

# Machine Learning Theory

## Lecture 11: Covering Numbers

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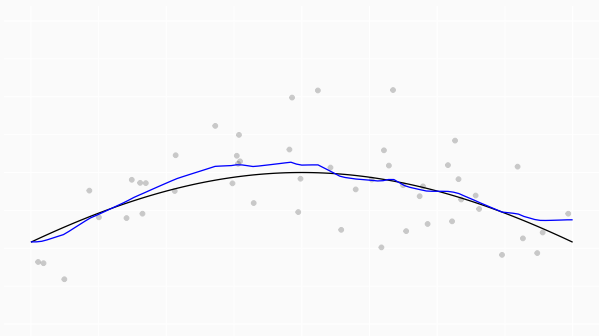
Emory University

## Review

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## Least squares with gaussian noise

We observe  $Y_i = \mu(X_i) + \epsilon_i$  for  $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$ .

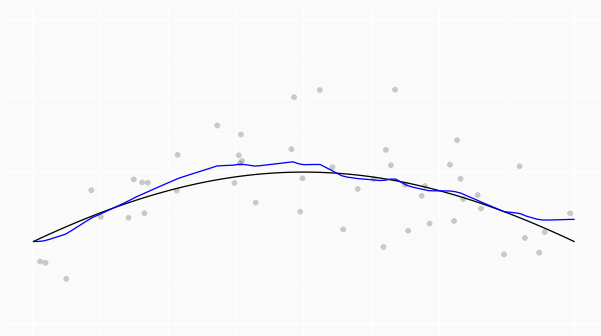


We've focused on [least squares estimators](#). That's the curve in your regression model that minimizes mean squared prediction error.

$$\hat{\mu} = \operatorname{argmin}_{m \in \mathcal{M}} \frac{1}{n} \sum_{i=1}^n \{Y_i - m(X_i)\}^2$$

## Least squares with gaussian noise

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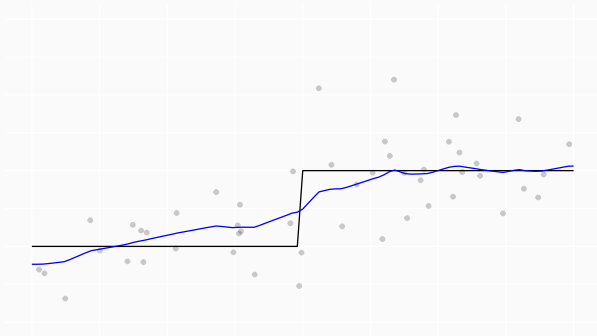
To think about how well this works, we've proven high probability bounds on the error.

$$\|\hat{\mu} - \mu\| < s \quad \text{with probability} \quad 1 - \delta \quad \text{where usually} \quad \|v\|^2 = \frac{1}{n} \sum_{i=1}^n v(X_i)^2$$

We've mostly talked about this error's *sample two norm*.

# Least squares with gaussian noise

We observe  $Y_i = \mu(X_i) + \epsilon_i$  for  $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$ .

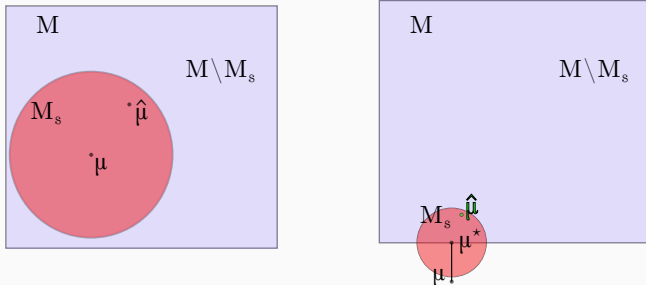


Or, more generally, on norms of the difference between our estimator and the model's best *approximation* to  $\mu$ .

$$\|\hat{\mu} - \mu^*\| < s \quad \text{with probability} \quad 1 - \delta \quad \text{where} \quad \mu^* = \underset{m \in \mathcal{M}}{\operatorname{argmin}} \|m - \mu\|$$

# What determines these bounds

It's the gaussian width of *neighborhoods* of this best approximation  $\mu^*$ .



In convex models, we can work with the width of neighborhood's *boundary*.  
And the bound *does not* depend on how good our approximation  $\mu^*$  is.

