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We can write it as

$$f_{X_{n+1}|X_n,\dots X_0}(x_{n+1}|x_n,\dots,x_0) = f_{X_{n+1}|X_n}(x_{n+1}|x_n), \ \forall x_i$$

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► For a Markov chain, given the current state, the future evolution is independent of the history of how you reached the current state

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- ightharpoonup Similarly, $\pi_n(x) = Pr[X_n = x]$

Recap: Chapman-Kolmogorov Equations

▶ *n*-step transition probabilities:

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► For a finite chain, the *n*-step transition probability matrix is *n*-fold product of the transition probability matrix

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$$P_x[N(y)=\infty]=\rho_{xy}, \quad \text{and} \quad G(x,y)=\left\{ \begin{array}{ll} 0 & \text{if} \quad \rho_{xy}=0 \\ \infty & \text{if} \quad \rho_{xy}>0 \end{array} \right.$$

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- ▶ We say, x leads to y if $\rho_{xy} > 0$ Theorem: If x is recurrent and x leads to y then y is recurrent and $\rho_{xy} = \rho_{yx} = 1$.

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- ▶ A set of states, $C \subset S$, is said to be closed if $x \in C$, $y \notin C$ implies $\rho_{xy} = 0$.
- Once the chain visits a state in a closed set, it cannot leave that set.

Recap: Partition of state space

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ightharpoonup We can calculate absorption probabilities for C_i using

$$\rho_C(x) = \sum_{y \in C} P(x, y) + \sum_{y \in S_T} P(x, y) \ \rho_C(y)$$

$$\pi(y) = \sum_{x \in S} \pi(x) P(x, y), \ \forall y \in S$$

lacktriangledown is said to be a stationary distribution for the Markov chain with transition probabilities P if

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Recap: Stationary distribution

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- ▶ If $\pi_n = \pi$, $\forall n$ then π is a stationary distribution
- For a finite chain, a stationary distribution always exists.
- ► The stationary distribution, when it exists, is related to expected fraction of time spent in different states.

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 \blacktriangleright We will first establish a limit for the above as $n \to \infty$

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- ► The expected fraction of time spent in a transient state is zero.
- ► This is intuitively obvious

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- ► Convergence would be with probability one.

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- ► We will prove this.
- ▶ Then T_u^r/r converges to m_y by law of large umbers

We have

$$Pr[W_y^3 = k_3 | W_y^2 = k_2, W_y^1 = k_1] =$$

$$Pr[W_y^3 = k_3 | W_y^2 = k_2, W_y^1 = k_1] = \\ Pr[X_{k_1+k_2+j} \neq y, \ 1 \leq j \leq k_3 - 1, \ X_{k_1+k_2+k_3} = y \mid B] \\ \text{where } B = [X_{k_1+k_2} = y, \ X_{k_1} = y, \ X_j \neq y, j < k_1 + k_2, j \neq k_1]$$

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▶ In general, we get

$$Pr[W_y^r = k_r \mid W_y^{r-1} = k_{r-1}, \cdots, W_y^1 = k_1] = P_y[W_y^1 = k_r]$$

$$P_y[W_y^2 = k_2] = \sum_{k} P_y[W_y^2 = k_2 \mid W_y^1 = k_1] P_y[W_y^1 = k_1]$$

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$$\begin{split} P_y[W_y^2 = k_2] &= \sum_{k_1} P_y[W_y^2 = k_2 \mid W_y^1 = k_1] \; P_y[W_y^1 = k_1] \\ &= \sum_{k_1} P_y[W_y^1 = k_2] \; P_y[W_y^1 = k_1] \\ &= P_y[W_y^1 = k_2] \end{split}$$

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⇒ identically distributed

$$P_{y}[W_{y}^{2} = k_{2}, W_{y}^{1} = k_{1}] = P_{y}[W_{y}^{2} = k_{2} | W_{y}^{1} = k_{1}]P_{y}[W_{y}^{1} = k_{1}]$$

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⇒ independent

▶ We have shown W_y^r , $r = 1, 2, \cdots$ are iid

- ▶ We have shown W_u^r , $r = 1, 2, \cdots$ are iid
- ▶ Since $E[W_u^1] = m_y$, by strong law of large numbers,

$$\lim_{k \to \infty} \frac{T_y^k}{k} = \lim_{k \to \infty} \frac{1}{k} \sum_{r=1}^k W_y^r = m_y, \quad (w.p.1)$$

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▶ Note that this is true even if $m_y = \infty$

$$T_y^{N_n(y)} \le n < T_y^{N_n(y)+1}$$

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- ▶ So, time of 9^{th} visit is beyond 50.

