Asymptomatic Alzheimer disease

Defining resilience

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ABSTRACT

Objective: To define robust resilience metrics by leveraging CSF biomarkers of Alzheimer disease (AD) pathology within a latent variable framework and to demonstrate the ability of such metrics to predict slower rates of cognitive decline and protection against diagnostic conversion.

Methods: Participants with normal cognition (n = 297) and mild cognitive impairment (n = 432) were drawn from the Alzheimer's Disease Neuroimaging Initiative. Resilience metrics were defined at baseline by examining the residuals when regressing brain aging outcomes (hippocampal volume and cognition) on CSF biomarkers. A positive residual reflected better outcomes than expected for a given level of pathology (high resilience). Residuals were integrated into a latent variable model of resilience and validated by testing their ability to independently predict diagnostic conversion, cognitive decline, and the rate of ventricular dilation.

Results: Latent variables of resilience predicted a decreased risk of conversion (hazard ratio < 0.54, p < 0.0001), slower cognitive decline (β > 0.02, p < 0.001), and slower rates of ventricular dilation (β < -4.7, p < 2 × 10⁻¹⁵). These results were significant even when analyses were restricted to clinically normal individuals. Furthermore, resilience metrics interacted with biomarker status such that biomarker-positive individuals with low resilience showed the greatest risk of subsequent decline.

Conclusions: Robust phenotypes of resilience calculated by leveraging AD biomarkers and baseline brain aging outcomes provide insight into which individuals are at greatest risk of short-term decline. Such comprehensive definitions of resilience are needed to further our understanding of the mechanisms that protect individuals from the clinical manifestation of AD dementia, especially among biomarker-positive individuals. **Neurology® 2016;87:2443-2450**

GLOSSARY

 $A\beta = \beta$ -amyloid; AD = Alzheimer disease; ADNI = Alzheimer's Disease Neuroimaging Initiative; ADNI-GO = Alzheimer's Disease Neuroimaging Initiative-Grand Opportunity; ICV = ICV

Limited empirical attention has focused on elucidating the factors that prevent individuals with asymptomatic Alzheimer disease (AD), those with autopsy-confirmed AD pathology (β-amyloid [Aβ] plaques and aggregation of neurofibrillary tangles) but without cognitive impairment, from expressing clinical AD.^{1,2} Both cognitive reserve, frequently operationalized by educational attainment, and brain reserve, frequently operationalized by cranial volume, confer resilience to cognitive deficits associated with AD pathology. However, it is quite likely that additional factors beyond traditional metrics of cognitive and brain reserve contribute to resilience observed in asymptomatic AD.

The well-established CSF biomarkers of AD (levels of A β -42 and total tau) may provide additional information that is useful for defining resilience. We therefore sought to leverage

Supplemental data at Neurology.org

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Interested investigative teams can contact Dr. Hohman for access to the calculated resilience phenotypes.

Data used in preparation of this article were obtained from the Alzheimer's Disease Neuroimaging Initiative (ADNI) database (adni.loni.usc.edu). As such, the investigators within the ADNI contributed to the design and implementation of ADNI and/or provided data but did not participate in analysis or writing of this report. A complete listing of ADNI investigators can be found in the coinvestigators list at Neurology.org.

Go to Neurology.org for full disclosures. Funding information and disclosures deemed relevant by the authors, if any, are provided at the end of the article.

a residual approach⁴ to quantify levels of resilience by examining the disparity between CSF biomarkers of AD pathology and brain aging outcomes (i.e., hippocampal atrophy and cognitive decline). With this approach, brain resilience reflects larger-than-predicted hippocampal volumes given biomarker levels of AD pathology, while cognitive resilience reflects better-than-predicted memory and executive function performance given biomarker levels of AD pathology. We then integrate these resilience metrics with established metrics into a latent variable model of reserve.^{2,3,5} This comprehensive approach will provide insight into the mechanisms of asymptomatic AD and demonstrate how multivariate information can be harnessed to improve in vivo estimates of resilience.

METHODS Participant data from the Alzheimer's Disease Neuroimaging Initiative (ADNI; adni.loni.usc.edu) were used for this article. ADNI began in 2003 and was subsequently followed up by ADNI-Grand Opportunity (ADNI-GO) and ADNI-2. In total, >1500 adults 55 to 90 years of age have been recruited. Participants were excluded for serious neurologic disease other than AD, a history of psychoactive medication use, or a history of brain lesion or head trauma (complete exclusion criteria can be found at www.adni-info.org).

