

# Variance from First Principles

CS 109  
Lecture 15  
April 29th, 2016

# Course Mean

$E[CS109]$

*This is actual midpoint of course  
(Just wanted you to know)*

# Review

# Bayes Theorem

- Ross's form:

$$\begin{aligned} P(E) &= P(EF) + P(EF^c) \\ &= P(E | F) P(F) + P(E | F^c) P(F^c) \end{aligned}$$

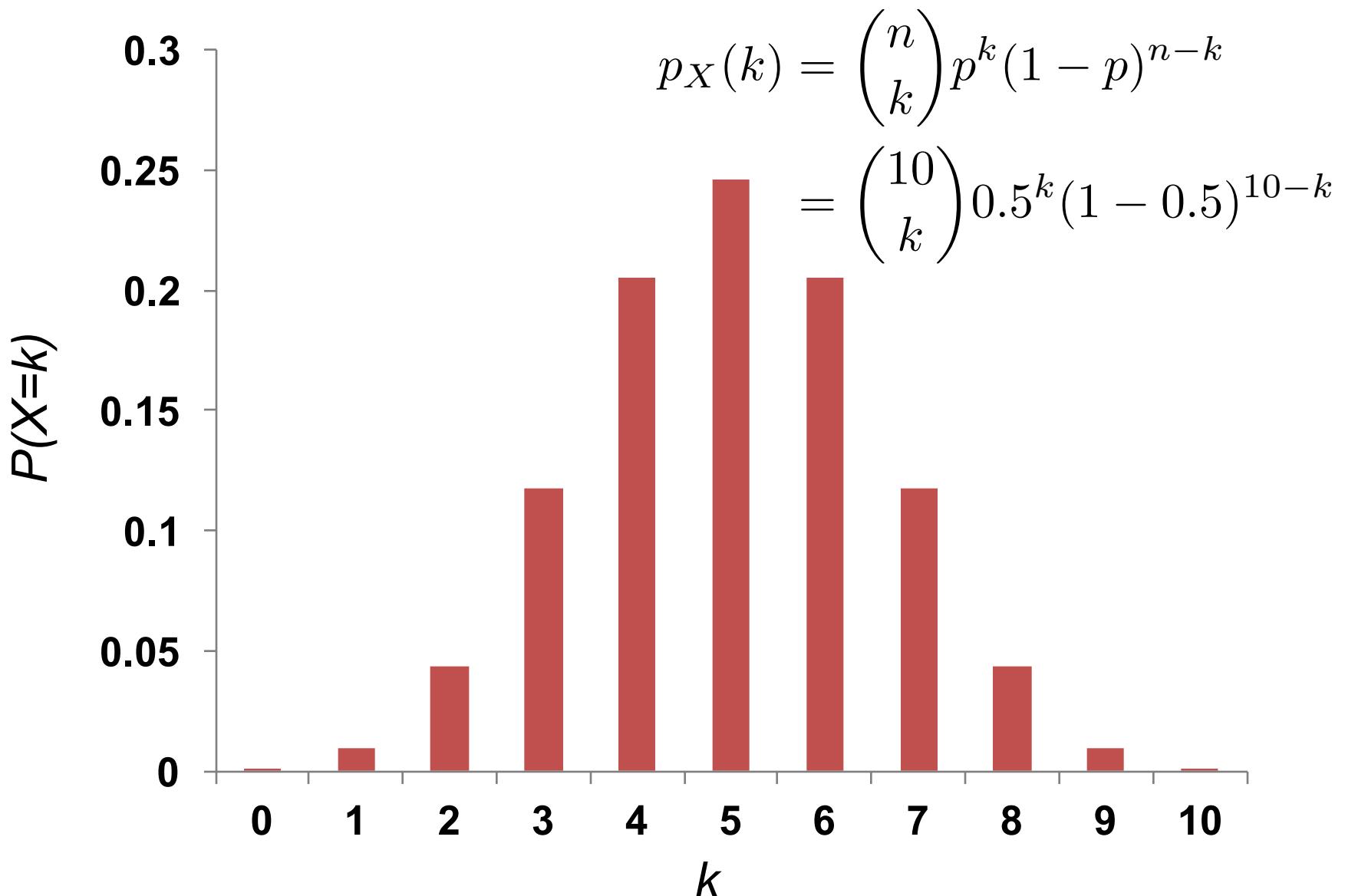
- Most common form:

$$\begin{aligned} P(F | E) &= P(EF) / P(E) \\ &= \frac{P(E | F) P(F)}{P(E)} \end{aligned}$$

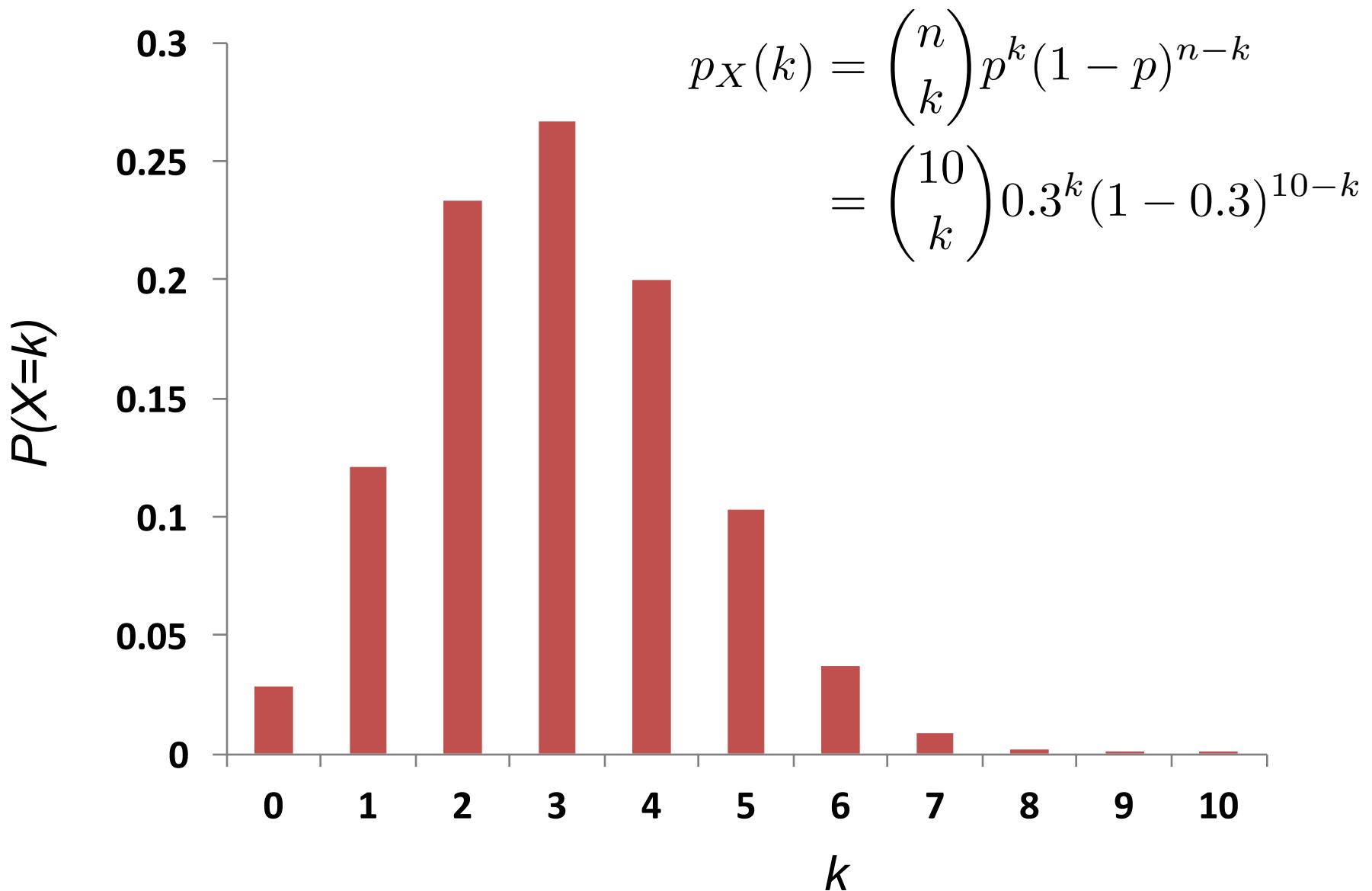
- Expanded form:

