

Continuous-Time Markov Chains

- ▶ Let $\{X(t), t \geq 0\}$ be a continuous-time discrete-state process

Continuous-Time Markov Chains

- ▶ Let $\{X(t), t \geq 0\}$ be a continuous-time discrete-state process
- ▶ Let $X(t)$ take non-negative integer values

Continuous-Time Markov Chains

- ▶ Let $\{X(t), t \geq 0\}$ be a continuous-time discrete-state process
- ▶ Let $X(t)$ take non-negative integer values
- ▶ It is called a continuous-time markov chain if

$$\begin{aligned} Pr[X(t+s) = j \mid X(s) = i, X(u) \in A_u, 0 \leq u < s] \\ = Pr[X(t+s) = j \mid X(s) = i] \end{aligned}$$

Continuous-Time Markov Chains

- ▶ Let $\{X(t), t \geq 0\}$ be a continuous-time discrete-state process
- ▶ Let $X(t)$ take non-negative integer values
- ▶ It is called a continuous-time markov chain if

$$\begin{aligned} Pr[X(t+s) = j \mid X(s) = i, X(u) \in A_u, 0 \leq u < s] \\ = Pr[X(t+s) = j \mid X(s) = i] \end{aligned}$$

- ▶ Only most recent past matters

Continuous-Time Markov Chains

- ▶ Let $\{X(t), t \geq 0\}$ be a continuous-time discrete-state process
- ▶ Let $X(t)$ take non-negative integer values
- ▶ It is called a continuous-time markov chain if

$$\begin{aligned} Pr[X(t+s) = j \mid X(s) = i, X(u) \in A_u, 0 \leq u < s] \\ = Pr[X(t+s) = j \mid X(s) = i] \end{aligned}$$

- ▶ Only most recent past matters
- ▶ It is called homogeneous chain if

$$Pr[X(t+s) = j \mid X(s) = i] = Pr[X(t) = j \mid X(0) = i], \forall s$$

► Define

$$P_{ij}(t) = \Pr[X(t) = j \mid X(0) = i]$$

► Define

$$P_{ij}(t) = Pr[X(t) = j \mid X(0) = i] = Pr[X(t+s) = j \mid X(s) = i]$$

► Define

$$P_{ij}(t) = Pr[X(t) = j \mid X(0) = i] = Pr[X(t+s) = j \mid X(s) = i]$$

It is the probability of going from i to j in time t

- Define

$$P_{ij}(t) = Pr[X(t) = j \mid X(0) = i] = Pr[X(t+s) = j \mid X(s) = i]$$

It is the probability of going from i to j in time t

- Analogous to transition probabilities in the discrete case

► Define

$$P_{ij}(t) = Pr[X(t) = j \mid X(0) = i] = Pr[X(t+s) = j \mid X(s) = i]$$

It is the probability of going from i to j in time t

- Analogous to transition probabilities in the discrete case
- Like in the discrete case, we can show that the Markov condition implies

$$\begin{aligned} Pr[X(s) \in B_s, s \in (t, t + \tau] \mid X(t) = i, X(s'), 0 \leq s' < t] \\ = Pr[X(s) \in B_s, s \in (t, t + \tau] \mid X(t) = i] \end{aligned}$$

► Define

$$P_{ij}(t) = \Pr[X(t) = j \mid X(0) = i] = \Pr[X(t+s) = j \mid X(s) = i]$$

It is the probability of going from i to j in time t

- Analogous to transition probabilities in the discrete case
- Like in the discrete case, we can show that the Markov condition implies

$$\begin{aligned} \Pr[X(s) \in B_s, s \in (t, t + \tau] \mid X(t) = i, X(s'), 0 \leq s' < t] \\ = \Pr[X(s) \in B_s, s \in (t, t + \tau] \mid X(t) = i] \end{aligned}$$

- Next we consider distribution of time spent in a state before leaving it

- ▶ By the Markov property and homogeneity we have

- By the Markov property and homogeneity we have

$$Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t]$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \end{aligned}$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- Let $X(0) = i$ and let T_i be time spent in i before leaving it for the first time

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- Let $X(0) = i$ and let T_i be time spent in i before leaving it for the first time

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[T_i > t + \tau \mid T_i > t] \end{aligned}$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- Let $X(0) = i$ and let T_i be time spent in i before leaving it for the first time

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[T_i > t + \tau \mid T_i > t] \end{aligned}$$

$$Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] = Pr[T_i > \tau]$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- Let $X(0) = i$ and let T_i be time spent in i before leaving it for the first time

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[T_i > t + \tau \mid T_i > t] \\ Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] = Pr[T_i > \tau] \\ \Rightarrow Pr[T_i > t + \tau \mid T_i > t] = Pr[T_i > \tau] \end{aligned}$$

- By the Markov property and homogeneity we have

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[X(s) = i, s \in [t, t + \tau] \mid X(t) = i] \\ = Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] \end{aligned}$$

- Let $X(0) = i$ and let T_i be time spent in i before leaving it for the first time

$$\begin{aligned} Pr[X(s) = i, s \in [t, t + \tau] \mid X(s') = i, 0 \leq s' \leq t] \\ = Pr[T_i > t + \tau \mid T_i > t] \end{aligned}$$

$$\begin{aligned} Pr[X(s) = i, s \in [0, \tau] \mid X(0) = i] &= Pr[T_i > \tau] \\ \Rightarrow Pr[T_i > t + \tau \mid T_i > t] &= Pr[T_i > \tau] \end{aligned}$$

$\Rightarrow T_i$ is memoryless and hence exponential

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.
- ▶ So, the chain can be viewed as follows

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.
- ▶ So, the chain can be viewed as follows
- ▶ Once you transit to a state, it spends time, say, $T_i \sim \text{exponential}(\nu_i)$ in it.

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.
- ▶ So, the chain can be viewed as follows
- ▶ Once you transit to a state, it spends time, say, $T_i \sim \text{exponential}(\nu_i)$ in it.
- ▶ Then, when it leaves i , it transits to state j with probability, say, z_{ij}

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.
- ▶ So, the chain can be viewed as follows
- ▶ Once you transit to a state, it spends time, say, $T_i \sim \text{exponential}(\nu_i)$ in it.
- ▶ Then, when it leaves i , it transits to state j with probability, say, z_{ij}
- ▶ We would have $z_{ij} \geq 0$, $\sum_j z_{ij} = 1$. Also, $z_{ii} = 0$

- ▶ Once you transit into a state, the time spent in it is exponentially distributed.
- ▶ So, the chain can be viewed as follows
- ▶ Once you transit to a state, it spends time, say, $T_i \sim \text{exponential}(\nu_i)$ in it.
- ▶ Then, when it leaves i , it transits to state j with probability, say, z_{ij}
- ▶ We would have $z_{ij} \geq 0$, $\sum_j z_{ij} = 1$. Also, $z_{ii} = 0$
- ▶ Note that $P_{ij}(t)$ is different from these z_{ij}

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time
- ▶ From i the process can only go to $i + 1$ or $i - 1$

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time
- ▶ From i the process can only go to $i + 1$ or $i - 1$
- ▶ A birth event takes it to $i + 1$ and a death event takes it to $i - 1$

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time
- ▶ From i the process can only go to $i + 1$ or $i - 1$
- ▶ A birth event takes it to $i + 1$ and a death event takes it to $i - 1$
- ▶ An example would be: $X(t)$ is number of people in a queuing system.

