

<p><b>EXPERIMENTAL METHODS</b></p> <p><b>treatment/indie variable</b> what we are able to control</p> <p><b>outcome/dependent variable</b> what we observe</p> <p><b>stratification</b> i.e. stratify by demographic vars before splitting into ctrl groups</p> <p><b>double-blind</b> even experimenters don't know who is placebo</p> <p><b>observational study</b> experiment not randomizable</p> <p><b>control variable</b> i.e. income level on education</p> <p><b>confounding variable</b> influences both treatment and outcome variables</p>	<p><b>LINEAR REGRESSION</b></p> <p><b>add first column of 1s</b></p> <pre>X.insert(0, "label", np.ones((100,1)))</pre> <p><b>LSE/Bayes estimator</b></p> <pre>beta = np.linalg.inv(X.T@X)@X.T@y # y is row vector, X is rows for sample</pre> <p><b>WELCH T TEST (<a href="#">ex.</a>)</b></p> <p><i>find features that are significantly different between two sample populations</i></p> <pre>result = stats.ttest_ind(data_aml, data_all, axis=1, equal_var=False, nan_policy='raise', alternative='two-sided')</pre> <p><b>MULTIPLE HYPOTHESES TESTING CORRECTIONS</b></p> <p><b>holm-bonferroni</b></p> <pre>pvalues = sorted(result.pvalue) count = 0 alpha = 0.05</pre> <pre>for i, v in enumerate(pvalues):     if v &lt;= alpha/(len(pvalues) - i):         count += 1</pre> <p>count</p> <p><b>bonferroni-hochberg</b></p> <pre>pvalues = sorted(result.pvalue) alpha = 0.05</pre> <pre>for i, v in enumerate(pvalues):     if v &lt;= (i+1)/len(pvalues) * alpha:         continue     else:         break</pre> <p>print(i)</p> <p><b>T TEST TO FIND UNNECESSARY LIN.REG. PARAMS (<a href="#">theory.</a>)</b></p> <pre>N = len(y) # nb of data points p = len(X[0]) # dimension of a sample feature vector var_hat = (np.linalg.norm(y - X@beta)**2)/(N - p) sigma_hat = np.sqrt(var_hat) beta_cov = np.linalg.inv(X.T@X) # T statistics Ts = beta/(sigma_hat*np.sqrt(np.diagonal(beta_cov))) pvalues = 2*scipy.stats.t.sf(np.abs(Ts), (N - p))</pre>	<p><b>PYTHON MISC</b></p> <p>convert row vector into column vector</p> <pre>Y = Y.reshape(-1,1)</pre> <p><b>sigmoid</b></p> <pre>from scipy.stats import logistic logistic.cdf(x)</pre> <p><b>build feature matrix X from individual arrays</b></p> <pre>X = np.vstack((np.ones((1,30)), A, B)).T</pre> <p><b>PANDAS MISC</b></p> <p><b>get first row</b></p> <pre>df.iloc[0]</pre> <p><b>get first 10 columns</b></p> <pre>df.iloc[:, 0:10]</pre> <p><b>iterate rows</b></p> <pre>for index, row in data.iterrows():     print(row["seconds"])</pre> <p><b>mean (of each row, over columns)</b></p> <pre>df.mean(axis=1)</pre> <p><b>QQ PLOTS</b></p> <p>here we were trying to figure out if we should apply a log transformation to some fields</p> <pre>stats.probplot(data[key], dist="norm", plot=plt) plt.show() stats.probplot(np.log(data[key]), dist="norm", plot=plt) plt.show()</pre> <p><b>GRADIENT DESCENT</b></p> <pre>def gradient_descent(X, y, step_size, precision):     beta = np.zeros((X.shape[1], 1))     losses = []      def loss(beta, X, y):         return float(np.sum(np.square(y - X@beta)))      def loss_grad(beta, X, y):         return (y - X@beta).T@X      while True:         last_loss = loss(beta, X, y)         beta += step_size * loss_grad(beta, X, y).T         this_loss = loss(beta, X, y)         if abs(last_loss - this_loss) &lt;= precision:             break      return beta  gradient_descent(X, y, 0.005, 10**(-6))</pre> <p><b>MLE (<a href="#">ex.</a>)</b></p> <p>derive MLE by hand, then implement as code</p>
<p><b>HYPOTHESIS TESTING WITH HYPERGEOMETRIC</b></p> <p>aka <b>Fisher's exact test</b></p> <p>does drug help? we assume <math>H_0</math> is <math>\mu_1 = \mu_2</math> so we use Hypergeometric distribution to check what is probability of drawing these numbers of deaths between study and control group under <math>H_0</math>.</p> <p>p-value: <math>\sum_{i=0}^{39} P_{H_0}(T=i)</math></p> <p><b>pro</b> doesn't assume knowledge of true value in ctrl group</p> <p><b>con</b> assumes knowledge of margins</p> <p><b>z-test</b> if we know the true value of <math>\sigma</math></p> <p><b>t-test</b> if we estimate <math>\hat{\sigma}</math>: <math>\frac{\bar{X}_n}{\frac{\hat{\sigma}}{\sqrt{n}}} \sim t_{n-1}</math></p> <p><b>LOAD CSV DATA</b></p> <pre>import pandas as pd data = pd.read_csv('gamma-ray.csv')</pre> <p>optional header=None and index_col</p> <p><b>LIKELIHOOD RATIO TEST (POISSON)</b></p> <pre>from scipy import stats  n = 1 d = 1 for row in data.data.iterrows():     G_i = row["count"]     t_i = row["seconds"]     lambda_i = G_i/t_i     n *= stats.poisson.pmf(G_i, lambda_mle)     d *= stats.poisson.pmf(G_i, lambda_i)  test_statistic = -2*np.log(n/d) test_statistic</pre>		

