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ightharpoonup We only need marginal distributions of individual X_n to decide whether a sequence converges to a constant in probability

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Weak law of large numbers states

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 Almost sure convergence is a stronger mode of convergence

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- ▶ Hence, $X_n \stackrel{a.s.}{\to} X$ iff

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- ▶ Since $\sum_{n} \frac{\sigma^2}{n\epsilon^2} = \infty$, the Chebyshev bound is not useful
- ▶ We need a bound: $P[|\frac{S_n}{n} \mu|] \le c_n$ such that $\sum_n c_n < \infty$.

$$\left(\sum_{i=1}^{n} (X_i - \mu)\right)^4 =$$

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Let us assume X_i have finite fourth moment

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▶ Hence we get

$$E\left[\left(\sum_{i=1}^{n} (X_i - \mu)\right)^4\right] = nE[(X_i - \mu)^4] + 3n(n-1)\sigma^4 \le C'n^2$$

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▶ Since $\sum_n \frac{C}{n^2} < \infty$, we get $\frac{S_n}{n} \stackrel{a.s.}{\to} \mu$

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- ► Strong law of large numbers says that sample mean converges to the expectation with probability one.

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- In this example X_n converges in r^{th} mean for all r

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- ▶ In general, neither of convergence almost surely and in r^{th} mean imply the other.
- We can generate counter examples for this easily.
- ▶ However, if all X_n take values in a bounded interval, then almost sure convergence implies r^{th} mean convergence

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- ▶ Take $a_n = \frac{1}{n^2}$ and $c_n = e^n$. Then $X_n \stackrel{a.s.}{\to} 0$ but the sequence does not converge in r^{th} mean for any r.

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- ▶ The proofs are straight-forward but we omit the proofs

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- ► The converse is not true. (e.g., sequence of iid random variables)



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► The sequence converges in distribution to an exponential rv

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- ▶ Hence $N_n \stackrel{P}{\rightarrow} \theta$
- ▶ Does it converge almost surely?



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- ▶ In general, almost sure convergence does not imply convergence in r^{th} mean and vice versa

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- Another useful result is the Central Limit Theorem (CLT)
- ► CLT is about (normalized) sums of of independent random variables converging to the Gaussian distribution

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- ▶ Central Limit Theorem states: $\tilde{S}_n \xrightarrow{d} \mathcal{N}(0,1)$

$$S_n = \sum_{i=1}^n X_i \Rightarrow ES_n = n\mu, \operatorname{Var}(S_n) = n\sigma^2$$

- ▶ Given any rv Y, let $Z = \frac{Y EY}{\sqrt{\mathsf{Var}(Y)}}$
- ▶ Then, EZ = 0 and Var(Z) = 1.
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$$\lim_{n \to \infty} P[\tilde{S}_n \le a] = \Phi(a) \triangleq \int_{-\infty}^a \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

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- ▶ We use characteristic functions for proving CLT

Characteristic Function

• Given rv X, its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E\left[e^{iuX}\right] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

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▶ Since $|e^{iux}| \le 1$, ϕ_X exists for all random variables

Properties of characteristic function

$$\phi_X(u) = E\left[e^{iuX}\right] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

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- ϕ is continuous; $|\phi(u)| \le \phi(0) = 1$; $\phi(-u) = \phi^*(u)$
- If Y = aX + b, $\phi_Y(u) = e^{iub}\phi_X(ua)$
- ▶ If $E|X|^r < \infty$, ϕ would be differentiable r times and

$$\phi^{(r)}(u) = E[(iX)^r e^{iuX}]$$

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▶ If all moments exist, then

$$\phi_X(u) = \sum_{s=0}^{\infty} \mu_s \, \frac{(iu)^s}{s!}$$

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 - If $F_n \to F$ then $\phi_{F_n} \to \phi_F$
 - ▶ If $\phi_{F_n} \to \psi$ and ψ is continuous at zero, then ψ would be characteristic function of some df, say, F, and $F_n \to F$

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 $\blacktriangleright \ \, \mathrm{Let} \,\, X \sim \mathcal{N}(0,1)$

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