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 - ightharpoonup As $n \to \infty$, $N_n(y) \to \infty$, w.p.1
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- ► Hence we get

$$\lim_{n \to \infty} \frac{n}{N_n(y)} = m_y, \quad w.p.1$$

or

$$\lim_{n \to \infty} \frac{N_n(y)}{n} = \frac{1}{m_n}, \quad w.p.1$$

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$$\lim_{n \to \infty} \frac{G_n(x, y)}{n} = \frac{\rho_{xy}}{m_y}$$

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- ▶ What if $m_y = \infty$, $\forall y$?
- ▶ Does not seem reasonable for a finite chain.
- ▶ But for infinite chains??
- ▶ Let us characterize y for which $m_y = \infty$

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► Thus the limiting fraction of time spent by the chain in transient and null recurrent states is zero.

► **Theorem:** Let *x* be positive recurrent and let *x* lead to *y*. Then *y* is positive recurrent.

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Proof

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- ▶ Since x is recurrent and x leads to y we know $\exists n_0, n_1$ s.t. $P^{n_0}(x,y)>0$, $P^{n_1}(y,x)>0$ and

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$$P^{n_1+m+n_0}(y,y) \ge P^{n_1}(y,x)P^m(x,x)P^{n_0}(x,y), \ \forall m$$

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Summing the above for $m=1,2,\cdots n$ and dividing by n

$$\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) \ge P^{n_1}(y,x) \quad \frac{1}{n} \sum_{m=1}^{n} P^m(x,x) \quad P^{n_0}(x,y), \ \forall n$$

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If we now let $n \to \infty$, the RHS goes to $P^{n_1}(y,x) \stackrel{1}{\xrightarrow{m}} P^{n_0}(x,y) > 0$.

$$\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) \ge P^{n_1}(y,x) \frac{1}{n} \sum_{m=1}^{n} P^m(x,x) P^{n_0}(x,y), \forall n$$

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▶ We can write the *LHS* of above as

$$\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) \ge P^{n_1}(y,x) \frac{1}{n} \sum_{m=1}^{n} P^m(x,x) P^{n_0}(x,y), \forall n$$

$$\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) = \frac{1}{n} \sum_{m=1}^{n_1+n+n_0} P^m(y,y) - \frac{1}{n} \sum_{m=1}^{n_1+n_0} P^m(y,y)$$

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$$\begin{split} \frac{1}{n} \sum_{m=1}^{n} P^{n_1 + m + n_0}(y, y) &= \frac{1}{n} \sum_{m=1}^{n_1 + n + n_0} P^m(y, y) - \frac{1}{n} \sum_{m=1}^{n_1 + n_0} P^m(y, y) \\ &= \frac{n_1 + n + n_0}{n} \frac{1}{n_1 + n + n_0} \sum_{m=1}^{n_1 + n + n_0} P^m(y, y) - \frac{1}{n} \sum_{m=1}^{n_1 + n_0} P^m(y, y) \end{split}$$

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$$\Rightarrow \lim_{n \to \infty} \frac{1}{n} \sum_{m=1}^{\infty} P^{n_1 + m + n_0}(y, y) = \frac{1}{m_y}$$

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$$= \frac{n_1+n+n_0}{n} \frac{1}{n_1+n+n_0} \sum_{m=1}^{n_1+n+n_0} P^m(y,y) - \frac{1}{n} \sum_{m=1}^{n_1+n_0} P^m(y,y)$$

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$$\Rightarrow \frac{1}{m_y} \ge P^{n_1}(y, x) \frac{1}{m_x} P^{n_0}(x, y) > 0$$

$$\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) \ge P^{n_1}(y,x) \quad \frac{1}{n} \sum_{m=1}^{n} P^m(x,x) \quad P^{n_0}(x,y), \ \forall n$$

