

# Math 275A: Homework 4

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## 1.7.4

Claim: Let  $\mu$  be a finite measure on  $\mathbb{R}$  and  $F(x) = \mu((-\infty, x])$ . Show that

$$\int (F(x+c) - F(x)) dx = c\mu(\mathbb{R})$$

Proof: Define the product measure  $\lambda \times \mu$  on  $\mathbb{R}^2$  where  $\lambda$  is the standard Lebesgue measure on  $\mathbb{R}$ . Also define a subset of  $\mathbb{R}^2$ ,

$$E = \{(x, y) : x < y \leq x + c\} = \{(x, y) : y - c < x \leq y\},$$

over which we will integrate. From Fubini's theorem, we can determine that

$$\int_{\mathbb{R}^2} \mathbf{1}_E d(\lambda \times \mu) = \int_{\mathbb{R}} \int_{x < y \leq x+c} d\mu dx = \int_{\mathbb{R}} F(x+c) - F(x) dx.$$

Likewise, if we flip the order of integration (and use the alternative expression for  $E$  above), we get

$$\int_{\mathbb{R}^2} \mathbf{1}_E d(\lambda \times \mu) = \int_{\mathbb{R}} \int_{y-c < x \leq y} dx d\mu = \int_{\mathbb{R}} c d\mu = c\mu(\mathbb{R}),$$

proving the claim.

## 2.1.5

Claim: (i) If  $X$  and  $Y$  are independent with distributions  $\mu$  and  $\nu$  then

$$P(X + Y = 0) = \sum_y \mu(\{-y\})\nu(\{y\}).$$

(ii) If  $X$  has a continuous distribution  $P(X = Y) = 0$ .

Proof: Since  $X$  and  $Y$  are independent, we know the vector  $(X, Y)$  has distribution  $\mu \times \nu$ . Thus can define the  $P(X + Y = 0)$  as an integral over the indicator function for  $x + y = 0$ , that is

$$P(X + Y = 0) = \int \mathbf{1}_{(x+y=0)} d(\mu \times \nu).$$

This integrand is nonzero only on a line, passing through the origin of  $\mathbb{R}^2$ . Thus we can rewrite the expression for the probability as

$$P(X + Y = 0) = \int \mathbf{1}_{(x+y=0)} d\mu d\nu = \int \mu(\{-y\}) d\nu = \sum_y \mu(\{y\})\nu(\{-y\}),$$

proving the first claim. Note here that the sum over  $y$  makes sense in that each distribution has only countably many discontinuities such that  $\mu$  and  $\nu$  only evaluate to zero at a finite number of points.

To prove the second claim, just note that if  $\mu$  (r.e.  $\nu$ ) has a continuous distribution function  $F(x)$ , then  $\mu(x) = F(x) - F(x^-) = 0$  by definition. Thus

$$P(X + Y = 0) = \sum_y \mu(\{y\})\nu(\{-y\}) = \sum_y 0 \cdot 0 = 0.$$

### 2.1.15

Claim: Let  $\Omega$  be the unit interval  $(0, 1)$  equipped with the Borel sets  $\mathcal{F}$  and Lebesgue measure  $P$ . Let  $Y_n(\omega) = 1$  if  $[2^n\omega]$  is odd and 0 if  $[2^n\omega]$  is even. Show that  $Y_1, Y_2, \dots$  are independent with  $P(Y_k = 0) = P(Y_k = 1) = 1/2$ .

Proof: For a given  $k$ , each successive  $h$ -interval  $[\frac{i-1}{2^k}, \frac{i}{2^k})$  for  $i \in 1, 2, \dots, 2^k$  maps alternately to either 0 or 1, and has an equivalent measure under  $P$ . Note that all intervals specified as  $[0, \frac{1}{2^k})$  are actually open intervals, but this doesn't affect their measure under  $P$  so I will not differentiate these intervals in the proof.

Given a specific  $y_1$ ,  $\{\omega : Y_1(\omega) = y_1\} = [\frac{i}{2}, \frac{i+1}{2})$  for some  $i \in \{1, 2\}$ . If we also specify  $y_2$ , then  $\{\omega : Y_1(\omega) = y_1, Y_2(\omega) = y_2\} = \{\omega : Y_1(\omega) = y_1\} \cap \{\omega : Y_2(\omega) = y_2\} = [\frac{i}{2}, \frac{i+1}{2})$  for some  $i \in \{0, 1, 2, 3\}$ . Proceeding by induction, it follows that for any given sequence of  $y_k$ 's of length  $n$ ,

$$\{\omega : Y_1(\omega) = y_1, \dots, Y_n(\omega) = y_n\} \cap_{i=k}^n \{\omega : Y_k(\omega) = y_k\} = [\frac{i}{2^n}, \frac{i+1}{2^n})$$

for some  $i \in \{0, 2^n - 1\}$ .

