)$$

$$\tilde{M}_j^q \equiv \sum_{l \in \mathcal{J}_G \setminus j} D_{lj} \left( C_l^q + \sum_{k \in \mathcal{J}_G} q_k \cdot \Omega_{kj} - \sum_m \tilde{u}_m^T \cdot T_{ml}^q - p_l \right). \quad (8)$$

*Proof.* See Appendix C.2 □

To understand expression (6), think first of an unconstrained social planner, in which case  $\lambda = 0$  and the final term drops out. Prices are equal to the Pigouvian solution—in addition to the direct costs, they are set to internalize environmental externalities, network effects, and an additional term that we call the *diversion term*, which we explain below.

Network effects are equal to the sum over modes of the product of  $\bar{u}_k^T$ , the derivative of gross utility with respect to time, and  $\tilde{T}_{kj}$ , the change in total time given an additional

trip using mode  $j$ . For  $j \neq k$ ,  $\tilde{T}_{kj}$  is positive due to congestion (or zero if the two modes are not interrelated), so those terms lead to a Pigouvian tax. But  $\tilde{T}_{jj}$  can be negative. For livery vehicles, for instance, increasing the number of trips and fleet size by the same factor results in a reduction in waiting times due to returns to scale (i.e., economies of density) in the matching process. In that case, there should be a Pigouvian subsidy, as noted by Arnott (1996).

We now explain the diversion term  $M_j^q$ . It accounts for the fact that modes that are not priced by the planner are suboptimally priced. If there is no road pricing, for instance, the price of driving a car will be too low and too many people will drive their car. As a second best, the government would want to lower the price of public transit to induce substitution away from cars. In the expression for  $M_j^q$  (equation 7), the term in brackets can be thought of as deviations of prices from a standard Pigouvian solution—marginal costs plus marginal externalities plus network effects.  $M_j^q$  is therefore the diversion-ratio weighted sum of these deviations for all other modes. It captures the extent to which an increase in the price of  $j$  induces substitution towards modes that are overpriced and, thus, are chosen by fewer travelers than is socially optimal. This form of substitution increases welfare if the diversion term is positive, and, hence, the optimal price is higher. In particular, note that this term is zero whenever all other modes are already priced at the Pigouvian solution.

We now also take into account the budget constraint. To stay on budget, the social planner behaves qualitatively like a monopolist because it needs to raise revenue. This introduces a market power markup in equation (6). For the same reason, the diversion term is distorted towards its revenue-motivated equivalent  $\tilde{M}_j^q$ . It is very similar to the social welfare-motivated diversion term  $M_j^q$ , except that it captures the extent to which an increase in the price of  $j$  induces substitution towards modes that are chosen by fewer travelers than is optimal *to maximize government revenues*, rather than to maximize social welfare.<sup>25</sup> The social planner now also under-weights environmental externalities. Finally, there is a Spence distortion: the government internalizes effects on other travelers' utility, but imperfectly: it accounts for changes in the utility of marginal travelers, rather

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<sup>25</sup>In this term, instead of the effect of a price change on all other modes, the city now only considers modes  $\mathcal{J}_G$  from which it can raise revenue. And the term in brackets now represents deviations in prices from the revenue maximizing prices. It no longer includes the marginal externality but instead includes a markup term  $\sum_{k \in \mathcal{J}_G} q_k \cdot \Omega_{kj}$ . Furthermore, rather than considering the effect of travel times on the average traveler  $u_m^T$ , the city now cares about the marginal traveler  $\tilde{u}_m^T$ .

than that of all travelers. As  $\lambda \rightarrow \infty$ , the social planner no longer cares about welfare and it only cares about its budget. These two expressions then become the first order conditions for profit maximization: terms related to environmental externalities cancel out, and there is a full markup and a full Spence distortion.

In Appendix C.1 we derive a similar decomposition for travel times, which follows the same logic as the price decomposition.

In our counterfactuals, we will come back to these results and empirically decompose the optimal prices and travel times into these constituent parts to explain the planner's competing motives when setting optimal prices and travel times. We use a version of this decomposition that applies to a multiple-markets setting, which we derive in Appendix C.3.

### 3.3 Empirical Model

Here we describe the empirical version of our demand model and of the transportation technology, which we use for  $U(\mathbf{p}, \mathbf{t})$  and  $\mathbf{t} = T(\mathbf{q}, \mathbf{k})$  in the above expression. The city is composed of Community Areas  $a \in \mathcal{A}$ , the main level of spatial aggregation we consider. We will also refer to them as *locations*. Time is divided into hours  $h \in \mathcal{H}$ .<sup>26</sup>

#### 3.3.1 Demand

First, we define a market  $m = (a, \tilde{a}, h) \in \mathcal{M}$  as the collection of people who make travel decisions from Community Area  $a$  to Community Area  $\tilde{a}$  at a particular time  $h$ .<sup>27</sup> In each market there is an exogenous number of potential travellers  $\lambda_m$ . They decide which mode  $j \in \mathcal{J}_m^i \cup \{0\}$  to use, where  $j = 0$  denotes the outside option of staying put, by solving the

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<sup>26</sup>Specifically, we define  $h$  as hour-of-the-week. We consider  $48 = 2 \times 24$  hours-of-the-week, corresponding to an average weekday day and an average weekend day. Daily variation is aggregated by taking averages across dates for the same hour-of-the-week.

<sup>27</sup>Given that our level of temporal variation averages across dates, traveling decisions should be thought as the choice over average trips at a given hour-of-the-week  $h$  rather than stemming from very short-run temporal variation coming from shocks or special occasions.

following problem:

$$U_m^i = \max_{j \in \mathcal{J}_m^i \cup \{0\}} \delta_{mj}^i + \varsigma_{mg(j)}^i + (1 - \rho) \epsilon_{mj}^i = \max_{j \in \mathcal{J}_m^i \cup \{0\}} \xi_{mj} + \alpha_T \cdot T_{mj} + \alpha_p^i \cdot p_{mj} + \varsigma_{mg(j)}^i + (1 - \rho) \epsilon_{mj}^i \quad (9)$$

where  $T_{mj}$  denotes the sum of the waiting and travel times for mode  $j$ ,  $p_{mj}$  is the price for mode  $j$ ,  $\alpha_T$  is the preference parameter over travel times, and  $\alpha_p^i$  is the person  $i$ -specific price coefficient.<sup>28</sup> Motivated by the disparities in mode choice across the income distribution from Section 2.3, we allow  $\alpha_p^i$  to vary by income  $y_i$  according to  $\alpha_p^i = \alpha_p / y_i^{1-\alpha_{py}}$ , so that  $\alpha_{py}$  captures the extent to which the price coefficient varies with income.<sup>29</sup> Differences in this coefficient then capture differences in the marginal utility of money across the income distribution, which lead to heterogeneity in the trade-off between time and money.

The parameters  $\varsigma_{mg(j)}^i$  and  $\epsilon_{mj}^i$  are idiosyncratic unobserved taste shocks. The taste shock  $\varsigma_{mg(j)}^i$  is common to all goods within the group  $g(j) \equiv \mathcal{J}_m^i$ , defined as all modes of transportation excluding the outside option. This error allows the unobserved utility shocks of all inside options to be correlated. The taste shock  $\epsilon_{mj}^i$  is specific to mode  $j$ . This model can be thought of as a sequential choice problem: First, travellers choose whether they want to travel or not, and second, conditional on travelling, travellers choose which mode they use.

We allow the choice sets  $\mathcal{J}_m^i$  to vary across markets and across consumers within the same market to capture that some modes of transportation may not be available within a market or to specific consumers. For instance, some Community Areas cannot be reached by train and some consumers do not own a car. To construct these choice sets, we first assume that the set of available modes in market  $m$  coincides with the set of modes we observe being used in that market.<sup>30</sup> We then allow for random choice set variation across consumers. In particular, for each consumer type  $i$  in a given market  $m$  we compute the probability  $p_{car,m}^i$  of owning a car using census data. We then assume that cars are in the choice set  $\mathcal{J}_m^i$  of consumers of type  $i$  only with probability  $p_{car,m}^i$ .

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<sup>28</sup>In this model, the VOT of consumer  $i$  is computed as  $\alpha_T / \alpha_p^i$ .

<sup>29</sup>This functional form corresponds to a first-order Maclaurin series approximation of utility that follows a Box-Cox transformation, as in Miravete *et al.* (2023).

<sup>30</sup>We also restrict to markets where we observe car trips so that cars are always an available mode. These markets capture 96% of observed trips.

We assume that  $\epsilon_{mj}^i$  follows a Type I Extreme Value distribution and that the *iid* shock for group  $g(j)$  follows the unique distribution such that  $\varsigma_{mg(j)}^i + (1-\rho)\epsilon_{mj}^i$  is also an extreme value random variable. The parameter  $\rho \in [0, 1]$  governs the within group correlation; in our case this means that a larger  $\rho$  implies that all inside options (e.g. all modes) are closer substitutes. The probability that person  $i$  chooses mode  $j$  in market  $m$  is therefore given by:<sup>31</sup>

$$\mathbb{P}_{mj}^i = \frac{\exp\left(\frac{\delta_{mj}^i}{1-\rho}\right)}{\sum_{j \in g} \exp\left(\frac{\delta_{mj}^i}{1-\rho}\right)^{\rho} \cdot \left[\sum_g \sum_{j \in g} \exp\left(\frac{\delta_{mj}^i}{1-\rho}\right)^{(1-\rho)}\right]}. \quad (10)$$

Integrating over  $\alpha_p^i$ , which follows a market-specific distribution  $F_m$ , we obtain that market shares and trips for mode  $j$  in market  $m$  are given by:

$$\mathbb{P}_{mj} = \int \mathbb{P}_{mj}^i dF_m(\alpha_p^i) \quad q_{mj} = \lambda_m \cdot \mathbb{P}_{mj}.$$

### 3.3.2 Transportation Technology

In this section, we describe the transportation technology, which determines travel times as a function of trips and fleet sizes (which determine frequencies). We describe the (effective) travel time as the sum of several components that vary by mode: walk time, wait time, and in-vehicle time:

$$T_{mj} = \gamma \cdot (T_{mj}^{\text{walk}} + T_{mj}^{\text{wait}}) + T_{mj}^{\text{vehicle}}.$$

The parameter  $\gamma$  represents the relative distaste for time spent walking or waiting relative to in-vehicle time. Throughout the empirical model, we set  $\gamma = 2$  following the survey on the empirical values of the different components of the value of time by Small (2012). For ride-hailing, walk times are assumed to be zero; for cars, both walk time and wait time are assumed to be zero. Walk times are taken as exogenous and given by market-level averages of the walk times according to our Google Maps queries. We now describe how we model in-vehicle time  $T_{mj}^{\text{vehicle}}$  as a function of vehicle flows, which affects all modes but trains. Then we describe how we model wait times for public transit and ride-hail.

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<sup>31</sup>This is the product of the conditional probability  $\mathbb{P}_{mj|g}^i = \frac{\exp\left(\frac{\delta_{mj}^i}{1-\rho}\right)}{D_g}$  and  $\mathbb{P}_g^i = \frac{D_g^{(1-\rho)}}{\sum_g D_g^{(1-\rho)}}$ , where  $D_g = \sum_{j \in g} \exp\left(\frac{\delta_{mj}^i}{1-\rho}\right)$ .

To model how traffic congestion affects travel times, we think of the city as a directed graph in which nodes are CAs and edges connect CAs that are spatially adjacent. Edge  $e = (a, a')$ , for instance, connects CAs  $a$  and  $a'$ . We assume that routes are exogenous and pre-determined. If a traveler uses mode  $j$  in market  $m = (a, \tilde{a}, h)$ , she follows a directed path  $r_{a, \tilde{a}, j} = (e_{a, a_1, j}, e_{a_1, a_2, j}, \dots, e_{a_{N_m-1}, \tilde{a}, j})$  over edges that connects  $a$  with  $\tilde{a}$ . During hour  $h$ , the total vehicle flow on edge  $e$  is

$$Q_{eh} = \sum_j w_j \cdot q_{ehj}, \quad (11)$$

where  $q_{ehj}$  is the total number of vehicles for mode  $j$  that go through  $e$  and can be defined as the sum across all routes that go through edge  $e$ ,  $\mathcal{R}_{hj}^e$ :

$$q_{ehj} \equiv \sum_{r \in \mathcal{R}_{hj}^e} q_{rhj}. \quad (12)$$

The weight  $w_j$  on vehicle flows captures the fact that some modes congest more than others—buses congest more than cars, and trains do not congest at all. We assume that the travel time over edge  $e$  at time  $h$  for mode  $j$  is given by the following function:

$$T_{ehj}^{\text{vehicle}} = \max\{T_{ej}^0, A_{ehj} \cdot Q_{eh}^{\beta_j}\}. \quad (13)$$

This functional form is directly motivated by the observed empirical patterns in Figure 6 in our descriptive section. For every segment in our directed graph, there is a range with low vehicle flows for which, regardless of the number of vehicles that are on the road, the time that it takes to travel through the edge is  $T_{ej}^0$ . We call this the *free-flow* time. As we have shown in Figure 6, free flow times vary by markets. Variation in free-flow traffic depends on existing infrastructure, for example the number of traffic lights, road quality, or the existence of a highway.

The second part of the function represents the range in which travel times increase with the flow of vehicles.  $A_{ehj}$  is a segment-mode specific parameter that describes the underlying road infrastructure, and  $Q_{eh}^{\beta_j}$  captures how traffic increases travel times.

Now that we have described how traffic flows map into travel times at the edge level, we define the travel time in market  $m$  as the sum of the travel times over all edges in the

path  $r_{mj}$ :

$$T_{mj}^{\text{vehicle}} = \sum_{e \in r_{mj}} T_{ehj}^{\text{vehicle}}. \quad (14)$$

Next, we define how public transit wait times are determined. Suppose the frequency of public transit is  $k_{mj}$ , which implies an average time between vehicles—buses or trains—of  $1/k_{mj}$ . However, we would also like to capture variability in realized wait times because of delays and traffic. We thus assume that the time between vehicles follows some distribution with density  $\phi(\cdot)$  that has mean  $1/k_{mj}$  and variance  $\omega^2/k_{mj}$ . The variance of this distribution captures the degree of unpredictability in arrivals.

The density of travelers arriving between two subsequent buses or trains with a time difference of  $t$  is  $t \cdot k_{mj} \cdot \phi(t)$ : the density  $\phi(t)$  is multiplied by  $t \cdot k_{mj}$  because the longer the gap between vehicles, the more riders arrive between them.<sup>32</sup> If the time difference is  $t$ , a rider arriving between these two vehicles needs to wait  $t/2$  in expectation until the second vehicle. Therefore, the expected waiting time is given by

$$T_{mj}^{\text{wait}} = \int \frac{1}{2} t \cdot (t \cdot k_{mj} \cdot \phi(t)) dt = \frac{1 + \omega^2}{2k_{mj}}.$$

This expression captures the fact that, for a given frequency, the expected waiting time increases with the unpredictability  $\omega$  of public transit. If vehicles are always on schedule,  $\omega = 0$  and waiting times are uniformly distributed between 0 and  $1/k_{mj}$ , with mean  $1/(2 \cdot k_{mj})$ . In the opposite extreme, vehicle arrivals are a Poisson process—arrivals are always equally likely, regardless of whether the last vehicle arrived recently or not—in which case  $\omega = 1$  and the expected waiting time is  $1/k_{mj}$ . Our functional form therefore nests both of these cases. Our data allows us to observe the deviations from schedules, which we use to estimate  $\omega$ . Trains follow schedules closely, so we set  $\omega = 0$ . For buses, we find that  $\hat{\omega} = 0.194$ , which means that there is substantially more variability than for trains.<sup>33</sup>

In the case of ride-hailing, wait time  $T_{mj}^{\text{wait}}$  arises from a simple model of matching and of driver movements that captures three main forces. First, waiting times are lower when many drivers are working: there will be a large number of idle drivers and, thus, the nearest driver will be close to the location where the rider requested a trip. Second,

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<sup>32</sup>This implicitly assumes that travelers arrive to the stop or station at times that are uniformly distributed.

<sup>33</sup>To estimate this number, we compute actual times between buses and divide them by their average at the hour by route level. The variance of this ratio is our estimate  $\hat{\omega} = 0.194$ .

waiting times are higher when many travelers demand ride-hailing trips, depleting the market of available drivers. Third, at a given point in time, waiting times are lowest in those areas where more idle drivers are located. We set up a model of driver movements that parsimoniously accounts for the fact that there will mechanically be more idle drivers in neighborhoods with a net inflows of trips, as well as for the fact that drivers tend to relocate towards areas with higher earnings opportunities. Appendix D.1 presents further details, and explains how we estimate our model using data on Uber waiting times.

### 3.4 Costs and environmental externalities

We assume that costs and environmental externalities are proportional to the number of miles driven by the vehicles involved in each mode. For cars and ride hailing, the number of miles driven depends on how many passengers choose to travel using these modes. For buses and trains, on the other hand, the number of miles driven depends on the frequency with which routes run; hence, the marginal cost and externality of an additional passenger is zero.

For all modes, the cost per mile accounts for fuel or energy, vehicle depreciation, and maintenance. For buses and trains, it also includes labor costs.<sup>34</sup> Environmental externalities account for the social cost of carbon, for which we use latest EPA proposal of \$190 per tonne as the baseline number, as well as for the social cost of local pollutants, which we obtain from Holland *et al.* (2016).<sup>35</sup> Appendix D.2 describes in detail the numbers that we use for all costs and externalities.

## 4 Estimation and Computation

### 4.1 Demand model

The utility of agent  $i$  for taking mode  $j$  in market  $m$  is given by:

$$U_{mj}^i = \xi_{mj} + \alpha_T \cdot T_{mj} + \alpha_p^i \cdot p_{mj} + \varsigma_{mg(j)}^i + (1 - \rho) \epsilon_{mj}^i.$$

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<sup>34</sup> For ride hailing, labor costs depend on the number of drivers that are working, which is an exogenous quantity that is independent of the number of people that request ride-hailing trips.

<sup>35</sup> See [EPA Issues Supplemental Proposal to Reduce Methane and Other Harmful Pollution from Oil and Natural Gas Operations](#).

In this section, we explain how to estimate parameters  $\alpha_T$  and  $\alpha_p^i$ . Note that to estimate such parameters we need to address two important endogeneity concerns.

First, while the prices of cars and public transit are fixed and not a function of time-varying market demand-shocks, ride-hailing prices adjust to demand conditions, and thus are potentially correlated with the demand shocks  $\xi_{mj}$ . Therefore, for car, bus, and train we can use directly their prices as instruments, following the logic pointed out by DellaVigna and Gentzkow (2019). On the other hand, to address potential endogeneity of ride-hailing prices, we use price variation introduced by a ride-hailing surcharge that applies to all trips that either originate or end in a downtown zone between 6AM and 10PM. In Appendix B we show the area affected by this policy and how we compute difference-in-difference estimates from this policy, which imply an own-price elasticity of  $-1.42$ . We then construct moments that capture the difference between the model-implied demand response and those derived by the treatment effect estimates.

The second endogeneity issue concerns travel times because these are an equilibrium object, and thus, a function of demand shocks. Concretely, a positive demand shock  $\xi_{jm}$  for road based modes of transportation leads to more travel in this market, which in turn induces congestion and increases travel times. As it would for prices, this type of endogeneity biases the travel time coefficient, which we expect to be negative, towards zero. The idea for our instrument is based on our observation of a hockey-stick functional form for congestion in Figure 6. There are large differences in free flow speed across markets, as captured by the flat portion of this function. These differences are not driven by intertemporal variation in vehicle flows but by permanent differences in road infrastructure. For instance, between two edges that are only connected by small roads with many traffic lights, traffic will flow slower than if they were connected by a highway. We therefore use free flow times  $T_{mj}^0$  as an instrument for  $T_{mj}$ .<sup>36</sup> Importantly, free flow times are not affected by demand shocks, and therefore are not subject to the endogeneity concerns we are worried about.

