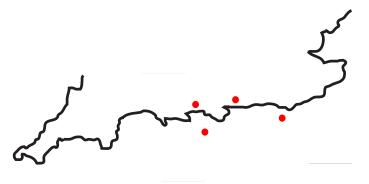
Mauricio A. Álvarez, PhD

Curso de entrenamiento ArcelorMittal

Sensor Network

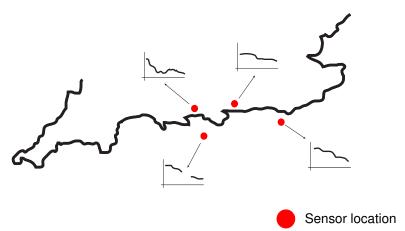
South Coast of England



Sensor location

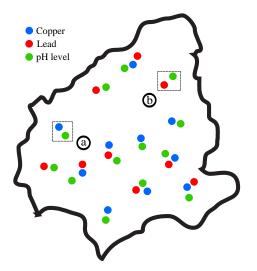
Sensor Network

South Coast of England



Jura Data Set

Region of Swiss Jura



Contents

Dependencies between processes

Intrinsic Coregionalization Model

Semiparametric Latent Factor Model

Linear Model of Coregionalization

Process convolutions

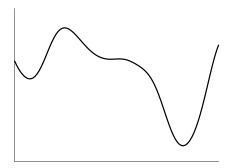
Covariance fitting and Prediction

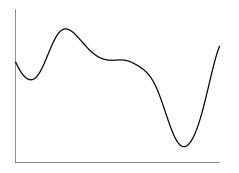
Cokriging

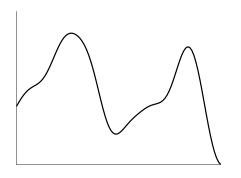
Extensions

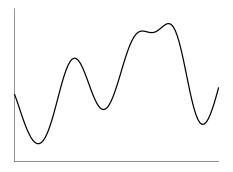
Computational complexity Variations of LMC Variations of PC

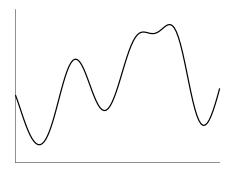
Summary





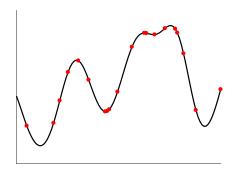






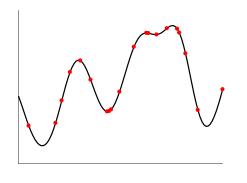
$$f(\boldsymbol{x}) \sim \mathcal{GP}(0, k(\boldsymbol{x}, \boldsymbol{x}'))$$

$$\mathcal{D} = \{(\mathbf{x}_i, f(\mathbf{x}_i)) | i = 1, \dots, N\}$$



$$f(\boldsymbol{x}) \sim \mathcal{GP}(0, k(\boldsymbol{x}, \boldsymbol{x}'))$$

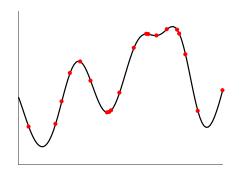
$$\mathcal{D} = \{(\mathbf{x}_i, f(\mathbf{x}_i)) | i = 1, \dots, N\}$$



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$$\begin{bmatrix} f(\mathbf{x}_1) \\ \vdots \\ f(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} \right)$$

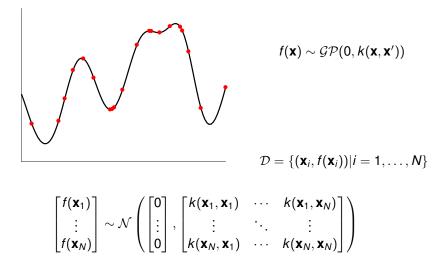


 $f(\boldsymbol{x}) \sim \mathcal{GP}(0, k(\boldsymbol{x}, \boldsymbol{x}'))$

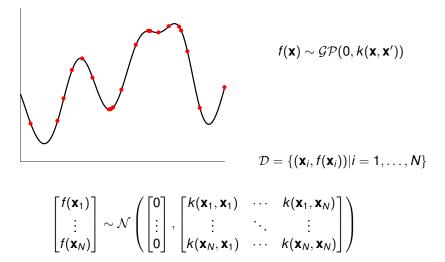
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$$\begin{bmatrix} f(\mathbf{x}_1) \\ \vdots \\ f(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} \right)$$

f



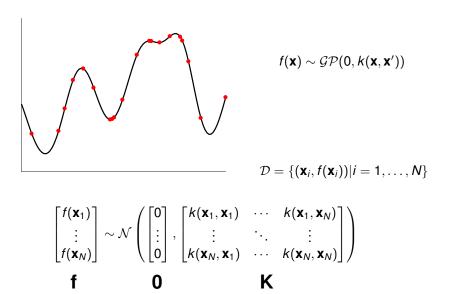
f



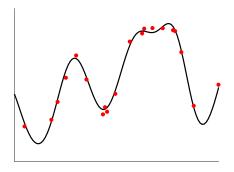
f

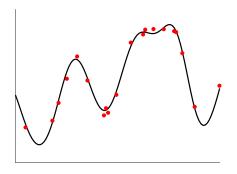
U

K



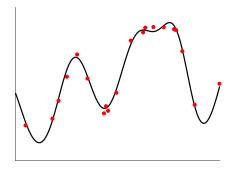
For prediction: $p(f(\mathbf{x}_*)|\mathbf{f})$



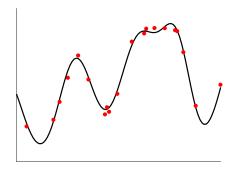


$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

 $y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i$



$$egin{aligned} f(\mathbf{x}) &\sim \mathcal{GP}(\mathbf{0}, k(\mathbf{x}, \mathbf{x}')) \ & y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i \ & \epsilon_i &\sim \mathcal{N}(\mathbf{0}, \sigma^2) \end{aligned}$$

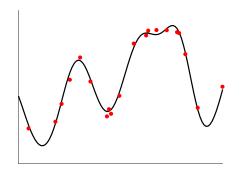


$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

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$$\mathcal{D} = \{(\mathbf{x}_i, y(\mathbf{x}_i)) | i = 1, \dots, N\}$$

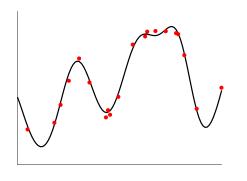


$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

 $y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i$
 $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$

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$$\begin{bmatrix} y(\mathbf{x}_1) \\ \vdots \\ y(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} + \sigma^2 \begin{bmatrix} 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 1 \end{bmatrix} \right)$$



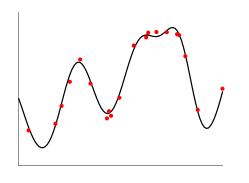
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 $y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i$
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$$\begin{bmatrix} y(\mathbf{x}_1) \\ \vdots \\ y(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} + \sigma^2 \begin{bmatrix} 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 1 \end{bmatrix} \right)$$

у



$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

 $y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i$
 $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$

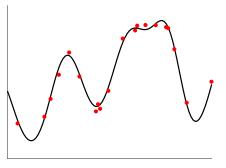
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у







$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

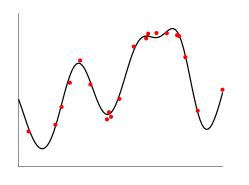
 $y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i$
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У





$$f(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$$

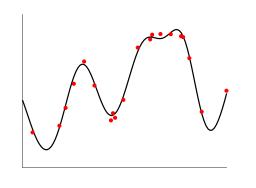
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$$\begin{bmatrix} y(\mathbf{x}_1) \\ \vdots \\ y(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} + \sigma^2 \begin{bmatrix} 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 1 \end{bmatrix} \right)$$

у

 \mathbf{K} + σ^2

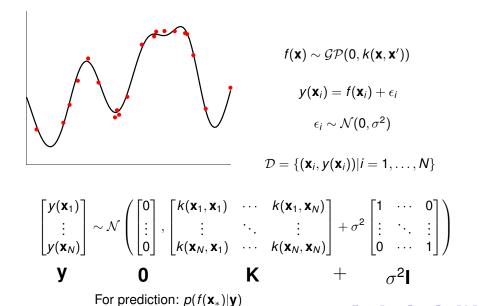


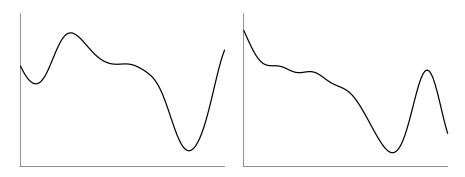
$$egin{aligned} f(\mathbf{x}) &\sim \mathcal{GP}(\mathbf{0}, k(\mathbf{x}, \mathbf{x}')) \ & y(\mathbf{x}_i) = f(\mathbf{x}_i) + \epsilon_i \ & \epsilon_i &\sim \mathcal{N}(\mathbf{0}, \sigma^2) \end{aligned}$$

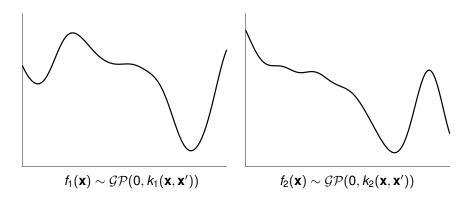
$$\mathcal{D} = \{(\mathbf{x}_i, y(\mathbf{x}_i)) | i = 1, \dots, N\}$$

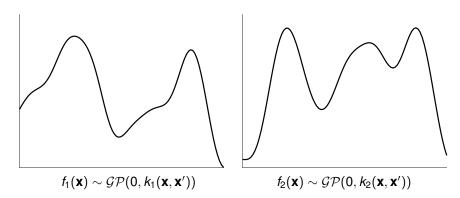
$$\begin{bmatrix} y(\mathbf{x}_1) \\ \vdots \\ y(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_N) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_N, \mathbf{x}_1) & \cdots & k(\mathbf{x}_N, \mathbf{x}_N) \end{bmatrix} + \sigma^2 \begin{bmatrix} 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 1 \end{bmatrix} \right)$$