$$P(F | E) = \frac{P(E | F) P(F)}{P(E | F) P(F) + P(E | F^c) P(F^c)}$$

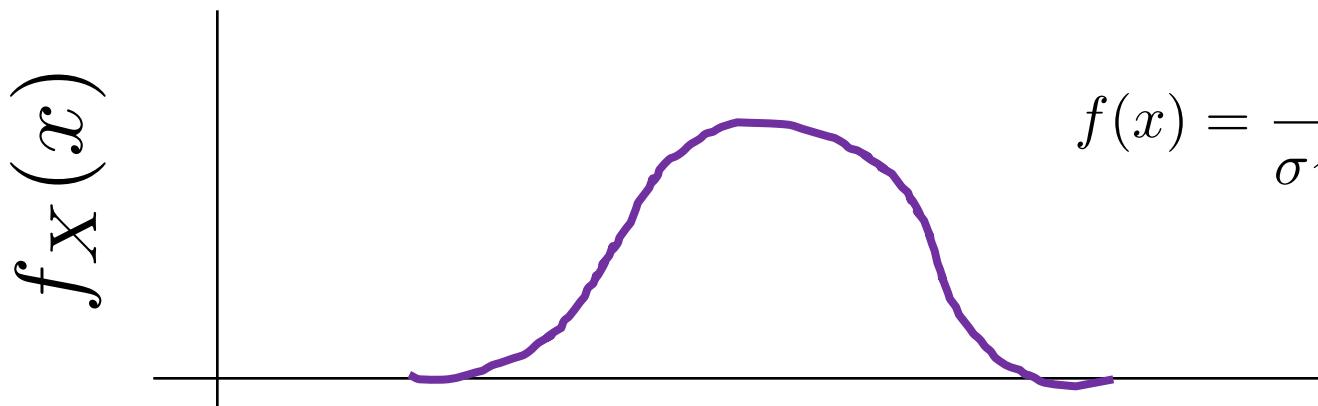
# PMF for $X \sim \text{Bin}(10, 0.5)$



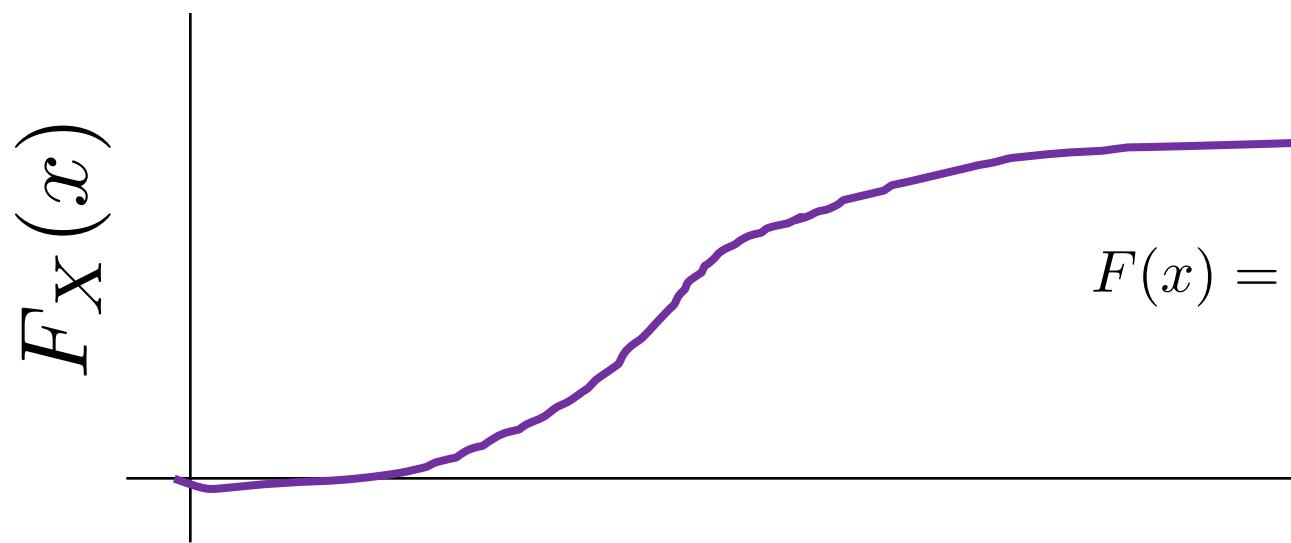
# PMF for $X \sim \text{Bin}(10, 0.3)$



# PDF and CDF of $X \sim N(4, 9)$



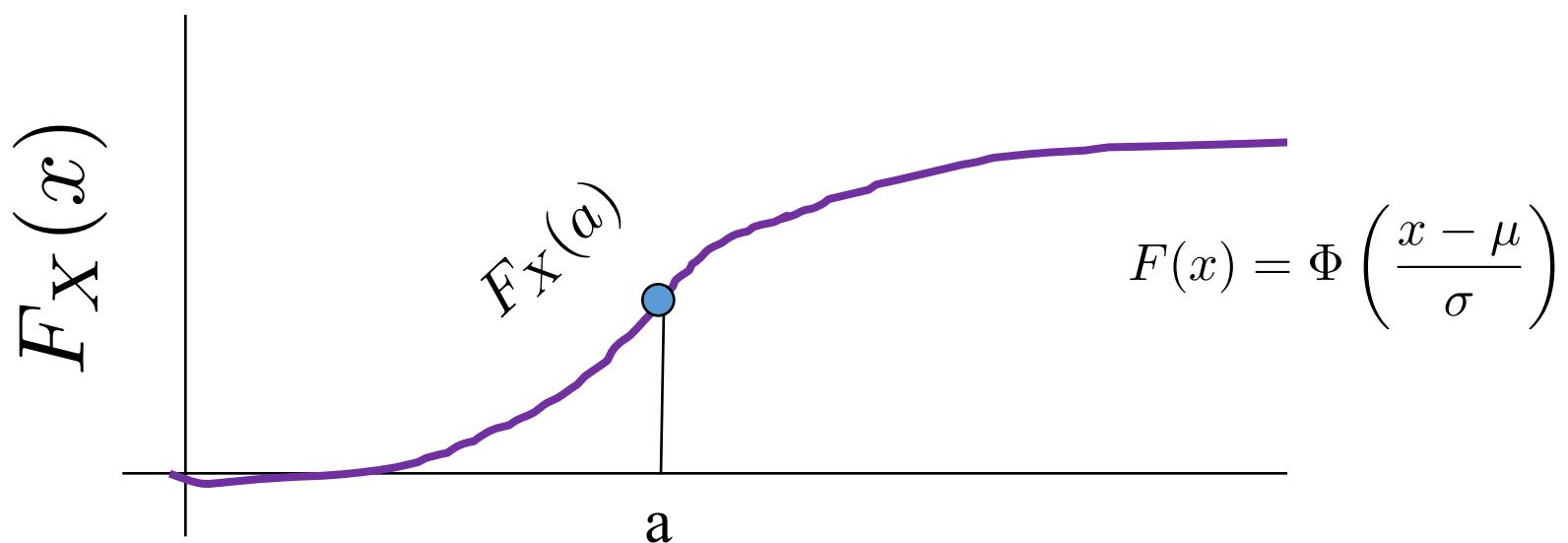
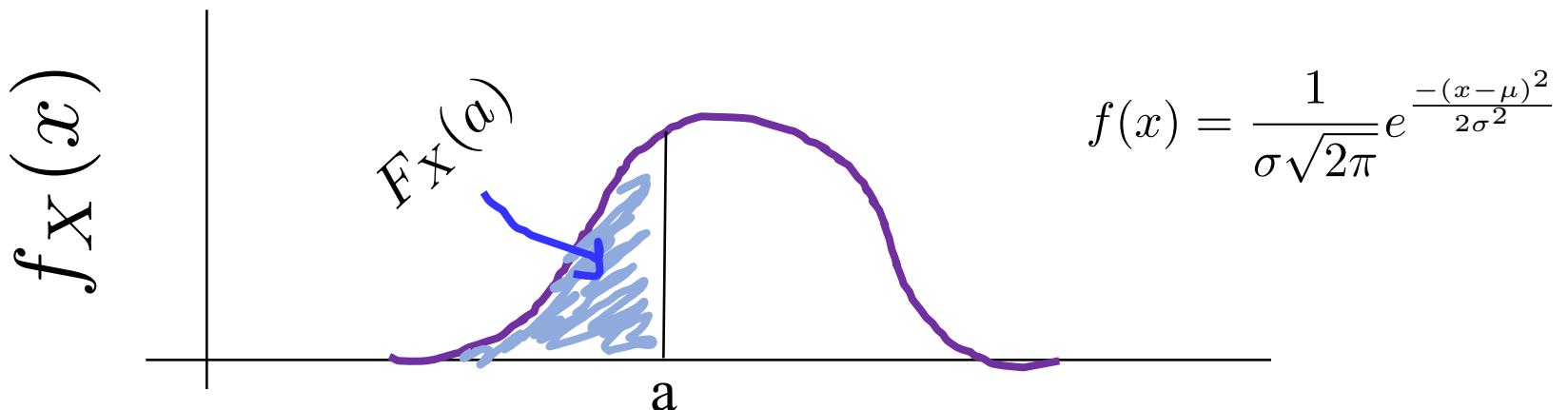
$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}$$



$$F(x) = \Phi\left(\frac{x - \mu}{\sigma}\right)$$

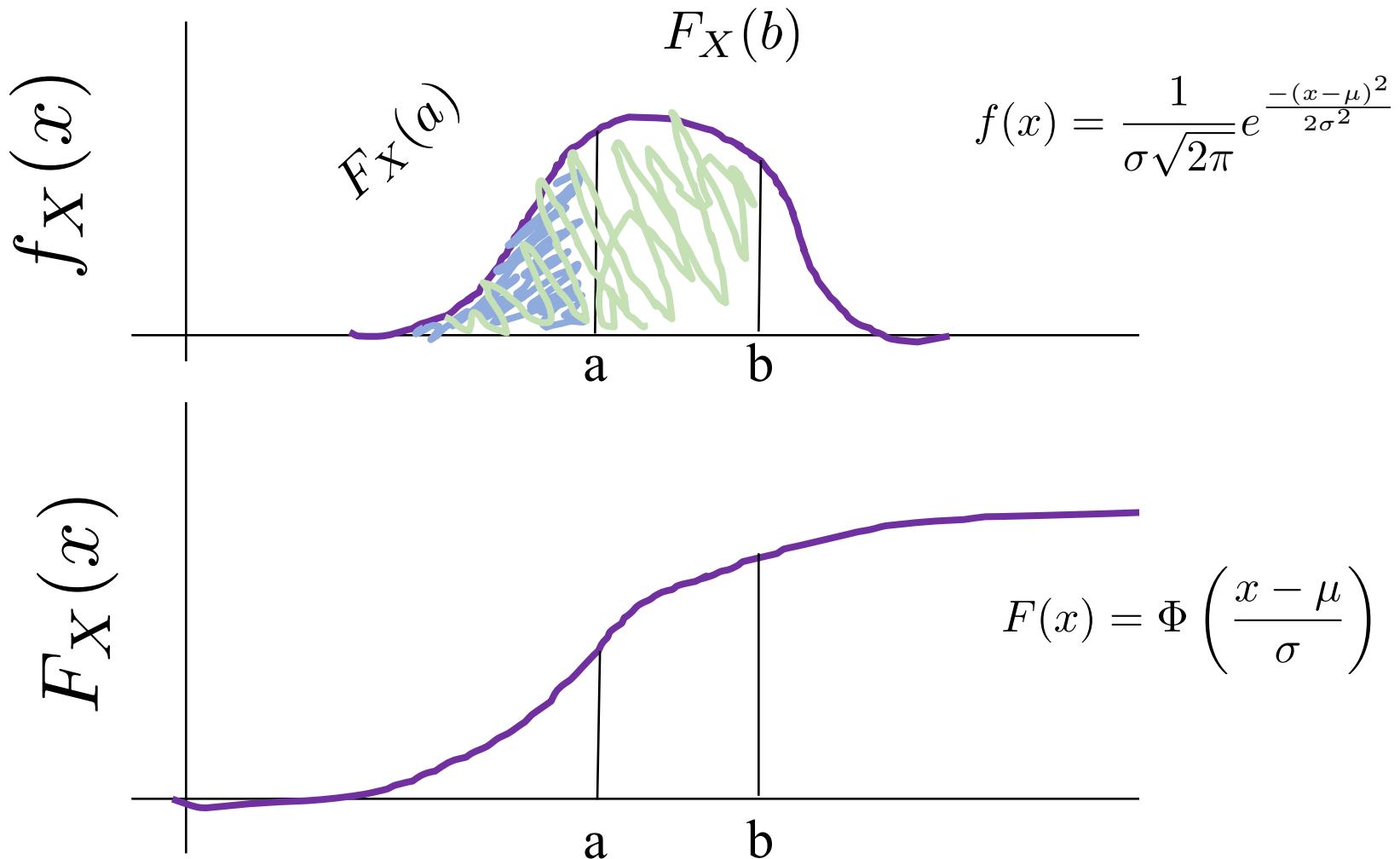
A CDF is the integral from  $-\infty$  to  $x$  of the PDF

# PDF and CDF of $X \sim N(4, 9)$



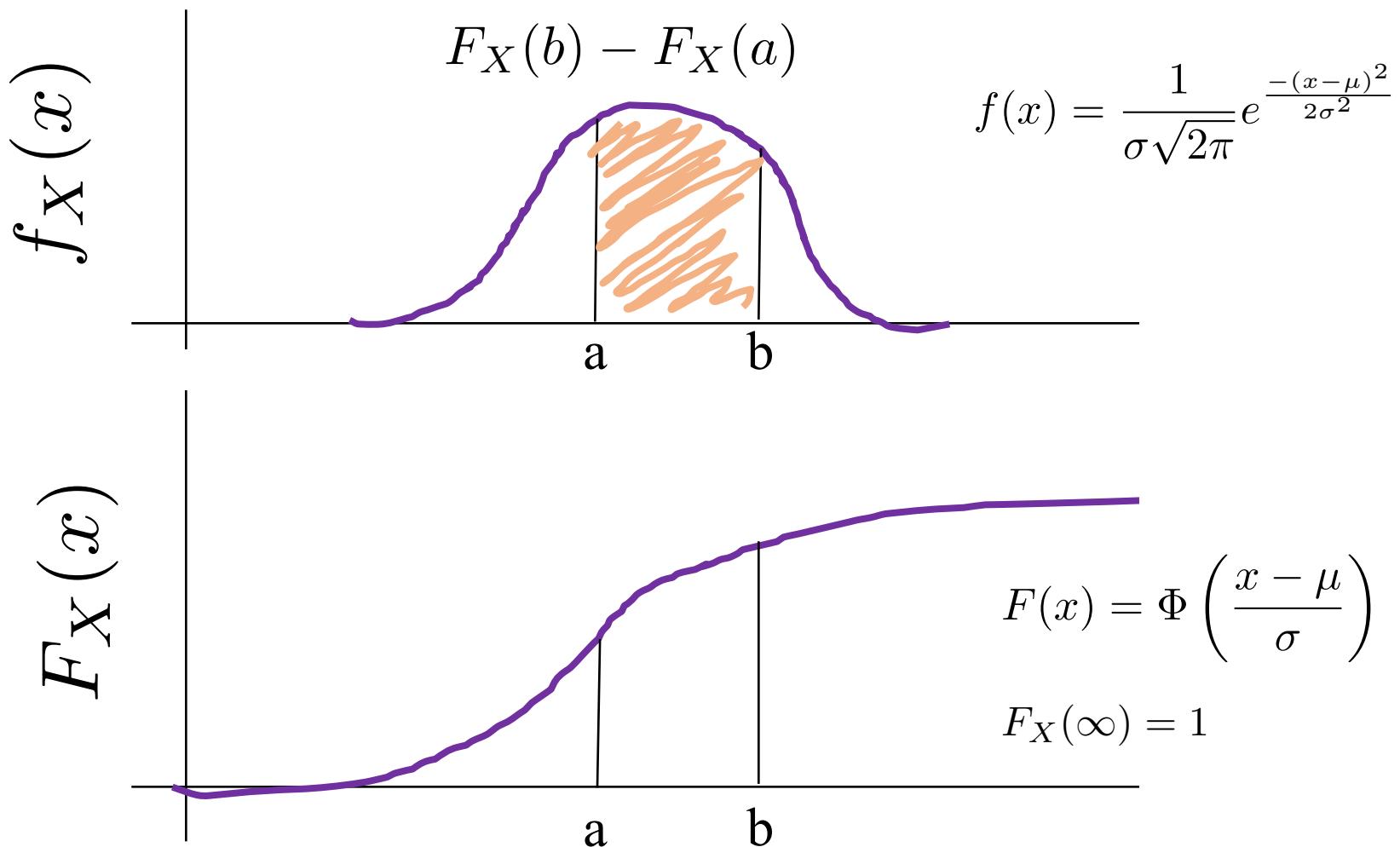
A CDF is the integral from  $-\infty$  to  $x$  of the PDF

