

Independence



CS 109
Lecture 5
April 6th, 2016

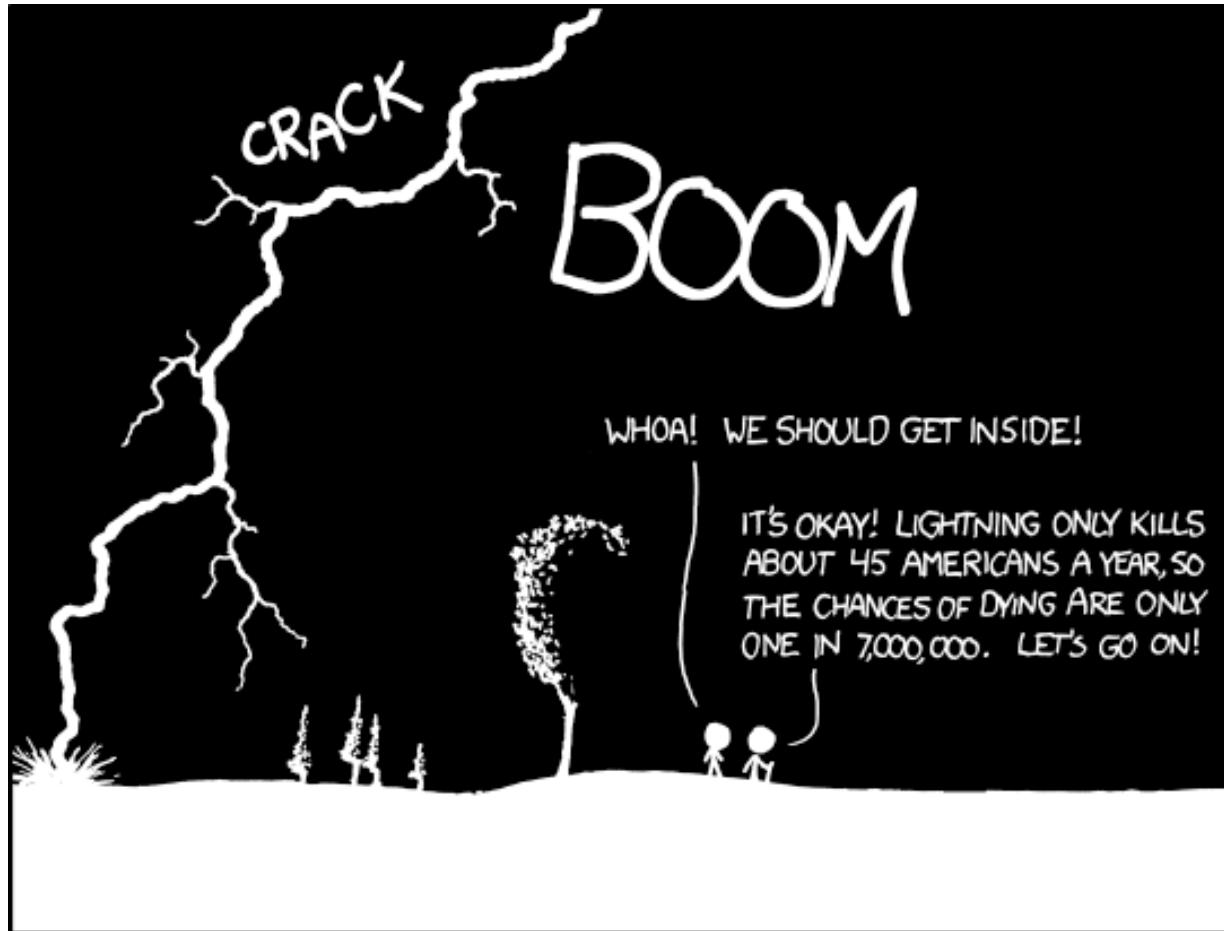
Today's Topics

Last time:
Conditional Probability
Bayes Theorem

Today:
Independence
Conditional Independence

Next time:
Random Variables

The Tragedy of Conditional Prob



THE ANNUAL DEATH RATE AMONG PEOPLE
WHO KNOW THAT STATISTIC IS ONE IN SIX.

Thanks xkcd! <http://xkcd.com/795/>

A Few Useful Formulas

- For any events A and B:

$$P(A \cap B) = P(B \cap A) \quad (\text{Commutativity})$$

$$\begin{aligned} P(A \cap B) &= P(A | B) P(B) \\ &= P(B | A) P(A) \end{aligned} \quad (\text{Chain rule})$$

$$P(A \cap B^c) = P(A) - P(AB) \quad (\text{Intersection})$$

$$P(A \cap B) \geq P(A) + P(B) - 1 \quad (\text{Bonferroni})$$

Generality of Conditional Probability

- For any events A, B, and E, you can condition consistently on E, and these formulas still hold:

$$P(A \cap B | E) = P(B | A \cap E)$$

$$P(A \cap B | E) = P(A | B \cap E) P(B | E)$$

$$P(A | B \cap E) = \frac{P(B | A \cap E) P(A | E)}{P(B | E)} \quad (\text{Bayes' Thm.})$$

BAE's Theorem?

$$P(A | B E) = \frac{P(B | A E) P(A | E)}{P(B | E)}$$



Generality of Conditional Probability

- For any events A, B, and E, you can condition consistently on E, and these formulas still hold:

$$P(A \cap B | E) = P(B | A \cap E)$$

$$P(A \cap B | E) = P(A | B \cap E) P(B | E)$$

$$P(A | B \cap E) = \frac{P(B | A \cap E) P(A | E)}{P(B | E)} \quad (\text{Bayes' Thm.})$$

- Can think of E as “everything you already know”
- Formally, $P(\cdot | E)$ satisfies 3 axioms of probability

