

Question 1.

Homework 3

Stat 764
Liangjian Tao
1253177

Group 1	Group 2	Group 3	Group 4	Group 5	Group 6
1 $y_{11} y_{12} y_{13} y_{14}$					
2	□ □ □	□ □ □	□ □ □	□ □ □	□ □ □
⋮	⋮	⋮	⋮	⋮	⋮
20 $y_{21} y_{22} y_{23} y_{24}$	□ □ □	□ □ □	□ □ □	□ □ □	□ □ □

$n=20$

$p=4$

$k=6$

37 | 40'

a. Assumptions:

- The observation vectors from each of the 6 populations come from a multivariate normal distributions.
 - Each of 6 populations has the same population covariance matrix (Σ). Assumptions not stated in context (-0.5)
 - 6 independent samples from 6 different populations.
 - No multicollinearity, 4 dependent variables moderately correlated with each other.
- ~~This is not an assumption~~ (-0.5)

b. Factor Effects Model

$$\underline{y}_{ij} = \underline{\mu} + \underline{\alpha}_i + \underline{\epsilon}_{ij}$$

~~$\underline{\mu} + \underline{\epsilon}_{ij}$~~

$$i=1, 2, 3, 4, 5, 6 \quad (k=6)$$

$$j=1, 2, \dots, 20 \quad (n=20)$$

$\underline{\mu}$: overall population mean vector

$\underline{\alpha}_i$: effect vector of i th population

$\underline{\mu}_i = \underline{\mu} + \underline{\alpha}_i$: population mean vector of i th population.

$\underline{\epsilon}_{ij} \sim N_p(\mathbf{0}, \Sigma)$: error vector for j th observation in i th sample.

Define \underline{y}_{ij} (-0.5)

Definitions not in context (-0.5)

C. Assumptions of Multivariate Normal Distributions.

By Q-Q plot for y_1, y_2, y_3, y_4 in each group of 6 groups,

All points are close to a straight line, \rightarrow There are clear outliers in some of the

There's no indication of departure from normality, Q-Q plots

Thus, all variables of y_1, y_2, y_3, y_4 in each group follow normality separately.

By Scatterplots of $y_1 \text{ vs } y_2, y_1 \text{ vs. } y_3, y_1 \text{ vs. } y_4$
 $y_2 \text{ vs. } y_3, y_2 \text{ vs. } y_4$
 $y_3 \text{ vs. } y_4$

There's no indication of a curved trend or other obvious nonnormality.

Although there are few outliers in group 1, group 4, But generally it is still can be inferred that each pair of variables has a bivariate normal distribution.

There's no indication of a clear deviation from MVN distribution. It's appropriate to assume that these observation vectors come from a multivariate Normal distribution.

d.

$$H_0: \underline{\mu}_1 = \underline{\mu}_2 = \underline{\mu}_3 = \underline{\mu}_4 = \underline{\mu}_5 = \underline{\mu}_6$$

$H_a:$ at least two population mean vectors differ. \checkmark

e.

See as attachments \checkmark

f.

Non-zero eigenvalues of $E^{-1}H$:

$$\lambda_1 = 1.0576$$

$$\lambda_2 = 0.3301$$

$$\lambda_3 = 0.1116$$

$$\lambda_4 = 0.0037$$

\checkmark

g.

$$\text{Wicks' Test: } \Lambda = \frac{|E|}{|E+H|} = \prod_{i=1}^5 \frac{1}{1+\lambda_i}$$

$$\begin{aligned} &= \frac{1}{1+1.0516} \cdot \frac{1}{1+0.3301} \cdot \frac{1}{1+0.1116} \cdot \frac{1}{1+0.0037} \\ &= \frac{1}{2.0516} \cdot \frac{1}{1.3301} \cdot \frac{1}{1.1116} \cdot \frac{1}{1.0037} \\ &= \frac{1}{3.0535} \doteq 0.3275 \quad \checkmark \end{aligned}$$

$$\Lambda_{\alpha, p, V_H, V_E} = \Lambda_{0.05, 4, 5, 114} \doteq 0.730 \quad \checkmark$$

Reject H_0 if $\Lambda < \Lambda_{0.05, 4, 5, 114}$

$$\Lambda = 0.3275 < 0.730$$

Thus, Reject H_0 .

At significance level of $\alpha=0.05$, there's sufficient evidence to claim H_a , which is that at least two population mean vectors differ.