# PDF and CDF of $X \sim N(4, 9)$



A CDF is the integral from  $-\infty$  to  $x$  of the PDF

# PDF and CDF of $X \sim N(4, 9)$



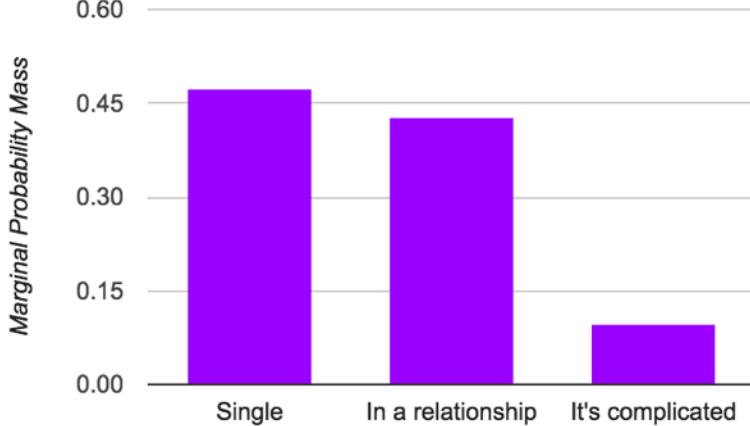
A CDF is the integral from  $-\infty$  to  $x$  of the PDF

F(-infinity)

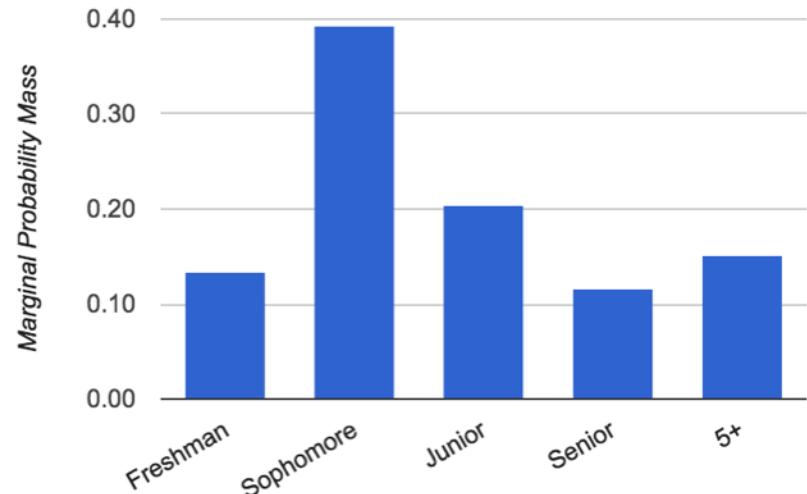
# Joint Probability Table

Joint Probability Table				
	Single	In a relationship	It's complicated	Marginal Year
Freshman	0.06	0.04	0.03	0.13
Sophomore	0.21	0.16	0.02	0.39
Junior	0.13	0.06	0.02	0.21
Senior	0.04	0.07	0.01	0.12
5+	0.04	0.09	0.03	0.15
<b>Marginal Status</b>	<b>0.47</b>	<b>0.43</b>	<b>0.10</b>	<b>1.00</b>

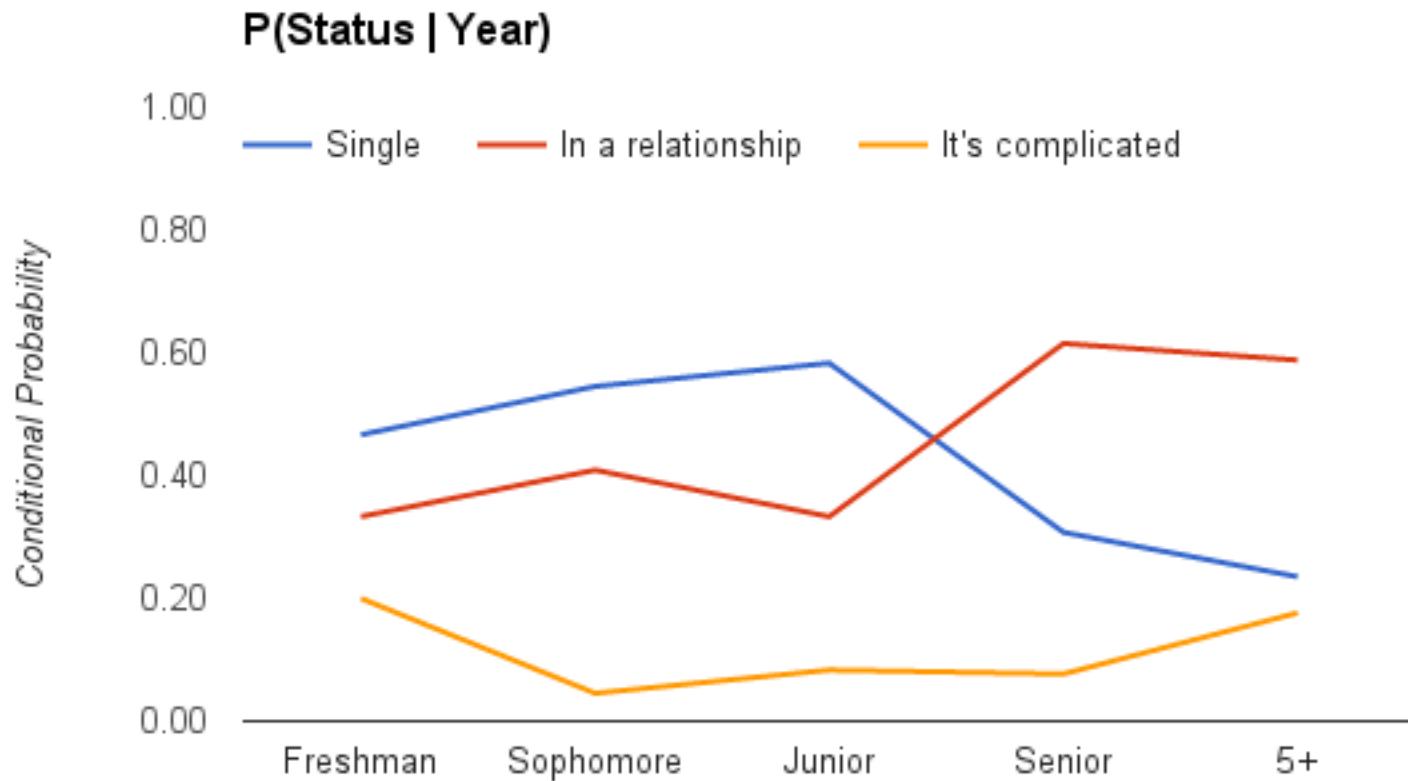
Marginal Status Probability



Marginal Year Probability



# Conditional Probability



# Joint Probability Density

- X and Y are continuous RVs with Joint PDF:

$$f_{X,Y}(x,y) = \begin{cases} \frac{12}{5}x(2-x-y) & \text{where } 0 < x, y < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$f_{X|Y}(x | y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$f_{X|N}(x | n) = \frac{p_{N|X}(n | x)f_X(x)}{p_N(n)}$$

Notation explosion!

## Probability Notation

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This handout maps between math notation used in CS109 and English. Note: “or” is not notation.

### Events and Sets

$E$ or $F$	Capital letters can denote events
$A$ or $B$	Sometimes they denote sets
$ E $ or $ A $	Size of an event or set
$E^C$ or $A^C$	Complement of an event or set
$EF$ or $AB$	Intersection of events or sets

⋮

$p_X(x)$	Probability mass function (PMF) of $X$
$p_{X,Y}(x,y)$	Joint probability mass function (PMF) of $X$ and $Y$
$p_{X Y}(x y)$	Conditional probability mass function (PMF) of $X$ given $Y$
$f_X(x)$	Probability density function (PDF) of $X$
$f_{X,Y}(x,y)$	Joint probability density function (PDF) of $X$ and $Y$
$f_{X Y}(x y)$	Conditional probability density function (PDF) of $X$ given $Y$
$F_X(x)$	Cumulative distribution function (CDF) of $X$
$F_{X,Y}(x,y)$	Joint cumulative distribution function (CDF) of $X$ and $Y$
$F_{X Y}(x y)$	Conditional cumulative distribution function (CDF) of $X$ given $Y$

⋮

# Flip a Coin With Unknown Probability

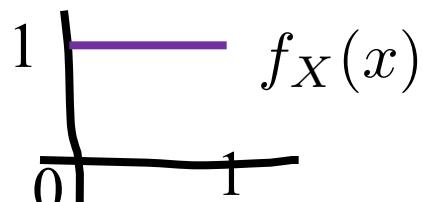
- Flip a coin  $(n + m)$  times, comes up with  $n$  heads
  - We don't know probability  $X$  that coin comes up heads
  - Our belief before flipping coins is that:  $X \sim \text{Uni}(0, 1)$
  - Let  $N$  = number of heads
  - Given  $X = x$ , coin flips independent:  $(N | X) \sim \text{Bin}(n + m, x)$

$$f_{X|N}(x|n) = \frac{P(N = n | X = x) f_X(x)}{P(N = n)}$$

Binomial

$$= \frac{\binom{n+m}{n} x^n (1-x)^m}{P(N = n)}$$

Constant

$$= \frac{\binom{n+m}{n}}{P(N = n)} x^n (1-x)^m$$
$$= \frac{1}{c} \cdot x^n (1-x)^m \quad \text{where } c = \int_0^1 x^n (1-x)^m dx$$


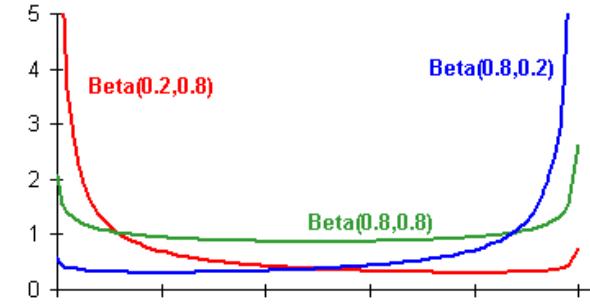
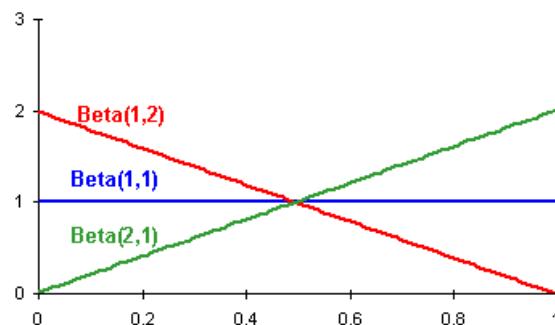
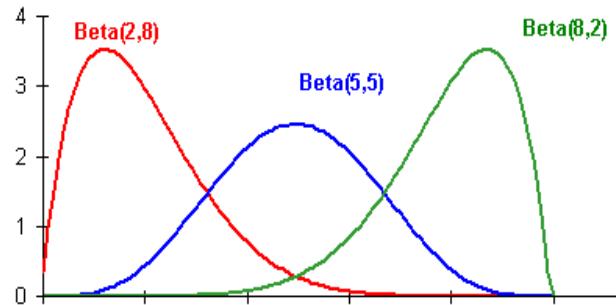
A graph showing the probability density function  $f_X(x)$  for a uniform distribution over the interval [0, 1]. The x-axis is labeled 'x' and ranges from 0 to 1. The y-axis is labeled 'f\_X(x)' and has a value of 1 at x=0. The function is a constant horizontal line at y=1 from x=0 to x=1, and then drops back to 0.