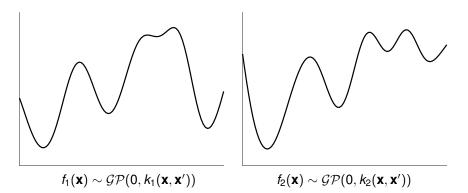
y 0 K + σ

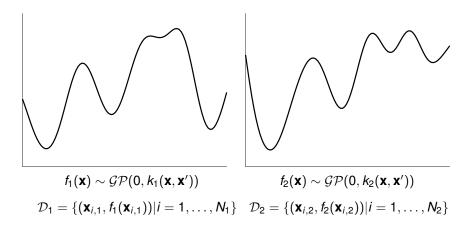


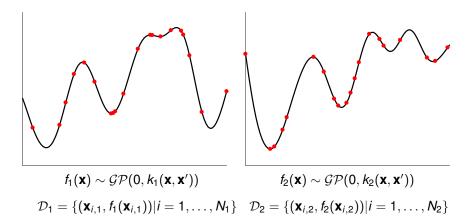


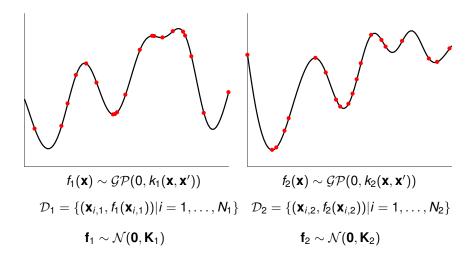


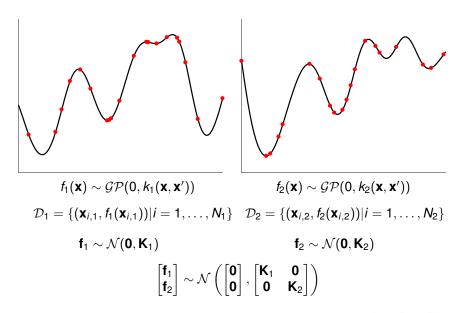


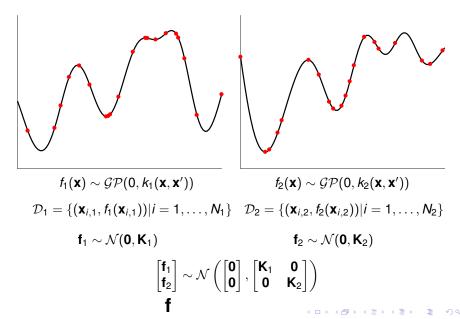


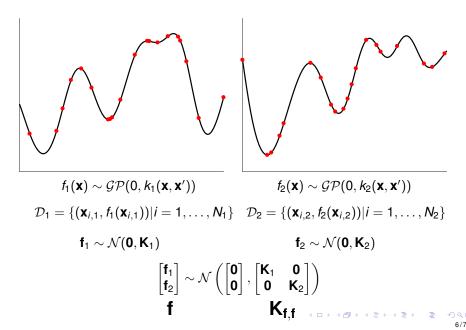


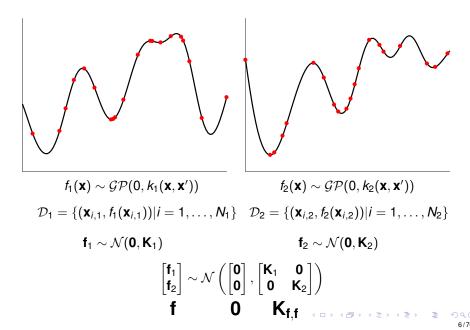


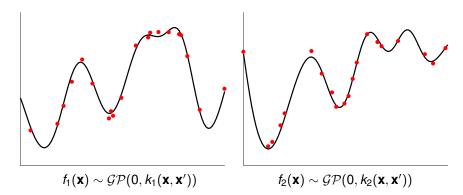


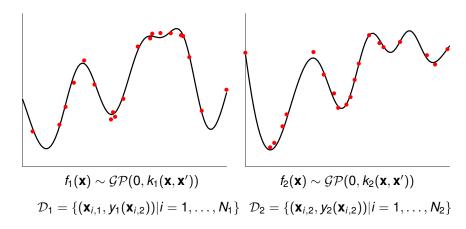


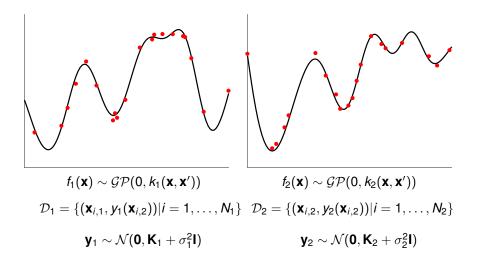


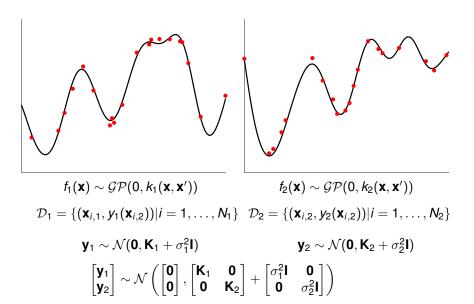


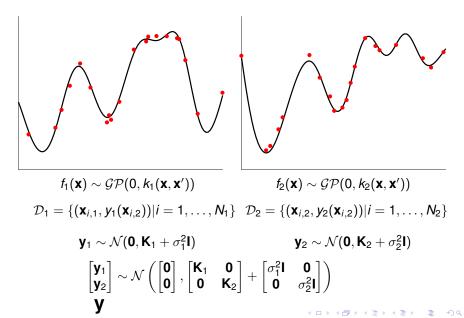


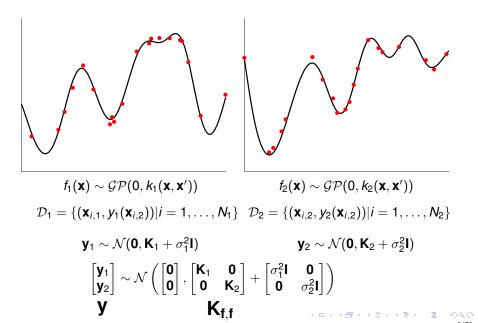


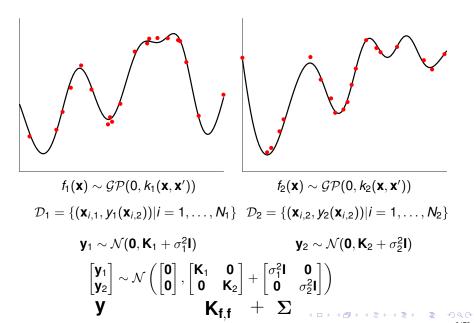


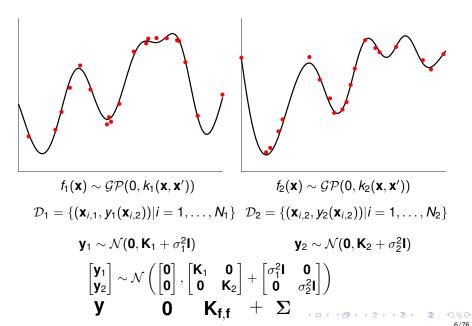


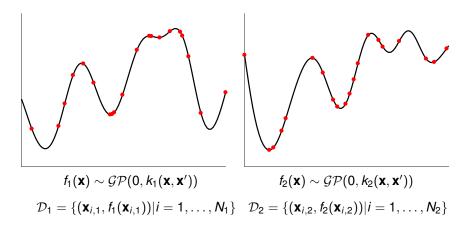


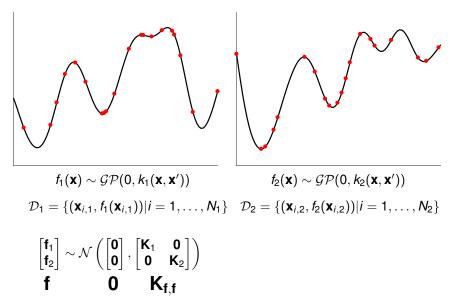


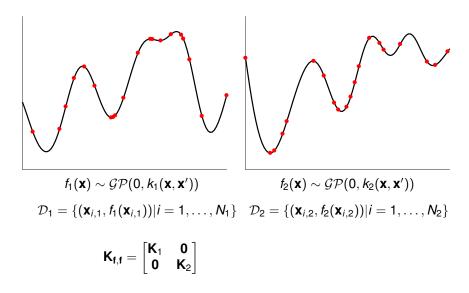


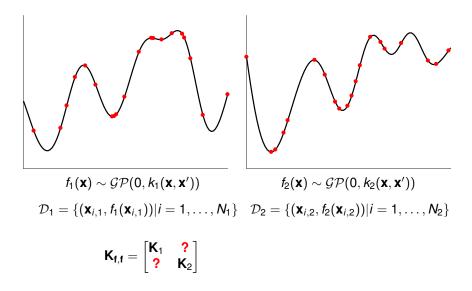


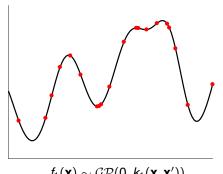








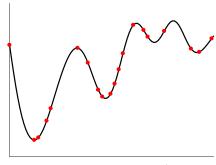




$$f_1(\mathbf{x}) \sim \mathcal{GP}(0, k_1(\mathbf{x}, \mathbf{x}'))$$

$$\mathcal{D}_1 = \{(\boldsymbol{x}_{i,1}, \mathit{f}_1(\boldsymbol{x}_{i,1})) | i = 1, \dots, N_1\} \quad \mathcal{D}_2 = \{(\boldsymbol{x}_{i,2}, \mathit{f}_2(\boldsymbol{x}_{i,2})) | i = 1, \dots, N_2\}$$

$$\mathbf{K}_{\mathbf{f},\mathbf{f}} = \begin{bmatrix} \mathbf{K}_1 & \textcolor{red}{?} \\ \textcolor{red}{?} & \mathbf{K}_2 \end{bmatrix}$$

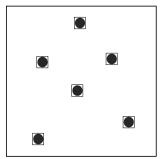


$$f_2(\mathbf{x}) \sim \mathcal{GP}(0, k_2(\mathbf{x}, \mathbf{x}'))$$

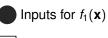
$$\mathcal{D}_2 = \{(\mathbf{x}_{i,2}, f_2(\mathbf{x}_{i,2})) | i = 1, \dots, N_2\}$$

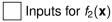
Build a cross-covariance function $cov[f_1(\mathbf{x}), f_2(\mathbf{x}')]$ such that $\mathbf{K}_{\mathbf{f},\mathbf{f}}$ is positive semi-definite.

Isotopic data

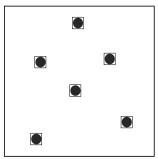


Sample sites are shared





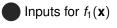
Isotopic data



Sample sites are shared

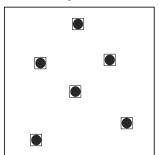
$$\mathcal{D}_1 = \{ (\mathbf{x}_i, f_1(\mathbf{x}_i))_{i=1}^N \}$$

$$\mathcal{D}_2 = \{ (\mathbf{x}_i, f_2(\mathbf{x}_i))_{i=1}^N \}$$



Inputs for
$$f_2(\mathbf{x})$$

Isotopic data

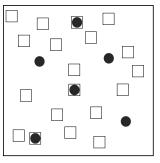


Sample sites are shared

$$\mathcal{D}_1 = \{ (\mathbf{x}_i, f_1(\mathbf{x}_i))_{i=1}^N \}$$

$$\mathcal{D}_2 = \{ (\mathbf{x}_i, f_2(\mathbf{x}_i))_{i=1}^N \}$$

Heterotopic data



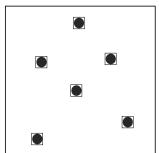
Sample sites may be different

Inputs for
$$f_1(\mathbf{x})$$

Inputs for
$$f_2(\mathbf{x})$$



Isotopic data

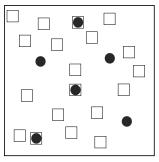


Sample sites are shared

$$\mathcal{D}_{1} = \{(\mathbf{x}_{i}, f_{1}(\mathbf{x}_{i}))_{i=1}^{N}\}$$

$$\mathcal{D}_{2} = \{(\mathbf{x}_{i}, f_{2}(\mathbf{x}_{i}))_{i=1}^{N}\}$$

Heterotopic data



$$\mathcal{D}_1 = \{(\mathbf{x}_{i,1}, f_1(\mathbf{x}_{i,1}))_{i=1}^{N_1}\}\$$

$$\mathcal{D}_2 = \{(\mathbf{x}_{i,2}, f_2(\mathbf{x}_{i,2}))_{i=1}^{N_2}\}\$$

Inputs for
$$f_1(\mathbf{x})$$

Inputs for
$$f_2(\mathbf{x})$$

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Summary

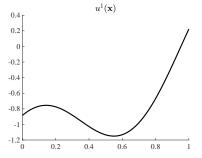
Intrinsic coregionalization model (ICM): two outputs

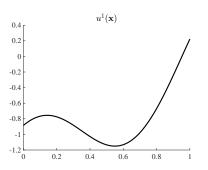
 \Box Consider two outputs $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ with $\mathbf{x} \in \mathbb{R}^p$.