Finally, we measure the overall size of the market as twice the number of trips that we observe in the cellphone data for that specific market.<sup>37</sup> We believe this margin of adjustment is important for our counterfactual simulations as some policies may induce

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<sup>36</sup> Specifically, we construct market-specific free-flow times for each mode by taking the minimum observed time for that mode across all hours for that origin-destination pair.

<sup>37</sup> For details about how we calibrate the size of the market please see Appendix A.4.1.

travelers not only to switch across modes but also to stop or start traveling.<sup>38</sup> In what follows, we define a traveler's outside option as staying put or walking.

We implement the estimation using a two-step GMM procedure with moment conditions in the form of:

$$\mathbb{E}[\xi_{mj} \cdot \mathbf{Z}_{mj}] = \mathbf{0},$$

for a given vector of instruments  $\mathbf{Z}_{mj}$ . Following the discussion above, for non-ride hail trips we use prices directly as instruments and free-flow times to instrument travel times. We also need additional instruments to identify the non-linear parameters  $\alpha_{py}$  and  $\rho$  that respectively govern income heterogeneity<sup>39</sup> and the correlation of taste shocks for the inside options. As suggested in Gandhi and Houde (2019), we construct both local and quadratic differentiation instruments based on free-flow times:

$$Z_{mj}^{\text{local}} = \sum_{j' \neq j} \mathbb{1}\{|T_{mj'}^0 - T_{mj}^0| < SD_{T^0}\} \quad \text{and} \quad Z_{mj}^{\text{quad}} = \sum_{j' \neq j} (T_{mj'}^0 - T_{mj}^0)^2.$$

With the idea that ride-hailing prices must be lower in markets where other modes offer similar travel times, we also interact these with an indicator for whether a mode is not ride-hail:  $\mathbb{1}_j^{rh} = \mathbb{1}\{j \neq \text{ride-hail}\}$ . Lastly, we interact the exogenous non-ride hail prices with income quintiles  $\pi_m^y$  in each market:

$$\mathbf{Z}_{mj} = (T_{mj}^0, Z_{mj}^{\text{local}}, Z_{mj}^{\text{quad}}, Z_{mj}^{\text{local}} \mathbb{1}_j^{rh}, Z_{mj}^{\text{quad}} \mathbb{1}_j^{rh}, p_{mj} \mathbb{1}_j^{rh}, \pi_m^y \cdot p_{mj} \mathbb{1}_j^{rh}).$$

Appendix E.1 validates the relevance of our instruments by showing the first-stage estimates.

Our GMM objective function also includes the following indirect inference moments, which match the demand response due to the surcharge:

$$\mathbb{E}[(\eta_{mj} - \tilde{\eta}_{mj}) \mathbb{1}\{j = \text{ride-hail}, m \in \mathcal{M}_\tau\}] = \mathbf{0},$$

where  $\eta_{mj}$  is the measured elasticity for ride-hailing trips as a response to the surcharge-

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<sup>38</sup>For example, Rosenblum *et al.* (2020) shows that a \$0.50 bus subsidy in Cambridge, MA increased ridership by 30%, without decreasing the shares of other modes for their surveyed population. Similarly, Brough *et al.* (2023) using experimental evidence show that the benefits of public transit subsidies primarily accrue from the creation of new trips to access services.

<sup>39</sup>For computational simplicity, we divide the population into five income bins according to their income quintile, with income levels set according to the median income of the corresponding quintile.

induced change in prices and  $\Delta\tilde{\eta}_{mj}$  is the model-implied elasticity.  $\mathcal{M}_\tau$  is the set of markets affected by the surcharge policy.

We compute the GMM objective function following the nested fixed-point algorithm outlined in Berry *et al.* (1995). First, we guess values for the parameter vector  $\theta \equiv (\alpha_p, \alpha_T, \alpha_{py}, \rho)$ . We can then recover mode-market mean utilities  $\hat{\delta}_{mj}(\theta) = \xi_{mj} + \alpha_T \cdot T_{mj}$  using an iterative contraction mapping.<sup>40</sup> Second, we construct the residuals  $\hat{\xi}_{mj}(\theta) = \hat{\delta}_{mj}(\theta) - \alpha_T \cdot T_{mj}$ , additionally concentrating out any fixed effects. Finally, we compute the value of the GMM objective function

$$J(\theta) = (\mathbf{Z} \cdot \hat{\xi}(\theta))' \cdot W \cdot (\mathbf{Z} \cdot \hat{\xi}(\theta)),$$

where  $\mathbf{Z}$  is the matrix whose columns are  $\mathbf{Z}_{mj}$  and  $\hat{\xi}(\theta)$  is the vector whose elements are  $\hat{\xi}_{mj}(\theta)$ .

In Table 1 we show estimates for several specifications for our model, gradually building up from a simple logit model to the main specification that we outline in Section 3.3. Columns (1) and (2) show estimates from a model without a nest or heterogeneity across consumers—a standard logit model—for which we show both OLS and IV estimates. These specifications also do not yet impose the indirect-inference moments that make use of the ride-hail surcharge policy. As expected, we see that the IV specification increases both the travel time and price elasticity in absolute value, roughly doubling both of them. Specification (3) makes use of the indirect-inference moments, which leads to small changes in the elasticity but a one dollar increase in the value of time.

Starting with specification (4), we incorporate heterogeneity in price sensitivity according to income. Introducing this heterogeneity leads to overall larger price and travel-time sensitivities and increases the value of time. In specification (5) we account for variation in car ownership by allowing for random choice set variation across travelers based on census car ownership data. Specification (6) also introduces a nest for all modes relative to the outside option, which is to not travel.

In the last two columns we present our main specification. We estimate the model from column (6) separately for peak hours—those during which the ride-hail surcharge is active—and non-peak hours. The value of time ranges from \$3.90 to \$45.30 when moving from the lowest to the highest income quintile during peak hours, and from \$3.40 to \$18 during off-peak hours. This large variation emphasizes the importance of accounting

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<sup>40</sup>Note that due to the Box-Cox functional form the price coefficient  $\alpha_p$  and the income-heterogeneity  $\alpha_{py}$  are not additively separable and so price enters only into the person-specific part of utility.

Table 1: Demand Estimation Results

	Pooled						Peak	Off-Peak
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
$\alpha_T$	-1.068 (0.011)	-1.692 (0.022)	-1.535 (0.018)	-2.345 (0.023)	-2.415 (0.023)	-1.928 (0.018)	-1.824 (0.022)	-1.872 (0.027)
$\alpha_p$	-0.058 (0.001)	-0.155 (0.002)	-0.129 (0.001)	-8.461 (0.492)	-3.416 (0.111)	-2.078 (0.09)	-2.388 (0.225)	-1.657 (0.068)
$\alpha_{py}$	.	.	.	-1.262 (0.039)	-0.588 (0.02)	-0.414 (0.022)	-0.696 (0.048)	-0.152 (0.022)
$\rho$	.	.	.	.	0.262 (0.012)	0.376 (0.017)	0.162 (0.017)	.
Estimator	OLS	IV	GMM	GMM	GMM	GMM	GMM	GMM
Policy Moment			✓	✓	✓	✓	✓	✓
Car Ownership				✓	✓	✓	✓	✓
Nest					✓	✓	✓	✓
Avg. VOT	18.41	10.89	11.9	23.88	14.65	13.47	19.47	9.81
VOT (Bot. Quintile)	.	.	.	2.44	3.26	3.62	3.9	3.42
VOT (Top Quintile)	.	.	.	64.24	32.36	27.94	45.32	18.09
Avg. Price Elast.	-0.2	-0.53	-0.44	-0.5	-0.61	-0.65	-0.55	-0.72
Avg. Time Elast.	-0.58	-0.91	-0.83	-1.26	-1.27	-1.29	-1.44	-1.07
M	92,284	92,284	92,284	91,908	91,561	91,561	42,989	48,572
N	281,755	281,755	281,755	281,042	280,185	280,185	136,337	143,848

Notes: This table presents demand estimation results from the specifications outlined in section 4.1. The average VOT is computed by first computing the within market average VOT as the weighted average of  $\alpha_T/\alpha_p^i$  and then averaging across markets, with weights given by market size. Similarly, the average elasticities are computed as the weighted average of own-price and own-time elasticities across all mode-market observations, with weights given by market size. Markets for which consumer incomes are missing are dropped in specifications that include income heterogeneity.

for income heterogeneity. When we weight peak and off-peak hours by trip counts, we find an average value of time of about \$15. Overall, the value of time is therefore quite stable across different specifications, ranging from \$10.89 to \$23.88. Appendix section E.2 presents a number of additional robustness checks and the value of time is again quite stable, ranging from \$8.30 to \$12.78.

We now present some of the key patterns implied by our main estimates. First, we examine how the value of time is distributed spatially. Second, we explore whether the model generates sensible diversion ratios across different modes of transportation.

Figure 7 shows the value of time that we infer for people in Chicago's different Community Areas. For the North Side, characterized by higher incomes, we infer values of

Table 2: Diversion Ratios

		(a) Overall				
From \ To		Bus	Car	Ridehail	Train	Outside
<b>Bus</b>	.	0.29	0.06	0.10	0.56	
<b>Car</b>	0.10	.	0.05	0.08	0.77	
<b>Ridehail</b>	0.14	0.31	.	0.09	0.46	
<b>Train</b>	0.16	0.25	0.07	.	0.52	
		(b) Car				
From \ To		Bus	Car	Ridehail	Train	Outside
<b>Bus</b>	.	0.57	0.03	0.06	0.35	
<b>Car</b>	0.10	.	0.04	0.08	0.78	
<b>Ridehail</b>	0.06	0.59	.	0.05	0.30	
<b>Train</b>	0.07	0.54	0.04	.	0.36	
		(c) No Car				
From \ To		Bus	Car	Ridehail	Train	Outside
<b>Bus</b>	.	.	0.10	0.14	0.76	
<b>Car</b>	.	.	.	.	.	
<b>Ridehail</b>	0.23	.	.	0.13	0.64	
<b>Train</b>	0.23	.	0.10	.	0.66	
		(d) Bottom Income Quintile				
From \ To		Bus	Car	Ridehail	Train	Outside
<b>Bus</b>	.	0.28	0.01	0.10	0.62	
<b>Car</b>	0.10	.	0.00	0.07	0.83	
<b>Ridehail</b>	0.15	0.25	.	0.09	0.51	
<b>Train</b>	0.16	0.25	0.01	.	0.58	
		(e) Top Income Quintile				
From \ To		Bus	Car	Ridehail	Train	Outside
<b>Bus</b>	.	0.43	0.13	0.07	0.36	
<b>Car</b>	0.10	.	0.16	0.07	0.67	
<b>Ridehail</b>	0.10	0.45	.	0.08	0.38	
<b>Train</b>	0.10	0.39	0.15	.	0.36	

*Notes:* These tables present average diversion ratios for various consumer types for the demand estimates in specification (6). To construct them, the diversion ratios are first averaged across markets, weighted by market size. Then, to account for the fact that not every mode is present in every market, they are normalized so that each row sums to 1. Element  $(m, m')$  of each table gives the diversion ratio from mode  $m$  to mode  $m'$ .

time between \$12 and \$18, on average. In the South Side we mostly observe values of time below \$10, with a few exceptions: Midway airport (center left) as well as the neighborhoods of Beverly, Mount Greenwood, and Morgan Park (bottom left), which were popular white-flight destinations during the 1950s and 1960s. These patterns correlate closely with the movement patterns in the bottom panel of Figure 1.

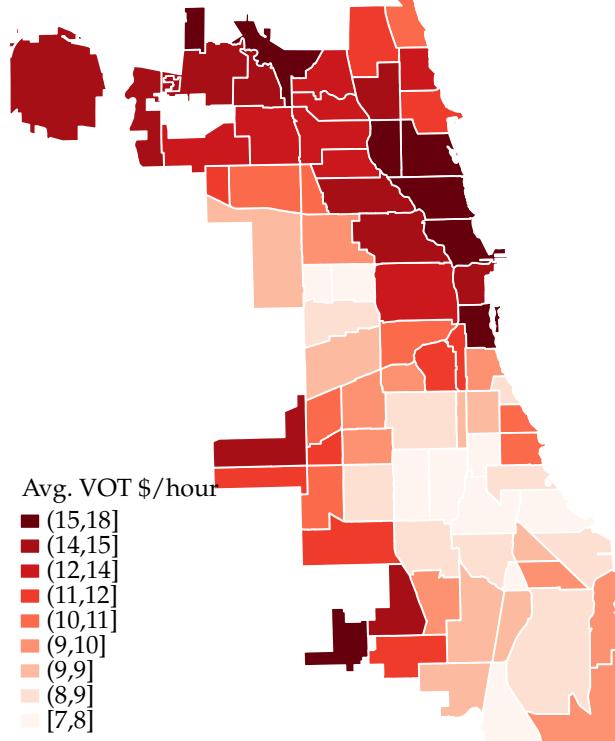


Figure 7: Value of Time across Space

*Notes:* This figure shows the VOT estimates across different regions in the city of Chicago under specification (6) of Table 1, which includes income heterogeneity, car ownership rates, and a nest for all inside goods. Income heterogeneity is driven by differences in the price coefficient in the utility function given by equation (9).

Table 2 presents substitution patterns in the form of diversion ratios. We can see that there is substantially more substitution to the outside option, which is to not travel at all, among travelers who own cars. Furthermore, travelers in the top income quintile substitute more often to cars and ride-hail than those in the bottom income quintile. By contrast, those in the lowest income quintile are more likely to substitute towards buses or the outside option. Among cars owners, car trips are twice as likely to substitute to the outside option as the remaining modes.

## 4.2 Traffic congestion

In this section we show how we estimate the traffic congestion model outlined in Section 3.3.2. Recall that we model time in vehicle for edge  $e$ , at hour  $h$ , for mode  $j$  as follows:

$$T_{ehj}^{\text{vehicle}} = \max\{T_{ej}^0, A_{ehj} \cdot Q_{eh}^{\beta_j}\}, \quad (15)$$

where  $Q_{eh} = \sum_j w_j q_{ehj}$ .<sup>41</sup>

As we can see in Figure 6, observations between 12 am and 5 am overwhelmingly lie in the region where travel times do not depend on traffic. For that reason, we define the free-flow time  $T_{ej}^0$  for mode  $j$  and edge  $e$  to be the average travel time between these early morning hours.

To estimate  $A_{ehj}$  and  $\beta_j$ , we focus on observations in which the time  $T_{ehj}^{\text{vehicle}}$  is above 110% of the free-flow time, which account for 70% of our sample. Since  $T_{ehj}^{\text{vehicle}} \geq T_{ej}^0$  for this subsample, our model becomes  $T_{ehj}^{\text{vehicle}} = A_{ehj} \cdot Q_{eh}^{\beta_j}$ . We take logs and assume that  $a_{ehj} = \log A_{ehj} = a_e + \varepsilon_{ehj}$ . Hence, our estimation equation is

$$\log T_{ehj}^{\text{vehicle}} = a_e + \beta_j \log Q_{eh} + \varepsilon_{ehj}. \quad (16)$$

The edge fixed effect  $a_e$  captures any edge-specific differences in infrastructure that determine travel speed outside of the free-flow region. The remaining error  $\varepsilon_{ehj}$  captures unobservable traffic shocks that vary across hours of the week  $h$  within edge  $e$ .

Tables 3 and 4 present results of regressions of the form (16) for car and bus in-vehicle travel times, respectively. These regressions include edge fixed effects, so the identification assumption is that, within an edge, shocks to the traffic congestion technology are uncorrelated with the number of vehicles. Since we aggregate data at the hour of the week level, the only threat to identification are shocks that repeat themselves every week, such as weather patterns or visibility due to day and night. When we control for such variables (temperature, visibility, and precipitation) in column (2), we estimate a similar elasticity. In column (3), we follow an instrumental variables approach to address the concern that people may re-optimize their route choices due to unobservable but expected traffic shocks, such as planned construction during certain hours of the day. Our

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<sup>41</sup>In our data, we observe edge travel times  $\tau_{ehj}$  as well as mode-specific vehicle flows  $q_{ehj}$ . To construct total-vehicle flow,  $Q_{eh}$ , we assume that  $w_{car} = w_{ride-hail} = 1$  and we calibrate  $w_{bus} = 2$ , a parameter that we take from the [Traffic Modelling Guidelines for London](#).

instrument is the potential number of travellers at the city-wide hourly level. We compute potential travelers by summing the market size across adjacent community areas conditional on the hour of the week. This city-wide market size is exogenous as long as the total demand for travel is driven by daily patterns—commuting to work in the morning, going out for dinner in the afternoon, etc.—rather than by unobserved traffic shocks like anticipated construction or public transit disruptions. We find elasticities of travel time with respect to traffic flows between 0.05 and 0.07 for buses and between 0.10 to 0.17 for cars. The main elasticities that we use for our model are those in the third column.

Table 3: Effect of road congestion on bus travel times

	<i>Dependent variable:</i> Log travel time for bus		
	(1) Fixed effect	(2) + Weather Controls	(3) + IV
Log Flow	0.054*** (0.004)	0.020*** (0.005)	0.069*** (0.006)
Edge FE	✓	✓	✓
Weather controls		✓	✓
IV			✓
within $R^2$	0.066	0.114	0.083
First-stage F			3229.320
Observations	9401	9401	9401

*Notes:* This table shows the regression estimates for the elastic portion of the congestion function separately for buses. The unit of observation is an edge. The dependent variable is the log of travel times for buses and the key independent variable is the log of all vehicles on the road. The first specification controls for edge fixed effects, the second specification adds weather controls and the third specification uses the potential market size as an instrument for potentially endogenous vehicle flows. Robust standard errors in parentheses. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

### 4.3 Solving for Equilibrium and the Planner's Problem

Before we move on to describe our counterfactuals, we spell out more concretely what an equilibrium for our model entails and how we compute both equilibria and the planner's problem. To remind you, a market equilibrium is a vector of travel times  $\mathbf{T}_m = (T_{1m}, \dots, T_{Jm})$  and quantities  $\mathbf{q}_m = (q_{1m}, \dots, q_{Jm})$  such that demand results under those travel times and travel times result from demand—i.e., equations (2) and (4) hold simultaneously.

Table 4: Effect of road congestion on car travel times

	Dependent variable: Log travel time for car		
	(1) Fixed effect	(2) + Weather Controls	(3) + IV
Log Flow	0.128*** (0.005)	0.100*** (0.005)	0.168*** (0.004)
Edge FE	✓	✓	✓
Weather controls		✓	✓
IV			✓
within $R^2$	0.411	0.529	0.444
First-stage F			4488.596
Observations	11739	11739	11739

*Notes:* This table shows the regression estimates for the elastic portion of the congestion function separately for cars. The unit of observation is an edge. The dependent variable is the log of travel times for cars and the key independent variable is the log of all vehicles on the road. The first specification controls for edge fixed effects, the second specification adds weather controls and the third specification uses the potential market size as an instrument for potentially endogenous vehicle flows. Robust standard errors in parentheses. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

**Definition 2.** A transportation equilibrium is a vector of quantities  $\{q_{mj}\}_{mj}$  and a vector of travel times  $\{T_{mj}\}_{mj}$  such that:

1. Quantities are determined from the demand model:  $q_{mj} = \lambda_m \cdot \mathbb{P}_{mj}(T_{mj})$  given by equation 9,  $\forall j \in \mathcal{J}, m \in \mathcal{M}$ .
2. Travel flows across adjacent edges  $e$ ,  $q_{ehj}$ , are determined as the sum of all travelers going through edge  $e$  following equation 12,  $\forall j \in \mathcal{J}, e \in \mathcal{E}$ .
3. Total travel times  $\forall j \in \mathcal{J}, m \in \mathcal{M}$  are the sum of three components  $T_{mj} = T_{mj}^{walk} + T_{mj}^{wait} + T_{mj}^{vehicle}$ , where
  - (a) Time in vehicle for road based modes result from the congestion function:  $T_{mj}^{vehicle} = \sum_{e \in r_{mj}} c_{ehj}(q_{eh, bus}, q_{eh, car})$  as determined by equation 13.
  - (b) Wait times for buses and trains are given by  $\frac{1+\omega^2}{2 \cdot k_{mj}}$ , where  $k_{mj}$  is the number of buses and trains running at the hour. Wait time for cars are zero.
  - (c) Walk times are fixed for all modes.

