

Recap: Modes of convergence

Recap: Modes of convergence

► $X_n \xrightarrow{d} X$ iff

$$F_n(x) \rightarrow F(x), \quad \forall x \text{ where } F \text{ is continuous}$$

Recap: Modes of convergence

- ▶ $X_n \xrightarrow{d} X$ iff

$$F_n(x) \rightarrow F(x), \quad \forall x \text{ where } F \text{ is continuous}$$

- ▶ $X_n \xrightarrow{P} X$ iff

$$\lim_{n \rightarrow \infty} P[|X_n - X| > \epsilon] = 0, \quad \forall \epsilon > 0$$

Recap: Modes of convergence

- ▶ $X_n \xrightarrow{d} X$ iff

$$F_n(x) \rightarrow F(x), \quad \forall x \text{ where } F \text{ is continuous}$$

- ▶ $X_n \xrightarrow{P} X$ iff

$$\lim_{n \rightarrow \infty} P[|X_n - X| > \epsilon] = 0, \quad \forall \epsilon > 0$$

- ▶ $X_n \xrightarrow{r} X$ iff

$$E[|X_n - X|^r] \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

Recap: Modes of convergence

- ▶ $X_n \xrightarrow{d} X$ iff

$$F_n(x) \rightarrow F(x), \quad \forall x \text{ where } F \text{ is continuous}$$

- ▶ $X_n \xrightarrow{P} X$ iff

$$\lim_{n \rightarrow \infty} P[|X_n - X| > \epsilon] = 0, \quad \forall \epsilon > 0$$

- ▶ $X_n \xrightarrow{r} X$ iff

$$E[|X_n - X|^r] \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

- ▶ $X_n \xrightarrow{a.s} X$ iff

$$P[X_n \rightarrow X] = 1 \quad \text{or} \quad P[\limsup |X_n - X| > \epsilon] = 0$$

Recap

- ▶ We have the following relations among different modes of convergence

Recap

- ▶ We have the following relations among different modes of convergence

$$X_n \xrightarrow{r} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

Recap

- ▶ We have the following relations among different modes of convergence

$$X_n \xrightarrow{r} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

$$X_n \xrightarrow{a.s.} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

Recap

- ▶ We have the following relations among different modes of convergence

$$X_n \xrightarrow{r} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

$$X_n \xrightarrow{a.s.} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

- ▶ All the implications are one-way and we have seen counter examples

Recap

- ▶ We have the following relations among different modes of convergence

$$X_n \xrightarrow{r} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

$$X_n \xrightarrow{a.s.} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$$

- ▶ All the implications are one-way and we have seen counter examples
- ▶ In general, almost sure convergence does not imply convergence in r^{th} mean and vice versa

Recap

Recap

- ▶ Given X_i are iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$

Recap

- ▶ Given X_i are iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Weak law of large numbers: $\frac{S_n}{n} \xrightarrow{P} \mu$

Recap

- ▶ Given X_i are iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Weak law of large numbers: $\frac{S_n}{n} \xrightarrow{P} \mu$
- ▶ strong law of large numbers: $\frac{S_n}{n} \xrightarrow{a.s.} \mu$

Recap

- ▶ Given X_i are iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Weak law of large numbers: $\frac{S_n}{n} \xrightarrow{P} \mu$
- ▶ strong law of large numbers: $\frac{S_n}{n} \xrightarrow{a.s.} \mu$
- ▶ Central Limit Theorem: $\frac{S_n - n\mu}{\sigma\sqrt{n}} \xrightarrow{d} \mathcal{N}(0, 1)$

Recap

- ▶ Take X_i iid, $EX_i = 0$, $\text{Var}(X_i) = 1$, $n = 1, 2, \dots$

Recap

- ▶ Take X_i iid, $EX_i = 0$, $\text{Var}(X_i) = 1$, $n = 1, 2, \dots$
- ▶ $S_n = \sum_{i=1}^n X_i$

Recap

- ▶ Take X_i iid, $EX_i = 0$, $\text{Var}(X_i) = 1$, $n = 1, 2, \dots$
- ▶ $S_n = \sum_{i=1}^n X_i$
- ▶ Strong law of large numbers implies

$$\frac{S_n}{n} \xrightarrow{a.s.} 0$$

Recap

- ▶ Take X_i iid, $EX_i = 0$, $\text{Var}(X_i) = 1$, $n = 1, 2, \dots$
- ▶ $S_n = \sum_{i=1}^n X_i$
- ▶ Strong law of large numbers implies

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} 0$$

- ▶ Central Limit Theorem implies

$$\frac{S_n}{\sqrt{n}} \xrightarrow{\text{a.s.}} \mathcal{N}(0, 1)$$

Recap: Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

Recap: Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

- ▶ Since $|e^{iux}| \leq 1$, ϕ_X exists for all random variables

Recap: Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

- ▶ Since $|e^{iux}| \leq 1$, ϕ_X exists for all random variables
 - ▶ ϕ is continuous; $|\phi(u)| \leq \phi(0) = 1$; $\phi(-u) = \phi^*(u)$

Recap: Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

- ▶ Since $|e^{iux}| \leq 1$, ϕ_X exists for all random variables
 - ▶ ϕ is continuous; $|\phi(u)| \leq \phi(0) = 1$; $\phi(-u) = \phi^*(u)$
 - ▶ If $Y = aX + b$, $\phi_Y(u) = e^{iub}\phi_X(ua)$

Recap: Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

- ▶ Since $|e^{iux}| \leq 1$, ϕ_X exists for all random variables
 - ▶ ϕ is continuous; $|\phi(u)| \leq \phi(0) = 1$; $\phi(-u) = \phi^*(u)$
 - ▶ If $Y = aX + b$, $\phi_Y(u) = e^{iub}\phi_X(ua)$
 - ▶ If $E|X|^r < \infty$, ϕ would be differentiable r times and

$$\phi^{(r)}(u) = E[(iX)^r e^{iuX}]$$

Recap

- ▶ Let $\mu_r = E[X^r]$ and let $\nu_r = E[|X|^r]$

Recap

- ▶ Let $\mu_r = E[X^r]$ and let $\nu_r = E[|X|^r]$
- ▶ If ν_r is finite, then

$$\phi_X(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}$$

where $|\rho(u)| \leq 1$ and $\rho(u) \rightarrow 1$ as $u \rightarrow 0$

Recap

- ▶ Let $\mu_r = E[X^r]$ and let $\nu_r = E[|X|^r]$
- ▶ If ν_r is finite, then

$$\phi_X(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}$$

where $|\rho(u)| \leq 1$ and $\rho(u) \rightarrow 1$ as $u \rightarrow 0$

- ▶ If all moments exist, then

$$\phi_X(u) = \sum_{s=0}^{\infty} \mu_s \frac{(iu)^s}{s!}$$

Recap

- ▶ We denote by ϕ_F characteristic function of df F

Recap

- ▶ We denote by ϕ_F characteristic function of df F
- ▶ Let F_n be a sequence of distribution functions

