Pattern Recognition 2015 Linear Models for Regression

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Linear Regression Model

The central assumption of linear regression is

$$\mathbb{E}[t|x] = y(x) = w_0 + w_1 x$$

Or, alternatively

$$t = w_0 + w_1 x + \varepsilon$$

with $\mathbb{E}[\varepsilon|x] = 0$.

Usually, we also assume that $var[t|x] = \sigma^2$, i.e. t has the same variance for each value of x.

For ML estimation, we typically assume $\varepsilon \sim \mathcal{N}(0, \sigma^2)$.



Minimizing empirical loss

Given training data

$$D = \{(x_1, t_1), (x_2, t_2), \dots, (x_N, t_N)\},\$$

find the values of w_0 and w_1 such that the sum of squared errors

$$E_D(w_0, w_1) = \sum_{n=1}^{N} (t_n - \underbrace{(w_0 + w_1 x_n)}^{\text{prediction for } t_n})^2$$
$$= \sum_{n=1}^{N} (t_n - w_0 - w_1 x_n)^2$$

is minimized.

Partial derivative with respect to intercept:

$$\frac{\partial E_D}{\partial w_0} = \sum_{n=1}^{N} 2(t_n - w_0 - w_1 x_n)(-1)$$
$$= -2 \sum_{n=1}^{N} (t_n - w_0 - w_1 x_n)$$

Equate to zero

$$\sum_{n=1}^{N} (t_n - w_0 - w_1 x_n) = \sum_{n=1}^{N} e_n = 0$$

Partial derivative with respect to slope:

$$\frac{\partial E_D}{\partial w_1} = \sum_{n=1}^{N} 2(t_n - w_0 - w_1 x_n)(-x_n)$$
$$= -2 \sum_{n=1}^{N} x_n (t_n - w_0 - w_1 x_n)$$

Equate to zero

$$\sum_{n=1}^{N} x_n (t_n - w_0 - w_1 x_n) = \sum_{n=1}^{N} x_n e_n = 0$$



Expand and collect terms:

$$\sum_{n=1}^{N} t_n = Nw_0 + w_1 \sum_{n=1}^{N} x_n$$
 (1)

$$\sum_{n=1}^{N} x_n t_n = w_0 \sum_{n=1}^{N} x_n + w_1 \sum_{n=1}^{N} x_n^2$$
 (2)

To solve for w_0 divide (1) by N:

$$w_0 = \bar{t} - w_1 \bar{x}$$

Hence, the least squares fitted line goes through the point of means (\bar{x},\bar{t}) .



To solve for w_1 , multiply (1) by $\sum x_n$ and (2) by N

$$\sum x_n \sum t_n = Nw_0 \sum x_n + w_1 \left(\sum x_n\right)^2 \tag{3}$$

$$N\sum x_n t_n = Nw_0 \sum x_n + Nw_1 \sum x_n^2$$
 (4)

Subtract (3) from (4) and solve for w_1 :

$$w_1 = \frac{N \sum x_n t_n - \sum x_n \sum t_n}{N \sum x_n^2 - (\sum x_n)^2}$$

Example

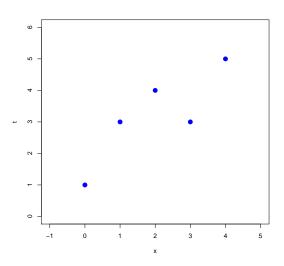
n	Xn	tn	$x_n t_n$	x_n^2
1	0	1	0	0
2	1	3	3	1
3	2	4	8	4
4	3	3	9	9
5	4	5	20	16
\sum	10	16	40	30

$$w_{1} = \frac{N \sum x_{n} t_{n} - \sum x_{n} \sum t_{n}}{N \sum x_{n}^{2} - (\sum x_{n})^{2}}$$
$$= \frac{5 \times 40 - 10 \times 16}{5 \times 30 - 10^{2}} = \frac{4}{5}$$

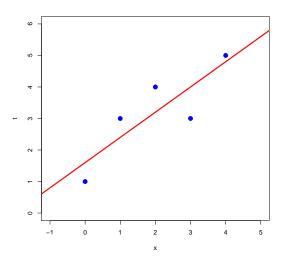
$$w_0 = \bar{t} - w_1 \bar{x} = \frac{16}{5} - \left(\frac{4}{5}\right) \left(\frac{10}{5}\right) = \frac{8}{5}$$



