

Solutions to Bain and Engelhardt's Introduction to Probability and Mathematical Statistics

All math should go into a math environment. $X + 1 = Y$ There are two types of math environments. Single dollar sign and double dollar sign.

$$X + 1 = Y$$

Another thing to keep in mind, there are no spaces.

helloeveryonehowareyou?

If you want to start a new paragraph, use a double backslash

This should be a new paragraph.

Now to go through some common commands used in 5080. All commands begin with a backslash.

To create a fraction $\frac{1}{X}$

To create a subscript: $F_{x,y}$

To integrate with bounds $\int_x^1 f(x,y)dx$

To specify a distribution: $Y \sim DE(\theta, 0)$ To write a branched function:

$$f(x : \theta, \eta) = \begin{cases} \frac{1}{2\theta} e^{x-\theta} & 0 > x \\ 0 & otherwise \end{cases}$$

One other useful thing is an equation array.

$$\begin{aligned} f(x) &= x + 3y \\ &= x + 3(x + 2) \\ &= x + 3x + 6 \\ &= 4x + 6 \end{aligned}$$

06.01

Given: the pdf of x $f_x(x) = \begin{cases} 4x^3 & , \quad 0 < x < 1 \\ 0 & , \quad o/w \end{cases}$

Find: PDF of a) $Y = X^4$

Setup: Use the CDF technique to get the CDF of Y in terms of a CDF of X

$$F_Y(y) = P[Y \leq y] = P[X^4 \leq Y] = P[-y^{\frac{1}{4}} \leq X \leq y^{\frac{1}{4}}] = F_X(y^{\frac{1}{4}}) - F_X(-y^{\frac{1}{4}})$$

Steps: i) Differentiate with respect to y to find an equation given in terms of the pdf of x:

$$f_y(y) = \frac{d}{dy} F_X(y^{\frac{1}{4}}) - \frac{d}{dy} F_X(-y^{\frac{1}{4}}) = f_x(y^{\frac{1}{4}}) \frac{d}{dy} y^{\frac{1}{4}} - f_x(-y^{\frac{1}{4}}) \frac{d}{dy} -y^{\frac{1}{4}} = f_x(y^{\frac{1}{4}}) \frac{y^{-\frac{3}{4}}}{4} - f_x(-y^{\frac{1}{4}}) \frac{-y^{-\frac{3}{4}}}{4}$$

ii) Plug in the original limits and function for the pdf of x, and compute the cdf for y

Result: $f_y(y) = \begin{cases} 4y^{\frac{3}{4}} \frac{1}{4y^{\frac{3}{4}}} & , \quad 0 < x < 1 \\ 0 & , \quad o/w \end{cases} = \begin{cases} 1 & , \quad 0 < x < 1 \\ 0 & , \quad o/w \end{cases}$

Find: PDF of b) $W = e^X$

Setup: Use the CDF technique to get the CDF of W in terms of a CDF of X
 $F_W(w) = P[W \leq w] = P[e^X \leq W] = P[X \leq \ln W] = F_X(\ln W)$

Steps: i) Differentiate with respect to w to find an equation given in terms of the pdf of x:
 $f_w(w) = \frac{d}{dw} F_X(\ln W) \frac{d}{dw} (\ln w) = f_x(\ln w) \frac{1}{w}$

ii) Plug in the original limits and function for the pdf of x, and compute the cdf for y

Result: $f_W(w) = \begin{cases} \frac{4(\ln w)^3}{w} & , \quad 1 < w < e \\ 0 & , \quad o/w \end{cases}$

Find: PDF of c) $Z = \ln x$

Setup: Use the CDF technique to get the CDF of Z in terms of a CDF of X
 $F_Z(z) = P[Z \leq z] = P[\ln x \leq z] = P[X \leq e^z] = F_X(e^z)$

Steps: i) Differentiate with respect to z to find an equation given in terms of the pdf of x:
 $f_z(z) = \frac{d}{dz} F_X(e^z) = f_x(e^z) \frac{de^z}{dz}$

ii) Plug in the original limits and function for the pdf of x, and compute the cdf for y

Result: $f_Z(z) = \begin{cases} 4e^{4z} & , \quad -\infty \leq z < 0 \\ 0 & , \quad o/w \end{cases}$

Find: PDF of d) $U = (X - 0.5)^2$

Setup: Use the CDF technique to get the CDF of U in terms of a CDF of X
 $F_U(u) = P[U \leq u] = P[(X - 0.5)^2 \leq u] = P[|X - 0.5| \leq u^{0.5}] = F_X(u^{1/2} + 1/2) - F_X(-u^{1/2} + 1/2)$

