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- ▶ If  $X, Y$  are independent,  $E[h(X)|Y] = E[h(X)]$

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$$P(A) = E[I_A] = E[ E[I_A|Y] ]$$

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- Actually, we did not use independence of  $X_i$ .

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If  $E[X_i]$  is same for all  $i$ ,  $ES_N = EX_1 EN$ .
- ▶ Assume  $X_i$  are iid. Suppose the event  $[N \leq n - 1]$  depends only on  $X_1, \dots, X_{n-1}$ .

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- ▶ With iid  $X_i$ , the formula  $ES = EN EX_1$  is valid even under some dependence between  $N$  and  $X_i$ .
- ▶ Here is one version of Wald's formula. We assume
  1.  $E[|X_i|] < \infty, \forall i$  and  $EN < \infty$ .
  2.  $E[X_n I_{[N \geq n]}] = E[X_n]P[N \geq n], \forall n$
- ▶ Let  $S_N = \sum_{i=1}^N X_i$  and let  $T_N = \sum_{i=1}^N E[X_i]$ .
- ▶ Then,  $ES_N = ET_N$ .

If  $E[X_i]$  is same for all  $i$ ,  $ES_N = EX_1 EN$ .
- ▶ Assume  $X_i$  are iid. Suppose the event  $[N \leq n - 1]$  depends only on  $X_1, \dots, X_{n-1}$ .
- ▶ Then the event  $[N \leq n - 1]$  and hence its complement  $[N \geq n]$  is independent of  $X_n$  and the assumption above is satisfied.

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- ▶ Such an  $N$  is an example of what is called a stopping time.

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- ▶ Now we can calculate  $P[X = k]$  using the conditioning argument.



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- ▶ So, we get:  $P[X = k] = \frac{1}{n+1}$ ,  $k = 0, 1, \dots, n$





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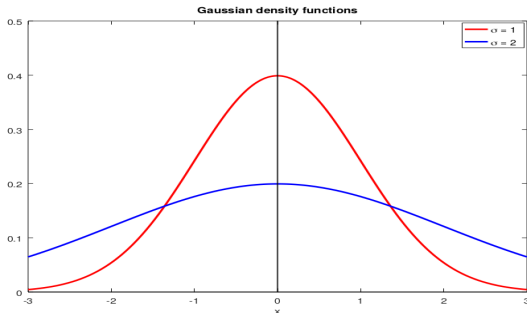
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- ▶ We will now show that this is a joint density function.

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- ▶ This is the multidimensional Gaussian distribution

- Consider  $\mathbf{Y}$  with joint density

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This shows that  $Z_i \sim \mathcal{N}(0, \frac{1}{m_i})$  and  $Z_i$  are independent.



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# Multi-dimensional Gaussian density

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- ▶ This implies  $X_i$  are independent.

# Multi-dimensional Gaussian density

- ▶  $\mathbf{X} = (X_1, \dots, X_n)^T$  are said to be jointly Gaussian if

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \Sigma^{-1}(\mathbf{x}-\boldsymbol{\mu})}$$

- ▶  $E\mathbf{X} = \boldsymbol{\mu}$  and  $\Sigma_X = \Sigma$ .
- ▶ Suppose  $\text{Cov}(X_i, X_j) = 0, \forall i \neq j$ .
- ▶ Then  $\Sigma_{ij} = 0, \forall i \neq j$ . Let  $\Sigma = \text{diag}(\sigma_1^2, \dots, \sigma_n^2)$ .

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} \sigma_1 \cdots \sigma_n} e^{-\frac{1}{2} \sum_{i=1}^n \left(\frac{x_i - \mu_i}{\sigma_i}\right)^2} = \prod_{i=1}^n \frac{1}{\sigma_i \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x_i - \mu_i}{\sigma_i}\right)^2}$$

- ▶ This implies  $X_i$  are independent.
- ▶ If  $X_1, \dots, X_n$  are jointly Gaussian then uncorrelatedness implies independence.



- Let  $\mathbf{X} = (X_1, \dots, X_n)^T$  be jointly Gaussian:

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- ▶ Then we saw that  $Z_i \sim \mathcal{N}(0, \frac{1}{m_i})$  and  $Z_i$  are independent.
- ▶ If  $X_1, \dots, X_n$  are jointly Gaussian then there is a ‘linear’ transform that transforms them into independent random variables.

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- ▶ This is the moment generating function of multi-dimensional Normal density



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$$f_{XY}(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left(\frac{x^2}{\sigma_x^2} + \frac{y^2}{\sigma_y^2} - \frac{2\rho xy}{\sigma_x\sigma_y}\right)}$$

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- ▶ This is the bivariate Gaussian density

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- ▶ Exercise for you – show all this starting with the joint density we have

- ▶ Suppose  $X, Y$  are jointly Gaussian (with the density above)
- ▶ Then, all the marginals and conditionals would be Gaussian.
- ▶  $X \sim \mathcal{N}(0, \sigma_x^2)$ , and  $Y \sim \mathcal{N}(0, \sigma_y^2)$
- ▶  $f_{X|Y}(x|y)$  would be a Gaussian density with mean  $y\rho\frac{\sigma_x}{\sigma_y}$  and variance  $\sigma_x^2(1 - \rho^2)$ .
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- ▶ Note that  $X, Y$  are individually Gaussian does not mean they are jointly Gaussian (unless they are independent)

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- ▶ We will prove this using moment generating functions