Variational inference for Gaussian processes

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Curso de entrenamiento ArcelorMittal

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Motivation

Variational inference for sparse GPs

Stochastic variational inference for sparse GPs

"Collapsed" variational inference for sparse GPs

Posterior inference

■ When using Bayesian inference, we need to compute the posterior distribution of f given the data.

 We then use that posterior distribution to compute the predictive distribution.

- Reasons as why computing the posterior distribution is an issue for GPs.
 - Computational complexity.
 - Non-Gaussian likelihood.
 - Both of the above.

Computational complexity

□ To compute the predictive mean and the predictive covariance we need to compute $\left[\mathbf{K}(\mathbf{X},\mathbf{X}) + \sigma_n^2 \mathbf{I}\right]^{-1}$

The usual way to do this is using the Cholesky decomposition which costs $\mathcal{O}(n^3)$.

If n = 1000, then we need to perform 10^9 operations.

Non-Gaussian likelihoods

In Bayesian inference we want to compute

$$p(\mathbf{f}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{f})}{p(\mathbf{y})},$$

where $p(\mathbf{y}) = \int p(\mathbf{y}|\mathbf{f})p(\mathbf{f})d\mathbf{f}$.

When p(y|f) is a Gaussian likelihood, then we can compute p(y) and p(f|y) analytically.

When p(y|f) is non-Gaussian (e.g. Bernoulli with a sigmoid link function) both p(y) and p(f|y) are intractable.

How to address these issues?

 A successful approach uses the idea of inducing variables or pseudo-variables.

The idea in itself was quite well know in the GP literature. See for example Chapter 8 of the GPML book and in the paper Quiñonero-Candela and Rasmussen (2005).

 However, if we couple this idea with a variational inference approach, we have a powerful tool to build complex GP models.

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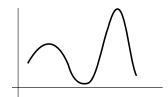
- We introduce a new set of M variables $\mathbf{u} = \{u(\mathbf{z}_m)\}_{m=1}^M$ that we refer to as inducing variables or pseudo-variables.
- □ The set of points $\mathbf{Z} = \{\mathbf{z}_m\}_{m=1}^M$ is usually known as inducing inputs.
- We augment the original prior $p(\mathbf{f})$ to $p(\mathbf{f}, \mathbf{u})$ such that

$$p(\mathbf{f}) = \int p(\mathbf{f}, \mathbf{u}) d\mathbf{u} = \int p(\mathbf{f}|\mathbf{u}) p(\mathbf{u}) d\mathbf{u},$$

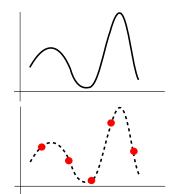
where $p(\mathbf{u})$ and $p(\mathbf{f}, \mathbf{u})$ are both Gaussians.

- The auxiliary variables \mathbf{u} can be part of the GP $f(\mathbf{x})$ or they can be linearly related to $f(\mathbf{x})$ (sometimes known as interdomain inducing variables).
- □ In the former case, $p(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\mathbf{0}, \mathbf{K}(\mathbf{Z}, \mathbf{Z}))$.

A sample from p(f)

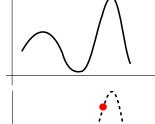


A sample from p(f)

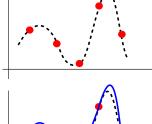


Inducing variables ${\bf u}$

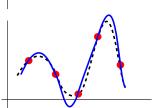
A sample from p(f)



Inducing variables ${\bf u}$



A sample from $p(f|\mathbf{u})$



Variational lower-bound for the marginal likelihood (I)

□ We can write the log marginal probability for **y** using

$$\log p(\mathbf{y}) = \mathcal{L}(q(\mathbf{f})) + \mathsf{KL}(q(\mathbf{f}) \| p(\mathbf{f} | \mathbf{y})),$$

where

$$\mathcal{L}(q(\mathbf{f})) = \int q(\mathbf{f}) \log \left\{ rac{p(\mathbf{y}, \mathbf{f})}{q(\mathbf{f})}
ight\} d\mathbf{f},$$
 $\mathsf{KL}(q(\mathbf{f}) || p(\mathbf{f} | \mathbf{y})) = - \int q(\mathbf{f}) \log \left\{ rac{p(\mathbf{f} | \mathbf{y})}{q(\mathbf{f})}
ight\} d\mathbf{f},$

with $q(\mathbf{f})$ the approximated posterior, KL(q||p) is the Kullback-Leibler divergence between q and p and $p(\mathbf{f}|\mathbf{y})$ is the true posterior.