LINEAR CLASSIFIERS	STOCHASTIC GRADIENT DESCENT	KERNEL PERCEPTRON	MISC
boundary line: $\theta \cdot x + \theta_0 = 0$  $\theta \cdot x$ is positive if $x$ is on the right side of the decision boundary	<b>update:</b> sample $i$ at random <b>decreasing learning rate</b> $\eta_T = \frac{1}{1+T}$  $\theta \leftarrow \theta - \eta_T \nabla_{\theta} \left[ Loss_h \left( y^{(i)} \theta \cdot x^{(i)} \right) + \frac{\lambda}{2} \ \theta\ ^2 \right]$  $= \theta \left[ \begin{cases} 0 & \text{if } Loss_h = 0 \\ -y^{(i)} \cdot x^{(i)} & \text{if } Loss_h > 0 \end{cases} + \lambda \theta \right]$	$\theta = 0$ for $i = 1..n$ <div><math>\text{if } y^{(i)} \cdot \theta \cdot \phi \left( x^{(i)} \right) \leq 0</math> <math>\theta \leftarrow \theta + y^{(i)} \phi \left( x^{(i)} \right)</math></div> $\theta \cdot \phi \left( x^{(i)} \right) = \sum_j^n \alpha_j y^{(j)} \phi \left( x^{(j)} \right) \phi \left( x^{(i)} \right)$ where $\alpha_j$ is nb of mistakes made on $j^{\text{th}}$ example	$\vec{a} \cdot \vec{b} = \sum a_i b_i$ $a \cdot a = \ a\ ^2$ $\theta_z \cdot Ax = \theta_z^T Ax = \left( A^T \theta_z \right) \cdot x$  <b>distance from point to line</b> $d = \frac{ \theta \cdot x + \theta_0 }{\ \theta\ }$  <b>derivative of norm squared</b> $\left( \ h(x)\ ^2 \right)' = 2h(x)h'(x)$  <b>derivative</b> $[y(\theta \cdot x)]' = yx$ because we derive for each $\theta_i$
<b>PERCEPTRON</b>  $y^{(i)} \in \{-1, 1\}$  training error: $\varepsilon(\theta, \theta_0) = \frac{1}{n} \sum 1 \left\{ y^{(i)} \left( \theta \cdot x^{(i)} + \theta_0 \right) \leq 0 \right\}$  <b>algorithm</b> $\theta = \vec{0}$ for $t = 1...T$ for $i = 1...n$ if $y^{(i)} \left( \theta \cdot x^{(i)} + \theta_0 \right) \leq 0$ then $\theta = \theta + y^{(i)} x^{(i)}$ $\theta_0 = \theta_0 + y^{(i)}$	<b>LINEAR REGRESSION</b>  <b>empirical risk</b>  $R_n(\theta) = \frac{1}{n} \sum \frac{\left( y^{(t)} - \theta x^{(t)} \right)^2}{2}$ small deviations less penalised than large ones  $\nabla_{\theta} R_n(\theta) = - \left( y^{(i)} - \theta x^{(t)} \right) \cdot x^{(t)}$  <b>gradient descent</b> $\theta = 0$ for: random $t = \{1..n\}$ # stochastic, could also loop $\theta = \theta + \eta_k \left( y^{(t)} - \theta \cdot x^{(t)} \right) \cdot x^{(t)}$  with $\eta_k = \frac{1}{1+k}$	$\theta = 0$ for $i = 1..n$ <div><math>\text{if } y^{(i)} \sum_j^n \alpha_j y^{(j)} K \left( x^{(j)}, x^{(i)} \right) \leq 0</math> <math>\alpha_i = \alpha_i + 1</math></div> <b>kernel composition</b> $\cdot f: \mathbb{R}^d \rightarrow R$ and $K(x, x')$ $\widehat{K}(x, x') = f(x) K(x, x') f(x')$ $\cdot K(x, x') = K_1(x, x') + K_2(x, x')$ $\cdot K(x, x') = K_1(x, x') + K_2(x, x')$  <b>radial basis kernel</b> $K(x, x') = \exp \left( -\frac{1}{2} \ x - x'\ ^2 \right)$  $\Rightarrow$ correctly classifies every point $\Rightarrow$ radial in $\phi(x)$ space  <b>graph classification</b> look at number of bends in figure to determine dimension	<b>L1 norm</b> $\ w\ _1 = \sum_i  w_i $  <b>L2 norm</b> $\ w\ _2 = \sqrt{\sum_i w_i^2}$  <b>cosine similarity</b> $\cos \left( x^{(i)}, x^{(j)} \right) = \frac{x^{(i)} \cdot x^{(j)}}{\ x^{(i)}\  \ x^{(j)}\ }$  <b>euclidean distance</b> $dist \left( x^{(i)}, x^{(j)} \right) = \ x^{(i)} - x^{(j)}\ $
