

Hierarchical and Mixed Effects Models in R

Marcus Becker

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Chapter 1: Overview and Introduction to Hierarchical and Mixed Models

- lme4 package

```
library(lme4)
```

```
## Loading required package: Matrix
```

```
##
```

```
## Attaching package: 'Matrix'
```

```
## The following objects are masked from 'package:tidyr':
```

```
##
```

```
##     expand, pack, unpack
```

```
library(lmerTest)
```

```
## Warning: package 'lmerTest' was built under R version 3.6.3
```

```
##
```

```
## Attaching package: 'lmerTest'
```

```
## The following object is masked from 'package:lme4':
```

```
##
```

```
##     lmer
```

```
## The following object is masked from 'package:stats':
```

```
##
```

```
##     step
```

```
library(broom.mixed)
```

```
## Warning: package 'broom.mixed' was built under R version 3.6.3
```

```
## Registered S3 methods overwritten by 'broom.mixed':
```

```
##   method      from
```

```
##   augment.lme  broom
```

```
##   augment.merMod broom
```

```
##   glance.lme    broom
```

```
##   glance.merMod broom
```

```
##   glance.stanreg broom
```

```
##   tidy.brmsfit  broom
```

```
##   tidy.gamlss   broom
```

```
##   tidy.lme      broom
```

```
##   tidy.merMod   broom
```

```

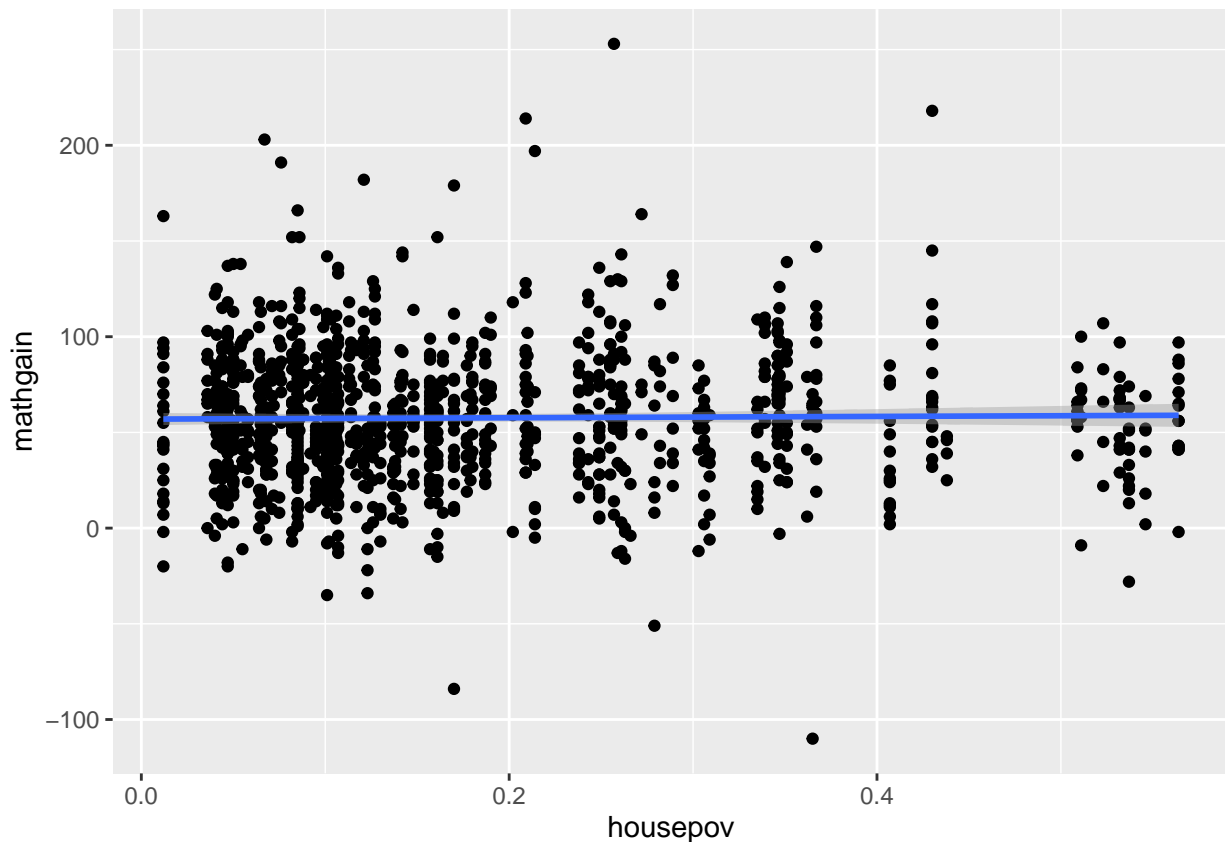
## tidy.rjags      broom
## tidy.stanfit    broom
## tidy.stanreg    broom

studentData <- read_csv("https://assets.datacamp.com/production/repositories/1803/datasets/975fe2b01908

## Parsed with column specification:
## cols(
##   sex = col_double(),
##   minority = col_double(),
##   mathkind = col_double(),
##   mathgain = col_double(),
##   ses = col_double(),
##   yearstea = col_double(),
##   mathknow = col_double(),
##   housepov = col_double(),
##   mathprep = col_double(),
##   classid = col_double(),
##   schoolid = col_double(),
##   childid = col_double()
## )

# Visualize the data first
ggplot(data = studentData, aes(x = housepov, y = mathgain)) +
  geom_point() +
  geom_smooth(method = "lm", se = TRUE)

