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- ► All stationary distributions of the chain are convex combinations of these

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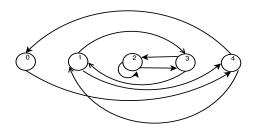
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- An irreducible chain is called aperiodic if the period is 1
- For an irreducible aperiodic positive recurrent chain, π_n converges to π , the unique stationary distribution, irrespective of what π_0 is.
- ▶ Also, for an irreducible, aperiodic, positive recurrent chain, $P^n(x,y)$ converges to $\frac{1}{m_n}$

Consider the umbrella problem

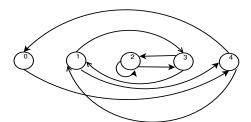
$$P = \begin{bmatrix} \begin{array}{c|ccccc} & 0 & 1 & 2 & 3 & 4 \\ \hline 0 & 0 & 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 & 1 - p & p \\ 2 & 0 & 0 & 1 - p & p & 0 \\ 3 & 0 & 1 - p & p & 0 & 0 \\ 4 & 1 - p & p & 0 & 0 & 0 \end{array} \end{bmatrix}$$

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This is an irreducible, aperiodic positive recurrent chain

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- We are using the fact that this chain converges to the stationary distribution starting with any initial probabilities.

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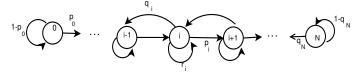
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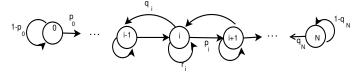
$$\pi(4) = \pi(0) + p\pi(1) \Rightarrow \pi(4) = \pi(1)$$

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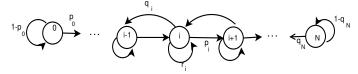
$$\pi(i) = \frac{1}{5-p} \ i = 1,2,3,4 \quad \text{and} \quad \pi(0) = \frac{1-p}{5-p}$$
 PS Sastry, IISc, Bangalore, 2020 8/36



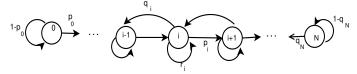
▶ The following is a finite birth-death chain



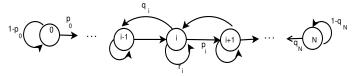
• We assume $p_i, q_i > 0, \forall i$.



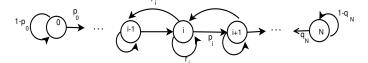
- We assume $p_i, q_i > 0, \forall i$.
- ▶ Then the chain is irreducible, positive recurrent

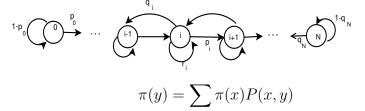


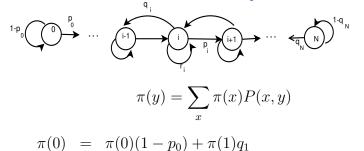
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- ▶ If we assume $r_i > 0$ at least for one i, it is also aperiodic

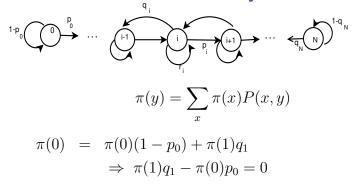


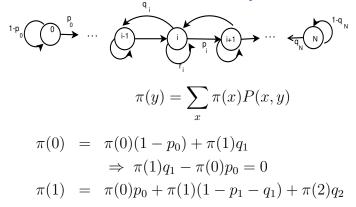
- We assume $p_i, q_i > 0, \forall i$.
- ▶ Then the chain is irreducible, positive recurrent
- ▶ If we assume $r_i > 0$ at least for one i, it is also aperiodic
- We can derive a general form for its stationary probabilities

