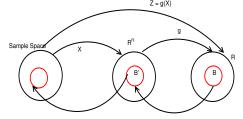
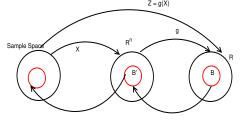
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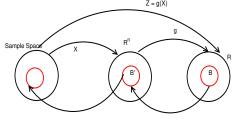


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Sum of independent exponential random variables has gamma density.

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Remaining details are left as an exercise for you!!

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▶ Is this correct for all values of z, w?

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- We calculated the order statistics for the case n=2.
- It can be shown that

$$f_{X_{(1)}\cdots X_{(n)}}(x_1, \cdots x_n) = n! \prod_{i=1}^n f(x_i), \ x_1 < x_2 < \cdots < x_n$$

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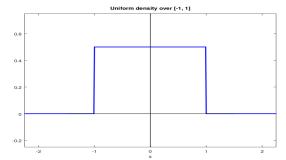
$$F_{X_{(k)}}(y) = \sum_{j=-k}^{n} {}^{n}C_{j}(F(y))^{j}(1-F(y))^{n-j}$$

We can get the density by differentiating this.

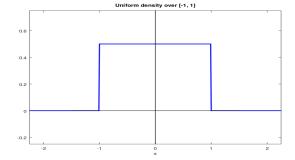
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•  $f_Z$  is convolution of this density with itself.

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- ▶ Thus we get

$$f_Z(z) = \begin{cases} \int_{-1}^{z+1} \frac{1}{4} dt = \frac{z+2}{4} & \text{if } -2 \le z < 0 \end{cases}$$

- $f_X(x) = 0.5, -1 < x < 1.$   $f_Y$  is also same
- ▶ Note that Z takes values in [-2, 2]

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(t) \ f_Y(z-t) \ dt$$

- For the integrand to be non-zero we need
  - $-1 < t < 1 \implies t < 1, t > -1$
  - $-1 < z t < 1 \implies t < z + 1, t > z 1$
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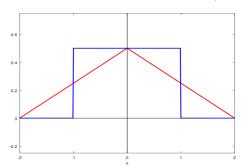
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- ► Exercise for you: Find density of  $X_1 + X_2 + X_3$  where  $X_1, X_2, X_3$  are iid uniform over (0, 1).

▶ Gamma density with parameters  $\alpha > 0$  and  $\lambda > 0$  is given by

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$$\begin{split} f_Z(z) &= \int_{-\infty}^{\infty} f_X(x) \ f_Y(z-x) \ dx \\ &= \int_{0}^{z} \frac{1}{\Gamma(\alpha_1)} \ \lambda^{\alpha_1} \ x^{\alpha_1-1} \ e^{-\lambda x} \frac{1}{\Gamma(\alpha_2)} \ \lambda^{\alpha_2} \ (z-x)^{\alpha_2-1} \ e^{-\lambda(z-x)} \ dx \\ &= \frac{\lambda^{\alpha_1+\alpha_2} \ e^{-\lambda z}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \int_{0}^{z} z^{\alpha_1-1} \left(\frac{x}{z}\right)^{\alpha_1-1} z^{\alpha_2-1} \left(1-\frac{x}{z}\right)^{\alpha_2-1} \ dx \\ &= \text{change the variable:} \quad t = \frac{x}{z} \ (\Rightarrow \ z^{-1} dx = dt) \end{split}$$

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Because

$$\int_0^1 t^{\alpha_1 - 1} (1 - t)^{\alpha_2 - 1} dt = \frac{\Gamma(\alpha_1) \Gamma(\alpha_2)}{\Gamma(\alpha_1 + \alpha_2)}$$

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- The algebra is a little involved.
- First take the two gaussians to be zero-mean.
- There is a calculation trick that is often useful with Gaussian density

$$I = \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2K} \left[x^2 - 2bx + c\right]\right) dx$$

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$$= \exp\left(-\frac{(c-b^2)}{2K}\right) \sqrt{2\pi K}$$

because

$$\frac{1}{\sqrt{2\pi K}} \int_{-\infty}^{\infty} \exp\left(-\frac{(x-b)^2}{2K}\right) dx = 1$$

► We next look at a general theorem that is quite useful in dealing with functions of multiple random variables.