```



```

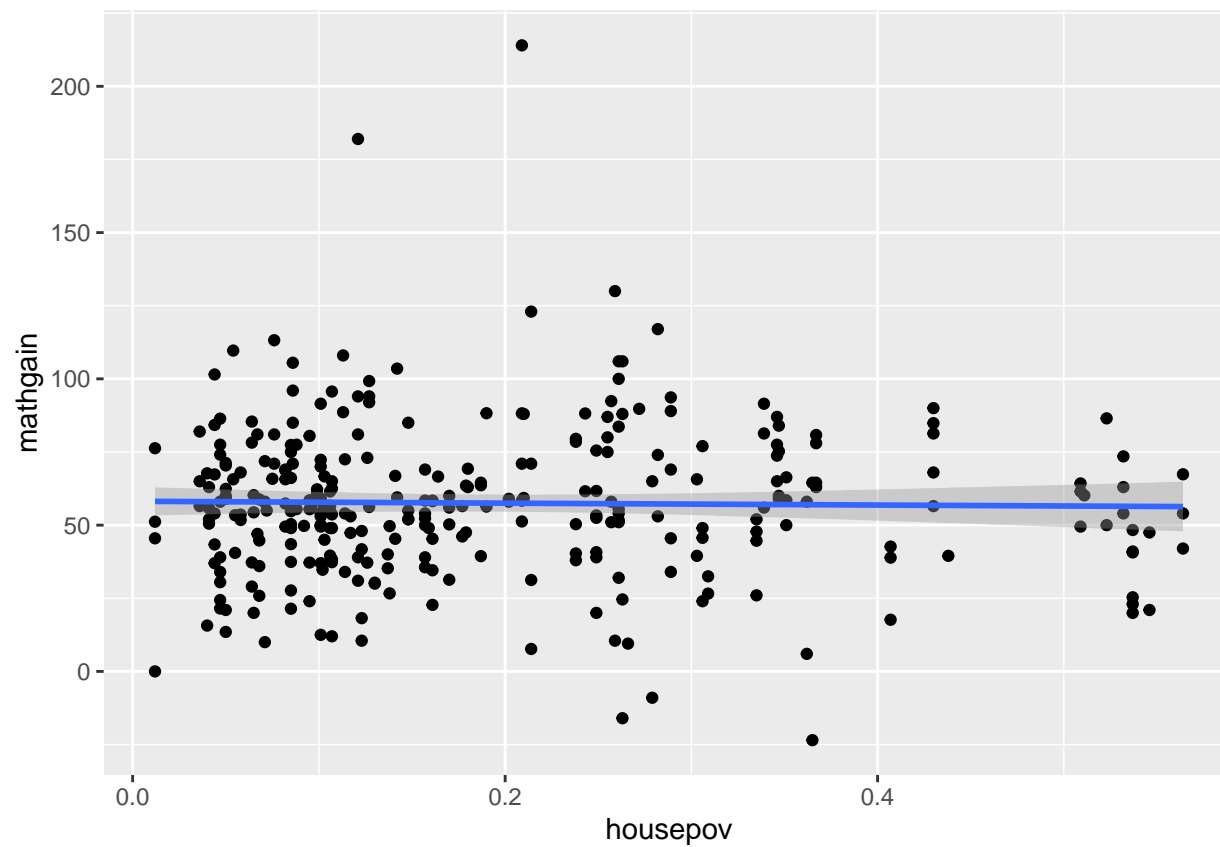
# Fit a linear model
summary(lm(formula = mathgain ~ housepov, data = studentData)) # Not predictive. We haven't accounted f

##
## Call:
## lm(formula = mathgain ~ housepov, data = studentData)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -168.226  -22.222   -1.306   19.763  195.156
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   56.937      1.674   34.02  <2e-16 ***
## housepov       3.531       7.515    0.47   0.639
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 34.63 on 1188 degrees of freedom
## Multiple R-squared:  0.0001858, Adjusted R-squared:  -0.0006558
## F-statistic: 0.2208 on 1 and 1188 DF,  p-value: 0.6385

# Class
classData <- studentData %>%
  group_by(classid) %>%
  summarise(mathgain = mean(mathgain),
            housepov = mean(housepov))

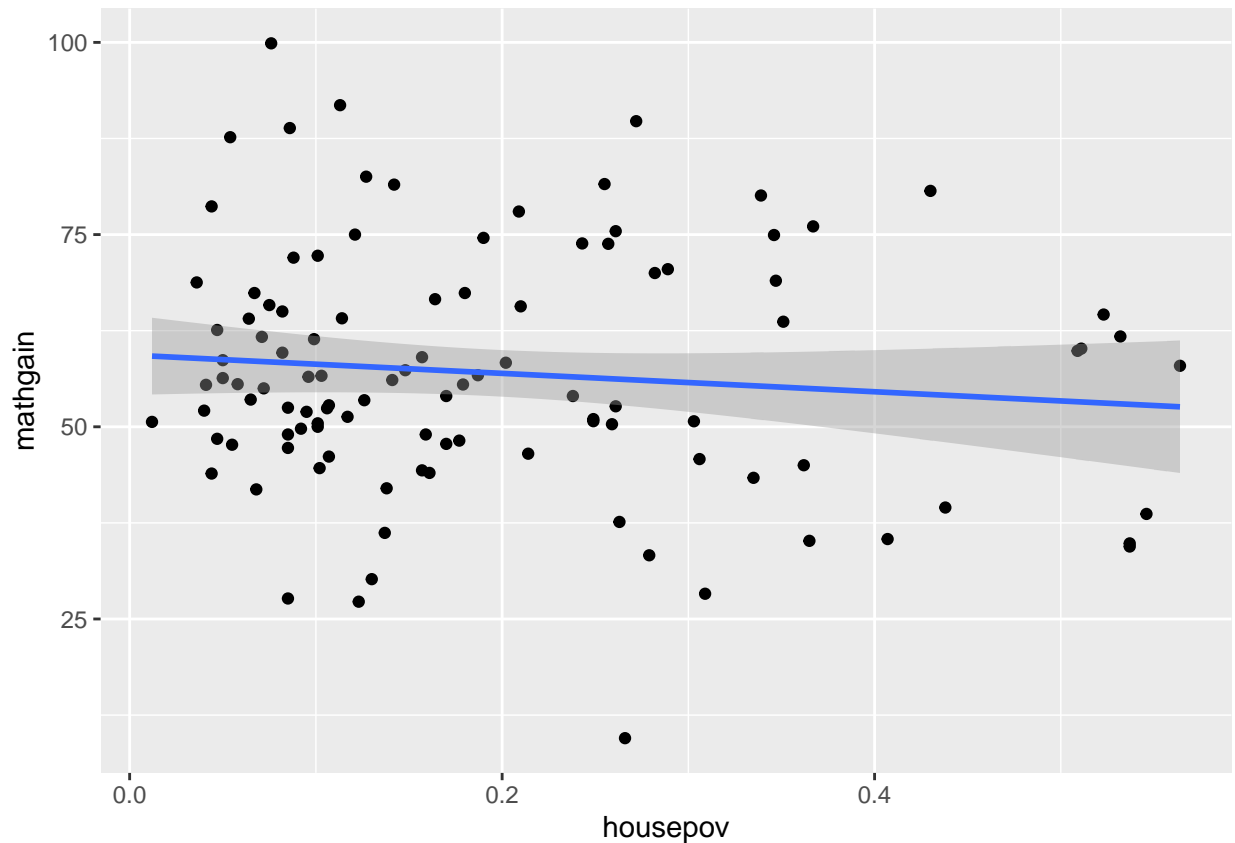
ggplot(data = classData, aes(x = housepov, y = mathgain)) +
  geom_point() +
  geom_smooth(method = "lm", se = TRUE)