Our Still Misunderstood Friend

- Roll two 6-sided dice, yielding values D_1 and D_2
 - Let E be event: $D_1 = 1$
 - Let F be event: $D_2 = 1$
- What is $P(E)$, $P(F)$, and $P(EF)$?
 - $P(E) = 1/6$, $P(F) = 1/6$, $P(EF) = 1/36$
 - $P(EF) = P(E) P(F)$ \rightarrow E and F independent
- Let G be event: $D_1 + D_2 = 5$ $\{(1, 4), (2, 3), (3, 2), (4, 1)\}$
- What is $P(E)$, $P(G)$, and $P(EG)$?
 - $P(E) = 1/6$, $P(G) = 4/36 = 1/9$, $P(EG) = 1/36$
 - $P(EG) \neq P(E) P(G)$ \rightarrow E and G dependent

Independence

- Two events E and F are called independent if:
 $P(EF) = P(E) P(F)$
Or, equivalently: $P(E | F) = P(E)$
- Otherwise, they are called dependent events
- Three events E, F, and G independent if:
 $P(EFG) = P(E) P(F) P(G)$, and
 $P(EF) = P(E) P(F)$, and
 $P(EG) = P(E) P(G)$, and
 $P(FG) = P(F) P(G)$

Independence

- Given independent events E and F , prove that E and F^C are independent

Independence

- Given independent events E and F, prove that E and F^c are independent
- Proof:

$$P(E \cap F^c)$$

We want to show
that this is equal
to $P(E)P(F^c)$

Independence

- Given independent events E and F, prove that E and F^C are independent
- Proof:

$$P(E \cap F^C) = P(E) - P(EF)$$

Since $P(E) = P(EF^C) + P(EF)$

Independence

- Given independent events E and F, prove that E and F^C are independent
- Proof:

$$\begin{aligned} P(E \cap F^C) &= P(E) - P(EF) \\ &= P(E) - P(E) P(F) \end{aligned}$$

Since we are told
E and F are
independent

Independence

- Given independent events E and F, prove that E and F^C are independent
- Proof:

$$\begin{aligned} P(E \cap F^C) &= P(E) - P(EF) \\ &= P(E) - P(E) P(F) \\ &= P(E) [1 - P(F)] \end{aligned} \quad \text{Factoring!}$$

Independence

- Given independent events E and F, prove that E and F^c are independent
- Proof:

$$\begin{aligned} P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E) P(F) \\ &= P(E) [1 - P(F)] \\ &= P(E) P(F^c) \end{aligned}$$

Yep, that's the complement

Independence

- Given independent events E and F, prove that E and F^c are independent
- Proof:

$$\begin{aligned} P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E) P(F) \\ &= P(E) [1 - P(F)] \\ &= P(E) P(F^c) \end{aligned}$$

So, E and F^c independent, implying that:

$$P(E | F^c) = P(E) = P(E | F)$$

Independence

- Given independent events E and F, prove that E and F^c are independent
- Proof:

$$\begin{aligned} P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E) P(F) \\ &= P(E) [1 - P(F)] \\ &= P(E) P(F^c) \end{aligned}$$

So, E and F^c independent, implying that:

$$P(E | F^c) = P(E) = P(E | F)$$

- Intuitively, if E and F are independent, knowing whether F holds gives us no information about E

Generalized Independence

- General definition of Independence:
Events E_1, E_2, \dots, E_n are independent if for every subset with r elements (where $r \leq n$) it holds that:
$$P(E_1, E_2, E_3, \dots, E_r) = P(E_1)P(E_2)P(E_3)\dots P(E_r)$$
- Example: outcomes of n separate flips of a coin are all independent of one another
 - Each flip in this case is called a “trial” of the experiment

Two Dice

- Roll two 6-sided dice, yielding values D_1 and D_2
 - Let E be event: $D_1 = 1$
 - Let F be event: $D_2 = 6$
 - Are E and F independent? Yes!
- Let G be event: $D_1 + D_2 = 7$
 - Are E and G independent? Yes!
 - $P(E) = 1/6, P(G) = 1/6, P(E \cap G) = 1/36$ [roll (1, 6)]
 - Are F and G independent? Yes!
 - $P(F) = 1/6, P(G) = 1/6, P(F \cap G) = 1/36$ [roll (1, 6)]
 - Are E, F and G independent? No!
 - $P(E \cap F \cap G) = 1/36 \neq 1/216 = (1/6)(1/6)(1/6)$

Generating Random Bits

- A computer produces a series of random bits, with probability p of producing a 1.
 - Each bit generated is an independent trial
 - $E = \text{first } n \text{ bits are 1's, followed by a single 0}$
 - What is $P(E)$?
- Solution
 - $P(\text{first } n \text{ 1's}) = P(1^{\text{st}} \text{ bit}=1) P(2^{\text{nd}} \text{ bit}=1) \dots P(n^{\text{th}} \text{ bit}=1)$
 $= p^n$
 - $P(n+1 \text{ bit}=0) = (1 - p)$
 - $P(E) = P(\text{first } n \text{ 1's}) P(n+1 \text{ bit}=0) = p^n (1 - p)$

Coin Flips

- Say a coin comes up heads with probability p
 - Each coin flip is an independent trial
- $P(n \text{ heads on } n \text{ coin flips}) = p^n$
- $P(n \text{ tails on } n \text{ coin flips}) = (1 - p)^n$
- $P(\text{first } k \text{ heads, then } n - k \text{ tails}) = p^k (1 - p)^{n-k}$
- $P(\text{exactly } k \text{ heads on } n \text{ coin flips}) = ?$

Explain...

P(exactly k heads on n coin flips)?

$$\binom{n}{k} p^k (1-p)^{n-k}$$

Think of the flips as ordered:

Ordering 1: T, H, H, T, T, T....

The coin flips are independent!

Ordering 2: H, T, H, T, T, T....

And so on...

$$P(F_i) = p^k (1-p)^{n-k}$$

Let's make each ordering with k heads an event... F_i

P(exactly k heads on n coin flips) = P(any one of the events)

P(exactly k heads on n coin flips) = P(F_1 or F_2 or F_3 ...)

Those events are mutually exclusive!

