## STA 235 - Multiple Regression: Statistical Adjustment

Spring 2021

McCombs School of Business, UT Austin

### **Quick reminders**

sta235.netlify.app

- Slides are posted on the website before the class.
- Required readings are posted in the Classes folder at least a week before.
- Check the code for each class.

#### Last week

- Quick multiple regression review
- Comparing **effect sizes**: Standardizing variables (i.e. all  $\hat{\beta}$ 's in the same scale)
- **Uncertainty quantification** in regression: Adj-R<sup>2</sup> and RSE.



## **Today**



#### • Statistical adjustment in regressions:

- o How do we interpret coefficients?
- What are those standard errors?
- Multicollinearity?
- Regression models with binary outcomes

#### **But first... JITTs!**

- (Almost) **everyone** answered the JITT.
- Answers are very useful and we will use them in today's class.
- People want more plots!
  - Ask and you shall receive.

Remember to ask questions!

## **Multiple Regression**

$$Y=eta_0+eta_1X_1+eta_2X_2+\ldots+eta_pX_p+arepsilon$$

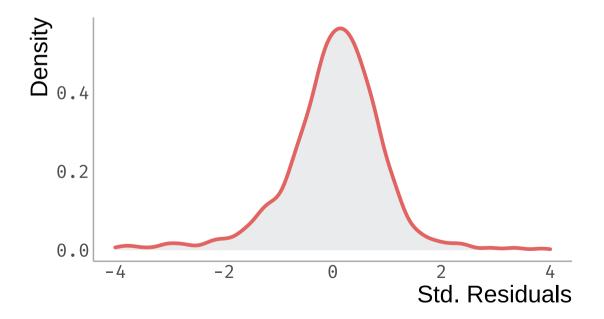
Assumptions of the linear regression model:

- (i) The conditional mean of Y is  $oxed{linear}$  in the  $X_j$  variables
- (ii) The error terms are:
  - normally distributed
  - independent from each other
  - identically distributed (i.e. constant variance)
  - ullet Last two assumptions are the ones that we refer to when saying  $arepsilon \sim iid$

## How can we check these assumptions?

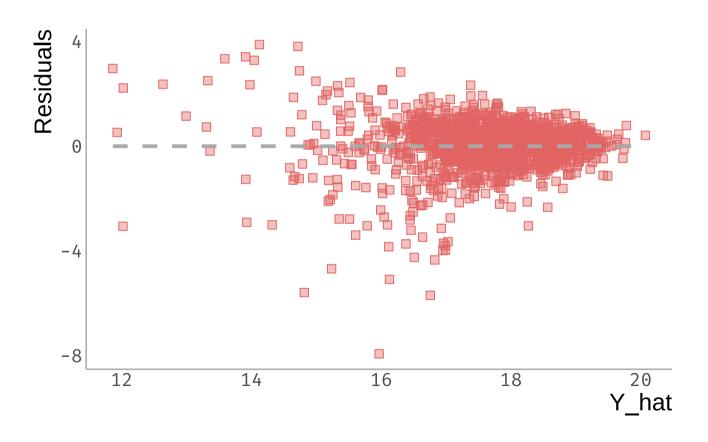
Remember last week's Bechdel test example:

```
ggplot(data = bechdel_fitted, aes(x = .std.resid)) +
  geom_density()
```



## How can we check these assumptions?

Remember last week's Bechdel test example:



### What happens if the assumptions break?



- Don't panic! There's still much to be done.
- Heteroskedasticity (non-constant variance) does not bias your estimates.
- Can be fixed many times with robust standard errors

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- Don't panic! There's still much to be done.
- Heteroskedasticity (non-constant variance) does not bias your estimates.
- Can be fixed many times with robust standard errors:
  - Easy to implement in R!

## Statistical Adjusment

- Remember that an important part of our job is to **correctly estimate** the  $\beta$  parameters:
  - Point estimates
  - Standard errors
- Two ways to estimate standard errors:
  - $\circ$  Direct estimation: Use probability theory (e.g.  $SE(ar{x})=rac{\hat{\sigma}}{\sqrt{n}}$ )
  - Simulation: Repeat the sampling process and estimates how much our estimate changes from one sample to the next (e.g. bootstrapping)

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- Remember that an important part of our job is to **correctly estimate** the  $\beta$  parameters:
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Important to understand R output!

