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This shows: $f_Z(z) = \int_{-\infty}^{\infty} \left| \frac{1}{x} \right| f_{XY} \left(x, \frac{z}{x} \right) dx$

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$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x) f_{XY}(x,y) dy dx$$

$$= \int_{-\infty}^{\infty} h(x) f_X(x) dx$$

$$= E[h(X)]$$

$$E[h_1(X) h_2(Y)|Y] = h_2(Y)E[h_1(X)|Y]$$

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$$= h_2(y) E[h_1(X)|Y = y]$$

Example

lackbox Let X,Y be random variables with joint density given by

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

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Thus, X is exponential and Y is gamma.

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► Hence we have

$$EX = 1; Var(X) = 1;$$



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► Hence we have

$$EX = 1$$
; $Var(X) = 1$; $EY = 2$; $Var(Y) = 2$



$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

lacktriangle Let us calculate covariance of X and Y

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \, f_{XY}(x,y) \, dx \, dy$$

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \, f_{XY}(x,y) \, dx \, dy$$
$$= \int_{0}^{\infty} \int_{0}^{y} xy e^{-y} \, dx \, dy$$

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \ f_{XY}(x,y) \ dx \ dy$$
$$= \int_{0}^{\infty} \int_{0}^{y} xy e^{-y} \ dx \ dy = \int_{0}^{\infty} \frac{1}{2} y^{3} e^{-y} \ dy = 3$$

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

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- $\rho_{XY} = \frac{1}{\sqrt{2}}$

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

$$f_X(x) = e^{-x}, \ x > 0; \quad f_Y(y) = ye^{-y}, \ y > 0$$

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 $f_X(x) = e^{-x}, \ x > 0; \ f_Y(y) = ye^{-y}, \ y > 0$

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

$$f_{XY}(x,y) = e^{-y}, \ 0 < x < y < \infty$$

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$$f_{Y|X}(y|x) = \frac{f_{XY}(x,y)}{f_{X}(x)} = \frac{e^{-y}}{e^{-x}} = e^{-(y-x)}, \quad 0 < x < y < \infty$$

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▶ We can now calculate the conditional expectation

$$E[X|Y = y] = \int x f_{X|Y}(x|y) dx$$

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$$E[X|Y=y] = \int x \, f_{X|Y}(x|y) \, dx = \int_0^y x \, \frac{1}{y} \, dx = \frac{y}{2}$$

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$$E[X|Y] = \frac{Y}{2}$$

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$$= e^x \left(-y e^{-y} \Big|_x^\infty + \int_x^\infty e^{-y} \, dy \right)$$

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$$= e^{x} \left(-y e^{-y} \Big|_{x}^{\infty} + \int_{x}^{\infty} e^{-y} dy \right)$$
$$= e^{x} \left(x e^{-x} + e^{-x} \right) = 1 + x$$

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Thus, E[Y|X] = 1 + X

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Using this we can verify:

$$E[E[X|Y]] = E\left[\frac{Y}{2}\right]$$

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Using this we can verify:

$$E[E[X|Y]] = E\left[\frac{Y}{2}\right] = \frac{EY}{2} = \frac{2}{2} = 1$$

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$$E[E[X|Y]] = E\left[\frac{Y}{2}\right] = \frac{EY}{2} = \frac{2}{2} = 1 = EX$$

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$$E[E[X|Y]] = E[X]$$

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$$EX = \sum_{y} E[X|Y = y] f_Y(y)$$
 or $\int E[X|Y = y] f_Y(y) dy$

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▶ Can be used to calculate probabilities of events too

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▶ Can be used to calculate probabilities of events too

$$P(A) = E[I_A] = E[E[I_A|Y]]$$

• Let X be geometric and we want EX.

- ightharpoonup Let X be geometric and we want EX.
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$$E[X] = E[E[X|Y]]$$

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$$E[X] = E[E[X|Y]]$$

= $E[X|Y=1] P[Y=1] + E[X|Y=0] P[Y=0]$

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$$= 1 p +$$

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$$\Rightarrow EX (1 - (1 - p)) = p + (1 - p)$$

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$$\Rightarrow EX (1 - (1 - p)) = p + (1 - p) = 1$$

- \blacktriangleright Let X be geometric and we want EX.
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$$\begin{split} E[X] &= E[\ E[X|Y]\] \\ &= E[X|Y=1]\ P[Y=1] + E[X|Y=0]\ P[Y=0] \\ &= E[X|Y=1]\ p + E[X|Y=0]\ (1-p) \\ &= 1\ p + (1+EX)(1-p) \\ &\Rightarrow EX\ (1-(1-p)) = p + (1-p) = 1 \\ &\Rightarrow EX = \frac{1}{p} \end{split}$$