```



```
# School
schoolData <- studentData %>%
  group_by(schoolid) %>%
  summarise(mathgain = mean(mathgain),
            housepov = mean(housepov))

ggplot(data = schoolData, aes(x = housepov, y = mathgain)) +
  geom_point() +
  geom_smooth(method = "lm", se = TRUE)
```



Multiple regression in R tips:

- `lm(y ~ x -1)` - the -1 estimates an intercept for each x (group), rather than relative to the first group.
- numeric vs factors - R automatically assumes a numeric predictor is a slope
- scaling parameters and slopes
- shortcut for interaction is `x1*x2`

Without other coefficients, a single intercept is the global mean of the data. Multiple intercepts allow you to estimate the mean for each group as long as other coefficients are not estimated (when the groups are discrete...).

What about continuous predictor variables? -> slopes.

```
# Mixed effects model (i.e. has both fixed and random effects)
me_mod <- lmer(formula = mathgain ~ sex + mathprep + mathknow + (1|classid) + (1|schoolid),
               data = studentData,
               na.action = "na.omit",
               REML = TRUE)

summary(me_mod)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: mathgain ~ sex + mathprep + mathknow + (1 | classid) + (1 | schoolid)
## Data: studentData
##
## REML criterion at convergence: 10677.9
##
```

```
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -4.3203 -0.6146 -0.0294  0.5467  5.5331
##
## Random effects:
##   Groups   Name      Variance Std.Dev.
##   classid  (Intercept) 103.57   10.177
##   schoolid (Intercept)  85.44    9.244
##   Residual                1019.47  31.929
## Number of obs: 1081, groups:  classid, 285; schoolid, 105
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)   52.250      3.838   233.946  13.613  <2e-16 ***
## sex           -1.526      2.041  1030.557  -0.747   0.4550
## mathprep       2.426      1.298   181.813   1.869   0.0632 .
## mathknow       2.405      1.299   206.425   1.851   0.0656 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) sex    mthprp
## sex          -0.268
## mathprep     -0.878  0.001
## mathknow     -0.003  0.011  0.005
```

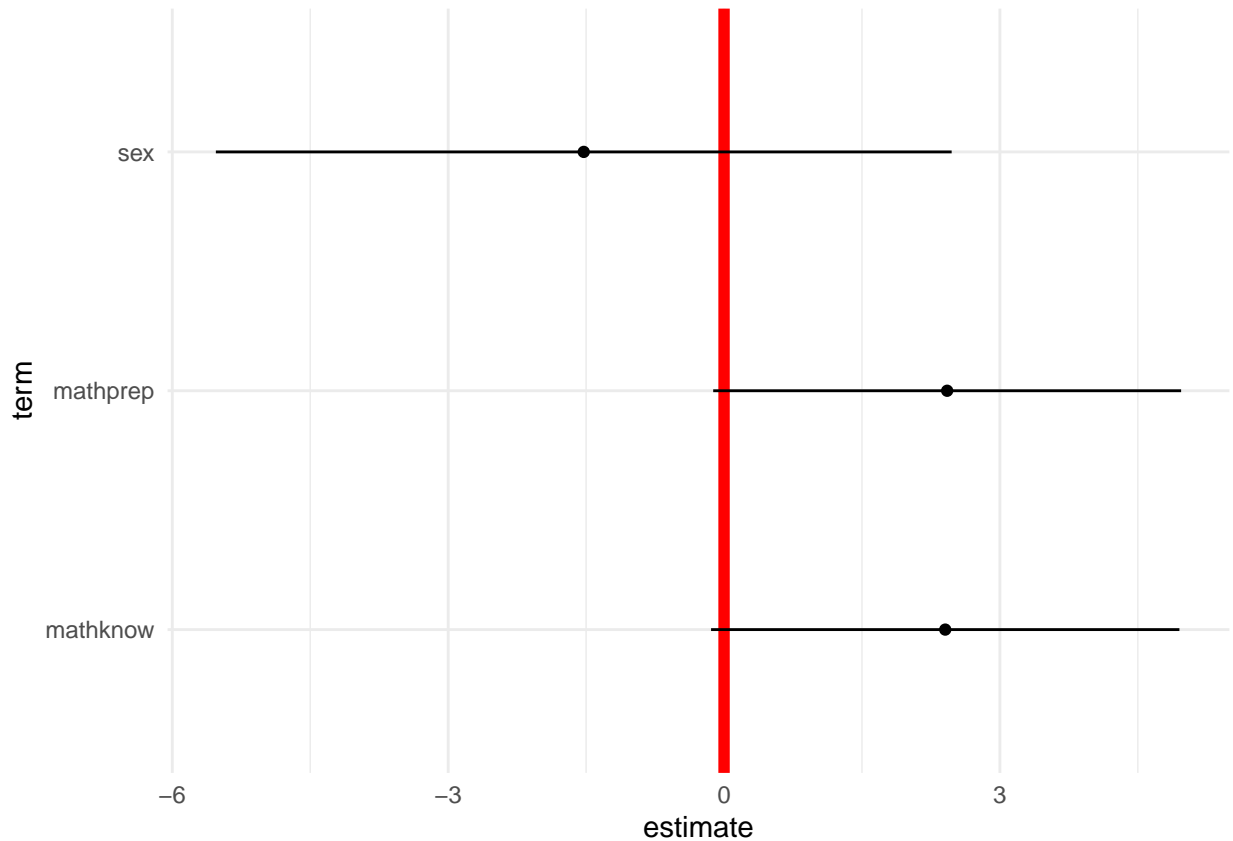
```
# Extract out the coefficients
```

```
modelOutPlot <- broom.mixed::tidy(me_mod, conf.int = TRUE) # note: have to use broom.extra for me model.
```

```
modelOutPlot <- modelOutPlot[modelOutPlot$effect == "fixed" &
                             modelOutPlot$term != "(Intercept)", ]
```

```
# Plot the coefficients of interest
```

```
ggplot(data = modelOutPlot, aes(x = term, y = estimate, ymin = conf.low, ymax = conf.high)) +
  theme_minimal() +
  geom_hline(yintercept = 0.0, color = "red", size = 2) +
  geom_point() +
  geom_linerange() +
  coord_flip()
```



```
# ~ Really like this plot!
```

