## Recap: Random Variables

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- ▶ Given a probability space  $(\Omega, \mathcal{F}, P)$ , a random variable is a real-valued function on  $\Omega$ .
- It essentially results in an induced probability space

$$(\Omega, \mathcal{F}, P) \stackrel{X}{\to} (\Re, \mathcal{B}, P_X)$$

where  ${\cal B}$  is the Borel  $\sigma$ -algebra and

$$P_X(B) = P[X \in B] = P(\{\omega \in \Omega : X(\omega) \in B\})$$

- ▶ An  $\mathcal{F} \subset 2^{\Omega}$  is called a  $\sigma$ -algebra (also called  $\sigma$ -field) on  $\Omega$  if it satisfies
  - 1.  $\Omega \in \mathcal{F}$
  - 2.  $A \in \mathcal{F} \Rightarrow A^c \in \mathcal{F}$
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- We also have  $\mathcal{B} = \sigma(\{(-\infty, x] : x \in \Re\})$

# Recap: Distribution function of a random variable

▶ Let X be a random variable. It distribution function,  $F_X: \Re \to \Re$ , is defined by

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▶ The distribution function,  $F_X$ , completely specifies the probability measure,  $P_X$ .

- The distribution function satisfies
  - 1.  $0 \le F_X(x) \le 1, \ \forall x$
  - 2.  $F_X(-\infty) = 0$ ;  $F_X(\infty) = 1$
  - 3.  $F_X$  is non-decreasing:  $x_1 \le x_2 \implies F_X(x_1) \le F_X(x_2)$
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 $P[a < X \le b] = F_X(b) - F_X(a).$ 

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- Note that the distribution function is defined for all random variables.

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$$\begin{split} [X \leq 1.57] &= \{\omega \ : \ X(\omega) \leq 1.57\} \\ &= \{\omega \ : \ X(\omega) = 0\} \cup \{\omega \ : \ X(\omega) = 1\} = [X = 0 \text{ or } 1] \end{split}$$

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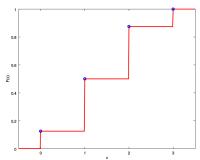
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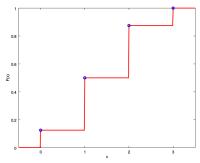
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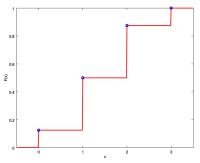
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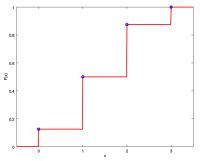




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- ▶ It has jumps at x = 0, 1, 2, 3, which are the values that X takes. In between these it is constant.
- ▶ The jump at, e.g., x = 2 is 3/8 which is the probability of X taking that value.

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- All discrete random variables would have this general form of distribution function.

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- If  $x < a_1$  then  $[X \le x] = \phi$ .
- If  $a_1 \le x < a_2$  then  $[X \le x] = [X = a_1] = B_1$

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- If  $x < a_1$  then  $[X \le x] = \phi$ .
- If  $a_1 \le x < a_2$  then  $[X \le x] = [X = a_1] = B_1$
- ▶ If  $a_2 \le x < a_3$  then  $[X \le x] = [X = a_1] \cup [X = a_2] = B_1 + B_2$

▶ Hence we can write the distribution function as

$$F_X(x) = \begin{cases} 0 & x < a_1 \\ P(B_1) & a_1 \le x < a_2 \\ P(B_1) + P(B_2) & a_2 \le x < a_3 \end{cases}$$

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▶ Note that all this holds even when *X* takes countably infinitely many values.

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▶ We can get pmf from df and df from pmf.

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- Please remember that we have defined distribution function for any random variable. But pmf is defined only for discrete random variables

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- As we saw this is how we can specify a probability assignment on any countable sample space.

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 We next look at some standard discrete random variable models

▶ Bernoulli random variable:  $X \in \{0, 1\}$  with

 $f_X(1) = p; \ f_X(0) = 1 - p; \quad \text{where } 0$ 

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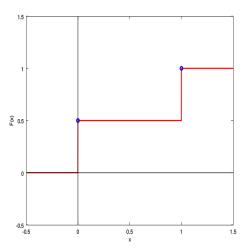
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- $P[I_B = 1] = P(\{\omega : I_B(\omega) = 1\}) = P(B)$
- ▶ Thus, this indicator rv has Bernoulli distribution with p = P(B)

#### One of the df examples we saw earlier is that of Bernoulli



#### Binomial Distribution

•  $X \in \{0, 1, \dots, n\}$  with pmf

$$f_X(k) = {}^{n}C_k p^k (1-p)^{n-k}, k = 0, 1, \dots, n$$

where n, p are parameters (n is a +ve integer and 0 ).

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► Consider *n* independent tosses of coin whose probability of heads is *p*. If *X* is the number of heads then *X* has the above binomial distribution.

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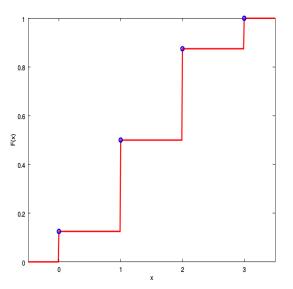
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  - (Number of successes in n bernoulli trials)
- Any one outcome (a seq of length n) with k heads would have probability  $p^k(1-p)^{n-k}$ . There are  ${}^nC_k$  outcomes with exactly k heads.

