802 Project Summary

Maksuda Aktar Toma, Jo Charbonneau, Ryan Lalicker December 15, 2024

Introduction

This paper summarizes the consulting that was done for our assigned STAT 802 group. For more information on the experiment, the data, or any other files used in this paper see our Github page which can be found at https://github.com/maksudatoma/Stat-802-Project. The coding languages used in the paper are R and SAS. The corresponding code can be found in Appendix A - R Code and Appendix B - SAS Code and Additional Output respectively.

Initial Meetings

The first meeting with our clients was on September 13th. We discussed their project and what kind of data they were going to be looking at. They detailed to us their project, which is looking at the levels of salmonella in beef jerky at different inoculations and thicknesses. Prior to the meeting they sent us what their variables would be, which gave us a good idea of what might be the best experimental design. The group informed us they were avoiding a completely randomized design (CRD) at the request of their professor. With that in mind, we suggested other possible models.

Later, after receiving feedback from Dr. Howard and several PhD students within the statistics department, we suggested adding a time component to the experiment and creating multiple batches to replicate each treatment combination. This lead to us suggesting a mixed model for the analysis.

In both the initial meeting and the follow-up session the clients were more than happy to implement our suggestions. In the end the experiment involved two thickness levels (one-fourth and one-eighth of an inch), two inoculation methods (dry and wet), and five evenly spaced time points where measurements were taken (weeks 1-5) creating twenty entries per batch. The exact number of batches would not be known until after the power analysis, found in the *Power Analysis* section. We provided the client an example dataset we created to give them a better idea of what the end product may look like. This dataset had five batches.

Study Objectives and Proposed Model

The clients were most interested in the effect of the thickness levels, the inoculation method, and their interaction had on the salmonella levels. In the final model we included the week effect and subsequent interactions as well. These variables are the fixed effects in the proposed mixed model.

The other variable included in the experiment is the batch number. This is therefore treated as a random variable. As mentioned above, the exact number of batches needed was unknown prior to the power analysis, but five was used as a starting value.

Overall the study employs a 2×2 factorial design with two main factors: Inoculation Method (Dry, Wet) and Thickness (1/4-inch, 1/8-inch). Repeated measurements are taken over five equally spaced time points (Weeks 1 to 5), allowing the analysis of both main effects, their interaction, and changes over time.

The model can be written in the form

$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \tau_k + (\alpha\beta)_{ij} + (\alpha\tau)_{ik} + (\beta\tau)_{jk} + (\alpha\beta\tau)_{ijk} + u_l + e_{ijkl}$$

Here, Y_{ijkl} is the salmonella level and μ is the overall mean. The fixed effects are represented by α_i for the effect of the ith inoculation method, β_j for the effect of the jth thickness level, and τ_k for the effect of the kth week. The interaction effect of the ith inoculation method and the jth thickness level is represented by $(\alpha\beta)_{ij}$, with the other two-way interactions following this form. The three-way interaction between all fixed effects is represented as $(\alpha\beta\tau)_{ijk}$. The random effect for batches is represented by u_l , which we assume are distributed as $u_l \sim N(0, \sigma_u^2)$. Lastly, the residuals are represented by e_{ijkl} , which we assume can be distributed as $e_{ijkl} \sim N(0, \sigma^2)$.

Power Analysis

Before power analysis, We reached out to the client later on in the process to determine what contrasts they were most interested in testing. They expressed they wanted to see the difference between the two levels of the inoculation method, the two levels of the thickness, and the orthogonal contrasts these variable. This resulted in six contrasts being tested.

To determine the necessary number of batches needed to increase the likelihood of detecting a true treatment effect, we performed a power analysis. To do this, probable treatment mean estimates across all five weeks and variance estimates were needed. The clients provided these metrics from Brown et al. (2024). We then used these metrics to create a dataset with five batches where the response variable was identical across the batches. This dataset was then evaluated to determine the power. The results of the power analysis perfomed in SAS are shown below.

Obs	Label	NumDF	DenDF	FValue	ProbF	Effect	ncparm	alpha	fcrit	power
1	Dry vs Wet	1	76	9.23	0.0033		9.226	0.05	3.96676	0.85058
2	1/4 vs 1/8 inches	1	76	103.28	<.0001		103.276	0.05	3.96676	1.00000
3	Dry vs Wet at 1/4 Inches	1	76	5.70	0.0195		5.695	0.05	3.96676	0.65406
4	Dry vs Wet at 1/8 Inches	1	76	44.65	<.0001		44.651	0.05	3.96676	1.00000
5	1/4 vs 1/8 inches for Dry inoculation	1	76	137.36	<.0001		137.364	0.05	3.96676	1.00000
6	1/4 vs 1/8 inches for Wet inoculation	1	76	7.03	0.0097		7.031	0.05	3.96676	0.74475
7		1	76	9.23	0.0033	Inoculation_Method	9.226	0.05	3.96676	0.85058
8		1	76	103.28	<.0001	Thickness	103.276	0.05	3.96676	1.00000
9		1	76	41.12	<.0001	Inoculatio*Thickness	41.120	0.05	3.96676	0.99999
10		4	76	7.82	<.0001	Week	31.280	0.05	2.49205	0.99659
11		4	76	13.77	<.0001	Inoculation_Met*Week	55.092	0.05	2.49205	1.00000
12		4	76	20.78	<.0001	Thickness*Week	83.132	0.05	2.49205	1.00000
13		4	76	4.04	0.0050	Inocula*Thickne*Week	16.162	0.05	2.49205	0.89525

Figure 1: Results of power analysis.

The first six rows of the table correspond to the contrasts the clients were interested in testing, while the bottom seven rows are measuring the fixed effects of the model. Many of the terms have more than 80% power. Specifically the fixed effects were all high enough for both the clients and ourselves to feel comfortable using five batches. Two of the orthogonal contrasts, Dry vs Wet at 1/4 Inches and 1/4 vs 1/8 inches for Wet inoculation did have lower power scores, but after talking with both the clients and Dr. Howard about them, we felt comfortable to proceed.