Standard protocol approvals, registrations, and patient consents. Informed written consent was obtained from all participants at each site. The Vanderbilt Institutional Review Board approved data access and analysis.

Participants. We accessed publicly available data from ADNI on June 13, 2015. Participant enrollment criteria are outlined in the ADNI protocol (http://www.adni-info.org/Scientists/ ADNIStudyProcedures.html). For the present analyses, we included all participants with MRI data, memory and executive function composite measures, CSF biomarkers, and American National Adult Reading test score. Participants with AD were excluded. These criteria resulted in 729 participants with baseline diagnoses of normal cognition (NC; n = 297) and mild cognitive impairment (MCI; n = 432). Participants included individuals from ADNI-1 and follow-up ADNI-GO (initiated in 2009) and ADNI-2 (initiated in 2011). Additional information about the ADNI protocols can be found online (http://adni.loni.usc.edu/about/). In the present analyses, the average follow-up period for individuals from ADNI-1 was 4.6 years and from ADNI-2/GO was 1.9 years. Given the difference in follow-up period between ADNI-1 and ADNI-2/GO and the use of 3T in ADNI-2/GO vs 1.5T in ADNI-1, we reran all analyses stratified by scanner strength (which coincided with ADNI phase) in appendix e-1 at Neurology.org (section S.3).

CSF biomarker processing. Previous work has outlined the quantification of AD biomarkers, including tau and $A\beta$ -42.^{6,7} For the present analyses, we used the first measure of $A\beta$ -42 and total tau, which coincided with baseline neuropsychological measures. These continuous measures were used as quantitative predictors in statistical analyses.

Composite neuropsychological measurements. Episodic memory composite scores and executive function composite scores were calculated and reported previously.^{8,9} To summarize, item-level data from the Logical Memory Test, Mini-Mental State Examination, Rey Auditory Verbal Learning Test, and AD Assessment Scale-Cognitive Subscale were included in a single-factor solution to quantify the ADNI episodic memory composite. Item-level data from the Vegetable Naming test, Trail Making Test A, Trail Making Test B, Digit Symbol, Backward Digit Span, Animal Naming, and Clock Drawing Test were used to quantify the ADNI executive function composite. The harmonization of these metrics across ADNI-1 and ADNI-2/GO was completed by Gibbons and colleagues9 and is available online (http://adni.bitbucket.org/docs/UWNPSYCHSUM/ADNI_ Methods_UWNPSYCHSUM.pdf). The composite measures were built to have a mean of 0 and standard deviation of 1.

Quantification of regional brain volume. The neuroimaging protocol in ADNI has been outlined previously. The present analyses included uncorrected 1.5T T1-weighted structural data for ADNI-1 and 3T data for ADNI-2/GO. FreeSurfer was used for reconstruction and segmentation using version 4.3 in ADNI-1 and 5.1 in ADNI-2/GO (http://surfer.nmr.mgh.harvard.edu/)¹¹⁻¹³ as previously described. The baseline neuroimaging visit was required to be within 6 months of the first cognitive visit. The mean interval between the first cognitive visit and first neuroimaging visit was 22 days.

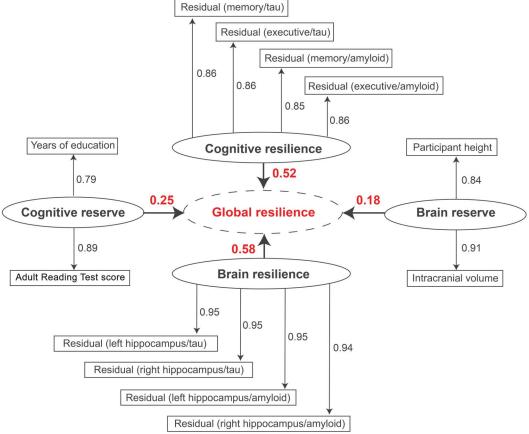
We used FreeSurfer-defined left hippocampal and right hippocampal volumes as our primary outcome measurements in brain resilience analyses.¹⁵ Lateralized hippocampal volumes were used on the basis of evidence that aging and disease do not occur symmetrically.¹⁶ Thus, lateralized analyses provide more sensitivity to the subtle changes in the preclinical stages of AD. Free-Surfer was also used to define left and right inferior lateral ventricle volumes (a neuroimaging marker that is sensitive to neurodegeneration)¹⁷ for follow-up longitudinal analyses.