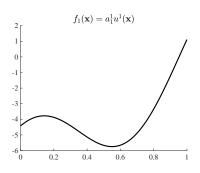
- We assume the following generative model for the outputs
 - 1. Sample from a GP $u(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$ to obtain $u^1(\mathbf{x})$
 - 2. Obtain $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ by linearly transforming $u^1(\mathbf{x})$

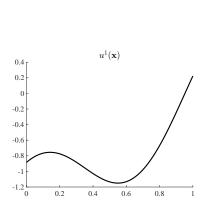
$$f_1(\mathbf{x})=a_1^1u^1(\mathbf{x})$$

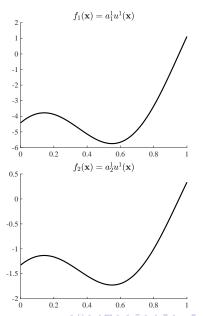
$$f_2(\mathbf{x}) = a_2^1 u^1(\mathbf{x})$$

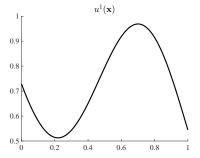


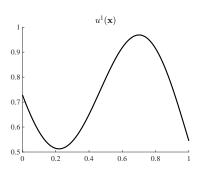


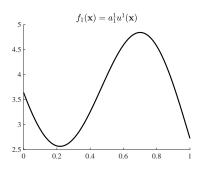


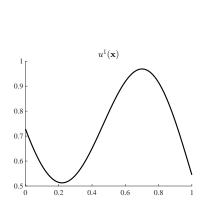


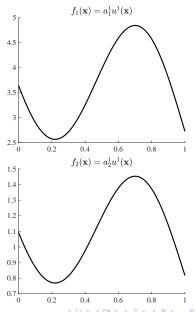












ICM: covariance (I)

For a fixed value of \mathbf{x} , we can group $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ in a vector $\mathbf{f}(\mathbf{x})$

$$\mathbf{f}(\mathbf{x}) = \begin{bmatrix} f_1(\mathbf{x}) \\ f_2(\mathbf{x}) \end{bmatrix}$$

- □ We refer to this vector as a *vector-valued function*.
- \Box The covariance for f(x) is computed as

$$\mathsf{cov}\left(f(\boldsymbol{x}), f(\boldsymbol{x}')\right) = \mathbb{E}\left\{f(\boldsymbol{x})[f(\boldsymbol{x}')]^\top\right\} - \mathbb{E}\left\{f(\boldsymbol{x})\right\} \left[\mathbb{E}\left\{f(\boldsymbol{x}')\right\}\right]^\top.$$

 $lue{}$ We compute first the term $\mathbb{E}\left\{\mathbf{f}(\mathbf{x})[\mathbf{f}(\mathbf{x}')]^{\top}\right\}$

$$\mathbb{E}\left\{\begin{bmatrix}f_1(\mathbf{x})\\f_2(\mathbf{x})\end{bmatrix}\begin{bmatrix}f_1(\mathbf{x}') & f_2(\mathbf{x}')\end{bmatrix}\right\} = \begin{bmatrix}\mathbb{E}\left\{f_1(\mathbf{x})f_1(\mathbf{x}')\right\} & \mathbb{E}\left\{f_1(\mathbf{x})f_2(\mathbf{x}')\right\}\\\mathbb{E}\left\{f_2(\mathbf{x})f_1(\mathbf{x}')\right\} & \mathbb{E}\left\{f_2(\mathbf{x})f_2(\mathbf{x}')\right\}\end{bmatrix}$$



ICM: covariance (II)

We compute the expected values as

$$\begin{split} \mathbb{E}\left\{f_{1}(\boldsymbol{x})f_{1}(\boldsymbol{x}')\right\} &= \mathbb{E}\left\{a_{1}^{1}u^{1}(\boldsymbol{x})a_{1}^{1}u^{1}(\boldsymbol{x}')\right\} = (a_{1}^{1})^{2}\mathbb{E}\left\{u^{1}(\boldsymbol{x})u^{1}(\boldsymbol{x}')\right\} \\ \mathbb{E}\left\{f_{1}(\boldsymbol{x})f_{2}(\boldsymbol{x}')\right\} &= \mathbb{E}\left\{a_{1}^{1}u^{1}(\boldsymbol{x})a_{2}^{1}(\boldsymbol{x}')\right\} = a_{1}^{1}a_{2}^{1}\mathbb{E}\left\{u^{1}(\boldsymbol{x})u^{1}(\boldsymbol{x}')\right\} \\ \mathbb{E}\left\{f_{2}(\boldsymbol{x})f_{2}(\boldsymbol{x}')\right\} &= \mathbb{E}\left\{a_{2}^{1}u^{1}(\boldsymbol{x})a_{2}^{1}u^{1}(\boldsymbol{x}')\right\} = (a_{2}^{1})^{2}\mathbb{E}\left\{u^{1}(\boldsymbol{x})u^{1}(\boldsymbol{x}')\right\} \end{split}$$

□ The term $\mathbb{E}\left\{\mathbf{f}(\mathbf{x})[\mathbf{f}(\mathbf{x}')]^{\top}\right\}$ follows as

$$\begin{split} \mathbb{E}\left\{\mathbf{f}(\mathbf{x})[\mathbf{f}(\mathbf{x}')]^{\top}\right\} &= \begin{bmatrix} (a_{1}^{1})^{2}\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} & a_{1}^{1}a_{2}^{1}\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} \\ a^{1}a^{2}\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} & (a_{2}^{1})^{2}\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} \end{bmatrix} \\ &= \begin{bmatrix} (a_{1}^{1})^{2} & a_{1}^{1}a_{2}^{1} \\ a_{1}^{1}a_{2}^{1} & (a_{2}^{1})^{2} \end{bmatrix} \mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} \end{split}$$

□ The term $\mathbb{E}\{\mathbf{f}(\mathbf{x})\}$ is computed as

$$\mathbb{E}\left\{\begin{bmatrix}f_1(\mathbf{x})\\f_2(\mathbf{x})\end{bmatrix}\right\} = \begin{bmatrix}\mathbb{E}\left\{f_1(\mathbf{x})\right\}\\\mathbb{E}\left\{f_2(\mathbf{x})\right\}\end{bmatrix} = \begin{bmatrix}\mathbb{E}\left\{a_1^1u^1(\mathbf{x})\right\}\\\mathbb{E}\left\{a_2^1u^1(\mathbf{x})\right\}\end{bmatrix} = \begin{bmatrix}a_1^1\\a_2^1\end{bmatrix}\mathbb{E}\left\{u^1(\mathbf{x})\right\}$$



ICM: covariance (III)

 \Box Putting the terms together, the covariance for f(x') follows as

$$\begin{bmatrix} (a_1^1)^2 & a_1^1 a_2^1 \\ a_1^1 a_2^1 & (a_2^1)^2 \end{bmatrix} \mathbb{E} \left\{ u^1(\mathbf{x}) u^1(\mathbf{x}') \right\} - \begin{bmatrix} a_1^1 \\ a_2^1 \end{bmatrix} \begin{bmatrix} a_1^1 & a_2^1 \end{bmatrix} \mathbb{E} \left\{ u^1(\mathbf{x}) \right\} \mathbb{E} \left\{ u^1(\mathbf{x}') \right\}$$

□ Defining $\mathbf{a} = [a_1^1 \ a_2^1]^\top$,

$$\begin{aligned} \text{cov}\left(\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')\right) &= \mathbf{a}\mathbf{a}^{\top}\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} - \mathbf{a}\mathbf{a}^{\top}\mathbb{E}\left\{u^{1}(\mathbf{x})\right\}\mathbb{E}\left\{u^{1}(\mathbf{x}')\right\} \\ &= \mathbf{a}\mathbf{a}^{\top}\underbrace{\left[\mathbb{E}\left\{u^{1}(\mathbf{x})u^{1}(\mathbf{x}')\right\} - \mathbb{E}\left\{u^{1}(\mathbf{x})\right\}\mathbb{E}\left\{u^{1}(\mathbf{x}')\right\}\right]}_{k(\mathbf{x},\mathbf{x}')} \\ &= \mathbf{a}\mathbf{a}^{\top}k(\mathbf{x},\mathbf{x}') \end{aligned}$$

■ We define $\mathbf{B} = \mathbf{a}\mathbf{a}^{\top}$, leading to

$$\operatorname{cov}\left(\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')\right) = \mathbf{B}k(\mathbf{x},\mathbf{x}') = \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \end{bmatrix} k(\mathbf{x},\mathbf{x}')$$

Notice that B has rank one.



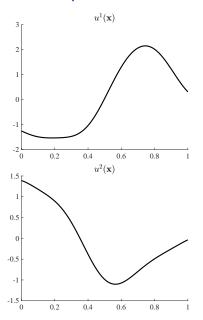
ICM: two outputs and two latent samples

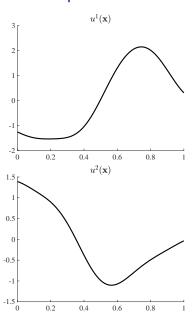
- We can introduce a bit more of complexity in the model before as follows.
- □ Consider again two outputs $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ with $\mathbf{x} \in \mathbb{R}^p$.
- We assume the following generative model for the outputs
 - 1. Sample twice from a GP $u(\mathbf{x}) \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$ to obtain $u^1(\mathbf{x})$ and $u^2(\mathbf{x})$
 - 2. Obtain $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ by adding a scaled transformation of $u^1(\mathbf{x})$ and $u^2(\mathbf{x})$

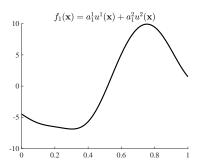
$$f_1(\mathbf{x}) = a_1^1 u^1(\mathbf{x}) + a_1^2 u^2(\mathbf{x})$$

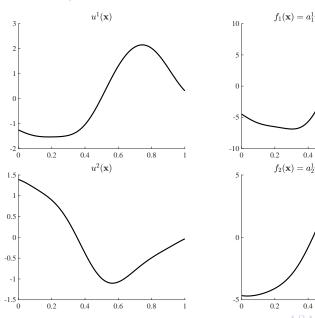
$$f_2(\mathbf{x}) = a_2^1 u^1(\mathbf{x}) + a_2^2 u^2(\mathbf{x})$$

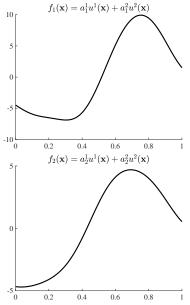
Notice that $u^1(\mathbf{x})$ and $u^2(\mathbf{x})$ are independent, although they share the same covariance $k(\mathbf{x}, \mathbf{x}')$.

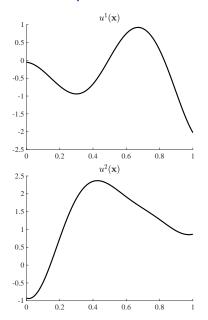


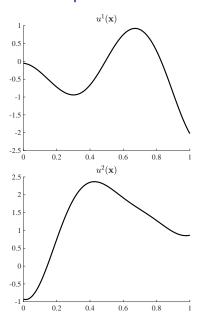


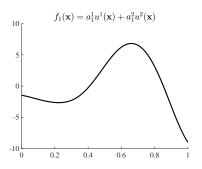




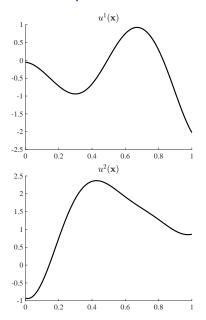


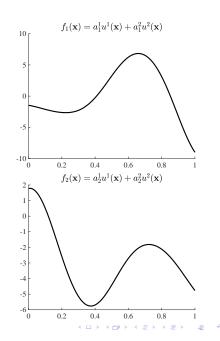






ICM: samples





ICM: covariance

 \Box The vector-valued function can be written as f(x)

$$\mathbf{f}(\mathbf{x}) = \mathbf{a}^1 u^1(\mathbf{x}) + \mathbf{a}^2 u^2(\mathbf{x})$$

where
$$\mathbf{a}^1 = [a_1^1 \ a_2^1]^{\top}$$
 and $\mathbf{a}^2 = [a_1^2 \ a_2^2]^{\top}$.