Simulating Data

After finding the necessary number of batches, which was five, we proceeded with simulating the data. The estimated treatment means and variances provided by the client were used in the simulation as well. We then reviewed the simulated dataset for major issues, such as negative response values, and reran the power analysis on the new data set to ensure everything was working properly. After finding no problems with the dataset, we sent it to the clients. Note, the simulation was performed in SAS.

Data Analysis

Summary Statistics

As part of the project, we analyzed the simulated dataset. Before fitting out model to the dataset, we first wanted to explore some of the variables. Figure 2 shows the mean values and

standard deviations for each treatment combination. We can see the changes in mean values are small, so further exploration and analysis are needed.

Table 1: Summary of response rariable across factors and weeks

Inoculation Method	Thickness	Week	Mean Response	SD Response	Count
Dry	1/4-inch	1	4.199	0.187	5
Dry	1/4-inch	2	4.137	0.163	5
Dry	1/4-inch	3	4.187	0.294	5
Dry	1/4-inch	4	4.306	0.272	5
Dry	1/4-inch	5	4.645	0.178	5
Dry	1/8-inch	1	4.801	0.173	5
Dry	1/8-inch	2	4.801	0.238	5
Dry	1/8-inch	3	4.596	0.235	5
Dry	1/8-inch	4	4.425	0.206	5
Dry	1/8-inch	5	4.876	0.177	5
Wet	1/4-inch	1	4.240	0.307	5
Wet	1/4-inch	2	4.523	0.236	5
Wet	1/4-inch	3	4.534	0.173	5
Wet	1/4-inch	4	4.357	0.421	5
Wet	1/4-inch	5	4.299	0.200	5
Wet	1/8-inch	1	4.697	0.181	5
Wet	1/8-inch	2	4.809	0.327	5
Wet	1/8-inch	3	4.530	0.215	5
Wet	1/8-inch	4	4.223	0.361	5
Wet	1/8-inch	5	4.151	0.295	5

Figure 2: ?(caption)

Distribution of response variable

Before continuing our investigation into the relationships among the treatment variables, we want to look into the response variable (salmonella levels). Specifically, we want to see how it is distributed. Figure 3 shows a histogram and Q-Q plot of the response variable in the left and right plots respectively. While the histogram shows a slight potential skew, this is not enough for us say the distribution is non-normal. Furthermore, the Q-Q plot indicates the response variable follows a relatively normal distribution.

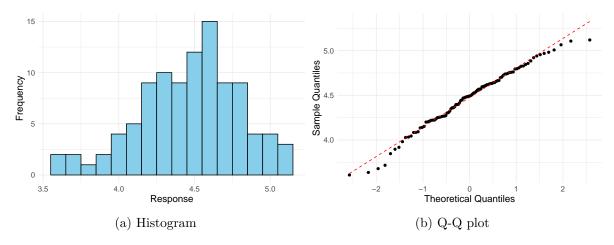


Figure 3: Histogram and Q-Q plot of the response variable.

Exploring the Data

Now we will graphically look at how the different variables of the model impact the response variable. Figure 4 shows two interaction plots with one being for samples cut to 1/4-inches and the other for 1/8-inch samples. Each one shows the relationship between the two inoculation methods over the five weeks for the respective thickness. At the 1/4-inch level, the dry and wet methods start at a similar level but then change over time. The dry method appears to stay a lower level for the first few weeks, but then rise sharply. The opposite trend occurs for the wet method. At the 1/8-inch level, the two methods are similar until the fifth week when the measurement of the wet method decreases and the dry method increases. In both plots we seem to have interaction among the the variables, so that will need to be investigated further.

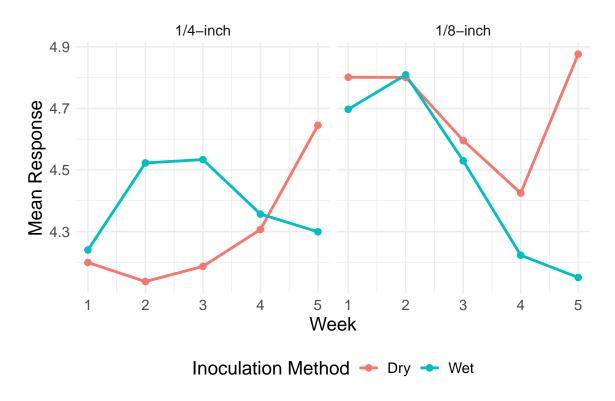


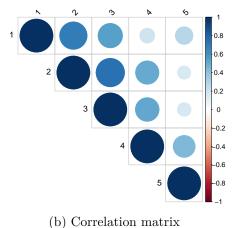
Figure 4: Interaction plots of inoculation method and time for each thickness level.

Another element to consider is correlation over time. Since this is a repeated measures experiment, we need to account for this correlation by selecting a type of covariance structure. There are several types of structures, such as variance components (VC), unstructured (UN), compound symmetry (CS), heterogeneous compound symmetry (CSH), p-order auto-regressive (AR(p)), and heterogeneous auto-regressive (ARH(p)), p-order ante-dependence (ANTE(p)), and Toeplitz (TOEP) among others. For more information on these, see Lipka and Tyner (2004) and USDA (n.d.). The AR(p) structure fits data that is ordered equally spaced through time. For that reason, our initial plan was to use this as the covariance structure with p=1.

To see if this first-order auto-regressive structure might fit the data, let's consider the table and plot in Figure 5. The correlation matrix (left) and plot (right) show the relationships between repeated measurements over weeks one to five. Strong correlations are observed between adjacent weeks (e.g., Week 1 vs. Week 2, r=0.69, Week 2 vs. Week 3, r=0.74), indicating temporal dependency. Correlations weaken as the time gap increases (e.g., Week 1 vs. Week 5, r=0.29), leading us to believe orders of p > 1 are unnecessary. This can be seen visually in the plot, which uses circle size and color to model the correlation metrics seen in the matrix. We can see as the gap between weeks increases, the circles become smaller and lighter. This pattern supports the use of models like AR(1).