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time
- ▶ From i the process can only go to $i + 1$ or $i - 1$
- ▶ A birth event takes it to $i + 1$ and a death event takes it to $i - 1$
- ▶ An example would be: $X(t)$ is number of people in a queuing system.
- ▶ A birth event would be a new person joining the queue.

Example: Birth-Death process

- ▶ This is generalization of birth-death chains we saw earlier to continuous time
- ▶ From i the process can only go to $i + 1$ or $i - 1$
- ▶ A birth event takes it to $i + 1$ and a death event takes it to $i - 1$
- ▶ An example would be: $X(t)$ is number of people in a queuing system.
- ▶ A birth event would be a new person joining the queue.
- ▶ A death event would be a person leaving after finishing service

- ▶ Suppose, in state n , time till next arrival or birth event is $\text{exponential}(\lambda_n)$.

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)

We assume that these two are independent

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).

- ▶ Let time till next departure or death event be exponential(μ_n)

We assume that these two are independent

- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$
- ▶ Hence it is the probability that a birth event occurs before a death event.

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$
- ▶ Hence it is the probability that a birth event occurs before a death event.
- ▶ Let $W_1 \sim \text{exponential}(\lambda_i)$ and $W_2 \sim \text{exponential}(\mu_i)$ be independent.

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$
- ▶ Hence it is the probability that a birth event occurs before a death event.
- ▶ Let $W_1 \sim \text{exponential}(\lambda_i)$ and $W_2 \sim \text{exponential}(\mu_i)$ be independent. Then

$$z_{i,i+1} = Pr[W_1 < W_2]$$

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$
- ▶ Hence it is the probability that a birth event occurs before a death event.
- ▶ Let $W_1 \sim \text{exponential}(\lambda_i)$ and $W_2 \sim \text{exponential}(\mu_i)$ be independent. Then

$$z_{i,i+1} = Pr[W_1 < W_2] = \frac{\lambda_i}{\lambda_i + \mu_i};$$

- ▶ Suppose, in state n , time till next arrival or birth event is exponential(λ_n).
- ▶ Let time till next departure or death event be exponential(μ_n)
We assume that these two are independent
- ▶ Now, these λ_n and μ_n completely determine ν_n and z_{ij} and hence completely specify the chain
- ▶ $z_{i,i+1}$ is the probability that when the system changes state it goes to $i + 1$
- ▶ Hence it is the probability that a birth event occurs before a death event.
- ▶ Let $W_1 \sim \text{exponential}(\lambda_i)$ and $W_2 \sim \text{exponential}(\mu_i)$ be independent. Then

$$z_{i,i+1} = Pr[W_1 < W_2] = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad \Rightarrow \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ The time spent in state i , T_i , is exponential(ν_i)

- ▶ The time spent in state i , T_i , is exponential(ν_i)
- ▶ The chain would be in state i till either a birth event or a death event occurs

- ▶ The time spent in state i , T_i , is exponential(ν_i)
- ▶ The chain would be in state i till either a birth event or a death event occurs
- ▶ Hence, $T_i = \min(W_1, W_2)$

- ▶ The time spent in state i , T_i , is $\text{exponential}(\nu_i)$
- ▶ The chain would be in state i till either a birth event or a death event occurs
- ▶ Hence, $T_i = \min(W_1, W_2)$
- ▶ Hence, $T_i \sim \text{exponential}(\lambda_i + \mu_i)$.

- ▶ The time spent in state i , T_i , is $\text{exponential}(\nu_i)$
- ▶ The chain would be in state i till either a birth event or a death event occurs
- ▶ Hence, $T_i = \min(W_1, W_2)$
- ▶ Hence, $T_i \sim \text{exponential}(\lambda_i + \mu_i)$.
- ▶ Thus, $\nu_i = \lambda_i + \mu_i$

- ▶ The time spent in state i , T_i , is $\text{exponential}(\nu_i)$
- ▶ The chain would be in state i till either a birth event or a death event occurs
- ▶ Hence, $T_i = \min(W_1, W_2)$
- ▶ Hence, $T_i \sim \text{exponential}(\lambda_i + \mu_i)$.
- ▶ Thus, $\nu_i = \lambda_i + \mu_i$
- ▶ We have taken state space to be non-negative integers.

- ▶ The time spent in state i , T_i , is exponential(ν_i)
- ▶ The chain would be in state i till either a birth event or a death event occurs
- ▶ Hence, $T_i = \min(W_1, W_2)$
- ▶ Hence, $T_i \sim \text{exponential}(\lambda_i + \mu_i)$.
- ▶ Thus, $\nu_i = \lambda_i + \mu_i$
- ▶ We have taken state space to be non-negative integers.
- ▶ Hence, $\mu_0 = 0$ and $\nu_0 = \lambda_0$ and $z_{01} = 1$

- ▶ Suppose $\lambda_n = \lambda, \forall n$ and $\mu_n = 0, \forall n$

- ▶ Suppose $\lambda_n = \lambda, \forall n$ and $\mu_n = 0, \forall n$
- ▶ It is called pure birth process

- ▶ Suppose $\lambda_n = \lambda, \forall n$ and $\mu_n = 0, \forall n$
- ▶ It is called pure birth process
- ▶ The process spend time $T_i \sim \text{exponential}(\lambda)$ in state i and then moves to state $i + 1$

- ▶ Suppose $\lambda_n = \lambda, \forall n$ and $\mu_n = 0, \forall n$
- ▶ It is called pure birth process
- ▶ The process spend time $T_i \sim \text{exponential}(\lambda)$ in state i and then moves to state $i + 1$
- ▶ This is the Poisson process

- ▶ Consider a queuing system

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .
- ▶ We assume that the arrival and service processes are independent.

