

The Geographic Effects of Carbon Pricing

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Abstract

This paper studies the heterogeneous impacts of carbon pricing on European regions. I follow the approach of [Känzig \(2023\)](#) and identify carbon policy shocks from changes in carbon futures price around regulatory events. The shock series is then combined with granular data on economic activity at the city- and county-level in Europe. I document that poorer regions are significantly more exposed to these shocks. Two years after a carbon policy shock the output of regions at the bottom quartile of the gross value added per capita distribution decreases 0.4 percentage points more relative to the output of regions in the top quartile. I investigate which channels might explain this result and find that the most important driver is across-country variation rather than sectoral compositions or within-country variation. The empirical evidence provided strongly encourages better coordination among European countries to avoid the economic costs of carbon pricing being unequally borne.

Keywords: Carbon pricing, macroeconomic effects, regional heterogeneity, sectoral heterogeneity

JEL classification: E32, H23, Q54, R11

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1 Introduction

The mitigation of climate change can be considered one of the most important challenges of our generation. Several policies have been adopted over the last decades to tackle the worldwide long-term shifts in temperatures and weather patterns. Among those policies, carbon pricing, i.e., setting the price of carbon to capture the external costs of greenhouse gas (GHG) emissions paid by the public, is one of the most important tools currently available. Nonetheless, the empirical evidence on the impact of this policy on the economy is rather limited. This is especially true at the regional level.

This paper studies the regional effects of carbon pricing in the Euro Area. A carbon policy shock tightening the carbon pricing regime causes a persistent fall in overall GHG emissions but also results in a strong contraction of the aggregate economy. The impact is extremely heterogeneous across regions. I document that the output of poorer regions in terms of gross value added (GVA) per capita is more affected by changes in the carbon price. I then evaluate plausible channels that can explain this finding. I show that variation across- rather than within-country or sectoral composition is the main driver behind the heterogeneous responses along the GVA per capita distribution.

To measure exogenous changes in the carbon price, I use the carbon policy shock series recently developed by [Känzig \(2023\)](#). He identifies 126 regulatory events that influenced the supply of emission allowances in the European Union Emissions Trading System (EU ETS). The series of carbon policy surprises is then computed from the change in the carbon futures price in a tight window around the regulatory news. The surprise series can be used as an instrument to estimate the dynamic causal effects of a carbon policy shock on the economy.

I evaluate whether regions are heterogeneously exposed to carbon pricing policies. To do so, I combine the carbon policy shock series with regional data at an extremely granular level (city- and county-level). The geographic data on gross value added and employment are taken from the European Regional Database (ERD) by Cambridge Econometrics and the responses of region-level economic activity are computed adopting a panel local projection à la [Jordà \(2005\)](#). I document that the output of poorer regions responds significantly more to carbon policy shocks than the output of rich regions.

Different channels that might explain this result are tested. On the one hand, the output of poorer regions tends to be produced more by labor-intensive sectors. On the other hand, capital- and labor-intensive sectors do not significantly differ in their sensitivity to carbon policy shocks so the heterogeneous regional sectoral composition cannot account for the

different responses of rich and poor regions. Neither does within-country heterogeneity. No significant difference is detected between the responses of rich and poor regions within the four largest countries in the sample, i.e., France, Germany, Italy, and Spain. Therefore, the across-country variation is the most important driver behind the main finding. In particular, I find the share of hand-to-mouth households to be extremely correlated with the countries' responsiveness to carbon policy shocks.

Carbon pricing has been found to be extremely effective in reducing GHG emissions. At the same time, some empirical evidence suggests that it can have detrimental effects on the economy. Different regions are heterogeneously exposed to changes in the price of carbon. Correctly understanding how and through which channels the regional characteristics influence the transmission of carbon policy shocks is of pivotal importance for policymakers to reduce the economic costs of carbon pricing. The findings of this paper suggest that fiscal policies coordinated at the European level to support poorer countries through the transition towards a greener economy might help to mitigate these adverse effects.

Related literature. This paper contributes to two strands of the literature. First, the results complement the large body of empirical evidence on the effects of carbon pricing on the economy. On the one hand, the effectiveness of carbon pricing for emission reductions is extensively supported by empirical evidence ([Ralf et al., 2014](#), [Andersson, 2019](#)). On the other hand, the impact on macroeconomic variables has been found to be negligible.

[Metcalf \(2019\)](#) and [Bernard and Kichian \(2021\)](#) focus on the effects of the British Columbia carbon tax on the GDP finding no significant impacts on GDP. Similarly, [Metcalf and Stock \(2020b\)](#) and [Metcalf and Stock \(2020a\)](#) do not find a negative relationship between the carbon taxes in European countries and employment or GDP growth. The result is extended to inflation by [Konradt and di Mauro \(2023\)](#) who focus on carbon taxes in Europe and Canada. [Benmir and Roman \(2022\)](#) find that carbon pricing shocks in the California cap-and-trade market have sizable effects on the economy, they result in an increase in the price of energy and in a decrease in energy consumption, wages, and asset returns.

The impact of carbon policies might not be limited to aggregate variables. [Ohlendorf et al. \(2021\)](#) adopt a meta-analysis approach on 53 empirical studies containing covering 39 countries and document that carbon pricing in the EU has affected substantially more lower-income households than richer ones. The carbon policy shocks used in this paper are developed by [Känzig \(2023\)](#). He shows that exogenous variation in the carbon price due to regulatory events leads to a fall in economic activity. On top of that, the consumption of

poorer households decreases significantly more than those of richer households and this is mainly due to general equilibrium effects. Combining the same carbon policy shocks with a French survey, [Hensel et al. \(2023\)](#) document that changes in the price of carbon lead to an increase in firms' inflation and own price expectation as well as realised price growth. Similarly, [Berthold et al. \(2022\)](#) study how different countries and firms respond to them. By adopting a panel VAR with 30 countries they document that more carbon-intensive countries are generally more affected. Moreover, brown sectors do not respond differently than the green sector but within a sector, brown firms tend to suffer more. I contribute by documenting how different regions are heterogeneously exposed to carbon pricing. Poorer regions are more sensitive to changes in carbon price mainly due to different country characteristics.

The second strand is the literature that answers macroeconomic questions with granular data at the regional level. Most macroeconomic variables are often limited over the time dimension. Therefore, more and more researchers have compensated for the lack of time variation by exploiting the cross-sectional variation in geographical data. [Auerbach and Gorodnichenko \(2012\)](#) and [Cloyne et al. \(2020\)](#) use cross-country panel data to estimate the fiscal multipliers. Similarly, [Nakamura and Steinsson \(2014\)](#) estimate the effects of government spending using regional variation across U.S. states in military build-ups.

Regional heterogeneity has been exploited as well to estimate the slope of the Phillips curve ([Hazell et al., 2021](#)) or to evaluate which regional characteristics matter the most in the transmission of monetary policy shocks. [Gallegos et al. \(2022\)](#) show that the share of financially constrained households strongly correlates with the output responsiveness to monetary shocks across Euro Area country. [Leahy and Thapar \(2022\)](#) and [Mangiante \(2023\)](#) document that the economic activity of U.S. states with an older demographic structure is more sensitive to shocks. [Herreno and Pedemonte \(2022\)](#) show that poorer U.S. cities respond more to monetary shocks. [Hauptmeier et al. \(2020\)](#) find a similar result for poorer Euro Area regions. I extend this literature by exploiting the geographical variation across Euro Area regions to study which country characteristics matter the most in the transmission of carbon policy shocks.

Road map. The remaining paper is organized as follows. Section 2 describes the data used in this paper. Section 3 shows the results of the main analysis. In Section 4, I perform a battery of robustness checks to strengthen the validity of the baseline results. Finally, Section 5 concludes.

2 Data

2.1 Regional data

The source of the geographical data is the European Regional Database (ERD) by Cambridge Econometrics¹. The ERD is based on Eurostat’s REGIO database and uses national statistics from the European Commission’s AMECO database and interpolation methods to fill some of the gaps. The data are based on Eurostat’s Nomenclature of Territorial Units for Statistics (NUTS), a hierarchical system that divides the economic territory of the EU into four levels. The highest level (NUTS0) corresponds to the nation-state and the lowest (NUTS3), roughly corresponds to the city and county levels. I will mainly focus on data at the NUTS3 level.