Recap

- ▶ We denote by ϕ_F characteristic function of df F
- ▶ Let F_n be a sequence of distribution functions
- ▶ **Continuity theorem**

Recap

- ▶ We denote by ϕ_F characteristic function of df F
- ▶ Let F_n be a sequence of distribution functions
- ▶ **Continuity theorem**
 - ▶ If $F_n \rightarrow F$ then $\phi_{F_n} \rightarrow \phi_F$

Recap

- ▶ We denote by ϕ_F characteristic function of df F
- ▶ Let F_n be a sequence of distribution functions
- ▶ **Continuity theorem**
 - ▶ If $F_n \rightarrow F$ then $\phi_{F_n} \rightarrow \phi_F$
 - ▶ If $\phi_{F_n} \rightarrow \psi$ and ψ is continuous at zero, then ψ would be characteristic function of some df, say, F , and $F_n \rightarrow F$

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

Proof:

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

Proof:

- ▶ Without loss of generality let us assume $\mu = 0$.

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

Proof:

- ▶ Without loss of generality let us assume $\mu = 0$.
- ▶ We use characteristic function of \tilde{S}_n for the proof.

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

Proof:

- ▶ Without loss of generality let us assume $\mu = 0$.
- ▶ We use characteristic function of \tilde{S}_n for the proof.
- ▶ Let ϕ be the characteristic function of X_i .

- ▶ Given X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ Let $\tilde{S}_n = \frac{S_n - ES_n}{\sqrt{\text{var}(S_n)}} = \frac{S_n - n\mu}{\sigma\sqrt{n}}$
- ▶ **(Lindberg-Levy) Central Limit Theorem**

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

Proof:

- ▶ Without loss of generality let us assume $\mu = 0$.
- ▶ We use characteristic function of \tilde{S}_n for the proof.
- ▶ Let ϕ be the characteristic function of X_i . Then

$$\phi_{S_n}(t) = (\phi(t))^n \quad \text{and} \quad \phi_{\tilde{S}_n}(t) = \left(\phi \left(\frac{t}{\sigma\sqrt{n}} \right) \right)^n$$

- Recall that we can expand ϕ in a Taylor series

$$\phi(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}, \quad \rho(u) \rightarrow 1, \text{ as } u \rightarrow 0$$

- Recall that we can expand ϕ in a Taylor series

$$\phi(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}, \quad \rho(u) \rightarrow 1, \text{ as } u \rightarrow 0$$

- Here we assume: $EX_i = 0$ and $EX_i^2 = \sigma^2$

$$\phi(t) = 1 + 0 - \frac{1}{2} \rho(t) \sigma^2 t^2$$

- Recall that we can expand ϕ in a Taylor series

$$\phi(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}, \quad \rho(u) \rightarrow 1, \text{ as } u \rightarrow 0$$

- Here we assume: $EX_i = 0$ and $EX_i^2 = \sigma^2$

$$\phi(t) = 1 + 0 - \frac{1}{2} \rho(t) \sigma^2 t^2$$

$$\phi\left(\frac{t}{\sigma\sqrt{n}}\right) = 1 - \frac{1}{2} \rho\left(\frac{t}{\sigma\sqrt{n}}\right) \sigma^2 \frac{t^2}{\sigma^2 n}$$

- Recall that we can expand ϕ in a Taylor series

$$\phi(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}, \quad \rho(u) \rightarrow 1, \text{ as } u \rightarrow 0$$

- Here we assume: $EX_i = 0$ and $EX_i^2 = \sigma^2$

$$\phi(t) = 1 + 0 - \frac{1}{2} \rho(t) \sigma^2 t^2$$

$$\begin{aligned} \phi\left(\frac{t}{\sigma\sqrt{n}}\right) &= 1 - \frac{1}{2} \rho\left(\frac{t}{\sigma\sqrt{n}}\right) \sigma^2 \frac{t^2}{\sigma^2 n} \\ &= 1 - \frac{1}{2} \frac{t^2}{n} + \frac{1}{2} \frac{t^2}{n} \left(1 - \rho\left(\frac{t}{\sigma\sqrt{n}}\right)\right) \end{aligned}$$

- Recall that we can expand ϕ in a Taylor series

$$\phi(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}, \quad \rho(u) \rightarrow 1, \text{ as } u \rightarrow 0$$

- Here we assume: $EX_i = 0$ and $EX_i^2 = \sigma^2$

$$\phi(t) = 1 + 0 - \frac{1}{2} \rho(t) \sigma^2 t^2$$

$$\begin{aligned} \phi\left(\frac{t}{\sigma\sqrt{n}}\right) &= 1 - \frac{1}{2} \rho\left(\frac{t}{\sigma\sqrt{n}}\right) \sigma^2 \frac{t^2}{\sigma^2 n} \\ &= 1 - \frac{1}{2} \frac{t^2}{n} + \frac{1}{2} \frac{t^2}{n} \left(1 - \rho\left(\frac{t}{\sigma\sqrt{n}}\right)\right) \\ &= 1 - \frac{1}{2} \frac{t^2}{n} + o\left(\frac{1}{n}\right) \end{aligned}$$

► Hence we get

► Hence we get

$$\lim_{n \rightarrow \infty} \phi_{\tilde{S}_n}(t) = \lim_{n \rightarrow \infty} \left(\phi \left(\frac{t}{\sigma \sqrt{n}} \right) \right)^n$$

► Hence we get

$$\begin{aligned}\lim_{n \rightarrow \infty} \phi_{\tilde{S}_n}(t) &= \lim_{n \rightarrow \infty} \left(\phi \left(\frac{t}{\sigma \sqrt{n}} \right) \right)^n \\ &= \lim_{n \rightarrow \infty} \left(1 - \frac{1}{2} \frac{t^2}{n} + o \left(\frac{1}{n} \right) \right)^n\end{aligned}$$

