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- ▶ Then the density of  $\mathbf{Y}$  is

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- ▶ Called multidimensional change of variable formula

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$$f_{(X/Y)}(z) = \int_{-\infty}^{\infty} |t| f_{XY}(zt, t) dt = \int_{-\infty}^{\infty} \left| \frac{t}{z^2} \right| f_{XY} \left( t, \frac{t}{z} \right) dt$$

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- ▶ If the random variables are exchangeable then the joint distribution function remains the same on permutation of arguments.
- ▶ Exchangeable random variables are identically distributed but they may not be independent.

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- ▶ Let  $Z = g(X_1, \dots, X_n) = g(\mathbf{X})$ . Then

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- ▶ This gives us:  $E[X + Y] = E[X] + E[Y]$
- ▶ In general,  $E[g_1(\mathbf{X}) + g_2(\mathbf{X})] = E[g_1(\mathbf{X})] + E[g_2(\mathbf{X})]$

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where we define **covariance** between  $X, Y$  as

$$\text{Cov}(X, Y) = E \left[ (X - EX)(Y - EY) \right]$$

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- ▶ Note that  $E[X + Y] = E[X] + E[Y]$  for all random variables.



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- ▶ Hence,  $\text{Cov}(X, Y) = E[XY] - EX \cdot EY = \frac{1}{4} - \frac{2}{9} = \frac{1}{36}$

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- ▶ Then,  $\text{Cov}(X, Y) = E[XY] - EX EY = 0$ .
- ▶  $X, Y$  independent  $\Rightarrow X, Y$  uncorrelated



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- ▶  $X, Y$  are uncorrelated does not imply they are independent.

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- ▶ We will show that  $|\rho_{XY}| \leq 1$
- ▶ Hence  $-1 \leq \rho_{XY} \leq 1, \forall X, Y$

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$$\frac{(E[XY])^2}{E[X^2]} + \beta^2 E[Y^2] - 2\beta \frac{(E[XY])^2}{E[X^2]} \geq 0, \quad \forall \beta \in \mathfrak{R}$$

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$$\frac{(E[XY])^2}{E[X^2]} + \beta^2 E[Y^2] - 2\beta \frac{(E[XY])^2}{E[X^2]} \geq 0, \quad \forall \beta \in \Re$$

$$\Rightarrow 4 \left( \frac{(E[XY])^2}{E[X^2]} \right)^2 - 4E[Y^2] \frac{(E[XY])^2}{E[X^2]} \leq 0$$

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- ▶ Informally, covariance captures the ‘linear dependence’ between the two random variables.



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$$\Sigma_{\mathbf{X}} = \begin{bmatrix} \text{Cov}(X_1, X_1) & \text{Cov}(X_1, X_2) & \cdots & \text{Cov}(X_1, X_n) \\ \text{Cov}(X_2, X_1) & \text{Cov}(X_2, X_2) & \cdots & \text{Cov}(X_2, X_n) \\ \vdots & \vdots & \vdots & \vdots \\ \text{Cov}(X_n, X_1) & \text{Cov}(X_n, X_2) & \cdots & \text{Cov}(X_n, X_n) \end{bmatrix}$$

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- ▶ If  $X_i$  are independent, variance of sum is sum of variances.

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- ▶ Hence,  $\Sigma_X$  fails to be positive definite only if there is a non-zero linear combination of  $X_i$ 's that is a constant.



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- ▶ We consider one simple example.

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- ▶ Hence the direction is the eigen vector corresponding to the highest eigen value.

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- ▶ We can similarly define joint moments of multiple random variables

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- ▶ More generally

$$\left. \frac{\partial^{m+n}}{\partial s_i^n \partial s_j^m} M_{\mathbf{X}}(\mathbf{s}) \right|_{\mathbf{s}=\mathbf{0}} = EX_i^n X_j^m$$

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- ▶ We will now define it formally

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- ▶ What this means is that we define  $E[h(X)|Y] = g(Y)$  where

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$$\begin{aligned} E[h(X)|Y = y] &= \sum_x h(x) f_{X|Y}(x|y) \\ &= \sum_x h(x) P[X = x|Y = y] \end{aligned}$$

- ▶ What this means is that we define  $E[h(X)|Y] = g(Y)$  where

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- ▶ The last property above follows directly from the definition.

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- ▶ Similarly

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