



## Informal notes and discussion points

### **ROADMAP: RandOmised Arthroplasty infection worldWide Multidomain Adaptive Platform trial - notes on design, specification and simulations**

Investigator initiated, Randomised Embedded Multifactorial Adaptive Platform (REMAP) trial, conducted across multiple hospitals in several regions of the world.

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Authors: Mark Jones and James Totterdell

Sponsor:	University of Newcastle, NSW, Australia
Registration (ANZCTR):	todo
HREC	todo
Chief Investigator(s):	Professor Joshua Davis Professor Laurens Manning

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These notes represent items relating to the study specification that should be considered and

discussed prior to finalising the protocol, DSAs and SAP.

Asterisks (\*, \*\*, \*\*\*) as used to denote importance by section, more asterisks implied higher importance.

## **1 Section 1.1 - Background \*\***

Previously this section included a statement on how we should interpret missing elements in the composite as a failure case (or alternative approach required) but it would be good to know and to document (1) how we should be handling this and (2) how likely it is to happen.

The core protocol, see section “6.6 Trial endpoints” does not presently indicate what happens when elements of the composite cannot be identified. It probably should make some statement regarding how missingness is dealt with from an operational perspective or in terms of the impact on interpretation.

## **2 Section 1.3 - Estimands**

### **2.1 ICEs**

Where the treatment policy strategy is used, any ICEs are effectively ignored and/or considered part of the treatment protocol. Where other methods of handling ICEs are used, it is important to recognise that treatment-specific ICEs may occur in isolation for a given unit or may occur in conjunction with ICEs applicable to treatments from other domains. For example, a patient may switch from DAIR to two-stage revision due to having a loose joint but also discontinue extended prophylaxis due to adverse reaction. Both of these events may impact the interpretation of the treatment effect.

### **2.2 Language**

In the core protocol section “6.7.1 Primary estimand” it has “To evaluate, within each relevant cell or combination of cells (as per the DSA), the effect of revealed randomised interventions compared to the domain control, on the probability of treatment success at 12 months after platform entry, in platform eligible participants” Just noting that the phrase “Within each relevant cell” is probably redundant.

### **2.3 Interpretation of effects \*\***

In the Surgical late acute dsa section “9.1.1 Primary estimand” we have “To evaluate, within the domain, the effect of revealed randomised intervention (revision) compared to the domain control (DAIR) on the probability of treatment success at 12 months after platform entry, in domain eligible participants.”

The above is silent on the status of the other variables that contribute to the probability of treatment success, my interpretation is that the underlying goal as one of trying to obtain a marginal effect (averaged over the sample).

## 2.4 Conditional vs marginal OR

Conditional OR: Among two hypothetical individuals with identical covariates  $X = x$ , one given  $A = 1$  and one  $A = 0$ , the odds of death differ by a factor of  $\exp(\beta_A)$ . Preference when making individual level inference or when covariates are effect modifiers.

Marginal OR: If the entire study population were switched from  $A = 1$  to  $A = 0$ , the overall odds of death would change by  $OR_{\text{marg}}$ . More relevant to a policy context or when comparing to an unadjusted analysis.

## 2.5 Redundant text \*\*\*

The estimands are introduced in the core protocol as general outlines and then made specific in the DSAs. For me, I feel like the estimands are not currently providing much value on account of the fact that there isn't enough detail and they do not really seem to consider the implications of ICEs on the existence or interpretation of the outcome. The purpose of an estimand specification is to ensure that the trial answers a very precise clinical question in a way that is clear and interpretable to all stakeholders (e.g., clinicians, regulators, patients), aligns with the trial objectives, and is both statistically valid and estimable from the collected data. If the estimands are inadequately specified, this could lead to criticism.

However, given that the majority of the analyses follow a treatment-policy, I am bit torn as to whether spending the time on clarifying things is going to make any real difference to anyone. One thing we could do to make things easier on ourselves is to minimise duplicating the estimands across documents.

For example, we might remove the estimands in the core protocol, which are not really adding anything, and just retain the DSA definitions. My thinking is that the core protocol specifications do not relate to any single intervention and therefore do not really address the issues that estimands are intended to address, i.e. they do not provide clarity on what effects we are targeting nor on the ICEs as noted above.