Scatter plot of Training Data



Fitted Line: y(x) = 1.6 + 0.8 x



Decomposition of total sample variation in t

- ② $\sum (y_n \bar{t})^2$ = explained sum of squares = SSR

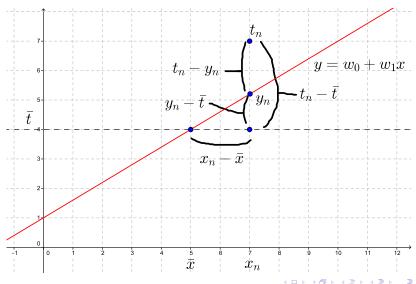
We have

$$SST = SSR + SSE$$

Proportion of variation in t explained by x:

$$R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$$

Decomposition of variation in t





Example: Computation of R^2

Fitted model

$$y(x) = 1.6 + 0.8x$$

n	Xn	tn	Уn	e _n	e_n^2	$(t_n-\bar{t})^2$	$(y_n-\bar{t})^2$
1	0	1	8/5	-3/5	9/25	121/25	64/25
2	1	3	12/5	3/5	9/25	1/25	16/25
3	2	4	16/5	4/5	16/25	16/25	0
4	3	3	20/5	-1	25/25	1/25	16/25
5	4	5	24/5	1/5	1/25	81/25	64/25
\sum	10	16	16	0	60/25	220/25	160/25

$$220/25 = 60/25 + 160/25$$

$$(SST) (SSE) (SSR)$$

$$R^{2} = \frac{SSR}{SST} = \frac{160}{220} \approx 0.73$$

Linear regression through the origin

Suppose we know that the population regression line goes through the origin, i.e.

$$\mathbb{E}[t|x] = wx$$

Find the value of w such that the sum of squared errors

$$E_D(w) = \sum_{n=1}^{N} (t_n - wx_n)^2$$

is minimized.



Regression through the origin: calculus

Take the derivative

$$\frac{dE_D}{dw} = -2\sum (t_n - wx_n)x_n$$

and equate to zero

$$\sum x_n t_n - w \sum x_n^2 = 0$$

so we get

$$w = \frac{\sum x_n t_n}{\sum x_n^2}$$

Regression through the origin: geometry

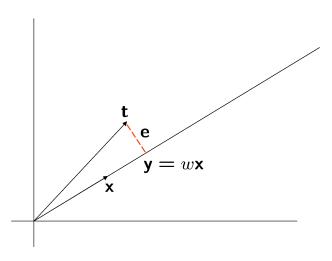
Regression through the origin: $y_n = wx_n$

 $D = \{(2,5), (1,3)\}$ contains only two observations.

$$\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} 2 \\ 1 \end{bmatrix} \text{ and } \mathbf{t} = \begin{bmatrix} t_1 \\ t_2 \end{bmatrix} = \begin{bmatrix} 5 \\ 3 \end{bmatrix}$$
$$\mathbf{e} = \begin{bmatrix} e_1 \\ e_2 \end{bmatrix} \text{ and } \mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}$$

$$\mathbf{y} = w\mathbf{x}$$
 and $\mathbf{e} = \mathbf{t} - w\mathbf{x}$

Least Squares Solution (*N* dimensional space!)



Length of
$$\mathbf{e} = \sqrt{\mathbf{e} \cdot \mathbf{e}} = \sqrt{e_1^2 + e_2^2}$$
.

Least Squares Solution

To minimize the length of \mathbf{e} , it must be perpendicular to \mathbf{x} so $\mathbf{x} \cdot \mathbf{e} = 0$.

$$\mathbf{x} \cdot \mathbf{e} = \mathbf{x} \cdot (\mathbf{t} - w\mathbf{x}) = \mathbf{x} \cdot \mathbf{t} - w\mathbf{x} \cdot \mathbf{x} = 0$$

Therefore

$$w = \frac{\mathbf{x} \cdot \mathbf{t}}{\mathbf{x} \cdot \mathbf{x}}$$

Matrix notation

$$w = \frac{\mathbf{x}^{\top} \mathbf{t}}{\mathbf{x}^{\top} \mathbf{x}}$$
 or $w = (\mathbf{x}^{\top} \mathbf{x})^{-1} \mathbf{x}^{\top} \mathbf{t}$