In words, travellers take as given travel times, prices, and wait times for buses, taxis, and ride-hailing when making their travel decisions. These decisions generate flows

across community areas and therefore congestion at a given edge, which is a function of all the vehicle flows through an edge. Road-based travel times have to be consistent with edge level congestion. Wait times for public transit have to be consistent with the fleet size of public transit and wait times for for hire vehicles with the driver movements. Driver movements in turn have to be consistent with consumer demand and road based travel times. The main computational challenge arises from requirements 2 and 5 in this list: given the linkages of markets through the congestion technology and driver movements, an equilibrium requires the plans of agents in thousands of markets to be consistent with each other.

To find an equilibrium, we write the two conditions equations (2) and (4) as a fixed point problem with prices  $p$  and capacities  $k$ . Let  $f^{p,k}(q) = q(p, T(q, k))$ . The problem of finding an equilibrium is equivalent to finding a fixed point of  $f^{p,k}$ . In words,  $q^*$  is an equilibrium if people would demand  $q^*$  trips when (wait and in-vehicle) times are those that arise when  $q^*$  trips take place. Naive algorithms to find fixed points—such as simple or damped fixed point iteration—often diverge. We thus rewrite the fixed point problem as a root-finding problem (i.e., finding a root  $f^{p,k}(q) - q = 0$ ) and we use a limited-memory version of Broyden’s method to solve it. Appendix D.3 describes in detail the algorithm that we use.

Appendix F.1 presents results that show that our model for the status quo resembles the data closely. We present scatterplots of travel times and market shares for each market by mode, and we find that most observations lie close to a 45 degree line.

Once we find an equilibrium, we can compute all quantities that go into the city government’s objective function. To solve the social planner’s problem—which involves a budget constraint—we follow the augmented Lagrangian method (Nocedal and Wright, 2006), where we iteratively maximize problems that approximate the Lagrangian of the main problem until convergence. We provide further details in Appendix D.4.

## 5 Optimal Counterfactual Policy Design

In what follows, we explore counterfactual policy designs based on the optimality conditions that we described previously in equation (5). We first separately analyze the case where the planner only sets public transit prices and frequencies. To quantify the additional distortions due to budget considerations, we also consider the case of a budget

constrained planner who cannot exceed its status quo deficit. We then separately analyze the effect of *Road Pricing*. To explore whether these two different sets of policies act as complements or substitutes in the planner's goal to maximize welfare, we next analyze the case where the planner has all of these policy tools available at once (*Transit + Road pricing*). It turns out that in this case the budget constraint is never binding and we therefore drop it. Lastly, as a benchmark, we explore the case where the planner also controls the price of ride-hailing services (*Social Planner*).

Table 5 presents a broad overview of the counterfactual results. We report results relative to the status quo, which we present in the first column. Figure 8 shows the resulting substitution patterns.

Table 5: Counterfactual Results

		Status quo	Transit	Transit, budget	Road pricing	Transit + Road pricing	Social planner
		(1)	(2)	(3)	(4)	(5)	(6)
Avg. Price (\$)	Bus	1.09	-0.33	0.44	1.09	0.07	-0.01
	Train	1.33	-0.45	0.66	1.33	0.10	-0.05
	Ride-hailing	9.10	9.10	9.10	9.10	9.10	8.09
Road Price (\$/km)		0	0	0	0.37	0.34	0.31
$\Delta$ Frequency	Bus	0%	-19.55%	-28.4%	0%	-19.06%	-18.22%
	Train	0%	19.6%	9.17%	0%	19.24%	20.3%
Avg. Wait (min)	Bus	6.47	7.87	8.61	6.57	7.92	7.84
	Train	4.29	3.93	4.10	4.41	3.99	3.96
$\Delta$ Welfare (\$M/week)		0	2.16	1.45	4.19	5.60	6.00
$\Delta$ CS (\$M/week)		0	10.37	0.74	-25.12	-16.20	-12.20
$\Delta$ City Surplus (\$M/week)		0	-8.56	0	24.91	17.40	14.87
$\Delta$ Transit Surplus (\$M/week)		0	-8.56	0	0.66	-4.94	-6.04
Taxes (\$M/week)		0	0	0	24.25	22.34	20.91
Externalities (\$M/week)		17.26	16.45	16.57	14.22	13.81	14.04

*Notes:* This table presents the changes in prices, frequencies, and welfare relative to the status quo (column 1) across different counterfactual scenarios. Column 3 only changes public transit prices and frequencies without budget considerations. Column 3 repeats the same exercise subject to a budget constraint. Column 3 changes road pricing. Column 4 combines both. Column 5 also sets optimal ride-hailing prices.

In all counterfactuals where the planner controls public transit, prices for both buses and trains are considerably reduced. The planner also reduces the frequency of buses

and increases the frequency of trains. The reason for this divergence is that buses serve more price sensitive travelers with a lower value of time and are poorly utilized. Trains, on the other hand, are more likely used by higher income travelers with a high value of time. One possible explanation for such inefficiently high bus frequencies in the status quo is that the government puts a larger weight on low-income travelers, who are more reliant on buses. We explore under which welfare weights for different income groups one would obtain the observed frequencies. We find that the government must weight travelers in the lowest income quintile 2.5 times more than travelers in the highest income quintile.<sup>42</sup>

*Transit* raises overall welfare by \$2.16 million per week, whereas *Budget Constrained Transit* only increases it by \$1.45 million per week. Although the budget constraint does not decrease overall welfare by much, it significantly decreases the surplus of travelers. Without the constraint, consumer surplus increases by \$10.4 million per week, but it only increases by \$0.74 million per week when the government must satisfy the budget constraint. This difference arises because, to stay on budget, the government must make two adjustments that hurt travelers: it increases the prices for both modes by about one dollar, and it reduces the frequency of both buses and trains by around ten percentage points. Lastly, while an unconstrained planner reduces the cost of environmental externalities by \$0.81 million per week, a constrained planner only does so by \$0.69 million per week. In terms of the total number of trips taken, we find that when the city does not face a budget constraint, these policies increase the total number of trips taken by 1,150,000 per week and by about 200,000 per week when the city is budget-constrained. The former presents a significant increase given the 17.3 Million trips at baseline.

We now turn to *Road Pricing*. When it is the only policy lever available to the government, the optimal per-km surcharge is 37 cents. The overall welfare gains, \$4.12 million per week, are much larger than from transit policies alone. These welfare gains predominantly result from a reduction in externalities. Travelers, however, are worse off. In the absence of rebates, consumer surplus decreases by a whopping \$25.1 million per week. Even if the government fully rebated the revenue it collected from road pricing, consumers would lose \$0.85 million in weekly surplus. The total number of trips goes down

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<sup>42</sup>Concretely, we assume that the government gives the same weight to its own surplus as to high-income travelers' surplus, and that it gives a weight to travelers in each income quintile that proportional to income to some power  $\phi$ . The status quo frequencies are consistent with  $\phi = -0.642$ , which means that the weights given to the five income quintiles, from lowest to highest, are 2.53, 1.89, 1.59, 1.34, and 1.

by 2.5 million per week (see Figure 8).

Using *Transit + Road Pricing* jointly has advantages over implementing either policy in isolation. By setting road prices at 34 cents per km, the government can generate substantial revenue, effectively eliminating budget considerations. This financial flexibility allows for public transit prices and frequencies that resemble closely those under *Transit*, in which the government faced no budget constraint. Specifically, the government offers nearly free public transit, with bus fares at \$0.07 and train fares at \$0.10. These policies lead to a welfare gain of \$5.60 million per week and a decrease of 1.8 million in weekly trips. As with *Pricing*, consumer surplus decreases without rebates—by \$16.20 million dollars per week. On the other hand, if the government rebates its surplus, consumer surplus can increase by as much as \$1.20 million per week. This joint policy exploits cross-subsidization to price roads efficiently while at the same time benefiting travelers.

As a benchmark, we also explore the case where the planner can set prices for all modes, including ride hailing. Interestingly, in this case the planner wants to reduce the price of ride hailing by around ten percent, which results in a welfare gain of \$6 million per week. This indicates that the markup that these platforms are charging more than offsets the congestion and environmental externalities that they cause. We can thus conclude that, in ride-hailing markets, the market power distortion is stronger than the distortion due to unpriced externalities.

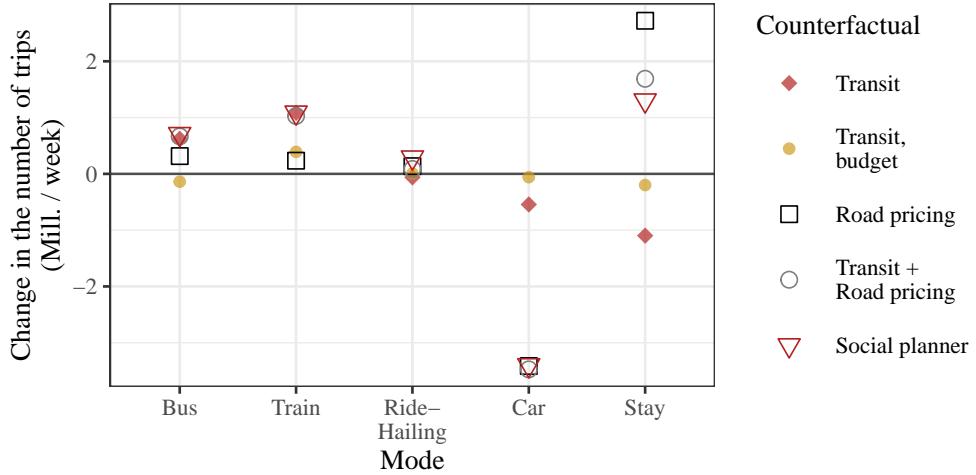


Figure 8: Substitution patterns across different policies

*Notes:* This figure presents changes in the number of trips by mode relative to the status quo under optimal policies across different counterfactual scenarios.

We now decompose the forces that give rise to these results. We present empirical analogues of our theoretical results from Section 3.2, where we decompose optimal prices and frequencies. The exact expressions that we use for this higher-dimensional problem are derived in Appendix C.3. Recall that these prices account for the direct cost as well as environmental and network externalities. In addition, there is a diversion term that arises due to the planner’s desire to internalize externalities of modes not directly under her control. Lastly, there is a term that accounts for budget considerations, akin to a markup term. The budget term summarizes all terms that multiply  $\lambda/(1 + \lambda)$  in equation 6: the market power markup, underweighting externalities, and the Spence distortion.

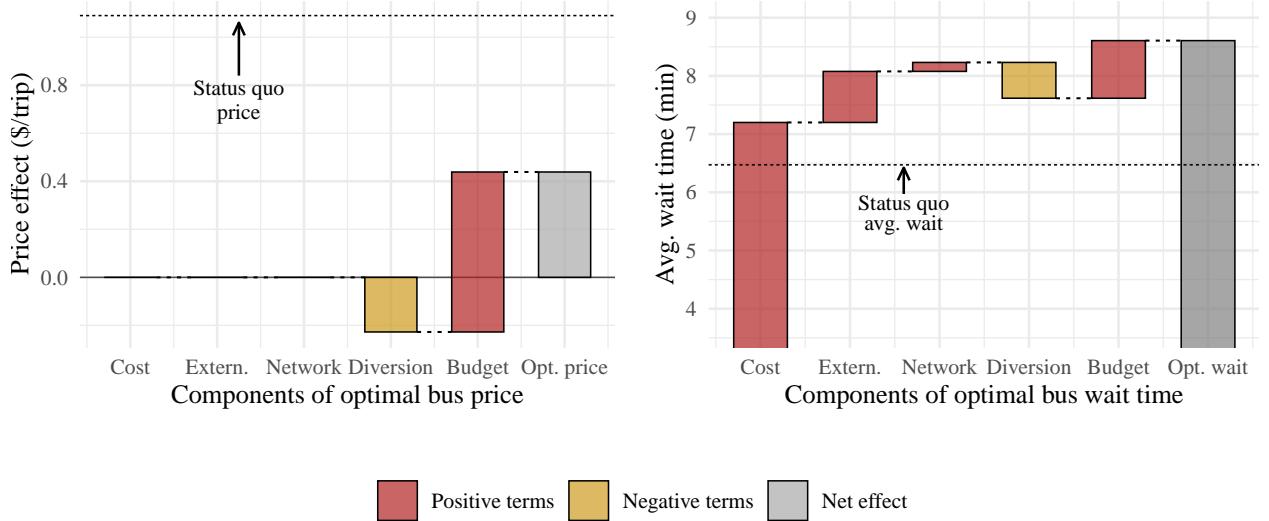


Figure 9: Decomposition of optimal price and waiting times for buses

*Notes:* This graph shows the a decomposition of the optimal prices and travel times for buses corresponding to our theoretical decomposition in Section 3.2. Red bars indicate terms that lead prices and travel times to be higher and yellow bars indicate terms that lead prices to be lower.

Figure 9 shows the decomposition of the optimal prices and travel times for buses under *Budget Constrained Transit* (Appendix F.2 presents a similar figure for trains). In the left panel, we observe that the optimal prices are lower than at the status quo, and that the optimal wait times are higher. If we focus on prices, we can see that the diversion term is negative. This is due to the planner’s inability to choose the prices of other road-based modes of transportation. The price of buses is therefore lowered to divert travelers away from these modes, which are priced too low and do not fully internalize their negative

externalities. This distortion is absent if the planner can also set a road price. Finally, the budget term is the only term that is positive: the only reason why the government should charge for riding a bus is the need to raise revenues to stay on budget.

The right panel shows that the government also decreases the frequency of buses, which increases wait times. The first bar—for "Cost"—shows what the optimal wait time would be if the government only accounted for the tradeoff between the marginal cost of buses and the direct benefit to the travelers riding those buses. Even in that case, the frequency of buses should still decrease so that the average wait time goes from 6.5 minutes in the status quo to 7.2 minutes. The optimal wait time only increases if we account for all the other forces in our model. Interestingly, the provision of additional bus-hours leads to an increase in environmental and congestion externalities, which also pushes towards reducing the frequency of buses. Budget considerations act as an additional force for fewer buses. The only counteracting force is again the planner's motive to divert passengers away from under-priced private modes of transportation, but this effect is relatively small. Consequently, the net effect is an increase in wait times to around 8.6 minutes.

Figure 10 provides a similar decomposition for the optimal price for cars. We observe that the optimal per-km price roughly doubles the status quo price, which is simply the cost of operating a car. About half of the price increase comes from road congestion externalities. Another third of the price increase is due to environmental externalities. The remaining fraction of the price increase arises because of diversion—effectively due to ride hailing, whose price the government has no direct control over. The budget term is zero because the ability to obtain road pricing revenue makes the government's budget constraint non-binding.

Appendix F.3 decomposes the welfare effects in Table 5 into different channels. It shows the extent to which consumer surplus changes due to prices and time. It also shows the extent to which changes in travel times and environmental externalities are driven by fleet size changes or by travelers' mode choices.

**Robustness** Figure 11 shows the extent to which our main results are sensitive to some of the key parameters of our model, focusing on the *Transit + Road pricing* counterfactual. Each panel shows how a 10% increase in several parameters of the model affect the five choice variables of the city government. In the first two panels, we see that the finding that the price of public transit should be close to zero is very robust: the optimal prices

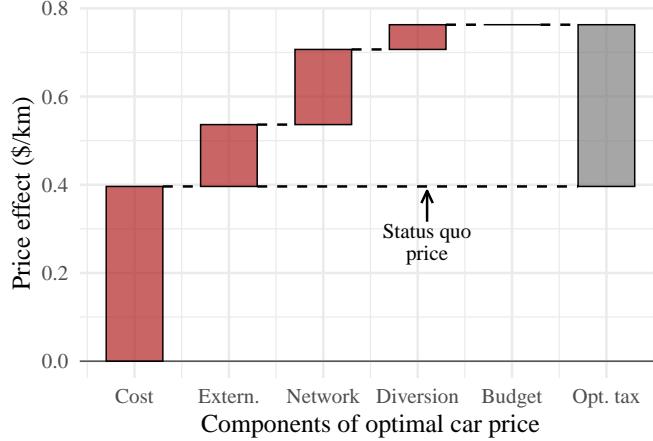


Figure 10: Decomposing the Optimal Price of Operating a Car

*Notes:* This figure shows the price decomposition for cars, following our theoretical derivations in Section 3.2. We follow the convention from previous figures that red bars indicate terms that lead prices to be higher and yellow bars indicate terms that lead prices to be lower; however, all terms in this figure are positive.

are always within 1.5 cents of our baseline results. On the other hand, our results about the optimal waiting time for public transit—in other words, for the frequencies of public transit—are more sensitive to parameters, as can be seen in the next two panels. For four of the six parameters (the marginal cost of public transit, the price and time sensitivity of travelers, the relative disutility of walking and waiting, and the variability of bus arrivals), a 10% change in the value of the parameter results in changes to the optimal bus wait time on the order of half a minute and in the optimal train wait time on the order of 0.2 minutes. These changes correspond to changes in frequencies of around 5%. Finally, the last panel shows that the road price is also quite robust: in every case, the optimal value is within 1.5 cents per km of the baseline value.

One important model parameter over which there is substantial disagreement in the literature is the social cost of carbon. This figure shows that the only result on which it has an important impact is the optimal road price. If instead of the latest recommendation from the EPA (\$190 per tonne of CO<sub>2</sub>) we used the previous figure of \$51 per tonne, the optimal road price would drop by around 9 cents to \$0.25 per km.