#### Let's look at some data

- I have some data for price and sales of product 1, but also I'm tracking the prices of its competitor, product 2.
- The data looks like this:

```
## p1 p2 Sales
## 1 5.14 5.20 144.49
## 2 3.50 8.06 637.25
## 3 7.28 11.68 620.79
## 4 4.66 8.36 549.01
## 5 3.58 2.15 20.43
## 6 5.17 10.15 713.01
```

#### Let's fit a model

$$Sales_i = eta_0 + eta_1 p 1_i + eta_2 p 2_i + arepsilon_i$$

```
summary(lm(Sales ~ p1 + p2, data = sales))
```

```
##
## Call:
## lm(formula = Sales ~ p1 + p2, data = sales)
##
## Residuals:
##
      Min
              10 Median
                                   Max
  -66.916 -15.663 -0.509 18.904 63.302
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 115.717 8.548 13.54 <2e-16 ***
             -97.657 2.669 -36.59 <2e-16 ***
## p1
                        1.409 77.20 <2e-16 ***
## p2
             108.800
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 28.42 on 97 degrees of freedom
## Multiple R-squared: 0.9871, Adjusted R-squared: 0.9869
## F-statistic: 3717 on 2 and 97 DF, p-value: < 2.2e-16
```

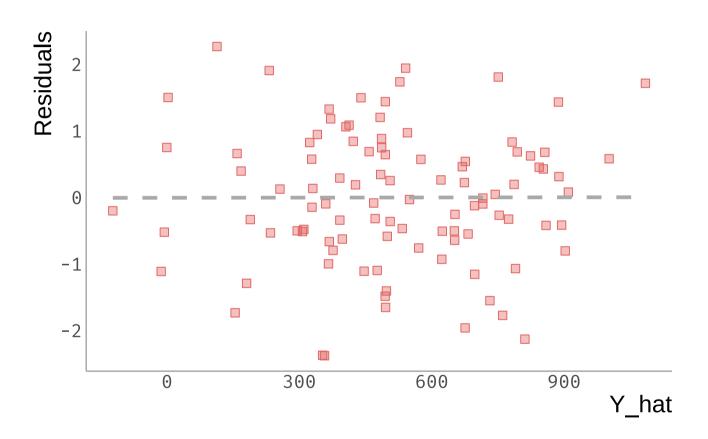
## Do assumptions hold?

$$arepsilon_i \sim N(0,\sigma^2)$$

```
sales_fitted <- augment(lm(Sales ~ p1 + p2, data = sales))

ggplot(data = sales_fitted, aes(x = .std.resid)) +
    geom_density()</pre>
```

## Do assumptions hold? (cont.)

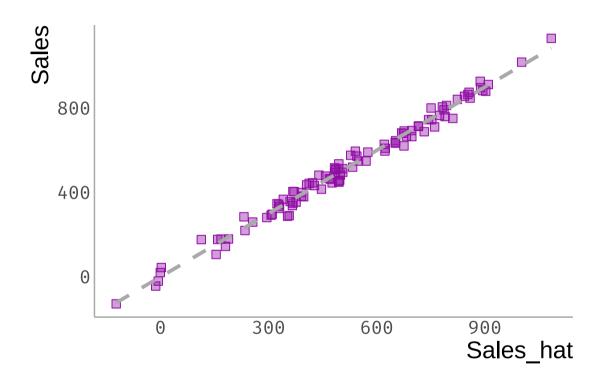


# Let's go back to our model

$$Sales_i = \beta_0 + \beta_1 p 1_i + \beta_2 p 2_i + \varepsilon_i$$

	Model 1			
(Intercept)	115.717***			
	(8.548)			
p1	-97.657***			
	(2.669)			
p2	108.800***			
	(1.409)			
Num.Obs.	100			
F	3717.292			
* p < 0.1, ** p < 0.05, *** p < 0.01				

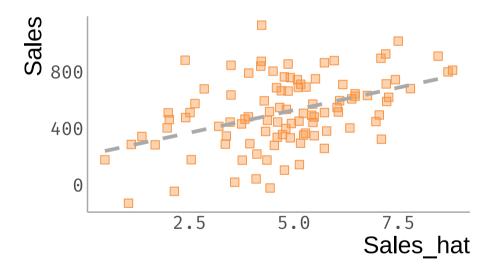
## How is the prediction working?



• Can you guess the slope?

## What if we only had p1 and not p2?

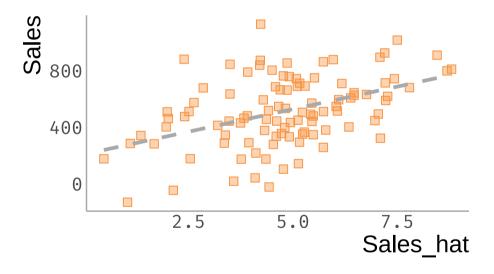
```
ggplot(data = sales, aes(x = p1, y = Sales)) +
  geom_point() +
  geom_smooth(method = "lm")
```



• If I increase the price sales go up?