Statistical analyses. Statistical analyses were performed in R version 3.0.1 (http://www.r-project.org/) using a multivariate partial least squares (PLS) path model. PLS path models are the least-squares equivalent to structural equation models but make fewer assumptions about variable distributions and are better positioned to handle highly correlated indicator variables.¹⁸ The single PLS path model (figure 1) quantified 4 first-order latent composite measures (cognitive reserve, brain reserve, cognitive resilience, and brain resilience) and one second-order (meaning a latent trait derived from other latent traits) latent composite measure (global resilience). This model incorporated established proxy measures for cognitive reserve (i.e., reading ability from the American National Adult Reading Test and education level in years) and brain reserve (i.e., height in inches and intracranial volume [ICV] in mm3). While brain reserve is aimed at evaluating premorbid brain size based on proxy measures defined in the literature,3,5 brain resilience is aimed at measuring the degree to which atrophic processes are already present in response to normal aging or neuropathologic processes. For cognitive resilience and brain resilience, residuals were used from linear regression models outlined below.

PLS path model step 1: Calculating residuals. The primary model inputs were residuals from individual linear regressions covarying for baseline age and ICV (for brain analyses only). Separate regressions were run for each particular type of resilience. For example, memory-specific resilience to tau was calculated as residuals from a regression model with memory as the outcome and tau as the predictor, represented as the Residual (memory/tau)

Figure 1 Partial least squares (PLS) path model results

Residual (memory/fau)



PLS path model results are presented; the goodness of fit was 0.76. Each first-order latent variable is presented as an oval. The variables included in each latent trait are presented as rectangles, and the factor loadings are presented as numbers above each arrow. Arrows are pointing away from the latent trait because we used reflective measurement. For the resilience metrics, each rectangle represents the residuals from a single linear regression model relating the given biomarker to the given outcome. The second-order latent variable (global resilience) is presented as a dotted oval. The loadings for each first-order latent variable are presented numerically above the bold arrows pointing to global resilience.

oval in figure 1. The 8 residuals used in the cognitive and brain resilience latent traits came from regressions relating amyloid and tau levels to baseline memory, executive function, left hippocampus volume, and right hippocampus volume.

PLS path model step 2: Running the PLS path model. A PLS path model requires both an outer model and an inner model. The outer model is the derivation of a latent trait from the indicator variables, and the inner model includes relations between the latent traits. For the outer model, reflective measurement was used to derive latent variables from the residuals calculated in step 1 and the proxy measures of cognitive and brain resilience described above. For the inner model, brain reserve, cognitive reserve, cognitive resilience, and brain resilience contributed to a global resilience variable (figure 1). A repeated indicators approach was used to derive the global resilience second-order latent trait from the 4 other latent variables. The goodness-of-fit metric in the plspm package was calculated from the average communality score and the average r^2 value; a goodness of fit >0.70 is generally accepted as a good fit. 19

PLS path model hypothesis testing: Evaluating resilience measures. To evaluate the resilience measures derived with baseline data in the PLS path model, we tested whether they each predicted diagnostic conversion or longitudinal decline in neuropsychological performance. All analyses evaluating resilience

metrics were performed outside the PLS framework. Scores from the PLS model were used in subsequent regression models as outlined below.

Right-censored Cox proportional hazards regression was used for conversion analyses covarying for baseline age, sex, baseline diagnosis, and *APOE* ε4. A participant was considered to have converted if the individual had a change in diagnosis code from NC to MCI, NC to AD, or MCI to AD over the course of the follow-up period. Diagnostic groups were defined according to the ADNI protocol and are presented as a footnote in table 1.

Mixed effects regression was used for longitudinal cognitive analyses, conducted separately for memory and executive function. Fixed effects included baseline age, sex, baseline diagnosis, $APOE\ \epsilon 4$, the resilience trait of interest, time (years from baseline), and a time \times resilience interaction term. Random effects included time and the intercept. The time \times resilience interaction term tested whether the resilience variable of interest related to change in cognition over time.

