

Notes - 11 Mar

State  $i$  is recurrent:  $P(X_n = i \text{ some } n \geq 1 | X_0 = i) = 1$ .  $i$  is transient:  $P(X_n = i \text{ some } n \geq 1 | X_0 = i) < 1$ .  
 $f_{ij}(n) = P(X_1 \neq j, X_2 \neq j, \dots, X_n \neq j, X_{n+1} = j | X_0 = i)$ .  $f_{ij} = \sum_{n=1}^{\infty} f_{ij}(n)$ .  $j$  is recurrent  $\iff f_{jj} = 1$ .

Theorem -  $j$  is recurrent if  $\sum_n P_{jj}^n = \infty$ .  $j$  is transient if  $\sum_n P_{jj}^n < \infty$ .

$T_j = \min\{n \geq 1 : X_n = j\}$ . time of first visit to  $j(X_0 = i)$ .  $P(T_i = \infty | X_0 = i) > 0 \iff i$  is transient.  
 $\mu_i = E(T_i | X_0 = i) = \{\sum_{n=1}^{\infty} f_{ii}(n) \text{ for } i \text{ recurrent}, \infty \text{ for } i \text{ transient}\}$ . Recurrent state is null if  $\mu_i = \infty$ .  
Recurrent state is positive if  $\mu_i < \infty$ .

Theorem - Recurrent state is null iff  $P_{ii}^n \rightarrow 0$  as  $n \rightarrow \infty$ .

etc, more review...

New stuff:

Example 3.2.8 -  $S = \{0, 1, 2, 3, 4, 5\}$ .

$$P = \begin{pmatrix} 1/2 & 1/2 & 0 & 0 & 0 & 0 \\ 1/4 & 3/4 & 0 & 0 & 0 & 0 \\ 1/4 & 1/4 & 1/4 & 1/4 & 0 & 0 \\ 1/4 & 0 & 1/4 & 1/4 & 0 & 1/4 \\ 0 & 0 & 0 & 0 & 1/2 & 1/2 \\ 0 & 0 & 0 & 0 & 1/2 & 1/2 \end{pmatrix}$$

$\{0, 1\}, \{4, 5\}$  irreducible and closed. Therefore contain positive recurrent states. 2, 3 are transient:  $2 \rightarrow 3 \rightarrow 5$ . But return to 2 or 3 from 5 is impossible.  $T = \{2, 3\}, C_1 = \{0, 1\}, C_2 = \{4, 5\}$ . All states have period 1 since  $P_{ii} > 0$  for all  $i$  (all entries on diagonal  $> 0$ ). 0, 1, 4, 5 are ergodic. We can compute  $f_0(1) = P_{00} = 1/2, f_{00}(n) = P_{01}(P_{11})^{n-2}P_{10} = 1/2(3/4)^{n-2}1/4, n \geq 2. \mu_0 = \sum_n f_{00}(n) * n = 3$ .

Example 3.2.9 - Success Runs -  $S = \{0, 1, \dots\}$ .  $P = (q_0 \ p_0 \ 0 \ \dots \ \& \ q_1 \ 0 \ p_1 \ 0 \ \dots \ \& \ q_2 \ 0 \ 0 \ p_2 \ 0 \ \dots \ \& \ \dots)$ .  $q_{ii}p_i \geq 0, q_i + p_i = 1$  for all  $i$ . This is a success run chain. Intuition: assume  $p_i = p$  for all  $i$ . We attempt independent Bernoulli trials with probability  $p$  of success. We count the number of successful trials in a row. If we had  $n$  successes in a row, we can extend the run to  $n+1$  if we have success on the next trial or we start over with a run of 0 if we fail in the next trial. This gives the row  $(q(0\text{th}) \ 0 \ \dots \ p((n+1)\text{st}) \ 0 \ \dots)$ . We assume  $0 < p_i < 1$  for all  $i$  so the chain is irreducible. This means state  $i$  is recurrent iff state 0 is recurrent. We have  $f_{00}(1) = q_0$ , and for  $n \geq 2, f_{00}(n) = P(X_1 = 1, X_2 = 2, \dots, X_{n-1} = n-1, X_n = 0 | X_0 = 0) = P_0 P_1 P_2 \dots P_{n-2} * q_{n-1}$ . Set  $U_n = \prod_{i=0}^n P_i, n \geq 0$  since  $q_{n-1} = 1 - P_{n-1}, f_{00}(n) = U_{n-2} - U_{n-1} = \prod_{i=0}^{n-2} P_i (1 - P_{n-1})$ . So  $\sum_{n=1}^{\infty} f_{00}(n) = q_0 + (U_0 - U_1) + \dots + (U_{N-1} - U_N) = q_0 + U_0 - U_N = 1 - U_N$ . 0 is recurrent iff  $U_N = \prod_{i=0}^N P_i \rightarrow 0$  as  $N \rightarrow \infty$ . L'Hopital's rule implies that if  $0 < P_i < 1$  for all  $i, U_N = \prod_{i=0}^N P_i \rightarrow 0 \iff \sum_{i=0}^{\infty} (1 - P_i) = \infty. \prod_{i=0}^{\infty} P_i > 0 \iff \sum_{i=0}^{\infty} (1 - P_i) < \infty$ . 0 is recurrent iff  $\sum_{i=0}^{\infty} (1 - P_i) = \infty$ , or the  $P_i$ 's cannot be too close to 1. If  $P_i = 1 - (1/2)^i$ , not recurrent.  $P_i$  constant then recurrent. (Chapter IV, section §3 in text.)