- □ The KL divergence is zero when q = p. In that case, $\log p(\mathbf{y}) = \mathcal{L}(q(\mathbf{f}))$.
- □ If this is not the case $KL(q(\mathbf{f})||p(\mathbf{f}|\mathbf{y}))$ and $\log p(\mathbf{y}) > \mathcal{L}(q(\mathbf{f}))$.



Variational lower-bound for the marginal likelihood (II)

- \Box We have two ways to find the optimal $q(\mathbf{f})$
 - 1. We find $q(\mathbf{f})$ by minimising $KL(q(\mathbf{f})||p(\mathbf{f}|\mathbf{y}))$.
 - 2. We find $q(\mathbf{f})$ by maximising $\mathcal{L}(q(\mathbf{f}))$.
- Option 1 is not possible since $p(\mathbf{f}|\mathbf{y})$ is unknown.
- So, in general, we appeal to option 2

$$\log p(\mathbf{y}) \ge \mathcal{L}(q(\mathbf{f})).$$

Lower-bound with inducing variables

For our augmented model we want to find an approximated posterior $q(\mathbf{f}, \mathbf{u})$ by maximising

$$\mathcal{L}(q(\mathbf{f}, \mathbf{u})) = \int \int q(\mathbf{f}, \mathbf{u}) \log \left\{ \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} \right\} d\mathbf{u} d\mathbf{f},$$

where $p(\mathbf{y}, \mathbf{f}, \mathbf{u}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})p(\mathbf{u})$ and we will approximate the posterior $q(\mathbf{f}, \mathbf{u})$ as $q(\mathbf{f}, \mathbf{u}) \approx p(\mathbf{f}|\mathbf{u})q(\mathbf{u})$.

Since we know that $q(\mathbf{f}) = \int_{\mathbf{u}} p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) d\mathbf{u}$, the bound above really only depends on $q(\mathbf{u})$

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{f}|\mathbf{u}) \rho(\mathbf{u})}{\rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}. \end{split}$$



Two approaches for optimising $\mathcal{L}(q(\mathbf{u}))$

- □ There are two approaches for optimising $q(\mathbf{u})$ in $\mathcal{L}(q(\mathbf{u}))$.
- First approach (Hensman et al., 2013):
 - We assume a multi-variate Gaussian form for $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\mu, \mathbf{S})$ with $\mu \in \mathbb{R}^{M \times 1}$ and $\mathbf{S} \in \mathbb{R}^{M \times M}$.
 - We then find μ and **S** by numerically optimising $\mathcal{L}(q(\mathbf{u}))$.
- Second approach (Titsias, 2009):
 - We marginalise q(u) from the bound and then compute it by using Jensen's inequality.
 - We then find μ and **S** by using the rules of probability.

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Stochastic variational inference

 Stochastic variational inference (SVI) allows (Hoffman et al., 2013) the use of stochastic gradients over variational lower bounds.

Hensman et al. (2013) proposed the use of SVI for sparse GPs.

The idea is to use stochastic gradients for optimising $\mathcal{L}(q(\mathbf{u}))$ with respect to $q(\mathbf{u})$, this is, μ and \mathbf{S} .

Lower bound $\mathcal{L}(q(\mathbf{u}))$ (I)

From a previous slide,

$$\mathcal{L}(q(\mathbf{u})) = \int \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \left\{ \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u}d\mathbf{f}.$$