Mixed effects regression was also used to evaluate longitudinal change in lateral inferior ventricle volume, again calculated separately for the right and left hemispheres given established lateralization differences. ²⁰ Fixed effects included the same terms used in the cognitive analyses but also included baseline ICV. Random effects were the same as for the cognitive analyses.

Table 1 Sample characteristics Baseline clinical diagnosis^a NC MCI Statistical results Sample size, n 297 432 NA APOE ε4 carriers, % 28 50 $\chi^2 = 34.78$, p < 0.001Female, % 42 $\chi^2 = 7.90, p = 0.003$ 54 $\chi^2 = 23.95$, p < 0.001Diagnostic conversion, b % 11 26 74 ± 5.8 72 ± 6.7 t(721.7) = 2.56, p = 0.011Baseline age, y Education, y 16 ± 2.6 16 ± 2.8 t(721.7) = 0.93, p = 0.353t(721.7) = 18.75, p < 0.001Composite memory, z score 0.94 ± 0.47 0.20 ± 0.59 Composite executive function, z score 0.79 ± 0.71 0.25 ± 0.80 t(721.7) = 9.61, p < 0.001Left hippocampal volume, mm³ $3,716 \pm 445$ $3,375 \pm 598$ t (721.7) = 8.82, p < 0.001Right hippocampal volume, mm³ 3.776 ± 479 3.443 ± 600 t (721.7) = 8.33, p < 0.001t(721.7) = -8.13, p < 0.001Left inferior lateral ventricle volume, mm³ 642 ± 418 $1,000 \pm 763$ Right inferior lateral ventricle volume, mm³ 542 ± 401 872 ± 730 t (721.7) = -7.82, p < 0.001t(721.7) = 7.45, p < 0.001CSF Aβ-42, pg/mL 201 ± 52 $171\,\pm\,53$ CSF total tau, pg/mL 67 ± 30 92 ± 55 t(721.7) = -7.79, p < 0.001

Abbreviations: MCI = mild cognitive impairment; NA = not applicable; NC = normal cognition.

^a Diagnosis was determined in accordance with the Alzheimer's Disease Neuroimaging Initiative protocol. NC individuals were required to score between 24 and 30 on the Mini-Mental State Examination, to have a 0 on the Clinical Dementia Rating scale, and to score <6 on the Geriatric Depression Scale. Individuals with MCI were required to score between 24 and 30 on the Mini-Mental State Examination, to show signs of objective and subjective memory impairment, and to score 0.5 on the Clinical Dementia Rating. Individuals with Alzheimer disease were required to meet clinical criteria for dementia, to score between 20 and 26 on the Mini-Mental State Examination, and to score ≥0.5 on the Clinical Dementia Rating.

^b Diagnosis based on the above criteria was calculated at each visit. Participants were considered to have converted if the diagnostic code changed from NC to MCI, NC to Alzheimer disease, or MCI to Alzheimer disease.

Next, we reran all conversion and longitudinal analyses, restricting to participants with NC. Then, given that cut points for CSF measures are often applied in clinical settings, we also reran analyses using binary CSF metrics (rather than continuous CSF metrics in the original regression analyses) based on previously defined cut points 7 of total tau $\geq\!93$ pg/mL (tau positive) and A β -42 $\leq\!192$ pg/mL (amyloid positive).

Models were also recalculated with adjustment for baseline hippocampal volume, baseline memory performance, and baseline biomarker levels to assess whether the observed resilience effects explained additional variance beyond these established predictors of decline. The longitudinal model also included a time interaction with each of these additional covariates.

Finally, to evaluate whether the resilience measures influenced risk of decline among individuals with biomarker evidence of AD pathology, we tested for a statistical interaction between biomarker positivity (tau \geq 93pg/mL or A β -42 \leq 192 pg/mL) and resilience.

RESULTS Diagnostic groups demonstrated expected differences in cognitive performance, hippocampal volume, and CSF AD biomarker levels (p < 0.05). For longitudinal analyses, the mean follow-up interval was 2.69 years among participants with NC and 2.92 years among participants with MCI (p = 0.15). Participant characteristics are presented in table 1.