**Distributional Effects** We now investigate the distributional effects of these policies and compute consumer surplus changes across income quintiles. Results can be seen on Figure 12, which shows both absolute and percentage changes across the income distri-

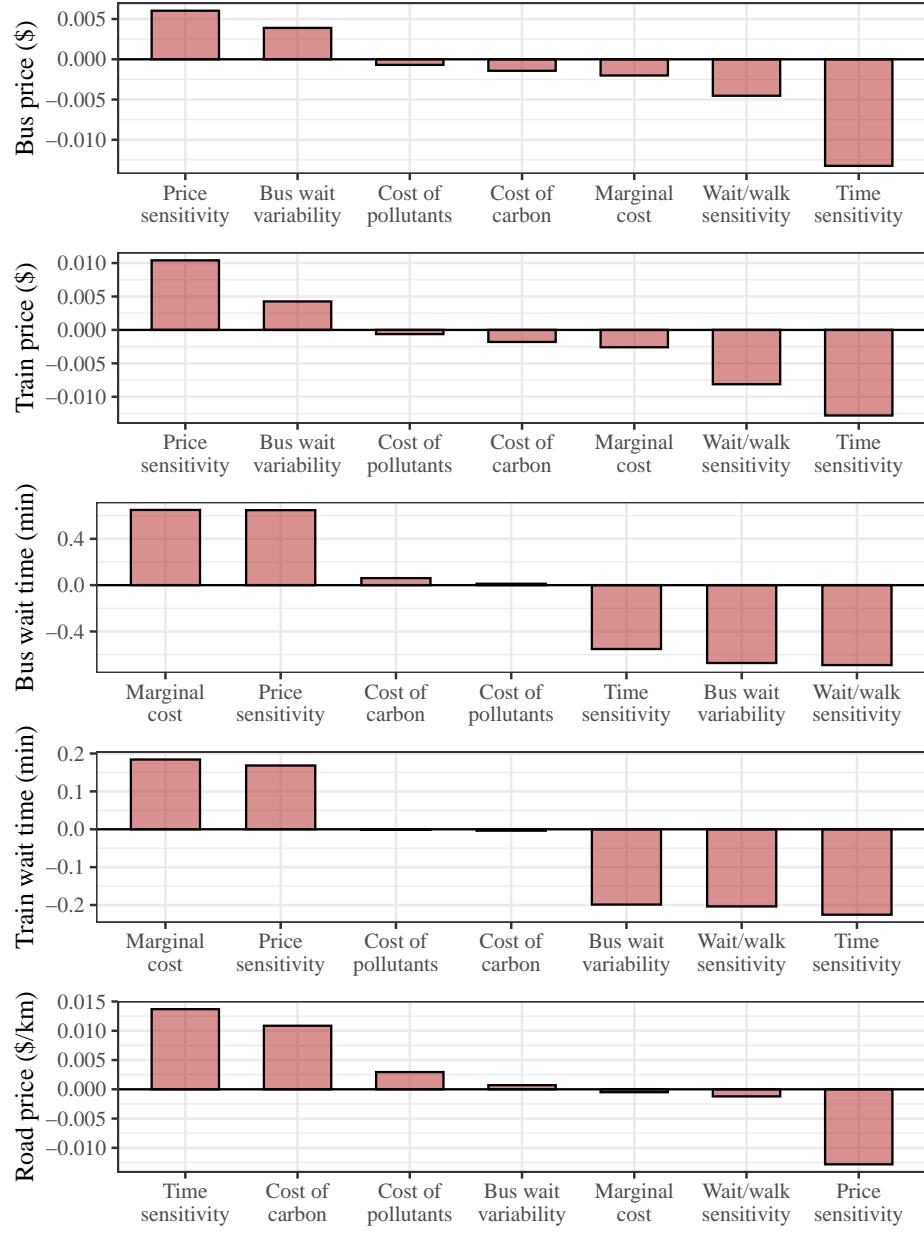


Figure 11: Robustness of counterfactual results

*Notes:* This figure presents how the choice variables of the social planner change in response to changes in some of the model parameters. We focus on the transit + road pricing counterfactual. In each panel, we show how a 10% increase in the model parameter specified in the x-axis affects the choice variable in the x-axis. Within each panel, we order model parameters from the one that causes the largest increase to the one that causes the largest decrease.

bution. Optimal public transit policies in isolation lead to gains in consumer surplus for all travelers except for the highest income group. They are also progressive both in abso-

lute terms and in relative terms. The three remaining policies, which include a congestion surcharge all exhibit a similar pattern. All income groups are now worse off and absolute losses are U-shaped, which is related to the U-shaped pattern in car market shares across income levels as shown in panel (a) of Figure 4. However, when we measure those losses as a percentage of consumer surplus, we find that these policies are highly regressive. To summarize, our counterfactual analysis reveals that the policies that deliver the largest efficiency gains are also the ones that are the most regressive in terms of consumer surplus.

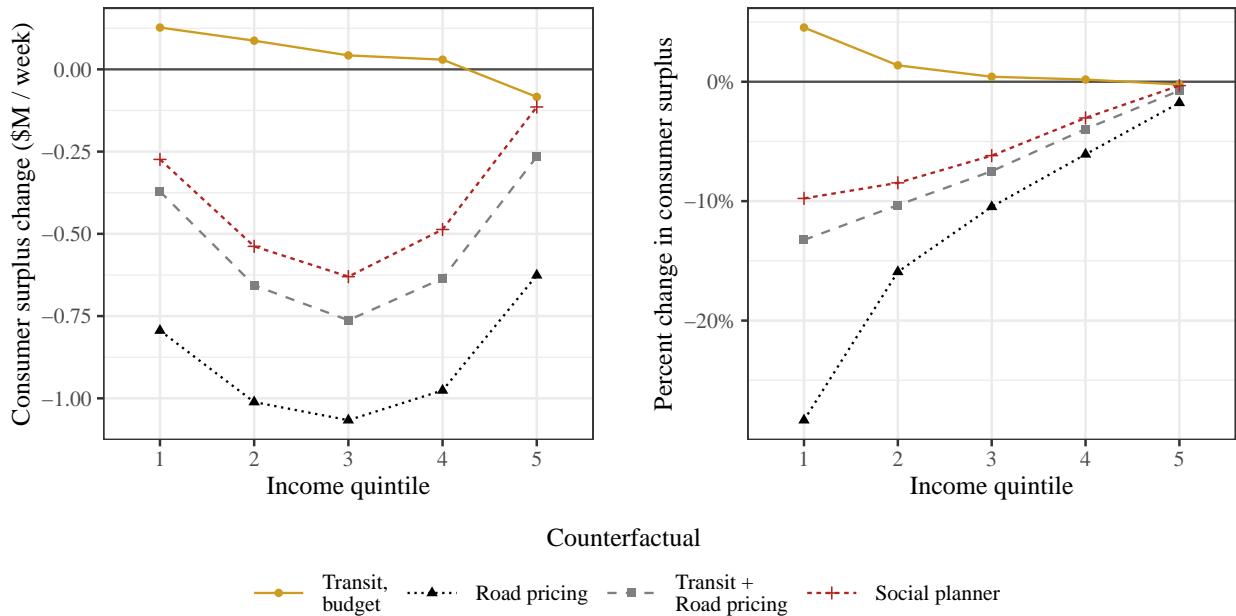


Figure 12: Change in Consumer Surplus Across Income Quintiles

*Notes:* This figure presents changes in consumer surplus relative to the status quo under optimal policies across different counterfactual scenarios. We calculate consumer surplus for five groups according to traveler's income quintiles. Panel (a) displays net changes per trip in dollar values across the five quintiles of the income distribution. Panel (b) displays changes as a percent.

## 5.1 Discussion

Although our model incorporates many ingredients, such as rich heterogeneity in substitution patterns and differences in marginal congestion across vehicle types, we make some simplifying assumptions to keep our model tractable. In this section, we discuss some of these assumptions and their implications.

First, our model does not allow travelers to choose when to travel. Intertemporal substitution is, instead, modeled indirectly: decisions not to travel at certain times are captured as substitution towards the outside option. Although this allows us to model elasticities to own prices and own travel times correctly, we are not able to capture spillover effects between different hours of the week. Although a more realistic model would capture those spillover effects, it would come at the cost of adding substantial complexity to our model. We thus prefer not to model intertemporal substitution directly. Furthermore, recent empirical evidence shows that this simplifying assumption may not have a major impact on our findings. Kreindler (2023) finds that, after accounting for individual constraints and heterogeneity, inter-temporal choices are rather inelastic and peak-spreading policies have a limited impact.

Second, although travelers often decide jointly the mode of transportation for the out-bound and return trips, our model only accounts for decisions over individual trips. This arises from a data limitation: we are not always able to link together two trips from the same rider. Once again, the main challenge this brings to our model is that we are not able to capture spillover effects between different hours of the week.

We think of our model as quantifying short-run responses, before the spatial distribution of the economic activity starts responding to policy interventions. One can thus interpret our model as the bottom layer of an economy that incorporates such long-run responses. Although we believe those longer-run responses are important, prior research has found that the long-run adjustments of residents Barwick *et al.* (2021) and firms (Herzog, 2021) to transportation policy is limited.<sup>43</sup> We defer their investigation for future research.

Finally, routes are fixed in our model. In reality, travelers could adjust routes in response to congestion. However, our spatial areas are relatively large—with a median size of 7 km<sup>2</sup>. Therefore, our model does capture small re-routing deviations, the most common ones in urban transit.

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<sup>43</sup>Part of the reason for this muted response is that the annual costs of these policies are relatively small compared to moving costs. For example, (Bayer *et al.*, 2016) find that within-city residential moving costs are on the order of \$100,000-\$200,000.

## 6 Conclusions

In this paper, we measure the welfare effects of urban transportation policies, and we explore how a budget-constrained planner should choose among a portfolio of policies. Based on a theoretical framework, we find expressions for optimal policies that show that budget considerations introduce additional sources of inefficiency as the planner starts behaving somewhat like a monopolist. We then move on to empirically quantify the welfare effects of different policies in Chicago. We find that the city government should decrease public transit prices, increase the frequency of trains and reduce the frequency of buses, and it should charge a road price on the order of 30 cents per kilometer. We also find that it is important for the government to use these policies jointly, as it can use road pricing revenue to subsidize public transit and achieve higher overall welfare and consumer surplus.

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# A Data Construction

## A.1 Cellphone location records

This subsection details how we construct our sample of trips based on the raw cellphone data. The raw data is composed of a sequence of pings. Each ping contains a timestamp, latitude, longitude, and a device identifier. The final output from this process is a dataset with a fraction of the universe of trips that took place in Chicago. A sequence of filtering steps leaves us with 5% of devices. We verify that the owners of these devices are representative and then scale up the number of trips by a factor such that the aggregate number of car trips is consistent with what is reported by the Chicago Metropolitan Agency for Planning (CMAP) 2019 Household Travel Survey.<sup>44</sup>

### A.1.1 Data filtering

We start by subsetting cellphone pings to a rectangle around the city of Chicago (i.e., latitude between 41.11512 and 42.494693, longitude between -88.706994 and -87.527174) for the months of July 2019 - March 2020 and for the months of June 2020 - August 2020.

Next, using the cellphone device identifier, the timestamp and geolocation of each ping, we calculate the time between two consecutive pings as well as the geodesic distance. These distances allow us to obtain the speed between consecutive pings. We then filter out “noisy” pings by using distance, time, and speed variables. In particular, we remove pings that are moving at an excessive speed since these pings are likely to be GPS “jumps” resulting from noise in the measurement of the GPS coordinates of the device.<sup>45</sup> We also drop “isolated” pings since they are not helpful for identifying whether people are moving. Additionally, we only keep pings belonging to a “stream” of pings.<sup>46</sup> We define a stream of pings as a sequence of pings for the same cellphone identifier such that a ping always has another ping within the next 15 minutes and within 1,000 meters. We drop streams with less than 3 pings. Finally, we aggregate pings to the minute of the day by taking the average location and timestamp across pings within each minute for a given

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<sup>44</sup> Source: My Daily Travel survey (website)

<sup>45</sup> 40 meters per second, i.e. about 145 kilometers per hour

<sup>46</sup> In particular, we only keep pings that satisfy the following two conditions: (i) no more than ten minutes to either the next or the previous ping, (ii) no more than 5,000 meters to either the next or the previous ping.

cellphone identifier. In what follows, we focus on the remaining filtered pings aggregated at the minute level.

### A.1.2 Defining movements, stays, and trips

We identify two consecutive (aggregated) pings as a “movement” for a given cellphone identifier if their distance is at least 50 meters or if their implied speed is at least 3 meters per second (6.7 miles per hour or 10.8 kilometers per hour). We then define a “stay” as a sequence of two or more successive pings with no movement.

Finally, we take all streams of pings and define trips as being a stream (i) with movement, (ii) that starts with a stay, and (iii) that ends with a stay. We remove all trips with a total geodesic trip distance between the starting and ending point below 0.25 miles (about 400 meters).

### A.1.3 Estimation of home locations and traveler’s income

This subsection details how we assign a home location and an income level to each individual cellphone identifier.

**Home locations** We start by assigning all cellphone pings to census block polygons for the subset of pings within Chicago during our sample period.<sup>47</sup> Next, we focus on pings during night hours, defined as between 10pm and 8am, when individuals are more likely to be at home.

Using this subset of pings, we attribute a score system for each hour between 10pm and 8am. Specifically, regardless of the number of pings, scores are assigned as follows:

- A value of 10 to all census blocks that were pinged between 1 am and 5 am.
- A value of 5 to all census blocks that were pinged between 11 pm and 1am or between 5 am and 7 am.
- A value of 2 to all census blocks that were pinged between 10pm and 11pm, or between 7am and 8am.

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<sup>47</sup> See Appendix A.1 for the sample restrictions.

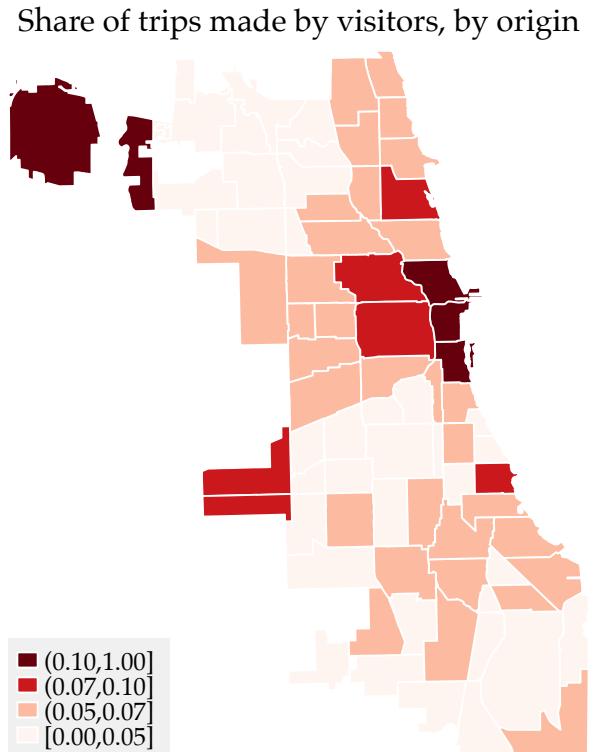


Figure 13: Share of visitors by origin location

*Notes:* This figure shows the share of trips at the origin Community Area level made by visitors. In our cellphone trips data, each market (origin-destination-hour triple) has a share of trips made by visitors. To construct the shares displayed in the figure, we take the weighted average of the share of trips made by visitors across destinations and hours of the week, for each origin Community Area, using inside market size (number of cellphone trips per market) as weight.

The basic idea is to assign a higher score to blocks where the cellphone owner is more likely to be at home. Finally, we sum the scores across all census blocks for each cellphone ID - month combination and keep the census block with highest score. If this highest-score census block appears on at least 3 or more separate nights during the month, we assign it as the cellphone's home census block for that month. Otherwise, we consider the cellphone as having an unknown home location, which we believe captures occasional Chicago visitors such as tourists. Throughout the text, we refer to these devices as *visitors*. Figure 13 plots the share of visitors by origin locations. We see that for trips done by visitors, the most common origin locations are the city center (center right), both airports (top left and center left), as well as Hyde Park the neighborhood home to the University of Chicago (right, south of the center).

**Individual cellphone's income** For all cellphones with an assigned home location, we impute their income by using the census tract median household income.<sup>48</sup>

**Market-level income quintile shares** Next, for each market, we estimate traveler's income distribution.<sup>49</sup> First, we take median income by tracts and divide tracts according to quintiles.<sup>50</sup> Next, assign an income quintile to each device according to their home location. Since we can follow how devices travel across space and over time, for each market we can measure the quintile from each traveler departing from its destination. Finally, for each market, we construct shares of traveler's income quintile. For markets with less than 5 trips, we impute market-level income shares using the underlying distribution of census tract-level income for the origin Community Area of that market.

#### A.1.4 Cellphone data validation

In this section, we provide evidence that the cellphone geo-location data is representative of the population and that it measures traveling patterns with precision.

First, we measure the coverage of our cellphone devices across the income distribution. To do so, for each census tract, we count the number of devices whose home location falls within the tract boundaries. We divide those counts by the census population. Figure 14 displays the share of the tract population covered by the cellphone data. We order tracts by income percentiles. We see that our coverage is fairly constant and hovers 5% across all percentiles of the income distribution.<sup>51</sup> We take this as evidence that our cellphone location records cover a representative sample of the population.

Next, we show that the cellphone location records can accurately represent travelling patterns. To do so, we plot the travel (geodesic) distance distribution for the universe of trips for both the cellphone data and the survey data. The two distributions, plotted in Figure 15, present a high degree of overlap and similarity. We conclude that movements measured using cellphone location records with the procedure described in Section A.1 are a good representations of the travel patterns in the city of Chicago.

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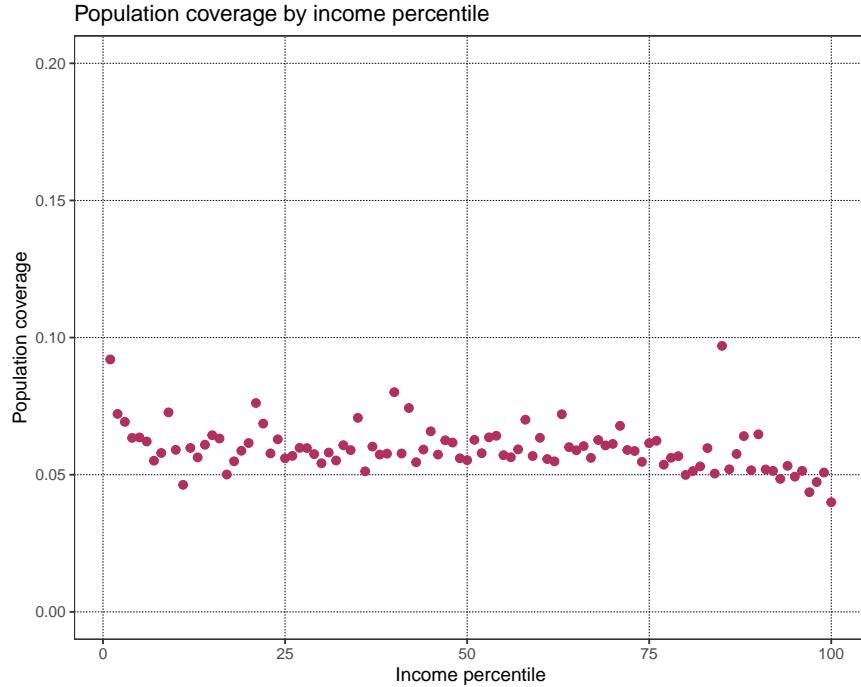
<sup>48</sup> We compute the census-tract median income percentile using the 2010 Census data.

<sup>49</sup> Recall, a market is defined as an (origin Community Area, destination Community Area, hour of the week)-tuple.

<sup>50</sup> For 2010, income quintiles are defined using the following cut-offs: 34,875, 46,261, 60,590 and 85,762 U.S. Dollars.

<sup>51</sup> A perfect representative sample should show a straight horizontal line.

Figure 14: Representativeness across income groups



*Notes:* This figure plots a binscatter of the fraction of the population in each income percentile covered by the mobile phone data. We define the census-tract specific population coverage as the ratio between (i) the number of cellphones whose home location is assigned to that specific census tract, and (ii) the he number of inhabitants of the census tract according to the 2010 Census data. Income percentiles are defined by the census tract median household income.