### **3 Section 1.3.1.1 - Estimand B.1 (surgical intervention)**

#### **3.1 Update required**

On (Page 19, ROADMAP Surgical Late Acute DSA. V1.1\_01Aug2024.pdf) there is a note in parentheses that states “this will be estimated within hip versus knee strata”. If the intention is to have joint specific focus on the surgical treatment effects then this should be explicitly stated as part of the estimand objective.

Follow up - at the stats meeting dated 9/12/2024, we decided to apply stopping rules to the aggregated hip and knee groups rather than adapting the trial based on joint specific effect. However, both overall and joint specific effects are to be reported (see later). This is not yet noted in a released protocol.

#### **3.2 Surgical treatment switch \*\***

If a patient switches to one-stage (or two-stage) does this patient then enter into the respective duration domain?

#### **3.3 Use of non-randomised duration for inference on surgical \***

Currently, we calculate the effect of revision when d2 and d3 are set to the non-randomised level. This is the only way to ensure bias is induced on the surgical treatment effects.

#### **3.4 ICE**

Other ICEs are possibly relevant but not yet considered. Withdrawal based on clinician opinion on patient's interests with regards to their involvement in roadmap. Even if a patient



is withdrawn by request of the clinician, it may be acceptable to continue to collect data on such a patient, in which case the treatment policy would still be relevant.

Deviations from assigned treatment as variations in the accepted procedures (i.e. different from the protocol definitions) but not leading to switching also constitute an ICE that could impact the interpretation of the treatment effects. An example could be secondary surgery that does not lead to replacement of prosthetic, i.e. not handled via the composite strategy. Deviations from any protocol timings such as a requirement to have a surgical procedure within a given time window from randomisation. Also, it should be considered whether switchover should be interpreted as a failure rather than an ICE. Currently, any additional unspecified ICE would also be handled via the treatment policy strategy

### **3.5 OR vs risk \***

Risk differences are maybe preferable as they are more directly interpretable than odds ratios, albeit a bit more work to the analyst.

The implication is that we need to be explicit about which population we are marginalising over but all that entails is saying that we are averaging over the sample. Typically, this would be the patients in the trial but given our comparisons are very specific, the levels which we use in the g-computation are also specific for some terms. The random effects will probably be taken to be zeros and so our inference will be interpretable for the 'typical' site (or whatever the random effect pertains to).

### **3.6 Switchover \*\***

I am wondering what the implications of switchover are on entry into the other domains. For example, a patient switches from DAIR to 2-stage revision. In theory, after the switch, they are eligible for entry into extended prophylaxis, but randomisation has already decided that they will receive non-randomised treatment for this domain.

#### **4 Section 1.3.1.2 - Estimand D.1 (antibiotic duration intervention) \***

The DSA indicates “Participants will be randomised at platform entry in a fixed 1:1 ratio across the domain.” where “platform entry” is defined as “the timepoint when the patient has met core eligibility criteria, given informed consent for the platform, and been randomised” in the core protocol.

However, for all silos, the decision on the classification of surgery type appears to be made on completion of the surgery when the surgeon indicates that there are no plans for a second stage.

In the DSA for abx duration it states: “Platform eligible participants within any of the silos who have had a single stage revision will be randomised to receive either 6 weeks (short-course antibiotics) or 12 weeks (long-course antibiotics).”

Does that mean that someone who was in the late-acute silo and that was randomised to DAIR, but switched to one-stage revision due to a clinical decision (perhaps because of a loose joint) will then enter into the abx duration domain?

This setup suggests that allocations are generated across all domains for all patients irrespective of whether their dependencies negate the possibility of entry? For example, someone in late-acute is allocated to DAIR, but nevertheless they are also allocated to both the duration domains and the choice domain. I believe this is the case, but very much need to confirm.

So, patients can only be revealed to randomised assignment in the abx duration domain if they survive surgery. Implies start of follow up and treatment initiation not coincident, implies analysis is actually a landmark analysis and should be reflected as such for the relevant domains.

## **4.1 Death during surgery \***

If participants die during surgery (assuming we can obtain a treatment designation as one-stage), they are failure cases and would not be revealed to either of the abx duration.