PLS path model. The data fit the model well with a goodness-of-fit score of 0.76 (figure 1). All indicators loaded

strongly on the given latent trait (loading >0.79), and no indicator cross-loaded onto another latent trait (all cross-loadings <0.37). Each first-order latent trait also demonstrated clear unidimensionality (Dillon-Goldstein $\rho>0.80$, first eigenvalue >1.40, second eigenvalue <1.0). Within the individual regression models that were included in the PLS models, we observed significant effects of CSF biomarkers in predicting both cognition (absolute value of $\beta>0.0029$, change in $R^2>0.03$, p<0.0001) and hippocampal volume (absolute value of $\beta>2.58$, change in $R^2>0.05$, p<0.0001).

Evaluating resilience measures. Results evaluating resilience metrics are presented in table 2. The 3 metrics of resilience (cognitive resilience, brain resilience, and global resilience) successfully predicted protection against diagnostic conversion, slower rates of cognitive decline (figure 2), and slower rates of ventricular dilation (all p < 0.001). All 3 resilience metrics also interacted with biomarker positivity (all p < 0.01) whereby the strongest effect of biomarker positivity was observed at lower levels of global resilience (figure 3). The 2 established reserve metrics showed null or negative association with protection against conversion and cognitive decline.

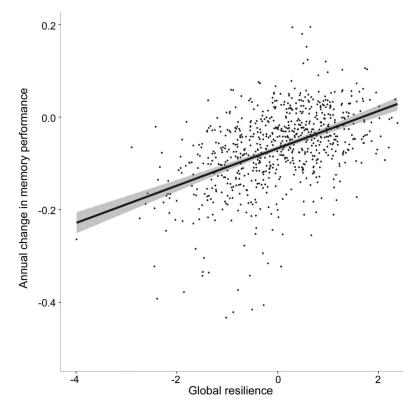
	Cognitive reserve				Brain reserve				Cognitive resilience				Brain resilience				Global resilience		
	HR	95% (CI	p Value	HR !	95% CI		p Value	HR 9	5% CI	ı	p Value	HR	95% CI	ı	p Value	HR 95	% CI	p Value
Main effects																			
Diagnostic conversion	0.90	0.78 t	o 1.04	0.15	0.95	0.76 to 1.1	L8	0.63	0.42 0	.34 to C).51	2×10^{-16}	0.54	0.46 to	0.64	6×10^{-14}	0.41 0.3	4 to 0.49	2×10^{-1}
	Cognitive reserve				Brain reserve				Cognitive resilience				Brain resilience				Global resilience		
	β (95%	CI)	t (df)	p Value	β (95% C	I) t (c	df)	p Value	β (95% CI)	t (df)	p Value	β (95%	CI)	t (df)	p Value	β (95% CI)	t (df)	p Value
Main effects																			
Longitudinal outcomes																			
Memory performance	0.01 (0. 0.02)	.00 to	1.42 (2,650)	0.16	-0.01 (-0 to 0.00)		83 650)	0.07	0.02 (0.01 0.04)	. to	3.34 (2,650)	0.0008	0.05 (0. 0.06)	03 to	7.16 (2,650)	$\begin{array}{c} 1\times \\ 10^{-12} \end{array}$	0.04 (0.03 to 0.05)	6.09 (2,650	1 × 10 ⁻⁹
Executive function performance	0.01 (0 0.03)	.00 to	1.27 (2,644)	0.20	-0.02 (-0 to 0.00)		96 644)	0.05	0.03 (0.02 0.05)	to !	3.60 (2,644)	0.0003	0.05 (0. 0.07)	04 to	6.17 (2,644)	7×10^{-10}	0.05 (0.03 to 0.06)	5.35 (2,644	9 × 10 ⁻⁸
Left ILV volume	−15.0 (to −4.9		-2.90 (2,109)	0.0038	31.9 (22. 41.7)		35 109)	3 × 10 ⁻¹⁰	-47.8 (-5 -38.2)	7.4 to	-9.75 (2,109)	5 × 10 ⁻²²	-41.0 (-31.1)	-51.0 to	-8.08 (2,109)	1 × 10 ⁻¹⁵	-46.8 (-56.5 -37.0)	to -9.40 (2,109	
Right ILV volume	−13.3 (to −3.0		-2.52 (2,109)	0.0117	26.1 (16. 36.3))2 109)	5 × 10 ⁻⁷	-44.9 (-5 -35.0)	4.8 to	-8.86 (2,109)	2 × 10 ⁻¹⁸	-45.4 (-35.4)	-55.5 to	-8.87 (2,109)	2 × 10 ⁻¹⁸	-48.5 (-58.4 -38.6)	to -9.61 (2,109	
	Cognitive reserve \times biomarker positivity ^a													Global resilience \times biomarker positivity $^{\mathbf{a}}$					
	β (95% C	(1)	t (df)	p Value	β (95% CI)	t (df)		p Value	β (95% CI)	t	(df)	p Value	β (95% C	I) 1	t (df)	p Value	β (95% CI)	t (df)	p Valu
Biomarker interaction effects																			
Longitudinal outcomes																			
Memory performance	0.00 (-0. 0.03)	02 to	0.27 (2,648)		0.00 (-0.0) to 0.03)	2 0.38 (2,64	8)	0.70	0.06 (0.03 t 0.08)		1.33 2,648)		0.03 (0.01 0.06)		2.69 (2,648)	0.0073	0.05 (0.03 to 0.08)	3.97 (2,648)	7 × 10 ⁻⁵
Executive function performance	0.00 (-0. 0.03)	04 to	-0.09 (2,642)		0.00 (-0.0 to 0.03)	4 -0.16 (2,64		0.87	0.07 (0.03 t 0.10)	-	3.80 2,642)		0.05 (0.01 0.08)		2.64 (2,642)	0.0083	0.06 (0.02 to 0.09)	3.34 (2,642)	0.000
Left ILV volume	-6.83 (-2 to 12.8)		-0.68 (2,107)		37.5 (18.6 56.5)	to 3.88 (2,10	7)	0.0001	-48.0 (-66 -29.5)		-5.09 2,107)	4 × 10 ⁻⁷	-27.5 (-4 -8.1)		-2.77 (2,107)	0.0056	-34.1 (-53.3 -14.8)	to -3.47 (2,107)	0.000
Right ILV volume	-10.2 (-:		-1.00	0.32	32.8 (13.0	to 3.24		0.0012	-40.6 (-59	٥.	4.13	4 ×	-33.5 (-5	-04.	-3.34	0.0008	-36.0 (-55.5	to -3.63	0.000