The main variables of interest are the gross value added, deflated to 2005 price levels, total employment, and population size. I will also focus on the breakdown of regional gross value added and employment into six sectors of the economy, corresponding to the disaggregation in NACE Rev.2 as agriculture, forestry, and fishing; industry less construction; construction; financial and business services; wholesale, retail, transport, accommodation, and food services, information and communication; and last, non-market services. I define the capital-intensive sector as the merging of the industry and construction sectors. Similarly, I merge the sectors of financial and business services, wholesale, retail, transport and food services, information and communication, and non-market services and define this as the labor-intensive sector.

The sample includes NUTS3 regions from the EA-19 member states, over the period 1999-2019. The countries considered are Austria (AT), Belgium (BR), Estonia (EE), Germany (DE), Greece (EL), Finland (FI), France (FR), Ireland (IE), Italy (IT), Latvia (LV), Lithuania (LT), Luxembourg (LU), Netherlands (NL), Portugal (PT), Slovenia (SI), Slovakia (SK), Spain (ES). The data are available only at an annual frequency and the final sample consists of 964 NUTS3 regions.

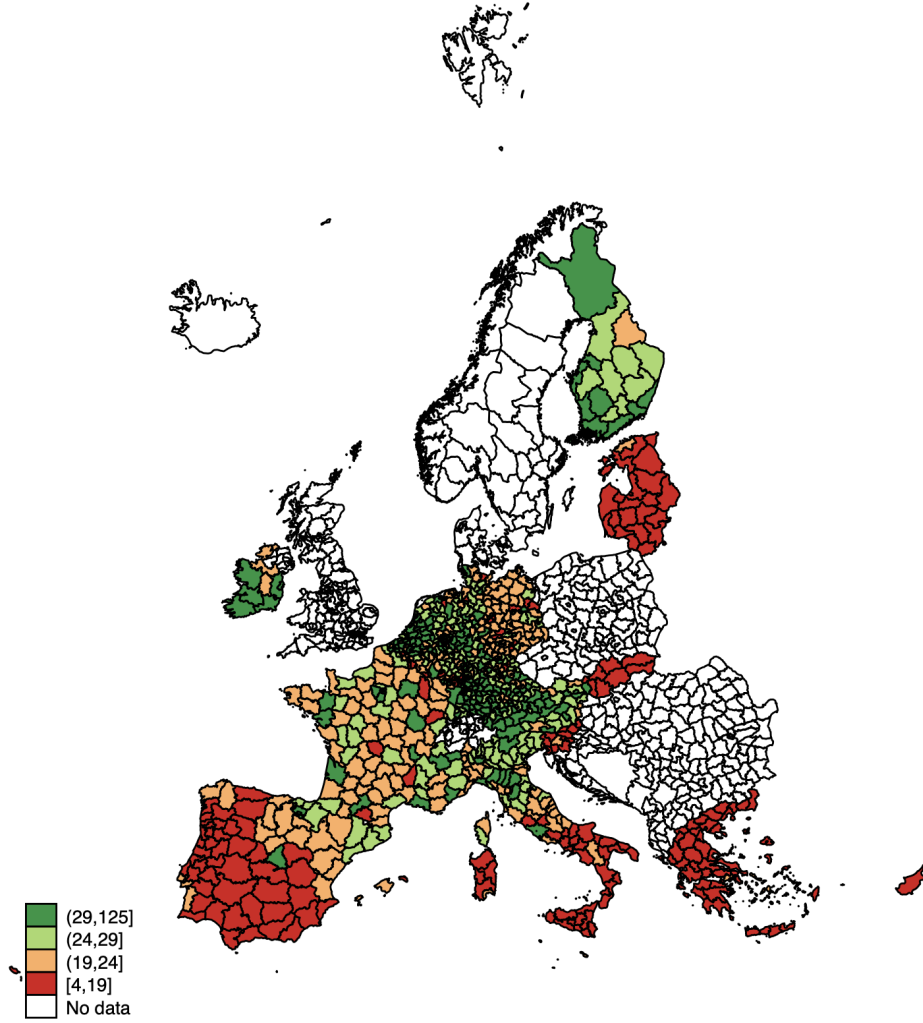
Figure 1 reports the real GVA per capita in 2015 at the NUTS3 level. As can be noticed, the geographic data displays an extremely high level of heterogeneity. This is true within as well as across countries.

2.2 Country-level data

The shares of financially constrained households at the country level are taken from [Gallegos et al. \(2022\)](#). The authors rely on the Eurosystem Household Finance and Consumption

¹Further details can be found in [Econometrics \(2017\)](#).

Figure 1: Heterogeneity in GVA per capita (thousands of euro)



Notes: The figure reports the real GVA per capita in 2015 at the NUTS3 level.

Survey (HFCS) to compute a measure of the share of Hand-to-Mouth (*HtM*) households for each country. They categorize a household as HtM if its liquid wealth is smaller than a certain share of monthly income. They further divide households into wealthy and poor HtM. A household is categorized as (*Wealthy HtM*) if it also has positive illiquid wealth otherwise is categorized as (*Poor HtM*). The average age of household heads is labeled *Age*.

Other country characteristics include a measure of trade openness defined as the sum of imports and exports as a share of GDP (*Trade openness*), a measure of how regulated labor markets (*ROL*) are, and the house price growth (*HP Growth*) calculated as the average quarterly year-on-year change in the house price index.

The remaining country characteristics are taken from Eurostat. They include the amount of CO2 emission per capita (*CO2 Emiss.*), the share of natural gas in gross available energy (*Gas Share in Energy*), the share of oil and petroleum product in gross available energy (*Petroleum Share in Energy*), and the ratio between gross available energy and GDP (*Energy intensity*).

2.3 Carbon policy shocks

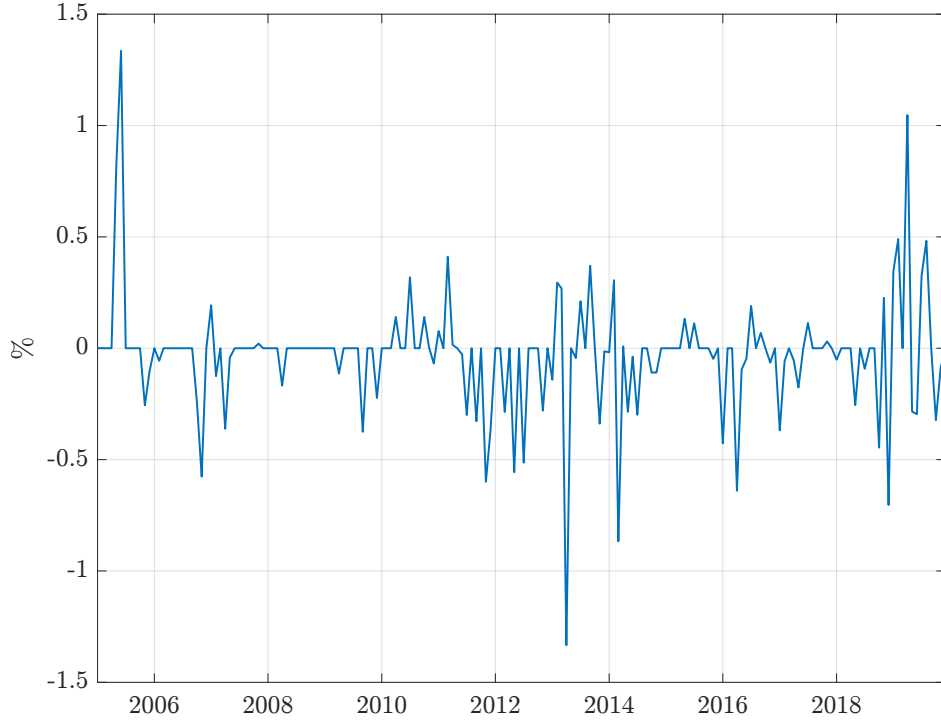
The carbon policy shocks are computed as in [Känzig \(2023\)](#). I here briefly summarise his approach but I refer to the original paper for a detailed description. The author exploits the fact that the European carbon market, established in 2005, operates under the cap and trade principle. This means that a cap is set on the overall amount of certain greenhouse gases that can be emitted by, for instance, power plants and industry factories. Within the cap, emission allowances are auctioned off and traded in different organized markets. Each year the cap decreases ensuring that total emissions fall.