Time	1	2	3	4	5
1	1.00	0.69	0.55	0.21	0.29
2	0.69	1.00	0.74	0.50	0.16
3	0.55	0.74	1.00	0.52	0.17
4	0.21	0.50	0.52	1.00	0.45
5	0.29	0.16	0.17	0.45	1.00

(a) Correlation matrix



()

Figure 5: Table and plots to see the correlation across time.

Model Comparison

After exploring the data we can move on to fitting the model. While we were confident in using the AR(1), we chose to fit the model using other covariance structures as well so we could see how the fit compares. The results are shown in Figure 6. For each of the fit statistics in this table a lower score is better, even when looking at negative values. (Bobbit (2021)). This means the model fit using an AR(1) structure had the best AIC and AICC scores and a respectable BIC score. This verifies our choice in the AR(1) covariance structure.

Model	AIC	AICC	BIC
VC	-12.25	-12.09	-13.03
UN	-12.22	-3.58	-18.47
CS	-11.60	-11.29	-12.77
AR(1)	-12.86	-12.55	-14.03
ARH(1)	-12.56	-11.01	-15.30
ANTE(1)	-8.57	-5.39	-12.48
TOEP	-7.34	-6.19	-9.68

Figure 6: Model comparison table.

After verifying the fit of the AR(1) covariance structure for the repeated measures, we needed to see if the assumptions for a linear mixed model are violated or not. These include the residuals being normally distributed and homogeneous. The plots in Figure 7 allow us to evaluate these assumptions. To test normality graphically, we can look at the histogram (top right) and the Q-Q plot (bottom left). These both appear approximately normal, indicating the assumption holds. The boxplot (bottom right) can also show normality and potential outliers. It appears there is one outlier, but the normality assumption still holds. The residual plot (top left) allows us to check if the homogeneous assumption holds, and it appears to be since the points seem somewhat randomly distributed with no clear pattern. Since the assumptions are holding, we can proceed with the linear mixed model.

Model Output

We can now look at some of the output from fitting the linear mixed model in SAS. Figure 8 shows three tables from this output. The first is the *Fit Statistics* table, where the metrics used in Figure 6 come from. These metrics indicate the model's fit, which seems fine to us.

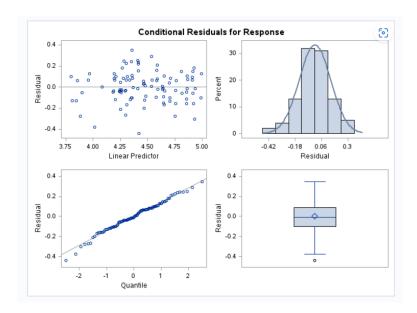


Figure 7: Residual plots for checking assumptions.

time points. The small values for the estimated residual variance (0.02664) are good news as it means there is minimal unexplained variability in this model.

			Fit Statist	ics					
		-2 Res Lo	og Likeliho	od	-18	.86			
		AIC (sma	iller is bette	ler is better)		.86			
		AICC (sn	naller is bet	aller is better)		.55			
		BIC (sma	iller is bette	-14	.03				
		CAIC (sn	naller is be	tter)	-11	.03			
		HQIC (sn	naller is be	tter)	-16	.01			
		Generali	zed Chi-Sq	uare	2	.13			
		Gener. C	hi-Square /	DF	0	.03			
	Covariance Parameter Estimates								
	Cov Parm	Subject		Estim		ate		dard Error	
	Intercept	Batches			0.038	302 0.0		2755	
	AR(1)	Batche*I	nocul*Thick	n -0.214		140	0.	1286	
	Residual				0.02664		0.004423		
		Type III	Tests of Fi	xed E	ffect	s			
Eff	fect		Num DF	Den	DF	F١	/alue	Pr>	F
Inc	oculation_M	ethod	1		76		4.98	0.02	86
Th	ickness		1		76	8	31.92	<.00	01
Inc	oculatio*Thic	ckness	1		76	3	32.65	<.00	01
We	eek		4		76		6.12	0.00	03
Inc	oculation_M	et*Week	4		76	1	17.90	<.00	01
IIIC		4		76	-	14.59	<.00	n 1	
	ickness*We	ек	4		70			~.00	٠.

Figure 8: Fit Statistics, Covariance Parameter Estimates, and Type III Tests of Fixed Effects tables.

The last table in Figure 8 is the Type III Tests of Fixed Effects table. This allows us to see if the fixed effects are significant by looking at the p-values reported in the Pr > F column. We must first consider the significance of the interaction effects and only look at the main effects if the interaction effects are insignificant. While the three-way interaction between the inoculation method, thickness, and week is not significant, all three two-way interactions between these variables are highly significant (p<0.0001). Therefore, we need to consider simple effects. Overall we can say each respective two-way interaction does have an impact on the data. This validates what we saw in Figure 4.

Figure 9 shows the output for the contrasts we evaluate at the client's request. While both the inoculation method and thickness are significant, this is just repeating the main effects for each variable that we saw in Figure 8. Therefore, we need to focus on the interaction terms. All interaction contrasts are significant, with Dry vs. Wet at 1/8 inches and 1/4 vs. 1/8 inches for Dry inoculation being highly significant. This requires further investigation.

Contrasts									
Label	Num DF	Den DF	F Value	Pr > F					
Dry vs Wet	1	76	4.98	0.0286					
1/4 vs 1/8 inches	1	76	81.92	<.0001					
Dry vs Wet at 1/4 Inches	1	76	6.06	0.0161					
Dry vs Wet at 1/8 Inches	1	76	31.56	<.0001					
1/4 vs 1/8 inches for Dry inoculation	1	76	108.99	<.0001					
1/4 vs 1/8 inches for Wet inoculation	1	76	5.57	0.0209					

Figure 9: Contrasts table.

Since we found three significant two-way interactions we will need to look at the *Least Squares Means* output for each interaction. The first one we are looking at is for the interaction between inoculation method and thickness. This can be seen in Figure 10. These tables allow us to see how the variables interact at each level.