## A.2 Travel times, routes, and schedules

### A.2.1 Travel times and routes

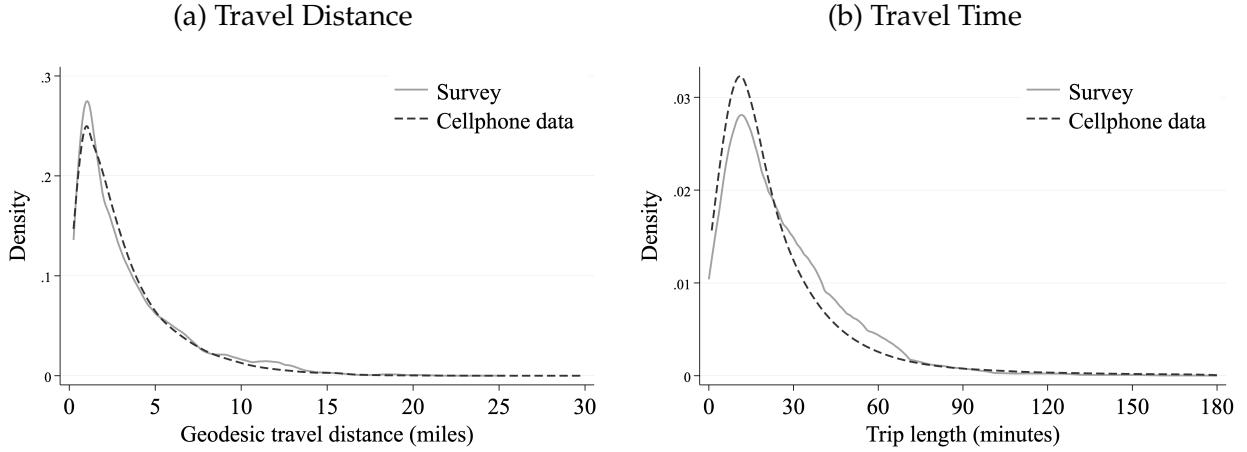
Similar to Akbar *et al.* (2023), we query and geocode trips using Google Maps. For each mode of transportation, we query 30,796,848 counterfactual trips and obtain their distance, duration and path.<sup>52</sup> Importantly, we can measure trip duration for the same origin-destination tuple over the time of the weekday (or weekend) and how this varies with traffic conditions. Moreover, using the detailed “steps” of the public transit Google Maps queries, we obtain walk times from the origin latitude/longitude to the “best” train or bus station.<sup>53</sup>

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<sup>52</sup>One trip for each (origin census tract, destination census tract, hour of day, weekend dummy) combination. We use all the 801 Chicago census tracts boundaries for the year 2010 from the [Chicago Data City portal website](#).

<sup>53</sup>The “best” bus or train station is not necessarily the closest one, depending on the destination and/or the time of the day.

Figure 15: Travel Distance Histograms



*Notes:* This figure plots kernel densities of the distribution of travel distances (Panel a) and travel times (Panel b) using trips in the survey data as well as in the cellphone data. Our level of observation is a trip. Trips in the cellphone data are constructed following the steps in Appendix A.1. Trips in the survey data do not include walking, biking or multi-modal trips.

We also obtain Google Maps data on train trip times by querying Google Maps three times for each pair of train stations in Chicago. These times represented three broad time categories: weekday peak, weekday non-peak, and weekend. In particular, the first query requested a trip time of 8am on Wednesday July 6th, 2022, the second query requested a trip time of 11am on Wednesday July 6th, 2022, and the third query requested a trip time of 11am on Saturday July 9th 2022.

### A.2.2 Public transit schedules

We obtain historical GTFS data from [Open Mobility Data](#). These data contain bus and train schedules for December 2019 through February 2020.

## A.3 Constructing Market Shares

The market shares are constructed using five main sources: (1) Taxi and TNP trips data from the City of Chicago, (2) Google Maps data, (3) cellphone trips data, (4) historical GTFS data containing public transit route schedules, and (5) Chicago public transit data from the MIT Transit Lab and the CTA.

### A.3.1 Raw data processing

#### Taxi and Transportation Network Provider (TNP) data

We obtain trip times, distances, and origin-destination census tracts for both Taxi and Transportation Network Provider (TNP) trips from the [City of Chicago's Data Portal](#).<sup>54</sup>

#### Cellphone trips data

We construct cellphone trips from cellphone pings using the procedure detailed in Appendix A.1. This procedure results in a trip-level dataset. Since our cellphone data only captures a portion of the total trips, we adjust for this by assigning an inflation factor to each trip. To account for varying rates of unobserved trips across different city areas, we allow inflation factors to vary by the neighborhood of the trip's origin.<sup>55</sup> Specifically, we calibrate these factors to ensure that the number of car trips beginning in each neighborhood in our dataset matches the corresponding number in the Chicago Metropolitan Agency for Planning (CMAP) Household Travel Survey.<sup>56</sup>

#### Public transit data

We obtain individual public transit trips via a partnership between the MIT Transit Lab and the CTA. Specifically, we use the near-universe of CTA bus and train trips within the city of Chicago. This data notably excludes trips taken via the Metra, which is a suburban rail system operating in and around Chicago. Metra is managed by a different agency, the Regional Transportation Authority. Each observations corresponds to a passenger swiping in to access the bus or the train station. For buses, we observe the specific bus stop, bus line, and boarding time. For trains, we observe the station and swiping time. Drop-off locations are given to us and imputed following Zhao *et al.* (2007). Unfortunately, these data are missing travel times for train trips, and so we are forced to impute these travel times. To do so, we first match each train trip to the historical GTFS data. To compute the match for a given train trip we first find all scheduled trips between the origin and destination stops of that trip. We then take the match to be the scheduled trip

<sup>54</sup>For privacy reasons, during periods of the day and for locations with very few trips, only the origin and/or destination Community Area of a trip is reported. See [this page](#) for a discussion of the approach to privacy in this data set.

<sup>55</sup>Each neighborhood is a group of about 8-9 Community Areas. The exact make-up of neighborhoods can be found on [Wikipedia](#).

<sup>56</sup>Source: My Daily Travel survey ([website](#))

whose boarding time is closest to the observed boarding time. We then take the scheduled travel time as the travel time. This matching process enables us to compute travel times for close to 90% of train trips.

For trips that have no matches in the schedule data, we impute travel times using Google Maps data.<sup>57</sup> In particular, we first assign each trip one of three time categorizations: weekend (if Saturday or Sunday), peak weekday (if between 5-9:59am or 2-6:59pm on a weekday), or non-peak weekday (otherwise). We then take the time to be the travel time of the matching train trip from the Google Maps data.

As an additional cleaning step, we drop any duplicate trips along with any trips that have missing coordinates, mode information, or time (for bus trips). We also compute travel distances for each trip. We use the Harversine formula to compute distances, with radius equal to 6371.0088, which is the mean radius of Earth in km. For bus trips, we compute the travel distances as the Manhattan distance between the boarding and alighting coordinates, while for train trips we compute the travel distances as the Euclidean distance between the boarding and alighting coordinates. Finally, since there could be measurement error, we account for trips that may be missing in our dataset by assigning each trip an inflation factor. This inflation factor is computed at the day-mode level such that

$$\text{infl}_{dm} T_{dm} = R_{dm},$$

where  $dm$  indexes the day-mode,  $T$  is the total number of observed trips, and  $R$  is the observed ridership, which we obtain from the [City of Chicago's Data Portal](#).

## A.4 Market Share Calculations

We first append together the transit, TNP, taxi, and cellphone trips data. We incorporate walk times to bus/train stations from the Google Maps data. We drop any trips that have a negative trip time, trip time exceeding 6 hours, negative prices, or missing values for origin, destination, distance, duration, mode, trip time, or price. Since our trip data is at the vehicle level, we account for unobserved vehicle occupancy by scaling trip numbers and prices using the average vehicle occupancy for that mode, which we report in appendix D.2.

We calculate market shares at the (origin Community Area, destination Community

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<sup>57</sup> Manual inspection suggests these trips typically involve an unobserved transfer between two lines.

Area, hour-of-the-week) level using a two-step process. First, we aggregate trips at the (origin Community Area, destination Community Area, hour-of-the-week, date) level. We then let the number of car trips be the residual after subtracting public transit, taxi, TNP, and shared trips from the cellphone trips.<sup>58</sup> Car prices are computed as 0.6374 U.S. Dollars per trip mile, which is AAA's estimate of per mile driving costs for an average 2020 model.<sup>59</sup> Finally, we obtain trip counts at the (origin Community Area, destination Community Area, hour-of-the-week, date) level by averaging across dates.

#### A.4.1 Market Size Calibration

To compute market shares, we need to take a stance on the size of the market, which captures how many people could be traveling at a given moment in time. For simplicity, we assume that market sizes are proportional to the total number of observed trips. To calibrate the factor of proportionality, we compare the population of each community area to the total number of trips originating from that community area in the morning hours (5-9:59am) on weekdays. The median ratio across community areas is 2.61. Implicitly, this factor assumes that the number of potential travelers in each community area in these morning hours is given by the total population, which is likely an upper bound. We also compute a more conservative factor by assuming the set of potential travelers is made up of commuters and school-age children, which gives a median factor of 1.48. Corresponding to roughly the midpoint of these two factors, we set our proportionality factor to 2.

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<sup>58</sup>If the residual is negative we assume that there are no car trips.

<sup>59</sup>Source: [AAA brochure "Your driving costs"](#).

## A.5 Sparsity of Our Data versus Survey

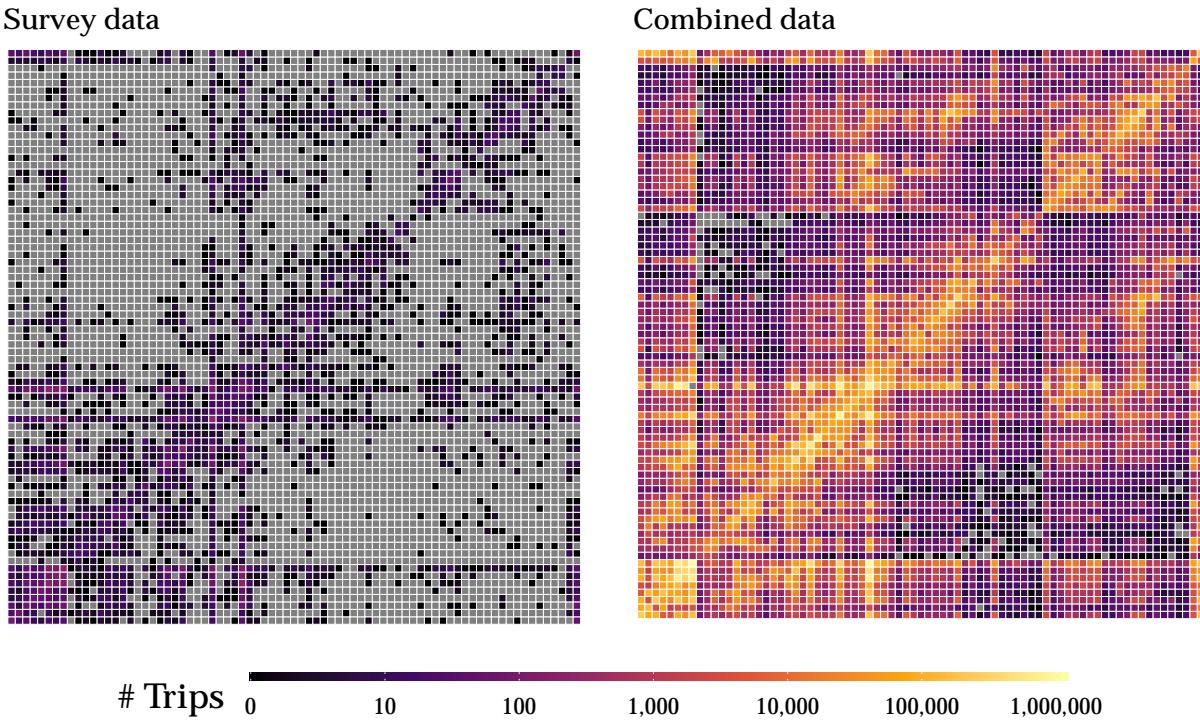


Figure 16: Combined vs. Survey Data: Flows Across Community Areas

*Notes:* These figures show the number of trips from every origin Community Area to every destination Community Area in our combined data (right panel) and in the survey data (left panel). Each row represents an origin Community Area and each column represents a destination Community Area. Grey points represent empty cells.

## B Downtown Surcharge

Effective January 6, 2020, the City of Chicago implemented a new Ground Transportation Tax structure for trips on Transportation Network Providers (TNPs) such as Uber, Lyft, and Via. The tax structure includes a \$0.02 per trip administrative fee and a \$0.10 per trip Accessibility Fund Fee, except for trips on Wheelchair Accessible Vehicles (WAVs).

For single TNP trips, the tax is set at \$1.25 without a Downtown Zone Surcharge and \$3.00 with the surcharge. If a single trip starts or ends in designated special zones, which include airports, Navy Pier, and McCormick Place, the tax is \$6.25 or \$8.00, respectively. For trips on WAVs, the tax is \$0.55, regardless of the zone in which the trip starts or ends.

The Downtown Zone Surcharge applies to any trip that starts or ends within a specified Downtown Zone Area during peak times, which are weekdays between 6 AM and 10 PM. For shared TNP trips, the tax is \$0.65 without the Downtown Zone Surcharge and \$1.25 with it. If a shared trip starts or ends in the designated special zones, the tax is \$5.65 or \$6.25, respectively.

Before January 6, 2020, the surcharge was \$0.72 for all TNC rides, at all times and in all Chicago areas.<sup>60</sup> This implies a \$0.53 basic increase in single rides, a \$0.07 decrease in pooled rides outside the zone, and extra charges in the surcharge zone at peak hours of \$1.75 and \$0.60 for single and pooled rides respectively .

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<sup>60</sup>See <https://abc7chicago.com/uber-lyft-chicago-congestion-tax-taxes/5818233/>

Figure 17: Downtown TNC Surcharge Area



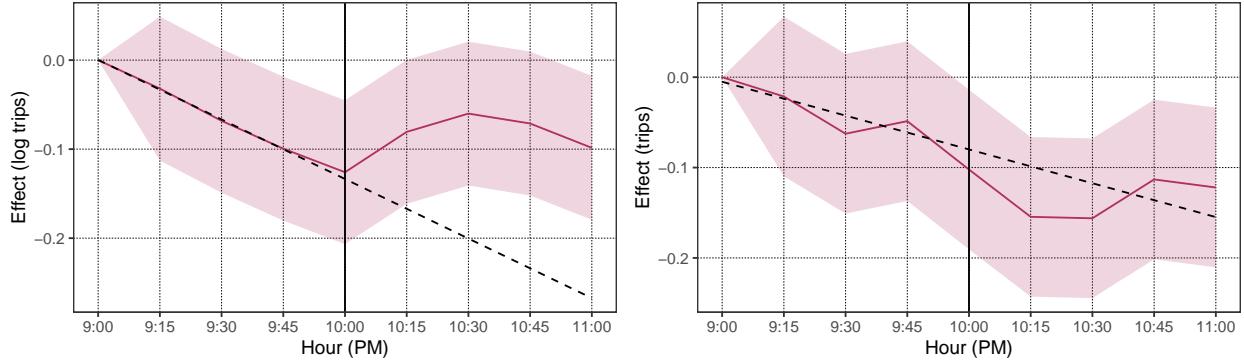
*Notes:* This figure shows the downtown surcharge zone. The surcharge of \$3 applies to any trip that starts or ends within this zone and which are weekdays between 6 AM and 10 PM. For shared TNP trips, the tax is \$0.65 without the Downtown Zone Surcharge and \$1.25 with it.

We use the policy to identify a price elasticity by comparing trips that originate from end in the zone to those that originate from or end in adjacent areas around 10PM, when the policy is phased out. Concretely, our specification is

$$y_{o,d,t} = \mu_{o,d} + \alpha_t + \beta_t \cdot treat_{o,d} + \epsilon_{o,d,t}$$

where  $y_{o,d,t}$  is either log price or log trips,  $o, d$  refers to origin/destination Community

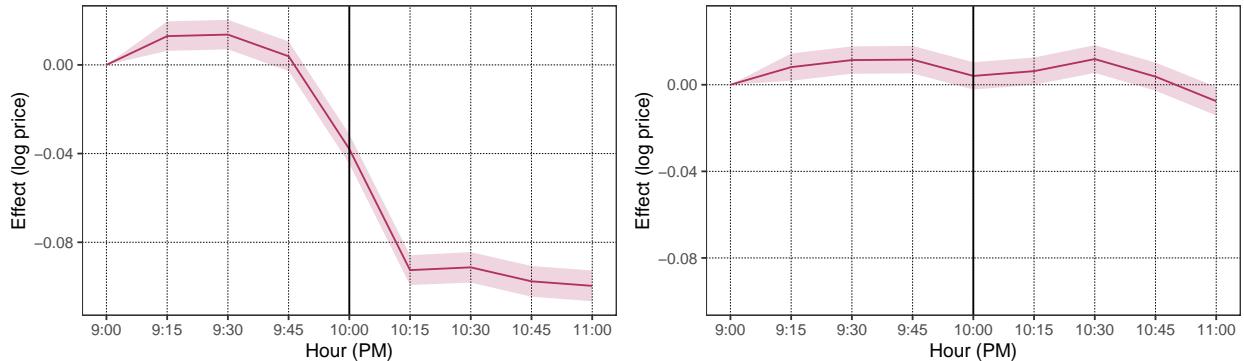
Figure 19: Evening Price Response, Surcharge Zone, 2020 (top) and 2019 (bottom)



*Notes:* The top panel shows how ride hail trips of areas affected by the ride hail surcharge of \$3 relative to unaffected adjacent areas around 10PM, after which the surcharge no longer applies. One can see an increase relative to a downwards trend. The bottom panel shows the same Figure in 2019, when the surcharge policy was not in place yet and we see that the downwards trend continues.

Area. Time  $t$  is measured in 15-min intervals.  $treat_{o,d}$  refers to all trips  $o \rightarrow d$  trips subject to surcharge. We plot the coefficients of these treatment effects in Figure 18 and Figure 19. Taking both estimates, we recover an implied price elasticity of  $-1.42$ .

Figure 18: Evening Price Response, Surcharge Zone, 2020 (top) and 2019 (bottom)



*Notes:* The top panel shows ride hail prices of areas affected by the ride hail surcharge of \$3 relative to unaffected adjacent areas around 10PM, after which the surcharge no longer applies. One can see a drop in prices, which implies an incidence of ... The bottom panel shows the same Figure in 2019, when the surcharge policy was not in place yet. We see that prices are flat.

## C Proofs and Additional Theoretical Results

### C.1 Optimality condition for fleet size

**Proposition 2.** *The first order conditions for the social planner's problem (5) with respect to fleet sizes can be written as*

$$\begin{aligned} \overbrace{u_j^T T_{jj}^k}^{\text{Direct benefit of fleet size}} &= \overbrace{C_j^k}^{\text{Mg. cost of fleet size}} + \overbrace{E_j^k}^{\text{Mg. env. externality of fleet size}} - \underbrace{\sum_l u_l^T \cdot \tilde{T}_{lj}^k}_{\text{Network effects}} + \overbrace{M_j^k}^{\text{Diversion}} + \\ &\quad \frac{\lambda}{1+\lambda} \underbrace{\left( E_j^k + \sum_k (\tilde{u}_k^T - u_k^T) \cdot T_{lj}^k \right)}_{\text{Spence distortion}}. \end{aligned} \quad (17)$$

where  $M_j^k$  and  $\tilde{M}_j^k$  are defined as:

$$\begin{aligned} M_j^k &= \sum_l \frac{\partial q_l}{\partial k_j} \left( C_l^q + E_l^q - \sum_m u_m^T \cdot T_{ml}^q - p_l \right) \\ \tilde{M}_j^k &= \sum_{l \in \mathcal{J}_G} \frac{\partial q_l}{\partial k_j} \left( C_l^q + \sum_{k \in \mathcal{J}_G} q_k \cdot \Omega_{kj} - \sum_m \tilde{u}_m^T \cdot T_{ml}^q - p_l \right) \end{aligned}$$

*Proof.* See Appendix C.2 □

This result takes a very similar form to equation (6). The left hand side is the direct benefit of an increase in the fleet size—on those riders taking that mode—instead of the price (which can be thought of as the direct benefit of an additional trip). The marginal cost, marginal externality, and network effects terms are almost identical, except that they are derivatives with respect to fleet sizes.