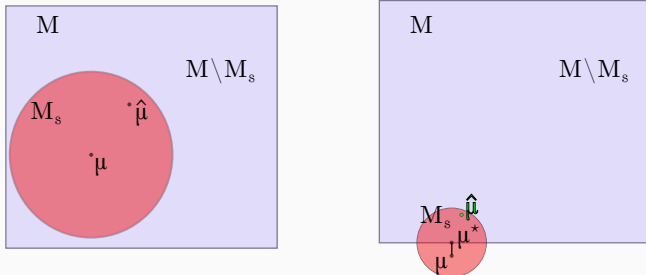
$$\|\hat{\mu} - \mu^*\| < s \quad \text{w.h.p. if} \quad s^2 \geq 2\sigma c_\delta \text{w}(\mathcal{M}_s^\circ)$$

where

$$\mathcal{M}_s^\circ = \{m \in \mathcal{M} : \|m - \mu^*\| = s\}.$$

# What determines these bounds

It's the gaussian width of *neighborhoods* of this best approximation  $\mu^*$ .



More generally, we work with the width of the neighborhood itself.  
And the bound can depend on the quality of our approximation.

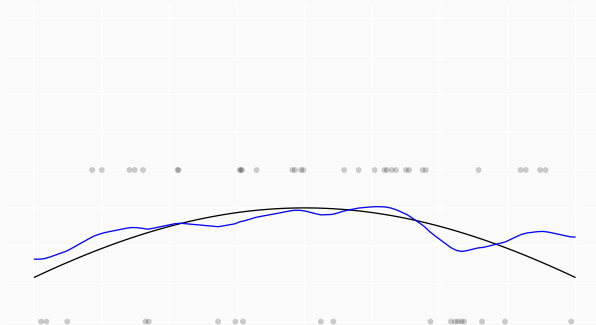
$$\|\hat{\mu} - \mu^*\| < s \quad \text{w.h.p. if} \quad s^2 \geq 2\sigma c_\delta w(\mathcal{M}_s) + 2\|\mu^* - \mu\|$$

where

$$\mathcal{M}_s = \{m \in \mathcal{M} : \|m - \mu^*\| \leq s\}.$$

# Generalization–Non-Gaussian Noise

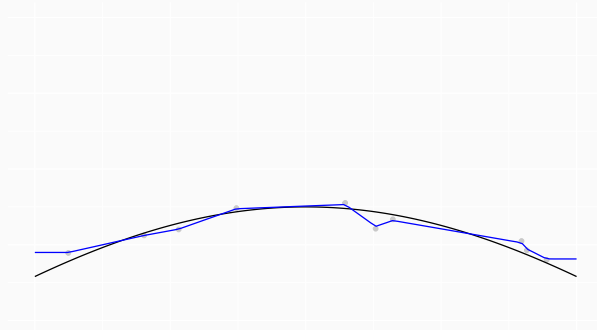
These bounds more or less work with non-gaussian noise, too.  
For example, bounded noise like what we get in *probabilistic classification*





# Generalization–Population Squared Error

Same deal when we're interested in the population two-norm of our error.  
Sampling from our population acts like subgaussian noise.



# To use this, we need to bound gaussian width

We've done this in a few models using specialized techniques.

1. Finite models using the Union Bound and the Gaussian Tail Bound.

$$s^2 \geq cs\sqrt{\log(K)/n} \quad \text{for} \quad s \geq c\sqrt{\log(K)/n}$$

2. Finite-dimensional models using Projection and the Cauchy-Schwarz Inequality.

$$s^2 \geq s\sqrt{K/n} \quad \text{for} \quad s \geq \sqrt{K/n}$$

3. Sobolev models using Fourier Analysis and the Cauchy-Schwarz Inequality.

$$s^2 \geq cs^{1-d/2p}/\sqrt{n} \quad \text{for} \quad s \geq c'n^{-1/(2+d/p)}$$

There are two essential ideas here.

1. Approximating many curves by combinations of a few.
2. Counting.

This week, we'll talk about a completely general technique for bounding width.

We'll use the same two ideas, but our approximations will be subtler.