[Känzig \(2023\)](#) identify 126 events during the period between 2005 and 2019 concerning the overall cap in the European Union Emissions Trading System (EU ETS), the free allocation of allowances, the auctioning of allowances as well as the use of international credits. He then defines carbon policy surprises from the changes in the futures price of the EU emission allowances (EUA) in the ICE since it is the most liquid market. In particular, the surprises are defined as the EUR change in carbon prices relative to the prevailing wholesale electricity price on the day before the event² The daily surprises are then aggregated into a monthly series by summing over the daily surprises in a given month. In months without any regulatory events, the series takes zero value. The resulting carbon policy surprise series is shown in Figure 2.

The carbon policy surprise series can be considered only a partial measure of the shock of interest. Indeed, it may not capture all relevant instances of regulatory news in the carbon market and could be subject to measurement errors. Therefore, to isolate the carbon policy shocks, [Känzig \(2023\)](#) uses the surprises as an external instrument in a VAR model with eight variables: the energy component of the HICP (*LHCPI Energy*), total GHG emissions, the headline HICP (*LHCPI*), industrial production (*LIP*), the unemployment rate (*UNEMP*), the policy rate (*E3M*), a stock market index (*LEuroStoxx50*), as well as the real effective exchange rate (*LREER*). The sample spans the period from January 1999 to December

²As alternative measures I also consider the difference in the settlement price and its percentage change. The results are not affected by the choice of the surprise measure.

Figure 2: The carbon policy surprise series



Notes: This figure shows the carbon policy surprise series, constructed by measuring the percentage change (blue solid line, left axis) as well as the change (red dashed line, right axis) of the EUA futures price around regulatory policy events.

2019. The VARs are estimated in levels using as controls six lags of all variables. Apart from unemployment and the policy rate, all variables enter in log levels. The carbon policy shocks are then extracted from the residuals of the monthly VAR (Stock and Watson, 2018) and are normalized to increase the energy component of the HICP by one percent on impact.

3 The heterogeneous impact of carbon policy shocks on regional output

I focus on how different European regions are heterogeneously exposed to carbon policy shocks. The carbon policy shock series extracted from the monthly VAR is combined it with regional data at NUTS3 level *European Regional Database* (ERD) by Cambridge Econometrics. The shocks are aggregated to an annual frequency by summing them over to match the frequency of the regional data.

The average region-level response to a climate policy shock is estimated using local projection à la [Jordà \(2005\)](#):

$$y_{i,t+h} = \alpha_{i,h} + \beta_h CPShock_t + \sum_{l=1}^L \theta_{i,h}^l y_{i,t-l} + \gamma_{i,h} X_{i,t-1} + \epsilon_{i,t+h}, \quad (1)$$

for $h = 0, \dots, 5$. $y_{i,t}$ is the dependent variable for region i at time t , $CPShock_t$ are the carbon policy shocks extracted from Proxy-VAR. I control for lags of the dependent variable, $\sum_{l=1}^L \theta_{i,h}^l y_{i,t-l}$, as well as regional variables, $X_{i,t}$. In the baseline specification L is equal to 2 and region fixed effects and population size are included. Standard errors are clustered at the country and year level. The main dependent variable is the log of real gross value added (GVA) at aggregate as well as sectoral levels. The coefficient of interest is β_h which captures the average response across regions of a carbon policy shock on the dependent variable for each horizon h .

I extend the baseline specification by introducing a categorical variable $D_{i,t}^p$ which identifies different percentiles P of the GVA per capita distribution. The categorical variable is then interacted with the climate policy shock $CPShock_t$:

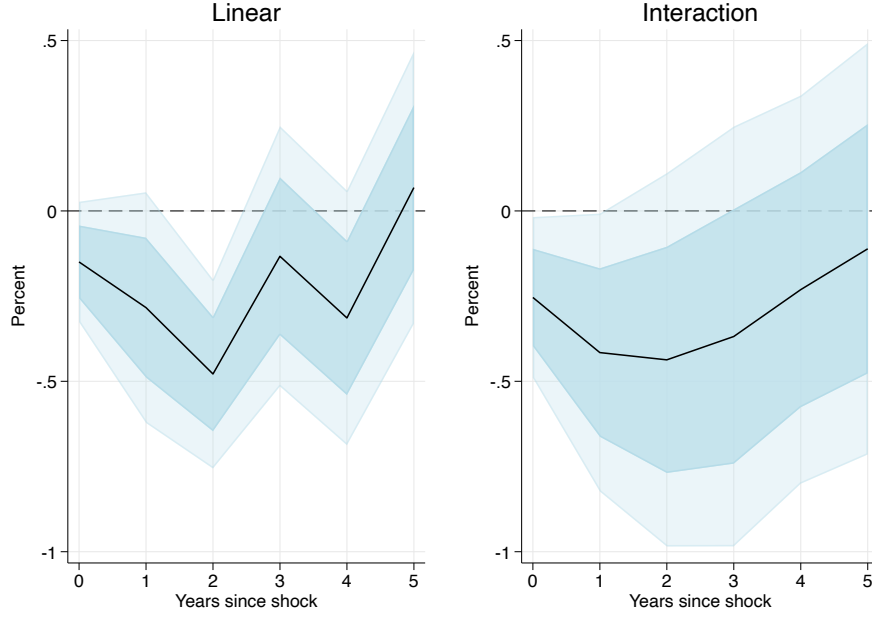
$$y_{i,t+h} = \alpha_{i,h} + \delta_{i,h} + \gamma_h D_{i,t}^p + \beta_h^p D_{i,t}^p CPShock_t + \sum_{l=1}^L \theta_{i,h}^l y_{i,t-l} + \gamma_{i,h} X_{i,t} + \epsilon_{i,t+h}, \quad (2)$$

where $\delta_{i,h}$ is year fixed effects. The coefficients β_h^p capture how regions are heterogeneously affected by carbon policy shocks according to their position in the GVA per capita distribution. The interaction coefficients can be interpreted as the differential response to a carbon policy shock of the different percentiles relative to the baseline group. I first compare the responses of the regions for which the GVA per capita belongs to the bottom 25% of the distribution relative to those of the top 25%.

The responses of regional real gross value added to a climate policy shock are reported in Figure 3. The left panel plots the estimated β_h coefficient at different horizons h from equation (1). The dark and light shaded areas are the 68 and the 90 percent confidence intervals respectively. Following a carbon policy shock that results in a one percent increase of the HICP energy component on impact, the gross value added significantly and persistently decreases.

The right panel of Figure 3 plots the estimated β_h^p coefficients from equation (2) for the bottom 25% regions in terms of GVA per capita relative to the top 25%. The negative

Figure 3: Impact of carbon policy shocks on log real GVA



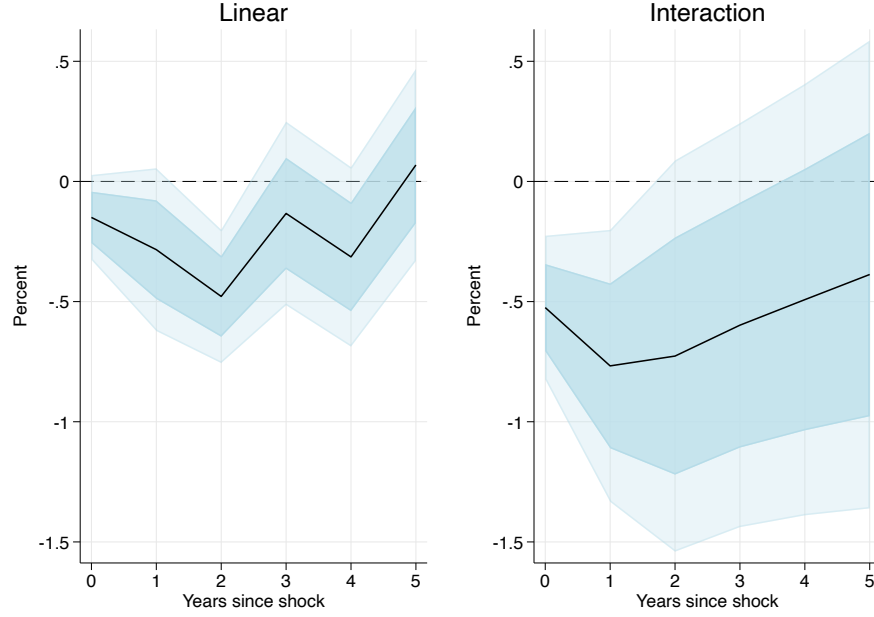
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

coefficients imply that the response of the “poor” regions is remarkably stronger than the response of the “rich” regions. After 2 years, the gross value added for the regions at the bottom of the GVA per capita distribution decreases by 0.4 percentage points more compared to that of the regions at the top. These results suggest there are sizable and economically meaningful differences in the extent to which poor and rich regions are impacted by carbon policy shocks.