We are not evaluating duration effects in the population, we are looking at duration effects based on a post-randomisation event, i.e. survival through surgery. We are selecting based on survival. Yes, he said (answering his own question) but the analysis is landmarked and we still have a randomised comparison.

## **5 Section 1.3.1.3 - Estimand E.1 (extended prophylaxis intervention)**

Get more detail on E.1(i2) and check my interpretation is correct. Would the ineligibility actually mean that they do not receive the extended prophylaxis but would still receive extended backbone antibiotic?

## **6 Section 1.3.1.4 - Estimand C.1 (antibiotic choice intervention)**

Clarify C.1(i1) - would this still constitute switchover after the delayed start (or is it rescue)?

As with other domains, patient discharge/relocation is probably relevant as another ICE.

## **7 Section 1.3.2.1 - Estimand B.2 (surgical intervention) \*\***

Secondary estimands have been removed. Principal stratum seems to be a pipe-dream here although throughout the dsa per protocol is conflated with principal stratum.

The manner in which the estimand is currently written, aligns with a per-protocol analysis whereby only those patients that were adherent to their assigned (observed) treatment are retained in the analysis. This doesn't consider whether those patients would also be adherent in their counterfactual state. The problem with this is that it violates the randomisation - the groups might be unbalanced on either known or unknown (or both) confounders. The traditional per-protocol effect is well known to produce biased estimates for treatment effects. What I believe is of interest is the effect of the surgical intervention in those that would have adhered to all arms. Given that the ICEs for the surgical domain negate the possibility of forcing the assignment on the patient, principal strata seems like the most well aligned approach. But this is very hard to implement in our setting. Are there any other options? I am still at a loss and have had no useful input on this. My thinking at the moment is that we just use per protocol in a traditional fashion and accept that the results are likely biased.

### **7.1 ICEs and treatment regimen \*\***

The DSA currently states that principle stratum strategy is to be used as the single mechanism for handling ICEs. Obviously, this is a placeholder as it was a fairly arbitrary inclusion. For the surgical domain currently specified ICEs are switchover to revision for DAIR where the joint is loose and similarly for revision when it is not possible to remove the existing implant. Additionally, revision might be abandoned if the patient becomes unstable during the operation. To handle these via desired strategy we first conceptualise strata. The subset associated with the compliers (however that is defined here) would be the one for which we would estimate the effect of DAIR vs revision. My thinking is that we get rid of these secondary estimands that we would struggle to define correctly anyway.

## **8 Section 1.4.3 - Data management \***

We need to check and write up how the data storage functionality is being decomposed.  
What are the redcap parts and what belongs to spiral?

## 9 Section 2.1.2 - Per-protocol \*\*

Principal stratum is referred to as a per-protocol in the DSAs but this is incorrect.

Per-protocol analyses traditionally exclude those patients that deviate from the protocol and conduct the primary analysis with this subset. The results from such analyses will only be valid where deviations from the protocol are non-informative.

Consider (1) exclusions can imbalance treatment groups on known or unknown confounders and prognostic factors (2) that noncompliance may relate to a specific intervention or disease severity producing differential dropout across arms.

Specific questions on the protocol definitions:

- For the 6wk and 12wk antibiotic duration assignment, is there a reqt for no abx after 16 wks?
- For the extended prophylaxis, is there a reqt for no abx after 90 days?
- For the Rifampicin, is there any requirement for the interval between the doses?
- Are we only considering per-protocol with respect to treatment?

Recommend removing the reference to principal stratum in the estimands.



## 10 Section 2.2 - Subgroups \*\*\*

Titles should be updated to subgroup analyses (rather than the current “secondary analyses”) in the following documents:

DSA Surgical Late-acute section 9.6 DSA for AB duration part A section 9.6 DSA AB Duration part B section 8.6 (toc does not link correctly) DSA AB Choice section 9.6 (this has both secondary and subgroup analyses, keeping a clear separation between the two would be helpful)

Duration of symptoms should be categorised for surgical domain.

For surgical domain, categorise time from implantation or get rid of.

For backbone duration, discuss removing ideal procedure vs not as this suggests a protocol deviation.

