

<sup>1</sup> Understanding cultural evolution in hummingbird leks through the  
<sup>2</sup> fossilized birth-death process

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<sup>16</sup> predictive simulation, sequence alignment

<sup>17</sup> **Abstract**

<sup>18</sup> **Introduction**

<sup>19</sup> The idea that culture changes and diversifies over time in a manner analogous to organic evolution and  
<sup>20</sup> phylogenetic diversification can be traced back to Darwin, who noted that “the formation of different languages  
<sup>21</sup> and of distinct species, and the proofs that both have been developed through a gradual process, are curiously  
<sup>22</sup> parallel” (Darwin 1871). The term culture can be used broadly to refer to socially transmitted information  
<sup>23</sup> that influences behavioural patterns within animal groups (Laland and Hoppitt 2003). While human language,  
<sup>24</sup> beliefs, norms and material artefacts are well-known cultural domains, many forms of culture exist among  
<sup>25</sup> other animals, such as vocal dialects (Catchpole and Slater 2003; Aplin 2019), navigation routes (Laland  
<sup>26</sup> and Williams 1997; Jesmer et al. 2018), and tool use traditions (Whiten et al. 2005; Luncz and Boesch  
<sup>27</sup> 2014). In the last century, the formal modelling of cultural change as evolution, whether in humans or  
<sup>28</sup> non-human animals, has been conducted by drawing analogies between cultural and population genetic  
<sup>29</sup> processes (Cavalli-Sforza and Feldman 1981; Boyd and Richerson 1985), or between cultural and phylogenetic  
<sup>30</sup> diversification (Gray et al. 2007; Mesoudi 2017). However, unlike genetic evolution, in which the most  
<sup>31</sup> fundamental units of transmission (nucleotides) are essentially universal, cultural evolution implies disparate  
<sup>32</sup> units of transmission across taxa and in different social contexts (e.g. tool design vs. language). In order to  
<sup>33</sup> address the long-standing question of whether cultural change is truly akin to evolution, we require means to  
<sup>34</sup> systematically assess the power of evolutionary methods, across the great variety of cultural forms that have  
<sup>35</sup> emerged in the history of animals.

<sup>36</sup> Some learnt behaviours can be described as sequences of ethological units. For example, visual displays can  
<sup>37</sup> be encoded as a string of stereotyped motor patterns (Ligon et al. 2018; Araya-Salas et al. 2019) and bird  
<sup>38</sup> and whale songs are typically structured as sequences of repeated and hierarchically nested sounds (Payne  
<sup>39</sup> and McVay 1971; Rivera-Cáceres et al. 2016; Kershenbaum et al. 2016; Garland et al. 2017). Encoding  
<sup>40</sup> behaviour as a sequence facilitates the adoption of phylogenetic approaches that take molecular data as their  
<sup>41</sup> main input. Such approaches have been developed within a strong theoretical framework that continues  
<sup>42</sup> to grow and increasingly accommodates biological realism (Yang and Rannala 2012). Substitution models  
<sup>43</sup> applied to molecular sequence evolution are routinely combined with clock models and tree priors, such as  
<sup>44</sup> the birth-death process, to understand the temporal dynamics of lineage diversification and turnover (Morlon  
<sup>45</sup> 2014; Bromham et al. 2018) Analogously, clock models, tree priors and substitution models may be used  
<sup>46</sup> to elucidate the temporal dynamics of cultural diversification, when culture can be adequately modeled as  
<sup>47</sup> behavioural sequences composed of discrete units. Estimating the absolute fit of such models to cultural data  
<sup>48</sup> is crucial to evaluate the utility of phylogenetic approaches for our understanding of cultural evolution and  
<sup>49</sup> diversification.

50 Cultural evolution poses special challenges to the application of phylogenetic models. Most implementations of  
51 phylogenetic models on molecular data start with a sequence alignment. Alignments represent assumptions of  
52 homology between characters in matched positions along a sequence, but are typically treated as observations  
53 for phylogenetic inference (Lutzoni et al. 2000; Redelings and Suchard 2005; Lunter et al. 2005). Numerous  
54 methods of sequence alignment have therefore been developed to capture the main features of molecular  
55 evolution, and in some cases to explicitly model substitution events (Yang and Rannala 2012; Chatzou et al.  
56 2016). Nonetheless, the accurate reconstruction of homology in sequence alignments is a pervasive challenge  
57 in molecular phylogenetics (Warnow 2021), that is only exacerbated when borrowing phylogenetic tools for  
58 the study of behavioural sequences (Caetano and Beaulieu 2020). We clearly have a better understanding of  
59 the basic rules that govern the rates of different nucleotide substitutions than we do for changes in the dance  
60 moves of a courtship display or changes in the sequence of sounds of a mating call. A crucial question for  
61 the nascent field of cultural phylogenetics (*sensu* Mesoudi 2017) is therefore whether alignment algorithms  
62 developed for molecular data can be suitably modified to represent the processes behind cultural change.

63 Despite these challenges, culture also poses unmatched opportunities for the application of phylogenetic  
64 inference. Culture can change very rapidly in comparison to molecular evolution (Perreault 2012), allowing  
65 researchers to document lineage diversification events as they occur, and during the span of one or a few  
66 academic lifetimes. Cultural phylogenetics can therefore capitalize on a relatively rich historical record,  
67 that markedly contrasts the sparse fossil record of many organismal groups (Kidwell and Holland 2002).  
68 Recently developed phylogenetic methods have shown that sampling ancestors of extant taxa and explicitly  
69 incorporating these data in the diversification process allow for more accurate estimation of divergence times  
70 (Gavryushkina et al. 2014, 2017; Zhang et al. 2016). This is accomplished by the fossilized birth-death process  
71 (FBDP) (Heath et al. 2014; Gavryushkina et al. 2014), a model that jointly describes the probabilities of  
72 lineage splitting, extinction and fossilization that give rise to the sampled taxa, whether extant or fossil. Of  
73 course, the fossilization rate estimated in the FBDP may represent actual fossilization events, but can also  
74 be used to describe serially sampled viral strains (Stadler and Yang 2013; Gavryushkina et al. 2014), or,  
75 as in this case, historical records of behavioural patterns that are socially learnt and transmitted (Rama  
76 2018; Ritchie and Ho 2019; Zhang et al. 2020). Thus, when culture evolves rapidly and learnt behaviours are  
77 sampled serially, a vast record of ancestral lineages can bolster inferences of cultural diversification dynamics  
78 through the FBDP.

79 Cultural phylogenetics research that builds on Bayesian estimation of origination, extinction and preservation  
80 rates is recently growing, but remains restricted to specific domains of human culture (Gjesfjeld et al. 2016,  
81 2020; Rama 2018; Ritchie and Ho 2019; Sagart et al. 2019; Zhang et al. 2020). Studies applying the FBDP

82 in particular have been focused on elucidating the history of diverse human language families (Rama 2018;  
83 Sagart et al. 2019; Zhang et al. 2020). Thus, a great untapped potential remains for investigating cultural  
84 diversification through the FBDP in non-human animals. Such an approach can help us determine whether  
85 the analogy between organic evolution and cultural change holds for other cultural phenomena. Because bird  
86 songs are often socially learnt and linearly composed of discrete subunits, they can be used to examine the  
87 suitability of the FBDP as a phylogenetic model of cultural diversification.

88 The Long-billed Hermit (*Phaethornis longirostris*; Fig. 1a) produces songs that can be represented as sequences  
89 of discrete sounds fused together into an unbroken signal (Fig. 1b; see ‘Methods’). Indeed, the most salient  
90 differences among song types reside in the composition and sequential order of their sounds (Araya-Salas and  
91 Wright 2013). Evidence of social learning in this species (*sensu* Ten Cate 2021), includes micro-geographic  
92 song variation decoupled from genetic structure (Araya-Salas et al. 2019) and adult replacement of crystallized  
93 songs (Araya-Salas and Wright 2013). Males sing a single song-type repertoire, which enables comparisons  
94 of individual songs as homologous traits (as opposed to multiple song-type repertoires). Courtship occurs  
95 within leks of 5-20 highly vocal males (Stiles and Wolf 1979), which facilitates longitudinal monitoring of all  
96 song types within a lek. Moreover, song types can be shared by sub-groups of males within leks, with no  
97 evidence of song type sharing across leks (Araya-Salas et al. 2019), suggesting that leks operate as relatively  
98 isolated cultural systems. Such independence across leks provides an unmatched opportunity to investigate  
99 the robustness of phylogenetic models across different iterations of an underlying cultural diversification  
100 process.

101 Here, we used the FBDP to model cultural diversification in five leks of Long-billed Hermits, using historical  
102 song surveys spanning up to five decades (Fig. 1c-d). We then investigated model reliability and absolute  
103 fit of phylogenetic models, using posterior predictive simulation and comparing features of empirical song  
104 sequences to sequences generated by models under the FBDP. We further asked how biologically informed  
105 assumptions during sequence alignment impact model reliability and estimates of diversification dynamics.  
106 Finally, we explored how the use and completeness of historical records (analogous to fossil records) affect  
107 model reliability, parameter estimation and the fit of alternative clock models to long-billed hermit song data.