Solution of Numerical Example

Solution of the numerical example

$$\mathbf{x}^{\mathsf{T}}\mathbf{t} = \begin{bmatrix} 2 & 1 \end{bmatrix} \begin{bmatrix} 5 \\ 3 \end{bmatrix} = 13$$

and

$$\mathbf{x}^{\mathsf{T}}\mathbf{x} = \begin{bmatrix} 2 & 1 \end{bmatrix} \begin{bmatrix} 2 \\ 1 \end{bmatrix} = 5$$

which yields

$$w = \frac{\mathbf{x}^{\mathsf{T}} \mathbf{t}}{\mathbf{x}^{\mathsf{T}} \mathbf{x}} = \frac{13}{5} = 2.6$$



Simple linear regression in matrix terms

We can write the observed t values as

$$t_n = w_0 + w_1 x_n + e_n \qquad \qquad n = 1, \dots, N$$

which is short for

$$t_1 = w_0 + w_1 x_1 + e_1$$

 $t_2 = w_0 + w_1 x_2 + e_2$
 \vdots
 $t_N = w_0 + w_1 x_N + e_N$

Matrix Notation

$$\mathbf{X} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots \\ 1 & x_N \end{bmatrix} \mathbf{t} = \begin{bmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{bmatrix} \mathbf{e} = \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_N \end{bmatrix} \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$

Then we can simply write

$$t = Xw + e$$



Check

$$\mathbf{t} = \mathbf{X}\mathbf{w} + \mathbf{e}$$

$$\begin{bmatrix} t_{1} \\ t_{2} \\ \vdots \\ t_{N} \end{bmatrix} = \begin{bmatrix} 1 & x_{1} \\ 1 & x_{2} \\ \vdots \\ 1 & x_{N} \end{bmatrix} \begin{bmatrix} w_{0} \\ w_{1} \end{bmatrix} + \begin{bmatrix} e_{1} \\ e_{2} \\ \vdots \\ e_{N} \end{bmatrix}$$

$$= \begin{bmatrix} w_{0} + w_{1}x_{1} \\ w_{0} + w_{1}x_{2} \\ \vdots \\ w_{0} + w_{1}x_{N} \end{bmatrix} + \begin{bmatrix} e_{1} \\ e_{2} \\ \vdots \\ e_{N} \end{bmatrix}$$

$$= \begin{bmatrix} w_{0} + w_{1}x_{1} + e_{1} \\ w_{0} + w_{1}x_{2} + e_{2} \\ \vdots \\ w_{0} + w_{1}x_{N} + e_{N} \end{bmatrix}$$



Least Squares Solution

y is a linear combination of the columns of **X**:

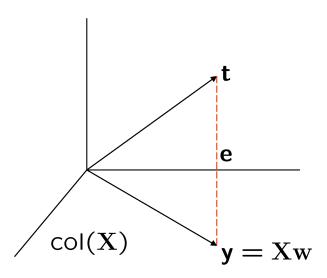
$$y = Xw$$

Typically, \mathbf{t} is not in the column space of \mathbf{X} . Find the value of \mathbf{y} that is closest to \mathbf{t} . For this to be the case, the error vector

$$e = t - Xw$$

must be orthogonal to *all columns* of **X**.

Least Squares Solution (N dimensional space)



Least Squares Solution

In other words, we should have

$$\mathbf{X}^{\top}\mathbf{e}=\mathbf{0}.$$

Since

$$e = (t - Xw)$$

we should have

$$\mathbf{X}^{\top}(\mathbf{t} - \mathbf{X}\mathbf{w}) = \mathbf{X}^{\top}\mathbf{t} - \mathbf{X}^{\top}\mathbf{X}\mathbf{w} = \mathbf{0}.$$

It follows that

$$\boldsymbol{X}^{\top}\boldsymbol{X}\boldsymbol{w} = \boldsymbol{X}^{\top}\boldsymbol{t}$$

Least Squares Solution

So we have

$$\mathbf{X}^{\top}\mathbf{X}\mathbf{w} = \mathbf{X}^{\top}\mathbf{t}$$

Premultiply both sides by the inverse of $\mathbf{X}^{\top}\mathbf{X}$:

$$(\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{X}\mathbf{w} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{t}$$

