# TDDE07 Bayesian Learning - Lab 4

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### 1. Poisson regression - the MCMC way.

(a)

The  $\beta_{MLE}$  found by fitting a Poisson regression model with glm can be seen in Table 1. Significant covariates are MinBidShare, Sealed, VerifyID and MajBlem.

| _ |   | Const | PowerSeller | VerifyID | Sealed | Minblem | MajBlem | LargNeg | LogBook | MinBidShare |
|---|---|-------|-------------|----------|--------|---------|---------|---------|---------|-------------|
|   | 1 | 1.072 | -0.021      | -0.395   | 0.444  | -0.052  | -0.221  | 0.071   | -0.121  | -1.894      |

Table 1: MLE of beta by glm model fitting

(b)

By approximating the posterior distribution of beta as a multivariate normal and determining the value for  $\beta_{MLE}$  by numerical optimization I got the values seen in Figure 2. They closely resemble the values I got by fitting a Poisson regression model using glm in (a). My implementation of the log of the poisson model with a normal prior can be found in Appendix A.

|   | Const | PowerSeller | VerifyID | Sealed | Minblem | MajBlem | LargNeg | LogBook | MinBidShare |
|---|-------|-------------|----------|--------|---------|---------|---------|---------|-------------|
| 1 | 1.070 | -0.021      | -0.393   | 0.443  | -0.053  | -0.221  | 0.071   | -0.120  | -1.892      |

Table 2: MLE of beta by numerical optimization

(c)

After having implemented the Metropolis Hastings simulation method and simulated 20000 draws from the posterior distribution from (b) I calculated the  $\beta$  coefficients as the mean of the draws, while omitting the first 10% of the draws because of the burn-in phase. These values can be seen in Table 3.

|   | Const | PowerSeller | VerifyID | Sealed | Minblem | MajBlem | LargNeg | LogBook | MinBidShare |
|---|-------|-------------|----------|--------|---------|---------|---------|---------|-------------|
| 1 | 1.070 | -0.021      | -0.394   | 0.442  | -0.054  | -0.224  | 0.067   | -0.122  | -1.895      |

Table 3: Mean values of betas drawn during Metropolis Hastings simulation

The convergence of the parameters can be seen in Figure 1, where I have taken the mean value of every two sequental beta drawn from the posterior and plotted it to visualize the auto-correlation between draws, and to see how the draws asymptotically approaches somewhat stationary values. My implementation of the Metropolis Hastings algorithm can be found in Appendix A.

### **Convergence of beta during Metropolis Hastings**

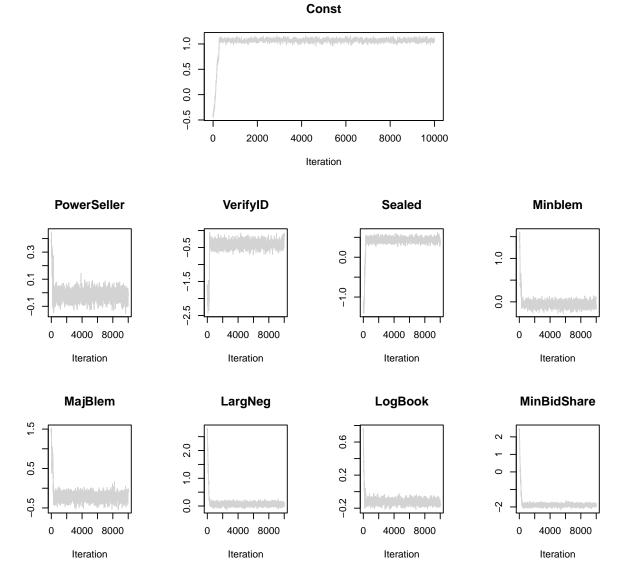


Figure 1: Convergence of betas drawn during Metropolis Hastings simulation

(d)

After having determined the predictive distribution of the sample  $\hat{y}$  as  $p(\hat{y}|\lambda) \sim Poisson(\lambda)$ , where  $\lambda = e^{\hat{X}\beta}$  using the betas draws during the simulation in (c) and  $\hat{X}$  as the sample to predict, I got the distribution seen in Figure 2. The probability that the sample has zero bidders by this distribution was  $p(\hat{y} = 0|\lambda) \approx 0.36$ .