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .
- ▶ We assume that the arrival and service processes are independent.
- ▶ Then this is a birth death process with

$$\lambda_n = \lambda, \quad n \geq 0 \quad \text{and} \quad \mu_n = \mu, \quad n \geq 1$$

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .
- ▶ We assume that the arrival and service processes are independent.
- ▶ Then this is a birth death process with

$$\lambda_n = \lambda, \quad n \geq 0 \quad \text{and} \quad \mu_n = \mu, \quad n \geq 1$$

- ▶ This is known as an $M/M/1$ queue

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .
- ▶ We assume that the arrival and service processes are independent.
- ▶ Then this is a birth death process with

$$\lambda_n = \lambda, \quad n \geq 0 \quad \text{and} \quad \mu_n = \mu, \quad n \geq 1$$

- ▶ This is known as an $M/M/1$ queue
- ▶ A variation: $M/M/K$ queue

- ▶ Consider a queuing system
- ▶ Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parameter μ .
- ▶ We assume that the arrival and service processes are independent.
- ▶ Then this is a birth death process with

$$\lambda_n = \lambda, \quad n \geq 0 \quad \text{and} \quad \mu_n = \mu, \quad n \geq 1$$

- ▶ This is known as an $M/M/1$ queue
- ▶ A variation: $M/M/K$ queue

$$\lambda_n = \lambda, \quad n \geq 0 \quad \text{and} \quad \mu_n = \begin{cases} n\mu & 1 \leq n \leq K \\ K\mu & n > K \end{cases}$$

- ▶ Consider an example of some calculations with continuous Markov chains

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.
- ▶ We want to calculate $E[Y_i]$.

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.
- ▶ We want to calculate $E[Y_i]$. (Note that $E[Y_0] = 1/\lambda_0$)

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.
- ▶ We want to calculate $E[Y_i]$. (Note that $E[Y_0] = 1/\lambda_0$)
- ▶ The chain may directly go to $i + 1$ or it may go to $i - 1$ and then to i and then to $i + 1$ or ...

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.
- ▶ We want to calculate $E[Y_i]$. (Note that $E[Y_0] = 1/\lambda_0$)
- ▶ The chain may directly go to $i + 1$ or it may go to $i - 1$ and then to i and then to $i + 1$ or ...
- ▶ Define

$$I_i = \begin{cases} 1 & \text{if first transition out of } i \text{ is to } i + 1 \\ 0 & \text{if first transition out of } i \text{ is to } i - 1 \end{cases}$$

- ▶ Consider an example of some calculations with continuous Markov chains
- ▶ Consider a Birth-Death process. Let Y_i be the time that a chain currently in i takes to reach state $i + 1$ for the first time.
- ▶ We want to calculate $E[Y_i]$. (Note that $E[Y_0] = 1/\lambda_0$)
- ▶ The chain may directly go to $i + 1$ or it may go to $i - 1$ and then to i and then to $i + 1$ or ...
- ▶ Define

$$I_i = \begin{cases} 1 & \text{if first transition out of } i \text{ is to } i + 1 \\ 0 & \text{if first transition out of } i \text{ is to } i - 1 \end{cases}$$

- ▶ We can find $E[Y_i]$ by conditioning on I_i .

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$
- ▶ If this transition is to $i + 1$ then that is the expected time to reach $i + 1$

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$
- ▶ If this transition is to $i + 1$ then that is the expected time to reach $i + 1$

$$E[Y_i \mid I_i = 1] = \frac{1}{\lambda_i + \mu_i}$$

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$
- ▶ If this transition is to $i + 1$ then that is the expected time to reach $i + 1$

$$E[Y_i \mid I_i = 1] = \frac{1}{\lambda_i + \mu_i}$$

- ▶ Suppose this transition is to $i - 1$.

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$
- ▶ If this transition is to $i + 1$ then that is the expected time to reach $i + 1$

$$E[Y_i \mid I_i = 1] = \frac{1}{\lambda_i + \mu_i}$$

- ▶ Suppose this transition is to $i - 1$.
- ▶ Then the expected time to reach $i + 1$ is this time plus expected time to reach i from $i - 1$ plus expected time to reach $i + 1$ from i

- ▶ Time spent in i is exponential with rate $\lambda_i + \mu_i$.
- ▶ Hence, expected time till transition out of i is $1/(\lambda_i + \mu_i)$
- ▶ If this transition is to $i + 1$ then that is the expected time to reach $i + 1$

$$E[Y_i \mid I_i = 1] = \frac{1}{\lambda_i + \mu_i}$$

- ▶ Suppose this transition is to $i - 1$.
- ▶ Then the expected time to reach $i + 1$ is this time plus expected time to reach i from $i - 1$ plus expected time to reach $i + 1$ from i

$$E[Y_i \mid I_i = 0] = \frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i]$$

- We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i};$$

- We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

$$E[Y_i] = Pr[I_i = 1] E[Y_i | I_i = 1] + Pr[I_i = 0] E[Y_i | I_i = 0]$$

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

$$\begin{aligned} E[Y_i] &= Pr[I_i = 1] E[Y_i | I_i = 1] + Pr[I_i = 0] E[Y_i | I_i = 0] \\ &= \frac{\lambda_i}{\lambda_i + \mu_i} \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} \left(\frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i] \right) \end{aligned}$$

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

$$\begin{aligned} E[Y_i] &= Pr[I_i = 1] E[Y_i | I_i = 1] + Pr[I_i = 0] E[Y_i | I_i = 0] \\ &= \frac{\lambda_i}{\lambda_i + \mu_i} \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} \left(\frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i] \right) \\ &= \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} (E[Y_{i-1}] + E[Y_i]) \end{aligned}$$

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

$$\begin{aligned} E[Y_i] &= Pr[I_i = 1] E[Y_i | I_i = 1] + Pr[I_i = 0] E[Y_i | I_i = 0] \\ &= \frac{\lambda_i}{\lambda_i + \mu_i} \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} \left(\frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i] \right) \\ &= \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} (E[Y_{i-1}] + E[Y_i]) \end{aligned}$$

$$E[Y_i] \left(1 - \frac{\mu_i}{\lambda_i + \mu_i} \right) = \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} (E[Y_{i-1}])$$

- ▶ We also have

$$Pr[I_i = 1] = z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}; \quad Pr[I_i = 0] = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate $E[Y_i]$ as

$$\begin{aligned} E[Y_i] &= Pr[I_i = 1] E[Y_i | I_i = 1] + Pr[I_i = 0] E[Y_i | I_i = 0] \\ &= \frac{\lambda_i}{\lambda_i + \mu_i} \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} \left(\frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i] \right) \\ &= \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} (E[Y_{i-1}] + E[Y_i]) \end{aligned}$$

$$\begin{aligned} E[Y_i] \left(1 - \frac{\mu_i}{\lambda_i + \mu_i} \right) &= \frac{1}{\lambda_i + \mu_i} + \frac{\mu_i}{\lambda_i + \mu_i} (E[Y_{i-1}]) \\ E[Y_i] &= \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}] \end{aligned}$$