<b>REGULARIZATION</b>  we want large boundaries: $\vdots \quad   \quad \vdots$ where $\vdots$ are $\theta \cdot x^{(i)} + \theta_0 = \{1, -1\}$ and $ $ is $\theta \cdot x^{(i)} + \theta_0 = 0$ $\cdot$ prevents overfitting $\cdot$ the bigger $\ \theta\ $ the faster (= shorter distance) these reach -1 and 1 $\cdot$ we can control the margins distance with $\ \theta\ $  <b>signed distance</b> $\gamma_i = y^{(i)} \frac{\theta \cdot x^{(i)} + \theta_0}{\ \theta\ } = \frac{1}{\ \theta\ }$ or $-\frac{1}{\ \theta\ }$ to maximise $\frac{1}{\ \theta\ }$ we minimise $\frac{1}{2} \ \theta\ ^2$  <b>hinge loss</b>  $Loss_h(z) = Loss_h(y^{(i)}(\theta \cdot x^{(i)} + \theta_0)) = \max(0, 1 - y^{(i)}(\theta \cdot x^{(i)} + \theta_0))$ $= \begin{cases} 0 & \text{if } z \geq 1 \quad \text{correct side, outside margins} \\ 1 - z & \text{if } z < 1 \end{cases}$  hinge loss improves by moving $\theta$ towards example	<b>NONLINEAR CLASSIFICATION</b>  not linearly separable $\Rightarrow$ map $x$ to higher dimension  <b>kernels</b>  i.e. $K(x, x') = \phi(x) \cdot \phi(X') = (x \cdot x') + (x \cdot x')^2$  where $\phi(x) = (x_1, x_2, x_1^2, \sqrt{2}x_1x_2, x_2^2)^T$ $\Rightarrow$ dot product is cheap to compute	<b>K-NEAREST NEIGHBOURS</b>  $\hat{Y}_{ai} = \frac{1}{K} \sum_{b \in \text{KNN}(a,i)} Y_{bi}$  or weigh by similarity  $\hat{Y}_{ai} = \frac{\sum_{b \in \text{KNN}(a,i)} \text{sim}(a, b) Y_{bi}}{\sum_{b \in \text{KNN}} \text{sim}(a, b)}$  <b>euclidian distance</b> $\ x_a - x_b\ $ <b>cosine similarity</b> $\cos(\theta) = \frac{x_a \cdot x_b}{\ x_a\  \ x_b\ }$  <b>limitations</b> my combination of tastes maybe uncommon	
<b>SOLVING QUADRATIC</b>  if linearly separable and don't allow any errors, can solve quadratic: $\Rightarrow J(\theta, \theta_0) = \frac{1}{2} \ \theta\ ^2$	<b>COLLABORATIVE FILTERING</b>  assumes X has low rank  $J(u, v) = \sum_{a,i \in D} \frac{(Y_{ai} - u_a v_i)^2}{2} + \frac{\lambda}{2} \sum u_a^2 + \frac{\lambda}{2} \sum v_i^2$  $X = \begin{bmatrix} \text{user1rating1} & \text{user1rating2} \\ \text{user2rating1} & \text{user2rating2} \end{bmatrix} = u \cdot v^T$  $\Rightarrow$ u: user's rating tendency $\Rightarrow$ v: info about movie  <b>alternating projections</b> set v, optimise u, set u... continue until local convergence ( <a href="#">exercise link</a> )	<b>ACTIVATION FUNCTIONS</b>  <b>ReLU</b> $f(z) = \max(0, z)$ <b>tanh</b> $f(z) = \tanh(z)$ <b>linear</b> $f(z) = z$	