**Figure 1.** Socially transmitted songs in the Long-billed Hermit. **a)** A male Long-billed Hermit. **b)** Spectrograms of two songs from different males in the SUR lek, sampled in 2019. The *colour* arrow shows a pure tone and the *another colour* arrow show a vibratory sound. **c)** Locations of the study leks in the Caribbean lowlands of Costa Rica. **d)** Historical sampling of song records in each lek.

## 108 Methods

### 109 Data collection and song structure coding

110 Sound recordings of Long-billed Hermits were registered from 2008 to 2019, in five leks, distributed across  
 111 four sites in the Caribbean slope of Costa Rica: La Selva Biological Station (leks SUR and CCE), Finca las  
 112 Brisas (BR1), Hitoy Cerere Biological Reserve (HC1) and La Tirimbina Lodge (TR1) (Fig. 1c). We also  
 113 included historical recordings available for three of the studied leks at La Selva and Hitoy Cerere (Stiles and  
 114 Wolf 1979). Recordings were gathered with different equipment at different points in time (i.e. shotgun or  
 115 parabolic microphones, analog or digital recorders). Nonetheless, the spectrographic structure of the signals  
 116 (used for determining signal structure, see below) is not affected by the recording equipment in a detectable  
 117 manner.

118 Long-billed Hermit songs are composed of two basic sound types: tonal and vibratory sounds (trills). Pure  
 119 tones can vary in the degree of modulation (i.e. changes in frequency through time), while trills vary in the

number of oscillations per unit of time (i.e. rate) (Fig. 1b). We subdivided these two basic sound types into six categories (Fig. S1): slow trill, medium-paced trill, fast trill, downward pure tone, upward pure tone and flat pure tone. Songs were split into 20 equal-length segments, and each segment was assigned to one of these six categories, based upon visual inspection of spectrograms (Fig. S1). The choice of 20 segments per song captured a compromise between our ability to discriminate sound types, which increased with segment length, and the probability that a single sound type occurred in each segment, which decreased with segment length (Supporting Text 1). We validate this choice by assessing inter-observer repeatability of sound classification (Supporting Text 2; Fig. S2).

## Sequence alignment

Alignment of behavioural sequences is complicated by the challenge of establishing homology between ethological segments or units (Caetano and Beaulieu 2020). Here, we implemented and compared three alignment strategies based on two methods originally developed for multiple sequence alignment (MSA) of molecular data. In alignments of nucleotide and protein sequences, gaps represent insertion or deletion mutations, so that characters at gapped sites lack homology across the data set. Commonly used MSA methods differ in their treatment of insertion and deletion events in ways that can impact homology inferences in cultural as well as in molecular characters (Löytynoja 2012). MAFFT (Katoh et al. 2002; Katoh and Standley 2013) uses a progressive alignment algorithm with a default gap-opening penalty (1.53) and no gap extension penalty by default, in versions > 6.626. The L-INS-i method follows the progressive alignment by iterative refinement, based on consistency and weighted sum-of-pairs scores. In MAFFT versions > 7.371 user-defined alphabets and scoring matrices can be implemented in addition of nucleotide and amino acid alternatives. MAFFT is therefore a flexible program to align behavioural sequences in which changes analogous to multi-site insertions and deletions have occurred, and which are composed by a variable number of characters and character states. In our first alignment strategy, which we hereafter refer to as ‘MAFFT-agnostic’, we used the MAFFT L-INS-i method with default gap penalties and a customized scoring matrix in which all transitions between alternative character states were equally likely.

Our second alignment strategy also used the MAFFT L-INS-i method and default gap penalties, but we made the assumption that when hummingbirds modify pre-existing songs they are more likely to replace a trill by a different type of trill and a tone by a different type of tone than to change from vibratory to pure sounds or *vice versa*. We implemented this assumption by enforcing a higher cost of mismatches between sound categories than within either trills or pure tones. To determine an appropriate difference in mismatch scores, we made two further assumptions, namely that cultural evolution is independent between leks and

151 also between individuals within leks (see Discussion). If insertions and deletions of song segments occur  
152 independently among individuals, alignment length should increase as sequences are more distantly related  
153 (Löytynoja 2012). We would thus expect longer alignments in data sets composed of sequences from different  
154 leks than in data sets composed of sequences from the same lek, as these sequences have a more recent  
155 common ancestor. Following this logic, we selected mismatch scores for substitutions within and between  
156 sound categories (trill vs. pure tone) that maximize the alignment length for pools of sequences from different  
157 leks relative to the alignment length for the same number of sequences originating from the same lek. We  
158 hereafter refer to this alignment strategy as ‘MAFFT-optimal’.

159 For our third alignment strategy, we used the phylogenetically informed alignment program PRANK (Löytynoja  
160 and Goldman 2005, 2008). PRANK also uses a progressive algorithm but handles the placement of insertions  
161 and deletions differently, by using outgroup information in the subsequent alignment step. PRANK thus  
162 uses the sequence phylogeny to differentiate insertions from deletions, and thereby avoids site overmatching  
163 by penalizing insertions in a single stage of the alignment (Löytynoja and Goldman 2005). Unlike MAFFT,  
164 PRANK is an evolutionary aware program in that insertions, deletions and substitutions are modelled  
165 explicitly on a phylogenetic tree. However, PRANK does not support customized alphabets and substitution-  
166 rate matrices. To use PRANK, we assumed that pure tones can be treated as ambiguous between upward  
167 and downward tones, and medium-speed trills can similarly be treated and ambiguous between fast and  
168 slow trills. We therefore used IUPAC ambiguity codes for DNA nucleotides to rename song segments, with  
169 tones as purines and trills as pyrimidines. As per PRANK’s defaults we used a TN93 nucleotide substitution  
170 model with empirical base frequencies and transition/transversion rate ratio ( $\kappa$ ) = 2. Therefore, as in the  
171 ‘MAFFT-optimal’ alignment, the ‘PRANK-TN93’ alignment explicitly assumed a higher transition rate  
172 within vibratory and tonal sound categories than between them. For this alignment, we used the default  
173 gap-opening rate and extension probabilities (0.025 and 0.75 respectively), and we omitted the -F option that  
174 fixes inferred insertions but increases sensitivity to guide-tree accuracy.

## 175 Phylogenetic analysis

176 All phylogenetic analyses were conducted in RevBayes v. 1.0.12 and v. 1.1.0 , a computation environment  
177 that uses probabilistic graphical models for Bayesian inferences in phylogenetics and evolution (Höhna et  
178 al. 2016). Our phylogenetic model was a fossilized birth-death process (FBDP) which describes the joint  
179 prior distribution of the tree topology, divergence times and lineage sampling times before the present (Heath  
180 et al. 2014). In the FBDP, extant taxa and lineages sampled before the present are part of the same  
181 macroevolutionary process. For many applications of the FBDP, extinct and ancestral taxa can only be

sampled through fossils. However, in the case of fast-evolving songs that are culturally transmitted among individuals, historical records of songs are equivalent to fossil data. Historical records contain the character sequences of songs that existed in the past and may be ancestral to extant songs or may represent lineages that have gone extinct. As in the FBDP with fossil data, the probability that a historical song is an ancestor of extant songs depends on the rates of lineage turnover and the rate of recovery of historical records. This recovery rate is the rate at which ancestral songs are sampled from the lineage diversification process and it is a random variable drawn from a prior distribution, such as the birth and death rates of a traditional birth-death model. Sampling ancestors as part of the same evolutionary process as we have done here improves estimation of diversification and clock rates (Gavryushkina et al. 2014) and sampling character data from ancestors further improves estimates of divergence times in simulated data (Luo et al. 2020).

Our data set of historical records of songs in hummingbird leks has three advantages in comparison to most fossil data sets used in phylogenetic analyses. First, there is no stratigraphic uncertainty. We can be certain that historical songs occurred in the year when they were recorded. Second, there are no partial fossils. Songs recorded in the past are just as complete as the most recent ones, creating no additional ambiguity in character states of historical songs. Third, the historical record is relatively rich. In all leks, there are multiple years sampled consecutively, and in two leks (SUR and CCE) historical records go back to 1969 (Fig. 1d). Because leks are small and hummingbirds are actively displaying their calls (Stiles and Wolf 1979), we can assume detection is nearly perfect and thus there is no missing taxa in any of the sampled years.

A possible complication in our analysis is that long gaps without sampling are interspersed in the three leks with deeper historical records (HC1, CCE and SUR). The temporal distribution of these ancestral samples is not unlike that of fossils, which are typically aggregated in discrete strata of exposed rocks (Holland 2016). The FBDP is robust to some forms of bias in fossil sampling, including non-continuous recovery (Heath et al. 2014). Nonetheless, to better understand the effects of deep, yet discontinuous historical sampling, we conducted all analyses for these leks both with the complete data set, including long gaps without lineage sampling, and with the more recent and continuously sampled data. Finally, to investigate the general impact of sampling historical records on phylogenetic inference of song diversification, we conducted an additional set of analyses, including only terminal tips (i.e. songs observed in the last year of sampling). For these analyses without historical records, we used the three leks (BR1, SUR and TR1) that had 3 or more distinct songs in their last year of sampling.