► Hence we get

$$\begin{aligned}\lim_{n \rightarrow \infty} \phi_{\tilde{S}_n}(t) &= \lim_{n \rightarrow \infty} \left(\phi \left(\frac{t}{\sigma \sqrt{n}} \right) \right)^n \\ &= \lim_{n \rightarrow \infty} \left(1 - \frac{1}{2} \frac{t^2}{n} + o \left(\frac{1}{n} \right) \right)^n \\ &= e^{-\frac{t^2}{2}}\end{aligned}$$

which is the characteristic function of standard normal

- ▶ Hence we get

$$\begin{aligned}\lim_{n \rightarrow \infty} \phi_{\tilde{S}_n}(t) &= \lim_{n \rightarrow \infty} \left(\phi \left(\frac{t}{\sigma \sqrt{n}} \right) \right)^n \\ &= \lim_{n \rightarrow \infty} \left(1 - \frac{1}{2} \frac{t^2}{n} + o \left(\frac{1}{n} \right) \right)^n \\ &= e^{-\frac{t^2}{2}}\end{aligned}$$

which is the characteristic function of standard normal

- ▶ By Continuity theorem, distribution function of \tilde{S}_n converges to that of standard Normal rv

$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's
- ▶ Let X_i be iid and $S_n = \sum_{i=1}^n X_i$

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's
- ▶ Let X_i be iid and $S_n = \sum_{i=1}^n X_i$

$$P[S_n \leq x] =$$

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's
- ▶ Let X_i be iid and $S_n = \sum_{i=1}^n X_i$

$$P[S_n \leq x] = P\left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq \frac{x - n\mu}{\sigma\sqrt{n}}\right]$$

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's
- ▶ Let X_i be iid and $S_n = \sum_{i=1}^n X_i$

$$P[S_n \leq x] = P\left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq \frac{x - n\mu}{\sigma\sqrt{n}}\right] \approx \Phi\left(\frac{x - n\mu}{\sigma\sqrt{n}}\right)$$

- ▶ What CLT says is that sums of iid random variables, when appropriately normalized, would always approach the Gaussian distribution.
- ▶ It allows one to approximate distribution of sums of independent rv's
- ▶ Let X_i be iid and $S_n = \sum_{i=1}^n X_i$

$$P[S_n \leq x] = P\left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq \frac{x - n\mu}{\sigma\sqrt{n}}\right] \approx \Phi\left(\frac{x - n\mu}{\sigma\sqrt{n}}\right)$$

- ▶ Thus, S_n is well approximated by a normal rv with mean $n\mu$ and variance $n\sigma^2$, if n is large

Example

- ▶ Twenty numbers are rounded off to the nearest integer and added. What is the probability that the sum obtained differs from true sum by more than 3.

Example

- ▶ Twenty numbers are rounded off to the nearest integer and added. What is the probability that the sum obtained differs from true sum by more than 3.
- ▶ A reasonable assumption is round-off errors are independent and uniform over $[-0.5, 0.5]$

Example

- ▶ Twenty numbers are rounded off to the nearest integer and added. What is the probability that the sum obtained differs from true sum by more than 3.
- ▶ A reasonable assumption is round-off errors are independent and uniform over $[-0.5, 0.5]$
- ▶ Take $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid.

Example

- ▶ Twenty numbers are rounded off to the nearest integer and added. What is the probability that the sum obtained differs from true sum by more than 3.
- ▶ A reasonable assumption is round-off errors are independent and uniform over $[-0.5, 0.5]$
- ▶ Take $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid.
- ▶ Then Z represents the error in the sum.

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$P[|Z| \leq 3] = P[-3 \leq Z \leq 3]$$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$\begin{aligned}
 P[|Z| \leq 3] &= P[-3 \leq Z \leq 3] \\
 &= P\left[\frac{-3}{\sqrt{\frac{5}{3}}} \leq \frac{Z - EZ}{\sqrt{\text{Var}(Z)}} \leq \frac{3}{\sqrt{\frac{5}{3}}}\right]
 \end{aligned}$$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$\begin{aligned} P[|Z| \leq 3] &= P[-3 \leq Z \leq 3] \\ &= P\left[\frac{-3}{\sqrt{\frac{5}{3}}} \leq \frac{Z - EZ}{\sqrt{\text{Var}(Z)}} \leq \frac{3}{\sqrt{\frac{5}{3}}}\right] \\ &\approx \Phi\left(\frac{3}{\sqrt{\frac{5}{3}}}\right) - \Phi\left(\frac{-3}{\sqrt{\frac{5}{3}}}\right) \end{aligned}$$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$\begin{aligned} P[|Z| \leq 3] &= P[-3 \leq Z \leq 3] \\ &= P\left[\frac{-3}{\sqrt{\frac{5}{3}}} \leq \frac{Z - EZ}{\sqrt{\text{Var}(Z)}} \leq \frac{3}{\sqrt{\frac{5}{3}}}\right] \\ &\approx \Phi\left(\frac{3}{\sqrt{\frac{5}{3}}}\right) - \Phi\left(\frac{-3}{\sqrt{\frac{5}{3}}}\right) \\ &\approx \Phi(2.3) - \Phi(-2.3) \end{aligned}$$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$\begin{aligned} P[|Z| \leq 3] &= P[-3 \leq Z \leq 3] \\ &= P\left[\frac{-3}{\sqrt{\frac{5}{3}}} \leq \frac{Z - EZ}{\sqrt{\text{Var}(Z)}} \leq \frac{3}{\sqrt{\frac{5}{3}}}\right] \\ &\approx \Phi\left(\frac{3}{\sqrt{\frac{5}{3}}}\right) - \Phi\left(\frac{-3}{\sqrt{\frac{5}{3}}}\right) \\ &\approx \Phi(2.3) - \Phi(-2.3) \\ &= 0.9893 - 0.0107 \approx 0.98 \end{aligned}$$

- ▶ $Z = \sum_{i=1}^{20} X_i$, $X_i \sim U[-0.5, 0.5]$, X_i iid
- ▶ $EX_i = 0$ and $\text{Var}(X_i) = \frac{1}{12}$.
- ▶ Hence, $EZ = 0$ and $\text{Var}(Z) = \frac{20}{12} = \frac{5}{3}$

$$\begin{aligned}
 P[|Z| \leq 3] &= P[-3 \leq Z \leq 3] \\
 &= P\left[\frac{-3}{\sqrt{\frac{5}{3}}} \leq \frac{Z - EZ}{\sqrt{\text{Var}(Z)}} \leq \frac{3}{\sqrt{\frac{5}{3}}}\right] \\
 &\approx \Phi\left(\frac{3}{\sqrt{\frac{5}{3}}}\right) - \Phi\left(\frac{-3}{\sqrt{\frac{5}{3}}}\right) \\
 &\approx \Phi(2.3) - \Phi(-2.3) \\
 &= 0.9893 - 0.0107 \approx 0.98
 \end{aligned}$$