## DEEP LEARNING

- intermediate layer is like feature representation: "learning  $\phi$ "
- activation layer acts like linear classifier

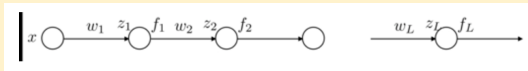
### backpropagation

compute partial derivatives using chain rule to do gradient descent

$$w_1 \leftarrow w_1 - \eta \left( \frac{\partial \text{Loss}}{\partial w_1} \right)$$

$$\frac{\partial \text{Loss}}{\partial w_1} = \frac{\partial \text{Loss}}{\partial f_1} \cdot \frac{\partial f_1}{\partial z_1} \cdot \frac{\partial z_1}{\partial w_1}$$

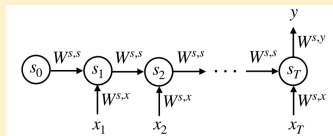
risks are vanishing or exploding gradients with long chains



## RNN

- same params for each layer
- variable number of layers
- ⇒ encodes sentences

$$s_t = \tanh(W^{s,s} s_{t-1} + W^{s,x} x_t)$$



RNN have some gradient issues

### gating network

in default RNN we overwrite  $s_t$  but sometimes it's useful to retain control over what's overwritten at each step

$$g_t = \text{sigmoid}(W^{g,s} s_{t-1} + W^{g,x} x_t)$$

$$s_t = (1 - g_t) \odot s_{t-1} + g_t \odot \tanh(W^{s,s} s_{t-1} + W^{s,x} x_t)$$

where  $\odot$  is element-wise multiplication

$$(1 - g_t) \odot \begin{pmatrix} 0 \\ 1 \end{pmatrix} \rightarrow \begin{pmatrix} 1 \\ 0 \end{pmatrix} \text{ i.e. } g_t \text{ is use to control what we retain from previous layer}$$