- Thus we get

$$E[Y_i] = \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}], \quad i \geq 1$$

- ▶ Thus we get

$$E[Y_i] = \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}], \quad i \geq 1$$

- ▶ Since $E[Y_0] = 1/\lambda_0$, we have a formula for $E[Y_i]$

- ▶ Thus we get

$$E[Y_i] = \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}], \quad i \geq 1$$

- ▶ Since $E[Y_0] = 1/\lambda_0$, we have a formula for $E[Y_i]$
- ▶ For example,

$$E[Y_1] = \frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0}; \quad E[Y_2] = \frac{1}{\lambda_2} + \frac{\mu_2}{\lambda_2} \left(\frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0} \right)$$

- ▶ Thus we get

$$E[Y_i] = \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}], \quad i \geq 1$$

- ▶ Since $E[Y_0] = 1/\lambda_0$, we have a formula for $E[Y_i]$
- ▶ For example,

$$E[Y_1] = \frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0}; \quad E[Y_2] = \frac{1}{\lambda_2} + \frac{\mu_2}{\lambda_2} \left(\frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0} \right)$$

- ▶ Expected time to go from i to j , $i < j$ can now be computed as

$$E[Y_i] + E[Y_{i+1}] + \cdots + E[Y_{j-1}]$$

- ▶ Thus we get

$$E[Y_i] = \frac{1}{\lambda_i} + \frac{\mu_i}{\lambda_i} E[Y_{i-1}], \quad i \geq 1$$

- ▶ Since $E[Y_0] = 1/\lambda_0$, we have a formula for $E[Y_i]$
- ▶ For example,

$$E[Y_1] = \frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0}; \quad E[Y_2] = \frac{1}{\lambda_2} + \frac{\mu_2}{\lambda_2} \left(\frac{1}{\lambda_1} + \frac{\mu_1}{\lambda_1 \lambda_0} \right)$$

- ▶ Expected time to go from i to j , $i < j$ can now be computed as

$$E[Y_i] + E[Y_{i+1}] + \cdots + E[Y_{j-1}]$$

- ▶ Note that these are only for birth-death processes

- ▶ Consider the transition probabilities, $P_{ij}(t)$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$P_{ij}(t + s) = Pr[X(t + s) = j \mid X(0) = i]$$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$\begin{aligned} P_{ij}(t+s) &= Pr[X(t+s) = j \mid X(0) = i] \\ &= \sum_k Pr[X(t+s) = j \mid X(s) = k, X(0) = i] Pr[X(s) = k \mid X(0) = i] \end{aligned}$$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$\begin{aligned} P_{ij}(t+s) &= \Pr[X(t+s) = j \mid X(0) = i] \\ &= \sum_k \Pr[X(t+s) = j \mid X(s) = k, X(0) = i] \Pr[X(s) = k \mid X(0) = i] \\ &= \sum_k \Pr[X(t+s) = j \mid X(s) = k] \Pr[X(s) = k \mid X(0) = i] \end{aligned}$$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$\begin{aligned}P_{ij}(t+s) &= Pr[X(t+s) = j \mid X(0) = i] \\&= \sum_k Pr[X(t+s) = j \mid X(s) = k, X(0) = i] Pr[X(s) = k \mid X(0) = i] \\&= \sum_k Pr[X(t+s) = j \mid X(s) = k] Pr[X(s) = k \mid X(0) = i] \\&= \sum_k Pr[X(t) = j \mid X(0) = k] Pr[X(s) = k \mid X(0) = i]\end{aligned}$$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$\begin{aligned}P_{ij}(t+s) &= Pr[X(t+s) = j \mid X(0) = i] \\&= \sum_k Pr[X(t+s) = j \mid X(s) = k, X(0) = i] Pr[X(s) = k \mid X(0) = i] \\&= \sum_k Pr[X(t+s) = j \mid X(s) = k] Pr[X(s) = k \mid X(0) = i] \\&= \sum_k Pr[X(t) = j \mid X(0) = k] Pr[X(s) = k \mid X(0) = i] \\&= \sum_k P_{kj}(t) P_{ik}(s)\end{aligned}$$

- ▶ Consider the transition probabilities, $P_{ij}(t)$
- ▶ These satisfy Chapman-Kolmogorov equation

$$\begin{aligned}
 P_{ij}(t+s) &= Pr[X(t+s) = j \mid X(0) = i] \\
 &= \sum_k Pr[X(t+s) = j \mid X(s) = k, X(0) = i] Pr[X(s) = k \mid X(0) = i] \\
 &= \sum_k Pr[X(t+s) = j \mid X(s) = k] Pr[X(s) = k \mid X(0) = i] \\
 &= \sum_k Pr[X(t) = j \mid X(0) = k] Pr[X(s) = k \mid X(0) = i] \\
 &= \sum_k P_{kj}(t) P_{ik}(s)
 \end{aligned}$$

- ▶ For finite chain, P is a matrix and $P(t+s) = P(t) P(s)$

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Hence we get

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Hence we get

$$P_{ij}(t+h) - P_{ij}(t) = \sum_k P_{ik}(h) P_{kj}(t) - P_{ij}(t)$$

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Hence we get

$$\begin{aligned} P_{ij}(t+h) - P_{ij}(t) &= \sum_k P_{ik}(h) P_{kj}(t) - P_{ij}(t) \\ &= \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t) \end{aligned}$$

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Hence we get

$$\begin{aligned} P_{ij}(t+h) - P_{ij}(t) &= \sum_k P_{ik}(h) P_{kj}(t) - P_{ij}(t) \\ &= \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t) \end{aligned}$$

- ▶ Define

$$q_{ik} = \lim_{h \rightarrow 0} \frac{P_{ik}(h)}{h}, \quad i \neq k, \quad \text{and} \quad q_{ii} = \lim_{h \rightarrow 0} \frac{1 - P_{ii}(h)}{h}$$

- ▶ Chapman-Kolmogorov equation gives

$$P_{ij}(t+s) = \sum_k P_{ik}(s) P_{kj}(t)$$

- ▶ Hence we get

$$\begin{aligned} P_{ij}(t+h) - P_{ij}(t) &= \sum_k P_{ik}(h) P_{kj}(t) - P_{ij}(t) \\ &= \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t) \end{aligned}$$

- ▶ Define

$$q_{ik} = \lim_{h \rightarrow 0} \frac{P_{ik}(h)}{h}, \quad i \neq k, \quad \text{and} \quad q_{ii} = \lim_{h \rightarrow 0} \frac{1 - P_{ii}(h)}{h}$$