 $\blacktriangleright P[X=k|Y=1]=1 \ \ \mbox{if} \ \ k=1$ (otherwise it is zero) and hence E[X|Y=1]=1

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$$P[X = k | Y = 0] = \begin{cases} 0 & \text{if } k = 1\\ \frac{(1-p)^{k-1}p}{(1-p)} & \text{if } k \ge 2 \end{cases}$$

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$$E[X|Y=0] = \sum_{k=0}^{\infty} k (1-p)^{k-2} p$$

▶ P[X = k|Y = 1] = 1 if k = 1 (otherwise it is zero) and hence E[X|Y = 1] = 1

$$P[X = k | Y = 0] = \begin{cases} 0 & \text{if } k = 1\\ \frac{(1-p)^{k-1}p}{(1-p)} & \text{if } k \ge 2 \end{cases}$$

$$E[X|Y=0] = \sum_{k=2}^{\infty} k (1-p)^{k-2} p$$
$$= \sum_{k=2}^{\infty} (k-1) (1-p)^{k-2} p + \sum_{k=2}^{\infty} (1-p)^{k-2} p$$

▶ P[X = k|Y = 1] = 1 if k = 1 (otherwise it is zero) and hence E[X|Y = 1] = 1

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$$E[X|Y=0] = \sum_{k=2}^{\infty} k (1-p)^{k-2} p$$

$$= \sum_{k=2}^{\infty} (k-1) (1-p)^{k-2} p + \sum_{k=2}^{\infty} (1-p)^{k-2} p$$

$$= \sum_{k=2}^{\infty} k' (1-p)^{k'-1} p + \sum_{k=2}^{\infty} (1-p)^{k'-1} p$$

▶ P[X = k | Y = 1] = 1 if k = 1 (otherwise it is zero) and hence E[X | Y = 1] = 1

$$P[X = k | Y = 0] = \begin{cases} 0 & \text{if } k = 1\\ \frac{(1-p)^{k-1}p}{(1-p)} & \text{if } k \ge 2 \end{cases}$$

$$E[X|Y=0] = \sum_{k=2}^{\infty} k (1-p)^{k-2} p$$

$$= \sum_{k=2}^{\infty} (k-1) (1-p)^{k-2} p + \sum_{k=2}^{\infty} (1-p)^{k-2} p$$

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If we can guess value of $E[R_n]$ then we can prove it using mathematical induction

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- We verify it using mathematical induction. We know $E[R_1] = 1$

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 $\Rightarrow E[R_n] = n$

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- ► So, final number of comparisons depends on the 'number of rounds'

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▶ Suppose r_n is the number of comparisons. If we get (roughly) equal parts, then

$$r_n \approx n + 2r_{n/2} = n + 2(n/2 + 2r_{n/4}) = n + n + 4r_{n/4} = \dots = n \log_2(n)$$

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- ▶ If unlucky, in the worst case, $O(n^2)$ comparisons
- Question: 'on the average' how many comparisons?



Average case complexity of quicksort

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$$= (n-1) + \frac{2}{n} \sum_{j=1}^{n-1} E[M_k], \text{ (taking } M_0 = 0)$$

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$$\begin{split} E[M_n] &= E[\ E[M_n|X]\] = \sum_{j=1}^n E[M_n|X=j]\ P[X=j] \\ &= \sum_{j=1}^n E[(n-1) + M_{j-1} + M_{n-j}]\ \frac{1}{n} \\ &= (n-1) + \frac{2}{n} \sum_{j=1}^{n-1} E[M_k], \quad \text{(taking } M_0 = 0\text{)} \end{split}$$

► This is a recurrence relation. (A little complicated to solve)

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$$g^*(X) = E[Y|X]$$

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Let us prove this.

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$$= (g(X) - E[Y | X])^{2} + (E[Y | X] - Y)^{2}$$

$$+ 2(g(X) - E[Y | X])(E[Y | X] - Y)$$

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We have

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- We first show that expectation of last term on RHS above is zero.

$$E[(g(X) - E[Y \mid X])(E[Y \mid X] - Y)]$$

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= $E[E[(g(X) - E[Y \mid X])(E[Y \mid X] - Y) \mid X]$

$$\begin{split} &E\left[(g(X)-E[Y\mid X])(E[Y\mid X]-Y)\right]\\ =&\ E\left[\begin{array}{cc} E\left\{(g(X)-E[Y\mid X])(E[Y\mid X]-Y)\mid X\right\} \end{array}\right]\\ &\text{because}\quad E[Z]=E[\ E[Z|X]\] \end{split}$$