Moment of Crystallization

Add vs Multiply?



Batman vs Superman



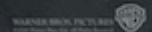
COMING SOON
#BATMAN v SUPERMAN

SEE IT IN 3D



VULCANI UNUSUAL FILMS

TM & © DC COMICS



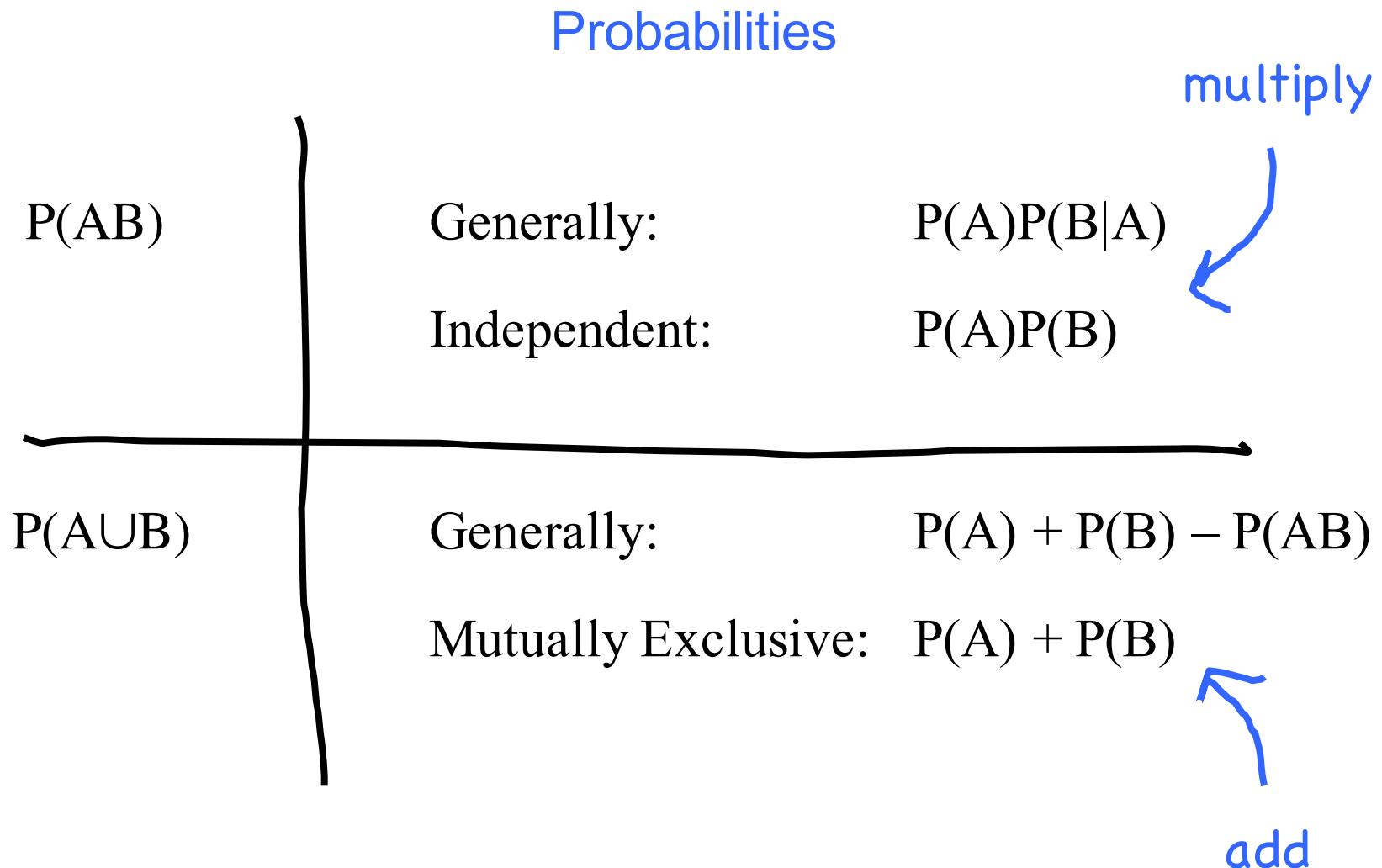
Add vs Multiply

+

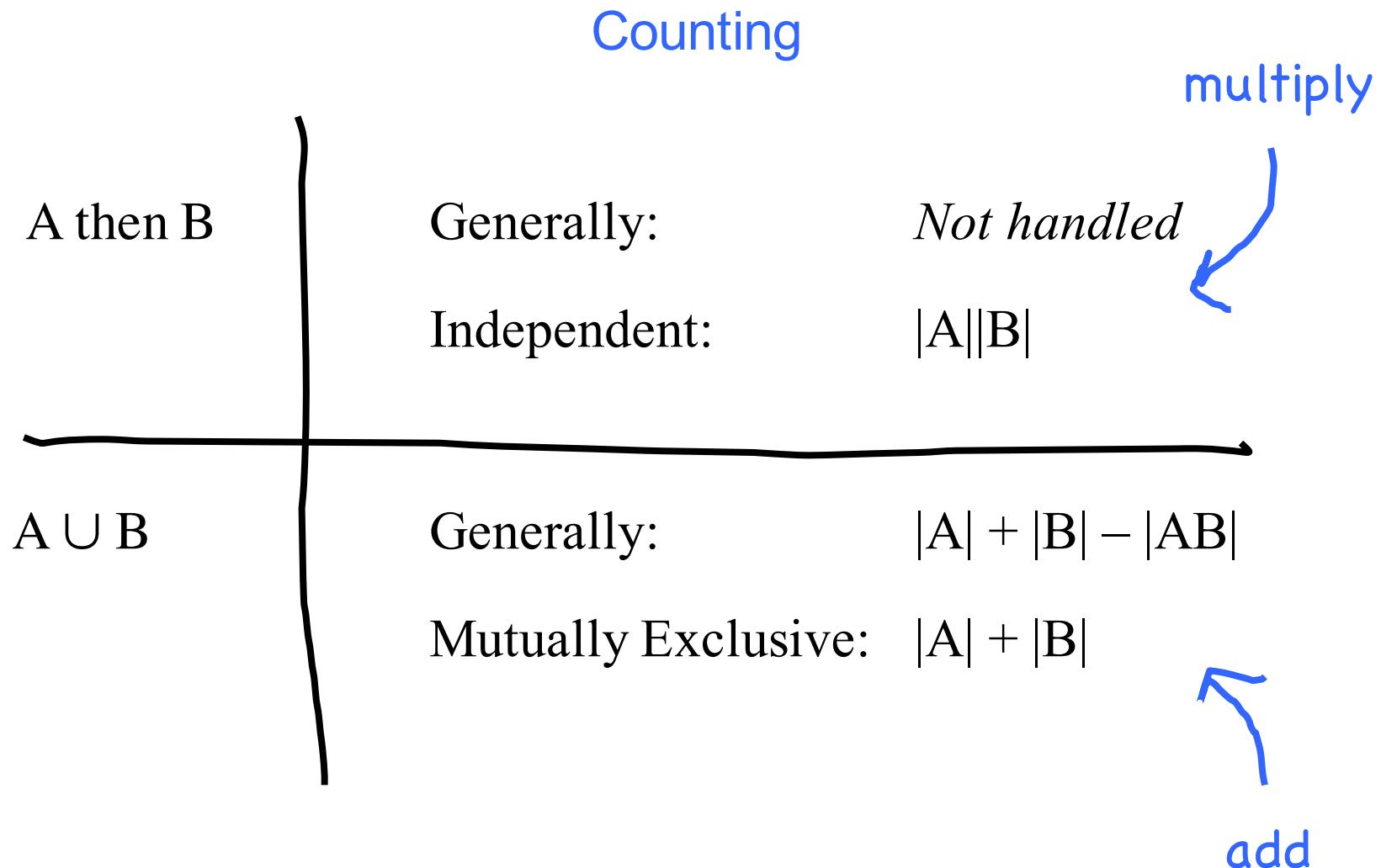
vs

×

Add vs Multiply



Add vs Multiply



Next up...

And vs Condition

$P(A \cap B)$ vs $P(A | B)$

$$P(A \cap B) = P(A | B) P(B)$$

Hash Tables

- m strings are hashed (unequally) into a hash table with n buckets
 - Each string hashed is an independent trial, with probability p_i of getting hashed to bucket i
 - $E =$ at least one string hashed to first bucket
 - What is $P(E)$?
- Solution

To the chalk board!

Yet More Hash Tables

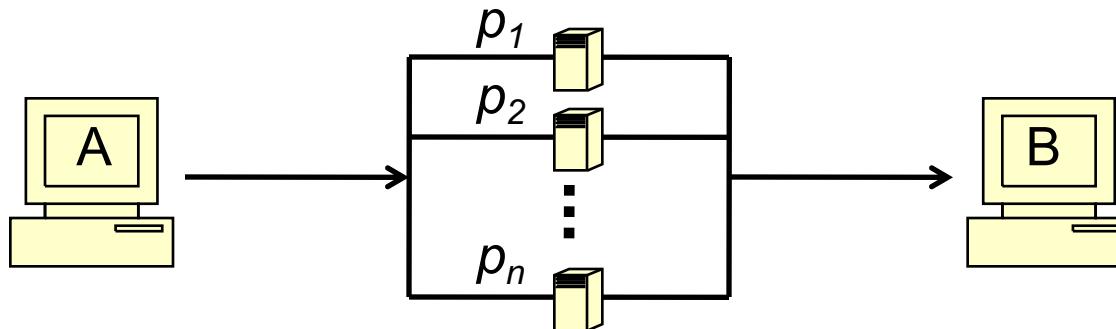
- m strings are hashed (unequally) into a hash table with n buckets
 - Each string hashed is an independent trial, with probability p_i of getting hashed to bucket i
 - $E = \text{At least 1 of}$ buckets 1 to k has ≥ 1 string hashed to it
- Solution
 - $F_i = \text{at least one string hashed into } i\text{-th bucket}$
 - $P(E) = P(F_1 \cup F_2 \cup \dots \cup F_k) = 1 - P((F_1 \cup F_2 \cup \dots \cup F_k)^c)$
 $= 1 - P(F_1^c F_2^c \dots F_k^c)$ (DeMorgan's Law)