- ▶ Then, assuming limit and sum can be interchanged,

$$\lim_{h \rightarrow 0} \frac{P_{ij}(t+h) - P_{ij}(t)}{h} = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .
Also, two or more transitions in h is $o(h)$

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .
Also, two or more transitions in h is $o(h)$

- ▶ Hence

$$1 - P_{ii}(h) = \nu_i h + o(h)$$

- ▶ Thus $q_{ii} = \nu_i$. It is rate of transition out of i

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .
Also, two or more transitions in h is $o(h)$

- ▶ Hence

$$1 - P_{ii}(h) = \nu_i h + o(h)$$

- ▶ Thus $q_{ii} = \nu_i$. It is rate of transition out of i
- ▶ We also have

$$\nu_i = q_{ii} = \lim_{h \rightarrow 0} \frac{1 - P_{ii}(h)}{h}$$

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .
Also, two or more transitions in h is $o(h)$

- ▶ Hence

$$1 - P_{ii}(h) = \nu_i h + o(h)$$

- ▶ Thus $q_{ii} = \nu_i$. It is rate of transition out of i
- ▶ We also have

$$\nu_i = q_{ii} = \lim_{h \rightarrow 0} \frac{1 - P_{ii}(h)}{h} = \lim_{h \rightarrow 0} \frac{\sum_{j \neq i} P_{ij}(h)}{h}$$

- ▶ By definition, $1 - P_{ii}(h)$ is the probability that the chain that started in i is not in i at h .
- ▶ This is equivalent to there being a transition in the time h and transitions out of i occur at the rate of ν_i .
Also, two or more transitions in h is $o(h)$

- ▶ Hence

$$1 - P_{ii}(h) = \nu_i h + o(h)$$

- ▶ Thus $q_{ii} = \nu_i$. It is rate of transition out of i
- ▶ We also have

$$\nu_i = q_{ii} = \lim_{h \rightarrow 0} \frac{1 - P_{ii}(h)}{h} = \lim_{h \rightarrow 0} \frac{\sum_{j \neq i} P_{ij}(h)}{h} = \sum_{j \neq i} q_{ij}$$

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.
- ▶ Transitions out of i occur with rate ν_i and z_{ij} fraction of these are into j

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.
- ▶ Transitions out of i occur with rate ν_i and z_{ij} fraction of these are into j
- ▶ Hence, $q_{ij} = \nu_i z_{ij}$, $i \neq j$

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.
- ▶ Transitions out of i occur with rate ν_i and z_{ij} fraction of these are into j
- ▶ Hence, $q_{ij} = \nu_i z_{ij}$, $i \neq j$
- ▶ Thus, we got

$$\nu_i = \sum_{j \neq i} q_{ij}, \quad z_{ij} = \frac{q_{ij}}{\sum_{j \neq i} q_{ij}}, \quad q_{ii} = \sum_{j \neq i} q_{ij}$$

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.
- ▶ Transitions out of i occur with rate ν_i and z_{ij} fraction of these are into j
- ▶ Hence, $q_{ij} = \nu_i z_{ij}$, $i \neq j$
- ▶ Thus, we got

$$\nu_i = \sum_{j \neq i} q_{ij}, \quad z_{ij} = \frac{q_{ij}}{\sum_{j \neq i} q_{ij}}, \quad q_{ii} = - \sum_{j \neq i} q_{ij}$$

- ▶ The $\{q_{ij}\}$ are called the infinitesimal generator of the process.

- ▶ By definition, $P_{ij}(h) = q_{ij}h + o(h)$, $i \neq j$
- ▶ Hence q_{ij} is the rate at which transitions out of i into j are occurring.
- ▶ Transitions out of i occur with rate ν_i and z_{ij} fraction of these are into j
- ▶ Hence, $q_{ij} = \nu_i z_{ij}$, $i \neq j$
- ▶ Thus, we got

$$\nu_i = \sum_{j \neq i} q_{ij}, \quad z_{ij} = \frac{q_{ij}}{\sum_{j \neq i} q_{ij}}, \quad q_{ii} = -\sum_{j \neq i} q_{ij}$$

- ▶ The $\{q_{ij}\}$ are called the infinitesimal generator of the process.
- ▶ A continuous time Markov Chain is specified by these q_{ij}

- ▶ Consider a Birth-Death process.

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate q_{ij}

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate q_{ij}

$$q_{i,i+1} = (\lambda_i + \mu_i) \frac{\lambda_i}{\lambda_i + \mu_i} = \lambda_i,$$

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate q_{ij}

$$q_{i,i+1} = (\lambda_i + \mu_i) \frac{\lambda_i}{\lambda_i + \mu_i} = \lambda_i, \quad q_{i,i-1} = (\lambda_i + \mu_i) \frac{\mu_i}{\lambda_i + \mu_i} = \mu_i$$

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate q_{ij}

$$q_{i,i+1} = (\lambda_i + \mu_i) \frac{\lambda_i}{\lambda_i + \mu_i} = \lambda_i, \quad q_{i,i-1} = (\lambda_i + \mu_i) \frac{\mu_i}{\lambda_i + \mu_i} = \mu_i$$

- ▶ This is intuitively obvious

- ▶ Consider a Birth-Death process.
- ▶ We got earlier

$$\nu_i = \lambda_i + \mu_i, \quad z_{i,i+1} = \frac{\lambda_i}{\lambda_i + \mu_i}, \quad z_{i,i-1} = \frac{\mu_i}{\lambda_i + \mu_i}$$

- ▶ Now we can calculate q_{ij}

$$q_{i,i+1} = (\lambda_i + \mu_i) \frac{\lambda_i}{\lambda_i + \mu_i} = \lambda_i, \quad q_{i,i-1} = (\lambda_i + \mu_i) \frac{\mu_i}{\lambda_i + \mu_i} = \mu_i$$

- ▶ This is intuitively obvious
- ▶ We specify a birth-death chain by
birth rate (rate of transition from i to $i + 1$), λ_i and
death rate (rate of transition from i to $i - 1$), μ_i .

- ▶ The Chapman-Kolmogorov equations give us

$$P_{ij}(t+h) - P_{ij}(t) = \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t)$$

- ▶ The Chapman-Kolmogorov equations give us

$$P_{ij}(t+h) - P_{ij}(t) = \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t)$$

- ▶ Using this we derived

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The Chapman-Kolmogorov equations give us

$$P_{ij}(t+h) - P_{ij}(t) = \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t)$$

- ▶ Using this we derived

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

Called Kolmogorov Backward equation

- ▶ The Chapman-Kolmogorov equations give us

$$P_{ij}(t+h) - P_{ij}(t) = \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t)$$

- ▶ Using this we derived

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

Called Kolmogorov Backward equation

- ▶ We can solve these ODEs to get $P_{ij}(t)$

- ▶ The Chapman-Kolmogorov equations give us

$$P_{ij}(t+h) - P_{ij}(t) = \sum_{k \neq i} P_{ik}(h) P_{kj}(t) - (1 - P_{ii}(h)) P_{ij}(t)$$

- ▶ Using this we derived

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

Called Kolmogorov Backward equation

- ▶ We can solve these ODEs to get $P_{ij}(t)$
- ▶ For a birth-death chain the equation becomes

$$P'_{ij}(t) = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t)$$

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0. \forall i.$

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0, \forall i$.
- ▶ This would be a Poisson process with rate λ .