The diversion terms follow a similar intuition to those for equation (6). They are, once again, weighted sums of deviations from Pigouvian prices, but the weights are now given by the increase in mode- $l$  trips caused by a change in  $k_j$ . This can be thought of as the mode substitution caused by an increase in the capacity of mode  $j$ .

Finally, the budget causes two monopoly-like distortions: underweighting the environmental externality, and a Spence distortion.

## C.2 Proof of Propositions 1 and 2

*Proof.* The Lagrangian for the social planner's problem is:

$$U(q, T(q, k)) - C(q, k) - E(q, k) - \lambda \left( C(q, k) - \sum_j p_j(q, T(q, k))q_j - B \right)$$

The first order condition for  $q_j$  is:

$$\frac{\partial U}{\partial q_j} + \sum_k \frac{\partial U}{\partial t_k} \frac{\partial T_k}{\partial q_j} - \frac{\partial C}{\partial q_j} - \frac{\partial E}{\partial q_j} + \lambda \left( p_j + \sum_k q_k \frac{dp_k}{dq_j} - \frac{\partial C}{\partial q_j} \right) = 0$$

The first order condition for  $k_j$  is:

$$\sum_k \frac{\partial U}{\partial t_k} \frac{\partial T_k}{\partial k_j} - \frac{\partial C}{\partial k_j} - \frac{\partial E}{\partial k_j} + \lambda \left( \sum_k q_k \frac{dp_k}{dk_j} - \frac{\partial C}{\partial k_j} \right) = 0$$

We now show that  $\partial U / \partial q_j = p_j$ . Let  $\partial\Theta_j(p, t)$  be the boundary between  $\Theta_j(p, t)$  and  $\Theta_0(p, t)$ , and let  $\partial\Theta_{jk}(p, t)$  be the boundary between  $\Theta_j(p, t)$  and  $\Theta_k(p, t)$ . Gross utility can be written as

$$U(q, t) = \int_{\Theta_j(q, t)} u(t_j, \theta) f(\theta) d\theta$$

Using the Leibniz integral rule, we get that

$$\frac{\partial}{\partial q_j} U(q, t) = \sum_k \int_{\partial\Theta_k(q, t)} u_k(t_k, \theta) e_k(\theta) f(\theta) d\theta + \sum_{kl} \int_{\partial\Theta_{kl}(q, t)} u_k(t_j, \theta) e_k(\theta) f(\theta) d\theta$$

(the interior term from the integral rule is zero because  $t$  is fixed), where  $e_k(\theta)$  denotes by how much  $\Theta_k(q, t)$  expands at  $\theta$  as  $q_j$  increases. This can also be written as:

$$\sum_k \int_{\partial\Theta_k(q, t)} u_k(t_k, \theta) e_k(\theta) f(\theta) d\theta + \sum_{k, l > k} \int_{\partial\Theta_{kl}(q, t)} (u_k(t_k, \theta) - u_l(t_l, \theta)) e_l(\theta) f(\theta) d\theta$$

Since agents in the boundaries are indifferent between two choices,  $u_k(t_k, \theta) = p_k$  for the first sum and  $u_k(t_k, \theta) - u_l(t_l, \theta) = p_k - p_l$  for the second sum. After substituting and

rearranging terms, our main expression can be written as:

$$\sum_k p_k \left( \int_{\partial\Theta_k(q,t)} e_k(\theta) f(\theta) d\theta + \sum_l \int_{\partial\Theta_{kl}(q,t)} e_k(\theta) f(\theta) d\theta \right)$$

The term in parentheses is how much  $\Theta_k(p,t)$  expands in total into all other regions, so it is equal to  $\partial q_k / \partial q_j$ . It is thus equal to 1 for  $j$  and 0 for  $k \neq j$ . We can thus conclude that  $\partial U(q,t) / \partial q_j = p_j$ . Substituting  $\partial U / \partial q_j = p_j$ ,  $dp_k / dq_j = \partial p_k / \partial q_j + \partial p_k / \partial t_j \cdot \partial T_k / \partial q_j$ , and  $dp_k / dk_j = \partial p_k / \partial t_j \cdot \partial T_k / \partial k_j$  on the first order conditions and a few steps of algebra yield

$$p_j = \frac{\partial C}{\partial q_j} - \frac{1}{1+\lambda} \sum_k \frac{\partial U}{\partial T_k} \frac{\partial T_k}{\partial q_j} + \frac{1}{1+\lambda} \frac{\partial E}{\partial q_j} - \frac{\lambda}{1+\lambda} \sum_k \left( q_k \frac{\partial p_k}{\partial q_j} + q_k \sum_{k'} \frac{\partial p_k}{\partial T_{k'}} \frac{\partial T_{k'}}{\partial q_j} \right), \quad (18)$$

$$\frac{\partial C}{\partial k_j} = \frac{1}{1+\lambda} \sum_k \frac{\partial U}{\partial T_k} \frac{\partial T_k}{\partial k_j} - \frac{1}{1+\lambda} \frac{\partial E}{\partial k_j} + \frac{\lambda}{1+\lambda} \sum_k q_k \sum_{k'} \frac{\partial p_k}{\partial T_{k'}} \frac{\partial T_{k'}}{\partial k_j}. \quad (19)$$

The final term in equation (18) can be written as

$$\sum_k q_k \sum_{k'} \frac{\partial p_k}{\partial T_{k'}} \frac{\partial T_{k'}}{\partial q_j} = \sum_k \left( \sum_{k'} q_{k'} \frac{\partial p_{k'}}{\partial T_k} \right) \frac{\partial T_k}{\partial q_j}$$

We now show that  $\sum_{k'} q_{k'} \left( \frac{\partial p_{k'}}{\partial T_k} \right)$  can be written as a weighted average of the change in gross utility among marginal travelers. By Leibniz's integral rule,

$$\frac{\partial q_j}{\partial p_j} = -W_j(p,t) - \sum_{k \neq j} W_{jk}(p,t)$$

where

$$W_j(p,t) = \int_{\partial\Theta_j(p,t)} v_j(\theta) \cdot \hat{n}_j(p,t,\theta) f(\theta) d\theta$$

and

$$W_{jk}(p,t) = \int_{\partial\Theta_{jk}(p,t)} v_{jk}(\theta) \cdot \hat{n}_{jk}(p,t,\theta) f(\theta) d\theta$$

are integrals over boundaries  $\partial\Theta_j(p,t)$  and  $\partial\Theta_{jk}(p,t)$ , where the integrand is the density of riders that are willing to switch modes in response to an increase in utility. That density is given by the dot product of  $v_{jk}(\theta)$ , the vector whose elements are the inverse of  $\partial u_j / \partial \theta - \partial u_k / \partial \theta$  (and the inverse of  $\partial u_j / \partial \theta$  for  $v_j(\theta)$ ), and  $\hat{n}_x(p,t,\theta)$ , the unit normal component of

the boundary  $\partial\Theta_x(p, t)$  at  $\theta$ . Also by Leibniz's integral rule,

$$\frac{\partial q_j}{\partial t_j} = V_j(p, t) + \sum_{k \neq j} V_{jk}(p, t)$$

where

$$V_j(p, t) = \int_{\partial\Theta_j(p, t)} \frac{\partial u_j(t, \theta)}{\partial t} v_j(\theta) \cdot \hat{n}_j(p, t, \theta) f(\theta) d\theta$$

and

$$V_{jk}(p, t) = \int_{\partial\Theta_{jk}(p, t)} \frac{\partial u_j(t, \theta)}{\partial t} v_{jk}(\theta) \cdot \hat{n}_{jk}(p, t, \theta) f(\theta) d\theta$$

These are similar integrals as before, only that the integrand is the density of riders that are willing to switch modes in response to an increase in pickup times.

Let  $\Lambda(p, t)$  be the matrix whose  $j$ -th diagonal element is  $W_j(p, t) + \sum_k W_{jk}(p, t)$ , and whose non-diagonal element  $(j, k)$  is  $V_{jk}(p, t)$ . Let  $\Sigma(p, t)$  be a matrix that is defined similarly, but whose elements arise from  $V_{jk}(p, t)$  instead of  $W_{jk}(p, t)$ . Then, by the implicit function theorem, the matrix of derivatives  $\partial p_j / \partial t_k$  is given by

$$\Psi(p, t) = \Lambda^{-1}(p, t) \Sigma(p, t)$$

From the definition of  $W$  and  $V$ , it is clear that this is a weighted average of  $\partial u_j(t, \theta) / \partial t$  over sets of marginal agents. We define

$$\tilde{u}_j^T \equiv \sum_k q_k \frac{\partial p_k}{\partial t_j} = \sum_k q_k \Psi_{kj}$$

the sum of such weighted averages, weighted by the number of agents in each market.

Substituting  $\tilde{u}_j^T \equiv \sum_k q_k \partial p_k / \partial t_j$  and the definitions of  $\epsilon_{jk}^{T,q}$ ,  $\epsilon_{jk}^{T,k}$ ,  $\bar{u}_j^T$ ,  $\Omega_{kj}$ ,  $\tilde{C}_j^q$ , and  $\tilde{E}_j^q$  into expressions (18) and (19) yields expressions (6) and (17).  $\square$

### C.3 Generalizing Propositions 1 and 2 to multiple markets

Consider a city government that faces many markets  $m$ , which can represent people going between different geographical areas at different times. The market can still be described as in Section 3.2, where the vectors  $\mathbf{q}$ ,  $\mathbf{p}$ , and  $\mathbf{t}$  represent quantities, prices, and times for all modes  $j$  and markets  $m$ . The vector of capacities  $\mathbf{k}$  can represent the capacities of

different bus or train routes at different times. We index it by  $r$ .

In this setting, it may not be realistic to think of a government that sets the price of every single mode in every single market and that sets the frequency of every route at every time. We therefore consider coarser policy levers, such as the price of buses for the whole market, the price of trains during rush hour, a per kilometer carbon tax for the whole city, the frequency of one particular bus route, or an overall factor for the frequency with which all trains run.

Consider one such policy lever, which we represent by some parameter  $\sigma$ . The government chooses the level that maximizes its objective function subject to the budget constraint, which can be written as

$$\max_{\sigma} U(\mathbf{q}(\sigma), T(\mathbf{q}(\sigma), \mathbf{k}(\sigma))) - C(\mathbf{q}(\sigma), \mathbf{k}(\sigma)) - E(\mathbf{q}(\sigma), \mathbf{k}(\sigma)) + \lambda \left[ \sum_{mj} p_{mj}(\sigma) q_{mj}(\sigma) - C(\mathbf{q}(\sigma), \mathbf{k}(\sigma)) \right], \quad (20)$$

where  $\mathbf{q}(\sigma)$  is taken to be the equilibrium vector of trips.

The first-order condition for this Lagrangian can be written out as

$$0 = \sum_{mj} p_{mj} \frac{dq_{mj}}{d\sigma} + \sum_{nkmj} \frac{\partial U}{\partial t_{nk}} \frac{\partial t_{nk}}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} + \sum_{nkr} \frac{\partial U}{\partial t_{nk}} \frac{\partial t_{nk}}{\partial k_r} \frac{dk_r}{d\sigma} - \sum_{mj} \frac{\partial C}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} - \sum_{mj} \frac{\partial E}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} - \sum_r \frac{\partial C}{\partial k_r} \frac{dk_r}{d\sigma} - \sum_r \frac{\partial E}{\partial k_r} \frac{dk_r}{d\sigma} \quad (21)$$

$$+ \lambda \left\{ \sum_{mjk} q_{nk} \frac{\partial p_{nk}}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} + \sum_{mjk} q_{nk} \frac{\partial p_{nk}}{\partial t_{ol}} \frac{\partial t_{ol}}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} + \sum_{mj} p_{mj} \frac{dq_{mj}}{d\sigma} - \sum_{mj} \frac{\partial C}{\partial q_{mj}} \frac{dq_{mj}}{d\sigma} - \sum_r \frac{\partial C}{\partial k_r} \frac{dk_r}{d\sigma} \right\} \quad (22)$$

Suppose that  $\sigma$  is a price instrument, in which case  $\frac{dk_r}{d\sigma}$  is equal to zero for all  $r$ . Then, after some algebra, this first order condition can be written as

$$p_j^\sigma = C_j^\sigma + E_j^\sigma - U_j^{A,\sigma} + M_j^{W,\sigma} + \frac{\lambda}{1+\lambda} \{ \mu_j^\sigma - E_j^\sigma - \Delta U_j^\sigma + \Delta M_j^\sigma \} \quad (23)$$

where

$$\begin{aligned}
w_{mj}^\sigma &= \frac{\frac{dq_{mj}}{d\sigma}}{\sum_n \frac{dq_{nj}}{d\sigma}} & D_{kj}^\sigma &= \frac{\sum_m \frac{dq_{mk}}{d\sigma}}{\sum_m \frac{dq_{mj}}{d\sigma}} \\
p_j^\sigma &= \sum_m p_{mj} w_{mj}^\sigma & C_j^\sigma &= \sum_m \frac{\partial C}{\partial q_{mj}} w_{mj}^\sigma & E_j^\sigma &= \sum_m \frac{\partial E}{\partial q_{mj}} w_{mj}^\sigma \\
U_j^{A,\sigma} &= \sum_{nkm} \frac{\partial U}{\partial t_{nk}} \frac{\partial t_{nk}}{\partial q_{mj}} w_{mj}^\sigma & U_j^{M,J,\sigma} &= \sum_{mnkol} q_{nk} \Phi_{nkol}^{J,\sigma} \frac{\partial t_{ol}}{\partial q_{mj}} w_{mj}^\sigma & \mu_j^\sigma &= \sum_{nkm} q_{nk} \Omega_{nkmj}^{J,\sigma} w_{mj}^\sigma \\
M_j^{W,\sigma} &= \sum_{k \neq j} D_{kj}^\sigma (C_k^\sigma + E_k^\sigma - U_k^{A,\sigma} - p_k^\sigma) & M_j^{\Pi,\sigma} &= \sum_{k \in J \setminus \{j\}} D_{kj}^\sigma (C_k^\sigma + \mu_k^\sigma - U_k^{M,\sigma} - p_k^\sigma) \\
\Delta U_j^\sigma &= U_j^{M,J,\sigma} - U_j^{A,\sigma} & \Delta M_j^\sigma &= M_j^{\Pi,\sigma} - M_j^{W,\sigma}
\end{aligned}$$

This equation resembles very closely equation (6). The key insight when generalizing it to this setting with many markets and a coarse policy lever  $\sigma$  is that the relevant price, marginal cost, marginal externality, network effects, and diversion ratios are weighted averages of individual-market quantities across markets. The weight given to market  $m$  is  $w_{mj}^\sigma = \frac{\frac{dq_{mj}}{d\sigma}}{\sum_n \frac{dq_{nj}}{d\sigma}}$ : to what extent does a change in  $\sigma$  affect the number of trips in market  $m$ .

For the specific case of a per-km road price, one can find a related expression that describes the road price in per-km form. The key point is to note that the price faced by travelers taking the taxed mode is given by  $p_{mj} = \frac{\partial C}{\partial q_{mj}} + r_{mj}\tau$ , where  $r_{mj}$  is the trip distance and  $\tau$  is the per km road price. One can substitute this expression on the above FOC, isolate  $\tau$ , and do some algebra to write it as follows:

$$\tau = \frac{1}{r_j^\sigma} \left( E_j^\sigma - U_j^{A,\sigma} + M_j^{W,\sigma} \right), \quad (24)$$

where  $r_j^\sigma = \sum_m r_{mj} w_{mj}^\sigma$  is the average distance per trip. This expression takes a standard Pigouvian form, where the optimal price is equal to the average per-km externality plus the average per-km network effects, plus a misallocation term. This expression doesn't account for the budget constraint because it is unlikely to be binding with the optimal road price.

If we now consider a policy lever that does affect  $k$ , we can rewrite the first-order

condition as

$$-\tilde{U}^{A,k,\sigma} = C^{k,\sigma} + E^{k,\sigma} - U^{A,k,\sigma} + M^{W,k,\sigma} + \frac{\lambda}{1+\lambda} \{-E^{k,\sigma} - \Delta U^{k,\sigma} + \Delta M^{k,\sigma}\} \quad (25)$$

where

$$\begin{aligned} C^{k,\sigma} &= \sum_r \frac{\partial C}{\partial k_r} \frac{dk_r}{d\sigma} & E^{k,\sigma} &= \sum_r \frac{\partial E}{\partial k_r} \frac{dk_r}{d\sigma} & U^{A,k,\sigma} &= \sum_{nkr} \frac{\partial U}{\partial t_{nk}} \frac{\partial t_{nk}}{\partial k_r} \frac{dk_r}{d\sigma} \\ \Delta q_k^\sigma &= \sum_m \frac{dq_{mk}}{d\sigma} & \tilde{U}^{M,k,J,\sigma} + U^{M,k,J,\sigma} &= \sum_{nkolr} q_{nk} \Phi_{nkol}^{J,\sigma} \frac{\partial t_{ol}}{\partial k_r} \frac{dk_r}{d\sigma} \\ M^{W,\sigma} &= \sum_{k,m} \Delta q_k^\sigma (C_k^\sigma + E_k^\sigma - U_k^{A,\sigma} - p_k^\sigma) & M^{\Pi,\sigma} &= \sum_{k \in J,m} \Delta q_k^\sigma (C_k^\sigma + \mu_k^\sigma - U_k^{M,\sigma} - p_k^\sigma) \\ \Delta U^{k,\sigma} &= U_j^{M,k,J,\sigma} - U_j^{A,k,\sigma} & \Delta M^{k,\sigma} &= M_j^{\Pi,k,\sigma} - M_j^{W,k,\sigma} \end{aligned}$$

and all other terms are defined as before. As in the main text, we decompose the effects of  $k$  on times into an effect due to waiting  $\tilde{U}^{M,k,J,\sigma}$  and an effect due to in-vehicle time  $U^{M,k,J,\sigma}$ .

Once again, this equation resembles equation (17) very closely. Quantities are also aggregated across markets through a weighted average in which the weight given to market  $m$  is  $w_{mj}^\sigma = \frac{\frac{dq_{mj}}{d\sigma}}{\sum_n \frac{dq_{nj}}{d\sigma}}$ .

## D Model Details

### D.1 Model of Waiting Times for Ride-hailing and Taxis

Consider mode  $j$  (either taxi or ride hailing). Let  $q_{ahj}$  be the number of mode- $j$  trips with origin  $a$  during hour  $h$ , and let  $I_{ahj}$  be the number of drivers working for mode  $j$  that are idle in location  $a$  during this time. We assume that there is a matching technology such that the expected waiting time for riders before their trip starts is given by

$$T_{ahj}^W = A_{aj}^W I_{ahj}^{-\phi_j} \quad (26)$$

$A_{aj}^W$  is a scale factor that measures the overall matching inefficiency for mode  $j$  in location  $a$ . The parameter  $\phi_j$  is an elasticity that determines how quickly waiting times decrease

with the number of idle drivers. This flexible specification nests simple models of matching in taxi and ride hailing markets.<sup>61</sup>

To determine the number of idle drivers in every location, we assume that the distribution of drivers across the city arises from a parsimonious model that captures the spatial dynamics of the market. Let  $L_{hj}$  be the total number of drivers working for mode  $j$  during hour  $h$ . The number of drivers that are busy is given by  $B_{hj} = \sum_{od} T_{odh}^{\text{vehicle}} q_{odhj}$ , where  $T_{odh}^{\text{vehicle}}$  are the travel times from the traffic congestion model, and  $q_{odhj}$  is the number of people taking mode  $j$  from  $o$  to  $d$ . Thus, the total number of idle drivers is given by  $I_{hj} = L_{hj} - B_{hj}$ .