Chapter 2: Linear Mixed-Effect Models

Random-effect syntax (for using lme4):

- (1 | group): random intercept with a fixed mean
- (1 | g1/g2): intercepts vary among g1 and g2 within g2
- (1 | g1) + (1 | g2): random intercepts for two variables
- x + (x | g): correlated random slope and intercept
- x + (x || g): uncorrelated random slope and intercept

```
# Using birth data
```

```
countyBirthsData <- read_csv("https://assets.datacamp.com/production/repositories/1803/datasets/eb95cb69")
```

```
## Warning: Missing column names filled in: 'X1' [1]
```

```
## Parsed with column specification:
```

```
## cols(
```

```
##   X1 = col_double(),
```

```
##   Year = col_double(),
```

```
##   TotalPopulation = col_double(),
```

```
##   BirthRate = col_double(),
```

```
##   AverageBirthWeight = col_double(),
```

```
##   AverageAgeofMother = col_double(),
```

```
##   CountyName = col_character(),
```

```
##   State = col_character()
```

```
## )
```

```

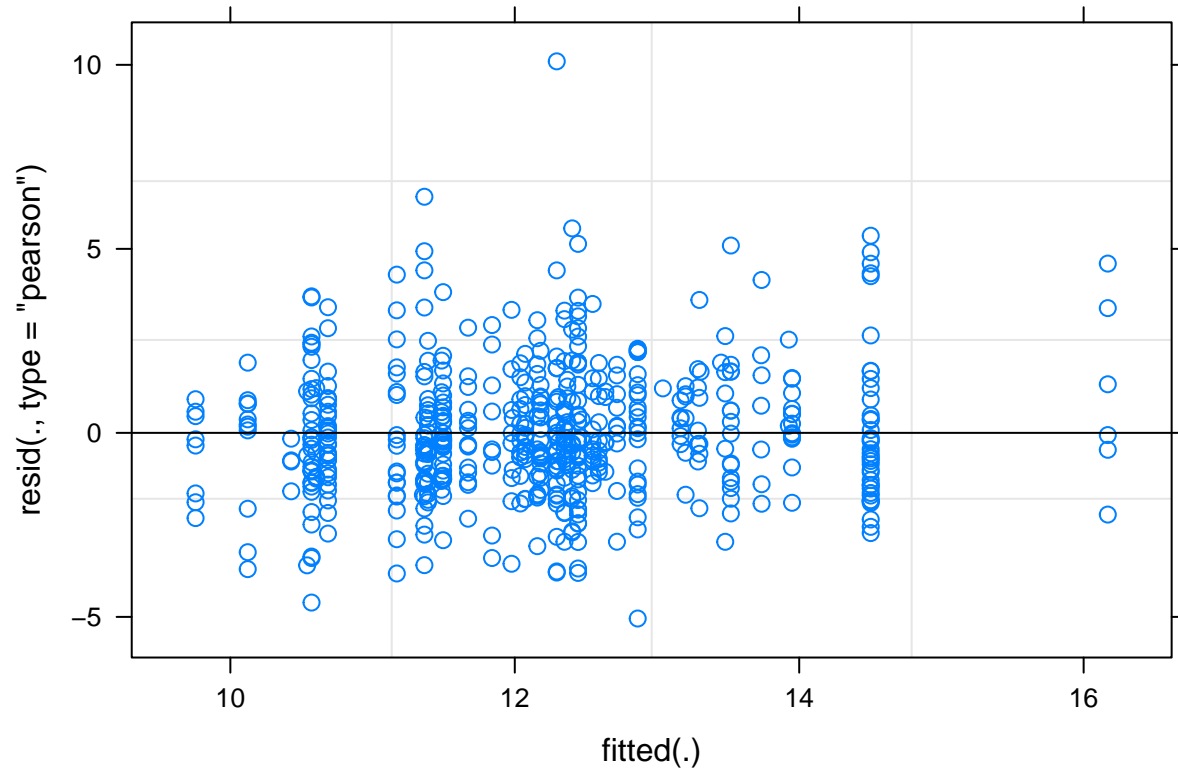
# Counties exist within states, and perhaps states contribute to variability. Hence, the need for random
# To start, we build a hierarchical model with a global intercept (fixed-effect) and random-effect for
mod1 <- lmer(formula = BirthRate ~ (1 | State),
             data = countyBirthsData)

summary(mod1)

## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: BirthRate ~ (1 | State)
## Data: countyBirthsData
##
## REML criterion at convergence: 2411
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.7957 -0.6056 -0.1063  0.5211  5.5948
##
## Random effects:
## Groups   Name      Variance Std.Dev.
## State    (Intercept) 1.899    1.378
## Residual                3.256    1.804
## Number of obs: 578, groups: State, 50
##
## Fixed effects:
##              Estimate Std. Error    df t value Pr(>|t|)
## (Intercept)  12.3362    0.2216 43.3830  55.67  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

# plot the residuals
plot(mod1)