The difference in responses is even more striking if we focus on the top and bottom 10% of the distribution as in Figure 4. The gap between the responses of gross value added of poor and rich regions is larger under this specification. Two years after the shocks the difference in GVA responses between the bottom 25% and top 25% regions in terms of GVA per capita is 0.4 percentage points whereas for the bottom 10% and top 10% regions is 0.65. This reinforces the conclusion that the responsiveness of output to carbon policy shocks is decreasing in the GVA per capita.

What could explain the sizable difference in the regional responses? Two channels might drive the result: sectoral composition and across-country variations. It might be the case that

Figure 4: Impact of carbon policy shocks on log real GVA, top/bottom 10%



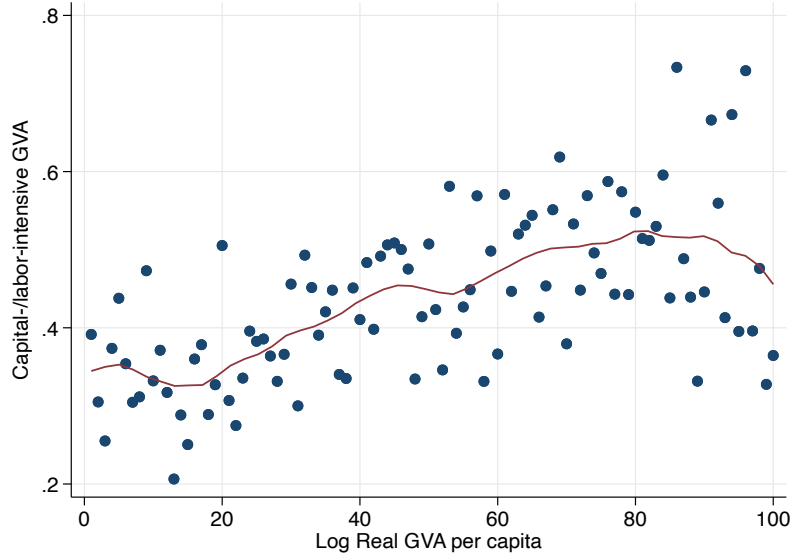
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 10% of the GVA per capita distribution relative to the top 10%.

the gross value added of the regions along the GVA per capita distribution is produced by different sectors. If the different sectors do not homogenously respond to carbon policy shocks, the different regional sectoral compositions might result in heterogeneous responses. A second potential explanation is that what I am capturing is variation across countries rather than across regions: poorer countries might be more exposed to these shocks than rich ones and the output of their regions respond more on average. I explore both channels in the next sections.

3.1 The role of sectoral composition

I now evaluate whether regions along the GVA per capita distribution specialize in different sectors. Figure 5 shows the relationship between the ratio of output in the capital-intensive sectors over the output in the labor-intensive sectors across different percentiles of the regional per-capita GVA distribution for 2015. There is a clear positive relationship between the two variables: the richer the region the more its output is produced by capital-intensive sectors. The industry ratio increases from 0.35% for the regions at the bottom of the distribution to 0.55% for those at the top.

Figure 5: Industry structure across percentiles



Notes: The figure shows the relationship between the ratio of output in the capital-intensive sectors over the output in the labor-intensive sectors and different percentiles of the regional per-capita GVA distribution for 2015.

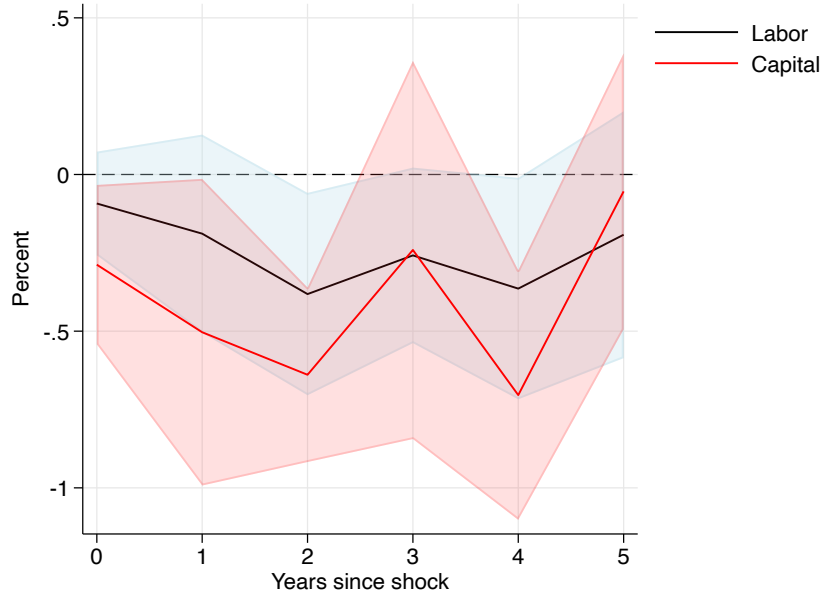
Therefore, if the labor-intensive sectors were to react more to carbon policy shocks, this would partially explain the stronger response observed for poorer regions. However, two pieces of evidence speak against industry structure as an explanation for the greater sensitivity of poorer regions.

First, the output of the labor-intensive sectors does not respond more to a climate policy shock than the output of capital-intensive sectors. Second, even within each type of industry, the differential response across the distribution remains intact: both the capital- and the labor-intensive production decrease more strongly and more persistently in the lower than in the upper part of the distribution.

Figure 6 shows the responses of the real gross value added for the capital- and labor-intensive sectors. In the first 2 years after the shock, the output from the capital-intensive sectors decreases slightly more than the output from the labor-intensive sectors although the difference is not statistically significant. This is most likely due to the higher energy use in the former group of sectors which is directly affected by the carbon policy shocks.

Figure 7 and Figure 8 report the output responses for the rich and the poor regions within each type of industry. As it can be noticed, conditioning for the type of industry does not change the fact that the gross value added of the poorer regions decreases remarkably more

Figure 6: Impact of carbon policy shocks on output across industries



Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region level log real GVA from the capital- and labor-intensive sectors. The solid lines are the point estimate and the shaded areas are the 90 percent confidence bands, respectively. The horizontal axis is in years.

in response to a carbon policy shock. Overall, we can conclude that sectoral composition does not explain the higher sensitivity to the shock of the poorer regions.

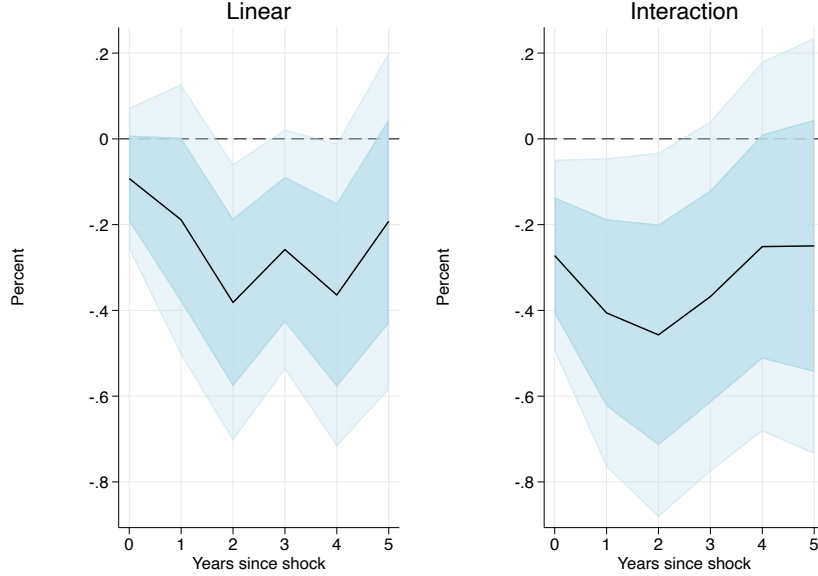
3.2 Across- vs within-country variation

Since sectoral composition at the regional level does not explain the different responses of output along the GVA per capita distribution, I now evaluate whether across-country heterogeneity might. Indeed, it might be the case that across- rather than within-country variation is driving the results.