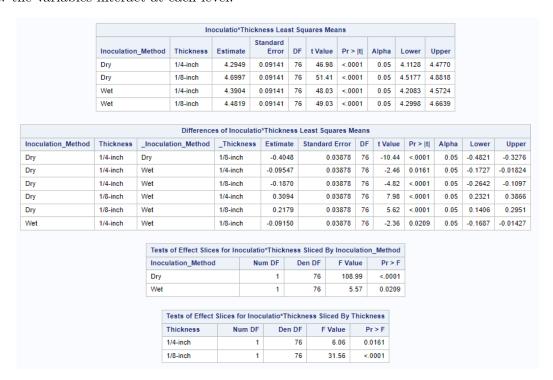


Figure 10: Least Squares Means output for inoculation method and thickness interaction.

We can see from the *Least Squares Means* table that all the combinations are significant, but we want to look closer at the differences in the second table. Once again, all the differences between each possible combination of the inoculation method and thickness are significant since the largest p-value is 0.0209. We can see the largest difference is between the combinations

of 1/4-inch with dry inoculation and 1/8-inch with dry inoculation. The negative difference implies the 1/8-inch pieces have significantly lower values, which is what we want. The smallest difference involved the two thicknesses with wet inoculation. This indicates both perform in a similar way, with the 1/8-inch pieces performing slightly but significantly better.

Innoviation Mathed	Week	Estimate	Standard Error	DF	t Value	Pr > Iti	Alaba	Lower	Unna
Inoculation_Method	vveek	Estimate	EIIOI		t value	PLSIU	Alpha	Lowel	Uppe
Dry	1	4.5002	0.1013	76	44.41	<.0001	0.05	4.2984	4.702
Dry	2	4.4690	0.1013	76	44.10	<.0001	0.05	4.2671	4.670
Dry	3	4.3912	0.1013	76	43.33	<.0001	0.05	4.1894	4.593
Dry	4	4.3655	0.1013	76	43.08	<.0001	0.05	4.1637	4.567
Dry	5	4.7607	0.1013	76	46.98	<.0001	0.05	4.5589	4.962
Wet	1	4.4685	0.1013	76	44.10	<.0001	0.05	4.2666	4.670
Wet	2	4.6659	0.1013	76	46.05	<.0001	0.05	4.4641	4.867
Wet	3	4.5317	0.1013	76	44.72	<.0001	0.05	4.3299	4.733
Wet	4	4.2898	0.1013	76	42.33	<.0001	0.05	4.0879	4.491
Wet	5	4.2248	0.1013	76	41.69	<.0001	0.05	4.0230	4.426

Figure 11: Least Squares Means table for inoculation method and week interaction.

Moving on to the inoculation method and week interaction, Figure 11 shows us that the estimated response variables differ for the treatment combinations. The largest estimate is for the dry samples in week 5 (4.7607), while that same week is the lowest estimate for the wet samples (4.2249). Additionally, week 2 has the highest estimate for the wet samples. Similarly, Figure 12 shows that the interaction between the two variables is most significant at week 5 (p-value<0.0001), followed by week 2 (p-value=0.0086). This suggests the impact of the inoculation method depends on the week of measurement.

Figure 15, which shows the differences among the different combinations of inoculation methods and weeks, can be found in *Appendix B - SAS Code and Additional Output*. Note that the second table in Figure 12 comes from the comparison of the dry versus wet method for a given week found in Figure 15. For the comparisons involving different weeks, most of the significant differences involve cases where there are two or more weeks between the treatment combinations, but some of the largest significant estimated differences occurred in weeks 4 and 5

The last of the three two-way interactions is for the thickness and the weeks. Figure 13 shows that the largest estimated response variable is 1/8-inch samples in week 2 (4.8049), followed by the same thickness the week prior (4.7491). Meanwhile, the 1/4-inch samples over the five weeks account for five of the six lowest estimates. This indicates that 1/4-inch samples are generally better at slowing bacterial growth over time and are also more consistent than the 1/8-inch pieces.

Adding to this, Figure 16 in Appendix B - SAS Code and Additional Output shows the differences among the various combinations of the thicknesses and the weeks. The differences

Inocu	ulation_Meth	od	Num DF	Den DF	F Value	Pr > F				
Dry			4	76	9.04	<.0001				
Wet			4	76	14.99	<.0001				
				BB-441871-1	Ni					
	Tests of Effect Slices for Inoculation_Met*Week Sliced By Week									
	Tests of E	ffect Slices for	Inoculation	_lviet~vveek :	sliced by we	ek				
	Tests of E	fect Slices for Num DF	Den D							

76

76

76

76

1

1

1

1

0.0086

0.0579

0.3026

<.0001

7.28

3.71

1.08

53.91

2

3

5

Figure 12	: Test of Effect Slices table for the two-way interaction with respect to the inoculation
	method and the week respectively.

	Thickness*Week Least Squares Means											
Thickness	Week	Estimate	Standard Error	DF	t Value	Pr > t	Alpha	Lower	Uppe			
1/4-inch	1	4.2195	0.1013	76	41.64	<.0001	0.05	4.0177	4.4214			
1/4-inch	2	4.3300	0.1013	76	42.73	<.0001	0.05	4.1282	4.5318			
1/4-inch	3	4.3601	0.1013	76	43.03	<.0001	0.05	4.1583	4.5619			
1/4-inch	4	4.3315	0.1013	76	42.75	<.0001	0.05	4.1297	4.5333			
1/4-inch	5	4.4721	0.1013	76	44.13	<.0001	0.05	4.2702	4.6739			
1/8-inch	1	4.7491	0.1013	76	46.87	<.0001	0.05	4.5473	4.9510			
1/8-inch	2	4.8049	0.1013	76	47.42	<.0001	0.05	4.6030	5.0067			
1/8-inch	3	4.5628	0.1013	76	45.03	<.0001	0.05	4.3610	4.7646			
1/8-inch	4	4.3238	0.1013	76	42.67	<.0001	0.05	4.1220	4.525			
1/8-inch	5	4.5135	0.1013	76	44.54	<.0001	0.05	4.3116	4.715			

Figure 13: $Least\ Squares\ Means$ table for thickness and week interaction.

involving combinations that hold one of the two variables constant (eg. 1/4-inch at week 1 versus 1/8-inch at week 1 where the week is constant) are shown in Figure 13. The second table in Figure 13 shows us the interaction of these variables is most significant in the first two weeks (p-value<0.0001), followed by the third week (p-value=0.0069). When looking at these same cases in Figure 16, we can see negative estimates, implying the 1/4-inch thickness returns a lower response value for the first three weeks. Weeks 4 and 5 are not significant though. Additionally, in the first table, we can see the interaction at the 1/8-inch level is highly significant (p-value<0.0001), while the 1/4-inch level has a higher p-value (0.0179) but is still significant.