We then find, since $(\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{X} = \mathbf{I}$ and $\mathbf{I}\mathbf{w} = \mathbf{w}$:

$$\mathbf{w} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{t} \tag{3.15}$$

Numeric example

$$\begin{aligned} \mathbf{D} &= \{(0,1),(1,1),(2,2),(3,2)\} \\ \mathbf{X} &= \begin{bmatrix} 1 & 0 \\ 1 & 1 \\ 1 & 2 \\ 1 & 3 \end{bmatrix} \quad \mathbf{t} = \begin{bmatrix} 1 \\ 1 \\ 2 \\ 2 \end{bmatrix} \\ \mathbf{X}^{\top}\mathbf{X} &= \begin{bmatrix} 4 & 6 \\ 6 & 14 \end{bmatrix} \quad \mathbf{X}^{\top}\mathbf{t} = \begin{bmatrix} 6 \\ 11 \end{bmatrix} \end{aligned}$$

Numeric Example

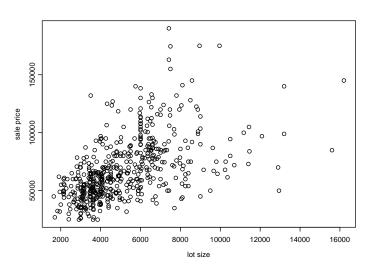
Now, since

$$\begin{bmatrix} a & b \\ c & d \end{bmatrix}^{-1} = \frac{1}{ad - bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$$

we get

$$\mathbf{w} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{t} = \frac{1}{20} \begin{bmatrix} 14 & -6 \\ -6 & 4 \end{bmatrix} \begin{bmatrix} 6 \\ 11 \end{bmatrix}$$
$$= \frac{1}{20} \begin{bmatrix} 18 \\ 8 \end{bmatrix} = \begin{bmatrix} 9/10 \\ 4/10 \end{bmatrix} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$

Scatter plot of lot size and sale price





Least Squares fitted line

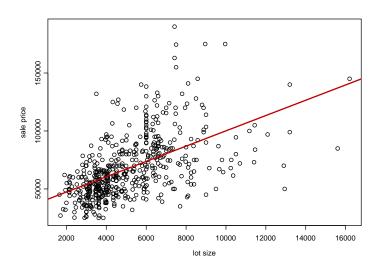
Using R we find:

sale price
$$=34136.1916+6.5988\times lot\ size$$

$$R^2 = 0.2871$$

Model explains only about 30% of variation in sale price.

Least Squares fitted line



Multiple Linear Regression

Usually, you want to use more than one input variable to predict t.

The basic assumption is

$$\mathbb{E}[t|\mathbf{x}] = w_0 + w_1 x_1 + w_2 x_2 + \ldots + w_{M-1} x_{M-1}$$

Multiple Linear Regression

We can write the observed t values as

$$t_n = w_0 + w_1 x_{n,1} + w_2 x_{n,2} + \ldots + w_{M-1} x_{n,M-1} + e_n$$

which is short for

$$t_1 = w_0 + w_1 x_{1,1} + w_2 x_{1,2} + \ldots + w_{M-1} x_{1,M-1} + e_1$$

$$t_2 = w_0 + w_1 x_{2,1} + w_2 x_{2,2} + \ldots + w_{M-1} x_{2,M-1} + e_2$$

$$\vdots$$

$$t_N = w_0 + w_1 x_{N,1} + w_2 x_{N,2} + \ldots + w_{M-1} x_{N,M-1} + e_N$$

Notation and Least Squares Solution

$$\mathbf{X} = \begin{bmatrix} 1 & x_{1,1} & x_{1,2} & \dots & x_{1,M-1} \\ 1 & x_{2,1} & x_{2,2} & \dots & x_{2,M-1} \\ \vdots & & & & \\ 1 & x_{N,1} & x_{N,2} & \dots & x_{N,M-1} \end{bmatrix}$$

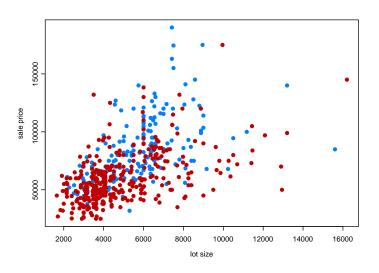
$$\mathbf{t} = \begin{bmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{bmatrix} \mathbf{e} = \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_N \end{bmatrix} \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_{M-1} \end{bmatrix}$$

Then we can write

$$\mathbf{t} = \mathbf{X}\mathbf{w} + \mathbf{e}, \quad \mathbf{w} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{t}$$



Scatter plot of lot size, airco and sale price





Fitted Equation

sale.price =
$$32692.9 + 5.6 \times lot.size + 20174.5 \times air.cond$$

Or, since air.cond is binary:

sale.price =
$$32692.9 + 5.6 \times lot.size$$

when air.cond=0.