- ▶ Hence probability that the sum differs from true sum by more than 3 is 0.02

- ▶ We can approximate binomial rv with Gaussian for large n

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:

$$S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$$

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:

$$S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$$

- ▶ Hence we can approximate distribution of S_n by

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:
 $S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$
- ▶ Hence we can approximate distribution of S_n by

$$P[S_n \leq x] = P\left[\frac{S_n - np}{\sqrt{np(1-p)}} \leq \frac{x - np}{\sqrt{np(1-p)}}\right]$$

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:
 $S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$
- ▶ Hence we can approximate distribution of S_n by

$$\begin{aligned} P[S_n \leq x] &= P \left[\frac{S_n - np}{\sqrt{np(1-p)}} \leq \frac{x - np}{\sqrt{np(1-p)}} \right] \\ &\approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right) \end{aligned}$$

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:
 $S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$
- ▶ Hence we can approximate distribution of S_n by

$$P[S_n \leq x] = P\left[\frac{S_n - np}{\sqrt{np(1-p)}} \leq \frac{x - np}{\sqrt{np(1-p)}}\right]$$

$$\approx \Phi\left(\frac{x - np}{\sqrt{np(1-p)}}\right)$$

- ▶ For large n , binomial rv is like a Gaussian rv with mean np and variance $np(1-p)$

- ▶ We can approximate binomial rv with Gaussian for large n
- ▶ Binomial random variable with parameters n, p is a sum of n independent Bernoulli variables:
 $S_n = \sum_{i=1}^n X_i; \quad X_i \in \{0, 1\}, \quad P[X_i = 1] = p, \quad X_i \text{ ind}$
- ▶ Hence we can approximate distribution of S_n by

$$P[S_n \leq x] = P\left[\frac{S_n - np}{\sqrt{np(1-p)}} \leq \frac{x - np}{\sqrt{np(1-p)}}\right]$$

$$\approx \Phi\left(\frac{x - np}{\sqrt{np(1-p)}}\right)$$

- ▶ For large n , binomial rv is like a Gaussian rv with mean np and variance $np(1-p)$
- ▶ The approximation is quite good in practice

- ▶ S_n be binomial with parameters n, p

- ▶ S_n be binomial with parameters n, p

$$P[S_n \leq x] \approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right)$$

- ▶ S_n be binomial with parameters n, p

$$P[S_n \leq x] \approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right)$$

- ▶ For example, with $p = 0.95$

$$P[S_{110} \leq 100] \approx \Phi \left(\frac{100 - 110 * 0.95}{\sqrt{110 * 0.05 * 0.95}} \right)$$

- ▶ S_n be binomial with parameters n, p

$$P[S_n \leq x] \approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right)$$

- ▶ For example, with $p = 0.95$

$$P[S_{110} \leq 100] \approx \Phi \left(\frac{100 - 110 * 0.95}{\sqrt{110 * 0.05 * 0.95}} \right) \approx \Phi(-1.97) = 0.025$$

- ▶ S_n be binomial with parameters n, p

$$P[S_n \leq x] \approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right)$$

- ▶ For example, with $p = 0.95$

$$P[S_{110} \leq 100] \approx \Phi \left(\frac{100 - 110 * 0.95}{\sqrt{110 * 0.05 * 0.95}} \right) \approx \Phi(-1.97) = 0.025$$

- ▶ Since S_n is integer-valued, the LHS above is same for all x between two consecutive integers; but RHS changes

- ▶ S_n be binomial with parameters n, p

$$P[S_n \leq x] \approx \Phi \left(\frac{x - np}{\sqrt{np(1-p)}} \right)$$

- ▶ For example, with $p = 0.95$

$$P[S_{110} \leq 100] \approx \Phi \left(\frac{100 - 110 * 0.95}{\sqrt{110 * 0.05 * 0.95}} \right) \approx \Phi(-1.97) = 0.025$$

- ▶ Since S_n is integer-valued, the LHS above is same for all x between two consecutive integers; but RHS changes
- ▶ To get a good approximation, to calculate $P[S_n \leq m]$ one uses $P[S_n \leq m + 0.5]$ in the above approximation formula

- ▶ CLT allows one to get rate of convergence of law of large numbers

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.
- ▶ Now, by CLT

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.
- ▶ Now, by CLT

$$P \left[\left| \frac{S_n}{n} - \mu \right| > \epsilon \right] = P [|S_n - n\mu| > n\epsilon]$$

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.
- ▶ Now, by CLT

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - \mu \right| > \epsilon \right] &= P [|S_n - n\mu| > n\epsilon] \\ &= P \left[\left| \frac{S_n - n\mu}{\sigma\sqrt{n}} \right| > \frac{n\epsilon}{\sigma\sqrt{n}} \right] \end{aligned}$$

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.
- ▶ Now, by CLT

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - \mu \right| > \epsilon \right] &= P [|S_n - n\mu| > n\epsilon] \\ &= P \left[\left| \frac{S_n - n\mu}{\sigma\sqrt{n}} \right| > \frac{n\epsilon}{\sigma\sqrt{n}} \right] \\ &\approx 1 - \left(\Phi \left(\frac{n\epsilon}{\sigma\sqrt{n}} \right) - \Phi \left(-\frac{n\epsilon}{\sigma\sqrt{n}} \right) \right) \end{aligned}$$

- ▶ CLT allows one to get rate of convergence of law of large numbers
- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$
- ▶ By Law of large numbers, $\frac{S_n}{n} \rightarrow \mu$.
- ▶ Now, by CLT

$$\begin{aligned}
 P \left[\left| \frac{S_n}{n} - \mu \right| > \epsilon \right] &= P [|S_n - n\mu| > n\epsilon] \\
 &= P \left[\left| \frac{S_n - n\mu}{\sigma\sqrt{n}} \right| > \frac{n\epsilon}{\sigma\sqrt{n}} \right] \\
 &\approx 1 - \left(\Phi \left(\frac{n\epsilon}{\sigma\sqrt{n}} \right) - \Phi \left(-\frac{n\epsilon}{\sigma\sqrt{n}} \right) \right) \\
 &= 2 \left(1 - \Phi \left(\frac{n\epsilon}{\sigma\sqrt{n}} \right) \right)
 \end{aligned}$$

(because $\Phi(-x) = (1 - \Phi(x))$)