Thus the probability of any length- $n$  sequence occurring is  $P([\frac{i}{2^n}, \frac{i+1}{2^n})) = \frac{1}{2^n}$ . This combined with the fact that each  $Y_k$  has an equal probability of being 0 or 1 lets us write

$$P(Y_1 = y_1, Y_2 = y_2, \dots, Y_n = y_n) = \frac{1}{2^n} = \prod_{i=1}^n P(Y_i = y_i) \quad \text{for all } n \in \mathbb{N},$$

confirming that  $Y_1, Y_2, \dots$  are independent.

### 2.2.3

Claim: (i) Let  $f$  be a measurable function on  $[0, 1]$  with  $\int_0^1 |f(x)| dx < \infty$ . Let  $U_1, U_2, \dots$  be independent and uniformly distributed on  $[0, 1]$ , and let

$$I_n = n^{-1}(f(U_1) + \dots + f(U_n))$$

Show that  $I_n \rightarrow I := \int_0^1 f dx$  in probability. (ii) Suppose  $\int_0^1 |f(x)|^2 dx < \infty$ . Use Chebyshev's inequality to estimate  $P(|I_n - I| > a/n^{1/2})$ .

Proof: Each function  $f(U_i)$  is a random variable since each  $U_i$  is a random variable and  $f$  is measurable. They also inherit the i.i.d. property from  $U_i$ . That  $f(x)$  is integrable implies  $Ef(U_i) = \int_0^1 f dU_i(x) = \int_0^1 f dx = I$  is finite. Thus we apply the  $L^1$  weak law of large numbers (2.2.14) to determine that  $I_n \rightarrow I$  in probability.

Letting  $\phi(x) = x^2$ , we can apply Chebyshev's inequality to the random variables  $|I_n - I|$  to determine that

$$\frac{a^2}{n} P(|I_n - I| > a/n^{1/2}) \leq E|I_n - I|^2$$

$I$  being equal to the expectation of each  $f(U_i)$  means

$$E(I_n - I)^2 = E(n^{-1}(\sum_{i=1}^n f(U_i) - nI))^2 = n^{-2} \text{var}(\sum_{i=1}^n f(U_i)) = n^{-1} \text{var}(f(U_i)).$$

Combining this result with the previous inequality implies

$$P(|I_n - I| > a/n^{1/2}) \leq a^{-2} \left( \int_0^1 f(x)^2 dx - \left( \int_0^1 f(x) dx \right)^2 \right) < \infty.$$

Where the boundedness of  $\text{var}(f(U_i))$  follows from  $f \in L^2$ .

## 2.2.4

Claim: Let  $X_1, X_2, \dots$  be i.i.d with  $P(X_i = (-1)^k k) = C/k^2 \log k$  for  $k \geq 2$  where  $C$  is chosen to make the sum of the probabilities = 1. Show that  $E|X_i| = \infty$ , but there is a finite constant  $\mu$  so that  $S_n/n \rightarrow \mu$  in probability.

Proof: To see that  $E|X_i| = \infty$  observe that

$$E|X_i| = \sum_{k=2}^{\infty} k \cdot C/k^2 \log k = \sum_{k=2}^{\infty} C/k \log k = C \sum_{k=2}^{\infty} 1/k \log k.$$

The sum multiplying  $C$  in the last expression is a divergent series of nonnegative terms so  $E|X_i| = \infty$ .

To show convergence in probability, we apply Durrett's weak law of large numbers (2.2.12). First note that

$$nP(|X_n| > n) = nC \sum_{k=n+1}^{\infty} 1/k^2 \log k \leq \frac{nC}{\log n} \sum_{k=n+1}^{\infty} \frac{1}{k^2} \leq \frac{nC}{\log n} \int_n^{\infty} \frac{x^2}{x^3} dx = \frac{C}{\log n}.$$

Thus  $nP(|X_n| > n) \rightarrow 0$  as  $n \rightarrow \infty$ , allowing us to apply the WLLN. Furthermore,

$$\mu_n = E(X_1 \mathbf{1}_{(|X_1| \leq n)}) = \sum_{k=2}^n \frac{(-1)^k C}{k \log k} \rightarrow \sum_{k=2}^{\infty} \frac{(-1)^k C}{k \log k} = \mu$$

Where the fact that  $\mu$  is an alternating series implies convergence. So by the WLLN  $S_n/n \rightarrow \mu_n \rightarrow \mu$  as  $n \rightarrow \infty$  proving the claim.