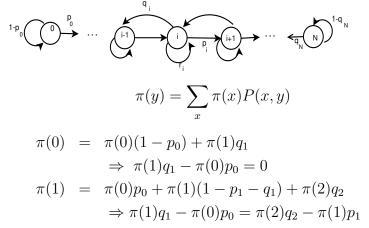


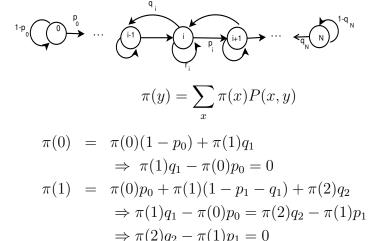


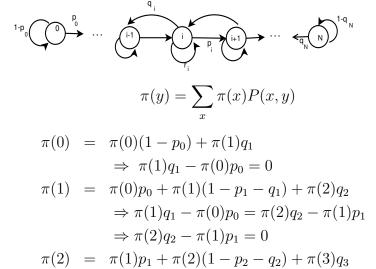


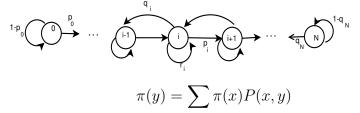












$$\pi(0) = \pi(0)(1 - p_0) + \pi(1)q_1$$

$$\Rightarrow \pi(1)q_1 - \pi(0)p_0 = 0$$

$$\pi(1) = \pi(0)p_0 + \pi(1)(1 - p_1 - q_1) + \pi(2)q_2$$

$$\Rightarrow \pi(1)q_1 - \pi(0)p_0 = \pi(2)q_2 - \pi(1)p_1$$

$$\Rightarrow \pi(2)q_2 - \pi(1)p_1 = 0$$

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▶ Iterating like this, we get

$$\pi(n) = \eta_n \; \pi(0), \; \text{ where } \; \eta_n = \frac{p_0 p_1 \cdots p_{n-1}}{q_1 q_2 \cdots q_n}, \; n = 1, 2, \cdots, N$$

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▶ With $\eta_0 = 1$, we get $\pi(0) \sum_{j=0}^N \eta_j = 1$ and hence

$$\pi(0) = \frac{1}{\sum_{i=0}^{N} \eta_i}$$
 and $\pi(n) = \eta_n \, \pi(0), \ n = 1, \cdots, N$

$$\pi(1)q_1 - \pi(0)p_0 = 0 \Rightarrow \pi(1) = \frac{p_0}{q_1} \pi(0)$$

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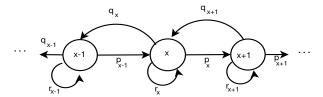
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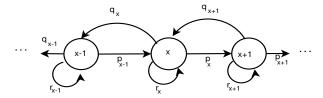
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Note that this process is applicable even for infinite chains with state space $\{0, 1, 2, \cdots\}$ (but there may not be a solution)

► Consider a birth-death chain

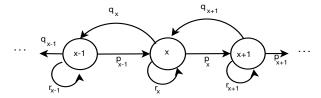


Consider a birth-death chain



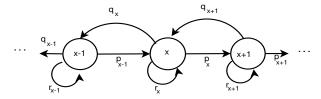
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Consider a birth-death chain



- ▶ The chain may be infinite or finite
- ▶ Let $a, b \in S$ with a < b. Assume $p_x, q_x > 0$, a < x < b.

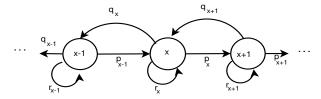
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- ▶ The chain may be infinite or finite
- ▶ Let $a, b \in S$ with a < b. Assume $p_x, q_x > 0$, a < x < b.
- Define

$$U(x) = P_x[T_a < T_b], \ a < x < b, \ U(a) = 1, \ U(b) = 0$$

Consider a birth-death chain

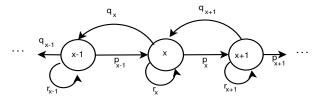


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$$U(x) = P_x[T_a < T_b], \ a < x < b, \ U(a) = 1, \ U(b) = 0$$

• We want to derive a formula for U(x)

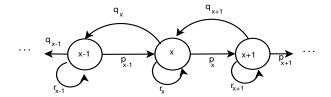
Consider a birth-death chain

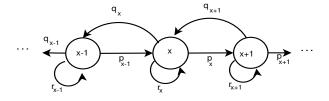


- The chain may be infinite or finite
- ▶ Let $a, b \in S$ with a < b. Assume $p_x, q_x > 0$, a < x < b.
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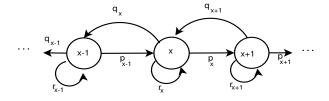
$$U(x) = P_x[T_a < T_b], \ a < x < b, \ U(a) = 1, \ U(b) = 0$$