State in context of problem

h. Roy's Largest Root:

$$\Theta = \frac{\lambda_1}{1+\lambda_1} = \frac{1.0576}{1+1.0516} \doteq 0.5140 \quad \checkmark$$

$$\Theta_{\alpha, s, m, N} = \Theta_{0.05, 4, 0, 54.5} \doteq 0.171 \quad \checkmark$$

Reject H_0 if $\Theta \geq \Theta_{0.05, 4, 0, 54.5}$

$$\Theta = 0.5140 > 0.171 \quad \checkmark$$

Thus, Reject H_0 . \checkmark

At significance level of $\alpha=0.05$, there's sufficient evidence to claim H_a , which is that at least two population mean vectors differ.

$$P=4$$

$$k=6$$

$$n=20$$

$$V_H = k-1 = 5$$

$$V_E = k(n-1) = 6 \times 19 = 114$$

$$S = \text{rank}(H)$$

$$= \min(p, V_H)$$

$$= \min(4, 5)$$

$$= 4$$

$$S = \min(p, V_H) = \min(4, 5) = 4$$

$$m = \frac{1}{2}(|V_H - P| - 1) = \frac{1}{2}(|5-4| - 1) = 0$$

$$N = \frac{1}{2}(V_E - P - 1) = \frac{1}{2}(114 - 4 - 1) = 54.5$$

i. Pillai's test

$$V^{(S)} = \sum_{i=1}^S \frac{\lambda_i}{1+\lambda_i} = \frac{1.0546}{1+1.0546} + \frac{0.3301}{1+0.3301} + \frac{0.1116}{1+0.1116} + \frac{0.0031}{1+0.0031} \quad (S=4)$$

$$= 0.8662 \quad \checkmark$$

$$V_{\alpha, m, N}^{(S)} = V_{0.05, 0, 54.5}^{(4)} \doteq 0.602 \quad \checkmark$$

Reject H_0 if $V^{(S)} > V_{\alpha, m, N}^{(S)}$

$$V^{(4)} = 0.8662 > 0.602$$

Thus, Reject H_0 . \checkmark

At significance level of $\alpha=0.05$, there's sufficient evidence to claim H_a , which is that at least two population mean vectors differ.

j. Lawley - Hotelling test

$$U^{(S)} = \sum_{i=1}^S \lambda_i$$

$$= 1.0546 + 0.3301 + 0.1116 + 0.0031$$

$$= 1.503 \quad \checkmark$$

$$\frac{V_E}{V_H} U^{(S)} = \frac{114}{5} \times 1.503 = 34.26 \quad \checkmark$$

$$\frac{V_E}{V_H} U_{\alpha, p}^{(S)} = \frac{114}{5} \times U_{0.05, 4}^{(4)} \doteq 3.8557 \quad \textcircled{0.05}$$

Reject H_0 if $\frac{V_E}{V_H} U^{(S)} \geq \frac{V_E}{V_H} U_{\alpha, p}^{(S)}$.

$$\frac{V_E}{V_H} U^{(S)} = 34.268 > 3.8557$$

Thus, Reject H_0 . \checkmark

At significance level of $\alpha=0.05$, there's sufficient evidence to claim H_a , which is that at least two population mean vectors differ.

k.

Based on g. h. i. j., H_0 is not true.

Thus, the power of these tests differ in different situations.

The mean vectors are collinear. Because $E^{-1}H$ has one large eigenvalue and the rest of the eigenvalues are small. I recommend λ , being at least 80% of sum of eigenvalues to indicate multicollinearity.

The power of the tests are ordered as:

$$\Theta \geq U^{(s)} \geq \Lambda \geq V^{(s)}$$

Roy's Largest Root \geq Lawley-Hotelling test \geq Wilks' test \geq Pillai's test.

There are few outliers in group 1 and group 4, any of the other three tests other than Roy's test will still perform reasonably well.

Λ remains the most popular test statistic for MANOVA because of its flexibility and historical dominance.

Based on power of the tests in this case,

I recommend Lawley-Hotelling test first, then Wilks' test second.

l.

Variables	Univariate F-Statistics	P-value	coefficient in α_1
width	4.36	0.0012	-0.1299
thickness	8.09	< 0.0001	-0.1268
Lwratio	5.86	< 0.0001	0.0960
Volumn	19.28	< 0.0001	0.0163

significant	Volumn
↓	Thickness
Less significant	Lwratio
	Width

by results of the univariate one-way ANOVAs.

Significant	width
↓	thickness
Less significant	Lwratio
	Volumn

by coefficients of the discriminant function.

Using the (unstandardized) discriminant function is only appropriate if the variances are similar. (-0.5)

m. Ranking variables based on the coefficients of the discriminant function takes the

correlations of the variables into account, it generally provides more pertinent information than ranking variables by a univariate ANOVA if the sample variances are similar among the different variables. I recommend coefficients of the discriminant function.