Thickness	Num Di	Den Di	F Value	e Pr>
1/4-inch		4 70	8 3.19	0.017
1/8-inch	-	4 70	8 17.53	<.000
1	Num DF	Den DF	52.64	<.0001
Week	Num DF	Den DF	F Value	Pr > F
2	1	76	42.32	<.0001
3	1	76	7.71	0.0069
4	1	76	0.01	0.9164

Figure 14: Test of Effect Slices table for the two-way interaction with respect to the thickness and the week, respectively.

Recommendation

In terms of how the results of this experiment would affect a producer looking to sell beef jerky we need to make a key assumption. Since they would obviously not be inoculating it with salmonella, let's assume the dry and wet inoculation methods reflect storage conditions. With that in mind, our recommendation depends on a few factors.

If a producer will have both dry and wet storage methods, we suggest prioritizing cutting the jerky to 1/4-inch. This length was more consistent and had generally lower estimates than the 1/8-inch slices did. On the other hand, if a producer wants to use both thicknesses, we recommend more wet storage conditions, as the wet inoculation method saw lower estimates of bacterial growth over time. Of course, this is simulated data so these recommendations are not meant for real-world application.

Conclusion

This report highlights the work done in the consultation of our STAT 802 clients. We first met with our group to learn more about the experiment they wished to perform. After consulting with others we were able to provide recommendations for the framework of the experiment, including a model. By completing a power analysis using data provided by the client we were able to verify the necessary requirements for the size of the experiment in terms. Then we carried out a simulation to create a dataset that mimics the results of a real experiment. We were then able to explore and analyze the data using R and SAS.

Acknowledgement

We want to thank Dr. Reka Howard, Malith Premarathna, and other members of the SC3L staff for helping us with the project. We're also grateful to the people of the STAT 802 group for sharing their study ideas and letting us work with them on a project in their area of expertise.

References

- Bobbit, Zach. 2021. "How to Interpret Negative AIC Values." Statology. 2021. https://www.statology.org/negative-aic/.
- Brown, Jessica, Taylor Bannon, Arie Havelaar, Chad Carr, Barbara B. Kowalcyk, and Jason M. Scheffler. 2024. "Inactivation of Salmonella Enterica, Escherichia Coli O157:H7, and Campylobacter Jejuni in a Restructured Beef Jerky Developed for Production in Ethiopia." Meat and Muscle Biology 8 (April): 1–12. https://doi.org/10.22175/mmb.16091.
- Lipka, Alex, and Benjamin Tyner. 2004. "Repeated Measures Covariance Structure." Statistical Consulting Service Purdue University. 2004. https://www.stat.purdue.edu/scs/help/method/Repeated%20measures%20analysis.pdf.
- USDA. n.d. "Appendix e: Selecting an Appropriate Covariance Structure," n.d., E.1–8. https://www.ars.usda.gov/ARSUserFiles/80000000/StatisticsGroupWebinars/Appendix%20E%20-%20Selecting%20a%20Covariance%20Structure.pdf.

Appendix A - R Code

```
library(knitr)
library(dplyr)
library(ggplot2)
library(naniar)
library(reshape2)
library(GGally)
library(janitor)
library(emmeans)
library(MASS)
library(multcomp)
library(lme4)
library(nnet)
library(tidyr)
library(knitr)
library(kableExtra)
library(corrplot)
data <- read.csv("rptm_simulation.csv")</pre>
response_summary <- data %>%
  group_by(Inoculation_Method, Thickness, Week) %>%
  summarise(
    Mean_Response = mean(Response, na.rm = TRUE),
    SD_Response = sd(Response, na.rm = TRUE),
    Count = n()
  )
knitr::kable(
  response_summary,
  caption = "Summary of response variable across factors and weeks",
  digits = 3,
  col.names = c("Inoculation Method", "Thickness", "Week", "Mean Response",
                "SD Response", "Count"),
  format = "markdown"
)
ggplot(data, aes(x = Response)) +
  geom_histogram(binwidth = 0.1, color = "black", fill = "skyblue") +
  labs(
```

```
x = "Response",
    y = "Frequency"
  theme_minimal() +
  theme(text = element_text(size = 14))
ggplot(data, aes(sample = Response)) +
  stat_qq(color = "black") +
  stat_qq_line(color = "red", linetype = "dashed") +
    x = "Theoretical Quantiles",
    y = "Sample Quantiles"
  ) +
  theme minimal() +
  theme(text = element_text(size = 14))
ggplot(data, aes(x = Week, y = Response,
                 color = Inoculation_Method,
                 group = interaction(Inoculation_Method, Thickness))) +
  geom_line(stat = "summary", fun = mean, size = 1) +
  geom_point(stat = "summary", fun = mean, size = 2) +
  facet_wrap(~ Thickness) +
  labs(
    x = "Week",
    y = "Mean Response",
    color = "Inoculation Method"
  theme_minimal() +
  theme(
    text = element_text(size = 14),
    legend.position = "bottom"
  )
wide_data <- pivot_wider(data, names_from = Week, values_from = Response)</pre>
time_data <- wide_data[ , -1]</pre>
time_data <- time_data[sapply(time_data, is.numeric)]</pre>
time_cor_matrix <- cor(time_data, use = "pairwise.complete.obs")</pre>
cor_df <- as.data.frame(time_cor_matrix)</pre>
```

```
cor_df <- cbind(Time = rownames(cor_df), cor_df)</pre>
kable(
  cor_df,
  digits = 2,
  col.names = c("Time", colnames(time_cor_matrix)))
corrplot(time_cor_matrix, method = "circle", type = "upper",
         tl.col = "black", tl.srt = 45)
# Model comparison data
covstruct <- data.frame(</pre>
  Model = c("VC", "UN", "CS", "AR(1)", "ARH(1)", "ANTE(1)", "TOEP"),
  AIC = c(-12.25, -12.22, -11.6, -12.86, -12.56, -8.57, -7.34),
 AICC = c(-12.09, -3.58, -11.29, -12.55, -11.01, -5.39, -6.19),
 BIC = c(-13.03, -18.47, -12.77, -14.03, -15.30, -12.48, -9.68)
knitr::kable(
 covstruct,
 format = "markdown",
  align = "c"
```