 \Box The covariance for f(x) is computed as

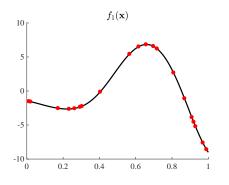
$$\begin{aligned} \text{cov}\left(\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')\right) &= \mathbf{a}^{1}(\mathbf{a}^{1})^{\top} \operatorname{cov}(u^{1}(\mathbf{x}),u^{1}(\mathbf{x}')) + \mathbf{a}^{2}(\mathbf{a}^{2})^{\top} \operatorname{cov}(u^{2}(\mathbf{x}),u^{2}(\mathbf{x}')) \\ &= \mathbf{a}^{1}(\mathbf{a}^{1})^{\top} k(\mathbf{x},\mathbf{x}') + \mathbf{a}^{2}(\mathbf{a}^{2})^{\top} k(\mathbf{x},\mathbf{x}') \\ &= \left[\mathbf{a}^{1}(\mathbf{a}^{1})^{\top} + \mathbf{a}^{2}(\mathbf{a}^{2})^{\top}\right] k(\mathbf{x},\mathbf{x}') \end{aligned}$$

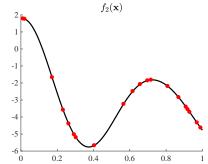
□ We define $\mathbf{B} = \mathbf{a}^1(\mathbf{a}^1)^\top + \mathbf{a}^2(\mathbf{a}^2)^\top$, leading to

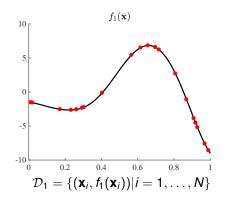
$$\operatorname{cov}\left(\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')\right) = \mathbf{B}k(\mathbf{x},\mathbf{x}') = \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \end{bmatrix} k(\mathbf{x},\mathbf{x}')$$

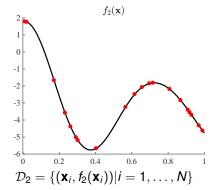
Notice that B has rank two.

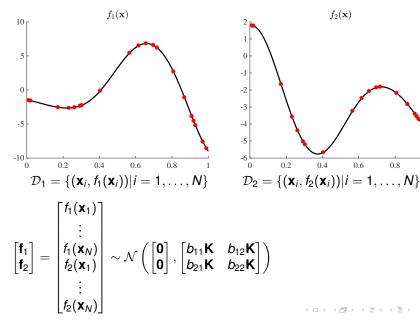


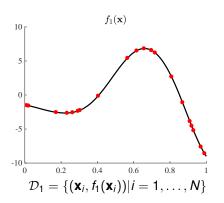










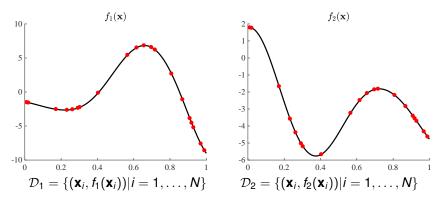


$$f_2(\mathbf{x})$$
 $f_2(\mathbf{x})$
 $f_2(\mathbf{x})$

$$\begin{bmatrix} \mathbf{f}_1 \\ \mathbf{f}_2 \end{bmatrix} = \begin{bmatrix} f_1(\mathbf{x}_1) \\ \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_2(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} b_{11}\mathbf{K} & b_{12}\mathbf{K} \\ b_{21}\mathbf{K} & b_{22}\mathbf{K} \end{bmatrix} \right)$$

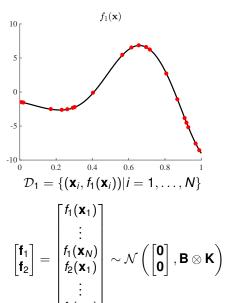
$$\sim \mathcal{N}\left(\begin{bmatrix}\mathbf{0}\\\mathbf{0}\end{bmatrix},\begin{bmatrix}b_{11}\mathbf{K}&b_{12}\mathbf{K}\\b_{21}\mathbf{K}&b_{22}\mathbf{K}\end{bmatrix}\right)$$

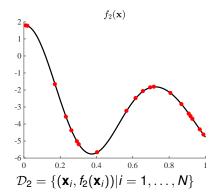
The matrix $\mathbf{K} \in \mathbb{R}^{N \times N}$ has elements $k(\mathbf{x}_i, \mathbf{x}_i)$.



The Kronecker product between matrices $\mathbf{C} \in \mathbb{R}^{c_1 \times c_2}$ and $\mathbf{G} \in \mathbb{R}^{g_1 \times g_2}$ with

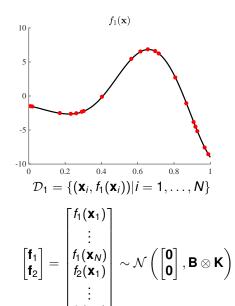
$$\mathbf{C} = \begin{bmatrix} c_{1,1} & \cdots & c_{1,c_2} \\ \vdots & \vdots & \vdots \\ c_{c_1,1} & \cdots & c_{c_1,c_2} \end{bmatrix} \quad \text{is} \quad \mathbf{C} \otimes \mathbf{G} = \begin{bmatrix} c_{1,1}\mathbf{G} & \cdots & c_{1,c_2}\mathbf{G} \\ \vdots & \vdots & \vdots \\ c_{c_1,1}\mathbf{G} & \cdots & c_{c_1,c_2}\mathbf{G} \end{bmatrix}$$

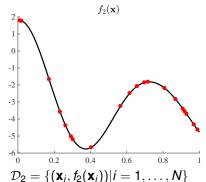




$$,\mathbf{B}\otimes\mathbf{K}\Big)$$







The matrix
$$\mathbf{K} \in \mathbb{R}^{N \times N}$$
 has elements $k(\mathbf{x}_i, \mathbf{x}_j)$.

ICM: general case

- Consider a set of functions $\{f_d(\mathbf{x})\}_{d=1}^D$.
- In the ICM

$$f_d(\mathbf{x}) = \sum_{i=1}^{H} a_d^i u^i(\mathbf{x}),$$

where the functions $u^i(\mathbf{x})$ are GPs sampled independently, and share the same covariance function $k(\mathbf{x}, \mathbf{x}')$.

For $\mathbf{f}(\mathbf{x}) = [f_1(\mathbf{x}) \cdots f_D(\mathbf{x})]^{\top}$, the covariance $\text{cov}[\mathbf{f}(\mathbf{x}), \mathbf{f}(\mathbf{x}')]$ is given as

$$cov[\mathbf{f}(\mathbf{x}), \mathbf{f}(\mathbf{x}')] = \mathbf{A}\mathbf{A}^{\top} k(\mathbf{x}, \mathbf{x}') = \mathbf{B} k(\mathbf{x}, \mathbf{x}'),$$

where $\mathbf{A} = [\mathbf{a}^1 \ \mathbf{a}^2 \cdots \mathbf{a}^R].$

 $lue{}$ The rank of $\mathbf{B} \in \mathbb{R}^{D imes D}$ is given by R.



ICM: autokrigeability

If the outputs are considered to be noise-free, prediction using the ICM under an isotopic data case is equivalent to independent prediction over each output.

This circumstance is also known as autokrigeability.

Contents

Dependencies between processes

Intrinsic Coregionalization Model

Semiparametric Latent Factor Model

Linear Model of Coregionalization

Process convolutions

Covariance fitting and Prediction

Cokriging

Extensions

Computational complexity Variations of LMC Variations of PC

Summary

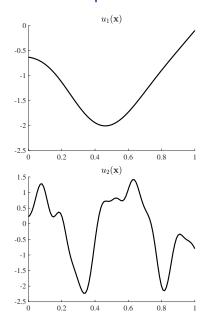
Semiparametric Latent Factor Model (SLFM)

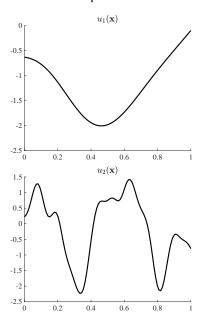
- □ ICM uses R samples $u^i(\mathbf{x})$ from $u(\mathbf{x})$ with the same covariance function.
- SLFM uses Q samples from $u_q(\mathbf{x})$ processes with different covariance functions.
- □ The SLFM with Q = 1 is the same to the ICM with R = 1.
- □ Consider two outputs $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ with $\mathbf{x} \in \mathbb{R}^p$.
- □ Suppose we have Q = 2.
- We assume the following generative model for the outputs
 - 1. Sample from a GP $\mathcal{GP}(0, k_1(\mathbf{x}, \mathbf{x}'))$ to obtain $u_1(\mathbf{x})$.
 - 2. Sample from a GP $\mathcal{GP}(0, k_2(\mathbf{x}, \mathbf{x}'))$ to obtain $u_2(\mathbf{x})$.
 - 3. Obtain $f_1(\mathbf{x})$ and $f_2(\mathbf{x})$ by adding a scaled versions of $u_1(\mathbf{x})$ and $u_2(\mathbf{x})$

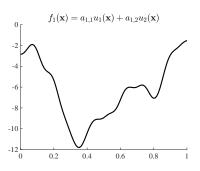
$$f_1(\mathbf{x}) = a_{1,1}u_1(\mathbf{x}) + a_{1,2}u_2(\mathbf{x})$$

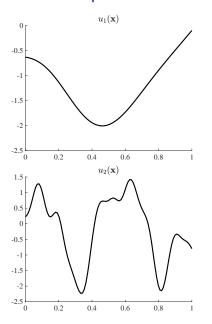
$$f_2(\mathbf{x}) = a_{2,1}u_1(\mathbf{x}) + a_{2,2}u_2(\mathbf{x})$$

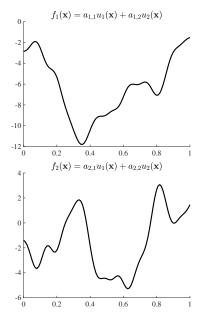


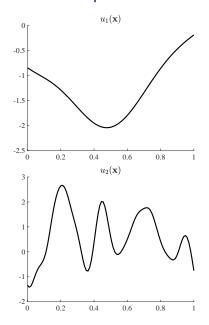


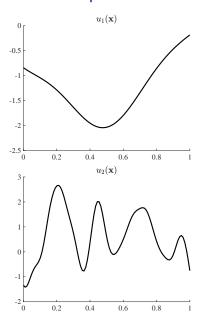


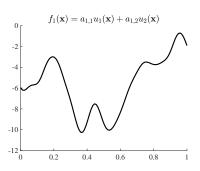


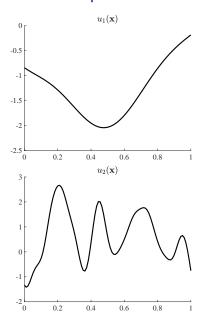


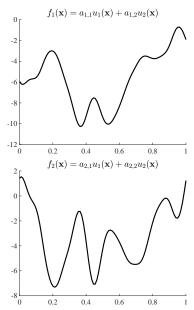












SLFM: covariance

 \Box The vector-valued function can be written as $\mathbf{f}(\mathbf{x})$

$$\mathbf{f}(\mathbf{x}) = \mathbf{a}_1 u_1(\mathbf{x}) + \mathbf{a}_2 u_2(\mathbf{x})$$

where
$$\mathbf{a}_1 = [a_{1,1} \ a_{2,1}]^{\top}$$
 and $\mathbf{a}_2 = [a_{1,2} \ a_{2,2}]^{\top}$.