## Finite Approximations and Gaussian Width

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# Finite Models

- In finite models, bounding width is easy.
- It's the maximum of gaussians with standard deviation  $\leq s/\sqrt{n}$ .

$$\begin{aligned} \mathbb{E} \langle g, m - \mu^\star \rangle^2 &= \mathbb{E} \left( \frac{1}{n} \sum_{i=1}^n g_i \{m(X_i) - \mu^\star(X_i)\} \right)^2 \\ &= \frac{1}{n^2} \sum_{i=1}^n \mathbb{E} g_i^2 \{m(X_i) - \mu^\star(X_i)\}^2 = \frac{\|m - \mu^\star\|^2}{n}. \end{aligned}$$

Q: What happened to the cross terms in the square?

- We can bound this via union bound. It's down to counting the curves in the model.

$$w(\mathcal{M}_s) \leq cs \sqrt{\frac{\log(K)}{n}} \quad \text{if } \mathcal{M} \text{ contains } K \text{ curves } v_1 \dots v_K, \text{ all with } \|v - \mu^\star\|_{L_2(\mathbb{P}_n)} \leq s.$$

- We may be overcounting. This bounds the max of  $K$  totally different gaussians.
- That's kind of the worst case, so if there's correlation we're overcounting.
- And our gaussians are as correlated as the curves in our neighborhood.

$$\mathbb{E} \langle g, v_k \rangle \langle g, v_{k'} \rangle = n^{-2} \mathbb{E} v_k^T g g^T v_{k'} = n^{-2} v_k^T (\mathbb{E} g g^T) v_{k'} = n^{-1} \langle v_k, v_{k'} \rangle.$$

- This definitely won't work for models with infinitely many curves.
- How do we take advantage of this correlation to tackle infinite models?

# Counting Curves in Infinite Models

$$w(\mathcal{M}_s) = \mathbb{E} \max_{v \in \mathcal{M}_s} \langle g, v \rangle \quad \text{for } g \sim N(0, I_{n \times n}).$$

The difference between many of these gaussians  $\langle g, v \rangle$  will be small.

- So small, sometimes, that we don't need to 'pay probability' to bound them all using the union bound. They needn't contribute to  $K$ .
- We can just use the Cauchy-Schwarz inequality to bound differences.

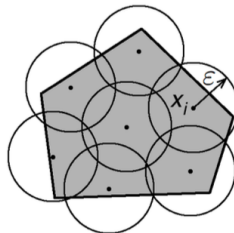
$$|\langle g, u \rangle - \langle g, v \rangle| = |\langle g, u - v \rangle| \leq \|g\| \|u - v\| \approx \|u - v\|.$$

If the curves  $u$  and  $v$  are *close enough*, by bounding  $\langle g, u \rangle$ , we bound  $\langle g, v \rangle$  *for free*.

- This means we can take  $K$  above to be smaller than the total number of curves.
- It's enough that some set  $u_1 \dots u_K$  gets close enough to all curves  $v \in \mathcal{M}$ .

This means we have to talk about how many *meaningfully different* curves we have.

We call a set  $\mathcal{M}^\epsilon$  an  $\epsilon$ -cover for the set  $\mathcal{M}$  if every curve in the set  $\mathcal{M}$  is within a distance  $\epsilon$  of some curve in  $\mathcal{M}^\epsilon$ .



If we have an  $\epsilon$ -cover  $\mathcal{M}_s^\epsilon$  of size  $K_\epsilon$  for  $\mathcal{M}_s$ , then we've got a bound on our width.

$$\begin{aligned}
 w(\mathcal{M}_s) &= \mathbb{E} \left[ \max_{v \in \mathcal{M}_s} \langle g, v \rangle \right] \\
 &= \mathbb{E} \left[ \max_{v \in \mathcal{M}_s} \min_{u \in \mathcal{M}_s^\epsilon} \langle g, v - u \rangle + \langle g, u \rangle \right] \\
 &\lesssim \underbrace{\max_{v \in \mathcal{M}_s} \min_{u \in \mathcal{M}_s^\epsilon} \|v - u\|}_{\epsilon} + \underbrace{\max_{u \in \mathcal{M}_s^\epsilon} \|u\|}_s \sqrt{\frac{\log(K_\epsilon)}{n}}.
 \end{aligned}$$

And this works for infinite models just as well as it does for finite ones. We can think of  $K_\epsilon$  as the size of the neighborhood  $\mathcal{M}_s$  at resolution  $\epsilon$ .