QUESTION 2 → Graded on completion

	Group 1	Group 2	Group 3	Group 4	Group 5	Group 6
	3H	5H	3H	5H	3H	5H
y_1	-0.120	0.156	0.212	0.041	0.189	0.031
y_2	-0.102	0.340	0.483	0.316	0.602	0.464
y_3	0.033	0.064	0.138	0.182	0.250	0.339
y_4	-3.335	10.040	16.451	10.488	21.511	14.745

a. $H_0: \underline{M}_1 + \underline{M}_3 + \underline{M}_5 - \underline{M}_2 - \underline{M}_4 - \underline{M}_6 = 0$ $\sum c_i = 1+1+1-1-1-1=0$

$H_a: \underline{M}_1 + \underline{M}_3 + \underline{M}_5 - \underline{M}_2 - \underline{M}_4 - \underline{M}_6 \neq 0$

b. $\hat{\delta} = \sum_{i=1}^k c_i \bar{y}_i$.

$$= \begin{pmatrix} -0.120 \\ -0.102 \\ 0.033 \\ -3.335 \end{pmatrix} + \begin{pmatrix} 0.212 \\ 0.483 \\ 0.138 \\ 16.451 \end{pmatrix} + \begin{pmatrix} 0.189 \\ 0.602 \\ 0.250 \\ 21.511 \end{pmatrix} - \begin{pmatrix} 0.156 \\ 0.340 \\ 0.064 \\ 10.040 \end{pmatrix} - \begin{pmatrix} 0.041 \\ 0.316 \\ 0.182 \\ 10.488 \end{pmatrix} - \begin{pmatrix} 0.031 \\ 0.464 \\ 0.339 \\ 14.745 \end{pmatrix}$$

$$= \begin{pmatrix} -0.120 + 0.212 + 0.189 - 0.156 - 0.041 - 0.031 \\ -0.102 + 0.483 + 0.602 - 0.340 - 0.316 - 0.464 \\ 0.033 + 0.138 + 0.250 - 0.064 - 0.182 - 0.339 \\ -3.335 + 16.451 + 21.511 - 10.040 - 10.488 - 14.745 \end{pmatrix} = \begin{pmatrix} 0.053 \\ -0.197 \\ -0.164 \\ -0.646 \end{pmatrix}$$

c. $\text{Cov}(\hat{\delta}) = \sum \frac{1}{n} \cdot \frac{k}{k-1} c_i^2$

$$= \frac{E}{V_E} \cdot \frac{1}{n} \cdot \frac{k}{k-1} c_i^2$$

$$= E \cdot \frac{1}{V_E} \cdot \frac{1}{n} \cdot \frac{k}{k-1} c_i^2$$

$$= \left[\begin{array}{cccc} 8.290 & 1.170 & -4.029 & 154.497 \\ 1.170 & 16.813 & -0.446 & 284.411 \\ -4.029 & -0.446 & 5.124 & -11.415 \\ 154.497 & 284.411 & -11.415 & 8475.735 \end{array} \right] \cdot \frac{1}{14} \cdot \frac{1}{20} \cdot 6$$

$$= \left[\begin{array}{cccc} 8.290 & 1.170 & -4.029 & 154.497 \\ 1.170 & 16.813 & -0.446 & 284.411 \\ -4.029 & -0.446 & 5.124 & -11.415 \\ 154.497 & 284.411 & -11.415 & 8475.735 \end{array} \right] \times 0.00263$$

d. See Attach File

contrast "contrast"

group 1 -1 1 -1 1 -1;

e. Non-zero eigenvalue of $E^{-1}H_1$,

$$\lambda = 0.2582$$

f. Wilkes' test:

$$\Lambda = \frac{|E|}{|E+H_1|} = \prod_{i=1}^S \frac{1}{1+\lambda_i} = \frac{1}{1+0.2582} = 0.795$$

$$\Lambda_{\alpha, p, 1, V_2} = \Lambda_{0.05, 4, 1, 114} \doteq 0.962$$

Reject H_0 if $\Lambda < \Lambda_{\alpha, p, 1, V_2}$

$$\Lambda = 0.795 < 0.962,$$

Thus, Reject H_0 .

At significance level of $\alpha = 0.05$, there's sufficient evidence to claim H_a , which is that the average of the population mean vectors for groups with a charring time of 3 hours differs from the average of the population mean vectors for groups with a charring time of 5 hours.