Move terms around

# Beta Random Variable

- $X$  is a **Beta Random Variable**:  $X \sim \text{Beta}(a, b)$ 
  - Probability Density Function (PDF): (where  $a, b > 0$ )

$$f(x) = \begin{cases} \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1} & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases} \quad \text{where } B(a,b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$$



- Symmetric when  $a = b$

$$\bullet E[X] = \frac{a}{a+b}$$

$$Var(X) = \frac{ab}{(a+b)^2(a+b+1)}$$

# Random Variable for $p$

No flips

12 flips, 8 heads

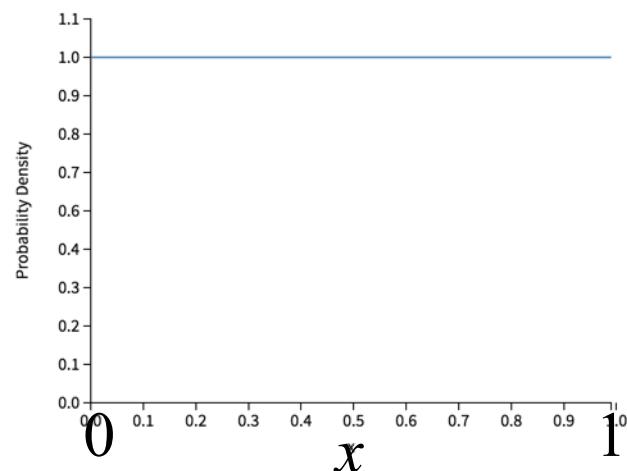
199 flips, 99 heads

$f_X|(0 \text{ heads, 0 tails})$

$f_X|(8 \text{ heads, 4 tails})$

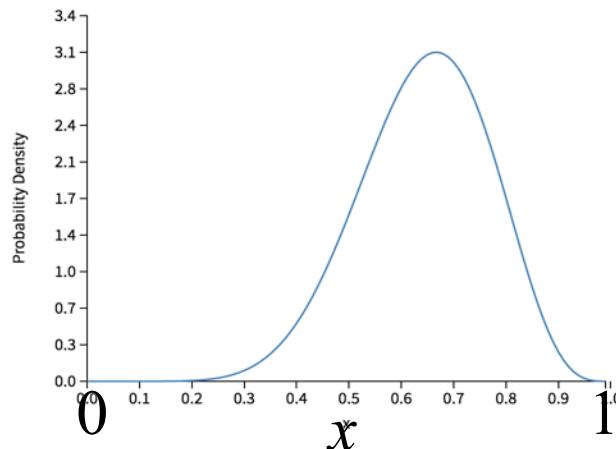
$f_X|(100 \text{ heads, 101 tails})$

Beta PDF



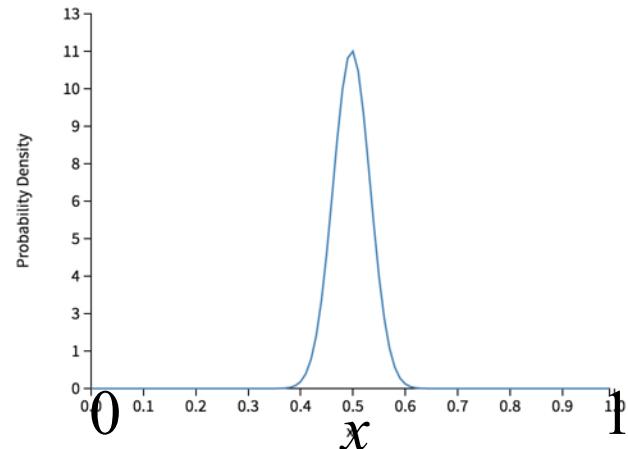
$$X \sim \text{Beta}(1,1)$$

Beta PDF



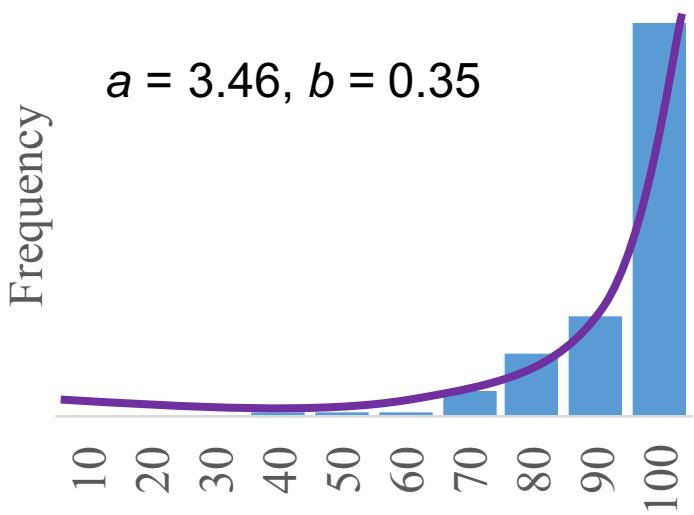
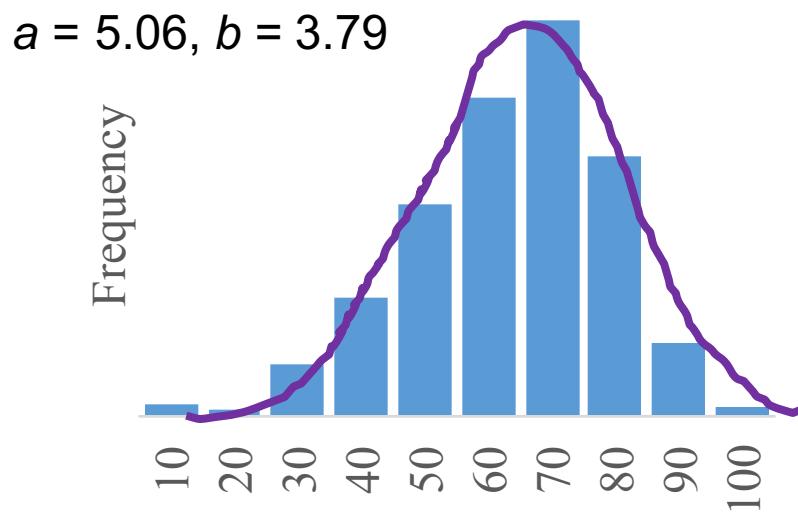
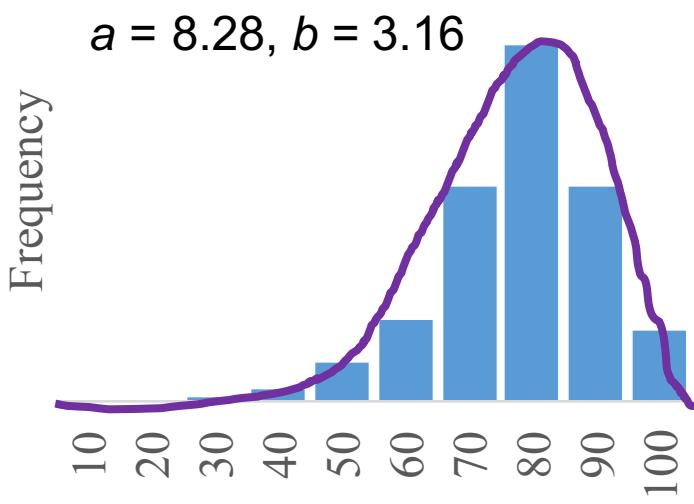
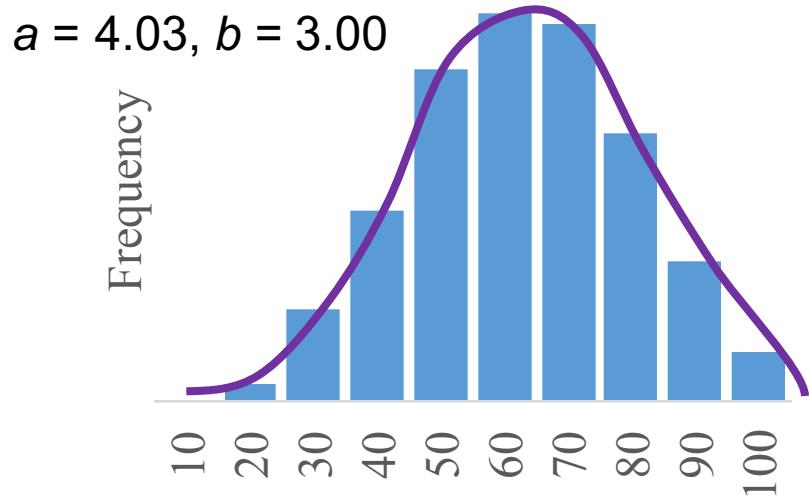
$$X \sim \text{Beta}(9,5)$$

Beta PDF



$$X \sim \text{Beta}(101,102)$$

# Assignment Grades



We have 2055 assignment distributions from grade scope

End Review

Expectation -> Covariance

# Expected Values of Sums

Big deal lemma: first  
stated without proof

$$E[X + Y] = E[X] + E[Y]$$

Generalized:  $E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$

Holds regardless of dependency between  $X_i$ 's



# Bool Was Cool

- Let  $E_1, E_2, \dots, E_n$  be events with indicator RVs  $X_i$ 
  - If event  $E_i$  occurs, then  $X_i = 1$ , else  $X_i = 0$
  - Recall  $E[X_i] = P(E_i)$
  - Why?

$$E[X_i] = 0 \cdot (1 - P(E_i)) + 1 \cdot P(E_i)$$

Bernoulli aka Indicator Random Variables were studied extensively by George Bool

Bool died of being too cool

# Expectation of Binomial

- Let  $Y \sim \text{Bin}(n, p)$ 
  - $n$  independent trials
  - Let  $X_i = 1$  if  $i$ -th trial is “success”, 0 otherwise
  - $X_i \sim \text{Ber}(p) \quad E[X_i] = p$

$$Y = X_1 + X_2 + \cdots + X_n = \sum_{i=1}^n X_i$$

$$E[Y] = E\left[\sum_{i=1}^n X_i\right]$$

$$= \sum_{i=1}^n E[X_i]$$

$$= E[X_1] + E[X_2] + \dots + E[X_n]$$

$$= np$$

# Expectation of Negative Binomial

- Let  $Y \sim \text{NegBin}(r, p)$ 
  - Recall  $Y$  is number of trials until  $r$  “successes”
  - Let  $X_i = \# \text{ of trials to get success after } (i - 1)\text{st success}$
  - $X_i \sim \text{Geo}(p)$  (i.e., Geometric RV)  $E[X_i] = \frac{1}{p}$

$$Y = X_1 + X_2 + \cdots + X_r = \sum_{i=1}^r X_i$$

$$E[Y] = E\left[\sum_{i=1}^r X_i\right]$$

$$= \sum_{i=1}^r E[X_i]$$

$$= E[X_1] + E[X_2] + \dots + E[X_r]$$

$$= \frac{r}{p}$$

# Hash Tables (aka Toy Collecting)

- Consider a hash table with  $n$  buckets
  - Each string equally likely to get hashed into any bucket
  - Let  $X = \#$  strings to hash until each bucket  $\geq 1$  string
  - What is  $E[X]$ ?
  - Let  $X_i = \#$  of trials to get success after  $i$ -th success
    - where “success” is hashing string to previously empty bucket
    - After  $i$  buckets have  $\geq 1$  string, probability of hashing a string to an empty bucket is  $p = (n - i) / n$
    - $P(X_i = k) = \frac{n-i}{n} \left(\frac{i}{n}\right)^{k-1}$  equivalently:  $X_i \sim \text{Geo}((n - i) / n)$
    - $E[X_i] = 1 / p = n / (n - i)$
  - $X = X_0 + X_1 + \dots + X_{n-1} \Rightarrow E[X] = E[X_0] + E[X_1] + \dots + E[X_{n-1}]$ 
$$E[X] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1} = n \left[ \frac{1}{n} + \frac{1}{n-1} + \dots + 1 \right] = O(n \log n)$$

This is your final answer

# Let's Do Some Sorting!

5	3	7	4	8	6	2	1
---	---	---	---	---	---	---	---

# QuickSort

5	3	7	4	8	6	2	1
---	---	---	---	---	---	---	---



select  
“pivot”

# Recursive Insight

5	3	7	4	8	6	2	1
---	---	---	---	---	---	---	---

Partition array so:

- everything smaller than pivot is on left
- everything greater than or equal to pivot is on right
- pivot is in-between

# Recursive Insight



Partition array so:

- everything smaller than pivot is on left
- everything greater than or equal to pivot is on right
- pivot is in-between

# Recursive Insight



Now recursive sort “red” sub-array

# Recursive Insight



Now recursive sort “red” sub-array

# Recursive Insight



Now recursive sort “red” sub-array

Then, recursive sort “blue” sub-array

# Recursive Insight



Now recursive sort “red” sub-array

Then, recursive sort “blue” sub-array

# Recursive Insight

1	2	3	4	5	6	7	8
---	---	---	---	---	---	---	---

Everything is sorted!

```
void Quicksort(int arr[], int n)
{
    if (n < 2) return;

    int boundary = Partition(arr, n);

    // Sort subarray up to pivot
    Quicksort(arr, boundary);

    // Sort subarray after pivot to end
    Quicksort(arr + boundary + 1, n - boundary - 1);
}
```

“boundary” is the index of the pivot

This is equal to the number of elements before pivot

```
int Partition(int arr[], int n)
{
    int lh = 1, rh = n - 1;

    int pivot = arr[0];
    while (true) {
        while (lh < rh && arr[rh] >= pivot) rh--;
        while (lh < rh && arr[lh] < pivot) lh++;
        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}
```