```

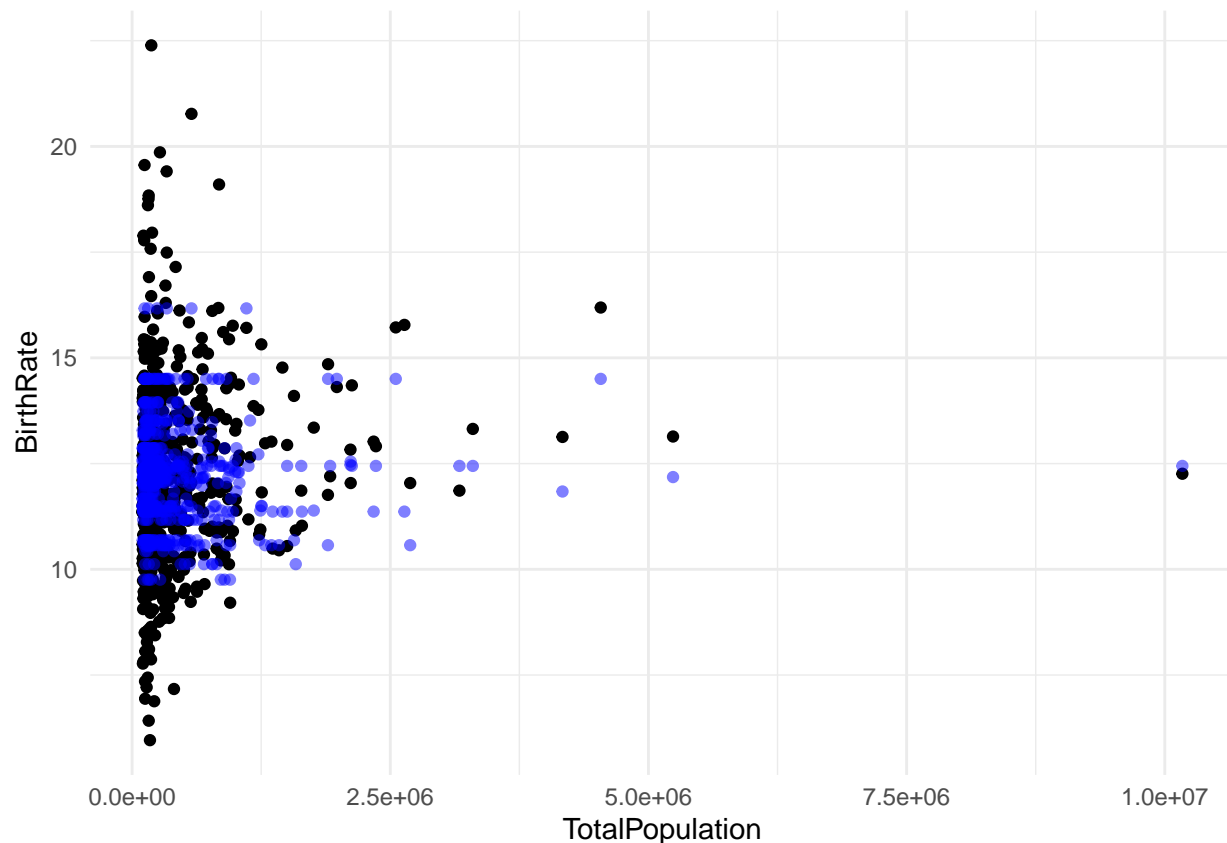



```
countyBirthsData$predictState <- predict(mod1, countyBirthsData)
```

```
ggplot(data = countyBirthsData) +  
  theme_minimal() +  
  geom_point(aes(x = TotalPopulation, y = BirthRate)) +  
  geom_point(aes(x = TotalPopulation, y = predictState),  
            color = "blue", alpha = 0.5)
```

```
## Warning: Removed 2 rows containing missing values (geom_point).
```

```
## Warning: Removed 2 rows containing missing values (geom_point).
```



Random-effects intercepts estimated for each state. This allowed us to account for each state having

Let's include a fixed effect: average age of mother.

```
mod2 <- lmer(formula = BirthRate ~ AverageAgeofMother + (1 | State),
             data = countyBirthsData)
```

```
summary(mod2)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: BirthRate ~ AverageAgeofMother + (1 | State)
## Data: countyBirthsData
##
## REML criterion at convergence: 2347.6
##
## Scaled residuals:
##    Min       1Q   Median       3Q      Max
## -2.9602 -0.6086 -0.1042  0.5144  5.2686
##
## Random effects:
## Groups Name Variance Std.Dev.
## State (Intercept) 1.562 1.250
## Residual 2.920 1.709
## Number of obs: 578, groups: State, 50
##
```

```

## Fixed effects:
##               Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)    27.57033    1.81801 575.96198  15.165 < 2e-16 ***
## AverageAgeofMother -0.53549    0.06349 574.18338  -8.434 2.71e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr)
## AvrgAgfMthr -0.994

# Now we want a random-effects slope for each state. A random-effect slope may be estimated for each gr

# Adding total population of each state as a random effect (it's numeric, hence slope not intercept)

countyBirthsData <- countyBirthsData %>%
  mutate(LogTotalPop = log10(TotalPopulation))

mod3 <- lmer(formula = BirthRate ~ AverageAgeofMother + (LogTotalPop | State),
             data = countyBirthsData)

## Warning in checkConv(attr("opt", "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.0260043 (tol = 0.002, component 1)

# Uncorrelated random effects model (use the || syntax):
mod4 <- lmer(formula = BirthRate ~ AverageAgeofMother + (LogTotalPop || State),
             data = countyBirthsData)

## boundary (singular) fit: see ?isSingular

summary(mod4) # Not good. Need to the more complex model (mod3)

## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: BirthRate ~ AverageAgeofMother + (LogTotalPop || State)
## Data: countyBirthsData
##
## REML criterion at convergence: 2347.6
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.9602 -0.6086 -0.1042  0.5144  5.2686
##
## Random effects:
## Groups   Name                Variance Std.Dev.
## State    (Intercept)  1.562      1.250
## State.1   LogTotalPop  0.000      0.000
## Residual                    2.920      1.709
## Number of obs: 578, groups: State, 50
##
## Fixed effects:
##               Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)    27.57033    1.81801 575.96198  15.165 < 2e-16 ***
## AverageAgeofMother -0.53549    0.06349 574.18338  -8.434 2.71e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

##
## Correlation of Fixed Effects:
##      (Intr)
## AvrgAgfMthr -0.994
## convergence code: 0
## boundary (singular) fit: see ?isSingular

# A predictor can be both a fixed-effect and a random-effect (e.g., mothers age)
mod5 <- lmer(formula = BirthRate ~ AverageAgeofMother + (AverageAgeofMother | State),
             data = countyBirthsData)

## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.0133555 (tol = 0.002, component 1)

summary(mod5) # This is helpful for prediction, as you can correct the effect of age of mother by state

## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: BirthRate ~ AverageAgeofMother + (AverageAgeofMother | State)
## Data: countyBirthsData
##
## REML criterion at convergence: 2337.5
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.8402 -0.5965 -0.1132  0.5233  5.1817
##
## Random effects:
## Groups Name Variance Std.Dev. Corr
## State (Intercept) 78.33144 8.8505
##      AverageAgeofMother 0.08433 0.2904 -0.99
## Residual 2.80345 1.6744
## Number of obs: 578, groups: State, 50
##
## Fixed effects:
##              Estimate Std. Error    df t value Pr(>|t|)
## (Intercept) 27.21961    2.41010 39.91438 11.294 5.31e-14 ***
## AverageAgeofMother -0.52344    0.08293 39.42721 -6.312 1.83e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##      (Intr)
## AvrgAgfMthr -0.997
## convergence code: 0
## Model failed to converge with max|grad| = 0.0133555 (tol = 0.002, component 1)

REML - restricted maximum likelihood method

# Extracting coefficients

# Fixed effects:
fixef(mod5)

##      (Intercept) AverageAgeofMother
##      27.219605      -0.523442