The covariance for f(x) is computed as

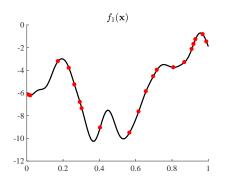
$$\begin{aligned} \operatorname{cov}\left(\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')\right) &= \mathbf{a}_{1}(\mathbf{a}_{1})^{\top} \operatorname{cov}(u_{1}(\mathbf{x}),u_{1}(\mathbf{x}')) + \mathbf{a}_{2}(\mathbf{a}_{2})^{\top} \operatorname{cov}(u_{2}(\mathbf{x}),u_{2}(\mathbf{x}')) \\ &= \mathbf{a}_{1}(\mathbf{a}_{1})^{\top} k_{1}(\mathbf{x},\mathbf{x}') + \mathbf{a}_{2}(\mathbf{a}_{2})^{\top} k_{2}(\mathbf{x},\mathbf{x}') \end{aligned}$$

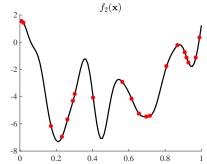
 $lue{lue}$ We define ${f B}_1={f a}_1({f a}_1)^{ op}$ and ${f B}_2={f a}_2({f a}_2)^{ op},$ leading to

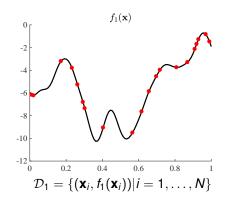
$$\mathsf{cov}\left(\boldsymbol{f}(\boldsymbol{x}),\boldsymbol{f}(\boldsymbol{x}')\right) = \boldsymbol{B}_1 k_1(\boldsymbol{x},\boldsymbol{x}') + \boldsymbol{B}_2 k_2(\boldsymbol{x},\boldsymbol{x}')$$

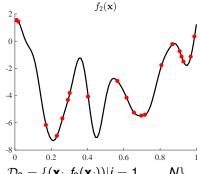
□ Notice that \mathbf{B}_1 and \mathbf{B}_2 have rank one.

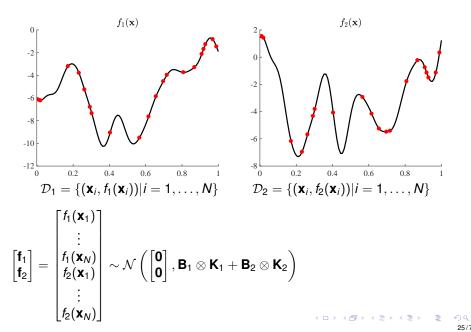


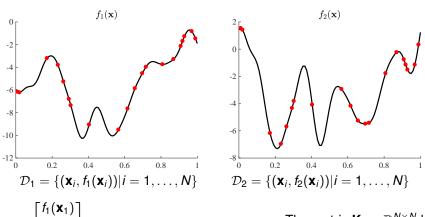








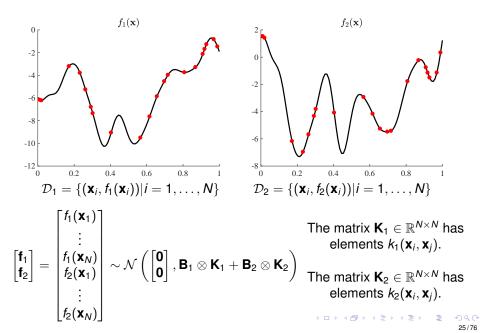




$$\begin{vmatrix} \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_2(\mathbf{x}_N) \end{vmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \mathbf{B}_1 \otimes \mathbf{K}_1 + \mathbf{B}_2 \otimes \mathbf{K}_2 \right)$$

The matrix $\mathbf{K}_1 \in \mathbb{R}^{N \times N}$ has elements $k_1(\mathbf{x}_i, \mathbf{x}_j)$.





SLFM: general case

- Consider a set of functions $\{f_d(\mathbf{x})\}_{d=1}^D$.
- In the SLFM

$$f_d(\mathbf{x}) = \sum_{q=1}^{Q} a_{d,q} u_q(\mathbf{x}),$$

where the functions $u_q(\mathbf{x})$ are GPs with covariance functions $k_q(\mathbf{x}, \mathbf{x}')$.

□ For $\mathbf{f}(\mathbf{x}) = [f_1(\mathbf{x}) \cdots f_D(\mathbf{x})]^{\top}$, the covariance $cov[\mathbf{f}(\mathbf{x}), \mathbf{f}(\mathbf{x}')]$ is given as

$$\operatorname{cov}[\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')] = \sum_{q=1}^{Q} \mathbf{A}_q \mathbf{A}_q^{\top} k_q(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} \mathbf{B}_q k_q(\mathbf{x},\mathbf{x}'),$$

where $\mathbf{A}_q = \mathbf{a}_q$.

lacktriangle The rank of each $oldsymbol{\mathsf{B}}_q \in \mathbb{R}^{D imes D}$ is one.

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Linear model of coregionalization (LMC)

- The LMC generalizes the ICM and the SLFM allowing several independent samples from GPs with different covariances.
- □ Consider a set of functions $\{f_d(\mathbf{x})\}_{d=1}^D$.
- In the LMC

$$f_d(\mathbf{x}) = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} a_{d,q}^i u_q^i(\mathbf{x}),$$

where the functions $u_q^i(\mathbf{x})$ are GPs with zero means and covariance functions

$$\operatorname{cov}[u_q^i(\mathbf{x}), u_{q'}^{i'}(\mathbf{x}')] = k_q(\mathbf{x}, \mathbf{x}'),$$

if
$$i = i'$$
 and $q = q'$.

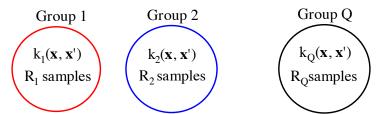


LMC: interpretation

In the LMC

$$f_d(\mathbf{x}) = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} a^i_{d,q} u^i_q(\mathbf{x}).$$

- There are Q groups of samples.
- For each group, there R_q samples obtained independently from the same GP with covariance $k_q(\mathbf{x}, \mathbf{x}')$.



- The LMC corresponds to the sum of Q ICMs.
- Suppose we have D = 2, Q = 2 and $R_q = 2$. According to the LMC

$$\begin{split} f_1(\mathbf{x}) &= a_{1,1}^1 u_1^1(\mathbf{x}) + a_{1,1}^2 u_1^2(\mathbf{x}) + a_{1,2}^1 u_2^1(\mathbf{x}) + a_{1,2}^2 u_2^2(\mathbf{x}), \\ f_2(\mathbf{x}) &= a_{2,1}^1 u_1^1(\mathbf{x}) + a_{2,1}^2 u_1^2(\mathbf{x}) + a_{2,2}^1 u_2^1(\mathbf{x}) + a_{2,2}^2 u_2^2(\mathbf{x}), \end{split}$$

- The LMC corresponds to the sum of Q ICMs.
- Suppose we have D = 2, Q = 2 and $R_q = 2$. According to the LMC

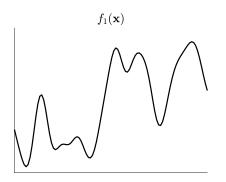
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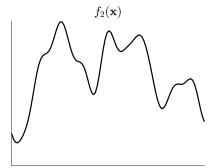
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- The LMC corresponds to the sum of Q ICMs.
- Suppose we have D = 2, Q = 2 and $R_q = 2$. According to the LMC

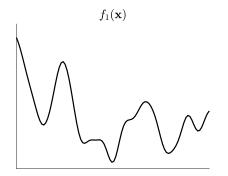
$$\begin{split} f_1(\mathbf{x}) &= a_{1,1}^1 u_1^1(\mathbf{x}) + a_{1,1}^2 u_1^2(\mathbf{x}) + a_{1,2}^1 u_2^1(\mathbf{x}) + a_{1,2}^2 u_2^2(\mathbf{x}), \\ f_2(\mathbf{x}) &= a_{2,1}^1 u_1^1(\mathbf{x}) + a_{2,1}^2 u_1^2(\mathbf{x}) + a_{2,2}^1 u_2^1(\mathbf{x}) + a_{2,2}^2 u_2^2(\mathbf{x}), \end{split}$$

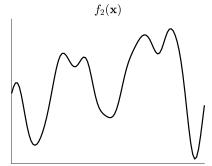




- The LMC corresponds to the sum of Q ICMs.
- Suppose we have D = 2, Q = 2 and $R_q = 2$. According to the LMC

$$\begin{split} f_1(\boldsymbol{x}) &= a_{1,1}^1 u_1^1(\boldsymbol{x}) + a_{1,1}^2 u_1^2(\boldsymbol{x}) + a_{1,2}^1 u_2^1(\boldsymbol{x}) + a_{1,2}^2 u_2^2(\boldsymbol{x}), \\ f_2(\boldsymbol{x}) &= a_{2,1}^1 u_1^1(\boldsymbol{x}) + a_{2,1}^2 u_1^2(\boldsymbol{x}) + a_{2,2}^1 u_1^2(\boldsymbol{x}) + a_{2,2}^2 u_2^2(\boldsymbol{x}), \end{split}$$





LMC: covariance for f(x)

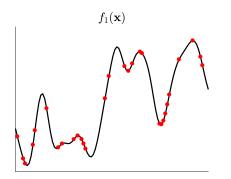
For $\mathbf{f}(\mathbf{x}) = [f_1(\mathbf{x}) \cdots f_D(\mathbf{x})]^{\top}$, the covariance $\text{cov}[\mathbf{f}(\mathbf{x}), \mathbf{f}(\mathbf{x}')]$ is given as

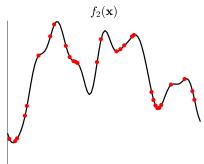
$$\mathsf{cov}[\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')] = \sum_{q=1}^Q \mathbf{A}_q \mathbf{A}_q^\top \, k_q(\mathbf{x},\mathbf{x}') = \sum_{q=1}^Q \mathbf{B}_q \, k_q(\mathbf{x},\mathbf{x}'),$$

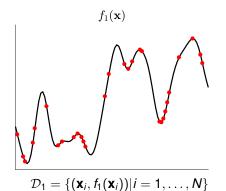
where
$$\mathbf{A}_q = [\mathbf{a}_q^1 \ \mathbf{a}_q^2 \cdots \mathbf{a}_q^{R_q}].$$

□ The rank of each \mathbf{B}_q is R_q .

 \Box The matrices \mathbf{B}_a are known as the *coregionalization matrices*.

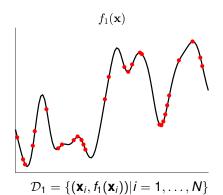


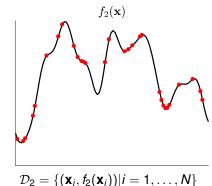




$$f_2(\mathbf{x})$$

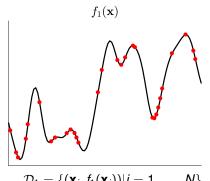
$$\mathcal{D}_2 = \{(\mathbf{x}_i, f_2(\mathbf{x}_i)) | i = 1, \dots, N\}$$

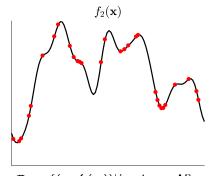




$$\begin{bmatrix} \mathbf{f}_1 \\ \mathbf{f}_2 \end{bmatrix} = \begin{bmatrix} f_1(\mathbf{x}_1) \\ \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_n(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \sum_{q=1}^Q \mathbf{B}_q \otimes \mathbf{K}_q \right)$$





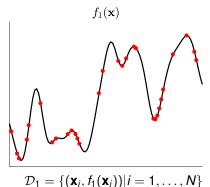


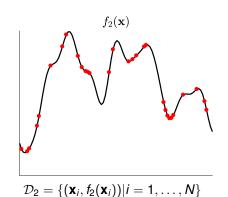
$$\mathcal{D}_1 = \{(\mathbf{x}_i, f_1(\mathbf{x}_i)) | i = 1, \ldots, N\}$$

$$\mathcal{D}_2 = \{(\mathbf{x}_i, f_2(\mathbf{x}_i)) | i = 1, \dots, N\}$$

$$\begin{bmatrix} \mathbf{f}_1 \\ \mathbf{f}_2 \end{bmatrix} = \begin{bmatrix} \mathbf{f}_1 \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \sum_{q=1}^{Q} \mathbf{B}_q \otimes \mathbf{K}_q \right)$$

The matrix $\mathbf{K}_q \in \mathbb{R}^{N \times N}$ has elements $k_q(\mathbf{x}_i, \mathbf{x}_j)$.





$$\mathcal{D}_1 = \{(\mathbf{x}_i, I_1(\mathbf{x}_i)) | I = 1, \dots, N\}$$

$$\left[f_1(\mathbf{x}_1) \right]$$

The matrix
$$\mathbf{K}_q \in \mathbb{R}^{N \times N}$$
 has elements $k_q(\mathbf{x}_i,\mathbf{x}_j)$.

$$\begin{bmatrix} \mathbf{f}_1 \\ \mathbf{f}_2 \end{bmatrix} = \begin{bmatrix} \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_n(\mathbf{x}_n) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \sum_{q=1}^Q \mathbf{B}_q \otimes \mathbf{K}_q \right)$$

The matrix $\mathbf{B}_q \in \mathbb{R}^{D imes D}$ has elements b_{ij}^q .