- We want to derive a formula for U(x)
- ▶ This can be useful, e.g., in the gambler's ruin chain



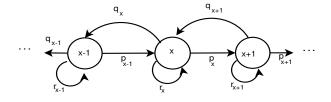


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$$= U(x-1)q_x + U(x)r_x + U(x+1)p_x$$

$$\cdots \qquad q_{x-1} \qquad x-1 \qquad p_{x-1} \qquad x \qquad p_{x} \qquad x+1 \qquad p_{x+1} \qquad \cdots$$

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$$\Rightarrow$$
 $q_x[U(x) - U(x-1)] = p_x[U(x+1) - U(x)]$

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Let
$$\gamma_y = \frac{q_y q_{y-1} \cdots q_{a+1}}{p_y p_{y-1} \cdots p_{a+1}}, \ a < y < b, \ \gamma_a = 1$$

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Now we get

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$$\Rightarrow U(a)-U(a+1)=\frac{\gamma_a}{\sum_{x=a}^{b-1}\gamma_x}$$

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▶ Putting $x = b - 1, b - 2, \dots, y$ in the above

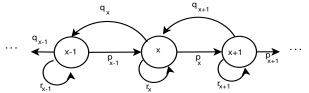
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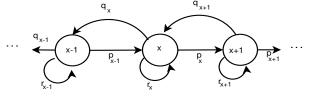
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Adding these we get $U(y) - U(b) = U(y) = \frac{\sum_{x=y}^{b-1} \gamma_x}{\sum_{x=a}^{b-1} \gamma_x}, \ a < y < b$ ▶ We are considering birth-death chains



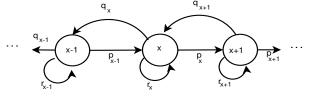
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• We have derived, for a < y < b,

$$U(y) = P_y[T_a < T_b] = \frac{\sum_{x=y}^{b-1} \gamma_x}{\sum_{x=a}^{b-1} \gamma_x}, \quad \gamma_x = \frac{q_x q_{x-1} \cdots q_{a+1}}{p_x p_{x-1} \cdots p_{a+1}}$$

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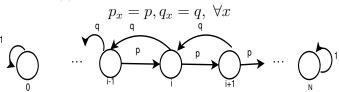
Hence we also get

$$P_y[T_b < T_a] = \frac{\sum_{x=a}^{y-1} \gamma_x}{\sum_{x=a}^{b-1} \gamma_x}$$

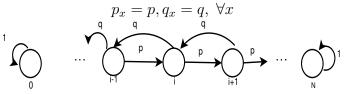
► Suppose this is a Gambler's ruin chain:

$$p_x = p, q_x = q, \ \forall x$$

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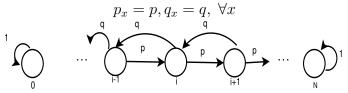


► Suppose this is a Gambler's ruin chain:



▶ Then, $\gamma_x = \left(\frac{q}{p}\right)^x$

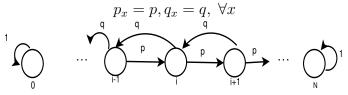
► Suppose this is a Gambler's ruin chain:



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- ► Hence, for a Gambler's ruin chain we get, e.g.,

$$P_i[T_N < T_0] = \frac{\sum_{x=0}^{i-1} \gamma_x}{\sum_{x=0}^{N-1} \gamma_x} = \frac{\left(\frac{q}{p}\right)^i - 1}{\left(\frac{q}{p}\right)^N - 1}$$

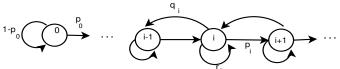
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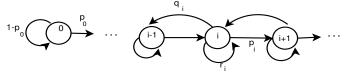


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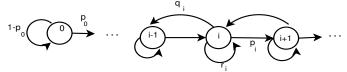
$$P_i[T_N < T_0] = \frac{\sum_{x=0}^{i-1} \gamma_x}{\sum_{x=0}^{N-1} \gamma_x} = \frac{\left(\frac{q}{p}\right)^i - 1}{\left(\frac{q}{p}\right)^N - 1}$$

▶ This is the probability of gambler being successful

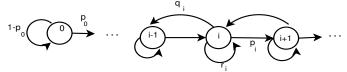




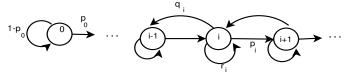
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- We want to know whether the chain is transient or recurrent etc.