Q: Does the  $\epsilon$ -cover  $\mathcal{M}_s^\epsilon$  have to be a subset of  $\mathcal{M}_s$  for this?

$$\begin{aligned} w(\mathcal{M}_s) &= \mathbb{E} \left[ \max_{v \in \mathcal{M}_s} \langle g, v \rangle \right] \\ &= \mathbb{E} \left[ \max_{v \in \mathcal{M}_s} \min_{u \in \mathcal{M}_s^\epsilon} \langle g, v - u \rangle + \langle g, u \rangle \right] \\ &\lesssim \underbrace{\max_{v \in \mathcal{M}_s} \min_{u \in \mathcal{M}_s^\epsilon} \|v - u\|}_{\epsilon} + \underbrace{\max_{u \in \mathcal{M}_s^\epsilon} \|u\|}_s \sqrt{\frac{\log(K_\epsilon)}{n}}. \end{aligned}$$

# Consequences

Suppose our log covering number grows like  $1/\epsilon$ .

$$\log(K_\epsilon) \leq \epsilon^{-1}$$

We know that  $\hat{\mu}$  is in a neighborhood of  $\mu^*$  of radius  $s$  satisfying

$$s^2 \geq 2c_\delta \sigma w(\mathcal{M}_s) \quad \text{for} \quad w(\mathcal{M}_s) \leq c\epsilon + s\sqrt{\log(K_\epsilon)/n} \approx \epsilon + sn^{-1/2}\epsilon^{-1/2}$$

This width bound holds for all  $\epsilon > 0$ , so we can choose  $\epsilon$  to minimize it.

$$0 = \frac{d}{d\epsilon} \left( \epsilon + sn^{-1/2}\epsilon^{-1/2} \right) = 1 - sn^{-1/2}\epsilon^{-3/2}/2 \quad \text{for} \quad \epsilon = \left( \frac{s}{2\sqrt{n}} \right)^{2/3} \approx s^{2/3}n^{-1/3}$$

And this tells us we're in a neighborhood of radius  $s$  like this.

$$s^2 \geq c\sigma s^{2/3}n^{-1/3} \quad \text{for} \quad s^{4/3} \geq \sigma n^{-1/3} \quad \text{i.e.} \quad s \geq \sigma^{3/4}n^{-1/4}.$$

$\geq c\sigma w(\mathcal{M}_s^\circ)$



# Dissatisfying Results

- We'll show, momentarily, that  $\log(K_\epsilon) \approx 1/\epsilon$  for the Lipschitz model.

$$w(\mathcal{M}_s) \lesssim \epsilon + s \sqrt{\frac{\log(K_\epsilon)}{n}} \approx \epsilon + \frac{s}{\sqrt{\epsilon n}} \approx s^{2/3} n^{-1/3} \quad \text{at optimal} \quad \epsilon \approx s^{2/3} n^{-1/3}.$$

- That gives us a  $n^{-1/4}$  rate.

$$s^2 \geq w(\mathcal{M}_s) \quad \text{if} \quad s^2 \gtrsim s^{2/3} n^{-1/3} \quad \text{i.e. if} \quad s \approx n^{-1/4}.$$

- But we know it converges at a faster rate.
- The Lipschitz model is contained in the Sobolev model of order 1.
- And we proved the rate of convergence  $s \approx n^{-1/3}$  for that using Fourier series.

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Has the covering idea failed us?

No. We just have to make better use of it. We'll do that next class.  
When we do that, we'll see a rough connection to Fourier series.

## Refined Bounds in terms of $\epsilon$ -nets

By working with  $\epsilon$ -nets at different resolutions, we can prove a refined upper bound.

$$w(\mathcal{M}_s^\circ) \lesssim \frac{1}{\sqrt{n}} \int_0^\infty \sqrt{\log(K_\epsilon)} d\epsilon \quad \text{where } K_\epsilon \text{ is the size of the smallest } \epsilon\text{-net for } \mathcal{M}_s^\circ.$$

This multi-resolution argument is called *chaining*.

The bound, *Dudley's Integral Bound*.

The Lipschitz Regression Case

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The bound, *Dudley's Integral Bound*.