Abbreviations: CI = confidence interval; HR = hazard ratio; ILV = inferior lateral ventricle.

Diagnostic conversion analyses were analyzed using a cox proportional hazard model in which the event of interest was defined as a change in diagnosis to either mild cognitive impairment or Alzheimer disease. Longitudinal analyses of cognition and brain volume were performed using mixed-effects regression models.

^a Participants were considered biomarker positive if they were tau positive (CSF total tau \geq 93 pg/mL) or amyloid positive (CSF A β -42 \leq 192 pg/mL).

Figure 2 Association between global resilience phenotype and annual change in memory performance



The second-order latent trait global resilience is along the x-axis, and annual change in composite memory performance is along the y-axis. Increased global resilience is associated with a slower trajectory of memory decline.

Similar cross-sectional and longitudinal effects were observed when limiting the sample to NC; when correcting for baseline cognition, hippocampal volume, and CSF biomarker levels; when using binarized rather than continuous CSF metrics; when stratifying the sample by scanner strength; and when removing memory performance or executive function performance from the cognitive resilience calculation (see appendix e-1).

DISCUSSION This study leveraged latent variable modeling to incorporate AD biomarkers into a model of reserve. All resilience phenotypes were protective against diagnostic conversion, cognitive decline, and neurodegeneration. The global resilience phenotype calculated across all resilience and reserve metrics provided the best predictive performance. We observed more significant associations for the 3 resilience metrics than for the established reserve metrics, highlighting the value of including AD biomarkers within statistical models of resilience.

The present results demonstrate that our innovative metrics of resilience predict slower rates of longitudinal cognitive decline. There is a tremendous need to identify individuals at greatest risk

for cognitive decline, particularly within the context of clinical trial participant selection and enrollment. Our results suggest that biomarker levels, hippocampal volume, and cognitive performance can be integrated into a powerful metric of resilience that can be applied to improve clinical trial screening procedures. Moreover, the multivariate technique implemented herein extracts valuable information about current risk for future decline that outperformed the independent predictive value of biomarker levels, brain volume, and cognitive performance alone.