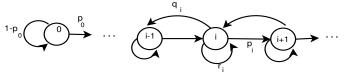


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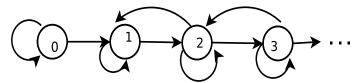
$$P_1[T_0 < T_n] = \frac{\sum_{x=1}^{n-1} \gamma_x}{\sum_{x=0}^{n-1} \gamma_x}, \ \forall n > 1$$

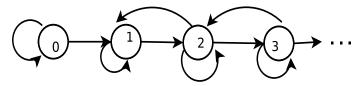


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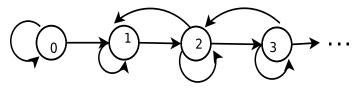
$$P_{1}[T_{0} < T_{n}] = \frac{\sum_{x=1}^{n-1} \gamma_{x}}{\sum_{x=0}^{n-1} \gamma_{x}}, \forall n > 1$$

$$= \frac{\sum_{x=0}^{n-1} \gamma_{x} - \gamma_{0}}{\sum_{x=0}^{n-1} \gamma_{x}} = 1 - \frac{1}{\sum_{x=0}^{n-1} \gamma_{x}}$$

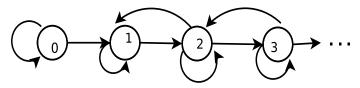




$$[T_0 < T_n] \subset [T_0 < T_{n+1}], \quad n = 2, 3, \cdots$$



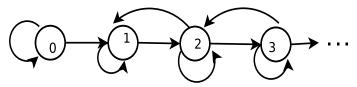
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since the chain cannot hit n+1 without hitting n.

▶ Also, $1 \le T_2 < T_3 < \cdots < T_n$ and $T_n \ge n$.



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$$P_1[T_0 < T_n, \ \text{ for some } n] \ = \ P_1\left(\cup_{n>1} \left[T_0 < T_n\right]\right)$$

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$$[T_0 < T_n] \subset [T_0 < T_{n+1}], \quad n = 2, 3, \cdots$$

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- ▶ Note that we have used the fact that the chain is infinite only to the right.

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- ▶ This chain with q < p is an example of an irreducible chain that is wholly transient

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- The earlier equations that we derived earlier hold for this infinite case also.

$$\pi(n) = \eta_n \; \pi(0), \; \text{ where } \; \eta_n = \frac{p_0 p_1 \cdots p_{n-1}}{q_1 q_2 \cdots q_n}, \; n = 1, 2, \cdots,$$

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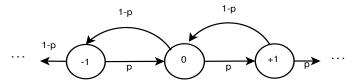
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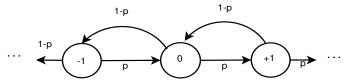
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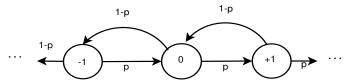


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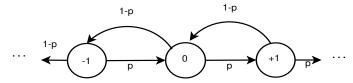
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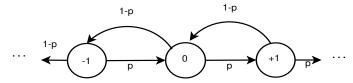
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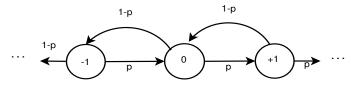
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Let P be the transition probabilities of a homogeneous irreducible Markov chain with state space S. Let

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ightharpoonup For this to be true for infinite S, we need some extra conditions

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- ightharpoonup One way to generate samples is to design an ergodic markov chain with stationary distribution π
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- lacktriangle We can also use the chain to generate samples from distribution π

• $\{X_n\}$: Markov chain with stationary dist $\pi(x) = \frac{b(x)}{Z}$

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▶ For all these, we need to design a Markov chain with π as stationary distribution

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- We want to choose $\alpha(i,j)$ to make X_n an ergodic markov chain with stationary probabilities π

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