To test this channel, I estimate the same equation (2) but for one country at a time. I focus in particular on the four largest countries in the sample and which have the highest within-country regional variations in GVA per capita: France, Germany, Italy, and Spain.

The aggregate and percentile-level responses of the GVA for each country are reported in Figure 9. The four countries are different in the magnitude and shape of the responses. However, looking at the right panel of each plot, we do not observe particularly different responses along the GVA distribution once we condition for the individual country. This

Figure 7: Impact of carbon policy shocks on log real GVA from labor-intensive sectors



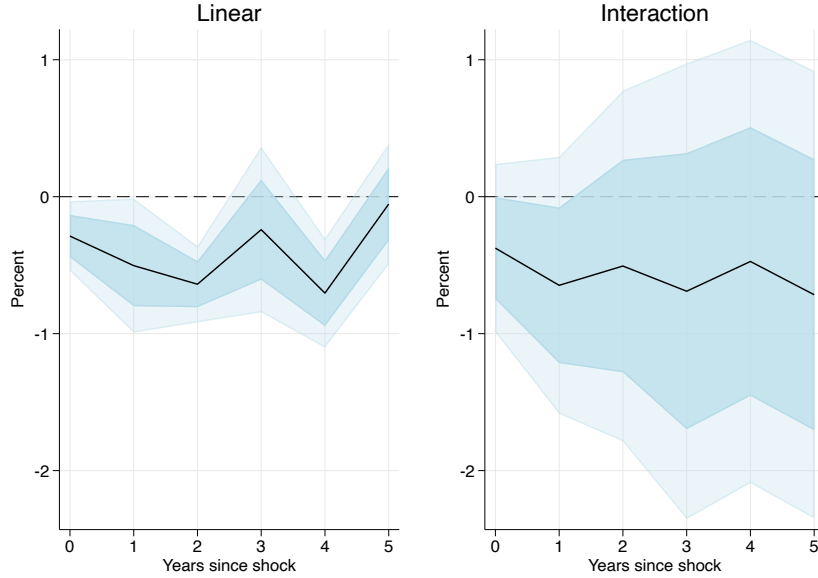
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA from labor-intensive sectors. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

suggests that across-country rather than within-country or across-region variation is what matters the most in terms of regional sensitivity to carbon policy shocks.

The same result can be more formally evaluated by considering an alternative definition of the GVA per capita percentiles. Instead of defining the categorical variable $D_{i,t}^p$ for the different quartiles of the overall GVA per capita distribution, I define it for the within-country distribution. In this way, the coefficients β_h^p of equation (2) compare the responses to carbon policy shocks of rich and poor regions *within* the same country.

Figure 10 shows the results of the within-country analysis. There is basically no difference in the responses of the regions at the bottom 25% and the top 25% of GVA per capita once we control for the country to which they belong. This striking result confirms that whether a region is poor or rich does not matter in terms of the sensitivity to carbon policy shocks. What matter is whether the region is part of a rich or poor country.

Figure 8: Impact of carbon policy shocks on log real GVA from capital-intensive sectors



Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA from capital-intensive sectors. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

3.3 The importance of country characteristics

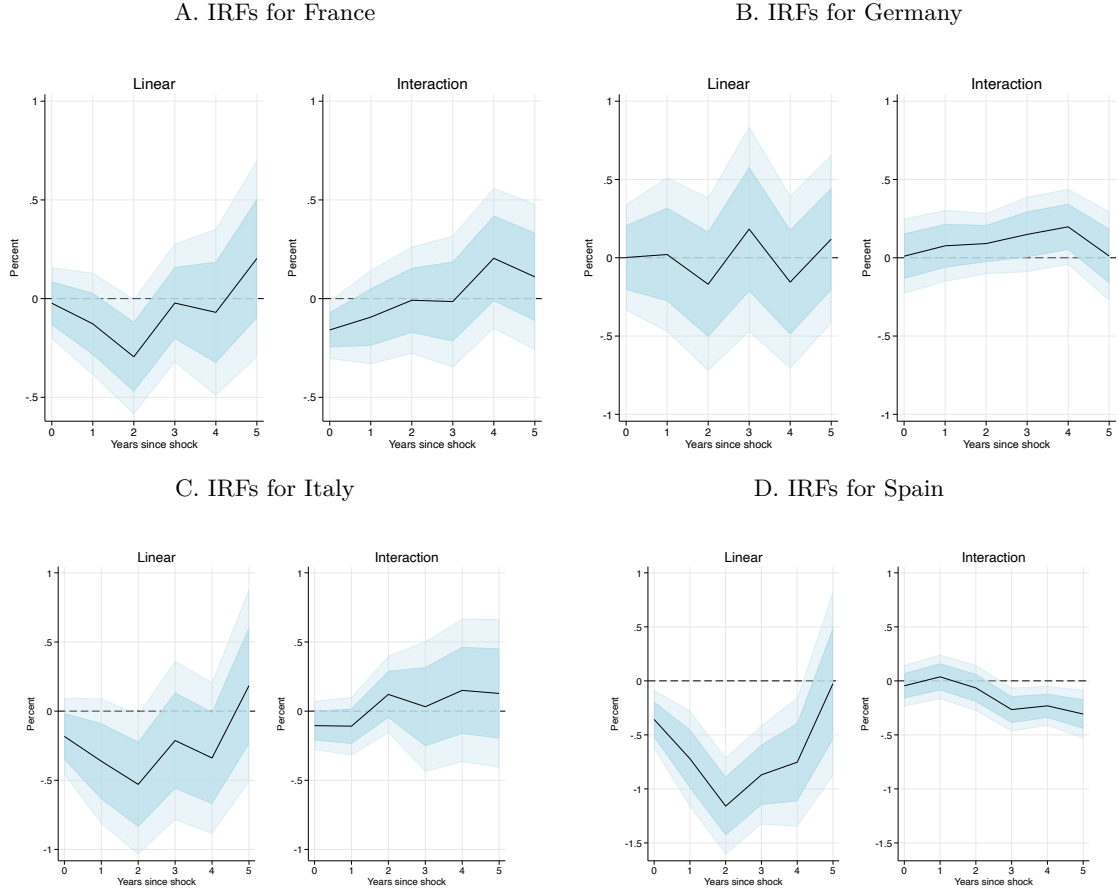
In the previous section, I show that across-country variation is the main driver behind the heterogeneous sensitivity of regions along the GVA per capita distribution. I now study which country characteristics correlate the most with the responsiveness to carbon policy shocks.

To do so, first I compute the impulse responses of real gross value added to a carbon policy shock for each country in the sample. Second, I define two measures of country responsiveness: The value of the impulse responses after two years, i.e., when the aggregate response is at its minimum, and the cumulative value to capture the persistency of the responses. Third, I relate the responsiveness measures to country characteristics commonly studied in the literature.

Figure 11 presents impulse responses of real gross value added for each country. On the one hand, the estimated impulse response functions reveal that carbon policy shocks lead to a significant decrease in output for all countries. On the other hand, the shape and the magnitude of the responses considerably differ across countries.

As in [Gallegos et al. \(2022\)](#), I gather country-level data on variables that are likely to be relevant for the sensitivity to carbon policy shocks. Subsequently, for each variable, I

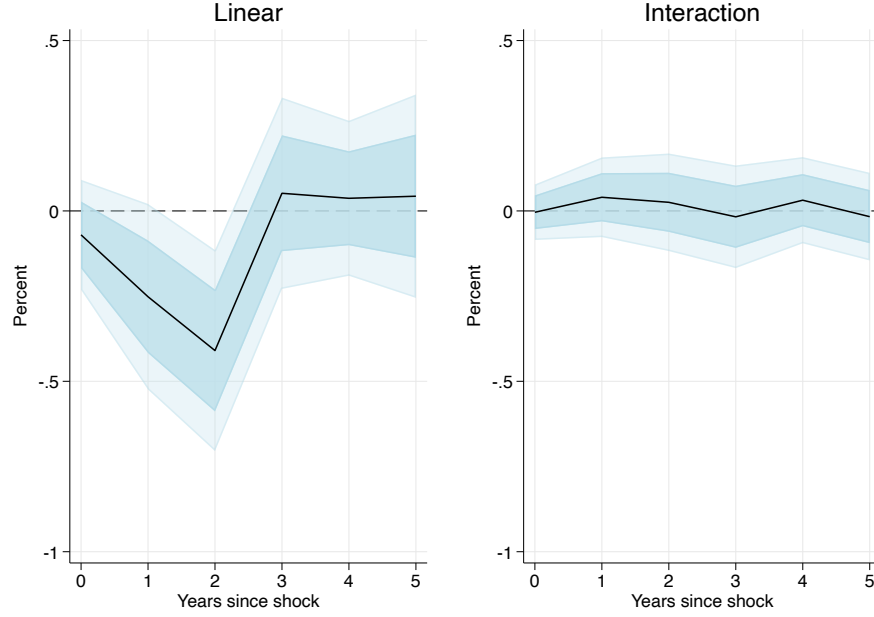
Figure 9: Impact of carbon policy shocks on the log real GVA for different countries