Appendix B - SAS Code and Additional Output

Power Analysis

```
data rptm_means;
input Inoculation_Method $ Thickness $ @@;
do Week=1 to 5 by 1;
    input mu @@;
    output;
end;
datalines;
Dry 1/4 4.26 4.25 4.47 4.33 4.54
Dry 1/8 4.91 4.95 4.67 4.56 4.97
Wet 1/4 4.21 4.57 4.65 4.49 4.38
Wet 1/8 4.86 4.78 4.62 4.32 4.22
data rptm_design;
 set rptm_means;
do Batches = 1 to 5; /* Creating 3 blocks (batches) */
  output;
 end;
run;
proc print data=rptm_design;
run;
/* Creating Model */
proc glimmix data=rptm_design;
    class Batches Inoculation_Method Thickness Week;
    model mu = Inoculation Method|Thickness|Week;
    random intercept / subject=Batches;
    random Week / subject=Batches*Inoculation_Method*Thickness type=ar(1) residual;
    parms (.029)(0.017)(.028)/hold=1,2,3;
    /* Provide 3 parameters for variance components */
    lsmeans Inoculation_Method*Thickness*Week / slicediff=Week cl;
    /* Define main effect contrasts */
    contrast 'Dry vs Wet'
        Inoculation_Method 1 -1;
    contrast '1/4 vs 1/8 inches'
```

```
Thickness 1 -1;
    /* Define interaction contrasts */
   contrast 'Dry vs Wet at 1/4 Inches'
        Inoculation Method 1 -1 Inoculation Method*Thickness 1 0 -1 0;
    contrast 'Dry vs Wet at 1/8 Inches'
        Inoculation_Method 1 -1 Inoculation_Method*Thickness 0 1 0 -1;
    contrast '1/4 vs 1/8 inches for Dry inoculation'
        Thickness 1 -1 Inoculation_Method*Thickness 1 -1 0 0;
    contrast '1/4 vs 1/8 inches for Wet inoculation'
        Thickness 1 -1 Inoculation_Method*Thickness 0 0 1 -1;
    ods output contrasts=f_contrast tests3=f_anova;
run;
/*Power*/
data power;
    set f_contrast f_anova;
    ncparm = numdf * fvalue;
    alpha = 0.05;
    fcrit = finv(1-alpha, numdf, dendf, 0);
    power = 1 - probf(fcrit, numdf, dendf, ncparm);
run;
proc print data=power;
run;
```

Simulation

```
do j = 1 to T;
            cov[i, j] = sqrt(sigma2[i] * sigma2[j]) * rho**abs(i - j);
        end;
    end;
    /* Print covariance matrix */
    print "Covariance Matrix:", cov;
    /* Generate simulated data using the covariance matrix */
    call randseed(12349):
                               /* Set random seed */
    x = randnormal(n, mean, cov); /* Simulate AR(1) correlated data */
    cname = {"t1", "t2", "t3", "t4", "t5"};
    /* Print the simulated data matrix directly */
    print "Simulated Data Matrix (x):", x;
    /* Print Sample mean */
    samplemean = x[:,];
    print samplemean n;
    /* Create dataset from simulated data */
    create inputdatacb from x[colname=cname];
    append from x;
close inputdatacb;
quit;
/* Step 2: Display the Simulated Data as a SAS Table */
proc print data=inputdatacb label;
    title "Simulated Data with AR(1) Covariance Structure";
run;
/* Step 3: Define Treatment Structure and Random Effects */
data rptm_simulation;
    retain Subject 0;
   keep Inoculation_Method Thickness Week Batches Response;
    array weeks[5] t1-t5;
    /* Define mean values for each combination of factors and week */
    if n = 1 then do;
        array mean_values[4,2,5] _temporary_ (
            /* Dry, 1/4 inch */
```

```
4.26, 4.25, 4.47, 4.33, 4.54,
            /* Dry, 1/8 inch */
            4.91, 4.95, 4.67, 4.56, 4.97,
            /* Wet, 1/4 inch */
            4.21, 4.57, 4.65, 4.49, 4.38,
            /* Wet, 1/8 inch */
            4.86, 4.78, 4.62, 4.32, 4.22
        );
    end;
    /* Simulation parameters */
    sigma_batch = sqrt(0.029); /* Batch variance */
    sigma_resid = sqrt(0.017); /* Residual variance */
    /* Loop through each combination of factors */
    do Batches = 1 to 5; /* Number of batches */
        batch_effect = rand("Normal", 0, sigma_batch); /*Random batch effect*/
        do Inoculation_Method = "Dry", "Wet";
            do Thickness = "1/4-inch", "1/8-inch";
                Subject + 1;
                set inputdatacb;
                /* Generate response for each week with AR(1) structure */
                do Week = 1 to 5;
                    Mean_Value = mean_values[
                      (Inoculation_Method="Dry")*1+(Inoculation_Method="Wet")*2,
                      (Thickness="1/4-inch")*1 + (Thickness="1/8-inch")*2,
                        Week
                    ];
                    Response = Mean_Value + batch_effect + weeks[Week];
                    output;
                end;
            end;
        end;
    end;
run;
/* Step 4: Display the Simulated Data in a Structured Format */
proc print data=rptm_simulation label;
    title "Simulated Data for 2x2 Factorial Design with Repeated Measures";
```

```
run;
```