Example: Opinion Polls

- ▶ let p denote the fraction of population that prefers product A to product B

Example: Opinion Polls

- ▶ let p denote the fraction of population that prefers product A to product B
- ▶ We want to estimate p

Example: Opinion Polls

- ▶ let p denote the fraction of population that prefers product A to product B
- ▶ We want to estimate p
- ▶ We conduct a sample survey by asking n people

Example: Opinion Polls

- ▶ let p denote the fraction of population that prefers product A to product B
- ▶ We want to estimate p
- ▶ We conduct a sample survey by asking n people
- ▶ We want to make a statement such as

$$p = 0.34 \pm 0.07 \text{ with a confidence of } 95\%$$

Example: Opinion Polls

- ▶ let p denote the fraction of population that prefers product A to product B
- ▶ We want to estimate p
- ▶ We conduct a sample survey by asking n people
- ▶ We want to make a statement such as
$$p = 0.34 \pm 0.07 \text{ with a confidence of } 95\%$$
- ▶ Here, the 0.34 would be the sample mean. The other two numbers can be fixed using CLT

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = P [|S_n - np| > n\epsilon]$$

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] &= P [|S_n - np| > n\epsilon] \\ &= 2 \left(1 - \Phi \left(\frac{n\epsilon}{\sqrt{np(1-p)}} \right) \right) \end{aligned}$$

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] &= P [|S_n - np| > n\epsilon] \\ &= 2 \left(1 - \Phi \left(\frac{n\epsilon}{\sqrt{np(1-p)}} \right) \right) \end{aligned}$$

- ▶ Suppose we want to satisfy

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = \delta$$

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] &= P [|S_n - np| > n\epsilon] \\ &= 2 \left(1 - \Phi \left(\frac{n\epsilon}{\sqrt{np(1-p)}} \right) \right) \end{aligned}$$

- ▶ Suppose we want to satisfy

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = \delta$$

- ▶ We can calculate any one of ϵ , δ or n given the other two using the earlier equation.

- ▶ $X_i \in \{0, 1\}$ iid, $EX_i = p$, $S_n = \sum_{i=1}^n X_i$
- ▶ Now, by CLT, we have

$$\begin{aligned} P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] &= P [|S_n - np| > n\epsilon] \\ &= 2 \left(1 - \Phi \left(\frac{n\epsilon}{\sqrt{np(1-p)}} \right) \right) \end{aligned}$$

- ▶ Suppose we want to satisfy

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = \delta$$

- ▶ We can calculate any one of ϵ , δ or n given the other two using the earlier equation.
- ▶ But we need value of p for it!

- ▶ Fortunately, $\sqrt{p(1-p)}$ does not change too much with p

- ▶ Fortunately, $\sqrt{p(1-p)}$ does not change too much with p
- ▶ It attains its maximum value of 0.5 at $p = 0.5$

- ▶ Fortunately, $\sqrt{p(1-p)}$ does not change too much with p
- ▶ It attains its maximum value of 0.5 at $p = 0.5$
- ▶ It is 0.458 at $p = 0.3$ and is 0.4 at $p = 0.2$

- ▶ Fortunately, $\sqrt{p(1-p)}$ does not change too much with p
- ▶ It attains its maximum value of 0.5 at $p = 0.5$
- ▶ It is 0.458 at $p = 0.3$ and is 0.4 at $p = 0.2$
- ▶ One normally fixes this variance as 0.5 or 0.45 to calculate the sample size, n .

- ▶ Fortunately, $\sqrt{p(1-p)}$ does not change too much with p
- ▶ It attains its maximum value of 0.5 at $p = 0.5$
- ▶ It is 0.458 at $p = 0.3$ and is 0.4 at $p = 0.2$
- ▶ One normally fixes this variance as 0.5 or 0.45 to calculate the sample size, n .
- ▶ There are other ways of handling it

► We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.
Let us approximate $\sqrt{p(1-p)} = 0.45$.

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.

Let us approximate $\sqrt{p(1-p)} = 0.45$. Then

$$2 \left(1 - \Phi \left(\frac{0.025 * 30}{0.45} \right) \right) = 2(1 - \Phi(1.66)) \approx 0.1$$

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.

Let us approximate $\sqrt{p(1-p)} = 0.45$. Then

$$2 \left(1 - \Phi \left(\frac{0.025 * 30}{0.45} \right) \right) = 2(1 - \Phi(1.66)) \approx 0.1$$

- If we took $\sqrt{p(1-p)} = 0.5$ we get the value as 0.14

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.

Let us approximate $\sqrt{p(1-p)} = 0.45$. Then

$$2 \left(1 - \Phi \left(\frac{0.025 * 30}{0.45} \right) \right) = 2(1 - \Phi(1.66)) \approx 0.1$$

- If we took $\sqrt{p(1-p)} = 0.5$ we get the value as 0.14
- If we use Chebyshev inequality with variance as 0.5 we get the bound as 0.8

- We have

$$P \left[\left| \frac{S_n}{n} - p \right| > \epsilon \right] = 2 \left(1 - \Phi \left(\frac{\epsilon \sqrt{n}}{\sqrt{p(1-p)}} \right) \right)$$

- Suppose $n = 900$ and $\epsilon = 0.025$.

Let us approximate $\sqrt{p(1-p)} = 0.45$. Then

$$2 \left(1 - \Phi \left(\frac{0.025 * 30}{0.45} \right) \right) = 2(1 - \Phi(1.66)) \approx 0.1$$

- If we took $\sqrt{p(1-p)} = 0.5$ we get the value as 0.14
- If we use Chebyshev inequality with variance as 0.5 we get the bound as 0.8
- If we change ϵ to 0.05, then at variance equal to 0.5 the probability becomes about 0.02 while the Chebyshev bound would be about 0.2

Confidence intervals

Confidence intervals

- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$.

Confidence intervals

- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$.
- ▶ Using CLT, we get

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

Confidence intervals

- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$.
- ▶ Using CLT, we get

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ If the RHS above is δ , then we can say that $\frac{S_n}{n} \in [\mu - c, \mu + c]$ with probability $(1 - \delta)$

Confidence intervals

- ▶ Let X_i iid, $EX_i = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$.
- ▶ Using CLT, we get

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ If the RHS above is δ , then we can say that $\frac{S_n}{n} \in [\mu - c, \mu + c]$ with probability $(1 - \delta)$
- ▶ This interval is called the $100(1 - \delta)\%$ confidence interval.