We can re-arrange the expression above using

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \left[\log \rho(\mathbf{y}|\mathbf{f}) + \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} \right] d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \rho(\mathbf{y}|\mathbf{f}) d\mathbf{u} d\mathbf{f} + \int q(\mathbf{u}) \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u}, \\ &= \int \log \rho(\mathbf{y}|\mathbf{f}) \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) d\mathbf{u} d\mathbf{f} + \int q(\mathbf{u}) \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u}, \\ &= \int \log \rho(\mathbf{y}|\mathbf{f}) q(\mathbf{f}) d\mathbf{f} - \mathrm{KL}(q(\mathbf{u}) || p(\mathbf{u})) d\mathbf{u}, \\ &= \mathbb{E}_{q(\mathbf{f})}[\log \rho(\mathbf{y}|\mathbf{f})] - \mathrm{KL}(q(\mathbf{u}) || p(\mathbf{u})). \end{split}$$

Lower bound $\mathcal{L}(q(\mathbf{u}))$ (II)

The lower bound is

$$\mathcal{L}(q(\mathbf{u})) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f})] - \mathsf{KL}(q(\mathbf{u})||p(\mathbf{u})).$$

 \Box We first compute $q(\mathbf{f})$, which is given as

$$q(\mathbf{f}) = \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u})d\mathbf{u},$$

where

$$\begin{split} & \rho(\mathbf{f}|\mathbf{u}) = \mathcal{N}(\mathbf{f}|\mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{u}, \mathbf{K}(\mathbf{X},\mathbf{X}) - \mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{K}^{\top}(\mathbf{X},\mathbf{Z})) \\ & q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\boldsymbol{\mu},\mathbf{S}). \end{split}$$

Leading to

$$q(\mathbf{f}) = \mathcal{N}(\mathbf{f}|\mathbf{K}(\mathbf{X},\mathbf{Z})[\mathbf{K}(\mathbf{Z},\mathbf{Z})]^{-1}\boldsymbol{\mu},\boldsymbol{\Lambda}),$$

with

$$\boldsymbol{\Lambda} = \boldsymbol{K}(\boldsymbol{X},\boldsymbol{X}) + \boldsymbol{K}(\boldsymbol{X},\boldsymbol{Z})\boldsymbol{K}^{-1}(\boldsymbol{Z},\boldsymbol{Z})\left(\boldsymbol{S} - \boldsymbol{K}(\boldsymbol{Z},\boldsymbol{Z})\right)\boldsymbol{K}^{-1}(\boldsymbol{Z},\boldsymbol{Z})\boldsymbol{K}^{\top}(\boldsymbol{X},\boldsymbol{Z})\right)$$

Lower bound $\mathcal{L}(q(\mathbf{u}))$ (III)

The first term of the lower bound $\mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f})]$ is then given as

$$\mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f})] = \mathbb{E}_{q(\mathbf{f})}\left[-\frac{n}{2}\log 2\pi - \frac{n}{2}\log \sigma_n^2 - \frac{1}{2\sigma_n^2}(\mathbf{y} - \mathbf{f})^\top(\mathbf{y} - \mathbf{f})\right].$$

■ We are only interested in the terms involving f, then

$$\begin{split} \mathbb{E}_{q(\mathbf{f})}[\log \rho(\mathbf{y}|\mathbf{f})] &= \frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbb{E}_{q(\mathbf{f})}(\mathbf{f}) - \frac{1}{2\sigma_n^2} \mathbb{E}_{q(\mathbf{f})}(\mathbf{f}^\top \mathbf{f}) + \text{const} \\ &= \frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mu - \frac{1}{2\sigma_n^2} \mathbb{E}_{q(\mathbf{f})} [\text{tr}(\mathbf{f}\mathbf{f}^\top)] + \text{const} \\ &= \frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mu - \frac{1}{2\sigma_n^2} \operatorname{tr}(\mathbb{E}_{q(\mathbf{f})} [\mathbf{f}\mathbf{f}^\top]) + \text{const} \\ &= \frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mu - \frac{1}{2\sigma_n^2} \operatorname{tr}(\Lambda + \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mu \mu^\top [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mathbf{K}(\mathbf{Z}, \mathbf{X})) + \text{const}. \end{split}$$

Lower bound $\mathcal{L}(q(\mathbf{u}))$ (IV)

