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▶ If moment of order k is finite then so is moment of order s for s < k.

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In general

$$\frac{d^k M_X(t)}{dt^k}\bigg|_{t=0} = E[X^k]$$

$$M_X(t) = Ee^{tX} = E\left[1 + \frac{tX}{1!} + \frac{t^2X^2}{2!} + \frac{t^3X^3}{3!} + \frac{t^4X^4}{4!} + \cdots\right]$$

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$$\frac{d^3 M_X(t)}{dt^3} = 0 + 0 + 0 + \frac{3 * 2 * 1 * t^0}{3!} EX^3 + \frac{4 * 3 * 2 * t}{4!} EX^4 + \cdots$$

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▶ Hence we get

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(Exercise: Differentiate it twice to find EX^2 and hence show that variance is λ).

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- We are not saying moments uniquely determine the distribution; we are saying mgf uniquely determines the distribution

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► Characteristic function always exists because $|e^{itx}| = 1, \forall t, x$

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- For example,

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Characteristic Function

▶ The characteristic function of *X* is defined by

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• We would consider ϕ_X later in the course

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Hence, we get

$$f_X(0) = P_X(0); \ f_X(1) = \frac{P_X'(0)}{1!}; \ f_X(2) = \frac{P_X''(0)}{2!}$$

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► For (positive integer valued) discrete random variables, it is more convenient to deal with generating functions than mgf.

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▶ Note that for a given *p* there can be multiple values for *x* to satisfy the above.

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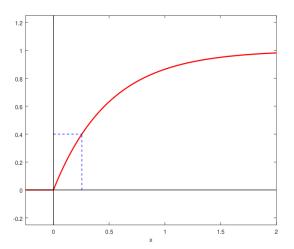
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- Let us see some examples.

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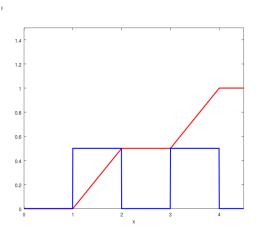


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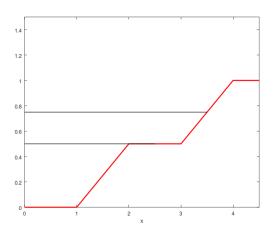
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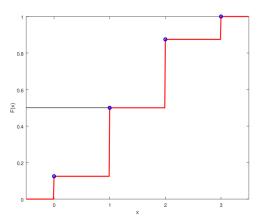
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- ▶ For $x_i \le x \le x_{i+1}$, we have $p \le F_X(x) \le p + P[X = x]$
- ▶ So, quantile of order p is not unique and all such x qualify.

▶ This situation is illustrated below



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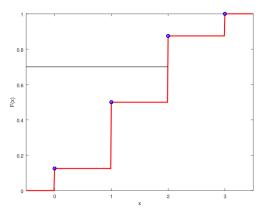
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- ▶ Similarly, for $x \ge x_{i+1}$ we have $F_X(x) > p + P[X = x]$.
- ▶ Thus quantile of order p is unique here.

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- Recall that the (standard) Cauchy density is given by

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▶ One can show that $\int_{-\infty}^{0} f_X(x) dx = 0.5$ and hence the median is at the origin.



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Markov inequality is often used in this form.



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This is known as the Chebyshev inequality.

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▶ This is true for all random variables and the RHS above does not depend on the distribution of *X*.

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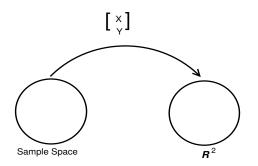
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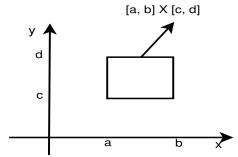
$$\mathcal{B}^2 = \sigma (\{B_1 \times B_2 : B_1, B_2 \in \mathcal{B}\})$$

where \mathcal{B} is the Borel σ -algebra we considered earlier, and \mathcal{B}^2 is the set of Borel sets of \Re^2 .

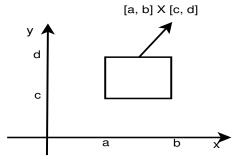
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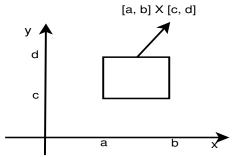


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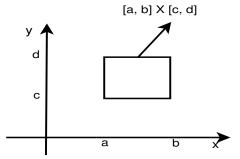
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- ▶ We saw that \mathcal{B} is also the smallest σ -algebra containing all intervals of the form $(-\infty, x]$.
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- ▶ F_{XY} fixes the probability of all cylindrical sets of the form $(-\infty, x] \times (-\infty, y]$ and hence uniquely determines the probability of all Borel sets of \Re^2 .

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▶ The joint distribution function is the probability of the intersection of the events $[X \le x]$ and $[Y \le y]$.

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- ▶ But there is another crucial property satisfied by F_{XY} .

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- ► We will now derive a similar expression in the case of two random variables.

- Recall that, for the case of a single rv, the probability of X being in any interval is given by the difference of F_X values at the end points of the interval.
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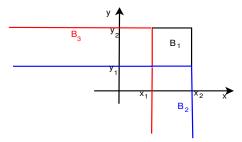
$$P[x_1 < X \le x_2] = F_X(x_2) - F_X(x_1)$$

- ► The LHS above is a probability. The RHS is non-negative because F_X is non-decreasing.
- ► We will now derive a similar expression in the case of two random variables.
- ► Here, the probability we want is that of the pair of rv's being in a cylindrical set.

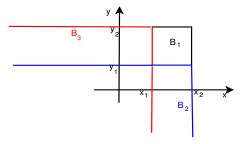
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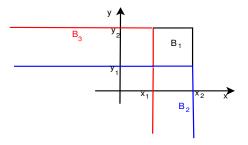


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$$B \triangleq (-\infty, x_2] \times (-\infty, y_2] = B_1 + (B_2 \cup B_3)$$

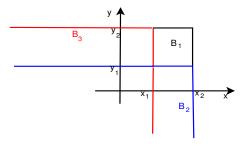
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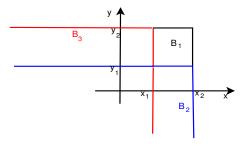


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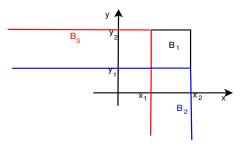
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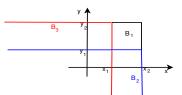
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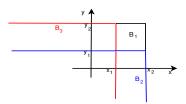
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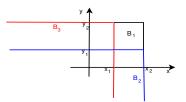
$$B_3 = (-\infty, x_1] \times (-\infty, y_2]$$

$$B_2 \cap B_3 = (-\infty, x_1] \times (-\infty, y_1]$$

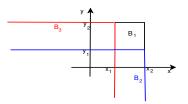




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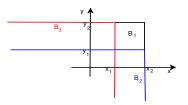
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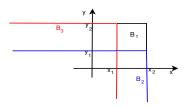


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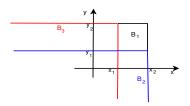
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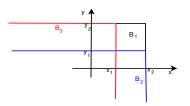
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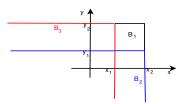
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➤ This is an additional condition that a function has to satisfy to be the joint distribution function of a pair of random variables

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

▶ Joint distribution function: $F_{XY}: \Re^2 \to \Re$

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It satisfies

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▶ Any $F: \Re^2 \to \Re$ satisfying the above would be a joint distribution function.