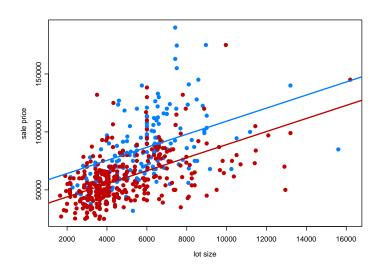
sale.price =
$$(32692.9 + 20174.5) + 5.6 \times lot.size$$

when air.cond=1.

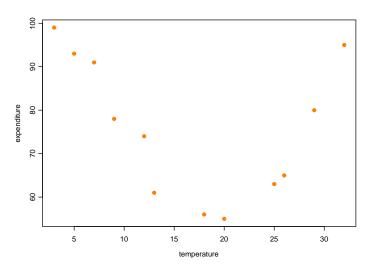
$$R^2 = 0.4048$$



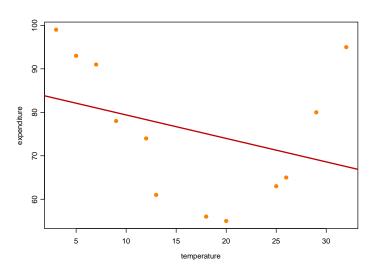
Fitted Equation



Scatter plot of Temperature and Energy Use



Fitting a Linear Function



Linear Equation

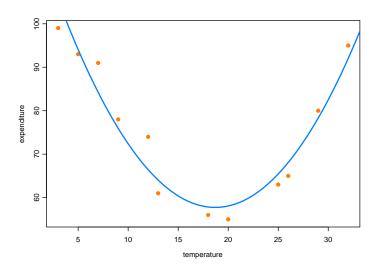
Fitted equation:

$$expenditure = 84.78 - 0.54 \times temperature$$

$$R^2 \approx 0.11$$

Bad fit!

Fitting a Quadratic Function



Quadratic Equation

Fitted equation:

$$expenditure = 125.44 - 7.24 \times temp + 0.19 \times temp^2$$

$$R^2 \approx 0.93$$

Spectacular improvement for only one extra parameter!

General Linear Model (OK, so I lied ...)

The term *linear* in linear regression means linear in the *parameters*, not linear in the *input variables*!

For example:

$$t = w_0 + w_1 x_1 + w_2 x_2 + w_3 x_1^2 + w_4 x_1 x_2 + \varepsilon$$

is still a linear regression model.



General Linear Model

Linear in the features, but not necessarily linear in the input variables that generate them.

$$\phi(t_n) = w_0\phi_0(\mathbf{x}_n) + w_1\phi_1(\mathbf{x}_n) + \ldots + w_{M-1}\phi_{M-1}(\mathbf{x}_n) + \varepsilon_n$$

- $\mathbf{x}_n = (x_{n,1}, \dots, x_{n,D})$
- $w_0, w_1, \ldots, w_{M-1}$: unknown parameters to be estimated;
- $\phi(\cdot), \phi_0(\cdot), \dots, \phi_{M-1}(\cdot)$: functions that do not involve unknown parameters; "basis functions".

Modeling Pizza Expenditure

Linear Model

$$\mathbb{E}(\mathsf{pizza}) = w_0 + w_1 imes \mathsf{inc} + w_2 imes \mathsf{age}$$
 $rac{\partial \mathbb{E}(\mathsf{pizza})}{\partial \mathsf{inc}} = w_1$ $rac{\partial \mathbb{E}(\mathsf{pizza})}{\partial \mathsf{age}} = w_2$

Fitted equation:

$$\mathsf{pizza} = 342.88 + 0.0024 \times \mathsf{inc} - 7.58 \times \mathsf{age}$$

$$R^2 \approx 0.33$$

Modeling Pizza Expenditure

Model with interaction between income and age:

$$\mathbb{E}(\mathsf{pizza}) = w_0 + w_1 \times \mathsf{inc} + w_2 \times \mathsf{age} + w_3 \times (\mathsf{age} \times \mathsf{inc})$$