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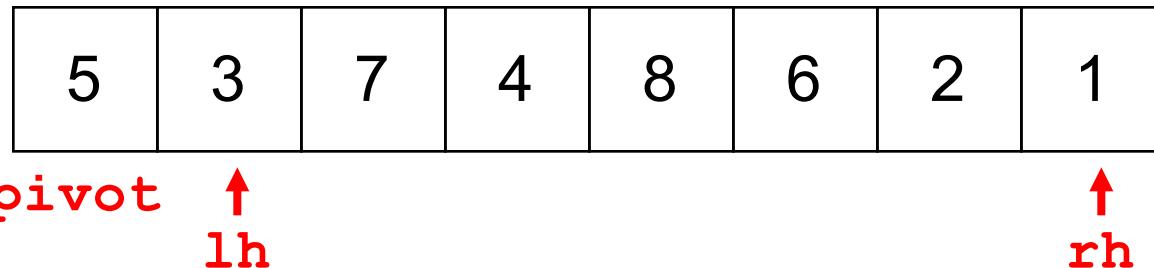
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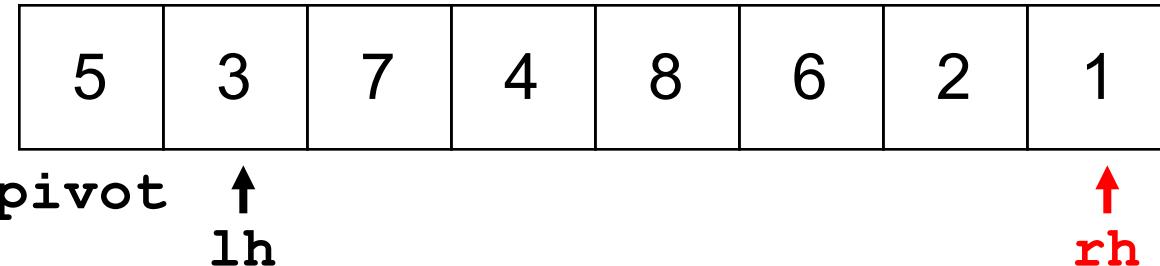


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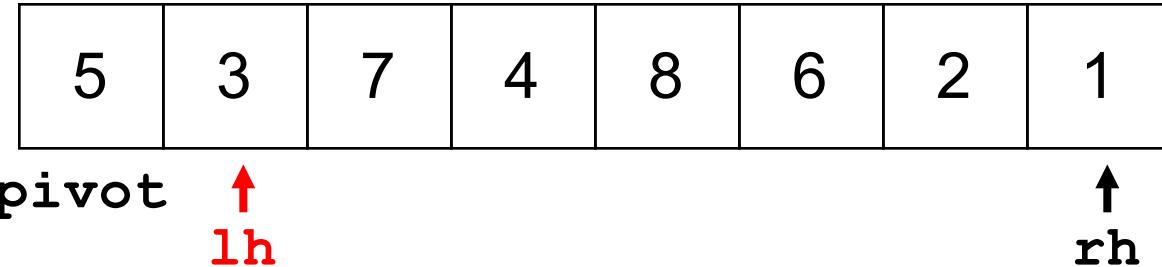


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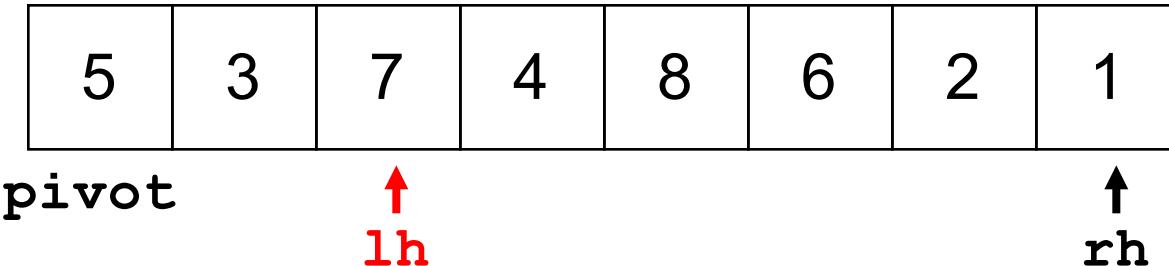


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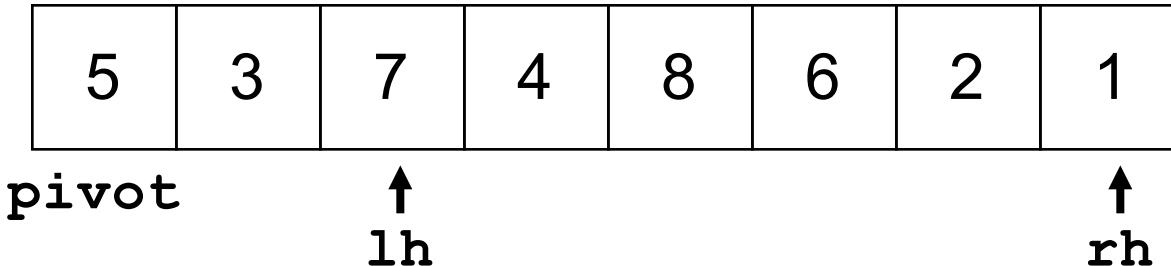
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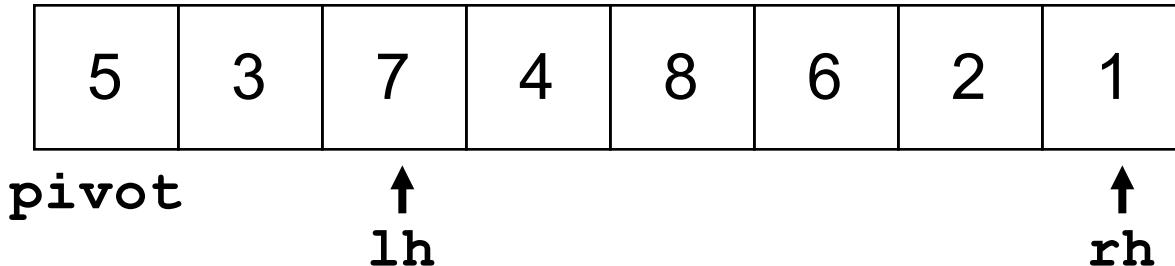


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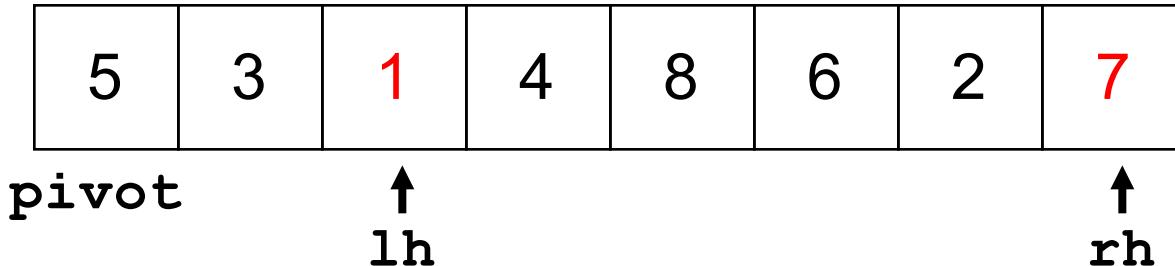


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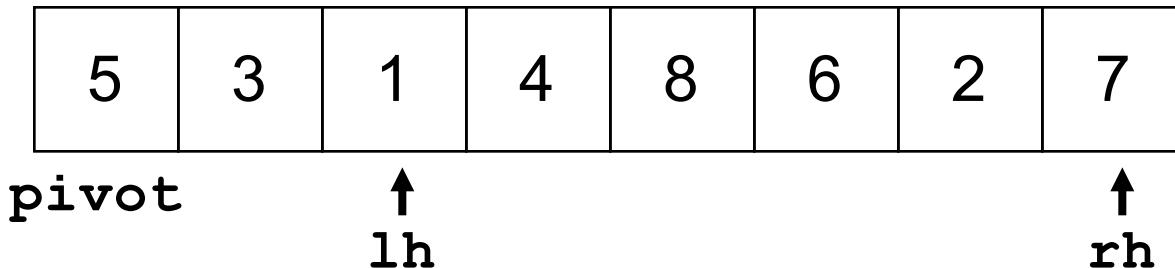


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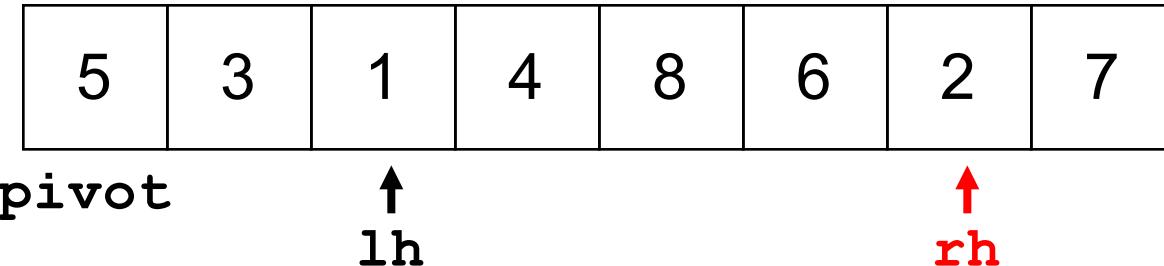


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}

```

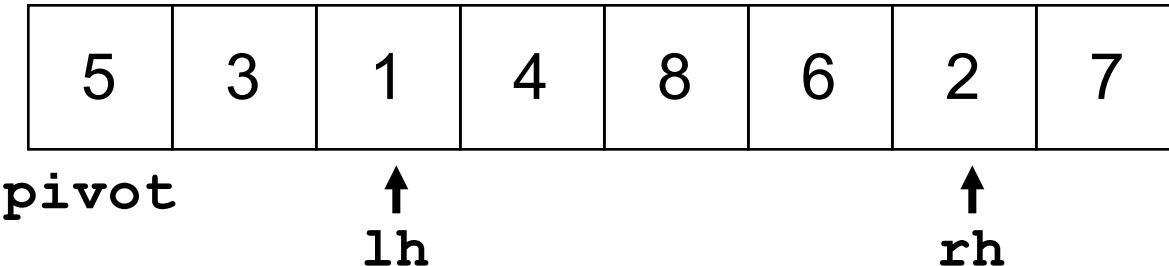


```

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    while (true) {
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        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}

```

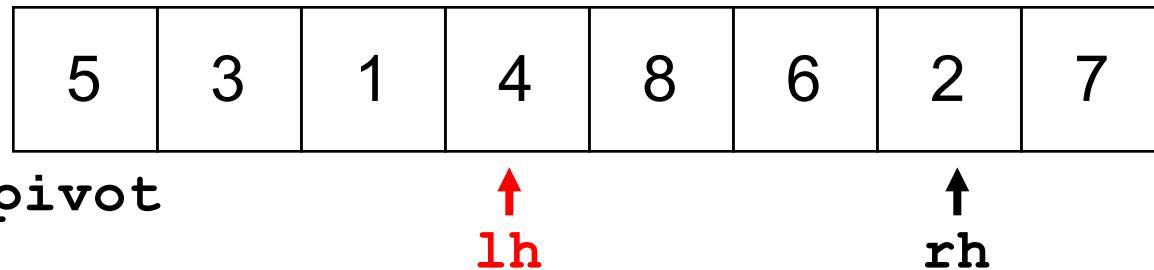


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    }
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}