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Moving average function

- □ Consider again a set of *D* functions $\{f_d(\mathbf{x})\}_{d=1}^D$.
- Each function could be expressed through a convolution integral between a kernel, $\{G_d(\mathbf{x})\}_{d=1}^D$, and a function $u(\mathbf{x})$,

$$f_d(\mathbf{x}) = \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} = G_d(\mathbf{x}) * u(\mathbf{x}).$$

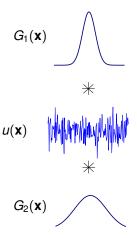
- \Box For the integral to exist, it is assumed that the kernel $G_d(\mathbf{x})$ is a continuous function with compact support or square-integrable.
- □ The kernel $G_d(\mathbf{x})$ is also known as the moving average function or the smoothing kernel.
- □ In Dependet Gaussian processes (DGP) the latent function $u(\mathbf{x})$ is white Gaussian noise (WGN).

A pictorial representation



 $u(\mathbf{x})$: latent function.

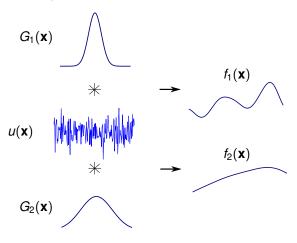
A pictorial representation



 $u(\mathbf{x})$: latent function.

 $G_1(\mathbf{x}), G_2(\mathbf{x})$: smoothing kernels.

A pictorial representation



 $u(\mathbf{x})$: latent function.

 $G_1(\mathbf{x}), G_2(\mathbf{x})$: smoothing kernels.

 $f_1(\mathbf{x}), f_2(\mathbf{x})$: output functions.

The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, $\operatorname{cov}[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] - \\
\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}\right] \mathbb{E}\left[\int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] d\mathbf{z}' d\mathbf{z} - \\
\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) \mathbb{E}\left[u(\mathbf{z})\right] d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}')\right] d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \times \\
\{\mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] - \mathbb{E}\left[u(\mathbf{z})\right] \mathbb{E}\left[u(\mathbf{z}')\right]\} d\mathbf{z} d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

□ In the DGP $k(\mathbf{z}, \mathbf{z}') = \sigma^2 \delta(\mathbf{z} - \mathbf{z}')$.



The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] - \\
\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}\right] \mathbb{E}\left[\int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] d\mathbf{z}' d\mathbf{z} - \\
\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) \mathbb{E}\left[u(\mathbf{z})\right] d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}')\right] d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \times \\
\{\mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] - \mathbb{E}\left[u(\mathbf{z})\right] \mathbb{E}\left[u(\mathbf{z}')\right]\} d\mathbf{z} d\mathbf{z}' \\
= \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

□ In the DGP $k(\mathbf{z}, \mathbf{z}') = \sigma^2 \delta(\mathbf{z} - \mathbf{z}')$.



The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] - \\
\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}\right] \mathbb{E}\left[\int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] d\mathbf{z}' d\mathbf{z} - \\
\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) \mathbb{E}\left[u(\mathbf{z})\right] d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}')\right] d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \times \\
\{\mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] - \mathbb{E}\left[u(\mathbf{z})\right] \mathbb{E}\left[u(\mathbf{z}')\right]\} d\mathbf{z} d\mathbf{z}' \\
= \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

In the DGP $k(\mathbf{z}, \mathbf{z}') = \sigma^2 \delta(\mathbf{z} - \mathbf{z}')$.



The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] - \\
\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}\right] \mathbb{E}\left[\int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] d\mathbf{z}' d\mathbf{z} - \\
\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) \mathbb{E}\left[u(\mathbf{z})\right] d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}')\right] d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \times \\
\{\mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] - \mathbb{E}\left[u(\mathbf{z})\right] \mathbb{E}\left[u(\mathbf{z}')\right]\} d\mathbf{z} d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$



The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] - \\
\mathbb{E}\left[\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}\right] \mathbb{E}\left[\int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') u(\mathbf{z}') d\mathbf{z}'\right] \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] d\mathbf{z}' d\mathbf{z} - \\
\int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) \mathbb{E}\left[u(\mathbf{z})\right] d\mathbf{z} \int_{\mathcal{X}} G_{d'}(\mathbf{x}' - \mathbf{z}') \mathbb{E}\left[u(\mathbf{z}')\right] d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') \times \\
\{\mathbb{E}\left[u(\mathbf{z}) u(\mathbf{z}')\right] - \mathbb{E}\left[u(\mathbf{z})\right] \mathbb{E}\left[u(\mathbf{z}')\right]\} d\mathbf{z} d\mathbf{z}' \\
= \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

In the DGP $k(\mathbf{z}, \mathbf{z}') = \sigma^2 \delta(\mathbf{z} - \mathbf{z}')$.

Example of cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$ (I)

The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, $cov[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

$$\operatorname{cov}\left[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \sigma^2 \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}) d\mathbf{z}$$

Example. Assume that the smoothing kernels follow a Gaussian form

$$G_d(\mathbf{x} - \mathbf{z}) = rac{S_d |\mathbf{P}_d|^{1/2}}{(2\pi)^{p/2}} \exp\left[-rac{1}{2}(\mathbf{x} - \mathbf{z})^{\top}\mathbf{P}_d(\mathbf{x} - \mathbf{z})
ight],$$

We use the identity of the product of two Gaussians

$$\mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_1,\boldsymbol{P}_1^{-1})\mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_2,\boldsymbol{P}_2^{-1}) = \mathcal{N}(\boldsymbol{\mu}_1|\boldsymbol{\mu}_2,\boldsymbol{P}_1^{-1}+\boldsymbol{P}_2^{-1})\mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_c,\boldsymbol{P}_c^{-1}),$$

where
$$\mu_c = (\mathbf{P}_1 + \mathbf{P}_2)^{-1} (\mathbf{P}_1 \mu_1 + \mathbf{P}_2 \mu_2)$$
 and $\mathbf{P}_c^{-1} = (\mathbf{P}_1 + \mathbf{P}_2)^{-1}$.

Example of cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$ (II)

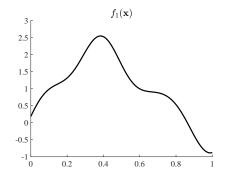
The cross-covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$, $\operatorname{cov}[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$, is

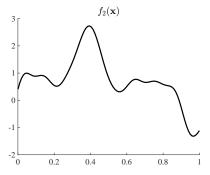
$$\begin{aligned} \operatorname{cov}\left[f_{d}(\mathbf{x}), f_{d'}(\mathbf{x}')\right] &= \sigma^{2} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}) d\mathbf{z} \\ &= \frac{\sigma^{2} S_{d} S_{d'}}{(2\pi)^{p/2} |\mathbf{P}_{\text{eqv}}|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \mathbf{x}')^{\top} \mathbf{P}_{\text{eqv}}^{-1}(\mathbf{x} - \mathbf{x}')\right], \end{aligned}$$

where
$$\mathbf{P}_{qv} = \mathbf{P}_{d}^{-1} + \mathbf{P}_{d'}^{-1}$$
.

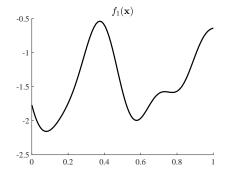
■ **Exercise**. Show how to obtain the expression above

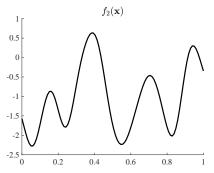
PC: samples

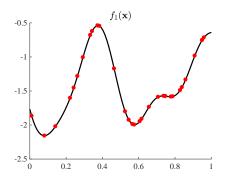


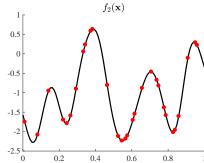


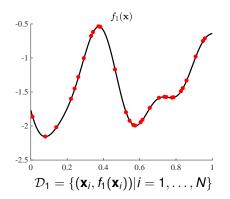
PC: samples

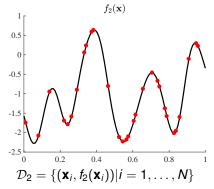


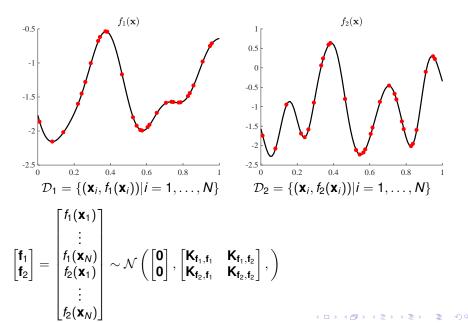


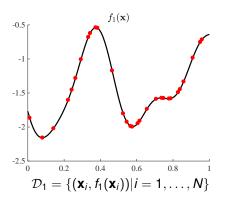












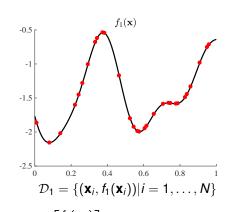
$$\begin{array}{c}
f_{2}(\mathbf{x}) \\
0.5 \\
0 \\
-0.5 \\
-1 \\
-1.5 \\
-2 \\
-2.5 \\
0.2 \\
0.4 \\
0.6 \\
0.8 \\
1
\end{array}$$

$$\mathcal{D}_{2} = \{(\mathbf{x}_{i}, f_{2}(\mathbf{x}_{i})) | i = 1, \dots, N\}$$

$$\begin{bmatrix} f_1(\mathbf{x}_1) \\ \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_2(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{K}_{\mathbf{f}_1, \mathbf{f}_1} & \mathbf{K}_{\mathbf{f}_1, \mathbf{f}_2} \\ \mathbf{K}_{\mathbf{f}_2, \mathbf{f}_1} & \mathbf{K}_{\mathbf{f}_2, \mathbf{f}_2} \end{bmatrix}, \right)$$

$$\begin{bmatrix} \mathbf{K}_{\mathbf{f}_1,\mathbf{f}_2} \\ \mathbf{K}_{\mathbf{f}_2,\mathbf{f}_2} \end{bmatrix},$$

The matrix $\mathbf{K}_{\mathbf{f}_d,\mathbf{f}_d} \in \mathbb{R}^{N \times N}$ has elements cov [$f_d(\mathbf{x}), f_d(\mathbf{x}')$].



$$\begin{bmatrix} \mathbf{f}_1 \\ \mathbf{f}_2 \end{bmatrix} = \begin{bmatrix} f_1(\mathbf{x}_1) \\ \vdots \\ f_1(\mathbf{x}_N) \\ f_2(\mathbf{x}_1) \\ \vdots \\ f_2(\mathbf{x}_N) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{K}_{\mathbf{f}_1, \mathbf{f}_1} & \mathbf{K}_{\mathbf{f}_1, \mathbf{f}_2} \\ \mathbf{K}_{\mathbf{f}_2, \mathbf{f}_1} & \mathbf{K}_{\mathbf{f}_2, \mathbf{f}_2} \end{bmatrix}, \right)$$

The matrix $\mathbf{K}_{\mathbf{f}_d,\mathbf{f}_d} \in \mathbb{R}^{N \times N}$ has elements $\operatorname{cov}\left[f_d(\mathbf{x}), f_d(\mathbf{x}')\right]$.

The matrix $\mathbf{K}_{\mathbf{f}_d,\mathbf{f}_{d'}} \in \mathbb{R}^{N \times N}$ has elements $\operatorname{cov}\left[f_d(\mathbf{x}),f_{d'}(\mathbf{x}')\right]$.