The SAS System

The GLM Procedure

Class Level Information		
Class	Levels	Values
group	6	1 2 3 4 5 6

Number of Observations Read	120
Number of Observations Used	120

The GLM Procedure

Dependent Variable: width

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	1.58390667	0.31678133	4.36	0.0012
Error	114	8.29029000	0.07272184		
Corrected Total	119	9.87419667			

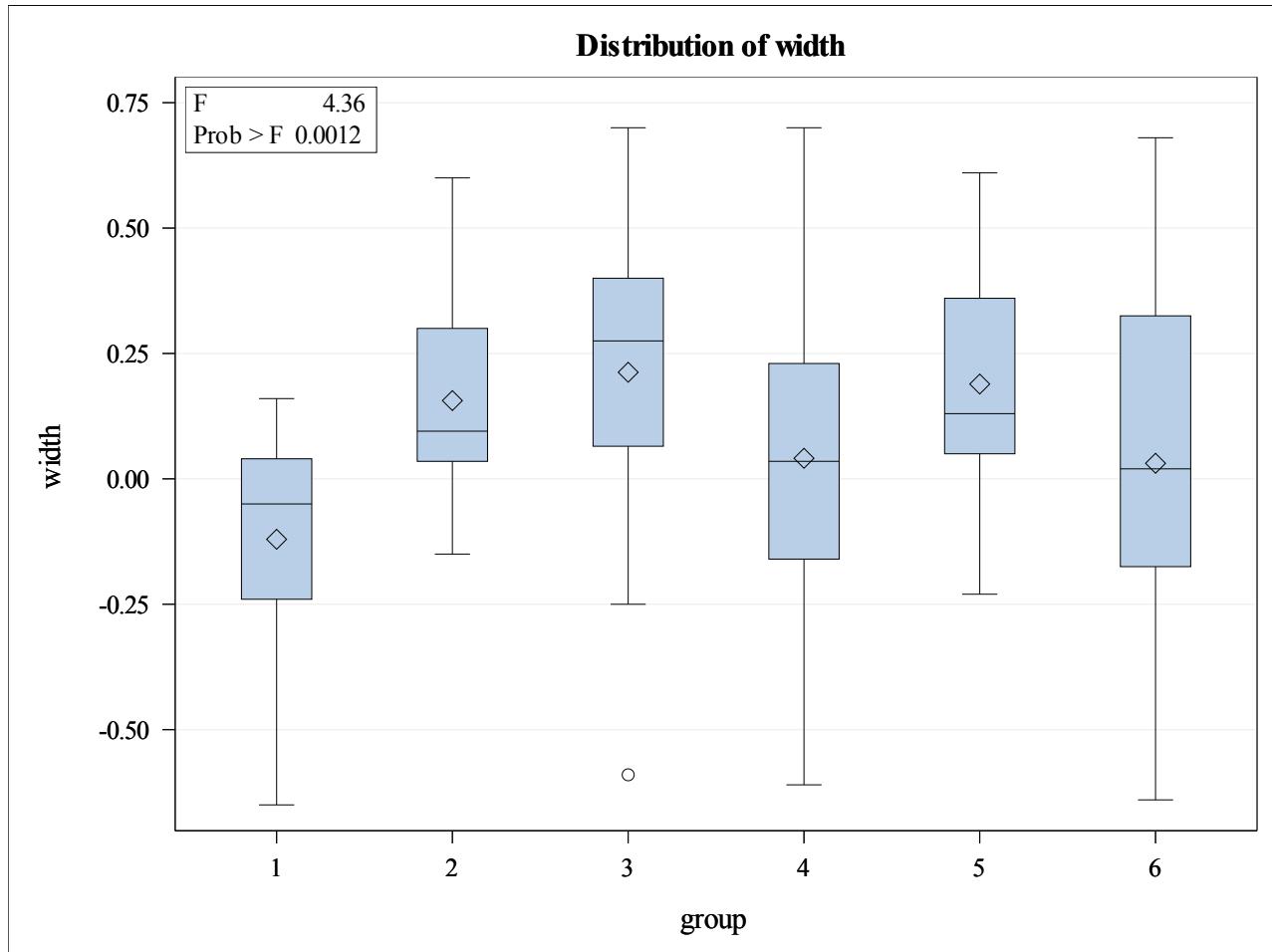
R-Square	Coeff Var	Root MSE	width Mean
0.160409	317.8820	0.269670	0.084833

Source	DF	Type I SS	Mean Square	F Value	Pr > F
group	5	1.58390667	0.31678133	4.36	0.0012

Source	DF	Type III SS	Mean Square	F Value	Pr > F
group	5	1.58390667	0.31678133	4.36	0.0012