## MARKOV MODELS

<beg>  $w_1$   $w_2$   $w_3$  <end>  
 $w_0$   $w_L$

$$P(w_1, \dots, w_L) = P(w_1 | w_0) \dots P(w_L | w_{L-1})$$

$$\text{with } P(w' | w) = \frac{\text{count}(w, w')}{\sum_{\tilde{w}} \text{count}(w, \tilde{w})}$$

### RNNs for sequences

variable history

$$\phi(\text{word}) : X^W : s_t X : p$$

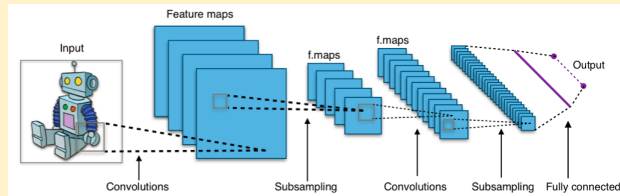
$$s_{t-1} : X$$

i.e. we're feeding the previous state as input to the next

## CNN

### convolution

- classify by small patches
  - slide over input
  - use the same weights for each patch to build feature map
- pooling (subsampling)**
- stores *if* there's an activation for feature in that area
  - stride



### cross-correlation/convolution

we don't actually need to flip  $g(t)$  as we would with convolution, because we're learning the filter, the values are just flipped

$$1d \quad f \odot g(t) = \sum f(T)g(T+t)$$

$$md \quad f \odot g(x, y) = \sum_{T_1} \sum_{T_2} f(T_1, T_2)g(T_1 + x, T_2 + y)$$

## K-MEANS

- 1) randomly assign  $z^{(1)} \dots z^{(2)}$
- 2) iterate
  - a) assign  $x$  to closest  $z$ 

$$\text{Cost}(z^{(1)} \dots z^{(k)}) = \sum_{i=1}^n \min_{j=1 \dots k} \|x^{(i)} - z^{(j)}\|^2$$
  - b) given  $C_1 \dots C_k$  find best rep.  $z$  in each  $C$ 

$$\text{Cost}(C_1 \dots C_k) = \min_{z^{(1)} \dots z^{(k)}} \sum_{j=1}^k \sum_{i \in C_j} \|x^{(i)} - z^{(j)}\|^2$$

### problems with K-means

- points are not in set of  $X$ s
- (euclidean distance)<sup>2</sup> only
- K-means is  $O(n \cdot k \cdot d)$  for one update step (2)

### find best $z_j$ for $C_j$

$$\frac{\delta}{\delta z_j} \sum_{i \in C_j} \|x^{(i)} - z^{(j)}\|^2 = 0$$

$$z^{(j)} = \frac{\sum x^{(i)}}{|C_j|}$$

where  $|C_j|$  is size of  $C_j$

### impact of init

- init matters
- we can end up with undesirable convergence

## K-MEDIANS

Like K-means but with medians. Compute median in each single dimension in the Manhattan distance formulation (see [ex.](#))

## K-MEDOIDS

- 1) pick init  $z^{(i)} \dots z^{(j)}$  from points  $X$
- 2) iterate until no change in cost
  - a) for  $i = 1 \dots n$ 

$$C_j = \left\{ i : z^{(j)} \text{ is closest to } x^{(i)} \right\}$$

- b) pick the most central  $x^{(i)}$  in  $C_j$  to be  $z^{(j)}$

for  $j = 1 \dots k$

$$z_j = \left\{ x^{(i)} \dots x^{(n)} : \sum_{i \in C_j} \text{dist}(x^{(i)}, z^{(j)}) \text{ is minimal} \right\}$$

· same idea as K-means but  $z_j$  is in  $x$  and can use any distance measure

· K-medoids is  $O(n^2 \cdot k \cdot d)$  for update step (2)