Steps: i) Differentiate with respect to u to find an equation given in terms of the pdf of x:
 $f_U(u) = \frac{d}{du} F_X(u^{1/2} + 1/2) = f_x(u^{1/2} + 1/2) \frac{d}{du} (u^{1/2} + 1/2) - f_x(-u^{1/2} + 1/2) \frac{d}{du} (-u^{1/2} + 1/2)$
 $f_x(u^{1/2} + 1/2) \frac{1}{2} u^{-1/2} - f_x(-u^{1/2} + 1/2) \frac{1}{2} u^{-1/2}$

ii) INCOMPLETE

Result: $f_Z(z) = \begin{cases} 4e^{4z} & , \quad -\infty \leq z < 0 \\ 0 & , \quad o/w \end{cases}$

06.02

Given: $X \sim Unif(0, 1)$

Find: a) PDF of $Y = X^{1/4}$

Setup: $F_Y(y) = P[Y \leq y] = P[X^{1/4} \leq y] = P[X \leq y^4] = F_X(y^4)$

Steps: i) find the pdf of x. Because X is a Uniform distribution with parameters 1 and 0, the pdf, which for Unif(a,b) is $1/(b-a)$ where $a < x < b$. Here, Unif(0,1) gives $1/1-0 = 1$

ii) Differentiate with respect to y to find an equation given in terms of the pdf of x:
 $f_Y(y) = \frac{d}{dy} F_X(y^4) = 4y^3$

Result: $f_Y(y) = \begin{cases} 4y^3 & , \quad 0 < y < 1 \\ 0 & , \quad o/w \end{cases}$

Find: b) PDF of $W = e^{-X}$

Setup: $F_W(z) = P[W \leq w] = P[e^{-X} \leq w] = P[-X \leq \ln w] = P[X \geq -\ln w] = 1 - F_X(-\ln w)$

Steps: i) find the pdf of x. See part a) for an explanation of why it is 1 when $a < x < b$

ii) Differentiate with respect to w to find an equation given in terms of the pdf of x:
 $f_W(w) = -\frac{d}{dw} F_X \frac{d}{dw} (-\ln w) = -f_X(-\ln w) \frac{-1}{w} \quad \text{for} \quad e^{-1} < w < 1 = -\frac{1}{w}$

Result: $f_W(w) = \begin{cases} \frac{1}{w} & , \quad e^{-1} < w < 1 \\ 0 & , \quad o/w \end{cases}$

Find: c) PDF of $Z = 1 - e^{-X}$

Setup: $F_Z(z) = P[Z \leq z] = P[1 - e^{-X} \leq z] = P[-e^{-X} \leq z - 1] = P[e^{-X} \geq 1 - z] = P[-X \geq \ln(1 - z)] = P[X \leq -\ln(1 - z)] = F_X(-\ln(1 - z))$

Steps: i) find the pdf of x. See part a) for an explanation of why it is 1 when $a < x < b$

ii) Differentiate with respect to w to find an equation given in terms of the pdf of x:
 $f_Z(z) = -\ln(1 - z) = -\frac{-1}{1-z} = \frac{1}{1-z} \quad \text{for} \quad 0 < z < 1 - e^{-1}$

Result: $f_Z(z) = \begin{cases} \frac{1}{1-z} & , \quad 0 < z < 1 - e^{-1} \\ 0 & , \quad o/w \end{cases}$

Find: d) PDF of $U = X(1 - X)$

Setup: $F_U(u) = P[U \leq u] = P[X(1 - x) \leq u] = P[-X^2 + X \leq u] = P[-(X - 1/2)^2 \leq u - 1/4] = P[(X - 1/2)^2 \geq 1/4 - u] = P[|(X - 1/2)| \geq (1/4 - u)^{1/2}] =$

Steps: i) find the pdf of x. See part a) for an explanation of why it is 1 when $a < x < b$

ii) INCOMPLETE:

$f_Z(z) = -\ln(1 - z) = -\frac{-1}{1-z} = \frac{1}{1-z} \quad for \quad 0 < z < e^{-1}$

Result: $f_W(w) = \begin{cases} \frac{1}{1-z} & , \quad 0 < z < e^{-1} \\ 0 & , \quad o/w \end{cases}$

06.03

Given: PDF $f_R(r) = \begin{cases} 6r(1 - r) & , \quad 0 < r < 1 \\ 0 & , \quad o/w \end{cases}$

Find: Distribution of the circumference

Setup: The circumference is $c = 2\pi r$. We have the pdf in terms of x, so this is the transformation:

$F_C(c) = P[C \leq c] = P[2\pi r \leq c] = P[r \leq c/2\pi] = F_x(c/2\pi)$

Steps: i) Differentiate with respect to c to find an equation given in terms of the pdf of x.