Contrast	DF	Contrast SS	Mean Square	F Value	Pr > F
contrast	1	0.00936333	0.00936333	0.13	0.7204

The SAS System



The GLM Procedure

Dependent Variable: thickness

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	5.98379417	1.19675883	8.09	<.0001
Error	114	16.87306500	0.14800934		
Corrected Total	119	22.85685917			

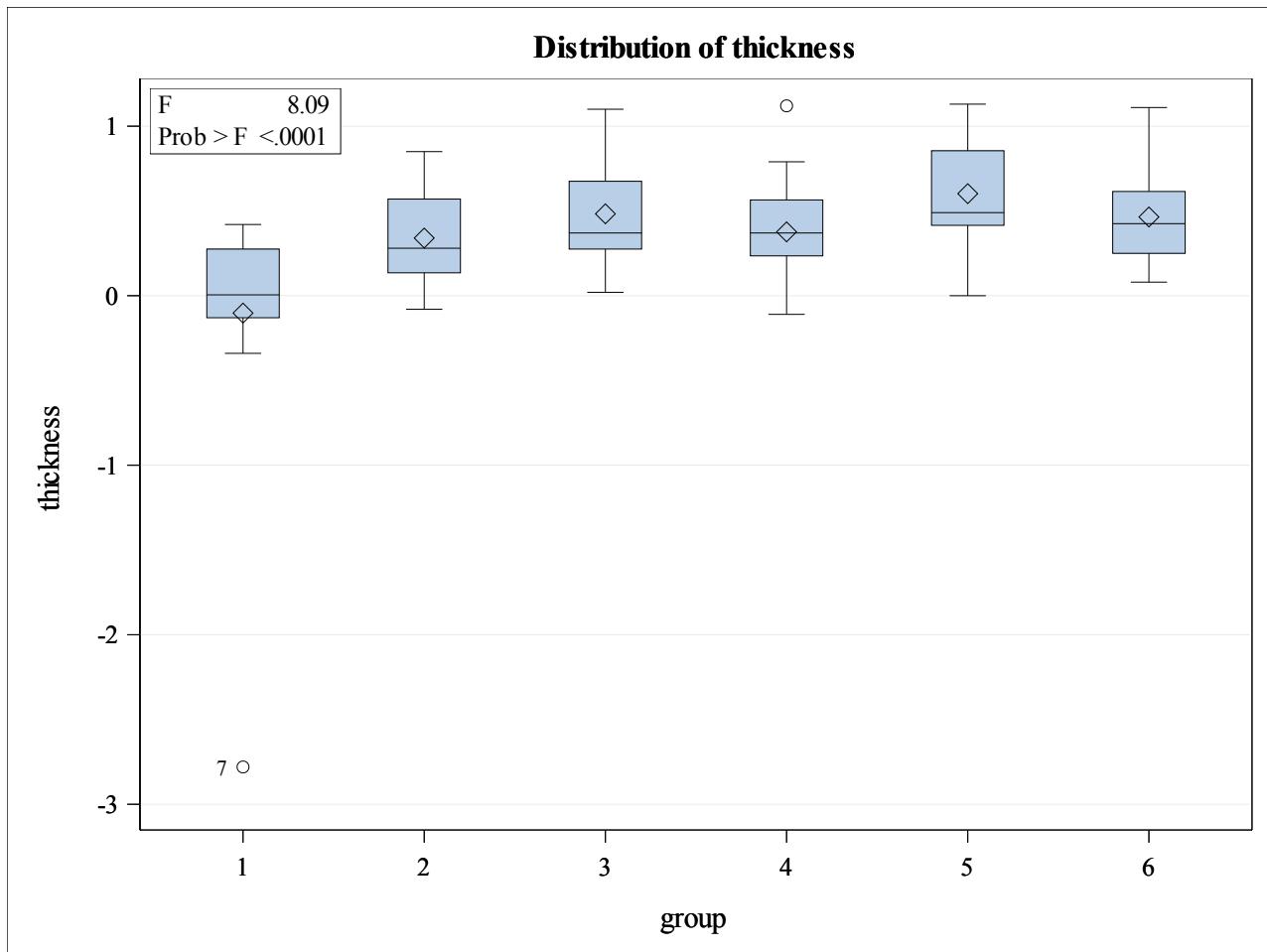
R-Square	Coeff Var	Root MSE	thickness Mean
0.261794	106.6937	0.384720	0.360583

Source	DF	Type I SS	Mean Square	F Value	Pr > F
group	5	5.98379417	1.19675883	8.09	<.0001

The SAS System

Source	DF	Type III SS	Mean Square	F Value	Pr > F
group	5	5.98379417	1.19675883	8.09	<.0001