Beyond $u(\mathbf{x})$ as a white Gaussian noise

- □ Consider again a set of *D* functions $\{f_d(\mathbf{x})\}_{d=1}^D$.
- Each function could be expressed through a convolution integral between a kernel, $\{G_d(\mathbf{x})\}_{d=1}^D$, and a function $u(\mathbf{x})$,

$$f_d(\mathbf{x}) = \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z} = G_d(\mathbf{x}) * u(\mathbf{x}).$$

- Assuming $u(\mathbf{x})$ is a GP with zero mean and covariance $k(\mathbf{x}, \mathbf{x}')$.
- The cross-covariance is now given as

$$\operatorname{\mathsf{cov}}\left[f_{d}(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') \mathrm{d}\mathbf{z} \mathrm{d}\mathbf{z}'$$



A process $u(\mathbf{x})$ with covariance $k(\mathbf{x}, \mathbf{x}')$

☐ The cross-covariance is

$$\operatorname{\mathsf{cov}}\left[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \int_{\mathcal{X}} \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') \mathrm{d}\mathbf{z} \mathrm{d}\mathbf{z}'$$

Example. Assume that the smoothing kernels and the covariance for $u(\mathbf{x})$ follow a Gaussian form

$$\begin{split} G_d(\mathbf{x} - \mathbf{z}) &= \frac{S_d |\mathbf{P}_d|^{1/2}}{(2\pi)^{p/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \mathbf{z})^\top \mathbf{P}_d(\mathbf{x} - \mathbf{z})\right], \\ k(\mathbf{z}, \mathbf{z}') &= \frac{|\mathbf{\Lambda}|^{1/2}}{(2\pi)^{p/2}} \exp\left[-\frac{1}{2}(\mathbf{z} - \mathbf{z}')^\top \mathbf{\Lambda} (\mathbf{z} - \mathbf{z}')\right], \end{split}$$

Using again the identities of products of two Gaussians, we get

$$\operatorname{cov}\left[f_{d}(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}) d\mathbf{z} \\
= \frac{S_{d} S_{d'}}{(2\pi)^{p/2} |\mathbf{P}_{eqv}|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \mathbf{x}')^{\top} \mathbf{P}_{eqv}^{-1} (\mathbf{x} - \mathbf{x}')\right],$$

where $P_{eqv} = P_{d}^{-1} + P_{d'}^{-1} + \Lambda^{-1}$.



A process $u(\mathbf{x})$ with covariance $k(\mathbf{x}, \mathbf{x}')$

■ The cross-covariance is

$$\operatorname{\mathsf{cov}}\left[f_{d}(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') \mathrm{d}\mathbf{z} \mathrm{d}\mathbf{z}'$$

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ight],$$
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A process $u(\mathbf{x})$ with covariance $k(\mathbf{x}, \mathbf{x}')$

The cross-covariance is

$$\operatorname{cov}\left[f_{d}(\mathbf{x}), f_{d'}(\mathbf{x}')\right] = \int_{\mathcal{X}} \int_{\mathcal{X}} G_{d}(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}') k(\mathbf{z}, \mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

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Using again the identities of products of two Gaussians, we get

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where $\mathbf{P}_{eqv} = \mathbf{P}_{d}^{-1} + \mathbf{P}_{d'}^{-1} + \mathbf{\Lambda}^{-1}$.



More general process convolutions

We can include more latent processes $u_1(\mathbf{x}), u_2(\mathbf{x}), \dots, u_Q(\mathbf{x})$

$$f_d(\mathbf{x}) = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x} - \mathbf{z}) u_q^i(\mathbf{z}) \mathrm{d}\mathbf{z},$$

where $\operatorname{cov}[u_q^i(\mathbf{z}), u_{q'}^{i'}(\mathbf{z}')] = k_q(\mathbf{z}, \mathbf{z}')\delta_{i,i'}\delta_{q,q'}$.

□ A general expression for $cov[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$ follows as

$$\sum_{q=1}^{Q} \sum_{i=1}^{n_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x} - \mathbf{z}) \int_{\mathcal{X}} G_{d',q}^i(\mathbf{x}' - \mathbf{z}') k_q(\mathbf{z}, \mathbf{z}') d\mathbf{z}' d\mathbf{z}.$$

More general process convolutions

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 \Box A general expression for cov $[f_d(\mathbf{x}), f_{d'}(\mathbf{x}')]$ follows as

$$\sum_{q=1}^{Q} \sum_{i=1}^{H_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x} - \mathbf{z}) \int_{\mathcal{X}} G_{d',q}^i(\mathbf{x}' - \mathbf{z}') k_q(\mathbf{z}, \mathbf{z}') d\mathbf{z}' d\mathbf{z}.$$

Assume we have D outputs, $\{f_d(\mathbf{x})\}_{d=1}^D$. The covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$ follows

$$k_{f_d,f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x}-\mathbf{z}) \int_{\mathcal{X}} G_{d',q}^i(\mathbf{x}'-\mathbf{z}') k_q(\mathbf{z},\mathbf{z}') d\mathbf{z}' d\mathbf{z}.$$

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Some particular cases:

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Some particular cases:

Intrinsic Coregionalization Model [Goovaerts, 1997] or Multi-task Gaussian Processes [Bonilla et al., 2008]

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$$G_{d,q}^i(\mathbf{x}-\mathbf{z})=a_{d,q}^i\delta(\mathbf{x}-\mathbf{z}), \quad Q=1, \quad R_q>1$$

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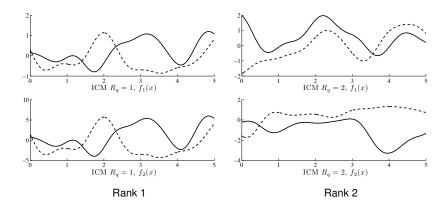
$$G_{d,q}^{i}(\mathbf{x} - \mathbf{z}) = a_{d,q}^{i} \delta(\mathbf{x} - \mathbf{z}), \quad Q = 1, \quad R_{q} > 1$$
 $k_{f_{d},f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{i=1}^{R_{1}} a_{d,1}^{i} a_{d',1}^{i} k_{1}(\mathbf{x},\mathbf{x}').$

Intrinsic Coregionalization Model

$$\textbf{K}_{\textbf{f},\textbf{f}} = \textbf{B} \otimes \textbf{K}$$

Intrinsic Coregionalization Model

$$\textbf{K}_{\textbf{f},\textbf{f}} = \textbf{B} \otimes \textbf{K}$$



Assume we have D outputs, $\{f_d(\mathbf{x})\}_{d=1}^D$. The covariance between $f_d(\mathbf{x})$ and $f_{d'}(\mathbf{x}')$ follows [Higdon, 2002, Boyle and Frean, 2005, Álvarez et al., 2012]

$$k_{f_d,f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x}-\mathbf{z}) \int_{\mathcal{X}} G_{d',q}^i(\mathbf{x}'-\mathbf{z}') k_q(\mathbf{z},\mathbf{z}') d\mathbf{z}' d\mathbf{z}.$$

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Some particular cases:

Semiparametric Latent Factor Model [Teh et al., 2005]

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Some particular cases:

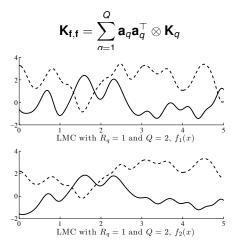
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 $k_{f_{d},f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} a_{d,q}^{1} a_{d',q}^{1} k_{q}(\mathbf{x},\mathbf{x}').$

Semiparametric Latent Factor Model

$$\mathbf{K}_{\mathbf{f},\mathbf{f}} = \sum_{q=1}^{Q} \mathbf{a}_q \mathbf{a}_q^{ op} \otimes \mathbf{K}_q$$

Semiparametric Latent Factor Model



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Some particular cases:

Linear Model of Coregionalization [Journel and Huijbregts, 1978, Goovaerts, 1997, Wackernagel, 2003].

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Linear Model of Coregionalization [Journel and Huijbregts, 1978, Goovaerts, 1997, Wackernagel, 2003].

$$G_{d,q}^i(\mathbf{x} - \mathbf{z}) = a_{d,q}^i \delta(\mathbf{x} - \mathbf{z}), \quad R_q > 1, \quad Q > 1,$$
 $k_{f_d,f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} a_{d,q}^i a_{d',q}^i k_q(\mathbf{x},\mathbf{x}').$

Linear Model of Coregionalization

$$\mathbf{K}_{\mathbf{f},\mathbf{f}} = \sum_{q=1}^{Q} \mathbf{B}_q \otimes \mathbf{K}_q$$

Linear Model of Coregionalization

$$\mathbf{K_{f,f}} = \sum_{q=1}^{Q} \mathbf{B_q} \otimes \mathbf{K_q}$$

$$\overset{5}{\underset{\text{LMC with } R_q = 2 \text{ and } Q = 2, f_1(x)}{\underset{\text{LMC with } R_q = 2 \text{ and } Q = 2, f_2(x)}{\underset{\text{LMC with } R_q = 2$$

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$$k_{f_d,f_{d'}}(\mathbf{x},\mathbf{x}') = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} \int_{\mathcal{X}} G_{d,q}^i(\mathbf{x}-\mathbf{z}) \int_{\mathcal{X}} G_{d',q}^i(\mathbf{x}'-\mathbf{z}') k_q(\mathbf{z},\mathbf{z}') d\mathbf{z}' d\mathbf{z}.$$

Some particular cases:

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$$Q=1, \quad R_q=1 \quad k_1(\mathbf{z},\mathbf{z}')=\sigma^2\delta(\mathbf{z},\mathbf{z}'),$$

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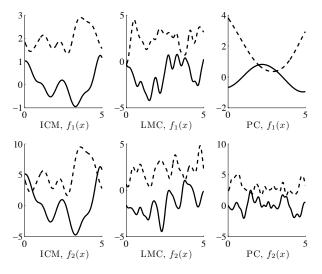
Some particular cases:

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$$\begin{split} Q &= 1, \quad R_q = 1 \quad k_1(\mathbf{z}, \mathbf{z}') = \sigma^2 \delta(\mathbf{z}, \mathbf{z}'), \\ k_{f_d, f_{d'}}(\mathbf{x}, \mathbf{x}') &= \sigma^2 \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) G_{d'}(\mathbf{x}' - \mathbf{z}) \mathrm{d}\mathbf{z}. \end{split}$$

Comparison

Comparison



Kernels for vector-valued functions

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Kernels for Vector-Valued Functions: A Review

By Mauricio A. Álvarez, Lorenzo Rosasco and Neil D. Lawrence

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Gaussian process priors for vector-valued functions

We saw a series of models for the set of outputs $\{f_d(\mathbf{x})\}_{d=1}^D$, that led to a valid covariance function for the vector $\mathbf{f}(\mathbf{x})$.

□ For a finite number of inputs, $\mathbf{X} = \{\mathbf{x}_n\}_{n=1}^N$, the prior distribution over the vector $\mathbf{f} = [\mathbf{f}_1^\top, \dots, \mathbf{f}_D^\top]^\top$ is given as

$$\begin{bmatrix} f_1 \\ f_2 \\ \vdots \\ f_D \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \\ \vdots \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{K}_{f_1,f_1} & \mathbf{K}_{f_1,f_2} & \cdots & \mathbf{K}_{f_1,f_D} \\ \mathbf{K}_{f_2,f_1} & \mathbf{K}_{f_2,f_2} & \cdots & \mathbf{K}_{f_2,f_D} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{K}_{f_D,f_1} & \mathbf{K}_{f_D,f_2} & \cdots & \mathbf{K}_{f_D,f_D} \end{bmatrix} \right).$$

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$$\begin{bmatrix} f_1 \\ f_2 \\ \vdots \\ f_D \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} K_{f_1,f_1} & K_{f_1,f_2} & \cdots & K_{f_1,f_D} \\ K_{f_2,f_1} & K_{f_2,f_2} & \cdots & K_{f_2,f_D} \\ \vdots & \vdots & \cdots & \vdots \\ K_{f_D,f_1} & K_{f_D,f_2} & \cdots & K_{f_D,f_D} \end{bmatrix} \right).$$

$$f \qquad \qquad K_{f,f}$$

Noisy observations

In practice, we usually have access to noisy observations, so we model the outputs $\{y_d(\mathbf{x})\}_{d=1}^D$ using

$$y_d(\mathbf{x}) = f_d(\mathbf{x}) + \epsilon_d(\mathbf{x}),$$

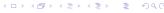
where $\{\epsilon_d(\mathbf{x})\}_{d=1}^D$ are independent white Gaussian noise processes with variance σ_d^2 .