```

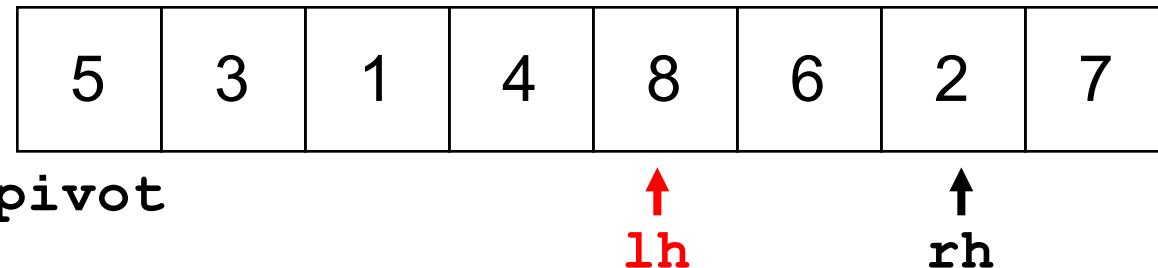


```

int Partition(int arr[], int n)
{
    int lh = 1, rh = n - 1;

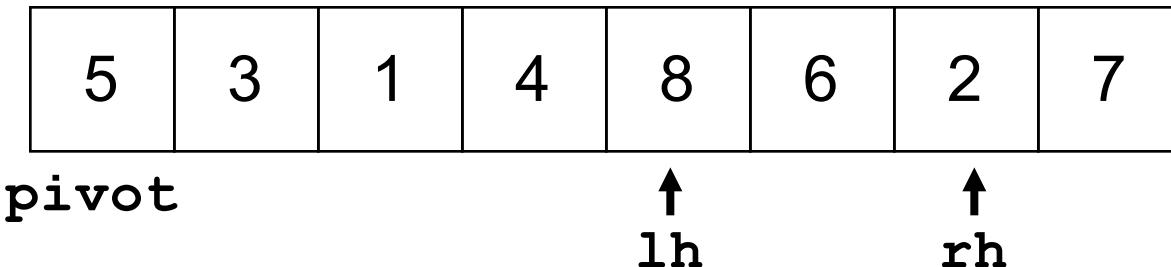
    int pivot = arr[0];
    while (true) {
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        Swap(arr[lh], arr[rh]);
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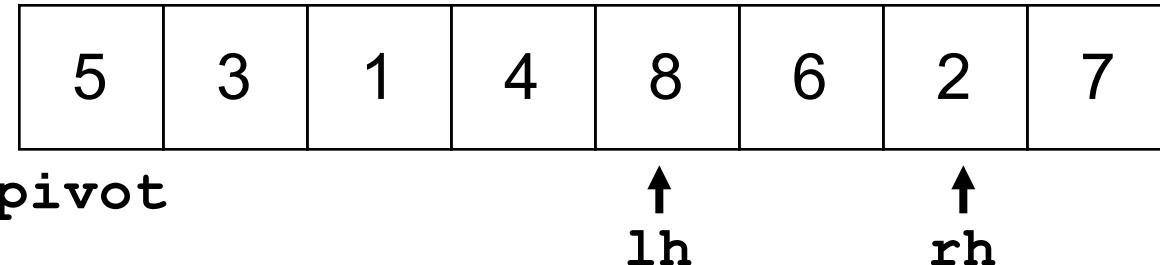


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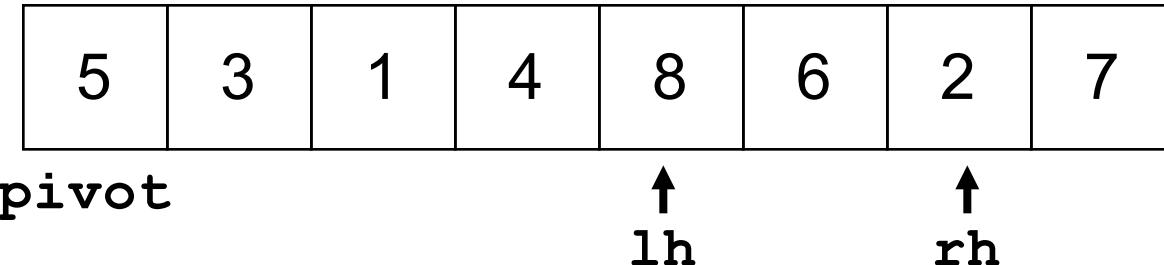


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    }
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    return lh;
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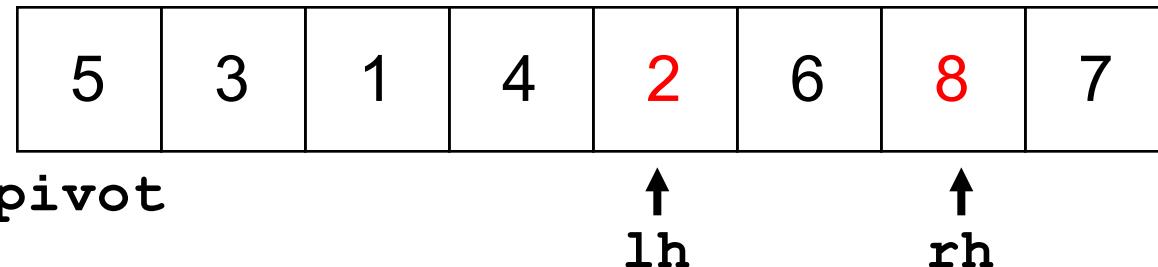


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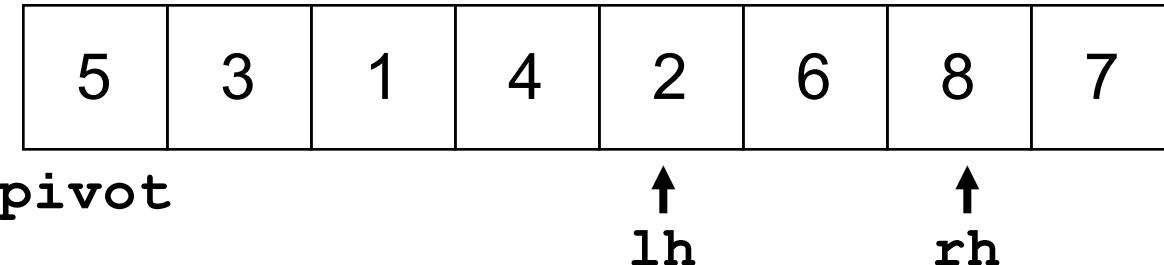


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    }
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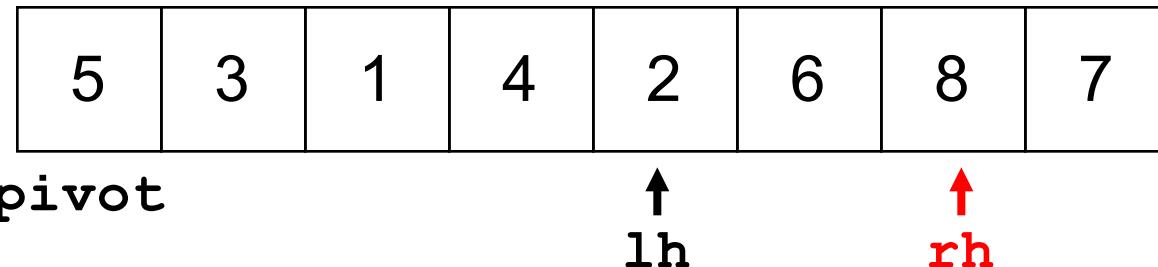


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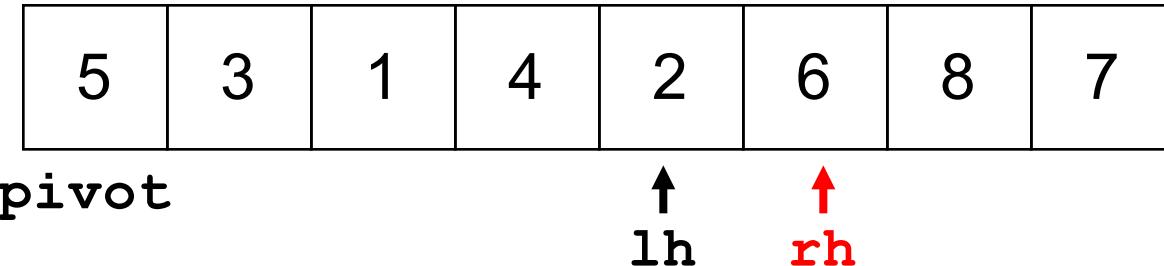


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        if (lh == rh) break;
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        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}

```

5	3	1	4	2	6	8	7
---	---	---	---	---	---	---	---

pivot

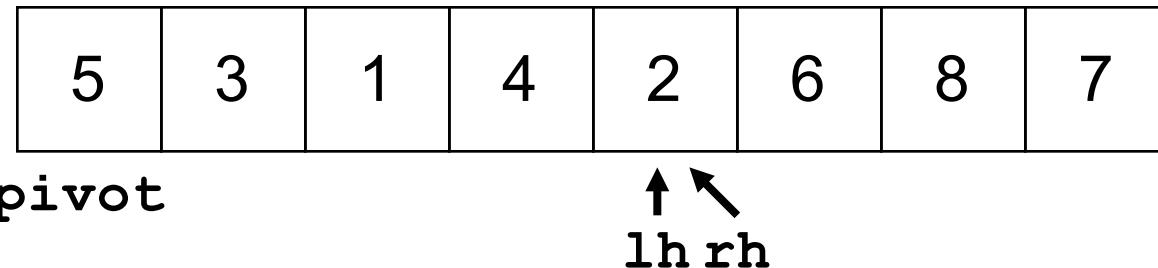
↑  
lh rh

```

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{
    int lh = 1, rh = n - 1;

    int pivot = arr[0];
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        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}

```

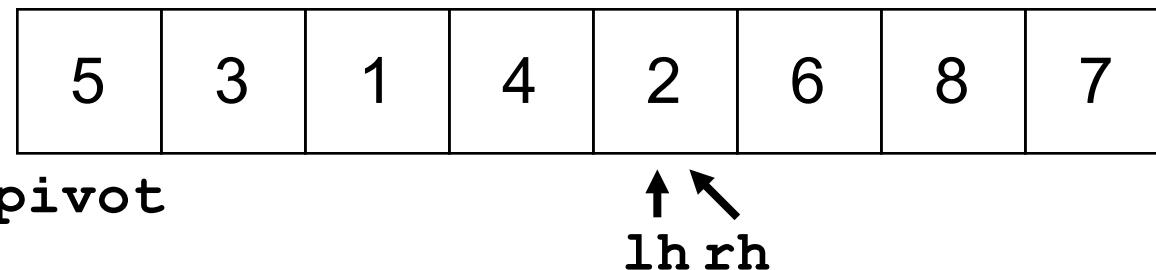


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    int pivot = arr[0];
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```

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---	---	---	---	---	---	---	---

pivot

↑ ↗  
lh rh

```

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    return lh;
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```

5	3	1	4	2	6	8	7
---	---	---	---	---	---	---	---

pivot

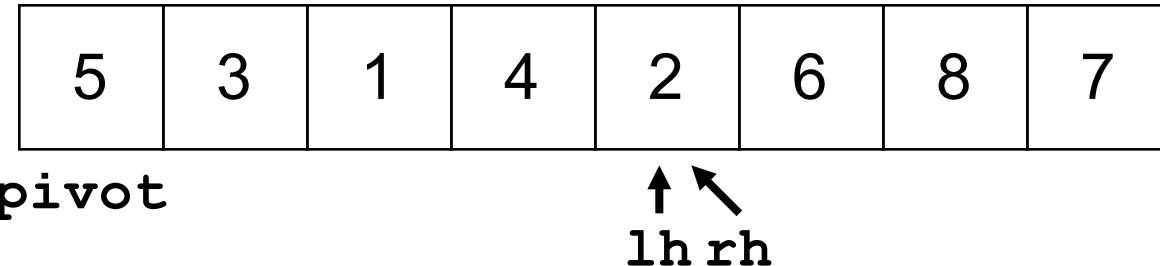
↑ ↗  
lh rh

```

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{
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        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}

```

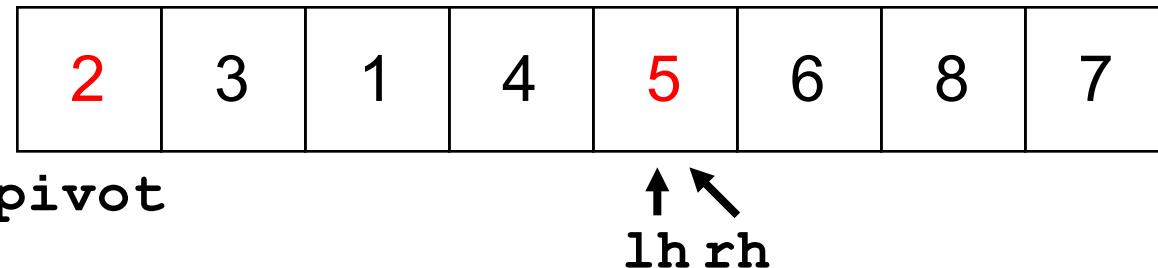


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    }
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    return lh;
}

```

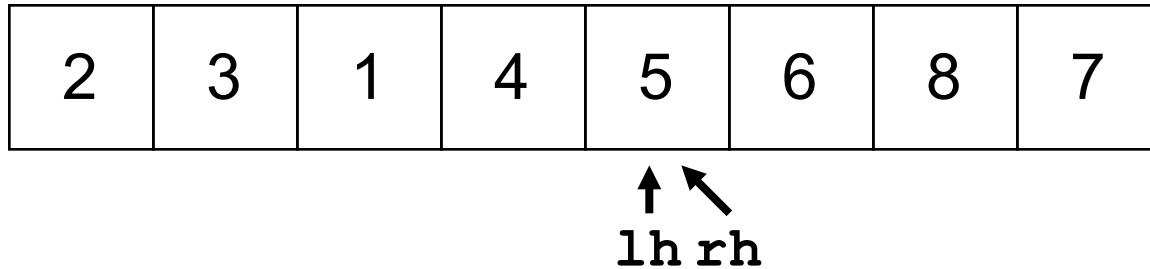


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    return lh;
}

```

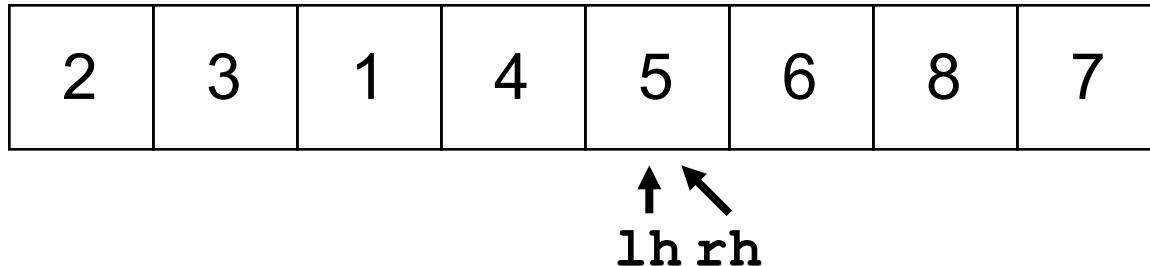


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        while (lh < rh && arr[lh] < pivot) lh++;
        if (lh == rh) break;
        Swap(arr[lh], arr[rh]);
    }
    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;      Returns 4 (index of pivot)
}