Contrast	DF	Contrast SS	Mean Square	F Value	Pr > F
contrast	1	0.13134083	0.13134083	0.89	0.3482



The GLM Procedure

Dependent Variable: lwratio

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	1.31706575	0.26341315	5.86	<.0001
Error	114	5.12399023	0.04494728		
Corrected Total	119	6.44105598			

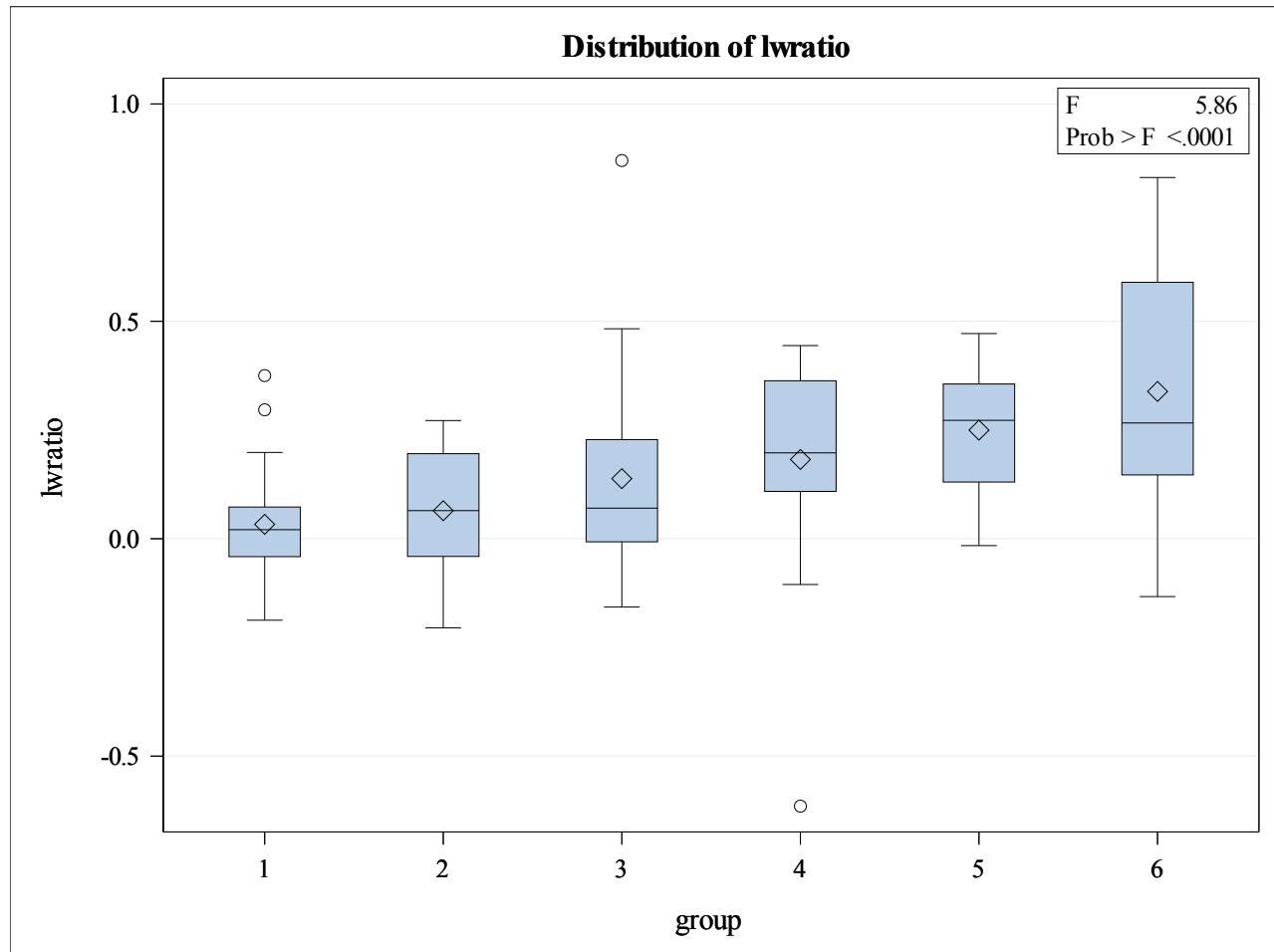
The SAS System

R-Square	Coeff Var	Root MSE	Iwratio Mean
0.204480	126.4071	0.212008	0.167718

Source	DF	Type I SS	Mean Square	F Value	Pr > F
group	5	1.31706575	0.26341315	5.86	<.0001