The marginal likelihood is given as

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}_{\mathsf{f},\mathsf{f}} + \boldsymbol{\Sigma}),$$

where
$$\mathbf{y} = \begin{bmatrix} \mathbf{y}_1^\top, \mathbf{y}_2^\top \dots, \mathbf{y}_D^\top \end{bmatrix}^\top$$

The vector θ refers to the hyperparameters and $\Sigma = \Sigma \otimes I_N$.



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The vector θ refers to the hyperparameters and $\Sigma = \Sigma \otimes I_N$.



Hyperparameter Learning

- Let $\mathcal{D} = \{\mathbf{X}_n, \mathbf{y}_n\}_{n=1}^N$ represents the data, and θ represents the hyperparameters of the covariance function.
- □ The marginal likelihood for the outputs can be written as

$$p(\mathbf{y}|\mathbf{X}, oldsymbol{ heta}) = \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}_{\mathsf{f},\mathsf{f}} + oldsymbol{\Sigma}),$$

where $\mathbf{K}_{\mathbf{f},\mathbf{f}} \in \mathbb{R}^{ND \times ND}$ with each element given by $\text{cov}[f_d(\mathbf{x}_n), f_{d'}(\mathbf{x}_{n'})]$.

- The matrix Σ represents the covariance associated with some independent processes.
- Hyperparameters are estimated by maximizing the logarithm of the marginal likelihood.

Predictive distribution

Prediction for a set of test inputs X_{*} is done using standard Gaussian process regression techniques.

The predictive distribution is given by

$$\rho(\mathbf{y}_*|\mathbf{y},\mathbf{X},\boldsymbol{\theta}) = \mathcal{N}(\mathbf{y}_*|\boldsymbol{\mu}_*,\mathbf{K}_{\mathbf{y}_*,\mathbf{y}_*}),$$

with

$$\begin{split} \boldsymbol{\mu}_* &= \mathbf{K}_{\mathsf{f}_*,\mathsf{f}} \left(\mathbf{K}_{\mathsf{f},\mathsf{f}} + \boldsymbol{\Sigma} \right)^{-1} \mathbf{y}, \\ \mathbf{K}_{\mathsf{y}_*,\mathsf{y}_*} &= \mathbf{K}_{\mathsf{f}_*,\mathsf{f}_*} - \mathbf{K}_{\mathsf{f}_*,\mathsf{f}} \left(\mathbf{K}_{\mathsf{f},\mathsf{f}} + \boldsymbol{\Sigma} \right)^{-1} \mathbf{K}_{\mathsf{f}_*,\mathsf{f}}^\top + \boldsymbol{\Sigma}_*. \end{split}$$

Can you prove autokrigeability?

The predictive distribution is given by

$$p(\mathbf{y}_*|\mathbf{y},\mathbf{X},oldsymbol{ heta}) = \mathcal{N}(\mathbf{y}_*|oldsymbol{\mu}_*,\mathbf{K}_{\mathbf{y}_*,\mathbf{y}_*}),$$

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Exercise: Prove that if the outputs are considered to be noise-free, prediction using the ICM under an isotopic data case is equivalent to independent prediction over each output.

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The cokriging estimator

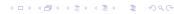
- In geostatistics, the framework that allows for optimal predictions in the multivariate case is known by the general name of *cokriging* [Goovaerts, 1997].
- \Box In general, the output value for f_d evaluated at \mathbf{x}_* is estimated as

$$\hat{f}_{d}(\mathbf{x}_{*}) - \mu_{d}(\mathbf{x}_{*}) = \sum_{s=1}^{D} \sum_{\alpha_{s}=1}^{n_{s}(\mathbf{x}_{*})} \lambda_{\alpha_{s}}(\mathbf{x}_{*}) \left[f_{s}(\mathbf{x}_{\alpha_{s}}) - \mu_{s}(\mathbf{x}_{\alpha_{s}}) \right],$$

where $\lambda_{\alpha_s}(\mathbf{x}_*)$ are the weights assigned to the output data $f_s(\mathbf{x}_{\alpha_s})$, $\mu_s(\mathbf{x}_{\alpha_s})$ are the expected values of $f_s(\mathbf{x}_{\alpha_s})$, and $n_s(\mathbf{x}_*) \leq N$.

Cokriging estimators need to be unbiased $(E[f_d(\mathbf{x}_*) - \hat{f}_d(\mathbf{x}_*)] = 0)$ and minimize the error variance σ_E^2 ,

$$\sigma_E^2(\mathbf{x}_*) = \operatorname{var}\left[f_d(\mathbf{x}_*) - \hat{f}_d(\mathbf{x}_*)\right].$$



Cokriging assumes a model for f_d

 \Box Cogriking estimators differ in the form they assume for $f_d(\mathbf{x})$.

In general, each output function is decomposed into a residual $R_d(\mathbf{x})$ and a trend $\mu_d(\mathbf{x})$,

$$f_d(\mathbf{x}) = R_d(\mathbf{x}) + \mu_d(\mathbf{x}), \quad \forall d$$

Residuals are assumed to be Gaussian processes with zero mean.

The covariance for the residuals is denoted as $k_{d,d}(\mathbf{x}, \mathbf{x}')$ and the cross-covariance between residuals as $k_{d,d'}(\mathbf{x}, \mathbf{x}')$.

Simple cokriging

The simple cokriging estimator is given as

$$\hat{f}_{d}(\mathbf{x}_{*}) - \mu_{d} = \sum_{s=1}^{D} \sum_{\alpha_{s}=1}^{n_{s}(\mathbf{x}_{*})} \lambda_{\alpha_{s}}(\mathbf{x}_{*}) \left[f_{s}(\mathbf{x}_{\alpha_{s}}) - \mu_{s} \right) \right].$$

- It can be shown that this is an unbiased estimator.
- Coefficients $\lambda_{\alpha_s}(\mathbf{x}_*)$ can be obtained by minimizing the variance $\sigma_E^2(\mathbf{x}_*)$, leading to

$$\begin{bmatrix} \boldsymbol{\lambda}_1(\boldsymbol{x}_*) \\ \vdots \\ \boldsymbol{\lambda}_D(\boldsymbol{x}_*) \end{bmatrix} = \begin{pmatrix} \begin{bmatrix} \boldsymbol{K}_{1,1} & \cdots & \boldsymbol{K}_{1,D} \\ \vdots & \vdots & \vdots \\ \boldsymbol{K}_{D,1} & \cdots & \boldsymbol{K}_{D,D} \end{bmatrix} \end{pmatrix}^{-1} \begin{bmatrix} \boldsymbol{k}_{1,1} \\ \vdots \\ \boldsymbol{k}_{D,1} \end{bmatrix}$$

where
$$\mathbf{K}_{d,d'} = [k_{d,d'}(\mathbf{x}_{\alpha_d}, \mathbf{x}_{\beta_{d'}})]$$
 and $\mathbf{k}_{d,1} = [k_{d,1}(\mathbf{x}_{\alpha_d}, \mathbf{x}_*)].$

 $lue{}$ The predictor is then $\hat{f}_d(\mathbf{x}_*) = oldsymbol{\lambda}^ op \mathbf{f}$.



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Efficient approximations (I)

Learning θ through marginal likelihood maximization involves the inversion of the matrix $\mathbf{K}_{\mathbf{f},\mathbf{f}} + \Sigma$.

The inversion of this matrix scales as $\mathcal{O}(D^3N^3)$.

□ If only a few number K < N of values of $u(\mathbf{x})$ are known, then the set of outputs are uniquely determined.

Efficient approximations (II)

Sample from
$$p(u)$$

$$f_d(\mathbf{x}) = \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}$$

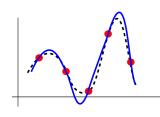
Efficient approximations (II)

Sample from p(u)



$$f_{\mathcal{J}}(\mathbf{x}) = \int_{\mathcal{X}} G_{\mathcal{J}}(\mathbf{x} - \mathbf{z}) u(\mathbf{z}) d\mathbf{z}$$

Sample from $p(u|\mathbf{u})$



$$f_d(\mathbf{x}) pprox \int_{\mathcal{X}} G_d(\mathbf{x} - \mathbf{z}) \, \mathsf{E} \left[u(\mathbf{z}) | \mathbf{u} \right] \mathrm{d}\mathbf{z}$$

Efficient approximations

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Computationally Efficient Convolved Multiple Output Gaussian Processes

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Cross-coregionalization matrices

In the LMC

$$f_d(\mathbf{x}) = \sum_{q=1}^{Q} \sum_{i=1}^{R_q} a_{d,q}^i u_q^i(\mathbf{x}).$$

The basic processes $u_q^i(\mathbf{x})$ [Guzmán et al., 2002] are assumed to be nonorthogonal, leading to the following covariance function

$$\mathsf{cov}[\mathbf{f}(\mathbf{x}),\mathbf{f}(\mathbf{x}')] = \sum_{q=1}^Q \sum_{q'=1}^Q \mathbf{B}_{q,q'} k_{q,q'}(\mathbf{x},\mathbf{x}'),$$

where $\mathbf{B}_{q,q'}$ are *cross-coregionalization* matrices. matrices.

Non-stationarity LMC

 \Box We can write the vector-valued function $\mathbf{f}(\mathbf{x})$ as

$$f(x) = Au(x),$$

where
$$\mathbf{A} = [\mathbf{a}_1 \cdots \mathbf{a}_Q]$$
 and $\mathbf{u}(\mathbf{x}) = [u_1(\mathbf{x}) \cdots u_Q(\mathbf{x})]^{\top}$.

 A non-stationary version allows A to change with x [Gelfand et al., 2004, Wilson et al., 2012]

$$f(x) = A(x)u(x).$$

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Extensions [Calder and Cressie, 2007]

A more general form

$$f_d(\mathbf{x}) = \int G_d(\mathbf{x}, \mathbf{z}) u(\mathbf{z}) d\mathbf{z}$$

 $f_d(\mathbf{x}) = \sum_j G_d(\mathbf{x}, \mathbf{z}_j) u(\mathbf{z}_j)$

Non-stationary models

$$f_d(\mathbf{x}) = \int G_{d,\theta(\mathbf{x})}(\mathbf{x},\mathbf{z})u(\mathbf{z})\mathrm{d}\mathbf{z},$$
 $f_d(\mathbf{x}) = \int G_d(\mathbf{x},\mathbf{z})u_{\theta(\mathbf{z})}(\mathbf{x})\mathrm{d}\mathbf{z}$

Latent force models [Álvarez et al., 2009]

Mechanistically inspired kernel smoothing functions.

$$\begin{array}{ll} G_d(t,t') \propto \exp{\left[-D_q\left(t-t'\right)\right]} & \text{first ODE} \\ G_d(t,t') \propto \exp{\left[-\alpha_q\left(t-t'\right)\right]} \sin{\left[\omega_q\left(t-t'\right)\right]} & \text{second ODE} \\ G_d(\mathbf{x},\mathbf{x}') = \exp{\left[-\sum_i \frac{(x_i-x_i')^2}{4C}\right]} & \text{PDE} \end{array}$$

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- We can do multi-task learning or transfer learning with GPs.
- Different ways to build meaningful cross-covariance functions.
- Once defined, we can do all the things we know to do with a single-output GP.
- Cokriging is just prediction with GPs (with a quadratic loss function).
- Several extensions of LMC and PCs.
- Current research: spectral representations for the joint covariance function.

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