Effects of income and age:

$$\frac{\partial \mathbb{E}(\mathsf{pizza})}{\partial \mathsf{inc}} = w_1 + w_3 \times \mathsf{age}$$

$$\frac{\partial \mathbb{E}(\mathsf{pizza})}{\partial \mathsf{age}} = w_2 + w_3 \times \mathsf{inc}$$

Modeling Pizza Expenditure

Fitted equation:

$$\begin{array}{rl} \mbox{pizza} &=& 161.47 + 0.01 \times \mbox{inc} - 2.98 \times \mbox{age} \\ &-& 0.0002 \times \mbox{(age} \times \mbox{inc)} \end{array}$$

$$R^2 \approx 0.39$$

Effect of income on pizza expenditure

$$\begin{array}{ll} \frac{\partial \mathbb{E}(\mathsf{pizza})}{\partial \mathsf{inc}} &=& w_1 + w_3 \times \mathsf{age} = 0.01 - 0.0002 \times \mathsf{age} \\ &=& \left\{ \begin{array}{ll} 0.006 & \mathsf{for age} = 20 \\ 0 & \mathsf{for age} = 50 \end{array} \right. \end{array}$$

So on average, a 20 year old will spend 60 cents on pizza of every 100 dollar of extra income.

How to in R

```
# read data (put header=T if first row in data file contains
# names of variables)

> pizza.dat <- read.table("C:/PR/pizza.txt", header=T)

# show first 5 rows

> pizza.dat[1:5,]

pizza sex edu1 edu2 edu3 income age
```

```
109
               0
                    0
                            15000
                                   25
2
                            30000 45
               0
                    0
3
               0
                    0
                          12000 20
   108 1
               0
                    0
                         0 20000 28
5
   220
                    0
                            15000 25
```

How to in R

```
# fit model with interaction between age and income
> pizza.model1 <- lm(pizza \sim age + income + age:income,
data=pizza.dat)
# show results (stuff deleted)
> summary(pizza.model1)
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.615e+02 1.207e+02 1.338
                                         0.1892
      -2.977e+00 3.352e+00 -0.888 0.3803
age
income 9.074e-03 3.670e-03 2.473 0.0183 *
age:income -1.602e-04 8.673e-05 -1.847 0.0730.
```

Multiple R-Squared: 0.3873

Interaction with a binary variable: house prices

$$\mathbb{E}(\mathsf{sale.price}) = w_0 + w_1 \times \mathsf{lot.size} + w_2 \times (\mathsf{air.cond} \times \mathsf{lot.size})$$

Price per square foot depends on presence of airco:

$$\mathbb{E}(\text{sale.price}) = w_0 + w_1 \times \text{lot.size}$$

if no airco (air.cond=0), and

$$\mathbb{E}(\mathsf{sale.price}) = w_0 + (w_1 + w_2) \times \mathsf{lot.size}$$

if air.cond=1.



Fitted Equation

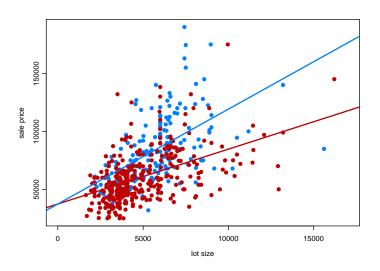
Fitted equation:

sale.price =
$$37341.69 + 4.73 \times lot.size +$$

 $3.45 \times (air.cond \times lot.size)$

$$R^2 \approx 0.41$$
.

Graph of fitted equation



- Weekly counts for 50 million of the most common search queries in the US between 2003 and 2008.
- Separate aggregate weekly counts were kept for every query in each state.
- Divide by total number of queries during that week to obtain a query fraction.
- CDC publishes national and regional data on influenza-like illness (ILI) physician visits on a weekly basis, typically with a 1-2 week reporting lag.

Fit linear regression model:

$$logit(P) = w_0 + w_1 logit(Q) + \varepsilon$$

where P is the percentage of ILI physician visits, Q is the ILI-related query fraction, and

$$logit(P) = ln\left\{\frac{P}{1-P}\right\}$$

Note that P and Q are fractions, i.e. numbers between 0 and 1.