$f_C(c) = \frac{d}{dc} F_R(c/2\pi) = f_R(c/2\pi) \frac{d}{dc} (c/2\pi) = f_R(c/2\pi) (1/2\pi)$

ii) Plug the original pdf back into this new form:

$f_C(c) = \frac{6c}{2\pi} (1 - (c/2\pi)) (1/2\pi) = \frac{6c(2\pi - c)}{(2\pi)^3} \quad if \quad 0 < c < 2\pi$

Result: $f_C(c) = \begin{cases} \frac{6c(2\pi - c)}{(2\pi)^3} & , \quad 0 < c < 2\pi \\ 0 & , \quad o/w \end{cases}$

Find: Distribution of the area

Setup: The area is $a = \pi r^2$ so the cdf $F_A(a) = P[A \leq a] = P[\pi r^2 \leq a] = P[r^2 \leq a/\pi] = P[|r| \leq (a/\pi)^{1/2}] = P[-(a/\pi)^{1/2} \leq c \leq (a/\pi)^{1/2}] = F_R((a/\pi)^{1/2}) - F_R(-(a/\pi)^{1/2})$

Steps: i) Differentiate with respect to a to find an equation in terms of the pdf of x.

$f_A(a) = \frac{d}{da} F_R((a/\pi)^{1/2}) - \frac{d}{da} F_R(-(a/\pi)^{1/2}) = f_R((a/\pi)^{1/2}) \frac{d}{da} (a/\pi)^{1/2} - f_R(-(a/\pi)^{1/2}) \frac{d}{da} (-(a/\pi)^{1/2})$

Result: $f_A(a) = \begin{cases} \frac{3(\sqrt{\pi}-\sqrt{a})}{\pi^{3/2}}, & 0 < a < \pi \\ 0, & o/w \end{cases}$

06.04 Please double check the results of this solution

For $X \sim WEI(\theta, \beta)$ we have the CDF as $F_X = 1 - e^{-\frac{x}{\theta}^\beta}$ and the pdf is $f(x) = \frac{\beta}{\theta^\beta} x^{\beta-1} e^{-\frac{x}{\theta}^\beta}$

a) We make the transformation by the CDF method:

$$\begin{aligned} \Pr(Y \leq y) &= \Pr\left(\frac{X^\beta}{\theta} \leq y\right) \\ &= \Pr\left(X \leq \theta y^{\frac{1}{\beta}}\right) \\ &= F_X\left(\theta y^{\frac{1}{\beta}}\right) \\ &= 1 - e^{-\frac{\theta y^{\frac{1}{\beta}}}{\theta}^\beta} \\ &= 1 - e^{-y}, \text{ where } 0 < y \end{aligned}$$

So we have our CDF. For the pdf we simply take the derivative of the above. So $pdf = e^{-y}$ where $0 < y$

b) $W = \ln X$. Again, the most simply method to get the CDF, and in turn the pdf is the CDF method.

$$\Pr(W \leq w) = \Pr(\ln X \leq w) \tag{1}$$

$$= \Pr(X \leq e^w) \tag{2}$$

$$= F(e^w) \tag{3}$$

$$= 1 - e^{-\frac{e^w}{\theta}^\beta} \text{ where } 0 < w \tag{4}$$

$$\tag{5}$$

Again we simply differentiate to get the pdf. which turns out to be $\beta e^{\beta w} \theta^{-\beta} e^{-\frac{e^w}{\theta}^\beta}$, $0 < w$

c)

06.10 Suppose X has pdf $f_X(x) = \frac{1}{2}e^{-|x|}$ for all real x .

(a) Find the pdf of $Y = |X|$.