Notes: The left side of each panel plot for different countries the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%. Panel A: The plot shows the responses for France. Panel B: The plot displays the responses for Germany. Panel C: The plot reports the responses for Italy. Panel D: The plot shows the responses for Spain.

investigate whether (i) it is correlated with the response values two years after the shock as well as the cumulative responses, (ii) it is correlated with the gross value added per capita, and (iii) whether after controlling for the variable, the gross value added per capita still explains a significant part of the output responses we observe. The variables I consider are the gross value added per capita (*GVA per capita*), the ratio of output in the capital-intensive sectors over the output in the labor-intensive sectors (*Industry Ratio*), the ratio between gross available energy and GDP (*Energy intensity*), the share of natural gas in gross available energy (*Gas Share in Energy*), the share of oil and petroleum product in gross available energy (*Petroleum Share in Energy*), the amount of CO2 emission per capita (*CO2 Emiss.*), the share of Hand-to-Mouth households (*HtM*), the share of wealthy Hand-to-Mouth households

Figure 10: Impact of carbon policy shocks on log real GVA, within-country percentiles



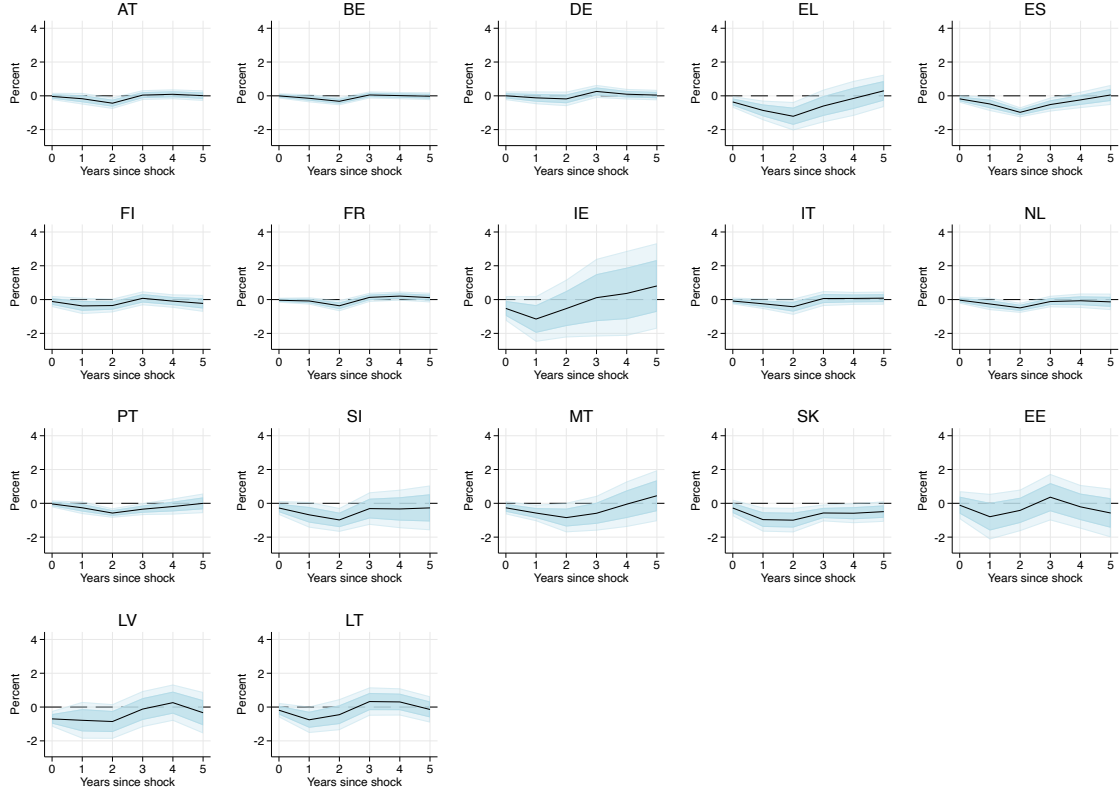
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

(*Wealthy HtM*), the share of poor Hand-to-Mouth households (*Poor HtM*), the average age of household heads *Age*, a measure of trade openness defined as the sum of imports and exports as a share of GDP (*Trade openness*), a measure of how regulated labor markets (*ROL*) are and the house price growth (*HP Growth*). See Section 2 for a description of the source of the data and how the variables are computed.

All results are summarized in Table 1. The first and second columns in the table present raw correlations between the two responsiveness measures and the different variables. In the third column, I report the correlations between the gross value added per capita and the variables that vary across the rows.

Both measures of responsiveness are positively and significantly correlated with the GVA per capita in line with the findings of the previous sections. This confirms that the poorer the country the more negative and more persistent the response to carbon policy shocks. Most of the other coefficients are close to zero and not significant. The only exceptions are the shares of Hand-to-Mouth and wealthy Hand-to-Mouth households which negatively correlate with the response values after two years, the cumulative responses as well as the GVA per capita.

Figure 11: Impact of carbon policy shocks on log real GVA at country-level



Notes: Each panel plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the log of real GVA at the country level. The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years.

The relationship implies that the higher the share of financially constrained households the stronger the output response.

In the theoretical literature, it is well established that the share of households who are Hand-to-Mouth plays a crucial role in determining the responsiveness of output to shocks (see, among others, [Bilbiie, 2008](#), [Auclert, 2019](#) and [Bilbiie, 2020](#)). As shown by [Bilbiie \(2019\)](#), the output response to shocks is amplified if the income elasticity of constrained agents with respect to aggregate income is larger than one and a larger fraction of constrained agents amplifies this channel. [Gallegos et al. \(2022\)](#) document empirically the relationship between financially constrained households and output responsiveness by focusing on monetary policy shocks. I extend their result by showing that the share of HtM households is a critical determinant of the country's sensitivity to carbon policy shocks as well.

Table 1: Correlations table

X	(1) $\rho(\text{IRF value after 2 y., } X)$	(2) $\rho(\text{Cum. IRF, } X)$	(3) $\rho(\text{GVA pc, } X)$	(4) $\rho(\text{IRF value after 2 y., GVA pc - } X)$	(5) $\rho(\text{Cum. IRF, GVA pc - } X)$
GVA per capita	0.528**	0.637***	1	NA	NA
Industry Ratio	-0.121	-0.250	0.00451	0.528**	0.638***
Energy Intensity	-0.0448	-0.236	-0.425*	0.561**	0.593**
Petroleum Share in Energy	-0.135	0.174	0.172	0.559**	0.616***
Gas Share in Energy	0.192	0.0718	0.162	0.503**	0.634***
CO2 Emiss.	0.289	0.194	0.506**	0.442*	0.625***
HtM	-0.607**	-0.595**	-0.586**	0.339	0.463**
Wealthy HtM	-0.719***	-0.740***	-0.747***	0.140	0.257*
Poor HtM	-0.00569	0.0939	0.148	0.638**	0.713***
Age	-0.246	-0.152	-0.383	0.580**	0.720***
Trade Openess	-0.206	-0.209	-0.146	0.502**	0.613***
ROL	-0.0459	-0.0301	-0.360	0.581**	0.670**
HP Growth	0.0558	-0.0962	-0.381	0.591**	0.642***

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Notes: The first column shows the correlation coefficients between estimated values of the IRFs two years after the shock and the country characteristics. The second column shows the correlation coefficients between the cumulative IRFs and the country characteristics. The third column reports the correlation coefficients between the GVA per capita and the country's characteristics. The fourth and fifth columns report the semipartial correlations. See the main text for information about the source and construction of the variables.

