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$$P[(X,Y) \in B] = \sum_{\substack{i,j:\\ (\bar{x}_i,y_i) \in B}} f_{XY}(x_i,y_j)$$
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- Any function $f_{XY}: \Re^2 \to \Re$ satisfying the above two is a joint density function. (Then the above F_{XY} can be shown to be a joint df).
- We also have

$$P[x_1 \le X \le x_2, y_1 \le Y \le y_2] = \int_{x_1}^{x_2} \int_{y_1}^{y_2} f_{XY} \, dy \, dx$$

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$$P[(X,Y)\in B]=\int_{B}f_{XY}(x,y)\,dx\;dy,\;\forall B\in\mathcal{B}^2$$
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$$f_X(x) = \sum_{y} f_{XY}(x, y); \quad f_Y(y) = \sum_{x} f_{XY}(x, y)$$

• If X, Y have joint pdf f_{XY} then the marginal pdf are

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) \, dy; \quad f_Y(y) = \int_{-\infty}^{\infty} f_{XY}(x, y) \, dx$$

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- If X, Y have a joint density or if X is continuous and Y is discrete, F_{X|Y} would be absolutely continuous and would have a density.

▶ Let X be a discrete random variable. Then

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This exists if X, Y have a joint density or when Y is discrete.

Recap

▶ When *X,Y* are both discrete or they have a joint density

$$f_{XY}(x,y) = f_{X|Y}(x|y)f_Y(y) = f_{Y|X}(y|x)f_X(x)$$

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► The above relation gives rise to the total probability rules and Bayes rule for rv's

▶ If *Y* is discrete

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

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▶ If X is continuous, the f_X , $f_{X|Y}$ are densities; If X is also discrete, they are mass functions (Where needed we assume the conditional density exists)

Recap Bayes rule

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This gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) f_Y(y)}{f_X(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

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lackbox We need to interpret f_X or $f_{X|Y}$ as mass functions when X is discrete and as densities when X is a continuous and so on

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- ▶ This also implies $F_{X|Y}(x|y) = F_X(x)$ and $f_{X|Y}(x|y) = f_X(x)$

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If they are continuous, they have a joint density if

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- ► The properties of joint distribution function such as it being non-decreasing in each argument etc are easily seen to hold here too.
- Generalizing the special property of the df (relating to probability of cylindrical sets) is a little more complicated. (An exercise for you!)

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty);$$

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► Similarly we get

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- We obtain the marginal mass functions for a subset of the rv's also similarly where we sum over the remaining variables.

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- ► With these we can generally calculate most quantities of interest.

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- \triangleright For example when Z is continuous

$$F_{XY|Z}(x,y|z) = \lim_{\delta \to 0} P[X \le x, Y \le y | Z \in [z,z+\delta]]$$

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▶ For example, the first one above follows from

$$P[X = x, Y = y | Z = z] = \frac{P[X = x, Y = y, Z = z]}{P[Z = z]}$$

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$$f_{XYZ}(x, y, z) = f_{Z|XY}(z|x, y) f_{XY}(x, y) = f_{Z|XY}(z|x, y) f_{Y|X}(y|x) f_{X}(x)$$

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- We use similar notation for marginal and conditional distributions

Independence of multiple random variables

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 - (Recall definition of independence of a set of events)
- ► Independence implies that the marginals would determine the joint distribution.

Let a joint density be given by

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▶ Suppose we want to find the (marginal) joint distribution of X and Z.

$$f_{XZ}(x,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) \, dy$$
$$= \int_{z}^{x} K \, dy, \quad 0 < z < x < 1$$
$$= 6(x-z), \quad 0 < z < x < 1$$

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$$= \int_{0}^{1} \left(6x \, z \big|_{0}^{x} - 6 \, \frac{z^{2}}{2} \bigg|_{0}^{x} \right) \, dx$$

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$$= \int_{0}^{1} \left(6x^{2} - 6 \frac{x^{2}}{2} \right) dx$$

$$= 3 \frac{x^{3}}{3} \Big|_{0}^{1} = 1$$

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► Hence,

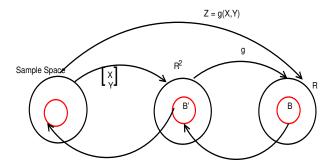
$$f_{Y|XZ}(y|x,z) = \frac{f_{XYZ}(x,y,z)}{f_{XZ}(x,z)} = \frac{1}{x-z}, \quad 0 < z < y < x < 1$$