Another potential issue arises from the years with highly frequent sampling, in which identical songs could be sampled at multiple time points. This is uncommon for fossil data, as it would entail the discovery of fossils with the same character state combination in multiple horizons. Here, we focus on the results of analyses in

214 which all historical occurrences are considered in the evolutionary process, including identical songs sampled  
215 in consecutive years. However, we also conducted all analyses accounting only once for each unique song, at  
216 its earliest occurrence.

217 Phylogenetic analyses were conducted with all three alignment strategies (MAFFT-agnostic, MAFFT-optimal  
218 and PRANK-TN93) for each lek. We used an exponential prior with rate parameter = 10 for the speciation,  
219 extinction and historical sampling rates, and a broad uniform prior, bounded between 1000 and 0 years, on  
220 the root age of all leks. Song sequences were assumed to evolve under a generalised time-reversible (GTR)  
221 model with exchangeability rates and stationary frequencies drawn from a flat Dirichlet prior. Site-rate  
222 heterogeneity was modelled with a discretised gamma distribution with four rate categories and with equal  
223 shape and scale parameters, in turn drawn from an exponential prior with rate = 10.

224 We tested both global and relaxed clocks for song evolution. Branch rates under the global clock were drawn  
225 from an exponential prior with rate = 10. Branch rates under the relaxed clock were uncorrelated and drawn  
226 from an exponential prior, with mean in turn drawn from an exponential hyperprior with rate = 10. We  
227 compared clock models using marginal likelihood approximation via the stepping stone algorithm (Xie et  
228 al. 2010). Clock-model comparisons were conducted for each lek (BR1, CCE, HC1, SUR, TR1), alignment  
229 (MAFFT-agnostic, MAFFT-optimal, PRANK-TN93), historical dataset (oldest records included, recent  
230 records only, no fossils) and use of historical records per song (using all, using earliest). For diversification  
231 dynamics, tree comparisons and tests of model reliability (see below), we present results under the preferred  
232 clock model.

233 We conducted two independent MCMC runs for each analyses, with 150 000 iterations and an additional 15  
234 000 of burn-in and parameter tuning every 200. To improve mixing, we used the Metropolis-Coupled MCMC  
235 sampler with three heated chains and default swapping parameters. To avoid autocorrelation in the posterior  
236 we saved samples every 100th iteration. We assessed MCMC performance using the package *coda* v. 0.19-4  
237 (Plummer et al. 2006) in R v. 4.0.4 (R Core Team 2021). We checked for convergence between independent  
238 runs visually and using the Gelma-Rubin potential scale-reduction factor (psrf). We assumed convergence if  
239 psrf < 1.05 for all variables, as well as the multivariate estimate. We also inspected autocorrelations between  
240 draws (targeted below 0.1) and effective sample sizes (targeted above 200) for all model variables. We plotted  
241 trees for visual inspection using the package *ggtree* (Yu et al. 2017) and our general results using the package  
242 *ggplot2* (Wickham 2016) in R.

243 **Model reliability**

244 We used predictive data simulations to test for absolute model fit, also implemented in RevBayes (Höhna et  
245 al. 2018). During parameter inference, a Stochastic-Variable-Monitor stored the stochastic variable values  
246 for each posterior sample. Then, these values were used to simulate new data sets based on the inference  
247 model. We specified a thinning of 2 iterations for the stochastic variable trace, thus simulating 3 000 datasets  
248 for the ‘large’ leks (CCE, SUR) and 1 500 datasets for the ‘small’ leks (BR1, HC1, TR1).

249 We present data-based test statistics comparing simulated to empirical datasets, as tests of absolute model  
250 fit. We calculated 10 such statistics: 1) the number of invariant sites in the alignment, 2) the number of  
251 segregating sites in the alignment, 3) the maximum length of invariant blocks, 4) the maximum length of  
252 variable blocks, 5) the number of invariant blocks, 6) the maximum pairwise difference between two sequences  
253 in an alignment, 7) the minimum pairwise difference between two sequences in an alignment, and three  
254 measurements of genetic diversity: 8) Watterson’s  $\theta$ , an estimate of “population mutation rate” (Watterson  
255 1975), 9) Tajima’s D, a measurement of whether a population evolves neutrally (Tajima 1989), 10)  $\pi$ , the  
256 average number of pairwise differences in the alignment, used to calculate Tajima’s D, and a measure of  
257 character diversity. For more details about these statics and how they are calculated see Höhna et al. (2018).

258 For each test, we report a posterior predictive effect size (PPES) and a two-tailed posterior predictive p-value  
259 (Höhna et al. 2018). The PPES of each statistic corresponds to the difference between the median of the  
260 posterior distribution of simulated data sets and the empirical value, normalized by the SD of the posterior  
261 distribution (Höhna et al. 2018). The two-tailed posterior predictive p-value is calculated by first obtaining  
262 a lower- and upper-tail p-value and multiplying the smaller of the one-tailed p-values by two. The lower  
263 one-tailed p-value is the proportion of simulated data sets in which the value for the test statistic is lower  
264 than or equal to the observed value. The upper one-tailed test is the proportion of simulated data sets in  
265 which the value for the test statistic is greater than or equal to the observed value. Especially with small  
266 data sets, it is possible that test statistics in mutliple simulated data sets are exactly equal to test statistics  
267 in the empirical data. In these cases the smaller of the two one-tailed p-values could be greater than 0.5. In  
268 these cases, the posterior predictive p-value was set to 1.

269 **Treespace and parameter sensitivity**

270 We explored tree topology congruence of different models by comparing topological distances between high  
271 posterior probability trees. Topologies were compared with the Robinson-Foulds distance (Robinson and  
272 Foulds 1981) with the R package *phangorn* v.2.11.1 (Schliep 2011). Topological distances were projected  
273 in a bidimensional space using Classic Multidimentional Scaling in order to quantify topological space.

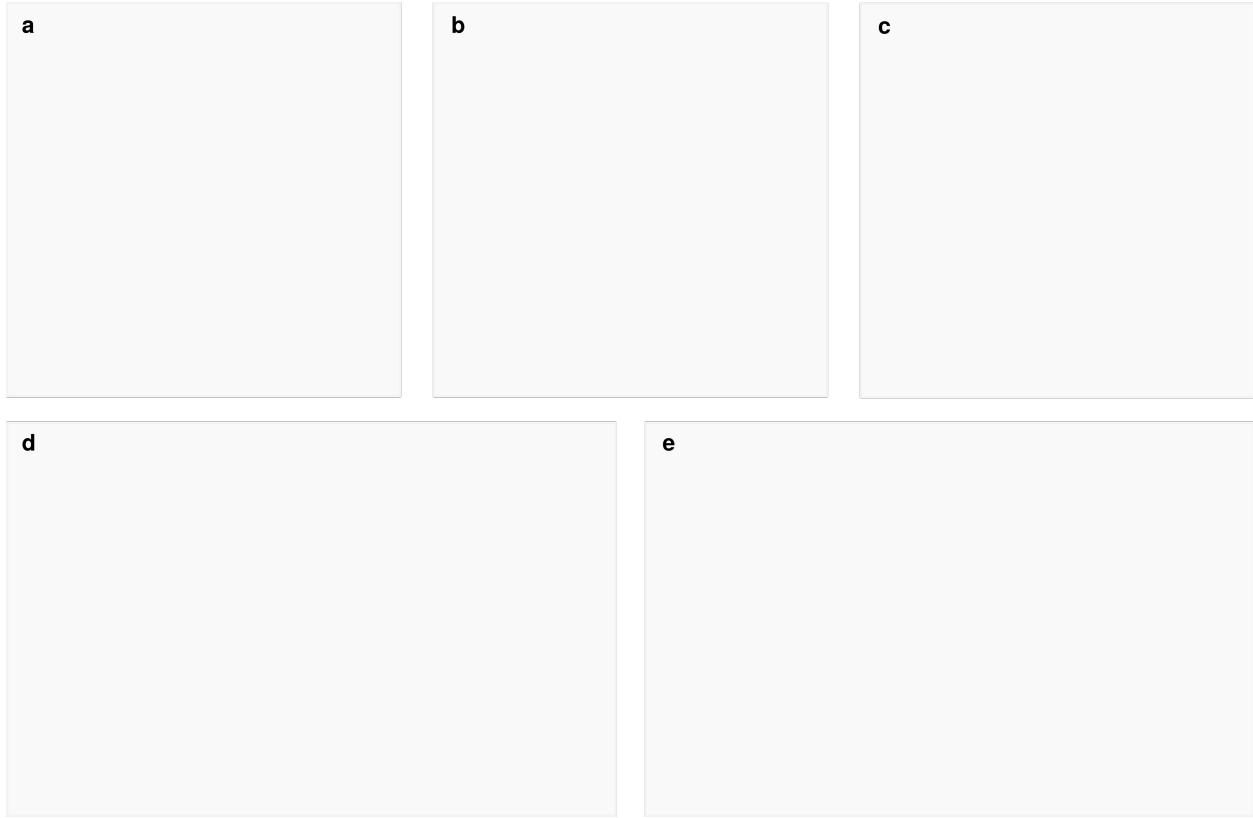
274 Only tree tips pairwise shared between trees were considered when calculating distances. We estimated  
275 the overall spread of the topological space (i.e. space size) for different models as a metric of within-model  
276 topological congruence. Topological space size was quantified as the 95% kernel density area. We also  
277 calculated between-model topological congruence as the overlap of the topological space of a model to the  
278 spaces from other models. Space overlap was estimated as the intersection-over-union of the two 95% kernel  
279 density spaces (i.e. proportion of the joint area of two spaces that was shared). Topological congruence  
280 descriptors were calculated using the R package *PhenotypeSpace* v.0.1.0 (Araya-Salas and Odom 2022).