Source	DF	Type III SS	Mean Square	F Value	Pr > F
group	5	1.31706575	0.26341315	5.86	<.0001

Contrast	DF	Contrast SS	Mean Square	F Value	Pr > F
contrast	1	0.09009865	0.09009865	2.00	0.1596



The GLM Procedure

Dependent Variable: volume

The SAS System

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	7167.15091	1433.43018	19.28	<.0001
Error	114	8475.73547	74.34856		
Corrected Total	119	15642.88637			

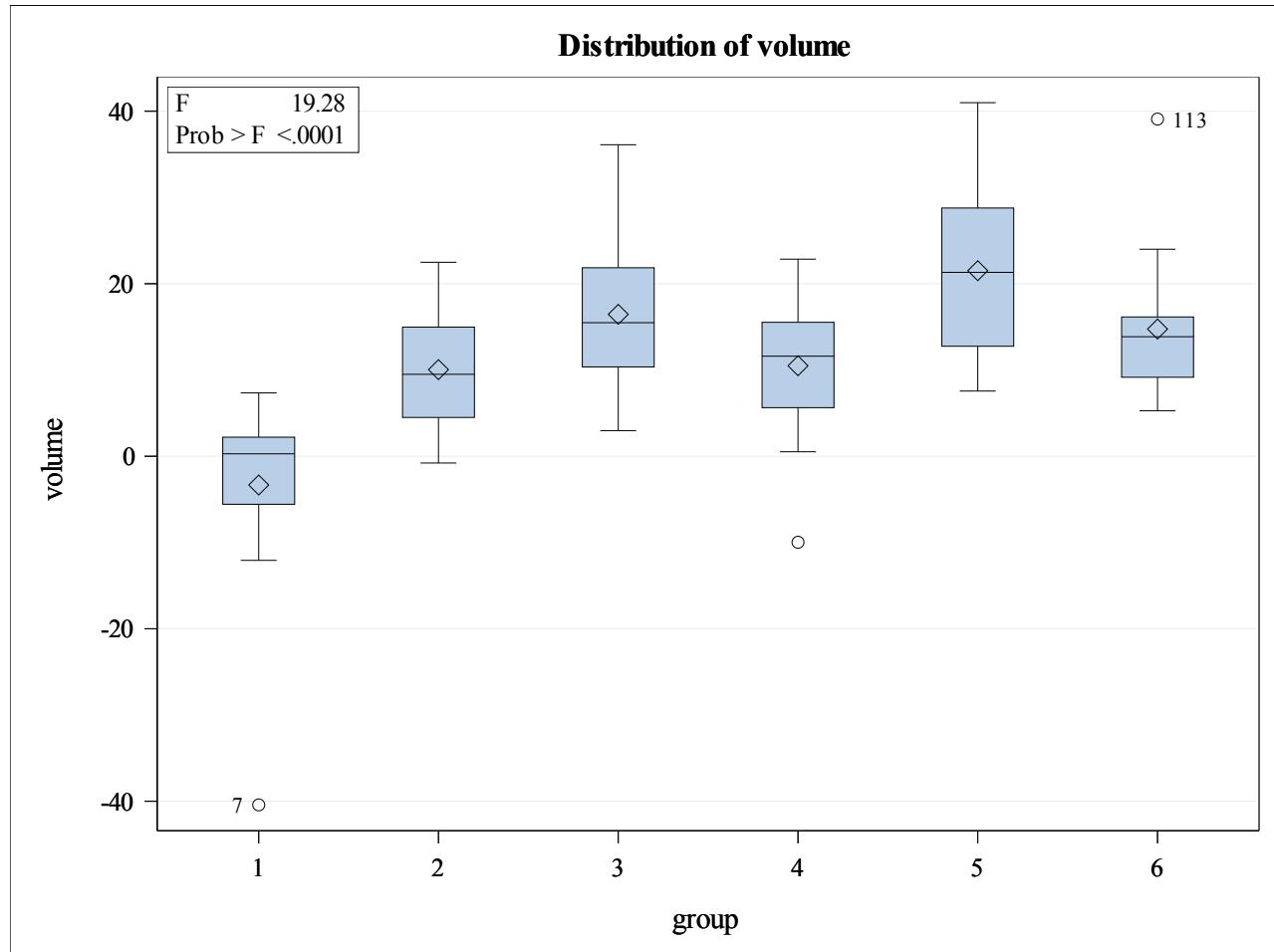
R-Square	Coeff Var	Root MSE	volume Mean
0.458173	74.01224	8.622561	11.65018

Source	DF	Type I SS	Mean Square	F Value	Pr > F
group	5	7167.150905	1433.430181	19.28	<.0001