Given fitted values for w_0 and w_1 , and a value for Q, predict

$$logit(P) = w_0 + w_1 logit(Q),$$

or

$$P = \frac{e^{w_0 + w_1 \operatorname{logit}(Q)}}{1 + e^{w_0 + w_1 \operatorname{logit}(Q)}}$$

- First this model is fitted for 50 million candidate queries.
- Sort queries on correlation between prediction and true target values.
- **3** Compare models including top N queries (N = 1, ..., 100).



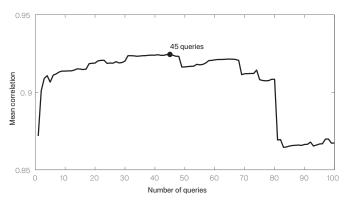


Figure 1: An evaluation of how many top-scoring queries to include in the ILI-related query fraction. Maximal performance at estimating out-of-sample points during cross-validation was obtained by summing the top 45 search queries. A steep drop in model performance occurs after adding query 81, which is "oscar nominations".

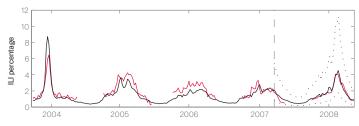


Figure 2: A comparison of model estimates for the Mid-Atlantic Region (black) against CDC-reported ILI percentages (red), including points over which the model was fit and validated. A correlation of 0.85 was obtained over 128 points from this region to which the model was fit, while a correlation of 0.96 was obtained over 42 validation points. 95% prediction intervals are indicated.

	Top 45 Queries		Next 55 Queries	
Search Query Topic	Ν	Weighted	Ν	Weighted
Influenza Complication	11	18.15	5	3.40
Cold/Flu Remedy	8	5.05	6	5.03
General Influenza Symptoms	5	2.60	1	0.07
Term for Influenza	4	3.74	6	0.30
Specific Influenza Symptom	4	2.54	6	3.74
Symptoms of an Influenza Complication	4	2.21	2	0.92
Antibiotic Medication	3	6.23	3	3.17
General Influenza Remedies	2	0.18	1	0.32
Symptoms of a Related Disease	2	1.66	2	0.77
Antiviral Medication	1	0.39	1	0.74
Related Disease	1	6.66	3	3.77
Unrelated to Influenza	0	0.00	19	28.37
	45	49.40	55	50.60

Table 1: Topics found in search gueries which were found to be most correlated with CDC ILI data. The top 45 queries were used in our final model; the next 55 queries are presented for comparison purposes. The number of queries in each topic is indicated, as well as query volume-weighted counts, reflecting the relative frequency of queries in each topic.

Regularized Least Squares

Add regularization term to control overfitting

$$E(\mathbf{w}) = E_D(\mathbf{w}) + \lambda E_W(\mathbf{w}) \tag{3.24}$$

Ridge regression

$$E_W(\mathbf{w}) = \sum_i w_i^2 = \mathbf{w}^\top \mathbf{w}$$
 (3.25)

The ridge regression error function is minimized by:

$$\mathbf{w} = (\mathbf{X}^{\top}\mathbf{X} + \lambda \mathbf{I})^{-1}\mathbf{X}^{\top}\mathbf{t}$$
 (3.28)

Regularization

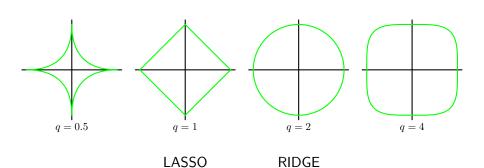
A more general regularizer

$$\sum_{n=1}^{N} \{t_n - \mathbf{w}^{\top} \mathbf{x}_n\}^2 + \lambda \sum_{j=0}^{M-1} |w_j|^q$$
 (3.29)

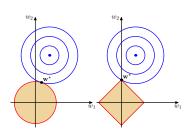
where q = 2 corresponds the ridge regression.

The case of q = 1 is known as the LASSO.

Contours of Regularization Term



LASSO gives sparse solution



Minimize

$$E_D(\mathbf{w}) = \sum_{n=1}^{N} \{t_n - \mathbf{w}^{\top} \mathbf{x}_n\}^2$$
 (3.12)

subject to

$$\sum_{j=0}^{M-1} |w_j|^q \le \eta \tag{3.30}$$

for an appropriate value of the parameter $\eta.$