$$E[(g(X) - E[Y | X])(E[Y | X] - Y)]$$

$$= E[E[(g(X) - E[Y | X])(E[Y | X] - Y) | X]]$$

$$= E[(g(X) - E[Y | X]) E\{(E[Y | X] - Y) | X]]$$

$$E[(g(X) - E[Y \mid X])(E[Y \mid X] - Y)]$$

$$= E[E\{(g(X) - E[Y \mid X])(E[Y \mid X] - Y) \mid X\}]$$

$$= E[(g(X) - E[Y \mid X]) E\{(E[Y \mid X] - Y) \mid X\}]$$
because $E[h_1(X)h_2(Z)|X] = h_1(X) E[h_2(Z)|X]$

$$E[(g(X) - E[Y | X])(E[Y | X] - Y)]$$

$$= E[E\{(g(X) - E[Y | X])(E[Y | X] - Y) | X\}]$$

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$$= E[(g(X) - E[Y | X]) (E[Y | X] - E[Y | X))]$$

$$E[(g(X) - E[Y | X])(E[Y | X] - Y)]$$

$$= E[E[(g(X) - E[Y | X])(E[Y | X] - Y) | X]]$$

$$= E[(g(X) - E[Y | X]) E\{(E[Y | X] - Y) | X\}]$$

$$= E[(g(X) - E[Y | X]) (E\{(E[Y | X])|X\} - E\{Y | X\})]$$

$$= E[(g(X) - E[Y | X]) (E[Y | X] - E[Y | X))]$$

$$= 0$$

$$\begin{array}{rcl} (g(X) - Y)^2 & = & \left(g(X) - E[Y \mid X]\right)^2 + \left(E[Y \mid X] - Y\right)^2 \\ & & + 2 \Big(g(X) - E[Y \mid X]\Big) \Big(E[Y \mid X] - Y\Big) \end{array}$$

$$(g(X) - Y)^{2} = (g(X) - E[Y \mid X])^{2} + (E[Y \mid X] - Y)^{2} + 2(g(X) - E[Y \mid X])(E[Y \mid X] - Y)$$

► Hence we get

$$(g(X) - Y)^{2} = (g(X) - E[Y \mid X])^{2} + (E[Y \mid X] - Y)^{2} + 2(g(X) - E[Y \mid X])(E[Y \mid X] - Y)$$

▶ Hence we get

$$E[(g(X) - Y)^{2}] = E[(g(X) - E[Y | X])^{2}] + E[(E[Y | X] - Y)^{2}]$$

$$(g(X) - Y)^{2} = (g(X) - E[Y \mid X])^{2} + (E[Y \mid X] - Y)^{2} + 2(g(X) - E[Y \mid X])(E[Y \mid X] - Y)$$

▶ Hence we get

$$E \left[(g(X) - Y)^{2} \right] = E \left[(g(X) - E[Y \mid X])^{2} \right]$$

$$+ E \left[(E[Y \mid X] - Y)^{2} \right]$$

$$\geq E \left[(E[Y \mid X] - Y)^{2} \right]$$

$$(g(X) - Y)^{2} = (g(X) - E[Y \mid X])^{2} + (E[Y \mid X] - Y)^{2} + 2(g(X) - E[Y \mid X])(E[Y \mid X] - Y)$$

▶ Hence we get

$$E\left[\left.(g(X)-Y)^2\right.\right] = E\left[\left.(g(X)-E[Y\mid X])^2\right.\right] \\ + E\left[\left.(E[Y\mid X]-Y)^2\right.\right] \\ \ge E\left[\left.(E[Y\mid X]-Y)^2\right.\right]$$

 \triangleright Since the above is true for all functions q, we get

$$q^*(X) = E[Y \mid X]$$

Let X_1, X_2, \cdots be iid rv on the same probability space. Suppose $EX_i = \mu, \ \forall i.$

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$$E[S] = E[\; E[S|N] \;]$$

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▶ We have

$$E[S|N=n] = E\left[\sum_{i=1}^{N} X_i \mid N=n\right]$$

$$\begin{split} E[S|N=n] &= E\left[\sum_{i=1}^N X_i \mid N=n\right] \\ &= E\left[\sum_{i=1}^n X_i \mid N=n\right] \\ &= \operatorname{since} E[h(X,Y)|Y=y] = E[h(X,y)|Y=y] \end{split}$$

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$$E[S|N] = N\mu$$

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▶ Hence we get

$$E[S|N] = N\mu \quad \Rightarrow \quad E[S] = E[N]E[X_1]$$

$$\begin{split} E[S|N = n] &= E\left[\sum_{i=1}^{N} X_{i} \mid N = n\right] \\ &= E\left[\sum_{i=1}^{n} X_{i} \mid N = n\right] \\ &= \operatorname{since} E[h(X, Y)|Y = y] = E[h(X, y)|Y = y] \\ &= \sum_{i=1}^{n} E[X_{i} \mid N = n] = \sum_{i=1}^{n} E[X_{i}] = n\mu \end{split}$$

► Hence we get

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• Actually, we did not use independence of X_i .