## What if we only had p1 and not p2?

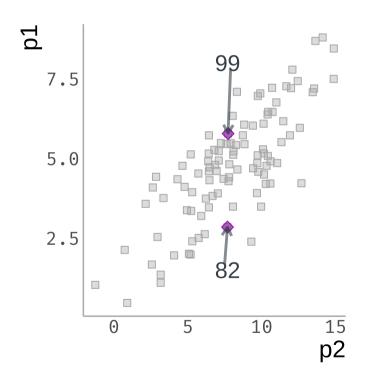
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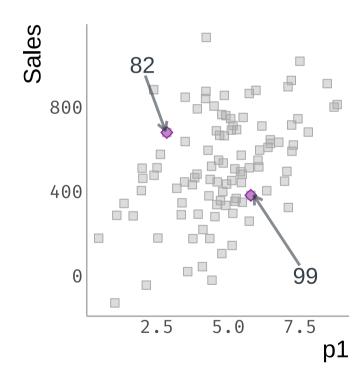


• If I increase the price sales go up? NO!

## Relationship between p1 and p2

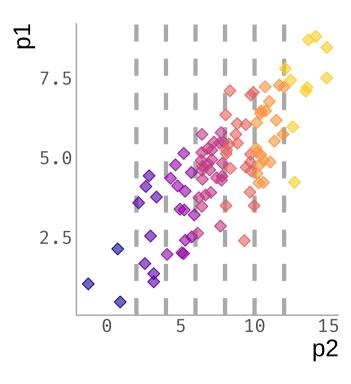
• Let's compare two different observations: Week 82 and week 99.





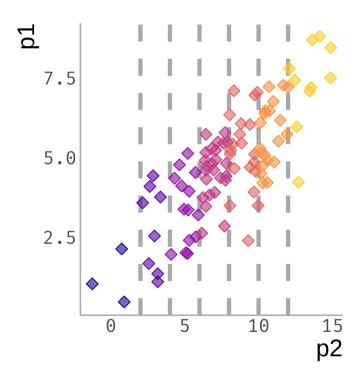
## Relationship between p1 and p2 (cont.)

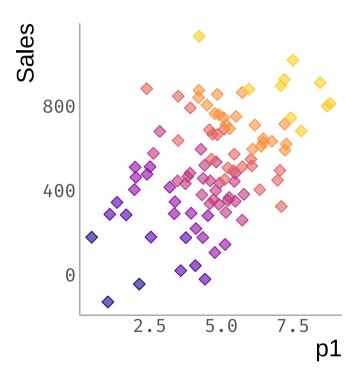
ullet Same thing happens when looking at a set of obs where  $p2\sim$  constant.



## Relationship between p1 and p2 (cont.)

• Same thing happens when looking at a set of obs where  $p2\sim$  constant.





#### **Conclusions?**

A larger p1 is associated with a larger p2, and overall, with more sales!

If we keep p2 constant, a larger p1 is associated with lower sales.

#### Let's look at more data: Beer limit

- From the JITT assignment, we have beers data
  - **nbeer**: Number of beers before getting tipsy
  - height, weight, age
  - female: Whether the student is female or not

##		nbeer	weight	height	age	female
##	1	12	192	72	26	0
##	2	12	160	66	27	0
##	3	5	155	65	25	0
##	4	5	120	66	28	0
##	5	7	150	67	28	0
##	6	13	175	71	31	0

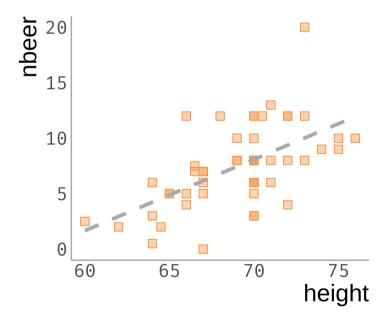
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$$nbeers_i = eta_0 + eta_1 \cdot height_i + arepsilon$$

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$$nbeers_i = eta_0 + eta_1 \cdot height_i + arepsilon$$

```
summary(lm(nbeer ~ height, data = beers))
```

```
##
## Call:
## lm(formula = nbeer ~ height, data = beers)
##
## Residuals:
     Min
            10 Median
                               Max
## -6.164 -2.005 -0.093 1.738 9.978
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -36.9200 8.9560 -4.122 0.000148 ***
## height 0.6430
                       0.1296 4.960 9.23e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.109 on 48 degrees of freedom
## Multiple R-squared: 0.3389, Adjusted R-squared: 0.3251
## F-statistic: 24.6 on 1 and 48 DF, p-value: 9.23e-06
```