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- Suppose $c = \frac{1.96\sigma}{\sqrt{n}}$

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ Suppose $c = \frac{1.96\sigma}{\sqrt{n}}$
- ▶ Then

$$P \left[\left| \frac{S_n}{n} - \mu \right| > \frac{1.96\sigma}{\sqrt{n}} \right] = 2 (1 - \Phi (1.96)) = 0.05$$

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ Suppose $c = \frac{1.96\sigma}{\sqrt{n}}$
- ▶ Then

$$P \left[\left| \frac{S_n}{n} - \mu \right| > \frac{1.96\sigma}{\sqrt{n}} \right] = 2 (1 - \Phi (1.96)) = 0.05$$

- ▶ Denoting $\bar{X} = \frac{S_n}{n}$, the 95% confidence interval is $\left[\bar{X} - \frac{1.96\sigma}{\sqrt{n}}, \bar{X} + \frac{1.96\sigma}{\sqrt{n}} \right]$

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ Suppose $c = \frac{1.96\sigma}{\sqrt{n}}$
- ▶ Then

$$P \left[\left| \frac{S_n}{n} - \mu \right| > \frac{1.96\sigma}{\sqrt{n}} \right] = 2 (1 - \Phi (1.96)) = 0.05$$

- ▶ Denoting $\bar{X} = \frac{S_n}{n}$, the 95% confidence interval is $\left[\bar{X} - \frac{1.96\sigma}{\sqrt{n}}, \bar{X} + \frac{1.96\sigma}{\sqrt{n}} \right]$
- ▶ One generally uses an estimate for σ obtained from X_i

$$P \left[\left| \frac{S_n}{n} - \mu \right| > c \right] = 2 \left(1 - \Phi \left(\frac{c\sqrt{n}}{\sigma} \right) \right)$$

- ▶ Suppose $c = \frac{1.96\sigma}{\sqrt{n}}$
- ▶ Then

$$P \left[\left| \frac{S_n}{n} - \mu \right| > \frac{1.96\sigma}{\sqrt{n}} \right] = 2 (1 - \Phi (1.96)) = 0.05$$

- ▶ Denoting $\bar{X} = \frac{S_n}{n}$, the 95% confidence interval is $\left[\bar{X} - \frac{1.96\sigma}{\sqrt{n}}, \bar{X} + \frac{1.96\sigma}{\sqrt{n}} \right]$
- ▶ One generally uses an estimate for σ obtained from X_i
- ▶ In analyzing any experimental data the confidence intervals or the variance term is important

central limit theorem

- ▶ CLT essentially states that sum of many independent random variables behaves like a Gaussian random variable

central limit theorem

- ▶ CLT essentially states that sum of many independent random variables behaves like a Gaussian random variable
- ▶ It is very useful in many statistics applications.

central limit theorem

- ▶ CLT essentially states that sum of many independent random variables behaves like a Gaussian random variable
- ▶ It is very useful in many statistics applications.
- ▶ We stated CLT for iid random variables.

central limit theorem

- ▶ CLT essentially states that sum of many independent random variables behaves like a Gaussian random variable
- ▶ It is very useful in many statistics applications.
- ▶ We stated CLT for iid random variables.
- ▶ While independence is important, all rv need not have the same distribution.

central limit theorem

- ▶ CLT essentially states that sum of many independent random variables behaves like a Gaussian random variable
- ▶ It is very useful in many statistics applications.
- ▶ We stated CLT for iid random variables.
- ▶ While independence is important, all rv need not have the same distribution.
- ▶ Essentially, the variances should not die out.

- ▶ We have been considering sequences: $X_n, n = 1, 2, \dots$

- ▶ We have been considering sequences: X_n , $n = 1, 2, \dots$
- ▶ We have so far considered only the asymptotic properties or limits of such sequences.

- ▶ We have been considering sequences: X_n , $n = 1, 2, \dots$
- ▶ We have so far considered only the asymptotic properties or limits of such sequences.
- ▶ Any such sequence is an example of what is called a random process or stochastic process

- ▶ We have been considering sequences: X_n , $n = 1, 2, \dots$
- ▶ We have so far considered only the asymptotic properties or limits of such sequences.
- ▶ Any such sequence is an example of what is called a random process or stochastic process
- ▶ Given n rv, they are completely characterized by their joint distribution.

- ▶ We have been considering sequences: X_n , $n = 1, 2, \dots$
- ▶ We have so far considered only the asymptotic properties or limits of such sequences.
- ▶ Any such sequence is an example of what is called a random process or stochastic process
- ▶ Given n rv, they are completely characterized by their joint distribution.
- ▶ How do we specify or characterize an infinite collection of random variables?

- ▶ We have been considering sequences: X_n , $n = 1, 2, \dots$
- ▶ We have so far considered only the asymptotic properties or limits of such sequences.
- ▶ Any such sequence is an example of what is called a random process or stochastic process
- ▶ Given n rv, they are completely characterized by their joint distribution.
- ▶ How do we specify or characterize an infinite collection of random variables?
- ▶ We need the joint distribution of every finite subcollection of them.

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

- ▶ We say it is a Markov chain if

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1} \cdots X_0 = x_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall$$

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

- ▶ We say it is a Markov chain if

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1} \cdots X_0 = x_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall$$

- ▶ We can write it as

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall x_i$$

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

- ▶ We say it is a Markov chain if

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1} \cdots X_0 = x_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall$$

- ▶ We can write it as

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall x_i$$

- ▶ Conditioned on X_n , X_{n+1} is independent of X_{n-1}, X_{n-2}, \dots

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

- ▶ We say it is a Markov chain if

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1} \cdots X_0 = x_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall$$

- ▶ We can write it as

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall x_i$$

- ▶ Conditioned on X_n , X_{n+1} is independent of X_{n-1}, X_{n-2}, \dots
- ▶ We think of X_n as state at n

Markov Chains

- ▶ Let X_n , $n = 0, 1, \dots$ be a sequence of discrete random variables taking values in S .

Note that S would be countable

- ▶ We say it is a Markov chain if

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1} \cdots X_0 = x_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall$$

- ▶ We can write it as

$$P[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = P[X_{n+1} = x_{n+1} | X_n = x_n], \forall x_i$$

- ▶ Conditioned on X_n , X_{n+1} is independent of X_{n-1}, X_{n-2}, \dots
- ▶ We think of X_n as state at n
- ▶ For a Markov chain, given the current state, the future evolution is independent of the history of how you reached the current state

Example

- ▶ Let X_i be iid discrete rv taking integer values.