 - $P(F_1^c F_2^c \dots F_k^c) = P(\text{no strings hashed to buckets 1 to } k)$
 $= (1 - p_1 - p_2 - \dots - p_k)^m$
 - $P(E) = 1 - (1 - p_1 - p_2 - \dots - p_k)^m$

No, Really, More Hash Tables

- m strings are hashed (unequally) into a hash table with n buckets
 - Each string hashed is an independent trial, with probability p_i of getting hashed to bucket i
 - $E = \text{Each of}$ buckets 1 to k has ≥ 1 string hashed to it
 - Solution
 - $F_i = \text{at least one string hashed into } i\text{-th bucket}$
 - $P(E) = P(F_1 F_2 \dots F_k) = 1 - P((F_1 F_2 \dots F_k)^c)$
 $= 1 - P(F_1^c \cup F_2^c \cup \dots \cup F_k^c)$ (DeMorgan's Law)
 $= 1 - P\left(\bigcup_{i=1}^k F_i^c\right) = 1 - \sum_{r=1}^k (-1)^{(r+1)} \sum_{i_1 < \dots < i_r} P(F_{i_1}^c F_{i_2}^c \dots F_{i_r}^c)$
- where $P(F_{i_1}^c F_{i_2}^c \dots F_{i_r}^c) = (1 - p_{i_1} - p_{i_2} - \dots - p_{i_r})^m$

Sending a Message Through Network

- Consider the following parallel network:



- n independent routers, each with probability p_i of functioning (where $1 \leq i \leq n$)
 - $E =$ functional path from A to B exists. What is $P(E)$?
- Solution:
 - $P(E) = 1 - P(\text{all routers fail})$
 $= 1 - (1 - p_1)(1 - p_2)\dots(1 - p_n)$
 $= 1 - \prod_{i=1}^n (1 - p_i)$

Phew...

