

Notes - 3 April

(Review) Theorem 3.4.1 - irreducible, aperiodic MC.  $P_{ij}^n \rightarrow_{n \rightarrow \infty} \frac{1}{\mu_j}$  for all  $i, j$ .  $\mu_j$  is mean recurrent time.

Theorem 3.4.4 - An ergodic MC has both stationary and limiting distributions and these are equal.

Recurrent cases in proof:  $X_n, Y_n$  have common  $S, P$ .  $Z_n = (X_n, Y_n) \in S \times S, Z_n$  M.C.  $P_{ij,kl} = P_{ik}P_{jl}$ . Assume  $X_n$  positive recurrent.  $\Pi$  is unique stationary distribution. (With  $Y_n = X_n$ ),  $Z_n$  has stationary distribution  $\nu = (\nu_{ij}, i, j \in S), \nu_{ij} = \Pi_i \Pi_j$ .  $Z_n$  is positive recurrent. Choose  $X_0 = i, Y_0 = j, Z_0 = (i, j)$ . Choose  $s \in S$ .  $T = \min\{n \geq 1 : Z_n = (s, s)\}$ .  $Z_n$  recurrent  $\Rightarrow P(T < \infty) = 1$ .

Observation: Suppose  $m \leq n$  and  $X_m = Y_m$ . Then  $X_n$  and  $Y_n$  are identically distributed. Thus conditional on  $\{T \leq n\}$ ,  $X_n$  and  $Y_n$  have the same distribution. We use this observation and the fact that  $T$  is finite to prove that in large time, the distributions of  $X_n$  and  $Y_n$  are independent of the initial values.

Computation: Start from  $Z_0 = (i, j)$ .  $P_{ik}^n = P(X_n = k) = P(X_n = k, T \leq n) + P(X_n = k, T > n) = P(Y_n = k, T \leq n) + P(Y_n = k, T > n) \leq P(Y_n = k) + P(T > n) = P_{jk}^n + P(T > n)$ . The symmetric argument implies  $P_{jk}^n \leq P_{ik}^n + P(T > n)$ . Hence,  $|P_{ik}^n - P_{jk}^n| \leq P(T > n) \rightarrow_{n \rightarrow \infty} 0$  ( $T$  finite!), for all  $i, j, k \in S$ . So (3.4.2)  $P_{ik}^n - P_{jk}^n \rightarrow 0$  as  $n \rightarrow \infty$  for all  $i, j, k \in S$ . If  $\lim_{n \rightarrow \infty} P_{ik}^n$  exists, then it is independent of  $i$ . We show the limit exists. We write  $\Pi_k - P_{jk}^n = \sum_i \Pi_i (P_{ik}^n - P_{jk}^n)$  (we can write this because  $\Pi$  is stationary ( $\Pi_k, P_{ik}^n$ ) and a pmf ( $P_{jk}^n$ )). For any finite set  $F \in S, \sum_{i \in S} \Pi_i |P_{ik}^n - P_{jk}^n| \leq \sum_{i \in F} |P_{ik}^n - P_{jk}^n| + 2 \sum_{i \notin F} \Pi_i$ . As  $n \rightarrow \infty$ , this converges to  $2 \sum_{i \notin F} \Pi_i$ . This converges to 0 as  $F \uparrow S$ . So (3.4.3)  $\Pi_k - P_{jk}^n = \sum_i \Pi_i (P_{ik}^n - P_{jk}^n) \rightarrow_{n \rightarrow \infty} 0$ . Read in notes about when  $X_n$  is null recurrent, 185-187.

More general version of theorem 3.4.1 (no proof given) that drops irreducibility. Theorem 3.4.6 - For any aperiodic state  $j$  of a Markov chain,  $P_{jj}^n \rightarrow \frac{1}{\mu_j}$  as  $n \rightarrow \infty$ . If  $i$  is any other state,  $P_{ij}^n \rightarrow \frac{f_{ij}}{\mu_j}$  as  $n \rightarrow \infty$ . More to the theorem in notes.

### §3.5 Reversibility

Some physical situations have the property that observations of the system taken at some times look the same if time runs forward or backward. Let  $X_n$  be a Markov chain,  $\{X_n, 0 \leq n \leq N\}$  irreducible, positive recurrent Markov chain, prob transition matrix  $P$  and stationary distribution  $\Pi$ .

Definition 3.5.1 - The reversed chain or time reversal  $Y_n$  is  $Y_n = X_{N-n}, 0 \leq n \leq N$ .

Theorem 3.5.1 -  $Y_n$  is a Markov chain with  $P(Y_{n+1} = j | Y_n = i) = \frac{\Pi_j}{\Pi_i} P_{ji}$ . Proof:  $P(Y_{n+1} = i_{n+1} | Y_n = i_n, \dots, Y_0 = i_0) = \frac{P(Y_k = i_j, 0 \leq k \leq n+1)}{P(Y_k = i_k, 0 \leq k \leq n)} = \frac{P(X_{N-n-1} = i_{n+1}, X_{N-n} = i_n, \dots, X_N = i_0)}{P(X_{N-n} = i_n, \dots, X_N = i_0)} = \frac{\Pi_{i_{n+1}} P_{i_{n+1}, i_n} P_{i_n, i_{n-1}} \dots P_{i_1, i_0}}{\dots} = \frac{\Pi_{i_{n+1}} P_{i_{n+1}, i_n}}{\Pi_{i_n}}$ .

Definition 3.5.2 - The chain is reversible if the probability transition matrices of  $X_n$  and its time reversal  $Y_n$  are the same, (3.5.1)  $\Pi_i P_{ij} = \Pi_j P_{ji}$  for all  $i, j$ . (3.5.1) are called the detailed balance equations. A transition matrix  $P$  and a probability distribution  $\lambda$  are in detailed balance if  $\lambda_i P_{ij} = \lambda_j P_{ji}$  for all  $i, j$ . An irreducible chain  $X_n$  with a stationary distribution  $\Pi$  is reversible in equilibrium if its probability transition matrix is in detailed balance with  $\Pi$ .

Theorem 3.5.1 - Let  $P$  be the probability transition matrix of an irreducible chain  $X_n$  and suppose there is a distribution  $\Pi$  with  $\Pi_i P_{ij} = \Pi_j P_{ji}$  for all  $i, j \in S$ . Then  $\Pi$  is the stationary distribution of  $X_n$  and  $X_n$  is reversible in equilibrium. Proof:  $\sum_i \Pi_i P_{ij} = \sum_i \Pi_j P_{ji} = \Pi_j \sum_i P_{ji} = \Pi_j$  or  $\Pi = \Pi P$ .