281 The effect of different model specifications on these two topological space descriptors was evaluated using  
282 Bayesian lineal regression models with each descriptor as the response variable (modeled with a gaussian  
283 distribution), model alignment strategy, use of historical data, historical record completeness, and clock  
284 model as predictors and lek as a varying intercept effect (i.e. random effect). Regression models were run in  
285 Stan (Carpenter et al. 2017) through the R platform (R Core Team 2021) using the package *brms* v.2.22.0  
286 (Bürkner 2017). We also computed multiple comparisons of alignment strategies (similar to post hoc tests in  
287 frequentist statistics) using the joint posterior distribution of the model parameters. We present effect size  
288 point estimates as median posterior estimates, and report 95% highest posterior density intervals (HPDIs)  
289 as a measure of uncertainty. We evaluated the predictive performance of all models, measured by the LOO  
290 Information Criterion (LOOIC; Vehtari et al. 2017), by comparing each model against its correspondent  
291 null model (i.e., a model with no fixed effects). Models were compared using the expected log predictive  
292 density (ELPD). Models were run on three chains for 2500 iterations, following a warm-up of 2500 iterations  
293 using weakly informative priors. Effective sample size was kept above 3000 for all parameters. Potential  
294 scale reduction factor was used to assess model convergence and kept below 1.05 for all parameter estimates.  
295 Performance was checked visually by plotting the trace and distribution of posterior estimates for all chains.  
296 We also plotted the autocorrelation of successive sampled values to evaluate independence of posterior samples.  
297 Posterior predictive check plots were also used to inspect if model predictions aligned with the distribution  
298 and variability of the observed data.

299 Finally, we asked if inferences of diversification dynamics and song evolution were influenced by the use of  
300 different alignment strategies. To do this we compared the posterior distributions of parameter estimates  
301 between the MAFFT-agnostic, MAFFT-optimal and PRANK-TN93 alignment strategies. For diversification  
302 dynamics we compared speciation, extinction, and net diversification rates, as well as the age of the MRCA  
303 of all songs (hereafter ‘root age’, including extinct lineages) and the age of the MRCA of only extant songs  
304 (hereafter ‘crown age’, including song present in the last year of sampling). Following Muff et al. (2021), we  
305 communicate our statistical results in the language of evidence.

306 **Results**

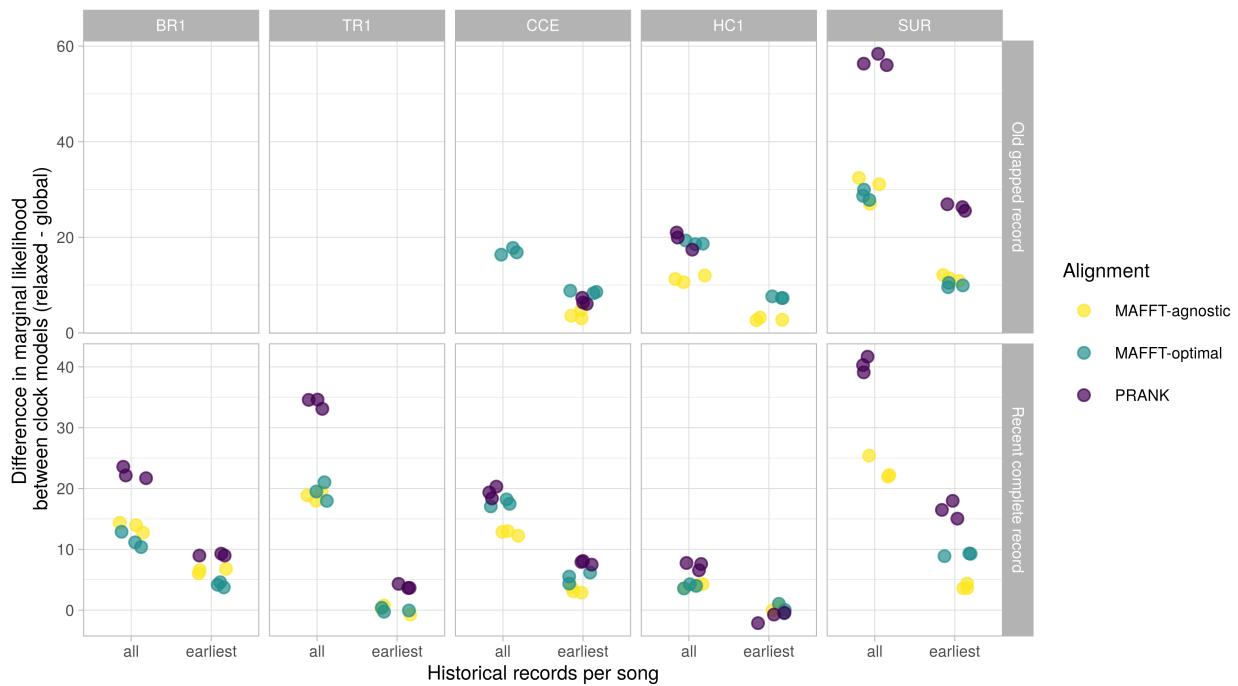
307 We used the FBDP to model cultural change and diversification in five independent leks of the Long-billed  
308 Hermit. These natural leks differ markedly in size and in the length and completeness of their historical  
309 sampling record (Fig. 1d; 2). MCMC model diagnostics reflect this disparity. The MCMC analysis of the  
310 smallest lek (BR1) under the global clock model was the only scenario in which all model parameters met out  
311 mixing (autocorrelation < 0.1) and convergence ( $\text{psrf} < 1.05$ ) criteria (Table S1). In larger leks and more  
312 complex models (i.e. relaxed clock) more model parameters, especially branch rate parameters, failed the  
313 mixing, and to a lesser extent, the convergence criteria (Table S1). Nonetheless, in all but one analysis (SUR  
314 lek using the MAFFT-optimal alignment and all observations of all historical records and a global clock),  
315 estimation of the most relevant model parameters (likelihood, root age, crown age and diversification rates)  
316 converged between independent runs (Table S1). We therefore used these models to investigate absolute  
317 model fit, relative fit of alternative clock models, and the effects of alignment strategies and the historical  
318 record on topological convergence and diversification inferences. We return to the caveats in parameter  
319 estimation in the Discussion.



**Figure 2.** Maximum *a posteriori* (MAP) trees for each lek under a relaxed clock model and utilizing the entire historical song record and a phylogenetically informed alignment algorithm (PRANK). **a)** BR1, **b)** TR1, **c)** HC1, **d)** CCE, **e)** SUR.

### 320 Clock model selection

321 A relaxed clock model of song evolution (in which different song lineages evolve at different rates) was  
 322 generally supported, but the strength of this support depended on the historical record and sampling strategy.  
 323 When historical data was entirely excluded, there was no increase in ML by relaxing the clock model (Fig 3).  
 324 However, the use of historical songs akin to fossils resulted in a higher fit of the relaxed model, particularly  
 325 when all historical records, including identical song sequences sampled in consecutive years, were incorporated  
 326 in the macroevolutionary process (Fig 3). While this trend was present in most leks and data sets, it tended  
 327 to be stronger under the phylogenetically-informed PRANK-TN93 alignment, especially in the historically  
 328 largest lek (SUR). Hereafter, we present the results of analyses using the relaxed clock model.



**Figure 3.** Relative fit of relaxed and global clock models in phylogenetic inference of song evolution under the FBDP. The fit of alternative clock models was compared in five leks (BR1, TR1, HC1, CCE, and SUR) using marginal likelihood approximation. Two leks (BR1 and TR1) lacked deep historical data (no sampling "including oldest records") and three leks contained enough extant songs to estimate clock model fit using a BDP (without historical data).