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- What do you think will happen when I control for weight?

```
beers_fitted_weight <- augment(lm(nbeer ~ weight, data = beers))</pre>
```

Height can be a proxy for "bigger" people.

```
summary(lm(nbeer ~ weight + height, data = beers))
##
## Call:
## lm(formula = nbeer ~ weight + height, data = beers)
##
## Residuals:
      Min
               10 Median
                                    Max
  -8.5080 -2.0269 0.0652 1.5576 5.9087
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) -11.18709 10.76821 -1.039 0.304167
## weight
               0.08530 0.02381 3.582 0.000806 ***
## height
               0.07751 0.19598 0.396 0.694254
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.784 on 47 degrees of freedom
## Multiple R-squared: 0.4807, Adjusted R-squared: 0.4586
## F-statistic: 21.75 on 2 and 47 DF, p-value: 2.056e-07
```

• Height can be a **proxy** for "bigger" people.

#### Let's look at the two models closer!

	Model 1	Model 2			
(Intercept)	-7.021***	-11.187			
	(2.213)	(10.768)			
weight	0.093***	0.085***			
	(0.014)	(0.024)			
height		0.078			
		(0.196)			
Num.Obs.	50	50			
R2	0.479	0.481			
R2 Adj.	0.468	0.459			
F	44.119	21.750			
* p < 0.1, ** p < 0.05, *** p < 0.01					

• Which model do you prefer?

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- Which model do you prefer?
- What happened to the SE for weight? Why?

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Multicolinearity

# Multicollinearity

- If variable x1 and x2 are highly correlated, it is difficult to disentangle their effects
  - Context matters!

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- Limit scenario:  $|Cor(x_1,x_2)|=1$ 
  - Cannot estimate both parameters: One is dropped!

# Multicollinearity

- If variable x1 and x2 are highly correlated, it is difficult to disentangle their effects
  - Context matters!
- ullet Limit scenario:  $|Cor(x_1,x_2)|=1$ 
  - Cannot estimate both parameters: One is dropped!

Can I add both binary variables US\_born and Foreign\_born to a regression?

# Does gender matter?

	Model 1	Model 2	Model 3	Model 4	
(Intercept)	-7.021***	-11.187	-7.830**	-12.067	
	(2.213)	(10.768)	(3.013)	(11.084)	
weight	0.093***	0.085***	0.097***	0.090***	
	(0.014)	(0.024)	(0.018)	(0.027)	
height		0.078		0.079	
		(0.196)		(0.198)	
female			0.528	0.536	
			(1.320)	(1.333)	
R2	0.479	0.481	0.481	0.482	
R2 Adj.	0.468	0.459	0.459	0.449	
* p < 0.1, ** p < 0.05, *** p < 0.01					

#### • Interactions:

• E,g, The relationship between weight and nbeers is different for males and females.

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```
lm(nbeer ~ weight*female, data = beers)
##
## Call:
## lm(formula = nbeer ~ weight * female, data = beers)
##
  Coefficients:
##
     (Intercept)
                         weight
                                        female
                                                 weight:female
       -7.790193
##
                       0.097234
                                      0.225748
                                                      0.002465
```

• How do we interpret these results?

- Other polynomial terms:
  - E,g, The relationship between weight and nbeers is quadratic.

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```
lm(nbeer ~ weight + I(weight^2), data = beers)

##
## Call:
## lm(formula = nbeer ~ weight + I(weight^2), data = beers)
##
## Coefficients:
## (Intercept) weight I(weight^2)
## 0.1078784 -0.0016255 0.0003033
```

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- Categorical variables:
  - E,g, You want to include a factor variable for year the student is in.

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```
table(beers$year)
##
## freshmen junior senior sophmore
##
                  14
                           14
                                    15
lm(nbeer ~ weight + factor(year), data = beers)
##
## Call:
## lm(formula = nbeer ~ weight + factor(year), data = beers)
##
  Coefficients:
##
            (Intercept)
                                       weight
                                                 factor(year)junior
               -6.85394
                                      0.09237
                                                            -0.35244
##
     factor(year)senior factor(year)sophmore
##
               -0.03144
                                      0.07474
##
```

• If your goal is **prediction**:

Overfitting

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• If your goal is **description**:

**Confidence Intervals** 

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Overfitting

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**Confidence Intervals** 

• If your goal is **causality**:

Bias

• If your goal is **prediction**:

Overfitting

• If your goal is **description**:

**Confidence Intervals** 

• If your goal is **causality**:

Bias

### All of them matter!

### References

• Hanck, C. et al. (2020). "Econometrics with R". The Multiple Regression Model