

- 1) Consider the probability space  $(\Omega, \mathcal{G}, P)$ , where  $\Omega = [0, 2]$  and  $\mathcal{G} = \{\emptyset, \Omega, [0, 1], (1, 2]\}$ . Let  $X$  and  $Y$  be two functions on  $\Omega$  defined as

$$X(\omega) = \begin{cases} 1 & \text{if } \omega \in [0, 1] \\ 2 & \text{if } \omega \in (1, 2] \end{cases}$$

and

$$Y(\omega) = \begin{cases} 2 & \text{if } \omega \in [0, 1.5] \\ 3 & \text{if } \omega \in (1.5, 2]. \end{cases}$$

Then which one of the following statements is true?

- a)  $X$  is a random variable with respect to  $\mathcal{G}$ , but  $Y$  is not a random variable with respect to  $\mathcal{G}$
  - b)  $Y$  is a random variable with respect to  $\mathcal{G}$ , but  $X$  is not a random variable with respect to  $\mathcal{G}$
  - c) Neither  $X$  nor  $Y$  is a random variable with respect to  $\mathcal{G}$
  - d) Both  $X$  and  $Y$  are random variables with respect to  $\mathcal{G}$
- 2) Let  $\Phi(\cdot)$  denote the cumulative distribution function of a standard normal random variable. If the random variable  $X$  has the cumulative distribution function

$$F(x) = \begin{cases} \Phi(x) & \text{if } x < -1 \\ \Phi(x+1) & \text{if } x \geq -1, \end{cases}$$

then which one of the following statements is true?

- a)  $P(X \leq -1) = \frac{1}{2}$
  - b)  $P(X = -1) = \frac{1}{2}$
  - c)  $P(X < -1) = \frac{1}{2}$
  - d)  $P(X \leq 0) = \frac{1}{2}$
- 3) Let  $X$  be a random variable with probability density function

$$f(x) = \begin{cases} \alpha \lambda x^{\alpha-1} e^{-\lambda x^\alpha} & \text{if } x > 0 \\ 0 & \text{otherwise,} \end{cases}$$

where  $\alpha > 0$  and  $\lambda > 0$ . If the median of  $X$  is 1 and the third quantile is 2, then  $(\alpha, \lambda)$  equals

- a)  $(1, \log_e 2)$
- b)  $(1, 1)$
- c)  $(2, \log_e 2)$
- d)  $(1, \log_e 3)$

- 4) Let  $X$  be a random variable having Poisson distribution with mean  $\lambda > 0$ . Then  $E\left(\frac{1}{X+1} \mid X > 0\right)$  equals
- $\frac{1-e^{-\lambda}-\lambda e^{-\lambda}}{\lambda(1-e^{-\lambda})}$
  - $\frac{1-e^{-\lambda}}{\lambda}$
  - $\frac{1-e^{-\lambda}-\lambda e^{-\lambda}}{1-e^{-\lambda}}$
  - $\frac{1-e^{-\lambda}}{\lambda+1}$

- 5) Suppose that  $X$  has the probability density function

$$f(x) = \begin{cases} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} & \text{if } x > 0 \\ 0 & \text{otherwise,} \end{cases}$$

where  $\alpha > 0$  and  $\lambda > 0$ . Which one of the following statements is NOT true?

- $E(X)$  exists for all  $\alpha > 0$  and  $\lambda > 0$
  - Variance of  $X$  exists for all  $\alpha > 0$  and  $\lambda > 0$
  - $E\left(\frac{1}{X}\right)$  exists for all  $\alpha > 0$  and  $\lambda > 0$
  - $E(\log_e(1+X))$  exists for all  $\alpha > 0$  and  $\lambda > 0$
- 6) Let  $(X, Y)$  have joint probability density function

$$f(x, y) = \begin{cases} 8xy & \text{if } 0 < x < y < 1 \\ 0 & \text{otherwise.} \end{cases}$$

If  $E(X|Y = y_0) = \frac{1}{2}$ , then  $y_0$  equals

- $\frac{3}{4}$
  - $\frac{1}{2}$
  - $\frac{1}{3}$
  - $\frac{2}{3}$
- 7) Suppose that there are 5 boxes, each containing 3 blue pens, 1 red pen and 2 black pens. One pen is drawn at random from each of these 5 boxes. If the random variable  $X_1$  denotes the total number of blue pens drawn and the random variable  $X_2$  denotes the total number of red pens drawn, then  $P(X_1 = 2, X_2 = 1)$  equals
- $\frac{5}{36}$
  - $\frac{5}{18}$
  - $\frac{5}{12}$
  - $\frac{5}{9}$
- 8) Let  $\{X_n\}_{n \geq 1}$  and  $\{Y_n\}_{n \geq 1}$  be two sequences of random variables and  $X$  and  $Y$  be two random variables, all of them defined on the same probability space. Which one of the following statements is true?
- If  $\{X_n\}_{n \geq 1}$  converges in distribution to a real constant  $c$ , then  $\{X_n\}_{n \geq 1}$  converges in probability to  $c$
  - If  $\{X_n\}_{n \geq 1}$  converges in probability to  $X$ , then  $\{X_n\}_{n \geq 1}$  converges in 3<sup>rd</sup> mean to  $X$
  - If  $\{X_n\}_{n \geq 1}$  converges in distribution to  $X$  and  $\{Y_n\}_{n \geq 1}$  converges in distribution to  $Y$ , then  $\{X_n + Y_n\}_{n \geq 1}$  converges in distribution to  $X + Y$
  - If  $\{E(X_n)\}_{n \geq 1}$  converges to  $E(X)$ , then  $\{X_n\}_{n \geq 1}$  converges in 1<sup>st</sup> mean to  $X$