The Lipschitz Regression Case

$\log(K_\epsilon) = 0$  for  $\epsilon > s$ . Why? And  $\log(K_\epsilon) \lesssim \epsilon^{-1}$  generally.

$$w(\mathcal{M}_s^\circ) \leq \frac{1}{\sqrt{n}} \int_0^s \epsilon^{-1/2} d\epsilon = n^{-1/2} \epsilon^{1/2} / 2 \Big|_0^s = n^{-1/2} s^{1/2} / 2.$$

and consequently

$$s^2 \geq w(\mathcal{M}_s^\circ) \quad \text{if} \quad s^{3/2} = n^{-1/2} / 2 \quad \text{i.e.} \quad s \propto n^{-1/3}$$

This approach to bounding gaussian width is almost optimal.

There's also a *lower bound*, *Sudakov's Minoration Inequality*, in terms of the size  $K_\epsilon$ .

$$w(\mathcal{M}_s^\circ) \gtrsim \frac{1}{\sqrt{n}} \max_{\epsilon > 0} \epsilon \sqrt{\log(K_\epsilon)}.$$

These bounds are close: the upper bound is no more than  $\log(n)$  times the lower.

The accuracy of our estimator is determined by the rate at which the gaussian width of our model's neighborhood boundary grows.

$$\|\hat{\mu} - \mu^*\| < s \quad \text{with high probability} \quad \text{if} \quad s^2 \gtrsim \sigma w(\mathcal{M}_s^\circ).$$

That gaussian width is a measure of the boundary's size at multiple resolutions.

$$\frac{1}{\sqrt{n}} \max_{\epsilon > 0} \epsilon \sqrt{\log(K_\epsilon)} \underset{\approx}{\approx} w(\mathcal{M}_s^\circ) \underset{\approx}{\approx} \frac{1}{\sqrt{n}} \int_0^\infty \sqrt{\log(K_\epsilon)} d\epsilon.$$

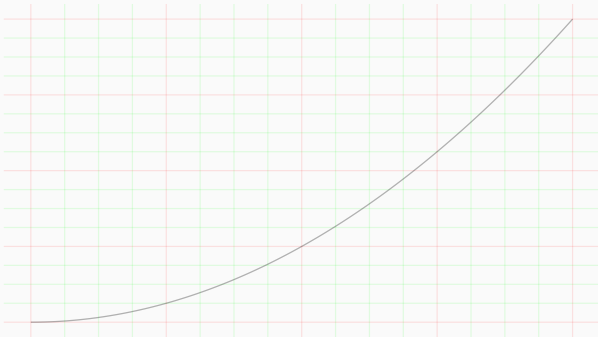
## Finite Approximations and Gaussian Width

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Bounding Our Covering Number in the  
Lipschitz Model

- Think of an  $\epsilon$ -cover of  $\mathcal{U}$  as the set of  $\epsilon$ -approximations  $\pi(u)$  for each  $u$  in  $\mathcal{U}$ .
- Often we base these approximations on a grid. Let's do the 1-Lipschitz case.

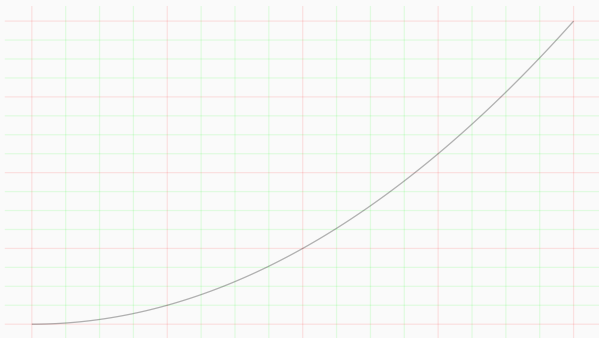
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- Often we base these approximations on a grid. Let's do the 1-Lipschitz case.

$$\mathcal{U} = \{u : |u(x') - u(x)| \leq |x' - x|, |u(x)| \leq 1\}.$$



1. Draw an  $\epsilon$ -spaced grid.
2. At each x-coordinate on the grid, snap to the closest grid point.
3. Because our function is 1-Lipschitz, it can't jump by more than  $\epsilon$  between points.

How many of these are there? Consider  $\epsilon = 1/M$  for an integer  $M$ .

$$(\text{starting points}) \cdot (\text{options per step})^{\text{steps}} = 1/\epsilon \cdot 2^{1/\epsilon}.$$

Some things borrowed from Vershynin's *High Dimensional Probability*.

- The presentation of the refined bounds
- The  $\epsilon$ -net picture.

Its chapters 7-8 are a good, although relatively sophisticated, reference for this stuff.