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0, \forall i$.
- ▶ This would be a Poisson process with rate λ .
- ▶ Taking $i = 0$, the differential equation becomes

$$P'_{0j}(t) = \lambda P_{1j}(t) - \lambda P_{0j}(t)$$

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0, \forall i$.
- ▶ This would be a Poisson process with rate λ .
- ▶ Taking $i = 0$, the differential equation becomes

$$P'_{0j}(t) = \lambda P_{1j}(t) - \lambda P_{0j}(t)$$

- ▶ $P_{0j}(t)$ is the probability of j events in an interval of length t which is same as what we had called $P_j(t)$.

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0, \forall i$.
- ▶ This would be a Poisson process with rate λ .
- ▶ Taking $i = 0$, the differential equation becomes

$$P'_{0j}(t) = \lambda P_{1j}(t) - \lambda P_{0j}(t)$$

- ▶ $P_{0j}(t)$ is the probability of j events in an interval of length t which is same as what we had called $P_j(t)$.
- ▶ Similarly, $P_{1j}(t)$ is same as what we called $P_{j-1}(t)$ there

Poisson process as a special case

- ▶ Consider the case: $\lambda_i = \lambda$ and $\mu_i = 0, \forall i$.
- ▶ This would be a Poisson process with rate λ .
- ▶ Taking $i = 0$, the differential equation becomes

$$P'_{0j}(t) = \lambda P_{1j}(t) - \lambda P_{0j}(t)$$

- ▶ $P_{0j}(t)$ is the probability of j events in an interval of length t which is same as what we had called $P_j(t)$.
- ▶ Similarly, $P_{1j}(t)$ is same as what we called $P_{j-1}(t)$ there
- ▶ Now one can see that the above ODE is what we got for Poisson process.

- ▶ Consider a two-state Birth-Death chain.

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$
- ▶ The two states can be a machine working or failed.

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$
- ▶ The two states can be a machine working or failed.
- ▶ λ is rate of failure. Time till next failure is $\text{exponential}(\lambda)$

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$
- ▶ The two states can be a machine working or failed.
- ▶ λ is rate of failure. Time till next failure is $\text{exponential}(\lambda)$
- ▶ μ is rate of repair. Time for repair is $\text{exponential}(\mu)$

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$
- ▶ The two states can be a machine working or failed.
- ▶ λ is rate of failure. Time till next failure is exponential(λ)
- ▶ μ is rate of repair. Time for repair is exponential(μ)
- ▶ We may want to calculate $P_{00}(T)$, the probability that the machine would be working at a time T units later given it is in working condition now

- ▶ Consider a two-state Birth-Death chain.
- ▶ We would have $\mu_0 = \lambda_1 = 0$. Let $\lambda_0 = \lambda$ and $\mu_1 = \mu$
- ▶ The two states can be a machine working or failed.
- ▶ λ is rate of failure. Time till next failure is exponential(λ)
- ▶ μ is rate of repair. Time for repair is exponential(μ)
- ▶ We may want to calculate $P_{00}(T)$, the probability that the machine would be working at a time T units later given it is in working condition now
- ▶ We can calculate it by solving the ODE's

$$P'_{ij}(t) = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t)$$

- For the two state chain, these equations are

$$P'_{00}(t) = \lambda_0 P_{10}(t) - \lambda_0 P_{00}(t)$$

$$P'_{01}(t) = \lambda_0 P_{11}(t) - \lambda_0 P_{01}(t)$$

$$P'_{10}(t) = \mu_1 P_{00}(t) - \mu_1 P_{10}(t)$$

$$P'_{11}(t) = \mu_1 P_{01}(t) - \mu_1 P_{11}(t)$$

$$P'_{ij}(t) = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t)$$

- For the two state chain, these equations are

$$P'_{00}(t) = \lambda_0 P_{10}(t) - \lambda_0 P_{00}(t)$$

$$P'_{01}(t) = \lambda_0 P_{11}(t) - \lambda_0 P_{01}(t)$$

$$P'_{10}(t) = \mu_1 P_{00}(t) - \mu_1 P_{10}(t)$$

$$P'_{11}(t) = \mu_1 P_{01}(t) - \mu_1 P_{11}(t)$$

- As is easy to see, we get a system of equations like this for any finite chain.

$$P'_{ij}(t) = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t)$$

- ▶ For the two state chain, these equations are

$$P'_{00}(t) = \lambda_0 P_{10}(t) - \lambda_0 P_{00}(t)$$

$$P'_{01}(t) = \lambda_0 P_{11}(t) - \lambda_0 P_{01}(t)$$

$$P'_{10}(t) = \mu_1 P_{00}(t) - \mu_1 P_{10}(t)$$

$$P'_{11}(t) = \mu_1 P_{01}(t) - \mu_1 P_{11}(t)$$

- ▶ As is easy to see, we get a system of equations like this for any finite chain.
- ▶ Solving these we can show

$$P_{00}(t) = \frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)t} + \frac{\mu}{\lambda + \mu}$$

- ▶ Consider a finite chain

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix
- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix
- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Differentiating the above with respect to t

$$P'(t + s) = P'(t)P(s)$$

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix
- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Differentiating the above with respect to t

$$P'(t + s) = P'(t)P(s)$$

- ▶ Putting $t = 0$ in the above we get

$$P'(s) = P'(0) P(s)$$

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix
- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Differentiating the above with respect to t

$$P'(t + s) = P'(t) P(s)$$

- ▶ Putting $t = 0$ in the above we get

$$P'(s) = P'(0) P(s) = \bar{Q} P(s), \text{ where } \bar{Q} = P'(0)$$

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix
- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Differentiating the above with respect to t

$$P'(t + s) = P'(t)P(s)$$

- ▶ Putting $t = 0$ in the above we get

$$P'(s) = P'(0) P(s) = \bar{Q} P(s), \quad \text{where } \bar{Q} = P'(0)$$

- ▶ The solution for this is

$$P(t) = e^{t\bar{Q}}, \quad \text{because } P(0) = I$$

- ▶ Consider a finite chain
- ▶ Then the transition probabilities can be represented as a matrix