### 329 Model reliability

330 Simulated data under our inference models accurately reflected certain features of most empirical data sets.  
 331 Namely, the number of segregating and invariant sites, the maximum length of variable and invariant blocks  
 332 and the minimum pairwise distance between sequences were generally in agreement between simulated and  
 333 empirical data (Table 1; Fig. 4; Fig. S3-S32). The two measures of song element diversity (Watterson's  $\theta$

334 and  $\pi$ ), analogous to population genetic diversity, were reliably modeled with the exception of about a third  
 335 of the PRANK alignments, and four MAFFT-agnostic alignments, in which simulated data underestimated  
 336 diversity (Fig. 5; Fig. S23-S29). Consequently, several of these alignments also resulted in relatively low  
 337 values of Tajima's D in simulated data, consistent with our models overestimating the effect of drift (Fig.  
 338 S30-32). Finally, the maximum pairwise distance between sequences was the data feature most often in  
 339 conflict between simulated and empirical data sets (Fig. 6), and overall underestimated in simulated data  
 340 (Fig. S20-S22).

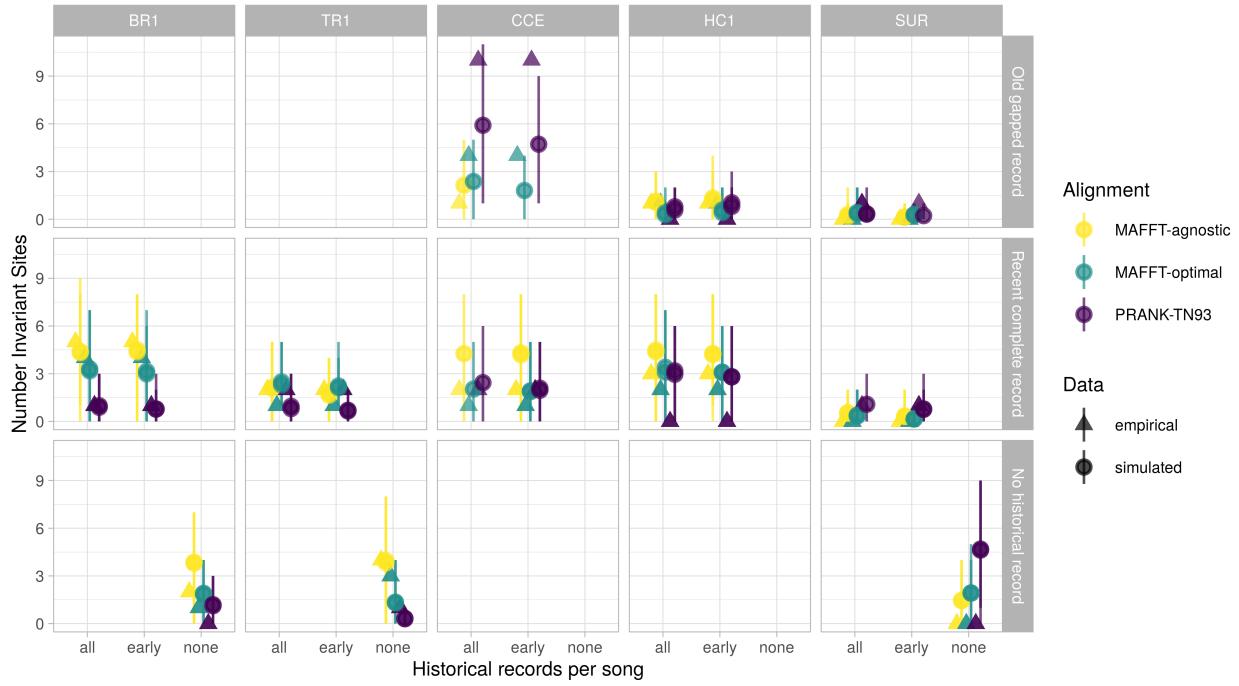
341 Discrepancies between empirical and simulated data were found in all leks and alignment strategies (Fig.  
 342 S3-S32). For the same lek and historical record, PRANK-TN93 alignments tended to result in relatively high  
 343 song-element diversity ( $\pi$ ), in both empirical and simulated data sets. Nonetheless, such high song diversity  
 344 was often underestimated, while the strength of drift in these alignments was overestimated, particularly in  
 345 the TR1 lek (Fig. 5; Fig. S27-32). Similarly, where estimates of the maximum pairwise distance between  
 346 songs were highest, mainly in PRANK-TN93 and MAFFT-optimal alignments, such song divergence tended  
 347 to be underestimated (Fig. 6; Fig. S20-22).

348 **Table 1.** Summary of absolute model fit statistics obtained by comparing properties of empirical data to  
 349 posterior predictive data simulations.

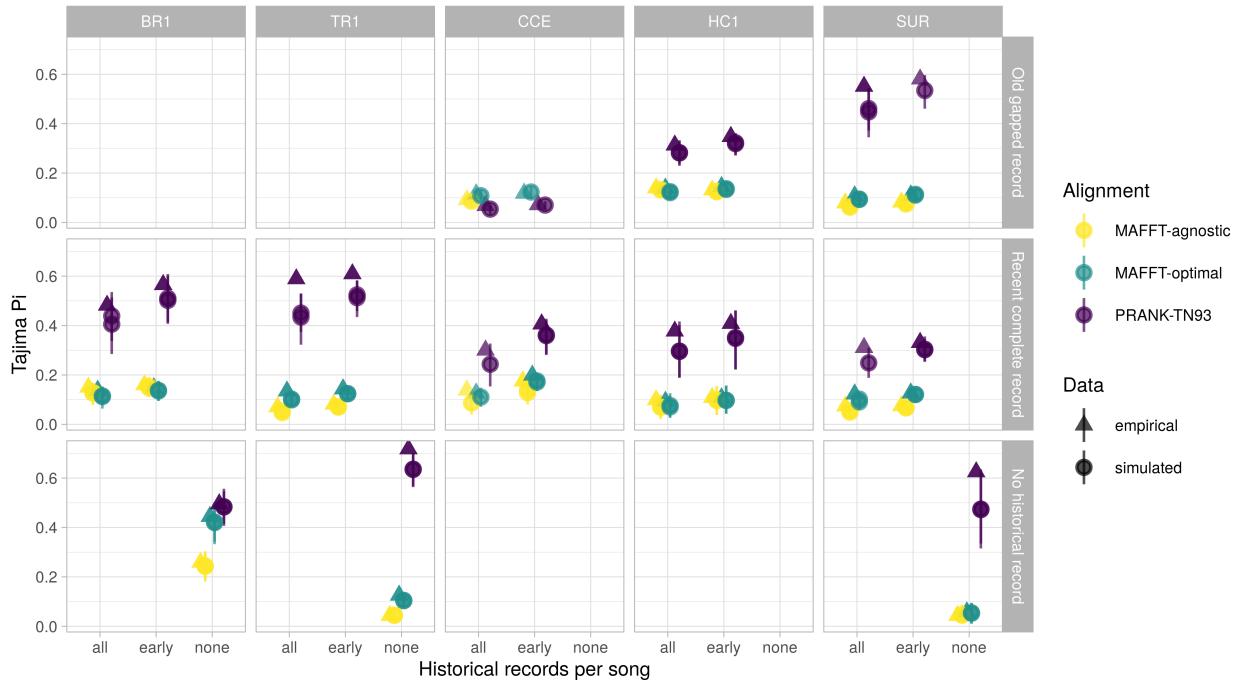
Alignment	Historical record	Mean PPE	SD	Min	Max	Percent P-value > 0.05
MAFFT-agnostic	Old gapped	0.63	0.79	0	3.23	0.97
MAFFT-agnostic	Recent complete	0.82	0.86	0	3.94	0.91
MAFFT-agnostic	None	0.53	0.55	0	2.10	1.00
MAFFT-optimal	Old gapped	0.95	0.84	0	3.09	0.95
MAFFT-optimal	Recent complete	0.85	0.83	0	3.93	0.88
MAFFT-optimal	None	1.05	0.74	0	2.89	0.96
PRANK-TN93	Old gapped	1.00	0.87	0	4.37	0.92
PRANK-TN93	Recent complete	0.81	0.88	0	4.06	0.93
PRANK-TN93	None	1.36	1.04	0	3.68	0.70

