

$$P(X_n = i \text{ for some } n \geq 1 \mid X_0 = i) = 1,$$

which says that the probability of returning to state i having started in i is 1.

If

$$P(X_n = i \text{ for some } n \geq 1 \mid X_0 = i) < 1,$$

i is called transient.

start #11 2/26

A recurrent state has the property that the chain returns to the initial state in finite time. For a transient state, there is a positive probability of no return.

Example 3.1.1

Consider the roulette wheel in Ex. 2.1.6. State 0 is trivially recurrent since if $X_0 = 0, X_1 = 0, X_2 = 0, \dots$. State $i \neq 0$ has the property that if we jump to 0, then we cannot return to i . Hence, for $i \neq 0$,

$P(X_n = i \text{ some } n \geq 1 \mid X_0 = i) < 1$,
and i is transient.

Example 3.1.2

Consider the genotype example Ex 2.1.7.
 We assume $X_0 = Aa$. If $X_1 = 0$ or $X_1 = 2$
 there can be no return to X_0 , hence
 $P(X_n = Aa \text{ some } n \geq 1 \mid X_0 = Aa) = P(X_1 = Aa \mid X_0 = Aa) = 1/2$.

Aa is transient.

Example 3.1.3

Consider the off/on system in Ex 2.2.3
 with

$$P = \begin{pmatrix} 1-p & p \\ g & 1-g \end{pmatrix}.$$

Assume $0 < p, g < 1$.

Consider $X_0 = 0$:

$$X_0 = 0 \xrightarrow{P=1-p} X_1 = 0$$

$$P = \cancel{p} \rightarrow X_1 = 1 \xrightarrow{P=g} X_2 = 0$$

$$\cancel{P=1-g} \rightarrow X_2 = 1 \xrightarrow{P=g} X_3 = 0$$

$$\xrightarrow{P=1-g} X_3 = 1 \rightarrow \dots$$

Each step is independent, hence

$$P(X_n = 0 \text{ some } n \geq 1 \mid X_0 = 0)$$

$$\begin{aligned}
 & (1-p) + p \cdot q + p \cdot (1-q) \cdot q + p \cdot (1-q)^2 \cdot q + \dots \\
 &= (1-p) + pq (1 + (1-q) + (1-q)^2 + \dots) \\
 &= (1-p) + pq \frac{1}{1-(1-q)} = (1-p) + pq \frac{1}{q} = 1.
 \end{aligned}$$

So 0 is recurrent. Now recall that in the limit of large time, X_n has probability $q/(p+q)$ of ending up in state 0. Even if $q \ll p$ and we are not likely to end up in state 0, with probability 1, we see $X_n = 0$ for some $n \geq 1$.

We are interested in

Definition 3.1.2

The first passage time from state i to state j is the smallest time it takes to move from state i to state j . In general, we are interested in the mean first passage times.

Definition 3.1.3

Define

$$f_{ij}(n) = P(X_1 \neq j, X_2 \neq j, \dots, X_{n-1} \neq j, X_n = j \mid X_0 = i)$$

to be the probability that the first visit to state j starting from state i takes place in the n^{th} step. We set

$$(3.1.1) \quad f_{ij} = \sum_{n=1}^{\infty} f_{ij}(n)$$

to be the probability that the chain ever visits j starting from i .

Note: j is recurrent iff $f_{jj} = 1$.

We look for a criterion for recurrence in terms of the n step transition probabilities

Recall our convention (pg 57), $P =$ prob. transition matrix with entries (P_{ij}) , and P^n is the n step prob. transition matrix with entries (P_{ij}^n) .

Definition 3.1.4

We define the prob. generating functions

$$P_{ij}(s) = \sum_{n=0}^{\infty} P_{ij}^n s^n$$

$$F_{ij}(s) = \sum_{n=0}^{\infty} f_{ij}(n) s^n,$$

where we set

$P_{ij}^0 = \delta_{ij}$, $f_{ij}(0) = 0$, for all i, j . Note

$f_{ij} = F_{ij}(1)$. We usually assume $|s| < 1$ so that $P_{ij}(s)$ is guaranteed to converge.

Theorem 3.1.1

$$(1) \quad P_{ii}(s) = 1 + F_{ii}(s) P_{ii}(s)$$

$$(2) \quad P_{ij}(s) = F_{ij}(s) P_{jj}(s), \quad i \neq j.$$

Proof

Fix $i, j \in S$. Set

$$A_m = \{X_m = j\}$$

and

$$B_m = \{X_k \neq j \text{ for } 1 \leq k < m, X_m = j\}$$

A_m is the event that X_m hits state j at the m^{th} step. It may have visited j earlier.
 B_m is the event that the first visit to j after 0 is at time m .