CDF Method

$$F_Y(y) = P[Y \leq y] = P[|X| \leq y] = P[-y \leq X \leq y] = F_X(y) - F_X(-y)$$

$$f_Y(y) = \frac{dF_X(y)}{dy} - \frac{dF_X(-y)}{dy}$$

$$f_Y(y) = f_X(y) \frac{dy}{dx} - f_X(-y) \left(\frac{-dy}{dy}\right)$$

$$f_y = \frac{1}{2}e^{-y} + \frac{1}{2}e^{-y} = e^{-y} \quad y > 0$$

(b) Let $W = 0$ if $X \leq 0$ and $W = 1$ if $X > 0$. Find the CDF of W

$$F_W(w) = P[W = 0] = \frac{1}{2}$$

$$F_W(w) = P[W = 1] = \frac{1}{2}$$

$$F_W(w) =$$

$$\begin{cases} 0 & w \leq 0 \\ \frac{1}{2} & 0 \leq w \leq 1 \\ 1 & w > 1 \end{cases}$$

06.13 X has pdf

$$f(x) = \begin{cases} \frac{x^2}{24} & -2 < x < 4 \\ 0 & \text{otherwise} \end{cases}$$

We want pdf of the CDF $Y = X^2$ with regions: $(-2, 0) \cup [0, 4)$

$$[F_x(\sqrt{y}) - F_x(-\sqrt{y})] = \left[f_x(\sqrt{y})\left(\frac{1}{2}\sqrt{y}\right) - f_x(-\sqrt{y})\left(-\frac{1}{2}\sqrt{y}\right) \right]$$

$$f_y(y) = \begin{cases} \frac{y}{48\sqrt{y}} + \frac{y}{48\sqrt{y}} & 0 < y < 4 \\ \frac{y}{48\sqrt{y}} & 4 \leq y \leq 16 \\ 0 & \text{otherwise} \end{cases}$$

$$f_y(y) = \begin{cases} \frac{\sqrt{y}}{24} & 0 < y < 4 \\ \frac{\sqrt{y}}{48} & 4 \leq y \leq 16 \\ 0 & \text{otherwise} \end{cases}$$

06.14

Given: Joint PDF $f(x, y) = \begin{cases} 4e^{-2(x+y)} & , \quad 0 < x < \infty, 0 < y < \infty \\ 0 & , \quad o/w \end{cases}$

Find: a) CDF of $W=X+Y$

Setup: $F_w(w) = P[W \leq w] = P[X + Y \leq w]$

Steps:

i) Express as a sum of probabilities, replace probabilities with binomials

ii) Simplify and Use Combinatorial Identity

Result: $\binom{n+m}{k}$

06.15 This is a simplified version of example 6.4.5.

$X_1, X_2 \sim POI(\lambda)$ so the MGF of both is $e^{\lambda(e^t-1)}$. Thus by theorem 6.4.4

$$M_Y(t) = e^{\lambda(e^t-1)} e^{\lambda(e^t-1)} = e^{2\lambda(e^t-1)} \sim POI(2\lambda)$$

The pdf then of Y is

$$f_Y(y) = \begin{cases} \frac{e^{-2\lambda(2\lambda)^y}}{y!} & y = 0, 1, 2, \dots \\ 0 & \text{otherwise.} \end{cases}$$

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06.16 Note: the pdf of $f_{x_1, x_2} = \frac{1}{x_1^2} \frac{1}{x_2^2}$

a) We need to find $f_{u,v} = f_{x_1, x_2}(x_1(u, v), x_2(u, v))|J|$ where J is our jacobian. First we let $u = x_1 x_2$ and $v = x_1$ thus $x_1 = v$ and $x_2 = \frac{u}{v}$, now we can find J.

$$J = \begin{vmatrix} 0 & 1 \\ \frac{1}{v} & -\frac{u}{v^2} \end{vmatrix} = \frac{1}{v}$$

Finally, our pdf is:

$$\begin{aligned} f_{U,V}(u, v) &= f_{x_1, x_2}\left(v, \frac{u}{v}\right) \left| \frac{1}{v} \right| \\ &= \frac{1}{v^2} \frac{1}{\left(\frac{u}{v}\right)^2} \left| \frac{1}{v} \right| \\ &= \frac{1}{u^2 v}, 1 < v < u < \infty \end{aligned}$$

b) We need to find $f_u(u)$ given $f_{U,V}(u, v) = \frac{1}{u^2 v}, 1 < v < u < \infty$

$$\begin{aligned} f_u(u) &= \int_1^u \frac{1}{u^2 v} dv \\ &= \frac{1}{u^2} \ln(v) \Big|_1^u \\ &= \frac{1}{u^2} (\ln(u) - 0) \\ &= \frac{1}{u^2} \ln(u), 1 < u < \infty \end{aligned}$$