- ▶ The Chapman-Kolmogorov equation gives

$$P(t + s) = P(t) P(s)$$

- ▶ Differentiating the above with respect to t

$$P'(t + s) = P'(t)P(s)$$

- ▶ Putting $t = 0$ in the above we get

$$P'(s) = P'(0) P(s) = \bar{Q} P(s), \quad \text{where } \bar{Q} = P'(0)$$

- ▶ The solution for this is

$$P(t) = e^{t\bar{Q}}, \quad \text{because } P(0) = I$$

- ▶ This is the expression for calculating $P_{ij}(t)$ for any t and i, j

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0)$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h}$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\text{for } k \neq j, \quad \bar{q}_{kj} = \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h}$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\text{for } k \neq j, \quad \bar{q}_{kj} = \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj}$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\begin{aligned} \text{for } k \neq j, \quad \bar{q}_{kj} &= \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj} \\ \bar{q}_{jj} &= \lim_{h \downarrow 0} \frac{P_{jj}(h) - 1}{h} \end{aligned}$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\text{for } k \neq j, \quad \bar{q}_{kj} = \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj}$$

$$\bar{q}_{jj} = \lim_{h \downarrow 0} \frac{P_{jj}(h) - 1}{h} = -q_{jj} = -\nu_j$$

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\begin{aligned} \text{for } k \neq j, \quad \bar{q}_{kj} &= \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj} \\ \bar{q}_{jj} &= \lim_{h \downarrow 0} \frac{P_{jj}(h) - 1}{h} = -q_{jj} = -\nu_j \end{aligned}$$

- ▶ Thus this \bar{Q} matrix has q_{ik} as off-diagonal entries and $-q_{jj}$ as diagonal entries

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\begin{aligned} \text{for } k \neq j, \quad \bar{q}_{kj} &= \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj} \\ \bar{q}_{jj} &= \lim_{h \downarrow 0} \frac{P_{jj}(h) - 1}{h} = -q_{jj} = -\nu_j \end{aligned}$$

- ▶ Thus this \bar{Q} matrix has q_{ik} as off-diagonal entries and $-q_{jj}$ as diagonal entries
- ▶ So, each row here sums to zero

- ▶ Let us examine the matrix $\bar{Q} = [\bar{q}_{ij}]$

$$\bar{Q} = P'(0) = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h} = \lim_{h \downarrow 0} \frac{P(h) - I}{h}$$

- ▶ This gives us

$$\begin{aligned} \text{for } k \neq j, \quad \bar{q}_{kj} &= \lim_{h \downarrow 0} \frac{P_{kj}(h) - 0}{h} = q_{kj} \\ \bar{q}_{jj} &= \lim_{h \downarrow 0} \frac{P_{jj}(h) - 1}{h} = -q_{jj} = -\nu_j \end{aligned}$$

- ▶ Thus this \bar{Q} matrix has q_{ik} as off-diagonal entries and $-q_{jj}$ as diagonal entries
- ▶ So, each row here sums to zero
- ▶ We normally write it as Q and call it the infinitesimal generator of the process

- ▶ The Kolmogorov backward equation is

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The Kolmogorov backward equation is

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The above can be written in a matrix form

$$P'(t) = QP(t)$$

- ▶ The Kolmogorov backward equation is

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The above can be written in a matrix form

$$P'(t) = QP(t)$$

- ▶ The off-diagonal entries of Q are q_{ik} and diagonal entries are $-q_{ii}$

- ▶ The Kolmogorov backward equation is

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The above can be written in a matrix form

$$P'(t) = QP(t)$$

- ▶ The off-diagonal entries of Q are q_{ik} and diagonal entries are $-q_{ii}$
- ▶ From the above equation, $P'(0) = Q$

- ▶ The Kolmogorov backward equation is

$$P'_{ij}(t) = \sum_{k \neq i} q_{ik} P_{kj}(t) - q_{ii} P_{ij}(t)$$

- ▶ The above can be written in a matrix form

$$P'(t) = QP(t)$$

- ▶ The off-diagonal entries of Q are q_{ik} and diagonal entries are $-q_{ii}$
- ▶ From the above equation, $P'(0) = Q$
- ▶ So, what we did is to write the backward equation in matrix form

- For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ Similar algebra as earlier gives us

$$P'_{ij}(t) = \sum_{k \neq j} P_{ik}(t) q_{kj} - q_{jj} P_{ij}(t)$$

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ Similar algebra as earlier gives us

$$P'_{ij}(t) = \sum_{k \neq j} P_{ik}(t) q_{kj} - q_{jj} P_{ij}(t)$$

(under some assumptions about interchanging limit and summation)

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ Similar algebra as earlier gives us

$$P'_{ij}(t) = \sum_{k \neq j} P_{ik}(t) q_{kj} - q_{jj} P_{ij}(t)$$

(under some assumptions about interchanging limit and summation)

- ▶ This is known as Kolmogorov forward equation

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ Similar algebra as earlier gives us

$$P'_{ij}(t) = \sum_{k \neq j} P_{ik}(t) q_{kj} - q_{jj} P_{ij}(t)$$

(under some assumptions about interchanging limit and summation)

- ▶ This is known as Kolmogorov forward equation
- ▶ For finite chains, both forward and backward equations are same

- ▶ For the backward equation, we started with

$$P_{ij}(t+h) = \sum_k P_{ik}(h) P_{kj}(t)$$

- ▶ The Chapman-Kolmogorov equation also gives us

$$P_{ij}(t+h) = \sum_k P_{ik}(t) P_{kj}(h)$$

- ▶ Similar algebra as earlier gives us

$$P'_{ij}(t) = \sum_{k \neq j} P_{ik}(t) q_{kj} - q_{jj} P_{ij}(t)$$

(under some assumptions about interchanging limit and summation)

- ▶ This is known as Kolmogorov forward equation
- ▶ For finite chains, both forward and backward equations are same
- ▶ For infinite chains there are some differences

- ▶ We can define transient and recurrent states as in the discrete case.