```

```
# Random effects:
ranef(mod5)
```

```
## $State
##      (Intercept) AverageAgeofMother
## AK    4.15003868      -0.109184783
## AL   -4.03378321       0.127790698
## AR    1.21082566      -0.026165323
## AZ   -5.11009217       0.148282653
## CA   11.43254356      -0.373285277
## CO    1.35768165      -0.043219507
## CT   -3.82970951       0.066100814
## DC    1.21115127      -0.012314396
## DE   -2.11510834       0.060412134
## FL  -11.31590854       0.334159606
## GA    6.93457372      -0.230274526
## HI   -0.51788167       0.020356335
## IA    3.41174428      -0.085895750
## ID    7.51123552      -0.222596318
## IL   -3.07183968       0.102977648
## IN   -3.02904144       0.097145469
## KS    5.15187643      -0.150220284
## KY    2.30557751      -0.058607045
## LA    5.32029471      -0.154658352
## MA   -5.72651427       0.144383430
## MD   -3.31971034       0.116256669
## ME   -5.76419680       0.145079276
## MI   -5.51723036       0.156052226
## MN   -1.41379992       0.077580405
## MO    0.46288340      -0.013492801
## MS   -0.77739331       0.021041676
## MT    0.07683320      -0.006560404
## NC    4.20869679      -0.159191892
## ND    3.20725583      -0.078130864
## NE    4.97980324      -0.122101047
## NH   -3.02333587       0.055641597
## NJ    6.33991764      -0.220034709
## NM    1.24087579      -0.066463156
## NV    0.35709749      -0.009233999
## NY   -5.34557647       0.162856972
## OH   -7.43833421       0.227092266
## OK    4.73462323      -0.140260716
## OR   -5.64958450       0.170021845
## PA   -8.17679196       0.230593543
## RI   -0.82967532      -0.014536489
## SC   -5.60511798       0.179511398
## SD    4.96670949      -0.128324308
## TN   -1.21831791       0.038711838
## TX   10.90259845      -0.331687671
## UT   14.12700916      -0.369723836
## VA   -8.76482204       0.350357075
## VT   -0.02787975      -0.014042974
## WA    0.30152985      -0.002489047
## WI   -0.35751943      -0.004946489
```

```
## WV -3.92421158      0.115236390
##
## with conditional variances for "State"
# Confidence intervals for fe:
confint(mod5)

## Computing profile confidence intervals ...

## Warning in nextpar(mat, cc, i, delta, lowcut, upcut): unexpected decrease in
## profile: using minstep

## Warning in FUN(X[[i]], ...): non-monotonic profile for .sig02

## Warning in confint.thpr(pp, level = level, zeta = zeta): bad spline fit
## for .sig02: falling back to linear interpolation

##              2.5 %      97.5 %
## .sig01         4.4104030 13.8610928
## .sig02        -0.9997023 -0.9607253
## .sig03         0.1333576  0.4630192
## .sigma         1.5761207  1.7821071
## (Intercept)    22.4263992 32.1350540
## AverageAgeofMother -0.6923128 -0.3576348

tidy(mod5, conf.int = TRUE)
```

```
## # A tibble: 6 x 10
##   effect group term estimate std.error statistic    df p.value conf.low
##   <chr>   <chr> <chr>   <dbl>    <dbl>    <dbl> <dbl>   <dbl>    <dbl>
## 1 fixed <NA> (Int~  27.2      2.41      11.3   39.9 5.31e-14    22.5
## 2 fixed <NA> Aver~ -0.523    0.0829     -6.31  39.4 1.83e- 7   -0.686
## 3 ran_p~ State sd__  8.85      NA         NA     NA  NA      NA
## 4 ran_p~ State sd__  0.290      NA         NA     NA  NA      NA
## 5 ran_p~ State cor__ -0.992      NA         NA     NA  NA      NA
## 6 ran_p~ Resi~ sd__  1.67      NA         NA     NA  NA      NA
## # ... with 1 more variable: conf.high <dbl>
```