```



```
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{
    int lh = 1, rh = n - 1;

    int pivot = arr[0];
    while (true) {
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    if (arr[lh] >= pivot) return 0;
    Swap(arr[0], arr[lh]);
    return lh;
}
```

- Complexity of algorithm determined by number of comparisons made to pivot

# Complexity QuickSort

- QuickSort is  $O(n \log n)$ , where  $n = \# \text{ elems to sort}$ 
  - But in “worst case” it can be  $O(n^2)$
  - Worst case occurs when every time pivot is selected, it is maximal or minimal remaining element
- What is  $P(\text{QuickSort worst case})$ ?
  - On each recursive call, pivot = max/min element, so we are left with  $n - 1$  elements for next recursive call
  - 2 possible “bad” pivots (max/min) on each recursive call

$$P(\text{Worst case}) = \frac{2}{n} \cdot \frac{2}{n-1} \cdot \dots \cdot \frac{2}{2} = \frac{2^{n-1}}{n!}$$

- Saw similar behavior for BSTs on problem set #1
  - $P(\text{Worst case})$  gets small very fast as  $n$  grows!

# Expected Running Time of QuickSort

- Let  $X = \#$  comparisons made when sorting  $n$  elems
  - $E[X]$  gives us expected running time of algorithm
  - Given  $V_1, V_2, \dots, V_n$  in random order to sort
  - Let  $Y_1, Y_2, \dots, Y_n$  be  $V_1, V_2, \dots, V_n$  in sorted order
  - Let  $I_{a,b} = 1$  if  $Y_a$  and  $Y_b$  are compared, 0 otherwise
  - Order where  $Y_b > Y_a$ , so we have:  $X = \sum_{a=1}^{n-1} \sum_{b=a+1}^n I_{a,b}$

# Expected Running Time of QuickSort

Aside:

$$X = \sum_{a=1}^{n-1} \sum_{b=a+1}^n I_{a,b}$$

---

When  $a = 1$

$$I_{1,2} + I_{1,3} + \dots + I_{1,n}$$

When  $a = 2$

$$+ I_{2,3} + \dots + I_{2,n}$$

When  $a = n-1$

$$+ I_{n-1,n}$$

Contains a comparison between each  $i$  and  $j$   
(where  $i$  does not equal  $j$ )  
exactly once

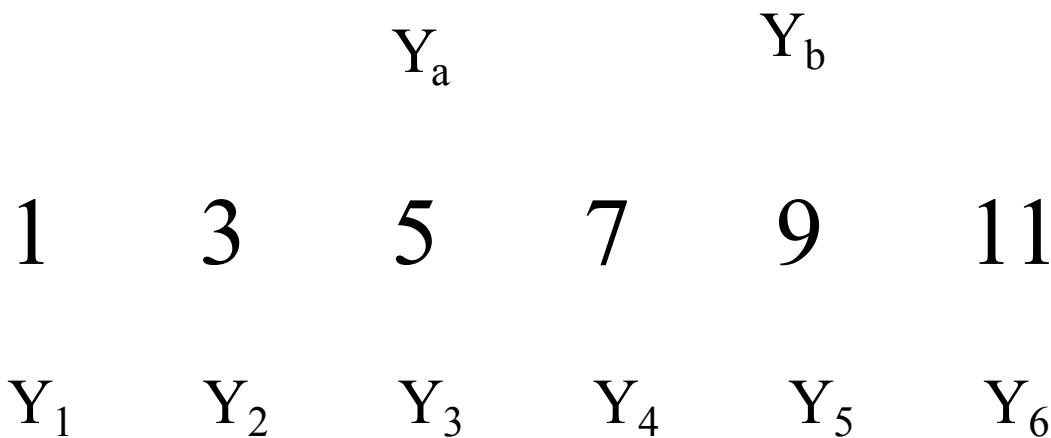
# Expected Running Time of QuickSort

- Let  $X = \#$  comparisons made when sorting  $n$  elems
  - $E[X]$  gives us expected running time of algorithm
  - Given  $V_1, V_2, \dots, V_n$  in random order to sort
  - Let  $Y_1, Y_2, \dots, Y_n$  be  $V_1, V_2, \dots, V_n$  in sorted order
  - Let  $I_{a,b} = 1$  if  $Y_a$  and  $Y_b$  are compared, 0 otherwise
  - Order where  $Y_b > Y_a$ , so we have:  $X = \sum_{a=1}^{n-1} \sum_{b=a+1}^n I_{a,b}$

$$E[X] = E\left[\sum_{a=1}^{n-1} \sum_{b=a+1}^n I_{a,b}\right] = \sum_{a=1}^{n-1} \sum_{b=a+1}^n E[I_{a,b}] = \sum_{a=1}^{n-1} \sum_{b=a+1}^n P(Y_a \text{ and } Y_b \text{ ever compared})$$

What is the probability  
that  $Y_a$  and  $Y_b$  are compared?

# Lets Imagine Our Array in Sorted Order



Whether or not they are compared  
depends on pivot choice

# Lets Imagine Our Array in Sorted Order

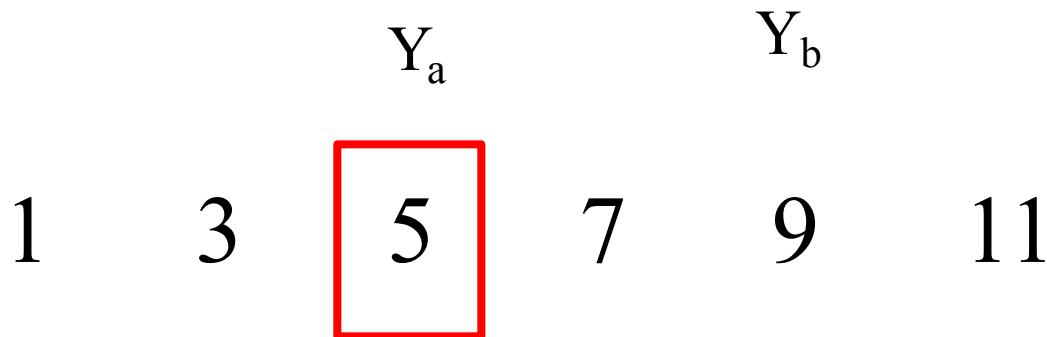
$Y_a$

$Y_b$

1      3      5      7      9      11

Whether or not they are compared  
depends on pivot choice

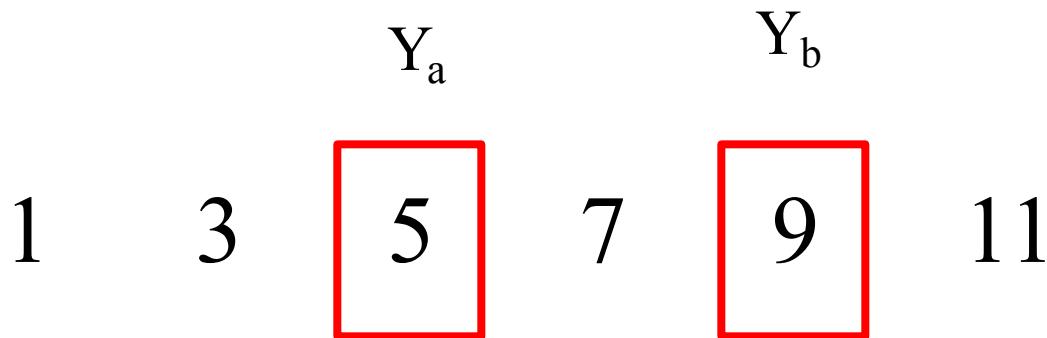
# $P(Y_a$ and $Y_b$ ever compared)



Consider pivot choice:  $Y_a$

They are compared

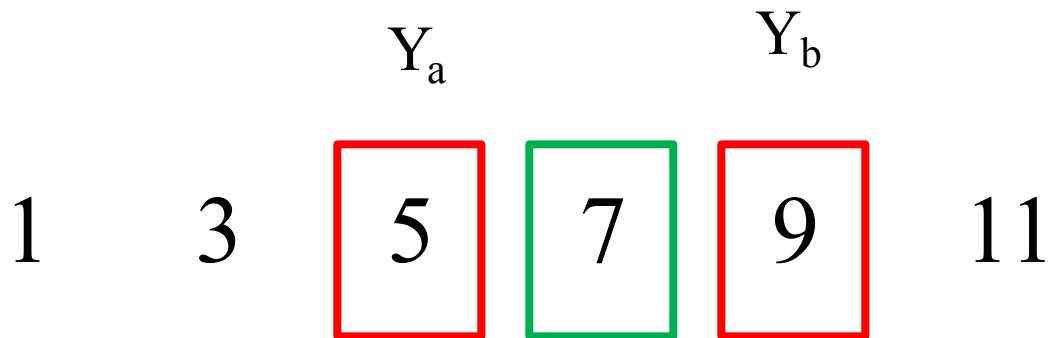
# $P(Y_a \text{ and } Y_b \text{ ever compared})$



Consider pivot choice:  $Y_b$

They are compared

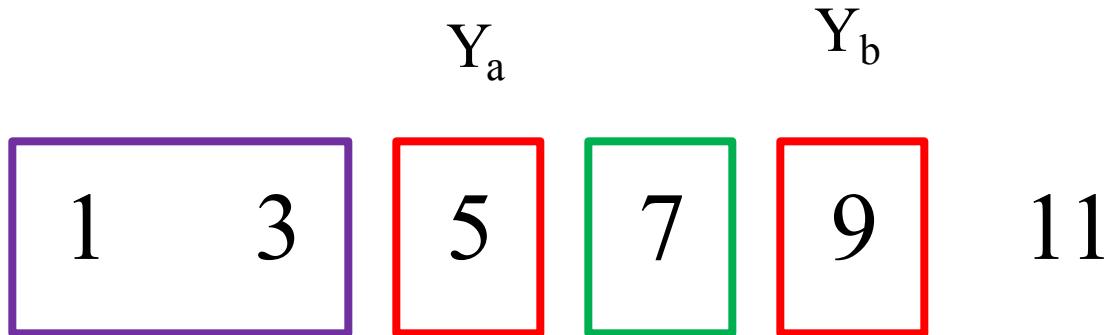
# $P(Y_a \text{ and } Y_b \text{ ever compared})$



Consider pivot choice: 7

They are **not** compared

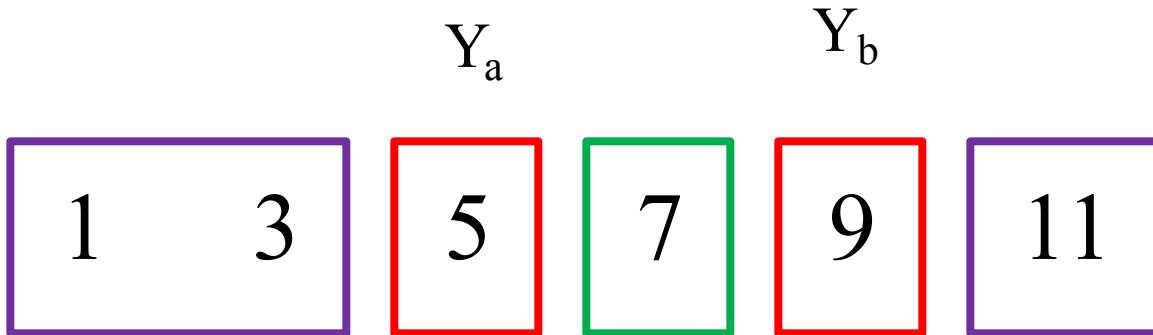
# $P(Y_a \text{ and } Y_b \text{ ever compared})$



Consider pivot choice:  $< Y_a$

Whether or not they are compared  
depends on future pivots

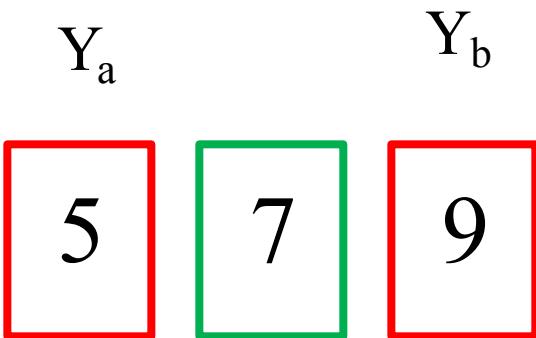
# $P(Y_a \text{ and } Y_b \text{ ever compared})$



Consider pivot choice:  $> Y_b$

Whether or not they are compared  
depends on future pivots

# $P(Y_a \text{ and } Y_b \text{ ever compared})$



Are  $Y_a$  and  $Y_b$  compared?