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\}$$

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\} \quad f_i = \min\{t : t > T_i, X(t) = i\}$$

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\} \quad f_i = \min\{t : t > T_i, X(t) = i\}$$

- ▶ For a chain started in i we take f_i as first return time to i

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\} \quad f_i = \min\{t : t > T_i, X(t) = i\}$$

- ▶ For a chain started in i we take f_i as first return time to i
- ▶ A state i is said to be

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\} \quad f_i = \min\{t : t > T_i, X(t) = i\}$$

- ▶ For a chain started in i we take f_i as first return time to i
- ▶ A state i is said to be
 - ▶ Transient if $Pr[f_i < \infty \mid X(0) = i] < 1$

- ▶ We can define transient and recurrent states as in the discrete case.
- ▶ However, we need to be careful about defining hitting times or first passage times
- ▶ We define

$$T_i = \min\{t > 0 : X(t) \neq i\} \quad f_i = \min\{t : t > T_i, X(t) = i\}$$

- ▶ For a chain started in i we take f_i as first return time to i
- ▶ A state i is said to be
 - ▶ Transient if $Pr[f_i < \infty \mid X(0) = i] < 1$
 - ▶ Recurrent if $Pr[f_i < \infty \mid X(0) = i] = 1$

- ▶ Most of the other definitions are also similar to the case of discrete chains

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite: $E[f_i \mid X(0) = i] < \infty$

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite: $E[f_i \mid X(0) = i] < \infty$
Otherwise it is null recurrent

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite: $E[f_i \mid X(0) = i] < \infty$
Otherwise it is null recurrent
- ▶ An irreducible positive recurrent chain would have a unique stationary distribution

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite: $E[f_i \mid X(0) = i] < \infty$
Otherwise it is null recurrent
- ▶ An irreducible positive recurrent chain would have a unique stationary distribution
- ▶ There is no concept of periodicity in the continuous time case

- ▶ Most of the other definitions are also similar to the case of discrete chains
- ▶ The chain is said to be irreducible if for all i, j there is a positive probability of going from i to j in some finite time: $P_{ij}(t) > 0$ for some t
- ▶ A recurrent state is positive recurrent if mean time to return is finite: $E[f_i \mid X(0) = i] < \infty$
Otherwise it is null recurrent
- ▶ An irreducible positive recurrent chain would have a unique stationary distribution
- ▶ There is no concept of periodicity in the continuous time case
- ▶ An irreducible positive recurrent chain would be called an ergodic chain

► Define

$$\pi_j(t) = Pr[X(t) = j]$$

► Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

► Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

This also analogous to the discrete case

- ▶ Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

This also analogous to the discrete case

- ▶ The above equation is true for general infinite chains.

- ▶ Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

This also analogous to the discrete case

- ▶ The above equation is true for general infinite chains.
- ▶ In the finite case, we can get a more compact expression

- ▶ Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

This also analogous to the discrete case

- ▶ The above equation is true for general infinite chains.
- ▶ In the finite case, we can get a more compact expression
- ▶ For a finite chain, taking π as a row vector,

$$\pi(t) = \pi(0) P(t)$$

- ▶ Define

$$\pi_j(t) = Pr[X(t) = j] = \sum_i \pi_i(0) P_{ij}(t)$$

This also analogous to the discrete case

- ▶ The above equation is true for general infinite chains.
- ▶ In the finite case, we can get a more compact expression
- ▶ For a finite chain, taking π as a row vector,

$$\pi(t) = \pi(0) P(t) = \pi(0) e^{Qt}$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution,
 $\pi'(t) = 0$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution,
 $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t)$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution, $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t) \quad \text{and hence} \quad \pi'_j(t) = \sum_i \pi_i(0) P'_{ij}(t)$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution, $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t) \quad \text{and hence} \quad \pi'_j(t) = \sum_i \pi_i(0) P'_{ij}(t)$$

- ▶ Using the forward equation for $P'_{ij}(t)$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution,
 $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t) \quad \text{and hence} \quad \pi'_j(t) = \sum_i \pi_i(0) P'_{ij}(t)$$

- ▶ Using the forward equation for $P'_{ij}(t)$

$$\sum_i \pi_i(0) \left(\sum_{k \neq j} q_{kj} P_{ik}(t) - q_{jj} P_{ij}(t) \right) = 0$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution,
 $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t) \quad \text{and hence} \quad \pi'_j(t) = \sum_i \pi_i(0) P'_{ij}(t)$$

- ▶ Using the forward equation for $P'_{ij}(t)$

$$\begin{aligned} \sum_i \pi_i(0) \left(\sum_{k \neq j} q_{kj} P_{ik}(t) - q_{jj} P_{ij}(t) \right) &= 0 \\ \Rightarrow \sum_{k \neq j} q_{kj} \pi_k - \pi_j \sum_{k \neq j} q_{jk} &= 0 \end{aligned}$$

- ▶ We say π is a stationary distribution if

$$\pi(0) = \pi \quad \Rightarrow \quad \pi(t) = \pi, \quad \forall t$$

- ▶ Hence, if we start the chain in the stationary distribution, $\pi'(t) = 0$
- ▶ We get from the earlier equation

$$\pi_j(t) = \sum_i \pi_i(0) P_{ij}(t) \quad \text{and hence} \quad \pi'_j(t) = \sum_i \pi_i(0) P'_{ij}(t)$$

- ▶ Using the forward equation for $P'_{ij}(t)$

$$\begin{aligned} \sum_i \pi_i(0) \left(\sum_{k \neq j} q_{kj} P_{ik}(t) - q_{jj} P_{ij}(t) \right) &= 0 \\ \Rightarrow \sum_{k \neq j} q_{kj} \pi_k - \pi_j \sum_{k \neq j} q_{jk} &= 0 \end{aligned}$$

when π is a stationary distribution and $\pi(0) = \pi$

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)
- ▶ q_{kj} is the rate of transition from k to j and π_k is the fraction present in k .

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)
- ▶ q_{kj} is the rate of transition from k to j and π_k is the fraction present in k .
- ▶ Hence $\sum_{k \neq j} q_{kj} \pi_k$ is the net flow into j

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)
- ▶ q_{kj} is the rate of transition from k to j and π_k is the fraction present in k .
- ▶ Hence $\sum_{k \neq j} q_{kj} \pi_k$ is the net flow into j
- ▶ $\pi_j \sum_{k \neq j} q_{jk}$ is the net flow out of j

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)
- ▶ q_{kj} is the rate of transition from k to j and π_k is the fraction present in k .
- ▶ Hence $\sum_{k \neq j} q_{kj} \pi_k$ is the net flow into j
- ▶ $\pi_j \sum_{k \neq j} q_{jk}$ is the net flow out of j
- ▶ At steady state the flows have to be balanced

- ▶ What we showed is that any stationary distribution π has to satisfy

$$\sum_{k \neq j} q_{kj} \pi_k = \pi_j \sum_{k \neq j} q_{jk}$$

- ▶ We can interpret this (as we did in discrete case)
- ▶ q_{kj} is the rate of transition from k to j and π_k is the fraction present in k .
- ▶ Hence $\sum_{k \neq j} q_{kj} \pi_k$ is the net flow into j
- ▶ $\pi_j \sum_{k \neq j} q_{jk}$ is the net flow out of j
- ▶ At steady state the flows have to be balanced
- ▶ The above equation is known as a balance equation