For ext proph, categorise duration between first and second and duration of antibiotic treatment between first stage and reimplant.

For ab choice, include silo or was that purposefully excluded as rif is expected to work by the same action?

There was a review question at the top of the subgroup section for which clarification is needed - *is it obvious what this means (in the parameter estimation)?*

For one-stage vs two-stage - this is a post-randomisation subgroup, so could be influenced by previously revealed assignments, eg rifampicin. AB Choice is revealed and administration commenced following the first procedure, i.e. after a DAIR, one-stage or between the first and second stage of a two-stage revision and therefore, the reveal of AB choice will not influence selection of surgery type.

## **11 Section 2.3 - Descriptive summaries \***

Query how the consent process works for say AB choice where eligibility is not assessed until after the initial surgical treatment? Do we need to think through consort...?

## 12 Section 2.4 - Sequential analyses

Revisit the timing of interim analyses as enrolment always seems to be a major challenge.

I think it is probably more operationally practical to run interims on a regular timing rather than for fixed sample size, but I still generally prefer the latter for simplicity of understanding in the design operating characteristics.

The subgroup analysis at each interim used to determine whether the trial should continue is problematic because

1. it opens us up to severe criticism of un-specified decisions occurring
2. operating characteristics will deviate from the existing plan

## **13 Section 2.6 - Primary analysis \***

Revisit the baseline characteristics to be included in the model. Reference some of the sub-groups for ideas.

## **14 Section 2.9 - Supportive analyses**

For logistic regression models, update treatment effect to risk difference so that we are consistent with the primary analysis.

## **15 Section 2.9.1 - Treatment success \***

Should we move treatment success from domain specific to domain agnostic analyses and assumed to replicate the primary analysis with the exception of using a per protocol population (per protocol across all domains is what has been assumed)?

## 16 Section 2.9.2 - DOOR

The PI (door probability) measure is NOT the probability that a patient will benefit from having the experimental treatment rather than control, which can generally only be obtained from a cross-over trial. A PI of 0.7 means that if you randomly select a patient from the experimental group and another from the control group, there is a 70% chance the experimental group patient will have a better outcome. This is not the same as saying that 70% of patients will benefit from the experimental treatment. For further cautions, limitations and critique on interpretation see Senn ([2011](#)).

A Bayesian approach to the MWU does exist, but has not been widely adopted to date and so we have opted for the above approach. Chechile ([2020](#)) provides details on the Bayesian perspective.

For the strata specific derivations, should the surgical domain revision group be restricted to those that entered duration and extended prophylaxis as non-randomised? Not doing this may lead to bias in the surgical domain in the scenario where there is no effect of revision but an effect within the duration or extended prophylaxis domain that will only be apparent for those that receive revision.

*The estimand in the core protocol assumes a cumulative logit analysis whereas in the sap we assume a 'standard' door analysis.*

## **17 Section 2.9.3 - Patient reported joint function \***

Would a conversion to a mean difference be conceivable based on the distribution of scores or is this too complicated given the conditional model? Currently the treatment effect is in terms of the cumulative odds ratio.



## **18 Section 2.9.7 - Clinical cure to 12 months \***

Update the estimand measure of effect to a risk difference.

Should this be approached as stated or was the intention to restrict this to the subset that don't die and don't get their initial prosthetic replaced. Personally, I think it should be based on the full analysis set because this allows the result to be compared with the primary analysis and it preserves randomisation.

## **19 Section 2.9.12 - Time alive and free from revision procedure to 24 months \*\***

The M-spline should include an intercept so that  $h(t)$  is not constrained to zero at the origin and the coefficients are constrained to a simplex for identifiability.

Note to self - go back and review M-spline vs B-spline differences in approach, stability, fit for purpose and so on.

The sap deviates from the current estimand specification A.12 (Pg 39 ROADMAP Core Protocol\_V1.1\_01Aug2024\_clean.pdf) median difference in time to revision. Was this intended to be median survival?

## **20 Section 2.10 - Supportive (domain specific) analyses**

Get the DSA estimand specifications update to indicate that these are all up to 12 months if they haven't already. Double check anyway.

## **21 Section 2.10.1.1 - Unplanned re-operation \***

Get update on whether this outcome is to 12 months.

