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- CLT is also important to get information on rate of convergence of law of large numbers.

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We can write it as

$$f_{X_{n+1}|X_n,\dots X_0}(x_{n+1}|x_n,\dots,x_0) = f_{X_{n+1}|X_n}(x_{n+1}|x_n), \ \forall x_i$$

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$$= \pi_0(0) (1 - p) p + \pi_0(1) q p$$

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▶ This can easily be seen through a graphical notation.

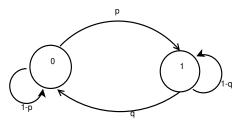
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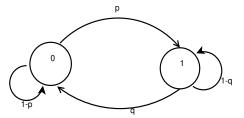
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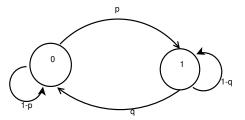


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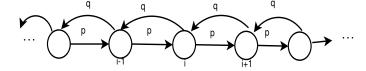
$$Pr[X_0 = 0, X_1 = 1, X_2 = 1, X_3 = 0] = \pi_0(0)p(1-q)q$$

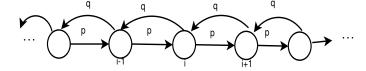
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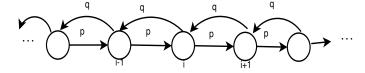
- ▶ A man has 4 umbrellas. carries them from home to office and back when needed. Probability of rain in the morning and evening is same, namely, p.
- What should be the state?
- $S = \{0, 1, \dots, 5\}$. The transition probabilities are

$$P = \begin{bmatrix} 0 & 1 & 2 & 3 & 4 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 & 1 - p & p \\ 2 & 0 & 0 & 1 - p & p & 0 \\ 3 & 0 & 1 - p & p & 0 & 0 \\ 4 & 1 - p & p & 0 & 0 & 0 \end{bmatrix}$$

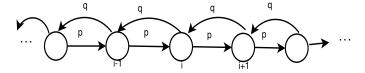




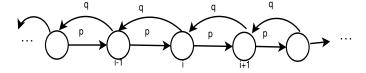
► The following Markov chain is known as a birth-death chain



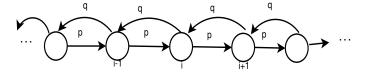
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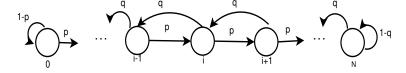


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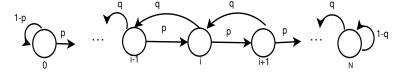


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- Queuing chains can also be birth-death chains

 We can have birth-death chains with finite state space also



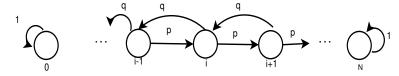
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▶ This chain keeps visiting all the states again and again

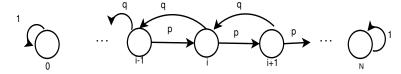
Gambler's Ruin chain

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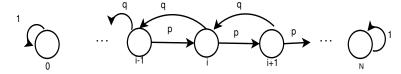
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- ▶ Here state can be the current funds that the gambler has

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- ▶ We first look at one consequence of markov property
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$$Pr[X_{n+m} = y | X_n = x, X_0] = Pr[X_{n+m} = y | X_n = x]$$

► We consider a simple case

$$Pr[X_3 = y | X_1 = x, X_0 = z]$$

$$Pr[X_3 = y | X_1 = x, X_0 = z] = \frac{Pr[X_3 = y, X_1 = x, X_0 = z]}{Pr[X_1 = x, X_0 = z]}$$

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► Thus we get

$$Pr[X_3 = y | X_1 = x, X_0 = z] = Pr[X_3 = y | X_1 = x]$$

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► Or, in general,

$$f_{X_{m+n}|X_m,\cdots,X_0}(y|x,\cdots) = f_{X_{m+n}|X_m}(y|x)$$

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- ightharpoonup That is why we use P^n for n-step transition probabilities

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- We often need conditional probability conditioned on the initial state.
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- We write the above as $P_x(T_y = 1) = P(x, y)$



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 $z\neq y$

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- ▶ Intuitively, all transient states would be visited only finitely many times while recurrent states are visited infinitely often.
- ► For any state y define

$$I_y(X_n) = \begin{cases} 1 & \text{if } X_n = y \\ 0 & \text{otherwise} \end{cases}$$

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$$N_y = \sum_{n=1}^{\infty} I_y(X_n)$$

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$$P_{x}(N_{y} = m) = P_{x}(N_{y} \ge m) - P_{x}(N_{y} \ge m + 1)$$

$$= \rho_{yy}^{m-1} \rho_{xy} - \rho_{yy}^{m} \rho_{xy} = \rho_{xy} \rho_{yy}^{m-1} (1 - \rho_{yy})$$

 \triangleright Now, the total number of visits to y is given by

$$N_y = \sum_{n=1}^{\infty} I_y(X_n)$$

• We can get distribution of N_u as

$$P_{x}(N_{y} \ge 1) = P_{x}(T_{y} < \infty) = \rho_{xy}$$

$$P_{x}(N_{y} \ge 2) = \sum_{m} P_{x}(T_{y} = m)P_{y}(T_{y} < \infty)$$

$$= \rho_{yy} \sum_{m} P_{x}(T_{y} = m) = \rho_{yy} \rho_{xy}$$

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$$P_{x}(N_{y} = 0) = 1 - P_{x}(N_{y} \ge 1) = 1 - \rho_{xy}$$

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▶ G(x,y) is the expected number of visits to y for a chain that is started in x.

(i). Let y be transient. Then

$$P_x(N_y < \infty) = 1, \ \forall x \ \text{ and } \ G(x,y) = \frac{\rho_{xy}}{1 - \rho_{yy}} < \infty, \ \forall x$$

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(ii) Let y be recurrent. Then

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$$P_x[N_y=\infty]=\rho_{xy}, \quad \text{and} \quad G(x,y)=\left\{ \begin{array}{ll} 0 & \text{if} \quad \rho_{xy}=0 \\ \infty & \text{if} \quad \rho_{xy}>0 \end{array} \right.$$