The probability that an idle driver is in location  $a$  during hour  $h$  is given by

$$\frac{\exp(\mu_a + \sum_b B_{ab} F_{hb})}{\sum_{a'} \exp(\mu_{a'} + \sum_b B_{a'b} F_{hb})}, \quad (27)$$

where  $F_{ha} = \sum_b (q_{bahj} - q_{abhj})$  represents the net inflow of mode- $j$  trips into location  $a$ ,  $B_{ab} = \lambda r_{ab}^{-\rho}$  is a factor for each pair of locations  $a$  and  $b$  that decays with the distance  $r_{ab}$  between them. This probability takes the form of a multinomial logit model that depends on two terms. First,  $\mu_a$ , which are fixed effects that capture the fact that drivers tend to work in certain locations of the city. Second,  $\sum_b B_{ab} F_b$ , which models the extent to which idle drivers are more likely to be located near areas where net inflows are high. The latter term is driven by two opposing forces: areas with high net inflows of trips are also areas with a high net inflow of drivers, so they tend to have many idle drivers; however, these areas have an oversupply of drivers so earnings go down, and drivers will try to move away from them.

Putting all these pieces together, the number of idle drivers in every location is given by

$$I_{ahj} = (L_{hj} - B_{hj}) \frac{\exp(\mu_a + \sum_b B_{ab} F_{hb})}{\sum_{a'} \exp(\mu_{a'} + \sum_b B_{a'b} F_{hb})}. \quad (28)$$

This expression, coupled with equation (26), determines the waiting times for taxis and ride hailing.

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<sup>61</sup> In the taxi model in Lagos (2003), for instance,  $\phi_j = 1$ . In the simplest ride-hailing model described by Castillo *et al.* (2018),  $\phi_j = 1/n$  in  $n$ -dimensional space.

**Estimation** We first estimate the parameters  $A_{ahj}^W$  and  $\phi_j$  that map the number of idle drivers into waiting times. Consider Community Area  $a$ . We make the simple assumption that the  $I_{ahj}$  available drivers are distributed homogeneously across  $a$  and that the pickup time conditional on distance is  $t(x) = M_{aj}x^{c_j}$ . That implies that the pickup time has a distribution whose expectation is<sup>62</sup>

$$T_{ahj}^W = M_{aj}\Gamma\left(1 + \frac{c_j}{2}\right)\left(\frac{1}{\pi I_{ahj}}\right)^{\frac{c_j}{2}}. \quad (29)$$

This takes the desired form  $A_{aj}^W I_{ahj}^{-\phi_j}$ , where  $A_{aj}^W = M_{aj}\Gamma\left(1 + \frac{c_j}{2}\right)\left(\frac{1}{\pi}\right)^{\frac{c_j}{2}}$  and  $\phi_j = \frac{c_j}{2}$ .

We obtain  $M_{aj}$  and  $c_j$  from a regression of the log of the travel time on the log of the travel distance for all car trips in our Google Maps dataset originating and ending within the same Community Area, where we include Community Area fixed effects. The main coefficient from this regression is  $c_j = 0.730$  (s.e.=0.0022), and  $M_{aj}$  are the fixed effects that we estimate. Based on those results, we can conclude that  $\phi_j = \frac{c_j}{2} = 0.365$ , and we back out  $A_{ahj}^W$  from the expression above.

We then move on to estimate the parameters of the driver location model ( $\mu_a$ ,  $\lambda$ , and  $\rho$ ). We do not observe drivers directly, but we use Uber data for the average waiting time at the Community Area by hour of the week level—i.e.,  $T_{ahj}^W$ . Inverting equation (29) allows us to compute all values of  $I_{ahj}$ . We can then estimate ( $\mu_a$ ,  $\lambda$ , and  $\rho$ ) by maximum likelihood, based on equation (27). Maximizing this likelihood is not a simple problem since the vector of  $\mu_a$  has 77 elements. We simplify the task by splitting the problem into an inner loop that computes the optimal vector of  $\mu_a$  given  $\lambda$  and  $\rho$  using a contraction mapping, as in Berry *et al.* (1995), and an outer loop that maximizes over  $\lambda$  and  $\rho$ . Table 6 presents our main estimates.

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<sup>62</sup>With a density of idle drivers  $I_{ahj}$ , the pdf of the distance to the nearest driver is given by  $2\pi x I_{ahj} e^{-\pi I_{ahj} x^2}$ , a Weibull distribution with parameters  $k = 2$  and  $\lambda = 1/\pi L$ . We integrate the travel time over this density to obtain equation (29).

Table 6: Driver Movement Estimates

	Coefficient	Standard Error
$\lambda$	0.0419	0.00007
$\rho$	-0.1312	0.0101

Notes: Standard errors are computed using a sandwich estimator.

## D.2 Additional Parameters and Assumptions

**Marginal costs:** We take a marginal cost of \$0.35 per km for all car-based modes (including taxis and ride-hailing) from the [AAA cost of driving](#). Taxis and ride-hailing also incur labor costs of \$10 per hour.

We combine several sources to obtain the marginal costs of public transit. For buses, we take the sum of several elements. First, capital costs of \$900,000 per bus that lasts 250,000 miles, which we take compute based on recent [diesel](#) and [electric bus](#) purchases made by Chicago Transit Board. Second, fuel costs of \$3.26 per gallon with a fuel efficiency of 3.38 mpg, which we take from the [National Transit Database \(NTD\)](#)'s records for the CTA in 2020. Third, wages of \$33 per hour, which are also obtained from the NTD, assuming the average driving speed is 20 km/h, times a factor of 2 to account for benefits and the wages of supervisors, schedulers, etc. Finally, we use maintenance costs of \$2.76 per km reported in the NTD. These numbers add up to \$7.528 per km.

We also sum several costs to obtain the marginal costs for trains. First, capital costs of \$11M per train that lasts 2 million miles, based on the [purchase price](#) of the trains operated currently by the CTA and assuming each train has 10 rail cars. The [CTA states](#) that trains last approximately 43 years, make around 15 trips a day and each trip is approximately 12.1 miles on average which provides us with our estimate of lifetime mileage. Second, energy costs that are twice the fuel costs of a bus. Third, wage costs of \$9.06, which we compute as the CTA's energy operating expenses divided by the total mileage of trains. Finally, we estimate maintenance costs to be about \$5.00 per km by dividing the total operational expenses by the total miles travelled by trains using [CTA's 2020 budget](#). These numbers add up to \$18.68 per km.

As a sanity check, we compare these numbers relative to the quantities implied by the [CTA's financial statements](#) for 2019 values. One challenge is that those statements report

operating expenses, some of which are not marginal costs (such as the wages of staff), and they do not account for the cost of capital. Nevertheless we can obtain upper and lower bounds for marginal costs. These statements imply that the marginal cost of buses is between \$5.17 and \$12.51 per km, and the marginal costs of trains is between \$9.07 and \$40.38 per km. Reassuringly, both ranges include the values we use in our model.

**Environmental externalities:** For the social cost of carbon, we use the latest EPA proposal of \$190 per tonne as the baseline number.<sup>63</sup> For local pollutants, we obtain estimates based on Holland *et al.* (2016). We use their findings for Cook County to obtain a cost of 44.93 cents per gallon of gasoline and a cost of 41.32 cents per gallon of diesel fuel. They provide damages incurred due to use of different vehicle types in the United States based on pollutants emitted by vehicle type. We aggregate these values for Cook County, weighing vehicle types by the number of miles travelled. For gasoline-related damages, we restrict the sample to non-truck vehicles, and for diesel-related damages, we use the sample of diesel-only trucks.

**Vehicle occupancy:** We take the average occupancy of private cars to be 1.5 people following the estimates for Chicago in Krile *et al.* (2019), and the average occupancy of ride-hailing and taxi trips to be 1.3 passengers following Hou *et al.* (2020).

### D.3 Equilibrium computation

Given prices and capacities  $(\mathbf{p}, \mathbf{k})$ , an equilibrium is a set  $(\mathbf{q}, \mathbf{t})$  that satisfies  $\mathbf{q} = q(\mathbf{p}, \mathbf{t})$  and  $\mathbf{t} = T(\mathbf{q}, \mathbf{k})$ , the demand and transportation technology equations. By plugging in the technology equation in the demand equation, the equilibrium condition can alternatively be written as  $\mathbf{q} = q(\mathbf{p}, T(\mathbf{q}, \mathbf{k}))$ . Thus, if we define the function  $f^{\mathbf{p}, \mathbf{k}}(\mathbf{q}) = q(\mathbf{p}, T(\mathbf{q}, \mathbf{k}))$ , an equilibrium is characterized by a vector of flows  $\mathbf{q}^{\mathbf{p}, \mathbf{k}}$  that is a fixed point of  $f^{\mathbf{p}, \mathbf{k}}$ . After finding a fixed point, the equilibrium vector of travel times can then be computed as  $\mathbf{t}^{\mathbf{p}, \mathbf{k}} = T(\mathbf{q}^{\mathbf{p}, \mathbf{k}}, \mathbf{k})$ .

One naive way to search for an equilibrium is by fixed point iteration. However, this procedure typically diverges. We, instead, find a root of  $f^{\mathbf{p}, \mathbf{k}}(\mathbf{q}) - \mathbf{q} = 0$  using a limited-memory version of Broyden's method. We use the actual vector of trips in the data as the

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<sup>63</sup>See [EPA Issues Supplemental Proposal to Reduce Methane and Other Harmful Pollution from Oil and Natural Gas Operations](#).

initial point, and we use an identity matrix as the initial guess for the Jacobian. The full Broyden algorithm is:

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**Algorithm 1** Equilibrium computation using Broyden's method

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Set initial value of trips  $\mathbf{q}$ .

Compute initial times  $\mathbf{t} = T(\mathbf{q}, \mathbf{k})$ .

Compute deviation  $\mathbf{d} = q(\mathbf{p}, \mathbf{t}) - \mathbf{q}$ .

Set new vector of trips  $\mathbf{q}' = \mathbf{q} + \gamma \mathbf{d}$  for a small step size  $\gamma > 0$ .

Compute new vector of times  $\mathbf{t}' = T(\mathbf{q}', \mathbf{k})$ .

Compute deviation  $\mathbf{d}' = q(\mathbf{p}, \mathbf{t}') - \mathbf{q}'$ .

Set initial approximation to inverse Jacobian  $\mathbf{A} = \mathbf{I}$ .

**while**  $\|\mathbf{d}'\| > \text{tolerance}$  **do**

Define differences  $\Delta \mathbf{q} = \mathbf{q}' - \mathbf{q}$  and  $\Delta \mathbf{d} = \mathbf{d}' - \mathbf{d}$ .

Update vectors of trips  $\mathbf{q} = \mathbf{q}'$  and deviation  $\mathbf{d} = \mathbf{d}'$ .

Compute new approximation to inverse Jacobian  $\mathbf{A} = \mathbf{A} + \frac{\Delta \mathbf{q} - \mathbf{A} \Delta \mathbf{d}}{\Delta \mathbf{q}^T \mathbf{A} \Delta \mathbf{d}} \Delta \mathbf{q}^T \mathbf{A}$ .

Compute new vector of trips  $\mathbf{q}' = \mathbf{q} - \mathbf{A} \mathbf{d}$ .

Compute new vector of times  $\mathbf{t}' = T(\mathbf{q}', \mathbf{k})$ .

Compute new deviation  $\mathbf{d}' = q(\mathbf{p}, \mathbf{t}') - \mathbf{q}'$ .

**end**

---

We make two adjustments to the above algorithm. First, we compute the approximation to the inverse Jacobian  $\mathbf{A}$  with the limited-memory approach in Byrd *et al.* (1994). Second, when we compute the new vector  $\mathbf{q}'$  we often obtain an infeasible vector of trips (the number of Uber or taxi drivers is not enough to satisfy demand). Whenever that is the case, we iteratively update  $\mathbf{q}' = \mathbf{q} + 1/2(\mathbf{q}' - \mathbf{q})$  until we get back to a feasible value.

## D.4 Optimization

Based on the procedure we describe in Appendix D.3, which we use to compute an equilibrium, we can compute welfare  $W(\mathbf{p}, \mathbf{t})$  and the net revenue of the city  $\Pi(\mathbf{p}, \mathbf{t})$ . The welfare maximization problem is

$$\max_{\mathbf{p}, \mathbf{t}} W(\mathbf{p}, \mathbf{t}). \quad (30)$$

We solve this problem in two steps. First, we approximate the solution with a Nelder-Mead optimizer, starting from the true prices and capacities and stopping after 100 iterations. Second, we run a quasi-Newton method starting from the Nelder-Mead optimum. This method differs from Newton's method in two ways, both of which greatly reduce

the computational cost of our procedure. First, to avoid computing the Hessian of the objective function, we use the BFGS approximation (Nocedal and Wright, 2006), which only requires computing the gradient. Second, we approximate the gradient with central differences. Every time we compute a finite difference, instead of fully running Broyden’s method until convergence to an equilibrium, we only take a few steps (typically three) starting from the central point, which allows us to obtain a good approximation to the gradient at a small fraction of the computational cost.

The social planner’s problem is

$$\max_{\mathbf{p}, \mathbf{t}} W(\mathbf{p}, \mathbf{t}) \quad \text{s.t.} \quad \Pi(\mathbf{p}, \mathbf{t}) = -B, \quad (31)$$

where  $B$  is the city’s transportation budget. To solve this problem, we use the augmented Lagrangian method. We iteratively solve the following approximation to the Lagrangian:

$$\max_{\mathbf{p}, \mathbf{t}} W(\mathbf{p}, \mathbf{t}) - \lambda_n (\Pi(\mathbf{p}, \mathbf{t}) + B) + \mu_n (\Pi(\mathbf{p}, \mathbf{t}) + B)^2. \quad (32)$$

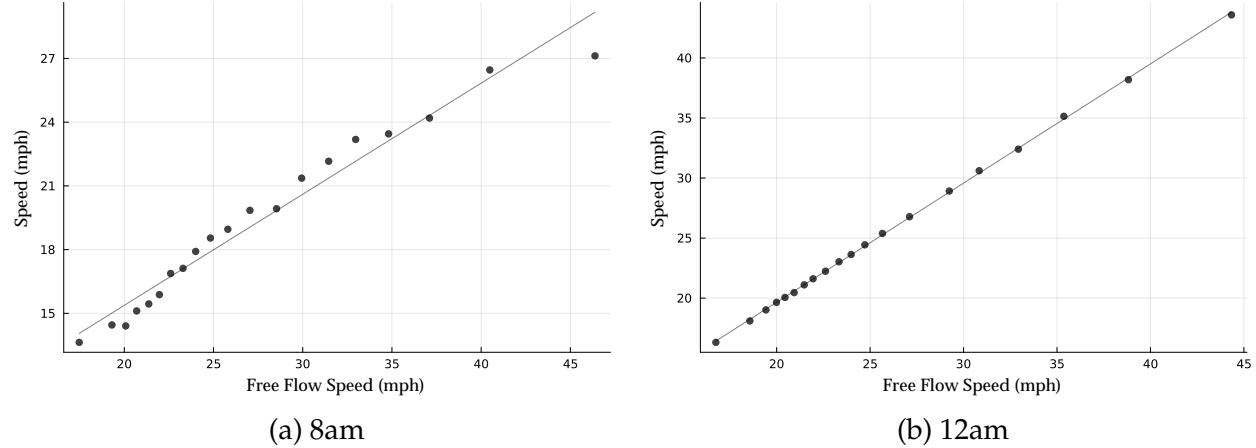
We initialize this iterative procedure by setting  $\mu_0 = 10^{-6}$  and  $\lambda_0 = 0$ . In every step  $n$  we use the method we described above to maximize the objective function, and we set  $\mu_{n+1} = 2\mu_n$  and  $\lambda_{n+1} = \lambda_n + \mu_n(\Pi^n + B)$ , where  $\Pi^n$  is the net revenue at the  $n$ -th step optimum. In this algorithm  $\lambda_n$  converges to the Lagrange multiplier that results in the budget constraint being satisfied with equality (Nocedal and Wright, 2006). This means that (32) converges to the true Lagrangian plus an extra penalty for deviations from the budget constraint—and thus, the sequence of solutions converge to the solution of (31).

## E Additional Empirical Results

### E.1 First-Stage Coefficients

Figure 20 shows a visualization of the first-stage for our free flow instrument. The left panel shows free flow speeds relative to actual travel speeds at 8am for cars. While actual travel speeds are typically lower than free flow speeds due to congestion, there is still a large and positive correlation between the two. That is, markets with high free flow speeds still have relatively higher actual travel speeds even when there is congestion. The right panel additionally shows free flow speeds relative to actual travel speeds at 12am

for cars. Once again there is a strong and positive correlation between the two. Moreover, the fact that these two speeds are very close provides a measure of validation for our free flow speeds, as we would expect there to be very little congestion at this time in most markets. Finally, table 7 shows the first-stage coefficients for the rest of our instruments.



*Notes:* This figure shows a binscatter of car free flow speeds vs. travel speeds across markets at 8am (left panel) and 12am (right panel).

Figure 20: Free Flow Speeds vs. Actual Travel Speeds

Table 7: First-Stage Coefficients

	Time	Price
	(1)	(2)
Free Flow Time	0.970*** (0.007)	-9.228*** (0.090)
Non-TNP Price	0.004*** (0.001)	-0.632*** (0.018)
Frac. Transfers	0.280*** (0.003)	-0.306*** (0.020)
Frac. Multimodal	0.166*** (0.004)	0.597*** (0.029)
Local Diff. x TNP Indic.	-0.014*** (0.001)	-1.445*** (0.016)
Quad. Diff. x TNP Indic.	0.013*** (0.004)	1.946*** (0.081)
Local Diff.	-0.024*** (0.002)	-0.336*** (0.013)
Quad. Diff.	0.055*** (0.006)	3.063*** (0.072)
$\pi^1 \times$ Non-TNP Price	-0.020*** (0.001)	0.691*** (0.021)
$\pi^2 \times$ Non-TNP Price	-0.011*** (0.001)	0.308*** (0.025)
$\pi^3 \times$ Non-TNP Price	-0.013*** (0.001)	0.234*** (0.024)
$\pi^4 \times$ Non-TNP Price	-0.006*** (0.001)	0.076** (0.028)
Mode Fixed Effects	Yes	Yes
Market Fixed Effects	Yes	Yes
<i>F</i>	9,074.944	5,190.973
<i>N</i>	273,833	273,833

*Notes:* This table presents the first-stage coefficients for the instruments used to estimate demand in section 4.1. In particular, we regress times and prices on the full vector of instruments as well as mode and market fixed effects. Singleton observations (markets with only a single mode) are dropped.

## E.2 Demand Robustness

To assess the robustness of our demand estimation results, we estimate a number of additional specifications. Results are shown in Table 8. We first relax the assumption that travelers have utility that is linear in time by including the square of time. The estimated coefficient on the square of time is  $-.312$ , implying that the disutility of travel time is increasing in the length of the trip. In particular, the marginal disutility of the first minute

is about half as much as the marginal disutility of the 60th minute.<sup>64</sup> Measured at the average trip length, the average VOT and time elasticity are both lower than in our main estimates. The average price elasticity is similar. It follows from equation 6 that counterfactuals using this alternative specification would therefore lead to larger frequency reductions and higher congestion prices (to a first order). Therefore, our results should be interpreted as conservative lower bounds on frequency reductions and road prices.

Next we allow for travellers to not only care about average travel times, but also about reliability, in particular for public transit. To do so we include the standard deviation of travel time for public transit modes (train and bus). We find that travelers are relatively insensitive to at least this measure of reliability, and our estimated coefficients imply a similar average VOT, price elasticity, and time elasticity as in our main specification.