350 **Treespace congruence**

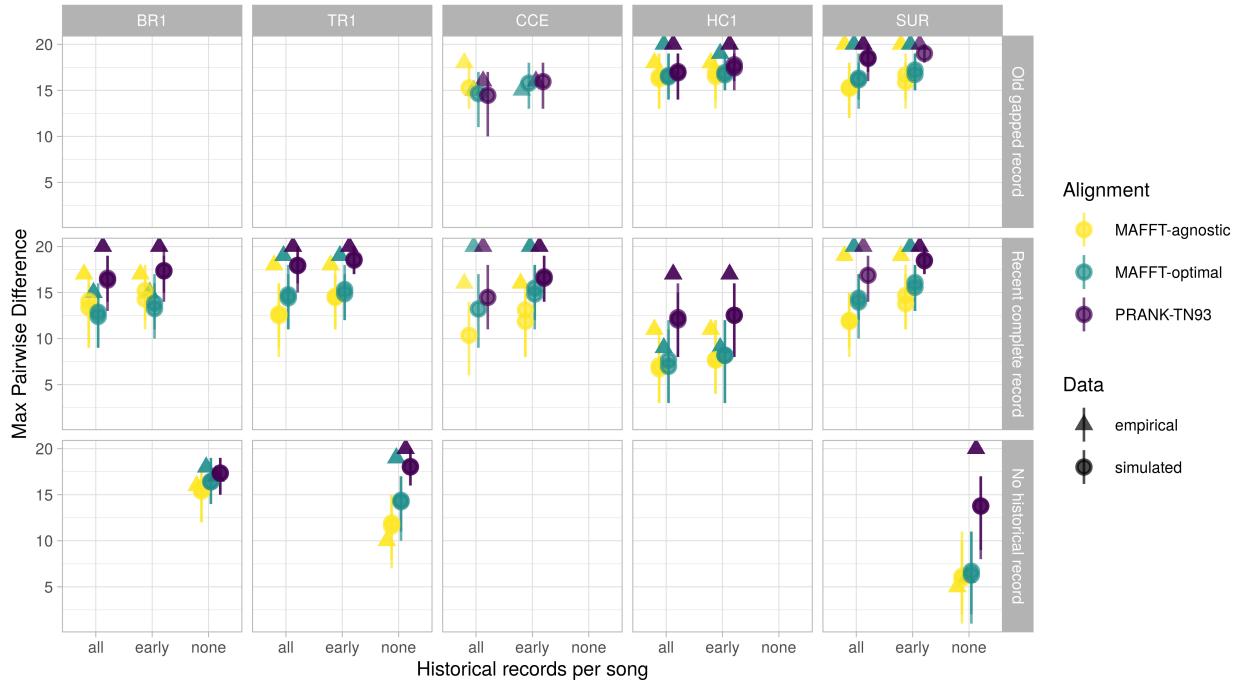
351 The spread of the topological space increased when including all historical records (effect size: 0.031,  
 352 uncertainty interval (UI): 0.023 - 0.040) and when including fossils (effect size: 0.012, UI: 0.005 - 0.019), but



**Figure 4.** Number of invariant sites in empirical song data and posterior predictive data simulations (see Methods). For simulated data, we report the mean and 95% HPD interval of 1500 (BR1, TR1, HC1) or 3000 (CCE and SUR) simulations.



**Figure 5.** Song element diversity ( $\pi$ ) in empirical song data and posterior predictive data simulations (see Methods). For simulated data, we report the mean and 95% HPD interval of 1500 (BR1, TR1, HC1) or 3000 (CCE and SUR) simulations.



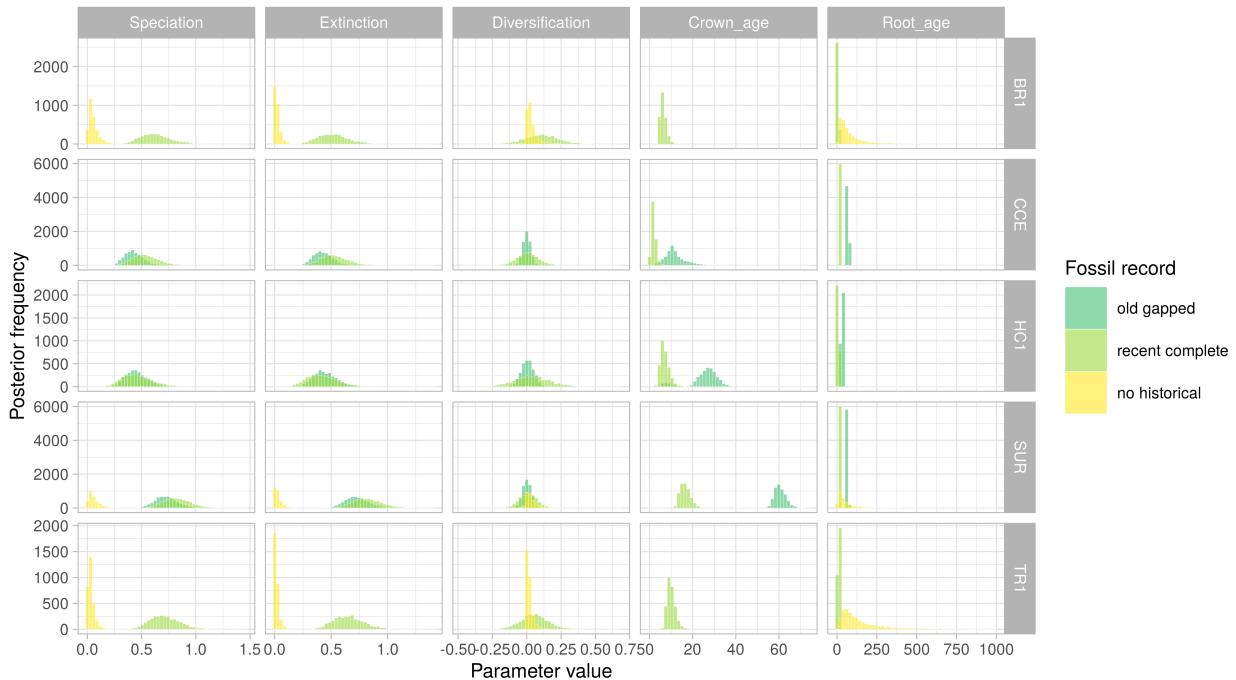
**Figure 6.** Maximum pairwise distance between songs in empirical song data and posterior predictive data simulations (see Methods). For simulated data, we report the mean and 95% HPD interval of 1500 (BR1, TR1, HC1) or 3000 (CCE and SUR) simulations.

353 was not affected by the use of different clock models (effect size: -0.002, UI: -0.009 - 0.005). For the alignment  
 354 strategy only the MAFFT-optimal alignment increased the spread of the topological space compared to  
 355 the PRANK-TN93 alignment (effect size: -0.009, UI: -0.018 - 0.000) no other differences in space size were  
 356 detected between alignment strategies (PRANK-TN93 VS MAFFT-agnostic, effect size: -0.005, UI:-0.014 -  
 357 0.003; MAFFT-optimal VS MAFFT-agnostic, effect size: 0.004, UI: -0.005 - 0.013). The dissimilarity of the  
 358 topological space between models increased by the inclusion of all historical records (effect size: 0.174, UI: 0.130  
 359 - 0.224) and fossils (effect size: 0.176, UI: 0.119 - 0.235), but not by the use of different clock models (effect  
 360 size: 0.030, UI: -0.015 - 0.076). The PRANK-TN93 alignment increased topological space dissimilarity when  
 361 compared to both MAFFT-optimal (effect size: 0.102, UI: 0.045 - 0.159) and MAFFT-agnostic alignments  
 362 (effect size: 0.110, UI: 0.043 - 0.180), but did not differ between the MAFFT alignments (effect size: 0.008,  
 363 UI: -0.049 - 0.067).

#### 364 Diversification dynamics

365 Diversification rates and divergence times were sensitive to historical information. In the three leks in  
 366 which analyses without fossils could be conducted (BR1, TR1, SUR), fossil-free inference resulted in much  
 367 slower speciation and extinction rates and markedly uncertain root ages, in comparison to analyses including