#### One of the df examples we considered was that of Binomial



 $X \in \{0, 1, 2, \cdots\}$  with pmf

$$f_X(k) = \frac{\lambda^k e^{-\lambda}}{k!}, \ k = 0, 1, 2, \dots$$

where  $\lambda > 0$  is a parameter.

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Poisson distribution is also useful in many applications

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- ► Consider tossing a coin (with prob of H being p) repeatedly till we get a head. X is the toss number on which we got the first head.
- ► In general waiting for 'success' in independent Bernoulli trials.

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(Does this also tell us what is df of geometric rv?)

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- Same as

$$P[X > m + n] = P[X > m]P[X > n]$$

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- ▶ Does it say that [X > m] is independent of [X > n]
- ► NO!

▶ If X is a geometric random variable, it satisfies

$$P[X > m + n | X > m] = P[X > n]$$

This is same as

$$P[X > m + n] = P[X > m]P[X > n]$$

- ▶ Does it say that [X > m] is independent of [X > n]
- $\blacktriangleright$  NO! Because [X>m+n] is not equal to intersection of [X>m] and [X>n]

$$P[X>m+n] = P[X>m]P[X>n]$$

▶ Suppose  $X \in \{0, 1, \cdots\}$  is a discrete rv satisfying, for all non-negative integers, m, n

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▶ We will show that *X* has geometric distribution

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- Let us take P[X > 0] = 1 (and hence P[X = 0] = 0).

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▶ This is pmf of geometric (with q = (1 - p))

▶ A rv, X, is said to be continuous (or of continuous type) if its distribution function,  $F_X$  is absolutely continuous.

▶ A function  $g:\Re\to\Re$  is absolutely continuous on an interval, I, if given any  $\epsilon>0$  there is a  $\delta>0$  such that for any finite sequence of pair-wise disjoint subintervals,  $(x_k,y_k)$ , with  $x_k,y_k\in I,\ \forall k$ , satisfying  $\sum_k(y_k-x_k)<\delta$ , we have  $\sum_k|f(y_k)-f(x_k)|<\epsilon$ 

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▶ In the above, g would be differentiable almost everywhere and h would be its derivative (wherever g is differentiable).

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- ► As mentioned earlier, there would be many random variables that are neither discrete nor continuous.

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▶ Hence, if X is a continuous random variable then

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- ▶ This shows the the  $F_X$  is a df and hence  $f_X$  is a pdf



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- We found that the df for this is

$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ x & \text{if } 0 \le x < 1 \\ 1 & \text{if } x \ge 1 \end{cases}$$

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This is absolutely continuous and we can get the pdf as

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- ▶ Let us find  $F_Y$  and  $f_Y$ .

$$[Y \le y] = \{\omega : Y(\omega) \le y\}$$

$$[Y \le y] = \{\omega : Y(\omega) \le y\} = \{\omega \in [0, 1] : 1 - \omega \le y\}$$

$$[Y \le y] = \{\omega : Y(\omega) \le y\} = \{\omega \in [0, 1] : 1 - \omega \le y\}$$
$$= \{\omega \in [0, 1] : \omega > 1 - y\}$$

$$\begin{split} [Y \leq y] &= & \{\omega \ : \ Y(\omega) \leq y\} = \{\omega \in [0, \ 1] \ : \ 1 - \omega \leq y\} \\ &= & \{\omega \in [0, \ 1] \ : \ \omega \geq 1 - y\} \\ &= & \left\{ \begin{array}{ccc} \phi & \text{if} \ y < 0 \end{array} \right. \end{split}$$

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▶ We have  $F_X = F_Y$  and thus  $f_X = f_Y$ . (However, note that  $X(\omega) \neq Y(\omega)$  except at  $\omega = 0.5$ ).

▶ Let *X* be a continuous rv.

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- ▶ Writing  $f_X(X=5)$  when  $f_X$  is a pdf, is particularly bad. Note that for a continuous rv, P[X=5]=0 and  $f_X(5) \neq P[X=5]$ .

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- We next consider a few standard continuous random variables.

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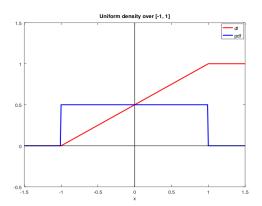
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- ► The earlier examples we considered are uniform over [0, 1].

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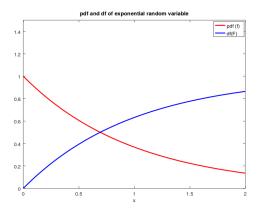
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► This also gives us:  $P[X > x] = 1 - F_X(x) = e^{-\lambda x}$  for x > 0.

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- ▶ If the distribution of a non-negative continuous random variable is memory less then it must be exponential.

▶ The pdf of Gaussian distribution is given by

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$$

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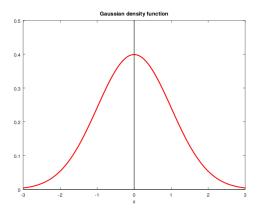
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Now converting the above integral into polar coordinates would allow you to show I=1.

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- ▶ Take  $\mu = 0, \sigma = 1$ . Let  $I = \int_{-\infty}^{\infty} f_X(x) dx$ . Then

$$I^{2} = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-0.5x^{2}} dx \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-0.5y^{2}} dy$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-0.5(x^{2} + y^{2})} dx dy$$

Now converting the above integral into polar coordinates would allow you to show I=1. (Left as an exercise for you!)