In our third robustness specification we additionally allow for heterogeneity in time sensitivity by income. We again adopt a Box-Cox function form:  $\alpha_T^i = \alpha_T + \frac{\alpha_{Ty}}{y_i^{1-\lambda_T}}$ .<sup>65</sup> The estimated coefficients imply a similar average VOT, price sensitivity, and time sensitivity as in our main results. However, the dispersion in VOT is compressed because we estimate that low-income individuals who are more price elastic are also more time elastic. While this would mute the dispersion in distributional consequences that we estimate in our counterfactual results, the results would remain qualitatively unchanged since lower-income individuals still exhibit significantly lower VOT than higher-income individuals.

Finally, we allow for more flexible fixed effects by including a mode-destination fixed effect. This fixed-effect controls for additional unobserved factors that vary at the mode-destination level, including factors such as varying parking costs. We find that once again the estimated average VOT, price elasticity, and time elasticity are similar to our main specification, suggesting such factors are not biasing our estimation.

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<sup>64</sup> Note that for estimation time is measured in hours.

<sup>65</sup> We also include an additional set of instruments that interacts free flow times with indicators for each income quintile.

Table 8: Demand Estimation Robustness

	(1)	(2)	(3)	(4)
$\alpha_T$	-0.574 (0.048)	-1.702 (0.025)	-1.286 (0.028)	-1.929 (0.018)
$\alpha_p$	-2.169 (0.115)	-3.080 (0.158)	-1.173 (0.032)	-1.000 (0.041)
$\alpha_{py}$	-0.508 (0.026)	-0.643 (0.026)	-0.089 (0.014)	-0.022 (0.022)
$\rho$	0.359 (0.012)	0.314 (0.015)	0.335 (0.009)	0.191 (0.010)
$\alpha_{T^2}$	-0.312 (0.014)	.	.	.
$\alpha_{std(T)}$	.	-0.114 (0.046)	.	.
$\alpha_{Ty}$	.	.	-25.611 (2.060)	.
$\lambda_T$	.	.	-1.457 (0.073)	.
Mode FEs	✓	✓	✓	
Mode-Destination FEs				✓
Market FEs	✓	✓	✓	✓
Transfer & Multimodal Controls	✓	✓	✓	✓
Policy Moment	✓	✓	✓	✓
Car Ownership	✓	✓	✓	✓
Nest	✓	✓	✓	✓
Avg. VOT	8.30	13.00	10.82	12.78
VOT (Bot. Quintile)	2.01	2.68	8.98	5.15
VOT (Top Quintile)	17.82	28.86	15.89	22.59
Avg. Price Elast.	-0.64	-0.64	-0.70	-0.63
Avg. Time Elast.	-0.76	-1.14	-1.25	-1.21
M	91,561	74,512	91,561	91,561
N	280,185	222,142	280,185	280,185

*Notes:* This table presents a number of robustness checks for our main specification in section 4.1. The average VOT is computed by first computing the within market average VOT as the weighted average of  $\alpha_T/\alpha_p^i$  and then averaging across markets, with weights given by market size. Similarly, the average elasticities are computed as the weighted average of own-price and own-time elasticities across all mode-market observations, with weights given by market size. In specification (2), markets for which we cannot compute the standard deviation of time are dropped.

### E.3 Bus Utilization

While our model does not consider capacity constraints for buses when solving for the optimal policy, we can consider *ex-post* the extent to which this constraint might bind. Our results imply frequency reductions for buses ranging from approximately 20-30%.

We consider whether these frequency reductions would result in binding capacity constraints, holding ridership levels fixed, by computing the fraction of buses that exceed 70% and 80% utilization across hours of the day. Figure 21 shows that this constraint is unlikely to make a first-order impact on our results as only 10% of buses reach even 70% utilization, and only during the morning and afternoon rush hours.

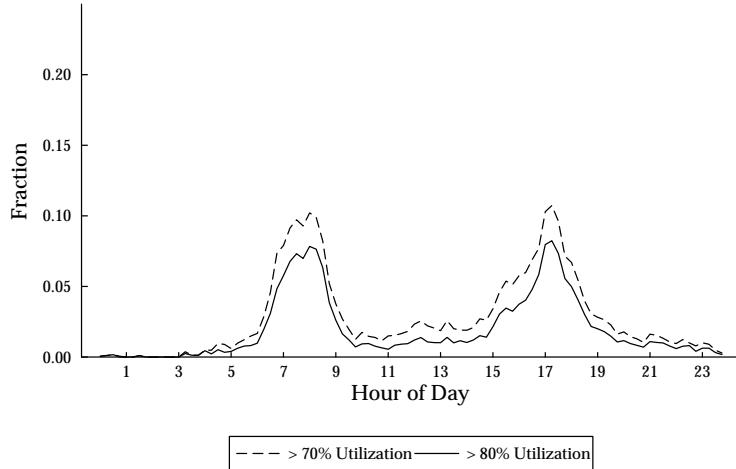


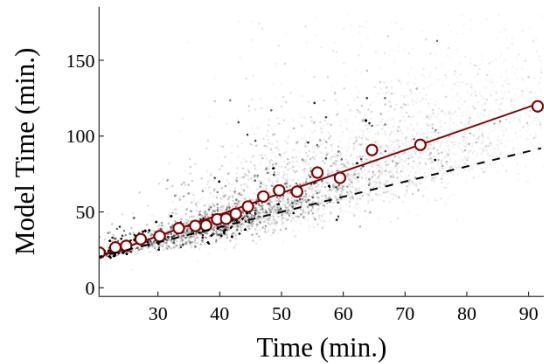
Figure 21: Bus Capacity

*Notes:* This figure shows the fraction of buses that exceed 80% (solid) and 70% (dashed) utilization over the course of the day.

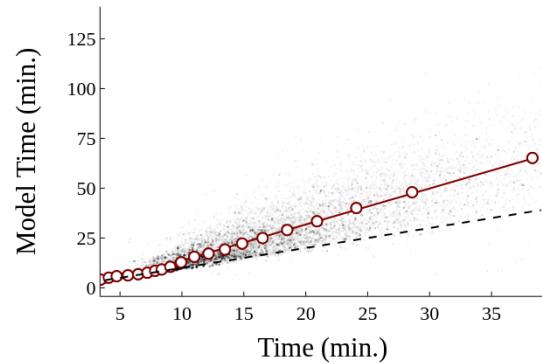
## F Additional Counterfactual Results

### F.1 Model Fit

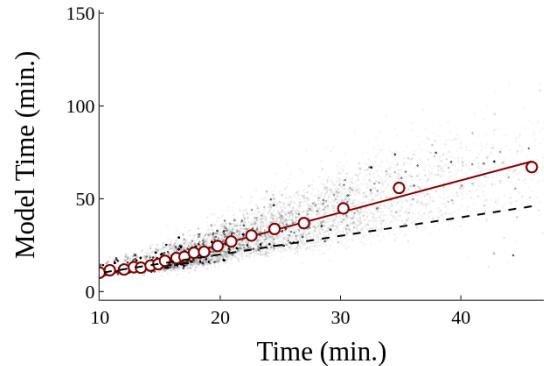
#### F.1.1 Trip Times



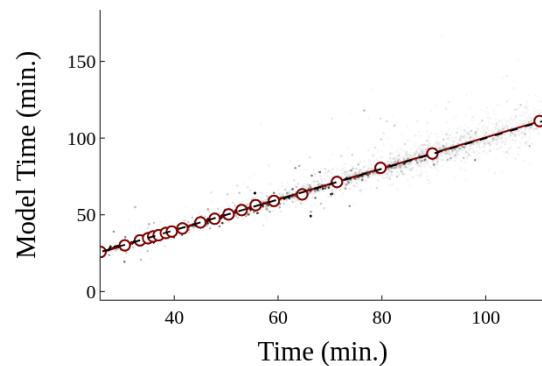
(a) Bus



(b) Car



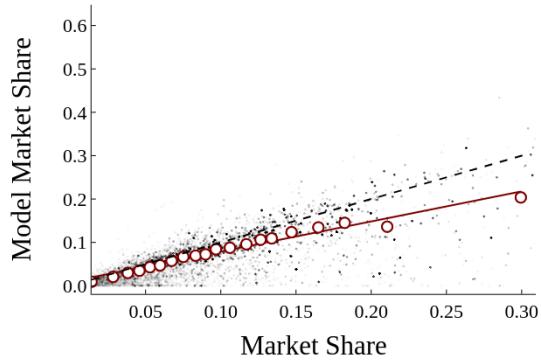
(c) Ride-hail



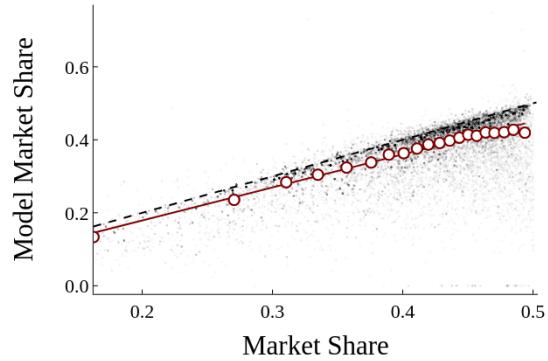
(d) Train

*Notes:* This figure compares observed trip times to model trip times separately for each mode. Each panel displays both a binscatter and a scatterplot of observed vs. model trip times for a sample of 25,000 markets, where markets are drawn randomly with replacement and sample weights are given by trip counts. The dashed line shows the 45 degree line.

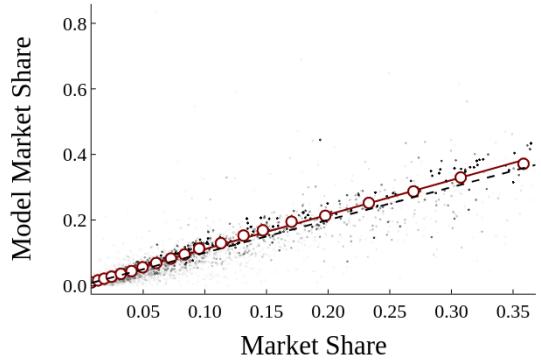
## F.1.2 Market Shares



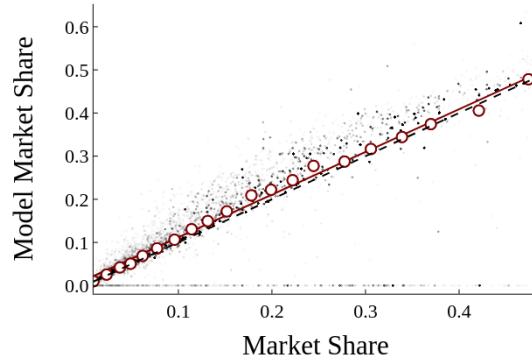
(a) Bus



(b) Car



(c) Ride-hail



(d) Train

*Notes:* This figure compares observed market shares to model market shares separately for each mode. Each panel displays both a binscatter and a scatterplot of observed vs. model trip times for a sample of 25,000 markets, where markets are drawn randomly with replacement and sample weights are given by trip counts. The dashed line shows the 45 degree line.

## F.2 Decomposition of train prices and waiting times

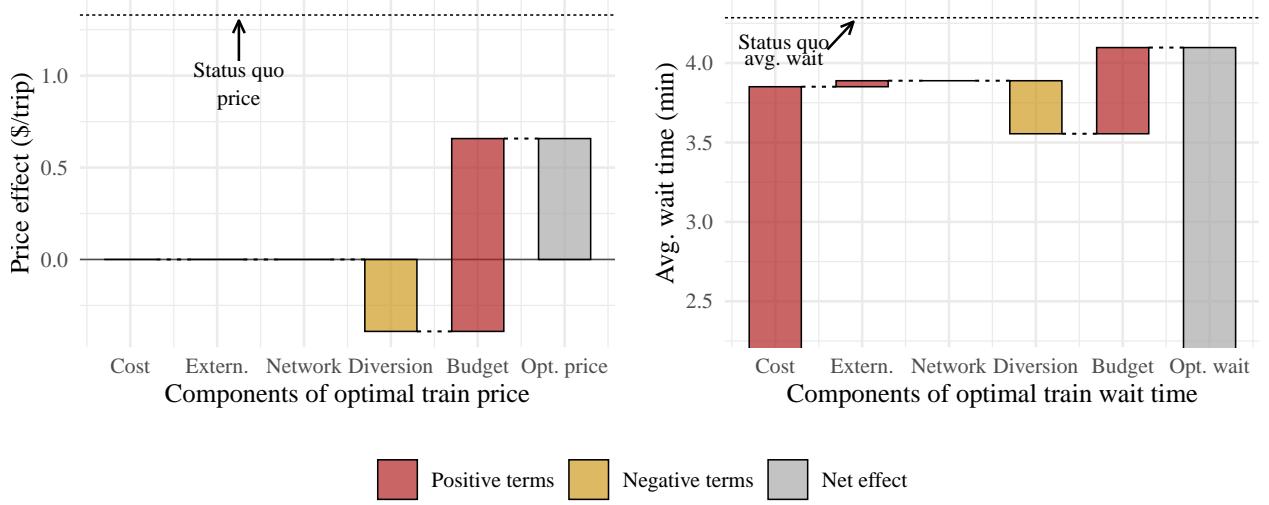


Figure 24: Decomposition of optimal price and waiting times for trains

*Notes:* This graph shows the a decomposition of the optimal prices and travel times for buses corresponding to our theoretical decomposition in Section 5. Red bars indicate terms that lead prices and travel times to be higher and yellow bars indicate terms that lead prices to be lower.

## F.3 Decomposition of welfare effects

In this section, we decompose the change in consumer surplus and in environmental externalities attributed to different channels. The change in consumer surplus is a product of two forces: the direct change in prices and the indirect effect in time due to changes in mode choices. The change in the environmental externalities is also due to two channels: the change in frequencies and the change in travelers' mode choices (substitution). Table 9 shows how each of these channels contribute to the overall aggregate effects across different scenarios.

Focusing on the counterfactual where the planner only sets public transit prices and frequencies subject to a budget constraint, column 3, we see that consumers face two opposing effects. On the one hand, lower prices means an increase in consumer surplus of \$3.8M per week. On the other hand, lower frequencies increase the overall travel times and, in turn, decreases consumer surplus by \$3.4M per week. In terms of externalities,

most of the reduction accrues through the reduction in frequencies and fewer vehicles running throughout the city.

When the planner only set road pricing, we see a large reduction in consumer surplus of \$25.5M per week. The reason is that consumers face an increase of prices for the most common mode of transportation, namely private cars. Because due to this increase in prices, consumers stop traveling by car, traveling speed goes up, which translates into lower overall travel times and an increase in consumer surplus of \$2M per week.

Simultaneously setting public transit prices and frequencies as well as road pricing can be viewed as the combination of the previous two cases. However, in this case we have two opposing effects for both prices and travel times that net each other out in the aggregate overall results.

Finally, when the planner sets all prices and reduces ride-hailing prices by 45%, we see some interactions of these policies accruing through two channels. First, consumer surplus increases by \$14.5M per week relative to the previous scenario. However, as travelers start substituting toward ride-hail, travel speed decreases and overall travel time increases, which partially undoes the effect of congestion of car surcharges.

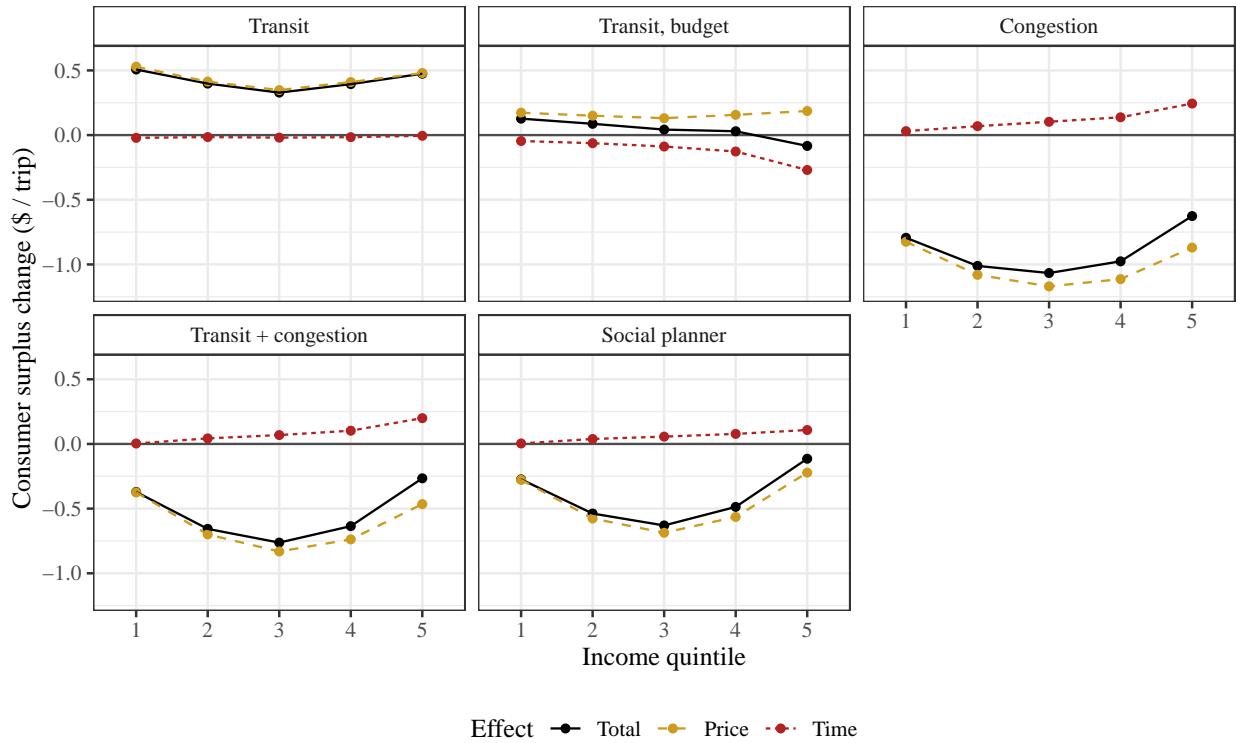
Next, we zoom in on how consumer surplus changes across the income distribution. The results, in percentage terms, can be seen in Figure 25. Observe that the absolute effect of prices for the public transit counterfactuals is larger for lower income consumers, as they are the ones who are more likely to use those modes of transportation. Conversely, the effect of time is more pronounced for higher income consumers, as they are the ones with the highest VOT.

Table 9: Decomposition of Consumer Surplus and Environmental Externalities

	Status quo	Transit	Transit, budget	Road pricing	Transit + Road pricing	Social planner
	Total	0	10.369	0.742	-25.119	-16.198
$\Delta$ CS (\$M/week)	Price	0	10.825	3.973	-28.058	-18.236
	Time	0	-0.455	-3.231	2.939	2.038
	Capacity	0	-1.247	-3.325	0	-1.213
	Substitution	0	0.791	0.094	2.939	3.251
	Total	0	-0.808	-0.686	-3.040	-3.450
$\Delta$ Externality (\$M/week)	Capacity	0	-0.374	-0.526	0	-0.365
	Substitution	0	-0.434	-0.160	-3.040	-3.085
$\Delta$ Average speed (km/h)	0.00%	0.52%	0.16%	2.68%	2.87%	2.59%

*Notes:* This table represent the change in consumer surplus and environmental externalities attributed to different channels. Changes in consumer surplus (first row) are divided into changes in prices (second row) and times (third row). Changes in times are a product in changes in fleet size (fourth row) and substitution of consumers across modes (fifth row). Total changes in externalities (sixth row) are decomposed into changes in fleet size (seventh row) and substitution across consumer (eighth row).

Figure 25: Decomposition of consumer surplus through different channels



*Notes:* These graphs presents changes in consumer surplus across income quintiles for four different counterfactual scenarios scenarios. Each of the lines represent the change in consumer surplus from each of the channels that affect traveler's utility.