2 min pedagogical pause



Digging Deeper on Independence

- Recall, two events E and F are called independent if

$$P(EF) = P(E) P(F)$$

- If E and F are independent, does that tell us whether the following is true or not:

$$P(EF | G) = P(E | G) P(F | G),$$

where G is an arbitrary event?

- In general, No!

Not So Independent Dice

- Roll two 6-sided dice, yielding values D_1 and D_2
 - Let E be event: $D_1 = 1$
 - Let F be event: $D_2 = 6$
 - Let G be event: $D_1 + D_2 = 7$
- E and F are independent
 - $P(E) = 1/6$, $P(F) = 1/6$, $P(EF) = 1/36$
- Now condition both E and F on G:
 - $P(E|G) = 1/6$, $P(F|G) = 1/6$, $P(EF|G) = 1/6$
 - $P(EF|G) \neq P(E|G) P(F|G)$ $\rightarrow E|G$ and $F|G$ dependent
- Independent events can become dependent by conditioning on additional information

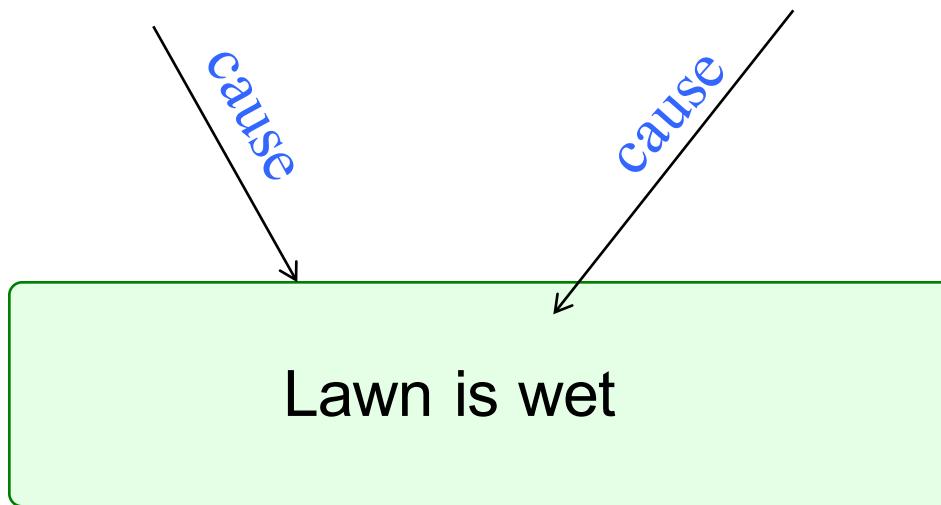
Explaining Away

- Say you have a lawn
 - It gets watered by rain or sprinklers
 - $P(\text{rain})$ and $P(\text{sprinklers were on})$ are independent
 - Now, you come outside and see the grass is wet
 - You know that the sprinklers were on
 - Does that lower probability that rain was cause of wet grass?
 - This phenomena is called “explaining away”
 - One cause of an observation makes other causes less likely