The present model provided 2 notable additions to previous models. First, the brain resilience metric based on residual variance in hippocampal volume provided a strong predictor of future cognitive impairment. Brain resilience may indeed represent a distinct pathway to resilience in which individuals are resistant to brain atrophy in the presence of AD biomarkers. Second, the inclusion of CSF biomarker predictors in the calculation of cognitive resilience provided additional predictive power. When comparing cognitive resilience calculated with CSF predictors to cognitive resilience calculated with structural MRI predictors, we found the CSF-based metric to be a slightly stronger predictor of protection against conversion (appendix e-1). Thus, the inclusion of CSF biomarkers may enhance the definitions of resilience, likely because of the sensitivity of CSF biomarkers to prodromal changes before notable neurodegeneration.

Past work has demonstrated that proxy measures of reserve predict improved outcomes in biomarker-positive individuals, ^{21,22} and our results suggest that factors beyond these proxies are important contributors to sustained cognition in the presence of AD pathology. As demonstrated by the biomarker-positive results, the residual method provides a continuous metric of resilience that is particularly sensitive to better-than-expected cognitive performance among individuals at highest risk of future cognitive decline. Thus, these metrics may be a valuable tool for uncovering molecular drivers of resilience that predispose some individuals to a more optimal innate response to the damaging effects of AD neuropathology.

The present results also have strong implications for how AD biomarkers may be best applied in clinical settings. For example, the low-resilience individuals identified with the present technique may need to be more aggressively managed given their elevated risk for rapid decline. Thus, moving back the treatment window may indeed be possible by developing simplified methods that integrate biomarker, imaging, and cognitive data into clinically meaningful metrics of risk and resilience.

The second-order latent trait global resilience is along the x-axis, and annual change in composite memory function performance is along the y-axis. Lines are colored by biomarker group where biomarker positive is defined as having a CSF A β -42 \leq 192 pg/mL or a CSF total tau level \geq 93 pg/mL. The association between global resilience and slowed cognitive decline is strongest in biomarker-positive individuals.

This study takes a step toward fully integrating biomarker data into models of resilience. Although we present lateralized hippocampal volumes in our models, a combined hippocampal volume provided comparable performance (data not shown) and could be applied within the proposed framework. Evidence from recent longitudinal applications of comparable residual models suggests that change in resilience may be most predictive of cognitive impairment.4 Future work will seek to integrate change in AD biomarkers and brain aging outcomes into longitudinal models of resilience. The present model could be further improved by including additional regions in the brain resilience metric such as those identified in the Spatial Pattern of Abnormalities for Recognition of Early AD score calculation²³ or brain regions known to atrophy during normal aging²⁴ or MCI.²⁵

The present results must be interpreted within the framework of the cohort. ADNI is not the ideal cohort for studying reserve and resilience given the skew toward white, non-Hispanic, highly educated individuals with an underrepresentation of cardiovascular risk and disease compared to the general

population. Future work is needed to extend these models in more diverse cohorts and to determine whether this type of model would extend to individuals with dementia. The present results highlight that variability in resilience exists even in a highly educated and homogeneous cohort. It should also be noted that the resilience metrics were calculated outside the latent variable framework, leaving open the possibility that resilience may in part reflect psychometric error. Thus, it may be beneficial to vet whether additional power can be garnered with the use of a more complex latent variable framework.

In the present framework, it is difficult to fully separate brain resilience and cognitive resilience from the cognitive performance and hippocampal volume metrics that were used in the resilience metric calculation. We have attempted to reduce circularity concerns by evaluating multiple independent outcomes that were not used in model calculation (i.e., diagnostic conversion and lateral inferior ventricle volume), but it remains possible that subtle collinearity or residual circularity remains.

This study calculated robust resilience phenotypes by integrating AD biomarkers and residual measures into a latent variable model of reserve. Our results provide a foundation from which the molecular basis of reserve and resilience can be explored in future genomic and proteomic analyses.

AUTHOR CONTRIBUTIONS

Dr. Hohman: study concept, design, analysis, interpretation, manuscript preparation. Dr. McLaren, Dr. Mormino, and Dr. Gifford: design, interpretation, and critical revision. Dr. Libon: interpretation and critical revision. Dr. Jefferson: study supervision, design, interpretation, critical revision.

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DISCLOSURE

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