- □ The second term of the lower bound is $KL(q(\mathbf{u})||p(\mathbf{u}))$.
- The KL divergence between two multivariate Gaussians $\mathcal{N}_0(\mu_0, \Sigma_0)$ and $\mathcal{N}_1(\mu_1, \Sigma_1)$, KL $(\mathcal{N}_0 || \mathcal{N}_1)$ is given as

$$\frac{1}{2}\left[\mathsf{tr}(\boldsymbol{\Sigma}_1^{-1}\boldsymbol{\Sigma}_0) + (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)^{\top}\boldsymbol{\Sigma}_1^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0) - k + \log\frac{|\boldsymbol{\Sigma}_1|}{|\boldsymbol{\Sigma}_0|}\right].$$

□ In this case, $q(\mathbf{u}) \sim \mathcal{N}(\mu, \mathbf{S})$ and $p(\mathbf{u}) \sim \mathcal{N}(\mathbf{0}, \mathbf{K}(\mathbf{Z}, \mathbf{Z}))$, and $\mathsf{KL}(q(\mathbf{u}) \| p(\mathbf{u}))$ is then given as

$$\frac{1}{2}\left[\operatorname{tr}(\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{S}) + \boldsymbol{\mu}^{\top}\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\boldsymbol{\mu} - n + \log\frac{|\mathbf{K}(\mathbf{Z},\mathbf{Z})|}{|\mathbf{S}|}\right]$$



Lower bound $\mathcal{L}(q(\mathbf{u}))$ (V)

Putting both terms together, we get the following expression for the lower bound $\mathcal{L}(q(\mathbf{u}))$,

$$\begin{split} &\frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \boldsymbol{\mu} \\ &- \frac{1}{2\sigma_n^2} \operatorname{tr}(\mathbf{K}(\mathbf{X}, \mathbf{X}) + \mathbf{K}(\mathbf{X}, \mathbf{Z}) \mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z}) \left(\mathbf{S} - \mathbf{K}(\mathbf{Z}, \mathbf{Z}) \right) \mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z}) \mathbf{K}^\top (\mathbf{X}, \mathbf{Z}) \right) \\ &+ \frac{1}{2\sigma_n^2} \operatorname{tr}(\mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \boldsymbol{\mu} \boldsymbol{\mu}^\top [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mathbf{K}(\mathbf{Z}, \mathbf{X})) \\ &- \frac{1}{2} \operatorname{tr}(\mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z}) \mathbf{S}) - \frac{1}{2} \boldsymbol{\mu}^\top \mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z}) \boldsymbol{\mu} - \frac{1}{2} \log \frac{|\mathbf{K}(\mathbf{Z}, \mathbf{Z})|}{|\mathbf{S}|} + \operatorname{const.} \end{split}$$

- □ We can find an estimate for μ and **S** by maximising $\mathcal{L}(q(\mathbf{u}))$ using numerical optimisation.
- We need to compute the gradients

$$rac{\partial \mathcal{L}}{\partial oldsymbol{\mu}}, \; rac{\partial \mathcal{L}}{\partial oldsymbol{\mathsf{S}}}$$



Partial derivative $\frac{\partial \mathcal{L}}{\partial \mu}$

 \Box The terms in $\mathcal{L}(q(\mathbf{u}))$ that depend on μ are

$$\begin{split} &\frac{1}{\sigma_n^2} \mathbf{y}^\top \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \boldsymbol{\mu} \\ &+ \frac{1}{2\sigma_n^2} \boldsymbol{\mu}^\top [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mathbf{K}(\mathbf{Z}, \mathbf{X}) \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \boldsymbol{\mu} \\ &- \frac{1}{2} \boldsymbol{\mu}^\top \mathbf{K}^{-1} (\mathbf{Z}, \mathbf{Z}) \boldsymbol{\mu}. \end{split}$$

 $lue{}$ Taking the derivative with respect to μ leads to

$$\begin{split} \frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}} &= \frac{1}{\sigma_n^2} \mathbf{K}(\mathbf{Z}, \mathbf{Z})^{-1} \mathbf{K}(\mathbf{Z}, \mathbf{X}) \mathbf{y} \\ &+ \frac{1}{\sigma_n^2} [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \mathbf{K}(\mathbf{Z}, \mathbf{X}) \mathbf{K}(\mathbf{X}, \mathbf{Z}) [\mathbf{K}(\mathbf{Z}, \mathbf{Z})]^{-1} \boldsymbol{\mu} - \mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z}) \boldsymbol{\mu}. \end{split}$$