- 9) Let  $X$  be a random variable with probability density function

$$f(x; \lambda) = \begin{cases} \frac{1}{\lambda} e^{-\frac{x}{\lambda}} & \text{if } x > 0 \\ 0 & \text{otherwise,} \end{cases}$$

where  $\lambda > 0$  is an unknown parameter. Let  $Y_1, Y_2, \dots, Y_n$  be a random sample of size  $n$  from a population having the same distribution as  $X^2$ .

If  $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$ , then which one of the following statements is true?

- $\sqrt{\frac{\bar{Y}}{2}}$  is a method of moments estimator of  $\lambda$
  - $\sqrt{\bar{Y}}$  is a method of moments estimator of  $\lambda$
  - $\frac{1}{2} \sqrt{\bar{Y}}$  is a method of moments estimator of  $\lambda$
  - $2 \sqrt{\bar{Y}}$  is a method of moments estimator of  $\lambda$
- 10) Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n (\geq 2)$  from a population having probability density function

$$f(x; \theta) = \begin{cases} \frac{2}{\theta x} (-\log_e x) e^{-\frac{(\log_e x)^2}{\theta}} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise,} \end{cases}$$

where  $\theta > 0$  is an unknown parameter. Then which one of the following statements is true?

- $\frac{1}{n} \sum_{i=1}^n (\log_e X_i)^2$  is the maximum likelihood estimator of  $\theta$
  - $\frac{1}{n-1} \sum_{i=1}^n (\log_e X_i)^2$  is the maximum likelihood estimator of  $\theta$
  - $\frac{1}{n} \sum_{i=1}^n \log_e X_i$  is the maximum likelihood estimator of  $\theta$
  - $\frac{1}{n-1} \sum_{i=1}^n \log_e X_i$  is the maximum likelihood estimator of  $\theta$
- 11) Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n$  from a population having uniform distribution over the interval  $(\frac{1}{3}, \theta)$ , where  $\theta > \frac{1}{3}$  is an unknown parameter. If  $Y = \max \{X_1, X_2, \dots, X_n\}$ , then which one of the following statements is true?
- $\left(\frac{n+1}{n}\right)\left(Y - \frac{1}{3}\right) + \frac{1}{3}$  is an unbiased estimator of  $\theta$
  - $\left(\frac{n}{n+1}\right)\left(Y - \frac{1}{3}\right) + \frac{1}{3}$  is an unbiased estimator of  $\theta$
  - $\left(\frac{n+1}{n}\right)\left(Y + \frac{1}{3}\right) - \frac{1}{3}$  is an unbiased estimator of  $\theta$
  - $Y$  is an unbiased estimator of  $\theta$
- 12) Suppose that  $X_1, X_2, \dots, X_n, Y_1, Y_2, \dots, Y_n$  are independent and identically distributed random vectors each having  $N_p(\mu, \Sigma)$  distribution, where  $\Sigma$  is non-singular,  $p > 1$  and  $n > 1$ . If  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$  and  $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$ , then which one of the following statements is true?
- There exists  $c > 0$  such that  $c(\bar{X} - \mu)^T \Sigma^{-1} (\bar{X} - \mu)$  has  $\chi^2$ -distribution with  $p$  degrees of freedom
  - There exists a  $c > 0$  such that  $c(\bar{X} - \bar{Y})^T \Sigma^{-1} (\bar{X} - \bar{Y})$  has  $\chi^2$ -distribution with  $(p - 1)$  degrees of freedom.
  - There exists  $c > 0$  such that  $c \sum_{i=1}^n (X_i - \bar{X})^T \Sigma^{-1} (X_i - \bar{X})$  has  $\chi^2$ -distribution with  $p$  degrees of freedom
  - There exists  $c > 0$  such that  $c \sum_{i=1}^n (X_i - Y_i - \bar{X} + \bar{Y})^T \Sigma^{-1} (X_i - Y_i - \bar{X} + \bar{Y})$  has  $\chi^2$ -distribution with  $p$  degrees of freedom

13) Consider the following regression model

$$y_k = \alpha_0 + \alpha_1 \log_e k + \epsilon_k, \quad k = 1, 2, \dots, n,$$

where the  $\epsilon_k$ 's are independent and identically distributed random variables each having probability density function  $f(x) = \frac{1}{2}e^{-|x|}$ ,  $x \in \mathbb{R}$ . Then which of the following statements is true?

- a) The maximum likelihood estimator of  $\alpha_0$  does not exist
- b) The maximum likelihood estimator of  $\alpha_1$  does not exist
- c) The least squares estimator of  $\alpha_0$  exists and is unique
- d) The least squares estimator of  $\alpha_1$  exists, but it is not unique