## **22 Section 2.10.2.1 - Acute liver injury following platform entry to 90 days \***

Check whether this was meant to be 90 days **post commencement of antibiotics** rather than post platform entry, the former requiring a different approach.

## **23 Section 2.10.2.3 - Laboratory-proven *Clostridium difficile* diarrhoea to x days \*\***

Check whether this should have a time limit on it. Presumably also 90 days...

## **24 Section 2.10.2.4 - Antibiotics ceased due to suspected adverse reaction to 90 days**

Check whether this was meant to be 90 days post commencement of antibiotics.

What is meant by “ceased due to other suspected adverse reaction”? Specifically, what is with the ‘other’ prefix? This question also applies to the equivalent estimand in the extended proph domain.

## **25 Section 2.10.3.1 - Time alive and free from any revision procedure to 12-months \*\***

This seems like it would be domain agnostic, why is it specific to the extended prophylaxis domain?



## **26 Section 2.10.3.2 - Time alive and free from any revision procedure to 24-months \*\***

As above, this seems like it would be domain agnostic. Get clarification.

## **27 Section 2.10.3.4 - Antibiotics ceased due to suspected adverse reaction to 12 months**

Again, what is meant by other?

## **28 Section 2.10.4.1 - Acute liver injury to day 100**

Check whether this should be **post commencement of rif.**

**29 Section 2.10.4.3 - Laboratory-proven *Clostridium difficile* diarrhoea to x days \*\***

Check whether this should have a time limit on it.

### **30 Section 2.10.4.4 - Antibiotics ceased due to suspected adverse reaction to 100 days**

Check why this is 100 days and others are 90. Should these all be made the same? Also check which antibiotics we are talking about here. Is it just rif under consideration or does it apply across all antibiotics, in which case isn't this a domain agnostic question?

## 31 Section 3.1 - Treatment effects

### 31.1 Reference notes to self

To expand on the above, assume that when a unit is randomised to revision and the clinical decision of what type of revision to do is based solely on disease severity with more severe units receiving one-stage and those that are less severe receive two-stage. This violates the exchangeability assumption in that if  $E[Y|T = 1] = y_1$  and  $E[Y|T = 2] = y_2$  (with  $Y$  denoting the observed outcomes and  $T = 1, T = 2$  denoting receipt of one and two-stage) then we would not obtain the same expectations if the sets of treatment units were switched. In other words,  $E[Y(1)|T = 1] \neq E[Y(1)|T = 2] \neq E[Y(1)]$  and  $E[Y(2)|T = 1] \neq E[Y(2)|T = 2] \neq E[Y(2)]$ . To remedy this, we adjust for a covariate set (in this example, disease severity but in the primary analysis we adjust for revision type preference and additional patient characteristics) with the goal of achieving conditional exchangeability.

## 32 Section 4.1 - Decision procedures overview

Unsure whether it is sensible to be examining all subgroups in order to determine the outcome of futility, especially given the number of subgroups to be evaluated.

Currently there is a placeholder that states: *For a futility decision to hold definitively, all subgroups within the relevant domain should show futility, consistent with the decision thresholds as detailed above. Any subgroups that are not futile may be permitted to continue to enrol into the study. For example, if revision is shown to be futile relative to DAIR in the main comparison, but the subgroup analysis indicates that revision is not futile for those participants where the infection site is the hip (rather than knee), then randomised assignment in the surgical domain may continue for participants with hip infections. Variations to this are permitted at the DSMC disgression.*

Personally, I think this is a bad idea. Subgroup effects are more variable due to the small cell size so it is reasonable to think that we will see random highs and lows that are not consistent with the result overall. This could mean that we see futility on average but neither of the subgroups is futile indepently. This would create ambiguity for the dsmc that they may not be equipped to resolve. Such an approach also somewhat negates the simulation results of the operating characteristics.

Chechile, R., 2020. A bayesian analysis for the mann-whitney statistic. Communications in Statistics 49, 670–696. <https://doi.org/10.1080/03610926.2018.1549247>

Senn, S., 2011. U is for unease. American Statistical Association 3. <https://doi.org/10.1198/sbr.2010.10024>