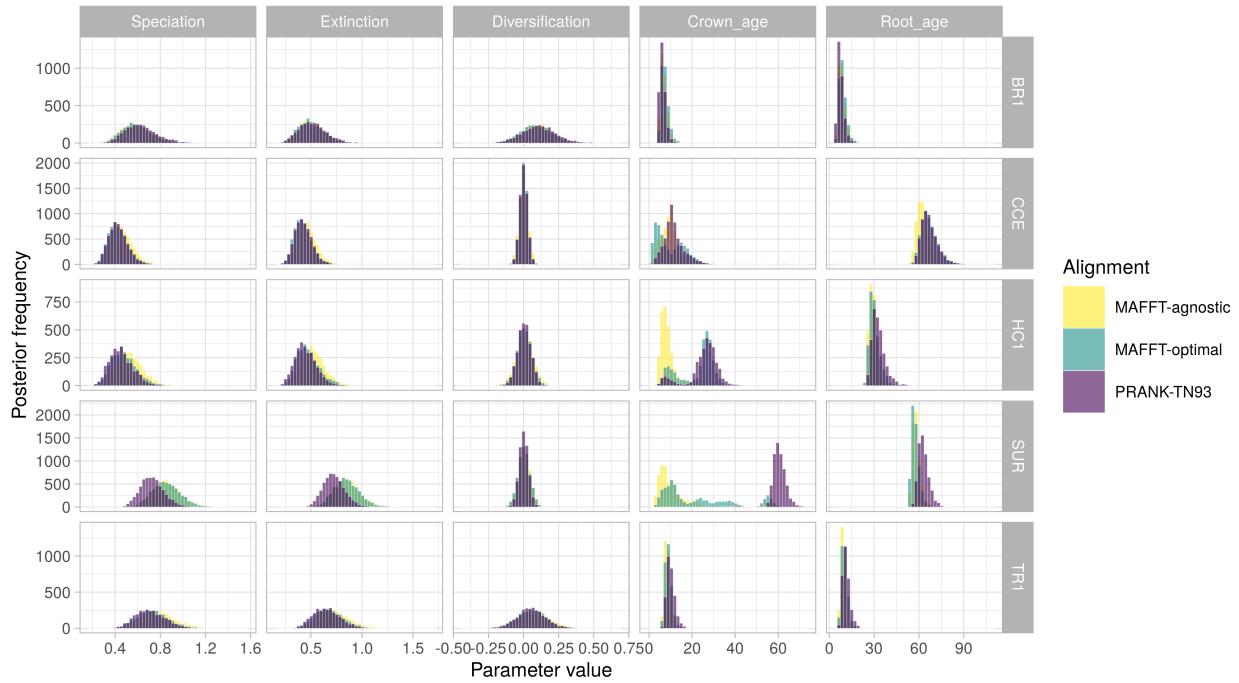
historical songs (Fig. 7; S33-S35). For the three leks in which a long but gapped historical record could be contrasted with a shorter but complete record (CCE, HC1, and SUR), the inclusion of more ancient records resulted in older crown and root age estimates (Fig. 7; S33-S35). In contrast, strong effects of alignment strategies were found in only two leks. In the largest lek (SUR), the PRANK-TN93 alignment led to a moderate decrease in speciation and extinction rate and a relatively strong increase in divergence times (Fig. 8). In the HC1 lek both the PRANK-TN93 and MAFFT-optimal alignments caused older crown age estimates (Fig. 8), compared to the MAFFT-agnostic alignment. Finally, sampling all historical occurrences, as opposed to only the earliest occurrence of each unique song, resulted in increased speciation and extinction rates and slightly more precise estimates of divergence times (Fig. 9; Fig. S36-S38).



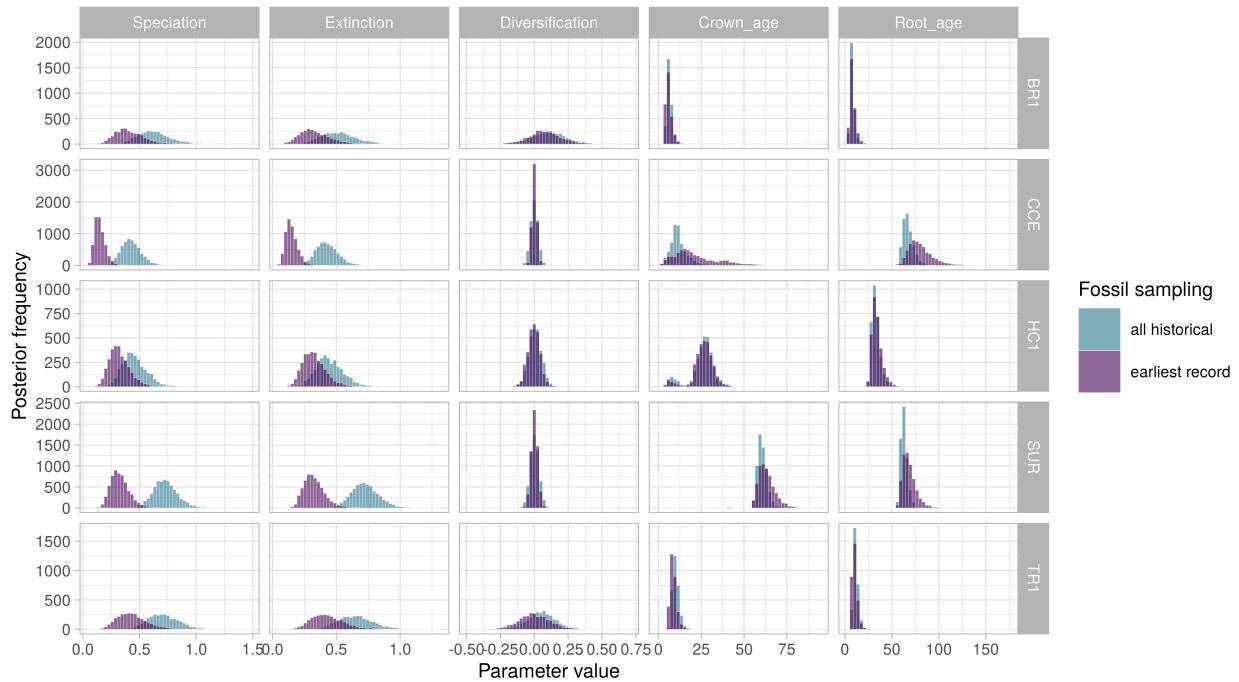
**Figure 7.** Age and diversification estimates under different sets of historical song records in five leks of the Long-billed Hermit. Histograms show posterior distributions of parameter estimates in models using a relaxed molecular clock and the PRANK-TN93 alignment strategy. For data sets including historical song records, all occurrences of such record are included in the analysis.

## 377 Discussion

Many researchers view cultural change as a process akin to organic evolution [REF]. However, a basic prediction of this parallel has largely evaded scrutiny. We typically do not know if evolutionary models and especially phylogenetic models applied to cultural change capture the most prominent features of this process, resulting absolute fit to cultural data [REF exceptions?]. Here, we address this question using the socially learnt songs of the Long-billed Hermit in multiple independent leks, and a phylogenetic model, the fossilized



**Figure 8.** Age and diversification estimates under different alignment strategies in five leks of the Long-billed Hermit. Histograms show posterior distributions of parameter estimates in models using a relaxed molecular clock and all historical records. All available historical records are included in the analysis.



**Figure 9.** Age and diversification estimates of songs in five leks of the Long-billed Hermit. Histograms show posterior distributions of parameter estimates in models using a relaxed molecular clock and the PRANK-TN93 alignment strategy. We compare analyses using all occurrences of historical songs, including identical songs sampled serially, and analyses including each unique song at its earliest occurrence.

383 birth-death process, that is particularly suitable for serially-sampled data. We found that these phylogenetic  
384 models are overall reliable, although they tend to underestimate song element diversity and song divergence  
385 in some scenarios. We also asked how the historical song record, analogous to the fossil record, and alignment  
386 strategies, ported from molecular phylogenetics, influence parameter estimation. In line with theoretical  
387 expectations [REF Heath?], inclusion of historical data resulted in more precise root-age estimates (Fig.  
388 7), a higher rate of lineage divergence (“speciation”) and extinction (Fig. 7). Our analyses with historical  
389 song records also lent support for heterogeneous branch rates (Fig. 3). If available, older historical records  
390 prompted deeper divergence times (Fig. 7). On the other hand, alignment strategies had very strong effects  
391 on tree topology (Fig. ?), and weak to moderate effects on diversification dynamics (Fig. 8).

392 Posterior predictive simulations suggested that the FBDP is a plausible model for cultural change in the  
393 Long-billed Hermit. Nearly 90% of posterior predictive effect sizes (PPEs) between empirical and simulated  
394 data under the FBDP were below 2.00 (mean PPE = 0.84, sd = 0.85, min = 0, max = 4.37; Table 1).  
395 Moreover, only one scenario, in which all historical records were excluded (SUR, PRANK-TN93 alignment),  
396 produced consistent evidence for model inadequacy in data-based posterior predictive checks (Fig. S3-S32).  
397 Even though tests of model adequacy are important safeguards against estimation biases (Carstens et al.  
398 2022), tools for their implementation in Bayesian phylogenetics are relatively new [REFs incl Hohna 2018 and  
399 some Duchene]. To our knowledge, these tools have been advocated [BEAST REf] but not yet utilized in  
400 cultural phylogenetics [but see REFs for other models of cultural evolution]. When applied to molecular and  
401 morphological phylogenetics, posterior predictive checks often reveal a wide range of model fits, depending on  
402 the locus (or trait) and clade analyzed (Richards et al. 2018; May et al. 2021; Khouri et al. 2022). We were  
403 therefore surprised to find that FBD models were similarly adequate in most of our data sets, ranging from  
404 8 ‘taxa’ (TR1 without historical records) to 100 ‘taxa’ (SUR with all available records). Nonetheless, our  
405 phylogenetic models failed to reproduce some diversity features of the data ( $\pi$  and Tajima’s D) in data sets in  
406 which such diversity was relatively high. Interestingly, the main source of variation in song-element diversity  
407 was not the lek identity or the use of historical data, but the alignment strategy (Fig. 5). Phylogenetically  
408 informed PRANK-TN93 alignments were generally more compact (Table S1), thus increasing the average  
409 number of mismatched positions in song alignments, and thus also increasing  $\pi$  and Tajima’s D, but reducing  
410 model fit, compared to MAFFT strategies.