Explaining Away

Sprinklers

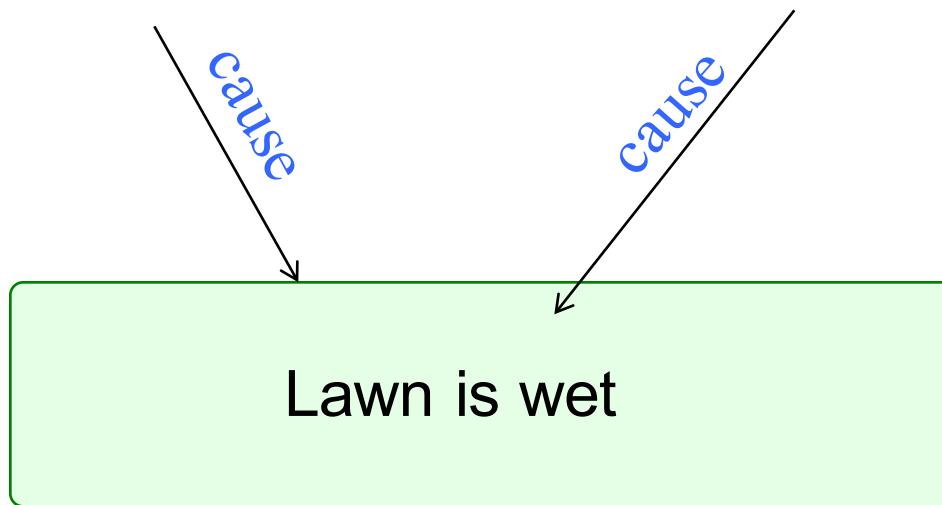
Rain



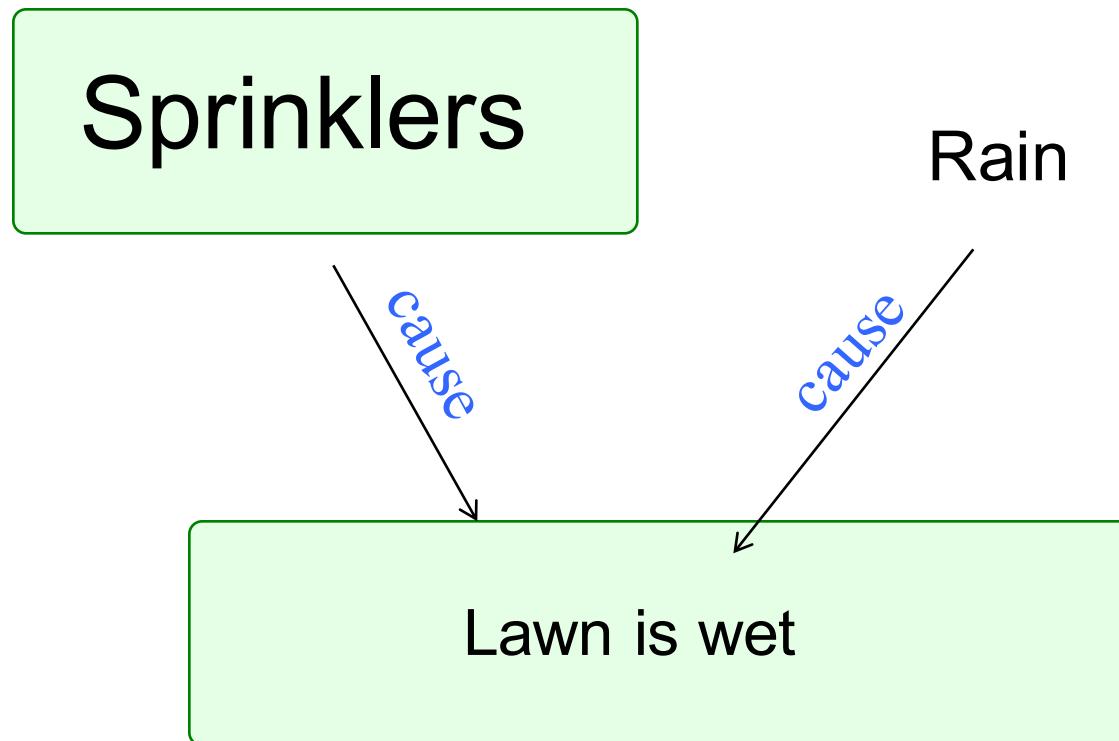
Explaining Away

Sprinklers

Rain



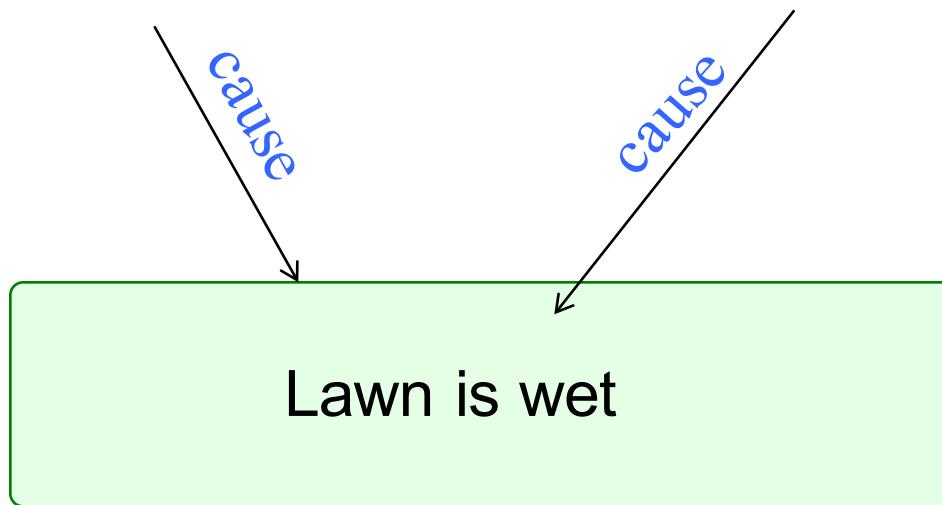
Explaining Away



Explaining Away

Sprinklers

Rain



Explaining Away

Sprinklers

Rain

cause

cause

Lawn is wet

Explaining Away

Sprinklers

Rain

cause

cause

Lawn is wet

Conditioning Can Make Independence

- Consider a randomly chosen day of the week
 - Let A be event: It is not Monday
 - Let B be event: It is Saturday
 - Let C be event: It is the weekend
- A and B are dependent
 - $P(A) = 6/7$, $P(B) = 1/7$, $P(AB) = 1/7 \neq (6/7)(1/7)$
- Now condition both A and B on C:
 - $P(A|C) = 1$, $P(B|C) = 1/2$, $P(AB|C) = 1/2$
 - $P(AB|C) = P(A|C) P(B|C) \rightarrow A|C \text{ and } B|C \text{ independent}$
- Dependent events can become independent by conditioning on additional information

Conditional Independence

- Two events E and F are called conditionally independent given G, if

$$P(E \cap F | G) = P(E | G) P(F | G)$$

Or, equivalently: $P(E | F \cap G) = P(E | G)$

NETFLIX

And Learn

Netflix and Learn

What is the probability
that a user will like
Life is Beautiful?

$$P(E)$$



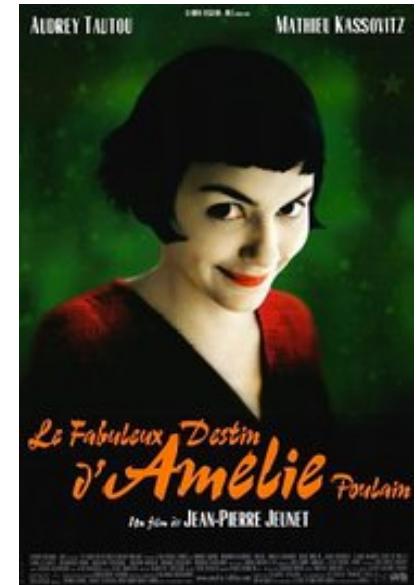
$$P(E) = \lim_{n \rightarrow \infty} \frac{n(E)}{n} \approx \frac{\# \text{ people who liked movie}}{\# \text{ people who watched movie}}$$

$$P(E) = 50,234,231 / 50,923,123 = 0.97$$

Netflix and Learn

What is the probability
that a user will like
Life is Beautiful, given
they liked Amelie?