The B_m are disjoint, so

$$P(A_m | X_0 = i) = \sum_{k=1}^m P(A_m \cap B_k | X_0 = i)$$

$A_m \cap B_k$ = event where X_m visits state j at the k^{th} and m^{th} steps.

The Markov property implies

$$\begin{aligned} P(A_m \cap B_k | X_0 = i) &= P(A_m | B_k, X_0 = i) P(B_k | X_0 = i) \\ &\quad \downarrow \text{Markov Property to} \\ &= P(A_m | X_k = j) P(B_k | X_0 = i) \\ &\quad \uparrow \\ &\quad \text{initial condition for} \\ &\quad \text{chain for times} \\ &\quad \text{larger than } k \end{aligned}$$

This means

$$(3.1.2) \quad P(A_m | X_0 = i) = p_{ij}^m = \sum_{k=1}^m p_{ij}^{m-k} f_{ij}(k), \quad m=1,2,\dots$$

We multiply by s^m , $|s| < 1$, and sum over $m \geq 0$ to obtain

$$P_{ij}(s) - \delta_{ij} = F_{ij}(s) P_{jj}(s).$$

Theorem 3.1.2

(1) State j is recurrent if $\sum_n P_{jj}^n = \infty$

and if this holds then $\sum_n P_{ij}^n = \infty$ for all i such that $f_{ij} > 0$.

(2) State j is transient if $\sum_n P_{jj}^n < \infty$

and if this holds then $\sum_n P_{ij}^n < \infty$ for all i

(3) If j is transient, then $P_{ij}^n \rightarrow 0$ as $n \rightarrow \infty$ for all i .

Recall that P_{ij}^n is the probability of revisiting state j starting at j in the n th step. So

A condition on $\sum_n P_{ij}^n$ is a condition on the size of the probabilities of revisiting j on any step.

$\sum_n P_{ij}^n$ has to do with the probabilities of visiting state j from state i on any step n .

Proof

We show that j is recurrent if and only if $\sum_n P_{jj}^n = \infty$. We know

$$P_{jj}(s) = \frac{1}{1 - F_{jj}(s)} \quad |s| < 1.$$

As $s \uparrow 1$, $P_{jj}(s) \rightarrow \infty$ if and only if $f_{jj} = F_{jj}(1) = 1$.

We now use Abel's theorem

Theorem 3.1.3

Let $G(s) = \sum_{j=0}^{\infty} a_j s^j$. If $a_j \geq 0$ for all j

and $G(s)$ is finite for $|s| < 1$, then

$\lim_{s \uparrow 1} G(s) = \sum_{j=0}^{\infty} a_j$ (whether the sum is finite or infinite).

This implies

$$\lim_{s \uparrow 1} P_{jj}(s) = \sum_n P_{jj}^n,$$

which shows the claim. The results in the theorem follow (exercise).

Example 3.1.4

Consider the genotype example Ex 2.1.7, where from Ex 2.2.1

$$P^n = \begin{pmatrix} 1 & 0 & 0 \\ \frac{1}{2}(1+(\frac{1}{2})^n) & (\frac{1}{2})^n & \frac{1}{2}(1-(\frac{1}{2})^n) \\ 0 & 0 & 1 \end{pmatrix}$$

For Aa ,

$$\sum_n P_{11}^n = \sum_n (\frac{1}{2})^n = 1 < \infty$$

so Aa is transient.

Example 3.1.5

Consider the off/on system in Ex 2.2.3,

$$P^n = \frac{1}{p+q} \begin{pmatrix} q & p \\ q & p \end{pmatrix} + \frac{(1-p-q)^n}{p+q} \begin{pmatrix} p & -p \\ -q & q \end{pmatrix}$$

For state 0

$$\sum_n P_{00}^n = \sum_n \left(\frac{q}{p+q} + \frac{(1-p-q)^n}{p+q} p \right) = \infty$$

so 0 is recurrent.

#12 3/4

Example 3.1.6Random Walk

We consider the simple random walk in Ex. 2.2.2,

$$X_n = X_0 + \sum_{k=1}^n B_k,$$

$\{B_k\}$ iid. Bernoulli variables with
 $P(B_k=1)=p, P(B_k=-1)=1-p=q.$

Consider any state j . We note that

$P_{jj}^{2n-1} = 0$ for $n=1, 2, \dots$ since we cannot return to j in an odd number of steps.