411 In addition to some measures of model adequacy, alignment strategies strongly influenced tree topology (Fig.  
412 ?) and age estimates for the ancestor of extant taxa, in the two leks with deeper historical records (Fig. SUR  
413 and CCE). In leks with historical records that spanned multiple decades (SUR, CCE, HC1), topologies based  
414 on PRANK-TN93 data were markedly distinct from topologies based on phylogenetically naive MAFFT

415 alignments (Fig ?). In smaller leks with more shallow historical records (BR1 and TR1), the two alignment  
416 strategies that enforced a lower substitution rate between vibratory and tonal sounds (PRANK-TN93 and  
417 MAFFT-optimal) resulted in relatively similar topologies (Fig ?). In a recent study that compared alternative  
418 MSA programs for ancestral protein sequence reconstruction, variations of PRANK and MAFFT algorithms  
419 had the highest overall performance (Vialle et al. 2018). However, the two methods were affected in different  
420 ways by the underlying substitution process and characteristics of the tree (Vialle et al. 2018). Because leks  
421 evolve independently, diverse cultural substitution processes and varying tree topologies are also expected to  
422 result in a mixture of alignment performance even within a single species. Our conflicting results between  
423 alignment strategies thus underscore a foremost challenge for cultural phylogenetics with behavioral sequences:  
424 establishing homology in sequence subunits.

425 In most phylogenetic analyses, a multiple sequence alignment is treated as an observation, and alignment  
426 uncertainty is ignored (REF bali-phy for an exception). However, for phenotypic sequences establishing  
427 homology can be a daunting challenge, even when characters are unambiguously coded [Caetano REF].  
428 Furthermore, because culture can change independently of genetic evolution [REF], guiding phenotypic  
429 sequence alignments with independently estimated molecular trees might be problematic. Comparing  
430 alternative alignment strategies across multiple cultural data sets as we have done here, is a step towards  
431 understanding the robustness of cultural phylogenetic analyses. However, as for phylogenetic models,  
432 developing tools to assess the adequacy of alignment methods is crucial for a broader application of phylogenetic  
433 approaches to cultural sequences. In the case of socially learnt bird songs, alignment methods can be informed  
434 by a growing mechanistic understanding of learning, sound production and song function [REFs]. In humans,  
435 alignment methods have been tailored to particular cultural domains and used to identify large-scale patterns  
436 in melody evolution (Savage et al. 2022) and to resolve deep ancestral relationships between language  
437 families (Jäger 2015). In addition to developing specific alignment approaches for different cultural domains,  
438 phylogenetic methods that incorporate alignment uncertainty can be implemented to reduce biases in homology  
439 hypotheses of behavioural sequences (Caetano and Beaulieu 2020).

440 Our analyses of socially learnt bird songs make three further assumptions, dictated by the phylogenetic  
441 models we sought to apply. First, we discretised song spectrograms of variable duration (X - Xs) into 20  
442 characters, each belonging to one of six categories. This conveniently allowed us to model song evolution  
443 as a substitution process, and was justified by the distinctiveness sound categories and high repeatability  
444 of character coding (Fig. Supporting text?). However, focusing on substitutions may obscure some **and**  
445 **perhaps important?** cultural changes in song composition. When learning birds modify their song template  
446 by extending or contracting particular segments, these changes will introduce indels in the alignment which

447 do not contribute to the modelled evolutionary process. **We could use a sentence here about how**  
448 **common these changes in song length are, both on LBH and other birds to then say something**  
449 **about how big of an issue this could be.**

450 Second, we assumed that song elements along a sequence evolve independently of one another. Mechanical,  
451 cognitive and functional constraints on song sequence may falsify this assumption. For example, it is likely  
452 that if a bird slows down the first trill of a series of consecutive trills, subsequent trills will also be replaced  
453 as **biological reason and REF here if there is one**. Similarly, rapid alternations between tonal and  
454 vibratory sounds may be strenuous [REF if this makes any sense], thus biasing substitution rates based  
455 on the identity of nearby sites. The consequences of such site epistasis have been studied mainly in the  
456 context of protein evolution, where structure and function generate coevolutionary dynamics among sites  
457 [REFs]. In protein evolution, incorrectly assuming site independence can result in biased estimation of  
458 site-wise substitution rates [REF], with potential consequences for ancestral sequence reconstruction, clock  
459 rate estimation and topological inference. Properties of songs and other behavioral sequences that arise from  
460 interactions between subunits may create transmission biases for particular subunit combinations [REF?],  
461 potentially resulting in model misspecification of traditional substitution models. Evolutionary models that  
462 relax the assumption of site independence may thus also be of applicability for cultural phylogenetics.

463 Finally, our study assumes that songs diversify in a tree-like manner, with no horizontal transfer. For the  
464 Long-billed Hermit system, this means that juvenile males introduce changes on their learning templates,  
465 independently of other songs that they may have heard in the same lek. We do not know if males of the  
466 Long-billed Hermit indeed copy and potentially modify a single song template, or if new songs can and are  
467 often formed by combining elements from multiple templates. **Maybe a sentence here about what we**  
468 **do know about the learning process.** Nonetheless, it is possible that cultural phylogenetic models are  
469 robust to a degree of horizontal transfer, as suggested by the relatively high absolute fit of these models  
470 to song data (Table 1). An early study based on cladistics showed that data sets across different cultural  
471 domains had a similar fit to a tree-like diversification process as biological data sets [REF Collard]. Further  
472 assessments of the robustness of cultural phylogenetics to observed levels of horizontal transfer and reticulate  
473 evolution are warranted, as well as the exploration of recently developed methods that can accommodate  
474 such processes [REF to that french paper].

475 Culture can change rapidly compared to morphology [REF]. Thus, cultural phylogenetics studies can often tap  
476 on a relatively rich historical record, obtained from archeological preservation (in humans) or from longitudinal  
477 studies. Historical records in our longitudinal sampling of Long-billed Hermit songs have some advantages  
478 over traditional fossils, such as complete character sequences and no stratigraphic uncertainty. Our study also

479 shares some of the challenges of phylogenetic inference with fossils, like time-heterogeneous preservation. We  
480 thus expected that historical song records would impact estimates of node ages and diversification rates to at  
481 least the same extent as total-evidence approaches under the FBDP influence the same parameters in organic  
482 lineages. We found that incorporating historical data resulted in more precise root age estimates, as seen  
483 across total-evidence dating analyses under the FBDP [REFs to Total evidence]. Furthermore, by sampling  
484 ancestral lineages that formed and went extinct before the present, the historical song record also accelerated  
485 estimates of both speciation and extinction, but without effects on net diversification (Fig. 7). Our results  
486 therefore suggest that unlike organic evolution, which typically follows a trend of increasing diversity over  
487 time, social factors, such as sexual selection and kin recognition, may impose net-zero diversification in stable  
488 cultural systems [REF?].

489 We were also interested in understanding how decisions on how to use available historical records would  
490 impact estimation of node ages. We found that including older but sporadically sampled records resulted in  
491 older estimates of crown and root ages for all leks with such data (CCE, HC1 and SUR; Fig. 7). In contrast,  
492 accounting for all historical occurrences, including identical sequences sampled on different years, increased  
493 only the precision of age estimates (crown and origin) in the two largest leks (CCE and SUR; Fig. 9). In  
494 fact, the main consequence of using all historical occurrences was to increase support for a relaxed clock  
495 model (Fig. 3), as clock rates must differ between branches separating identical songs and branches in which  
496 substitutions occurred.

497 Simulation studies show that accounting for all possible fossil occurrences [REF OReily] and sampling fossils at  
498 regular time intervals [REF 2020Sim] bolster the accuracy of phylogenetic inference of organic evolution under  
499 the FBDP. These sampling practices, especially the latter, may be prohibitive for some organic clades. Here,  
500 we based our analyses on pre-existing data, most of which was collected before the methods we used were even  
501 developed. However, future cultural phylogenetics studies based on longitudinal data can incorporate these  
502 best sampling strategies in their design, and more systematically assess how the availability and regularity of  
503 historical records impact age estimates of cultural lineages.

## 504 **Conclusions**

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508 **Author contributions**

509 **B.W.** and **M.S.A.** conceived the study. **M.S.A.** collected the data with support from **A.R.G.. M.S.A.**  
510 and **B.W.** analyzed the data and wrote the manuscript with contributions from **A.R.G..**

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676 **Supporting Material**

677 Figure S1. Examples of songs and coding?

678 **Supporting text 1**

679 How do we arrive to 20 segments

680 **Supporting text 2**

681 Repeatability of sound classification

682 Figure S2. Main repeatability result.

683 Table S1. MCMC diganostics, add alignment legnht and no. of sequences to the table we have

684 Figures S3-S32. All model reliability results

685 Figures S33-S35. Histograms of age and diversification parameters per record type for all alignment strategies.

686 Figures S36-S38. Histograms of age and diversification parameters per sampling strategy for all alignment  
687 strategies.