- ▶ We next look at a general theorem that is quite useful in dealing with functions of multiple random variables.
- ▶ This result is only for continuous random variables.

Let  $X_1, \dots, X_n$  be continuous random variables with joint density  $f_{X_1 \dots X_n}$ .

$$Y_1 = g_1(X_1, \dots, X_n) \quad \dots \quad Y_n = g_n(X_1, \dots, X_n)$$

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- ▶ Let h be the inverse of g. That is

$$X_1 = h_1(Y_1, \cdots, Y_n) \quad \cdots \quad X_n = h_n(Y_1, \cdots, Y_n)$$

$$Y_1 = g_1(X_1, \dots, X_n) \quad \dots \quad Y_n = g_n(X_1, \dots, X_n)$$

We think of  $q_i$  as components of  $q: \Re^n \to \Re^n$ .

- $\blacktriangleright$  We assume g is continuous with continuous first partials and is invertible.
- ▶ Let *h* be the inverse of *q*. That is

$$X_1 = h_1(Y_1, \dots, Y_n) \quad \dots \quad X_n = h_n(Y_1, \dots, Y_n)$$

▶ Each of  $g_i, h_i$  are  $\Re^n \to \Re$  functions and we can write them as

$$y_i = g_i(x_1, \cdots, x_n); \quad \cdots \quad x_i = h_i(y_1, \cdots, y_n)$$

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$$y_i = g_i(x_1, \cdots, x_n); \quad \cdots \quad x_i = h_i(y_1, \cdots, y_n)$$

We denote the partial derivatives of these functions by  $\frac{\partial x_i}{\partial u_i}$  etc.

$$J = \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \dots & \frac{\partial x_1}{\partial y_n} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \dots & \frac{\partial x_2}{\partial y_n} \\ \vdots & \vdots & \vdots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \frac{\partial x_n}{\partial y_2} & \dots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

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- ▶ **Theorem**: Under the above conditions, we have

$$f_{Y_1\cdots Y_n}(y_1,\cdots,y_n) = |J|f_{X_1\cdots X_n}(h_1(y_1,\cdots,y_n),\cdots,h_n(y_1,\cdots,y_n))$$

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- ► We assume that *J* is non-zero in the range of the transformation
- ▶ **Theorem**: Under the above conditions, we have

$$f_{Y_1\cdots Y_n}(y_1,\cdots,y_n) = |J|f_{X_1\cdots X_n}(h_1(y_1,\cdots,y_n),\cdots,h_n(y_1,\cdots,y_n))$$

Or, more compactly,  $f_{\mathbf{Y}}(\mathbf{y}) = |J| f_{\mathbf{X}}(h(\mathbf{y}))$ 

▶ Let  $B = (-\infty, y_1] \times \cdots \times (-\infty, y_n] \subset \Re^n$ .

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Then we have

$$F_{Y_1...Y_n}(y_1, \dots, y_n) = P[g_i(X_1, \dots, X_n) \le y_i, i = 1, \dots n]$$

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$$F_{Y_1 \cdots Y_n}(y_1, \cdots y_n) = P[g_i(X_1, \cdots, X_n) \le y_i, \ i = 1, \cdots n]$$
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$$= \int_{g^{-1}(B)} f_{X_1\cdots X_n}(x_1',\cdots,x_n') \ dx_1' \ \cdots \ dx_n'$$
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$$(x_1',\cdots x_n') \in g^{-1}(B) \Rightarrow (y_1',\cdots,y_n') \in B$$

$$x_i' = h_i(y_1',\cdots,y_n'), \quad dx_1'\cdots dx_n' = |J|dy_1'\cdots dy_n'$$

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$$(x'_1, \dots x'_n) \in g^{-1}(B) \Rightarrow (y'_1, \dots, y'_n) \in B$$

$$x' = h(x'_1, \dots, x'_n) \quad dx'_1 \dots dx'_n = |I|dx'_1$$

$$x'_i = h_i(y'_1, \dots, y'_n), \quad dx'_1 \dots dx'_n = |J|dy'_1 \dots dy'_n$$

$$F_{\mathbf{Y}}(y_1, \cdots, y_n) = \int_B f_{X_1 \cdots X_n}(h_1(\mathbf{y}'), \cdots, h_n(\mathbf{y}')) |J| dy'_1 \cdots dy'_n$$