Partial derivative $\frac{\partial \mathcal{L}}{\partial \mathbf{S}}$

lacktriangle The terms in $\mathcal{L}(q(\mathbf{u}))$ that depend on **S** are

$$\begin{split} &-\frac{1}{2\sigma_n^2}\operatorname{tr}(\mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{S}\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{K}^{\top}(\mathbf{X},\mathbf{Z})))\\ &-\frac{1}{2}\operatorname{tr}(\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{S})+\frac{1}{2}\log|\mathbf{S}|. \end{split}$$

Taking the derivative with respect to S leads to

$$\frac{\partial \mathcal{L}}{\partial \boldsymbol{S}} = -\frac{1}{2\sigma_n^2} \boldsymbol{K}^{-1}(\boldsymbol{Z}, \boldsymbol{Z}) \boldsymbol{K}(\boldsymbol{Z}, \boldsymbol{X}) \boldsymbol{K}(\boldsymbol{X}, \boldsymbol{Z}) \boldsymbol{K}^{-1}(\boldsymbol{Z}, \boldsymbol{Z}) - \frac{1}{2} \boldsymbol{K}^{-1}(\boldsymbol{Z}, \boldsymbol{Z}) + \frac{1}{2} \boldsymbol{S}^{-1}.$$

Stochastic gradient descent

It can be shown that the lower bound can be written as

$$\mathcal{L}(q(\mathbf{u})) = \sum_{i=1}^{n} \ell(y_i, \mathbf{x}_i, \theta) - \mathsf{KL}(q(\mathbf{u}) || p(\mathbf{u})),$$

where $\ell(y_i, \mathbf{x}_i, \theta)$ is a function that depends on the data, the variational parameters μ , **S**, and any other (hyper) parameters in the model (e.g. the hyperparameters of the kernel).

- □ For *n* large, we could only use a subset of the data to compute the gradients to be used in numerical optimisation.
- This is usually known as stochastic gradient descent.
- The computational complexity is this model is $\mathcal{O}(nM^2)$, where M is the number of inducing points.

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Marginalising $q(\mathbf{u})$ from the bound (I)

- □ Instead of finding particular parameters μ and \mathbf{S} as before, we can marginalise $q(\mathbf{u})$ from the lower bound and then use probability rules to compute $q(\mathbf{u})$.
- Let us go back to the general expression for the bound

$$\mathcal{L}(q(\mathbf{u})) = \int \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \left\{ \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u}d\mathbf{f}.$$

- Let us assume that $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\boldsymbol{\mu},\mathbf{S})$ and find expressions for \mathbf{u} and \mathbf{S} .
- We first integrate over f.

Marginalising $q(\mathbf{u})$ from the bound (II)

The bound can be expressed as

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f} \\ &= \int q(\mathbf{u}) \int \rho(\mathbf{f}|\mathbf{u}) \left\{ \log \rho(\mathbf{y}|\mathbf{f}) + \log \left[\frac{\rho(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{f} d\mathbf{u}. \end{split}$$

Let us focus on the integral over f

$$\log T(\mathbf{y}, \mathbf{u}) = \int \log p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})d\mathbf{f}.$$

Marginalising $q(\mathbf{u})$ from the bound (III)

The bound can now be expressed as

$$\mathcal{L}(q(\mathbf{u})) = \int q(\mathbf{u}) \left\{ \log \mathcal{N}(\mathbf{y}|\alpha, \sigma_n^2 \mathbf{I}) - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}}) + \log \left[\frac{\rho(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u},$$

where

$$\begin{split} \boldsymbol{\alpha} &= K(\boldsymbol{X},\boldsymbol{Z})K^{-1}(\boldsymbol{Z},\boldsymbol{Z})\boldsymbol{u} \\ \widetilde{K} &= K(\boldsymbol{X},\boldsymbol{X}) - K(\boldsymbol{X},\boldsymbol{Z})K^{-1}(\boldsymbol{Z},\boldsymbol{Z})K(\boldsymbol{X},\boldsymbol{Z})^{\top}. \end{split}$$

It follows that

$$\mathcal{L}(q(\mathbf{u})) = \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|lpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight]
ight\} d\mathbf{u} - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}})$$