Analysis

```
proc glimmix data=work.import plots=residualpanel;
    class Batches Inoculation_Method Thickness Week;
    model Response = Inoculation_Method|Thickness|Week;
    random intercept / subject=Batches;
    random Week / subject=Batches*Inoculation Method*Thickness type=ar(1) residual;
    lsmeans Inoculation_Method*Thickness / diff slicediff = week;
    lsmeans Inoculation_Method*Week / diff slicediff = Thickness;
    lsmeans Thickness*Week / diff slicediff = Inoculation_Method;
    lsmeans Inoculation Method*Thickness / slice=week;
    lsmeans Inoculation_Method*Week / slice=Thickness;
    lsmeans Thickness*Week / slice=Inoculation_Method;
    /* Define main effect contrasts */
    contrast 'Dry vs Wet'
        Inoculation_Method 1 -1;
    contrast '1/4 vs 1/8 inches'
        Thickness 1 -1;
    /* Define interaction contrasts */
   contrast 'Dry vs Wet at 1/4 Inches'
        Inoculation_Method 1 -1 Inoculation_Method*Thickness 1 0 -1 0;
    contrast 'Dry vs Wet at 1/8 Inches'
        Inoculation_Method 1 -1 Inoculation_Method*Thickness 0 1 0 -1;
    contrast '1/4 vs 1/8 inches for Dry inoculation'
        Thickness 1 -1 Inoculation_Method*Thickness 1 -1 0 0;
    contrast '1/4 vs 1/8 inches for Wet inoculation'
        Thickness 1 -1 Inoculation_Method*Thickness 0 0 1 -1;
    ods output contrasts=f_contrast tests3=f_anova;
run;
```

Additional Output

Inoc	Week	Inoc_	Week	Estimate	Std Error	DF	t Value	Prt
Dry	1	Dry	2	0.031250	0.08042	76	0.39	0.6987
$\overline{\text{Dry}}$	1	Dry	3	0.109000	0.07130	76	1.53	0.1304
$\overline{\text{Dry}}$	1	Dry	4	0.134700	0.07335	76	1.84	0.0702
$\overline{\text{Dry}}$	1	Dry	5	-0.260500	0.07291	76	-3.57	6e-04
$\overline{\text{Dry}}$	1	Wet	1	0.031740	0.07299	76	0.43	0.6649
$\overline{\text{Dry}}$	1	Wet	2	-0.165700	0.07299	76	-2.27	0.026
$\frac{1}{\text{Dry}}$	1	Wet	3	-0.031500	0.07299	76	-0.43	0.6673
$\overline{\text{Dry}}$	1	Wet	4	0.210400	0.07299	76	2.88	0.0051
Dry	1	Wet	5	0.275400	0.07299	76	3.77	3e-04
Dry	2	Dry	3	0.077780	0.08042	76	0.97	0.3365
Dry	2	Dry	4	0.103400	0.07130	76	1.45	0.151
$\overline{\text{Dry}}$	2	Dry	5	-0.291800	0.07335	76	-3.98	2e-04
$\overline{\text{Dry}}$	2	Wet	1	0.000498	0.07299	76	0.01	0.9946
$\overline{\text{Dry}}$	2	Wet	2	-0.197000	0.07299	76	-2.70	0.0086
$\overline{\text{Dry}}$	2	Wet	3	-0.062740	0.07299	76	-0.86	0.3927
$\frac{1}{\text{Dry}}$	2	Wet	4	0.179200	0.07299	76	2.46	0.0164
$\frac{1}{\text{Dry}}$	2	Wet	5	0.244200	0.07299	76	3.35	0.0013
Dry	3	Dry	4	0.025660	0.08042	76	0.32	0.7506
$\frac{\text{Dry}}{\text{Dry}}$	3	Dry	5	-0.369500	0.07130	76	-5.18	<.0001
$\frac{\text{Dry}}{\text{Dry}}$	3	Wet	1	-0.077280	0.07299	76	-1.06	0.293
Dry	3	Wet	2	-0.274800	0.07299	76	-3.76	3e-04
Dry	3	Wet	3	-0.140500	0.07299	76	-1.93	0.0579
$\frac{\text{Dry}}{\text{Dry}}$	3	Wet	4	0.101400	0.07299	76	1.39	0.1687
$\frac{\text{Dry}}{\text{Dry}}$	3	Wet	5	0.166400	0.07299	76	2.28	0.0254
Dry	4	Dry	5	-0.395200	0.08042	76	-4.91	<.0001
Dry	4	Wet	1	-0.102900	0.07299	76	-1.41	0.1625
$\frac{1}{\text{Dry}}$	4	Wet	2	-0.300400	0.07299	76	-4.12	<.0001
$\overline{\text{Dry}}$	4	Wet	3	-0.166200	0.07299	76	-2.28	0.0256
$\overline{\text{Dry}}$	4	Wet	4	0.075760	0.07299	76	1.04	0.3026
$\overline{\text{Dry}}$	4	Wet	5	0.140700	0.07299	76	1.93	0.0576
$\overline{\text{Dry}}$	5	Wet	1	0.292300	0.07299	76	4.00	1e-04
$\overline{\text{Dry}}$	5	Wet	2	0.094800	0.07299	76	1.30	0.198
$\frac{v}{\text{Dry}}$	5	Wet	3	0.229000	0.07299	76	3.14	0.0024
$\overline{\text{Dry}}$	5	Wet	4	0.471000	0.07299	76	6.45	<.0001
$\overline{\text{Dry}}$	5	Wet	5	0.535900	0.07299	76	7.34	<.0001
Wet	1	Wet	2	-0.197500	0.08042	76	-2.46	0.0164
Wet	1	Wet	3	-0.063240	0.07130	76	-0.89	0.3779
Wet	1	Wet	4	0.178700	0.07335	76	2.44	0.0172
Wet	1	Wet	5	0.243700	0.07291	76	3.34	0.0013
Wet	2	Wet	3	0.134200	0.08042	76	1.67	0.0992
Wet	2	Wet	4	0.376200	0.07130	76	5.28	<.0001
Wet	2	Wet	5	0.441100	0.07335	76	6.01	<.0001
Wet	3	Wet	4	0.241900	0.08042	76	3.01	0.0036
Wet	3	Wet	5	0.306990	0.07130	76	4.30	<.0001
Wet	4	Wet	5	0.064970	0.08042	76	0.81	0.4217
		1		1 3.33 2010	1 0.00012		1 3.01	

Figure 15: Differences of Inoculation Method * Week Least Squares Means table.