### § 3.3 - Stationary distributions and the limit theorem

We consider behavior as  $n \rightarrow \infty$ . Does the distribution of  $X_n$  converge to something?

Example 3.3.1 - ON/OFF system - ex 2.2.3 -  $P = (1-p, p \ \& \ q, 1-q)$ .  $P^n$  as before.  $0 < p < 1, 0 < q < 1, P^n \rightarrow \frac{1}{p+q} * (q \ p \ \& \ q \ p)$  [as before]. We choose the initial state  $X_0$  according to the probabilities  $P(X_0 = 0) = \nu_0, P(X_0 = 1) = \nu_1 = 1 - \nu_0$ .

Definition 3.3.1 - An initial distribution is a probability distribution for the initial state of a Markov chain. The probability distribution of  $X_1$ , conditioned on  $X_0$  is  $P(X_1 = j | X_0) = P_{0j}\nu_0 + P_{1j}\nu_1, j = 0, 1$ . Matrix notation  $(P(X_1 = 0 | X_0)P(X_1 = 1 | X_0)) = \nu P$ . Suppose we take  $\nu_0 = \frac{q}{q+p}, \nu_1 = \frac{p}{q+p}$ . If we compute,  $P(X_1 = 0) = (1-p)\frac{q}{q+p} + q\frac{p}{q+p} = \frac{q}{q+p} = \nu_0$  and  $P(X_1 = 1) = \nu_1$ . In matrix notation  $\nu = \nu P$ . That particular initial distribution does not change over time.

Definition 3.3.2 - Let  $S$  = state space, The vector  $\Pi$  is a stationary distribution if  $\Pi = (\Pi_i)_{i \in S}$  satisfies (1)  $\Pi_i \geq 0$  all  $i, \sum_{i \in S} \Pi_i = 1$ , (2)  $\Pi = \Pi P (\Pi_j = \sum_{i \in S} \Pi_i P_{ij} \text{ all } j \in S)$ .  $P$  = probability transition matrix. These are also called invariant distributions and equilibrium distributions.

Theorem 3.3.1 - If  $\Pi$  is a stationary distribution, (3.3.1)  $\Pi P^n = \Pi$  for all  $n \geq 0$ . If  $X_0$  has distribution  $\Pi$ , then so does  $X_n$  for  $n \geq 0$ . Proof: exercise.

Aside: long time behavior of ODE's  $\dot{y} = f(y)$ . Stead-state/equilibrium solutions  $f(y_s) = 0 \Rightarrow y_s$  constant,  $\dot{y}_s = f(y_s) = 0$ .

We assume the chain is irreducible and explore the existence of stationary distributions.

Example 3.3.2 - Consider ex 3.3.1 (ON/OFF),  $\Pi = \Pi P = (\Pi_0 \Pi_1)(1 - p, p \text{ \& q, } 1 - q) = (\Pi_0 \Pi_1) \Rightarrow \Pi_1 = p/q \Pi_0$  (1st equation),  $\Pi_1 = p/q \Pi_0$  (2nd equation),  $\Pi_0 + \Pi_1 = 1 \Rightarrow \Pi = (\frac{q}{p+q}, \frac{p}{p+q})$ .