△ step 2 gets executed even if no change and cost, and order we cycle in matters and can update the medoid (see [ex.](#))

## GENERATIVE MODELS

$p(w | \theta) = \theta w$ : probability of generating word  $w$  given all possibilities`

### probability of generating doc D

$$p(D | \theta) = \prod_{i=1}^n \theta_{w_i} = \prod_{w \in W} \theta_w^{\text{count}(w)}$$

### MLE

$$\max P(D | \theta) \rightarrow \sum_{w \in W} \text{count}(w) \log(\theta_w)$$

$$\Rightarrow \tilde{\theta}_w = \frac{\text{count}(w)}{\sum_{w' \in W} \text{count}(w')}$$

### prediction

essentially acts as a linear classifier through origin

$$\log \frac{p(D | \theta^+)}{p(D | \theta^-)} = \begin{cases} \geq 0 & + \\ < 0 & - \end{cases}$$

### prior, posterior, likelihood

for when we have same prior knowledge of document classification

$$P(y = + | D) = \frac{P(D | \theta^+) P(y = +)}{P(D)}$$

⇒ linear classifier with offset influenced by prior

△ N-1 params in multinomial (i.e. remember the binomial uses 1 param for two values)

<div>GAUSSIAN GENERATIVE MODELS</div> <div> <math display="block">X \mid \mu, \sigma^2 \sim N(\mu, \sigma^2)</math> <p>want to find <math>\mu</math> and <math>\sigma^2</math> that give highest likelihood to training data.</p> <math display="block">\Rightarrow \frac{\delta l_n}{\delta \mu} = 0 \text{ and } \frac{\delta l_n}{\delta \sigma^2} = 0</math> <math display="block">\widehat{\mu} = \frac{1}{n} \sum x^{(i)} \text{ and } \widehat{\sigma^2} = \frac{1}{nd} \sum \ x^{(i)} - \mu\ ^2</math> </div>	<div>MARKOV DECISION PROCESS</div> <div> <ul style="list-style-type: none"> <li>states <math>s \in S</math></li> <li>actions <math>a \in A</math></li> <li>action dependent on transition probabilities</li> </ul> <math display="block">T(s, a, s') = P(s' \mid a, s)</math> <math display="block">\sum_{s' \in S} T(s, a, s') = 1</math> <math display="block">R(s, a, s'):</math> reward for starting s, doing a, ending s' </div>	<div>RL</div> <div> <p>typically won't have T and R ahead of time</p> <p><b>Q-value iteration</b>  want to calculate estimates for Q but don't have T and R.  we'll use <b>exponential running average</b> to get estimates for <math>Q(s, a)</math>:</p> <math display="block">Q_{i+1}(s, a) = \alpha \cdot \text{sample} + (1 - \alpha)Q_i(s, a)</math> <p>sample: <math>R(s, a, s') + \gamma \max_{a'} Q_i(s', a')</math></p> </div>
<div>MIXTURE MODELS</div> <div> <p><b>mixture components</b> <math>K: N(x, \mu_j, \sigma_j^2)</math></p> <p><b>mixture weights</b> <math>p_1 \dots p_k \quad \sum p_j = 1</math></p> <math display="block">\theta = p_1 \dots p_k, \mu_1 \dots \mu_k, \sigma_1^2 \dots \sigma_k^2</math> <math display="block">p(x \mid \theta) = \sum_{j=1}^k p_j N(x, \mu_j, \sigma_j^2)</math> <math display="block">p(S_n \mid \theta) = \prod_{i=1}^n \prod_{j=1}^k p_j N(x, \mu_j, \sigma_j^2)</math> </div>	<div>UTILITY FUNCTIONS</div> <div> <p><b>finite horizon (not important)</b></p> <p>count reward for next n steps: <math>U(s_0 \dots s_n) = \sum^n R(s_i)</math></p> <p><b>infinite horizon/discounted reward</b></p> <math display="block">U(s_0 \dots s_n) = \sum_{k=0}^\infty \gamma^k R(s_k) \leq \frac{R_{\max}}{1 - \gamma}</math> <p><b>value function</b></p> <p><math>V^*(s)</math>: expected reward if agent acts optimally starting at s</p> <p><math>\pi: S \rightarrow A</math>: assigns action <math>\pi(s)</math> to state s</p> <p>optimal policy <math>\pi^*</math> assigns action at every state that maximises expected utility</p> </div>	<div> <p>1) <b>initialisation</b> <math>Q(s, a) = 0 \forall a</math></p> <p>2) <b>iterate until convergence</b></p> <div> <p>a) collect samples: <math>s, a, s', R(s, a, s')</math></p> <p>b) <math>Q_{i+1}(s, a) = \alpha \left[ R(s, a, s') + \gamma \max_{a'} Q_i(s', a') \right] + (1 - \alpha)Q_i(s, a)</math></p> </div> <p><math>\epsilon</math> <b>greedy</b></p> <p>balance exploration and exploitation</p> <ul style="list-style-type: none"> <li>random action w.p. <math>\epsilon</math></li> <li>best current action w.p. <math>(1 - \epsilon)</math></li> </ul> <p><math>\epsilon</math> should decay over time</p> </div>
<div>EM algorithm</div> <div> <p>"random" init of <math>\theta: \mu^{(i)}, \sigma^{(i)}, p_i</math></p> <p>init matters for speed and outcome</p> <ul style="list-style-type: none"> <li>converges locally</li> <li>good idea to use k-means for init</li> </ul> <p><b>expectation step</b></p> <math display="block">p(j \mid i) = \frac{p_j N(x_i, \mu_j, \sigma_j^2 I)}{p(x \mid \theta)}</math> <p><math>\Rightarrow</math> probability that point i in cluster j</p> <p><math>\Rightarrow p(x \mid \theta) = \sum_{j=1}^k p_j N(x_i, \mu_j, \sigma_j^2 I)</math></p> <p><b>maximisation step</b></p> <p>re-estimate parameters, (nb: cluster j, point i)</p> <math display="block">\widehat{n}_j = \sum_{i=1}^n p(j \mid i)</math> <math display="block">\widehat{p}_j = \frac{\widehat{n}_j}{n}</math> <math display="block">\widehat{\mu}_j = \frac{1}{\widehat{n}_j} \sum_{i=1}^n p(j \mid i) x^{(i)}</math> <math display="block">\widehat{\sigma}_j^2 = \frac{1}{\widehat{n}_j d} \sum_{i=1}^n p(j \mid i) \ x^{(i)} - \mu^{(j)}\ ^2</math> </div>	<div>BELLMAN EQUATIONS</div> <div> <math display="block">V^*(s) = \max_a Q^*(s, a) = Q^*(s, \pi^*(s))</math> <math display="block">Q^*(s, a) = \sum_{s'} T(s, a, s') [R(s, a, s') + \gamma V^*(s')]</math> </div> <div>VALUE ITERATION ALGORITHM</div> <div> <math display="block">V_k^*(s) \rightarrow_{k \rightarrow \infty} V(s)</math> <ul style="list-style-type: none"> <li>init: <math>V_0^*(s) = 0</math></li> <li>iterate until <math>V_k^*(s) \approx V_{k+1}^*(s) \forall s</math></li> </ul> <math display="block">V_{k+1}^* = \max_a \left[ \sum_{s'} T(s, a, s') [R(s, a, s') + \gamma V_k^*(s')] \right]</math> <ul style="list-style-type: none"> <li>compute <math>Q^*(s, a)</math></li> </ul> <p>then <math>\pi^*(s) = \arg \max_a Q^*(s, a)</math></p> </div> <div>Q-VALUE ITERATION</div> <div> <p>same as above but with Q and the update is:</p> <math display="block">Q_{k+1}^* = \sum_{s'} T(s, a, s') \left[ R(s, a, s') + \gamma \max_a Q_k^*(s', a') \right]</math> <p><math>\Rightarrow</math> we're summing over all possible states we may land in</p> </div>	