Using Maryland Crime data

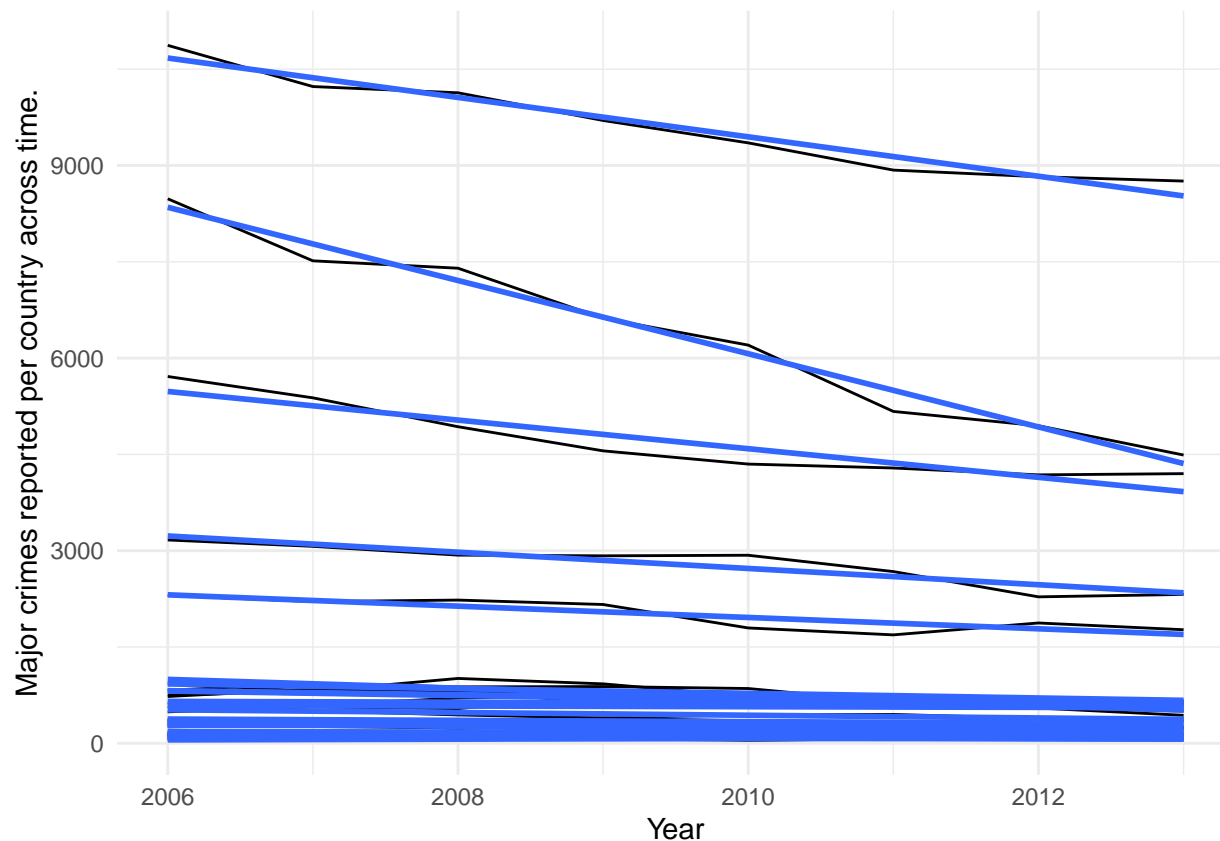
```
MDcrime <- read_csv("https://assets.datacamp.com/production/repositories/1803/datasets/e5e076efd3c3b7661")
```

```
## Warning: Missing column names filled in: 'X1' [1]
```

```
## Parsed with column specification:
```

```
## cols(
##   X1 = col_double(),
##   County = col_character(),
##   Year = col_double(),
##   Crime = col_double(),
##   Year2 = col_double()
## )
```

```
ggplot(data = MDcrime, aes(x = Year, y = Crime, group = County)) +
  geom_line() +
  geom_smooth(method = "lm", se = FALSE) +
  theme_minimal() +
  ylab("Major crimes reported per country across time.")
```



Looks like we will require a random-effect intercept, and likely a random-effect slope.

Try a linear model:

```
mod1 <- lm(formula = Crime ~ Year, data = MDcrime)
summary(mod1) # Year(2) not significant. Note: need to use Year2 variable.
```

```
##
## Call:
## lm(formula = Crime ~ Year, data = MDcrime)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1514.3 -1156.6  -930.9  -511.5   9293.7
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 136642.97  147510.24   0.926   0.355
## Year         -67.33     73.41  -0.917   0.360
##
## Residual standard error: 2331 on 190 degrees of freedom
## Multiple R-squared:  0.004408,    Adjusted R-squared:  -0.0008315
## F-statistic: 0.8413 on 1 and 190 DF,  p-value: 0.3602
```

Fit the model with Year(2) as both a fixed and random-effect:

```
mod2 <- lmer(formula = Crime ~ Year2 + ( 1 + Year2 | County), data = MDcrime)
summary(mod2)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: Crime ~ Year2 + (1 + Year2 | County)
## Data: MDcrime
##
## REML criterion at convergence: 2535.7
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.8080 -0.2235 -0.0390  0.2837  3.0767
##
## Random effects:
## Groups Name Variance Std.Dev. Corr
## County (Intercept) 7584514 2754.00
##      Year2      16940  130.15 -0.91
## Residual      8425   91.79
## Number of obs: 192, groups: County, 24
##
## Fixed effects:
##              Estimate Std. Error    df t value Pr(>|t|)
## (Intercept)  1577.28    562.29   23.02  2.805  0.0100 *
## Year2        -67.33    26.72   23.01 -2.519  0.0191 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##      (Intr)
## Year2 -0.906

# ANOVA - Analysis of Variance
# Build a null model with only County as a random-effect
null_model <- lmer(Crime ~ (1 | County), data = MDcrime)
# Build alternative model, with Year2 as a fixed and random slope and County as a random effect
alt_model <- lmer(Crime ~ Year2 + (1 + Year2 | County), data = MDcrime)

# Compare models - look at Chi square test.
anova(null_model, alt_model)

## refitting model(s) with ML (instead of REML)

## Data: MDcrime
## Models:
## null_model: Crime ~ (1 | County)
## alt_model: Crime ~ Year2 + (1 + Year2 | County)
##              Df      AIC      BIC logLik deviance Chisq Chi Df Pr(>Chisq)
## null_model   3 2954.4 2964.2 -1474.2  2948.4
## alt_model    6 2568.9 2588.4 -1278.4  2556.9 391.52      3 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Chapter 3: Generalized Linear Mixed-Effects Models

Topics covered in this chapter (for future reference):

- Logistic regression with `glm()` - family = "binomial"
- Poisson regression with `glm()` - family = "poisson" -> for count data.
- Plotting logistic regression results with `ggplot` - `stat_smooth(method = "glm", method.args = list(family = "binomial"))`
- Using `glmer()` to estimate glms with mixed effects.
- Handling different data inputs (e.g. matrix with `cbind`)
- Calculating odds-ratios with `exp(fixef())` and `exp(confint())`. Can also be done via `tidy()`
- Using `ggplot` to visualize random-effects poisson models

Chapter 4: Repeated Measures

Paired t-test -> special case of a t-test

`t.test(paired = TRUE)` -> does not assume equal variance of both groups.

Repeated measures ANOVA -> tests if means are constant across time. ^ simply a special type of mixed effects model.