$$P(E|F)$$



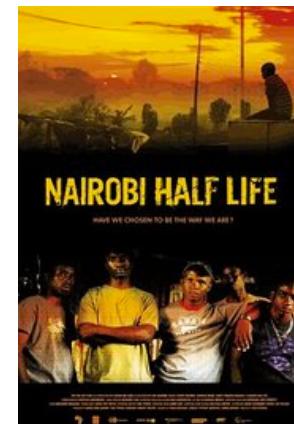
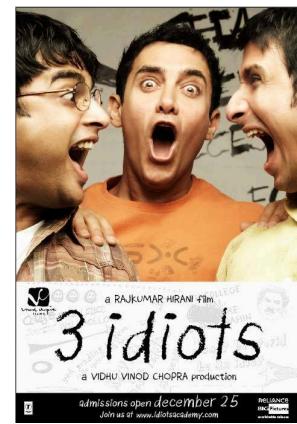
$$P(E|F) = \frac{P(EF)}{P(F)} = \frac{\text{people who liked both}}{\text{people who watched both}} = \frac{\text{people who liked amelie}}{\text{people who watched amelie}}$$

$$P(E|F) = 0.99$$

Conditioned on liking a set of movies?

Netflix and Learn

Each event corresponds to liking a particular movie



E_1

E_2

E_3

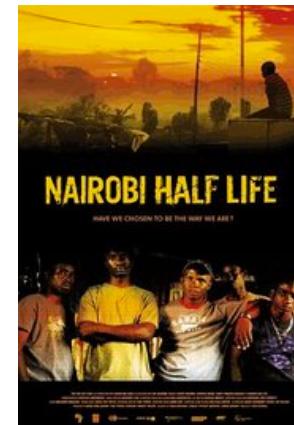
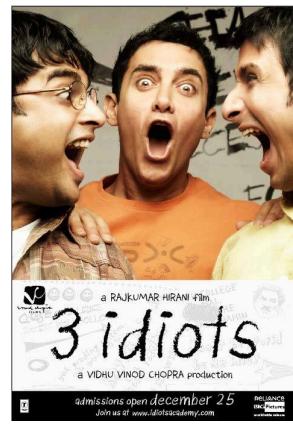
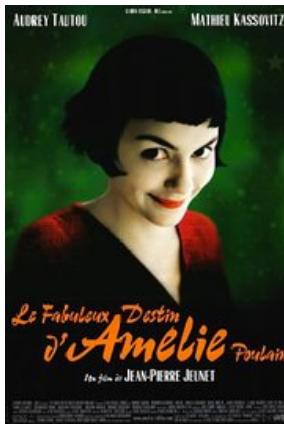
E_4

$$P(E_4 | E_1, E_2, E_3) ?$$

Is E_4 independent of E_1, E_2, E_3 ?

Netflix and Learn

Is E_4 independent of E_1, E_2, E_3 ?



E_1

E_2

E_3

E_4

$$P(E_4 | E_1, E_2, E_3) \stackrel{?}{=} P(E_4)$$

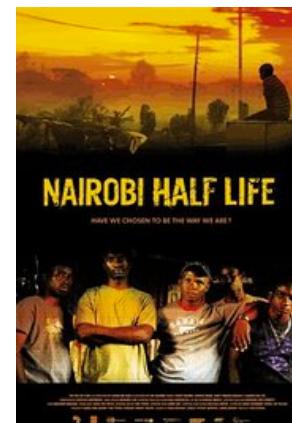
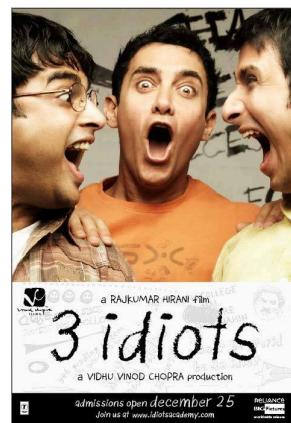
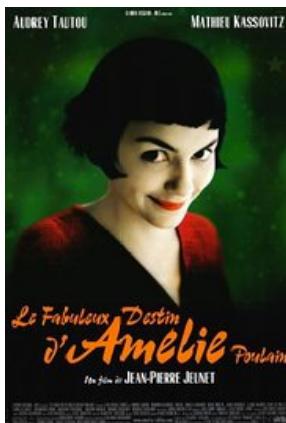
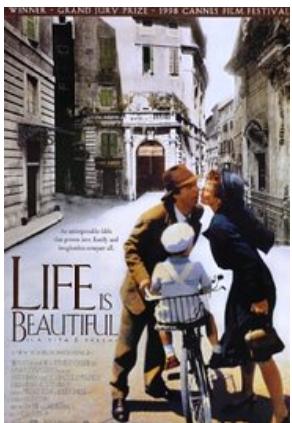
Netflix and Learn

- What is the probability that a user watched four particular movies?
 - There are 13,000 titles on Netflix
 - The user watches 30 random titles
 - E = movies watched include the given four.
- Solution:

$$P(E) = \frac{\binom{4}{4} \binom{12996}{24}}{\binom{13000}{30}} = 10^{-11}$$

Watch those four Choose 24 movies
not in the set
Choose 30 movies
from netflix

Netflix and Learn



E_1

E_2

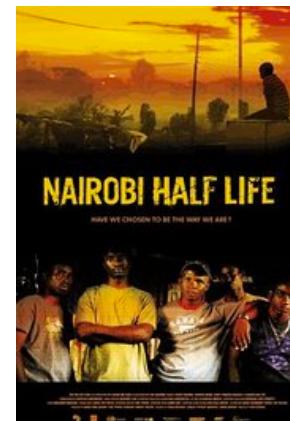
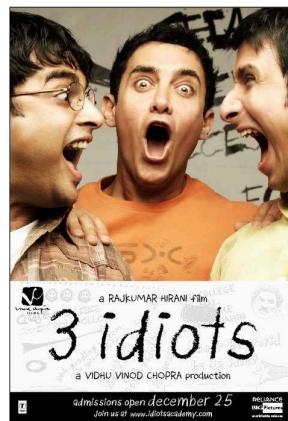
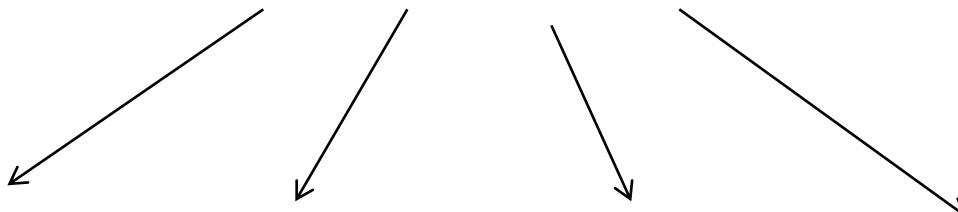
E_3