▶ We can write the *LHS* of above as

$$1^{n_1+n+n_0}$$

 $\frac{1}{n} \sum_{m=1}^{n} P^{n_1+m+n_0}(y,y) = \frac{1}{n} \sum_{m=1}^{n_1+n+n_0} P^m(y,y) - \frac{1}{n} \sum_{m=1}^{n_1+n_0} P^m(y,y)$

 $=\frac{n_1+n+n_0}{n}\frac{1}{n_1+n+n_0}\sum_{m=1}^{n_1+n+n_0}P^m(y,y)-\frac{1}{n}\sum_{m=1}^{n_1+n_0}P^m(y,y)$

$$\Rightarrow \lim_{n \to \infty} \frac{1}{n} \sum_{m=1}^{n} P^{n_1 + m + n_0}(y, y) = \frac{1}{m_y}$$

$$\Rightarrow \frac{1}{m_y} \geq P^{n_1}(y,x) \, \frac{1}{m_x} \, P^{n_0}(x,y) > 0$$
 which implies y is positive recurrent

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- We next show that a finite chain cannot have any null recurrent states.

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where we could take the limit inside the sum because ${\cal C}$ is finite.

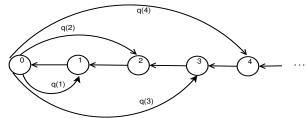
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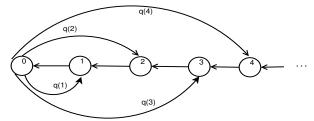
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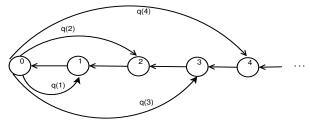
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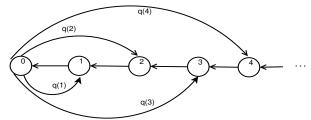


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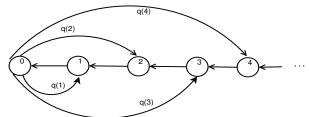
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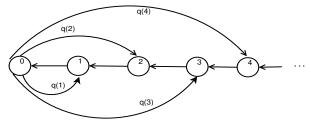
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 (Note that $P_0[T_0=1]=0$)

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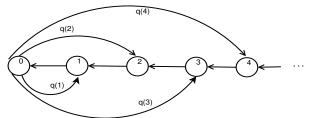
▶ Here, $q(k) \ge 0, \forall k \text{ and } \sum_{k=1}^{\infty} q(k) = 1$. We have

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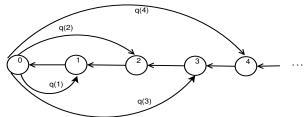
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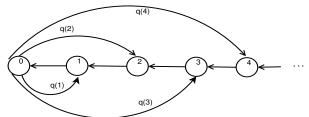
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► The proof is complete if we can take the limit inside the sum

► Bounded Convergence Theorem: Suppose

 $a(x) \geq 0, \ \forall x \in S \ \text{and} \ \sum_{x} a(x) < \infty. \ \text{Let} \ b_n(x), \ x \in S$ be such that $|b_n(x)| \leq K, \ \forall x, n \ \text{and suppose}$ $\lim_{n \to \infty} b_n(x) = b(x), \forall x \in S. \ \text{Then}$

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- ► An irreducible finite chain has a unique stationary distribution

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- ► This answers all questions about existence and uniqueness of stationary distributions

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 - For example, $a_n = (-1)^n$

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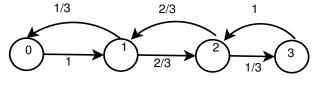
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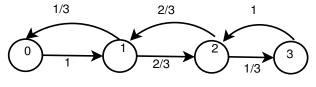
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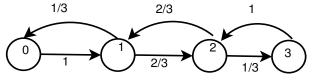
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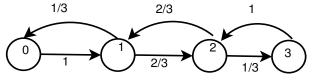
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- ► Such a chain is called a periodic chain

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- ▶ If the period is 1 then chain is called aperiodic

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- ▶ For an aperiodic, irreducible, positive recurrent chain, there is a unique stationary distribution and π_n converges to it irrespective of what π_0 is.
- ► An aperiodic, irreducible, positive recurrent chain is called an ergodic chain