## Proof of Theorem

- $B = (-\infty, y_1] \times \cdots \times (-\infty, y_n].$
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$$x'_i = h_i(y'_1, \dots, y'_n), \quad dx'_1 \dots dx'_n = |J|dy'_1 \dots dy'_n$$

$$F_{\mathbf{Y}}(y_1,\cdots,y_n) = \int_{\mathbb{R}} f_{X_1\cdots X_n}(h_1(\mathbf{y}'),\cdots,h_n(\mathbf{y}')) |J| dy'_1\cdots dy'_n$$

$$\Rightarrow f_{Y_1\cdots Y_n}(y_1,\cdots,y_n) = f_{X_1\cdots X_n}(h_1(\mathbf{y}),\cdots,h_n(\mathbf{y})) |J|$$

•  $X_1, \cdots X_n$  are continuous rv with joint density

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► The transformation is continuous with continuous first partials and is invertible and

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► Called multidimensional change of variable formula

▶ Let X, Y have joint density  $f_{XY}$ . Let Z = X + Y.

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$$J = \begin{vmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{vmatrix} = -\frac{1}{2}$$

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ightharpoonup Now we get density of Z as

$$f_Z(z) = \int_{-\infty}^{\infty} \frac{1}{2} f_{XY}\left(\frac{z+w}{2}, \frac{z-w}{2}\right) dw$$

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$$\text{change the variable: } t = \frac{z+w}{2} \implies dt = \frac{1}{2} dw$$

$$\implies w = 2t - z \implies z - w = 2z - 2t$$

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 $=\int_{-\infty}^{\infty} f_{XY}(z-t,t) dt$ , by using  $t=\frac{z-w}{2}$  above

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We get same result as earlier.

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• We get same result as earlier. If, X, Y are independent

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(t) f_Y(z-t) dt$$

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$$= \int_{-\infty}^{\infty} f_{XY}(t+w, t) dt, \quad \text{using } t = \frac{z-w}{2} \quad \text{above}$$

## Example

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▶ This we have when  $X, Y \sim U(0,1)$  iid

$$f_{X-Y}(z) = 1 - |z|, -1 < z < 1$$



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Let X, Y be iid U(0, 1). Let Z = XY.

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$$f_Z(z) = \int_z^1 \left| \frac{1}{w} \right| dw = \int_z^1 \frac{1}{w} dw = -\ln(z), \ \ 0 < z < 1$$

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► We cannot always interchange density and mass functions!!

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▶ The  $f_Z$  should be same in both cases.

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We can show that the density of quotient is same in both these approches.

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- ▶ Take n = 3. Suppose  $F_{X_1X_2X_3}(a, b, c) = g(a, b, c)$ . If they are exchangeable, then

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=  $P[X_1 \le c, X_2 \le a, X_3 \le b]$ 

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- ▶ let  $(i_1, \dots, i_n)$  be a permutation of  $(1, 2, \dots, n)$ . Then joint df of  $(X_{i_1}, \dots, X_{i_n})$  should be same as that  $(X_1, \dots, X_n)$
- ▶ Take n = 3. Suppose  $F_{X_1X_2X_3}(a, b, c) = g(a, b, c)$ . If they are exchangeable, then

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$$\begin{aligned} F_{X_2X_3X_1}(a,b,c) &= & P[X_2 \le a, X_3 \le b, X_1 \le c] \\ &= & P[X_1 \le c, X_2 \le a, X_3 \le b] \\ &= & g(c,a,b) = g(a,b,c) \end{aligned}$$

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$$= g(c, a, b) = g(a, b, c)$$

► The df or density should be "symmetric" in its variables if the random variables are exchangeable.

$$f(x, y, z) = \frac{2}{3}(x + y + z), \quad 0 < x, y, z < 1$$

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▶ So, the joint density is not the product of marginals

# Expectation of functions of multiple rv

▶ **Theorem**: Let  $Z = g(X_1, \dots X_n) = g(\mathbf{X})$ . Then

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ightharpoonup Similarly, if all  $X_i$  are discrete

$$E[Z] = \sum_{\mathbf{x}} g(\mathbf{x}) \ f_{\mathbf{X}}(\mathbf{x})$$

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- Expectation is a linear operator.
- ► This is true for all random variables.