E_4

Netflix and Learn

K_1

Like foreign emotional comedies



E_1

E_2

E_3

E_4

Assume E_1, E_2, E_3 and E_4 are conditionally independent given K_1

Netflix and Learn

K_1

Like foreign emotional comedies

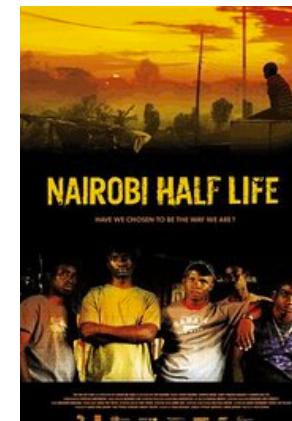
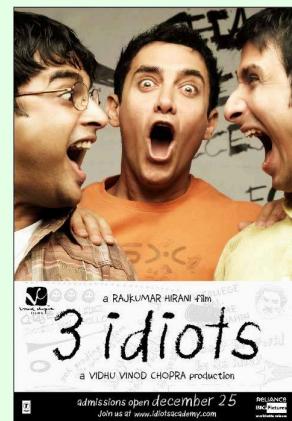
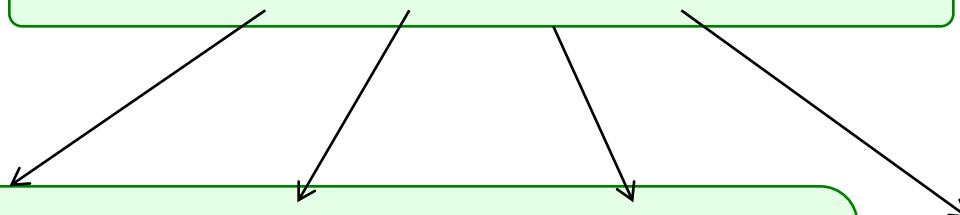


Assume E_1 , E_2 , E_3 and E_4 are conditionally independent given K_1

Netflix and Learn

K_1

Like foreign emotional comedies



E_1

E_2

E_3

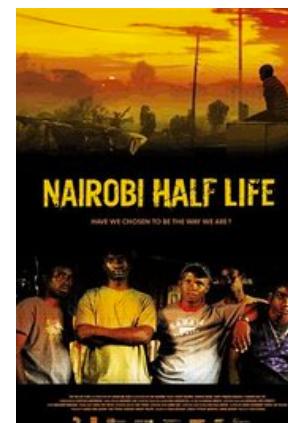
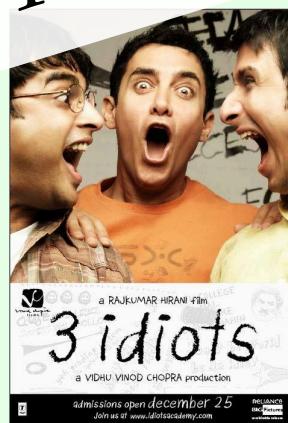
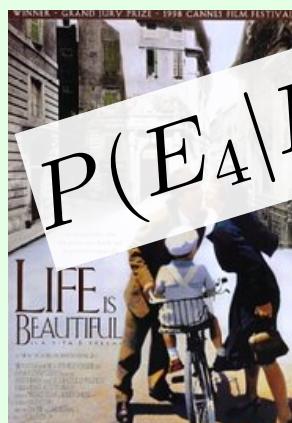
E_4

Assume E_1, E_2, E_3 and E_4 are conditionally independent given K_1

Netflix and Learn

K_1

Like foreign emotional comedies



E_1

E_2

E_3

E_4

Assume E_1, E_2, E_3 and E_4 are conditionally independent given K_1

Conditional independence is a practical, real world way of decomposing hard probability questions.

Big Deal

“Exploiting conditional independence to generate fast probabilistic computations is one of the main contributions CS has made to probability theory”

-Judea Pearl wins 2011 Turing Award, “*For fundamental contributions to artificial intelligence through the development of a calculus for probabilistic and causal reasoning*”

Extra problem given time

Reminder: Geometric Series

- Geometric series: $x^0 + x^1 + x^2 + x^3 + \dots + x^n = \sum_{i=0}^n x^i$
- If x is greater than 0 and less than 1:

$$\sum_{i=0}^{\infty} x^i = \frac{1}{1-x}$$

- From your “Calculation Reference” handout:

Simplified Craps

- Two 6-sided dice repeatedly rolled (roll = ind. trial)
 - E = 5 is rolled before a 7 is rolled
 - What is $P(E)$?
- Solution
 - F_n = no 5 or 7 rolled in first $n - 1$ trials, 5 rolled on n^{th} trial
 - $P(E) = P\left(\bigcup_{n=1}^{\infty} F_n\right) = \sum_{n=1}^{\infty} P(F_n)$
 - $P(5 \text{ on any trial}) = 4/36 \quad P(7 \text{ on any trial}) = 6/36$
 - $P(F_n) = (1 - (10/36))^{n-1} (4/36) = (26/36)^{n-1} (4/36)$
 - $P(E) = \frac{4}{36} \sum_{n=1}^{\infty} \left(\frac{26}{36}\right)^{n-1} = \frac{4}{36} \sum_{n=0}^{\infty} \left(\frac{26}{36}\right)^n = \frac{4}{36} \frac{1}{1 - \frac{26}{36}} = \frac{2}{5}$