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- ▶ Let *X,Y* be random variables on the same probability space.
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- ► This is analogous to functions of a single rv



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- $\blacktriangleright \text{ let } Z = g(X,Y)$
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$$F_Z(z) = P[Z \le z] = P[g(X, Y) \le z]$$

 \blacktriangleright For example, if X,Y are discrete, then

$$f_Z(z) = P[Z = z] = P[g(X, Y) = z] = \sum_{\substack{x_i, y_j: \ g(x_i, y_j) = z}} f_{XY}(x_i, y_j)$$

$$f_Z(z) = P[\min(X, Y) = z]$$

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= $P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$

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$$+P[X = z, Y = z]$$

$$f_{Z}(z) = P[\min(X, Y) = z]$$

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$$+ P[X = z, Y = z]$$

$$= \sum_{y>z} f_{XY}(z, y) + \sum_{x>z} f_{XY}(x, z) + f_{XY}(z, z)$$

$$f_{Z}(z) = P[\min(X, Y) = z]$$

$$= P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

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▶ Now suppose *X,Y* are independent and both of them have geometric distribution with the same parameter, *p*.

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- Now suppose X, Y are independent and both of them have geometric distribution with the same parameter, p.
- ► Such random variables are called **independent and identically distributed** or **iid** random variables.

▶ Now we can get pmf of Z as (note $Z \in \{1, 2, \dots\}$)

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

= $P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

= $P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$
= $p(1-p)^{z-1}$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

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$$= p(1-p)^{z-1}(1-p)^z * 2 +$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 + (p(1-p)^{z-1})^2$$

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$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 + (p(1-p)^{z-1})^2$$

$$= 2p(1-p)^{z-1}(1-p)^z + (p(1-p)^{z-1})^2$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

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$$= 2p(1-p)^{2z-1} + p^2(1-p)^{2z-2}$$

$$f_{Z}(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

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 $= (2-p)p(1-p)^{2z-2}$

$$\sum_{1}^{\infty} f_Z(z) = \sum_{1}^{\infty} (2 - p)p(1 - p)^{2z - 2}$$

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$$= F_{X_{1} \dots X_{n}}(z, \dots, z)$$

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▶ For example if all X_i are uniform over (0,1) and ind, then $F_Z(z) = z^n, \ 0 < z < 1$

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= $P[X>z]P[Y>z]$, using independence

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$$\begin{split} P[Z>z] &= P[\min(X,Y)>z] \\ &= P[X>z,Y>z] \\ &= P[X>z]P[Y>z], \quad \text{using independence} \\ &= (1-F_X(z))(1-F_Y(z)) \end{split}$$

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Hence,
$$F_Z(z) = 1 - (1 - F_X(z))(1 - F_Y(z))$$

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 We can once again find density of Z if X, Y are continuous ▶ Suppose X, Y are iid uniform (0, 1).

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 \blacktriangleright Hence, when X_i are iid, the df of Z is

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Remaining details are left as an exercise for you!!

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= $\sum_{x}^z P[X = k, Y = z - k]$

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▶ Z is Poisson with parameter $\lambda_1 + \lambda_2$

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Let
$$X,Y$$
 have a joint density f_{XY} . Let $Z=X+Y$
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Density of sum of independent random variables is the convolution of their densities.

$$f_{X+Y} = f_X * f_Y$$
 (Convolution)

$$f_X(x) = \lambda e^{-\lambda x}, \ x > 0$$

▶ Suppose *X*, *Y* are iid exponential rv's.

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▶ Suppose *X*, *Y* are iid exponential rv's.

$$f_X(x) = \lambda e^{-\lambda x}, \quad x > 0$$

▶ Let Z = X + Y. Then, density of Z is

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z - x) dx$$
$$= \int_0^z \lambda e^{-\lambda x} \lambda e^{-\lambda (z - x)} dx$$
$$= \lambda^2 e^{-\lambda z} \int_0^z dx = \lambda^2 z e^{-\lambda z}$$

► Thus, sum of independent exponential random variables has gamma distribution:

$$f_Z(z) = \lambda z \ \lambda e^{-\lambda z}, \ z > 0$$