Example

- ▶ Let X_i be iid discrete rv taking integer values.
- ▶ Let $Y_0 = 0$ and $Y_n = \sum_{i=1}^n X_i$

Example

- ▶ Let X_i be iid discrete rv taking integer values.
- ▶ Let $Y_0 = 0$ and $Y_n = \sum_{i=1}^n X_i$
- ▶ $Y_n, n = 0, 1, \dots$ is a Markov chain with state space as integers

Example

- ▶ Let X_i be iid discrete rv taking integer values.
- ▶ Let $Y_0 = 0$ and $Y_n = \sum_{i=1}^n X_i$
- ▶ $Y_n, n = 0, 1, \dots$ is a Markov chain with state space as integers
- ▶ Note that $Y_{n+1} = Y_n + X_{n+1}$ and X_{n+1} is independent of Y_0, \dots, Y_n .

Example

- ▶ Let X_i be iid discrete rv taking integer values.
- ▶ Let $Y_0 = 0$ and $Y_n = \sum_{i=1}^n X_i$
- ▶ $Y_n, n = 0, 1, \dots$ is a Markov chain with state space as integers
- ▶ Note that $Y_{n+1} = Y_n + X_{n+1}$ and X_{n+1} is independent of Y_0, \dots, Y_n .

$$P[Y_{n+1} = y | Y_n = x, Y_{n-1}, \dots] = P[X_{n+1} = y - x]$$

Example

- ▶ Let X_i be iid discrete rv taking integer values.
- ▶ Let $Y_0 = 0$ and $Y_n = \sum_{i=1}^n X_i$
- ▶ $Y_n, n = 0, 1, \dots$ is a Markov chain with state space as integers
- ▶ Note that $Y_{n+1} = Y_n + X_{n+1}$ and X_{n+1} is independent of Y_0, \dots, Y_n .

$$P[Y_{n+1} = y | Y_n = x, Y_{n-1}, \dots] = P[X_{n+1} = y - x]$$

- ▶ Thus, Y_{n+1} is conditionally independent of Y_{n-1}, \dots conditioned on Y_n

- ▶ In this example, we can think of X_n as the number of people or things arriving at a facility in the n^{th} time interval.

- ▶ In this example, we can think of X_n as the number of people or things arriving at a facility in the n^{th} time interval.
- ▶ Then Y_n would be total arrivals till end of n^{th} time interval.

- ▶ In this example, we can think of X_n as the number of people or things arriving at a facility in the n^{th} time interval.
- ▶ Then Y_n would be total arrivals till end of n^{th} time interval.
- ▶ Number of packets coming into a network switch, number people joining the queue in a bank, number of infections till date are all Markov chains.

- ▶ In this example, we can think of X_n as the number of people or things arriving at a facility in the n^{th} time interval.
- ▶ Then Y_n would be total arrivals till end of n^{th} time interval.
- ▶ Number of packets coming into a network switch, number people joining the queue in a bank, number of infections till date are all Markov chains.
- ▶ This is a useful model for many dynamic systems or processes

- ▶ The Markov property is: given current state, the future evolution is independent of the history of how we came to current state.

- ▶ The Markov property is: given current state, the future evolution is independent of the history of how we came to current state.
- ▶ It essentially means the current state contains all needed information about history

- ▶ The Markov property is: given current state, the future evolution is independent of the history of how we came to current state.
- ▶ It essentially means the current state contains all needed information about history
- ▶ We are considering the case where states as well as time are discrete.

- ▶ The Markov property is: given current state, the future evolution is independent of the history of how we came to current state.
- ▶ It essentially means the current state contains all needed information about history
- ▶ We are considering the case where states as well as time are discrete.
- ▶ It can be more general and we discuss some of them

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

- ▶ Define function $P : S \times S \rightarrow [0, 1]$ by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

- ▶ Define function $P : S \times S \rightarrow [0, 1]$ by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ P is called the state transition probability function. It satisfies

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

- ▶ Define function $P : S \times S \rightarrow [0, 1]$ by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ P is called the state transition probability function. It satisfies
 - ▶ $P(x, y) \geq 0, \forall x, y \in S$

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

- ▶ Define function $P : S \times S \rightarrow [0, 1]$ by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ P is called the state transition probability function. It satisfies
 - ▶ $P(x, y) \geq 0, \forall x, y \in S$
 - ▶ $\sum_{y \in S} P(x, y) = 1, \forall x \in S$

Transition Probabilities

- ▶ Let $\{X_n, n = 0, 1, \dots\}$ be a Markov Chain with (countable) state space S

$$Pr[X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} \cdots X_0] = Pr[X_{n+1} = x_{n+1} | X_n = x_n], \forall x$$

(Notice change of notation)

- ▶ Define function $P : S \times S \rightarrow [0, 1]$ by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ P is called the state transition probability function. It satisfies
 - ▶ $P(x, y) \geq 0, \forall x, y \in S$
 - ▶ $\sum_{y \in S} P(x, y) = 1, \forall x \in S$
- ▶ If S is finite then P can be represented as a matrix

- ▶ The state transition probability function is given by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ The state transition probability function is given by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ In general, this can depend on n though our notation does not show it

- ▶ The state transition probability function is given by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ In general, this can depend on n though our notation does not show it
- ▶ If the value of that probability does not depend on n then the chain is called homogeneous

- ▶ The state transition probability function is given by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ In general, this can depend on n though our notation does not show it
- ▶ If the value of that probability does not depend on n then the chain is called homogeneous
- ▶ For a homogeneous chain we have

$$Pr[X_{n+1} = y | X_n = x] = Pr[X_1 = y | X_0 = x], \forall n$$

- ▶ The state transition probability function is given by

$$P(x, y) = Pr[X_{n+1} = y | X_n = x]$$

- ▶ In general, this can depend on n though our notation does not show it
- ▶ If the value of that probability does not depend on n then the chain is called homogeneous
- ▶ For a homogeneous chain we have

$$Pr[X_{n+1} = y | X_n = x] = Pr[X_1 = y | X_0 = x], \forall n$$

- ▶ In this course we will consider only homogeneous chains

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space \mathcal{S}

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

- ▶ It is the pmf of the rv X_0

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

- ▶ It is the pmf of the rv X_0
- ▶ Hence it satisfies

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

- ▶ It is the pmf of the rv X_0
- ▶ Hence it satisfies
 - ▶ $\pi_0(x) \geq 0, \forall x \in S$

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

- ▶ It is the pmf of the rv X_0
- ▶ Hence it satisfies
 - ▶ $\pi_0(x) \geq 0, \forall x \in S$
 - ▶ $\sum_{x \in S} \pi_0(x) = 1$