Jensen's inequality

 \Box A function φ is *convex* if

$$\varphi(\lambda a + (1 - \lambda)b) \le \lambda \varphi(a) + (1 - \lambda)\varphi(b).$$

 \Box A function φ is *concave* if

$$\varphi(\lambda a + (1 - \lambda)b) \ge \lambda \varphi(a) + (1 - \lambda)\varphi(b).$$

Let φ be a convex function. It can be shown that

$$arphi(\mathbb{E}(\mathbf{x})) \leq \mathbb{E}(arphi(\mathbf{x}))$$
 $arphi\left(\int \mathbf{x} p(\mathbf{x}) d\mathbf{x}\right) \leq \int arphi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x}.$

This inequality is known as the Jensen's inequality.

 \Box If φ is a concave function then

$$arphi(\mathbb{E}(\mathbf{x})) \geq \mathbb{E}(arphi(\mathbf{x}))$$
 $arphi\left(\int \mathbf{x} p(\mathbf{x}) d\mathbf{x}\right) \geq \int arphi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x}.$

Jensen's inequality applied to $\mathcal{L}(q(\mathbf{u}))$

Reversing Jensen's inequality, we can write

$$\begin{split} &\log\left[\int q(\mathbf{u})\left[\frac{\mathcal{N}(\mathbf{y}|\boldsymbol{\alpha},\sigma_n^2\mathbf{I})\boldsymbol{p}(\mathbf{u})}{q(\mathbf{u})}\right]d\mathbf{u}\right] \geq \\ &\int q(\mathbf{u})\left\{\log\left[\frac{\mathcal{N}(\mathbf{y}|\boldsymbol{\alpha},\sigma_n^2\mathbf{I})\boldsymbol{p}(\mathbf{u})}{q(\mathbf{u})}\right]\right\}d\mathbf{u} \end{split}$$

The expression above can be simplified as

$$\log \left[\int \mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] \ge \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u}$$

A tighter bound $\mathcal{L}(q(\mathbf{u}))$

Reversing Jensen's inequality, we can write

$$egin{aligned} \log \left[\int q(\mathbf{u}) \left[rac{\mathcal{N}(\mathbf{y} | lpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight] d\mathbf{u}
ight] & \geq \ \int q(\mathbf{u}) \left\{ \log \left[rac{\mathcal{N}(\mathbf{y} | lpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight]
ight\} d\mathbf{u} \end{aligned}$$

The expression above can be simplified as

$$\log \left[\int \mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] \ge \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u}$$

If we define

$$\mathcal{L}_2 = \log \left[\int \mathcal{N}(\mathbf{y}|\alpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}}),$$

then

$$\mathcal{L}_2 \geq \mathcal{L}(q(\mathbf{u})).$$

And \mathcal{L}_2 is closer to $\log p(\mathbf{y})$ than $\mathcal{L}(q(\mathbf{u}))$.



Expression for $q(\mathbf{u})$

We can compute $q(\mathbf{u})$ using Bayes theorem and the properties of Gaussians.

Starting with $p(\mathbf{y}|\mathbf{u})$ and $p(\mathbf{u})$, we use Bayes theorem to compute $q(\mathbf{u})$ $q(\mathbf{u}|\mathbf{y}) \propto p(\mathbf{y}|\mathbf{u})p(\mathbf{u}).$

Using the properties of the Gaussian distribution, we get

$$q(\mathbf{u}|\mathbf{y}) = \mathcal{N}(\mathbf{u}|\sigma_n^{-2}\mathbf{K}(\mathbf{Z},\mathbf{Z})\mathbf{A}^{-1}\mathbf{K}(\mathbf{Z},\mathbf{X})\mathbf{y},\mathbf{K}(\mathbf{Z},\mathbf{Z})\mathbf{A}^{-1}\mathbf{K}(\mathbf{Z},\mathbf{Z})),$$

where $\mathbf{A} = \mathbf{K}(\mathbf{Z}, \mathbf{Z}) + \sigma_n^{-2} \mathbf{K}(\mathbf{Z}, \mathbf{X}) \mathbf{K}(\mathbf{X}, \mathbf{Z})$.

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