It is important to assess whether the strong correlation between GVA per capita and output responsiveness disappears once we control for these other variables. To get a sense of whether this could be the case, I calculate semipartial correlations between the measures of responsiveness and the GVA per capita. The semipartial correlation measures the strength of the linear relationship between the responsiveness measure and GVA per capita holding a third variable, varying across the rows in the table, constant for the GVA per capita. The semipartial correlations are reported in the fourth and fifth columns of Table 1. The size and the significance of the correlation coefficients are not remarkably affected by controlling for a third variable.

4 Robustness

In this section, I perform some robustness checks to strengthen the validity of the main results. First, I use total employment as a measure of economic activity instead of the gross value added. Second, I extend the sample to more countries that are part of the EU ETS. Third, I repeat the empirical analysis using geographical data at NUTS1 and NUTS2 level of aggregation. The plots are reported in Appendix A.

4.1 Employment

As an alternative measure of economic activity to gross value added I use total employment. The European Regional Database provides data on total employment at the NUTS3 level at

aggregate as well as sectoral levels. I then compute the same regional impulse responses to carbon policy shocks using employment as dependent variable.

Figure 12 reports the aggregate response (left panel) as well as the responses for the regions at the bottom 25% of the GVA per capita distribution relative to those at the top (right panel). Aggregate employment strongly decreases following the shock down to 0.25% after 2 years.

The response of the poorer regions is stronger than those of the rich ones. After 2 years, poor regions observe a decreases in the employment of 0.4 percentage points stronger than rich regions.

Figure 13 shows the sectoral responses of employment. Similarly to the baseline specification with gross value added, employment in the capital-intensive sectors responds slightly more than employment in the labor-intensive sectors. The difference in the responses across sectors is not statistically significant. In conclusion, using gross value added or employment as a measure of economic activity delivers similar results.

4.2 Extra countries

In the baseline analysis, the sample includes some of the EA-19 member states. However, there are more countries that are part of the EU Emissions Trading System and for which the geographical data from the European Regional Database are available. Therefore, I extend the sample by including the NUTS3 regions of Bulgaria, Cyprus, Croatia, the Czech Republic, Denmark, Hungary, Iceland, Liechtenstein, Malta, Norway, Poland, Romania, and Sweden.

The responses of regional real gross value added to a climate policy shock are reported in Figure 14. The inclusion of 13 extra countries does not affect either the magnitude or the shape of the responses. Following a carbon policy shock, the gross value added significantly decreases, and the response of the “rich” regions (top 25%) is remarkably more muted than the response of the “poor” regions.

4.3 Regional data at NUTS1 and NUTS2 level

The baseline sample consists of 964 NUTS3 regions. I repeat the same analysis using data at NUTS1 (80 regions) as well as NUTS2 (188 regions) level. The results are reported in Figure 15 and Figure 16 respectively.

The responses to a climate policy shock are extremely similar across the three NUTS levels of aggregation. This is not surprising since I have shown that most of the heterogeneity in exposure to the shocks comes from across-country rather than across-region variations.

5 Conclusion

Climate change cannot be solved by a single country or by implementing a specific policy. Only through coordination across countries and by adopting an appropriate policy mix we can hope to reduce its negative effects on the environment. Assessing the economic impact that these policies can have is crucial for the design of complementary fiscal policies targeted at mitigating the potential negative spillovers.

In this paper, I study the heterogeneous effects of carbon pricing across regions. This is done by combining the carbon policy shocks developed by [Känzig \(2023\)](#) with regional-level data for Europe. I document that the regions in poorer countries are significantly more exposed to these shocks. Following a tightening carbon policy shock, the gross value added of regions at the bottom quartile of the GVA per capita distribution decreases more than twice as much relative to the regions at the top quartile.

I show that different sectoral compositions or within-country variations do not explain this result. The main driver of the heterogeneous responses along the GVA per capita distribution is the across-country variation.

Since no single country can solve climate change by itself, no country should bear the economic costs of climate policies alone. The empirical findings I provide suggest that countries are heterogeneously impacted by variations in carbon price due to changes in regulation with poorer countries significantly more affected. Therefore, I believe more coordination is necessary among European governments to effectively allocate the burden of carbon pricing across countries.

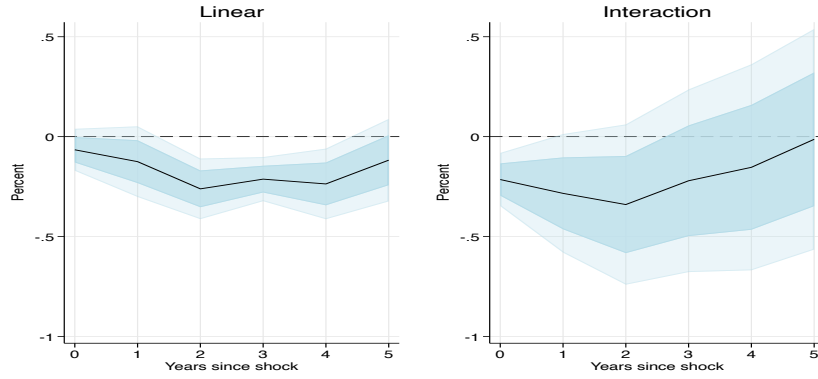
References

- Andersson, J. J. (2019). "Carbon Taxes and CO2 Emissions: Sweden as a Case Study". *American Economic Journal: Economic Policy*, 11(4): 1–30.
- Auclert, A. (2019). "Monetary policy and the redistribution channel". *American Economic Review*, 109(6): 2333(67).
- Auerbach, A. J. and Gorodnichenko, Y. (2012). "Fiscal Multipliers in Recession and Expansion". *Fiscal Policy After the Financial Crisis*, University of Chicago Press, pp. 63–98.
- Benmir, G. and Roman, J. (2022). "The Distributional Costs of Net-Zero: A HANK Perspective". *Working paper*.
- Bernard, J.-T. and Kichian, M. (2021). "The Impact of a Revenue-Neutral Carbon Tax on GDP Dynamics: The Case of British Columbia". *The Energy Journal*, 42(3).
- Berthold, B., Cesa-Bianchi, A., and Pace, F. D. (2022). "The economic effects of climate policies: an empirical investigation". *Forthcoming*.
- Bilbiie, F. (2008). "Limited asset markets participation, monetary policy and (inverted) aggregate demand logic". *Journal of Economic Theory*, Elsevier, vol. 140(1), pages 162-196.
- Bilbiie, F. (2019). "Monetary Policy and Heterogeneity: An Analytical Framework". *Mimeo*.
- Bilbiie, F. (2020). "The New Keynesian cross". *Journal of Monetary Economics*, 114: 90-108.
- Cloyne, J., Jorda, O., and Taylor, A. (2020). "Decomposing the Fiscal Multiplier". *Federal Reserve Bank of San Francisco Working Paper*, 2020-12.
- Econometrics, C. (2017). "European Regional Database: Technical Note".
- Gallegos, J.-E., Almgren, M., Kramer, J., and Lima, R. (2022). "Monetary Policy and Liquidity Constraints: Evidence from the Euro Area". *AEJ: Macroeconomics (forthcoming)*.
- Hauptmeier, S., Holm-Hadulla, F., and Nikalixi, K. (2020). "Monetary Policy and Regional Inequality". *ECB Working Paper No. 2385*.
- Hazell, J., Herreno, J., Nakamura, E., and Steinsson, J. (2021). "The Slope of the Phillips Curve: Evidence from U.S. States". *Working paper*.