## Predictive distribution of sample

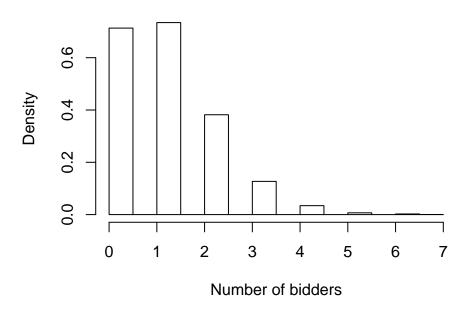


Figure 2: Predictive distribution of sample

### Appendix A

#### Code for Lab 4

```
require(MASS)
require(geoR)
require(mvtnorm)
require(LaplacesDemon)
# -----
# Lab 4
# -----
data = read.table("data/eBayNumberOfBidderData.dat", header=TRUE)
n = length(data)
n_features = ncol(data) - 1 # Except y and const
feature_labels = colnames(data[,2:ncol(data)])
y = data$nBids
X = as.matrix(data[,2:ncol(data)])
X_X = t(X)%*%X
# ----
# (a)
# ----
glm_model = glm(nBids ~ 0 + ., data = data, family = poisson)
pdf("./plots/4_1_1_mle_beta.pdf", width=7, height=7)
par(oma = c(0, 0, 3, 0))
layout(matrix(c(0,1,1,0,2,3,4,5,6,7,8,9), 3, 4, byrow = TRUE))
for (i in 1:ncol(X)){
  mean = glm_model$coefficients[i]
  std_dev = summary(glm_model)[["coefficients"]][,2][i]
  x_grid = seq(mean-4*std_dev, mean+4*std_dev, 0.001)
  plot(x_grid,
       dnorm(x_grid, mean=mean, sd=std_dev),
       type="1",
       ylab="Density",
       xlab=expression(beta),
       main=feature_labels[i])
title("Normal approximation of MLE of beta", outer=TRUE, cex=1.5)
dev.off()
# ----
# (b)
```

```
# Beta prior (Zellner's g-prior)
mu0 = rep(0, n_features)
covar0 = 100 * ginv(X_X)
init_beta = mvrnorm(n=1, mu0, covar0)
# This is the log of the Poisson model
logPostPoiNorm <- function(betas, X, y){</pre>
  log_prior = dmvnorm(betas, mu0, covar0, log=TRUE)
  lambda = exp(X%*%betas)
  \# Assume independence among samples and take the sum of
  # log(p(y_i|lambda)), where lambda is exp(X.dot(beta)) and p \sim Poisson
  log_lik = sum(dpois(y, lambda, log=TRUE))
 return (log_lik + log_prior)
log_post = logPostPoiNorm
opt_results = optim(init_beta,
                    log_post,
                    gr=NULL,
                    Х,
                    у,
                    method=c("BFGS"),
                    control=list(fnscale=-1),
                    hessian=TRUE)
# MLE beta
post_mode = opt_results$par
# Covariance (J^-1(beta hat))
post_cov = -solve(opt_results$hessian)
# ----
# (c)
# ----
Sigma = post_cov
c = .5
n_draws = 20000
metropolisHastings = function(logPostFunc, theta, c, ...){
  theta_draws = matrix(0, n_draws, length(theta))
  # Set initial
  theta_c = mvrnorm(n=1, theta, c*Sigma)
  for(i in 1:n_draws){
    # 1: Draw new proposal theta
    theta_p = mvrnorm(n=1, theta_c, c*Sigma)
```

```
# 2: Determine the acceptance probability
    p_prev = logPostFunc(theta_c, ...)
    p_new = logPostFunc(theta_p, ...)
    acc_prob = min(c(1, exp(p_new - p_prev)))
    # 3: Set new value with prob = acc_prob
    if(rbern(n=1, p=acc_prob)==1){
      theta_c = theta_p
    }
    theta_draws[i,] = theta_c
  }
  return (theta_draws)
init_beta = mvrnorm(n=1, mu0, covar0)
beta_draws = metropolisHastings(logPostPoiNorm, init_beta, c, X, y)
# Calculate mean of batches of 2 draws to visualize the
# auto correlation between sequential draws
mean_draws = matrix(0, n_draws/2, n_features)
for (i in 1:n draws){
  if(i\%2 == 0){
   f = i-1
   t = i
    mean_draws[i/2,] = colMeans(beta_draws[f:t,])
  }
}
# Avoid first 10% of the draws
burn_in = floor(n_draws / 10)
beta_draws = beta_draws[burn_in:nrow(beta_draws),]
beta_means = colMeans(beta_draws)
pdf("./plots/4_1_2_beta_conv.pdf", width=7, height=7)
par(oma = c(0, 0, 3, 0))
layout(matrix(c(0,1,1,0,2,3,4,5,6,7,8,9), 3, 4, byrow = TRUE))
x_grid = 1:nrow(mean_draws)
for (i in 1:ncol(X)){
  plot(x_grid,
       mean_draws[,i],
       type="1",
       ylab="",
       xlab="Iteration",
       col="lightgray",
       main=feature_labels[i])
}
title ("Convergence of beta during Metropolis Hastings", outer=TRUE, cex=1.5)
dev.off()
# ----
```

```
# (d)
sample = c(
  Constant = 1,
  PowerSeller = 1,
 VerifyID = 1,
 Sealed = 1,
 MinBlem = 0,
 MajBlem = 0,
 LargNeg = 0,
 LogBook = 1,
 MinBidShare = 0.5
lambda = exp(beta_draws%*%sample)
pred_draws = rpois(10000, lambda)
# Probability that the sample has O bidders
prob = length(pred_draws[pred_draws == 0]) / length(pred_draws)
pdf("./plots/4_1_3_pred_distr.pdf", width=5, height=4)
# Plot the predictive distribution
plot(hist(pred_draws, right=FALSE, plot=FALSE),
     freq=FALSE,
    xaxt="n",
    xlab="Number of bidders",
    ylab="Density",
     main="Predictive distribution of sample")
axis(1,
     at=0:max(pred_draws),
     labels=0:max(pred_draws)
  )
dev.off()
```