Source	DF	Type III SS	Mean Square	F Value	Pr > F
group	5	7167.150905	1433.430181	19.28	<.0001

Contrast	DF	Contrast SS	Mean Square	F Value	Pr > F
contrast	1	1.38997817	1.38997817	0.02	0.8915

The SAS System



The GLM Procedure
Multivariate Analysis of Variance

E = Error SSCP Matrix

	width	thickness	lwratio	volume
width	8.29029	1.170045	-4.029202104	154.49735915
thickness	1.170045	16.873065	-0.44628408	284.4113197
lwratio	-4.029202104	-0.44628408	5.1239902254	-11.47549504
volume	154.49735915	284.4113197	-11.47549504	8475.7354672

The SAS System

Partial Correlation Coefficients from the Error SSCP Matrix / Prob > r 				
DF = 114	width	thickness	lwratio	volume
width	1.000000	0.098928 0.2929	-0.618201 <.0001	0.582838 <.0001
thickness	0.098928 0.2929	1.000000	-0.047997 0.6105	0.752075 <.0001
lwratio	-0.618201 <.0001	-0.047997 0.6105	1.000000	-0.055065 0.5589
volume	0.582838 <.0001	0.752075 <.0001	-0.055065 0.5589	1.000000

The GLM Procedure
Multivariate Analysis of Variance

Characteristic Roots and Vectors of: E Inverse * H, where H = Type III SSCP Matrix for group E = Error SSCP Matrix						
Characteristic Root	Percent	Characteristic Vector V'EV=1				
		width	thickness	lwratio	volume	
1.05758634	70.37	-0.12987286	-0.12681120	0.09598563	0.01633480	
0.33007068	21.96	-1.60599251	-0.96010780	-1.25300712	0.05899748	
0.11160307	7.43	0.12443219	0.02337911	-0.30090176	0.00121369	
0.00373024	0.25	0.32207304	-0.17711522	0.26723234	-0.00148523	

MANOVA Tests for the Hypothesis of No Overall group Effect H = Type III SSCP Matrix for group E = Error SSCP Matrix		
S=4 M=0 N=54.5		
Statistic	Value	P-Value
Wilks' Lambda	0.32749183	<.0001
Pillai's Trace	0.86626861	<.0001
Hotelling-Lawley Trace	1.50299033	<.0001
Roy's Greatest Root	1.05758634	<.0001

The SAS System

Characteristic Roots and Vectors of: E Inverse * H, where
H = Contrast SSCP Matrix for contrast
E = Error SSCP Matrix

Characteristic Root	Percent	Characteristic Vector V'EV=1			
		width	thickness	lwratio	volume
0.25815117	100.00	-1.54049148	-0.94110022	-1.27949481	0.05765378
0.00000000	0.00	0.45018911	0.06238517	0.08481721	-0.00382185
0.00000000	0.00	-0.32426570	0.07472546	-0.23663604	0.01066277
0.00000000	0.00	-0.18536681	-0.27323805	0.20223140	0.01729011

MANOVA Tests for the Hypothesis of No Overall contrast Effect
H = Contrast SSCP Matrix for contrast
E = Error SSCP Matrix

S=1 M=1 N=54.5

Statistic	Value	P-Value
Wilks' Lambda	0.79481705	<.0001
Pillai's Trace	0.20518295	<.0001
Hotelling-Lawley Trace	0.25815117	<.0001
Roy's Greatest Root	0.25815117	<.0001

```
ods rtf file='hw3 with contrast.rtf';

DATA rice;
  INFILE 'grain.csv' firstobs=2 dsd;
  INPUT group width thickness lwratio volume;
RUN;

ODS STARTPAGE = NO;

/*Add SAS code for Homework 3 below this line and above green line of stars below*/
PROC GLM;
  CLASS group;
  MODEL width thickness lwratio volume = group;
  CONTRAST 'contrast'
    group 1 -1 1 -1 1 -1;
  MANOVA H = group/PRINTE PRINTE MSTAT=EXACT;
RUN;

*****
```

```
ods rtf file='hw3 with contrast.rtf';

DATA rice;
  INFILE 'grain.csv' firstobs=2 dsd;
  INPUT group width thickness lwratio volume;
RUN;

ODS STARTPAGE = NO;

/*Add SAS code for Homework 3 below this line and above green line of stars below*/
PROC GLM;
  CLASS group;
  MODEL width thickness lwratio volume = group;
  CONTRAST 'contrast'
    group 1 -1 1 -1 1 -1;
  MANOVA H = group/PRINTE PRINTE MSTAT=EXACT;
RUN;

*****
```