- Hensel, J., Mangiante, G., and Moretti, L. (2023). "Carbon Pricing and Inflation Expectations: Evidence from France". *Working paper*.
- Herreno, J. and Pedemonte, M. (2022). "The Geographic Effects of Monetary Policy". *Federal Reserve Bank of Cleveland Working Paper Series*.
- Jordà, O. (2005). "Estimation and Inference of Impulse Responses by Local Projections". *American Economic Review*, 95 (1), 161–182.
- Känzig, D. R. (2023). "The unequal economic consequences of carbon pricing". *Working paper*.
- Konradt, M. and di Mauro, B. W. (2023). "Carbon Taxation and Inflation: Evidence from the European and Canadian Experience". *Journal of the European Economic Association*.
- Leahy, J. and Thapar, A. (2022). "Age Structure and the Impact of Monetary Policy". *American Economic Journal: Macroeconomics*, 14 (4): 136-73.
- Mangiante, G. (2023). "Demographic Trends and the Transmission of Monetary Policy". *Working paper*.
- Metcalf, G. E. (2019). "On the economics of a carbon tax for the United States". *Brookings Papers on Economic Activity*, 2019(1): 405–484.
- Metcalf, G. E. and Stock, J. H. (2020a). "Measuring the Macroeconomic Impact of Carbon Taxes". *AEA Papers and Proceedings*, 110: 101–06.
- Metcalf, G. E. and Stock, J. H. (2020b). "The Macroeconomic Impact of Europe’s Carbon Taxes". *NBER Working Paper*.
- Nakamura, E. and Steinsson, J. (2014). "Fiscal Stimulus in a Monetary Union: Evidence from US Regions". *American Economic Review* 104(3): 753–92.
- Ohlendorf, N., Jakob, M., Minx, J. C., Schroder, C., and Steckel, J. C. (2021). "Distributional impacts of carbon pricing: A meta-analysis". *Environmental and Resource Economics*, 78(1): 1–42.
- Ralf, M., Preux, L. B. D., and Wagner, U. J. (2014). "The impact of a carbon tax on manufacturing: Evidence from microdata". *Journal of Public Economics*, 117: 1–14.
- Stock, J. H. and Watson, M. W. (2018). "Identification and Estimation of Dynamic Causal Effects in Macroeconomics Using External Instruments". *The Economic Journal*, Volume 128 Issue 610.

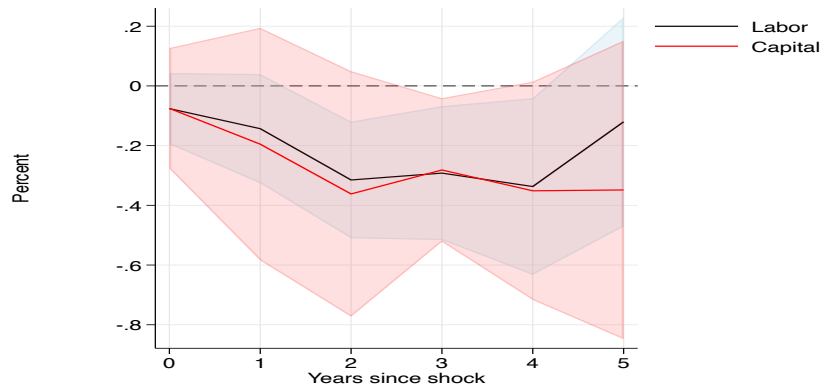
A Additional figures and tables

Figure 12: Impact of carbon policy shocks on log employment



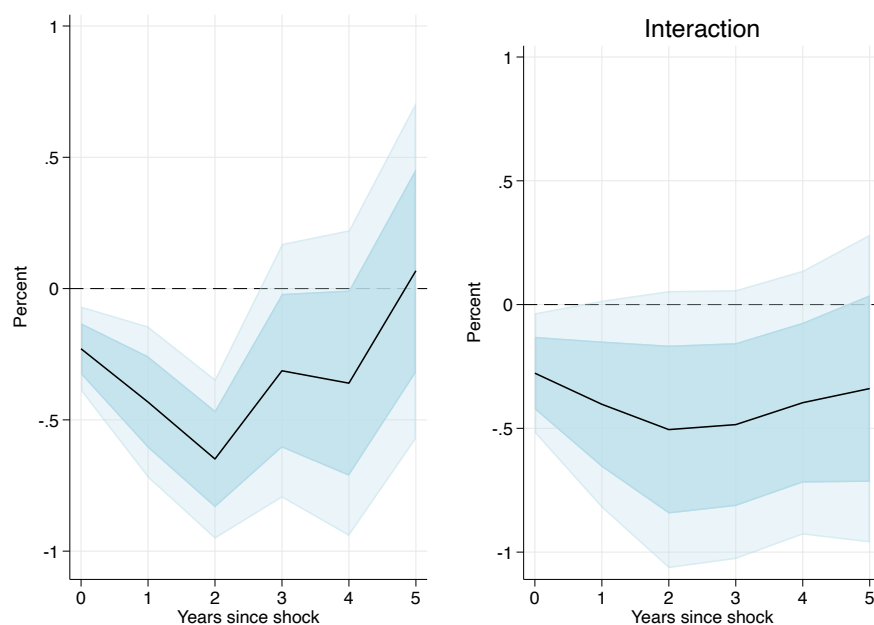
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log employment. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

Figure 13: Impact of carbon policy shocks on employment across industries



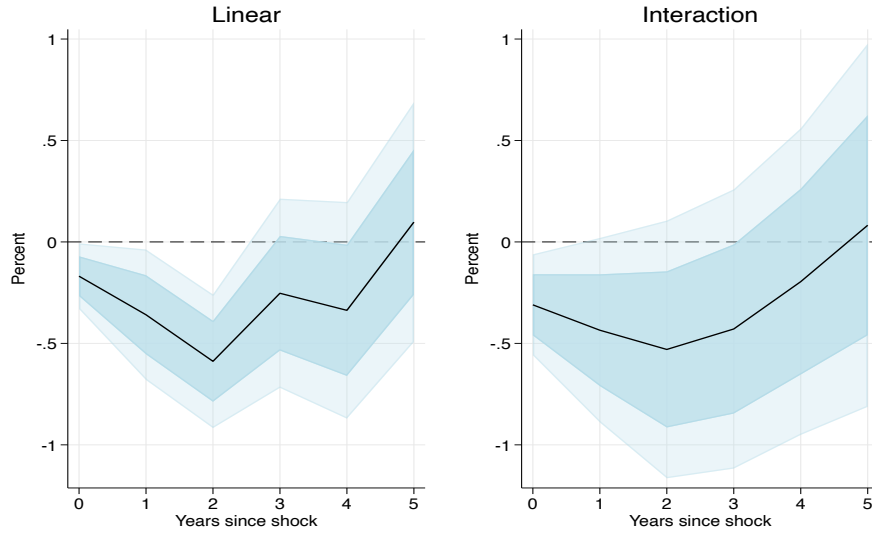
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region level log employment from the capital- and labor-intensive sectors. The solid lines are the point estimate and the shaded areas are the 90 percent confidence bands, respectively. The horizontal axis is in years

Figure 14: Impact of carbon policy shocks on log real GVA, extra countries



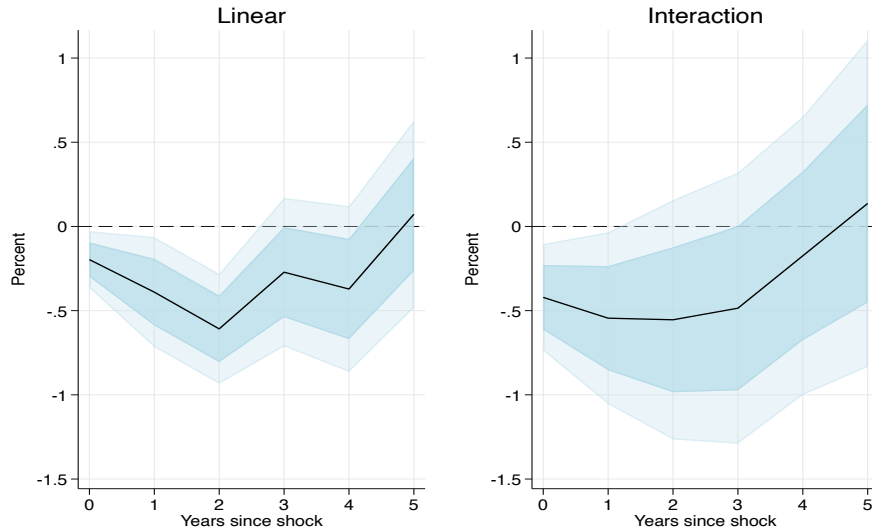
Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

Figure 15: Impact of carbon policy shocks on log real GVA, NUTS1



Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.

Figure 16: Impact of carbon policy shocks on log real GVA, NUTS2



Notes: The left panel of the figure plots the response to a carbon policy shock, normalized to increase the HICP energy by 1 percent on impact, for the region-level log of real GVA. The solid lines are the point estimate and the dark and light shaded areas are the 68 and 90 percent confidence bands, respectively. The horizontal axis is in years. The right panel reports the interaction coefficients between the carbon policy shock and the categorical variable identifying the bottom 25% of the GVA per capita distribution relative to the top 25%.