Keep repeating pivot choice until you get a pivot  
In the range  $[Y_a, Y_b]$  inclusive

# P(Y<sub>a</sub> and Y<sub>b</sub> ever compared)

- Consider when Y<sub>a</sub> and Y<sub>b</sub> are directly compared
  - We only care about case where pivot chosen from set: {Y<sub>a</sub>, Y<sub>a+1</sub>, Y<sub>a+2</sub>, ..., Y<sub>b</sub>}
  - From that set either Y<sub>a</sub> and Y<sub>b</sub> must be selected as pivot (with equal probability) in order to be compared
  - So,

$$P(Y_a \text{ and } Y_b \text{ ever compared}) = \frac{2}{b - a + 1}$$

$$E[X] = \sum_{a=1}^{n-1} \sum_{b=a+1}^n P(Y_a \text{ and } Y_b \text{ ever compared}) = \sum_{a=1}^{n-1} \sum_{b=a+1}^n \frac{2}{b - a + 1}$$

# Bring it on Home (i.e. Solve the Sum)

$$E[X] = \sum_{a=1}^{n-1} \sum_{b=a+1}^n \frac{2}{b-a+1}$$
$$\sum_{b=a+1}^n \frac{2}{b-a+1} \approx \int_{a+1}^n \frac{2}{b-a+1} db$$

Recall:  $\int \frac{1}{x} dx = \ln(x)$

$$= 2 \ln(b-a+1) \Big|_{a+1}^n = 2 \ln(n-a+1) - 2 \ln(2)$$
$$\approx 2 \ln(n-a+1) \text{ for large } n$$

$$E[X] \approx \sum_{a=1}^{n-1} 2 \ln(n-a+1) \approx 2 \int_{a=1}^{n-1} \ln(n-a+1) da$$

Let  $y = n-a+1$

$$= -2 \int_{y=n}^2 \ln(y) dy$$

Recall:  
 $\int \ln(x) dx = x \ln(x) - x$

$$= -2(y \ln(y) - y) \Big|_n^2$$

$$= -2[(2 \ln(2) - 2) - (n \ln(n) - n)] \approx 2n \ln(n) - 2n = O(n \log n)$$

Thanks  
Riemann

Ahhh ☺

# Variance from first principles

# Indicators: Now with pair-wise flavor!

- Recall  $I_i$  is indicator variable for event  $A_i$  when:

$$I_i = \begin{cases} 1 & \text{if } A_i \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$$

- Let  $X = \#$  of events that occur:  $X = \sum_{i=1}^n I_i$

$$E[X] = E\left[\sum_{i=1}^n I_i\right] = \sum_{i=1}^n E[I_i] = \sum_{i=1}^n P(A_i)$$

- Now consider pair of events  $A_i A_j$  occurring
  - $I_i I_j = 1$  if both events  $A_i$  and  $A_j$  occur, 0 otherwise
  - Number of pairs of events that occur is  $\binom{X}{2} = \sum_{i < j} I_i I_j$

I remember you!

# From Event Pairs to Variance

- Expected number of pairs of events:

$$E\left[\binom{X}{2}\right] = E\left[\sum_{i < j} I_i I_j\right] = \sum_{i < j} E[I_i I_j] = \sum_{i < j} P(A_i A_j)$$

$$E\left[\frac{X(X-1)}{2}\right] = \frac{1}{2}(E[X^2] - E[X]) = \sum_{i < j} P(A_i A_j)$$

$$E[X^2] - E[X] = 2 \sum_{i < j} P(A_i A_j) \Rightarrow E[X^2] = 2 \sum_{i < j} P(A_i A_j) + E[X]$$

- Recall:  $\text{Var}(X) = E[X^2] - (E[X])^2$

$$\text{Var}(X) = 2 \sum_{i < j} P(A_i A_j) + E[X] - (E[X])^2$$

$$= 2 \sum_{i < j} P(A_i A_j) + \sum_{i=1}^n P(A_i) - \left( \sum_{i=1}^n P(A_i) \right)^2$$

# Let's Try it With the Binomial

- $X \sim \text{Bin}(n, p)$   $E[X] = \sum_{i=1}^n P(A_i) = np$ 
  - Each trial:  $X_i \sim \text{Ber}(p)$   $E[X_i] = p$
  - Let event  $A_i = \text{trial } i \text{ is success}$  (i.e.,  $X_i = 1$ )

$$\sum_{i < j} P(A_i A_j) = \sum_{i < j} p^2 = \binom{n}{2} p^2 = \frac{n(n-1)}{2} p^2$$

$$\text{Var}(X) = 2 \sum_{i < j} P(A_i A_j) + E[X] - (E[X])^2$$

$$\text{Var}(X) = 2 \frac{n(n-1)}{2} p^2 + np - (np)^2$$

Substitute in for each term

$$\begin{aligned} &= n^2 p^2 - np^2 + np - n^2 p^2 \\ &= np - np^2 = np(1 - p) \end{aligned}$$

Expand and simplify

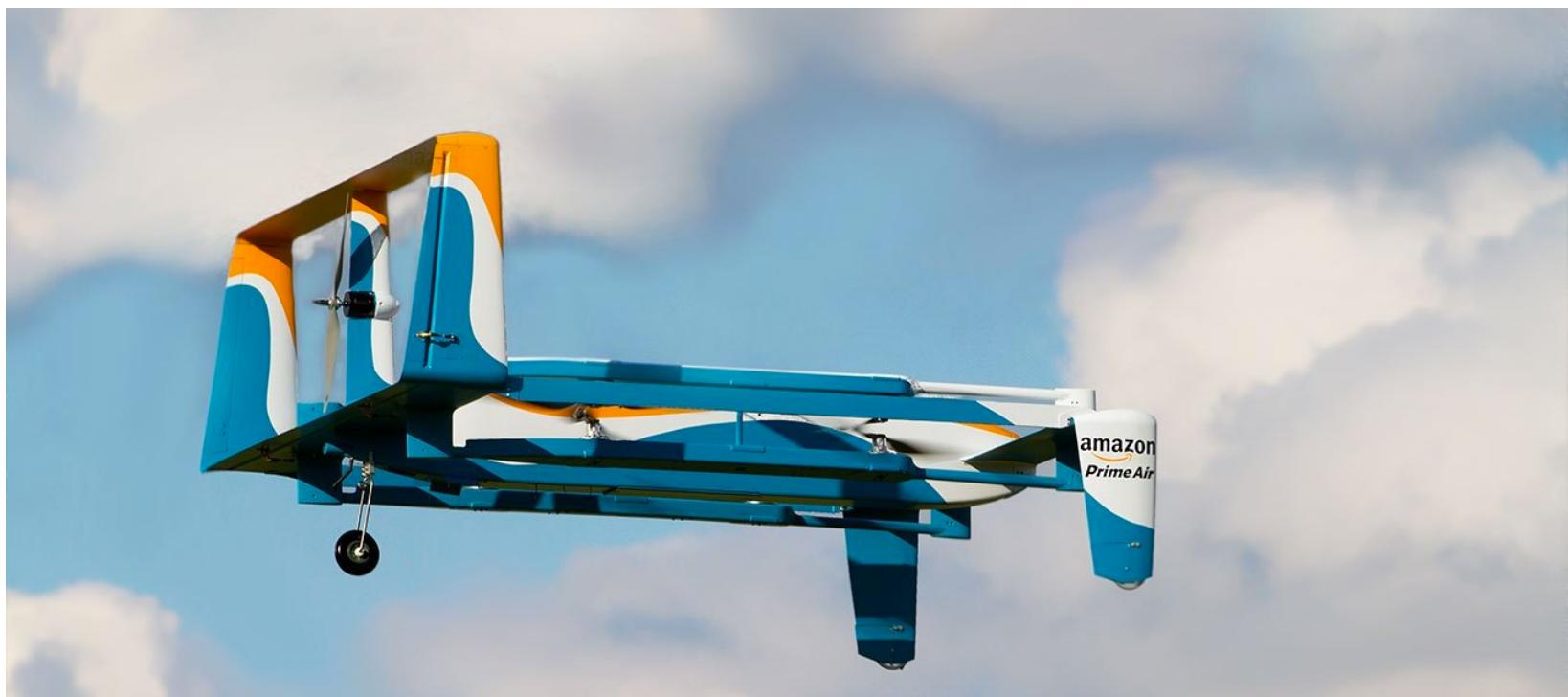
# Computer Cluster Utilization

- Computer cluster with  $k$  servers
  - Requests independently go to server  $i$  with probability  $p_i$
  - Let event  $A_i =$  server  $i$  receives no requests
  - $X = \#$  of events  $A_1, A_2, \dots, A_k$  that occur
  - $Y = \#$  servers that receive  $\geq 1$  request  $= k - X$
  - $E[Y]$  after first  $n$  requests?
  - Since requests independent:  $P(A_i) = (1 - p_i)^n$

$$E[X] = \sum_{i=1}^k P(A_i) = \sum_{i=1}^k (1 - p_i)^n$$

$$E[Y] = k - E[X] = k - \sum_{i=1}^k (1 - p_i)^n$$

when  $p_i = \frac{1}{k}$  for  $1 \leq i \leq k$ ,  $E[Y] = k - \sum_{i=1}^k \left(1 - \frac{1}{k}\right)^n = k \left(1 - \left(1 - \frac{1}{k}\right)^n\right)$





\* 52% of Amazons Profits

\*\*As profitable as Amazon's North America commerce operations

# Computer Cluster Utilization

- Computer cluster with  $k$  servers
  - Requests independently go to server  $i$  with probability  $p_i$
  - Let event  $A_i =$  server  $i$  receives no requests
  - $X = \#$  of events  $A_1, A_2, \dots, A_k$  that occur
  - $Y = \#$  servers that receive  $\geq 1$  request  $= k - X$
  - $\text{Var}(Y)$  after first  $n$  requests? ( $= (-1)^2 \text{Var}(X) = \text{Var}(X)$ )
  - Independent requests:  $P(A_i A_j) = (1 - p_i - p_j)^n, i \neq j$

$$\begin{aligned}\text{Var}(X) &= 2 \sum_{i < j} (1 - p_i - p_j)^n + E[X] - (E[X])^2 & E[X] &= \sum_{i=1}^k (1 - p_i)^n \\ &= 2 \sum_{i < j} (1 - p_i - p_j)^n + \sum_{i=1}^k (1 - p_i)^n - \left( \sum_{i=1}^k (1 - p_i)^n \right)^2 = \text{Var}(Y)\end{aligned}$$

# Computer Cluster = Coupon Collecting

- Computer cluster with  $k$  servers
  - Requests independently go to server  $i$  with probability  $p_i$
  - Let event  $A_i$  = server  $i$  receives no requests
  - $X = \#$  of events  $A_1, A_2, \dots, A_k$  that occur
  - $Y = \#$  servers that receive  $\geq 1$  request =  $k - X$
- This is really another “Coupon Collector” problem
  - Each server is a “coupon type”
  - Request to server = collecting a coupon of that type
- Hash table version
  - Each server is a bucket in table
  - Request to server = string gets hashed to that bucket

A lemma to start your weekend

# Product of Expectations

- Say  $X$  and  $Y$  are independent random variables, and  $g(\bullet)$  and  $h(\bullet)$  are real-valued functions

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

- Proof:

$$\begin{aligned} E[g(X)h(Y)] &= \int_{y=-\infty}^{\infty} \int_{x=-\infty}^{\infty} g(x)h(y)f_{X,Y}(x,y) dx dy \\ &= \int_{y=-\infty}^{\infty} \int_{x=-\infty}^{\infty} g(x)h(y)f_X(x)f_Y(y) dx dy \\ &= \int_{x=-\infty}^{\infty} g(x)f_X(x) dx \cdot \int_{y=-\infty}^{\infty} h(y)f_Y(y) dy \\ &= E[g(X)]E[h(Y)] \end{aligned}$$

Next time:  
Covariance lemmade