

# 1 2023

## 1.1 Safety Stock

**Theorem 1.1.** *Let the lead-time  $T$  be normally distributed, i.e.  $T \sim N(\mu_\ell, \sigma_\ell^2)$  and denote the set of demand of period  $i$  as  $\{D_i\}_{i=1}^T$  where  $D_i \sim N(\mu_d, \sigma_d^2)$  for each  $1 \leq i \leq T$ . Then the safety stock level,  $SS$  satisfies the following equation*

$$SS = z_\alpha \sqrt{\sigma_d^2 \mu_\ell + \mu_d^2 \sigma_\ell^2},$$

where  $z_\alpha$  is the Z-value of a desired  $\alpha$  which is chosen.

*Proof.* Denote the demand between within lead-time as  $D(T) := \sum_{i=1}^T D_i$ . The safety stock level is then set at

$$SS = z_\alpha \sqrt{\text{Var}(D(T))}.$$

Thus we need only to prove  $\text{Var}(D(T))$ .

From Towering-property, note that

$$\begin{aligned} \mathbb{E}[D(T)] &= \mathbb{E}[\mathbb{E}[D(T)|T]] \\ &= \mathbb{E}[T\mu_d] = \mu_d \mathbb{E}[T] = \mu_d \mu_\ell. \end{aligned}$$

Furthemore, again with Towering-property, we have

$$\begin{aligned} \mathbb{E}[D^2(T)] &= \mathbb{E}[\mathbb{E}[D^2(T)|T]] \\ &= \mathbb{E}[\text{Var}[D(T)|T] + (\mathbb{E}[D(T)|T])^2] \\ &= \mathbb{E}[T\sigma_d^2 + (T\mu_d)^2] \\ &= \sigma_d^2 \mathbb{E}[T] + \mu_d^2 \mathbb{E}[T^2] \\ &= \sigma_d^2 \mu_\ell + \mu_d^2 (\text{Var}(T) + (\mathbb{E}[T])^2) \\ &= \sigma_d^2 \mu_\ell + \mu_d^2 (\sigma_\ell^2 + \mu_\ell^2). \end{aligned}$$

Finally,

$$\begin{aligned} \text{Var}(D(T)) &= \mathbb{E}[D^2(T)] - (\mathbb{E}[D(T)])^2 \\ &= \sigma_d^2 \mu_\ell + \mu_d^2 (\sigma_\ell^2 + \mu_\ell^2) - (\mu_d \mu_\ell)^2 \\ &= \sigma_d^2 \mu_\ell + \mu_d^2 \sigma_\ell^2. \end{aligned}$$

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## 1.2 Binomial approx to Poisson

**Theorem 1.2.** For  $\lambda = np$ ,

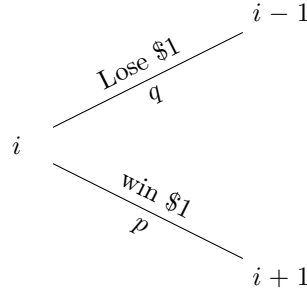
$$\lim_{n \rightarrow \infty} \binom{n}{k} p^k (1-p)^{n-k} = \frac{\lambda^k}{k!} e^{-\lambda}$$

## 1.3 Gambler's Ruin

Let  $p$  is the probability of winning \$1 and  $q = 1 - p$  is the probability of losing \$1. Suppose that a person starts with \$ $i$  amount of money where  $0 < i < N$  and  $N \in \mathbb{N}$  is the number he would want to walk away with. The probability of him winning from this starting position is given by the following equation:

$$a_k = \begin{cases} \frac{1 - (\frac{q}{p})^i}{1 - (\frac{q}{p})^N} & \text{where } p \neq q, \\ \frac{i}{N} & \text{where } p = q = \frac{1}{2}. \end{cases}$$

*Proof.* Note that



Then, we observe

$$\begin{aligned} a_i &= pa_{i+1} + qa_{i-1} \\ \underbrace{(p+q)}_{=1} a_i &= pa_{i+1} + qa_{i-1} \\ a_{i+1} - a_i &= \frac{p}{q}(a_i - a_{i-1}). \end{aligned}$$

Denote  $b_j := a_i - a_{i-1}$  and note that  $b_1 = a_1$  since  $a_0 = 0$  is the absorption edge, we deduce by recursion that

$$b_{i+1} = \left(\frac{p}{q}\right)^{i+1} a_1.$$

Moreover, note that

$$\begin{aligned} a_{i+1} - a_1 &= \sum_{k=1}^{i+1} (a_k - a_{k-1}) \\ &= \sum_{k=1}^{i+1} \left(\frac{p}{q}\right)^k a_1 \\ \Rightarrow a_{i+1} &= a_1 \left(1 + \sum_{k=1}^i \left(\frac{p}{q}\right)^k\right) \\ &= a_1 \left(\sum_{k=0}^i \left(\frac{p}{q}\right)^k\right). \end{aligned}$$

If  $p \neq q$ , by geometric series, we obtain

$$a_{i+1} = a_1 \frac{1 - \left(\frac{p}{q}\right)^{i+1}}{1 - \left(\frac{p}{q}\right)}.$$

If  $p = q = 0.5$ , then

$$a_{i+1} = a_1(i+1).$$

Now to solve for  $a_1$ , observe that since at absorption point  $a_N = 1$ , then letting  $i+1 = N$ , we get

$$1 = a_N = \begin{cases} \frac{1 - \left(\frac{q}{p}\right)^N}{1 - \left(\frac{q}{p}\right)} & \text{where } p \neq q, \\ a_1 N & \text{where } p = q = \frac{1}{2}, \end{cases}$$

which implies

$$a_1 = \begin{cases} \frac{1 - \left(\frac{q}{p}\right)^N}{1 - \left(\frac{q}{p}\right)} & \text{where } p \neq q, \\ \frac{1}{N} & \text{where } p = q = \frac{1}{2}. \end{cases}$$

Substituting this back, we obtained the desired equation. ■