```
set.seed(1234)

n_ind <- 10
before <- rnorm(n = n_ind, mean = 0, sd = 0.5)
after <- before + rnorm(n = n_ind, mean = 4.5, sd = 5)

t.test(before, after, paired = F)

##
## Welch Two Sample t-test
##
## data: before and after
## t = -2.3343, df = 9.1605, p-value = 0.04396
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -7.6873743 -0.1309186
## sample estimates:
## mean of x mean of y
## -0.1915787 3.7175678

t.test(before, after, paired = T) # Paired is more powerful.

##
## Paired t-test
##
## data: before and after
## t = -2.3164, df = 9, p-value = 0.04576
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -7.72678305 -0.09150987
## sample estimates:
## mean of the differences
## -3.909146

dat <- data.frame(y = c(before, after),
                  trial = rep(c("before", "after"), each = n_ind),
                  ind = rep(letters[1:n_ind], times = 2))
```

```

# Now do it with lmer:
mod <- lmer(y ~ trial + (1|ind), data = dat)
summary(mod) # Woot. Pretty much the same (same t value, p value)

## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: y ~ trial + (1 | ind)
## Data: dat
##
## REML criterion at convergence: 103.2
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.0871 -0.5599 -0.0249  0.1110  3.3158
##
## Random effects:
## Groups Name Variance Std.Dev.
## ind (Intercept) 9.047e-07 0.0009512
## Residual 1.402e+01 3.7446323
## Number of obs: 20, groups: ind, 10
##
## Fixed effects:
## Estimate Std. Error df t value Pr(>|t|)
## (Intercept) 3.718 1.184 18.000 3.139 0.00567 **
## trialbefore -3.909 1.675 18.000 -2.334 0.03137 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
## (Intr)
## trialbefore -0.707

htcrime <- read_csv("https://assets.datacamp.com/production/repositories/1803/datasets/45e88fe1bc8d1d76

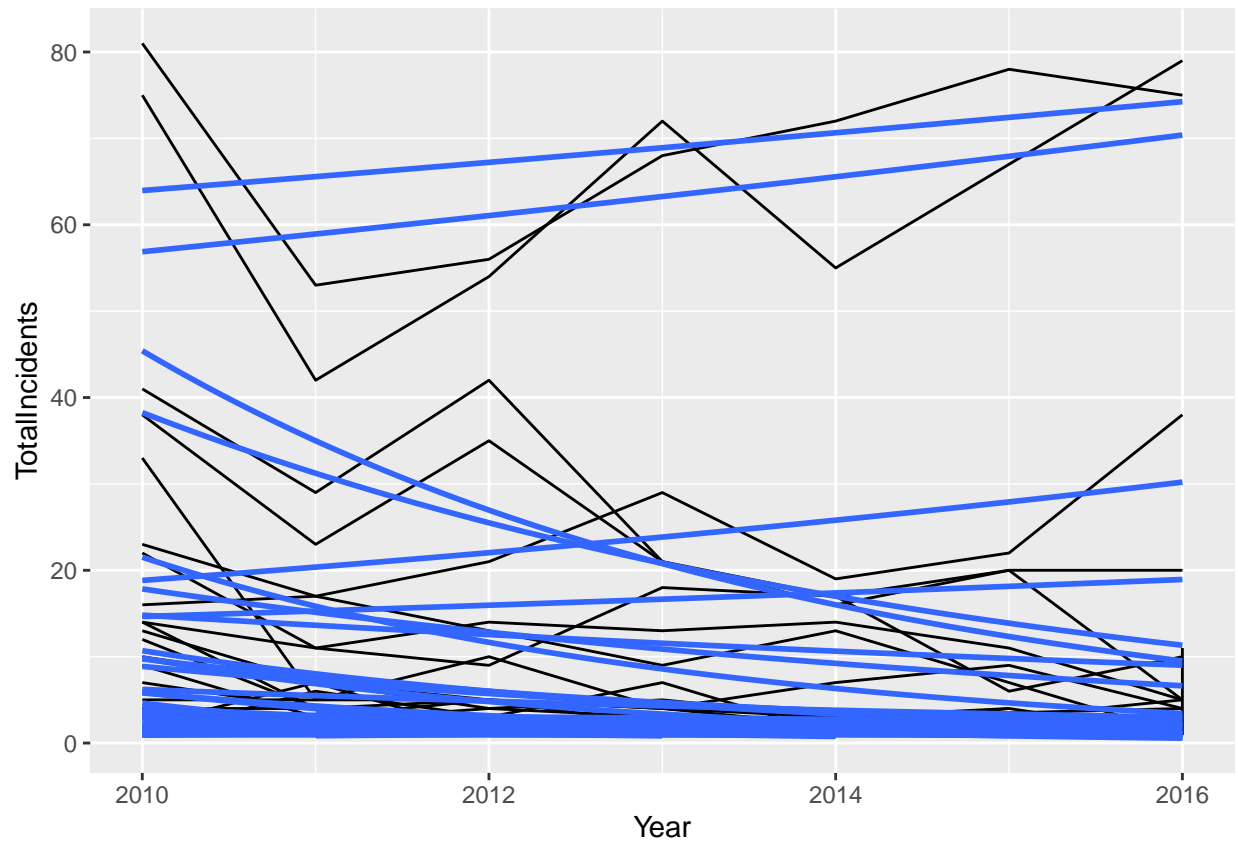
## Parsed with column specification:
## cols(
##   Year = col_double(),
##   County = col_character(),
##   TotalIncidents = col_double(),
##   Year2 = col_double()
## )

# Is the number of hate crimes changing over time in NY counties?
# 1) Is the state-wide number of hate crimes changing?
# 2) Are the number of hate crimes changing differently in each county?

# Step 1. Visualize

ggplot(data = htcrime, mapping = aes(x = Year, y = TotalIncidents, group = County)) +
  geom_line() +
  geom_smooth(method = "glm", method.args = c("poisson"), se = FALSE)

```



We can see that different countries have different trends. Therefore it would be appropriate to use di.