Thickness	Week	Thickness	Week_	Estimate	Std Error	DF	t Value	Prt
1/4-inch	1	1/4-inch	2	-0.110500	0.08042	76	-1.37	0.1735
1/4-inch	1	1/4-inch	3	-0.140600	0.07130	76	-1.97	0.0523
$\frac{7}{1/4\text{-inch}}$	1	1/4-inch	4	-0.111900	0.07335	76	-1.53	0.1311
$\frac{7}{1/4\text{-inch}}$	1	1/4-inch	5	-0.252500	0.07291	76	-3.46	9e-04
$\frac{7}{1/4\text{-inch}}$	1	1/8-inch	1	-0.529600	0.07299	76	-7.26	<.0001
$\frac{7}{1/4\text{-inch}}$	1	1/8-inch	2	-0.585300	0.07299	76	-8.02	<.0001
$\frac{7}{1/4\text{-inch}}$	1	1/8-inch	3	-0.343300	0.07299	76	-4.70	<.0001
$\frac{7}{1/4\text{-inch}}$	1	1/8-inch	4	-0.104300	0.07299	76	-1.43	0.1573
$\frac{7}{1/4\text{-inch}}$	1	1/8-inch	5	-0.293900	0.07299	76	-4.03	1e-04
$\frac{7}{1/4\text{-inch}}$	2	1/4-inch	3	-0.030060	0.08042	76	-0.37	0.7096
$\frac{7}{1/4\text{-inch}}$	2	1/4-inch	4	-0.001450	0.07130	76	-0.02	0.9838
$\frac{7}{1/4\text{-inch}}$	2	1/4-inch	5	-0.142000	0.07335	76	-1.94	0.0565
$\frac{7}{1/4\text{-inch}}$	2	1/8-inch	1	-0.419100	0.07299	76	-5.74	<.0001
$\frac{7}{1/4\text{-inch}}$	2	1/8-inch	2	-0.474800	0.07299	76	-6.51	<.0001
$\frac{7}{1/4\text{-inch}}$	2	1/8-inch	3	-0.232800	0.07299	76	-3.19	0.0021
$\frac{7}{1/4\text{-inch}}$	2	1/8-inch	4	0.006232	0.07299	76	0.09	0.9322
$\frac{7}{1/4\text{-inch}}$	2	1/8-inch	5	-0.183400	0.07299	76	-2.51	0.0141
$\frac{7}{1/4\text{-inch}}$	3	1/4-inch	4	0.028610	0.08042	76	0.36	0.723
$\frac{7}{1/4\text{-inch}}$	3	1/4-inch	5	-0.112000	0.07130	76	-1.57	0.1205
$\frac{7}{1/4\text{-inch}}$	3	1/8-inch	1	-0.389000	0.07299	76	-5.33	<.0001
$\frac{1}{4-inch}$	3	1/8-inch	2	-0.444800	0.07299	76	-6.09	<.0001
$\frac{1}{4-inch}$	3	1/8-inch	3	-0.202700	0.07299	76	-2.78	0.0069
$\frac{1}{4}$ -inch	3	1/8-inch	4	0.036290	0.07299	76	0.50	0.6205
$\frac{7}{1/4\text{-inch}}$	3	1/8-inch	5	-0.153400	0.07299	76	-2.10	0.0389
$\frac{1}{4}$ -inch	3	1/4-inch	5	-0.140600	0.08042	76	-1.75	0.0845
1/4-inch	4	1/8-inch	1	-0.417700	0.07299	76	-5.72	<.0001
1/4-inch	4	1/8-inch	2	-0.473400	0.07299	76	-6.49	<.0001
$\frac{1}{4\text{-inch}}$	4	1/8-inch	3	-0.231300	0.07299	76	-3.17	0.0022
$\frac{1}{4-inch}$	4	1/8-inch	4	0.007684	0.07299	76	0.11	0.9164
1/4-inch	4	1/8-inch	5	-0.182000	0.07299	76	-2.49	0.0148
1/4-inch	5	1/8-inch	1	-0.277100	0.07299	76	-3.80	3e-04
1/4-inch	5	1/8-inch	2	-0.332800	0.07299	76	-4.56	<.0001
1/4-inch	5	1/8-inch	3	-0.090740	0.07299	76	-1.24	0.2176
1/4-inch	5	1/8-inch	4	0.148300	0.07299	76	2.03	0.0457
1/4-inch	5	1/8-inch	5	-0.041400	0.07299	76	-0.57	0.5723
1/8-inch	1	1/8-inch	2	-0.055730	0.08042	76	-0.69	0.4904
1/8-inch	1	1/8-inch	3	0.186300	0.07130	76	2.61	0.0108
1/8-inch	1	1/8-inch	4	0.425300	0.07335	76	5.80	<.0001
1/8-inch	1	1/8-inch	5	0.235700	0.07291	76	3.23	0.0018
1/8-inch	2	1/8-inch	3	0.242100	0.08042	76	3.01	0.0035
1/8-inch	2	1/8-inch	4	0.481100	0.07130	76	6.75	<.0001
1/8-inch	2	1/8-inch	5	0.291400	0.07335	76	3.97	2e-04
1/8-inch	3	1/8-inch	4	0.239000	0.08042	76	2.97	0.004
1/8-inch	3	1/8-inch	5	2 5049340	0.07130	76	0.69	0.4911
1/8-inch	4	1/8-inch	5	-0.189700	0.08042	76	-2.36	0.0209

Figure 16: $Differences\ of\ Thickness\ *\ Week\ Least\ Squares\ Means\ table.$