Initial State Probabilities

- ▶ Let $\{X_n\}$ be a Markov Chain with state space S
- ▶ Define function $\pi_0 : S \rightarrow [0, 1]$ by

$$\pi_0(x) = Pr[X_0 = x]$$

- ▶ It is the pmf of the rv X_0
- ▶ Hence it satisfies
 - ▶ $\pi_0(x) \geq 0, \forall x \in S$
 - ▶ $\sum_{x \in S} \pi_0(x) = 1$
- ▶ From now on, without loss of generality, we take $S = \{0, 1, \dots\}$

- ▶ Let X_n be a (homogeneous) Markov chain

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$Pr[X_0 = x_0, X_1 = x_1] = Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) \end{aligned}$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Now we can extend this as

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Now we can extend this as

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1, X_2 = x_2] &= Pr[X_2 = x_2 | X_1 = x_1, X_0 = x_0] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \end{aligned}$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Now we can extend this as

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1, X_2 = x_2] &= Pr[X_2 = x_2 | X_1 = x_1, X_0 = x_0] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \\ &= Pr[X_2 = x_2 | X_1 = x_1] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \end{aligned}$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Now we can extend this as

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1, X_2 = x_2] &= Pr[X_2 = x_2 | X_1 = x_1, X_0 = x_0] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \\ &= Pr[X_2 = x_2 | X_1 = x_1] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \\ &= P(x_1, x_2) P(x_0, x_1) \pi_0(x_0) \end{aligned}$$

- ▶ Let X_n be a (homogeneous) Markov chain
- ▶ Then we have

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1] &= Pr[X_1 = x_1 | X_0 = x_0] Pr[X_0 = x_0], \forall x_0, x_1 \\ &= P(x_0, x_1) \pi_0(x_0) = \pi_0(x_0) P(x_0, x_1) \end{aligned}$$

- ▶ Now we can extend this as

$$\begin{aligned} Pr[X_0 = x_0, X_1 = x_1, X_2 = x_2] &= Pr[X_2 = x_2 | X_1 = x_1, X_0 = x_0] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \\ &= Pr[X_2 = x_2 | X_1 = x_1] \cdot \\ &\quad Pr[X_0 = x_0, X_1 = x_1] \\ &= P(x_1, x_2) P(x_0, x_1) \pi_0(x_0) \\ &= \pi_0(x_0) P(x_0, x_1) P(x_1, x_2) \end{aligned}$$

- ▶ This calculation is easily generalized to any number of time steps

- ▶ This calculation is easily generalized to any number of time steps

$$\Pr[X_0 = x_0, \dots, X_n = x_n] = \Pr[X_n = x_n | X_{n-1} = x_{n-1}, \dots, X_0 = x_0] \cdot \Pr[X_{n-1} = x_{n-1}, \dots, X_0 = x_0]$$

- This calculation is easily generalized to any number of time steps

$$\begin{aligned} Pr[X_0 = x_0, \dots X_n = x_n] &= Pr[X_n = x_n | X_{n-1} = x_{n-1}, \dots X_0 = x_0] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \dots X_0 = x_0] \\ &= Pr[X_n = x_n | X_{n-1} = x_{n-1}] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \dots X_0 = x_0] \end{aligned}$$

- This calculation is easily generalized to any number of time steps

$$\begin{aligned} Pr[X_0 = x_0, \cdots X_n = x_n] &= Pr[X_n = x_n | X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= Pr[X_n = x_n | X_{n-1} = x_{n-1}] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= P(x_{n-1}, x_n) Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \end{aligned}$$

- This calculation is easily generalized to any number of time steps

$$\begin{aligned} Pr[X_0 = x_0, \cdots X_n = x_n] &= Pr[X_n = x_n | X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= Pr[X_n = x_n | X_{n-1} = x_{n-1}] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= P(x_{n-1}, x_n) Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= P(x_{n-1}, x_n) Pr[X_{n-1} = x_{n-1} | X_{n-2} = x_{n-2}] \cdot \\ &\quad Pr[X_{n-2} = x_{n-2}, \cdots X_0 = x_0] \end{aligned}$$

- This calculation is easily generalized to any number of time steps

$$\begin{aligned} Pr[X_0 = x_0, \cdots X_n = x_n] &= Pr[X_n = x_n | X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= Pr[X_n = x_n | X_{n-1} = x_{n-1}] \cdot \\ &\quad Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= P(x_{n-1}, x_n) Pr[X_{n-1} = x_{n-1}, \cdots X_0 = x_0] \\ &= P(x_{n-1}, x_n) Pr[X_{n-1} = x_{n-1} | X_{n-2} = x_{n-2}] \cdot \\ &\quad Pr[X_{n-2} = x_{n-2}, \cdots X_0 = x_0] \\ &\quad \vdots \\ &= \pi_0(x_0) P(x_0, x_1) \cdots P(x_{n-1}, x_n) \end{aligned}$$

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ We showed

$$\Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.
- ▶ They give us the joint distribution of any finite subcollection of the rv's

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.
- ▶ They give us the joint distribution of any finite subcollection of the rv's
- ▶ Suppose you want joint distribution of X_{i_1}, \dots, X_{i_k}

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.
- ▶ They give us the joint distribution of any finite subcollection of the rv's
- ▶ Suppose you want joint distribution of X_{i_1}, \dots, X_{i_k}
- ▶ Let $m = \max(i_1, \dots, i_k)$

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.
- ▶ They give us the joint distribution of any finite subcollection of the rv's
- ▶ Suppose you want joint distribution of X_{i_1}, \dots, X_{i_k}
- ▶ Let $m = \max(i_1, \dots, i_k)$
- ▶ We know how to get joint distribution of X_0, \dots, X_m .

- ▶ We showed

$$Pr[X_0 = x_0, \dots, X_n = x_n] = \pi_0(x_0)P(x_0, x_1) \cdots P(x_{n-1}, x_n)$$

- ▶ This shows that the transition probabilities, P , and initial state probabilities, π_0 , completely specify the chain.
- ▶ They give us the joint distribution of any finite subcollection of the rv's
- ▶ Suppose you want joint distribution of X_{i_1}, \dots, X_{i_k}
- ▶ Let $m = \max(i_1, \dots, i_k)$
- ▶ We know how to get joint distribution of X_0, \dots, X_m .
- ▶ The joint distribution of X_{i_1}, \dots, X_{i_k} is now calculated as a marginal distribution from the joint distribution of X_0, \dots, X_m